

**Article** 

https://doi.org/10.1038/s43587-022-00293-x

# Immune system-wide Mendelian randomization and triangulation analyses support autoimmunity as a modifiable component in dementia-causing diseases

Received: 24 February 2022

Accepted: 5 September 2022

Published online: 14 October 2022

Check for updates

Joni V. Lindbohm ® <sup>1,2,3</sup> ⋈, Nina Mars ® <sup>1,4</sup>, Pyry N. Sipilä ® <sup>3</sup>, Archana Singh-Manoux<sup>2,5</sup>, Heiko Runz<sup>6</sup>, FinnGen\*, Gill Livingston ® <sup>7,8</sup>, Sudha Seshadri <sup>9,10,11,12</sup>, Ramnik Xavier <sup>1,13,14</sup>, Aroon D. Hingorani ® <sup>15,16,17</sup>, Samuli Ripatti ® <sup>1,3,4,18</sup> and Mika Kivimäki ® <sup>2,3,7,18</sup>

Immune system and blood-brain barrier dysfunction are implicated in the development of Alzheimer's and other dementia-causing diseases, but their causal role remains unknown. We performed Mendelian randomization for 1,827 immune system- and blood-brain barrier-related biomarkers and identified 127 potential causal risk factors for dementia-causing diseases. Pathway analyses linked these biomarkers to amyloid-β, tau and α-synuclein pathways and to autoimmunity-related processes. A phenome-wide analysis using Mendelian randomization-based polygenic risk score in the FinnGen study (n = 339,233) for the biomarkers indicated shared genetic background for dementias and autoimmune diseases. This association was further supported by human leukocyte antigen analyses. In inverse-probability-weighted analyses that simulate randomized controlled drug trials in observational data, antiinflammatory methotrexate treatment reduced the incidence of Alzheimer's disease in high-risk individuals (hazard ratio compared with no treatment, 0.64,95% confidence interval 0.49-0.88, P=0.005). These converging results from different lines of human research suggest that autoimmunity is a modifiable component in dementia-causing diseases.

Due to limited success in drug trials targeting the amyloid- $\beta$  pathway, recent dementia research has explored alternative therapeutic targets from biomarkers linked to immune system dysfunction  $^1$ . This new focus has been supported by epidemiological studies that have linked chronic inflammatory diseases (for example, diabetes, autoimmune diseases and severe infections) to increased risk of dementias  $^{2-4}$ . In healthy state, the blood–brain barrier (BBB) protects the central nervous system (CNS) from peripheral neurotoxic molecules and pathogens, keeping the CNS immune privileged  $^{2-4}$ . However, aging  $^5$  and peripheral inflammation that arises from low-grade systemic inflammation  $^{6.7}$  and

infections and isrupt this function A dysfunctional BBB may promote expression of endothelial adhesion molecules and chemokines, leading to migration of peripheral leukocytes to the CNS. These processes activate the central immune system and are hypothesized to expose the CNS to prolonged neuroinflammation and subsequent neurodegeneration, which is supported by recent plasma proteomics studies  $^{7,10}$ .

However, evidence on causal associations between a dysfunctional peripheral immune system, BBB and dementia-causing diseases remains limited. While some studies have observed that higher circulating C-reactive protein, IL-1, IL-6, tumor necrosis factor- $\alpha$  (TNF- $\alpha$ ) and

A full list of affiliations appears at the end of the paper. Me-mail: joni.lindbohm@helsinki.fi

CD4 cell count may increase the risk of dementias<sup>11-15</sup>, these datasets are relatively small and captured only a limited number of biomarkers. A Mendelian randomization (MR) approach that uses large Genome-Wide Association Study (GWAS) libraries for unconfounded genetic proxies for biomarkers would allow a more comprehensive examination of the immune system and BBB biomarkers. This method enables the integrated use of data from multiple independent studies, testing of causality (although under strong assumptions), and has informed drug development<sup>16</sup>. As such, MR based on GWAS libraries is appealing in regard to explorative analyses of potential therapeutic targets for dementia-causing diseases. A complementary approach to improvement of reliability is triangulation, in which alternative methods, study designs and biomarkers with different sources of bias are used to test a common hypothesis<sup>17</sup>. If these converge, the results are more robust.

Here, we combine six studies using MR and triangulation to gain new insights into dementia etiology and to identify drug targets and anti-inflammatory medications for repurposing for dementia-causing diseases (Fig. 1 and Table 1). In the first study (study 1), we perform MR analyses based on GWAS libraries for a total of 1,827 peripheral immune system and BBB biomarkers to identify causal associations with dementia-causing diseases (including Alzheimer's and Parkinson's disease, vascular dementia, frontotemporal dementia and cognitive performance). The findings suggest that autoimmune biomarkers may play an important role in disease etiology. The second study (study 2) uses pathway analyses to identify biological processes in which these biomarkers are enriched and provides additional support for the autoimmune hypothesis. Studies 3-6 lend further, consistent support for the autoimmune hypothesis from four analyses that are independent of studies 1 and 2: plasma proteomics, polygenic risk scores, human leukocyte antigen (HLA) allele analyses and inverse-probability-weighted (IPW) survival analysis that simulates a randomized controlled trial (RCT) design using observational data. We identify several potential new drug targets for dementia-causing diseases and provide evidence of a shared genetic background between dementia-causing and autoimmune diseases. Based on the different lines of research from the six studies, we propose that dementia-causing diseases may have an inflammatory autoimmune component that is modifiable with currently available anti-inflammatory medications and new therapeutics targeting the identified biomarkers.

### Results

### MR and plasma protein analyses

In the discovery step, we used MR to identify potential causal risk factors for dementias. The MR-Base search provided 1,140 biomarkers for the immune system (Supplementary Data 1) and 687 biomarkers for BBB (Supplementary Data 2). A total of 253 immune system and 130 BBB biomarkers passed the false discovery rate (FDR) correction of 5%. After removal of duplicates, 127 unique biomarkers remained: 69 were related to immune cells, five to tumor necrosis factors, five to immunoglobulins, five to interleukins, five to cell membranes, four to complement components, three to platelet characteristics, two to interferon, two to metabolites, two to adhesion molecules, one to chemokines, one to endothelium, one to erythrocyte characteristics and 22 to other immune system- or BBB-related processes. All the biomarkers associated with dementia-causing diseases at P < 0.00052 in MR analyses using either the Wald ratio (when only one single-nucleotide polymorphism (SNP) was available) or the inverse-variance-weighted method (IVW) (when two or more SNPs were available) (Figs. 2 and 3, Extended Data Figs. 1 and 2 and Supplementary Data 3 and 4). While for three outcomes there was evidence of horizontal pleiotropy, MR sensitivity analyses showed no strong evidence of reverse causality for any of the biomarkers (Supplementary Table 1 and Extended Data Figs. 3 and 4). Of the 127 biomarkers, 49 were proteins, and for these, we identified 25 cis and 71 trans protein quantitative loci (pQTLs) (Supplementary Data 5). The cis loci were for AZGP1, BIN1, C1R, C4B, CFB, CD33, CD40, CNTN2,

FCGR2A, GPNMB, IFNAR1, IL-27 and NEGR1. Four of these—AZGP1, CD33, FCGR2A and GPNMB—had three or more SNPs available and passed MR sensitivity analyses. In addition, several CD20- and CD33-expressing leukocytes increased the risk of Alzheimer's disease; CD11-expressing leukocytes increased, and CD27- and CXCR1-expressing leukocytes decreased, Parkinson's disease risk in MR sensitivity analyses. The few off-target associations for the 127 biomarkers (FDR < 5%) were mainly with type 1 diabetes and low-density lipoprotein cholesterol (Extended Data Fig. 5).

Eight proteins were associated with all-cause dementia outcome and their pQTLs were centered within 500 kilobases (kb) from the APOE gene, one of the strongest genetic risk factors for late-onset Alzheimer's disease (Supplementary Data 6) reduced. To examine whether these associations were attributable to the effects of APOE, we measured plasma proteins associated with these eight pQTLs in the Whitehall II cohort study (n = 6,545). The study included as an outcome a 20-year follow-up of all-cause dementia but did not have data on dementia subtypes. Of these eight proteins, two (LRRN1 and IFIT2) were associated with dementia and one (IFIT2) remained significantly associated with reduced risk of dementia after adjustment for APOE status (Fig. 4).

The remaining 119 non-*APOE*-linked biomarkers were more outcome specific and did not show similar enrichment around high-risk genes for Alzheimer's disease, including *APP*, *PSEN1*, *PSEN2*, *ADAM10*, *TREM2*, *PLD3* and *UNCSC*. Instead, these were characterized by inflammatory, chemokine, complement and adhesion processes (C1Q, C1R, C4B, CCL1, CDHR5, GPNMB, IL-1 $\beta$ , IL-17, IL-27, IL-37, LTBR, PTP1B and SIGIRR), antigen-presenting and immune checkpoints (HLA-DR, HLA-DQ, BAFFR, C1R, C1Q, CD11, CD19, CD20, CD33, CD40, CX3CR1, PD-1 and PDL-1) and BBB tight-junction-related biomarkers (TJP1, AIMP1 and BIN1).

### **Pathway analyses**

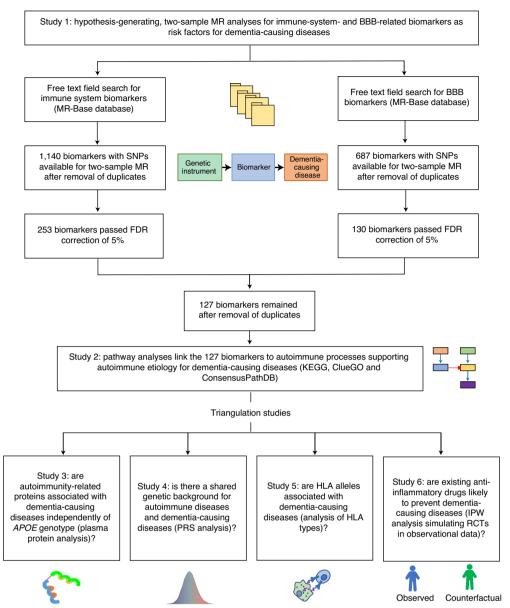
We then used ConsensusPathDB to test whether 42 proteins of the 127 biomarkers that were not bound to any cell and were associated with frontotemporal dementia, Alzheimer's or Parkinson's disease play a role in pathways leading to amyloid precursor protein, tau protein or  $\alpha$ -synuclein that characterize these diseases. These analyses showed that all of the proteins shared a common pathway and were within only zero to two molecules distance from these proteins, providing additional support for the link between proteins and dementia-causing diseases (Extended Data Figs. 6–10).

To identify other biological processes that may be regulated by the 127 biomarkers, we performed analyses based on Kyoto Encyclopedia of Genes and Genomes (KEGG), ClueGO and Consensus Path DB databases using the 78 biomarkers that were plasma proteins or receptors on a cell and thus had an ID applicable for analysis. These analyses suggested that the biomarkers are involved in several processes of autoimmunity, ranging from hematopoiesis to self-tolerance and antigen processing and presentation. These included increased HLA-DR expression (a risk allele for several autoimmune diseases) across all hematopoietic cell lines; MHC-II-mediated antigen presentation (a key mechanism that is dysfunctional in autoimmune diseases) in several processes, including autoimmune diseases and responses to infection; increased neuronal adhesion molecule CNTN2 and increased PD-L1 in T cell-antigen interactions that reduce self-tolerance. Furthermore, the biomarkers also altered expression of several cluster differentiation molecules on leukocytes and decreased barrier-protecting IL-17F, self-tolerance-increasing PDCD1 in T cell-antigen interactions and antiviral complement factor B and IFNAR1 (Supplementary Figs. 1 and 2).

Based on MR and pathway analyses, we hypothesized that diseases causing dementia have an inflammatory autoimmune component.

### PRS and HLA analysis

To further study autoimmunity and the combined effects of the 127 biomarkers, we created an MR-based polygenetic risk score (MR-PRS) using SNPs associated with biomarkers and then performed



**Fig. 1**| **Design and rationale of six complementary studies.** To study BBB- and immune system-related biology, biomarkers and drug targets for dementia-causing diseases, we conducted six separate studies. Study 1 used MR and MR-Base database to explore how BBB and immune system-related biomarkers associate with dementia-causing diseases . This hypothesis-generating study identified 127 biomarkers associated with dementia-causing diseases, many related to BBB, inflammation and self-tolerance, suggesting that inflammatory and autoimmune processes may play a role in these diseases. Study 2 is a pathway analysis on the associations of study 1. Providing additional support for the autoimmune hypothesis, the analysis showed that the biomarkers are enriched in several autoimmune-related biological processes and share pathways with amyloid-β, tau and α-synuclein proteins that characterize dementia-causing diseases. Study 3 examined the eight proteins that have protein quantitative

loci near the APOE gene. In line with the autoimmune hypothesis, this study showed that IFIT2, an anti-inflammatory protein, decreases risk for dementia-causing diseases independent of APOE. Study 4 examined which diseases are associated with a polygenic risk score constructed from SNPs associated with the 127 biomarkers. Using phenome-wide analysis, this study showed that several autoimmune diseases, especially type 1 diabetes and rheumatic arthritis, share a genetic background with dementia-causing diseases. Study 5 provided further support for the autoimmune hypothesis by identifying nine HLA alleles associated with dementia-causing diseases. Study 6 used IPW analyses to simulate randomized control trials in observational data. It examined whether the autoimmune component is modifiable with anti-inflammatory medication. These analyses showed that methotrexate and TNF- $\alpha$  inhibitors may be preventative medications for dementia-causing diseases.

phenome-wide association analysis (PheWAS) (Fig. 5). PRSs were created separately for each dementia-causing disease using only SNPs linked to outcome-specific biomarkers. After linkage disequilibrium (LD) pruning, excluding extreme SNPs with beta >1.34 and SNP matching in FinnGen (n = 339,233), 92 SNPs were available for PRS for Alzheimer's disease. The number of SNPs ( $\leq$ 25) for other outcomes was insufficient for PRS association analysis. In PheWAS analyses, Alzheimer's PRS was associated with increased risk of all types of

dementia-causing diseases and autoimmune diseases, especially rheumatic diseases and type 1 diabetes and its complications, but with reduced risk of cancers. The associations with dementia-causing diseases were largely attributable to three SNPs within 500 kb from the *APOE* region, whereas associations with autoimmune diseases were independent of *APOE* (Supplementary Table 2).

