Western University

Scholarship@Western

Paediatrics Publications

Paediatrics Department

5-1-2021

Predicting Adverse Outcomes for Shiga Toxin-Producing Escherichia coli Infections in Emergency Departments

Chu Yang Lin University of Alberta, Faculty of Medicine and Dentistry

Jianling Xie
Cumming School of Medicine

Stephen B. Freedman
Cumming School of Medicine

Ryan S. McKee University of Oklahoma Health Sciences Center

David Schnadower
University of Cincinnati College of Medicine

See next page for additional authors

Follow this and additional works at: https://ir.lib.uwo.ca/paedpub

Citation of this paper:

Lin, Chu Yang; Xie, Jianling; Freedman, Stephen B.; McKee, Ryan S.; Schnadower, David; Tarr, Phillip I.; Finkelstein, Yaron; Desai, Neil M.; and McKenzie, Roni D., "Predicting Adverse Outcomes for Shiga Toxin-Producing Escherichia coli Infections in Emergency Departments" (2021). *Paediatrics Publications*. 2357.

https://ir.lib.uwo.ca/paedpub/2357

Authors Chu Yang Lin, Jianling Xie, Stephen B. Freedman, Ryan S. McKee, David Schnadower, Phillip I. Tarr, Yaro Finkelstein, Neil M. Desai, and Roni D. McKenzie					



HHS Public Access

Author manuscript

J Pediatr. Author manuscript; available in PMC 2022 May 01.

Published in final edited form as:

J Pediatr. 2021 May; 232: 200–206.e4. doi:10.1016/j.jpeds.2020.12.077.

Predicting Adverse Outcomes for Shiga Toxin-Producing *E. coli* Infections in Emergency Departments

Chu Yang Lin, BMedSci¹, Jianling Xie, MD, MPH², Stephen B. Freedman, MDCM, MSc³, Ryan McKee, MD⁴, David Schnadower, MD, MPH⁵, Phillip I. Tarr, MD⁶, Yaron Finkelstein, MD⁷, Neil M. Desai, MBBCh⁸, Roni D. Lane, MD⁹, Kelly R. Bergmann, DO, MS¹⁰, Ron L. Kaplan, MD¹¹, Selena Hariharan, MD, MHSA⁵, Andrea T. Cruz, MD, MPH¹², Daniel M. Cohen, MD¹³, Andrew Dixon, MD¹⁴, Sriram Ramgopal, MD¹⁵, Elizabeth C. Powell, MD, MPH¹⁶, Jennifer Kilgar, MD¹⁷, Kenneth A. Michelson, MD, MPH¹⁸, Martin Bitzan, MD¹⁹, Kenneth Yen, MD, MS²⁰, Garth D. Meckler, MD, MSHS²¹, Amy C. Plint, MD, MSc²², Fran Balamuth, MD, PhD, MSCE²³, Stuart Bradin, DO²⁴, Serge Gouin, MDCM²⁵, April J. Kam, MD, MScPH²⁶, James Meltzer, MD, MS²⁷, Tracy E. Hunley, MD²⁸, Usha Avva, MD²⁹, Robert Porter, MD, MSc³⁰, Daniel M. Fein, MD³¹, Jeffrey P. Louie, MD³², Gillian A.M. Tarr, PhD³³ Pediatric Emergency Research Canada (PERC) and Pediatric Emergency Medicine Collaborative Research Committee (PEMCRC) STEC Study Group*

¹Faculty of Medicine and Dentistry, University of Alberta, Edmonton, Alberta, Canada ²Section of Pediatric Emergency Medicine, Department of Pediatric, Alberta Children Hospital, Cumming School of Medicine, University of Calgary, Calgary, Alberta ³Sections of Pediatric Emergency Medicine and Gastroenterology, Departments of Pediatrics and Emergency Medicine, Alberta Children's Hospital and Alberta Children's Hospital Research Institute, Cumming School of Medicine, University of Calgary, Calgary, Alberta ⁴Section of Pediatric Emergency Medicine, Department of Pediatrics, University of Oklahoma Health Sciences Center, Oklahoma City, OK ⁵Division of Emergency Medicine, Department of Pediatrics, Cincinnati Children's Hospital Medical Center, University of Cincinnati College of Medicine, Cincinnati, OH ⁶Division of Gastroenterology, Hepatology, & Nutrition, Department of Pediatrics, Washington University in St. Louis School of Medicine, St. Louis, MO ⁷Divisions of Emergency Medicine and Clinical Pharmacology & Toxicology, Hospital for Sick Children, University of Toronto, Toronto, Ontario ⁸British Columbia Children's Hospital Division of Pediatric Emergency Medicine ⁹Division of Pediatric Emergency Medicine, University of Utah School of Medicine, Salt Lake City, Utah ¹⁰Department of Emergency Medicine, Children's Minnesota ¹¹Department of Pediatrics, Division of Emergency Medicine, University of Washington School of Medicine, Seattle Children's Hospital, Seattle, WA ¹²Pediatric Emergency Medicine and Pediatric Infectious Diseases, Baylor College of Medicine ¹³Professor of Clinical Pediatrics, Nationwide Children's, Professor of Clinical

Corresponding author: Gillian A.M. Tarr, University of Minnesota, MMC 807, Room 1240, 420 Delaware St. SE, Minneapolis, MN, 55455; 612-626-9308; gtarr@umn.edu.

Publisher's Disclaimer: This is a PDF file of an unedited manuscript that has been accepted for publication. As a service to our customers we are providing this early version of the manuscript. The manuscript will undergo copyediting, typesetting, and review of the resulting proof before it is published in its final form. Please note that during the production process errors may be discovered which could affect the content, and all legal disclaimers that apply to the journal pertain.

