## Research Article

# Chronic Oral Anticoagulation Therapy and Prognosis of Patients Admitted to Hospital for COVID-19: Insights from the HOPE COVID-19 Registry

José Miguel Rivera-Caravaca, 1,2 Iván J. Núñez-Gil,3 Gregory Y. H. Lip, 2,4
Aitor Uribarri,5 María C. Viana-Llamas,6 Adelina Gonzalez,7 Alex F. Castro-Mejía,8
Berta Alonso González,9 Emilio Alfonso,10 Juan Fortunato García Prieto,11
Chiara Cavallino,12 Bernardo Cortese,13 Gisela Feltes,14 Inmaculada Fernández-Rozas,15
Jaime Signes-Costa,16 Jia Huang,17 Marcos García Aguado,18 Martino Pepe,19
Rodolfo Romero,20 Enrico Cerrato,21 Víctor Manuel Becerra-Muñoz,22
Sergio Raposeiras Roubin,23 Francesco Santoro,24,25 Rodrigo Bagur,26
Luciano Sposato,26,27,28,29 Ibrahim El-Battrawy,30 Alvaro López Masjuan,31
Antonio Fernandez-Ortiz,3 Vicente Estrada,3 Carlos Macaya,3 and Francisco Marín 10

<sup>&</sup>lt;sup>1</sup>Department of Cardiology, Hospital Clínico Universitario Virgen de la Arrixaca, University of Murcia, Instituto Murciano de Investigación Biosanitaria (IMIB-Arrixaca), CIBERCV, Murcia, Spain

<sup>&</sup>lt;sup>2</sup>Liverpool Centre for Cardiovascular Science, University of Liverpool, Liverpool Heart & Chest Hospital, Liverpool, UK

<sup>&</sup>lt;sup>3</sup>Hospital Clínico San Carlos Universidad Complutense de Madrid,

Instituto de Investigación Sanitaria del Hospital Clínico San Carlos (IdISSC), Madrid, Spain

<sup>&</sup>lt;sup>4</sup>Department of Clinical Medicine, Aalborg University, Aalborg, Denmark

<sup>&</sup>lt;sup>5</sup>Hospital Clínico Universitario de Valladolid, Valladolid, Spain

<sup>&</sup>lt;sup>6</sup>Hospital Universitario Guadalajara, Guadalajara, Spain

<sup>&</sup>lt;sup>7</sup>Hospital Universitario Infanta Sofía, San Sebastian de los Reyes, Madrid, Spain

<sup>&</sup>lt;sup>8</sup>Hospital General del Norte de Guayaquil IESS Los Ceibos, Guayaquil, Ecuador

<sup>&</sup>lt;sup>9</sup>Hospital Universitario la Paz, Madrid, Spain

<sup>&</sup>lt;sup>10</sup>Instituto de Cardiología y Cirugía Cardiovascular, La Habana, Cuba

<sup>&</sup>lt;sup>11</sup>Hospital de Manises, Valencia, Spain

<sup>&</sup>lt;sup>12</sup>Sant' Andrea Hospital, Vercelli, Italy

<sup>&</sup>lt;sup>13</sup>San Carlo Clinic, Milano, Italy

<sup>&</sup>lt;sup>14</sup>Hospital Nuestra Señora de América, Madrid, Spain

<sup>&</sup>lt;sup>15</sup>Hospital Severo Ochoa, Leganés, Spain

<sup>&</sup>lt;sup>16</sup>Hospital Clínico Universitario, INCLIVA, Universidad de Valencia, Valencia, Spain

<sup>&</sup>lt;sup>17</sup>The Second People's Hospital of Shenzhen, Shenzhen, China

<sup>&</sup>lt;sup>18</sup>Hospital Universitario Puerta de Hierro, Majadahonda, Spain

<sup>&</sup>lt;sup>19</sup>Azienda Ospedaliero Universitaria Consorziale Policlinico di Bari, Bari, Italy

<sup>&</sup>lt;sup>20</sup>Hospital Universitario de Getafe, Madrid, Spain

<sup>&</sup>lt;sup>21</sup>San Luigi Gonzaga University Hospital, Rivoli, Turin, Italy

<sup>&</sup>lt;sup>22</sup>Unidad de Gestión Clínica Área del Corazón, Instituto de Investigación Biomédica de Málaga (IBIMA), Hospital Universitario Virgen de la Victoria, Universidad de Málaga (UMA),

Centro de Investigación Biomédica en Red Enfermedades Cardiovasculares (CIBERCV), Málaga, Spain

<sup>&</sup>lt;sup>23</sup>University Hospital Álvaro Cunqueiro, Vigo, Spain

<sup>&</sup>lt;sup>24</sup>Azienda Sanitaria Locale della Provincia di Barletta-Andria-Trani, Andria, Italy

<sup>&</sup>lt;sup>25</sup>Department of Medical and Surgery Sciences, University of Foggia, Foggia, Italy

<sup>&</sup>lt;sup>26</sup>London Health Sciences Centre, London, Ontario, Canada

<sup>&</sup>lt;sup>27</sup>Schulich School of Medicine and Dentistry, Western University, London, Ontario, Canada

<sup>&</sup>lt;sup>28</sup>Lawson Health Research Institute, London, Ontario, Canada

<sup>&</sup>lt;sup>29</sup>Robarts Research Institute, London, Ontario, Canada

<sup>30</sup>University of Mannheim, Mannheim, Germany

Correspondence should be addressed to Francisco Marín; fcomarino@hotmail.com

Received 11 February 2022; Revised 21 March 2022; Accepted 21 April 2022; Published 26 May 2022

Academic Editor: Melanie Deutsch

Copyright © 2022 José Miguel Rivera-Caravaca et al. This is an open access article distributed under the Creative Commons Attribution License, which permits unrestricted use, distribution, and reproduction in any medium, provided the original work is properly cited.

Background. Most evidence regarding anticoagulation and COVID-19 refers to the hospitalization setting, but the role of oral anticoagulation (OAC) before hospital admission has not been well explored. We compared clinical outcomes and short-term prognosis between patients with and without prior OAC therapy who were hospitalized for COVID-19. Methods. Analysis of the whole cohort of the HOPE COVID-19 Registry which included patients discharged (deceased or alive) after hospital admission for COVID-19 in 9 countries. All-cause mortality was the primary endpoint. Study outcomes were compared after adjusting variables using propensity score matching (PSM) analyses. Results. 7698 patients were suitable for the present analysis (675 (8.8%) on OAC at admission: 427 (5.6%) on VKAs and 248 (3.2%) on DOACs). After PSM, 1276 patients were analyzed (638 with OAC; 638 without OAC), without significant differences regarding the risk of thromboembolic events (OR 1.11, 95% CI 0.59–2.08). The risk of clinically relevant bleeding (OR 3.04, 95% CI 1.92–4.83), as well as the risk of mortality (HR 1.22, 95% CI 1.01–1.47; log-rank p value = 0.041), was significantly increased in previous OAC users. Amongst patients on prior OAC only, there were no differences in the risk of clinically relevant bleeding, thromboembolic events, or mortality when comparing previous VKA or DOAC users, after PSM. Conclusion. Hospitalized COVID-19 patients on prior OAC therapy had a higher risk of mortality and worse clinical outcomes compared to patients without prior OAC therapy, even after adjusting for comorbidities using a PSM. There were no differences in clinical outcomes in patients previously taking VKAs or DOACs. This trial is registered with NCT04334291/EUPAS34399.

