# Relative age effect in elite swimmers in U14 Czech Championship

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#### Abstract

PURPOSE The issue of the Relative Age Effect (RAE) has been studied in the theory of sports for more than 30 years. Most studies concentrate on team sports, while the area of some individual sports like swimming can be considered still underexplored.

METHODS The aim of our study was to verify the RAE in young elite swimmers (n = 198) who participated in Czech Republic U14 Championship (1) in male and female samples (2) according to swimming disciplines and distances (3) and performance (times in individual disciplines) between individual quartiles / semesters of birth. The analysis was performed with the use of adequate statistical (chi--square test, Kruskal-Wallis H test, Mann-Whitney U test) and effect size (effect size w index, eta-square test, effect size r index) tests.

RESULTS The results showed a different intensity of RAE sex-differences (male: w = 0.033; female: w = 0.006). In the division by the swimming disciplines and swimming distances, statistically significant values with large effect size were found in males in 50 m freestyle, 200 m individual medley, 100 m butterfly and 200 m butterfly. However, this did not apply for girls. Analysis of differences in performance showed a significant difference between the dependent variables (sex, distance, discipline) by different independent variables of quartile / semester of birth with large effect size only in cases of male 100 m breaststroke and female 200 m individual medley.

CONCLUSIONS The issue of RAE should be circulated among the coaches working with youth, athletes, sports organizations, but also parents of athletes in order to avoid the termination of actively spent time or drop-outs.

Key words: birthdate, performance, sport, swimming, sex, adolescence

# INTRODUCTION

In the field of sports sciences, the Relative Age Effect (RAE) term refers to the deviation of the distribution in birthdates of selected athletes from the normal distribution in the population (Delorme, Boiché & Raspaud, 2010). This means that the birthdates for a selected sample of athletes are not distributed proportionally, as are the birthdates for a corresponding segment of normal population, i.e. approximately evenly throughout the year. A higher frequency of athletes is, on the contrary, cumulated at the beginning of the selected time period (year, season), which means that athletes born at the beginning of the year/season are represented more often in the selected sample than athletes born in later months (Agricola, Zháněl & Hubáček, 2013; Lames, Augste, Dreckmann, Görsdorf & Schimanski, 2008). This fact can significantly affect the level of physical and performance preconditions in previously born athletes, especially in youth categories and, in particular, during the period of puberty, when the differences in anthropometric, resp. motor characteristics get even deeper (Cobley, Baker, Wattie & McKenna, 2009; Lames et al., 2008; Nykodým, Bozděch, Agricola & Zháněl, 2020).

In general, most studies have confirmed RAE mainly in youth categories and in top-level / elite sports (Cobley et al., 2009). The main interest of the authors of RAE studies in sports is mostly

focused on team sports, especially ice hockey (Bezuglov et al., 2020; Fumarco, Gibbs, Jarvis & Rossi, 2017; Nykodým et al., 2020) and soccer (Götze & Hoppe, 2020; Li et al., 2020; Roberts et al., 2020). However, we can find numerous RAE studies also in some individual sports, for instance in tennis (Moreira, Lopes, Faria & Albuquerque, 2017; Ulbricht, Fernandez-Fernandez, Mendez-Villanueva & Ferrauti, 2015; Wendling & Mills, 2018) and skiing (Bjerke, Lorås, Vorland & Pedersen, 2020; Müller, Gonaus, Perner, Müller & Raschner, 2017; Steidl-Müller, Hildebrandt, Raschner & Müller, 2019). On the other hand, we record only a small number of RAE studies in some (mainly individual) sports including swimming. Nevertheless, we can state from the conclusions of available studies that the existence of RAE has been proven in most cases in elite junior male and female swimmers, similarly as in other sports (Baker, Schorer & Cobley 2010; Cobley et al., 2009; Lames et al., 2008). We know from the conclusions of various studies that the occurrence of RAE in swimming is less frequent in girls / females than in boys/males (Baker et al., 2010; Cobley et al., 2009; Romann, Rössler, Javet & Faude, 2018) and the approximate age of 12-13 can be marked as the period of its strongest influence; later the RAE effect gradually and irregularly weakens, it may disappear completely, or so called reverse RAE may appear (Cobley et al., 2018, 2019). However, this does not mean that the relatively older swimmers achieve better specific performance (time in a given discipline) than their relatively younger peers (Costa et al., 2013). Based on the assessment through effect size tests, we can state that RAE is more pronounced with a higher level of performance, especially in the categories of 13–15 years (Cobley et al., 2019). Some studies have demonstrated RAE also in senior categories (Ferreira et al., 2017), which is usually the result of drop-out of swimmers already in junior categories and age bias environment.

As mentioned above, we still do not find a sufficient number of studies in some sports to analyse the issue of RAE in individual disciplines, age categories or performance levels which prevents a comprehensive understanding of RAE in the specific sport.