Certain HLA alleles increase risk for autoimmune diseases, including type 1 diabetes and rheumatoid arthritis 18,19. We therefore ran an

### Table 1 | Characteristics of the main cohorts used in the six studies

Cohorts used	Baseline characteristics	Exposure measurement	Outcome measurement
Study 1: MR			
Exposures			
UK Biobank, UK BiLEVE, INTERVAL	173,480 participants; mean age 54 years; 48% men	Blood cells including leukocytes, erythrocytes and platelets	-
SardiNIA dataset, Italy	3,757 participants; mean age 45 years; 43% men	Flow cytometry of detailed leukocyte types	-
INTERVAL study, UK	3,301 participants; mean age 44 years; 51% men	Plasma proteins measured with SomaScan v.3	-
KORA F4 study, Germany	3,080 participants; mean age 56 years; 49% men	Plasma proteins measured with SomaScan v.3	-
Outcomes			
FinnGen, Finland	339,233 participants; mean age 54 years; 44% men	-	Alzheimer's disease and subtypes, Parkinson's disease, frontotemporal dementia, vascular dementia dementia outcomes from national hospital discharge (available from 1968), death (from 1969), cancer (from 1953) and medication reimbursement (from 1964) and purchase (from 1995) registries
IGAP (several cohorts)	54,162 participants; mean age 71 years; 41% men	-	Late-onset Alzheimer's disease from hospital discharge, death, autopsy and medication reimbursement registries
Multicohort study	1,131,881 participants; mean age 61 years; 46% men	-	Cognitive performance measured using immediate word recall task, a delayed word recall task, a naming task and a counting task, Henmon–Nelson test of mental ability, overall GPA, math, science and verbal GPA and educational attainment
ADGC, EADI, CHARGE, GERAD/PERADES consortium	63,926 participants; mean age 73 years; 41% men	-	Late-onset Alzheimer's disease from hospital discharge, death, autopsy and medication reimbursement registries
IPDGC consortium	1,474,097 participants; mean age 57 years; 45% men	-	Parkinson's disease from hospital discharge, death, autopsy and medication reimbursement registries
Study 2: pathway analyses			
KEGG, ConsensusPathDB databases	ConsensusPathDB-human integrates interaction networks in humans including 31 public databases; KEGG pathways are a collection of manually drawn pathway maps of known molecular interactions	Available UniProt IDs for the 127 biomarkers	Interaction path to amyloid precursor protein, tau protein or a-synuclein that characterize Alzheimer's and Parkinson's disease
Study 3: plasma proteomics			
Whitehall II, UK	6,545 participants; mean age 56 years; 71% men	Plasma proteins measured with SomaScan v.4	National Health Services Hospital Episode Statistics database, the British National mortality register and 5-yearly clinical screening
Study 4: PRS			-
FinnGen, Finland	339,233 participants; mean age 54 years; 44% men	Illumina and Affymetrix arrays; AxiomGT1 algorithm for Affymetrix data; imputation with population-specific SISu v.3	National hospital discharge (from 1968), death (from 1969), cancer (from 1953) and medication reimbursement (from 1964) and purchase (from 1995) registries
Study 5: HLA analyses			
FinnGen, Finland	339,233 participants; mean age 54 years; 44% men	rSSO-Luminex technology (Labtype, One Lambda); PCR-SSP (Micro SSP Generic HLA Class I/II DNA Typing Trays, One Lambda; Olerup SSP genotyping; AlleleSEQR PCR/Sequencing kits, Atria Genetics; BI 3130xl genetic analyzer (Applied Biosystems, Thermo Fisher Scientific); Immunochip array (Illumina); imputation HLA*IMP:0240 (The Oxford HLA Imputation Framework)	National hospital discharge (from 1968), death (from 1969), cancer (from 1953) and medication reimbursement (from 1964) and purchase (from 1995) registries
Cohort in study 6: IPW analyses			
FinnGen, Finland	117,773 participants; mean age 55 years; 55% men	ATC codes from medication reimbursement (1997–2019) and purchase (1997–2019) registries Illumina and Affymetrix arrays; AxiomGT1 algorithm for Affymetrix data; imputation with population-specific SISu v.3	National hospital discharge (available from 1968), death (from 1969), cancer (from 1953) and medication reimbursement (from 1964) and purchase (from 1995) registries

For consortium and multicohort studies, mean age and proportion of men are reported for each cohort. ADGC, Alzheimer Disease Genetics Consortium; CHARGE, Cohorts for Heart and Aging Research in Genomic Epidemiology Consortium; EADI, European Alzheimer's Disease Initiative; GERAD/PERADES, Genetic and Environmental Risk in AD/Defining Genetic, Polygenic and Environmental Risk for Alzheimer's Disease Consortium; GPA, grade point average; IGAP, International Genomics of Alzheimer's Project; IPDGC, International Parkinson Disease Genomics Consortium; KORA Kooperative Gesundheitsforschung in der Region Augsburg

HLA allele-wide analyses for dementia-causing diseases to study autoimmunity by an additional method independent of MR and PRS. We identified nine risk HLA alleles for dementia-causing diseases after FDR correction of 5% (P < 0.00085) (Table 2). These analyses supported the autoimmune hypothesis.

### **IPW** analyses

To evaluate the autoimmune hypothesis in relation to modifiability and drug repurposing, we examined whether commonly used anti-inflammatory and immunosuppressive medications are likely to reduce the risk of dementia-causing diseases. For this, we used IPW analyses in the FinnGen study (Table 3 and Supplementary Table 3). As a preliminary step, we validated the IPW protocol with two analyses. The first replicated the RCT effect between statin medication and myocardial infarction (a positive control), and the second replicated the null findings in anti-inflammatory medication trials for cardiovascular diseases (a negative control)<sup>20–23</sup>.

For the main analysis, we selected all anti-inflammatory medication categories if there were data for at least ten individuals who were treated with the medication and developed dementia-causing disease over the follow-up. Supporting the autoimmune hypothesis, these analyses including 117,773 participants showed that use of methotrexate or TNF- $\alpha$  inhibitors was associated with reduced risk of Alzheimer's and Parkinson's disease. Stratifying by APOE status, the risk of Alzheimer's disease was reduced only in individuals with high genetic risk, as indicated by above-median MR-PRS (50% cutoff) or who had at least one  $APOE\ \mathcal{E}4$  allele.

To examine whether the effect was dependent on *APOE* status, we repeated the analysis in participants with above-median MR– PRS or above-median more comprehensive PRS (including 1,092,011 SNPs) derived from the largest available Alzheimer's disease GWAS<sup>24</sup> and excluded the *APOE* area from these PRSs. These analyses showed null results, suggesting that the protective effect of methotrexate is specific for those with at least one *APOE*  $\varepsilon$ 4 allele. The numbers of individuals using TNF- $\alpha$  inhibitors or developing Parkinson's disease were too small for subgroup analyses.

To allow an additional analysis for rare medications, we searched the Open Targets database for medications that modify the levels of the 127 biomarkers. This search identified 64 drugs (mostly monoclonal antibodies) that target 18 of the 127 biomarkers, suggesting that these may also have potential for repurposing in the treatment of dementiacausing diseases (Supplementary Data 7).

### Discussion

Consistent evidence from six independent studies suggests that inflammatory autoimmunity may play a causal role in dementia-causing diseases. Our MR analyses identified causal support for 127 risk factors including inflammatory, self-tolerance and/or BBB tight-junctionrelated biomarkers. Pathway analyses linked these 127 biomarkers to autoimmunity via several alterations in processes, from hematopoiesis to antigen presentation and reduced self-tolerance. They also showed that all 42 circulating proteins associated with frontotemporal dementia, Alzheimer's or Parkinson's diseases among the 127 biomarkers are closely related to α-synuclein, amyloid precursor and tau protein pathways that characterize these diseases. A phenome-wide analysis of our MR-PRS, constructed from SNPs associated with the identified risk factors for Alzheimer's disease, indicated shared genetic background with autoimmune diseases, such as rheumatoid arthritis and type 1 diabetes. The autoimmune hypothesis was further supported by HLA analyses showing nine HLA-type associations with dementias. According to IPW analyses mimicking randomized controlled drug trials in observational data, repurposed use of anti-inflammatory or immunomodulatory medications may reduce the risk of Alzheimer's in individuals with an APOE  $\varepsilon 4$  allele. This finding suggests that the inflammatory autoimmune component may be modifiable with currently available medications.

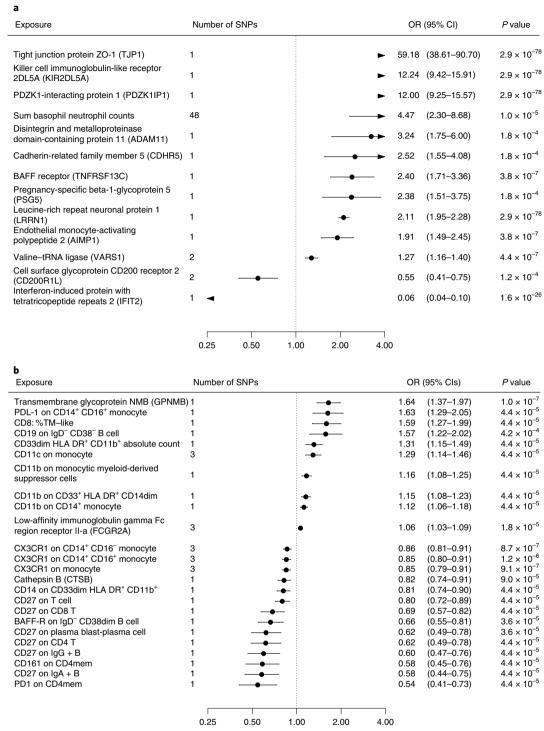
To our knowledge, this is the most comprehensive MR and triangulation study to date on immune system- and BBB-related biomarkers as risk factors of dementia-causing diseases. Our MR focused on 1,827 biomarkers whereas earlier MR analyses included fewer than 200 biomarkers specific to these systems 10,12,25-33. We obtained the strongest causal evidence for autoimmunity- and inflammation-related AZGP1 (ref. 34) and CD33 (ref. 35) for Alzheimer's disease, and for FCGR2A36 and GPNMB<sup>37</sup> in Parkinson's disease. These proteins had cis pQTLs available, and they passed MR sensitivity analyses. For CD33, a monoclonal antibody, gemtuzumab ozogamicin<sup>38</sup> is in routine clinical use, and it may have potential for drug repurposing in Alzheimer's disease. To our knowledge, associations of AZGP1 and FCGR2A with dementias have not been reported previously whereas an earlier MR study on CD33 and GPNMB exists<sup>33</sup>. As an additional supportive finding, our MR sensitivity analyses showed that several CD20- and CD33-expressing leukocytes increase the risk of Alzheimer's disease and that CD11-expressing leukocytes may increase, and CD27- and CXCR1-expressing leukocytes may decrease, Parkinson's disease risk.

In general, our results provide evidence on the role of BBB in the etiology of dementia-causing diseases by suggesting that higher plasma levels of the tight junction component TJP1 (ref.  $^{39}$ ) and proteins degrading the tight junction, such as AIMP1 (ref.  $^{40}$ ) and BIN1 (ref.  $^{41}$ ), increase—and higher levels of barrier-protecting IL-17F  $^{42}$  reduce—the risk of Alzheimer's disease. These findings suggest potential causal risk factors that support the BBB dysfunction and barrier breach hypothesis  $^{9,43}$ , linking BBB breakdown to subsequent inflammatory and autoimmune responses in the CNS. The results are also in line with experimental studies that have linked cerebral vascular dysfunction to cognitive decline, and with evidence linking BBB dysfunction in the hippocampal area with increased risk of Alzheimer's disease independent of amyloid- $\beta$  or tau  $^{44-46}$ .

In agreement with previous research, several proinflammatory biomarkers were associated with increased risk of dementia-causing diseases. IL-1β<sup>47</sup> increased the risk for vascular dementia; C1Q, C1R, CD20 and CDHR5 (refs. <sup>48-50</sup>) the risk for Alzheimer's disease; and GPNMB and CD11b<sup>37,51</sup> the risk for Parkinson's disease<sup>37,51</sup>. Anti-inflammatory biomarkers C4B, IL-27, IL-37, PTP1B and SIGGIRR<sup>52-56</sup> in turn improved cognitive performance. In addition, our MR analyses identified checkpoint regulators BAFFR, C1R, C1Q, CD11, CD19, CD20, CD22, CD33, CD40, CX3CR1, LTBR, PD-1 and PDL-1<sup>51,57-66</sup> as potential causal risk factors for poor cognitive performance and dementiacausing diseases, uncovering the importance of checkpoint control and potential sources of autoreactivity.

Previous MR studies on Alzheimer's disease<sup>10,12,25-33</sup> suggest potential causal associations with BIN1, CCL27, C3, CD33, CD4 T cells, GDF-15 and SVEP1, whereas MR studies on Parkinson's disease suggest a potential causal association with GPNMB, IL-6 and MIP1b. Compared with these studies, we used a stricter *P* value cutoff with multiple testing correction and were able to replicate associations between BIN1, CD33 and Alzheimer's disease and those between GPNMB and Parkinson's disease. Our MR analyses did not replicate the results of other biomarkers. Potential reasons for this discrepancy include the use of different sets of SNPs to test associations between these biomarkers and dementia-causing diseases, and differences in population characteristics.

Pathway analyses provided further understanding of processes that may be regulated by the identified biomarkers. In line with the MR results, these analyses revealed that the biomarkers were enriched in inflammatory and autoimmunity-related biological processes including altered hematopoiesis, cytokine–receptor interaction, responses to infections, self-tolerance, phagosome processing, cell adhesion and antigen presenting, transferring them towards increased autoreactivity. The analyses also showed that the biomarkers were involved in amyloid- $\beta$ , tau protein and  $\alpha$ -synuclein pathways that characterize diseases causing dementia. This converging support for the autoimmune



**Fig. 2** | **Biomarkers associated with general Alzheimer´s and Parkinson´s disease in Mendelian randomization analyses. a,b,** Odds ratios (ORs) and 95% confidence intervals (CIs) for an increase of 1 s.d. in biomarkers associated with general Alzheimer's disease outcome (**a**) and Parkinson's disease (**b**) in MR after

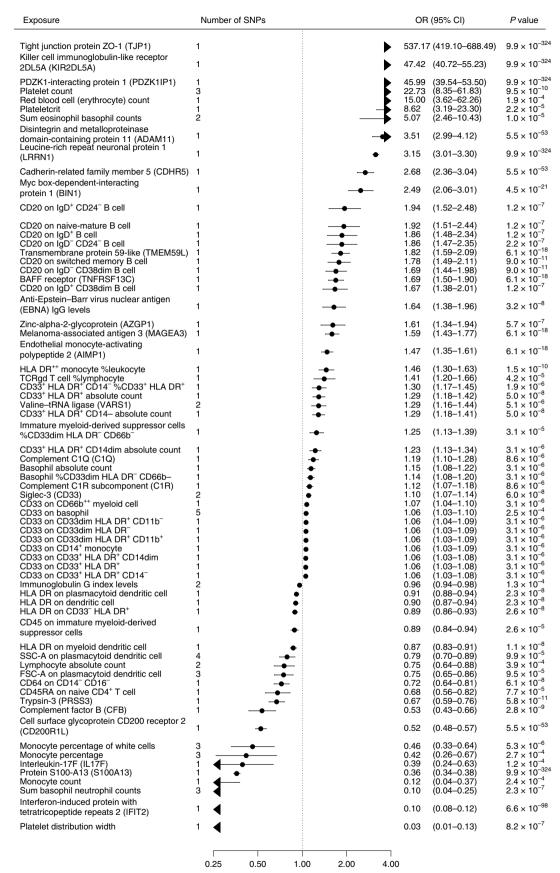
FDR correction of 5% (P< 0.00052). ORs were derived from Wald ratios when only one SNP was available, and from IVW estimates when two or more SNPs were available. All tests are two-sided.

hypothesis from the MR and pathway analyses adds to previous limited human evidence on this hypothesis  $^{43,67}\!.$ 

We obtained additional insights into autoimmune hypothesis from four further studies. Strengthening the supportive evidence, our PheWAS analyses showed that several autoimmune diseases share immune-related genetic background with dementia-causing diseases. In addition, MR-PRS for dementia-causing diseases was associated with

reduced risk of cancers, which is a well-described beneficial side effect of reduced self-tolerance commonly harnessed in immune-oncological cancer medications  $^{68}$ .

The association between MR-PRS and dementia-causing diseases was driven by SNPs in the *APOE* region that are associated with several BBB- and autoimmunity-related proteins such as IFIT2 (ref. <sup>69</sup>), LRRN1 (ref. <sup>70</sup>), TJP1 (ref. <sup>39</sup>), KIR2DL5A<sup>71</sup>, AIMP1 (ref. <sup>72</sup>) and BAFF receptor <sup>57</sup>,



**Fig. 3** | **Biomarkers associated with late-onset Alzheimer´s disease in Mendelian randomization analyses.** ORs and 95% CIs for an increase of 1 s.d. in biomarkers associated with late-onset Alzheimer's disease in MR after FDR

correction of 5% (P < 0.00052). ORs were derived from Wald ratios when only one SNP was available, and from IVW estimates when two or more SNPs were available. All tests are two-sided.

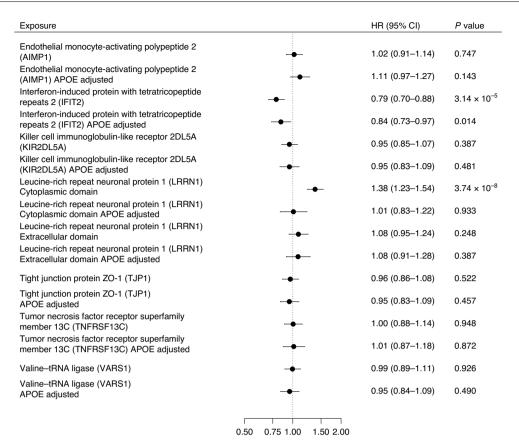


Fig. 4 | Association between plasma proteins that had pQTLs within 500 kb from APOE gene and dementia. Hazard ratios and 95% Cls for association between an increment of 1 s.d. in plasma protein levels and dementia in the Whitehall II cohort. The analyses included eight proteins with pQTLss clustered

around APOE and that were associated with at least three dementia subtypes in MR analyses. Analyses were first adjusted for age and sex and then additionally for APOE status. This analysis was not corrected for multiple testing, and all the tests are two-sided.

suggesting that the autoimmune component in APOE  $\varepsilon 4$  allele carriers could be related to these proteins. The findings on these proteins should be interpreted cautiously, because only LRRN1 and IFIT2 were replicated using plasma proteins and only the protective effect of IFIT2 was independent of the APOE gene. IFIT2 is protective of viral infections and may reduce the load of acute and chronic inflammation in the CNS 3, suggesting that it may be a promising APOE-independent drug target.