#### 1. Introduction

Vascular inflammation, hypercoagulable state, and endothelial dysfunction have been described in patients with severe acute respiratory syndrome coronavirus-2 (SARS-CoV-2) infection [1, 2]. As a result, thromboembolic complications are common in patients with coronavirus disease 2019 (COVID-19) [3–5]. Therefore, antithrombotic therapy, particularly anticoagulation, gained attention in the context of COVID-19. Indeed, some studies demonstrated that anticoagulation may be associated with improved outcomes among patients with COVID-19 [6, 7]. However, most of the evidence in relation to anticoagulation and COVID-19 refer to the acute hospitalization setting, whereas the role of stable oral anticoagulation (OAC) therapy before the admission for COVID-19 has not been well explored.

One study suggested a protective role of chronic directacting OAC (DOAC) therapy in elderly patients with COVID-19 [8]. In a preliminary analysis of the International COVID-19 Clinical Evaluation (HOPE COVID-19) Registry published previously, we observed that COVID-19 patients on OAC therapy at hospital admission had lower survival and higher mortality risk compared to patients without prior OAC [9].

In the present study, we aimed to compare clinical outcomes and in-hospital prognosis between patients on prior OAC therapy and patients not on OAC therapy who were admitted for COVID-19 and enrolled in the HOPE COVID-19 Registry, using a propensity score matching (PSM) approach. Second, we aimed to compare clinical outcomes and prognosis between patients on vitamin K

antagonist (VKA) therapy and patients on DOACs before admission.

#### 2. Methods

A detailed description of the HOPE COVID-19 Registry has been published elsewhere [10, 11]. Briefly, the HOPE COVID-19 is an ambispective international registry, real-life cohort "all comers" type, including more than 8100 patients from 9 countries (Canada, China, Chile, Colombia, Cuba, Ecuador, Germany, Italy, and Spain). The study was an initiative without conflicts of interest, no financial remuneration, and methodological support from the Institute for the Improvement of Health Care (IMAS) foundation (Madrid, Spain).

All patients discharged (deceased or alive) after hospital admissions for COVID-19 were suitable for the study. There were no exclusion criteria, except for patients' explicit refusal to participate. The first patient was included in February 2020. Clinical and demographic data were collected at inclusion and during the hospitalization in an anonymized database presented in the electronic format, to be filled in at each participating center (NCT04334291/EUPAS34399).

Reporting of the study conforms to broad EQUATOR guidelines. The study was performed according to the ethical principles of Declaration of Helsinki and Good Clinical Practice Guidelines and has been approved by Ethics Research Committee from the Hospital Clínico San Carlos (Madrid, Spain) (20/241-E) and the Spanish Agency for Medicines and Health Products (EPA-0D). Given the anonymous characteristics of the registry and the health

<sup>&</sup>lt;sup>31</sup>Hospital Universitario Juan Ramón Jiménez, Huelva, Spain

alarm situation generated by the virus, in principle, written informed consent was waived. However, at least verbal authorization from the patient (or familiar or caregiver, when unavailable) was required.

2.1. Laboratory Analyses. Laboratory parameters were considered elevated as defined by local laboratory cutoff levels. However, the HOPE COVID-19 Registry protocol suggested the following as "elevated:" for D-dimer (≥0.5 mg/L), for procalcitonin (≥0.5 ng/mL), for C-reactive protein (≥10 mg/L), for troponins (>99th percentile), for transaminases (≥40 U/L), for ferritin (≥336 ng/mL), and for lactate dehydrogenase (≥280 U/L).

2.2. Study Outcomes. The primary endpoint for this analysis was in-hospital all-cause mortality. Any thrombotic/thromboembolic event and any clinically relevant bleeding were the secondary outcomes. Bleeding was defined as "relevant" at the discretion of the attending medical team and classified using the BARC bleeding score as type 2, 3, or 5.

Although not classified as primary or secondary outcomes, other adverse events during hospitalization were recorded, including renal failure, respiratory insufficiency, upper respiratory tract infection, heart failure, sepsis, and systemic inflammatory response syndrome (SIRS).

Local researchers identified, confirmed, and recorded all adverse events. The clinical management was decided, in all cases, by the attending team and researchers had no role in this point.

2.3. Statistical Analysis. Quantitative variables were expressed as mean ± standard deviation (SD) or median and interquartile range (IQR), as appropriate according to the Kolmogorov–Smirnov test, whilst categorical variables were expressed as absolute frequencies and percentages. Pearson's chi-squared test was used to compare proportions. Differences between two groups regarding a quantitative variable were tested with Student's t or the Mann–Whitney U tests, as appropriate if normally or not normally distributed.

To compare the risk of the study outcomes among patients on prior OAC therapy and patients without prior OAC therapy, we conducted a propensity score matching (PSM) adjusting for demographics and baseline comorbidities. The risk of the study outcomes among patients on prior VKA therapy or DOACs was also evaluated by another PSM. In both PSMs, those variables that were significantly different between both cohorts were included in the model to adjust for differences. Patients were matched 1:1 across each cohort on a propensity score generated by logistic regressions using the nearest neighbour technique without replacement with a maximum caliper of 0.2, thus avoiding at least 98% of the bias due to the measured confounders. The value of absolute standardized mean difference <10% indicated balance of matched cohorts [12, 13].

Survival analyses by Kaplan-Meier estimates were performed after PSM to assess differences in event-free

survival of the primary outcome depending on the use (or not) of prior OAC therapy and depending on the use of prior VKA or DOAC therapy. The risk of suffering from the primary outcome was assessed by Cox proportional hazard regression, and results were reported as hazard ratio (HR) with 95% confidence interval (CI). The risk of suffering from other study outcomes was investigated by logistic regression analyses, since the exact date for these events was not recorded. In these analyses, results were reported as odds ratio (OR) with 95% confidence interval (CI).

Two-sided *p* values <0.05 were accepted as statistically significant. Statistical analyses were performed using SPSS v. 24.0 (SPSS, Inc., Chicago, IL, USA) and MedCalc v. 16.4.3 (MedCalc Software bvba, Ostend, Belgium) for Windows.