^{*}List of additional members of the Pediatric Emergency Research Canada (PERC) and Pediatric Emergency Medicine Collaborative Research Committee (PEMCRC) STEC Study Group is available at www.jpeds.com (Appendix 1)

Pediatrics, Nationwide Children's Hospital & The Ohio State University, Columbus, OH 14University of Alberta, Stollery Children's Hospital, Women's and Children's Health Research Institute ¹⁵Division of Emergency Medicine, Ann & Robert H. Lurie Children's Hospital of Chicago, Northwestern University Feinberg School of Medicine, Chicago, Illinois ¹⁶Professor of Pediatrics, Northwestern University Feinberg School of Medicine; Ann & Robert H. Lurie Children's Hospital of Chicago: Division of Emergency Medicine ¹⁷Department or Pediatrics & Division of Emergency Medicine, Children's Hospital, Schulich School of Medicine and Dentistry, Western University, London, Ontario ¹⁸Division of Emergency Medicine, Boston Children's Hospital ¹⁹Division of Nephrology, Montreal Children's Hospital, McGill University Health Centre, Montreal, Canada; A1 Jalila Children's Hospital, Kidney Centre of Excellence, Dubai, UAE ²⁰Pediatric Emergency Medicine, Children's Medical Center, UT Southwestern ²¹Pediatrics and Emergency Medicine; University of British Columbia, Vancouver, BC ²²Departments of Pediatrics and Emergency Medicine, University of Ottawa and the Children's Hospital of Eastern Ontario, Ottawa, Ontario, Canada ²³Department of Pediatrics, University of Pennsylvania Perelman School of Medicine; Division of Emergency Medicine, Children's Hospital of Philadelphia ²⁴Children's Emergency Services, Department of Emergency Medicine, University of Michigan Medical School ²⁵Professor, Departments of Pediatric Emergency Medicine & Pediatrics, CHU Sainte-Justine, Universite de Montreal, QC, Canada ²⁶Department of Pediatrics, McMaster Children's Hospital, McMaster University ²⁷Division of Emergency Medicine; Department of Pediatrics; Jacobi Medical Center ²⁸Division of Pediatric Nephrology, Monroe Carell Jr. Children's Hospital at Vanderbilt, Nashville, Tennessee ²⁹Department of Pediatrics, Joseph M Sanzari Women and Children's Hospital, Hackensack University Medical Center 30 Discipline of Pediatrics, Memorial University of Newfoundland ³¹Division of Pediatric Emergency Medicine, Department of Pediatrics, Children's Hospital at Montefiore, Albert Einstein College of Medicine, Bronx, NY 32 Division of Emergency Medicine, University of Minnesota, Masonic Children's Hospital 33 Division of Environmental Health Sciences, University of Minnesota, Minneapolis, MN

Abstract

Objective—To assess the performance of a hemolytic uremic syndrome (HUS) severity score among children with Shiga toxin-producing *Escherichia coli* (STEC) infections and HUS by stratifying them according to their risk of adverse events. The score has not been previously evaluated in a North American acute care setting.

Study design—We reviewed medical records of children <18 years old infected with STEC and treated in one of 38 participating EDs in North America between 2011 and 2015. The HUS severity score [hemoglobin (g/dL) plus two-times serum creatinine (mg/dL)] was calculated using first available laboratory results. Children with scores >13 were designated as high-risk. We assessed score performance to predict severe adverse events (ie, dialysis, neurologic complication, respiratory failure and death) using discrimination and net benefit (i.e. threshold probability), with subgroup analyses by age and day-of-illness.

Results—A total of 167 children had HUS, of whom 92.8% (155/167) had relevant data to calculate the score; 60.6% (94/155) experienced a severe adverse event. Discrimination was acceptable overall (AUC 0.71, 95% CI 0.63, 0.79) and better among children <5 years old (AUC

0.77, 95% CI 0.68, 0.87). For children <5 years, greatest net benefit was achieved for a threshold probability >26%.

Conclusions—The HUS severity score was able to discriminate between high- and low-risk children <5 years old with STEC-associated HUS at a statistically acceptable level; however, it did not appear to provide clinical benefit at a meaningful risk threshold.

Keywords

hemolytic uremic syndrome; prognostic index; stx1; stx2

Children with Shiga toxin-producing *Escherichia coli* (STEC) may progress to develop hemolytic uremic syndrome (HUS),^{1, 2} which is characterized by azotemia or renal failure, microangiopathic hemolytic anemia, and thrombocytopenia.³ Though a subset of STEC infections are devastating,^{4, 5} most resolve without significant complications. Emerging approaches such as early-in-illness intravascular volume expansion^{6, 7} may have the potential to alter the disease trajectory. Therefore, early diagnosis and risk stratification of STEC-infected children at the first point-of-contact may indicate the need for closer monitoring for disease evolution and in the future such approaches may enable the provision of therapeutic interventions to improve outcomes.

Although STEC infections are rarely confirmed during an initial healthcare encounter due to the need to obtain a stool specimen and perform diagnostic testing, rapid molecular multiplex polymerase chain reaction assays are increasingly being employed in high-income countries. Simultaneously, given that nearly one-in-five children with high-risk STEC infections (i.e. Stx2-producing strains of STEC) develop HUS and its associated complications, institutions are increasingly adopting more standardized approaches to baseline and ongoing laboratory monitoring. These evolving strategies have highlighted the importance of being able to identify high-risk children to facilitate the selective monitoring and the provision of interventions to improve outcomes.

