A Primer on the Conditional Mediation Analysis in PLS-SEM

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Abstract

Conditional mediation (CoMe) analysis integrates mediation and moderation analyses to examine and test hypotheses about how mediated relationships vary as a function of context, boundaries, or individual differences. Although CoMe analysis can be a crucial element of empirical studies that seek to advance theory in information systems, applications of such analysis are scarce, in general, and in partial least squares structural equation modeling (PLS-SEM), in paper particular. This clarifies conceptual fundamentals of and develops guidelines for CoMe analysis within the PLS-SEM context. Furthermore, the paper outlines the illustrative use of CoMe analysis in PLS-SEM and presents detailed step-by-step procedures to do so in the PLS-SEM setting. Overall, this paper provides researchers and practitioners with the required knowledge to properly carry out, report, and interpret CoMe analysis in PLS-SEM.

Keywords: PLS-SEM; Partial Least Squares; Structural Equation Modeling; Conditional Mediation Analysis; Conditional Mediation Model; Conditional Indirect Effect.

Introduction

Partial least squares structural equation modeling (PLS-SEM) has remained prevalent in information systems, business, and social science research (Hwang et al., 2020; Khan et al., 2019; Petter, 2018). Current advances in PLS-SEM enable a variety of nuanced analyses and encourage use of further robustness checks (Sarstedt et al., 2020b) and predictive assessments (Chin et al., 2020; Hair, 2020; Shmueli et al., 2019; Sharma et al. 2019, 2021) to better interpret PLS-SEM analysis. However, although mediation analysis (Nitzl et al., 2016; Sarstedt et al., 2020a) and moderation analysis (Becker et al., 2018) in PLS-SEM are frequently employed and well understood, the same does not apply to conditional mediation (CoMe) analysis that draws on both. Indeed, extant PLS-SEM applications of CoMe analysisoften referred to differently-vary in their execution and how they are interpreted. Hence, this paper seeks to provide guidance to proficiently conduct CoMe analysis in PLS-SEM.

Applications of CoMe analysis in a PLS-SEM context are scarce. Based on a search of papers published from 2000 to 2019, we could identify only 27 papers that in some way considered CoMe in PLS-SEM. Eleven of them appeared to disregard the use of a continuous variable when conducting CoMe analysis in PLS-SEM (see, e.g., Lee-Rodríguez et al., 2014; Ahadzadehet et al., 2018; Koay, 2018; Park et al., 2019). Instead, they employed a separate PROCESS

macro in their analysis. Meanwhile, 16 studies multigroup analysis (MGA) performed conducting CoMe analysis using a categorical variable. However, the majority of these studies (14 studies) did not examine the difference in the mediated effect before comparing each group's mediated effect values (see, e.g., Calvo-Mora et al., 2016; Fu et al., 2018; Willems et al., 2019; Limaj & Bernroider, 2019), except for two studies from Rezaei and Valaei (2017) and Hernandez-Ortega et al. (2017). Furthermore, many prior studies have remained vague in specifying theoretical reasoning for their whether the hypothesized CoMe relationship(s) concerns the rate of change (slope) or difference in strength (magnitude) of the effect(s). Clearly, the use of CoMe analysis in PLS-SEM is still in its infancy, which is of fundamental concern to researchers seeking to employ PLS-SEM when assessing theoretical arguments and deriving robust managerial recommendations that rest on relationships that are mediated but conditional on certain exogenous variables.

CoMe analysis can concern distinct types of CoMe models (see Hayes, 2018), which in turn require different analyses. However, some researchers appear to be not familiar with these requirements when conducting PLS-SEM analysis. Furthermore, many researchers still combine their PLS-SEM analyses with Hayes's (2017) PROCESS macros for SPSS or SAS. This is not surprising, because assessing whether a mediated relationship is moderated is relatively easy using a PROCESS macro. Recently, however, Sarstedt et al. (2020a) provided several reasons why this combination is not advisable. They stressed that simultaneous estimations in a nomological network, in contrast to conducting a supplementary regression analysis, are necessary to produce reliable and valid results, especially in the case of multi-item construct measurements. Notably, the sole use of PLS-SEM for estimating a CoMe model offers the following advantages: (1) it overcomes the limitations of traditional sequential approaches by enabling researchers to analyze complex latent variables interrelationships between simultaneously, (2) while accounting for measurement error inherent in the multi-item measurements (Edwards & Lambert, 2007; Hayes & Scharkow, 2013; Muller et al., 2005). Additionally, in PLS-SEM applications, computing the significance of the index of moderated mediation (the CoMe index) and evaluating it remain scarce (Hair et al., 2022). The CoMe index quantifies the effect that a moderator has on a mediated relationship (Hayes, 2015). If the CoMe index is significantly different from zero, a conditional mediated effect is smaller, larger, or opposite in sign at different levels of the moderator. One reason for the infrequent consideration of the CoMe index could relate to the lack of established and adopted

conceptual understanding about the nuances of CoMe analysis (i.e., distinguishing between distinct types of CoMe models), in general, and within the PLS-SEM context, in particular. Related, the issue of scarce and sometimes vague applications of CoMe analysis in PLS-SEM likely rests on lacking access to detailed step-by-step procedures for its conduct in the PLS-SEM setting. Summarizing these concerns, Sarstedt et al. (2020a) recently noted that clear guidelines for conducting CoMe analysis must be urgently provided; otherwise, its use in PLS-SEM will be hindered.

Overall, this study bridges this gap by offering researchers a guide for estimating and assessing CoMe models in PLS-SEM. This paper first clarifies CoMe analysis. In doing so, the paper illuminates conceptual fundamentals and develops guidelines for CoMe analysis within the PLS-SEM context. Subsequently, the paper outlines the illustrative use of CoMe analysis in and presents detailed step-by-step procedures for the PLS-SEM setting. Together, these conceptual guidelines and practical step-by-step procedures should enable a more proficient CoMe analysis in PLS-SEM applications. In turn, results from CoMe analysis using PLS-SEM in information systems studies should provide a better basis for drawing accurate conclusions to advance theory and produce practical insights.

What is Conditional Mediation Analysis?

A mediated relationship, also referred to as an indirect relationship, is defined by the presence of one or more variables that intervene to transmit the influence of, for example, variable X on variable Y (Baron & Kenny, 1986; Pirlott & MacKinnon, 2016; Zhao et al., 2010). In particular, in mediation analysis, researchers examine whether a change in an independent construct changes a mediator, which in turn results in changes in the dependent construct in the model (Demming et al., 2017). This analysis considers three relationships: a direct path between the independent and dependent variables, a first-stage path (representing the link between an independent variable and the mediator), and a second-stage path (signifying the relationship between the mediator and the dependent variable). In contrast, a moderating effect is defined by the presence of a variable that changes the influence of variable X on variable Y by impacting the nature, direction, or strength of this influence, which can vary under the value of the moderator (Aguinis & Gottfredson, 2010; Dawson, 2014). Although many quidelines for modeling mediating and moderating effects (see Aguinis et al., 2017; Gardner et al., 2017; Holland et al., 2017; Preacher & Hayes, 2008) and some directly concerning PLS-SEM (e.g., Becker et al., 2018; Carrión et al., 2017; Nitzl et al., 2016; Sarstedt et al., 2020a) have been widely adopted, explicit guidelines concerning the CoMe analysis in PLS-SEM do not exist.

As already indicated, CoMe analysis combines mediation and moderation analyses to examine and test hypotheses about how mediated relationships vary because of context, boundaries, or individual differences. It occurs when a moderator interacts with one or more of the paths of the mediated effect, such that the value of the mediated effect changes depending on the value of the moderator (Hayes, 2017, 2018). This type of modeling is well suited to investigate how relationships between cause and outcomes vary depending on the characteristics of their contexts (Bachl, 2017). For instance, "Does an effect operate through different relationships for certain individuals or under certain circumstances?" or "Is the mediated effect transmitted via some mediators stronger or weaker for some individuals or in some situations than others?" More formally, in a CoMe model, the independent variable of X influences the outcome variable of Y through one or more mediators of M, and the relationship is conditioned by the values of one or more moderators of W and/or Z (Hayes, 2017, 2018).

Scholars in various fields have encouraged researchers to go beyond the separate analysis of mediating and moderating effects as part of the quest to understand and predict theorized effects better (Edwards & Lambert, 2007; Hayes, 2018; Hayes & Scharkow, 2013). Two considerations explain why the integration of mediation and moderation can provide a more complete understanding of a phenomenon rather than focusing solely on either mediation or moderation alone as distinct processes (Karazsia et al., 2013). First, examining mediation is not always a universal process. Any bivariate or linear relationship of X on Y can be moderated by a third variable. Thus, researchers may be interested in understanding whether a mediated relationship is moderated by another variable such that the mediated effect value changes depending on the moderator's value (or the mediated effect value is conditional on the moderator's value). Second, the particular process(es) through which a moderating effect occurs may be of considerable interest. Thus, researchers may wish to examine whether a moderator conditions the firststage path or the second-stage path in the mediated relationship, or both. The literature has referred to this type of CoMe analysis also as conditional indirect effect analysis (Edwards & Lambert, 2007; Hayes, 2015; Preacher et al., 2007).

A CoMe effect represents an effect which is mediated through M, through which X exerts its influence on Y, and depends on the value of a moderator W (Hayes, 2017). In other words, a CoMe effect is present when the magnitude, size, or direction of the mediated effect

of variable X on variable Y via a mediator M varies according to the value of a moderator Z (Preacher et al., 2007). However, many previous studies that purported to test CoMe effects lacked precision in describing how analyses were conducted or draw on different terminology. Hence, the descriptions concerning CoMe analysis vary greatly in literature, making it difficult to fully comprehend what has been done or ought to be done.

In the earlier literature, Muller et al. (2005) discussed mediated moderation in which a single variable moderates the three paths in a simple mediation model. However, this view has been criticized for some of the same concerns that have been raised about the causal-step approach discussed by Baron and Kenny (1986). Subsequently, Morgan-Lopez and MacKinnon (2006) depicted a formal test for both moderated mediation and mediated moderation that researchers could use to establish that an interaction effect on some outcome is carried through a mediator. In their view, mediated moderation refers to the effect of a mediated interaction on a dependent variable, whereas moderated mediation occurs when the strength of the mediated effect depends on the level of a moderator.

Edwards and Lambert (2007) and Preacher et al. (2007) also referred to mediated moderation, but they went far beyond the single model delineated by Muller et al. (2005). Edwards and Lambert (2007) described eight models that could be established by allowing a single variable to moderate one or more causal paths in a mediation process and how the various direct and mediated effects can be assessed. Subsequently, they considered the term mediated moderation to be a specific case of moderated mediation. Nonetheless. this concept is subject to debate. Meanwhile, Preacher et al. (2007) introduced a formal definition of conditional mediation into the lexicon of statistical mediation analysis. They demonstrated how CoMe analysis is conducted to test hypotheses via the construction of standard errors and bootstrap confidence intervals (CIs). They also illustrated how moderated mediation could be conceptualized in a model with two moderators of different paths in the causal system.

Similarly, Hayes and Preacher (2013) and Hayes (2017) challenged the relevance of the very notion of mediated moderation by Edwards and Lambert (2007) and advised against using this term because the corresponding analysis provides no additional insights into the path model effects. Specifically, they argued that the interaction term in a mediated moderation model "has no substantive grounding in the measurement or manipulation process. The product does not quantify anything. And if ... [the interaction term] has no meaning and no substantive

interpretation, then what does the indirect effect of a product mean? The answer, in my opinion, is that it means nothing" (Hayes, 2017, p. 466). Furthermore, establishing theoretically reasonable underpinnings of mediated moderation models that can be empirically tested is often challenging. For these reasons, Hayes (2017) called to disregard the concept of mediated moderation. Instead, Hayes and Preacher (2013) and Hayes (2017) emphasize moderated (conditional) mediation. In their work, Hayes and Preacher (2013) took a covariance-based structural equation modeling approach in estimating conditional direct and mediated effects simultaneously to demonstrate the evidence of a CoMe effect. Since then, various methodological studies regarding CoMe analysis have appeared to address the means of conceptualizing and quantifying the contingencies of a mediated effect (Borau et al., 2015; Wang & Preacher, 2015).

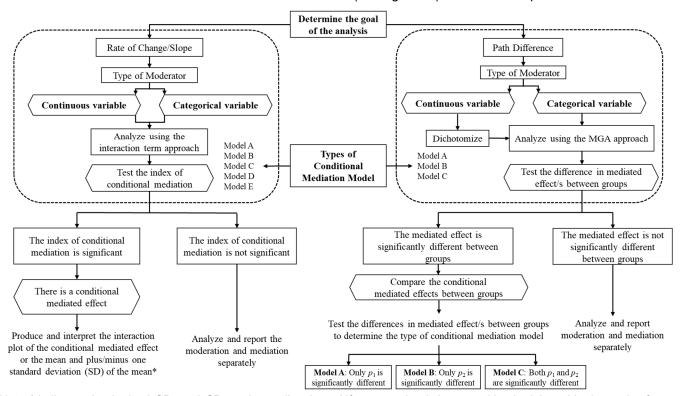
In this paper, we use to the term "CoMe model" when referring to the theoretical (research) model. The term "CoMe analysis" denotes the statistical assessment carried out when analyzing and estimating a CoMe model. And the term "CoMe effect" represents the value that is generated through CoMe analysis and quantifies the mediated effect conditional on a moderator.

Conceptual Fundamentals of and Guidelines for Conducting Conditional Mediation Analysis in PLS-SEM

To conduct CoMe analysis in PLS-SEM, this paper proposes a conceptual logic reflected in a decision tree (Figure 1). It captures conceptual fundamentals and encapsulates guidelines for CoMe analysis in PLS-SEM. Distinguishing between two goals that may guide the CoMe analysis, we outline two general paths of actions that ought to be followed.

The Analysis Goal

Commonly, researchers would be theory-driven, not data-driven, in their research design, especially when testing a hypothesized research model. This means that researchers would normally develop appropriate hypotheses that they seek to test drawing, in our case, on a CoMe model. Hair et al. (2022) stated that hypotheses are individual conjectures and are developed following the scientific method used to explain and predict outcomes. It allows researchers to follow a set of formal steps in conducting calculated analysis on their data. Theory-driven hypotheses can take on either of the following two forms when specifying a CoMe model, which concern either the rate of change (or slope) or difference in the strength (or magnitude) of a relationship:



Note: * indicates that both ±2 SD or ±3 SD can be easily adopted if a researcher is interested in obtaining wider intervals of SD.

Figure 1. Decision Tree of Conditional Mediation Analysis in PLS-SEM

- 1. The mediated relationship (M) between X and Y increases as W increases (*vice-versa*).
- 2. The mediated relationship (M) between X and Y is greater for high values of W (or Group A) than for low values of W (or Group B) (*vice-versa*).

Drawing from the above mentioned, any research design to test a CoMe model must be informed by underlying theory. Underlying theoretical reasoning has two crucial implications. First, it substantiates the distinct type of the CoMe model considered (please see latter sections for more details on distinct types of CoMe models). Second, the theoretical argument determines whether to assess either the rate of a slope's change or differences in the strength (i.e., magnitude) of a path. To answer this question, if researchers seek to assess the first type of hypothesis (i.e., rate of change/slope), they should assess and interpret the interaction term, whereas if researchers seek to assess the second type of hypothesis (i.e., differences in the strength/magnitude of a path), then they must compare the path values of two or more different groups. This means that the theoretical argument ought to drive the goal of the analysis; in turn, the choice of variable—regardless of continuous or categorical moderator—and the approach estimation (interaction term or MGA), which are explained in the latter parts of this paper. Importantly, when the analysis goal is pre-determined via a hypothesis, it should enable the researcher to specify and implement a suitable research design that, in turn, guides the collection of appropriate data to fit the required estimation approach to assess the hypothesis.

