# Improving the diagnosis of severe malaria in African children using platelet counts and plasma *Pf*HRP2 concentrations

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Severe falciparum malaria is difficult to diagnose accurately in children in high transmission settings. Using data from 2,649 patients enrolled in four studies of severe illness in three countries (Bangladesh, Kenya, and Uganda), we fitted Bayesian latent class models using two diagnostic biomarkers: the platelet count and the plasma *Pf*HRP2 concentration. In severely ill patients with clinical features consistent with severe malaria, a combined platelet count $\leq$ 150,000 per $\mu$L and a plasma *Pf*HRP2 concentration $\geq$ 1,000 ng/mL had an estimated sensitivity of 74% and specificity of 93% in identifying 'true' severe falciparum malaria. Patients with true severe malaria had higher parasite densities, lower hematocrit, lower rates of invasive bacterial disease, and a lower prevalence of both HbAS and HbSS than children misdiagnosed. We estimate one third of African children enrolled in clinical studies of severe malaria in high transmission settings had another cause of severe illness.

# 1 Introduction

Severe falciparum malaria is defined clinically as vital organ dysfunction in the presence of circulating *Plasmodium falciparum* parasites (*1*). The primary objective of this definition is to identify severely ill patients rapidly and provide life-saving clinical management (notably parenteral artesunate treatment). This prioritizes diagnostic sensitivity over specificity. In areas of moderate and high *P. falciparum* transmission, many apparently healthy children have malaria parasitemia if their blood is examined either by light microscopy or a rapid diagnostic test. It follows that many hospitalized children will also be parasitemic. As the major clinical features of severe malaria are not specific, it is difficult to differentiate clinically between severe falciparum malaria caused by extensive sequestration of malaria parasites in the microvasculature (i.e. 'true' severe malaria (*2*)) and other causes of severe febrile illness accompanied by either

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coincidental asymptomatic parasitemia or uncomplicated malaria (*3, 4*). We have shown previously that complete blood counts (platelet counts and total white blood cell counts) provide substantial discriminating value in distinguishing true severe malaria from other causes of severe illness (*5*). Approximately one third of a cohort of 2,220 Kenyan children diagnosed with severe malaria in a moderate transmission area were estimated to have had another cause of their severe illness (*5*). The discriminative value of complete blood counts was validated using genetic polymorphism data in two different ways. First, the distribution of sickle trait (HbAS) was correlated very strongly with the estimated probability of severe malaria. The prevalence of HbAS was around five times higher in the subgroup of patients likely to have been misdiagnosed compared to the subgroup with a likely correct diagnosis (*5*). Second, we showed that genome-wide false discovery rates could be reduced substantially in a case-control whole-genome association study in which the output probability weights were used in a 'data-tilting' framework to adjust for patient mis-classification (*5*).

The diagnostic value of complete blood counts in severe malaria has great operational utility as blood counts are widely measured in routine practice at low cost. Moderate thrombocytopenia is a consistent feature of all human malaria infections (*6–8*), although the diagnostic utility of the platelet count has been debated (*9, 10*). Platelets are activated in malaria infections and have increased turnover. In severe falciparum malaria there is endothelial activation with release of platelet aggregating activated high-multimeric von Willebrand Factor from specialized secretory vesicles in endothelial cells (the Weibel-Palade bodies). Platelets may also contribute to, and be consumed by, parasitised erythrocyte cytoadherence and autoagglutination (*11, 12*). Severe thrombocytopenia is associated with increased mortality in severe malaria (*10, 13, 14*). The total white blood cell count is also informative (although much less so than thrombocytopenia), with a greater prognostic than diagnostic value (very high or very low white counts in severe disease are associated with high case fatality ratios) (*5, 8*). In addition, total white counts

vary according to age and ethnicity (*15*), confounding cross-study assessments.

The main pathophysiological process in severe falciparum malaria is the extensive seques-
tration of parasitised erythrocytes in the vascular beds of vital organs (*1, 2*). These parasites,
which cause potentially lethal pathology, have stopped circulating and are therefore not repre-
sented in the peripheral blood smear. In African children the peripheral parasite density is a
poor indicator of disease severity and a poor diagnostic biomarker (*16*). The parasite protein
*Plasmodium falciparum* histidine-rich protein-2 (*Pf*HRP2), the basis for most rapid diagnostic
tests, is liberated mainly at schizont rupture - and so the amount released is proportional to
the extent of recent schizogony. Plasma *Pf*HRP2 is a much better discriminant of severe falci-
parum malaria than the peripheral blood parasite count (*17, 18*). An example of how *Pf*HRP2
can help interpret clinical trial data comes from the large multi-center AQUAMAT trial which
was a randomized comparison of parenteral artesunate versus parenteral quinine in African chil-
dren clinically diagnosed as having severe falciparum malaria. Overall, artesunate reduced the
mortality by approximately 22% (*19*). However, there was evidence for treatment effect het-
erogeneity across *Pf*HRP2 strata. In the subgroup of children in the highest tertile of *Pf*HRP2
artesunate reduced the mortality by one third (a similar proportion to that observed earlier in
the SEAQUAMAT randomized comparison conducted in Southeast Asia (*20*)), whereas there
was no substantial difference in mortality in the subgroup of children in the lowest tertile, i.e.
the group with likely other causes of severe illness (*18*).