The autoimmunity hypothesis was further supported by HLA allele-wide analyses that identified nine HLA types associated with dementia-causing diseases. By identification of the specific risk alleles from HLA classes, these results complement previous studies<sup>74–77</sup> that have identified HLA-DR and HLA-DQ as risk factors for dementia-causing diseases.

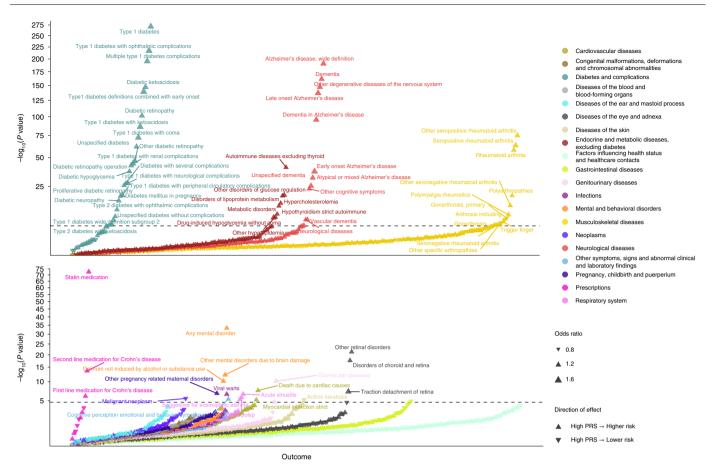
### Anti-inflammatory medications and dementias

We obtained evidence on the modifiability of the autoimmunity-dementia association in analyses simulating RCTs using observational data. The validity of these IPW survival analyses was supported by expected results from the positive and negative control analyses. To ensure sufficient data on key variables (positivity condition) and well-defined interventions and outcomes (consistency condition), we used linkage to electronic health records from high-quality, nation-wide registries of filled drug prescriptions and diseases outcomes. To simulate an RCT, we included only dementia-free individuals with no history of the investigated anti-inflammatory medications at baseline and followed them up over 20 years. Thus, the analyses simulated a two-decade randomized trial focusing on the effect of preventative

anti-inflammatory medication on the risk of dementia-causing diseases. The analyses supported the autoimmune hypothesis and suggested that inflammatory autoimmune processes are modifiable. More specifically, we were able to replicate the associations between the use of TNF- $\alpha$  inhibitors and <sup>78</sup> methotrexate <sup>79</sup> and reduced risk of Alzheimer's disease and to provide evidence of the potential benefits of methotrexate in Parkinson's disease.

As a previously unreported finding, we showed that the protective effect of methotrexate in Alzheimer's disease is observed only in those who carry at least one APOE £4 allele. This finding is in line with experimental studies on the effects of the APOE ε4 allele and methotrexate on the BBB and immune system. Individuals carrying the APOE £4 allele have been shown to have higher rates of BBB and immune system dysfunction<sup>80</sup>, whereas methotrexate targets both these vulnerabilities by protecting the BBB and being anti-inflammatory. These mechanisms are thought to improve endothelial integrity and regulatory T cell differentiation, inhibition of neutrophil adhesion and recruitment, cytokine expression in macrophages, T cell activation, T cell-mediated cell death and metalloproteinase production<sup>81</sup>, all mechanisms highlighted in our MR and pathway findings. These findings suggest that future RCTs on the preventative potential of methotrexate against Alzheimer's disease might be feasible and should include only people with at least one APOE & allele. To date, no RCTs are available for methotrexate as a therapeutic treatment for dementia-causing diseases.

TNF- $\alpha$  inhibitors are safe and well tolerated in patients with Alzheimer's disease dementia (phase 2 trial), but no evidence of benefits is available from studies with 1- to 2-year follow-up <sup>82</sup>. Similarly, RCTs with 1- to 2-year follow-up on corticosteroids, nonsteroidal



**Fig. 5** | **Phenome-wide association analyses for Mendelian randomization-based Alzheimer's diseases risk polygenic risk score.** Phenome-wide association analyses for MR-PRS constructed from SNPs associated with levels of causal Alzheimer's disease biomarkers in MR Wald ratio or IVW analyses. ORs and

 $-\log_{10}P$  values are presented. Upward- and downward-pointing triangles denote increasing and decreasing risk, while larger triangles indicate larger effect size. This analysis was not corrected for multiple testing, and all tests are two-sided.

anti-inflammatory drugs and anti-inflammatory minocycline that have included participants with cognitive decline or Alzheimer's disease have shown no benefit<sup>83-85</sup>. A key strength of our IPW analysis is the tenfold longer follow-up. To reduce tissue damage in autoimmune diseases, early initiation of anti-inflammatory medication is of paramount importance. The protective effect in long follow-up compared with a null effect in short follow-up is in line with studies of autoimmune diseases 19 and suggests that medication in trials for people with established dementia may come too late. In the future, trials on the effectiveness of anti-inflammatory medication in dementiacausing diseases should include high-risk individuals when they are still asymptomatic or present with only early symptoms of the disease. Future studies should also investigate autoimmunity in greater detail to determine the role of central tolerance in major immune organs and the subsequent escape of autoreactive B and T cells to the periphery <sup>59,86</sup> neoepitopes and autoimmune risk-increasing HLA alleles 19,87,88 as well as peripheral self-tolerance mechanisms, such as ignorance, anergy, suppression, inhibition and antigen presentation 59,86. Such studies may identify antigen-specific therapeutic strategies that offer new avenues in the search for treatment for dementia-causing diseases.

### Strengths and limitations

Combining multiple lines of research allowed us to examine the role of immune system and BBB biomarkers in the etiology of dementia-causing diseases and to identify potential new drug targets and opportunities to repurpose existing medications for these diseases. The use of an MR approach across 1,827 biomarkers contributed to the evaluation

Table 2 | ORs and 95% CIs for association between HLA alleles and dementia-causing diseases that survived FDR correction of 5% (P<0.00085). All tests are two-sided

Disease	HLA type	OR (95% CI)	Pvalue	FDR Pvalue
Alzheimer's disease	DQB1 05:01	1.11 (1.06–1.17)	6.7×10 <sup>-5</sup>	0.005
	DRB1 01:01	1.11 (1.05–1.17)	1.5×10 <sup>-4</sup>	0.007
	DQA1 01:01	1.11 (1.05–1.17)	9.9×10 <sup>-5</sup>	0.006
Parkinson's disease	A 03:01	1.13 (1.05–1.22)	8.4×10 <sup>-4</sup>	0.027
Dementia	DQA1 05:01	0.90 (0.85-0.95)	6.4×10 <sup>-5</sup>	0.005
	DQB1 02:01	0.90 (0.85-0.95)	4.9×10 <sup>-5</sup>	0.005
	DRB1 03:01	0.90 (0.85-0.95)	7.1×10 <sup>-5</sup>	0.005
	DRB1 04:01	0.90 (0.85-0.96)	4.9×10 <sup>-4</sup>	0.020
	DRB4 01:03	0.93 (0.90-0.97)	7.9×10 <sup>-4</sup>	0.027

of causality<sup>16</sup>. The findings were summarized with KEGG and ClueGO analyses, both pointing to autoimmune processes. ConsensusPathDB, one of the most comprehensive collections of databases on molecular pathways and interactions<sup>89</sup>, linked the biomarkers to proteinopathies in dementia-causing diseases. Plasma protein analyses allowed us to

Table 3 | Hazard ratios and 95% CIs for associations between dementias, methotrexate and TNF- $\alpha$  inhibitor medications from IPW Cox proportional-hazards survival analyses in the FinnGen study

Medication	Outcome	HR (95% CI)	Pvalue
Statins			
Positive control	Coronary heart disease IPW	0.38 (0.19-0.78)	0.009
	Coronary heart disease RCT	0.39 (0.29-0.49)	7.3×10 <sup>-12</sup>
Methotrexate	Alzheimer's disease (AD)	0.74 (0.59 -0.93)	0.013
	Alzheimer's disease including those with AD-MR-PRS≥50%	0.64 (0.47-0.88)	0.005
	Alzheimer's disease including those with AD-MR-PRS<50%	0.84 (0.59–1.19)	0.330
	Alzheimer's disease including those with AD-MR-PRS ≥50%(APOE region excluded)	0.75 (0.54–1.04)	0.085
	Alzheimer's disease including those with AD-MR-PRS ≥50% (APOE region excluded)	0.73 (0.52–1.02)	0.065
	Alzheimer's disease including those with Jansen's PRS≥50%	0.70 (0.51–0.97)	0.033
	Alzheimer's disease including those with Jansen's PRS < 50%	0.80 (0.57–1.12)	0.197
	Alzheimer's disease including those with Jansen's PRS ≥50% (APOE region excluded) ≥50%	0.74 (0.53–1.04)	0.081
	Alzheimer's disease including those with Jansen's PRS ≥50% (APOE region excluded)	0.75 (0.54–1.06)	0.102
	Alzheimer's disease including those with at least one APOE $\epsilon 4$ allele	0.69 (0.49-0.97)	0.032
	Alzheimer's disease with those carrying APOE ε4 allele excluded	0.79 (0.57–1.10)	0.168
	Vascular dementia	0.64 (0.30–1.37)	0.253
	Parkinson's disease	0.48 (0.29-0.81)	0.006
Negative control	Coronary heart disease IPW	0.97 (0.84–1.11)	0.647
	Coronary heart disease RCT	0.96 (0.79–1.16)	0.862
TNF-a inhibitors	Alzheimer's disease	0.32 (0.14–0.76)	0.010
	Vascular dementia	0.27 (0.04–1.91)	0.188
	Parkinson's disease	0.26 (0.04–1.82)	0.173
Negative control	Coronary heart disease IPW	1.03 (0.55–1.95)	0.916
	Coronary heart disease RCT	1.09 (0.77–1.56)	0.645

Positive-control analyses validate the IPW analysis protocol by replicating the established association between statin medication and reduced coronary heart disease risk. As a further validation step, negative-control IPW analyses replicate the null finding between anti-inflammatory medications and coronary heart disease. In the main IPW analyses, baseline variables were birth year, sex, ten principal components and time-varying variables statin, ACE-blocker, AT-blocker, renin-blocker, calcium channel blocker, any diuretic, insulin, metformin, other diabetes drug, antidepressant, antipsychotic and anticoagulant medication use; as well as time-varying disease diagnosis (any cancer, myocardial infarction, atrial fibrillation, heart failure, venous thromboembolism, ischemic stroke, intracerebral hemorrhage, subarachoid hemorrhage, obesity, sleep apnea or chronic obstructive pulmonary diseases), with informative censoring included. Analyses included only individuals who did not use anti-inflammatory studied at baseline. The only exception was TNF-a inhibitor analyses, where baseline users were included because of the small number of individuals using this medication in the FinnGen cohort. The analyses here were not corrected for multiple testing, and all tests are two-sided. The estimates for RCTs are from refs. <sup>20,22,23</sup>. A secondary prevention RCT was used for methotrexate, because no primary prevention trials were available. MR-PRS, MR-based PRS for Alzheimer's disease.

adjust effect estimates for the *APOE* genotype for certain proteins. PRS and HLA analyses used data from the FinnGen study, with a sample size of 340,000 providing sufficient statistical power. The medication analyses in 120,000 FinnGen participants relied on the IPW method, which is proposed to provide more reliable causal estimates than traditional survival analyses of observational data $^{90}$ .

Our study also has limitations. Rather than having a single dataset with complete information, we used an approach in which separate analyses were performed in separate cohorts. This heterogeneity in study samples and assessment of biomarkers and outcomes is a potential source of inconsistent results, but simultaneously, convergent findings across different studies and methodological approaches support the robustness and generalizability of the results. Although we explored 1,827 biomarkers related to the BBB and immune system, we may have missed some biomarkers due to limited numbers of immune system- and BBB-related biomarkers captured by our free text field searches or lack of SNPs available. In addition, MR provides unconfounded estimates if the genetic variants being used as an instrument

for exposure are associated with that exposure but not with confounding factors, and there is no independent pathway between the genetic variants and the outcome other than through the exposure. While the first assumption was confirmed in the present study, it is not possible to exclude potential violations of the latter two assumptions. Many MR analyses had a limited number of SNPs, which increased the probability of chance findings and did not allow MR sensitivity analyses for some biomarkers. However, for biomarkers with multiple SNPs, only three showed evidence of horizontal pleiotropy. Future studies with access to fine-mapping results from protein GWASs and to in-sample LD data to perform fine-mapping on the summary statistics should further examine causality and drug targets using colocalization analyses.

Ascertainment of dementia was based on linkage to electronic health records. Although this has the advantage of providing data for everyone recruited to the study, it misses participants with milder dementia and is not the gold standard method for assessment of dementia subtypes. Furthermore, because the onset of late-onset dementias in cohort studies is often at an older age than mean age at

death, our results may be subject to collider bias potentially underestimating the role of risk factors that affect longevity, including systemic inflammation, in the development of diseases causing dementia. Our plasma protein analyses were limited by lack of data on dementia subtypes. The PRS and PheWAS analyses were done on samples with European ancestry and may not apply across different ancestries. Due to the limited number of SNPs included in PRS, we may have missed important immune system- and BBB-related associations in our phenome-wide analyses. The IPW analyses on medications may include some bias due to a limited number of individuals in medication subgroups and to limitations in covariate data for simulation of RCTs. However, major bias is unlikely because the analysis protocol was supported by positive- and negative-control analyses.

In summary, this study provides new insights into autoimmunity, BBB and inflammatory dysfunction as contributors to the development of diseases causing dementias. These components are potentially modifiable with medications, suggesting that anti-inflammatory medications and antigen-specific prevention strategies may offer new avenues in the search for treatment for dementias. The present investigation generated new hypotheses on several specific drug targets for dementia-causing diseases, but these need to be validated in future experimental studies. In particular, RCTs assessing the benefits of early autoimmunity-targeted therapies for high-risk individuals are warranted.

### **Methods**

### MR

The SNPs for biomarkers and outcomes were searched from the MR-Base database<sup>91</sup>. Immune system and BBB search terms were identified using identifiers of cell types, receptors, proteins, metabolites and genes. Identifiers were searched from the literature 4,7,9,10,86,88,92-94 and the UniProt database95 using the search terms 'immune' and 'blood brain barrier'. A complete list is available in Supplementary Data 1 and 2. Outcomes were diseases causing dementia, including the following conditions: all types of Alzheimer's disease, Parkinson's disease, vascular dementia, frontotemporal dementia, dementia in general and progression of dementia. Cognitive performance was chosen as an intermediate outcome. Additional SNPs for sensitivity analyses were searched from full summary statistics of three additional plasma-proteome-wide studies<sup>96-98</sup>. Two-sample MR was used to analyze associations between biomarkers and outcomes 16. The first analyses estimated effects using the Wald ratio or IVW analyses<sup>91</sup>. We applied a FDR correction of 5% for the total number of tests conducted within each biomarker class, leading to cutoffs of P < 0.00043 and P < 0.00052 for immune system- and BBB-related biomarkers, respectively. For biomarker-outcome pairs that passed FDR of 5% but shared fewer than three SNPs, we performed sensitivity analyses with backward MR. For biomarker-outcome pairs with three or more shared SNPs, we performed additionally weighted median, weighted mode and MR Egger analyses<sup>16</sup> using the R packages TwoSampleMR and MRInstruments. To assess potential off-target effects for the observed causal biomarkers, we performed phenomewide MR analyses separately for each biomarker using Neale laboratory GWAS summary statistics for 210 UK Biobank endpoints. The phenomewide outcomes in these analyses also included recognized risk factors for dementia-causing diseases 93,99.

In all analyses, we used individuals of European ancestry, a clumping cutoff  $R^2$  of 0.01 and a 500-kb window. LD proxies were searched with a threshold of  $R^2$  = 0.6 and a proxy split size of 500. Biomarkers and outcome alleles were harmonized by inference from positive-strand alleles using allele frequencies for palindromes. For these analyses, we used statistical software R (3.6.0 and 4.1.0). The novelty of MR findings was examined by systematic PubMed search using the following search terms: (Mendelian randomization) AND (dementia OR Alzheim\* OR Parkin\* OR cognitive decline) AND (Entrez gene symbol OR UniProt protein name) without limitations.

### **Pathway analyses**

We used KEGG pathway analysis with Generally Applicable Gene-set Enrichment  $^{100}$  to study the effect of biomarkers on validated pathways. We used MR Wald ratios or IVW betas and P values as input for expression ratios. Gene Ontology term enrichment analyses were done with ClueGO v.2.5.8 (ref.  $^{101}$ ) in Cytoscape v.3.7.2 (ref.  $^{102}$ ). In the hypergeometric test, we used 78 of the 127 biomarkers that were plasma proteins or receptors on a cell and thus had an ID applicable for these analyses as input, and all immune system- and BBB-related proteins from UniProt  $^{95}$  as background and a correction for 5% FDR. The shortest interaction path analyses were done with ConsensusPathDB  $^{89}$ , a web-based analysis tool containing a range of biomedical databases. ConsensusPathDB was used to decipher potential common pathways between biomarkers and amyloid- $\beta$ , tau protein and  $\alpha$ -synuclein.