The ItalKid-HUS Network proposed an HUS severity score to predict severe adverse events in children with STEC-related thrombotic microangiopathy referred to tertiary HUS nephrology centers in Northern Italy. ¹¹ They proposed that severity could be predicted by the following equation: [(Hemoglobin in g/dL) + (serum creatinine in mg/dL [ISP]× 2)]. ¹¹ A cut-point of 13 identified those at high risk of severe adverse events with adequate discrimination [i.e. area under the curve (AUC) = 0.75]. ¹¹ However, the score was applied to a high-risk population and initial estimates of prognostic model performance are often overly optimistic. ¹² Our primary objective was to apply the ItalKID-HUS Network HUS severity score to children with STEC infections who developed HUS, who were enrolled in our pediatric emergency department (ED)-based study. As a secondary objective, we explored the performance of the severity score among all children with STEC, prior to HUS onset, to determine the score's potential for use earlier in illness.

Methods

For this retrospective study, participating sites were members of the Pediatric Emergency Medicine Collaborative Research Committee (PEMCRC) and/or Pediatric Emergency Research Canada (PERC). Study participants were children <18 years with microbiologic evidence of STEC infection, who visited one of 38 participating EDs in the United States and Canada between January 1, 2011 and December 31, 2015. Clinical findings, laboratory results, interventions, and complications from the ED visit, subsequent ED visits within 30 days, and inpatient data were extracted through chart review. The study received ethics approval at all institutions. ¹³

Cohort Definitions

For our primary objective, we defined the HUS cohort as participants with microbiologic evidence of STEC infection who met HUS criteria at any ED visit (i.e. index or follow-up) or during hospitalization. For the secondary objective, we defined the STEC cohort as participants with microbiologic evidence of STEC infection that had laboratory (i.e. hemoglobin and serum creatinine) testing performed and neither presented with nor developed HUS (Figure 1; available at www.jpeds.com). A child was considered to have HUS if, at any point in time, their platelets were <150,000/mm³ (<150 x 109/L), hematocrit was <30% (<0.3), and serum creatinine concentration was above the upper limit of normal for age. ¹⁴ We excluded participants whose serum hemoglobin and creatinine concentrations were not available.

Outcomes

As not all outcomes from the ItalKid-HUS Network study were available in our cohort, we focused on short-term outcomes that would be most applicable to the acute care and early decision-making context. An STEC-associated severe adverse event was defined by the initiation of renal replacement therapy, occurrence of a severe neurologic event (i.e. seizure or stroke), respiratory failure (i.e. intubation), or death. Patient-specific indications for utilizing interventions were not recorded.

HUS Severity Score

As employed in the ItalKid-HUS Network study, ¹¹ we calculated the severity score as the sum of the serum hemoglobin (g/dL) plus double the serum creatinine (mg/dL) (Supplementary Methods). Scores for participants in the HUS cohort were based on laboratory tests closest to, or on the day of development of HUS, but before dialysis commenced. Scores for participants in the STEC cohort were based on first-available laboratory collections. All serum creatinine and hemoglobin values used for score calculation were collected during an ED visit. Scores were dichotomized into high (>13) and low (13) risk groups per the ItalKid-HUS Network study. ¹¹ We excluded participants whose serum hemoglobin and creatinine concentrations were not available.

Statistical Analyses

The severity score was analyzed within the HUS and STEC cohorts using discrimination and decision curve analysis (DCA). DCA is an approach to evaluating prediction models that

balances trade-offs in clinical decision making. ¹⁵ For STEC-associated severe adverse events, we measured discrimination by sensitivity, specificity, receiver operating characteristics (ROC) curves, and AUC. An AUC >0.7 was *a priori* categorized as acceptable. ¹⁶

In the context of STEC infection, although treatment options are currently limited, early recognition of disease progression is vital to avoid children presenting with advanced renal failure and its associated complications (eg, electrolyte abnormalities, hypertension). There is also observational study evidence pointing towards benefits associated with early intravascular volume expansion^{17, 18} and other candidate interventions are undergoing evaluation. We used DCA to compare the net benefit of different clinical approaches to managing STEC patients; i.e. treating only patients the HUS severity score determines to be at high risk, treating all patients, and not treating any STEC patients. Net benefit was measured along a spectrum of risk strata (i.e. threshold probabilities)¹⁵, which is the risk level at which a clinician would treat to avoid one adverse event. For example, a threshold probability of 10% means that if a child has a 10% risk of a severe adverse outcome, the clinician would opt to treat. This method introduces the clinical context to the evaluation of prognostic indices and provides guidance to physicians with varying willingness for intervention. We identified the threshold probability levels for which the score provided a greater net benefit compared with the alternative approaches.¹⁵

In both cohorts, we conducted subgroup analyses by age, *a priori* specified as <5, 5 to <10, and 10 years (Supplementary Methods).²⁰ In the STEC cohort, we also analyzed score performance by day of illness, 3 days or >3 days to determine the applicability of the score to an acute care setting. Not all children with STEC infections are hospitalized. We did not extend the day-of-illness analysis to the HUS cohort because children identified with HUS would generally receive hospitalization on the day of diagnosis. If a child had serum creatinine and hemoglobin levels measured in the first three days of their illness, the first measure recorded was used to calculate a score, and the child was included in the 3 days of illness analysis. If a child had the necessary laboratory values measured after day 3 of their illness and HUS had not yet developed, the first values recorded on day 4 or after were used to calculate a score for inclusion in the >3 days of illness analysis. If a child had the necessary labs measured both within and after 3 days of illness, they were included in both day-of-illness analyses.

Analyses were performed using SPSS software (Version 25.0, IBM Corp.; Armonk, NY).

Results

In total, 927 children with microbiologically confirmed STEC were identified, 167 (18.0%) of whom met criteria for HUS; 92.8% (n=155) of children who developed HUS were included in the analysis of the HUS cohort (Figure 1). The HUS cohort had a mean score of 14.5 and a standard deviation of 3.3 (Figure 2). The remaining 7.2% (n=12) lacked sufficient laboratory data to calculate the severity score. The STEC cohort included 626 children and excluded 260.