Defining optimal cutoff values for diagnostic biomarkers in severe malaria is difficult be-
cause in high-transmission areas there is no 'gold-standard' for diagnosis against which to cali-
brate thresholds. In the absence of a gold-standard, latent class models can be used to assess the
sensitivity and specificity of threshold values for diagnostic indices (*21*). Latent class models
rely typically on having multiple biomarkers measured in the same individuals across different
populations with varying disease prevalences, thus allowing for triangulation (*21, 22*). We ap-

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plied Bayesian parametric latent class models to admission platelet counts and plasma *Pf*HRP2 concentrations in combination in order to estimate their diagnostic operating characteristics and to estimate the proportion of misdiagnosed patients in studies of severe malaria. We used data from four prospective studies of severe malaria or severe illness from Uganda, Kenya and Bangladesh, reflecting a range of *P. falciparum* transmission intensities from high to low, and thus a range of false diagnosis rates. The results suggest that the high rates of misdiagnosis could be reduced substantially by incorporating measurement of platelet counts and plasma *Pf*HRP2 concentrations in the diagnostic criteria.

## Results

## Platelet counts and *Pf*HRP2 concentrations in severe febrile illness

We pooled individual patient data from 2,649 severely ill African children and Asian adults, for whom platelet counts and measured plasma *Pf*HRP2 concentrations were available. The patients were from four separate studies in 3 countries:

- An observational study of severe falciparum malaria in Bangladesh ($n = 172$; all patients were clinically diagnosed with severe falciparum malaria).

- The Ugandan sites of the FEAST trial ($n = 567$, a randomized controlled trial of fluid resuscitation approaches in severe childhood illness not specific to severe malaria) (*23*).

- An observational study of cerebral malaria and severe malarial anemia in Kampala, Uganda ($n = 492$) (*24*).

- A large cohort of children diagnosed with severe malaria in Kilifi, Kenya ($n = 1,418$) (*25*).

Malaria transmission intensity is generally low in Bangladesh, moderate in Kilifi and Kampala, and high around the other Ugandan sites.

In total 27 patients (1%: 1 in Bangladesh, 8 in the FEAST sites in Uganda, and 18 from Kenya) had no detectable *Pf*HRP2 in the ELISA assay but had a parasite density $>1,000$ per $\mu$L by microscopy. These could have either been assay errors or parasites with *HRP2/3* gene deletions.They were removed from the analysis, leaving a total of 2,622 samples. A summary of the patient characteristics in the analyzed data set is shown in `Table 1`.

Log$_{10}$-transformed platelet counts and log$_{10}$-transformed *Pf*HRP2 concentrations were strongly inversely correlated in both the studies of severely ill African children ($\rho =-0.54$), and in the Bangladeshi adults with severe malaria ($\rho = -0.36$). Patients with platelet counts in the normal range ($>150,000$ per $\mu$L) generally had low to non-measurable plasma *Pf*HRP2 concentrations (median concentration was 269 ng/mL, IQR: 24 to 1,043), while patients with thrombocytopenia ($\leq 150,000$ per $\mu$L) had a median *Pf*HRP2 concentration of 3,031 ng/mL (IQR: 1,261 to 6,035).

## Discriminative value of platelet counts and plasma *Pf*HRP2 in patients diagnosed with severe malaria

We first assessed the diagnostic value of platelet counts and plasma *Pf*HRP2 concentrations in patients who were diagnosed clinically with severe malaria (a total of $n$=2,063; this excluded patients from the FEAST study which explicitly included non malarial causes of severe febrile illness). We fitted a two component parametric Bayesian latent class model (with the two latent classes representing 'severe malaria' and 'not severe malaria') using the log$_{10}$ transformed biomarkers (see Methods for description of the informative priors used and the key assumptions). `Fig. 1` shows the model estimated sensitivities and specificities for all cutoff values of the platelet count and the *Pf*HRP2 concentration. As expected, the platelet count and the

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| | Kilifi (Kenya) | Kampala (Uganda) | FEAST (Uganda) | Bangladesh |
|---|---|---|---|---|
| $n$ | 1,400 | 492 | 559 | 171 |
| Age (years, IQR) | 2.4 (1.4-3.7) | 3.3 (2.2-4.6) | 2.0 (1.2-3.3) | 30 (23-45) |
| Proportion parasite positive (%) | 100 | 100 | 59.4 | 100 |
| Parasite density* (per $\mu$L, IQR) | 69,824 (6,099-316,350) | 42,530 (10,635-198,540) | 37,600 (3,640-153,680) | 148,874 (23,550-348,540) |
| Platelet count ($10^3$/$\mu$L, IQR) | 111 (64-215) | 96 (49-170) | 165 (75-326) | 50 (27-139) |
| *Pf*HRP2 (ng/mL, IQR) | 2,207 (419-5,072) | 1,838 (588-4,097) | 175 (0-1,953) | 2,667 (1,083-6,128) |
| White blood cell count ($10^3$/$\mu$L, IQR) | 12.6 (8.9-19) | 10.4 (7.5-15.3) | 12.0 (8.4-18.7) | 9.0 (6.9-11.0) |
| Mortality (%) | 11.1 | 6.7 | 11.4 | 26.9 |
| HbAS ($n$, %) | 41 (2.9) | 4 (0.8) | 46 (8.2) | - |
| HbSS ($n$, %) | 7 (0.5) | 23 (4.7) | 21 (3.8) | - |

Table 1: Patient characteristics across the four studies. For age, parasite densities, platelet counts, total white blood cell counts and *Pf*HRP2 concentrations we show the median values and inter-quartile ranges. No hemoglobin S (HbS) genotyping was done for the patients from Bangladesh as HbS is absent from that population. *For FEAST, this excludes patients who had a negative RDT. IQR: inter-quartile range. HbAS: sickle trait; HbSS: homozygous sickle cell anemia.

