

1 **STATISTICAL ASSESSMENT OF IN-PLANE MASONRY**
2 **PANELS USING LIMIT ANALYSIS WITH SLIDING**
3 **MECHANISM**

4 Alejandro Jiménez Ríos ¹, Marco Pingaro ¹, Emanuele Reccia ², Patrizia Trovalusci ¹

5 **ABSTRACT**

6 *Historical masonry structures have a great interest in civil engineering since they constitute a large part of the world's building heritage. In this paper the effects that different geometrical (panel ratio, block ratio, bond type) and mechanical (friction ratio) parameters have on the in-plane structural response of brick masonry panels are investigated. A discrete modelling approach, based on a *Limit Analysis, capable of reproducing sliding mechanisms*, formulation by one of the Authors have been adopted, enhanced and implemented. Results, in terms of collapse multipliers and collapse mechanisms, are presented and analysed following a systematic statistical approach. Statistically significant effects have been found for each factor considered. Furthermore, the statistical model adopted included non-linear terms that allowed to identify whether the effect of one parameter on the response depended on the level of any other parameters. Thus, it was observed that two-way factor interactions played an important role on the in-plane response of masonry panels. The panel ratio-friction ratio two-way factor interaction was the one with a more significant effect.*

19 **Keywords:** Limit Analysis, Friction, Masonry, Panels, No-tension contacts, Statistical
20 assessment.

21 **INTRODUCTION**

22 Masonry is a non-homogeneous material constituted by blocks (stone, bricks or
23 adobes) and joints (mortar or dry) (Lourenço 1998). The structural behaviour of ma-
24 sonry is affected by geometry, disposition and mechanical properties of its constituents,
25 as well as by the aspect ratios of the panels and of the blocks, the arrangement of
26 blocks and scale factor of the units and the whole panel (Baggio and Trovalusci 1993;

¹Dept. of Structural and Geotechnical Engineering., Sapienza University of Rome, Via Gramsci 53, 00197, Rome, Italy. ORCID: 0000-0003-4470-255X, 0000-0002-7037-8661, 0000-0001-7946-3590 E-mail: alejandro.jimenezrios@uniroma1.it, marco.pingaro@uniroma1.it, patrizia.trovalusci@uniroma1.it

²Dept. of Civil and Environmental Engineering and Architecture, University of Cagliari, Via Marengo 2, 09123, Cagliari, Italy. ORCID: 0000-0003-0499-4295 E-mail: emanuele.reccia@unica.it

27 Trovalusci and Masiani 1999; Trovalusci and Masiani 2003; Pau and Trovalusci 2012;
28 Baraldi et al. 2018).

29 Through history several masonry block arrangements, also known as bond types or
30 textures, have been implemented perhaps aiming at enhancing the response of masonry
31 structures. The work of (Huerta Fernández 2004), based on the experimental results
32 obtained by Rondelet in a series of masonry specimens on the XIX century, highlighted
33 the fact that masonry joints negatively affect masonry resistance.

34 It was observed by (Vasconcelos and Lourenço 2009) that shape and arrangement
35 of units in a masonry panel clearly affect ductility and energy dissipation capabilities
36 of a series of stone masonry panels tested under hysteretic dynamic loading. The influ-
37 ence of bond type on the structural response of masonry panels has been further studied
38 by (Drougkas et al. 2015) who have reported an extensive list of masonry mechani-
39 cal properties based on their bond type. In (Taguchi and Cuadra 2015) a comparison
40 between *english* and *flemish* masonry walls has been carried out demonstrating that a
41 larger volume of joints in a *flemish* wall would lead to a weaker structure in comparison
42 with an *english* bond masonry wall. Furthermore, (Shrestha et al. 2020) implemented
43 a micro-modelling numerical approach to simulate the structural behaviour of masonry
44 panels composed with different bond types and concluded that the elastic response of
45 the panels was not influenced by the bond type, only the ultimate failure load was
46 affected by this parameter.

47 In particular, the in-plane structural response of masonry panels is **not only** influ-
48 enced by the panel aspect ratio **but also by block shape factor as well on scale factor**
49 as pointed out by (Giuffrè 1990; Ponte et al. 2019). In (Anthoine et al. 1995) a strong
50 negative relation between masonry lateral resistance and panel ratio has been reported.
51 This observation has been further verified by several authors (Kikuchi et al. 2003;
52 Drysdale Robert and Hamid Ahmed 2005; Haach et al. 2011). After analysing a se-
53 ries of masonry panels composed of blocks with different aspect ratios, (Baraldi et al.
54 2018) concluded that the in-plane collapse mechanism, and consequently the panel re-
55 sistance, would be affected by the blocks geometry. Some other recent studies focus
56 on the effects of the internal geometry of masonry walls (shape, size and textures of
57 brick/blocks) (Pepe 2020; Pepe et al. 2020c).

58 Due to masonry heterogeneity and to **the** effect **that** geometrical parameters have
59 on its behaviour, the simulation of masonry structural response and assessment is a
60 complicated task. Over the past decades a variety of numerical approaches have been
61 proposed by several authors trying to reproduce masonry structural behaviour at differ-
62 ent scales and levels of detail. Among the more suitable strategies to capture masonry
63 structural response are the so called block-based models (BBM). **These** are discrete
64 models in which every masonry block is modelled along with a suitable formulation
65 to represent the inter-block interactions. The main advantage of BBM is that masonry
66 bond can be represented and accounted for. Furthermore, BBM are usually charac-
67 terized by relatively simple mechanical constitutive models **which require few input**
68 parameters (friction, cohesion, etc.) and provide clear results in terms of easily inter-
69 preted collapse mechanism and failure modes. A further description of BBM and al-
70 ternative approaches available to reproduce masonry structural behaviour can be found

71 in well-known review papers (Lourenço 2002; Roca et al. 2010; D’Altri et al. 2020).

72 Micromechanical models take into account the constituents, that is units and in-
73 terfaces, made of mortar if present, between elements are separately modelled and to
74 each part is assigned a properly calibrated constitutive law such as in (Lotfi and Shing
75 1994; Lourenço and Rots 1997; Oliveira and Lourenço 2004; Cecchi and Sab 2004;
76 Alfano and Sacco 2006). In macromechanical models (Del Piero 1989; Gambarotta
77 and Lagomarsino 1997; Roca et al. 2005), the heterogeneous medium is modelled as
78 a continuum and the constitutive behaviour is usually described through phenomeno-
79 logically based mathematical relations, also including damage or friction phenomena.
80 Finally, the multiscale models represent a very promising approach for the analysis
81 of masonry structures since they can accurately keep track of the the mechanical and
82 geometrical properties of the material at the microstructure scale with a reduced com-
83 putational cost if compared to a fully micromechanical model.

84 They are continuum models derived from finer descriptions and generally based
85 on the classical homogenization strategies, (Addessi et al. 2018; Addessi et al. 2016;
86 Greco et al. 2016; Greco et al. 2017) or on other coarse-graining strategies, based
87 on the so-called Cauchy rule and its generalizations (Trovalusci 2014; Capecchi et al.
88 2011), also allowing the derivation of generalized continua such as micropolar continua
89 able to properly account for scale effects, that in masonry materials are significant
90 (Masiani and Trovalusci 1996; Trovalusci and Masiani 1999; Trovalusci and Masiani
91 2003; Trovalusci and Pau 2014; Leonetti et al. 2018; Reccia et al. 2018).

92 The more general approach is the Discrete Element Method (DEM) (Cundall and
93 Strack 1979; Cundall and Hart 1992) originally developed for granular materials and
94 then successfully applied to masonry (Lemos 2007) and its combination with Finite
95 Elements (FEM/DEM)(Reccia et al. 2012; Smoljanović et al. 2013) in which blocks
96 could be represented as a deformable or as a rigid bodies. Another very effective
97 approach is the non-smooth contact dynamic method (NSCD) (Dubois et al. 2018;
98 Clementi et al. 2020), in which blocks are modelled as rigid interacting bodies. In this
99 framework, the so-called rigid block models (RBM) (Portioli et al. 2013; Angelillo
100 et al. 2018; Baraldi et al. 2020; Casolo 2004; Casolo 2009) are particularly fit for
101 historical masonries, where mortar is much more deformable than blocks and joints
102 thickness is negligible.

103 Within the context of BBM, Limit Analysis permits the evaluation of the ultimate
104 load capacity of the structure and its corresponding failure mechanism, requiring a
105 limited number of material parameters, overcoming the common difficulties of obtain-
106 ing reliable experimental data for historical masonry structures. Furthermore, Limit
107 Analysis is largely recognized as a very effective tool to estimate collapse load and
108 collapse mechanisms for masonry structures (Baggio and Trovalusci 1998; Baggio
109 and Trovalusci 2000; Ferris and Tin-Loi 2001; Milani 2011; Portioli et al. 2014; Mi-
110 lani and Taliercio 2016; Rossi et al. 2020; Cascini et al. 2020; Grillanda et al. 2019) or
111 masonry structures in presence of settlements (Landolfo et al. 2020; Pepe et al. 2020c;
112 Tiberti et al. 2020).

113 The basic hypothesis, introduced by (Heyman 1966; Heyman 1969), upon which
114 Limit Analysis would be applicable to masonry structures are: (a) sliding cannot occur,

115 (b) masonry has no tensile strength, (c) masonry has an infinite compressive strength
116 and (d) failure occurs under small displacements. Under hypothesis (a) masonry can be
117 considered as a material with associative flow rules for which the normality rule holds.
118 For these structures, however, hypothesis (a) is strongly limitative, as it reduces the
119 collapse of a masonry structure only to the occurrence of hinging mechanisms, while
120 in also sliding can be observed.

121 The first known pioneering contribute concerning the possibility to consider in the
122 study of the collapse of masonry structures the presence of friction is due to Coulomb
123 (Coulomb 1776). Coulomb recognises that in the presence of sliding mechanisms the
124 solution is not unique and that the collapse load can be limited by minor and ma-
125 jor bounds. The theorems of Limit Analysis currently formulated for materials with
126 finite resistance to friction, which are described as non-standard materials with non-
127 associative flow rules, for which the normality rule is not satisfied, still confirm this
128 finding. In particular, suitable lower and upper bounds for the collapse load of non-
129 standard materials, as systems with frictional interfaces, has been respectively identi-
130 fied with the collapse load of a standard material having an ideal yield surface with
131 outward normal directed as the vector representing the plastic flow (plastic potential),
132 and another standard material with the actual yield surface and ideal plastic flow di-
133 rected as the normal to this surface, as a material with dilatant interfaces (Drucker
134 1953; Radenkovic 1961).

135 In this work, the strategy for tackling the problem proposed in (Baggio and Trovalusci
136 2000) was followed. This strategy is based on the solution of a linear programming
137 problem (LP) obtained by replacing friction with dilatancy and it assumes an associa-
138 tive collapse mechanism for which the normality rule holds. Other strategies involve,
139 in order to satisfy the normality rule, the modification of the yield surface such as in
140 (Gilbert et al. 2006).

