Extropy: Complementary Dual of Entropy

Frank Lad, Giuseppe Sanfilippo and Gianna Agrò

This article provides a completion to theories of information based on entropy, resolving a longstanding question in its axiomatization as proposed by Shannon and pursued by Jaynes. We show that Shannon's entropy function has a complementary dual function which we call "extropy." The entropy and the extropy of a binary distribution are identical. However, the measure bifurcates into a pair of distinct measures for any quantity that is not merely an event indicator. As with entropy, the maximum extropy distribution is also the uniform distribution, and both measures are invariant with respect to permutations of their mass functions. However, they behave quite differently in their assessments of the refinement of a distribution, the axiom which concerned Shannon and Jaynes. Their duality is specified via the relationship among the entropies and extropies of course and fine partitions. We also analyze the extropy function for densities, showing that relative extropy constitutes a dual to the Kullback–Leibler divergence, widely recognized as the continuous entropy measure. These results are unified within the general structure of Bregman divergences. In this context they identify half the L_2 metric as the extropic dual to the entropic directed distance. We describe a statistical application to the scoring of sequential forecast distributions which provoked the discovery.

Key words and phrases: Differential and relative entropy/extropy, Kullback–Leibler divergence, Bregman divergence, duality, proper scoring rules, Gini index of heterogeneity, repeat rate.

1. SCOPE, MOTIVATION AND BACKGROUND

The *entropy* measure of a probability distribution has had a myriad of useful applications in information sciences since its full-blown introduction in the extensive article of Shannon (1948). Prefigured by its usage in thermodynamics by Boltzmann and Gibbs, entropy has subsequently bloomed as a showpiece in theories of communication, coding, probability and statistics. So

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widespread is its application and advocacy, it is surprising to realize that this measure has a complementary dual which merits recognition and comparison, perhaps in many realms of its current application, a measure we term extropy. In this article we display several intriguing properties of this information measure, resolving a fundamental question that has surrounded Shannon's measure since its very inception. The results provide links to other notable information functions whose relation to entropy have not been recognized. In particular, the standard L_2 distance between two densities is identified as dual to the entropic measure of Kullback-Leibler, an understanding provoked by considering the extropy function as a Bregman function. We shall follow Shannon's original notation and extend it.

If X is an unknown but observable quantity with a finite discrete range of possible values $\{x_1, x_2, ..., x_N\}$ and a probability mass function (p.m.f.) vector $\mathbf{p}_N = (p_1, p_2, ..., p_N)$, the Shannon entropy measure de-

noted by H(X) or $H(\mathbf{p}_N)$ equals $-\sum_{i=1}^N p_i \log(p_i)$. Its complementary dual, to be denoted by J(X) or $J(\mathbf{p}_N)$, equals $-\sum_{i=1}^N (1-p_i)\log(1-p_i)$. We propose this as the measure of extropy. As is entropy, extropy is interpreted as a measure of the amount of uncertainty represented by the distribution for X. The duality of $H(\mathbf{p}_N)$ and $J(\mathbf{p}_N)$ will be found to derive formally from the symmetric relationship they bear with the sums of the (entropies, extropies) in the N crude event partitions defined by $[(X=x_i), (X\neq x_i)]$. The complementarity of H and J arises from the fact that the extropy of a mass function, $J(\mathbf{p}_N)$, equals a location and scale transform of the entropy of another mass function that is complementary to \mathbf{p}_N : that is,

$$J(\mathbf{p}_N) = (N-1)[H(\mathbf{q}_N) - \log(N-1)],$$

where $\mathbf{q}_N = (N-1)^{-1}(\mathbf{1}_N - \mathbf{p}_N)$. This p.m.f. \mathbf{q}_N is constructed by norming the probabilities of the events $\tilde{E}_1, \dots, \tilde{E}_N$ which are complementary to E_1, \dots, E_N . When N=2 this yields the standard p.m.f. for \tilde{E}_1 as opposed to the p.m.f. for E_1 . Together, these two relationships establish extropy as the complementary dual of entropy.

In his seminal article that characterized the entropy function, Shannon (1948) began by formulating three properties that might well be required of *any* function $H(\cdot)$ that is meant to measure the amount of information inhering in a p.m.f. \mathbf{p}_N . He suggested the following three properties as axioms for $H(\mathbf{p}_N)$:

- (i) $H(p_1, p_2, ..., p_N)$ is continuous in every argument;
- (ii) $H(\frac{1}{N}, \frac{1}{N}, \dots, \frac{1}{N})$ is a monotonic increasing function of the dimension N; and
- (iii) for any positive integer N, and any values of p_i and t each in [0, 1],

$$H(p_1, ..., p_{i-1}, tp_i, (1-t)p_i, p_{i+1}, ..., p_N)$$

= $H(p_1, p_2, ..., p_N) + p_i H(t, 1-t)$.

Shannon then proved that the entropy function $H(\mathbf{p}_N) = -\sum_{i=1}^N p_i \log(p_i)$ is the *only* function of \mathbf{p}_N that satisfies these axioms. It is unique up to an arbitrary specification of location and scale. Subsequently, the article of Rényi (1961) presented alternative characterizations of entropy due to Fadeev and himself. These involved alternating these axioms with various properties of Shannon's function, such as its invariance with respect to permutations of its arguments and its achieved maximum occurring at the uniform distribution.

Shannon's third axiom concerns the behavior of the function $H(\cdot)$ when any category of outcome for X is split into two distinguishable possibilities, and the probability mass function \mathbf{p}_N is thereby refined into a p.m.f. over (N+1) possibilities. It implies that the entropy in a joint distribution for two quantities equals the entropy in the marginal distribution for one of them plus the expectation for the entropy in the conditional distribution for the second given the first:

(1.1)
$$H(X,Y) = H(X) + \sum_{i=1}^{N} P(X = x_i) H(Y|X = x_i).$$

The appeal of this result was a motivation favoring Shannon's choice of his axiom (iii). However, in his original article Shannon slighted his own characterization theorem for entropy, noting in a discussion (page 393) that its motivation is unclear and that it is in no way necessary for the larger theory of communication he was developing. He viewed it merely as lending plausibility to some subsequent definitions. He considered the real justification of the three axioms for entropy to reside in the useful applications they support. In particular, he regarded the implication of equation (1.1) as welcome substantiation for considering $H(\cdot)$ as a reasonable measure of information.

While the relevance of entropy to a wide array of important applications has emerged over the subsequent half-century, Shannon's attitude toward the foundational basis for entropy has persisted. As one important example, the synthetic exposition of Cover and Thomas (1991) begins directly with now common definitions required for further developments and analysis, along with an unmotivated specification of the entropy axioms. The authors found it "irresistible to play with their relationships and interpretations, taking faith in their later utility" (page 12). They did so with flair, exposing various roles understood for entropy in the fields of electrical engineering, computer science, physics, mathematics, economics and philosophy of science. In a similar vein, the stimulating published lectures of Caticha (2012) reassert and clarify this standard take on axiomatic issues. Caticha writes (page 79) that "both Shannon and Jaynes agree that one should not place too much significance on the axiomatic derivation of the entropy equation, that its use can be fully justified a-posteriori by its formal properties, for example by the various inequalities it satisfies. Thus, the standard practice is to define 'information'

as a technical term using the entropy equation and proceed. Whether this meaning is in agreement with our colloquial meaning is another issue. . . . the difference is not about the equations but about what they mean, and ultimately, about how they should be used." Caticha considers such issues in his development of a conceptual understanding of physical theory.

Forthrightly, the thoughtful discussion of Jaynes [(2003), Section 11.3] explicitly recognized and addressed the discussable open status of Shannon's third axiom characterizing entropy. Should this axiom really be required of any measure of the amount of uncertainty in a distribution? Despite recognizing its crucial role in specifying Shannon's entropy function mathematically, Jaynes was not convinced that an adequate foundation for the uniqueness claims of entropy as an information measure had been found. He concluded this long section of his book by writing (Jaynes, 2003, page 351) "Although the above demonstration appears satisfactory mathematically, it is not yet in completely satisfactory form conceptually. The functional equation (Shannon's third axiom) does not seem quite so intuitively compelling as our previous ones did. In this case, the trouble is probably that we have not yet learned how to verbalize the argument leading to [axiom (iii)] in a fully convincing manner. Perhaps this will inspire others to try their hand at improving the verbiage that we used just before writing [axiom (iii)]."

In fact, Jaynes appended an "Exercise 11.1" to his discussion, concluding with an injunction to "Carry out some new research in this field by investigating this matter; try either to find a possible form of the new functional equations, or to explain why this cannot be done." Concerns with claims regarding the uniqueness of entropy (along with other matters regarding continuous distributions which we shall address in this article) had also been aired by Kolmogorov (1956), page 105.

Nonetheless, Jaynes clearly expected that a satisfactory motivation for the special status of entropy as a measure of information would be found, thinking that his "exercise" would be resolved with a solution explaining "why this cannot be done." In a direct sense, our construction and analysis of the extropy measure shows the exercise to be solved rather by an exhibition of the long sought "new functional equation." We shall specify this in our Result 3, which provides an alternative to Shannon's third axiom and yields a different information measure. The results of the present article show that the extropy measure, far from generating inconsistencies which Jaynes feared (page 350), is actually a complementary dual of the entropy function.

The two measures are clearly distinct, yet are fundamentally intertwined with each other. In tandem with Shannon's entropy measure denoted by $H(\cdot)$, we respectfully denote our extropy measure by $J(\cdot)$. It provides a resolution to Jaynes' insightful concerns and accomplishments.

Our recognition of extropy as the complementary dual of entropy emerged from a critical analysis and completion of the logarithmic scoring rule for distributions in applied statistics. Proper scoring rules are functions of forecast distributions and the realized observations of the quantities at issue. According to the subjectivist understanding of probability and statistics as promoted by Bruno de Finetti, the assessment of proper scoring rules for proposed forecasting distributions replaces the role of hypothesis testing in objectivist methods. None of an array of proposed probability distributions can be considered to be right or wrong. Each merely represents a different point of view regarding a sequence or collection of unknown but observable quantities. The applied assessment of proper scoring rules provides a method for evaluating the comparative qualities of the competing points of view in the face of actual observed values of the quantities as they come to be known. The scoring functions are intimately related to the theory of utility. Such rules can also be used to aid in the elicitation of subjective probabilities.