Since the conclusions of our detailed literary research imply that one of these sports is also swimming, we decided to turn our attention to this problem. The aim of our study was to find whether there existed any RAE: (1) in male and female participants of 2019 Czech swimming Championship in the U14 category, (2) in the division according to swimming disciplines and distances (3) and in relation to the performance between individual quartiles or semesters of birth, based on the times achieved in individual disciplines.

# **METHODS**

#### **Participants**

The participants of the presented descriptive research were swimmers (male n = 94; female n = 104) who participated in the Czech U14 Summer Masters Swimming Championship in Zlín (50 m pool) in 2019. Czech Masters Swimming Championship are organized by the Czech Swimming Federation (ČSPS) according to the valid rules of the FINA international swimming federation, adjusted for the situation of the Czech Republic. The condition of qualification is participation at regional Championship; for long distance swimming, at long-distance swimming regional pupil Championship, eventually participation in the Czech Long-Distance Swimming Cup (hereinafter referred to as qualifying races). At the Czech Masters Swimming Championship, the swimmers can compete only in the disciplines in which they have raced in qualifying races; and only in six disciplines. A maximum of 32 swimmers can participate in 50 m, 100 m, 200 m and 400 m disciplines; maximum of 24 swimmers in 800 m and 1500 m. Admission of a competitor to the race depends on qualifying time; the number of swimmers from one club is not limited.

# Procedures

Research variables (n = 26) consisted of 5 swimming disciplines (freestyle [FS], backstroke [BAS], butterfly [BF], breaststroke [BRS], individual medley [IM] and 6 different distances (50 m, 100 m, 200 m, 400 m, 800 m, 1500 m) by sex (male, female). As a categorical variable we chose, with regard to our research objective, dates of birth divided by quarters ( $Q_i$ ) into  $Q_1$  (January through March),  $Q_2$  (April through June),  $Q_3$  (July through September) and  $Q_4$  (October through December), see Table 1–3, or divided by semesters ( $S_i$ ) into  $S_1$  (January through June) and  $S_2$  (July through December), see Table 4.

# Statistical analysis

Research data were analysed using adequate methods of descriptive (absolute and relative frequency, median, lower and upper quartile) as well as inference (chi-square test, Kruskal-Wallis H test and Mann-Whitney U test) statistics and effect size tests (effect size w index, eta-square test and effect size r index). Chi-square goodness of fit test ( $\chi^2$ ) was used to verify the differences between the expected and observed birthdate distributions. The expected distribution was determined by days in  $Q_1$ :  $Q_1 = 90/365.25$  (24.6%);  $Q_2 = 91/365.25$  (24.9%);  $Q_3 = 92/365.25$ (25.2%); Q<sub>4</sub> = 92/365.25 (25.2%). Threshold values for small (w = 0.10), medium (w = 0.30), large (w = 0.50) effect were used to assess the effect size (ES) w index (Cohen, 1988). The significance of the differences between the independent variable (Q) and dependent variables (sex, swimming styles, and race distance) was assessed using Kruskal-Wallis H test and eta-squared effect size test  $(\eta^2)$ , together with 95 % confidence interval (CI). The threshold values for  $\eta^2$  (Cohen, 1988) are small ( $\eta^2 = 0.01$ ), medium ( $\eta^2 = 0.06$ ) and large ( $\eta^2 = 0.14$ ) effects. Mann-Whitney U test was used in case if independent variables  $(Q_i)$  contained less than 5 swimmers; in these cases, we chose semesters (S) instead of the Q independent variables. To assess effect size, we used effect size r index for Mann-Whitney U test, which can be interpreted as a small (r = 0.10), medium (r = 0.30), or large (r = 0.50) effect (Cohen, 1988). All the values of ES indexes (w,  $\eta^2$ , r) smaller than small effect were marked as trivial (Cohen, 1988) and we transformed them into the Common Language Effect Size, (CLES), for their better interpretability and generalizability (Cohen, 1988; Dunlap, Cortina, Vaslow, & Burke, 1996; McGraw & Wong, 1992). Statistical calculations were performed using licensed IBM SPSS Statistics for Windows software (IBM Corp, Armonk, New York, USA, v. 25.0). The threshold at  $\alpha = 0.05$  was chosen as the level of statistical significance.

# RESULTS

Table 1 contains the results of an analysis of birthdate distribution and the evaluation of RAE for the whole research group and its effect on sex. Table 2 further includes analyses of birthdate distribution and an assessment of the level of RAE effect for individual research variables, performed with the use of test in terms of statistical significance and magnitude of effect size. The following Table 3 and 4, however, do not give the numbers of frequencies as in the case of Table 1 and 2, but the final times of the competitors [MM:SS.SS]. This made it possible to verify the presumption whether the relatively older swimmers performed better, in the form of better time, because of their biological advantage.