141 Starting from the work of (Baggio and Trovalusci 2000), a new version of the
142 ALMA code (Analisi Limite Murature Attritive) has been developed based on the Limit
143 Analysis (Pepe 2020; Pepe et al. 2020a; Pepe et al. 2020b; Pepe et al. 2021) namely
144 ALMA 2.0. The new version of ALMA, by the adoption of the recent coding language
145 PythonTM and the advantages of the novel MOSEK library (www.mosek.com) opti-
146 mization subroutine, overcomes the limitation in terms of the number of blocks with
147 respect to the original version (Baggio and Trovalusci 2000) and it has been improved
148 in order to take into account foundation settlement (Pepe et al. 2020c), cohesion be-
149 tween the joints and the effects of a retrofitting chain (Pepe 2020).

150 In this paper the effect of different geometrical and mechanical parameters in the
151 in-plane structural response of brick masonry panels using a Limit Analysis approach
152 capable of reproducing sliding mechanisms is presented. The main difference of this
153 work with respect to parametric analysis performed by other authors (Bustamante
154 2003; Casapulla and Argiento 2018), is the fact that a systematic statistical approach
155 has been implemented which has enabled the authors not only to identify, but also
156 to quantify, the effect on the response of the different factors studied. This approach
157 consisted in the application of a design of experiments (DOE) to a series of determi-
158 nistic Limit Analysis simulations. With the data obtained from the DOE, a metamodel

159 (Montgomery 2019) was created and, subsequently, the effect that the studied parameters have in the response was analysed on the metamodel. First the formulation of the
 160 Limit Analysis implemented is described in section "Adopted model", followed by the
 161 description of the systematic parametric analysis used in section "Design of experiments (DOE)". Then, the results obtained in terms of collapse multipliers and collapse
 162 mechanisms are shown and discussed in section "Results and discussion". Finally, in
 163 section "Conclusions", the main conclusions drawn from the analysis and discussion
 164 of the results are summarized.
 165
 166

167 ADOPTED MODEL

168 In this study the framework of the Limit Analysis has been adopted in accordance
 169 to the notation used in (Baggio and Trovalusci 1998; Baggio and Trovalusci 2000).
 170 The masonry structures have been described as a system of n rigid blocks and m joints
 171 unable to carry tension and resistant to sliding by friction, $f = \tan(\phi)$, where ϕ is
 172 the friction angle. Limited to the in-plane problems, the blocks can translate and ro-
 173 tate about the edges of the contact blocks (hinging) as well as slide along the joints
 174 as shown in Figure 1 in which a single block is depicted. It is important to notice
 175 that in case of sliding (Figure 1b) we assume dilatant behaviour, such that the block
 176 slides going up of the friction angle, ϕ . This assumption is explained in the following
 Subsection "Limit Analysis".

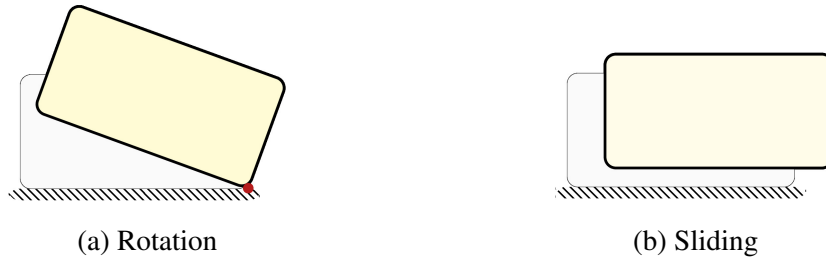


FIG. 1: Schematic representation of possible mechanisms for one-block structure

177

178 Limit Analysis (kinematic approach)

179 Let consider a system of n parallelepiped blocks in two-dimensional space with
 180 the orthonormal basis $e = \{e_1, e_2\}^T$. Over all single blocks the loads, applied in the
 181 respective centroid of mass of each i^{th} rigid block, is

$$\mathbf{f}^i = \mathbf{f}_0^i + \alpha \mathbf{f}_L^i, \quad \text{with } i = 1, \dots, n, \quad (1)$$

182 where $\mathbf{f}_0^i = \{f_{01}^i, f_{02}^i, m_0^i\}^T$ and $\mathbf{f}_L^i = \{f_{L1}^i, f_{L2}^i, m_L^i\}^T$ are the constant 'dead' and
 183 'live' generalized loads vectors, respectively. As usual in the Limit Analysis the load
 184 vector in Equation 1 is split into two parts in which live loads are proportional to the
 185 dead loads through a non-negative coefficient α , called load multiplier, as shown in
 186 Figure 2a. In both cases the vector \mathbf{f}_\star^i , with $\star = 0, L$ contains the two components of

187 the force f_{*j}^i with $j = 1, 2$ and, the moment m_*^i applied to i^{th} block. The global load
 188 vector \mathbf{f} is obtained collecting the single load vectors \mathbf{f}^i .

189 The vector $\mathbf{u}^i = \{u_1^i, u_2^i, \theta^i\}^T$, that contains the displacement components u_1, u_2 ,
 190 and the rotation θ (Figure 2b), represents the generalized displacement of the centre of
 191 the block. As previously, we define the collection of all single vector of generalized
 192 displacement in a global vector \mathbf{u} , which corresponds in the virtual work sense to the
 193 global load vector \mathbf{f} .

194 Over each j^{th} joint in that the contact surfaces between blocks is represented in
 195 the local system (Figure 2), we introduce the generalized stress and strain measures σ^j
 196 and ϵ , respectively (Figure 2b-2c).

197 The static variables $\sigma^j = \{N^j, T^j, M^j\}^T$, $j = 1, \dots, m$, are the internal forces
 198 acting at each j^{th} joint, where N^j , T^j and M^j are the components of the normal,
 199 shear force and the moment, respectively. The collection of the local generalized stress
 200 vector σ^j is the global vector of generalized stress σ .

201 The kinematic variables, or generalized strain, are the relative displacement rates at
 202 joints, that is normal displacement ξ^j , tangential displacement γ^j and rotation χ^j . For
 203 each joint $j = 1, \dots, m$ they are collected in the vector $\epsilon^j = \{\xi^j, \gamma^j, \chi^j\}^T$. The vector
 204 ϵ refers to the whole structure and corresponds in a virtual work sense to the vector of
 static variables σ .

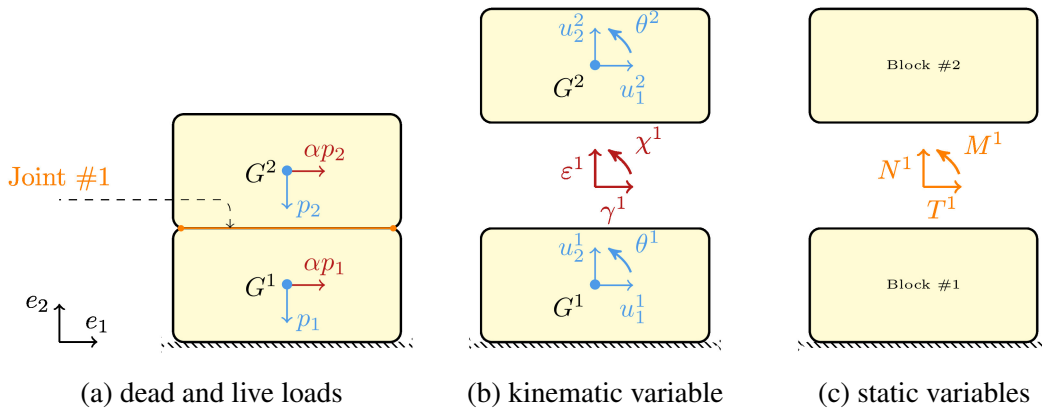


FIG. 2: Schematic representation of a two-block structure with one joint represented in the local reference system

205
 206 Within the framework of the holonomic perfect plasticity, the following relations
 207 govern the problem of a non-standard rigid-plastic discrete material. The kinematic
 208 compatibility and the equilibrium equations for the whole system are expressed as
 209 follow:

$$\epsilon = \mathbf{B} \mathbf{u} , \quad (2)$$

$$\mathbf{B}^T \sigma + \mathbf{f} = \mathbf{0} , \quad (3)$$

210 where \mathbf{B} represents the compatibility matrix defined in (Baggio and Trovalusci 2000).
 211 In case of joint k with arbitrary direction between two blocks, \mathbf{B} is constructed using
 212 the rotational matrix that maps the local joint to the global one.

213 The generalized yield domain of the system can be written as

$$\mathbf{y} = \mathbf{N}^T \boldsymbol{\sigma} \leq \mathbf{0}, \quad (4)$$

214 where \mathbf{N} is so-called gradient matrix referred to the adopted failure surface. For each
 215 i^{th} blocks, Equation (4) assumes the following form:

$$\begin{Bmatrix} y_1^i \\ y_2^i \\ y_3^i \\ y_4^i \end{Bmatrix} = \begin{bmatrix} l^k/2 & 0 & -1 \\ l^k/2 & 0 & 1 \\ \tan(\phi) & -1 & 0 \\ \tan(\phi) & 1 & 0 \end{bmatrix} \begin{Bmatrix} N^k \\ T^k \\ M^k \end{Bmatrix}, \quad (5)$$

216 where l^k is the length of k^{th} joint.

217 The flow rule expresses the vector $\boldsymbol{\epsilon}$ as a linear combination of non-negative coef-
 218 ficients ordered in the vector $\boldsymbol{\lambda}$, called plastic multiplier vector, and it can be written
 219 as

$$\boldsymbol{\epsilon} = \mathbf{M} \boldsymbol{\lambda}, \quad (6)$$

220 where \mathbf{M} is matrix of the modes of failures.

221 Finally, the complementarity condition and the non-negative work of the live loads,
 222 which cause the collapse mechanism, must be imposed by satisfying the following
 223 equations

$$\boldsymbol{\lambda}^T \mathbf{y} = 0, \quad (7)$$

$$\mathbf{f}_L^T \mathbf{u} = 1. \quad (8)$$

224 Resorting to the formal analogy between rigid perfectly plastic discrete systems and
 225 rigid blocks with no-tension and frictional interfaces, the collapse load for a masonry
 226 structure, under the hypothesis of proportional load with the factor $\alpha > 0$, can be
 227 determined. After some algebraic manipulations, for the sake of brevity the final non-
 228 linear and non-convex programming problem (NLNCP) is reported, that reads:

$$\begin{aligned} \alpha_C = \min \{ \alpha \} \text{ subjected to} \\ (\mathbf{A}\mathbf{M}_1 - \mathbf{M}_2) \boldsymbol{\lambda} = 0, \text{ kinematic compatibility} \\ (\mathbf{A}_0 \mathbf{N}_1^T) (\mathbf{f}_0 + \alpha \mathbf{f}_L) + [\mathbf{N}_2^T - (\mathbf{A}\mathbf{N}_1)^T] \boldsymbol{\sigma}_2 \leq 0, \text{ static admissibility} \\ \boldsymbol{\lambda}^T (\mathbf{A}_0 \mathbf{M}_1)^T \mathbf{f}_L - 1 = 0, \text{ normalized positive work of live loads} \\ \boldsymbol{\lambda}^T \left\{ (\mathbf{f}_0 + \alpha \mathbf{f}_L) + [\mathbf{N}_2^T - (\mathbf{A}\mathbf{N}_1)^T] \boldsymbol{\sigma}_2 \right\} = 0, \text{ complementarity condition} \end{aligned} \quad (9)$$

229 where α_C is the collapse multiplier, \mathbf{B}_1 is the kinematical submatrix of maximum rank
 230 of the compatibility matrix \mathbf{B} and \mathbf{B}_2 the rest of the kinematical matrix. The matrix

231 \mathbf{A}_0 is the inverse of \mathbf{B}_1 . The matrix \mathbf{A} is defined as $\mathbf{A} = \mathbf{B}_2\mathbf{B}_1^{-1}$ and \mathbf{N}_i , with
 232 $i = 1, 2$, are two submatrices of \mathbf{N} obtained after sharing the kinematical variables
 233 into two parts: the independent and the linear dependent ones (Baggio and Trovalusci
 234 2000).