*Pf*HRP2 concentration both had high discriminative value (illustrated by the receiver operating characteristic (ROC) curves). Overall, for thresholds giving the same sensitivity, the plasma *Pf*HRP2 concentration had a higher specificity. For example, in these populations, a lower limit of 1,000 ng/mL for the *Pf*HRP2 concentration had an estimated sensitivity of 87% and a specificity of 83%. In comparison, an upper limit for the platelet count of 150,000 per $\mu$L had an estimated sensitivity of 83% but a specificity of 71%. A series of sensitivity analyses (using non-informative priors, different choices for the parametric model, or using only data from the two studies of African children with severe malaria) showed near identical results (`Fig. S1`).

## Joint diagnostic thresholds for platelet counts and plasma *Pf*HRP2

For clinical and epidemiological studies, and in contrast to clinical practice, specificity is usually more important that sensitivity. We used both biomarkers in combination to improve the precision of the definition of severe malaria, in order to achieve low false positive rates. Multiple combinations of the platelet count and *Pf*HRP2 concentration have approximately the same operating characteristics for diagnosis, so the optimal choice is subjective and can be made based on operational simplicity (`Fig. S2`). In severe illness with clinical features consistent with severe malaria, platelet counts $\leq$ 150,000 per $\mu$L and plasma *Pf*HRP2 concentrations $\geq$ 1,000 ng/mL in combination have an estimated diagnostic specificity of 93% and a sensitivity of 74%. If this definition was applied to a hospital cohort with a prevalence of severe malaria of 60% (approximately the prevalence estimated for the Kenyan cohort of children with clinically diagnosed severe malaria), we would expect over 94% of the resulting population identified by these two biomarkers to have 'true' severe malaria (positive predictive value).

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## Estimating the probability of severe malaria

We fitted a three component parametric Bayesian latent class model to all available platelet count and *Pf*HRP2 concentration data from patients from the four studies of severe febrile illness (a total of $n$=2,622) in order to estimate the individual patient probabilities that severe malaria was the true cause of their severe illness, and to calculate the prevalences of 'true' severe malaria among those diagnosed. The additional third component captured a cluster of patients in FEAST trial who had no detectable plasma *Pf*HRP2 and normal platelet counts (Fig. 2). Under this model, we estimated that the prevalence of true severe malaria was 96% (95% CI: 91-99) in Bangladeshi adults; 37% (95% CI: 31-42) in the Ugandan children enrolled in the FEAST trial (FEAST intentionally enrolled severely ill children with and without malaria, although 66% of all patients had a clinical diagnosis of severe malaria (*23*)); 74% (95% CI: 67-79) in the children enrolled in Kampala, Uganda (*24*); and 66% (95% CI: 61-70) in the children diagnosed with severe malaria in Kilifi, Kenya.

Fig. 2 shows scatter plots of the platelet counts versus the plasma *Pf*HRP2 concentrations, colored by the probability that the patient had severe malaria and grouped by study (dark blue: high probability of severe malaria; dark red: low probability of severe malaria). The model estimated a geometric mean platelet count in severe malaria across the four studies of 74,000 per $\mu$L (95% of patients are predicted to have platelet counts between 17,000 and 312,000) and a geometric mean *Pf*HRP2 of 3,135 ng/mL (95% prediction interval: 402 to 24,452). The model-based probabilities of severe malaria were highly concordant with our previously published model which used platelet counts and total white blood cell counts ($\rho$=0.63, Fig. S3) (*5*).

We compared the estimated false diagnosis rates for cerebral malaria versus severe malaria without coma, and for severe malarial anemia versus severe malaria without severe anemia (where severe anemia is defined as a hematocrit $\leq$15%) in the two cohorts of children clinically diagnosed with severe malaria. In the Kenyan cohort, the estimated false diagnosis rate was

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slightly higher in the cerebral malaria group relative to the non-cerebral malaria groups: 36% of patients with coma, and 26% of patients without coma were classified as "falsely diagnosed" with severe malaria ($p$=0.001). For the subgroup with severe anemia, the false diagnosis rate was 15%, compared to 36% in the subgroup without severe anemia. In the Ugandan cohort, both of these trends were reversed. A false diagnosis of severe malaria was estimated for only 12% of patients with coma, but for 40% of patients without coma. For severe anemia, these proportions were 36% and 15%, respectively (the Ugandan cohort only recruited patients with either severe malarial anemia or patients with cerebral malaria).

Mortality rates varied substantially as a function of the estimated probability of severe malaria and across the studies (Fig. 3). In African children with a high probability of having severe malaria, the mean mortality was consistently around 10% for the three studies included. Apart from the Bangladeshi adults, all patients were treated with initial intravenous quinine. Mortality in severe malaria in adults from Bangladesh was substantially higher ($\sim$30%), consistent with previous studies (20). In Kenyan children, the mortality in the mis-classified group of patients was higher than in the correctly classified patients (14% versus 10%), whereas in the Ugandan study the trend was reversed (1% versus 10%). This is largely explained by the study populations: the group with a low probability of having severe malaria in the Ugandan study was predominantly composed of patients with severe anemia without other features of severity.