Sex -		Birthdate	n	$\chi^2$	р	W		
	Q <sub>1</sub> (%)	Q <sub>2</sub> (%)	Q <sub>3</sub> (%)	Q <sub>4</sub> (%)	-			
Male	35 (37.23)	22 (23.40)	22 (23.40)	15 (15.96)	94 (100%)	9.846	0.020	0.033
Female	28 (26.92)	27 (25.96)	25 (24.04)	24 (23.08)	104 (100%)	0.385	0.943	0.006
Total	63 (31.82)	49 (24.75)	47 (23.74)	39 (19.70)	198 (100%)	6.600	0.086	0.013

Table 1. Differences between the Expected and Observed Distribution of Swimmers' Birthda	late
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Note:  $Q_{1-4}$ : quarter of birth;  $\chi^2$ : chi-square goodness of fit test; w: effect size w index (Cohen's w)

**Table 2.** Differences between the Expected and Observed Distribution of Swimmers' Birthdates in Sex,

 Distance and Discipline

	0	Birthdate quarters								
Sex	Distanc	Discipli	Q <sub>1</sub> (%)	Q <sub>2</sub> (%)	Q <sub>3</sub> (%)	Q4(%)	n	χ <sup>2</sup>	р	W
М	50	FS	14 (45.16)	8 (25.81)	8 (25.81)	1 (3.23)	31 (100%)	10.625	0.014	0.585
М	100	FS	14 (43.75)	7 (21.88)	7 (21.88)	4 (12.50)	32 (100%)	6.750	0.080	0.459
М	100	BAS	11 (34.38)	8 (25.00)	7 (21.88)	6 (18.75)	32 (100%)	1.750	0.626	0.234
М	100	BF	16 (50.00)	11 (34.38)	4 (12.50)	1 (3.13)	32 (100%)	17.250	0.001	0.734
М	100	BRS	8 (34.78)	4 (17.39)	7 (30.43)	4 (17.39)	23 (100%)	2.167	0.539	0.307
М	200	FS	10 (31.25)	5 (15.63)	10 (31.25)	7 (21.88)	32 (100%)	2.250	0.522	0.265
М	200	BF	13 (40.63)	11 (34.38)	6 (18.75)	2 (6.25)	32 (100%)	9.250	0.026	0.538
М	200	IM	13 (40.63)	7 (21.88)	10 (31.25)	2 (6.25)	32 (100%)	8.250	0.041	0.508
М	200	BAS	11 (34.38)	6 (18.75)	6 (18.75)	9 (28.13)	32 (100%)	2.250	0.522	0.265
М	200	BRS	10 (31.25)	9 (28.13)	9 (28.13)	4 (12.50)	32 (100%)	2.750	0.432	0.293
М	400	FS	13 (40.63)	6 (18.75)	5 (15.63)	8 (25.00)	32 (100%)	4.750	0.191	0.385
М	400	IM	9 (29.03)	10 (32.26)	7 (22.58)	5 (16.13)	31 (100%)	1.875	0.600	0.246
М	1500	FS	6 (26.09)	6 (26.09)	4 (17.39)	7 (30.43)	23 (100%)	0.833	0.841	0.190
F	50	FS	9 (28.13)	11 (34.38)	7 (21.88)	5 (15.63)	32 (100%)	2.500	0.475	0.280
F	100	FS	11 (34.38)	8 (25.00)	9 (28.13)	4 (12.50)	32 (100%)	3.250	0.355	0.319
F	100	BAS	10 (31.25)	11 (34.38)	7 (21.88)	4 (12.50)	32 (100%)	3.750	0.290	0.342
F	100	BF	11 (34.38)	7 (21.88)	3 (9.38)	11 (34.38)	32 (100%)	5.500	0.139	0.415
F	100	BRS	7 (21.88)	7 (21.88)	9 (28.13)	9 (28.13)	32 (100%)	0.500	0.919	0.125
F	200	FS	11 (34.38)	7 (21.88)	8 (25.00)	6 (18.75)	32 (100%)	1.750	0.626	0.234
F	200	BF	8 (25.00)	10 (31.25)	3 (9.38)	11 (34.38)	32 (100%)	4.750	0.191	0.385
F	200	IM	7 (21.88)	9 (28.13)	8 (25.00)	8 (25.00)	32 (100%)	0.250	0.969	0.088
F	200	BAS	11 (34.38)	8 (25.00)	8 (25.00)	5 (15.63)	32 (100%)	2.250	0.522	0.265
F	200	BRS	7 (21.88)	6 (18.75)	11 (34.38)	8 (25.00)	32 (100%)	1.750	0.626	0.234
F	400	FS	11 (34.38)	7 (21.88)	10 (31.25)	4 (12.50)	32 (100%)	3.750	0.290	0.342
F	400	IM	7 (21.88)	9 (28.13)	9 (28.13)	7 (21.88)	32 (100%)	0.500	0.919	0.125
F	800	FS	6 (25.00)	5 (20.83)	6 (25.00)	7 (29.17)	24 (100%)	0.333	0.954	0.118

**Note:** M: male; F: female; FS: freestyle; BAS: backstroke; BF: butterfly; BRS: breaststroke; IM: individual medley;  $Q_{1-4}$ : quarter of birth;  $\chi^2$ : chi-square goodness of fit test; w: effect size w index (Cohen's w)