The so-called logarithmic score has long been touted for its uniqueness in a specific respect relative to other proper scoring rules. The application we shall introduce raises issues concerning its incompleteness in assessing asserted distributions. We shall discuss details after the analysis of the duality of entropy and extropy is exposed. It will then be clear that the expected logarithmic score of a distribution \mathbf{p}_N coincides with $-H(\mathbf{p}_N)$, which is called negentropy. The *completion* of the log score, which is motivated for a specific application, involves the assessment of negextropy as well.

After developing the formal dual structure of the paired (entropy, extropy) functions in Sections 2–5 of this article, we shall outline in Section 6 the role that extropy plays in the scoring of forecasting distributions, using the *Total* log scoring rule. We present the axiomatization of extropy relative to entropy in Section 2, focusing on an alternative to axiom (iii). In Section 3 we display graphically the contours of the dual measures for the case of N=3. Section 4 identifies the dual equations and the complementary contraction mapping. In Section 5 we develop the theory

for continuous density functions, formalizing differential and relative (entropy, extropy) in the context of general Bregman functions. We show how relative extropy arises as a second directed distance function that is a complementary dual to the Kullback–Leibler divergence, the standard formulation of relative entropy. Section 7 presents a concluding discussion.

2. THE CHARACTERIZATION OF EXTROPY

Context: Consider an observable quantity X with possible values contained in the range $R(X) = \{x_1, x_2, \ldots, x_N\}$. The vector $\mathbf{p}_N = (p_1, p_2, \ldots, p_N)$ is composed of probability mass function values asserted for X over the event partition $[(X = x_1), (X = x_2), \ldots, (X = x_N)]$. Though we typically refer to \mathbf{p}_N as a p.m.f., we sometimes use common parlance that is an abuse of *formal* terminology, referring to it as a "distribution." To begin our discussion, we recall the following:

DEFINITION 1. The *entropy* in X or in \mathbf{p}_N equals

(2.1)
$$H(X) = H(\mathbf{p}_N) \equiv -\sum_{i=1}^{N} p_i \log(p_i).$$

We note that we use natural logarithms as opposed to base 2, and we introduce the following:

DEFINITION 2. The *extropy* in X or in \mathbf{p}_N equals

(2.2)
$$J(X) = J(\mathbf{p}_N) \equiv -\sum_{i=1}^{N} (1 - p_i) \log(1 - p_i).$$

RESULT 1. If N = 2, so X denotes merely an event, then H(X) = J(X), but when $N \ge 3$, $H(\mathbf{p}_N) > J(\mathbf{p}_N)$ as long as \mathbf{p}_N contains three or more positive components.

Clearly, $H(\mathbf{p}_2) = -p_1 \log(p_1) - (1 - p_1) \log(1 - p_1) = J(\mathbf{p}_2)$. An algebraic proof of Result 1 appears in Appendix A. However, its truth is apparent easily from computational examples. Figure 1 displays the range of possibilities for the (entropy, extropy) pairs for probability mass functions within the unit-simplexes of dimensions 1 through 6 (values of N = 2 through 7).

Evidently, the range of possible (entropy, extropy) pairs at each successive value of N incorporates the range for the previous value of N, with another section merely attached to this range. Notice particularly that the range of possible (entropy, extropy) pairs is *not* convex. As viewed across the six examples shown in Figure 1, the range exhibits convex scallops along its upper boundary: there are (N-2) scallops and one flat edge along its upper boundary for the unit-simplex of dimension (N-1). The flat edge as the northwest boundary is the line defined by H(p, 1-p) = J(p, 1-p), running in the southwest to northeast direction from (0,0) to $(-\log(0.5), -\log(0.5))$. The

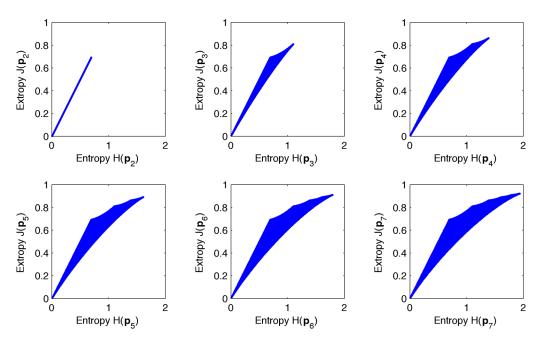


FIG. 1. The range of (entropy, extropy) pairs $(H(\cdot), J(\cdot))$ corresponding to all distributions within the unit-simplex of dimensions 1 through 6. The ranges of the quantities they assess have sizes N=2 through 7.

lower boundary of the range of pairs is a single concave scallop, ruling its own interior out of the range of possible (entropy, extropy) pairs.

RESULT 2. J(X) satisfies Shannon's axioms (i) and (ii).

The function $J(\cdot)$ is evidently continuous in its arguments [axiom (i)], and

$$J\left(\frac{1}{N}, \frac{1}{N}, \dots, \frac{1}{N}\right) = -N\left(1 - \frac{1}{N}\right)\log\left(1 - \frac{1}{N}\right)$$
$$= (N-1)\left[\log(N) - \log(N-1)\right]$$

is a monotonic increasing function of N [axiom (ii)].

2.1 Further Shared Properties of $H(\cdot)$ and $J(\cdot)$

As to other touted properties of entropy, extropy shares many of them. For example, the extropy measure is obviously permutation invariant. It is also invariant with respect to monotonic transformations of the variable X into Y = g(X). Moreover, for any size of N, the maximum extropy distribution is the uniform distribution. This can be proved by standard methods of constrained maximization using Lagrange multipliers. Let $L(\mathbf{p}_N, \lambda)$ be the Lagrangian expression for the extropy of \mathbf{p}_N subject to the constraint $\sum p_i = 1$:

$$L(\mathbf{p}_{N}, \lambda) = -\sum_{i=1}^{N} (1 - p_{i}) \log(1 - p_{i}) + \lambda \left(1 - \sum_{i=1}^{N} p_{i}\right).$$

The N partial derivatives have the form $\frac{\partial L}{\partial p_i} = \log(1 - p_i) + 1 - \lambda$. Setting each of these equal to 0 yields N equations of the form $\lambda = 1 + \log(1 - p_i)$. These N equations, together with $\frac{\partial L}{\partial \lambda} = 0$, ensure that all the p_i are equal, and thus they must each equal 1/N. Second order conditions for a maximum are satisfied at this first order solution. Analysis of the boundaries of the unit-simplex constraining \mathbf{p}_N yields the *minimum* values of extropy at the vertices: $J(\mathbf{e}_i) = 0$ for each echelon basis $\mathbf{e}_i \equiv (0, 0, \dots, 0, 1_i, 0, \dots, 0)$ with $i = 1, 2, \dots, N$.

As to differences in the two measures, notice that the scale of the maximum entropy measure is unbounded as N increases, because $H(\frac{1}{N}, \frac{1}{N}, \dots, \frac{1}{N}) = \log(N)$. In contrast, the scale of the maximum extropy is bounded by 1, for $J(\frac{1}{N}, \frac{1}{N}, \dots, \frac{1}{N}) = (N-1)\log[N/(N-1)]$. The limit of 1 can be determined by recognizing that

$$\lim_{N \to \infty} (N-1) \log \left(\frac{N}{N-1}\right)$$

$$= \lim_{N \to \infty} \log \left(1 + \frac{1}{N-1}\right)^{N-1} = \log(e) = 1.$$

2.2 The Extropy Measure of a Refined Distribution

We can now examine precisely how and why extropy does *not* satisfy Shannon's third axiom for entropy, and how it does behave with respect to measuring the refinement of a probability distribution. Algebraically, the refinement axiom for extropy arises from its definition, which yields the following result:

RESULT 3. For any positive integer N, and any values of p_i and t each in [0, 1],

$$J(p_1, ..., p_{i-1}, tp_i, (1-t)p_i, p_{i+1}, ..., p_N)$$

= $J(p_1, p_2, ..., p_N) + \Delta(p_i, t),$

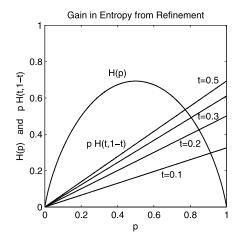
where

$$\Delta(p_i, t) = (1 - p_i) \log(1 - p_i)$$
$$- (1 - tp_i) \log(1 - tp_i)$$
$$- [1 - (1 - t)p_i] \log[1 - (1 - t)p_i].$$

This follows directly from the definition of $J(\mathbf{p}_N)$. The structure of the gain to a refined extropy, $\Delta(p_i,t)$, can be recognized by introducing a function $\varphi(p) \equiv (1-p)\log(1-p)$ and noting that $\Delta(p_i,t) = \varphi(p_i) - [\varphi(tp_i) + \varphi((1-t)p_i)]$. This difference can be shown to be always nonnegative.

Result 3 is easily interpreted visually when N = 2. The left panel of Figure 2 displays the difference between the entropies H(tp, (1-t)p, 1-p) and H(p, 1-p) according to Shannon's axiom (iii). The right panel displays the extropy J(p, 1-p) along with the difference between the extropies J(tp, (1-t)p, 1-p) and J(p, 1-p) according to Result 3. The important feature of the display is the difference between pH(t, 1-t) on the left and $\Delta(p,t)$ on the right, a difference which does not depend on the magnitude of N. In each panel, the differences are shown as functions of $p \in [0, 1]$ for the four values of t = 0.1, 0.2, 0.3 and 0.5. For any value of t, the difference functions $\Delta(p,t) = \Delta(p,t')$ for t' = (1-t).

According to Shannon's axiom (iii), the entropy for the refined mass function [tp, (1-t)p, 1-p] increases linearly with p at the rate of the entropy in the refining split factor, H(t, 1-t). In contrast, the extropy of the refined distribution increases at an increasing rate as a function of p. For small values of p, the extropy of the refined distributions increases more slowly with p than does entropy, while for large values of p it increases more quickly. When the value of p equals 1, the values of the entropy and extropy of the refined distribution equalize, for each $t \in [0, 1]$. This results from the fact



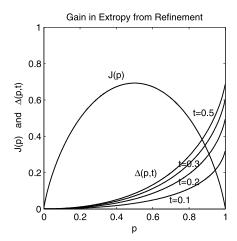


FIG. 2. Entropy and extropy for a refined distribution [tp, (1-t)p, 1-p] both equal the entropy or extropy for the base probabilities (p, 1-p) plus an additional component.

that when p = 1, the refined distribution is virtually a binary distribution (t, 1 - t, 0), for which entropy and extropy are equal. In this case the distribution being refined would be a degenerate distribution representing certainty.