HUS Severity Score

Of participants with HUS, 60.6% (94/155) had a severe adverse event, including neurologic complications in 26 (16.8%) and death in 2 (1.3%) (Table I; available at www.jpeds.com). Of the 26 children with neurologic complications, 23 (88.5%) received dialysis. The score classified 118 (76.1%) children as high-risk (Appendix 2, Table 1; available at www.jpeds.com).

Discrimination—For predicting the occurrence of at least one severe adverse event, the HUS severity score had an overall acceptable discrimination (AUC 0.71; 95% CI 0.63, 0.79) and sensitivity (89.4%, 95% CI 81.3%, 94.8%) (Table 2 and Figure 3, A). Stratifying by age, discrimination was greatest for children <5 years old (AUC 0.77; 95% CI 0.68, 0.87), with sensitivity of 86.8% (95% CI 74.7%, 94.5%) and specificity of 56.8% (95% CI 39.5%, 72.9%). Discrimination was greatly attenuated for children 10 years of age.

Net Benefit for Children <5 Years—The DCA for children <5 years of age showed that above a threshold probability of 26%, the highest net benefit is achieved by treating those with HUS severity score >13 (Figure 4, A). Clinicians wanting to treat at lower probabilities of a severe adverse event would find the greatest net benefit in a treat-all approach.

Extension to the STEC Cohort

Severe adverse events occurred in 11.8% (74/626) of children with STEC who did not have HUS at initial presentation (Table 1). Sensitivity (89.2%; 95% CI 79.8%, 95.2%) was comparable with that of the HUS cohort, but specificity (19.7%; 95% CI 16.5%, 23.3%) was lower (Table 2). The overall AUC for this group was poor (AUC 0.58; 95% CI 0.51, 0.65) but was acceptable for children <5 years old (AUC 0.75; 95% CI 0.66, 0.84) (Figure 3, B). DCA for those <5 years old showed the highest benefit is achieved using the HUS severity score between threshold probabilities of 6% and 22% (Figure 4, B). To justify a treat-none approach, clinicians would require a threshold probability of >22%, corresponding to a willingness to treat 4 or fewer children to prevent one severe event.

Day of Illness—There were 292 children with serum creatinine and hemoglobin measured on day 3 of illness, and 352 whose laboratory tests were performed on or after day four of illness. The score had the highest AUC when calculated using laboratory values measured on day 4 of illness or after among children <5 years old (AUC 0.86; 95% CI 0.79, 0.93) (Appendix 2, Table 2).

Discussion

This study evaluates the HUS severity score, which was developed in a tertiary care nephrology center, in an outpatient setting, to determine whether it can be adopted into practice in this setting as published. In our study population, the score had an overall high sensitivity (89%), which was offset by poor specificity (44%), yielding an AUC of 0.71. Discrimination was greater for children aged <5 years (AUC 0.77). Although the AUC showed statistical adequacy of distinguishing high- from low-risk children, our DCA analysis suggested that the clinical utility of the score may be limited in children <5 years of

age with HUS. The score only yielded greater net benefit than a treat-all approach if a clinician would be unwilling to intervene until a child had at least a 26% probability of experiencing a severe adverse event. If a child already has HUS, we find it unlikely many clinicians would wait till there is a one in 4 chance of a severe adverse event before acting, making a treat-all approach more beneficial in our population.

If a child is at high-risk of a severe adverse outcome, clinicians may consider several interventions, including admitting the child for observation and close laboratory monitoring or even intravascular volume expansion. Although there is some evidence that intravenous volume expansion^{17, 18} and avoidance of hemoconcentration⁶ may improve outcomes,⁷ there is not yet consensus on this approach and there exists the potential risk of fluid overload.²¹ Nonetheless, our DCA provides insights as to whether use of a severity score would aid in clinical decision-making regarding whether to utilize a more interventional approach. For example, a clinician may be willing to admit a child with even a small risk (e.g. 5%) of a severe adverse outcome, because the drawbacks of admission are minimal in comparison with the consequences of a seizure or death. However, the same clinician may be cautious when it comes to aggressive volume expansion and willing to act only if a child had at least a 20% risk of a severe adverse outcome.

The use of threshold probabilities has been explained in terms of the number of patients a clinician would be willing to treat to avoid one undesirable outcome. ²² Following this logic, in the context of STEC infection, a 5% threshold probability would correspond to a willingness to admit 20 children to prevent 1 child from suffering a severe adverse event under observation or treatment. In a recently conducted survey, pediatric emergency medicine physicians and nephrologists stated they would be willing to admit a median of 25 children to prevent HUS-associated dialysis in one child.²³ Given that the threshold probability at which the HUS severity score yielded the highest benefit was 26%, and that the surveyed clinicians would treat above a risk of 4%, ²³ a treat-all approach would be more beneficial than using the HUS severity score in children with HUS. However, the threshold probabilities of some clinicians may differ.