		Je		Time [M	M:SS.SS]		n	Н	р	η <sup>2</sup> (95 % CI)
Sex	Distance	Discipli	$Q_1 [med (x_{25} - x_{75})]$	$Q_2 [med (x_{25} - x_{75})]$	$Q_3 [med (x_{25}-x_{75})]$	$Q_4 [med (x_{25}-x_{75})]$	-			
М	100	BAS	1:11.09 (1:07:27-1:12.26)	1:11.07 (1:06.11-1:12.47)	1:08:00 (1:06:00-1:15:04)	1:10:03 (1:09:03-1:12:47)	32	0.306	0.959	0.096 (0.000-0.254)
М	200	FS	2:12.90 (2:10.20-2:16.60)	2:10.0 (2:04.7-2:11.6)	2:12.8 (02:09.8–2:15.2)	2:13.40 (2:10.50-2:17.00)	32	4.101	0.251	0.039 (0.000-0.150)
М	200	BAS	2:30.00 (02:26.00-02:36.00)	2:30.00 (2:22.00-2:40.00)	2:33.00 (2:21.00-2:40.00)	2:31.00 (2:28.00-2:37.00)	32	0.236	0.972	0.099 (0.000-0.259)
М	400	FS	4:42.00 (4:35.5-4:47.5)	4:37.00 (4:30.00-4:39.20)	4:33.00 (4:33.00-4:41.50)	4:43.50 (4:33.70-4:48.80)	32	5.931	0.115	0.105 (0.000-0.267)
М	400	IM	5:18.00 (5:11.50-05:36.50)	5:30.50 (5:17.50-5:42.30)	5:18.00 (5:09.00-5:28.00)	5:37.00 (5:16.50-5:43.69)	31	3.928	0.269	0.034 (0.000-0.137)
F	50	FS	0:29.00 (0:28.00-0:30.00)	0:30.00 (0:29.00-0:30.00)	0:29.00 (0:28.00-0:30.00)	0:30.00 (0:29.00-0:30.00)	32	0.903	0.825	0.075 (0.000-0.222)
F	100	BRS	1:23.00 (1:20.00-1:26.00)	1:23.00 (1:17.00-1:26.00)	1:22.00 (1:16.00-1:26.50)	1:22.00 (1:20.50-1:26.50)	32	0.225	0.973	0.099 (0.000-0.259)
F	200	FS	2:22.00 (2:16.00-2:27.00)	2:24.00 (2:21.00-2:24.00)	2:23.50 (2:17.80-2:27.80)	2:22.00 (2:20.50-2:25.80)	32	0.742	0.863	0.081 (0.000-0.232)
F	200	IM	2:35.00 (2:34.00-2:42.00)	2:41.00 (2:39.00-2:45.00)	0:02:36 (0:02:35-0:02:40)	0:02:41 (0:02:37-0:02:47)	32	6.973	0.073	0.142 (0.000-0.314)
F	200	BAS	2:40.00 (2:33.00-2:46.00)	2:39.00 (2:38.00-2:39.80)	2:39.00 (2:34.50-2:43.00)	2:43.00 (2:38.00-2:47.00)	32	2.336	0.506	0.024 (0.000-0.105)
F	200	BRS	3:04.00 (2:55.00-03:07.00)	2:58.00 (02:49.00-03:02.00)	2:58.00 (2:53.00-3:03.00)	03:05.00 (2:54.00-3:08.00)	32	3.250	0.355	0.009 (0.000-0.016)
F	400	IM	5:37.00 (5:22.00-5:59.00)	5:50.00 (5:40.00-5:56.50)	5:35.00 (5:26.50-5:46.00)	5:50.00 (5:46.00–5:59.00)	32	6.759	0.080	0.134 (0.000-0.305)
F	800	FS	10:13.00 (09:47.50-10:28.30)	10:32.00 (10:07.00-10:48.00)	10:00.00 (9:51.00-10:36.00)	10:32.00 (10:24.00-10:36.00)	24	4.121	0.249	0.040 (0.000-0.157)

**Table 3.** Differences between Swimmers' Time according to Quarters of Birthdate in Sex, Discipline and Distance

*Note:* M: male; F: female; FS: freestyle; BAS: backstroke; BF: butterfly; BRS: breaststroke; IM: individual medley; Med: median;  $x_{25}$ : lower quartile,  $x_{75}$ : upper quartile;  $Q_{1-4}$  quarter of birth; H: Kruskal-Wallis H test;  $\eta^2$ : Eta-square test; CI: confidence interval

Although the Table 1 shows a gradual decrease in the number of both absolute and relative frequencies from  $Q_1$  to  $Q_4$  in all the above given variables (male, female, total), which indicates the existence of RAE, we calculated – assessing the values of chi-square goodness of fit test – that comparing the occurrence of expected and observed birthdate (in  $Q_i$ ) showed a statistically significant deviation from the birthdate distribution only in male sample ( $\chi^2(3) = 9.846$ , p = 0.020, w = 0.033, ES = trivial).