As a gauge of the increase in uncertainty provided when a distribution is refined, this nonlinear feature of the extropy measure is appealing in its own right. Refining a larger probability with a splitting factor of size t may well be considered to increase the amount of uncertainty that is specified at a greater rate than when refining a smaller probability by this same factor. Consider two ways of refining a mass function $\mathbf{p}_2 =$ (0.04, 0.96), for example, into $\mathbf{p}_3 = (0.01, 0.03, 0.96)$ as opposed to $\mathbf{p}_3 = (0.04, 0.24, 0.72)$. In both cases, one of the probabilities is refined into two pieces in the ratio of 1:3. Examine the values of $\triangle(0.04, 0.25)$ and $\triangle(0.96, 0.25)$ in Figure 2(right). Although the rate of increase in entropy due to the refinement of either probability p_i is identical in the two cases, the rate of increase in extropy when refining the component $p_i = 0.04$ is nearly zero, while it is far greater when refining the larger probability component $p_i = 0.96$. It is a natural feature of the extropy function that this information measure adjusts toward the maximum entropy/extropy more quickly the more quickly the refined distribution adjusts toward the uniform.

Replacing Shannon's axiom (iii) with our Result 3 would complete an axiomatic characterization of extropy. When N = 1, the specifications of axiom (iii) and Result 3 are algebraically identical, yielding H(t, 1-t) = J(t, 1-t). When $N \ge 2$ the bifurcation first occurs. In this context, Result 3 can then be seen to

be a generator of the entire function $J(\mathbf{p}_N)$ for all values of N. The extropy function is the unique function that adheres to Shannon's axioms (i) and (ii) and to the content of Result 3, considered as an axiom.

3. ISOENTROPY, ISOEXTROPY CONTOURS IN THE UNIT-SIMPLEX

For the graphical displays that follow, we suppose that a quantity X has range $R(X) = \{1, 2, 3\}$ and that these possibilities are assessed with a probability mass function \mathbf{p}_3 in the unit-simplex \mathbf{S}^2 . Figure 3(left) displays some contours of constant entropy distributions in the 2-dimensional unit-simplex (N=3) to compare with some contours of constant extropy distributions in Figure 3(right). These contours exhibit a geometrical sense in which the extropy and entropy measures of a distribution are complementary. Whereas entropy contours sharpen into the vertices of the simplex and flatten along the faces, the extropy contours sharpen into the midpoints of the faces and flatten toward the vertices.

Further understanding can be gained from Appendix B which displays the single isoentropy contour at $H(\mathbf{p}_3) = 0.9028$ along with some members of the range of isoextropy contours that intersect with it. A computable application in astronomy is mentioned.

4. EXTROPY AS THE COMPLEMENTARY DUAL OF ENTROPY

Two behaviors identify the mathematical relation of extropy to entropy as its complementary dual. To begin, the duality is distinguished by a pair of symmetric

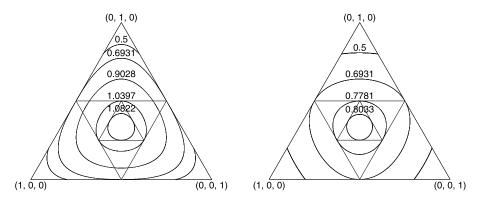


FIG. 3. At left are contours of equal entropy distributions within the 2-D unit-simplex, S^2 . At right are contours of equal extropy distributions. The relevance of the inscribed triangles shall become apparent in Section 4.

equations relating the sum of the entropy and extropy of a distribution to the entropies and extropies of their component probabilities.

RESULT 4.

$$H(\mathbf{p}_N) + J(\mathbf{p}_N) = \sum_{i=1}^{N} H(p_i, 1 - p_i)$$
$$= \sum_{i=1}^{N} J(p_i, 1 - p_i).$$

This equation for the sum of $H(\mathbf{p}_N)$ and $J(\mathbf{p}_N)$ derives from summing separately the two components of each $H(p_i, 1-p_i) = -p_i \log(p_i) - (1-p_i) \log(1-p_i) = J(p_i, 1-p_i)$ over values of $i=1,2,\ldots,N$. This simple result identifies the symmetric dual equations that relate extropy to entropy:

$$J(\mathbf{p}_N) = \sum_{i=1}^{N} H(p_i, 1 - p_i) - H(\mathbf{p}_N),$$

and symmetrically,

$$H(\mathbf{p}_N) = \sum_{i=1}^{N} J(p_i, 1 - p_i) - J(\mathbf{p}_N).$$

These two equations, symmetric in $H(\cdot)$ and $J(\cdot)$, display that the extropy of a distribution equals the difference between the sum of the entropies over the crudest partitions defined by the possible values of X, that is, $[(X = x_i), (X \neq x_i)]$, and the entropy in the finest partition they define, $[(X = x_1), (X = x_2), \ldots, (X = x_N)]$. Extropy and entropy can each be represented by the same function of the other. Since these two functions differ only in the refinement axioms that generate them, it is apparent that their symmetric duality is fundamentally related to the refinement characteristics inherent in their third axioms.

As to the complementarity of their relation, it is based on generalizing the notion of a complementary event to a complementary quantity. Relative to a probability mass function \mathbf{p}_N for a partition vector $[(X = x_1), (X = x_2), ..., (X = x_N)],$ define the complementary mass function as $\mathbf{q}_N = (N-1)^{-1}(\mathbf{1}_N \mathbf{p}_N$). The general complementary mass function \mathbf{q}_N can be considered to specify a "distribution of unlikeliness" of the possible values of X, as opposed to \mathbf{p}_N which distributes the assessed likeliness of the possible values. If N = 2, complementarity specifies $\mathbf{q}_2 =$ $(q_1, q_2) = (1 - p_1, 1 - p_2) = (p_2, p_1)$. This merely identifies the arbitrariness of analyzing an event in terms of E_1 and its complement $\tilde{E}_1 = E_2$, as opposed to $F_1 = \tilde{E}_1$ and its complement $\tilde{F}_1 = E_1$. For larger values of N, however, general complementarity generates \mathbf{q}_N from \mathbf{p}_N as a truly distinct mass function. In these terms, the general relation between H and J is that the extropy of a p.m.f. \mathbf{p}_N equals a linearly rescaled measure of entropy of its complementary p.m.f. \mathbf{q}_N .

RESULT 5.

$$J(\mathbf{p}_N) = (N-1)[H(\mathbf{q}_N) - \log(N-1)].$$

To be explicitly clear, the extropy of \mathbf{p}_N is *not* a rescaled value of the entropy of \mathbf{p}_N . It is a rescaled value of the entropy of the general complement of \mathbf{p}_N .

This result follows from simple algebra. Structurally, the entropy measure of a probability mass function has a complementary dual in its extropy measure, which derives from the entropy of a complementary mass function. In turn, this complementary mass function has its own extropy. However, this extropy value does *not* derive from the entropy of the original p.m.f., but from a further *complement of this complement*.

Most statisticians will be familiar with the notion of duality from the fact that any linear programming

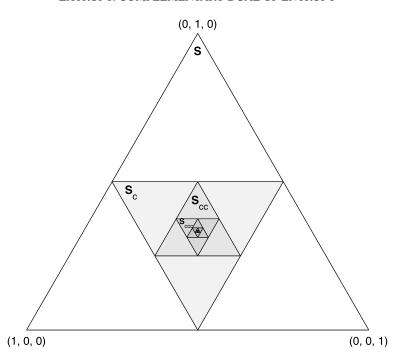


FIG. 4. The complementary distribution mapping contracts the unit-simplex S into the inscribed simplex S_c , which it contracts in turn into the inscribed S_{cc} , and then into S_{cc} and so on.

problem has a dual formulation in which the coefficient vector of the linear objective function has a dual relation with the vector of constraint values. The linear programming duality has the feature that the dual structure of a dual problem yields the original problem structure. Duals with this property are called "involutions." As we shall see now, the duality of extropy with entropy does *not* prescribe an involution, but rather a second distinct structure.

The mapping of a probability mass function \mathbf{p}_N to its complement $\mathbf{q}_N = (N-1)^{-1}(\mathbf{1}_N - \mathbf{p}_N)$ is a contraction mapping. Every mass function in a unit-simplex is mapped onto a complementary function lying within an inscribed simplex of the same dimension. In turn, this complementary mass function has its own complementary distribution lying within a simplex inscribed in that one. The fixed-point theorem for contraction mappings assures that the uniform distribution in the center of the unit-simplex is the unique mass function whose complementary mass function equals itself. Figure 4 displays the way this contraction works in two dimensions for mass functions p_3 . Notice that the points in the vertex triangles of the unit-simplex are not contraction images of any other point in the unit-simplex. Thus, the formal complementary duality of $H(\cdot)$ and $J(\cdot)$ with respect to \mathbf{p}_N and \mathbf{q}_N inheres in their forward and backward images rather than a cyclic image. The dual is not an involution.

A numerical example detailing how the isocontours of $H(\cdot)$ generate isocontours of $J(\cdot)$ appears in Appendix C.

5. DIFFERENTIAL EXTROPY AND RELATIVE EXTROPY FOR CONTINUOUS DISTRIBUTIONS

Devising the extropy measure of a continuous distribution admitting a density function yields a pleasant surprise. As to entropy, Shannon [(1948), page 628] had initially proposed that the entropy measure $-\sum p_i \log(p_i)$ has an analogue in the definition $-\int f(x) \log f(x) dx$ when the distribution function for a variable X admits a continuous density. He motivated this (page 623) by the idea that refining the categories for a discrete quantity X, with diminishing probabilities in each, yields this analogous definition in the limit. This definition has subsequently become known as "differential entropy." In a critical and constructive review, Kolmogorov (1956) concurred with Shannon's suggestion, but with qualifying reservations regarding its noninvariance with respect to monotonic transformations of the variable X and its relativity to a uniform dominating measure over the domain of X. His clarifications established a more general definition of "relative entropy" which includes differential entropy as a special case. Relative entropy was analyzed in measure theoretic detail in the classic work of Kullback (1959). Now known as the Kullback-Leibler divergence (or directed distance) between a density $f(\cdot)$ and a related absolutely continuous density $g(\cdot)$, this is defined for the continuous case as $D(f||g) \equiv \int f(x) \log \frac{f(x)}{g(x)} dx$. When g(x) is the special case of a uniform density, this reduces to Shannon's definition of differential entropy.