Type of Moderating Variable

Two types of moderating variables can be embedded in CoMe analysis: continuous and categorical moderating variables. When a moderator is continuous (e.g., switching costs, firm resources, or privacy concerns), researchers can generate an interaction term that expresses the joint influence of the exogenous construct and moderator on the endogenous construct (see Hair et al., 2022). This approach would suit estimating the rate of change (i.e., slope) of a relationship. Alternatively, when a moderator is categorical (e.g., industry, gender, or nationality), the variable may serve as a grouping variable that divides the data into subsamples. The theoretical model is then estimated for each distinct subsample. Researchers are usually interested in understanding any significant differences between subsamples; thus, the model estimates for the subsamples are usually compared using MGA (Sarstedt et al., 2011). The latter approach would enable assessments of differences in the strength (i.e., magnitude) of one or more paths.

Conditional Mediation Analysis with Continuous Moderators

An index term analysis should be conducted when the moderator is continuous and the goal of the CoMe analysis is to assess the rate of change (or slope) of a relationship. An "index term is known as the index of moderated mediation" that can be defined as a "direct quantification of the linear association between the indirect effect and the putative moderator of that effect" (Hayes, 2015, p. 3). As indicated in an earlier section of this paper, we refer to this index of moderated mediation as the CoMe index (ω). Researchers should be clear about the relationships on which the moderator exerts its influence when estimating this index term. In this vein, Figures 2-6 illustrate the five types of CoMe models (Models A-E) that allow a mediated relationship to be moderated by a continuous type of moderator. The four most popular forms of CoMe models are Model A (the first-stage conditional process model), Model B (the secondstage conditional process model), and both Models C and D (the first- and second-stage conditional process models); Model E is less common. 1 Notably, this present study does not illustrate the moderating effect on the mediated relationship between X and Y (also known as conditional direct effect)2 because even though a moderating effect may exist on the direct relationship of the CoMe model, the focus of the statistical estimation is only based on X and M on Y.

In CoMe Model A (see Hayes, 2017, 2018), moderator W is assumed to affect the first stage of the mediation process (Figure 2). W moderates the effect of X on M. but the effect of M on Y is independent of W. To substantiate that the mediation is moderated. researchers traditionally note the need for evidence of a significant moderation of the path linking X and M (Muller et al., 2005; Preacher et al., 2007), which is equivalent to testing the significance of path (p_5) in Figure 2. However, Hayes and Rockwood (2020) have emphasized that a nonsignificant moderation result does not necessarily imply that W does not moderate the mediated effect of X on Y. Additionally, researchers must consider the impact of the moderator on the mediated effect as a whole rather than on a single element of the mediated effect in isolation (in this case, path p_2) (Hayes, 2018). To formally test this effect in CoMe Model A, Hayes (2018) proposed the CoMe effect, which is defined as follows:

$$\omega = p_1 \cdot p_2 + p_2 \cdot p_5 \cdot W_{(1)}$$

Based on equation 1, the path coefficients of " $p_2 \cdot p_5$ " refer to the interrelationships of M between the constructs of X and Y (Figure 2), whereas W refers to the moderator. If the index " $p_2 \cdot p_5$ " is significantly different from zero, the mediated effect of X on Y through M depends on moderator W. In other words,

" $p_2 \cdot p_5$ " is the slope of the mediated effect of W contingent upon ω, which represents the CoMe index (Hayes, 2015; Morgan-Lopez & MacKinnon, 2006). The index is a product of path coefficients; hence, a bootstrap-based percentile confidence interval should be used for inference without any assumption about the shape of its sampling distribution. ³ The CoMe index (ω) can be easily obtained based on the path coefficient output generated by any software that enables running of PLS-SEM algorithms (e.g., R, SmartPLS, WarpPLS, ADANCO, etc.).

The second way to combine mediation and moderation is through the second-stage conditional process model (Figure 3). In *CoMe Model B*, moderator W affects only the second stage of the mediation process, in which the relationship between the mediator M and endogenous variable Y is assumed to be moderated by W. In other words, this second-stage conditional process model fixes the effect of X on M to be constant but allows the effect of M on Y to be moderated by W. To assess the occurrence of this second-stage conditional process model, Hayes (2018) defined the CoMe effect as follows:

Conceptual Model Model A Model A W M M M Y Y Y Y Y Statistical Model X·W W P₄ P₅ Y Y

Figure 2. CoMe Model A

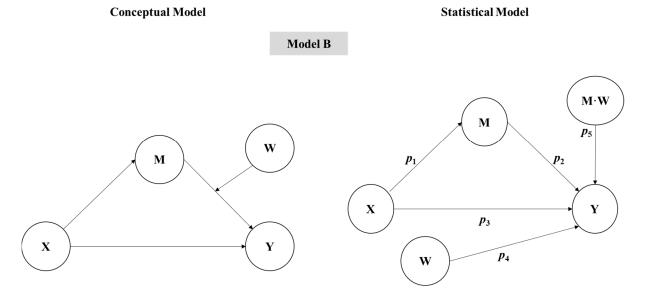


Figure 3. CoMe Model B

$$\omega = p_1 \cdot p_2 + p_1 \cdot p_5 \cdot W_{(2)}$$

Based on equation 2, if the index " $p_1 \cdot p_5$ " is significant, the mediated effect through M is significantly moderated by W. The index also quantifies the effect of W on the mediated effect of X on Y through M. However, at present, such an output of the CoMe index (ω) cannot yet be directly obtained from any type of PLS-SEM software (R, SmartPLS, WarpPLS, ADANCO, etc.). Hence, the path coefficient results must be examined using the bootstrap sample output. For example, a researcher can copy the bootstrap results into an Excel spreadsheet and multiply all lines (e.g., 5,000 or 10,000 subsample bootstraps) of p₁ and p₅, to calculate the bootstrap CIs from the data in this new calculated column. Alternatively, using other software packages such as the R software is an efficient way for generating this result (i.e., $p_1 \cdot p_5$ and the bootstrap CIs) (R Core Team, 2021). The third type of CoMe model, shown in CoMe Model C, is the combination of the first- and second-stage conditional process models, in which W simultaneously moderates the effects of both X on M and M on Y (Figure 4). The proposed conditional effect is specified like any mediation model $(p_1 \cdot p_2)$ but includes the product for testing the moderating effect of both X·W on M and M·W on Y. Hayes (2018) defined the following conditional effect equation for CoMe Model C:

$$\omega = p_1 \cdot p_2 + (p_2 \cdot p_5 + p_1 \cdot p_7)W + p_5 \cdot p_7W^2$$
 (3)

In this equation, " $p_2 \cdot p_5 + p_1 \cdot p_7$ " is the CoMe index (ω) for *CoMe Model C*. When this sum of the products is significant, the CoMe effect occurs; that is, the mediated effect of X on Y through M is moderated by W.

The fourth type of CoMe model, shown in *CoMe Model D* (Figure 5), is similar to *CoMe Model C*. The only difference is that *CoMe Model D* has two distinct moderators influencing the mediated effect. Particularly, W moderates the effect of X on M, and Z moderates the effect of M on Y. Additionally, the proposed conditional effect is specified like any mediation model ($p_1 \cdot p_2$) but includes both the product of X·W and M·Z in the models of the effect of X on Y through M (Hayes, 2018). This type of CoMe effect equation is defined as follows:

$$\omega = p_1 \cdot p_2 + (p_2 \cdot p_5)W + (p_1 \cdot p_7)Z + p_5 \cdot p_7WZ_{(4)}$$

In this equation, " $p_2 \cdot p_5 + p_1 \cdot p_7$ " is the CoMe index (ω) for CoMe Model D. When this sum of products is significant, the CoMe effect occurs; that is, the mediated effect of X on Y through M is moderated by W and Z respectively.

Finally, Figure 6 presents the fifth type of the CoMe model. In *CoMe Model E*, X moderates its own mediating effect on Y through M via the moderation of the effect of M on Y by X. Here, the independent variable is the moderator model, and the mediated effect of X on Y through M is the product of the effect of X on M. Hayes (2017) defined the following conditional effect equation for *CoMe Model E*:

$$\omega = p_1 \cdot p_2 + p_1 \cdot p_4 \cdot X_{(5)}$$

In this equation, " $p_1 \cdot p_4$ " is the CoMe index (ω) for CoMe Model E (Hayes, 2015). It quantifies the relationship between X and the mediated effect of X on Y through M. X functions as a moderator of its own mediated effect.

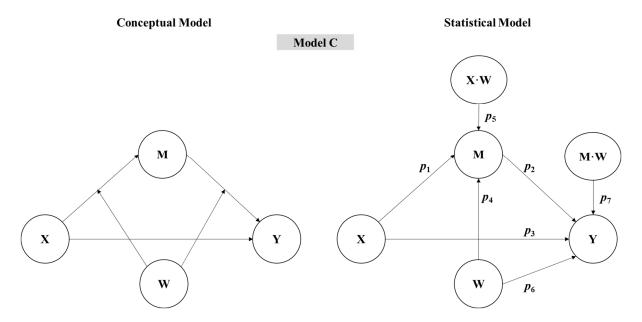


Figure 4. CoMe Model C

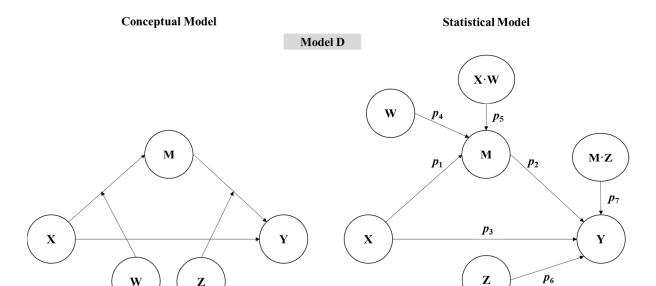


Figure 5. CoMe Model D

Conceptual Model Statistical Model

Model E

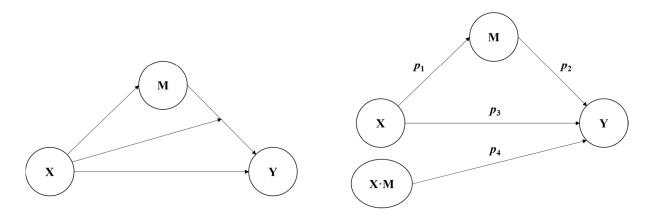


Figure 6. CoMe Model E

Conditional Mediation Analysis with (Multi)Categorical Moderators

When W is a categorical (or dichotomous) moderator (possibly transformed from a continuous variable^{4,5}), we can conduct an interaction analysis for a CoMe model if the research goal is to study the rate of change (or slope) of a relationship. These variables are usually dummy coded (i.e., 0 and 1), whereby 0 represents the reference category (Henseler & Fassott, 2010; Wooldridge, 2015). For instance, suppose W is gender, then the researcher could use a variable with 0 for "Male" and 1 for "Female" (vice-versa). Meanwhile, when W is a dichotomous variable

represented in the data with values that differ by one unit (e.g., W = 0 and W = 1), the CoMe index (ω) is the difference between the mediated effect of X on Y in the two groups. In referring to Equation 1 (the same applies similarly for equations 2–5), when W acts as a categorical moderator in the CoMe model, the indexes of " $p_2 \cdot p_5$ " will be tested as usual; that is, the two values of W chosen to code the two groups can be used in place of ω .

However, in many studies, W is multi-categorical, for example, (1) a control group versus two or more experimental conditions and (2) the positions on W could be three values on a continuous variable that are

operationalized "high," "medium," and "low" on that variable. In such a scenario, the researcher should rely on the fact that CoMe effect differences can be estimated with a linear model by representing groups with a set of k-1 variable, where k is the number of groups that should be transformed into as many dummy variables as there are levels (Henseler & Fassott, 2010; Wooldridge, 2015). For example, in the case of three groups suggested by Hair et al. (2022) (e.g., for a service-type variable: short-term contract, long-term contract, and prepaid), researchers can split the moderator into two dummy variables, which are simultaneously included in the model. Particularly, both dummy variables would take the value of 0 for the reference category (e.g., prepaid). The other two categories would be indicated by the value 1 in the corresponding dummy variable: dummy 1 (prepaid: 0 vs. short-term contract: 1) and dummy 2 (prepaid: 0 vs. long-term contract: 1).

Generating and Testing Interaction Terms in Conditional Mediation Analysis

To generate the interaction term, Becker et al. (2018) and Fassott et al. (2016) highlighted three potential procedures, namely, the product-indicator approach (Chin et al., 2003), the orthogonalizing approach (Lance, 1988), and the two-stage approach (Henseler & Fassott, 2010). The latter approach is more commonly considered as the more suitable one (see Hair et al., 2022). When interpreting the results of a moderation analysis, scholars are primarily interested in the significance of the interaction term. Suppose that the effect of the interaction term significantly moderates at least one path in the causal process linking X to Y via M. When the remaining unmoderated path is also significantly different from zero, the presence of a CoMe effect could exist (Muller et al., 2005). However, PLS-SEM researchers should conduct a formal test to assess the CoMe index (ω) because a mediated effect can be moderated even if one cannot show significant moderation of its causal process linking X to Y via M.

Importantly, before estimating these several types of CoMe models (Models A–E), approaches that gauge measurement error must be considered (Hair & Sarstedt, 2019). Doing so is important to monitor the quality of the variables' measurement models, in consideration of either reflective (Mode A estimation) or formative (Mode B estimation) measurement modes. ⁶ Particularly, if a reflective measurement mode is used, researchers must ensure that the indicator loading has a value of at least 0.708 with an associated significant result (Hair et al., 2022). Then,

the measurements must be considered reliable (i.e., indicator reliability, Cronbach's alpha, and composite reliability) and valid (i.e., convergent validity and discriminant validity) to detect a CoMe effect (see Hair et al., 2020). Alternatively, if we have a formative measurement mode, convergent validity (Cheah et al., 2018), variance inflation factor (Becker et al., 2015; Mason & Perreault Jr, 1991), and the statistical significance and relevance of the indicator weights must be monitored (Hair et al., 2022; Hair et al., 2019).

Once the properties of the measurement models for the constructs embedded in the CoMe model are validated (irrespective of whether reflective or formative), the researcher can then start to examine the CoMe index (ω) using the bootstrapping technique. Particularly, to establish the presence of the CoMe index (ω), the researcher must observe 90 percent or 95 percent CI of index ω . If the CoMe index (ω) is significantly different from zero (or statistically significant), researchers can conclude that the mediated effect of X on Y through M depends on W. This allows establishing a CoMe effect. Vice-versa, if the CoMe index (ω) is not statistically significant, the CoMe effect is not verified. Even though a mediated effect may exist, it does not depend on the level of the moderator. The lack of significance of the CoMe effects for low, medium, and high levels of the moderator means that the moderator cannot produce a differential effect. A higher level of the moderator leads to a similar effect as the lower level of the moderator. Therefore, PLS-SEM researchers should subsequently analyze the mediation and moderation separately because the results might be individually significant in explaining a mediated or interaction effect.

Conditional Mediation Analysis at Different Levels of the Moderator

Once the existence of a CoMe effect is supported by the CoMe index (ω), one may wish to investigate the mediated effect at representative levels (or values) of the moderator (depicted as the CoMe effect) to explore further the conditions under which mediation exists (Preacher et al., 2007). Specifically, researchers must determine how the moderator at different levels (or values) influences the mediated effect, which may be stronger, weaker, or opposite in sign. For a continuous moderator, by default, the CoMe effects are tested based on the mean of the moderator and at values of one or two standard deviations (SDs) below (or low), average (or medium), and above (or high) the mean using the following equations in Table 1.