In terms of assessing the effect size w index values, we can conclude that the influence of birthdate-showed in male (w = 0.033, CLES = 0.51) as well as in female swimmers (w = 0.006, CLES = 0.50) a trivial effect ( $w_{diff}$  = 0.027). Which are the smallest differences of ES values, and therefore the most homogenous effect in our study.

Since, in terms of statistical significance, different levels of RAE effect were found between the sexes, it was important to find whether and how this RAE effect would be reflected also in the division by disciplines and distances (Table 2), or by swimmers' performance (Table 3 and 4).

The chi-square goodness of fit test was calculated comparing the occurrence of expected and observed birth dates (in  $Q_i$ ). Uneven distribution of the dates of birth – i.e. the existence of RAE – was demonstrated in male swimmers using statistical significance in 4 of 13 cases (30.77 %;

2times in BF, once in FS and IM). In female swimmers, we found statistically insignificant differences in birthdate frequency, so in this case we rejected the existence of RAE.

The values of ES w index in male swimmers ranged from small (w = 0.190, CLES = 0.51) to large (w = 0.734, CLES = 0.61) effect ( $w_{diff}$  = 0.544) in 13 swimming disciplines, which is the largest difference of ES values in our study, and thus the most heterogeneous effect. More precisely, we found a large effect in two disciplines (15.38 %; FS and BF), medium effect in five disciplines (38.46 %; 2times in FS and once in BRS, BF and IM), small effect in 6 disciplines (46.15 %; 3times in FS, 2times in BAS, once in BRS) and trivial effect in two disciplines (15.38 %; FS and BF).

In female swimmers, the values of ES w index ranged from trivial (w = 0.088, CLES = 0.51) to medium (w = 0.415, CLES = 0.57) effect ( $w_{diff} = 0.327$ ). More strictly, we found a medium effect in five disciplines (38.46%; 2times in FS and BF and once in BAS), small effect in seven disciplines (53.85 %; 2times in FS and BRS, once in BAS) and trivial effect in one discipline (7.69 %; IM).

After assessing the existence of RAE (Table 1 and 2), we were interested in whether there existed performance differences (times achieved in individual disciplines) between quartiles ( $Q_i$ ) in individual swimming disciplines. In this case, there was not a statistically significant difference between the dependent variables (sex, distance, style). This means that there was not any statistical difference in swim time between swimmers born in different  $Q_i$ . Which can be interpreted that there is no difference between relatively older and younger swimmers in performance; not even in the groups where we confirmed the effect of RAE (Table 2). The values of ES  $\eta^2$  index in male swimmers in 5 swimming disciplines reached medium ( $\eta^2 = 0.034 - 0.105$ , CLES = 0.56 - 0.61) effect ( $\eta^2_{diff} = 0.071$ ). These are more homogenous ES values than in swimmers in Table 2.

In female swimmers, we found in 8 disciplines the values of ES  $\eta^2$  index ranging from trivial ( $\eta^2 = 0.009$ , CLES = 0.53) to large ( $\eta^2 = 0.142$ , CLES = 0.62) effect ( $\eta^2_{diff} = 0.133$ ), which were more homogenous results than in the case of female swimmers from Table 2. We found, more precisely, a large effect in one discipline (12.50 %; IM), medium effect in 4 disciplines (50.00 %; 2times in FS, once in BRS and IM), small effect in two disciplines (25.00 %; BAS and FS) and trivial effect in one discipline (12.5 %; BRS). As some research variables did not meet the conditions of Kruskal-Wallis H test (more than 5 swimmers in group  $[Q_i]$ ), the following Table 4 contains the results of Mann-Whitney U test, which we used to assess the categorial variable of birth semester (S<sub>i</sub>) instead of birth quartile (Q<sub>i</sub>).

The results of Mann-Whitney U test in male swimmers showed that the difference between the achieved times for  $S_1$  and  $S_2$  variables were statistically significant in two cases (BF and BRS, both on 100 m distance); once in case of girls (100 m BAS). In other cases, we did not prove any differences between times.

ES index r values in male swimmers in 8 swimming disciplines reached small (r = 0.116, CLES = 0.54) to large (r = 0.576, CLES = 0.70) effect ( $r_{diff}$  = 0.460). We found – more precisely – a large effect in one discipline (12.50 %; BRS), medium effect in three disciplines (37.50 %; BF, BRS and FS) and small effect in 4 disciplines (50.00 %; 2times in FS and once in BF and IM). ES r index values in female swimmers in 5 swimming disciplines reached trivial (r = 0.060, CLES = 0.52) to medium (r = 0.351, CLES = 0.61) effect ( $r_{diff}$  = 0.291). We found – more precisely – a medium effect in one discipline (20.00 %; BAS), small effect in two disciplines (40.00 %; FS and BF) and trivial effect also in two disciplines (40.00 %; BF and FS).