The dual complementarity of extropy with entropy for continuous densities can be derived in the context of relative entropy. The details are couched in the language of general Bregman functions, which unifies the discrete theory as well. We shall develop these results forthwith. For a novice reader of these ideas, the development of continuous differential entropy and extropy in the style suggested by Shannon is perhaps more instructive. It motivates the definition of differential extropy as $-\frac{1}{2} \int f^2(x) dx$. The role played by the uniform dominating measure in generating this integral will be apparent. We present an introductory analysis in Appendix D. We now begin directly by developing the more general formulation of relative extropy as the dual to relative entropy in a discrete context, and then pursuing the continuous analysis using Bregman functions.

5.1 (Relative Entropy, Relative Extropy) for Two Mass Functions: Kullback's Directed Distance and Its Complementary Dual

We continue to work in the context of a considered quantity whose possible values generate the finite partition vector $[(X = x_1), (X = x_2), ..., (X = x_N)]$. Suppose that the vector \mathbf{s}_N represents a second p.m.f., distinct from \mathbf{p}_N . In this context we recall the following:

DEFINITION 3. The *relative entropy* of \mathbf{p}_N with respect to \mathbf{s}_N is defined as the Kullback–Leibler divergence to equal

(5.1)
$$D(\mathbf{p}_N || \mathbf{s}_N) \equiv \sum_{i=1}^N p_i \log \left(\frac{p_i}{s_i} \right).$$

Notice that this definition does *not* involve a minus sign in front, as $D(\mathbf{p}_N || \mathbf{s}_N)$ is always nonnegative. It makes no difference whether the variable X is transformed by any monotone function to a new variable Y: the relative entropy in \mathbf{p}_N with respect to \mathbf{s}_N remains the same. We recall that this *directed* distance function is not symmetric in \mathbf{p}_N and \mathbf{s}_N , and thus its name.

To define the relative extropy of \mathbf{p}_N with respect to \mathbf{s}_N , we follow the same tack as in defining extropy itself:

DEFINITION 4. The relative extropy of \mathbf{p}_N with respect to \mathbf{s}_N is defined by a function complementary to

the Kullback-Leibler divergence as

$$D^{c}(\mathbf{p}_{N} || \mathbf{s}_{N}) \equiv \sum_{i=1}^{N} (1 - p_{i}) \log \left(\frac{1 - p_{i}}{1 - s_{i}} \right).$$

RESULT 6. When the p.m.f. \mathbf{s}_N happens to be the uniform p.m.f. $\mathbf{u}_N = N^{-1}\mathbf{1}_N$, the relative entropy and extropy measures return to rescaled values of the discrete entropy and extropy measures with which we are familiar:

$$D(\mathbf{p}_N || \mathbf{u}_N) = \sum_{i=1}^{N} p_i \log \left(\frac{p_i}{1/N} \right) = \log(N) - H(\mathbf{p}_N)$$
$$= H(\mathbf{u}_N) - H(\mathbf{p}_N),$$

and

$$D^{c}(\mathbf{p}_{N} || \mathbf{u}_{N}) = \sum_{i=1}^{N} (1 - p_{i}) \log \left(\frac{1 - p_{i}}{1 - 1/N} \right)$$

$$= \sum_{i=1}^{N} (1 - p_{i}) \log \left(\frac{N}{N - 1} \right)$$

$$+ \sum_{i=1}^{N} (1 - p_{i}) \log (1 - p_{i})$$

$$= (N - 1) \log \left(\frac{N}{N - 1} \right)$$

$$+ \sum_{i=1}^{N} (1 - p_{i}) \log (1 - p_{i})$$

$$= J(\mathbf{u}_{N}) - J(\mathbf{p}_{N}).$$

5.1.1 The complementary equation. It is straightforward to recognize that again, defining now two complementary mass functions $\mathbf{q}_N = (N-1)^{-1}(\mathbf{1}_N - \mathbf{p}_N)$ and $\mathbf{t}_N = (N-1)^{-1}(\mathbf{1}_N - \mathbf{s}_N)$, we find that a complementary equation identifies $D^c(\mathbf{p}_N || \mathbf{s}_N)$ as the K-L divergence between the p.m.f.'s complementary to \mathbf{p}_N and \mathbf{s}_N :

RESULT 7.

$$D^{c}(\mathbf{p}_{N}\|\mathbf{s}_{N}) = (N-1)D(\mathbf{q}_{N}\|\mathbf{t}_{N}).$$

Moreover, an alternative algebraic manipulation of Definition 4 provides that

$$D^{c}(\mathbf{p}_{N} || \mathbf{s}_{N}) = \sum_{i=1}^{N} (1 - p_{i}) \log(1 - p_{i})$$
$$- \sum_{i=1}^{N} (1 - p_{i} + s_{i} - s_{i}) \log(1 - s_{i})$$

$$= \sum_{i=1}^{N} (1 - p_i) \log(1 - p_i)$$

$$- \sum_{i=1}^{N} (1 - s_i) \log(1 - s_i)$$

$$+ \sum_{i=1}^{N} (p_i - s_i) \log(1 - s_i)$$

$$= J(\mathbf{s}_N) - J(\mathbf{p}_N) + \sum_{i=1}^{N} p_i \log\left(\frac{1 - s_i}{N - 1}\right)$$

$$- \sum_{i=1}^{N} s_i \log\left(\frac{1 - s_i}{N - 1}\right),$$

because $\sum_{i=1}^{N} (p_i - s_i) \log(N - 1) = 0$. This yields another interesting and useful representation:

RESULT 8.

$$D^{c}(\mathbf{p}_{N} || \mathbf{s}_{N}) = J(\mathbf{s}_{N}) - J(\mathbf{p}_{N}) + E_{\mathbf{p}_{N}} [\log(t^{o}(X))] - E_{\mathbf{s}_{N}} [\log(t^{o}(X))],$$

where
$$t^{o}(X) \equiv \sum_{i=1}^{N} (X = x_i)t_i$$
.

That is, $t^o(X)$ equals the component probability in the \mathbf{t}_N vector associated with the value of X that happens to be observed. This holds algebraically because one of the event indicators, $(X = x_i)$, equals 1 (since the equation it indicates is true) while the other (N-1) event indicators equal 0. The equations they indicate are false.

The relative extropy value of \mathbf{p}_N relative to \mathbf{s}_N equals the difference in their extropy values, adjusted by a difference in two expectations of a specific log mass function value: the mass function component of \mathbf{t}_N associated with the particular partition event that is found to occur. This is the mass function that is complementary to \mathbf{s}_N . The usefulness of Result 8 shall arise as a motivation for a definition of relative extropy between two densities.

The analogous result pertinent to the K-L divergence, deriving from (5.1) would be as follows:

RESULT 8'.

$$D(\mathbf{p}_N || \mathbf{s}_N) = H(\mathbf{s}_N) - H(\mathbf{p}_N)$$
$$- E_{\mathbf{p}_N} [\log(s^o(X))] + E_{\mathbf{s}_N} [\log(s^o(X))],$$

where
$$s^{o}(X) \equiv \sum_{i=1}^{N} (X = x_i) s_i$$
.

5.1.2 Relative (entropy, extropy) of complementary mass functions. A final note of interest concerns the pair of relative (entropy, extropy) assessments between complementary mass functions such as \mathbf{p}_N and \mathbf{q}_N . The relative entropy of \mathbf{p}_N with respect to \mathbf{q}_N equals a translated expected value of the asserted log odds ratio in favor of the occurring partition event: $D(\mathbf{p}_N || \mathbf{q}_N) = \sum_{i=1}^N p_i \log(\frac{p_i}{1-p_i}) + \log(N-1)$. Intriguingly, but again deriving easily from a direct application of Definition 4, their relative extropy also equals (N-1) times an expected log odds ratio in favor of the occurring partition event too. However, this odds ratio is assessed in terms of the complementary distribution of unlikeliness, \mathbf{q}_N , rather than in terms of the usual distribution of likeliness, \mathbf{p}_N :

$$D^{c}(\mathbf{p}_{N} \| \mathbf{q}_{N})$$

$$= (N-1) \left[\sum_{i=1}^{N} q_{i} \log \left(\frac{q_{i}}{1-q_{i}} \right) + \log(N-1) \right].$$

Both of these interpretations as expected log odds ratios are adjusted by an additive constant, $\log(N-1)$. This additive constant can be recognized as the expected log odds associated with a uniform distribution: $\sum_{i=1}^{N} u_i \log(\frac{u_i}{1-u_i}) = \sum_{i=1}^{N} (1/N) \log(\frac{1/N}{(1-1/N)}) = -\log(N-1)$. Thus, we have an interesting pair of representations for the relative (entropy, extropy) between complementary mass functions:

RESULT 9.

$$D(\mathbf{p}_N \| \mathbf{q}_N)$$

$$= E_{\mathbf{p}_N} \left[\log \left(\frac{p^o}{1 - p^o} \right) \right] - E_{\mathbf{u}_N} \left[\log \left(\frac{u^o}{1 - u^o} \right) \right],$$

and

$$D^{c}(\mathbf{p}_{N} \| \mathbf{q}_{N}) = (N-1) \left\{ E_{\mathbf{q}_{N}} \left[\log \left(\frac{q^{o}}{1-q^{o}} \right) \right] - E_{\mathbf{u}_{N}} \left[\log \left(\frac{u^{o}}{1-u^{o}} \right) \right] \right\},$$

where p^o , q^o and u^o are the probabilities assessed for the value of X that happens to be observed, as assessed according to the p.m.f.'s \mathbf{p}_N , \mathbf{q}_N and \mathbf{u}_N , respectively.