## 3. Results

A cohort of 8168 patients was included. After excluding patients with insufficient or not reliable data on previous OAC, 7698 patients remained in the study (4500 (58.5%) male; median age of 65 (IQR 51–77) years). Of these, 675 (8.8%) were on OAC therapy at hospital admission, 427 (5.6%) were on VKAs, and 248 (3.2%) were on DOACs.

3.1. Outcomes on Prior OAC Therapy. In the overall cohort of 7698 patients, we found that patients on prior OAC therapy were less commonly admitted in the intensive care unit (ICU) compared to patients not previously taking OACs (6.7% vs. 10.1%, p = 0.004). During hospitalization, the prognosis of patients on prior OAC therapy was also poor, and these patients had more incident heart failure, renal failure, sepsis, and SIRS (all with p value <0.001). As expected, the risk of any clinically relevant bleeding in patients with previous OAC therapy was higher compared to patients not taking OAC previously (11.6% vs. 3.4%, p < 0.001; OR 3.71, 95% CI 2.83-4.85), without differences in terms of thromboembolic events (3.1% vs. 2.7%, p = 0.493). The risk of mortality was found to be significantly increased in patients on prior OAC therapy (39.1% vs. 17.0%, p < 0.001; HR 2.45, 95% CI 2.14-2.79); however, there were significant differences between patients on prior and not on prior OAC in terms of several comorbidities. We therefore performed PSM to adjust these analyses (Table 1).

After PSM, 1276 patients remained in the study (638: 638 paired comparisons), with no significant differences regarding admission to the ICU in patients on prior OAC compared to patients not previously taking OACs (6.9% vs. 6.3%, p=0.652). The prognosis of patients on prior OAC therapy during hospitalization was still poor even after adjustment, and these patients suffered more commonly from heart failure, renal failure, and SIRS (all with p value <0.05). No significant differences were found in terms of respiratory insufficiency (67.2% vs. 64.7%; p=0.280), upper respiratory tract infection (13.9% vs. 14.1%; p=0.987), or sepsis (15.0% vs. 12.1%; p=0.299) (Table 2).

Similar to the finding observed before PSM, the risk of any clinically relevant bleeding was higher in patients with previous OAC therapy compared to patients not taking

Table 1: Comparison of clinical characteristics of the study cohort before and after propensity score matching.

| -   | Before propensity score matching  |                                   |                | After propensity score matching      |                                   |            |  |
|---|-----------------------------------|-----------------------------------|----------------|--------------------------------------|-----------------------------------|------------|--|
|   | Patients without prior OAC N=7023 | Patients with prior OAC $N = 675$ | <i>P</i> value | Patients without prior OAC $N = 638$ | Patients with prior OAC $N = 638$ | P<br>value |  |
| Demographic   |                                   |                                   |                |                                      |                                   |            |  |
| Male sex, n (%)                                       | 4097 (58.3)                       | 403 (59.7)                        | 0.491          | 386 (60.5)                           | 372 (58.3)                        | 0.425      |  |
| Age (years), median (IQR)                             | 63 (50–75)                        | 80 (72–86)                        | < 0.001        | 80 (72–86)                           | 80 (72–86)                        | 1.000      |  |
| Race (non-Caucasian), n (%)                           | 1603 (22.8)                       | 59 (8.7)                          | < 0.001        | 49 (7.7)                             | 59 (9.2)                          | 0.315      |  |
| Body mass index (kg/m²),<br>median (IQR)              | 27.1 (24.2–30.7)                  | 27.7 (25.0–31.2)                  | 0.011          | 26.9 (24.5–30.5)                     | 26.7 (25.0–31.3)                  | 0.168      |  |
| Baseline comorbidities, $n$ (%)                       |                                   |                                   |                |                                      |                                   |            |  |
| Hypertension  | 3176 (45.2)                       | 542 (80.3)                        | < 0.001        | 433 (68.0)                           | 516 (80.9)                        | 0.053      |  |
| Diabetes mellitus                                     | 1257 (17.9)                       | 198 (29.3)                        | < 0.001        | 168 (27.0)                           | 190 (30.3)                        | 0.198      |  |
| Heart failure   | 128 (1.8)                         | 46 (6.8)                          | < 0.001        | 35 (5.5)                             | 39 (6.1)                          | 0.632      |  |
| Stroke/TIA  | 439 (6.3)                         | 131 (19.4)                        | < 0.001        | 92 (14.4)                            | 122 (19.1)                        | 0.437      |  |
| Chronic kidney disease                                | 369 (5.3)                         | 115 (17.0)                        | < 0.001        | 69 (11.0)                            | 109 (17.0)                        | 0.487      |  |
| Vascular disease*                                     | 543 (7.7)                         | 102 (15.1)                        | < 0.001        | 93 (14.6)                            | 88 (13.8)                         | 0.688      |  |
| Hypercholesterolemia                                  | 2096 (29.8)                       | 344 (51.0)                        | < 0.001        | 288 (45.1)                           | 326 (51.1)                        | 0.085      |  |
| Current smoking habit                                 | 407 (5.8)                         | 35 (5.2)                          | 0.243          | 21 (3.3)                             | 31 (4.9)                          | 0.071      |  |
| COPD/SAHS   | 419 (6.0)                         | 104 (15.4)                        | < 0.001        | 81 (12.7)                            | 84 (13.2)                         | 0.802      |  |
| History of malignant disease                          | 822 (11.7)                        | 139 (20.6)                        | < 0.001        | 129 (20.2)                           | 129 (20.2)                        | 1.000      |  |
| Liver disease   | 238 (3.4)                         | 33 (4.9)                          | 0.001          | 30 (4.7)                             | 31 (4.9)                          | 0.795      |  |
| Dysthyroidism   | 334 (4.8)                         | 40 (5.9)                          | 0.177          | 37 (5.8)                             | 40 (6.3)                          | 0.724      |  |
| Any dependency level                                  | 819 (11.7)                        | 210 (31.1)                        | < 0.001        | 177 (28.2)                           | 194 (30.6)                        | 0.365      |  |
| Concomitant treatment at adm                          | nission, $n$ (%)                  |                                   |                |                                      |                                   |            |  |
| Beta-blockers   | 865 (12.3)                        | 328 (48.6)                        | < 0.001        | 132 (20.7)                           | 311 (48.7)                        | < 0.001    |  |
| ACEi/ARBs   | 2320 (33.0)                       | 369 (54.7)                        | < 0.001        | 311 (48.7)                           | 350 (54.9)                        | 0.086      |  |
| Antiplatelet therapy                                  | 1229 (17.5)                       | 74 (11.0)                         | < 0.001        | 199 (31.2)                           | 72 (11.3)                         | < 0.001    |  |
| Laboratory parameters at admissi                      |                                   | •                                 |                | , ,                                  | , ,                               |            |  |
| Creatinine (mg/dL), median (IQR)                      | 0.90 (0.72–1.17)                  | 1.19 (0.90–1.64)                  | <0.001         | 0.98 (0.78-1.42)                     | 1.20 (0.88–1.66)                  | <0.001     |  |
| Hemoglobin (g/dL), median (IQR)                       | 14.0 (12.0–15.0)                  | 13.0 (11.0–14.0)                  | <0.001         | 13.0 (12.0–15.0)                     | 13.0 (11.0–14.0)                  | < 0.001    |  |
| Platelet count (×10 <sup>9</sup> /L),<br>median (IQR) | 203.0 (155.0–265.8)               | 179.0<br>(136.0–240.0)            | <0.001         | 195.0 (145.0–260.8)                  | 181.0 (138.0-241.0)               | 0.019      |  |
| Elevated D-dimer, $n$ (%)                             | 3921 (55.8)                       | 358 (53.0)                        | 0.036          | 425 (66.6)                           | 342 (53.6)                        | < 0.001    |  |
| Elevated procalcitonin, $n$ (%)                       | 1048 (14.9)                       | 126 (18.7)                        | 0.001          | 103 (16.1)                           | 123 (19.3)                        | 0.299      |  |
| Elevated C-reactive protein, <i>n</i> (%)             | 5841 (83.2)                       | 608 (90.1)                        | <0.001         | 566 (88.7)                           | 576 (90.3)                        | 0.657      |  |
| Elevated troponins, $n$ (%)                           | 527 (7.5)                         | 107 (15.9)                        | < 0.001        | 54 (8.5)                             | 100 (15.7)                        | < 0.001    |  |
| Elevated transaminases, $n$ (%)                       | 2598 (37.0)                       | 220 (32.6)                        | 0.009          | 216 (33.9)                           | 210 (32.9)                        | 0.023      |  |
| Elevated ferritin, $n$ (%)                            | 2306 (32.8)                       | 207 (30.7)                        | 0.424          | 198 (31.0)                           | 198 (31.0)                        | 1.000      |  |
| Elevated lactate dehydrogenase, <i>n</i> (%)          | 4414 (62.9)                       | 464 (68.7)                        | 0.005          | 427 (66.9)                           | 440 (69.0)                        | 0.466      |  |