Table 1. CoMe Effect Testing at Different Levels of the Moderator

Model	Equation
Α	• CoMe effect (low) = $[(p_1 + (p_5 - SD \text{ of W})) \cdot p_2]_{(6)}$
	• CoMe effect (medium) = $[(p_1 + (p_5 \cdot SD \text{ of } 0 \text{ for } W)) \cdot p_2]_{(7)}$
	• CoMe effect (high) = $[(p_1 + (p_5 \cdot + SD \text{ of W})) \cdot p_2]_{(8)}$
В	• CoMe effect (low) = $[p_1 \cdot (p_5 \cdot - SD \text{ of W}) + p_2)]_{(9)}$
	• CoMe effect (medium) = $[p_1 \cdot (p_5 \cdot SD \text{ of 0 for W}) + p_2)]_{(10)}$
	• CoMe effect (high) = $[p_1 \cdot (p_5 \cdot + SD \text{ of W}) + p_2)]_{(11)}$
С	• CoMe effect (low) = $[(p_1 + (p_5 - SD \text{ of W})) \cdot ((p_7 - SD \text{ of W}) + p_2)]_{(12)}$
	• CoMe effect (medium) = $[(p_1 + (p_5 \cdot SD \text{ of 0 for W})) \cdot ((p_7 \cdot SD \text{ of 0 for W}) + p_2)]_{(13)}$
	• CoMe effect (high) = $[(p_1 + (p_5 + SD \text{ of W})) \cdot ((p_7 + SD \text{ of W}) + p_2)]_{(14)}$
D	• CoMe effect (low) = $[(p_1 + (p_5 - SD \text{ of W})) \cdot ((p_7 - SD \text{ of Z}) + p_2)]_{(15)}$
	• CoMe effect (medium) = $[(p_1 + (p_5 \cdot SD \text{ of 0 for W})) \cdot ((p_7 \cdot SD \text{ of 0 for Z}) + p_2)]_{(16)}$
	• CoMe effect (high) = $[(p_1 + (p_5 + SD \text{ of W})) \cdot ((p_7 + SD \text{ of Z}) + p_2)]_{(17)}$
Е	• CoMe effect (low) = $[p_1 \cdot (p_4 \cdot - SD \text{ of } X) + p_2)]_{(18)}$
	• CoMe effect (medium) = $[p_1 \cdot (p_4 \cdot SD \text{ of 0 for X}) + p_2)]_{(19)}$
	• CoMe effect (high) = $[p_1 \cdot (p_4 \cdot + SD \text{ of } X) + p_2)]_{(20)}$

By simplifying the algebraic expressions, we can derive the equations in Table 2 from equations 1-5 (see Hayes, 2017; Stride et al., 2015). To clarify these equations, we draw on explaining this for CoMe Model A. 8 In the initial step, the researcher must run bootstrapping for each path coefficient to obtain the bootstrap subsample result (i.e., 5,000 or 10,000 subsamples) (Hair et al., 2022; Streukens & Leroi-Werelds, 2016a). Then, the researcher can take the path coefficient from the interaction term to the mediator (p₅) and multiply it by the SD values of W (the moderator), which represent low (mean-1 SD of the moderator), medium (mean of the moderator), and high (mean + 1 SD of the moderator) levels. 9 Next, the researcher must take these three values and multiply them by the path coefficient from the independent variable to the mediator (p_1) . Finally, the product of these values will be multiplied by the path coefficient from the mediator to the dependent variable (p_2) to obtain CoMe effects for low, medium, and high levels of the moderator.

Conversely, if W is a (multi)categorical moderator, the researcher can follow the procedures by taking the path coefficient from the interaction term to the mediator (p_5) and multiplying it by the SD values of W (similar to Models B–E). The only difference is that the researcher cannot treat it as CoMe effects for low, medium, and high levels of the moderator; rather, the dummy is compared with the reference category.

Finally, the researcher can create a graphical plot as part of visualization the data (Spiller et al., 2013) to corroborate the findings of the CoMe effects, that is, to investigate the significance of each condition (CoMe effects for low, medium, and high levels of the moderator) at bounds of the CoMe index (ω) (Preacher et al., 2007). Particularly, the researcher can use the simple slope or Johnson-Neyman's plot suggested by Preacher et al. (2007) to apply the regions of significance approach or identify the ranges of the moderator for which a mediated effect is statistically significant.

Conditional Mediation Analysis Using MGA

In most cases, where the moderator (W) is categorical and the goal is to understand the differences in the strength (magnitude) or directionality of mediated effects, an MGA can be used to understand the CoMe effect more comprehensively. Notably, the MGA can also be applied drawing on a continuous moderator if the aim is to test the differences in the strength (magnitude) or directionality of mediated effects. In this case, the continuous variable must be dichotomized, however. Once the observations are grouped, models are estimated separately for each group of observations. The differences in the mediated effects between the data groups are interpreted as CoMe effects. Figure 7 depicts the CoMe effect based on comparing effects for two groups, in which the impact of W can be estimated using a nominal variable such as Group A versus Group B or high values versus low values of a particular moderator. Researchers are usually interested in analyzing group effects related to mediated relationships. More precisely, a population parameter p1·p2 is hypothesized to be different (d) across two subpopulations (i.e., $p_1^{(1)} \cdot p_2^{(1)}$ and $p_1^{(2)} \cdot p_2^{(2)}$), which are expressed by different modalities in the categorical moderating variable.

Additionally, an MGA enables the researcher to test for differences between identical models estimated for different groups. According to Shrout and Bolger (2002), a group comparison can be useful for analyzing mediated effects. For instance, a CoMe

effect may occur when the mediated effect of one group is significantly stronger than or even different from that of another group. Such groups can, for example, relate to observable characteristics (e.g., gender) or refer to experimental conditions (e.g., high value of W versus low value of W). For instance, a researcher may assume that X influences Y indirectly through M, but the effect is stronger for men (or the high value of W) than for women (or the low value of W); alternatively, the effect is perhaps present only for men or women (or only for the high or low value of W) (Hair et al., 2018). Figure 8 shows that full mediation exists in Group A: Male (or the high value of W), whereas it does not exist in Group B: Female (or the low value of W).

To test group effects in PLS-SEM, researchers must conduct a measurement invariance test to determine whether construct measurements are similarly suitable across groups or heterogeneous data (Henseler et al., 2016). 10 In other words, this test concerns the statistical property of an instrument, indicating that the same construct(s) are measured in the same way across subgroups of respondents. The reason for such differences may stem from (1) respondents embracing different cultural values who interpret a given measure in a conceptually different manner or gender, ethnic, or other individual differences that entail responding to instruments in systematically different ways, and (2) respondents who use the available options on a scale differently (or tendency to choose or not choose the extremes) (Hair et al., 2018; Vandenberg & Lance, 2000).

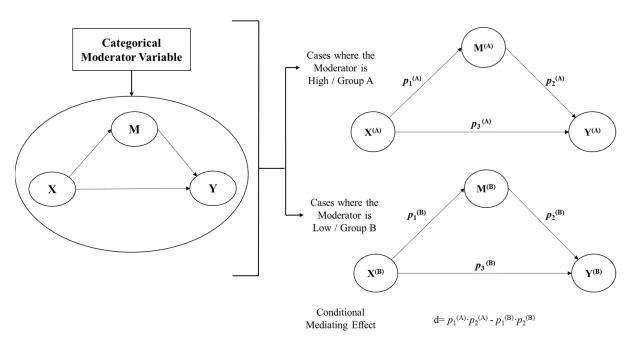
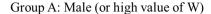
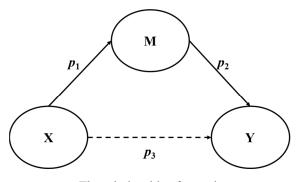


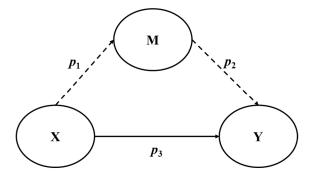
Figure 7. Detecting a Conditional Mediated Effect (d) through Group Comparisons





The relationship of $p_1 \cdot p_2$ is significant but p_3 is insignificant

Group B: Female (or low value of W)



The relationship of $p_1 \cdot p_2$ is nonsignificant but p_3 is significant

Figure 8. Full Mediation in Testing Group A (or the high value of W) and No Mediation in Testing Group B (or the low value of W)

Once this similarity is established, a group comparison can be estimated via five methods of MGA in PLS-SEM: the permutation test (Chin & Dibbern, 2010), Henseler's bootstrap-based MGA (Henseler et al., 2009), the non-parametric distance-based test (Klesel et al., 2019), the parametric test (Keil et al., 2000), or the Welch-Satterthwaite test (Welch, 1947). Drawing on these available MGA methods allows comparing a model across different groups to determine whether the strength (i.e., magnitude) of one or more relationships in the model is different between groupbased estimations and whether any differences are statistically significant. If the test does not support a significant difference in the mediated effect between groups, the researcher should reject the existence of a CoMe effect. In other words, if the CoMe effect is not significant in the MGA estimation, the PLS-SEM researcher should subsequently analyze mediation and moderation separately.

Before performing an MGA, the researcher should ensure that measurement invariance is achieved drawing on the measurement invariance of composite models (MICOM) procedure (see Henseler et al., 2016). By establishing MICOM, the researcher can be assured that the group differences in the model estimates do not result from distinct content and meanings of measurements used for the latent variables across groups. Notably, conducting an MGA requires partial measurement invariance, which is given when configural and compositional invariance hold. If partial measurement invariance is confirmed and the composites have equal mean values that vary across the groups, full measurement invariance is confirmed, which supports the pooled data analysis. In contrast, if partial measurement invariance is not present, this absence can deflate the power of statistical testing, influence the accuracy of estimators, and provide misleading results in further analyses.

In the initial step, a significant difference must exist in the mediated effect for the groups by subtracting the mediated effect for Group A from that for Group B (i.e., $d = p_1^{(A)} \cdot p_2^{(A)} - p_1^{(B)} \cdot p_2^{(B)}$). This procedure can be performed using the MGA estimation approach to determine whether the mediated effect is significantly different, noting that the CoMe effect is present (Ryu & Cheong, 2017). A p-value smaller than 0.05 suggests that the difference d between the group-specific mediated effect is present. Thus the researcher can continue with comparing the mediated effects for each of the two groups (comparing the β values of the mediated effect for Group A and Group B).

The comparison of the mediated effect for the two groups can be generated using any of the MGA techniques that we mentioned earlier (i.e., the permutation test, the non-parametric distance-based test, etc.). For example, if the mediated effect for the group with male individuals is large and significant (CI 90 percent of the mediated effects exclude zero), whereas it is small and significant (CI 90 percent of the mediated effects exclude zero) for the group with female individuals, one would then conclude that the mediated effect is significantly stronger (i.e., the magnitude of the relationship is greater) in the male group than in the female group. Alternatively, if the mediated effect for the group with male individuals is significant (CI 90 percent of the mediated effects exclude zero), but it is not significant for the group with female individuals, one would infer that the mediated effect only occurs in the male group but not in the female group. In other words, this approach allows the researcher to compare group differences in terms of the obtained significant mediated effect values. Finally, PLS-SEM researchers can compare the type of mediated effect such as full mediation versus partial mediation (complementary or competitive) (see Nitzl et al., 2016).

Assessment of Differences in Direct Effects to Determine the Type of Conditional Mediation Model

CoMe analysis that draws on the interaction term approach (as per Sections 3.1-3.4) to examine hypotheses that concern the rate of change (i.e., slope) is based on an a priori specified type of CoMe model. However, the type of CoMe model must be determined a posteriori when conducting MGA to assess hypotheses that concern differences in the strength (i.e., magnitude) of relationships. Hence and as per Figure 1, it is necessary to identify whether a CoMe Model is of the A, B, or C type, when carrying out MGA, in which the researcher analyzes whether $p_1 \cdot p_2$ in Group 1 is different from that in Group 2. Based on the results and the presence of a significant difference in the two groups' mediated effects, the researcher can now observe whether p_1 is also significantly different between groups 1 and 2; thus, the researcher will find a *Model A* type CoMe. If only p_2 is significantly different, the researcher will find a *Model B* type CoMe. Lastly, if both p_1 and p_2 are significantly different, the result will indicate that the researcher has a *Model C* type CoMe, where a particular categorical variable moderates both paths.

Empirical Illustrations of Conditional Mediation Analyses in PLS-SEM

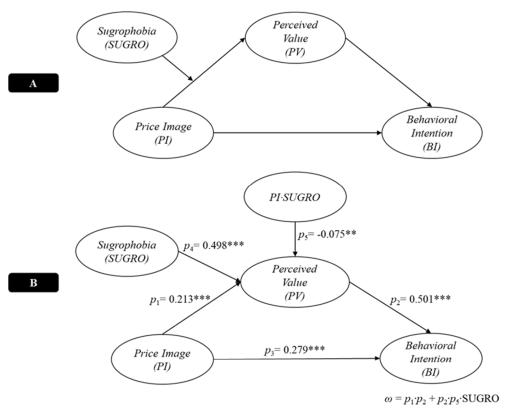
In the following we describe how the conceptual fundamentals and guidelines embedded in Figure 1, and described above, are empirically applied in PLS-SEM. In illustrating their application, we also refer to detailed step-by-step procedures for implementation in a PLS-SEM setting and outline these procedures in corresponding appendices to this paper.

The illustrative example of a CoMe model was extracted from the research model of Cheah et al. (2020b). The data were collected from 400 consumers who had experienced purchasing luxury products from a luxury retail store. The four main variables were modeled as composites: behavioral intention (BI), price image (PI), perceived value (PV), and sugrophobia (SUGRO). 12 BI was established as a composite estimated in Mode B, whereas PI, PV, and SUGRO were designed as a composite estimated in Mode A. Although their measurement models and their respective measurement modes were theoretically substantiated, to assess potential for misspecification in measurement mode, confirmatory tetrad analysis (CTA-PLS) (Gudergan et al. 2008) was used (see Appendix A). The CTA-PLS assessment for PI. PV. and SUGRO revealed that all tetrads vanished. providing support for the employed reflective measurement mode. Conversely, the CTA-PLS

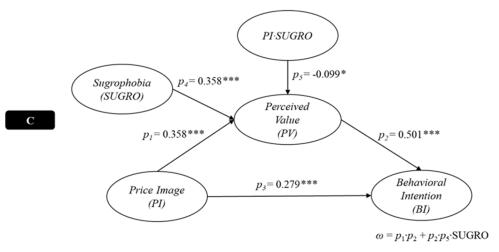
assessment showed that several tetrads failed to vanish for BI, implying that its measurement mode is indeed not reflective but rather formative. Importantly, the evaluation of the measurement models by using standard evaluation criteria (Hair et al., 2022) supports the measures' reliability and validity. This also holds for discriminant validity assessment using Henseler et al.'s (2015) heterotrait—monotrait (HTMT) criterion. Moreover, prior study has analyzed how PV positively mediates the relationship between PI and BI. In addition, SUGRO is posited to negatively moderate the link between PI and PV (see Cheah et al., 2020).

Based on the abovementioned, we extend past findings by illustrating the CoMe analysis for three types of CoMe models (Models A, B, and C) and for two types of moderators, either using a continuous moderator or a categorical moderator that has been transformed from a continuous one. 13 In regard to our empirical illustrations that draw on the transformation of a continuous moderator to a categorical moderator for CoMe analysis, we reiterate that researchers ought to be cautious in doing so because this transformation can produce a reduction in prediction power. Therefore, unless the researcher seeks to present findings in a particular way that serve a particular purpose (e.g., presentation of findings to a managerial audience to ease comprehension of actionable insights), or encounters data characterized by irregular or nonnormal distributions or outliers (DeCoster et al., 2009), estimations drawing on such transformed data ought to be avoided. Thus, the illustrations of the CoMe analysis for the three types of CoMe models (Models A, B, and C) with the transformation of the moderator (SUGRO) to a categorical variable serve merely as empirical illustrations on how to establish both the CoMe index (ω) and CoMe effects (see Section 4.1)¹⁴ when explicitly desired by the researcher.