## Relationship with other admission variables

We explored the relationship between the model estimated probability of severe malaria and the admission parasite densities, admission hematocrit, the total white blood cell counts, and the blood culture positivity rates (i.e. cultures growing a likely pathogen). In the three African studies, parasite densities were between 12 and 16 times higher in patients with a high probability of severe malaria versus those with a low probability of severe malaria (Fig. 4). After

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adjusting for study differences, the parasite densities were estimated to be 12.8-fold higher (95% C.I 9.9-16.5) in severe malaria versus not severe malaria.

The admission hematocrit were also highly correlated with the model estimated probability of severe malaria (Fig. S4). In the three African studies, children with a high probability of having severe malaria had median admission hematocrit between 16 and 20% (FEAST: 19%; Kampala: 16%; Kilifi: 20%). The hematocrit distributions in this group were unimodal (Fig. S5). In contrast, the hematocrit distributions in patients with low probabilities of having severe malaria ($<0.2$) were strongly bi-modal, with the majority of patients having higher hematocrit (the median hematocrit in this group were: FEAST: 30%; Kampala: 12%; Kilifi: 28%), but a substantial minority had low hematocrit of around 10%. The few Bangladeshi adults with a low probability of severe malaria also had higher hematocrit.

Blood cultures were done for all 1,400 Kenyan children and for 298 out of 332 (90%) Ugandan children in the FEAST study who had malaria parasitemia on admission. Overall 51 and 35 patients respectively had positive blood cultures after removing contaminants. The probability of having severe malaria was highly predictive of the blood culture result, with an adjusted odds-ratio of 0.43 (95% CI: 0.27-0.66; p=0.0002) for a positive culture in patients likely to have severe malaria versus those unlikely to have severe malaria (after adjustment for study). This difference in blood culture positivity rates was also reflected in the total white blood cell counts (Fig. S6). Across the four studies, after adjustment for age, in patients likely to have severe malaria (probability $>0.5$) compared to those unlikely to have severe malaria (probability $<0.5$) the odds-ratio for having a total white count $>15,000$ per $\mu$L was 0.50 (95% CI: 0.46-0.56; p=$10^{-12}$).

## Gene-dose relationship for hemoglobin S and severe malaria

In the three studies in African children, the prevalence of both HbAS and HbSS were strongly inversely correlated with the model estimated probability of severe malaria (data are shown in `Fig. 2`). Pooling the three African studies and adjusting for differences in the estimated prevalence of severe malaria, the odds-ratio for being classified as severe malaria (probability $>0.5$) for HbAS patients relative to HbAA patients was 0.24 (95% CI: 0.15 to 0.40, $p = 10^{-8}$), and for HbSS relative to HbAA the odds-ratio was 0.08 (95% CI: 0.04 to 0.18, $p = 10^{-10}$). Under an additive model of association, each additional hemoglobin S (HbS) allele was associated with an odds-ratio for severe malaria of 0.27 (95% CI: 0.19 to 0.37, $p = 10^{-16}$). This association between HbS genotypes and the biomarker-based severe malaria classification was highly concordant across the three studies.

# Discussion

The diagnosis of severe malaria in African children is imprecise (*3, 4, 26*). This is because it is difficult to distinguish clinically between severe illness caused by malaria from severe illness with incidental asymptomatic or uncomplicated malaria since they share many clinical characteristics. This is a substantial problem for studies of one of the most important life threatening infections in childhood. It dilutes and distorts the results of epidemiology (*27*), pathophysiology (*3*), genetic association (*5*) and therapeutic investigations (*18*). In areas of moderate and high levels of malaria transmission asymptomatic parasitemia is very common. At any time a high proportion of children have detectable malaria parasitemia. Malaria parasitemia is therefore a sensitive but not specific indicator that malaria is the cause of illness (*26*). Other simple laboratory tests provide valuable diagnostic information. Thrombocytopenia is a common feature of all symptomatic malarias (*6–8*). A low platelet count therefore supports (but does not

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prove) attributing malaria as the cause of illness. We estimate that less than one in five patients with severe malaria will have platelet counts in the normal range ($>$150,000 per $\mu$L). This analysis of data from large prospective studies of severe illness in African children shows that measurement of platelet counts and the plasma concentration of the parasite protein *Pf*HRP2 can together substantially improve the specificity of a diagnosis of severe falciparum malaria.

A key strength of this work is that we can validate the discriminant power of the platelet count and the plasma *Pf*HRP2 concentration by comparing the prevalence of HbS genotypes (HbAS and HbSS) and the blood culture positivity rates across in the inferred subgroups. HbAS is the genotype that provides the strongest known protection against severe falciparum malaria ([28, 29]). The prevalence of HbAS was four times lower in children considered likely to have severe malaria compared with those considered less likely to have severe malaria. It is important to note that HbAS also protects against complications of malaria, such as an increased risk of bacterial infections, so the HbAS prevalence in children with presumed bacterial infections is still expected to be lower than that in the healthy population ([30]). Recent work suggests that the protective effect of HbAS may vary according to parasite genotype ([31]). It may be that the few HbAS individuals who have "true" severe malaria have parasite genotypes that can evade the HbAS defense mechanisms. Future work will assess the relationship between these probabilistic classifications and the parasite genotypes. Whereas people with the AS genotype are essentially hematologically normal, those with homozygous SS suffer from sickle cell disease. This causes anemia, and may also cause thrombocytopenia and leukocytosis. This could confound the use of full blood count data in the probabilistic assessment of severe malaria. However, the interpretation of the plasma *Pf*HRP2 concentration (or other parasite biomass indicators) should not be affected by sickle cell disease. In this study there was strong evidence for an additive effect between the number of HbS alleles and a decreased probability of severe malaria under the platelet/HRP2 model. This suggests that HbSS is strongly protective against a

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high parasite biomass (*32*). However as sickle crises may be severe, and they are often triggered by infections, it is possible a low parasite biomass could trigger severe events. Whether or not this should be described as 'severe malaria' is a semantic question (*33*). A recent in-depth analysis of the Ugandan cohort from Kampala showed that the HbSS children considered to have severe malaria had a lower parasite biomass but higher levels of endothelial dysregulation (*32*). The main differential diagnosis in suspected severe malaria is bacterial sepsis. The relationship is complicated as severe malaria does predispose to bacterial sepsis (*34*). Blood culture has low sensitivity, however we observed a three fold increase in positivity rates in the patients likely have been mis-diagnosed compared to those with a likely correct diagnosis.