235 The unknowns of the problems are: α , $\boldsymbol{\sigma}_2$, $\boldsymbol{\lambda}$ with the bounds $\alpha \geq 0$ and $\boldsymbol{\lambda} \geq \mathbf{0}$;
 236 $\boldsymbol{\sigma}_2$ are the undetermined unknown of the system which represent the statically unde-
 237 termined term of the generalized stress $\boldsymbol{\sigma}$ (Baggio and Trovalusci 1998).

238 For systems with non-associated flow rules the Drucker stability postulate no longer
 239 holds, the solution loses its uniqueness and lower/upper bounds for the collapse multi-
 240 plipliers can be found (Drucker 1953; Radenkovic 1961). The solution of a non-linear
 241 and non-convex programming problem could not exist and when it is found, it can be
 242 locked in a local minimum rather than the global one (Kirsch 1993).

243 In order to deal with the NLNCP, authors (Baggio and Trovalusci 2000) developed
 244 a specific computational code (*ALMA*: Analisi Limite Murature Attritive), based on a
 245 two-step procedure: initially a linear programming problem (LP), obtained by adopt-
 246 ing the assumption of dilatancy, hypothesis which makes the problem governed by
 247 associative flow rule, is solved, followed by the attempt to approach the non-linear
 248 solution using as initial guess of NLNCP the solution previously obtained for such LP.

249 As the solutions obtained following the dilatancy assumption approach LP (first
 250 step), which in most cases provided results frequently quite close to the solutions of
 251 NLCP (second step), both in terms of collapse multipliers and mechanisms (Baggio
 252 and Trovalusci 2000), in this work we decided to focus on the linear programming
 253 optimization problem referred as the upper bound approach of Limit Analysis, for pro-
 254 viding collapse multipliers and the corresponding collapse mechanisms of analysed
 255 structures. It is known that when normality rule holds the static and kinematic theo-
 256 rems of Limit Analysis can be considered as two dual problems of linear programming
 257 optimization, which lead to a unique solution. In particular, the adopted kinematic
 258 upper bound problem is defined as

$$\begin{aligned} \alpha_c = \min \left\{ -\boldsymbol{\lambda}^T (\mathbf{A}_0\mathbf{N}_1)^T \mathbf{f}_0 \right\} \text{ subjected to} \\ (\mathbf{A}\mathbf{N}_1 - \mathbf{N}_2) \boldsymbol{\lambda} = \mathbf{0}, \text{ compatibility condition} \\ \boldsymbol{\lambda}^T (\mathbf{A}_0\mathbf{N}_1)^T \mathbf{f}_L - 1 = 0, \text{ normalized positive work of live loads} \end{aligned} \quad (10)$$

259 with the bounds on the unknowns $\boldsymbol{\lambda} \geq \mathbf{0}$.

260 DESIGN OF EXPERIMENTS (DOE)

261 As the aim of this work was to objectively determine the influence of various geo-
 262 metrical and mechanical parameters on the collapse multiplier and collapse mechanism
 263 of a masonry panel, a systematic methodology has been implemented. A general full
 264 factorial design was used to identify both the main and the two-way interaction effects
 265 of these parameters on the masonry panels response. The factors considered and their
 266 correspondent levels (in this context, a level refers to a particular value adopted by a
 267 parameter) are presented in Table 1.

268 Three levels were adopted for the panel ratio factor, namely, 2:1, 1:1 and 1:2, keep-
269 ing the length of the panels fixed, at a value of $B=1440$ mm, and varying their heights.
270 These panel ratios correspond to the ones studied by (Baraldi et al. 2018), and are
271 considered to be representative from typical masonry panels ratios present in historical
272 buildings. The different panel ratios studied are presented in Figure 3.

273 Regarding the block ratio factor, three different levels were assumed: 4:1, 2:1 and
274 1:1, being this time the block's height the fixed value and their length the varying di-
275 mension. Block ratios of 4:1 and 2:1 are typically found both in historical and modern
276 masonry typologies. Even if blocks with a 1:1 block ratio are rarely found in masonry
277 buildings, this ratio has been also considered, thus enhancing the comparison purpose
278 of this work

279 The different bond types studied in this work were: *running* (R), *flemish* (F), *en-*
280 *glish* (E) and *stack* (S) (see Figure 4). *Running*, *english* and *flemish* are implemented
281 both as coating and for structural purposes, due to their relative higher resistance cre-
282 ated by the good interlocking generated by the offset of their blocks, whereas the *stack*
283 bond type is generally used only as coating.

284 Mathematically speaking the values that could be adopted for the friction angle,
285 ϕ , could lay within the interval $0 < \phi < 90$ (in degrees). Masonry friction angles
286 have been experimentally determined by several researchers in the past and the values
287 reported oscillate between the 17 and the 63 degrees (Rahman and Ueda 2014). How-
288 ever, most commonly friction angle values for historical masonry vary between 15 and
289 45 degrees. Therefore, the different levels for the friction ratio, $\tan(\phi)$, studied were
290 0.27, 0.60 and 1.00 which correspond to 15, 30 and 45 degrees, respectively. The lower
291 friction adopted would represent a situation in which the blocks surface were relatively
292 smooth, whereas that a rough surface would be better represented by the higher friction
293 value assumed. Finally, the adopted value of 0.60 would correspond to an intermediate
294 level of block surface roughness.

295 Other important simulation parameters adopted in this work correspond to the
296 thickness of the panels, which was assumed to be fixed for all panels at a value of
297 120 mm, an also constant specific weight of 18 kN/m^3 was assumed throughout the
298 performed simulations as well as a null value for cohesion. The load condition applied
299 in all simulations consisted of a self-weight (dead load) and a horizontal body force
300 (live load) proportional to the self-weight. In summary, the full factorial DOE resulted
301 in the simulation of 108 different masonry panels.

302 The basic principles of statistical study require to perform multiple experiments
303 because the order in which they are run, the number of times they are run or the way
304 in which they are grouped may affect the answer obtained. This is usually the case
305 when laboratory experiments are performed as the response can be influenced by un-
306 controlled factors.

TABLE 1: Factors and their respective levels.

Factor	Level	Value
A. Panel ratio (length:height)	1. 2:1	1440x720 mm
	2. 1:1	1440x1440 mm
	3. 1:2	1440x2880 mm
B. Block ratio (length:height)	1. 4:1	240x60 mm
	2. 2:1	120x60 mm
	3. 1:1	60x60 mm
C. Bond type	1. Running	-
	2. Flemish	-
	3. English	-
	4. Stack	-
D. Friction ($\tan(\phi)$)	1. Low	0.27
	2. Medium	0.60
	3. High	1.00

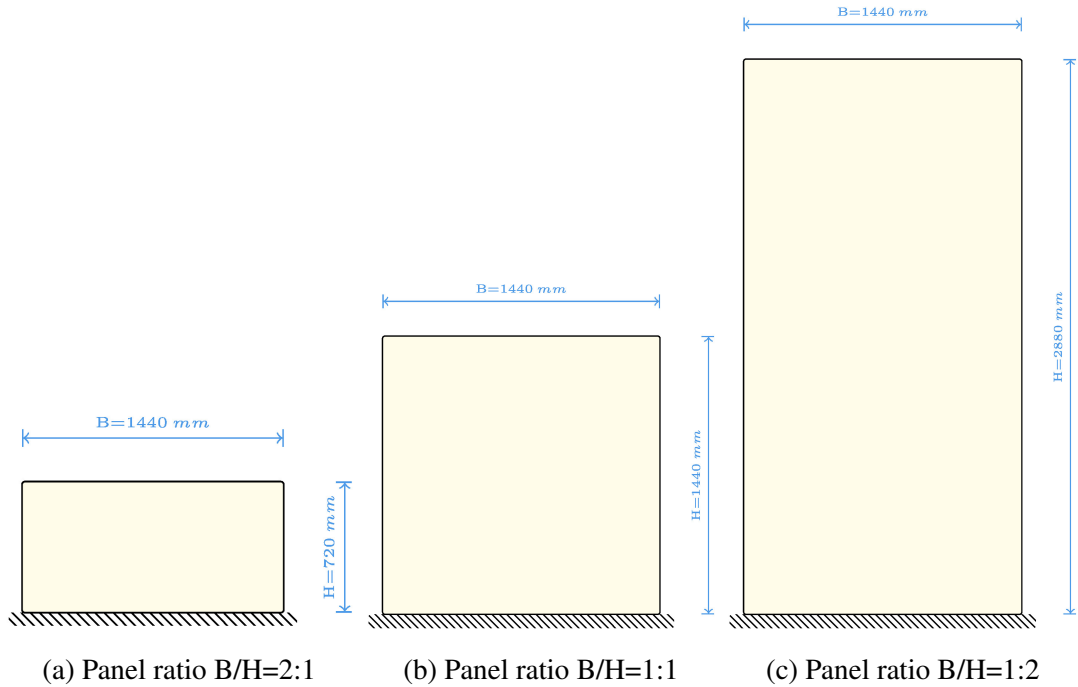


FIG. 3: Masonry panels with the different ratios adopted for the analysis.

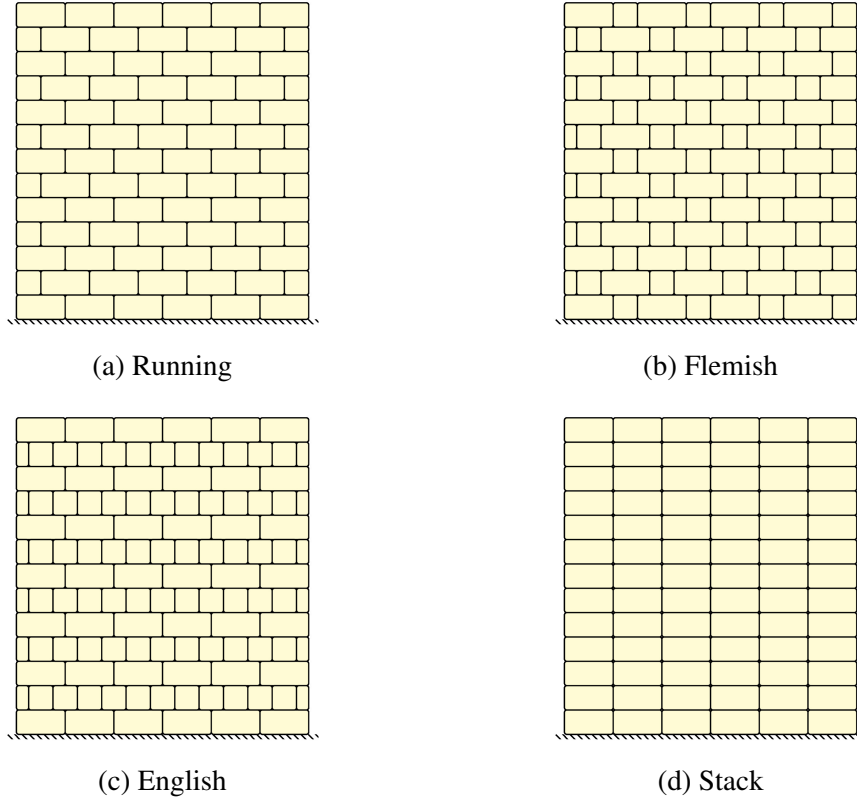


FIG. 4: Masonry panels with the different bond types tested.