ACEi, angiotensin-converting enzyme inhibitors; ARBs, angiotensin II receptor blockers; IQR, interquartile range; TIA, transient ischemic attack; COPD/SAHS, chronic obstructive pulmonary disease/sleep apnea-hypopnea syndrome. \*Coronary artery disease and/or peripheral artery disease.

OAC previously (11.4% vs. 4.1%, p < 0.001; OR 3.04, 95% CI 1.92–4.83), without differences in the risk of thromboembolic events (3.3% vs. 3.0%, p = 0.748; OR 1.11, 95% CI 0.59–2.08). There was increased mortality in patients who were on previous OAC therapy in comparison to patients who were not on previous OAC (38.1% vs. 30.9%, p = 0.007), with a significantly higher risk of death (HR 1.22, 95% CI 1.01–1.47), also confirmed by the Kaplan–Meier analysis (log-rank p value = 0.041) (Figure 1). There were no differences between patients on prior or non-prior OAC therapy regarding specific causes of death (cardiovascular death: 2.6% vs. 2.5%; respiratory-related: 59.7% vs. 62.9%;

SIRS-related: 4.9% vs. 3.6%; sepsis-related: 3.3% vs. 7.6%; other reasons or combined causes of death: 29.6% vs. 23.4%; p = 0.187).

3.2. Impact of OAC Type. In patients on prior OAC therapy, we observed significant differences regarding age and comorbid conditions between patients who were previously taking VKAs and those who were on prior DOAC therapy. We performed another PSM to balance these characteristics. This analysis demonstrated no differences in the remaining 464 subjects: 232 on VKAs and 232 on DOACs, as given in Table 3.

|   | Patients      | without prior OAC $(N = 638)$                | Patients      | s with prior OAC $(N = 638)$                 |                     | P       |
|---|---------------|--|---------------|--|---------------------|---------|
|   | N (%)         | Incidence per 100 patients-<br>days (95% CI) | N (%)         | Incidence per 100 patients-<br>days (95% CI) | OR (95% CI)         | value   |
| Intensive care unit admission           | 40 (6.3)      | 0.52 (0.37-0.71)                             | 44 (6.9)      | 0.58 (0.42-0.77)                             | 1.11<br>(0.71–1.73) | 0.652   |
| Renal failure                           | 151<br>(23.7) | 1.97 (1.67–2.31)                             | 212<br>(33.2) | 2.77 (2.41–3.17)                             | 1.61<br>(1.26–2.06) | 0.001   |
| Respiratory insufficiency               | 413<br>(64.7) | 5.39 (4.89–5.94)                             | 429<br>(67.2) | 5.60 (5.09–6.16)                             | 1.09<br>(0.86–1.38) | 0.280   |
| Upper respiratory tract infection       | 90<br>(14.1)  | 1.18 (0.95–1.44)                             | 89<br>(13.9)  | 1.16 (0.93–1.43)                             | 0.99<br>(0.72–1.35) | 0.987   |
| Heart failure                           | 65<br>(10.2)  | 0.85 (0.66–1.08)                             | 115<br>(18.0) | 1.50 (1.24–1.80)                             | 1.93<br>(1.39–2.68) | < 0.001 |
| Sepsis                                  | 77<br>(12.1)  | 1.01 (0.79–1.26)                             | 96<br>(15.0)  | 1.25 (1.02–1.53)                             | 1.29<br>(0.93–1.78) | 0.299   |
| Systemic inflammatory response syndrome | 129<br>(20.2) | 1.69 (1.41–2.00)                             | 181<br>(28.4) | 2.36 (2.03–2.73)                             | 1.55<br>(1.20–2.02) | 0.003   |
| All-cause mortality                     | 197<br>(30.9) | 2.57 (2.23–2.96)                             | 243<br>(38.1) | 3.17 (2.79–3.60)                             | 1.38<br>(1.09–1.74) | 0.007   |
| Any thrombotic/<br>thromboembolic event | 19 (3.0)      | 0.25 (0.15–0.39)                             | 21 (3.3)      | 0.27 (0.17–0.42)                             | 1.11<br>(0.59–2.08) | 0.748   |
| Any clinically relevant bleeding        | 26 (4.1)      | 0.34 (0.22-0.50)                             | 73<br>(11.4)  | 0.96 (0.75–1.20)                             | 3.04<br>(1.92–4.83) | < 0.001 |

TABLE 2: Clinical outcomes during hospitalization after propensity score matching.

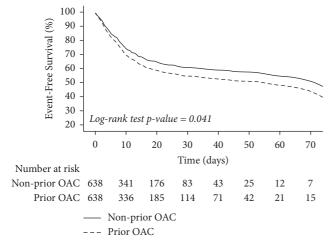


FIGURE 1: Comparison of survival curves between patients on prior OAC and nonprior OAC. Solid line, nonprior OAC; dashed line, prior OAC.