5.1.3 Unifying $D(\cdot||\cdot)$ and $D^c(\cdot||\cdot)$ as Bregman divergences. The theory of Bregman functions both unifies our understanding of the (entropy, extropy) duality and provides the basis for formalizing their functional representations for continuous densities. In this context it will yield still another surprise. The text of Censor and Zenios (1997) develops the general theory

of Bregman functions and a wide variety of applications. In the definition below we recall the notion of Bregman divergence from Banerjee et al. (2005):

DEFINITION 5. Let \mathcal{C} be a convex subset of \Re^N with a nonempty relative interior, denoted by $\operatorname{ri}(\mathcal{C})$. Let $\Phi: \mathcal{C} \to \Re$ be a strictly convex function, differentiable in $\operatorname{ri}(\mathcal{C})$. For $\mathbf{p}_N, \mathbf{s}_N \in \mathcal{C}$ the *Bregman divergence* $d_\Phi: \mathcal{C} \times \operatorname{ri}(\mathcal{C}) \to \Re$ corresponding to Φ is given by

$$d_{\Phi}(\mathbf{p}_{N}, \mathbf{s}_{N}) = \Phi(\mathbf{p}_{N}) - \Phi(\mathbf{s}_{N}) - \langle \nabla \Phi(\mathbf{s}_{N}), (\mathbf{p}_{N} - \mathbf{s}_{N}) \rangle,$$

where $\nabla \Phi(\mathbf{s}_N)$ is the gradient vector of Φ evaluated at \mathbf{s}_N and the angle brackets $\langle \cdot, \cdot \rangle$ denote "inner product." The function $\Phi(\cdot)$ is called a *Bregman function*.

An important special case of the Bregman function reduces its action to the sum of a common function applied to each of the components of a vector, that is, $\Phi(\mathbf{p}_N) = \sum_{i=1}^N \phi(p_i)$. In this case the Bregman divergence is said to be "separable" (Stummer and Vajda, 2012), with the form

(5.3)
$$d_{\Phi}(\mathbf{p}_{N}, \mathbf{s}_{N}) = \sum_{i=1}^{N} [\phi(p_{i}) - \phi(s_{i}) - \phi'(s_{i})(p_{i} - s_{i})].$$

A standard application of the separable case identifies the Shannon entropy as a Bregman divergence. Consider the component function $\phi(p) = \varphi_1(p)$, where $\varphi_1(p) \equiv p \log(p)$, which identifies the vector Bregman function as $\Phi(\mathbf{p}_N) = -H(\mathbf{p}_N)$. Since $\phi'(p) = \log(p) + 1$, a direct application of the separable Bregman divergence form (5.3) yields the following well-known result, which is reported in Banerjee et al. (2005):

RESULT 10. The Bregman divergence associated with $\Phi(\mathbf{p}_N) = -H(\mathbf{p}_N)$ is

$$d_{\Phi}(\mathbf{p}_N, \mathbf{s}_N) = \sum_{i=1}^N p_i \log \left(\frac{p_i}{s_i} \right) = D(\mathbf{p}_N || \mathbf{s}_N).$$

The same Bregman divergence results from the separable component function $\phi(p) = \varphi_2(p)$, where $\varphi_2(p) \equiv p \log(p) + (1-p)$.

As to extropy, again in the separable case consider the component function $\phi^c(p) = \varphi_1^c(p)$, where $\varphi_1^c(p) \equiv \varphi_1(1-p) = (1-p)\log(1-p)$. This identifies the vector Bregman function as $\Phi^c(\mathbf{p}_N) = -J(\mathbf{p}_N)$. Since $\phi^{c'}(p) = -\log(1-p) - 1$, another direct application of (5.3) yields a complementary result regarding $D^c(\cdot||\cdot)$:

RESULT 11. The Bregman divergence associated with $\Phi^c(\mathbf{p}_N) = -J(\mathbf{p}_N)$ is

$$d_{\Phi^c}(\mathbf{p}_N, \mathbf{s}_N) = \sum_{i=1}^N (1 - p_i) \log \left(\frac{1 - p_i}{1 - s_i} \right)$$
$$= D^c(\mathbf{p}_N || \mathbf{s}_N).$$

This same Bregman divergence also results from the Bregman function associated with $\phi^c(p) = \varphi_2(1-p)$, where $\varphi_2(1-p) \equiv (1-p)\log(1-p) + p$.

It is clear that the duality of entropy and extropy persists through the representation of relative (entropy, extropy) as complementary Bregman divergences for dual Bregman functions.

5.2 (Relative Entropy, Relative Extropy) for Continuous Densities

The unification of the general theory of directed distances formulated via Bregman functions provides the representations of entropy and extropy for continuous densities as well. Similar to the form of the separable Bregman divergence between two vectors, the Bregman directed distance between two density functions $f(\cdot)$ and $g(\cdot)$ defined on $[x_1, x_N]$, associated with a function $\phi(\cdot)$, is denoted by $B_{\phi}(f, g)$, defined to equal

$$\int_{x_1}^{x_N} \{ \phi(f(x)) - \phi(g(x)) - \phi'(g(x)) [f(x) - g(x)] \} dx.$$

The function $\phi:(0,\infty)\to\Re$ should be differentiable and strictly convex, and the limits $\lim_{x\to 0}\phi(x)$ and $\lim_{x\to 0}\phi'(x)$ must exist (in some topology), but not necessarily be finite. See Frigyik, Srivastava and Gupta (2008), page 1681, and Basseville (2013), page 623. Moreover, the integral operation is constrained to be an integration over the two functions' common domain.

It is well known that when $\phi(f) = \varphi_1(f) \equiv f \log(f)$, or $\phi(f) = \varphi_2(f) \equiv f \log(f) + (1 - f)$, specifying a convex function defined on $[0, +\infty)$ which satisfies these conditions, then

$$B_{\phi}(f,g) = \int_{x_1}^{x_N} f(x) \log \left(\frac{f(x)}{g(x)}\right) dx.$$

This Bregman directed distance is known as the relative entropy between the two densities, denoted by d(f||g).

To specify the relative extropy between two densities $f(\cdot)$ and $g(\cdot)$, we begin by recalling the relative extropy between the mass functions \mathbf{p}_N and \mathbf{s}_N as represented in the equality following (5.2):

(5.4)
$$D^{c}(\mathbf{p}_{N} || \mathbf{s}_{N}) = J(\mathbf{s}_{N}) - J(\mathbf{p}_{N}) + \sum_{i=1}^{N} (p_{i} - s_{i}) \log(1 - s_{i}).$$

On the basis of its Maclaurin series expansion, the function $(1-p_i)\log(1-p_i)\approx -p_i+\frac{1}{2}p_i^2$ when p_i is small and, thus, $J(\mathbf{p}_N)=-\sum_{i=1}^N (1-p_i)\log(1-p_i)\approx 1-\frac{1}{2}\sum_{i=1}^N p_i^2$ when max p_i is small. Of course, a similar result pertains to $J(\mathbf{s}_N)$. Moreover, the common recognition that $\log(1-s_i)\approx -s_i$ for small values of s_i yields $(p_i-s_i)\log(1-s_i)\approx -p_is_i+s_i^2$. Applying these two approximations (which agree with the bivariate Maclaurin series expansion through order 3) to equation (5.4) yields the surprising recognition that

(5.5)
$$D^{c}(\mathbf{p}_{N}||\mathbf{s}_{N}) \approx \frac{1}{2} \sum_{i} (p_{i} - s_{i})^{2}$$

when both max p_i and max s_i are small.

This is one-half the usual squared Euclidean distance between the vectors \mathbf{p}_N and \mathbf{s}_N ; moreover, it is also the Bregman divergence associated with the component function $\phi(p) = \varphi_3(p) \equiv -p + \frac{p^2}{2}$ or $\phi(p) = \varphi_4(p) \equiv \frac{p^2}{2}$.

A sensible definition for the relative extropy between two densities arises from each of two consequences of this fact. First, replacing the two component arguments of $D^c(\mathbf{p}_N || \mathbf{s}_N)$ in (5.5) by $p_i = f(x_i) \Delta x$ and $s_i = g(x_i) \Delta x$, as when motivating the definitions of differential (entropy, extropy) in Appendix D, we find that

$$\lim_{\Delta x \to 0} \frac{D^c(\mathbf{p}_N || \mathbf{s}_N)}{\Delta x} = \frac{1}{2} \int_{x_1}^{x_N} \left[f(x) - g(x) \right]^2 dx.$$

Second, this same formulation arises from evaluating the Bregman divergence between the densities $f(\cdot)$ and $g(\cdot)$ over a closed interval $[x_1, x_N]$ corresponding to either of the convex functions $\phi(f) = \varphi_3(f)$ or $\phi(f) = \varphi_4(f)$, where $\varphi_3(f) = -f + \frac{1}{2}f^2$ and $\varphi_4(f) = \frac{1}{2}f^2$, viz..

$$B_{\phi}(f,g) = \frac{1}{2} \int_{x_1}^{x_N} [f(x) - g(x)]^2 dx.$$

Motivated by these two results, we define the following:

DEFINITION 6. The *relative extropy in a density* $f(\cdot)$ relative to $g(\cdot)$ defined over $[x_1, x_N]$ is

$$d^{c}(f||g) \equiv \frac{1}{2} \int_{x_{1}}^{x_{N}} [f(x) - g(x)]^{2} dx.$$

The status of relative entropy and half the L_2 metric as Bregman divergences are well known. However, they have never been recognized heretofore as formulations of the complementary duals, entropy and extropy.

For example, Censor and Zenios [(1997), page 33] refer to these as "the most popular Bregman functions," without any hint how they are related.

We should expressly clarify that the duality of entropy and extropy we are touting is distinct from the Legendre duality between points and lines that underlies the general structure of Bregman divergences. See Boissonnat, Nielsen and Nock (2010), Section 2.2. Ours is a content-based duality that derives from their symmetric co-referential relation which we exposed following Result 4 in Section 4. In this regard it is quite surprising and provocative that half the squared L_2 distance (the relative extropy between two densities) arises as the dual of the entropic norm of Kullback–Leibler.