### Plasma protein analyses

Plasma protein measurements in the Whitehall II study were available for 6,545 individuals of whom 310 developed dementia<sup>7,103-105</sup>. The participants were linked to the National Health Service (NHS) Hospital Episode Statistics (HES) database and the UK national mortality register using individual NHS identification numbers for linkage<sup>103</sup>. The NHS provides almost complete health care coverage for all individuals legally resident in the UK. We defined incident dementia using the WHO International Classification of Diseases, revision 10 (ICD-10) codes F00, F01, F03, G30 and G31 and ICD-9 codes 290.0–290.4, 331.0–331.2, 331.82 and 331.9. We also conducted informant interviews and checked participants' medications at each screening (in 1996–1998, 2011–2013 and 2016–2017) for dementia-related medication. Sensitivity and specificity of dementia assessment based on HES records are 0.78 and 0.92, respectively<sup>104</sup>.

Plasma proteins were measured using SomaScan v.4.0 and v.4.1 assays  $^{7,106,107}$ . Assays were validated against an external reference population, and protein-specific conversion coefficients were used to balance technical differences between versions 4.0 and 4.1. The analyses used plasma samples measured in 1997/1999 and stored in 0.25-ml aliquots at  $-80\,^{\circ}\text{C}$ . Earlier studies have described in detail the performance of the SomaScan assay and the modified aptamer binding  $^{7,105-107}$ . In brief, the assay uses a mix of thousands of slow, off-rate modified aptamers that bind to proteins in participants' plasma samples, where specificity is ensured with a two-step process analogous to a conventional immunoassay. The specificity of aptamer reagents is good and has been confirmed in several ways  $^{7,108,109}$ . Median intra- and interassay coefficients of variation for SomaScan v.4 are -5% and assay sensitivity is comparable to that of typical immunoassays, with a median lower limit of detection in the femtomolar range.

In the Whitehall II study, standard self-administered questionnaires provided data on age and sex. Using DNA extracted from whole blood, a standard PCR assay determined *APOE* genotype using the salting-out method<sup>110,111</sup>. Two blinded independent observers read the genotype, and any discrepancies were resolved by repeating the PCR analysis.

In Whitehall II analyses, we studied the eight proteins associated with all-cause dementia in MR analyses. Dementia subtype data were not available in Whitehall. The distributions of protein values were skewed and therefore transformed to a normal distribution using inverse rank-based normal transformation. The follow-up started at clinical examination in 1997/1999 and ended at onset of dementia, death or 1 October 2019, whichever occurred first. Age, sex and *APOE*-adjusted Cox regression models estimated associations between proteins and diseases causing dementia<sup>112</sup>. The proportionality assumption in Cox models was assessed with Schoenfeld residuals and log–log plots<sup>112</sup>. We used statistical software R (3.6.0 and 4.1.0) for these analyses.

In the Whitehall II study, research ethics approvals were renewed at each wave; the most recent approval was obtained from the University College London Hospital Committee on the Ethics of Human Research (reference no. 85/0938). Written, informed consent from participants was obtained at each contact.

### PRSs and IPW analyses

FinnGen Data Freeze 8 comprises 339,233 individuals and represents approximately 7% of the adult Finnish population. FinnGen is a collection of prospective epidemiological and disease-based cohorts and hospital biobank samples that links genotypes by unique national personal identification numbers to nationwide health registries, including national hospital discharge (available from 1968 onwards), death (1969), cancer (1953) and medication reimbursement (1964) and purchase (1995) registries. The registry-based follow-up ended on 31 December 2020. Alzheimer's disease was defined with ICD-10 codes under F00 and G30, ICD-9 codes under 3310, ICD-8 code under 29010 and medication purchase Anatomical Therapeutic Chemical (ATC) code N06D; vascular dementia with ICD-10 codes under F01 and ICD-9 codes under 4378; and Parkinson's disease with ICD-10 codes under G20, ICD-9 codes under 3320 A, ICD-8 code under 34200 and medication reimbursement code 110.

FinnGen samples were genotyped with Illumina and Affymetrix (Thermo Fisher Scientific) arrays. Genotype calls were made with Gen-Call or zCall (for Illumina) and the AxiomGT1 algorithm (for Affymetrix data). Individuals with ambiguous gender, high genotype missingness (>5%), excess heterozygosity ( $\pm 4$  s.d.) or non-Finnish ancestry were excluded, as well as all variants with high missingness (>2%), low Hardy-Weinberg equilibrium ( $P < 1 \times 10^{-6}$ ) and minor allele count <3. Array data prephasing was carried out with Eagle 2.3.5 (ref. <sup>113</sup>) with the number of conditioning haplotypes set at 20,000. Genotype imputation was done using the population-specific SISu v.3 imputation reference with 3,775 high-coverage (25–30×), whole-genome sequences in Finns, described in detail at https://doi.org/10.17504/protocols.io.xbgfijw.

We constructed PRSs from SNPs associated with the 127 biomarkers identified from MR analyses. To ensure comparability of the SNPs, we used only studies that included participants of European ancestry from the MR-Base database; in this database, data are harmonized. To ensure interoperability, PRSs were designed to be outcome specific by creating a separate PRS for each outcome from the pool of SNPs for biomarkers associated with the outcome of interest. SNPs were LD pruned, with clumping cutoff  $R^2 = 0.01$  and a 500-kb window with the R package TwoSampleMR. The final PRS contained only SNPs available in FinnGen genotypes. Final scores were determined with PLINK v.2.00aLM3, by calculating the SNP biomarker beta-weighted sum of risk alleles for each SNP. PRSs were scaled to zero mean and one-unit variance. Of the outcome-specific PRSs, we analyzed phenome-wide associations across 2.401 disease endpoints for Alzheimer's disease PRS that was the only one with a sufficient number of SNPs available. For validity, we also applied a second genome-wide disease PRS for Alzheimer's disease, generated with PRS-CS<sup>114</sup> (PRS-CS-auto; LD reference panel 1000 G phase 3 Europeanancestry individuals), using a recent large GWAS on Alzheimer's disease by Jansen et al.<sup>24</sup> as input. For sensitivity analyses excluding the APOE region, we calculated Jansen PRS and MR-PRS by excluding SNPs in positions 35,000,000–70,000,000 on chromosome 19 (GRCh38 in PRS and GRCh37 in MR-PRS). The association between PRS and endpoints was studied with logistic regression, adjusting for birth year, sex and the first ten principal components of ancestry.

We used IPW analyses to simulate RCTs on the effect of anti-inflammatory medication on risk of dementias in the observational FinnGen study  $^{90}$ . These analyses included participants aged over 45 years and with no dementia-causing disease at baseline (n = 117,773). ATC codes for anti-inflammatory medication use were searched from medication purchase registry starting from 1995. To ensure powered analyses, we included only medications with at least ten users in participants who were diagnosed with dementia during follow-up. To simulate trial design and to avoid selection and immortal time bias, each analysis included only new medication users. In IPW analyses, we assumed that when medication is initiated, it is continued until the end of follow-up, to simulate intention-to-treat analyses and to provide conservative estimates. The baseline variables in IPW analyses were birth year, sex,

ten principal components of ancestry and the following time-varying variables: statin, ACE-blocker, AT-blocker, renin-blocker, calcium channel blocker, any diuretic, insulin, metformin, other diabetes drug, depression medication, antipsychotic and anticoagulant use, as well as time varying any diagnoses of cancer, myocardial infarction, atrial fibrillation, heart failure, venous thromboembolism, ischemic stroke, intracerebral hemorrhage, subarachnoid hemorrhage, obesity, sleep apnea and chronic obstructive pulmonary disease, with informative censoring included. The positive control analyses on statin medication used these same variables but did not include statin as a time-varying covariate. In IPW analyses, PRSs were categorized into individuals above and below the median (PRS  $\geq$  50% and <50%). IPW analyses used the weighted FinnGen data to estimate the causal effect of each medication compared to no use of the medication studied. For an RCT that did not report P values, these were estimated using a method described by Altman and Bland<sup>115</sup>. APOE alleles in FinnGen were inferred based on genotype (rs7412 with minor allele frequency (MAF) 0.054 in Finns, INFO 0.997; rs429358, MAF 0.18, INFO 0.999). We used R (4.1.2) for these analyses.

### **HLA** analyses

These analyses were done in FinnGen using HLA alleles and imputed with high accuracy using a Finnish-specific reference panel, as previously described in detail  $^{116}$ . After filtering based on an HLA carrier frequency of  $\geq 0.01$  and posterior probability of  $\geq 0.6$ , we assessed the association between HLA alleles and dementias and autoimmune diseases using logistic regression adjusted for birth year, sex and the first ten principal components of ancestry.

Patients and control participants in FinnGen provided informed consent for biobank research, based on the Finnish Biobank Act. Separate research cohorts, with data collected before the Finnish Biobank Act came into effect (in September 2013) and before the start of FinnGen (August 2017), were based on study-specific consents and were later transferred to the Finnish biobanks after ethical approval by Fimea (Finnish Medicines Agency), the National Supervisory Authority for Welfare and Health. Recruitment protocols followed the biobank protocols approved by Fimea. The Coordinating Ethics Committee of the Hospital District of Helsinki and Uusimaa (HUS) statement number for the FinnGen study is HUS/990/2017.

### **Open Targets analyses**

Medications that changed the levels of the 127 biomarkers were searched in the Open Targets database (https://www.opentargets.org/) using UniProt protein names and Entrez gene symbols.

### Statistics and reproducibility

To study BBB- and immune system-related biology, biomarkers and drug targets for dementia-causing diseases, we conducted six separate studies, the designs and data of which are described in Fig. 1 and Table 1. No statistical methods were used to predetermine sample sizes; instead, these were determined based on available data. Study 1 used the freely available MR-Base GWAS catalog and MR to examine associations between a range of biomarkers and dementia-causing diseases. The details are described in Github<sup>117</sup>, and the data used in these analyses are provided in Zenodo<sup>118</sup>. Study 2 examined the pathways regulated by the biomarkers identified in study 1 using publicly avail $able\,KEGG, ClueGO\, and\, Consensus PathDB\, databases, Cytoscape\, and\, Consensus PathDB\, database, Cytoscape\, and\, Cyto$ web-based analysis tools. Study 3 was an observational cohort study to investigate associations between plasma proteins and dementia in the Whitehall II cohort using Cox proportional-hazards models. Plasma proteins were available for 6,545 individuals (71% men) that participated in clinical screening between 1995 and 1997; 310 participants were excluded from the analyses due to missing data. Before the analyses, proteins were inverse rank based, normal transformed due to skewed distributions. None of the proteins violated proportionality

assumptions of the Cox models. The analyses of study 3 are described in Github<sup>117</sup>. Studies 4.5 and 6 were observational cohort studies and used the FinnGen dataset. Studies 4 and 5 used logistic regression to examine the associations of MR-Base polygenic risk score and HLA types with dementia-causing diseases. The participants included all 339,233 individuals (44% men) that were part of FinnGen Data Freeze 8; in studies 4 and 5, 0 and 24,788 participants, respectively, were excluded due to missing data. Study 6 used an IPW Cox proportional-hazards model to simulate RCTs on the effect of anti-inflammatory medications on risk of dementia-causing diseases. These analyses included 117,773 participants (55% men) aged over 45 years not treated with the medication investigated and without dementia-causing diseases at baseline. None of these analyses violated the assumptions of the IPW Cox proportional-hazards model. The analyses of studies 4. 5 and 6 are described in Github<sup>117</sup>. The R package versions used were data.table 1.14.2, dplyr 1.0.7, tidyr 1.1.4, survival 3.2.13, survminer 0.4.9, ggplot2 3.3.5 plyr 1.8.6, cluster 2.1.2, lubridate 1.8.0, stats 4.1.1, readxl 1.3.1, scales 1.1.1, tidyverse 1.3.1, Hmisc 4.6.0, devtools 2.4.2, TwoSampleMR 0.5.6, MRInstruments 0.3.2, ipw 1.0.11 and metafor 3.0.2. Other software used included ClueGO v.2.5.8 Cytoscape v.3.7.2, Cromwell 61, PLINK v.2.00aLM3, BCFtools 1.7 and 1.9, Eagle 2.3.5 and Beagle 4.1 (08Jun17.d8b).

### **Ethics statement**

In the Whitehall II study, research ethics approvals were renewed at each wave; the most recent approval was obtained from the University College London Hospital Committee on the Ethics of Human Research (reference no. 85/0938). Written, informed consent from participants was obtained at each contact. Patients and control subjects in FinnGen provided informed consent for biobank research, based on the Finnish Biobank Act. Alternatively, separate research cohorts, collected before the Finnish Biobank Act came into effect (in September 2013) and start of FinnGen (August 2017), were collected based on study-specific consents and later transferred to the Finnish biobanks after approval by Fimea, the National Supervisory Authority for Welfare and Health. Recruitment protocols followed the biobank protocols approved by Fimea. The Coordinating Ethics Committee of HUS statement number for the FinnGen study is HUS/990/2017. The FinnGen study is approved by Finnish Institute for Health and Welfare (permit nos. THL/2031/6.02.00/2017, THL/1101/5.05.00/2017, THL/341/6.02.00/2018. THL/2222/6.02.00/2018. THL/283/6.02.00/ 2019. THL/1721/5.05.00/2019 and THL/1524/5.05.00/2020), the Digital and population data service agency (permit nos. VRK43431/2017-3, VRK/6909/2018-3 and VRK/4415/2019-3), the Social Insurance Institution (permit nos. KELA 58/522/2017, KELA 131/522/2018, KELA 70/522/2019, KELA 98/522/2019, KELA 134/522/2019, KELA 138/ 522/2019, KELA 2/522/2020 and KELA 16/522/2020), Findata permit nos. THL/2364/14.02/2020, THL/4055/14.06.00/2020, THL/3433/14. 06.00/2020, THL/4432/14.06/2020, THL/5189/14.06/2020, THL/ 5894/14.06.00/2020, THL/6619/14.06.00/2020, THL/209/14.06. 00/2021, THL/688/14.06.00/2021, THL/1284/14.06.00/2021, THL/ 1965/14.06.00/2021, THL/5546/14.02.00/2020, THL/2658/14.06. 00/2021 and THL/4235/14.06.00/2021 and Statistics Finland (permit nos. TK-53-1041-17, TK/143/07.03.00/2020 (previously TK-53-90-20) and TK/1735/07.03.00/2021). The Biobank Access Decisions for FinnGen samples and data utilized in FinnGen Data Freeze 8 include THL Biobank BB2017\_55, BB2017\_111, BB2018\_19, BB\_2018\_34, BB\_2018\_67, BB2018\_71, BB2019\_7, BB2019\_8, BB2019\_26 and BB2020\_1, Finnish Red Cross Blood Service Biobank 7.12.2017, Helsinki Biobank HUS/359/2017, Auria Biobank AB17-5154 and amendment no. 1 (17 August 2020) and AB20-5926 and amendment no. 1 (23 April 2020), Biobank Borealis of Northern Finland 2017 1013, Biobank of Eastern Finland 1186/2018 and amendment 22 § /2020, Finnish Clinical Biobank Tampere MH0004 and amendments (21.02.2020 and 06.10.2020), Central Finland Biobank 1-2017 and Terveystalo Biobank STB 2018001.

### **Reporting summary**

Further information on research design is available in the Nature Research Reporting Summary linked to this article.

### **Data availability**

This study used publicly available data at https://www.mrbase.org/, https://www.uniprot.org/, http://cpdb.molgen.mpg.de/, https://www. genome.jp/kegg/and https://www.opentargets.org/. Data used in MR are deposited with Zenodo at https://zenodo.org/deposit/7042008. Data, protocols and other metadata of the Whitehall II and FinnGen studies are available according to the data-sharing policies of these studies. The pre-existing data access policy for the Whitehall II study specifies that research data requests can be submitted to the study steering committee, and these will be promptly reviewed for confidentiality or intellectual property restrictions and will not unreasonably be refused. Detailed information on data sharing can be found at https://www.ucl. ac.uk/epidemiology-health-care/research/epidemiology-and-publichealth/research/whitehall-ii/data-sharing Individual-level patient or protein data may further be restricted by consent, confidentiality or privacy laws/considerations. FinnGen data can be accessed through Finnish Biobanks' FinBB portal (www.finbb.fi). FinnGen summary statistics are freely available at https://www.finngen.fi/en/access\_results, with results for new data freezes updated every 6 months.