With this matched cohort, the rate of ICU admission between patients on VKAs (14, 6.0%) and patients on DOACs (14, 6.0%) was similar (p = 1.000). There were no differences in terms of respiratory insufficiency, heart failure, development of renal failure, upper respiratory tract infection, sepsis, or SIRS, in patients on prior VKAs or DOACs (all p > 0.005) (Supplementary Table 1).

No significant differences in the incidences of clinically relevant bleeding or thromboembolic events were observed in patients previously taking VKAs compared to DOACs (1.01 vs. 0.83 per 100 patient-days (p = 0.458) and 0.29 vs. 0.11 per 100 patient-days (p = 0.127), respectively) (Supplementary Table 1). Mortality rate between previous VKA and DOAC users was also similar (37.9% vs. 39.7%,

p = 0.703), with a non-significant difference in mortality risk amongst previous VKA users (HR 0.84, 95% CI 0.62–1.12; p = 0.233) (Figure 2).

3.3. Anticoagulation Management during Hospitalization. Regarding anticoagulation during hospitalization in the PSM cohort of previous vs. no previous OAC, most patients not taking OAC previously were prescribed heparin (79.9%, 382/478) and 19.2% (92/478) did not receive anticoagulation. In patients who were previously on OAC, 65.5% (330/504) were switched to heparin, 25% (126/504) continued on OAC, and 9.5% (48/504) did not receive any anticoagulation therapy. These proportions were significantly different (p<0.001).

In the PSM cohort of previous VKA vs. DOAC, most patients under either therapy received heparin during hospitalization, without differences between drug families (116 vs. 114; p = 0.567). Those patients who were maintained on OAC during admission were predominately treated with the same OAC that they were before (95.5% for previous VKAs users and 93.2% for previous DOACs users; p < 0.001).

## 4. Discussion

In this study of the HOPE COVID-19 Registry, including a large cohort of patients hospitalized for COVID-19, we demonstrate that the risk of in-hospital worse clinical outcomes was higher in patients with prior OAC therapy, even after adjustment by a PSM. Importantly, this study population showed a 22% higher risk of mortality and bleeding, without significant differences in the prognosis with regard to the particular anticoagulant drug, i.e., VKAs versus DOACs.

Table 3: Comparison of clinical characteristics of patients on VKA or DOAC prior admission after propensity score matching.

| · ————————————————————————————————————             |                                 |                                  |         |
|--|---------------------------------|----------------------------------|---------|
|  | Patients on prior VKA $N = 232$ | Patients on prior DOAC $N = 232$ | P value |
| Demographic  |                                 |                                  |         |
| Male sex, $n$ (%)                                  | 139 (59.9)                      | 127 (54.7)                       | 0.260   |
| Age (years), median (IQR)                          | 80 (72–87)                      | 81 (73–86)                       | 0.575   |
| Body mass index (kg/m <sup>2</sup> ), median (IQR) | 28.0 (25.1–31.6)                | 27.3 (24.3–31.0)                 | 0.445   |
| Baseline comorbidities, n (%)                      |                                 |                                  |         |
| Hypertension                                       | 189 (81.5)                      | 178 (76.7)                       | 0.209   |
| Diabetes mellitus                                  | 61 (26.3)                       | 74 (31.9)                        | 0.184   |
| Heart failure                                      | 15 (6.5)                        | 14 (6.0)                         | 0.848   |
| Stroke/TIA   | 42 (18.1)                       | 47 (20.3)                        | 0.555   |
| Chronic kidney disease                             | 31 (13.4)                       | 31 (13.4)                        | 1.000   |
| Vascular disease*                                  | 33 (14.2)                       | 34 (14.7)                        | 0.895   |
| Hypercholesterolemia                               | 112 (48.3)                      | 113 (48.7)                       | 0.926   |
| Current smoking habit                              | 13 (5.6)                        | 9 (3.9)                          | 0.143   |
| COPD/SAHS  | 30 (12.9)                       | 29 (12.5)                        | 0.889   |
| History of malignant disease                       | 38 (16.4)                       | 38 (16.4)                        | 1.000   |
| Dysthyroidism                                      | 18 (7.8)                        | 17 (7.3)                         | 0.860   |
| Any dependency level                               | 68 (29.3)                       | 72 (31.0)                        | 0.686   |
| Concomitant treatment at admission, n (%)          |                                 |                                  |         |
| Beta-blockers                                      | 103 (44.4)                      | 124 (53.4)                       | 0.042   |
| ACEi/ARBs  | 130 (56.0)                      | 123 (53.0)                       | 0.703   |
| Antiplatelet therapy                               | 24 (10.3)                       | 27 (11.6)                        | 0.656   |
| Laboratory parameters at admission                 |                                 |                                  |         |
| Creatinine (mg/dL), median (IQR)                   | 1.19 (0.87–1.56)                | 1.13 (0.87–1.56)                 | 0.628   |
| Hemoglobin (g/dL), median (IQR)                    | 13.0 (12.0–14.0)                | 13.0 (11.0–14.0)                 | 0.853   |
| Platelet count (×10 <sup>9</sup> /L), median (IQR) | 178.0 (138.0-244.8)             | 176.0 (134.0-233.0)              | 0.432   |
| Elevated D-dimer, n (%)                            | 121 (52.2)                      | 118 (50.9)                       | 0.954   |
| Elevated procalcitonin, $n$ (%)                    | 45 (19.4)                       | 40 (17.2)                        | 0.795   |
| Elevated C-reactive protein, n (%)                 | 209 (90.1)                      | 210 (90.5)                       | 0.984   |
| Elevated troponins, $n$ (%)                        | 35 (15.1)                       | 31 (13.4)                        | 0.711   |
| Elevated transaminases, n (%)                      | 86 (37.1)                       | 66 (28.4)                        | 0.138   |
| Elevated ferritin, n (%)                           | 75 (32.3)                       | 61 (26.3)                        | 0.344   |
| Elevated lactate dehydrogenase, n (%)              | 167 (72.0)                      | 150 (64.7)                       | 0.182   |

ACEi, angiotensin-converting enzyme inhibitors; ARBs, angiotensin II receptor blockers; IQR, interquartile range; TIA, transient ischemic attack; COPD/SAHS, chronic obstructive pulmonary disease/sleep apnea-hypopnea syndrome. \*Coronary artery disease and/or peripheral artery disease.

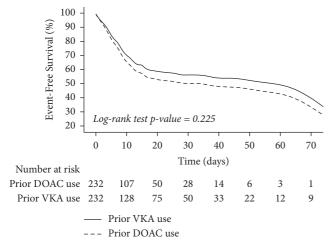


FIGURE 2: Comparison of survival curves between patients on prior VKAs or DOACs. Solid line, prior VKA use; dashed line, prior DOAC use.