The main limitation of this study is the absence of a gold-standard diagnosis, and thus the reliance on parametric latent class models. The application of latent class models to multiple populations with varying disease prevalence provides a powerful framework for estimating the diagnostic accuracy of imperfect tests, but the model outputs rely on the validity of key assumptions, notably the underlying parametric models used for the biomarker distributions. We used log-normal distributions for the platelet count and plasma *Pf*HRP2 concentration in both 'true' severe malaria and 'not severe malaria'. Visual data plots suggest that this is a good approximation of the true underlying distributions. Each individual biomarker has its own limitations. Thrombocytopenia can be caused by other infections (e.g. severe arbovirus infection) and may occur in sepsis, although it is much less prevalent in sepsis than malaria (*8*). The main limitation of plasma *Pf*HRP2 as a marker of total parasite biomass is that it requires specialist measurement (although rapid point-of-care tests can be applied to plasma (*35*) and modifications are under development). As plasma *Pf*HRP2 accumulates each asexual parasite cycle, an individual with a sustained low parasite multiplication rate at high parasite densities can have the same *Pf*HRP2 concentration as an individual with a fulminant infection and a high sequestered biomass (*17*). In addition mutations in the *Pf*HRP2 gene causing changes in antigenicity may

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affect immunoassays, although these are rare outside of the Horn of Africa (*36, 37*).

In these very large prospective series of patients hospitalised with a diagnosis of severe malaria, combining platelet counts and plasma *Pf*HRP2 concentrations provided good discrimination between "true" severe falciparum malaria and other severe illnesses (likely to be bacteraemia in many of the cases) with concomitant incidental malaria. A major strength of this study is in combining two measures which are mechanistically distinct (platelet counts are measuring the host response to acute malaria illness; *Pf*HRP2 is measuring the parasite biomass). We show that in combination these biomarkers provide a high level of discrimination between patients who were likely to have had true severe malaria and those with a different illness aetiology. This proportion ranged from one third in the FEAST study which intentionally studied children with both severe malaria and other severe illnesses (likely mainly sepsis) requiring fluid resuscitation to >90% in Bangladesh -a low transmission area where the diagnosis of falciparum malaria as the cause of illness is highly specific. Overall it suggests that approximately one third of African children diagnosed with severe malaria have another cause of their severe illness (*5*). Our results suggest that severe malaria with a normal platelet count is unusual and a platelet count >150,000 per $\mu$L in a child with suspected severe malaria should motivate further examination for a possible different cause of illness, while giving prompt antimalarial treatment. The high diagnostic accuracy of the plasma *Pf*HRP2 concentration should motivate future work on simple modifications to the currently used *Pf*HRP2 rapid tests, in order to estimate plasma PfHRP2 at the bedside. Finally, we recommend that future studies of severe falciparum malaria should include only children with platelet counts of $\leq$150,000 per $\mu$L and plasma *Pf*HRP2 concentrations of $\geq$1,000 ng/mL.

# Methods

## Study design

This study is a retrospective analysis of platelet counts and plasma *Pf*HRP2 concentrations in patients with severe febrile illness. We merged data from four separate studies, three of which were studies of severe malaria and one was a study of fluid resuscitation in severe febrile illness (FEAST trial). The main goal of the analysis was to use Bayesian latent class modelling to estimate the proportion of mis-diagnosed patients in the three severe malaria studies, and to assess the operating characteristics of different cutoff values for the platelet count and *Pf*HRP2 concentrations for future clinical studies. For the analysis and reporting we followed the EQUATOR guidelines for the reporting of diagnostic accuracy studies that use Bayesian latent class models (STARD-BLCM, see Supplementary Materials).

## Data

All the clinical studies were prospective studies of severe illness and had appropriate ethical approval.

### Bangladesh

We included data from observational studies in severe falciparum malaria conducted by the Mahidol Oxford Tropical Medicine Research Unit in Bangladesh between 2003 and 2019. These pooled data have been described previously (*38*). Malaria transmission is seasonal and of low intensity in this location. In brief, adults and children with microscopy evidence for asexual stage malaria parasites (thick and thin blood films) who met the WHO definition for severe malaria were enrolled in study after written informed consent was obtained from the patient or an attending relative.

Criteria for severe malaria included coma (Glasgow Coma Score <11 or Blantyre Coma

Score <3), pulmonary oedema, repeated convulsions (≥2 in 24 hours), severe anemia (hematocrit level <20%, plus a parasite count >100 000 parasites/$\mu$L) or jaundice (bilirubin level >3.0 mg/dL, plus a parasite count >100 000 parasites/$\mu$L), renal failure (serum creatinine level >3 mg/dL), hypoglycemia (blood glucose level <40 mg/dL), shock (systolic blood pressure <80 mm Hg with cool extremities), hyperparasitemia (peripheral asexual stage parasitemia >10%), hyperlactemia (venous plasma lactate >4 mmol/L), and/or acidemia (venous plasma bicarbonate level <15 mmol/L).