#### DISCUSSION

Based on the results of the presented study, we cannot describe the RAE influence in the examined group as unambiguous because the level of its effect depends very much on the nature of dependent variable. Similar conclusions were reached also by Buhre and Tschernij (2018) due to

Sex	Distance	Discipline	Time [MM:SS.SS]		$n_{s1}/n_{s2}$	U	р	r
			$S_{1}$ [Med $(x_{25} x_{75})$ ]	$S_{2} [Med (x_{25} x_{75})]$	-			
М	50	FS	0:28.10	0:27.30	22/9	79.0	0.384	0.156
М	100	FS	(0.20.90 - 0.28.30) 1:01.00 (0.59.00 + 1.02.00)	(0:27.10-0:27.80) 1:01.00 (1:00.00, 1:02.00)	21/11	99.0	0.513	0.116
М	100	BF	(0.39.00 - 1.02.00) 1:09.10 (1.06.20 - 1.12.10)	(1.00.00 - 1.02.00) 1:04.30 (1:01.70 - 1:07.60)	27/5	25.5	0.029	0.385
М	100	BRS	1:11.50 (1:09.50-1:12.70)	1:14.00 (1:13.00-1:16.00)	12/11	21.5	0.006	0.576
М	200	BF	2:40.5 (2:33.30-2:52.50)	2:30.00 (2:22.30–3:00.00)	24/8	72.0	0.296	0.185
М	200	IM	2:28.30 (2:27.00–2:33.52)	2:28.30 (2:25.00-2:31.00)	20/12	103.0	0.507	0.117
М	200	BRS	2:52.00 (2:42.00-2:59.00)	2:56.00 (2:48.50-3:06.00)	19/13	73.5	0.055	0.340
М	1500	FS	18:46.50 (18:15.30–19:12.70)	18:18.00 (17:53.00–18:44.00)	12/11	40.0	0.109	0.334
F	100	FS	1:05.00 (1:03.00-0:01:06)	1:06.00 (1:03.00-1:07.00)	19/13	96.0	0.291	0.186
F	100	BAS	1:13.00 (1:11.00-1:16.00)	1:15.00 (1:14.00-1:17.00)	21/11	65.5	0.047	0.351
F	100	BF	1:15.00 (1:11.00-1:19.00)	1:17.00 (1:14.00-1:18.00)	18/14	117.0	0.732	0.060
F	200	BF	2:51.50 (2:47.00-3:00.30)	2:55.00 (2:53.00-3:01.00)	18/14	93.0	0.209	0.222
F	400	FS	5:01.50 (4:55.50-5:05.20)	5:01.50 (4:55.00-5:07.50)	18/14	117.5	0.747	0.057

**Table 4**. Differences between Swimmers' Time according to Semesters of Birthdate in Different Sexes, Distances and Disciplines

*Note:* M: male; F: female; FS: freestyle; BAS: backstroke; BF: butterfly; BRS: breaststroke; IM: individual medley; Med: median;  $x_{25}$ : lower quartile;  $x_{75}$ : upper quartile;  $S_i$ : semester of birth; U: Mann-Whitney U test; r: effect size r index; CI: confidence interval

inconclusive evidence of RAE occurrence in elite Swedish swimmers. Although there are significantly less academic studies dealing with RAE issue in female/girls than number of studies dealing with male/boys, it is clear – despite this disparity – that the level of RAE in female/girl sample is lower (or none) than in the case of males/boys, which is also valid in all comparable sports (Cobley et al., 2009; Smith, Weir, Till, Romann & Cobley, 2018). We reached similar conclusions as well for our research group in most of the examined variables, like Cobley et al. (2018) in the group of Australian swimmers or Costa et al. (2013) in Portuguese swimmers. Most common reasoning explaining this trend is a lower competition in the membership base (so the coaches cannot give preference to the temporarily biologically more mature girls) and earlier completion of the adolescence process in girls in comparison with boys (Baxter-Jones, Helms, Maffulli, Baines-Preece & Preece, 1995). The only conclusion we found that the RAE was higher in female swimmers than in the group of males was in Ferreira et al. (2017) study, where the authors had found a larger and statistically significant RAE effect in the group of female Olympic swimmers, but a statistically insignificant RAE effect in the male group. The authors of this study did not find any other relationship between birthdates (in Q<sub>i</sub>) and winning medals or distribution of

athletes by continents (with the exception of Asia). These findings confirm the hypothesis that the RAE effect is most evident during adolescence and then gradually and irregularly weakens (Cobley et al., 2018; Ulbricht et al., 2015). However, we can find examples when RAE is also evident in Swimming Masters age categories, which is usually caused by a progressive increase of the RAE influence during each decade of life (Medic, Young, Starkes, Weir & Grove, 2009) and thus a possible drop-out of relatively younger athletes in junior categories, where each subsequent age category created a stronger environment of age bias because of the departure of relatively younger athletes.