Anticoagulation in the context of COVID-19 has been widely debated, with some studies showing that prophylactic and therapeutic anticoagulation might reduce mortality in

hospitalized COVID-19 patients [14]. Patients who received high-intensity prophylactic anticoagulation have a downtrend in D-dimer levels and improved 30-day mortality [15]. Indeed, a cross-sectional analysis showed that anticoagulation use was associated with delayed death, both at prophylactic (HR 0.29, 95% CI 0.15-0.58; p < 0.001) and therapeutic doses (HR 0.15, 95% CI 0.07-0.32; p < 0.001), compared with no anticoagulation [16]. In contrast, one retrospective analysis of hospitalized COVID-19 patients suggested that therapeutic anticoagulation provided no mortality benefit over thromboprophylaxis, independently of comorbidities or disease severity, and more adverse events were observed with therapeutic anticoagulation [17]. On the other hand, a large cohort study simulating an intention-totreat clinical trial analyzed the effect on mortality of anticoagulation therapy chosen in the first 48 hours of hospitalization showing that patients with moderate or severe illness benefited from anticoagulation and that apixaban had a similar efficacy to enoxaparin in decreasing mortality amongst these patients [18]. Another study showed that hospitalized COVID-19 patients suffered from more bleeding events in those on low-molecular-weight heparin (LMWH) compared to DOACs, and DOAC use may be associated with better survival and lower invasive respiratory support rate compared to LMWH [19]. Given such contradictory observations, there are a number of studies and clinical trials with the aim to assess the role of antithrombotic therapy on mortality and thromboembolic events [20–26].

OAC management in the setting of the COVID-19 pandemic is even more complex. VKAs have the limitation of routine monitoring and dose adjusting for maintaining good quality of anticoagulation. One study demonstrated a significant increase in high INR results during the COVID-19 pandemic, the majority of them after the introduction of a lockdown [27]. In addition, patients on VKA hospitalized with SARS-CoV-2 showed greater instability of PT INR due to the inflammatory state and the interactions with numerous drugs. On the other hand, DOACs avoid some of the VKA limitations, but DOAC-treated patients have an increase in DOAC plasma levels when treated with antiviral drugs for COVID-19 [28]. For these reasons, some groups have suggested replacing OAC with parenteral heparin during hospitalization to avoid the risk of over/under treatment [29, 30]. Nevertheless, other authors suggested that the indications for antiplatelet/anticoagulant use (prevention, prophylaxis, and therapy) should be guided by the clinical context and the COVID-19 severity and not based on a systematic change per protocol in all patients [31, 32].

Nevertheless, most of the evidence focused on hospitalized patients, whereas the potential effect of chronic antithrombotic therapies in COVID-19 progression and prognosis remains uncertain. The pathophysiology underlying the prothrombotic state elicited by SARS-CoV-2 outlines possible protective mechanisms of antithrombotic therapy for this viral disease. In particular, aspirin and FXa inhibitors have been postulated as potential prophylactic and therapeutic treatment for high-risk patients with COVID-19 [31, 33]. Unsurprisingly, ongoing clinical trials are comparing the effectiveness and safety of apixaban, aspirin, and rivaroxaban versus heparin, placebo, and other therapies on progression, arterial, and venous thromboembolic events and mortality in patients with COVID-19 not yet admitted to hospital [34–36].

To date, data in this particular context are scarce and limited, with positive, negative, and neutral results. One small study in an Italian cohort of elderly patients with COVID-19 concluded that chronic DOAC intake was an independent parameter associated with a decreased mortality risk (HR 0.38, 95% CI 0.17–0.58; p = 0.010) [8]. Similarly, another study in Italy showed that elderly patients with COVID-19 on chronic OAC treatment for atrial fibrillation had lower all-cause mortality rate ratio compared to their PSM non-anticoagulated counterpart [37]. However, Sivaloganathan et al. demonstrated that patients taking antithrombotic therapy (anticoagulant or antiplatelet agents) at the time of infection with COVID-19 did not have a significantly different mortality risk to those patients not taking these drugs [38]. Another study showed no difference in the risk of acute respiratory distress syndrome at admission or death during hospitalization between COVID-19

patients treated or not with antiplatelets or anticoagulants preadmission [39]. Likewise, anticoagulant use pre-COVID-19 diagnosis was not associated with a decreased risk for allcause mortality, mechanical ventilation, or hospital admission in a study from the New York City health system, suggesting that previous anticoagulant use did not protect against development of severe COVID-19 [40]. Also, our preliminary analysis of the HOPE COVID-19 Registry observed a significantly lower survival and higher mortality risk in COVID-19 patients on OAC therapy at hospital admission compared to patients without prior OAC at admission [9]. More recently, a nationwide register-based cohort study in Sweden demonstrated that ongoing DOAC use at the time of SARS-CoV-2 infection was not associated with reduced risk of COVID-19 hospitalization or the composite of ICU admission or death due to COVID-19, indicating that the evidence for DOACs in this context is controversial [41].

Our results in the present study confirm our previous observation about the higher risk of mortality in COVID-19 patients with OAC therapy before hospital admission. Of note, our analysis is balanced by PSM, and there were no differences regarding admission to the ICU in patients on prior and no prior OAC. However, not only mortality was increased in patients with prior OAC therapy but also other clinical outcomes. Despite an appropriate PSM adjusting for comorbidities, postadmission serum creatinine as a marker of renal function (and injury) and troponins as markers of myocardial damage were higher in these patients, thereby showing increased rates of heart failure and renal failure during hospitalization. This reinforces the hypothesis that OAC-treated patients are particularly vulnerable and still have an inherent proinflammatory state.

4.1. Limitations. We should acknowledge some limitations in relation to this study. First, the constraints of an observational registry study of this design need to be considered. Second, the HOPE Registry only included patients from the first wave of the pandemic, and therefore, our results probably require further investigation during the subsequent waves. A bias inherent in the first wave neither can be excluded, given that hospitalization services throughout the world were overwhelmed. We also recognize that including several different indications for OAC may hinder and dissipate the specific effect that each indication has, since patients presented different risk profiles.

In addition, the indication for OAC as a whole may have some influence on the risk of outcomes, but comparing patients with prior OAC and no OAC was actually our aim, so we cannot adjust for specific indications for OAC but only for demographics data and other comorbidities at baseline. The absence of INR determinations (and therefore the time in therapeutic range (TTR)) in VKA-treated patients is also a limitation since the efficacy and safety of VKA depend on the quality of anticoagulant control, as reflected by the average TTR of INRs 2.0-3.0, and therefore may be related to the risk of worse outcomes. In addition, the type of DOAC was unknown in some cases, and this prevented us for analyzing

drug types as separate. Finally, although this cohort was collected in a prospective manner, the results reported in this study are based on a post hoc analysis and should be regarded as hypothesis-generating.

