Sum of a random number of independent random variables

February 9, 2022

1/15

Sum of a random number of independent random variables

Let N and $X_1, X_2, \ldots, X_n, \ldots$ be independent random variables such that:

- ► *N* takes nonnegative integer values.
- ▶ The variables X_i , $i \ge 1$, are identically distributed.

We are interested in the sum

 $S = X_1 + X_2 + \cdots + X_N$

in which the number N of terms is random. (We take S = 0 if N = 0 happens.)

Sum of a random number of independent random variables

The characteristic function of the random variable S can be obtained as follows:

$$\begin{split} M_{S}(\omega) &= \mathbb{E}\left(e^{i\omega S}\right) \\ &= \sum_{k \geq 0} \mathbb{E}\left(e^{i\omega S} \mid N = k\right) \mathbb{P}(N = k) \end{split}$$

Sum of a random number of independent random variables

Contents

Examples

Expected values

2/15

Sum of a random number of independent random variables

Observe that
$$\mathbb{E} \left(e^{i\omega S} \mid N = 0 \right) = \mathbb{E} \left(e^{i\omega 0} \right) = 1.$$

If $N = k, \ k \ge 1$, then
 $\mathbb{E} \left(e^{i\omega S} \mid N = k \right) = \mathbb{E} \left(e^{i\omega(X_1 + X_2 + \dots + X_N)} \mid N = k \right)$
 $= \mathbb{E} \left(e^{i\omega(X_1 + X_2 + \dots + X_k)} \mid N = k \right)$
 $= \mathbb{E} \left(e^{i\omega(X_1 + X_2 + \dots + X_k)} \right)$
 $= \mathbb{E} \left(e^{i\omega X_1} \right) \mathbb{E} \left(e^{i\omega X_2} \right) \cdots \mathbb{E} \left(e^{i\omega X_k} \right)$
 $= (M_X(\omega))^k$

5 / 15

Sum of a random number of independent random variables

Remark. A shortest way to formulate the previous calculation is as follows:

 $M_{S}(\omega) = \mathbb{E}\left(e^{i\omega S}\right) = \mathbb{E}\left(\mathbb{E}\left(e^{i\omega S}|N\right)\right),$

where

$$\mathbb{E}\left(\left.e^{i\omega S}\right|N\right)=(M_{X}(\omega))^{N}$$

Therefore

$$M_{S}(\omega) = \mathbb{E}\left(\left(M_{X}(\omega)\right)^{N}\right) = G_{N}(M_{X}(\omega))$$

Sum of a random number of independent random variables

Let $G_N(z) = \sum_{k \ge 0} \mathbb{P}(N = k) z^k$ be the probability generating function of N (considered here as a function of a complex variable z).

Then

$$M_{\mathsf{S}}(\omega) = \sum_{k \ge 0} \left(M_{\mathsf{X}}(\omega) \right)^k \mathbb{P}(\mathsf{N} = k) = G_{\mathsf{N}}\left(M_{\mathsf{X}}(\omega) \right)$$

Notice that the composition G_N (M_X(ω)) is well-defined, because |M_X(ω)| ≤ 1 for all ω ∈ ℝ and G_N(z) converges for all z ∈ ℂ such that |z| ≤ 1.

6/15

Example 1

- Let the number N of costumers arriving at a service point be a Po(λ)-distributed random variable.
- Each arrival is randomly and independently served with probability *p*.
 (Hence an arrival is not served with probability *q* = 1 *p*.)

Let us determine the probability distribution of the number ${\bf S}$ of served customers.

Example 1

We have

$$S=X_1+X_2+\cdots+X_N,$$

where X_i is the indicator random variable of the event "the *i*-th arrival is served". So $X_i \sim Be(p)$.

Moreover,

$$egin{aligned} M_{X_i}(\omega) &= q + p e^{i \omega}, \quad \omega \in \mathbb{R} \ & G_N(z) = e^{\lambda(z-1)}, \quad z \in \mathbb{C}, \end{aligned}$$

Example 1

Therefore

$$M_{S}(\omega) = G_{N}(M_{X}(\omega)) = e^{\lambda(z-1)}\Big|_{z=q+pe^{i\omega}}$$

 $=e^{\lambda(q+pe^{i\omega}-1)}=e^{\lambda p(e^{i\omega}-1)}$

This is the characteristic function of a Poisson random variable with parameter λp . Thus

 $S \sim \text{Po}(\lambda p)$

9/15

Example 1

Analogously, if *R* denotes the number of non-served customers, then $R \sim Po(\lambda q)$.

- ▶ It can be proved that *S* and *R* are independent variables.
- ► Notice how the convolution theorem applies. Indeed,

$$M_{S}(\omega) M_{R}(\omega) = e^{\lambda p(e^{i\omega}-1)} e^{\lambda q(e^{i\omega}-1)}$$
$$= e^{\lambda(p+q)(e^{i\omega}-1)} = e^{\lambda(e^{i\omega}-1)} = M_{N}(\omega),$$

in accordance with the fact that S + R = N.

Example 2

As a second example consider the following scenario.

- The number N of costumers arriving at a service point is a Ge(p)-distributed.
- The service times X_i, i ≥ 1, are independent and Exp(µ)-distributed random variables.

Let S be total time of occupancy of the service point,

 $S = X_1 + X_2 + \cdots + X_N$

10 / 15

Now

$$M_X(\omega) = rac{\mu}{\mu - i\omega}, \quad G_N(z) = rac{pz}{1 - qz}$$

Therefore

$$\begin{split} M_{S}(\omega) &= G_{N}\left(M_{X}(\omega)\right) = \left.\frac{pz}{1-qz}\right|_{z=\frac{\mu}{\mu-i\omega}} \\ &= \frac{p \frac{\mu}{\mu-i\omega}}{1-q \frac{\mu}{\mu-i\omega}} = \frac{\mu p}{\mu p - i\omega} \end{split}$$

We conclude that

$$S \sim \text{Exp}(\mu p)$$

Expected values

We can obtain the expected value of $S = X_1 + X_2 + \cdots + X_N$ by conditioning on N, that is to say, $\mathbb{E}(S) = \mathbb{E}(\mathbb{E}(S \mid N))$.

We have

$$\mathbb{E}(S \mid N = k) = \mathbb{E}(X_1 + X_2 + \dots + X_k \mid N = k)$$
$$= \mathbb{E}(X_1 + X_2 + \dots + X_k) = k m,$$

where $m = \mathbb{E}(X)$ is the common expected value of the variables X_i . Therefore $\mathbb{E}(S \mid N) = mN$ and so

$$E(S) = \mathbb{E}(\mathbb{E}(S \mid N)) = \mathbb{E}(mN) = m\mathbb{E}(N) = \mathbb{E}(N)\mathbb{E}(X).$$

13 / 15

Expected values

In Example 1 we have X ~ Be(p), N ~ Po(λ), and S ~ Po(λp). Therefore

$$\mathbb{E}(S) = \lambda p = \mathbb{E}(N) \mathbb{E}(X)$$

In Example 2 we have X ~ Exp(µ), N ~ Ge(p), and S ~ Exp(µp). Hence

$$\mathbb{E}(S) = \frac{1}{\mu \rho} = \frac{1}{\rho} \cdot \frac{1}{\mu} = \mathbb{E}(N) \mathbb{E}(X)$$

14 / 15