Although Ardissino et al identified the score cutoff of 13 empirically, ¹¹ an application of the score in Argentina yielded an optimal cutoff of 12.6.²⁴ Our goal was to determine whether the score, as published, could be applied by ED or other front-line clinicians. Given the relative rarity of STEC-associated HUS, individual physicians cannot conduct their own studies to determine their setting's optimal cutoff, so the cutoff value, like the score itself, must be generalizable. Our DCA results suggest that it may not be, because it does not provide clinical benefit at a reasonable threshold probability. Future studies should consider alternative cutoffs to optimize both discrimination and net benefit. Application in settings where children with HUS first contact the healthcare system, such as the ED, should be considered. Incorporating demographic factors (e.g. age)¹³ into prognostic index design would accommodate individual patient features, rather than relying on laboratory values alone. For example, an alternative HUS risk score was recently developed, which stratifies patient risks based on features such as age, presenting symptoms and location.²⁵

The reduced performance of the severity score in our HUS cohort relative to the derivation study¹¹ is not surprising. Prognostic indices are often over-fit to the data from which they were developed, and performance metrics should be adjusted for the resulting 'optimism'. 12 Additionally, several differences between the ItalKid-HUS Network study and our own, preclude a true external validation of the HUS severity score. They classified all children with microangiopathy as having HUS.¹¹ We used an alternate, more commonly accepted definition of HUS based on the triad of anemia, thrombocytopenia and renal insufficiency.²⁶ Long-term outcomes were not available for our cohort; we instead used several in-hospital severe adverse outcomes that clinicians may be concerned about averting. Our cohort was also older on average, which may explain the poorer performance in children 5 and older. Younger children have lower serum creatinine at baseline, and this measure may be less sensitive in older children with higher baseline levels. Additionally, nearly half of the children in the ItalKid-HUS Network study received early volume expansion, 11 which may have averted some of the severe adverse outcomes they were measuring and artificially lowered their specificity. This treatment anecdotally was routinely used at two of the 38 hospitals in our study, likely introducing little bias but also making our study less comparable with the derivation study.

We did not examine our data as a function of STEC genotype. The performance of the HUS severity score could differ for STEC expressing different sets of Shiga toxins. STEC that do not encode Shiga toxin 2 have little, if any, likelihood of causing HUS, and the virulence of those that encode Shiga toxin 2 is attenuated if genes encoding Shiga toxin 1 are also present. ^{25, 27}

For exploratory purposes, we applied the score to STEC-infected children without, or prior to, development of HUS. In this group, although discrimination was poor overall (AUC 0.58), it was acceptable among children <5 years of age (AUC 0.75). The greatest net benefit was obtained by using the score between threshold values of 6% and 22%, a range that may be appropriate for several clinical decisions. Discrimination improved when scores were calculated from laboratory values obtained later in the disease course (AUC 0.86). For children with STEC without HUS, pediatric emergency medicine physicians and nephrologists were inclined to treat a median of ten children to prevent one case of HUS (i.e. when risk is above 10%).²³ We found that at this risk threshold, the HUS severity score had the highest net benefit in children <5 years old in our STEC cohort. Although our results therefore suggest that the score could be useful prior to the development of HUS, the score relies on serum creatinine concentration, which may not be elevated at this stage of illness. Similarly, hemoconcentration might evolve later in the pre-HUS interval, as capillary leak develops. Additionally, 29% of children in the STEC cohort were excluded as they did not have sufficient laboratory testing performed. Although there were limitations to this exploratory analysis, and the HUS severity score was not designed for children without HUS, we believe its performance after the third day of illness, particularly, warrants further study.

The HUS severity score has the potential to help guide care in the outpatient setting, particularly as rapid molecular diagnostics become widespread. Although the score discriminated best between high- and low-risk children <5 years old with HUS on a

statistical level, it had limited clinical benefit, as a greater net benefit than a treat-all approach was only achieved when the risk of a severe adverse event exceeds 26%. Further refinement of the score may be necessary prior to broad clinical application, including extension to children with STEC but without HUS.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

Acknowledgements

We acknowledge the contributions of PERC/PEMCRC STEC Study Groups Members (Appendix 1) for their involvement in the original study. We also acknowledge non-group members Dr Roger Zemek, Dr Samina Ali, Dr Mandi Newton, Dr Naveen Poonai, Dr Maala Bhatt, Dr Kate Maki, Candice McGahern, Rebecca Emerton, Dr Rakesh D. Mistry, Dr Lise E. Nigrovic, Dr Paul L. Aronson, Dr Todd A. Florin, Dr Christopher M. Pruitt, Dr Amy D. Thompson, Victor Gonzalez, Dr Colleen K. Gutman, and Dr Angelica DesPain for their help in facilitating the conduct of the study.

Supported by the Cumming School of Medicine-Alberta Health Services Clinical Research Fund. S.F. is supported by the Alberta Children's Hospital Foundation Professorship in Child Health and Wellness. P.T. is supported by the Administrative and Resource Access Core of the Washington University Digestive Diseases Research Core Center (National Institutes of health [NIH] grant number P30DK052574). K.M. was funded by award 1K08 HS026503 from the Agency for Healthcare Research and Quality. A.P. is supported by the University of Ottawa Research in Pediatric Emergency Medicine. G.T. was supported by a Canadian Institutes of Health Research Banting Postdoctoral Fellowship, Alberta Innovates Health Programs Postgraduate Fellowship, and University of Calgary Eyes High Postdoctoral Fellowship. The funder/sponsor did not participate in the work. S.F. receives in-kind research support from BioMerieux and Luminex Corporation and is a consultant to Eligo Bioscience. P.T. is an unpaid member of the Data Safety Monitoring Board of Inmunova, consultant to Takeda Pharmaceuticals on childhood digestive disorders, and a consultant to, member of the Scientific Advisory Board of, and holder of equity in MediBeacon Inc for work that is unrelated to this study. He is also an inventor of a patent for which he might receive royalties for technology that is also unrelated to this study. E.P. has received grants from the Health Resources and Services Administration Washington, Boston, Chicago Network, the NIH–Emergency Department Probiotics, and the NIH-RNA Biosignatures. The other authors declare no conflicts of interest.

Appendix I –: List of Additional Members in the Pediatric Emergency Research Canada (PERC) and Pediatric Emergency Medicine Collaborative Research Committee (PEMCRC) STEC Study Group

There are no conflicts of interests for the following additional members of the Study Group:

Annie Rominger, MD, Division of Pediatric Emergency Medicine, Department of Pediatrics, University of Louisville, Kentucky.