Platelet counts and *Pf*HRP2 concentrations were jointly measured in a total of 172 patients. The majority of these patients received intravenous artesunate.

**Kilifi (Kenya)**

The Kenyan case-control cohort has been described in detail previously (*25*). Severe malaria cases consisted of all children aged <14 years who were admitted with clinical features of severe falciparum malaria to the high dependency ward of Kilifi County Hospital between June 11th 1999 and June 12th 2008. Severe malaria was defined as a positive blood-film for *P. falciparum* along with: prostration (Blantyre Coma Score of 3 or 4); cerebral malaria (Blantyre Coma Score of <3); respiratory distress (abnormally deep breathing); severe anemia (hemoglobin <5 g/dL). The standard of care antimalarial treatment during this period was intravenous quinine.

**FEAST (Uganda)**

FEAST was a multi-center randomized controlled trial comparing fluid boluses for severely ill children with shock ($n = 3,161$) that was not specific to severe malaria (*23*). Children between 2 months and 12 years of age were eligible for the trial if they presented with a severe febrile illness complicated by impaired consciousness (prostration or coma), respiratory distress (increased work of breathing), or both, and with impaired perfusion, as evidenced by one or

more of the following: a capillary refill time of 3 or more seconds, lower-limb temperature gradient,19 weak radial-pulse volume, or severe tachycardia (>180 beats per minute in children younger than 12 months of age, >160 beats per minute in children 1 to 5 years of age, or >140 beats per minute in children older than 5 years of age). Exclusion criteria were severe malnutrition, gastroenteritis, noninfectious causes of shock (e.g., trauma, surgery, or burns).

Platelet counts and plasma *Pf* HRP2 concentrations were measured in 502 children in the Ugandan sites (Mulago National Referral Hospital, Mbale and Soroti Regional Referral Hospitals and St Mary's Hospital, Lacor). The standard of care antimalarial treatment during this period was intravenous quinine.

**Kampala (Uganda)**

The trial by Brand *et al* was an observational study of cerebral malaria and severe malarial anemia in Mulago Hospital, Kampala, Uganda (*24*). Children were enrolled if they were between 18 months and 12 years of age. Cerebral malaria was defined as coma (Blantyre Coma Score of <3 or Glasgow Coma Score <8) in presence of asexual parasites on a blood smear. Severe malarial anemia was defined as presence of *Plasmodium falciparum* parasites on a blood smear in children with a hemoglobin level ≤5g/dL. The standard of care antimalarial treatment during this study was intravenous quinine.

## Procedures

For the Kilifi cohort and the Ugandan sites in the FEAST trial, pediatric blood samples for bacterial cultures were collected in BACTEC Peds Plus bottles and processed with a BACTEC automated blood culture instrument (Becton Dickinson) for initial detection of bacteria in the blood. BACTEC-positive samples were subcultured on standard media by routine microbiological techniques. Either biochemical test kits (API, BioMérieux), serological tests, or both

were used to confirm suspected pathogens. Good Clinical Laboratory Practice was audited by Qualogy and external quality assurance was provided by the UK National External Quality Assessment service. The following organisms were considered as contaminants: Bacillus species, coryneforms, Micrococcus species, coagulase- negative Staphylococcus and citrobacter.

In all four studies plasma *Pf*HRP2 levels were quantitated using the same previously published methodology (*18*). The lower limit of detection of the ELISA plasma assay is approximately 2 ng/mL. Patients in the Kilifi study and the FEAST trial were genotyped for the rs334 SNP (HbS) using DNA extracted from fresh or frozen samples of whole blood as described in detail previously (*25, 39*).

## Statistical analysis

We fitted a series of Bayesian parametric latent class models to the available biomarker data (*21*). The key assumptions of the main models are summarized as follows:

1. The marginal distribution of each biomarker in each latent class is log-normal.

2. The marginal distribution of each biomarker in each latent class is the same across the different studies and countries.

3. Informative Bayesian priors on all parameters.

Sensitivity analyses relaxed assumptions 1-3 (using bivariate t-distributions, only using data from African children, and using weakly informative priors). We compared models with and without correlation between the biomarkers within each latent class. The models with correlation performed better and were more conservative.

In an initial exploratory analysis, we used the R package *mclust* (*40*) (Gaussian Mixture Modelling for Model-Based Clustering - this uses the EM algorithm for parameter estimation) to estimate the number of latent components in the data (Fig. S7). This suggested three

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major underlying clusters, one of which was only present in the FEAST trial (patients with no measurable plasma *Pf*HRP2 and normal platelet counts). We then used full Bayesian inference with informative priors to determine the distributions of the latent clusters. In this initial exploratory analysis, we also looked at the relationship between the peripheral parasite density and the two biomarkers (Fig. S8). This preliminary analysis confirmed that the parasite density is a poor biomarker for discriminating between 'true' severe malaria and other causes of illness as it was not possible to robustly fit clustering models with either the platelet count or the plasma *Pf*HRP2.

The first set of models used data from the three severe malaria studies only (Kenyan children from Kilifi, Ugandan children from Kampala, and Bangladeshi adults). For each biomarker, the marginal distribution in each latent class (severe malaria versus not severe malaria) was assumed to be log-normal with the following informative priors (weakly informative). Means and standard deviations are given on the $\log_{10}$ scale.