In drawing on the abovementioned research model of Cheah et al. (2020b), we first illustrate the CoMe analysis in a PLS-SEM setting to assess the rate of change (i.e., slope) in the mediated relationship(s) and focus on the following hypothesis: The indirect effect of the consumer's PI on his or her BI via PV decreases as SUGRO increases. Therefore, in Section 4.1, we illustrate CoMe analysis drawing on the interaction term approach to assess the conditional mediated relationship(s). Then, we demonstrate the CoMe analysis in a PLS-SEM setting to examine differences in strength (i.e., magnitude) in the conditional mediated relationship(s) and consider the following hypothesis: The indirect effect between PI and BI via PV is more pronounced for the group of female consumers than for the group of male consumers. In doing so, we outline in Section 4.2 CoMe analysis employing the MGA approach.



Note: SUGRO is treated as a continuous variable. *** p < 0.001, ** p < 0.01 (based on t(9999), one-tailed test)



Note: SUGRO is treated as a categorical variable. *** p < 0.001, ** p < 0.01 (based on t(9999), one-tailed test)

Figure 9. Conceptual Model (Panel A), Structural Model Results with SUGRO as Continuous Variable (Panel B), and Structural Model Results with SUGRO as Categorical Variable (Panel C) for CoMe Model A

Conditional Mediation Analysis-Interaction Term Approach

Conditional Mediation Analysis Model A– Continuous and Categorical Moderator

For CoMe Model A, ¹⁵ we focus on testing the influence of the moderator (SUGRO) on the mediated effect ($p_1 \cdot p_2$) (Figure 9). We first assess the CoMe index (ω),

which is derived from the following equation: $\omega = p_1 \cdot p_2 + p_2 \cdot p_5 \cdot \text{SUGRO}$ (Equation 1). Based on the estimation (Appendix B(i); see Figure B1, Panels A and B), the result shows that " $p_2 \cdot p_5$ " (or PI*SUGRO -> PV -> BI) for the use of the continuous moderator is supported, as the CI excludes zero (90 percent CI: -0.065 to -0.010). Similarly, the estimation (Appendix B(ii); see Figure B7, Panels A and B) also shows that the result of " $p_2 \cdot p_5$ " (or PI*SUGRO -> PV -> BI) for the

use of the categorical moderator is supported, as the CI excludes zero (90 percent CI: -0.093 to -0.004). Since the index " $p_2 \cdot p_5$ " of SUGRO is significant for both the continuous ($\beta = -0.037$) and categorical ($\beta = -0.049$) variable uses, we can conclude that the mediated effect of PI on BI through PV negatively depends on the moderator SUGRO (Table 2).

Then, we investigate the mediated effects at different levels of the moderator (Preacher et al., 2007) for both continuous and categorical moderators. Following the procedures derived from our guidelines (see Appendix B(i)), we initially assess the SD of the moderator (SUGRO) when using the continuous moderator, as it will denote that the moderator is one SD below and one SD above the mean. 16 To estimate the CoMe effects for low, medium, and high levels of the moderator, we use equations 6-8 for CoMe Model A (see Appendix B(i), Figures B2-B5 for the details of the analysis). Results in Table 2 show that the path coefficients for the CoMe effects become weaker when the SUGRO effect is one SD above the mean. Additionally, the CoMe effects for low, medium, and high levels of the moderator are significant because both extremes of the CI are positive. For the use of the categorical moderator, we illustrate this for the comparison between the low value (0) and the high value (1) (see Appendix B(ii), Figures B8-B11 for the details of the analysis). Moreover, the results in Table 2 show that the path coefficient for the CoMe effect for high SUGRO (β = 0.130) is weaker than for low SUGRO (β = 0.178), and both CoMe effects are

significant because both the extremes of the CI are positive. Therefore, we find some evidence supporting the hypothesis in *CoMe Model A* for both the categorical and continuous moderator use. That is, the indirect effect of PI on BI via PV decreases due to the negative impact of SUGRO (to corroborate this finding, also see Figure 10 that depicts the results using the Johnson-Neyman plot and simple slope for both the continuous and categorical moderators).¹⁷

Conditional Mediation Analysis Model B-Continuous and Categorical Moderator

For CoMe Model B, we focus on testing the influence of the moderator (SUGRO) on the mediated effect (p1 $\cdot p_2$) (Figure 11). We first assess the CoMe index (ω). which is derived from the following equation: $\omega = p_1 \cdot p_2$ + $p_1 \cdot p_5$ SUGRO (Equation 2). Based on our estimations (see Appendix C(i); Figure C4), the results show that " $p_1 \cdot p_5$ " for the use of the continuous moderator is not supported, as the CI does include zero (90 percent CI: -0.044 to 0.002). However, the estimations (see Appendix C(ii); Figure C11) show that the result of " $p_1 \cdot p_5$ " for the use of the categorical moderator is supported, as the CI does not include zero (90 percent CI: -0.117 to -0.040). The index " p_1 · p₅" of SUGRO is significant for the categorical moderator ($\beta = -0.078$); therefore, we can conclude that the mediated effect of PI on BI through PV decreases as the moderator SUGRO increases, but not when estimating with the continuous moderator (β = -0.021) (see Table 2).¹⁸

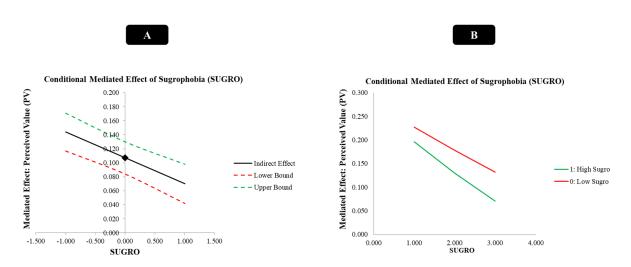


Figure 10. Continuous Variable (Panel A) and Categorical Variable (Panel B) for CoMe Effects Plot for CoMe Model A

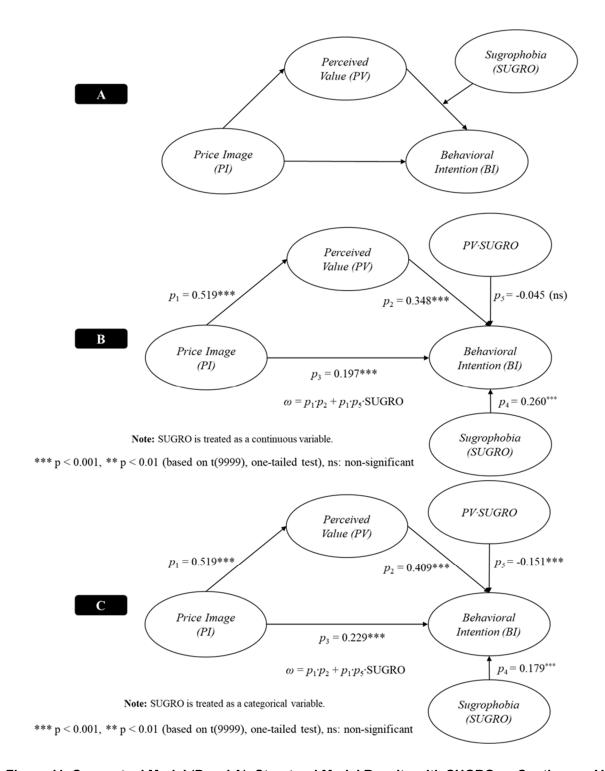


Figure 11. Conceptual Model (Panel A), Structural Model Results with SUGRO as Continuous Variable (Panel B), and Structural Model Results with SUGRO as Categorical Variable (Panel C) for CoMe Model B

Then, we investigate the mediated effects at different levels of the moderator (Preacher et al., 2007) for both the use of continuous and categorical moderators. Based on the procedures derived from our guidelines (please see details in Appendix C(i)), we initially assess the SD of the moderator (SUGRO) as

continuous moderator, as it will denote that the moderator is one SD below and one SD above the mean. To estimate the CoMe effects for low, medium, and high moderator levels, we use equations 9–11 for *CoMe Model B* (to conduct the analysis, also see Appendix C(i), Figures C5–C7). Results in Table 2

show that the path coefficients of the CoMe effect become weaker when the SUGRO effect is one SD above the mean. 19 Additionally, the CoMe effects for low, medium, and high moderator levels are significant because both extremes of the CI are positive. However, we could not reveal whether a CoMe effect exists for the use of a continuous moderator because of the insignificant result of the CoMe index (ω). Meanwhile, the use as categorical moderator concerns the comparison between the low value (0) and the high value (1) (for the analysis, also see Appendix C(ii), Figures C12-C14). Results in Table 2 show that the path coefficients of the CoMe effect for high SUGRO $(\beta = 0.134)$ are weaker than for low SUGRO ($\beta =$ 0.212), and both CoMe effects are significant because both extremes of the CI are positive. Therefore, we find some evidence supporting the hypothesis included in CoMe Model B for the use of the categorical moderator only; that is, the mediated effect of PI on BI via PV reduces as SUGRO increases (to corroborate this finding, also see Figure 12 that depicts the results using the simple slope for categorical moderator).

Conditional Mediation Analysis Model C-Continuous and Categorical Moderator

Based on Figure 13 for *CoMe Model C*, we focus on testing the influence of the moderator (SUGRO) on the mediated effect $(p_1 \cdot p_2)$. We first assess the CoMe index (ω) , which is derived from the following equation: $\omega = p_1 \cdot p_2 + (p_2 \cdot p_5 \cdot + p_1 \cdot p_7)$ SUGRO $+ p_5 \cdot p_7 \cdot SUGRO^2$ (Equation 3). Based on our estimations (see Appendix D(i); Figure D4), the results show that " $p_2 \cdot p_5 \cdot + p_1 \cdot p_7$ " for the continuous moderator is supported, as the CI does not include zero (90 percent CI: -0.057 to -0.013). Similarly, our estimations (see Appendix D(ii); Figure D13) show that the result of " $p_2 \cdot p_5 \cdot + p_1 \cdot p_7$ " for the categorical moderator is also supported, as the CI excludes zero (90 percent CI: -0.137 to -0.049). ²⁰

Moreover, the index " $p_2 \cdot p_5 \cdot + p_1 \cdot p_7$ " of SUGRO is significant for both the continuous ($\beta = -0.035$) and categorical ($\beta = -0.093$) moderators. Therefore, we can conclude that the mediated effect of PI on BI through PV decreases as the moderator SUGRO increases (Table 2).

Then, we investigate the mediated effects at different levels of the moderator (Preacher et al., 2007) for the use of both continuous and categorical moderators. Again, based on the procedures derived from our guidelines (see Appendix D(i)), we initially assess the SD of the moderator (SUGRO) for continuous moderator, as it will denote that the moderator is one SD below and one SD above the mean.²¹ To estimate the CoMe effects for low, medium, and high moderator levels, we use equations 12-14 for Model C (see Appendix D(i), Figures D5-D10). Results in Table 2 show that the path coefficients of the CoMe effect become weaker when the SUGRO effect is one SD above the mean. Additionally, the CoMe effects for low, medium, and high levels of the moderator are significant because both extremes of the CI are positive. Meanwhile, using the categorical moderator concerns the comparison between the low value (0) and the high value (1) (for the analysis, also see Appendix D(ii), Figures D14-D15). Results in Table 2 reveal that the path coefficients of the CoMe effect for high SUGRO (β = 0.067) are weaker than for low SUGRO (β = 0.146), and both CoMe effects are significant because both extremes of the CI are positive. Therefore, we find some evidence supporting the hypothesis in CoMe Model C for using both the categorical and continuous moderators. That is, the mediated effect of PI on BI via PV decreases as the moderator SUGRO increases (to corroborate this finding, also see Figure 14 that depicts the results using the Johnson-Neyman plot and simple slope for the use of both the continuous and categorical moderators).

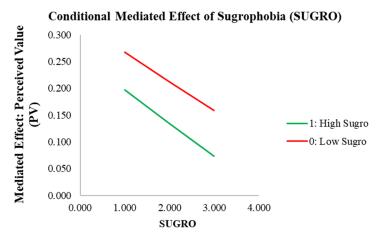


Figure 12. Categorical Variable for CoMe Effects Plot for CoMe Model B

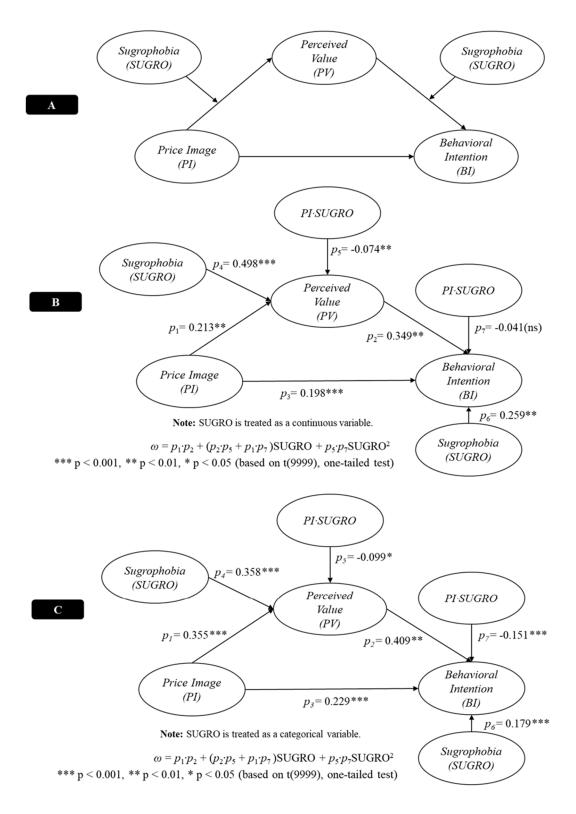


Figure 13. Conceptual Model (Panel A), Structural Model Results with SUGRO as Continuous Variable (Panel B), Structural Model Results with SUGRO as Categorical Variable (Panel C) for CoMe Model C

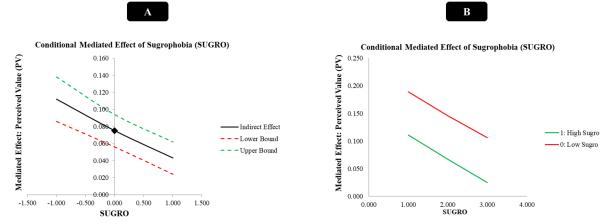


Figure 14. Continuous Variable (Panel A) and Categorical Variable (Panel B) for CoMe Effects Plot for CoMe Model C

Table 2. Conditional Mediated Effect Results for CoMe Models A–C (Continuous Data & Transformation from Continuous Data to Median Split)

Model	Moderator Type	Conditional mediated effect (CoMe effect)	Std Beta	Lower Percentile (5%)	Upper Percentile (95%)	SE	t-value
Α	Continuous	Index of Conditional Mediation	-0.037	-0.064	-0.011	0.016	2.295
		Sugrophobia (W)					
		Low (-1.0013)	0.144	0.101	0.189	0.027	5.378
		Medium (0.0000)	0.107	0.071	0.144	0.023	4.770
		High (1.0013)	0.070	0.023	0.116	0.028	2.474
	Dichotomous (Median Split)	Index of Conditional Mediation– Dummy Variable	-0.049	-0.093	-0.003	0.027	1.798
		Sugrophobia (W)					
		Low (0)	0.178	0.132	0.227	0.029	6.159
		High (1)	0.130	0.071	0.196	0.038	3.406
В	Continuous	Index of Conditional Mediation	-0.021	-0.044	0.002	0.014	1.518
		Sugrophobia (W)					
		Low (-1.0011)	0.201	0.144	0.261	0.036	5.657
		Medium (0.0000)	0.180	0.128	0.236	0.033	5.502
		High (1.0011)	0.159	0.103	0.220	0.036	4.465
	Dichotomous (Median Split)	Index of Conditional Mediation– Dummy Variable	-0.078	-0.117	-0.040	0.024	3.294
		Sugrophobia (W)					
		Low (0)	0.212	0.159	0.267	0.033	6.504
		High (1)	0.134	0.073	0.197	0.038	3.559
С	Continuous	Index of Conditional Mediation	-0.035	-0.057	-0.013	0.013	2.658
		Sugrophobia (W)					
		Low (-1.0012)	0.112	0.073	0.157	0.026	4.378
		Medium (0.0000)	0.075	0.047	0.107	0.019	4.035
		High (1.0012)	0.043	0.013	0.076	0.019	2.227
	Dichotomous (Median Split)	Index of Conditional Mediation– Dummy Variable	-0.093	-0.137	-0.048	0.027	3.443
		Sugrophobia (W)					
		Low (0)	0.146	0.106	0.189	0.026	5.658
		High (1)	0.067	0.031	0.111	0.025	2.699

Comparison of Prediction Ability between Continuous and Categorical Moderators

We extend the evaluation assessment of CoMe effects drawing on the interaction term approach by comparing the result of in-sample prediction (i.e., coefficient of determination, adjusted coefficient of determination, and model selection criteria) and outprediction of-sample (i.e., blindfolding PLS predict).²² This allows assessment of the efficacy of various quality criteria between the use of a continuous moderator and the transformation from continuous to categorical moderator. Results in Table 3 exhibit that the majority of the criteria for in-sample and out-of-sample predictions favor the use of a continuous moderator compared with the use of transformation from continuous to the categorical moderator.