During our extensive literary research, we had not found a study which would deal with the effect of RAE in various disciplines (n = 6), distances (n = 6) and sex (male, female), which – in our case – consisted of a total of 26 different research variables. If we compare our results with Australian U16 swimmers (Cobley et al., 2018), we find that the ES values of our research group were comparable (small ES) 2times for the same variables (male 200 m breaststroke, female 50 m freestyle); 3times we found higher ES values (male 400 m freestyle, male 100 m breaststroke, female 400 m freestyle) and 3-times significantly higher ES values (male 50 m freestyle, female 200 m breaststroke, fe

As Cobley et al. (2018) monitored the effect of RAE in several age categories (U13 to U19), we can state, based on their results, that the most obvious effect of RAE was in younger swimmers and its intensity both in boys and girls gradually and irregularly decreased. We can conclude then that the period with the greatest RAE influence is about the age of 12. Similar conclusions were reached also by Costa et al. (2013) for elite Portugal swimmers. One explanation of this phenomenon may be that swimmers' growth curve indicates swimmers as early matures, compared to other young gymnastic, soccer, and tennis athletes (Baxter-Jones et al., 1995).

# **CONCLUSIONS**

Our recommendation for practice are similar to the recommendations of other authors (Baker et al., 2010; Cobley et al., 2009, 2018; Smith et al., 2018), i.e. focus on the prevention of the struggles of late-born children and the protection from possible discontinuation of active spending of leisure time. Awareness of RAE issues should be spread among the coaches working with (not only talented) youth athletes themselves (e.g. in the form of workshops), representatives of sports organizations (who can modify the sports system), but also parents of athletes. This should be spread along with the emphasis on the information that relatively older individuals have only a temporary biological advantage, which does not guarantee them an excellent performance after the completion of their adolescence. This way, the drop-out of relatively younger athletes at the beginning of their adolescence process should be reduced (first wave of drop-out caused by RAE) as well as the drop-out of relatively older athletes after the period of adolescence who, in turn, lost the temporary biological advantage resulting from their date of birth (second wave of drop-out caused by RAE).