## 5. Conclusion

Hospitalized COVID-19 patients on prior OAC therapy had a higher risk of mortality and worse clinical outcomes compared to patients without prior OAC therapy, even after adjusting for comorbidities using PSM. There were no differences in clinical outcomes in patients previously taking VKAs versus DOACs.

## **Data Availability**

The data used to support this study cannot be shared for ethical/privacy reasons.

#### **Disclosure**

This nonprofit institution had no role in the study design; in the collection, analysis, and interpretation of data; in the writing of the report; nor in the decision to submit the article for publication.

## **Conflicts of Interest**

JMR-C has received an unrestricted educational grant from Bristol-Myers Squibb-Pfizer Alliance (BMS protocol number: CV185-805) for this study. GYHL is a consultant and speaker for BMS/Pfizer, Boehringer Ingelheim, and Daiichi-Sankyo. Other authors declare that they have no conflicts of interest.

## **Authors' Contributions**

JMR-C performed statistical analyses and drafted the study. IJN-G, GYHL, and FM supervised the research and draft the article. AU, MCV-L, AG, AFC-M, BAG, EA, JFG-P, CC, BC, GF, IF-R, JS-C, JH, MGA, MP, RR, EC, VMB-M, SR-R, FS, RB, LS, IE-B, ALM, AF-O, VE, and CM collaborated in data collection and made a critical revision of the article. IJN-G, is the guarantor of the article and takes responsibility for the integrity of the work as a whole, from inception to published article.

## Acknowledgments

The authors thank Cardiovascular Excellence SL for their essential support in the database and registry web page, all HOPE COVID-19 researchers (nonconditional grant (Fundación Interhospitalaria para la investigación Cardiovascular, FIC. Madrid, Spain)). This work was also supported by the Spanish Ministry of Economy, Industry, and Competitiveness, through the Instituto de Salud Carlos III after independent peer review (PI21/00607 cofinanced by the European Regional Development Fund and group CB16/11/00385 from CIBERCV).

## **Supplementary Materials**

On behalf of HOPE COVID-19 investigators, full list of investigators is shown in the Supplementary Material (HOPE participating hospitals, principal investigators, HOPE participating hospitals, coprincipal investigators, scientific committee, and list of collaborators). Supplementary Table 1. Clinical outcomes during hospitalization after propensity score matching in patients on prior oral anticoagulation therapy. (Supplementary Materials)

#### References

- [1] F. Violi, D. Pastori, R. Cangemi, P. Pignatelli, and L. Loffredo, "Hypercoagulation and antithrombotic treatment in coronavirus 2019: a new challenge," *Thrombosis & Haemostasis*, vol. 120, no. 6, pp. 949–956, 2020.
- [2] F. Del Sole, A. Farcomeni, L. Loffredo et al., "Features of severe COVID-19: a systematic review and meta-analysis," *European Journal of Clinical Investigation*, vol. 50, no. 10, Article ID e13378, 2020.
- [3] G. Piazza, U. Campia, S. Hurwitz et al., "Registry of arterial and venous thromboembolic complications in patients with COVID-19," *Journal of the American College of Cardiology*, vol. 76, no. 18, pp. 2060–2072, 2020.
- [4] B. Bikdeli, M. V. Madhavan, D. Jimenez et al., "COVID-19 and thrombotic or thromboembolic disease: implications for prevention, antithrombotic therapy, and follow-up," *Journal of the American College of Cardiology*, vol. 75, no. 23, pp. 2950–2973, 2020.
- [5] R. D. McBane, V. D. Torres Roldan, A. S. Niven et al., "Anticoagulation in COVID-19: a systematic review, metaanalysis, and rapid guidance from mayo clinic," *Mayo Clinic Proceedings*, vol. 95, no. 11, pp. 2467–2486, 2020.
- [6] I. Paranjpe, V. Fuster, A. Lala et al., "Association of treatment dose anticoagulation with in-hospital survival among hospitalized patients with COVID-19," *Journal of the American College of Cardiology*, vol. 76, no. 1, pp. 122–124, 2020.
- [7] N. Tang, H. Bai, X. Chen, J. Gong, D. Li, and Z. Sun, "Anticoagulant treatment is associated with decreased mortality in severe coronavirus disease 2019 patients with coagulopathy," *Journal of Thrombosis and Haemostasis*, vol. 18, no. 5, pp. 1094–1099, 2020.
- [8] R. Rossi, F. Coppi, M. Talarico, and G. Boriani, "Protective role of chronic treatment with direct oral anticoagulants in elderly patients affected by interstitial pneumonia in COVID-19 era," *European Journal of Internal Medicine*, vol. 77, pp. 158–160, 2020.
- [9] J. M. Rivera-Caravaca, I. J. Núñez-Gil, D. Vivas et al., "Clinical profile and prognosis in patients on oral anticoagulation before admission for COVID-19," European Journal of Clinical Investigation, vol. 51, no. 1, Article ID e13436, 2020.
- [10] I. J. Núñez-Gil, V. Estrada, C. Fernández-Pérez et al., "Health outcome predictive evaluation for COVID-19 international registry (HOPE COVID-19), rationale and design," *Con*temporary Clinical Trials Communications, vol. 20, Article ID 100654, 2020.
- [11] International COVID-19 Clinical Evaluation Registry and (HOPE COVID-19), "ClinicalTrials.gov Identifier: NCT04334291," 2020, https://clinicaltrials.gov/ct2/show/ NCT04334291.