- The mean platelet count: $N(\log_{10} 75, 0.1)$ [weakly informative: $N(\log_{10} 100, 0.5)$] for severe malaria and $N(\log_{10} 200, 0.1)$ [weakly informative: $N(\log_{10} 200, 0.5)$] for not severe malaria. The platelet count geometric mean of 100,000 per $\mu$L versus 200,000 per $\mu$L in severe malaria versus not severe malaria is informed by the results of (*3*) (autopsy study in patients who had died with a diagnosis of cerebral malaria: the patients with evidence of parasite sequestration all had very low platelet counts compared to those who did not; we expect slightly higher platelet counts in all patients as the platelet count is a prognostic factor as well as diagnostic).

- The mean *Pf*HRP2: $N(\log_{10} 2500, 0.2)$ for severe malaria and $N(\log_{10} 500, 0.2)$ for not severe malaria. The weakly informative prior used a standard deviation of 0.5 instead of 0.2. The mean plasma *Pf*HRP2 values are informed from previous modeling of data from

the AQUAMAT trial (*18*).

- The standard deviation on the $\log_{10}$ scale for each log-normal distribution was given a
  exponential prior with rate parameter set to 2 (i.e. a mean standard deviation of 0.5 on the
  $\log_{10}$ scale)

We used informative beta priors for the prevalence of severe malaria in the three studies: Beta(19,1) for Bangladesh; Beta(14,6) for Kilifi; Beta(14,6) for Kampala (the weakly infor- mative beat priors: Beta(4,1), Beta(2,1), and Beta(2,1), respectively).

To assess the robustness of the model outputs, we performed three sensitivity analyses: (i) changing the parametric form to a bivariate student *t*-distribution with 10 degrees of freedom (robustness to assumption 1); (ii) by fitting the model to data only from the studies in African children with severe malaria (Kampala and Kilifi; robustness to assumption 2 as the Bangladeshi adults could plausibly be different from African children in terms of parasite biomass and host response); (iii) by using weakly informative priors (larger standard deviations around the prior mean values and weaker beta priors on the prevalence parameters).

For the second set of models, we included data from the FEAST trial (*23*). FEAST inten- tionally enrolled severely ill patients with and without malaria, thus there is a subset of patients who are clearly distinct from all patients in the studies of severe malaria (this subset is char- acterized on average by a negative malaria blood slide, no measurable plasma *Pf*HPR2 and a normal platelet count). The latent class model needed to include a third component in order to fit the data. Under this model, the biomarker distributions were also assumed to be log-normal in the severe malaria class, but were a mixture of two log-normal distributions for the 'not severe malaria' class.

We used the same priors for the second set of models with the addition of the following priors for the third additional component (which can be summarized as severe illness with no

evidence of acute malarial infection):

- The mean platelet count: $N(\log_{10} 300, 0.1)$ for not severe malaria.

- The mean *Pf*HRP2: $N(\log_{10} 1, 0.2)$ for not severe malaria.

For the three component model, Dirichlet priors were used for the prevalence parameters for each latent class: Dirichlet(19,1,0.1) for the Bangladesh study (with hyperparameters corresponding to the severe malaria and the two not severe malaria classes, respectively); Dirichlet(3,3,3) for the FEAST trial; Dirichlet(14,7,1) for the Kampala and Kilifi studies.

In order for patients with non-measurable plasma *Pf*HRP2 concentrations to be included in this analysis, we set all non-measurable concentrations to 1 ng/mL (approximately half the lower limit of detection of the ELISA assay). Patients who had non-measurable plasma *Pf*HRP2 concentrations but parasite densities above 1,000 per $\mu$L were excluded from the analysis as these could represent *P. falciparum* parasites with *HRP2/3* gene deletions. The median platelet count in these samples was 142,000 per $\mu$L (range: 28,000 to 808,000 per $\mu$L).

All posterior distributions were estimated using Monte Carlo methods. All models were implemented in the *rstan* language (*41*). For each model we ran 4 independent chains for 2000 iterations, discarding half for burn-in. Convergence was checked by visually assessing traceplots. Convergence problems due to 'label switching' were avoided by using parameter constraints: the mean value of each likelihood distribution was set as increasing (for the platelet count) or decreasing (for the *Pf*HRP2 and parasite density). This uses the *ordered* parameter class in *stan*. The *stan* language does not support discrete parameters, however posterior distributions of latent class models can be sampled by using the *logsumexp* trick, computing the marginal likelihood over the different class combinations.

## Code and data availability

All code in the form of an RMarkdown document along with a minimal clinical dataset is available online via the following github repository: https://github.com/jwatowatson/SevereMalariaDiagnosis. A static version can be found at https://doi.org/10.5281/zenodo.5602924 (*42*).

# References

1. World Health Organisation, *Tropical Medicine & International Health* **19**, 7 (2014).

2. N. J. White, G. D. Turner, N. P. Day, A. M. Dondorp, *The Journal of Infectious Diseases* **208**, 192 (2013).

3. T. E. Taylor, *et al.*, *Nature Medicine* **10**, 143 (2004).

4. N. M. Anstey, R. N. Price, *PLoS Medicine* **4**, e267 (2007).

5. J. A. Watson, *et al.*, *eLife* **10**, e69698 (2021).

6. C. Naing, M. A. Whittaker, *Infectious Diseases of Poverty* **7**, 1 (2018).

7. S. Kho, *et al.*, *Blood* **132**, 1332 (2018).

8. S. Ladhani, B. Lowe, A. O. Cole, K. Kowuondo, C. R. Newton, *British Journal of Haematology* **119**, 839 (2002).

9. Y. Chimalizeni, K. Kawaza, T. Taylor, M. Molyneux, *The American Journal of Tropical Medicine and Hygiene* **83**, 48 (2010).