To reiterate, unless the researcher seeks to derive and present findings in a way that they might serve a particular purpose (e.g., presentation of findings to a managerial audience to ease comprehension of actionable insights) or encounters data characterized by irregular or non-normal distributions or outliers (DeCoster et al., 2009), when a continuous moderator is suitable and available in CoMe analysis, we recommend researchers to be very cautious in transforming continuous data to categorical one for CoMe analysis because it loses prediction power. In the case that the researcher chooses to work with data that are transformed to categorical one, all CoMe analyses ought to be based on sufficient statistical

power to warrant producing consequential results and drawing meaningful conclusions.

Conditional Mediation Analysis-MGA Approach Using Categorical Moderator

In this section, we continue drawing on the data from Cheah et al. (2020) to illustrate the use of MGA in CoMe analysis, especially when the moderator is categorical and the research aim is testing the difference in strength (magnitude) between mediated effects. Hence, we re-estimate the model by replacing SUGRO with gender as the categorical moderator.²³ As outlined above, the three constructs (PI and PV) are modeled as composites and estimated in Mode A, while BI is modeled as a composite and estimated in Mode B. Additionally, since gender is used as a potential moderator, the total of 400 consumers sample is split into two groups: female (n = 206) and male consumers (n = 194) (Figure 15, Panel A).

In this illustration, we evaluate whether the mediated effect of $p_1 \cdot p_2$ for the female group is significantly different from that for the male group. ²⁴ The equation for this significant difference is $d = p_1^{\text{(Female)}} \cdot p_2^{\text{(Female)}} - p_2^{\text{(Male)}}$. Thus, an MGA is carried out. After following the MICOM procedure and concluding that partial measurement invariance is established, we apply a permutation-based test (Chin & Dibbern, 2010), representing the most powerful nonparametric MGA testing approach (Klessel et al., 2021). ²⁵

Table 3. Causal-Prediction Criteria Values for CoMe Models A-C

		R2/A	dj R²	Blindfo	folding Q ² PLS_predict				Model Selection Criteria			
CoMe	Moderator	PV	BI	PV	BI	RMSE	MAE	Q ² _predict	AIC	BIC	GM	
Model	Type					(PV/BI)	(PV/BI)	(PV/BI)	(PV/BI)	(PV/BI)	(PV/BI)	
Α	Continuous	0.475/	0.474/	0.382	0.424	0.738/	0.588/	0.466/	-251.063/	235.097/	422.966/	
		0.471	0.471			0.775	0.595	0.409	-251.895	-239.920	416.974	
	Dichotomous	0.380/	0.474/	0.305	0.423	0.802/	0.633/	0.368/	184.344/	-168.378/	422.966/	
	(Median	0.376	0.471			0.808	0.615	0.357	-251.773	-239.799	416.974	
	Split)											
В	Continuous	0.270/	0.516/	0.217	0.458	0.865/	0.682/	0.261/	122.741/	-114.758/	410.983/	
		0.268	0.512			0.773	0.594	0.412	-281.632	-261.675	428.957	
	Dichotomous	0.270/	0.508/	0.217	0.447	0.867/	0.680/	0.261/	-122.737/	114.754/	410.983/	
	(Median	0.268	0.503			0.796	0.618	0.376	-274.957	-254.999	428.957	
	Split)											
С	Continuous	0.475/	0.516/	0.382	0.458	0.737/	0.588/	0.465/	-250.831/	-234.866/	422.966/	
		0.471	0.511			0.754	0.569	0.455	-281.444	-261.486	428.957	
	Dichotomous	0.380/	0.508/	0.305	0.447	0.801/	0.633/	0.367/	-184.321/	-168.355/	422.966/	
	(Median	0.375	0.503			0.783	0.597	0.397	-274.952	-254.995	428.957	
	Split)											

Note: Gray shading represents the best value

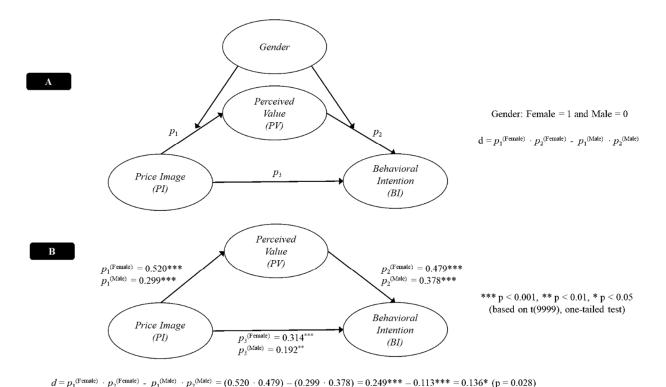


Figure 15. Conceptual Model (Panel A) and Structural Model Results (Panel B) with Categorical Moderator

Table 4. The Difference in Mediated Effects Between Female and Male Consumers

Gender		Female (n = 206)				Male (n = 194)				Male-Female		
Mediated Effect	β	<i>t</i> -value	<i>p</i> -value	PBCI 90%	β	<i>t</i> -value	<i>p</i> -value	PBCI 90%	d	Permutation p-value		
PI -> PV -> BI $(p_1 \cdot p_2)$	0.249	5.317	0.000	[0.172; 0.328]	0.113	3.355	0.000	[0.062; 0.170]	0.136	0.028		
Direct Effect	β	<i>t</i> -value	<i>p</i> -value	PBCI 90%	β	<i>t</i> -value	<i>p</i> -value	PBCI 90%	d	Permutation <i>p</i> -value		
PI -> PV (p ₁)	0.520	8.169	0.000	[0.399; 0.611]	0.299	4.305	0.000	[0.174; 0.403]	0.220	0.013		
PV -> BI (p ₂)	0.479	6.691	0.000	[0.357; 0.586]	0.378	5.149	0.000	[0.248; 0.493]	0.102	0.164		
PI -> BI (p ₃)	0.314	4.624	0.000	[0.204; 0.426]	0.192	2.370	0.009	[0.044; 0.312]	0.122	0.103		

Note: Gray shading represents significant value.

The permutation-based test shows how the difference in mediated effects between the female and male groups is significantly different from zero (Table 4). Therefore, the mediated effect of PI on BI through PV is conditional on gender as the moderator. Although we observe partial mediation in both groups (see the results in Figure 15, Panel B and Table 4), the mediated effect is significantly higher for the female group (β = 0.249) than for the male group (β = 0.113).²⁶ Therefore, we find support for a significant CoMe effect. ²⁷ Finally, given that the study shows a significant difference only between $p_1^{(Female)}$ and $p_1^{(Male)}$, the model described here corresponds to a *CoMe Model A* (Table 4).

Summary and Conclusions

CoMe analysis often opens up new avenues for analyzing new research questions. Such modeling

provides deeper insights into the intricacies of processes or under which conditions they occur, which is not the case when mediation and moderation are tested independently. However, the use application of CoMe analysis in PLS-SEM have been scarce and, in many cases, vague in their explanation. One reason could be the lack of experience and knowledge to handle such complexities in a PLS-SEM setting, especially when the estimation procedure is not straightforward. With these issues in mind, this paper illuminates conceptual fundamentals and develops guidelines for CoMe analysis within the PLS-SEM context. In doing so, clear distinction between the different types of CoMe models (i.e., Models A-E) is made. Furthermore, the paper outlines the illustrative use of CoMe analysis in PLS-SEM and presents detailed step-by-step procedures to estimate and assess CoMe effects in the PLS-SEM setting (as per corresponding appendices). In doing so, we explain

how to estimate the CoMe index (ω), and illustrate the applications of CoMe analysis in PLS-SEM for the most prominent CoMe models (i.e., Models A, B, and C). Such detailed illustrations will increase researchers' awareness of the procedures for and serve as a guide for estimating and interpreting CoMe results in PLS-SEM use. Thus, this paper responds to the recent claim by Sarstedt et al. (2020a) about the unnecessity of using PLS-SEM and PROCESS in tandem.

To recap the utility of conducting CoMe analysis in provide concludina PLS-SEM. we three recommendations concerning model specification, causal inferences, and issues regarding sample size. Although some of these recommendations apply to all research areas, they are particularly crucial for proficiently conducting CoMe analysis. First, a clear and robust theoretical basis must be established for any CoMe model (i.e., first-, second-, or first- and second-stage moderation) (Borau et al., 2015). Holland et al. (2017) have stressed the importance of theoretical rigor to substantiate the correct CoMe model and whether hypothesized arguments concern the rate of change (i.e., slope) or differences in the strength (i.e., magnitude) of mediated relationships that are conditional. The guidelines and procedures provided here, regardless of how robust and effective they are, in no way undermine the central importance of a theoretical foundation on which a CoMe model ought to rest. In other words, each moderated path in a mediation model must be founded and rooted clearly in a theoretical framework (see Holland et al., 2017). The same logic applies when multiple or serial mediation relationships are moderated in a model: each moderation must be justified. Hence, the approach to conducting CoMe analysis rests on sound theoretical substantiation and suitable data from a fitting research design.

Second, CoMe analysis implies sequential causation, which may occur simultaneously but can also happen over time, where changes in some variables affect changes in other variables over various points in time (Holland et al., 2017). Therefore, drawing on data from experimental and longitudinal research designs is often appropriate (see Mathieu et al., 2008; Roemer, 2016; Stone-Romero & Rosopa, 2011; Streukens & Leroi-Werelds, 2016b). However, this does not render the use of data from cross-sectional research designs useless when conducting CoMe analysis. Despite the limitations of such research designs (i.e., omitted variable problems that can cause spurious relationships), CoMe analysis resting on such research designs still allows drawing inferences about the direct and mediated effects and differences in the rate of change (i.e., slope) or in the strength (i.e., magnitude) in relationships produced by moderation.

Thus, one might use the approach provided by Hult et al. (2018) to assess and control endogeneity, particularly in the case of omitted variables (Sarstedt et al., 2020b; Spector, 2019), to conclude a causal effect from the CoMe effect estimation in PLS-SEM. Furthermore, when theoretical reasoning requires constructs being measured at certain points in time, such should be considered in a cross-sectional research design underlying the data when conducting CoMe analysis.

Third, when researchers seek to conduct MGA to detect CoMe effects, sufficient statistical power must be given to avoid type II errors (Dawson, 2014; Hayes, 2017). Researchers must ensure that the split samples for each group have close to equal and sufficient sample sizes to achieve sufficient statistical power (Aguinis et al., 2017). Likewise, if researchers choose to work with categorial data that are transformed from continuous one, all CoMe analyses ought to be based on sufficient statistical power to warrant producing consequential results. They can rely on power analyses that consider the model structure, the anticipated significance level, and the expected effect sizes (e.g., Marcoulides & Chin, 2013), Alternatively, Hair et al. (2022) have provided power tables that indicate the required sample sizes for various measurement and structural model constellations. Finally, Kock and Hadaya (2018) suggested the inverse square root method and the gamma exponential method as two new approaches for minimum sample size calculations assess whether sufficient statistical power exists.

Overall, we hope this paper, together with the detail provided in the appendices, will increase the number of studies applying CoMe analysis in PLS-SEM proficiently. Although our paper offers guidance for carrying out CoMe analysis in PLS-SEM, future research should seek to extend the detail provided here. For example, researchers could examine CoMe analysis for different types of CoMe models such as CoMe Model D or CoMe Model E that we had excluded in this paper but also others such as one in which two exogenous moderators (rather than one) simultaneously condition either the first-stage or second-stage of the mediated relationship.

Additionally, researchers could explore the efficacy of the PLS algorithm when estimating the CoMe effect against other composite-based SEM methods, such as consistent PLS (Dijkstra & Henseler, 2015a; 2015b), generalized structured component analysis (Hwang et al., 2010), and regularized canonical correlation analysis (Tenenhaus & Tenenhaus, 2011), with the adaptation of sampling weight adjustments in PLS-SEM (Becker & Ismail, 2016; Cheah et al., 2020a). Moreover, researchers could explore consideration of more complex CoMe research frameworks, such as

how a moderator influences parallel or serial mediating models simultaneously (Hayes, 2017).

In addition, a further interesting direction for future research is to examine how best to incorporate CoMe analyses into repeated-measures designs (experimental designs) given that repeated-measures data allow scholars to target statistical inference at between-subject and within-subject sources of variation (Curran & Bauer, 2011; Streukens & Leroi-Werelds, 2016b). Similarly, assessments of CoMe models that involve intricate research designs, such as multilevel modeling (Hwang et al., 2007) and longitudinal data (Roemer, 2016), although in their early stages of development, require further advancement within the PLS-SEM context, in general, and when assessing CoMe models, in particular.

Also, whereas the approaches to CoMe analysis presented and outlined in this paper concern *a priori*, theoretically derived determination of CoMe models and considerations of whether moderations concern the rate of change (i.e., slope) versus differences in the strength (i.e., magnitude) of mediated relationships, the guidance presented here can be

Notes

 1 The illustrative uses of the CoMe index (ω) for CoMe Model A, CoMe Model and CoMe Model C are explained in Section 4.1. As for the CoMe index (ω) for CoMe Model D and CoMe Model E, the present study does not illustrate the empirical procedures but researchers could refer to Hayes (2015) or contact the corresponding author.

² A "conditional direct effect" occurs when the direct effect of X on Y estimates how differences in X relate to differences in Y holding constant the proposed mediator (Hayes, 2017). In other words, the conditional direct effect quantifies the effect of X on Y independent of the mediator but conditioned on a value of a moderator.

³ To obtain the bootstrap confidence intervals, in line with Aguirre-Urreta and Rönkkö (2018), "researchers should generally use the percentile method. However, when the reliability coefficient's bootstrap distribution is skewed, the bias corrected and accelerated (BCa) bootstrap interval method" should be used (p. 9).

⁴ If using existing data—especially continuous ones—that do not align with requirements of the analysis goal (i.e., researchers may have hypothesized that the first- and/or second-stage paths in a mediated relationship differ in terms of strengths/magnitude or directionality), researchers could subsequently transform their data into categorical data to test the CoMe effect.

⁵ Researchers need to be very cautious when converting continuous variables to categorical variables (also known as artificially dichotomizing) for testing a CoMe model. Doing so can reduce the statistical power to detect a CoMe effect and can produce a downward bias in the size of estimated CoMe effects (Aguinis et al., 2017). Additionally, some researchers also throw up theoretical questions about what particular dividing points (i.e., mean, median, and percentile) should be used (Hayes & Preacher, 2014). Hence, researchers

expanded to also incorporate exploratory approaches that draw on the assessment of unobserved heterogeneity. For instance, approaches such as FIMIX-PLS in combination with, for example, PLS-POS, PLS-GAS, or PLS-IRRS (e.g., Ratzmann, Gudergan, & Bouncken, 2016; Sarstedt, Ringle, & Gudergan, 2016; Sarstedt, Ringle, & Hair, 2017) can be applied in a CoMe analysis context, especially when assessing differences in the strength (i.e., magnitude) of certain paths.