- [12] P. C. Austin, "Balance diagnostics for comparing the distribution of baseline covariates between treatment groups in propensity-score matched samples," *Statistics in Medicine*, vol. 28, no. 25, pp. 3083–3107, 2009.
- [13] P. C. Austin, "Optimal caliper widths for propensity-score matching when estimating differences in means and differences in proportions in observational studies," *Pharmaceutical Statistics*, vol. 10, no. 2, pp. 150–161, 2011.
- [14] A. M. Kamel, M. Sobhy, N. Magdy, N. Sabry, and S. Farid, "Anticoagulation outcomes in hospitalized Covid-19 patients: a systematic review and meta-analysis of case-control and cohort studies," *Reviews in Medical Virology*, vol. 31, no. 3, Article ID e2180, 2020.
- [15] A. Hsu, Y. Liu, A. S. Zayac, A. J. Olszewski, and J. L. Reagan, "Intensity of anticoagulation and survival in patients hospitalized with COVID-19 pneumonia," *Thrombosis Research*, vol. 196, pp. 375–378, 2020.
- [16] F. Ionescu, G. Grasso-Knight, E. Castillo et al., "Therapeutic anticoagulation delays death in COVID-19 patients: crosssectional analysis of a prospective cohort," *TH Open*, vol. 04, no. 03, pp. e263–e270, 2020.
- [17] L. Lynn, J. A. Reyes, K. Hawkins et al., "The effect of anticoagulation on clinical outcomes in novel Coronavirus (COVID-19) pneumonia in a USA cohort," *Thrombosis Re*search, vol. 197, pp. 65–68, 2021.
- [18] H. H. Billett, M. Reyes-Gil, J. Szymanski et al., "Anti-coagulation in COVID-19: effect of enoxaparin, heparin, and apixaban on mortality," *Thrombosis and haemostasis*, vol. 120, no. 12, pp. 1691–1699, 2020.
- [19] G. N. Nadkarni, A. Lala, E. Bagiella et al., "Anticoagulation, bleeding, mortality, and pathology in hospitalized patients with COVID-19," *Journal of the American College of Cardiology*, vol. 76, no. 16, pp. 1815–1826, 2020.
- [20] Bill-&-Melinda-Gates-Medical-Research-Institute, "A trial to evaluate safety and efficacy of rivaroxaban (COVID-19)," 2020, https://clinicaltrials.gov/ct2/show/NCT04504032.
- [21] U. Landmesser, "Effect of anticoagulation therapy on clinical outcomes in COVID-19 (COVID-prevent)," 2020, https:// clinicaltrials.gov/ct2/show/NCT04416048.
- [22] Fundació-Institut-de-Recerca-de-l'Hospital-de-la-Santa-Creu-i-Sant-Pau, "Evolution of COVID-19 in anticoagulated or antiaggregated patients (CORONA study) (CORONA)," 2020, https://clinicaltrials.gov/ct2/show/NCT04518735.
- [23] P. Horby, "Randomised evaluation of COVID-19 therapy (RECOVERY)," 2020, https://clinicaltrials.gov/ct2/show/ NCT04381936.
- [24] J. Hochman, "Anti-thrombotics for adults hospitalized with COVID-19 (ACTIV-4)," 2020, https://clinicaltrials.gov/ct2/ show/NCT04505774.
- [25] S. Webb, C. McArthur, M. Bonten, L. Derde, M. Marshall, and A. Derek, "Randomized, embedded, multifactorial adaptive platform trial for community-acquired pneumonia (REMAP-CAP)," 2020, https://clinicaltrials.gov/ct2/show/ NCT02735707.
- [26] P. R. Lawler, E. C. Goligher, and R. Zarychanski, "Antithrombotic therapy to ameliorate complications of COVID-19 (ATTACC)," 2020, https://clinicaltrials.gov/ct2/show/ NCT04372589.
- [27] V. Speed, R. K. Patel, R. Byrne, L. N. Roberts, and R. Arya, "A perfect storm: root cause analysis of supra-therapeutic anticoagulation with vitamin K antagonists during the COVID-19 pandemic," *Thrombosis Research*, vol. 192, pp. 73-74, 2020.
- [28] S. Testa, P. Prandoni, O. Paoletti et al., "Direct oral anticoagulant plasma levels' striking increase in severe COVID-19

- respiratory syndrome patients treated with antiviral agents: the cremona experience," *Journal of Thrombosis and Haemostasis*, vol. 18, no. 6, pp. 1320–1323, 2020.
- [29] S. Testa, O. Paoletti, M. Giorgi-Pierfranceschi, and A. Pan, "Switch from oral anticoagulants to parenteral heparin in SARS-CoV-2 hospitalized patients," *Internal and Emergency Medicine*, vol. 15, no. 5, pp. 751–753, 2020.
- [30] G. T. Gerotziafas, P. Van Dreden, M. P. Colgan et al., "Guidance for the management of patients with vascular disease or cardiovascular risk factors and COVID-19: position paper from VAS-European independent foundation in angiology/vascular medicine," *Sang Thrombose Vaisseaux*, vol. 32, no. 6, pp. 241–259, 2020.
- [31] C. Godino, A. Scotti, N. Maugeri et al., "Antithrombotic therapy in patients with COVID-19?-rationale and evidence," *International Journal of Cardiology*, vol. 324, pp. 261–266, 2020.
- [32] D. Vivas, V. Roldán, M. A. Esteve-Pastor et al., "Recommendations on antithrombotic treatment during the COVID-19 pandemic. position statement of the working group on cardiovascular thrombosis of the Spanish society of cardiology," Revista Espanola de Cardiologia, vol. 73, no. 9, pp. 749–757, 2020.
- [33] G. H. Frydman, M. B. Streiff, J. M. Connors, and G. Piazza, "The potential role of coagulation factor Xa in the pathophysiology of COVID-19: a role for anticoagulants as multimodal therapeutic agents," *TH Open*, vol. 4, no. 4, pp. e288–e299, 2020.
- [34] F. Sciurba, "COVID-19 positive outpatient thrombosis prevention in adults aged 40-80 2020," 2020, https://clinicaltrials.gov/ct2/show/NCT04498273.
- [35] R. Whitlock, E. Belley-Cote, and J. Eikelboom, "Anti-coronavirus therapies to prevent progression of coronavirus disease 2019 (COVID-19) trial (ACTCOVID19) 2020," 2020, https://clinicaltrials.gov/ct2/show/NCT04324463.
- [36] Janssen-Research-&-Development-L, "A study of rivaroxaban to reduce the risk of major venous and arterial thrombotic events, hospitalization and death in medically ill outpatients with acute, symptomatic coronavirus disease 2019 (COVID-19) infection (PREVENT-HD)," 2020, https://clinicaltrials. gov/ct2/show/NCT04508023.
- [37] G. Denas, N. Gennaro, E. Ferroni et al., "Reduction in all-cause mortality in COVID-19 patients on chronic oral anticoagulation: a population-based propensity score matched study," *International Journal of Cardiology*, vol. 329, pp. 266–269, 2020.
- [38] H. Sivaloganathan, E. E. Ladikou, and T. Chevassut, "COVID-19 mortality in patients on anticoagulants and antiplatelet agents," *British Journal of Haematology*, vol. 190, no. 4, pp. e192–e5, 2020.
- [39] V. Russo, M. Di Maio, E. Attena et al., "Clinical impact of preadmission antithrombotic therapy in hospitalized patients with COVID-19: a multicenter observational study," *Phar*macological Research, vol. 159, Article ID 104965, 2020.
- [40] D. Tremblay, M. van Gerwen, M. Alsen et al., "Impact of anticoagulation prior to COVID-19 infection: a propensity score-matched cohort study," *Blood*, vol. 136, no. 1, pp. 144–147, 2020.
- [41] B. Flam, V. Wintzell, J. F. Ludvigsson, J. Mårtensson, and B. Pasternak, "Direct oral anticoagulant use and risk of severe COVID-19," *Journal of Internal Medicine*, vol. 289, no. 3, pp. 411–419, 2020.