307 After all simulations were successfully run, the responses (collapse multipliers)
 308 obtained were visually analysed by the averages of main effects and two-way interac-
 309 tion effects plots. The main effect plots allow to see the actual effect that every single
 310 parameter has in the response. By computing the average values at every level of each
 311 parameter it is assumed that the response is independent of the other parameters. On
 312 the other hand, two-way (or higher order) interaction plots allow to study the possi-
 313 ble interaction between two (or more) parameters and how this affects the response.
 314 The points of a two-way interaction plot are computed by averaging the values of α
 315 obtained for a certain combination of two parameters' levels.

316 Furthermore, the collapse multipliers obtained were formally analysed through an
 317 analysis of variance (ANOVA). The ANOVA, sometimes referred to as significance
 318 of regression test, determines whether there is a relationship between the parameters
 319 of the statistical model (also known as regressor variables) and the response. The
 320 hypotheses of the ANOVA test are:

$$\begin{aligned}
 H_0 : \beta_1 = \beta_2 = \dots \beta_k = 0 , \\
 H_1 : \beta_j \neq 0 \text{ for at least one } j .
 \end{aligned}
 \tag{11}$$

321 Where H_0 and H_1 are the null and the alternative hypothesis respectively, and β_j
 322 represents every coefficient of the linear and two-way interaction terms of the statistical

323 model adopted. The rejection of H_0 implies that at least one of the terms contributes
324 significantly to the output of the statistical model. The statistical model adopted for the
325 masonry panels is composed by four linear terms (panel ratio, block ratio, bond type
326 and friction), six two-way interaction terms (combinations of the four linear terms
327 previously mentioned) and an error term. Linear terms correspond to the effect that
328 individual parameters have in the response whereas that two-way interaction terms
329 depict how the response is affected by a certain parameter in combination with the
330 different levels of a second one. The error term is related to the inherent variation of
331 the model and is assumed to be normally and independently distributed.

332 Further details about the statistical approach adopted and the ANOVA analysis are
333 reported in Appendix A. Additionally, the magnitude and importance of each one of the
334 main factors and factor interaction effects were obtained. Those results are presented as
335 Pareto charts of standardized effect. The Pareto chart of standardized effects is used to
336 compare the relative magnitude and the statistical significance of the main parameters
337 and of the two-way interaction terms in the response. Moreover, a reference line is also
338 plotted in the Pareto chart in order to simplify the identification of the significant terms
339 (every term with a standardized effect value higher than the reference line is considered
340 to be statistically significant).

341 The suitability of the adopted statistical model to describe the response, which is
342 basically a regression model, was measured as the values of the coefficient of determi-
343 nation, R^2 (equal to the regression sum of squares divided by the total sum of squares),
344 and of the predicted coefficient of determination, R_{pred}^2 . Finally, the assumptions that
345 the data was independent and normally distributed, in other words, that the analysed
346 data was not affected by non-controlled parameters and that it roughly presents the
347 shape of the Gauss curve, were visually validated by analysing the standardized resid-
348 ual plots of the response, α , its histogram and its normal probability plot.

349 RESULTS AND DISCUSSION

350 After running the 108 different simulations generated with the adopted DOE, re-
351 sults were obtained in terms of collapse multiplier values and collapse mechanisms.

352 Focusing our attention first on the analysis of the collapse multipliers, Table 5, in
353 Appendix B, presents the collapse multipliers obtained for every simulation. Figure 5
354 presents the main effects plot for the response, $\langle\alpha_c\rangle$, in which the effect that every
355 single parameter studied has in the response can be observed. Each curve corresponds
356 to one of the factors considered in this study. For instance, the left graph of Figure 5
357 provides mean $\langle\alpha_c\rangle$ values obtained while the variation of the rest of parameters is
358 neglected. In this context, to study the effect that the panel ratio parameter has in
359 the response it is necessary to compute the mean response at each one of its levels,
360 namely, 2:1, 1:1 and 1:2. From Table 5 it can be observed that a value of panel ratio
361 equal to 2:1 was implemented in a total of 36 simulations. Thus, the mean response
362 value is computed as the average of the 36 corresponding collapse multipliers. This
363 mean value is plotted as the point corresponding to a panel ratio of 2:1 in the leftmost
364 curve of Figure 5. Similarly, mean response values were computed and plotted for
365 panel ratio values of 1:1 and 1:2, as well as for the rest of levels of each individual

366 parameter studied in this paper.

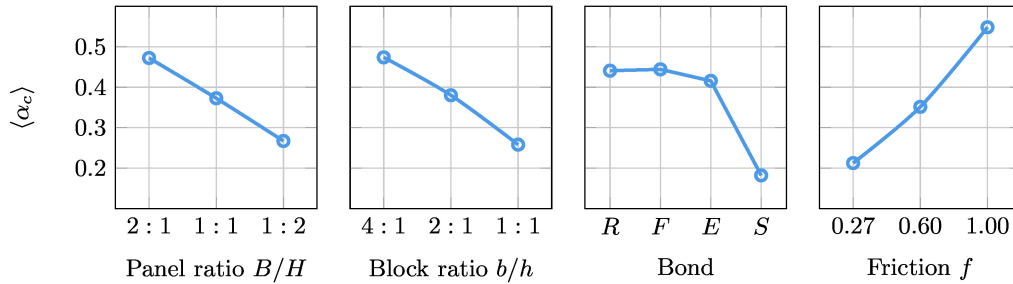


FIG. 5: Variation of mean value $\langle \alpha_c \rangle$ fixing one analysis parameters and varying the others.

367 It can be clearly noticed that all factors significantly influence the values of $\langle \alpha_c \rangle$.
368 There is a direct correspondence between the values of $\langle \alpha_c \rangle$ and panel ratio and block
369 ratio values, in the sense that the more slender the panel/brick the lower is the value
370 of the collapse multiplier. Differently, it can be observed that the higher the friction
371 coefficient, the greater the value of $\langle \alpha_c \rangle$.

372 A clear different response can be identified in terms of the bond type. While
373 *running*, *english* and *flemish* bond types present similar average values of $\langle \alpha_c \rangle$, the
374 *stack* bond type shows a relatively low value for the average collapse multiplier. This is
375 without a doubt the result of the lack of units interlocking of the stack masonry panels
376 simulated as shown in (Baraldi et al. 2018). After further analysing the bond type
377 main effect, by the means of a Tukey's multi comparison confidence intervals, it was
378 observed that there was effectively a significant difference between the mean collapse
379 multipliers of the *stack* bond panels and the other types of bond, but not between
380 *running*, *english* and *flemish* bond mean collapse multipliers. This fact may justify a
381 modelling geometry simplification when in real historic masonry structures it is not
382 possible to fully observe masonry texture, i.e. when the wall is partially rendered.

383 Furthermore, interesting information can be drawn from the observation of the re-
384 sponse interaction plots presented in Figure 6. In this figure different plots, correspond-
385 ing to the interaction of every pair of main factors analysed in this study, are organized
386 and presented as a lower matrix.

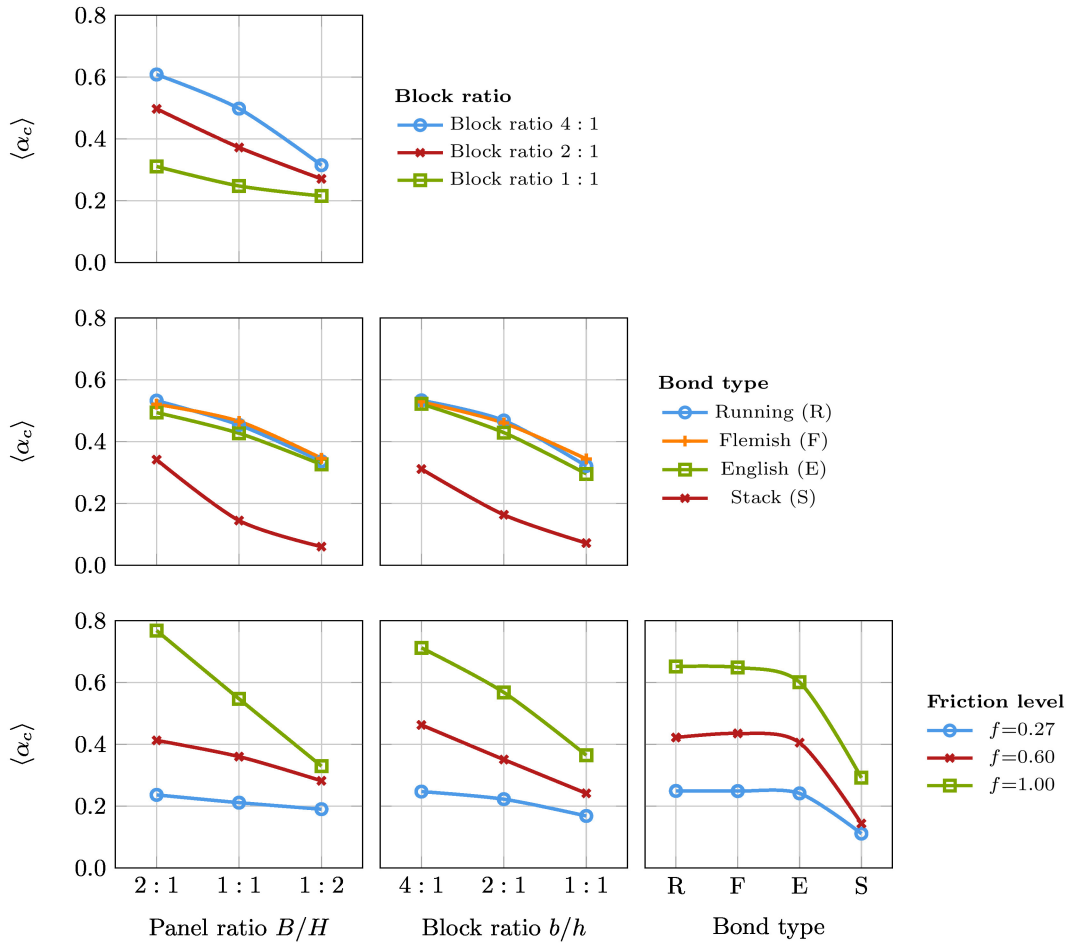


FIG. 6: Interaction plot for $\langle \alpha_c \rangle$.