Darcy Beer, MD, Division of Pediatric Emergency Medicine, Department of Pediatrics, Max Rady College of Medicine, University of Manitoba, Winnipeg, Canada.

Christopher M. Pruitt, MD, Division of Pediatric Emergency Medicine, Department of Pediatrics, University of Alabama at Birmingham.

Thomas J. Abramo, MD, Departments of Pediatrics and Emergency Medicine, University of Arkansas School of Medicine, Arkansas Children's Hospital Research Institute, Little Rock.

Abigail Schuh, MD, Division of Pediatric Emergency Medicine, Department of Pediatrics, Medical College of Wisconsin, Milwaukee.

John T. Kanegaye, MD, Department of Pediatrics, University of California, San Diego School of Medicine, La Jolla.

Nicholas E. Jones, MD, Division of Pediatric Emergency Medicine, Department of Pediatrics, Emory University, Children's Healthcare of Atlanta, Georgia.

Abbreviations:

AUC Area under the curve

CI Confidence interval

CKD Chronic kidney disease

DCA Decision curve analysis

ED Emergency department

Hb Serum hemoglobin

HUS Hemolytic uremic syndrome

LDH Lactate dehydrogenase

PEMCRC Pediatric Emergency Medicine Collaborative Research Committee

PERC Pediatric Emergency Research Canada

ROC Receiver operating characteristics

sCr Serum creatinine

STEC Shiga toxin-producing Escherichia coli

References

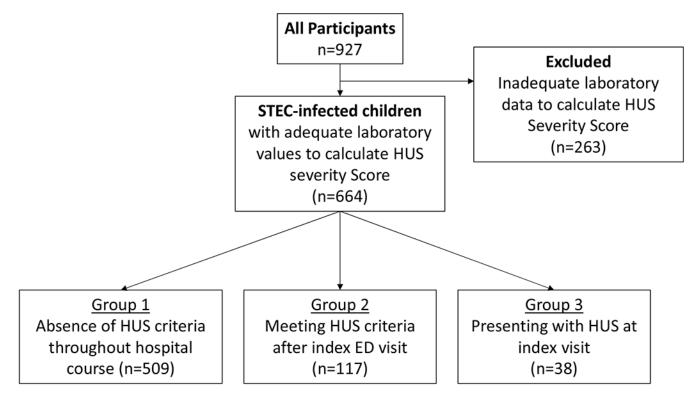
- Freedman SB, Xie J, Neufeld MS, Hamilton WL, Hartling L, Tarr PI, et al. Shiga Toxin–Producing Escherichia coli Infection, Antibiotics, and Risk of Developing Hemolytic Uremic Syndrome: A Meta-analysis. Clinical Infectious Diseases. 2016;62:1251–8. [PubMed: 26917812]
- 2. Noris M, Remuzzi G. Hemolytic Uremic Syndrome. Journal of the American Society of Nephrology. 2005;16:1035–50. [PubMed: 15728781]
- 3. Tarr GAM, Oltean HN, Phipps AI, Rabinowitz P, Tarr PI. Strength of the association between antibiotic use and hemolytic uremic syndrome following *Escherichia coli* O157:H7 infection varies with case definition. Int J Med Microbiol. 2018;308:921–6. [PubMed: 30257808]
- 4. Siegler R, Oakes R. Hemolytic uremic syndrome; pathogenesis, treatment, and outcome. Curr Opin Pediatr. 2005;17:200–4. [PubMed: 15800412]
- 5. Siegler RL, Milligan MK, Burningham TH, Christofferson RD, Chang SY, Jorde LB. Long-term outcome and prognostic indicators in the hemolytic-uremic syndrome. The Journal of pediatrics. 1991;118:195–200. [PubMed: 1993944]
- Grisaru S, Xie J, Samuel S, Hartling L, Tarr PI, Schnadower D, et al. Associations Between Hydration Status, Intravenous Fluid Administration, and Outcomes of Patients Infected With Shiga Toxin-Producing *Escherichia coli*: A Systematic Review and Meta-analysis. JAMA Pediatr. 2017;171:68–76. [PubMed: 27893870]

7. Ardissino G, Tel F, Possenti I, Testa S, Consonni D, Paglialonga F, et al. Early Volume Expansion and Outcomes of Hemolytic Uremic Syndrome. Pediatrics. 2016; 137. [PubMed: 27543009]