### **Code availability**

Code used in this manuscript can be found at Github<sup>117</sup>: https://github.com/JVLind/Dementias and autoimmunity.

### References

- Knopman, D. S. Lowering of amyloid-beta by β-secretase inhibitors – some informative failures. N. Engl. J. Med. 380, 1476–1478 (2019).
- Bettcher, B. M., Tansey, M. G., Dorothée, G. & Heneka, M. T. Peripheral and central immune system crosstalk in Alzheimer disease – a research prospectus. *Nat. Rev. Neurol.* 17, 689–701 (2021).
- Kinney, J. W. et al. Inflammation as a central mechanism in Alzheimer's disease. Alzheimers Dement. (N Y) 4, 575–590 (2018).
- Tan, E. K. et al. Parkinson disease and the immune system associations, mechanisms and therapeutics. *Nat. Rev. Neurol.* 16, 303–318 (2020).
- Barisano, G. et al. Blood-brain barrier link to human cognitive impairment and Alzheimer's Disease. Nat. Cardiovasc. Res. 1, 108–115 (2022).
- Franceschi, C., Garagnani, P., Parini, P., Giuliani, C. & Santoro, A. Inflammaging: a new immune-metabolic viewpoint for agerelated diseases. *Nat. Rev. Endocrinol.* 14, 576–590 (2018).
- Lindbohm, J.V. et al. Plasma proteins, cognitive decline, and 20year risk of dementia in the Whitehall II and Atherosclerosis Risk in Communities studies. Alzheimers Dement. (N Y) 18, 612–624 (2022).
- 8. Sipilä, P. N. et al. Hospital-treated infectious diseases and the risk of dementia: a large, multicohort, observational study with a replication cohort. *Lancet Infect. Dis.* **21**, 1557–1567 (2021).
- Sweeney, M. D., Zhao, Z., Montagne, A., Nelson, A. R. & Zlokovic,
   V. Blood-brain barrier: from physiology to disease and back.
   Physiol. Rev. 99, 21–78 (2019).
- 10. Walker, K. A. et al. Large-scale plasma proteomic analysis identifies proteins and pathways associated with dementia risk. *Nat. Aging* **1**, 473–489 (2021).
- Engelhart, M. J. et al. Inflammatory proteins in plasma and the risk of dementia: the Rotterdam Study. Arch. Neurol. 61, 668–672 (2004).
- Fani, L. et al. Circulating biomarkers of immunity and inflammation, risk of Alzheimer's disease, and hippocampal volume: a Mendelian randomization study. *Transl. Psychiatry* 11, 291 (2021).

- Koyama, A. et al. The role of peripheral inflammatory markers in dementia and Alzheimer's disease: a meta-analysis. J. Gerontol. A Biol. Sci. Med. Sci. 68, 433–440 (2013).
- Schmidt, R. et al. Early inflammation and dementia: a 25-year follow-up of the Honolulu-Asia Aging Study. Ann. Neurol. 52, 168–174 (2002).
- Tan, Z. S. et al. Inflammatory markers and the risk of Alzheimer disease: the Framingham Study. Neurology 68, 1902–1908 (2007).
- Holmes, M. V., Richardson, T. G., Ference, B. A., Davies, N. M. & Davey Smith, G. Integrating genomics with biomarkers and therapeutic targets to invigorate cardiovascular drug development. *Nat. Rev. Cardiol.* 18, 435–453 (2021).
- Munafo, M. R. & Davey Smith, G. Robust research needs many lines of evidence. *Nature* 553, 399–401 (2018).
- Roep, B. O., Thomaidou, S., van Tienhoven, R. & Zaldumbide, A. Type 1 diabetes mellitus as a disease of the β-cell (do not blame the immune system?). Nat. Rev. Endocrinol. 17, 150–161 (2021).
- Smolen, J. S. et al. Rheumatoid arthritis. Nat. Rev. Dis. Primers 4, 18001 (2018).
- Law, M. R., Wald, N. J. & Rudnicka, A. R. Quantifying effect of statins on low density lipoprotein cholesterol, ischaemic heart disease, and stroke: systematic review and meta-analysis. *Br. Med. J.* 326, 1423 (2003).
- Bally, M. et al. Risk of acute myocardial infarction with NSAIDs in real world use: Bayesian meta-analysis of individual patient data. Br. Med. J. 357, j1909 (2017).
- Thanigaimani, S., Phie, J., Krishna, S. M., Moxon, J. & Golledge, J. Effect of disease modifying anti-rheumatic drugs on major cardiovascular events: a meta-analysis of randomized controlled trials. Sci. Rep. 11, 6627 (2021).
- Ridker, P. M. et al. Low-dose methotrexate for the prevention of atherosclerotic events. N. Engl. J. Med. 380, 752–762 (2018).
- Jansen, I. E. et al. Genome-wide meta-analysis identifies new loci and functional pathways influencing Alzheimer's disease risk. Nat. Genet. 51, 404–413 (2019).
- Bottigliengo, D. et al. A Mendelian randomization study investigating the causal role of inflammation on Parkinson's disease. Brain https://doi.org/10.1093/brain/awac193 (2022).
- Pagoni, P. et al. Causal effects of circulating cytokine concentrations on risk of Alzheimer's disease and cognitive function. *Brain Behav. Immun.* 104, 54–64 (2022).
- Park, Y. H. et al. Association of blood-based transcriptional risk scores with biomarkers for Alzheimer disease. *Neurol. Genet.* 6, e517 (2020).
- Storm, C. S. et al. Finding genetically-supported drug targets for Parkinson's disease using Mendelian randomization of the druggable genome. *Nat. Commun.* 12, 7342 (2021).
- Wu, P. F. et al. Growth differentiation factor 15 is associated with Alzheimer's disease risk. Front. Genet. 12, 700371 (2021).
- Yeung, C. H. C. & Schooling, C. M. Systemic inflammatory regulators and risk of Alzheimer's disease: a bidirectional Mendelianrandomization study. *Int. J. Epidemiol.* 50, 829–840 (2021).
- Zhao, Y. et al. Genetically predicted levels of circulating inflammatory cytokines and the risk and age at onset of Parkinson's disease: a two-sample Mendelian randomization study. Front. Aging. Neurosci. 14, 811059 (2022).
- 32. Zhu, J., Liu, X., Yin, H., Gao, Y. & Yu, H. Convergent lines of evidence support BIN1 as a risk gene of Alzheimer's disease. *Hum. Genomics* **15**, 9 (2021).
- Png, G. et al. Mapping the serum proteome to neurological diseases using whole genome sequencing. Nat. Commun. 12, 7042 (2021)
- 34. Na, H. S. et al. Th17 and IL-17 cause acceleration of inflammation and fat loss by inducing α(2)-glycoprotein 1 (AZGP1) in rheumatoid arthritis with high-fat Diet. *Am. J. Pathol.* **187**, 1049–1058 (2017).

- 35. Zhao, L. CD33 in Alzheimer's disease biology, pathogenesis, and therapeutics: a mini-review. *Gerontology* **65**, 323–331 (2019).
- 36. Dahlqvist, J. et al. Systematic identification of genomic elements that regulate FCGR2A expression and harbor variants linked with autoimmune disease. *Hum. Mol. Genet.* **31**, 1946–1961 (2022).
- Saade, M., Araujo de Souza, G., Scavone, C. & Kinoshita, P. F. The role of GPNMB in inflammation. Front. Immunol. 12, 674739 (2021).
- 38. Castaigne, S. et al. Effect of gemtuzumab ozogamicin on survival of adult patients with de-novo acute myeloid leukaemia (ALFA-0701): a randomised, open-label, phase 3 study. *Lancet* **379**, 1508–1516 (2012).
- 39. Furuse, M. et al. Direct association of occludin with ZO-1 and its possible involvement in the localization of occludin at tight junctions. *J. Cell Biol.* **127**, 1617–1626 (1994).
- Li, C., Ma, W., Zhao, Y. & Wang, H. Changes in the expression of endothelial monocyte-activating polypeptide II in the rat hippocampus following status epilepticus. *Int. J. Mol. Med.* 47, 699–707 (2021).
- Thomas, S. et al. Novel colitis immunotherapy targets Bin1 and improves colon cell barrier function. *Dig. Dis. Sci.* 61, 423–432 (2016).
- 42. McGeachy, M. J., Cua, D. J. & Gaffen, S. L. The IL-17 family of cytokines in health and disease. *Immunity* **50**, 892–906 (2019).
- 43. Arshavsky, Y. I. Alzheimer's disease: from amyloid to autoimmune hypothesis. *Neuroscientist* **26**, 455–470 (2020).
- 44. Nation, D. A. et al. Blood-brain barrier breakdown is an early biomarker of human cognitive dysfunction. *Nat. Med.* **25**, 270–276 (2019).
- 45. Montagne, A. et al. Blood-brain barrier breakdown in the aging human hippocampus. *Neuron* **85**, 296–302 (2015).
- 46. Montagne, A. et al. APOE4 leads to blood-brain barrier dysfunction predicting cognitive decline. *Nature* **581**, 71–76 (2020).
- Migliorini, P., Italiani, P., Pratesi, F., Puxeddu, I. & Boraschi, D. The IL-1 family cytokines and receptors in autoimmune diseases. Autoimmun. Rev. 19, 102617 (2020).
- 48. Carmona, F. D. et al. Novel identification of the IRF7 region as an anticentromere autoantibody propensity locus in systemic sclerosis. *Ann. Rheum. Dis.* **71**, 114–119 (2012).
- Dalakas, M. C., Alexopoulos, H. & Spaeth, P. J. Complement in neurological disorders and emerging complement-targeted therapeutics. *Nat. Rev. Neurol.* 16, 601–617 (2020).
- 50. Meffre, E. & O'Connor, K. C. Impaired B-cell tolerance checkpoints promote the development of autoimmune diseases and pathogenic autoantibodies. *Immunol. Rev.* **292**, 90–101 (2019).
- Nowatzky, J., Manches, O., Khan, S. A., Godefroy, E. & Bhardwaj, N. Modulation of human Th17 cell responses through complement receptor 3 (CD11/b/CD18) ligation on monocyte-derived dendritic cells. J. Autoimmun. 92, 57–66 (2018).
- 52. Berdnikovs, S. et al. PTP1B deficiency exacerbates inflammation and accelerates leukocyte trafficking in vivo. *J. Immunol.* **188**, 874–884 (2012).
- Kamitaki, N. et al. Complement genes contribute sex-biased vulnerability in diverse disorders. Nature 582, 577–581 (2020).
- Meka, R. R., Venkatesha, S. H., Dudics, S., Acharya, B. & Moudgil, K. D. IL-27-induced modulation of autoimmunity and its therapeutic potential. *Autoimmun. Rev.* 14, 1131–1141 (2015).
- 55. Xu, W. D., Zhao, Y. & Liu, Y. Insights into IL-37, the role in autoimmune diseases. *Autoimmun. Rev.* 14, 1170–1175 (2015).
- Cavalli, G. & Dinarello, C. A. Suppression of inflammation and acquired immunity by IL-37. *Immunol. Rev.* 281, 179–190 (2018).
- 57. Brink, R. Regulation of B cell self-tolerance by BAFF. Semin. Immunol. **18**, 276–283 (2006).
- 58. Choi, J. W. et al. The prognostic significance of VISTA and CD33-positive myeloid cells in cutaneous melanoma and their relationship with PD-1 expression. *Sci. Rep.* **10**, 14372 (2020).

- von Boehmer, H. & Melchers, F. Checkpoints in lymphocyte development and autoimmune disease. *Nat. Immunol.* 11, 14–20 (2010).
- Wang, Q. et al. Transient BAFF blockade inhibits type 1 diabetes development in nonobese diabetic mice by enriching immunoregulatory B lymphocytes sensitive to deletion by anti-CD20 cotherapy. J. Immunol. 199, 3757–3770 (2017).
- Yamauchi, T. et al. T-cell CX3CR1 expression as a dynamic bloodbased biomarker of response to immune checkpoint inhibitors. Nat. Commun. 12, 1402 (2021).
- 62. Dörner, T., Shock, A. & Smith, K. G. CD22 and autoimmune disease. Int. Rev. Immunol. 31, 363–378 (2012).
- Gommerman, J. L. & Summers deLuca, L. LTβR and CD40: working together in dendritic cells to optimize immune responses. *Immunol. Rev.* 244, 85–98 (2011).
- Maldini, C. R., Ellis, G. I. & Riley, J. L. CAR T cells for infection, autoimmunity and allotransplantation. *Nat. Rev. Immunol.* 18, 605–616 (2018).
- Schittenhelm, L., Hilkens, C. M. & Morrison, V. L. β(2) Integrins as regulators of dendritic cell, monocyte, and macrophage function. Front. Immunol. 8, 1866 (2017).
- Wang, J. & Fu, Y. X. The role of LIGHT in T cell-mediated immunity. Immunol. Res. 30, 201–214 (2004).
- D'Andrea, M. R. Add Alzheimer's disease to the list of autoimmune diseases. Med. Hypotheses 64, 458–463 (2005).
- Leach, D. R., Krummel, M. F. & Allison, J. P. Enhancement of antitumor immunity by CTLA-4 blockade. Science 271, 1734–1736 (1996).
- Das Sarma, J. et al. Ifit2 deficiency restricts microglial activation and leukocyte migration following murine coronavirus (m-CoV) CNS infection. PLoS Pathog. 16, e1009034 (2020).
- Taguchi, A. et al. Molecular cloning of novel leucine-rich repeat proteins and their expression in the developing mouse nervous system. *Brain Res. Mol. Brain Res.* 35, 31–40 (1996).
- Estefanía, E. et al. Human KIR2DL5 is an inhibitory receptor expressed on the surface of NK and T lymphocyte subsets. J. Immunol. 178, 4402–4410 (2007).
- Adly, A. A. M., Ismail, E. A., Tawfik, L. M., Ebeid, F. S. E. & Hassan, A. A. S. Endothelial monocyte activating polypeptide II in children and adolescents with type 1 diabetes mellitus: relation to microvascular complications. Cytokine 76, 156–162 (2015).
- 73. Prüss, H. Autoantibodies in neurological disease. *Nat. Rev. Immunol.* **21**, 798–813 (2021).
- Zhang, X. et al. Regulation of the late onset Alzheimer's disease associated HLA-DQA1/DRB1 expression. Am. J. Alzheimers Dis. Other Demen. 37, 15333175221085066 (2022).
- Yan, Y. et al. Genetic association of FERMT2, HLA-DRB1, CD2AP, and PTK2B polymorphisms with Alzheimer's disease risk in the Southern Chinese population. Front. Aging Neurosci. 12, 16 (2020).
- Wang, Z. X. et al. Effects of HLA-DRB1/DQB1 genetic variants on neuroimaging in healthy, mild cognitive impairment, and Alzheimer's disease cohorts. Mol. Neurobiol. 54, 3181–3188 (2017).
- Mäkelä, M. et al. Alzheimer risk loci and associated neuropathology in a population-based study (Vantaa 85+). Neurol. Genet. 4, e211 (2018).
- Torres-Acosta, N., O'Keefe, J. H., O'Keefe, E. L., Isaacson, R. & Small, G. Therapeutic potential of TNF-α inhibition for Alzheimer's disease prevention. J. Alzheimers Dis. 78, 619–626 (2020).
- Judge, A. et al. Protective effect of antirheumatic drugs on dementia in rheumatoid arthritis patients. *Alzheimers Dement*. (N Y) 3, 612–621 (2017).
- 80. Yamazaki, Y., Zhao, N., Caulfield, T. R., Liu, C. C. & Bu, G. Apolipoprotein E and Alzheimer disease: pathobiology and targeting strategies. *Nat. Rev. Neurol.* **15**, 501–518 (2019).