10. J. Hanson, *et al.*, *BMC Medicine* **13**, 1 (2015).

11. D. Larkin, *et al.*, *PLOS Pathogens* **5**, 1 (2009).

23

477   12. E. C. Löwenberg, *et al.*, *Thrombosis and haemostasis* **103**, 181 (2010).

478   13. D. A. Lampah, *et al.*, *The Journal of Infectious Diseases* **211**, 623 (2015).

479   14. P. Gérardin, *et al.*, *The American Journal of Tropical Medicine and Hygiene* **66**, 686 (2002).

480   15. D. Reich, *et al.*, *PLoS Genetics* **5**, e1000360 (2009).

481   16. L. Von Seidlein, *et al.*, *Clinical Infectious Diseases* **54**, 1080 (2012).

482   17. A. M. Dondorp, *et al.*, *PLoS Medicine* **2** (2005).

483   18. I. C. Hendriksen, *et al.*, *PLoS Medicine* **9**, e1001297 (2012).

484   19. A. M. Dondorp, *et al.*, *The Lancet* **376**, 1647 (2010).

485   20. A. M. Dondorp, F. Nosten, K. Stepniewska, N. Day, N. White, *The Lancet* **366**, 717 (2005).

486   21. J. Collins, M. Huynh, *Statistics in Medicine* **33**, 4141 (2014).

487   22. S. L. Hui, S. D. Walter, *Biometrics* pp. 167–171 (1980).

488   23. K. Maitland, *et al.*, *New England Journal of Medicine* **364**, 2483 (2011).

489   24. N. R. Brand, R. O. Opoka, K. E. Hamre, C. C. John, *PloS one* **11**, e0163728 (2016).

490   25. C. M. Ndila, *et al.*, *The Lancet Haematology* **5**, e333 (2018).

491   26. P. Bejon, *et al.*, *PLoS Medicine* **4**, e251 (2007).

492   27. H. Reyburn, *et al.*, *British Medical Journal* **329**, 1212 (2004).

493   28. S. M. Taylor, C. M. Parobek, R. M. Fairhurst, *The Lancet Infectious Diseases* **12**, 457
494       (2012).

29. G. Band, *et al.*, *Nature Communications* **10**, 5732 (2019).

30. S. Uyoga, *et al.*, *Nature Communications* **10**, 856 (2019).

31. G. Band, *et al.*, *Nature* (2021).

32. R. C. Henrici, *et al.*, *Blood Advances* (2021).

33. T. N. Williams, S. K. Obaro, *Trends in Parasitology* **27**, 315 (2011).

34. J. Church, K. Maitland, *BMC Medicine* **12**, 1 (2014).

35. I. Sinha, N. Ekapirat, A. M. Dondorp, C. J. Woodrow, *Malaria Journal* **14**, 1 (2015).

36. S. M. Feleke, *et al.*, *Nature Microbiology* pp. 1–11 (2021).

37. A. B. Bosco, *et al.*, *PLOS ONE* **15**, 1 (2021).

38. S. J. Leopold, *et al.*, *PLoS Medicine* **16** (2019).

39. S. Wambua, *et al.*, *PLoS Medicine* **3**, e158 (2006).

40. L. Scrucca, M. Fop, T. B. Murphy, A. E. Raftery, *The R Journal* **8**, 289 (2016).

41. Stan Development Team, RStan: the R interface to Stan (2020). R package version 2.21.2.

42. J. Watson, jwatowatson/SevereMalariaDiagnosis: Preprint code version (2021).

# Acknowledgements

# Supplementary materials

EQUATOR STARD-BLCM checklist

Figs. S1 to S7

**Fig. 1.** The diagnostic value of platelet counts (pink) and plasma *Pf*HRP2 concentrations (green) estimated using data from 2,063 patients diagnosed with severe falciparum malaria in three studies. Sensitivities (dotted lines) and specificities (dashed lines) are estimated under a Bayesian parametric two-component latent class model which assumes the same sensitivity and specificity for each biomarker across the three studies (African children: Kilifi, Kenya and Kampala, Uganda; Asian adults: Bangladesh). For the platelet count, thresholds correspond to upper limits, whereas for the *Pf*HRP2 concentration the thresholds correspond to lower limits.

**Fig. 2.** Probabilistic model of severe falciparum malaria using platelet counts and plasma *Pf*HRP2 concentrations in 2,622 severely ill patients based on a Bayesian parametric latent class model with three latent classes (a severe malaria class and two 'not severe malaria' classes). The colors correspond to the probability of severe malaria under the model (dark blue: high probability; dark red: low probability). Triangles show the HbAS individuals; crosses show the HbSS individuals. In order to show data points with non-measurable plasma *Pf*HRP2, non-measurable concentrations were set to 1 ng/mL $\pm$ random jitter on the $\log_{10}$ scale (approximately half the lower limit of quantification of the assay).

**Fig. 3.** Mortality as a function of the probability of having severe malaria under the Bayesian latent class model (based on platelet counts and *Pf*HRP2 concentrations). The lines (shaded areas) show mean (95% confidence intervals) mortality estimates from logistic regression fits.

**Fig. 4.** Admission parasite densities as a function of the probability of severe malaria under the Bayesian latent class model. Data from the FEAST trial includes only the malaria RDT positive patients. The thick lines show the additive linear model fit (spline based).