- 8. Tarr GAM, Tarr PI, Freedman SB. Clinical interpretation of enteric molecular diagnostic tests. Clinical microbiology and infection: the official publication of the European Society of Clinical Microbiology and Infectious Diseases. 2019;25:1454–6.
- 9. Tarr GAM, Stokowski T, Shringi S, Tarr PI, Freedman SB, Oltean HN, et al. Contribution and Interaction of Shiga Toxin Genes to Escherichia coli O157:H7 Virulence. Toxins (Basel). 2019; 11.
- Shane AL, Mody RK, Crump JA, Tarr PI, Steiner TS, Kotloff K, et al. 2017 Infectious Diseases Society of America Clinical Practice Guidelines for the Diagnosis and Management of Infectious Diarrhea. Clinical Infectious Diseases. 2017;65:e45–e80. [PubMed: 29053792]
- 11. Ardissino G, Tel F, Testa S, Paglialonga F, Longhi S, Martelli L, et al. A simple prognostic index for Shigatoxin-related hemolytic uremic syndrome at onset: data from the ItalKid-HUS network. Eur J Pediatr. 2018;177:1667–74. [PubMed: 30094644]
- Harrell FE Jr., Lee KL, Mark DB. Multivariable prognostic models: issues in developing models, evaluating assumptions and adequacy, and measuring and reducing errors. Stat Med 1996;15:361– 87. [PubMed: 8668867]
- 13. McKee RS, Schnadower D, Tarr PI, Xie J, Finkelstein Y, Desai N, et al. Predicting Hemolytic Uremic Syndrome and Renal Failure in Shiga Toxin-Producing Escherichia coli Infected Children. Clin Infect Dis. 2019.
- 14. Meites S, Buffone GJ. Pediatric clinical chemistry: reference (normal) values. Washington, D.C.: AACC Press; 1989.
- 15. Vickers AJ, Van Calster B, Steyerberg EW. Net benefit approaches to the evaluation of prediction models, molecular markers, and diagnostic tests. Bmj. 2016;352:i6. [PubMed: 26810254]
- Mandrekar JN. Receiver Operating Characteristic Curve in Diagnostic Test Assessment. Journal of Thoracic Oncology. 2010;5:1315–6. [PubMed: 20736804]
- Ake JA, Jelacic S, Ciol MA, Watkins SL, Murray KF, Christie DL, et al. Relative nephroprotection during Escherichia coli O157:H7 infections: association with intravenous volume expansion. Pediatrics. 2005;115:e673–80. [PubMed: 15930195]
- 18. Hickey CA, Beattie TJ, Cowieson J, Miyashita Y, Strife CF, Frem JC, et al. Early volume expansion during diarrhea and relative nephroprotection during subsequent hemolytic uremic syndrome. Arch Pediatr Adolesc Med. 2011; 165:884–9. [PubMed: 21784993]
- Mühlen S, Dersch P. Treatment Strategies for Infections With Shiga Toxin-Producing Escherichia coli. Front Cell Infect Microbiol. 2020;10:169. [PubMed: 32435624]
- 20. Center for Disease Control. National Enteric Disease Surveillance: Shiga toxin-producing Escherichia coli (STEC) Annual Report, 2015. 2015.
- Alobaidi R, Morgan C, Basu RK, Stenson E, Featherstone R, Majumdar SR, et al. Association Between Fluid Balance and Outcomes in Critically Ill Children: A Systematic Review and Metaanalysis. JAMA pediatrics. 2018;172:257–68. [PubMed: 29356810]
- 22. Vickers AJ, Elkin EB. Decision curve analysis: a novel method for evaluating prediction models. Medical decision making: an international journal of the Society for Medical Decision Making. 2006;26:565–74. [PubMed: 17099194]
- 23. Freedman SB, Grisaru S, Xie J, Samuel S, Dixon A, Plint AC, et al. Management of Shiga toxin producing Escherichia coli-infected children: A multi-national, multi-specialty survey. Journal of Paediatrics and Child Health. 2018;54:390–7. [PubMed: 29111613]
- 24. Balestracci A, Meni Battaglia L, Toledo I, Martin SM, Alvarado C. Comment to: "A simple prognostic index for Shigatoxin-related hemolytic uremic syndrome at onset: data from the ItalKid-HUS network" by Ardissino et al. Eur J Pediatr. 2018 8 10. doi: 10.1007/s00431-018-3198-7.Eur J Pediatr 2018;177:1869–70. [PubMed: 30194526]
- Hamilton D, Cullinan J. A practical composite risk score for the development of Haemolytic Uraemic Syndrome from Shiga toxin-producing Escherichia coli. Eur J Public Health. 2019;29:861–8. [PubMed: 31326985]
- 26. Gerber A, Karch H, Allerberger F, Verweyen HM, Zimmerhackl LB. Clinical course and the role of shiga toxin-producing Escherichia coli infection in the hemolytic-uremic syndrome in pediatric

patients, 1997-2000, in Germany and Austria: a prospective study. J Infect Dis. 2002;186:493–500. [PubMed: 12195376]

27. Tarr GAM, Stokowski T, Shringi S, Tarr PI, Freedman SB, Oltean HN, et al. Contribution and Interaction of Shiga Toxin Genes to Escherichia coli O157:H7 Virulence. Toxins. 2019;11.



Cohort	Included	Excluded
HUS (Groups 2, 3)	155	12
STEC (Groups 1, 2)	626	260

Figure 1. Patient cohorts and definitions.

Of 927 participants with microbiologically confirmed STEC, 664 had adequate clinical information to calculate severity scores. They were binned into STEC (n=626), and/or HUS (n=155) cohorts. Children may have been included in both cohorts.

Abbreviations: ED, emergency department; HUS, hemolytic uremic syndrome; STEC, Shiga toxin-producing *Escherichia coli*

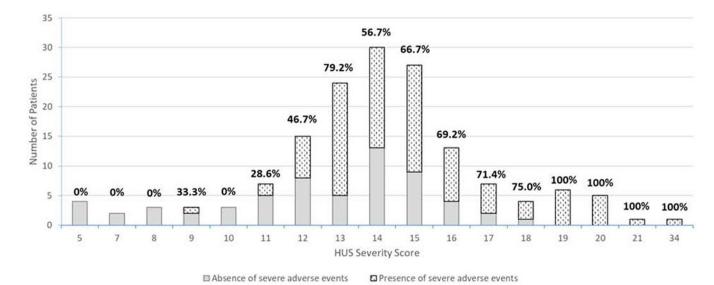


Figure 2.Distribution of HUS severity scores in the HUS cohort, with the frequency of STEC-associated severe adverse events in each score. The score was right skewed with a mean of 14.5 (SD of 3.3). Percentages denote the proportion of patients with severe adverse events. Abbreviations: HUS, hemolytic uremic syndrome; STEC, Shiga toxin-producing *Escherichia coli*