- 81. Cronstein, B. N. & Aune, T. M. Methotrexate and its mechanisms of action in inflammatory arthritis. *Nat. Rev. Rheumatol.* **16**, 145–154 (2020).
- 82. Butchart, J. et al. Etanercept in Alzheimer disease: a randomized, placebo-controlled, double-blind, phase 2 trial. *Neurology* **84**, 2161–2168 (2015).
- 83. Aisen, P. S. et al. A randomized controlled trial of prednisone in Alzheimer's disease. Alzheimer's Disease Cooperative Study. *Neurology* **54**, 588–593 (2000).
- 84. Aisen, P. S. et al. Effects of rofecoxib or naproxen vs placebo on Alzheimer disease progression: a randomized controlled trial. *JAMA* 289, 2819–2826 (2003).
- 85. Howard, R. et al. Minocycline at 2 different dosages vs placebo for patients with mild Alzheimer disease: a randomized clinical trial. *JAMA Neurol.* **77**, 164–174 (2020).
- 86. Theofilopoulos, A. N., Kono, D. H. & Baccala, R. The multiple pathways to autoimmunity. *Nat. Immunol.* **18**, 716–724 (2017).
- 87. Buzzetti, R., Zampetti, S. & Maddaloni, E. Adult-onset autoimmune diabetes: current knowledge and implications for management. *Nat. Rev. Endocrinol.* **13**, 674–686 (2017).
- 88. Dendrou, C. A., Petersen, J., Rossjohn, J. & Fugger, L. HLA variation and disease. *Nat. Rev. Immunol.* **18**, 325–339 (2018).
- Herwig, R., Hardt, C., Lienhard, M. & Kamburov, A. Analyzing and interpreting genome data at the network level with ConsensusPathDB. *Nat. Protoc.* 11, 1889–1907 (2016).
- Hernán, M. A. & Robins, J. M. Using big data to emulate a target trial when a randomized trial is not available. *Am. J. Epidemiol*. 183, 758–764 (2016).
- 91. Hemani, G. et al. The MR-Base platform supports systematic causal inference across the human phenome. eLife **7**, e34408 (2018).
- 92. Akiyama, H. et al. Inflammation and Alzheimer's disease. *Neurobiol. Aging* **21**, 383–421 (2000).
- 93. Livingston, G. et al. Dementia prevention, intervention, and care: 2020 report of the Lancet Commission. *Lancet (London, England)* **396**, 413–446 (2020).
- O'Brien, J. T. & Thomas, A. Vascular dementia. *Lancet* 386, 1698–1706 (2015).
- 95. UniProt, C. UniProt: a worldwide hub of protein knowledge. Nucleic Acids Res. 47. D506–D515 (2019).
- 96. Folkersen, L. et al. Genomic and drug target evaluation of 90 cardiovascular proteins in 30,931 individuals. *Nat. Metab.* **2**, 1135–1148 (2020).
- 97. Pietzner, M. et al. Mapping the proteo-genomic convergence of human diseases. Science **374**, eabj1541 (2021).
- 98. Sun, B. B. et al. Genomic atlas of the human plasma proteome. *Nature* **558**, 73–79 (2018).
- BioBank UK. GWAS results.http://www.nealelab.is/uk-biobank/ (2018).
- 100. Luo, W., Friedman, M. S., Shedden, K., Hankenson, K. D. & Woolf, P. J. GAGE: generally applicable gene set enrichment for pathway analysis. *BMC Bioinformatics* **10**, 161 (2009).
- 101. Bindea, G. et al. ClueGO: a Cytoscape plug-in to decipher functionally grouped gene ontology and pathway annotation networks. *Bioinformatics* **25**, 1091–1093 (2009).
- 102. Shannon, P. et al. Cytoscape: a software environment for integrated models of biomolecular interaction networks. *Genome Res.* **13**, 2498–2504 (2003).
- 103. Kivimaki, M. et al. Validity of cardiovascular disease event ascertainment using linkage to UK hospital records. *Epidemiology* **28**, 735–739 (2017).
- 104. Sommerlad, A. et al. Accuracy of general hospital dementia diagnoses in England: sensitivity, specificity, and predictors of diagnostic accuracy 2008-2016. *Alzheimers Dement.* **14**, 933–943 (2018).

- 105. Kivimäki, M. et al. Cognitive stimulation in the workplace, plasma proteins, and risk of dementia: three analyses of population cohort studies. BMJ 374. n1804 (2021).
- 106. Gold, L. et al. Aptamer-based multiplexed proteomic technology for biomarker discovery. *PLoS ONE* **5**, e15004 (2010).
- Williams, S. A. et al. Plasma protein patterns as comprehensive indicators of health. *Nat. Med.* 25, 1851–1857 (2019).
- 108. Emilsson, V. et al. Co-regulatory networks of human serum proteins link genetics to disease. *Science* **361**, 769–773 (2018).
- 109. Tin, A. et al. Reproducibility and variability of protein analytes measured using a multiplexed modified aptamer assay. J. Appl. Lab. Med. 4, 30–39 (2019).
- Bolla, M. K., Wood, N. & Humphries, S. E. Rapid determination of apolipoprotein E genotype using a heteroduplex generator. J. Lipid Res. 40, 2340–2345 (1999).
- Miller, S. A., Dykes, D. D. & Polesky, H. F. A simple salting out procedure for extracting DNA from human nucleated cells. *Nucleic Acids Res.* 16, 1215 (1988).
- 112. Cox, D. R. Regression models and life-tables. J. R. Stat. Soc. Series B Stat. Methodol. **34**, 187–220 (1972).
- Loh, P. R. et al. Reference-based phasing using the Haplotype Reference Consortium panel. *Nat. Genet.* 48, 1443–1448 (2016).
- 114. Ge, T., Chen, C.-Y., Ni, Y., Feng, Y.-C. A. & Smoller, J. W. Polygenic prediction via Bayesian regression and continuous shrinkage priors. *Nat. Commun.* 10, 1776 (2019).
- 115. Altman, D. G. & Bland, J. M. How to obtain the P value from a confidence interval. *Br. Med. J.* **343**, d2304 (2011).
- 116. Ritari, J., Hyvärinen, K., Clancy, J., Partanen, J. & Koskela, S. Increasing accuracy of HLA imputation by a population-specific reference panel in a FinnGen biobank cohort. NAR Genom. Bioinform. 2, Iqaa030 (2020).
- 117. Lindbohm, J. V. et al. Analysis codes for article: immune system wide Mendelian randomization and triangulation analyses support autoimmunity as a modifiable component in dementia causing diseases. *Nat. Aging* (2022). https://github.com/JVLind/Dementias\_and\_autoimmunity
- 118. Lindbohm, J. V. et al. Supplementary data file for article: immune system wide Mendelian randomization and triangulation analyses support autoimmunity as a modifiable component in dementia causing diseases. *Nat. Aging* (2022). https://zenodo.org/ deposit/7042008.

### **Acknowledgements**

We thank the participants and investigators of the FinnGen and Whitehall II studies. This study was supported by the Wellcome Trust (no. 221854/Z/20/Z) and the UK Medical Research Council (no. S011676). For the purpose of open access, the author has applied a CC BY public copyright license to any author-accepted manuscript version arising from this submission. The FinnGen project is funded by two grants from Business Finland (nos. HUS 4685/31/2016 and UH 4386/31/2016) and by the following industry partners: AbbVie Inc., AstraZeneca UK Ltd, Biogen MA Inc., Bristol Myers Squibb (and Celgene Corporation & Celgene International II Sàrl), Genentech Inc., Merck Sharp & Dohme Corp., Pfizer Inc., GlaxoSmithKline Intellectual Property Development Ltd., Sanofi US Services Inc., Maze Therapeutics Inc., Janssen Biotech Inc., Novartis AG and Boehringer Ingelheim. The following biobanks are acknowledged for delivering biobank samples to FinnGen: Auria Biobank (www. auria.fi/biopankki), THL Biobank (www.thl.fi/biobank), Helsinki Biobank (www.helsinginbiopankki.fi), Biobank Borealis of Northern Finland (https://www.ppshp.fi/Tutkimus-ja-opetus/Biopankki/ Pages/Biobank-Borealis-briefly-in-English.aspx), Finnish Clinical Biobank Tampere (www.tays.fi/en-US/Research\_and\_development/ Finnish Clinical Biobank Tampere), Biobank of Eastern Finland

(www.ita-suomenbiopankki.fi/en), Central Finland Biobank (www. ksshp.fi/fi-FI/Potilaalle/Biopankki), Finnish Red Cross Blood Service Biobank (www.veripalvelu.fi/verenluovutus/biopankkitoiminta) and Terveystalo Biobank (www.terveystalo.com/fi/Yritystietoa/Terveystalo-Biopankki/Biopankki/). All Finnish Biobanks are members of BBMRI. fi infrastructure (www.bbmri.fi). Finnish Biobank Cooperative (FINBB; https://finbb.fi/) is the coordinator of BBMRI-ERIC operations in Finland. Finnish biobank data can be accessed through Fingenious services (https://site.fingenious.fi/en/) managed by FINBB.

The Whitehall II study was supported by the Wellcome Trust (no. 221854/Z/20/Z), the UK Medical Research Council (no. R024227), the National Institute on Aging (National Institutes of Health, nos. R01AG056477 and RF1AG062553), the British Heart Foundation (no. RG/16/11/32334) and SomaLogic, Inc. In the Whitehall II study. some proteins were measured as an academic-industry partnership project beween UCL and SomaLogic, Inc., which provided expertise in plasma proteins and funded 2,240 SOMAscan assays. In this study, SomaLogic had no role in study design, data collection and analysis, decision to publish or preparation of the manuscript. J.V.L. was supported by the Academy of Finland (no. 339568) and the Päivikki and Sakari Sohlberg foundation. N.M. was supported by the Academy of Finland (no. 331671). P.N.S. was supported by the Emil Aaltonen Foundation. G.L. is supported by University College London Hospitals' National Institute for Health Research (NIHR) Biomedical Research Centre, North Thames NIHR Applied Research Collaboration, as an NIHR Senior Investigator, and by the Wellcome Trust (no. 221854/Z/20/Z) and the UK Medical Research Council (no. MR/ S011676,). A.D.H. was supported by the UCL British Heart Foundation Accelerator (no. AA/18/6/34223), the UCL NIHR Biomedical Research Centre and the UKRI/NIHR-funded Multimorbidity Mechanism and Therapeutics Research Collaborative (no. MR/V033867/1). A.D.H. is a NIHR Senior Investigator. S.R. was supported by the Academy of Finland (nos. 285380 and 312062), the Sigrid Jusélius Foundation and University of Helsinki HiLIFE Fellow grant nos. 2017-2020. M.K. was supported by the Wellcome Trust (no. 221854/Z/20/Z), the UK Medical Research Council (nos. MR/S011676 and MR/R024227), the US National Institute on Aging (nos. R01AG062553 and R01AG056477), NordForsk (no. 75021), the Academy of Finland (nos. 311492 and 350426), the Helsinki Institute of Life Science (no. H970) and the Finnish Work Environment Fund (no. 190424). The funders had no role in study design, data collection and analysis, decision to publish or preparation of the manuscript.

### **Author contributions**

J.V.L., N.M., S.R. and M.K. generated the hypothesis and designed the study. J.V.L. wrote the first draft of the report, did the primary analyses, with support from N.M., and performed literature searches. All authors interpreted the data and critically commented on and reviewed the report. J.V.L. and M.K. had full access to pseudonymized data from the Whitehall II study. J.V.L., N.M. and S.R. had full access to pseudonymized data from the FinnGen study.

### **Funding**

Open Access funding provided by University of Helsinki including Helsinki University Central Hospital.

### **Competing interests**

H.R. is a full-time employee at Biogen. Biogen had no role in study design, data collection and analysis, decision to publish or preparation of the manuscript. The remaining authors declare no competing interests.

### **Additional information**

**Extended data** is available for this paper at https://doi.org/10.1038/s43587-022-00293-x.

**Supplementary information** The online version contains supplementary material available at https://doi.org/10.1038/s43587-022-00293-x.

**Correspondence and requests for materials** should be addressed to Joni V. Lindbohm.

**Peer review information** *Nature Aging* thanks Sara Hägg, Berislav Zlokovic and Stephanie Debette for their contribution to the peer review of this work.

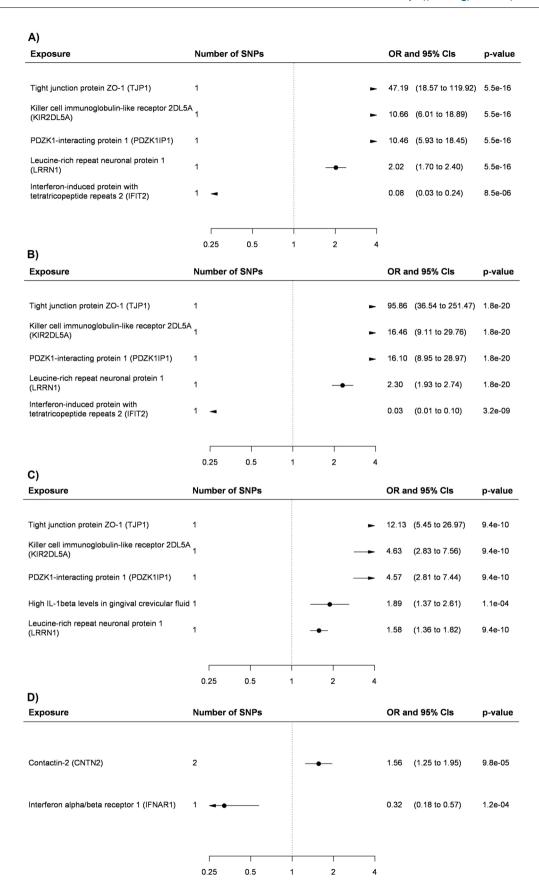
**Reprints and permissions information** is available at www.nature.com/reprints.

**Publisher's note** Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.

Open Access This article is licensed under a Creative Commons Attribution 4.0 International License, which permits use, sharing, adaptation, distribution and reproduction in any medium or format, as long as you give appropriate credit to the original author(s) and the source, provide a link to the Creative Commons license, and indicate if changes were made. The images or other third party material in this article are included in the article's Creative Commons license, unless indicated otherwise in a credit line to the material. If material is not included in the article's Creative Commons license and your intended use is not permitted by statutory regulation or exceeds the permitted use, you will need to obtain permission directly from the copyright holder. To view a copy of this license, visit <a href="https://creativecommons.org/licenses/by/4.0/">https://creativecommons.org/licenses/by/4.0/</a>.

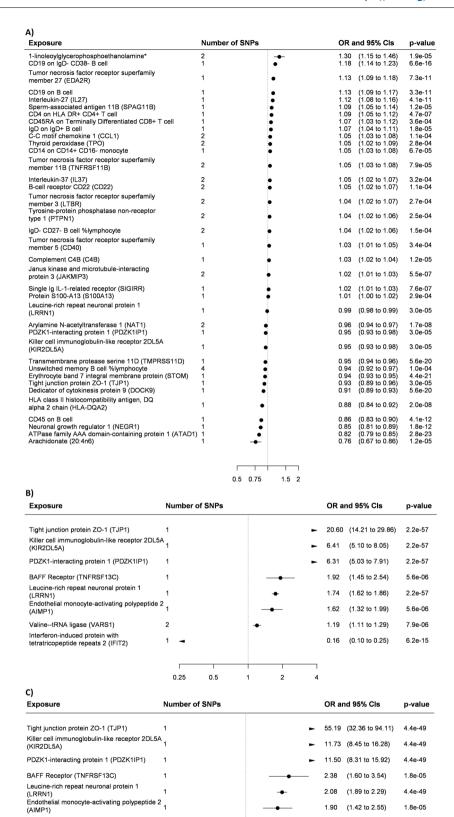
© The Author(s) 2022

Broad Institute of the Massachusetts Institute of Technology and Harvard University, The Klarman Cell Observatory, Cambridge, MA, USA. <sup>2</sup>Department of Epidemiology and Public Health, University College London, London, UK. <sup>3</sup>Clinicum, Department of Public Health, University of Helsinki, Helsinki, Finland. <sup>4</sup>Institute for Molecular Medicine Finland, HiLIFE, University of Helsinki, Helsinki, Finland. <sup>5</sup>Université de Paris, Inserm U1153, Epidemiology of Ageing and Neurodegenerative diseases, Paris, France. <sup>6</sup>Research & Development, Biogen Inc., Cambridge, MA, USA. <sup>7</sup>Division of Psychiatry, University College London, London, UK. <sup>8</sup>Camden and Islington NHS Foundation Trust, London, UK. <sup>9</sup>Glenn Biggs Institute of Alzheimer's and Neurodegenerative Diseases, University of Texas Health Science Center, San Antonio, TX, USA. <sup>10</sup>Boston University School of Public Health, Boston, MA, USA. <sup>11</sup>New York University Grossman School of Medicine, New York, NY, USA. <sup>12</sup>Boston University School of Medicine, Boston, MA, USA. <sup>13</sup>Center for Computational and Integrative Biology, Massachusetts General Hospital and Harvard Medical School, Boston, MA, USA. <sup>14</sup>Department of Molecular Biology, Massachusetts General Hospital and Harvard Medical School, Boston, MA, USA. <sup>15</sup>Institute of Cardiovascular Science, University College London, UK. <sup>16</sup>University College London, British Heart Foundation Research Accelerator, London, UK. <sup>17</sup>Health Data Research UK, London, UK. <sup>18</sup>These authors contributed equally: Samuli Ripatti, Mika Kivimäki. \*A list of members and their affiliations appears in the Supplementary Information. —e-mail: joni.lindbohm@helsinki.fi



Extended Data Fig. 1 | Odds ratios and 95% confidence intervals per one standard deviation increase in biomarker level derived from Wald ratios when only one SNPs was available and from inverse variance weighted Mendelian randomization when two or more SNPs were available. All

biomarkers passed false discovery rate correction of 5% (p-value < 0.00052) and all the tests were two-sided. (**A**) atypical or mixed Alzheimer's disease, (**B**) early onset Alzheimer's disease, (**C**) vascular dementia, (**D**) frontotemporal dementia.