Finally, although we advocate rigorous theoretical substantiation for any CoMe model, another direction for future research concerns developing approaches for model selection among different CoMe model types (e.g., Models A–B). Such advances can build on using the effectiveness of various quality criteria available in PLS-SEM, such as the model fit (Dijkstra & Henseler, 2015a, 2015b), PLS_predict (Shmueli et al., 2016; Shmueli et al., 2019), cross-validated predictive ability testing (Liengaard et al., 2020), and model selection criteria (Danks et al. 2020; Sharma et al., 2019; Sharma et al., 2020).

ought to exercise caution unless converting a continuous variable to a categorical variable has strong underlying reasoning and justification. Despite these well-known problems, dichotomizing from a continuous variable is frequently used by experimental consumer researchers. Furthermore, some of the benefits surrounding the continued use of dichotomization are: (1) it produces meaningful findings that are easy to understand for a wider, possibly less technically knowable, audience (i.e., practitioners), and (2) it allows dealing with irregular or nonnormal distributions, or the presence of outliers (DeCoster et al., 2009).

⁶ By default, the estimation of reflectively specified constructs draws on Mode A, whereas PLS-SEM uses Mode B for formatively specified constructs (Hair et al., 2022). When using Mode A (i.e., correlation weights), the bivariate correlation between each indicator and the construct determine the outer weights. In contrast, Mode B (i.e., regression weights) computes indicator weights by regressing each construct on its associated indicators.

 7 To test the significant of ω , the 95% percentile is often used for testing a two-sided hypothesis, in which the t-value must be greater than 1.96. However, if researchers are interested in testing a one-sided hypothesis, they can use the 90% percentile (Kock, 2015).

⁸ To obtain further clarification of these equations for CoMe Models B–E in Table 2, researchers can contact the corresponding author.

⁹ Hair et al. (2022) and Henseler and Fassot (2010) pointed out that when estimating a basic moderating effect, often researchers encounter the collinearity problem, especially when creating the interaction term from an exogenous construct and moderator. Thus, they typically standardize the indicators of the moderator before creating the interaction term. The standardization converts each variable

to a mean of 0 and a standard deviation of 1, which reduces collinearity introduced by indicator reuse. Furthermore, this standardization procedure helps facilitate the interpretation of the moderating effect. Based on these reasons, ±1 SD is used in estimating the CoMe effect. However, researchers can easily adopt the ±2 SD or ±3 SD because using a wider interval of SD will help researchers show how the data are scattered.

¹⁰ Before performing an MGA, the researcher should ensure that measurement invariance is achieved drawing on the measurement invariance of composite models (MICOM) procedure (see Henseler et al., 2016). By establishing MICOM, the researcher can be assured that the group differences in the model estimates do not result from distinct content and meanings of measurements used for the latent variables across groups. Notably, conducting an MGA requires partial measurement invariance, which is given when configural and compositional invariance hold. If partial measurement invariance is confirmed and the composites have equal mean values that vary across the groups, full measurement invariance is confirmed, which supports the pooled data analysis. In contrast, if partial measurement invariance is not present, this absence can deflate the power of statistical testing, influence the accuracy of estimators, and provide misleading results in further analyses.

¹¹ To test the significant difference in the mediated effect, the 95% percentile is often used for testing a two-sided hypothesis, in which the t-value must be greater than 1.96. For instance, the CoMe effect of A on B via M will be different in groups 1 and 2. However, if researchers are interested in testing a one-sided hypothesis, they can use the 90% percentile. For instance, the CoMe effect of A on B via M will be higher in group 1 than in group 2. Furthermore, if researchers are interested in examining the differences in the mediated effect of more than two groups, either Bonferroni correction or the Šidák procedure are the standard approach for controlling for the familywise error rate (Hair et al., 2018, p. 282).

¹² The term "sugrophobia" indicates consumers' fear of being taken advantage of, duped, and exploited by salespeople, especially when they are less savvy or experienced regarding a purchase (Vohs et al., 2007). Consumers experience sugrophobia when they are concerned about whether the salespeople judge them as less savvy or less experienced about a purchase. Moreover, consumers who are highly sugrophobic may show signs of paranoia and distrust in uncertainty during the purchasing, thus the effect could weaken the decision-making process from price image to behavioral intention (Cheah et al., 2020; Babin et al., 2016).

¹³ For the transformation of continuous data, we used the median split. Additionally, we have conducted a mean split, and the estimated CoMe results for Models A–C are similar to the median split (see Table 2).

 14 The descriptions of the procedures employed to assess the CoMe index (ω) and CoMe effects for CoMe Model A, CoMe Model B, and CoMe Model C (using Excel) are available from the corresponding author (or see Online Supplement OS.1).

¹⁵ The interaction effect (PI × SUGRO) was computed by applying the two-stage approach (Becker et al., 2018).

Bootstrapping based on 10,000 subsamples was used to test the significance of the parameters. The hypothesized effects were assessed by applying a one-tailed test for a t-distribution (90% CI).

¹⁶ The standard deviation of SUGRO is 1.0013.

¹⁷ If a continuous moderator is used in determining the different level of the CoMe effect, Preacher et al. (2007) suggest to use the Johnson-Neyman's plot rather than the simple slope that we use for a categorical moderator. Often, it is interpreted by examining the middle line (black color) and the upper and lower bound values. The middle line represents the mediated effect, in which the top point of the middle line represents a low conditional mediated effect and the bottom point of the middle line represents a high conditional mediated effect. Looking at Figure 10 (Panel A), the mediated effect becomes weaker based on the conditional effect of SUGRO. If a categorical moderator is used in determining the CoMe effect, one can use the simple slope suggested by Preacher et al. (2017), in which the researcher can compare the direction and steepness (or gradient) of the CoMe effects for low and high values of the moderator (in our case this is Low Sugro and High Sugro) (see Figure 10, Panel B).

 18 If the CoMe index (ω) is nonsignificant, researchers should stop assessing the CoMe effects for different levels of the moderator; rather, they should analyze moderation and mediation separately (see Figure 1). In our case, we continue with the empirical illustration by showing the steps on how to obtain the results of the CoMe effects for different levels of the moderator.

¹⁹ The standard deviation of SUGRO is 1.0011.

²⁰ When the analysis aims to examine differences in the strength (magnitude) of the mediated effects, the researcher could transform the data into categorical data for *CoMe Model C* and perform an MGA test. For the details of the steps for doing this, researchers can refer to Section 4.2.

²¹ The standard deviation of SUGRO is 1.0012.

²² For further explanations on in-sample and out-of-sample predictions, please refer to Chin et al. (2020).

²³ Many empirical studies have demonstrated differences between men and women's behavior on their information processing and purchase decision-making, such as in social media (Pivec & Maček, 2019), consumer purchase behavior (Lim et al., 2021), and in other m-commerce settings (Faqih & Jaradat, 2015). Therefore, examining gender as the potential moderator variable for any mediated effect is warranted.

²⁴ See Appendix E.

²⁵ This analysis was carried out using SmartPLS 3.3.3 (Ringle et al., 2015).

²⁶ The path model was re-estimated using a bootstrapping routine based on 10,000 samples to test the significance of the parameters. Regarding the assessment of the CoMe effect, because we provide a one-sided hypothesis, we applied a one-tailed test (90% CI).

 27 However, when the analysis aims to examine the rate of change (or slope) of the mediated relationship, researcher could refer to the steps in Section 4.1 to assess both the CoMe index (ω) and at different moderator levels for the CoMe effects.

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Appendix A. Assessment of Confirmatory Tetrad Analysis

					CI	adj*	
Construct	Model Implied Non-Redundant Vanishing Tetrad	Tetrad Value	Standard Error	t-value	Lower Bound	Upper Bound	Supported Reflective
Behavioral Intention	1: Attitude2, Attitude3, BI1, BI2	0.317	0.055	5.736	0.210	0.428	No
	2: Attitude2, Attitude3, BI2, BI1	0.293	0.055	5.300	0.186	0.402	
Price Image	1: PI1, PI2, PI3, PI4	0.103	0.066	1.561	-0.005	0.216	Yes
	2: PI1, PI2, PI4, PI3	-0.191	0.125	1.528	-0.411	0.040	
Perceived Value	1: PV1, PV2, PV3, Attitude1	-0.016	0.059	0.270	-0.129	0.102	Yes
	2: PV1, PV2, Attitude 2, PV3	-0.082	0.063	1.288	-0.201	0.048	
Sugrophobia	1: Inertia1, Sugro1, Sugro2, Sugro3	0.014	0.103	0.891	-0.193	0.210	
	2: Inertia1, Sugro1, Sugro3, Sugro2	0.082	0.093	0.376	-0.107	0.257	

Note: (i) * means that CI adj = 90% bias corrected and Bonferroni-adjusted bootstrap confidence interval and (ii) Since confirmatory tetrad analysis (CTA) does not function for latent variables with less than four items – in our case behavioral intention, perceived value, and sugrophobia – therefore, this study borrows other somewhat related items from attitude and inertia constructs with the highest correlation and highest loading (see Gudergan et al., 2008).

Appendix B(i). Step-by-Step Procedure for Conditional Mediation Analysis of *CoMe Model A* Using a Continuous Moderator

To obtain the index of " $p_2 \cdot p_5$ ", this result can be directly obtained from any type of PLS-SEM software (e.g., SmartPLS). Specifically, the researcher needs to estimate the path coefficients using the bootstrapping output of the specific mediated effect. After the convergence of the bootstrapping result, the researcher needs to examine the significance of " $PI*SUGRO \rightarrow PV \rightarrow BI$ " from the specific mediated effect output (see Figure B1, Panel A and Panel B).

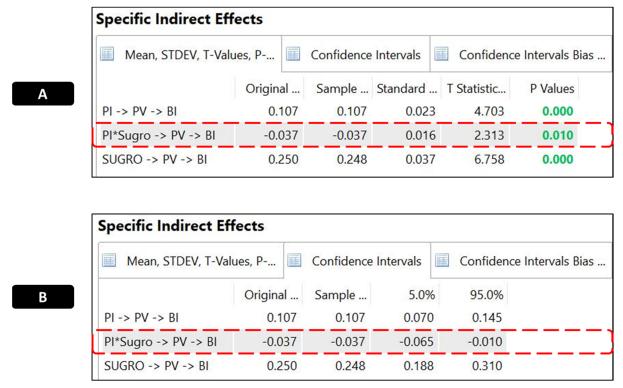


Figure B1. (Panel A) The p-value Result for the Index of " $p_2 \cdot p_5$ " and (Panel B) the Percentile-Corrected Result for the Index of " $p_2 \cdot p_5$ "

Next, we examine the mediated effect at different levels of the moderator. To formally test this for CoMe Model A, we initially assess the standard deviation of the moderator (SUGRO), as it will denote that the moderator is one standard deviation below and one standard deviation above the mean. The standard deviation for SUGRO is 1.0013. This standard deviation value can be obtained, for example, via the SPSS software or the use of an Excel spreadsheet. Subsequently, using a spreadsheet application (i.e., MS Excel), we suggest carrying out the following steps:

 Take the 10,000 path coefficients from all direct effects created by the bootstrap sampling and copy and paste them into a spreadsheet (see Figure B2).

	Α	В	С	D	E	F
1						
2						
3						
4						
5						
6						
7		p1	p2	p3	р4	p5
8		PI -> PV	PV -> BI	PI -> BI	SUGRO -> PV	PI*SUGRO -> PV
9	Sample 0	0.179	0.533	0.289	0.547	-0.076
10	Sample 1	0.249	0.503	0.295	0.472	-0.086
11	Sample 2	0.229	0.469	0.333	0.512	-0.056
12	Sample 3	0.221	0.513	0.263	0.469	-0.037
13	Sample 4	0.192	0.513	0.289	0.418	-0.151
14	Sample 5	0.216	0.397	0.370	0.484	-0.081

Figure B2. Step 1 Illustrative Example for Use of Continuous Moderator (Model A)

2) Based on columns B to F in Figure B2, create a new column for the CoMe effects at low, medium, and high levels of the moderator using the step-by-step illustrated equations suggested in Figure B3.

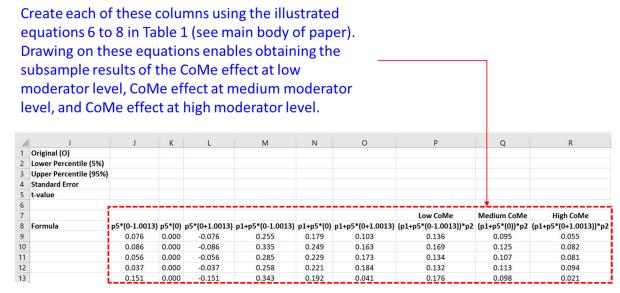


Figure B3. Step 2 Illustrative Example for Use of a Continuous Moderator (CoMe Model A)

3) Calculate the original values (O) as provided by the PLS-SEM path coefficients for each subsample (see Figure B4). The results show that the path coefficient effects of the CoMe effects are weaker when the SUGRO effect is one standard deviation above the mean.

Calculate the original values (O) provided by PLS for path coefficients.

The details are: =AVERAGE(P9:P10008), =AVERAGE(Q9:Q10008), and =AVERAGE(R9:R10008)

	ı	J	K	L	М	N	0	Р	Q	R
1	Original (O)							0.144	0.107	0.070
2	Lower Percentile (5%)									
3	Upper Percentile (95%)									
4	Standard Error							0.027	0.023	0.028
5	t-value							5.378	4.770	2.474
6										
7								Low CoMe	Medium CoMe	High CoMe
8	Formula	p5*(0-1.0013)	p5*(0)	p5*(0+1.0013)	p1+p5*(0-1.0013)	p1+p5*(0)	p1+p5*(0+1.0013)	(p1+p5*(0-1.0013))*p2	(p1+p5*(0))*p2	(p1+p5*(0+1.0013))*p2
9		0.076	0.000	-0.076	0.255	0.179	0.103	0.136	0.095	0.055
10		0.086	0.000	-0.086	0.335	0.249	0.163	0.169	0.125	0.082
11		0.056	0.000	-0.056	0.285	0.229	0.173	0.134	0.107	0.081
12		0.037	0.000	-0.037	0.258	0.221	0.184	0.132	0.113	0.094
13		0.151	0.000	-0.151	0.343	0.192	0.041	0.176	0.098	0.021

Figure B4. Step 3 Illustrative Example of for Use of Continuous Moderator (CoMe Model A)

4) Estimate the percentile bootstrap confidence interval (CI) using the function PERCENTILE (range, k), where k is the percentile value between 0 and 1 (see Figure B5). If the CI for the product "does not include a 0 value, this result means that the mediated effects at different levels of the moderator are significantly different from 0." In our example, the CoMe effects at low, medium, and high levels of the moderator are significant because both extremes of the CI are positive. Therefore, we find some evidence to support the hypothesis of CoMe Model A; that is, the mediated effect of the consumer's PI on BI via PV reduces as SUGRO increases.

Calculate both the lower percentile (5%) and the upper percentile (95%).

- The details of the lower percentile for CoMe effect at low moderator level, CoMe effect at medium moderator level, and CoMe effect at high moderator level are: =PERCENTILE(P9:P10008, 0.05), =PERCENTILE(Q9:Q10008, 0.05), and =PERCENTILE(R9:R10008, 0.05)
- The details of the upper percentile for CoMe effect at low moderator level, CoMe effect at medium moderator level, and CoMe effect at high moderator level are: =PERCENTILE(P9:P10008, 0.95), =PERCENTILE(Q9:Q10008, 0.95), and =PERCENTILE(R9:R10008, 0.95)

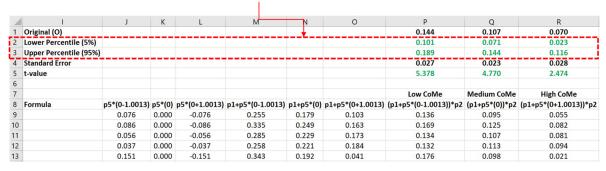


Figure B5. Step 4 Illustrative Example for Use of Continuous Moderator (CoMe Model A)

¹ Researchers could also in addition produce the t-value by first calculating the standard error (=STDEV(number 1, number 2, ...)). Next, use the original values (O) divided by the standard error to obtain the t-value.