387 In fact, this matrix is symmetric and the upper part has been omitted for the sake of
388 clarity. In every plot a specific factor level has been plotted against the mean value of
389 $\langle \alpha_c \rangle$, and the different curves in every plot correspond to the different levels of the sec-
390 ond main factor indicated in the legend at the right side of Figure 6. From these plots,
391 it can be observed that a significant correlation exist between the studied interaction
392 factors and the response of the different masonry panels. Regarding the Panel*Block
393 interaction plot it can be seen that slender panels (1:2 ratio) always produce lower mean
394 values of $\langle \alpha_c \rangle$ in comparison with square (1:1 ratio) and squat panels (2:1 ratio). More-
395 over, for the different levels of block ratio it can be observed that the more slender the
396 block (1:1 ratio) the lower would be the values of the response. These observations are
397 in agreement with results obtained from other parametric analysis (Bustamante 2003;
398 Casapulla and Argiento 2018). If the Panel*Friction and the Block*Friction plots are
399 analysed, it is clear that higher values of friction would give a higher strength to the
400 panel and result into higher collapse multiplier values. Furthermore, from these three
401 plots a wider variation on the response can be noticed for the squat panel/block ratios in
402 comparison with the slender levels of those factors. This indicates that the influence of

403 block ratio and friction ratio in the response of squat panels would be more significant
404 than in the response of slender panels where the response would be mainly influenced
405 by the geometry of the panel itself.

406 Analogously as it was observed in the main effects plots, the response obtained for
407 the *english*, *flemish* and *running* bonds was quite similar, whereas that a clear lower
408 response is obtained for the *stack* bond panels simulated. The Panel*Bond and the
409 Block*Bond plots follow the tendencies previously described. Since the bond type
410 factor is a categorical factor, no clear trend can be drawn from the Bond*Friction plot
411 except for the fact that at every factor levels combination, the *stack* bond resulted into
412 lower values of α_c in comparison with the other bond types. As per the main effects
413 plots, it can be said that a significant interaction exists between bond type and the rest
414 of the factors if the *stack* bond is included. On the other hand, no interaction would be
415 observed if only the *running*, *english* and *flemish* bonds were considered.

416 These visual assumptions are formally verified through an analysis of variance
417 (ANOVA) test of the results. Table 2 presents the ANOVA results performed using
418 the software *Minitab*[®] (<https://www.minitab.com/en-us/>). In the first col-
419 umn of Table 2 the statistical model terms are identified by their names, in the second
420 column the degrees of freedom (DoF) of every term are presented, in the third column
421 the adjusted sum of squares (Adj SS) corresponding to every term are shown and in
422 column four the adjusted mean squares (Adj MS) are listed. In the fifth column of the
423 ANOVA table we can see the value corresponding to the F statistical test and finally,
424 in the last column of Table 2, the corresponding P-Values to each term are presented.

425 From Table 2 it can be observed from the P-Values of the ANOVA table that besides
426 from the Block*Bond, all linear and two-way interactions are statistically significant
427 ($P - Value < 0.05$) at a confidence level of 95%.

TABLE 2: ANOVA.

Source	DoF	Adj SS	Adj MS	F-Value	P-Value
Model	39	6.0529	0.1552	51.98	0.000
Linear	9	4.9403	0.5489	183.86	0.000
Panel	2	0.7573	0.3787	126.83	0.000
Block	2	0.8426	0.4213	141.12	0.000
Bond	3	1.2930	0.4310	144.37	0.000
Friction	2	2.0474	1.0237	342.88	0.000
Two-Way Interactions	30	1.1126	0.0371	12.42	0.000
Panel*Block	4	0.1347	0.0337	11.28	0.000
Panel*Bond	6	0.0658	0.0110	3.67	0.003
Panel*Friction	4	0.5117	0.1279	42.85	0.000
Block*Bond	6	0.0234	0.0039	1.31	0.266
Block*Friction	4	0.2203	0.0551	18.45	0.000
Bond*Friction	6	0.1566	0.0261	8.74	0.000
Error	68	0.2030	0.0030		
Total	107	6.2559			

DoF=Degrees of freedom, Adj SS= Adjusted sum of squares, Adj MS = Adjusted mean of squares.

428 Furthermore, the magnitude and importance of each one of the main factors and
 429 factor interaction effects were obtained. Figure 7 presents the Pareto chart of standard-
 430 ized effect. The Pareto chart of standardized effects is used to compare the relative
 431 magnitude and the statistical significance of the main parameters and of the two-way
 432 interaction terms in the response. Moreover, in Figure 7 a reference line is also plot-
 433 ted in order to simplify the identification of the significant terms (every term with a
 434 standardized effect value higher than the reference line is considered to be statistically
 435 significant).

436 In Figure 7 it can be seen that all main factors as well as the Panel*Friction inter-
 437 action would have a bigger effect on the response of the masonry panels. These are fol-
 438 lowed in order of importance by the Block*Friction, the Panel*Block, the Bond*Friction
 439 and the Panel*Bond interaction effects. Finally, it can be observed that the standard-
 440 ized effect of the Block*Bond interaction is smaller than the reference value, in this
 441 case 1.995, and therefore is not considered to be statistically significant.

442 Regarding the suitability of the adopted statistical model to describe the response,
 443 the values of the coefficient of determination, R^2 (equal to the regression sum of
 444 squares divided by the total sum of squares), and of the predicted coefficient of de-
 445 termination, R^2_{pred} , obtained were of 96.75% and 91.81% respectively. This proves the
 446 good fit and the high prediction capabilities of the model adopted. Nevertheless, the
 447 statistical model could be further improved if a new statistical model were performed

448 without considering the Block*Bond interaction term, but that step is outside the scope
449 of this work.

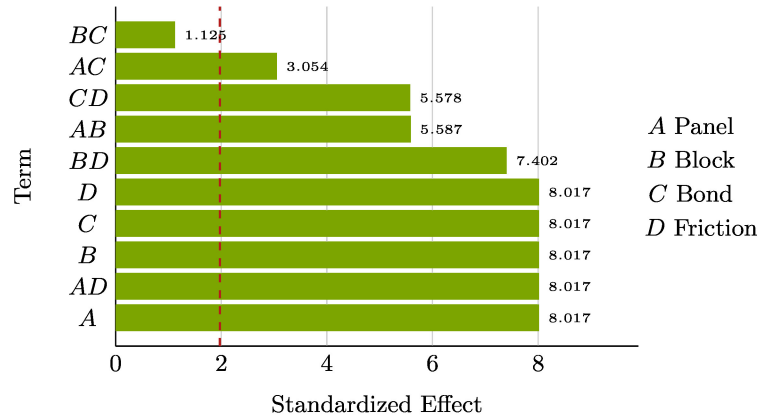


FIG. 7: Pareto chart of the standardized effects: magnitude and importance of the different main factors and factor interactions effects.

450 The statistical analysis performed was carried out under the assumption that the
451 data was independent and normally distributed, in other words, that the analysed data
452 was not affected by non-controlled parameters and that it roughly present the shape
453 of the Gauss curve. These assumptions were visually validated by analysing the stan-
454 dardized residual plots of the response, α_c , its histogram and its normal probability
455 plot. In the normal probability plot of Figure 8 (upper right plot) it could be observed
456 that most of the points were relatively close to the diagonal line. There was only one
457 standardized residual with a value larger than three, which represented an outlier, that
458 did not influence significantly the suitability of the adopted model. Furthermore, if this
459 point were disregarded it could be observed that the histogram on Figure 8 resembled
460 to a Gauss distribution and that no clear structure is present on the Versus Fits nor on
461 the Versus order plots of the standardized residuals in Figure 8.

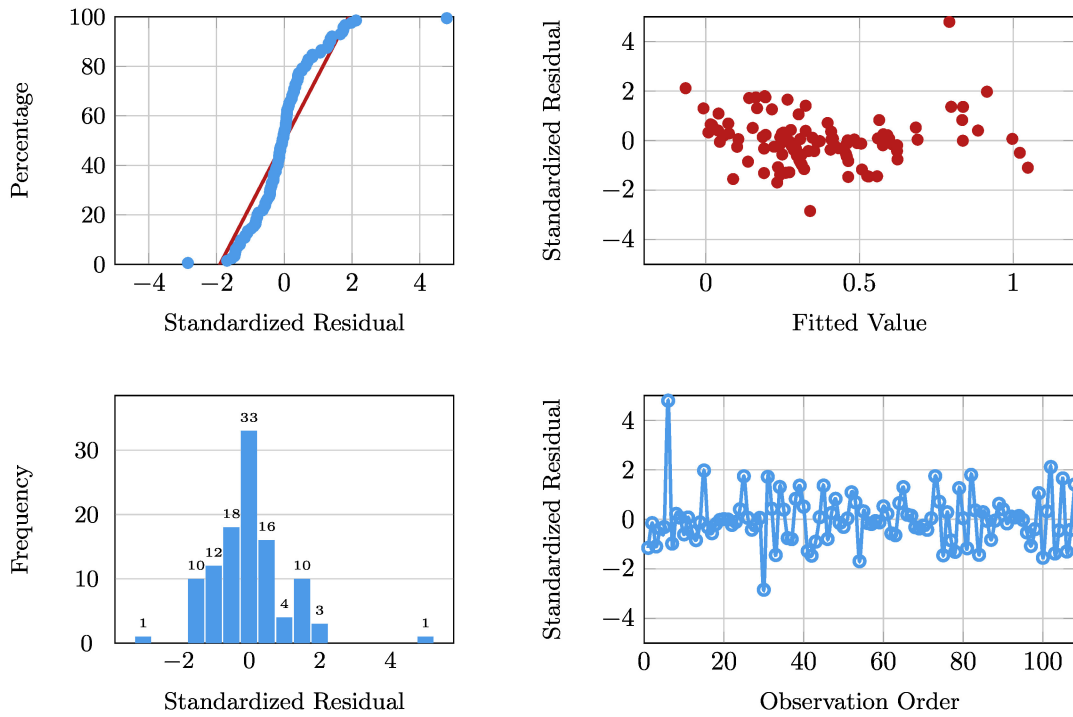


FIG. 8: Normal probability, histogram and standardized residual plots.

462 Although not explicitly included as one of the main factors in the DOE presented
 463 in this paper, the scale factor, meaning the ratio between block and panel dimension,
 464 could as well be analysed based on the results obtained. Figure 9 presents the mean val-
 465 ues of $\langle \alpha_c \rangle$ obtained as a function of the scale factor, H/b , for *running* (Figure 9(left)),
 466 *english* (Figure 9(center)) and *flemish* (Figure 9(right)) bond panels at different values
 467 of friction coefficient. In the group of performed analysis all of those with the same
 468 scale ratio H/b at fix bond type and level of friction coefficient have been selected for
 469 computing the mean value of $\langle \alpha_c \rangle$.

470 By comparing the plots, it can be observed that the panels response for the three
 471 type of textures well interlocked, R, F, E, is not affected by the choice of bond, same
 472 values of α_c were obtained for the three bond types, while for the *stack* bond panels
 473 lower collapse multiplier values are detected. Friction has greater influence for low
 474 value of ratio H/b and decreases as panel became slender.