Extended Data Fig. 2 | Odds ratios and 95% confidence intervals per one standard deviation increase in biomarker level derived from Wald ratios when only one SNPs was available and from inverse variance weighted Mendelian randomization when two or more SNPs were available. All

Valine--tRNA ligase (VARS1)

Interferon-induced protein with

tetratricopeptide repeats 2 (IFIT2)

0.25

0.5

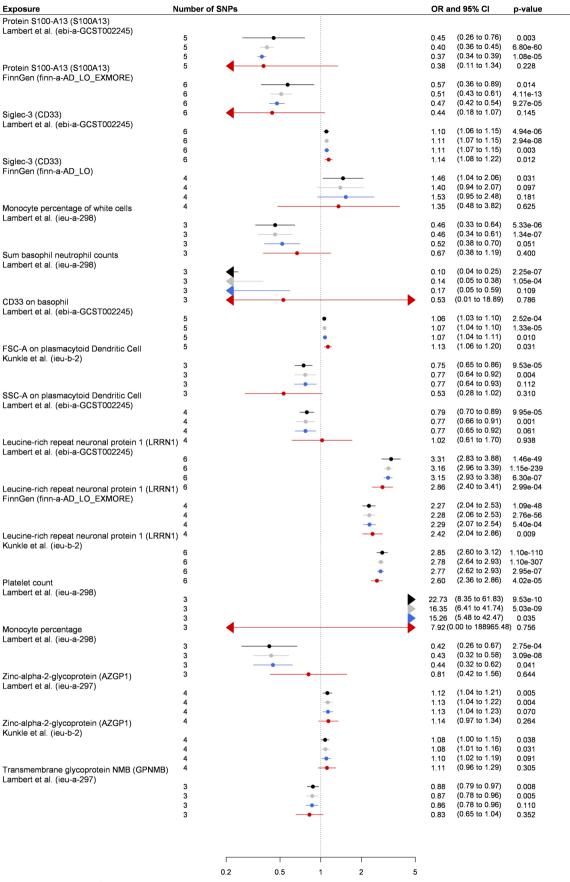
biomarkers passed false discovery rate correction of 5% (p-value < 0.00052) and all the tests were two-sided. (**A**) continuous cognitive performance, (**B**) general dementia outcome, (**C**) dementia in Alzheimer's disease.

2.5e-05

(1.14 to 1.41)

(0.03 to 0.11)

1.27



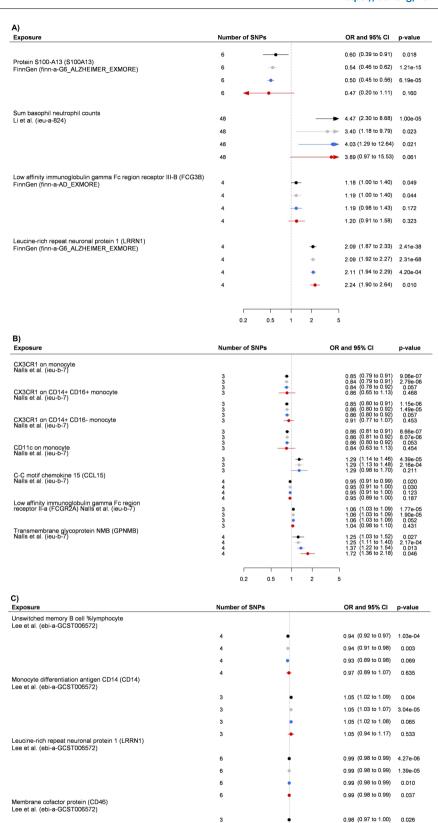
Extended Data Fig. 3 | See next page for caption.

Extended Data Fig. 3 | Odds ratios and 95% confidence intervals between one standard deviation change in biomarker levels and late onset Alzheimer's disease. Results are from Mendelian randomization sensitivity analyses when at least 3 SNPs were available. All biomarkers passed false discovery rate correction of 5% (p-value < 0.00052) in inverse variance weighted Mendelian randomization

and all the tests were two-sided. The source of outcome and MR-base outcome identifier is described below the Biomarker. Black = inverse variance weighted, grey = weighted median, blue = weighted mode, red = Egger Mendelian randomization derived estimate.

0.98 (0.97 to 0.99) 0.003 0.98 (0.96 to 0.99) 0.109 0.96 (0.93 to 0.99) 0.235

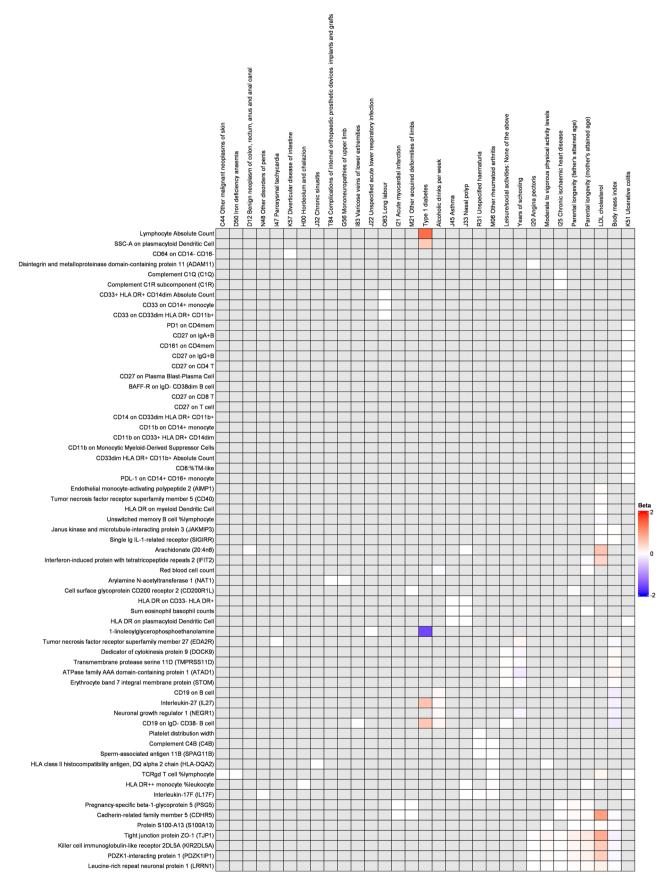
0.5



Extended Data Fig. 4 | See next page for caption.

Extended Data Fig. 4 | Odds ratios and 95% confidence intervals between one standard deviation change in biomarker levels and (A) Alzheimer's diseases, (B) Parkinson's disease, and (C) cognitive performance. Results are from Mendelian randomization sensitivity analyses when at least 3 SNPs were available. All biomarkers passed false discovery rate correction of 5% (p-value < 0.00052) in inverse variance weighted Mendelian randomization and all the

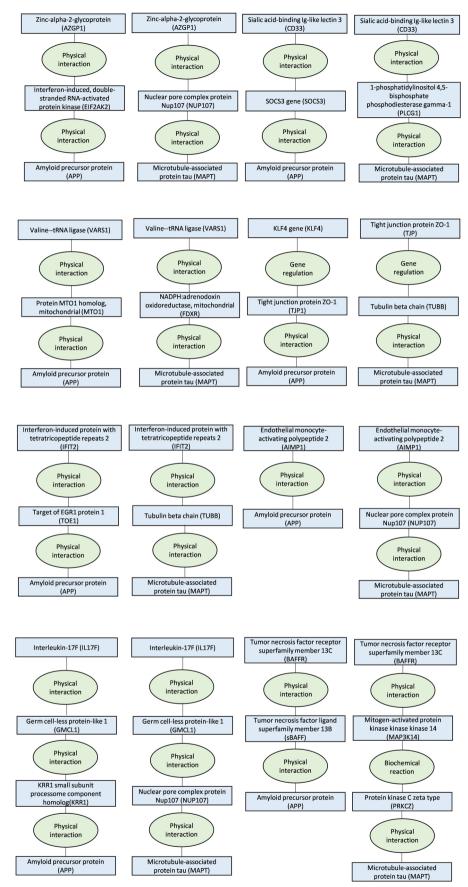
tests were two-sided. The source of outcome and MR-base outcome identifier is described below the Biomarker. Black = inverse variance weighted, grey = weighted median, blue = weighted mode, red = Egger Mendelian randomization derived estimate. Egger Mendelian randomization estimate for CD11c on monocyte is omitted because it did not converge.



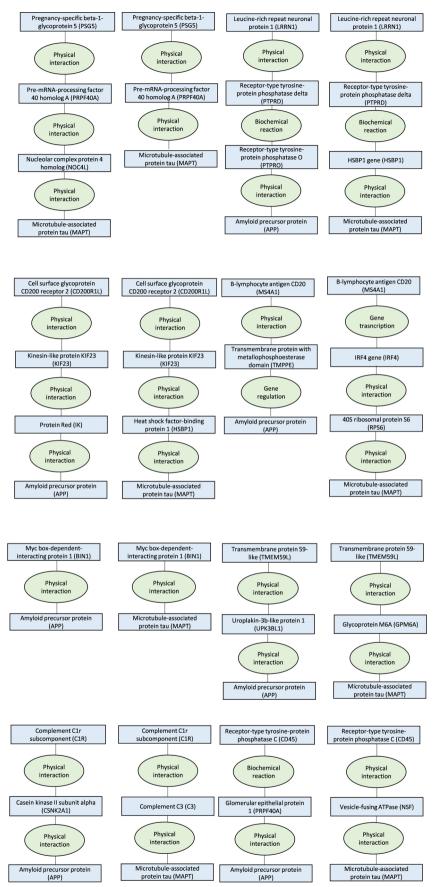
Extended Data Fig. 5 | See next page for caption.

Extended Data Fig. 5 | Phenome-wide Mendelian randomization analyses for the 127 biomarkers that associated with dementia causing diseases. Betas are derived from Wald ratios when only one SNP was available and from inverse variance weighted Mendelian randomization when two or more SNPs were available. Results are presented for the 63 of the 127 biomarkers that passed false

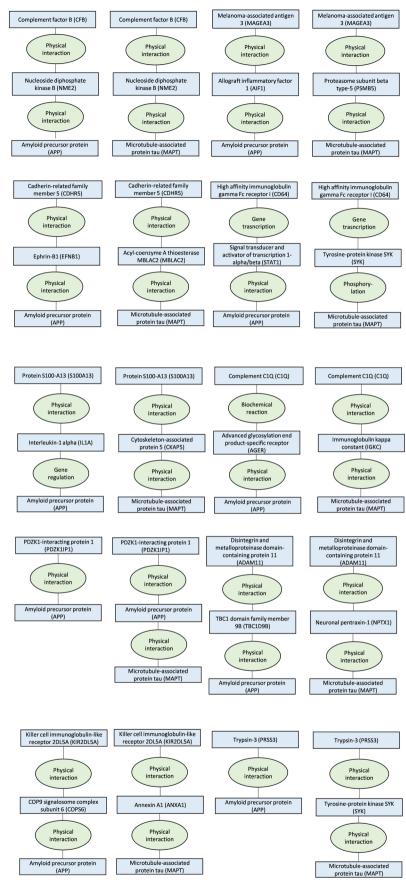
discovery rate correction of 5% (p-value < 0.00029). Most biomarkers associated with only few outcomes and the grey boxes indicate no association after false discovery rate correction of 5% (p-value < 0.00029) or lack of common SNPs. All the tests were two-sided.



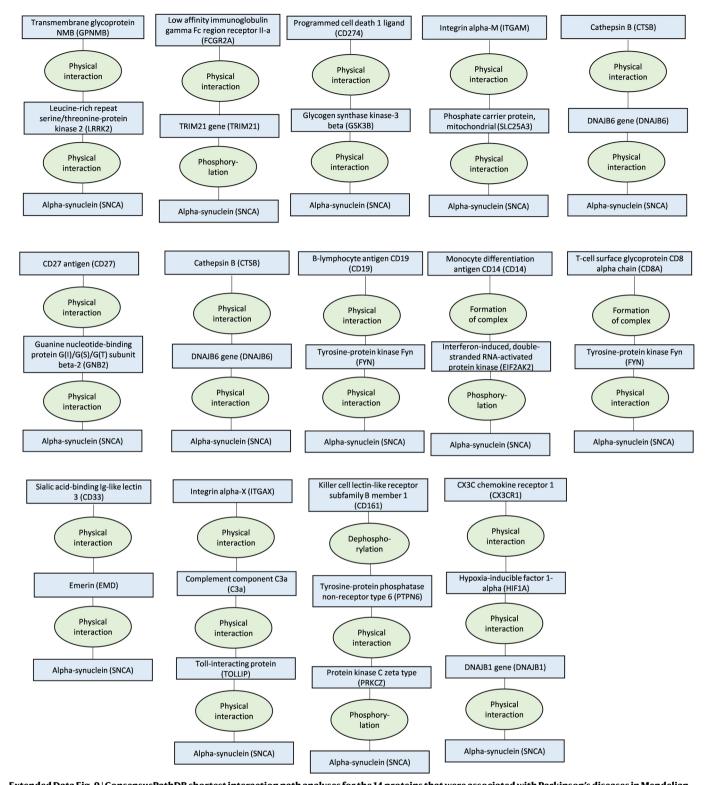
 $Extended \ Data \ Fig.\ 6 \ |\ Consensus \ Path DB\ shortest\ interaction\ path\ analyses\ for\ the\ first\ 8\ of\ the\ 26\ proteins\ that\ were\ associated\ with\ Alzheimer's\ diseases\ in\ Mendelian\ randomization\ analyses\ .$  The figure describes shortest interaction\ path\ between\ biomarkers\ and\ amyloid\ and\ tau.



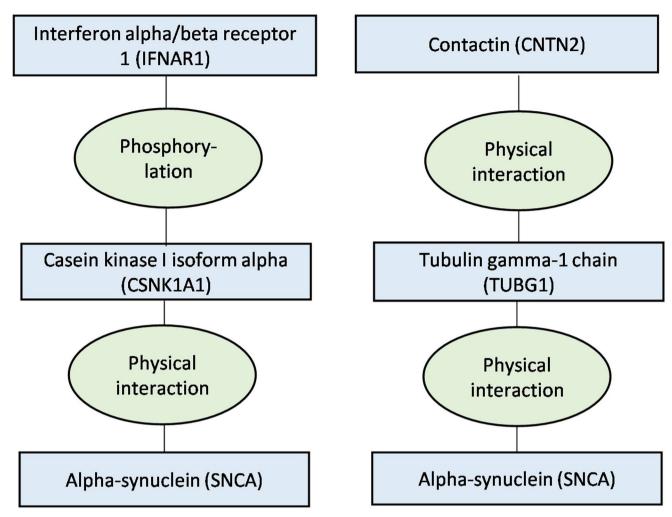
 $Extended \, Data \, Fig. \, 7 \, | \, Consensus Path DB \, shortest \, interaction \, path \, analyses \, for \, additional \, 8 \, of \, the \, 26 \, proteins \, that \, were \, associated \, with \, Alzheimer's \, diseases \, in \, Mendelian \, randomization \, analyses. \, The \, figure \, describes \, shortest \, interaction \, path \, between \, biomarkers \, and \, amyloid \, and \, tau.$ 



 $Extended \, Data \, Fig. \, 8 \, | \, Consensus Path DB \, shortest \, interaction \, path \, analyses \, for \, the \, last \, 10 \, of \, the \, 26 \, proteins \, that \, were \, associated \, with \, Alzheimer's \, diseases \, in \, Mendelian \, randomization \, analyses. \, The \, figure \, describes \, shortest \, interaction \, path \, between \, biomarkers \, and \, amyloid \, and \, tau.$ 



 $\label{lem:constraint} Extended Data Fig. 9 | Consensus Path DB shortest interaction path analyses for the 14 proteins that were associated with Parkinson's diseases in Mendelian randomization analyses. The figure describes shortest interaction path between biomarkers and $\alpha$-synuclein.$ 



 $\label{lem:extended} Extended Data Fig. 10 | Consensus Path DB shortest interaction path analyses for the 2 proteins that were associated with front otemporal dementia in Mendelian randomization analyses. The figure describes shortest interaction path between biomarkers and $\alpha$-synuclein.$ 

# nature research

Corresponding author(s):	Joni V. Lindbohm
Last updated by author(s):	Aug 31, 2022

# **Reporting Summary**

Nature Research wishes to improve the reproducibility of the work that we publish. This form provides structure for consistency and transparency in reporting. For further information on Nature Research policies, see our <u>Editorial Policies</u> and the <u>Editorial Policy Checklist</u>.