It is satisfying that a final result codifies the definitions of Shannon's "analogue" differential entropy function $h(f) \equiv -\int_{x_1}^{x_N} f(x) \log(f(x)) dx$ and our differential extropy function $j(f) \equiv -\frac{1}{2} \int f^2(x) dx$ (discussed in Appendix D) as a special case of their relative measures with respect to a uniform density:

RESULT 12. Suppose $f(\cdot)$ is any density defined on $[x_1, x_N]$ and that $u(x) = (x_N - x_1)^{-1}$ is a uniform density. Then the relative (entropy, extropy) pair identify the differential (entropy, extropy) forms

$$d(f||u) = h(u) - h(f)$$

and

$$d^{c}(f||u) = j(u) - j(f).$$

Recalling from Result 7 the relation of relative extropy $D^{c}(\mathbf{p}_{N} \| \mathbf{s}_{N})$ to the relative entropy in the complementary mass functions via $D(\mathbf{q}_N || \mathbf{t}_N)$, it would seem natural to search for the general relative extropy measure between any two densities by searching for an appropriate complementary density to a density. As it turns out, such a search would be chimeric because the complementary density to every density is identical ... the uniform density! This can be recognized by examining the complementary mass function $\mathbf{q}_N \equiv (\mathbf{1}_N - \mathbf{p}_N)/(N-1)$. In the limiting process we have devised, the value of N increases while the maximum value of the \mathbf{p}_N vector becomes small, with each component of \mathbf{p}_N converging toward zero. In the process, each of their complementary p.m.f. components becomes indistinguishable from $\frac{1}{N}$. Thus, the complementary density values become uniform everywhere.

This argument also implies that the values of the two expectations in the limiting equation of Result 8, $E_{\mathbf{p}_N}[\log(t^o(X))]$ and $E_{\mathbf{s}_N}[\log(t^o(X))]$, both become

indistinguishable from log(N) as N increases. This is the entropy of a uniform p.m.f. Thus, in the limit their *difference* equals 0.

6. STATISTICAL APPLICATION TO PROPER SCORING RULES

Our discovery of extropy was stimulated by a problem that arises in the application of the theory of proper scoring rules for alternative forecast distributions. These functions are the central construct of a subjectivist statistical practice used to evaluate the relative quality of different asserted distributions. A proper scoring rule $S(\mathbf{p}_N, X = x^o)$ is a function of both the p.m.f. assertion and the observation value, with the property that the expected scoring function value (with respect to the asserted p.m.f. \mathbf{p}_N) exceeds the expected score to be achieved by any other p.m.f. The application of such rules for theory comparison is said to promote honesty and accuracy in one's assessment of a p.m.f. to assert. There are many proper scoring functions. DeGroot (1984) discusses the relation of the various scoring functions to differing utility functions.

Proper scoring rules were the last applied statistical topic addressed in the publications of Savage (1971). They are presented systematically and promoted in the text of Lad (1996). Theory and applications over the past half century have been reviewed by Gneiting and Raftery (2007). The log probability for the observed outcome of $X = x^o$ is widely considered to be an eminent proper scoring rule and has been used extensively: $S_{\log}(\mathbf{p}_N, X = x^o) = \sum_{i=1}^{N} (X = x_i) \log p_i = \log(p^o)$. This score has long been recognized to be the unique proper scoring rule for distributions that are a function *only* of the observed value of $X = x^o$, irrespective of the probabilities assessed for the "unobserved" possibilities of X. See Shuford, Albert and Massengill (1966) and Bernardo (1979). The probability assessor's expected logarithmic score equals the negentropy in the assessed distribution:

$$E_{\mathbf{p}_N}[S_{\log}(\mathbf{p}_N, X = x^o)] = \sum_{i=1}^N p_i \log(p_i).$$

It now appears that the logarithmic score's claim to fame should be viewed as a weakness rather than a virtue, for it provides an *incomplete* assessment of the probabilities composing \mathbf{p}_N . The recognition of extropy as the complementary dual of entropy plays on the fact that the observation of $X = x^o$ is concomitant with the observations that $X \neq x_i$ for *every other* x_i in the range of X that is different from x^o . Probabilities

for these *observed* negated events are inherent in the assertion of \mathbf{p}_N , yet the logarithmic scoring function ignores them. The *total logarithmic scoring rule* has been proposed to address this issue:

$$S_{\text{Totallog}}(\mathbf{p}_N, X = x^o)$$

$$\equiv \sum_{i=1}^{N} (X = x_i) \log p_i + \sum_{i=1}^{N} (X \neq x_i) \log(1 - p_i).$$

Evidently, the expectation of this score equals the negentropy plus the negextropy of the distribution:

$$E_{\mathbf{p}_{N}}[S_{\text{Totallog}}(\mathbf{p}_{N}, X = x^{o})]$$

$$= \sum_{i=1}^{N} p_{i} \log(p_{i}) + \sum_{i=1}^{N} (1 - p_{i}) \log(1 - p_{i}).$$

Moreover, each component sum and any positive linear combination of the two components of the Total log score is a proper score as well.

A preliminary report by Lad, Sanfilippo and Agrò (2012) investigates the importance of this issue in an application scoring alternative forecasting distributions for daily stock prices (Agrò, Lad and Sanfilippo, 2010). The distributions considered differ in the attitudes they portray toward tail area probabilities, and the two components of the Total log score assess the expected price and the tail area probabilities in different ways. The international financial collapse of recent years has accentuated an awareness of the importance of evaluating probabilities for extreme events that seldom occur, even when they don't occur. One of the major insights the report provides is that the quadratic scoring rule for distributions should be considered not as an alternative to the usual log score but as a complement. For while the utility of a price forecast surely does derive from decisions that depend on the expected prices, it also hinges on the level of insurance cover suggested by the forecasting distribution to protect against extreme outcomes. It should become standard practice to evaluate the logarithmic score and the quadratic score in tandem. This conclusion derives from the same logic we have used in this article in identifying the squared L_2 distance as the extropic complement to the Kullback-Leibler formulation of relative entropy.

Further applications of this notion are already being promoted. An extension of the total log proper scoring rule for probability distributions to *partial probability assessments* has been given in Capotorti, Regoli and Vattari (2010) as a discrepancy measure between a conditional assessment and the class of unconditional probability distributions compatible with the

assessments that *are* made. Taking the work of Predd et al. (2009) as a starting point, Gilio and Sanfilippo (2011) use the extension of a scoring rule to partial assessments while analyzing the Total log score as a particular Bregman divergence. Biazzo, Gilio and Sanfilippo (2012) address the case of conditional prevision assessments.

7. CONCLUDING DISCUSSION

What's in a name? We are aware of prior uses of the word "extropy," documented in both the Online Oxford English Dictionary and in Wikipedia. In one usage it seems to have arisen as a metaphorical term rather than a technical term, naming a proposed primal generative natural force that stimulates order rather than disorder in both physical and informational systems. In the other usage within a technical context, "extropy" has apparently had some parlance being used interchangeably with the more commonly used "negentropy," the negative scaling of entropy. Neither usage of "extropy" appears to be very common. While we are not stuck on this particular word, the information measure we have introduced in this article seems aptly to merit the coinage of "extropy." Whereas entropy is recognized as minus the expected log probability of the occurring value of X (a measure which could be considered "interior" to the observation X), our proposed extropy is derived from the expected log nonoccurrence probability for the partition event that does occur less the sum of log nonoccurrence probabilities, that is, $\sum_{i=1}^{N} p_i \log(1 - p_i) - \sum_{i=1}^{N} \log(1 - p_i)$. This could be considered to be a measure "exterior" to the observation X. The exterior measure of all the nonoccurring quantity possibilities is complementary to the entropy measure of the unique occurring possibility. Together, in their joint assessment of the information inhering in a system of probabilities, entropy and extropy identify what many people think of as yin and yang, and what artists commonly refer to as positive and negative space.

A word is in order about concerns of mathematical statisticians regarding the limitations of the theory of continuous information measures. These typically revolve upon measurability conditions and the limitation of continuous extropy to L_2 densities. In our present digital age, the time has surely come for statistical theorists to come to grips with the fact that *every* statistical measurement procedure in any field whatsoever is actually limited to a finite and discrete set of possible measurement values. No one has ever observed a

real-valued measurement of anything. The actual application of statistics to inference or estimation problems involves only discrete finite quantities. Of course, continuous mathematics is useful for approximate computations in situations of fine measurements. However, such approximations need not require every imaginable feature of mathematical structures for real computational problems. This outlook stands in contrast to received attitudes from earlier centuries. These were based on the notion that reality is actually continuous and that numerical methods of applied mathematics can only yield discrete approximations. We ought to recognize that such notions are now outdated.

The statistical application to proper scoring rules that we outlined in Section 6 is one of many areas of possible relevance of our dual construction. In any commercial or scientific arena in which entropic computations have become standard, such as astronomical measurements of heat distribution in galaxies, the insights provided by extropic computations would be well worth investigating. Unrecognized heretofore, the relevance of the duality may lie hidden in applications already conducted and may become apparent more widely now that it is recognized. For example, terms comprising the difference of extropy from entropy arise in a representation of the Bethe free energy and the Bethe permanent in Vontobel [(2013), pages 7–8], though they are not recognized there as such. Even earlier, the Fermi-Dirac entropy function applied in nuclear physics specifies the sum of extropy and entropy as its Bregman divergence without recognizing the duality of the two components. See Furuichi and Mitroi (2012). Given the broad range of applications of entropy on its own over the past half century, we suspect that the awareness of extropy as its complementary dual will raise as many new interesting questions as it answers.

APPENDIX A: ENTROPY ≥ EXTROPY

Let X be a random quantity with a finite discrete realm of possibilities $\{x_1, x_2, ..., x_N\}$ with probability masses p_i , with $p_i = P(X = x_i)$, i = 1, ..., N. We recall that $H(X) = -\sum_{i=1}^{N} p_i \log(p_i)$ and $J(X) = -\sum_{i=1}^{N} (1 - p_i) \log(1 - p_i)$. We consider the following real functions defined on [0, 1]:

$$\varphi_1(p) = p \log(p), \quad \text{with } 0 \log(0) \equiv 0;$$

$$\varphi_1^c(p) = \varphi_1(1-p);$$

$$u(p) = -(\varphi_1(p) - \varphi_1^c(p))$$

$$= -p \log(p) + (1-p) \log(1-p).$$

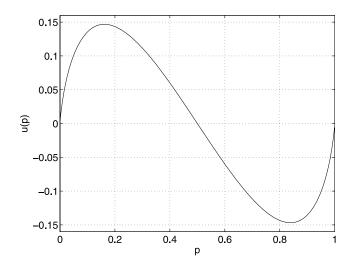


FIG. A.1. *The function* u(p).