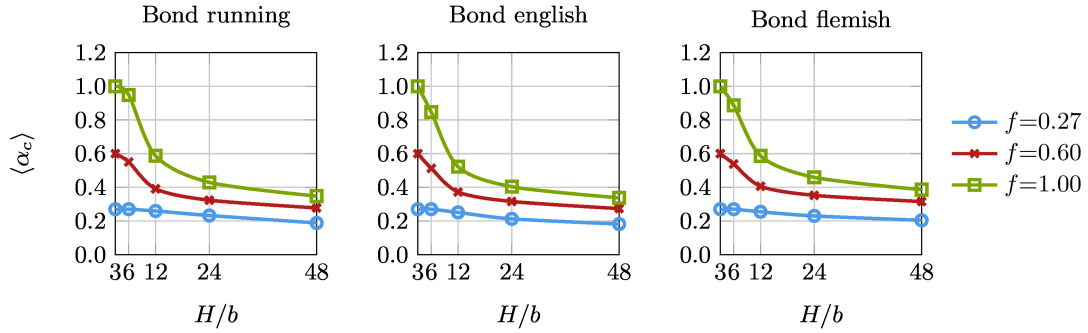


FIG. 9: Variation of the mean value of collapse multiplier $\langle \alpha_c \rangle$ vs the scale factor, H/b , for different friction ratios, $f = \tan(\phi)$.

475 A similar trend may be observed considering scale factor B/b . In Figure 10 presents
 476 the mean values of $\langle \alpha_c \rangle$ obtained as a function of the scale factor, B/b , for different
 477 panel ratio at different values of friction coefficient. It can be observed that at low
 478 friction levels the panel and the scale factor are both negligible as the panel response
 479 is mostly influenced by friction. For the case of slender panels ($B/H=0.5$) the scale
 480 factor seems to play a minor role on the average collapse multiplier values. On the
 481 other hand, for intermediate and high friction values (0.6 and 1.0 respectively), and
 482 panel ratios $B/H=1$ and $B/H=2$, an interaction effect can be observed where at lower
 483 scale factor, higher values of α_c are obtained.

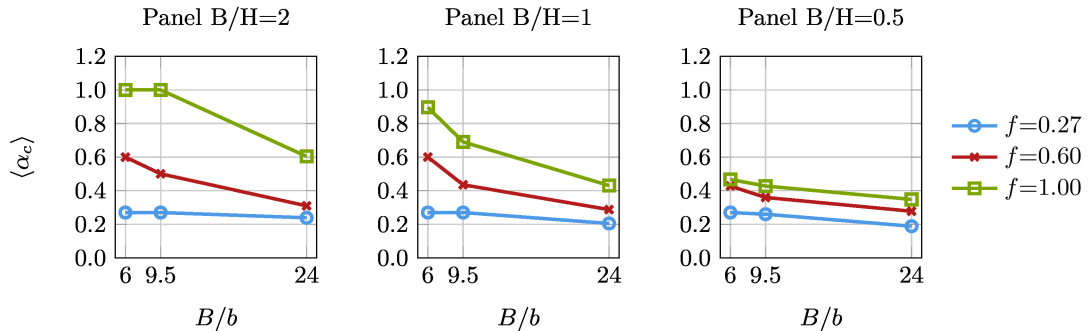


FIG. 10: Variation of the mean value of collapse multiplier $\langle \alpha_c \rangle$ vs the scale factor B/b for different friction coefficient: (left) Panel ratio $B/H=2$, (center) Panel ratio $B/H=1$ and (right) Panel ratio $B/H=0.5$.

484 Focusing our attention now on the failure type presented by the different masonry
 485 panels simulated, the collapse mechanism of each panel was also plotted and qualita-
 486 tively analysed. Considering a texture with high level of interlocking, which tends to
 487 behave as a monolithic assembly, in terms of collapse mechanism two different out-
 488 comes are detected: 1) a sliding mechanism for the squat panels and 2) a rotation
 489 mechanism for slender panels. From the results obtained in this parametric analysis

490 it was observed that those typical failures could be altered for certain factor combina-
491 tions.

492 In Figure 11, the collapse mechanism of three different panel ratio have been
493 shown. In this case the friction coefficient, block ratio and panel bond were fixed
494 to $f = 0.6$, 4:1 and *english*, respectively. In particular we can observe three different
495 results for the different panel ratio assumed:

- 496 • panel ratio B/H=2:1, Figure 11a, we observe a sliding mechanics, in fact the
497 collapse multiplier α_c is equal to the friction coefficient assumed, $\alpha_c = 0.60$;
- 498 • panel ratio B/H=1:1, Figure 11b we have a mixed mechanism in which appear
499 sliding and rotation; the collapse mechanism is close to the friction coefficient,
500 $\alpha_c \cong f$;
- 501 • panel ratio B/H=1:2, Figure 11c, rotational mechanism has been obtained and
502 the collapse multiplier is lower than the value of the friction coefficient, $\alpha_c < f$.

503 The previous observations are true only in case of panel with good interlocking, i.e
504 *running*, *english* and *flemish*. Similar collapse mechanisms were found on masonry
505 panels composed by slender blocks by (Baraldi et al. 2018).

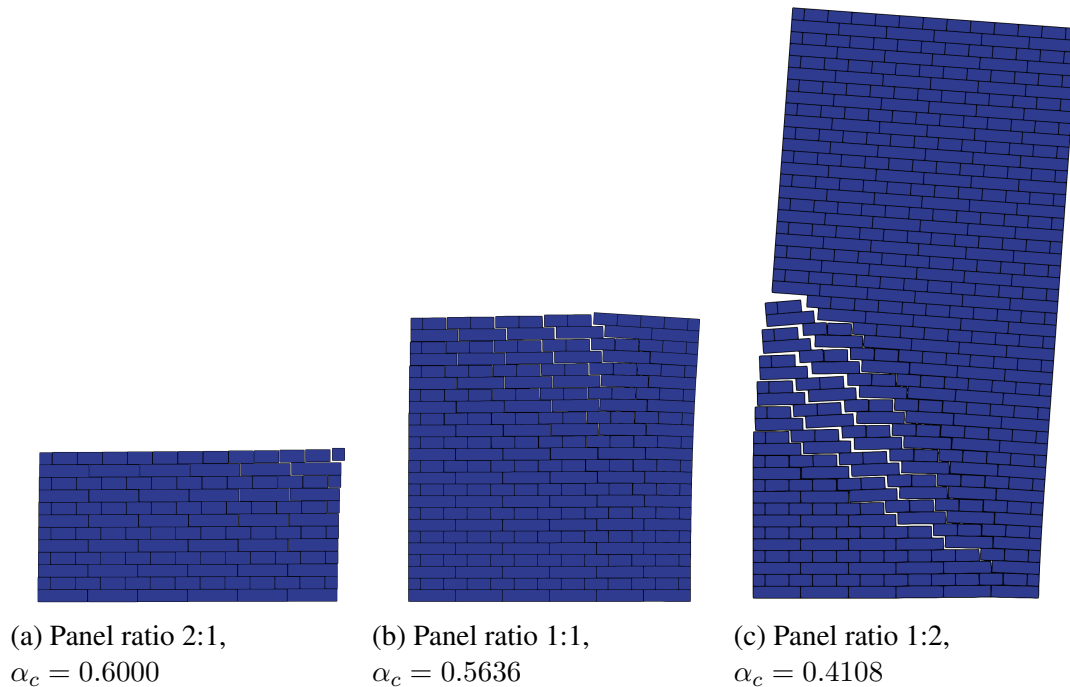


FIG. 11: Collapse mechanisms of *english* panels with friction coefficient $f = 0.60$ and block ratio 4:1

506 In order to show the effect of the block ratio, in Figure 12 is reported the slender
507 *flemish* panel with ratio 1:2 and friction coefficient equal to 0.6. Analysing different

508 types of block ratio, a rotation mechanism has been observed in accordance to the re-
 509 sults obtained in Figure 11c. A correlation between block ratio and collapse multiplier
 510 is highlighted: an increase of block ratio causes a correspondent growing of the panel
 511 portion involved in the collapse mechanism along with an increment of the collapse
 512 multiplier value.

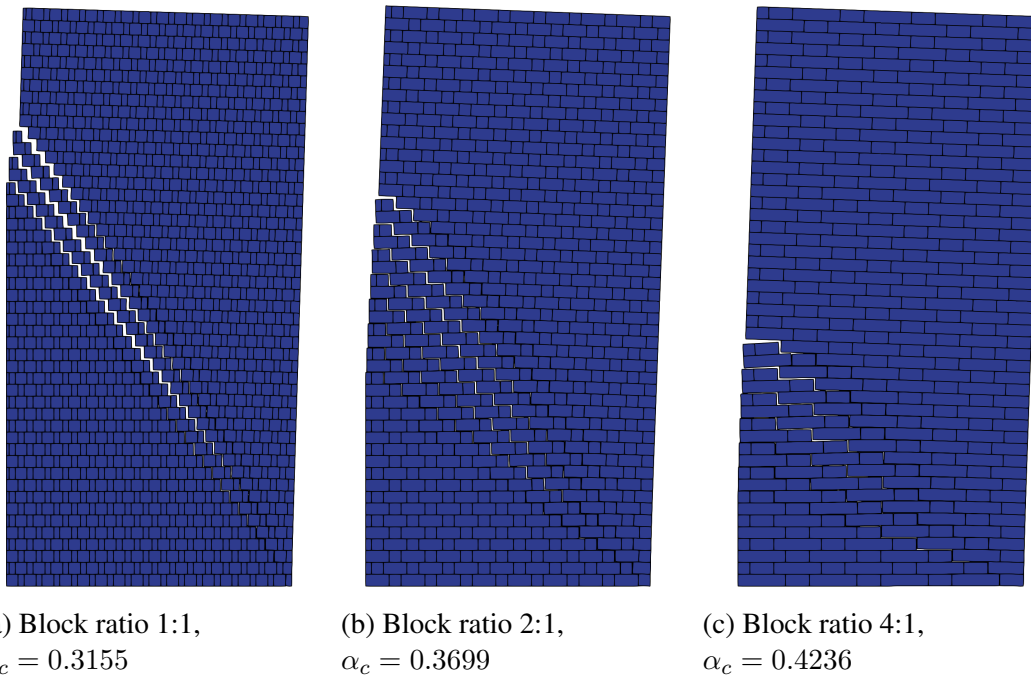
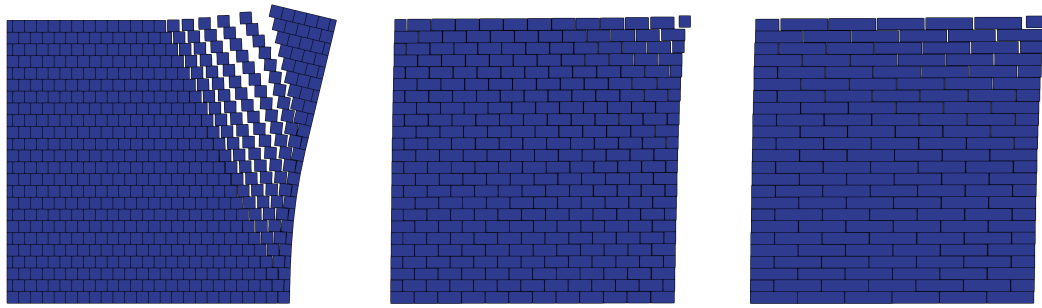


FIG. 12: Collapse mechanisms of *Flemish* panels with panel ratio 1:2 and medium value of friction, $f = 0.60$

513 In Figure 13 the collapse mechanisms for *running* panels with panel ratio 1:1 and
 514 low level of friction have been shown. In this case the same trend between block ratio
 515 and collapse multiplier is obtained. In particular a rotational mechanism is obtained
 516 for the low value of block ratio (Figure 13a) and sliding mechanisms for the other two
 517 level of block ratio (Figure 13b and Figure 13c).