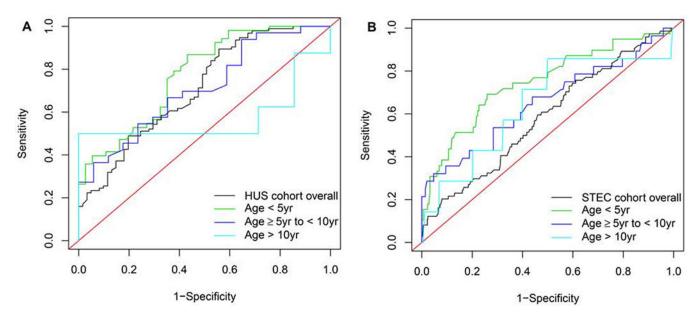


Figure 3. ROC curves for the HUS severity score in predicting any severe adverse event.(**A**) In the HUS cohort, the AUC was 0.71 (95% CI 0.63, 0.79) for all ages and 0.77 (95% CI 0.68, 0.87) for those <5 years old. (**B**) In the STEC cohort, overall, the AUC was 0.58 (0.51, 0.65), and for those <5 years old, the AUC was 0.75 (0.66, 0.84). Abbreviations: AUC, area under the curve; HUS, hemolytic uremic syndrome; ROC, receiver operating characteristics; STEC, Shiga toxin-producing *Escherichia coli*

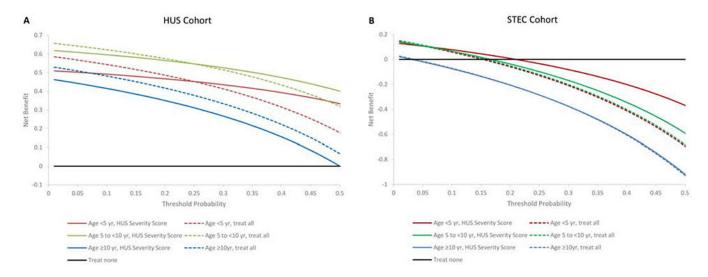


Figure 4. Decision curve analysis.

Net benefit is shown as a function of the risk of a severe adverse event that a clinician will tolerate (i.e. threshold probabilities) when deciding to treat. (A) For patients <5 years old in the HUS cohort, the HUS severity score had the greatest net benefit above a threshold of 26% (i.e. a clinician would treat once the probability of a severe adverse event reaches 26% in those treated). For clinicians who would want to treat at lower threshold probabilities, a treat-all approach had the greatest net benefit. (B) In the STEC cohort, the greatest net benefit is achieved for children <5 years old with the HUS severity score when clinicians would treat patients whose probability of a severe adverse event is 6-22%.

Abbreviations: HUS, hemolytic uremic syndrome; STEC, Shiga toxin-producing *Escherichia coli*

 Table 1.

 Demographic information and frequency of outcomes in the overall and individual cohorts.

	Overall Cohort	HUS Cohort	STEC Cohort
Characteristics (median (IQR) or n(%))			
Number of participants	927	155	626
Age (year)	6.0 (2.8, 11.0)	4.4 (2.4, 7.3)	6.4 (3.4, 11.7)
Male gender	490 (52.9%)	74 (47.7%)	319 (51.0%)
Hemoglobin (g/dL) ^a	13.4 (12.3, 14.6)	12.5 (10.0, 14.3)	13.6 (12.5, 14.7)
Hemoglobin (g/L) ^a	134.0 (123.0, 146.0)	125.0 (100.0, 143.0)	136.0 (125, 147)
Serum creatinine (mg/dL) ^a	0.45 (0.31, 0.64)	0.64 (0.37, 1.50)	0.42 (0.30, 0.60
Serum creatinine (micromol/L) ^a	39.8 (27.4, 56.5	56.5 (32.7, 132.6)	37.1 (26.5, 53.0)
Outcomes			
Any severe adverse event	100 (10.8%)	94 (60.6%)	74 (11.8%)
Need for dialysis	94 (10.1%)	89 (57.4%)	69 (11.0%)
Neurologic complications (seizure, stroke)	28 (3.0%)	26 (16.8%)	23 (3.7%)
Respiratory failure	26 (2.8%)	26 (16.8%)	22 (3.5%)
Death	2 (0.2%)	2 (1.3%)	2 (0.3%)
Neurologic complications and/or death	28 (3.0%)	26 (16.8%)	23 (3.7%)

Abbreviations: IQR, inter-quartile range; HUS, hemolytic uremic syndrome; STEC, Shiga toxin-producing Escherichia coli

^adata available for 664 patients.

 Table 2.

 Sensitivity and specificity for severe adverse events, stratified by age group in the HUS and STEC cohorts.

Age group (years)	Sensitivity (%) (95% CI)	Specificity (%) (95% CI)	AUC (95% CI)			
HUS Cohort						
Overall	89.4 (81.3, 94.8)	44.3 (31.5, 57.6)	0.71 (0.63, 0.79)			
0 to <5	86.8 (74.7, 94.5)	56.8 (39.5, 72.9)	0.77 (0.68, 0.87)			
5 to <10	93.9 (79.8, 99.3)	35.3 (14.2, 61.7)	0.71 (0.56, 0.86)			
10	87.5 (47.3, 99.7)	0.0 (0.0, 41.0)	0.57 (0.25, 0.89)			
STEC Cohort						
Overall	89.2 (79.8, 95.2)	19.7 (16.5, 23.3)	0.58 (0.51, 0.65)			
0 to <5	87.2 (72.6, 95.7)	40.7 (34.1, 47.6)	0.75 (0.66, 0.84)			
5 to <10	92.9 (76.5, 99.1)	12.2 (7.4, 18.5)	65.2 (52.8, 77.6)			
10	85.7 (42.1, 99.6)	1.6 (0.3, 4.6)	0.644 (0.41, 0.87)			

Abbreviations: AUC, area under the curve; CI, confidence interval; HUS, hemolytic uremic syndrome; STEC, Shiga toxin-producing Escherichia coli