The function u(p) satisfies the following properties (see Figure A.1):

- 1. u(p) = 0 iff $[p = 0, \text{ or } p = 1 \text{ or } p = \frac{1}{2}]$;
- 2. u(p) > 0 iff 0 ;
- 3. u(p) < 0 iff $\frac{1}{2} ;$
- 4. u(1-p) = -u(p), for all $p \in [0, 1]$;
- 5. u(p) is strictly concave in $[0, \frac{1}{2}]$, that is, for any given pair (p_1, p_2) with $0 \le p_1 < p_2 \in (0, \frac{1}{2}]$, and for any given $\alpha \in (0, 1)$, we have

$$u(\alpha p_1 + (1 - \alpha)p_2) > \alpha u(p_1) + (1 - \alpha)u(p_2).$$

By exploiting the function u(p), it is evident that

$$H(X) - J(X) = \sum_{i=1}^{N} u(p_i).$$

This difference is permutation invariant with respect to the components p_i .

We observe that for any N > 1, if there exist $i \in \{1, 2, ..., N\}$ such that $p_i = 0$, then by considering an arbitrary quantity Y with a realm of cardinality N - 1 and probability masses $(p_1, p_2, ..., p_{i-1}, p_{i+1}, ..., p_N)$ we are ensured that

$$H(X) = H(Y)$$
 and $J(X) = J(Y)$.

We have the following result:

Let *X* be a finite random quantity, with realm $\{x_1, x_2, ..., x_N\}$ and probability masses $(p_1, p_2, ..., p_N)$ such that $p_i > 0$, for i = 1, 2, ..., N, we have the following:

(a)
$$H(X) = J(X)$$
 if $N \le 2$;

(b)
$$H(X) > J(X)$$
 if $N > 2$.

Case (a). If N = 1, we trivially have H(X) = J(X) = 0 and, if N = 2, it is $H(X) = J(X) = -p_1 \log(p_1) - (1 - p_1) \log(1 - p_1)$.

Case (b). We distinguish two alternatives: (b1) $p_i \le \frac{1}{2}$, i = 1, 2, ..., N; and (b2) $p_i > \frac{1}{2}$ for only one index i

Case (b1). By the hypotheses, for each i, $0 < p_i \le \frac{1}{2}$ and $\sum_{i=1}^{N} p_i = 1$. It follows from Properties 1 and 2 of the function u(p) that

$$H(X) - J(X) = \sum_{i=1}^{N} u(p_i) > 0.$$

Case (b2). To begin, suppose that N=3. Without loss of generality, we can assume $p_3 > \frac{1}{2}$, because of the permutation invariance of $u(\cdot)$; consequently, $0 < p_1 + p_2 < \frac{1}{2}$. Now from Property 4 we deduce

$$u(p_3) = -u(1 - p_3) = -u(p_1 + p_2).$$

Then statement

$$H(X) - J(X) = u(p_1) + u(p_2) - u(p_1 + p_2) > 0$$

amounts to $u(p_1) + u(p_2) > u(p_1 + p_2)$. Since u(p) is strictly concave over the interval $[0, \frac{1}{2}]$ (see Property 5) and u(0) = 0, we have

$$u(p_1) = u\left(\frac{p_2}{p_1 + p_2}0 + \frac{p_1}{p_1 + p_2}(p_1 + p_2)\right)$$

$$> \frac{p_2}{p_1 + p_2}u(0) + \frac{p_1}{p_1 + p_2}u(p_1 + p_2)$$

$$= \frac{p_1}{p_1 + p_2}u(p_1 + p_2)$$

and

(A.2)
$$u(p_2) > \frac{p_1}{p_1 + p_2} u(0) + \frac{p_2}{p_1 + p_2} u(p_1 + p_2) = \frac{p_2}{p_1 + p_2} u(p_1 + p_2).$$

From (A.1) and (A.2) it follows $u(p_1) + u(p_2) > u(p_1 + p_2)$ and then H(X) - J(X) > 0.

Generally, let N > 2. Again without loss of generality, we can assume $p_N > \frac{1}{2}$. We have

$$u(p_N) = -u(1-p_N) = -u(p_1+p_2+\cdots+p_{N-1}).$$

For each i = 1, ..., N - 1, it is easy to see that

(A.3)
$$u(p_i) > \frac{p_i}{p_1 + p_2 + \dots + p_{N-1}} \cdot u(p_1 + p_2 + \dots + p_{N-1}),$$

because of the concavity of $u(\cdot)$.

Finally, we have

$$H(X) - J(X)$$

$$= \sum_{i=1}^{N} u(p_i)$$

$$= \sum_{i=1}^{N-1} u(p_i) - u(p_1 + p_2 + \dots + p_{N-1})$$
> 0

APPENDIX B: THE RANGE OF EXTROPY VALUES THAT SHARE AN ENTROPY

In the same observational context as Figure 3, Figure A.2 displays a single entropy contour at the value $H(\mathbf{p}_3) = 0.9028$. Inscribed and exscribed are the maximum and minimum extropy contours that intersect with it. Each of these extreme extropy contours has three intersection points with the entropy contour, and the p.m.f. that each of these points represents has two equal components. So the three triples constituting the mass function intersection points on both the max and the min J contours are permutations of one another. The intermediate extropy contour intersects the $H(\mathbf{p}_3) = 0.9028$ contour at six points, the six per-

mutations of a \mathbf{p}_3 vector with *three distinct* components. Both the $H(\cdot)$ and $J(\cdot)$ functions are permutation invariant. In higher dimensions, the intersection of $H(\mathbf{p}_N)$ and $J(\mathbf{p}_N)$ contours yields surfaces in (N-2) dimensions that are symmetric across the permutation kernels of the unit-simplex \mathbf{S}^{N-1} .

When the entropy is calculated for any assemblage such as the heat distribution for a galaxy of stars, a companion calculation of the extropy would allow us to complete our understanding of the variation inherent in its empirical distribution. The extropy value completes the measure of disorder in the array, placing it within the extremes that are possible for the calculated entropy value.

APPENDIX C: ISOCONTOURS OF $H(\cdot)$ GENERATE ISOCONTOURS OF $J(\cdot)$ VIA RESULT 5

As a numerical and geometrical example, consider again Figure 3 in the context of the following computational results. These need to be compared with the points they represent in the figure as you go. To begin, notice that $H(\frac{1}{4}, \frac{1}{2}, \frac{1}{4}) = 1.0397$ and $J(\frac{1}{4}, \frac{1}{2}, \frac{1}{4}) = 0.7781$ identify the points at the apex of specific isoentropy and isoextropy contours from your perspective as you view the left and right sides of Figure 3. Both of these contours lie precisely on and are

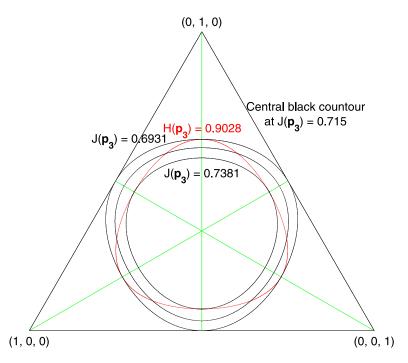


FIG. A.2. The iso-entropy contour, $H(\mathbf{p}_3) = 0.9028$, intersects with each of the inscribed and the exscribed iso-extropy contours at three points, and it intersects with any intermediate iso-extropy contour at six points. The three lines bisecting the vertex angles partition the unit-simplex into six symmetric permutation kernels.

tangent to the triangular sub-simplex S_c that is inscribed within the unit-simplex S^2 in Figure 3(left) and Figure 3(right). Result 5 tells us that the source of this isoextropy contour on the right is the higher level isoentropy contour H = 1.082 that contains the point $\mathbf{q}_3 = (\frac{3}{8}, \frac{1}{4}, \frac{3}{8})$ at the bottom of this entropy contour. This is the mass function complementary to $\mathbf{p}_3 = (\frac{1}{4}, \frac{1}{2}, \frac{1}{4})$. Computationally, $J(\mathbf{p}_3 = (\frac{1}{4}, \frac{1}{2}, \frac{1}{4})) =$ $0.7781 = 2[H(\mathbf{q}_3 = (\frac{3}{8}, \frac{1}{4}, \frac{3}{8})) - \log(2)] = 2[1.0822 -$ 0.6931], as prescribed by Result 5. Transformed into an isoextropy contour, this isoentropy contour containing $H(\mathbf{q}_3) = 1.0822$ is flipped and expanded to represent $J(\mathbf{p}_3) = 0.7781$. If we would *begin* with a consideration of the *entropy* contour containing $H(\frac{3}{8}, \frac{1}{4}, \frac{3}{8}) = 1.0822$, regarding *this* triple as \mathbf{p}_3 , we would find its dual extropy contour is denominated J = 0.8033, containing the member $J(\frac{3}{8}, \frac{1}{4}, \frac{3}{8}) = 0.8033$. These two contours are precisely inscribed in the sub-sub-simplex S_{cc} which is inlaid within S_c in Figure 3(left) and Figure 3(right). This visualization completes our understanding of extropy as the complementary dual of entropy.

APPENDIX D: DIFFERENTIAL ENTROPY AND EXTROPY FOR CONTINUOUS DENSITIES

We begin this exposition by reviewing how the analogical character of Shannon's differential entropy measure for a continuous density derives from its status as the limit of a linear transformation of the discrete entropy measure.