² If researchers are interested in creating a graphical plot of the mediated effects at different levels of the moderator, they need to continue generating the standard error using the following formula: =STDEV(P9:P10008), =STDEV(Q9:Q10008), and =STDEV(R9:R10008) (see the result in Figure B5).

5) Finally, the researcher can create a graphical plot of the mediated effects at different levels of the moderator by using the proposed Excel template (see https://github.com/cortrudolph/Conditional_Indirect_Effects or https://www.md2c.nl/). Once the researcher has inserted this result (see Figure B6), he or she can easily produce the graphical plot (see Figure 10 in the main body of this paper).

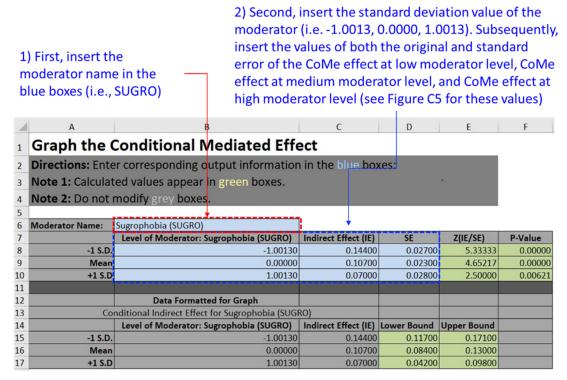


Figure B6. Step 5 Illustrative Example for Use of Continuous Moderator (CoMe Model A).

Appendix B(ii). Step-by-Step Procedure for Conditional Mediation Analysis of CoMe Model A When Moderator Is Converted from Continuous to Categorical

To obtain the index of " $p_2 \cdot p_5$ ", this result can be directly obtained from any type of PLS-SEM software even when continuous data has been transformed to categorical data via median split. Specifically, the researcher needs to estimate the path coefficients using the bootstrapping output of the specific mediated effect. After the convergence of the bootstrapping result, the researcher needs to examine the significance of " $PI*SUGRO \rightarrow PV \rightarrow BI$ " from the specific mediated effect output (see Figure B7, Panel A and Panel B).

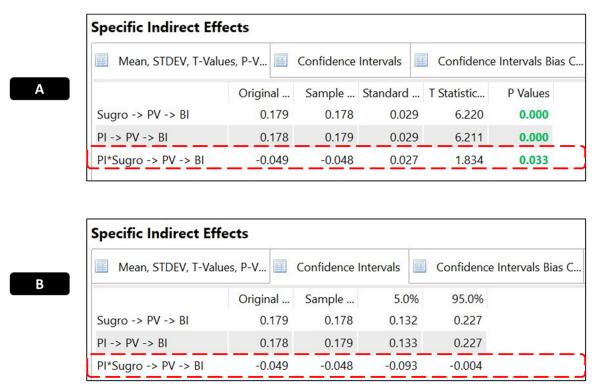


Figure B7. (Panel A) The p-value result for the index of " $p_2 \cdot p_5$ " and (Panel B) the percentile-corrected result for the index of " $p_2 \cdot p_5$ "

Next, we examine the mediated effects at different levels of the moderator. To formally test this effect for CoMe Model A, a researcher must be aware that there is no CoMe effects at low, medium, and high levels of the moderator as specified in the main body of the paper (see equations 6 to 8); rather, it requires them to insert the value of low SUGRO as 0 and high SUGRO as 1 into the formula of CoMe effects of Model A. Subsequently, using a spreadsheet application (i.e., MS Excel), we suggest carrying out the following steps:

1) Take the 10,000 path coefficients from all direct effects created by the bootstrap sampling and copy and paste them into a spreadsheet (see Figure B8).

	А	В	С	D	E	F
1	Original (O)			Ì		
2	Percentile Lower (5%)					
3	Percentile Upper (95%)					
4	Standard Error					
5	t-value					
6						
7		p1	p2	р3	p4	р5
8		PI -> PV	PV -> BI	PI -> BI	SUGRO -> PV	PI*SUGRO -> PV
9	Sample 0	0.347	0.505	0.255	0.344	-0.197
10	Sample 1	0.304	0.547	0.194	0.412	-0.058
11	Sample 2	0.489	0.440	0.311	0.293	-0.069
12	Sample 3	0.365	0.462	0.320	0.370	-0.024
13	Sample 4	0.409	0.487	0.320	0.343	-0.127
14	Sample 5	0.409	0.481	0.327	0.323	-0.068

Figure B8. Step 1 Illustrative Example for Use of Categorical Moderator (CoMe Model A)

2) Based on columns B to F, create a new column for the CoMe effect at low moderator level (or Low Sugro) and CoMe effect at high moderator level (or High Sugro) using the equations suggested in Figure B9.

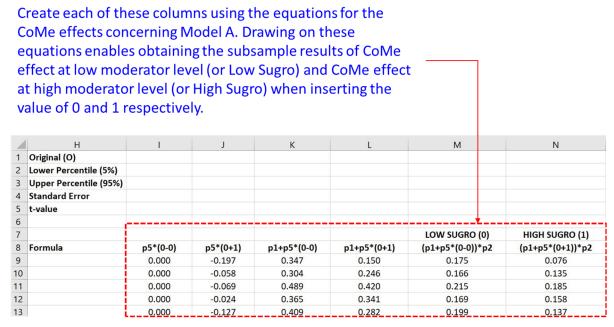


Figure B9. Step 2 Illustrative Example for Use of Categorical Moderator (CoMe Model A)

 Calculate the original values (O) provided by the PLS-SEM path coefficients for each subsample (see Figure B10). The results show that the path coefficient effects for CoMe effects of High Sugro are weaker than for Low Sugro. Calculate the original values (O) provided by PLS for path coefficients.

The details are: =AVERAGE(M9:M10008) and =AVERAGE(N9:N10008)

Z	Н		J	К	L	M	N
1	Original (O)					0.178	0.130
2	Lower Percentile (5%)						
3	Upper Percentile (95%)						
4	Standard Error					0.029	0.038
5	t-value					6.159	3.406
6							
7						LOW SUGRO (0)	HIGH SUGRO (1)
8	Formula	p5*(0-0)	p5*(0+1)	p1+p5*(0-0)	p1+p5*(0+1)	(p1+p5*(0-0))*p2	(p1+p5*(0+1))*p2
9		0.000	-0.197	0.347	0.150	0.175	0.076
10		0.000	-0.058	0.304	0.246	0.166	0.135
11		0.000	-0.069	0.489	0.420	0.215	0.185
12		0.000	-0.024	0.365	0.341	0.169	0.158
13		0.000	-0.127	0.409	0.282	0.199	0.137

Figure B10. Step 3 Illustrative Example for Use of Categorical Moderator (CoMe Model A)

4) Estimate the percentile bootstrap confidence interval (CI) using the function PERCENTILE (range, k), where k is the percentile value between 0 and 1 (see Figure B11). If the CI for the product "does not include a 0 value, this result means that the mediated effects at different levels of the moderator are significantly different from 0." In our example, the CoMe effects for Low Sugro and High Sugro are significant because both the extremes of the CI are positive. Therefore, we find some evidence to support the hypothesis of Model A; that is, the mediated effect of the consumer's PI on BI via PV reduces as SUGRO increases.

Calculate both the lower percentile (5%) and the upper percentile (95%).

- The details of the lower percentile for CoMe effect at low moderator level (or LOW SUGRO) and CoMe effect at high moderator level (or HIGH SUGRO) are: =PERCENTILE(M9:M10008, 0.05) and =PERCENTILE(N9:N10008, 0.05)
- The details of the upper percentile for CoMe effect at low moderator level (or LOW SUGRO) and CoMe effect at high moderator level (or HIGH SUGRO) are: =PERCENTILE(M9:M10008, 0.95) and =PERCENTILE(N9:N10008, 0.95)

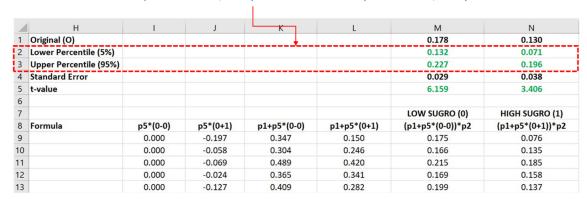


Figure B11. Step 4 Illustrative Example for Use of Categorical Moderator (CoMe Model A)

5) Finally, the researcher can create a graphical plot of the mediated effects at different levels of the moderator via an Excel spreadsheet (see Figure 10 in the main body of the paper).

³ To calculate the standard error and t-value, researchers can consider the comments in Footnote 1 and Footnote 2 in this appendix.

Appendix C(i): Step-by-Step Procedure for Conditional Mediation Analysis (CoMe Model B) Using a Continuous Moderator

Given that the index of " $p_1 \cdot p_5$ " cannot be directly obtained from any type of PLS-SEM software (e.g., SmartPLS), we need to examine the results of the path coefficients using the output of the bootstrapping sample. Consequently, once we run the estimations, we next perform a bootstrapping procedure with 10,000 subsamples. Then, using a spreadsheet and multiplying the bootstrapping outputs (i.e., $p_1 \cdot p_5$), we calculate the percentile bootstrap CI for the index of " $p_1 \cdot p_5$ ". Thus, using a spreadsheet application (i.e., MS Excel), we suggest carrying out the following steps:

1) Take the 10,000 path coefficients from all direct effects created by the bootstrap procedure and copy and paste them into a spreadsheet (Figure C1).

	Α	В	C	D	E	F
1	Original (O)					
2	Percentile Lower (5%)					
3	Percentile Upper (95%)					
4	Standard Error					
5	t-value					
6						
7		p1	p2	р3	p4	р5
8		PI -> PV	PV -> BI	PI -> BI	SUGRO -> BI	PV*Sugro -> BI
9	Sample 0	0.543	0.341	0.208	0.291	-0.046
10	Sample 1	0.55	0.309	0.14	0.377	-0.036
11	Sample 2	0.52	0.277	0.217	0.338	-0.041
12	Sample 3	0.548	0.325	0.259	0.21	-0.055
13	Sample 4	0.517	0.438	0.114	0.261	-0.068

Figure C1. Step 1 Illustrative Example for Use of Continuous Moderator (CoMe Model B)

2) Create a new column for the index of " $p_1 \cdot p_5$ " under assessment, and calculate the product of " $p_1 \cdot p_5$ " (see Figure C2).

Δ	Α	В	C	D	E	F	G
1	Original (O)						
2	Percentile Lower (5%)	C	reate a i	new col	umn "p1 * p	5", and calculat	e
3	Percentile Upper (95%)		the proc	luct of t	he paths tha	t form the inde	x —
4	Standard Error				ur	ider assessmen	t
5	t-value						
6							
7		p1	p2	р3	p4	p5	Index
8		PI -> PV	PV -> BI	PI -> BI	SUGRO -> BI	PV*Sugro -> BI	p1*p5
9	Sample 0	0.543	0.341	0.208	0.291	-0.046	-0.02498
10	Sample 1	0.55	0.309	0.14	0.377	-0.036	-0.0198
11	Sample 2	0.52	0.277	0.217	0.338	-0.041	-0.02132
12	Sample 3	0.548	0.325	0.259	0.21	-0.055	-0.03014
13	Sample 4	0.517	0.438	0.114	0.261	-0.068	-0.03516

Figure C2. Step 2 Illustrative Example for Use of Continuous Moderator (CoMe Model B)

3) Calculate the original values (O) to obtain the index value of " $p_1 \cdot p_5$ " (see Figure C3).

Calculate the original values (O) to obtain the index value of "p1 * p5".

The detail is: =AVERAGE(G9:G10008)

A	Α	В	С	D	E	F	6
1	Original (O)						-0.021
2	Percentile Lower (5%)						
3	Percentile Upper (95%)						
4	Standard Error						0.014
5	t-value						-1.518
6							
7		p1	p2	р3	p4	p5	Index
8		PI -> PV	PV -> BI	PI -> BI	SUGRO -> BI	PV*Sugro -> BI	p1*p5
9	Sample 0	0.543	0.341	0.208	0.291	-0.046	-0.02498
10	Sample 1	0.55	0.309	0.14	0.377	-0.036	-0.0198
11	Sample 2	0.52	0.277	0.217	0.338	-0.041	-0.02132
12	Sample 3	0.548	0.325	0.259	0.21	-0.055	-0.03014
13	Sample 4	0.517	0.438	0.114	0.261	-0.068	-0.03516

Figure C3. Step 3 Illustrative Example for Use of Continuous Moderator (CoMe Model B)

4) Estimate the percentile bootstrap CI, particularly for column " $p_1 \cdot p_5$ ", using the function PERCENTILE (range, k), where k is the percentile value between 0 and 1. In our case, given that our hypothesized effect is postulated with direction, we use a one-sided test, and we estimate a 90% CI (Figure C4). If the CI for the product of " $p_1 \cdot p_5$ " does not include a 0 value, this result means that the CoMe index is significantly different from 0.⁴ In our example, the index " $p_1 \cdot p_5$ " is insignificant because both extremes of the CI have mixed negative and positive sign; thus, there is change in sign between them. Therefore, we do not find evidence to support the hypothesis of Model B; that is, the mediated effect of a consumer's PI on BI via PV does not appear to decrease as SUGRO increases.

⁴ To calculate the standard error and t-value, researchers can consider the comments in Footnote 1 and Footnote 2 in this appendix.

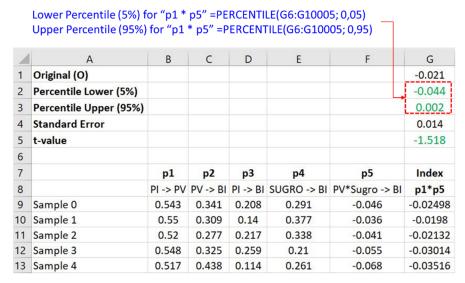


Figure C4. Step 4 Illustrative Example for Us of Continuous Moderator (CoMe Model B)

Next, we will continue to examine the mediated effects at different levels of the moderator.⁵ To formally test this effect for Model B, we initially assess the standard deviation of the moderator (SUGRO). The standard deviation for Sugro is 1.0011. Subsequently, we continue using MS Excel to execute the following steps.

5) Create a new column for the CoMe effects at low, medium, and high levels of the moderator using the illustrated formula suggested in Figure C5.

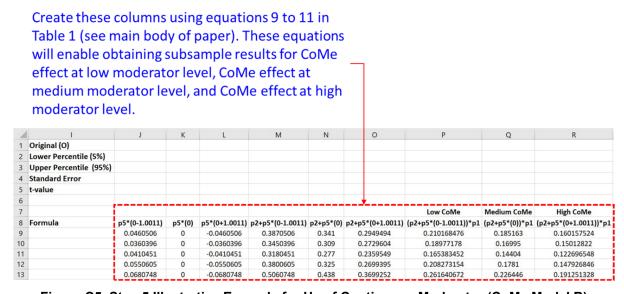


Figure C5. Step 5 Illustrative Example for Us of Continuous Moderator (CoMe Model B)

6) Calculate the original values (O) provided by the PLS-SEM path coefficients for each subsample (see Figure C6). The results show that the path coefficient effects of the CoMe effect become weaker when the Sugro effect is one standard deviation above the mean.

⁵ If the index of conditional mediation is nonsignificant, researchers should stop assessing the different level of CoMe effects; rather, they should analyze moderation and mediation separately (see Figure 1 in main body of this paper). In our case, we continue the empirical illustration merely to show the steps on how to obtain the CoMe effects for different levels of the moderator.