Stati	くさいへ	C

n/a Confirmed  The exact sample size (n) for each experimental group/condition, given as a discrete number and unit of measurement  A statement on whether measurements were taken from distinct samples or whether the same sample was measured repeatedly  The statistical test(s) used AND whether they are one- or two-sided  Only common tests should be described solely by name; describe more complex techniques in the Methods section.  A description of all covariates tested  A description of any assumptions or corrections, such as tests of normality and adjustment for multiple comparisons  A full description of the statistical parameters including central tendency (e.g. means) or other basic estimates (e.g. regression coefficient AND variation (e.g. standard deviation) or associated estimates of uncertainty (e.g. confidence intervals)  For null hypothesis testing, the test statistic (e.g. F, t, r) with confidence intervals, effect sizes, degrees of freedom and P value noted Give P values as exact values whenever suitable.  For Bayesian analysis, information on the choice of priors and Markov chain Monte Carlo settings  For hierarchical and complex designs, identification of the appropriate level for tests and full reporting of outcomes  For hierarchical sizes (e.g. Cohen's d, Pearson's r), indicating how they were calculated  Our web collection on statistics for biologists contains articles on many of the points above.	For	all st	atistical analyses, confirm that the following items are present in the figure legend, table legend, main text, or Methods section.
A statement on whether measurements were taken from distinct samples or whether the same sample was measured repeatedly  The statistical test(s) used AND whether they are one- or two-sided Only common tests should be described solely by name; describe more complex techniques in the Methods section.  A description of all covariates tested  A description of any assumptions or corrections, such as tests of normality and adjustment for multiple comparisons  A full description of the statistical parameters including central tendency (e.g. means) or other basic estimates (e.g. regression coefficient AND variation (e.g. standard deviation) or associated estimates of uncertainty (e.g. confidence intervals)  For null hypothesis testing, the test statistic (e.g. F, t, r) with confidence intervals, effect sizes, degrees of freedom and P value noted Give P values as exact values whenever suitable.  For Bayesian analysis, information on the choice of priors and Markov chain Monte Carlo settings  For hierarchical and complex designs, identification of the appropriate level for tests and full reporting of outcomes  Estimates of effect sizes (e.g. Cohen's d, Pearson's r), indicating how they were calculated	n/a	Cor	nfirmed
The statistical test(s) used AND whether they are one- or two-sided  Only common tests should be described solely by name; describe more complex techniques in the Methods section.  A description of all covariates tested  A description of any assumptions or corrections, such as tests of normality and adjustment for multiple comparisons  A full description of the statistical parameters including central tendency (e.g. means) or other basic estimates (e.g. regression coefficient AND variation (e.g. standard deviation) or associated estimates of uncertainty (e.g. confidence intervals)  For null hypothesis testing, the test statistic (e.g. F, t, r) with confidence intervals, effect sizes, degrees of freedom and P value noted Give P values as exact values whenever suitable.  For Bayesian analysis, information on the choice of priors and Markov chain Monte Carlo settings  For hierarchical and complex designs, identification of the appropriate level for tests and full reporting of outcomes  Estimates of effect sizes (e.g. Cohen's d, Pearson's r), indicating how they were calculated		$\boxtimes$	The exact sample size $(n)$ for each experimental group/condition, given as a discrete number and unit of measurement
<ul> <li>A description of all covariates tested</li> <li>A description of any assumptions or corrections, such as tests of normality and adjustment for multiple comparisons</li> <li>A full description of the statistical parameters including central tendency (e.g. means) or other basic estimates (e.g. regression coefficient AND variation (e.g. standard deviation) or associated estimates of uncertainty (e.g. confidence intervals)</li> <li>For null hypothesis testing, the test statistic (e.g. F, t, r) with confidence intervals, effect sizes, degrees of freedom and P value noted Give P values as exact values whenever suitable.</li> <li>For Bayesian analysis, information on the choice of priors and Markov chain Monte Carlo settings</li> <li>For hierarchical and complex designs, identification of the appropriate level for tests and full reporting of outcomes</li> <li>Estimates of effect sizes (e.g. Cohen's d, Pearson's r), indicating how they were calculated</li> </ul>		$\boxtimes$	A statement on whether measurements were taken from distinct samples or whether the same sample was measured repeatedly
A description of any assumptions or corrections, such as tests of normality and adjustment for multiple comparisons  A full description of the statistical parameters including central tendency (e.g. means) or other basic estimates (e.g. regression coefficient AND variation (e.g. standard deviation) or associated estimates of uncertainty (e.g. confidence intervals)  For null hypothesis testing, the test statistic (e.g. F, t, r) with confidence intervals, effect sizes, degrees of freedom and P value noted Give P values as exact values whenever suitable.  For Bayesian analysis, information on the choice of priors and Markov chain Monte Carlo settings  For hierarchical and complex designs, identification of the appropriate level for tests and full reporting of outcomes  Estimates of effect sizes (e.g. Cohen's d, Pearson's r), indicating how they were calculated		$\boxtimes$	The statistical test(s) used AND whether they are one- or two-sided Only common tests should be described solely by name; describe more complex techniques in the Methods section.
A full description of the statistical parameters including central tendency (e.g. means) or other basic estimates (e.g. regression coefficient AND variation (e.g. standard deviation) or associated estimates of uncertainty (e.g. confidence intervals)  For null hypothesis testing, the test statistic (e.g. <i>F</i> , <i>t</i> , <i>r</i> ) with confidence intervals, effect sizes, degrees of freedom and <i>P</i> value noted Give P values as exact values whenever suitable.  For Bayesian analysis, information on the choice of priors and Markov chain Monte Carlo settings  For hierarchical and complex designs, identification of the appropriate level for tests and full reporting of outcomes  Estimates of effect sizes (e.g. Cohen's <i>d</i> , Pearson's <i>r</i> ), indicating how they were calculated		$\boxtimes$	A description of all covariates tested
For null hypothesis testing, the test statistic (e.g. <i>F</i> , <i>t</i> , <i>r</i> ) with confidence intervals, effect sizes, degrees of freedom and <i>P</i> value noted <i>Give P values as exact values whenever suitable.</i> For Bayesian analysis, information on the choice of priors and Markov chain Monte Carlo settings  For hierarchical and complex designs, identification of the appropriate level for tests and full reporting of outcomes  Estimates of effect sizes (e.g. Cohen's <i>d</i> , Pearson's <i>r</i> ), indicating how they were calculated		$\boxtimes$	A description of any assumptions or corrections, such as tests of normality and adjustment for multiple comparisons
For Bayesian analysis, information on the choice of priors and Markov chain Monte Carlo settings  For hierarchical and complex designs, identification of the appropriate level for tests and full reporting of outcomes  Estimates of effect sizes (e.g. Cohen's <i>d</i> , Pearson's <i>r</i> ), indicating how they were calculated		$\boxtimes$	A full description of the statistical parameters including central tendency (e.g. means) or other basic estimates (e.g. regression coefficient AND variation (e.g. standard deviation) or associated estimates of uncertainty (e.g. confidence intervals)
For hierarchical and complex designs, identification of the appropriate level for tests and full reporting of outcomes  Estimates of effect sizes (e.g. Cohen's <i>d</i> , Pearson's <i>r</i> ), indicating how they were calculated		$\boxtimes$	For null hypothesis testing, the test statistic (e.g. <i>F</i> , <i>t</i> , <i>r</i> ) with confidence intervals, effect sizes, degrees of freedom and <i>P</i> value noted <i>Give P values as exact values whenever suitable.</i>
Estimates of effect sizes (e.g. Cohen's $d$ , Pearson's $r$ ), indicating how they were calculated	$\boxtimes$		For Bayesian analysis, information on the choice of priors and Markov chain Monte Carlo settings
	$\boxtimes$		For hierarchical and complex designs, identification of the appropriate level for tests and full reporting of outcomes
Our web collection on <u>statistics for biologists</u> contains articles on many of the points above.		$\boxtimes$	Estimates of effect sizes (e.g. Cohen's d, Pearson's r), indicating how they were calculated
			Our web collection on <u>statistics for biologists</u> contains articles on many of the points above.

### Software and code

Policy information about availability of computer code

Data collection No software was used for data collection.

Data analysis

We used following softwares for imputation, handling the genetic and phenotypic data:

Cromwell 61 PLINK v2.00aLM3 BCFtools 1.7 and 1.9 Eagle 2.3.5 Beagle 4.1 (08Jun17.d8b)

The full genotyping and imputation protocol for FinnGen is described at dx.doi.org/10.17504/protocols.io.nmndc5e For data transformations, visualization and plotting of the results, we used R (3.6.0, 4.1.0, 4.1.2), including packages:

data.table 1.14.2 dplyr 1.0.7 tidyr 1.1.4 survival 3.2.13

survminer 0.4.9 ggplot2 3.3.5 plyr 1.8.6

cluster 2.1.2 lubridate 1.8.0 stats 4.1.1 readxl 1.3.1

scales 1.1.1

tidyverse 1.3.1	
Hmisc 4.6.0	
devtools 2.4.2	
TwoSampleMR 0.5.6	
MRInstruments 0.3.2	
metafor 3.0.2.	
Other software include: ClueGO version 2.5.8, Cytoscape version 3.7.2.	
ggnlot2 surviviner TwoSampleMR MRInstruments and inw	

For manuscripts utilizing custom algorithms or software that are central to the research but not yet described in published literature, software must be made available to editors and reviewers. We strongly encourage code deposition in a community repository (e.g. GitHub). See the Nature Research guidelines for submitting code & software for further information.

### Data

Policy information about availability of data

All manuscripts must include a data availability statement. This statement should provide the following information, where applicable:

- Accession codes, unique identifiers, or web links for publicly available datasets
- A list of figures that have associated raw data
- A description of any restrictions on data availability

The updated data availability statement is here as it is described in the manuscript:

Data availability

Blinding

This study used publicly available data at https://www.mrbase.org/, https://www.uniprot.org/, https://cpdb.molgen.mpg.de/, https://www.genome.jp/kegg/, and https://www.opentargets.org/. Data used in Mendelian Randomization are deposited to Zenodo: https://zenodo.org/deposit/7042008. Data, protocols, and other metadata of the Whitehall II and FinnGen studies are available according to the data sharing policies of these studies. Pre-existing data access policy for Whitehall II specify that research data requests can be submitted to study steering committee; these will be promptly reviewed for confidentiality or intellectual property restrictions and will not unreasonably be refused. Detailed information on data sharing can be found here: https://www.ucl.ac.uk/epidemiology-health-care/research/epidemiology-and-public-health/research/whitehall-ii/data-sharing Individual-level patient or protein data may further be restricted by consent, confidentiality, or privacy laws/considerations. The FinnGen data can be accessed through Finnish Biobanks' FinBB portal (www.finbb.fi). FinnGen summary statistics are freely available at https://www.finngen.fi/en/access\_results, with results for new data freezes updated every 6 months.

Please select the one below that is the best fit for your research. If you are not sure, read the appropriate sections before making your selection.

	•		1		• •	•			100		
⊢.	וםו		_ C	മ	CIT		rΔ	ро	rtı	n	$\boldsymbol{\sigma}$
	ı	ı	כדו	$\cup$	CII	IL.	$I \subset$	$\mathbf{u}\mathbf{u}$	'I LI		⋍
		_	_	_			. –	-			U

	, , , , , , , , , , , , , , , , , , , ,
∠ Life sciences	Behavioural & social sciences Ecological, evolutionary & environmental sciences
For a reference copy of	the document with all sections, see <u>nature.com/documents/nr-reporting-summary-flat.pdf</u>
Life scier	nces study design
All studies must dis	close on these points even when the disclosure is negative.
Sample size	We used all available data on individuals 18 or older at the end of follow-up. The study is based on FinnGen Data Freeze 8 with 339,233 individuals. In the Whitehall II study we used the 6,545 individuals with plasma proteomic data available.
Data exclusions	Inverse-probability-weighted analyses in FinnGen used only 117,773 participants aged above 45 and with no medication studied at baseline. This was done to ensure that all participants were old enough to develop dementia over the follow-up and to simulate randomized control trial design using observational data.
Replication	The study did not include direct replication with same methods in different cohort. Instead it used six complimentary approaches to study the same research question and all provided consistent results.
Randomization	The study used inverse-probability-weighted analyses to simulate randomized control trial design using observational data. This is an observational study.

## Reporting for specific materials, systems and methods

No blinding was relevant for this observational study.

We require information from authors about some types of materials, experimental systems and methods used in many studies. Here, indicate whether each material, system or method listed is relevant to your study. If you are not sure if a list item applies to your research, read the appropriate section before selecting a response.

Matariala C avecrimo anto	Louatores				
Materials & experimenta	<del></del>				
n/a   Involved in the study 	n/a   Involved in the study				
	ChIP-seq				
Eukaryotic cell lines	Flow cytometry				
Palaeontology and archa					
Animals and other organ					
Human research particip	ants				
Clinical data					
Dual use research of con-	cern				
Human research par	ticipants				
Policy information about studie	s involving human research participants				
Population characteristics	This study included participants from different study cohorts, all of European ancestry. FinnGen Data Freeze 8 comprises 339,233 individuals (56.3% women; mean age 59.8 at the end of follow-up in 2019, with standard deviation, s.d. 17.3 years).				
	In Whitehall II the mean age was 56.1 (SD 5.9) and 73.0% were men and 92.3% of European ancestry.				
Recruitment	Random sample of subjects from Finnish population-based and clinical biobanks. Participation was voluntary. A proportion of FinnGen was ascertained through hospital biobanks and disease-based collections.				
	In 1985-1988, all civil servants aged 35 to 55 years based in 20 departments in London, UK, were invited to participate in the Whitehall II cohort study, and 73% (n=10 308) agreed. Blood samples for proteomic analyses were collected from a random subsample of 6,545 dementia-free individuals in 1997-99 which was the baseline for this study.				
Ethics oversight	Patients and control subjects in FinnGen provided informed consent for biobank research, based on the Finnish Biobank Act. Alternatively, separate research cohorts, collected prior the Finnish Biobank Act came into effect (in September 2013) and start of FinnGen (August 2017), were collected based on study-specific consents and later transferred to the Finnish biobanks after approval by Fimea (Finnish Medicines Agency), the National Supervisory Authority for Welfare and Health. Recruitment protocols followed the biobank protocols approved by Fimea. The Coordinating Ethics Committee of the Hospital District of Helsinki and Uusimaa (HUS) statement number for the FinnGen study is Nr HUS/990/2017.				
	The FinnGen study is approved by Finnish Institute for Health and Welfare (permit numbers: THL/2031/6.02.00/2017, THL/1101/5.05.00/2017, THL/341/6.02.00/2018, THL/2222/6.02.00/ 2018, THL/283/6.02.00/2019, THL/1721/5.05.00/2019, THL/1524/5.05.00/2020, and THL/2364/ 14.02/2020), Digital and population data service agency (permit numbers: VRK43431/2017-3, VRK/6909/2018-3, VRK/4415/2019-3), the Social Insurance Institution (permit numbers: KELA 58/522/2017, KELA 131/522/2018, KELA 70/522/2019, KELA 98/522/2019, KELA 138/522/2019, KELA 2/522/2020, KELA 16/522/2020, Findata THL/2364/14.02/2020 and Statistics Finland (permit numbers: TK-53-1041-17 and TK/143/07.03.00/2020 (earlier TK-53-90-20).				
	The Biobank Access Decisions for FinnGen samples and data utilized in FinnGen Data Freeze 7 include: THL Biobank BB2017_55, BB2017_111, BB2018_19, BB_2018_34, BB_2018_67, BB2018_71, BB2019_7, BB2019_8, BB2019_26, BB2020_1, Finnish Red Cross Blood Service Biobank 7.12.2017, Helsinki Biobank HUS/359/2017, Auria Biobank AB17-5154 and amendment #1 (August 17 2020), Biobank Borealis of Northern Finland_2017_1013, Biobank of Eastern Finland 1186/2018 and amendment 22 § /2020, Biobank Tampere MH0004 and amendments (21.02.2020 & 0.000) Control Finland Biobank 1.2017, and Tampere MH0004 and amendments (21.02.2020 & 0.000).				

06.10.2020), Central Finland Biobank 1-2017, and Terveystalo Biobank STB 2018001.

In Whitehall II study, participants provided written informed consent for inclusion at each examination. Research ethics approval was granted by the University College London Hospital Committee on the Ethics of Human Research.

Note that full information on the approval of the study protocol must also be provided in the manuscript.