#### REFERENCES

- Abbott, S., Moulds, K., Salter, J., Romann, M., Edwards, L., & Cobley, S. (2020). Testing the application of corrective adjustment procedures for removal of relative age effects in female youth swimming. *Journal of Sports Sciences*, 38(10), 1077–1084. https://doi.org/10.1080/02640414. 2020. 1741956
- Agricola, A., Zháněl, J., Hubáček, O. (2013). The Relative Age Effect in junior tennis (male). Acta Universitatis Palackianae Olomucensis Gymnica. 43(1), 27–33. https://doi.org/10.5507/ag. 2013.003
- Baker, J., Schorer, J., & Cobley, S. (2010). Relative age effects: An inevitable consequence of elite sport? Sportwissenschaft, 40(1), 26–30. https://doi.org/10.1007/s12662-009-0095-2
- Baxter-Jones, A. D. G., Helms, P., Maffulli, N., Baines-Preece, J. C., & Preece, M. (1995). Growth and development of Male gymnasts, swimmers, soccer and tennis players: A longitudinal study. *Annals of Human Biology*, 22(5), 381–394. https:// doi.org/10.1080/03014469500004072
- Buhre, T., & Tschernij, O. (2018). The impact of relative age on sampling and performance in Swedish age-group swimming system. *Sport Journal, October*, 1. http://search.ebscohost.com/login.aspx? direct=true&db=jlh&AN=132055052&sit e=ehost-live
- Bezuglov, E., Shvets, E., Lyubushkina, A., Lazarev, A., Valova, Y., Zholinsky, A., & Waskiewicz, Z. (2020). Relative Age Effect in Russian Elite Hockey. *The Journal of Strength & Conditioning Research*, 34(9), 2522–2527. https://doi.org/ 10.1519/ jsc.00000000003687
- Bjerke, Ø, Lorås, H., & Vorland Pedersen, A. (2020). Variations in the constituent year effect in Junior World Championship in alpine skiing: A window into relative development effects? *PloS one*, *15*(4), e0231384. https://doi.org/10.1371/ journal.pone.0231384
- Cobley, S., Abbott, S., Dogramaci, S., Kable, A., Salter, J., Hintermann, M., & Romann, M. (2018). Transient Relative Age Effects across annual age groups in National level Australian Swimming. *Journal of Science and Medicine in Sport*, 21(8), 839–845. https://doi.org/10.1016/j.jsams. 2017.12.008
- Cobley, S., Abbott, S., Eisenhuth, J., Salter, J., McGregor, D., & Romann, M. (2019). Removing relative age effects from youth swimming: The development and testing of corrective adjustment procedures. *Journal of science and medicine in sport*, 22(6), 735–740. https://doi.org/10.1016/j. jsams.2018.12.013
- Cobley, S., Baker, J., Wattie, N., & McKenna, J. (2009). Annual age-grouping and athlete development: A meta-analytical review of relative age effects in sport. *Sports Medicine*, *39*(3), 235–256. https://doi.org/10.2165/00007256-200939030-00005
- Cohen, J. (1988). Statictical power analysis for the behavioral sciences (2nd ed.). Lawrence Erbaum associates.
- Costa, A. M., Marques, M. C., Louro, H., Ferreira, S. S., & Marinho, D. A. (2013). The relative age effect among elite youth competitive swimmers. *European Journal of Sport Science*, 13(5), 437–444. https://doi.org/10.1080/17461391.2012.742571
- Delorme, N., Boiché, J., & Raspaud, M. (2010). Relative age effect in elite sports: Methodological bias or real discrimination? *European Journal of Sport Science*, 10(2), 91–96. https://doi.org/10.1080/17461390903271584
- Dunlap, W. P., Cortina, J. M., Vaslow, J. B., & Burke, M. J. (1996). Meta-analysis of experiments with matched groups or repeated measures designs. *Psychological Methods*, 1(2), 170–177. https://doi.org/10.1037/1082-989X.1.2.170
- Ferreira, R. M., Coelho, E. F., Morais, A. V. de, Werneck, F. Z., Tucher, G., & Lisboa, A. L. R. (2017). The relative age effect in olympic swimmers. Revista Portuguesa de Ciências Do Desporto, 2017(S2A), 104–114. https://doi.org/10.5628/rpcd.17.s2a.104
- Fumarco, L., Gibbs, B. G., Jarvis, J. A., & Rossi, G. (2017). The relative age effect reversal among the National Hockey League elite. *PloS one*, 12(8), e0182827. https://doi.org/10.1371/journal. pone.0182827
- Götze, M., & Hoppe, M. W. (2020). Relative age effect in elite German soccer: Influence of gender and competition level. Frontiers in Psychology, 11, 3725. https://doi.org/10.3389/fpsyg.2020.587023
- Lames, M., Augste, C., Dreckmann, Ch., Görsdorf, K., & Schimanski, M. (2008). Der "Relative Age Effect" (RAE): neue Hausaufgaben für den Sport. *Leistungssport*, 38(6), 4–9.
- Li, Z., Mao, L., Steingröver, C., Wattie, N., Baker, J., Schorer, J., & Helsen, W. F. (2020). Relative age effects in Elite Chinese soccer players: Implications of the 'one-child'policy. *Plos one*, *15*(2), e0228611. https://doi.org/10.1371/journal.pone.0228611
- McGraw, K. O., & Wong, S. P. (1992). A Common Language Effect Size Statistic. Psychological Bulletin, 111(2), 361–365. https:// doi.org/10.1037/0033-2909.111.2.361
- Medic, N., Young, B. W., Starkes, J. L., Weir, P. L., & Grove, J. R. (2009). Gender, age, and sport differences in relative age effects among US masters swimming and track and field athletes. *Journal of Sports Sciences*, 27(14), 1535–1544. https://doi. org/10.1080/02640410903127630
- Moreira, J. P. A., Lopes, M. C., Faria, L. O., & Albuquerque, M. R. (2017). Relative age effect and constituent year effect: An analysis of the International Tennis Federation ranking. *Journal of Physical Education*, 28. https://doi.org/10.4025/ jphyseduc.v28i1.2814
- Müller, L., Gonaus, C., Perner, C., Müller, E., & Raschner, C. (2017). Maturity status influences the relative age effect in national top-level youth alpine ski racing and soccer. *PloS One*, *12*(7), e0181810. https://doi.org/10.1371/journal.pone.0181810
- Nykodým, J., Bozděch, M., Agricola, A., & Zháněl, J. (2020). The Relative Age Effect, at the Ice Hockey World Championship (IHWC) in Years 2015 – 2017. *Journal of Human Kinetics*, 75(1), 150–159. https://doi.org/10.2478/hukin-2020-004
- Roberts, S. J., McRobert, A. P., Rudd, J., Enright, K., & Reeves, M. J. (2020). Research in Another un-Examined (RAE) context. A Chronology of 35 Years of Relative Age Effect Research in Soccer: Is it time to move on?. *Science and Medicine in Football*, 1–9. https://doi.org/10.1080/24733938. 2020.1841278

- Romann, M., Rössler, R., Javet, M., & Faude, O. (2018). Relative age effects in Swiss talent development–a nationwide analysis of all sports. *Journal of Sports Sciences*, *36*(17), 2025–2031. https://doi.org/10.1080/02640414.2018.1432964
- Smith, K. L., Weir, P. L., Till, K., Romann, M., & Cobley, S. (2018). Relative Age Effects Across and Within Female Sport Contexts: A Systematic Review and Meta-Analysis (1451–1478)). Sports Medicine, 48(8), 1989–1990. https://doi.org/10.1007/ s40279-018-0915-3
- Steidl-Müller, L., Hildebrandt, C., Raschner, C., & Müller, E. (2019). Challenges of talent development in alpine ski racing: a narrative review. *Journal of sports sciences*, 37(6), 601–612. https://doi.org/ 10.1080/02640414.2018.1513355
- Ulbricht, A., Fernandez-Fernandez, J., Mendez-Villanueva, A., & Ferrauti, A. (2015). The relative age effect and physical fitness characteristics in German male tennis players. *Journal of Sports Science and Medicine*, 14(3), 634–642.
- Wendling, E., & Mills, B. M. (2018). Mitigation of Relative Age Effects Through Targeted Policy Intervention: A Natural Experiment from Professional Tennis. Social Science Quarterly, 99(4), 1496–1509. https://doi.org/10.1111/ssqu.12516