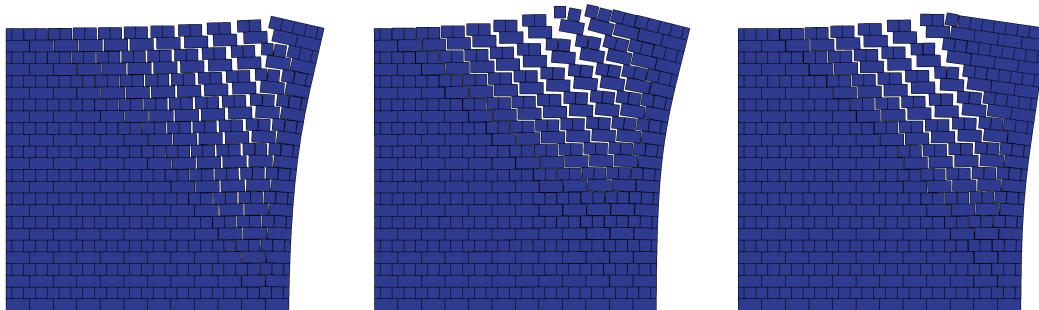
(a) Block ratio 1:1,
 $\alpha_c = 0.2049$

(b) Block ratio 2:1,
 $\alpha_c = 0.2670$

(c) Block ratio 4:1,
 $\alpha_c = 0.2670$

FIG. 13: Collapse mechanisms of *Running* panels with panel ratio 1:1 and low value of friction, $f = 0.27$

518 Collapse mechanisms of *english* panels with ratio 1:1 and block ratio 2:1 at dif-
519 ferent values of friction are shown in Figure 14. A relationship between the friction
520 coefficient and the collapse multiplier is evident in Figure 14.



(a) Friction $f = 0.27$,
 $\alpha_c = 0.2670$

(b) Friction $f = 0.60$,
 $\alpha_c = 0.4097$

(c) Friction $f = 1.00$,
 $\alpha_c = 0.6142$

FIG. 14: Collapse mechanisms of *English* panels with panel ratio 1:1 and block ratio 2:1

521 Finally, in Figure 15 the collapse mechanisms of panels with panel and block ratio
522 equal to 2:1 and friction $f = 0.60$ for different bond type have been reported.

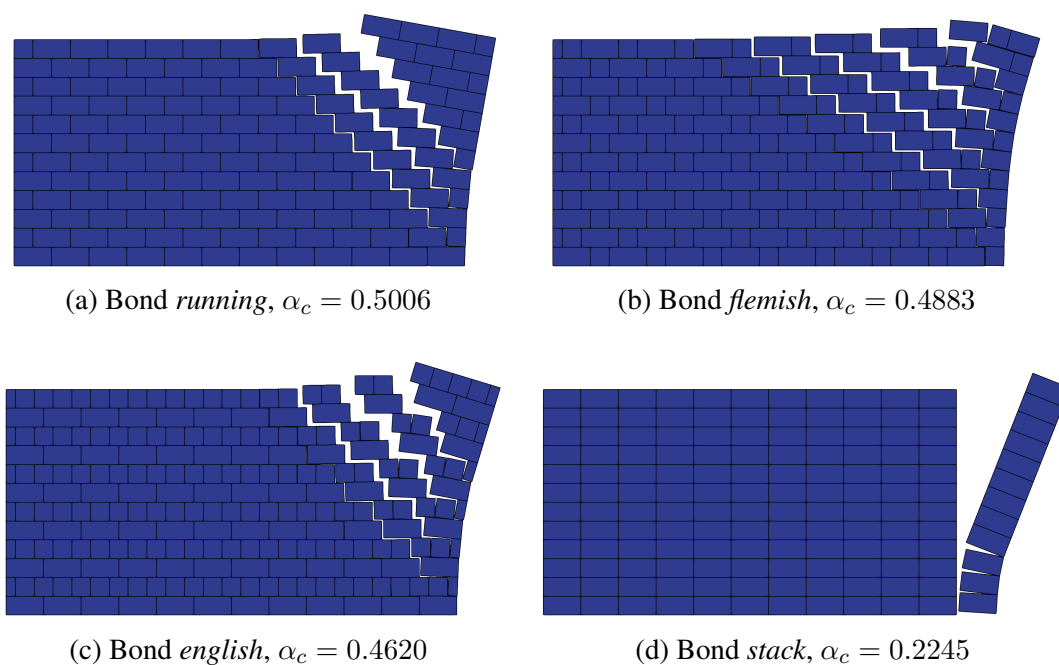


FIG. 15: Collapse mechanisms of panels with panel ratio 2:1, block ratio 2:1 and medium value of friction, $f = 0.60$

523 The different steps of the performed experiment and the main findings are summa-
 524 rized in Table 3.

TABLE 3: Parametric analysis summary.

Step	Information
Recognition of and statement of the problem.	What is the effect of panel ratio, block ratio, bond type and friction on the collapse multiplier α_c of a masonry panel?
Selection of response variable.	Collapse multiplier α_c .
Choice of factors and levels.	Panel ratio at 3 levels: 2:1, 1:1, 1:2 Block ratio at 3 levels: 4:1, 2:1, 1:1 Bond type at 4 levels: Running, Stack, English, Flemish Friction at 3 levels: 0.27, 0.60, 1.00
Choice of experimental design.	Full-composite.
Performing the experiment.	Trivial
Statistical analysis of the data.	$R^2 = 96.75\%$, $R^2_{pred} = 91.81\%$

525 CONCLUSIONS

526 The parametric analysis presented in this paper allowed to objectively identify the

527 effect that the panel ratio, block ratio, bond type and friction ratio parameters have in
528 the collapse multiplier value and in the collapse mechanism of a brick masonry panel
529 using a non-standard limit analysis approach. All the analyses performed considered
530 masonry walls made of bricks of different size and texture subjected to self-weight,
531 the dead load, and to an horizontal body force proportional to the weight through a
532 non negative load factor, the live load, which statically simulates a seismic action. The
533 main findings drawn from this work are:

- 534 • There is a strong correlation between the main factors and the response.
- 535 • With the exception of Block*Bond, all two-way interactions also resulted to be
536 statistically significant.
- 537 • The expected collapse mechanism could be modified under certain factor levels
538 combinations.
- 539 • The collapse multiplier for a panel that presents a sliding failure type cor-
540 responds with the value of the friction ratio, $\tan(\phi)$, whereas that the col-
541 lapse multiplier for panels that develop a rotation failure type would always
542 be smaller than this ratio.
- 543 • All the analyses show the importance in the collapse behaviour of the size and
544 the disposition of the bricks that determine the level of interlocking among
545 bricks, and then the cohesion of the whole.

546 The statistical model implemented in this work could be further refined if the non-
547 statistically significant terms were neglected. Moreover, the results obtained from this
548 study would be further exploited by the authors in a future work where a response
549 surface analysis (an statistical analysis that provides a series of equations that can be
550 used to predict the response of a model based on different levels of its input param-
551 eters) will be performed with the aim of providing a series of interpolation equations to
552 compute the approximate collapse multiplier of a masonry panel based on a random
553 combination of factor levels.

554 **APPENDIX A: ANOVA**

555 This appendix contains all details related to the statistical approach adopted and
556 the ANOVA analysis.

557 In an ANOVA context, the DoF are the amount of free data available to estimate
558 the coefficient of every statistical term in the model. For instance, the total DoF is
559 equal to the number of collapse multipliers obtained from the simulations minus one.
560 Every linear term in the model as a DoF equal to their number of levels minus one.
561 The addition of all DoF corresponding to the linear terms of the model provides the
562 total DoF for the linear part of the model. The DoF of the interaction terms is equal to
563 the product of the DoF from the corresponding linear terms. Similarly, the total DoF
564 corresponding to the two-way interaction part of the model is equal to the addition of
565 the DoF from the different two-way individual interaction terms. The DoF of the model
566 are equal to the sum of the DoF from its linear and two-way interaction parts. Finally,
567 the DoF that correspond to the error term of the model are equal to the total DoF

568 minus the DoF of the model. For further clarification, if the levels of every parameter
 569 studied: panel ratio (A), block ratio (B), bond type (C) and friction (D), are assigned
 570 respectively with the letters a , b , c and d , then the DoF are computed as presented in
 571 Table 4.

TABLE 4: Computation of DoF for the ANOVA.

Source	DoF
Model	$DoF_{Model} = DoF_{Linear} + DoF_{Two-WayInteractions}$
Linear	$DoF_{Linear} = DoF_A + DoF_B + DoF_C + DoF_D$
A	$DoF_A = a - 1$
B	$DoF_B = b - 1$
C	$DoF_C = c - 1$
D	$DoF_D = d - 1$
Two-Way Interactions	$DoF_{Two-WayInteractions} = DoF_{AB} + DoF_{AC} + DoF_{AD} + DoF_{BC} + DoF_{BD} + DoF_{CD}$
A*B	$DoF_{AB} = DoF_A * DoF_B$
A*C	$DoF_{AC} = DoF_A * DoF_C$
A*D	$DoF_{AD} = DoF_A * DoF_D$
B*C	$DoF_{BC} = DoF_B * DoF_C$
B*D	$DoF_{BD} = DoF_B * DoF_D$
C*D	$DoF_{CD} = DoF_C * DoF_D$
Error	$DoF_{Error} = DoF_{Total} - DoF_{Model}$
Total	$DoF_{Total} = a * b * c * d - 1$

572 Let y_{ijkl} represent the collapse multiplier obtained for every simulation where i is
 573 the level of factor A ($i = 1, 2 \dots a$), j the level of factor B ($j = 1, 2 \dots b$), k the level of
 574 factor C ($k = 1, 2 \dots c$) and l the level of factor D ($l = 1, 2 \dots d$). Then, $y_{i\dots}$, $y_{j\dots}$, $y_{\dots k}$ and
 575 $y_{\dots l}$ represent the total addition of the collapse multipliers corresponding to every level
 576 of factors A , B , C and D respectively. The totals for every two-way interaction are
 577 represented by $y_{ij\dots}$, $y_{i\dots k}$, $y_{i\dots l}$, $y_{\dots jk}$, $y_{\dots jl}$ and $y_{\dots kl}$. Finally, y_{\dots} represents the grand total.
 578 The formulas to compute each one of these terms are shown in Equation 12.

$$\begin{aligned}
y_{i...} &= \sum_{j=1}^b \sum_{k=1}^c \sum_{l=1}^d y_{ijkl}, \\
y_{.j.} &= \sum_{i=1}^a \sum_{k=1}^c \sum_{l=1}^d y_{ijkl}, \\
y_{..k.} &= \sum_{i=1}^a \sum_{j=1}^b \sum_{l=1}^d y_{ijkl}, \\
y_{...l} &= \sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^c y_{ijkl}, \\
y_{ij..} &= \sum_{k=1}^c \sum_{l=1}^d y_{ijkl}, \\
y_{i.k.} &= \sum_{j=1}^b \sum_{l=1}^d y_{ijkl}, \\
y_{i..l} &= \sum_{j=1}^b \sum_{k=1}^c y_{ijkl}, \\
y_{.jk.} &= \sum_{i=1}^a \sum_{l=1}^d y_{ijkl}, \\
y_{.j.l} &= \sum_{i=1}^a \sum_{k=1}^c y_{ijkl}, \\
y_{..kl} &= \sum_{i=1}^a \sum_{j=1}^b y_{ijkl}, \\
y_{....} &= \sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^c \sum_{l=1}^d y_{ijkl},
\end{aligned} \tag{12}$$

580

581 For every total term in Equation 12, a average can be computed as presented in
582 Equation 13.

$$\begin{aligned}
\bar{y}_{i...} &= \frac{y_{i...}}{b * c * d}, \\
\bar{y}_{.j..} &= \frac{y_{.j..}}{a * c * d}, \\
\bar{y}_{..k.} &= \frac{y_{..k.}}{a * b * d}, \\
\bar{y}_{...l} &= \frac{y_{...l}}{a * b * c}, \\
\bar{y}_{ij..} &= \frac{y_{ij..}}{c * d}, \\
\bar{y}_{i.k.} &= \frac{y_{i.k.}}{b * d}, \\
\bar{y}_{i..l} &= \frac{y_{i..l}}{b * c}, \\
\bar{y}_{.jk.} &= \frac{y_{.jk.}}{a * d}, \\
\bar{y}_{.j.l} &= \frac{y_{.j.l}}{a * c}, \\
\bar{y}_{..kl} &= \frac{y_{..kl}}{a * b}, \\
\bar{y}_{....} &= \frac{y_{....}}{a * b * c * d},
\end{aligned} \tag{13}$$

584

585 To describe the variability in the statistical model adopted, the total adjusted sum
586 of squares is computed according to Equation 14.