D.1 Shannon's Differential Entropy: $-\int f(x) \log f(x) dx$

For the following simple exposition of Shannon's considerations, presume again that the range of a quantity X is $\{x_1, \ldots, x_N\}$ and that the values of x_1 and x_N are fixed. For each larger value of N, presume that more elements are included uniformly in the interval between them and that the p_i values are refined in such a way that the maximum p_i value reduces toward 0. Now define $\Delta x \equiv (x_N - x_1)/(N - 1)$ for any specific N, and define $f(x_i) \equiv p_i/\Delta x$. In these terms, the entropy $H(\mathbf{p}_N)$ can be expressed as

$$H(\mathbf{p}_N) = -\sum p_i \log(p_i)$$
(D.1)
$$= -\sum f(x_i) \triangle x \log(f(x_i) \triangle x)$$

$$= -\sum f(x_i) \log(f(x_i)) \triangle x - \log(\triangle x).$$

Thus, the entropy measure $H(\mathbf{p}_N)$ is unbounded as N increases, with $\Delta x \to 0$. However, the summand $-\sum f(x_i)\log(f(x_i))\Delta x$ on (D.1) is merely a location transform of the entropy $-\sum p_i\log(p_i)$, shifting only by $\log(\Delta x)$ which is finite for any N. The limit of the relocated entropy expression suggests Shannon's definition of the continuous analogue:

DEFINITION D.1. The differential entropy of a density $f(\cdot)$ over the interval $[x_1, x_N]$ is defined as

$$h(f) \equiv -\int_{x_1}^{x_N} f(x) \log(f(x)) dx$$
$$= \lim_{\Delta x \to 0} [H(\mathbf{p}_N) + \log \Delta x].$$

Shannon himself noted that this analogous measure loses the absolute meaning that the finite measure enjoys, because its value must be considered relative to an assumed standard of the coordinate system in which the value of the variable is expressed. If the variable X were transformed into Y, then the continuous measure of the differential entropy $h_Y(f(\cdot))$ needs to be adjusted from $h_X(f(\cdot))$ by the Jacobian of the specific transformation. He suggested, however, that the continuous analogue retains its value as a *comparative* measure of the uncertainties contained in two densities because they would both be affected by the transformation in the same way. In any case, the characterization of *relative* entropy, which we address in Section 5.1, has been found to circumvent the invariance problem. See the discussion in Caticha (2012), page 85. We shall now examine differential extropy in the style suggested by Shannon's argument.

D.2 Motivating the Differential Extropy Measure as $-\frac{1}{2} \int f^2(x) dx$

At first sight, the extropy measure $-\sum (1-p_i) \times \log(1-p_i)$ appears problematic: if each p_i were simply replaced by a density value f(x), the measure would not be defined when f(x) > 1, which it may. However, the situation clarifies by expanding $(1-p_i)\log(1-p_i)$ through three terms of its Maclaurin series with remainder: $(1-p_i)\log(1-p_i) = -p_i + \frac{p_i^2}{2} + \frac{p_i^3}{6(1-r_i)^2}$ for some $r_i \in (0, p_i)$. Summing these expansion terms over $i = 1, \dots, N$ shows that when the range of possibilities for X increases (as a result of larger N) in such a way that $\Delta x \to 0$ and $\max_{i=1}^{N} p_i$ decreases toward 0, the extropy measure becomes closely approximated by $1 - \frac{1}{2} \sum_{i=1}^{N} p_i^2$.

Following the same tack as for entropy in representing p_i by $f(x_i)\Delta x$ suggests that for large N the extropy measure can be approximated by

$$J(\mathbf{p}_N) \approx 1 - \frac{1}{2} \sum_{i=1}^{N} p_i^2 \quad \text{(when max } p_i \text{ is small)}$$
$$= 1 - \frac{1}{2} \sum_{i=1}^{N} f^2(x_i) (\triangle x)^2$$
$$= 1 - \frac{\triangle x}{2} \sum_{i=1}^{N} f^2(x_i) \triangle x.$$

This approximation is merely a location and scale transformation of $-\frac{1}{2}\sum f^2(x_i)\Delta x$. In the same spirit as for differential entropy, the measure of differential extropy for a continuous density can well be defined via the limit of $J(\mathbf{p}_N)$ as N increases in the same context as Definition D.1:

DEFINITION D.2. The differential extropy of the density $f(\cdot)$ is defined as

$$j(f) \equiv -\frac{1}{2} \int f^2(x) dx = \lim_{\Delta x \to 0} \{ [J(\mathbf{p}_N) - 1]/\Delta x \}.$$

The sum of the squares of probability masses (as well as the integral of the square of a density) has received attention for more than a century for a variety of reasons, but never in a direct relation to the entropy of a distribution. Rather, it has merely been considered to be an *alternative* measure of uncertainty. Good (1979) referred to this measure as the "repeat rate" of a distribution, developing an original idea of Turing. Gini (1912, 1939) had earlier proposed this measure as an "index of heterogeneity" of a discrete distribution, via $1 - \sum_{i=1}^{N} p_i^2$, deriving from the sum of the individual event variances, $p_i(1 - p_i)$. We now find that in a discrete context, a rescaling of Gini's index is an approximation to the extropy of a distribution when the maximum probability mass is small. In a continuous context, half the negative expected value of a density function value is the continuous differential analogue of the extropy measure of a distribution that we are proposing.

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REFERENCES

- AGRÒ, G., LAD, F. and SANFILIPPO, G. (2010). Sequentially forecasting economic indices using mixture linear combinations of EP distributions. *J. Data Sci.* **8** 101–126.
- BANERJEE, A., MERUGU, S., DHILLON, I. S. and GHOSH, J. (2005). Clustering with Bregman divergences. *J. Mach. Learn. Res.* **6** 1705–1749. MR2249870
- BASSEVILLE, M. (2013). Divergence measures for statistical data processing—An annotated bibliography. *Signal Process.* **93** 621–633.
- BERNARDO, J.-M. (1979). Expected information as expected utility. Ann. Statist. 7 686–690. MR0527503
- BIAZZO, V., GILIO, A. and SANFILIPPO, G. (2012). Coherent conditional previsions and proper scoring rules. In *Advances in Computational Intelligence*. CCIS 300 146–156. Springer, Heidelberg.
- BOISSONNAT, J.-D., NIELSEN, F. and NOCK, R. (2010). Bregman Voronoi diagrams. *Discrete Comput. Geom.* 44 281–307. MR2671013
- CAPOTORTI, A., REGOLI, G. and VATTARI, F. (2010). Correction of incoherent conditional probability assessments. *Internat. J. Approx. Reason.* **51** 718–727. MR2653348
- CATICHA, A. (2012). *Entropic Inference and the Foundations of Physics*. USP Press, Sao Paulo. Available at http://www.albany.edu/physics/ACaticha-EIFP-book.pdf.
- CENSOR, Y. and ZENIOS, S. A. (1997). Parallel Optimization: Theory, Algorithms, and Applications. Oxford Univ. Press, New York. MR1486040
- COVER, T. M. and THOMAS, J. A. (1991). *Elements of Information Theory*. Wiley, New York. MR1122806
- DEGROOT, M. H. (1984). Changes in utility as information. *Theory and Decision* 17 287–303. [See errata noted in 18 (1985) 319. MR0810605] MR0781339
- FRIGYIK, B. A., SRIVASTAVA, S. and GUPTA, M. R. (2008). Functional Bregman divergence. In *ISIT* (F. R. Kschischang and E.-H. Yang, eds.) 1681–1685. IEEE, Piscataway, NJ.
- FURUICHI, S. and MITROI, F.-C. (2012). Mathematical inequalities for some divergences. *Phys. A* **391** 388–400. MR2851038
- GILIO, A. and SANFILIPPO, G. (2011). Coherent conditional probabilities and proper scoring rules. In ISIPTA'11: Proceedings of the Seventh International Symposium on Imprecise Probability: Theories and Applications (F. Coolen, G. de Cooman, Th. Fetz and M. Oberguggenberger, eds.) 189–198. SIPTA, Innsbruck.
- GINI, C. (1912). Variabilità e mutabilità. Studi economicogiuridici della Facoltà di Giurisprudenza dell'Università di Cagliari, Anno III, parte 2^a.
- GINI, C. (1939). Memorie di Metodologia Statistica, Vol. I, Variabilità e Concentrazione. Giuffrè, Milano.

- GNEITING, T. and RAFTERY, A. E. (2007). Strictly proper scoring rules, prediction, and estimation. J. Amer. Statist. Assoc. 102 359–378. MR2345548
- GOOD, I. J. (1979). Studies in the history of probability and statistics. XXXVII. A. M. Turing's statistical work in World War II. Biometrika 66 393–396. MR0548210
- JAYNES, E. T. (2003). Probability Theory: The Logic of Science. Cambridge Univ. Press, Cambridge. MR1992316
- KOLMOGOROV, A. (1956). On the Shannon theory of information transmission in the case of continuous signals. *IRE Trans. Inf. Theory* **2** 102–108.
- KULLBACK, S. (1959). *Information Theory and Statistics*. Wiley, New York. MR0103557
- LAD, F. (1996). Operational Subjective Statistical Methods: A Mathematical, Philosophical, and Historical Introduction. Wiley, New York. MR1421323
- LAD, F., SANFILIPPO, G. and AGRÒ, G. (2012). Completing the logarithmic scoring rule for assessing probability distributions. In *XI Brazilian Meeting on Bayesian Statistics: EBEB* 2012. *AIP Conf. Proc.* **1490** 13–30. AIP Publishing, Melville, NY.
- PREDD, J. B., SEIRINGER, R., LIEB, E. H., OSHERSON, D. N., POOR, H. V. and KULKARNI, S. R. (2009). Probabilistic co-

- herence and proper scoring rules. *IEEE Trans. Inform. Theory* **55** 4786–4792. MR2597577
- RÉNYI, A. (1961). On measures of entropy and information. In Proc. 4th Berkeley Sympos. Math. Statist. and Prob., Vol. 1547– 561. Univ. California Press, Berkeley, CA. MR0132570
- SAVAGE, L. J. (1971). Elicitation of personal probabilities and expectations. J. Amer. Statist. Assoc. 66 783–801. Reprinted in The Writings of Leonard "Jimmie" Savage, American Statistical Association, Washington, DC, 1981, 565–614. MR0331571
- SHANNON, C. E. (1948). A mathematical theory of communication. Bell Syst. Tech. J. 27 379–423, 623–656. Reprint (1949) The Mathematical Theory of Communication (C. E. Shannon and W. Weaver, eds.), Univ. Illinois Press, Urbana, IL. MR0026286
- SHUFORD, E. H., ALBERT, A. and MASSENGILL, E. H. (1966). Admissible probability measurement procedures. *Psychometrika* **31** 125–145.
- STUMMER, W. and VAJDA, I. (2012). On Bregman distances and divergences of probability measures. *IEEE Trans. Inform. The*ory 58 1277–1288. MR2932808
- VONTOBEL, P. O. (2013). The Bethe permanent of a nonnegative matrix. *IEEE Trans. Inf. Theory* **59** 1866–1901. MR3030759