587

$$adjSS_T = \sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^c \sum_{l=1}^d (y_{ijkl} - \bar{y}_{....})^2, \tag{14}$$

588

589 The total adjusted sum of squares is split to represent the variability for each one
590 of the terms in the model as presented in Equation 15.

591

$$\begin{aligned}
adjSS_T &= adjSS_A + adjSS_B + sdjSS_C + sdjSS_D + adjSS_{AB} + adjSS_{AC} + \\
&adjSS_{AD} + adjSS_{BC} + adjSS_{BD} + adjSS_{CD} + adjSS_{Error},
\end{aligned} \tag{15}$$

592

593

Where:

$$\begin{aligned}
adjSS_A &= b * c * d * \sum_{i=1}^a (\bar{y}_{i...} - \bar{y}_{....})^2, \\
adjSS_B &= a * c * d * \sum_{j=1}^b (\bar{y}_{.j..} - \bar{y}_{....})^2, \\
adjSS_C &= a * b * d * \sum_{k=1}^c (\bar{y}_{..k.} - \bar{y}_{....})^2, \\
adjSS_D &= a * b * c * \sum_{l=1}^d (\bar{y}_{...l} - \bar{y}_{....})^2, \\
adjSS_{AB} &= c * d * \sum_{i=1}^a \sum_{j=1}^b (\bar{y}_{ij..} - \bar{y}_{i...} - \bar{y}_{.j..} - \bar{y}_{....})^2, \\
adjSS_{AC} &= b * d * \sum_{i=1}^a \sum_{k=1}^c (\bar{y}_{i.k.} - \bar{y}_{i...} - \bar{y}_{..k.} - \bar{y}_{....})^2, \quad (16) \\
adjSS_{AD} &= b * c * \sum_{i=1}^a \sum_{l=1}^d (\bar{y}_{i..l} - \bar{y}_{i...} - \bar{y}_{...l} - \bar{y}_{....})^2, \\
adjSS_{BC} &= a * d * \sum_{j=1}^b \sum_{k=1}^c (\bar{y}_{.jk.} - \bar{y}_{.j..} - \bar{y}_{..k.} - \bar{y}_{....})^2, \\
adjSS_{BD} &= a * c * \sum_{j=1}^b \sum_{l=1}^d (\bar{y}_{.jl} - \bar{y}_{.j..} - \bar{y}_{...l} - \bar{y}_{....})^2, \\
adjSS_{CD} &= a * b * \sum_{k=1}^c \sum_{l=1}^d (\bar{y}_{..kl} - \bar{y}_{..k.} - \bar{y}_{...l} - \bar{y}_{....})^2,
\end{aligned}$$

$$\begin{aligned}
adjSS_{Error} &= adjSS_T - adjSS_A - adjSS_B - adjSS_C - adjSS_D - \\
&adjSS_{AB} - adjSS_{AC} - adjSS_{AD} - adjSS_{BC} - adjSS_{BD} - adjSS_{CD},
\end{aligned}$$

595

596 The adjusted average of squares of every term in an ANOVA table is computed by
597 dividing the corresponding adjusted sum of squares value by its number of DoF. The F-
598 Values are computed by dividing the corresponding value of every adjusted average of
599 squares term by the value of the adjusted average of squares of the error term. Finally,
600 the P-Values are obtained using a F-test statistic table with the adequate DoF for the
601 numerator and denominator of every statistical term tested.

602 APPENDIX B: COLLAPSE MULTIPLIERS

603 This appendix contains the results obtained in terms of collapse multipliers for all
604 the simulations performed.

TABLE 5: Collapse multipliers obtained for every simulation.

2:1	4:1	Running	0.27	0.2700
1:1	4:1	Running	0.27	0.2700
1:2	4:1	Running	0.27	0.2700
2:1	2:1	Running	0.27	0.2700
1:1	2:1	Running	0.27	0.2700
1:2	2:1	Running	0.27	0.2600
2:1	1:1	Running	0.27	0.2386
1:1	1:1	Running	0.27	0.2049
1:2	1:1	Running	0.27	0.1884
2:1	4:1	Stack	0.27	0.2700
1:1	4:1	Stack	0.27	0.1750
1:2	4:1	Stack	0.27	0.0880
2:1	2:1	Stack	0.27	0.1758
1:1	2:1	Stack	0.27	0.0889
1:2	2:1	Stack	0.27	0.0437
2:1	1:1	Stack	0.27	0.0912
1:1	1:1	Stack	0.27	0.0443
1:2	1:1	Stack	0.27	0.0218
2:1	4:1	English	0.27	0.2700
1:1	4:1	English	0.27	0.2700
1:2	4:1	English	0.27	0.2700
2:1	2:1	English	0.27	0.2700
1:1	2:1	English	0.27	0.2670
1:2	2:1	English	0.27	0.2335
2:1	1:1	English	0.27	0.2158
1:1	1:1	English	0.27	0.1915
1:2	1:1	English	0.27	0.1820
2:1	4:1	Flemish	0.27	0.2700
1:1	4:1	Flemish	0.27	0.2700
1:2	4:1	Flemish	0.27	0.2700
2:1	2:1	Flemish	0.27	0.2700
1:1	2:1	Flemish	0.27	0.2700
1:2	2:1	Flemish	0.27	0.2482
2:1	1:1	Flemish	0.27	0.2237

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TABLE 5 – *Continued from previous page*

Panel ratio	Block ratio	Bond type	Friction ratio	Collapse multiplier
1:1	1:1	Flemish	0.27	0.2115
1:2	1:1	Flemish	0.27	0.2040
2:1	4:1	Running	0.60	0.6000
1:1	4:1	Running	0.60	0.6000
1:2	4:1	Running	0.60	0.4273
2:1	2:1	Running	0.60	0.5004
1:1	2:1	Running	0.60	0.4352
1:2	2:1	Running	0.60	0.3597
2:1	1:1	Running	0.60	0.3098
1:1	1:1	Running	0.60	0.2873
1:2	1:1	Running	0.60	0.2776
2:1	4:1	Stack	0.60	0.4191
1:1	4:1	Stack	0.60	0.2165
1:2	4:1	Stack	0.60	0.1006
2:1	2:1	Stack	0.60	0.2245
1:1	2:1	Stack	0.60	0.1033
1:2	2:1	Stack	0.60	0.0477
2:1	1:1	Stack	0.60	0.1086
1:1	1:1	Stack	0.60	0.0491
1:2	1:1	Stack	0.60	0.0229
2:1	4:1	English	0.60	0.6000
1:1	4:1	English	0.60	0.5636
1:2	4:1	English	0.60	0.4108
2:1	2:1	English	0.60	0.4620
1:1	2:1	English	0.60	0.4097
1:2	2:1	English	0.60	0.3518
2:1	1:1	English	0.60	0.2954
1:1	1:1	English	0.60	0.2797
1:2	1:1	English	0.60	0.2736
2:1	4:1	Flemish	0.60	0.6000
1:1	4:1	Flemish	0.60	0.5886
1:2	4:1	Flemish	0.60	0.4236
2:1	2:1	Flemish	0.60	0.4883
1:1	2:1	Flemish	0.60	0.4535
1:2	2:1	Flemish	0.60	0.3699

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TABLE 5 – *Continued from previous page*

Panel ratio	Block ratio	Bond type	Friction ratio	Collapse multiplier
2:1	1:1	Flemish	0.60	0.3419
1:1	1:1	Flemish	0.60	0.3346
1:2	1:1	Flemish	0.60	0.3155
2:1	4:1	Running	1.00	1.0000
1:1	4:1	Running	1.00	0.8958
1:2	4:1	Running	1.00	0.4668
2:1	2:1	Running	1.00	1.0000
1:1	2:1	Running	1.00	0.6904
1:2	2:1	Running	1.00	0.4277
2:1	1:1	Running	1.00	0.6039
1:1	1:1	Running	1.00	0.4307
1:2	1:1	Running	1.00	0.3480
2:1	4:1	Stack	1.00	1.0000
1:1	4:1	Stack	1.00	0.3996
1:2	4:1	Stack	1.00	0.1326
2:1	2:1	Stack	1.00	0.5679
1:1	2:1	Stack	1.00	0.1591
1:2	2:1	Stack	1.00	0.0579
2:1	1:1	Stack	1.00	0.2161
1:1	1:1	Stack	1.00	0.0660
1:2	1:1	Stack	1.00	0.0263
2:1	4:1	English	1.00	1.0000
1:1	4:1	English	1.00	0.8574
1:2	4:1	English	1.00	0.4572
2:1	2:1	English	1.00	0.8358
1:1	2:1	English	1.00	0.6143
1:2	2:1	English	1.00	0.4172
2:1	1:1	English	1.00	0.4948
1:1	1:1	English	1.00	0.3899
1:2	1:1	English	1.00	0.3377
2:1	4:1	Flemish	1.00	1.0000
1:1	4:1	Flemish	1.00	0.8700
1:2	4:1	Flemish	1.00	0.4630
2:1	2:1	Flemish	1.00	0.9029
1:1	2:1	Flemish	1.00	0.7060

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TABLE 5 – Continued from previous page

Panel ratio	Block ratio	Bond type	Friction ratio	Collapse multiplier
1:2	2:1	Flemish	1.00	0.4329
2:1	1:1	Flemish	1.00	0.5906
1:1	1:1	Flemish	1.00	0.4853
1:2	1:1	Flemish	1.00	0.3861

605

606 **DATA AVAILABILITY STATEMENT**

607 Some or all data, models, or code generated or used during the study are available in
608 a repository or online in accordance with funder data retention policies (Jiménez Rios
609 et al. 2020b; Jiménez Rios et al. 2020a).

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