Calculate the original values (O) provided by PLS for path coefficients.

The details are: =AVERAGE(P9:P10008), =AVERAGE(Q9:Q10008), and =AVERAGE(R9:R10008)

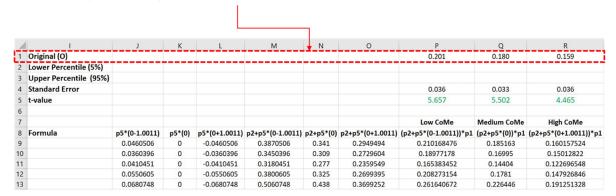


Figure C6. Step 6 Illustrative Example for Use of Continuous Moderator (CoMe Model B)

7) Estimate the percentile bootstrap CI using the function PERCENTILE (range, k), where k is the percentile value between 0 and 1 (see Figure C7). If the CI for the product "does not include a 0 value, this result means that the mediated effects at different levels of the moderator are significantly different from 0." In our example, the CoMe effects at low, medium, and high levels of the moderator are significant because both extremes of the CI are positive. However, we are unable to support the hypothesis of Model B; that is, the mediated effect of PI on BI via PV is weaker due to the negative impact of SUGRO. The reason is that the index of conditional mediation is nonsignificant (see Figure C4). Thus, it is preferable to analyze and report moderation and mediation separately. Notably, researchers should not proceed with creating the graphical plot of the mediated effects at different levels of the moderator.

Calculate both the lower percentile (5%) and the upper percentile (95%).

- The details of the lower percentile for CoMe effect at low moderator level, CoMe effect at medium moderator level, and CoMe effect at high moderator level are:
 =PERCENTILE(P9:P10008, 0.05), =PERCENTILE(Q9:Q10008, 0.05), and
 =PERCENTILE(R9:R10008, 0.05)
- The details of the upper percentile for CoMe effect at low moderator level, CoMe effect at medium moderator level, and CoMe effect at high moderator level are:
 =PERCENTILE(P9:P10008, 0.95), =PERCENTILE(Q9:Q10008, 0.95), and
 =PERCENTILE(R9:R10008, 0.95)

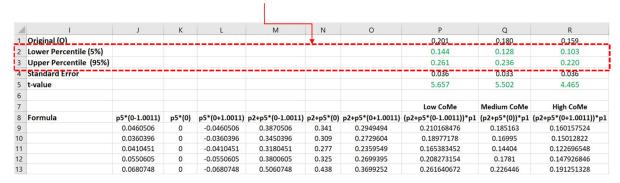


Figure C7. Step 7 Illustrative example for use of continuous moderator (CoMe Model B)

⁶ To calculate the standard error and t-value, researchers can consider the comments in Footnote 1 and Footnote 2 in this appendix.

Appendix C(ii): Step-by-Step Procedure for Conditional Mediation Analysis (CoMe Model B) When Moderator Is Converted from Continuous to Categorical

Given that the index of " $p_1 \cdot p_5$ " cannot be directly obtained from any type of PLS-SEM software, we need to examine the results of the path coefficients using the output of the bootstrapping sample, even when continuous data has been transformed to categorical data via median split. Consequently, once we estimate the model, we next perform a bootstrapping procedure with 10,000 subsamples. Then, using a spreadsheet and multiplying the bootstrapping outputs (i.e., $p_1 \cdot p_5$), we calculate the percentile bootstrap CI for the index of " $p_1 \cdot p_5$ ". Thus, using an Excel spreadsheet, we suggest carrying out the following steps:

1) Take the 10,000 path coefficients from all direct effects created by the bootstrap procedure and copy and paste them into a spreadsheet (Figure C8).

1	А	В	С	D	E	F	G
1	Original (O)						
2	Percentile Lower (5%)						
3	Percentile Upper (95%)						
4	Standard Error						
5	t-value						
6							
7		p1	p2	р3	p4	p5	Index
8		PI -> PV	PV -> BI	PI -> BI	Sugro -> B	PV*Sugro -> BI	p1*p5
9	Sample 0	0.543	0.412	0.232	0.115	-0.085	-0.046155
10	Sample 1	0.522	0.312	0.248	0.229	-0.181	-0.094482
11	Sample 2	0.45	0.455	0.166	0.137	-0.239	-0.10755
12	Sample 3	0.554	0.391	0.26	0.207	-0.185	-0.10249
13	Sample 4	0.488	0.47	0.183	0.173	-0.212	-0.103456
14	Sample 5	0.559	0.313	0.352	0.198	-0.091	-0.050869

Figure C8. Step 1 Illustrative Example for Use of Categorical Moderator (CoMe Model B)

2) Create a new column for the index of " $p_1 \cdot p_5$ " under assessment, and calculate the product of " $p_1 \cdot p_5$ " (see Figure C9).

4	Α	В	С	D	E	F	G
1	Original (O)						
2	Percentile Lower (5%)						
3	Percentile Upper (95%)					", and calculat	
4	Standard Error	the	product	of the p		form the inde	
5	t-value				und	der assessmen	It
6							
7		p1	p2	р3	p4	p5	Index
		PΤ	PZ	ps	РТ	ps	illuex
8		2000	200.00		ALC: 100 SA	PV*Sugro -> BI	p1*p5
	Sample 0	2000	200.00		ALC: 100 SA	PV*Sugro -> BI	CONTRACTOR OF THE PARTY OF THE
8	Sample 0 Sample 1	PI -> PV	PV -> BI	PI -> BI	Sugro -> B	PV*Sugro -> BI -0.085	p1*p5
8		PI -> PV 0.543	PV -> BI 0.412	PI -> BI 0.232	Sugro -> B 0.115	PV*Sugro -> BI -0.085	p1*p5 -0.046155 -0.094482
8 9 10	Sample 1	PI -> PV 0.543 0.522	PV -> BI 0.412 0.312	PI -> BI 0.232 0.248	Sugro -> B 0.115 0.229	PV*Sugro -> BI -0.085 -0.181	p1*p5 -0.046155 -0.094482 -0.10755
8 9 10 11	Sample 1 Sample 2	PI -> PV 0.543 0.522 0.45	PV -> BI 0.412 0.312 0.455	PI -> BI 0.232 0.248 0.166	Sugro -> B 0.115 0.229 0.137	PV*Sugro -> BI -0.085 -0.181 -0.239 -0.185	p1*p5 -0.046155 -0.094482 -0.10755

Figure C9. Step 2 Illustrative Example for Use of Categorical Moderator (CoMe Model B)

3) Calculate the original values (O) to obtain the index value of " $p_1 \cdot p_5$ " (see Figure C10).

Calculate the original values (O) to obtain the index value of "p1 * p5".

The detail is: =AVERAGE(G9:G10008)

1	Α	В	С	D	Е	F	G
1	Original (O)						-0.078
2	Percentile Lower (5%)						
3	Percentile Upper (95%)						
4	Standard Error						0.024
5	t-value						-3.294
6							
7		p1	p2	р3	p4	p5	Index
8		PI -> PV	PV -> BI	PI -> BI	Sugro -> B	PV*Sugro -> BI	p1*p5
9	Sample 0	0.543	0.412	0.232	0.115	-0.085	-0.046155
10	Sample 1	0.522	0.312	0.248	0.229	-0.181	-0.094482
11	Sample 2	0.45	0.455	0.166	0.137	-0.239	-0.10755
12	Sample 3	0.554	0.391	0.26	0.207	-0.185	-0.10249
13	Sample 4	0.488	0.47	0.183	0.173	-0.212	-0.103456

Figure C10. Step 3 Illustrative Example for Use of Categorical Moderator (CoMe Model B)

4) Estimate the percentile bootstrap CI, particularly for column " $p_1 \cdot p_5$ ", using the function PERCENTILE (range, k), where k is the percentile value between 0 and 1. In our case, given that our hypothesized effect is postulated with direction, we use a one-sided test, and we estimate a 90% CI (Figure C11). If the CI for

the product of " $p_1 \cdot p_5$ " does not include a 0 value, this result means that the conditional mediating effect is significantly different from 0.⁷ In our example, the index " $p_1 \cdot p_5$ " is significant because both extremes of the CI have negative sign, thus, there is change in sign between them. Therefore, we do find evidence to support the hypothesis of Model B; that is, the mediated effect of a consumer's PI on BI via PV reduces as SUGRO increases.

Lower Percentile (5%) for "p1 * p5" =PERCENTILE(G9:G10008; 0,05) Upper Percentile (95%) for "p1 * p5" = PERCENTILE(G9:G10008; 0,95) F A В C D E G Original (O) 1 -0.078Percentile Lower (5%) -0.1173 Percentile Upper (95%) -0.040Standard Error 0.024 5 t-value -3.2946 7 p1 p2 p3 p4 p5 Index PI -> PV PV -> BI PI -> BI Sugro -> B PV*Sugro -> BI 8 p1*p5 9 Sample 0 0.543 0.412 0.232 0.115 -0.085 -0.046155 10 Sample 1 0.522 0.312 0.248 0.229 -0.181 -0.094482 11 Sample 2 0.45 0.166 -0.2390.455 0.137 -0.10755 12 Sample 3 0.554 0.26 -0.185 0.391 0.207 -0.1024913 Sample 4 0.488 0.47 0.183 0.173 -0.212 -0.103456

Figure C11. Step 4 Illustrative Example for Use of Categorical Moderator (CoMe Model B)

5) Create a new column for values concerning CoMe effects at low (or Low Sugro) and CoMe effects at high (or High Sugro) levels of moderator using the illustrated formula suggested in Figure C12.

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⁷ To calculate the standard error and t-value, researchers can consider the comments in Footnote 1 and Footnote 2 in this appendix.

Create values for each column using Model B's equation of the CoMe effects. This will enable obtaining subsample results for CoMe effect at low moderator level (or Low Sugro) and CoMe effect at high moderator level (or High Sugro) when inserting the value of 0 and 1 respectively.

4	i	J	K	L	М	N	0
1	Original (O)						
2	Lower Percentile (5%)						
3	Upper Percentile (95%)						
4	Standard Error						
5	t-value						
6					<u> </u>		
7						Low SUGRO	High SUGRO
8	Formula	p5*(0)	p5*(1)	p2+p5*(0)	p2+p5*(1)	(p2+p5*(0))*p1	(p2+p5*(0+1))*p1
9		0	-0.085	0.412	0.327	0.223716	0.177561
10		0	-0.181	0.312	0.131	0.162864	0.068382
11		0	-0.239	0.455	0.216	0.20475	0.0972
12		0	-0.185	0.391	0.206	0.216614	0.114124
13		0	-0.212	0.47	0.258	0.22936	0.125904

Figure C12. Step 5 Illustrative example for use of categorical moderator (CoMe Model B)

6) Calculate the original values (O) provided by the PLS-SEM path coefficients for each subsample (see Figure C13). The results show that the path coefficient effects for CoMe effects of high Sugro become weaker than for low Sugro.

Calculate the original values (O) provided by PLS for path coefficients.

The details are: =AVERAGE(N9:N10008) and =AVERAGE(O9:O10008)

4	<u> </u>	J	K	↓ L	M	N.	0
1	Original (O)					0.212	0.134
2	Lower Percentile (5%)						
3	Upper Percentile (95%)						
4	Standard Error					0.033	0.038
5	t-value					6.504	3.559
6							
7						Low SUGRO	High SUGRO
8	Formula	p5*(0)	p5*(1)	p2+p5*(0)	p2+p5*(1)	(p2+p5*(0))*p1	(p2+p5*(0+1))*p1
9		0	-0.085	0.412	0.327	0.223716	0.177561
10		0	-0.181	0.312	0.131	0.162864	0.068382
11		0	-0.239	0.455	0.216	0.20475	0.0972
12		0	-0.185	0.391	0.206	0.216614	0.114124
13		0	-0.212	0.47	0.258	0.22936	0.125904

Figure C13. Step 6 Illustrative Example for Use of Categorical Moderator (CoMe Model B)

7) Estimate the percentile bootstrap confidence interval (CI) using the function PERCENTILE (range, k), where k is the percentile value between 0 and 1 (see Figure C14). If the CI for the product "does not include a 0 value, this result means that the mediated effects at different levels of the moderator are significantly different from 0." In our example, the CoMe effects for low Sugro and high Sugro are significant because both extremes of the CI are positive. Therefore, we find some evidence to support the hypothesis of Model B; that is, the mediated effect of the consumer's PI on BI via PV reduces as SUGRO increases.

Calculate both the lower percentile (5%) and the upper percentile (95%).

- The details of the lower percentile for CoMe effect at low moderator level (or Low Sugro) and CoMe effect at high moderator level (or High Sugro) are: =PERCENTILE(N9:N10008, 0.05) and =PERCENTILE(O9:O10008, 0.05)
- The details of the upper percentile for CoMe effect at low moderator level (or Low Sugro) and CoMe effect at high moderator level (or High Sugro) are: =PERCENTILE(N9:N10008, 0.95) and =PERCENTILE(O9:O10008, 0.95)

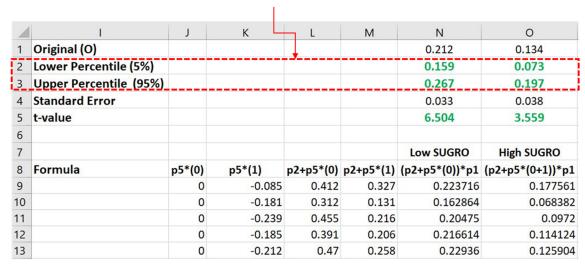


Figure C14. Step 7 Illustrative Example for Use of Categorical Moderator (CoMe Model B)

8) Finally, the researcher can create a graphical plot of the mediated effects at different levels of the moderator via an Excel spreadsheet (see Figure 12 in the main body of the paper).

Appendix D(i). Step-by-Step Procedure for Conditional Mediation Analysis (CoMe Model C) for continuous moderator

Given that the index of " $p_2 \cdot p_5$ " can be obtained except for " $p_1 \cdot p_7$ ", we decided to examine the results of the path coefficients using the output of the bootstrapping sample. Consequently, once we estimate the model, we next perform a bootstrapping procedure with 10,000 subsamples. Then, using a spreadsheet and multiplying the bootstrapping outputs, we calculate the percentile bootstrap CI for the index of " $p_2 \cdot p_5$ " and " $p_1 \cdot p_7$ ". Subsequently, we estimate the index of CoMe Model C by estimating the sum of both index of " $p_2 \cdot p_5$ " and " $p_1 \cdot p_7$ ". Thus, using an excel spreadsheet, we suggest carrying out the following steps:

1) Take the 10,000 path coefficients from all direct effects created by the bootstrap procedure and copy and paste them into a spreadsheet (Figure D1).

1	A	В	C	D	E	F	G	Н
1								
2	Original (O)							
3	Lower Percentile (5%)							
4	Upper Percentile (95%)							
5	Standard Error							
6	t-value							
7		p1	p2	р3	p4	р5	р6	р7
8		PI -> PV	PV -> BI	PI -> BI	SUGRO -> PV	PI*SUGRO -> PV	SUGRO -> BI	PV*SUGRO -> BI
9	Sample 0	0.168	0.305	0.207	0.506	-0.092	0.301	-0.02
10	Sample 1	0.263	0.349	0.23	0.44	-0.046	0.246	-0.029
11	Sample 2	0.251	0.388	0.246	0.491	-0.092	0.202	-0.015
12	Sample 3	0.229	0.406	0.197	0.481	-0.062	0.107	-0.081
13	Sample 4	0.209	0.268	0.215	0.514	-0.064	0.307	-0.079

Figure D1. Step 1 Illustrative Example for Use of Continuous Moderator (CoMe Model C)

2) Create a column for both the index of " $p_2 \cdot p_5$ ", " $p_1 \cdot p_7$ ", and " $p_2 \cdot p_5 + p_1 \cdot p_7$ " as well as calculate the product of the paths that form the index under assessment (see Figure D2).

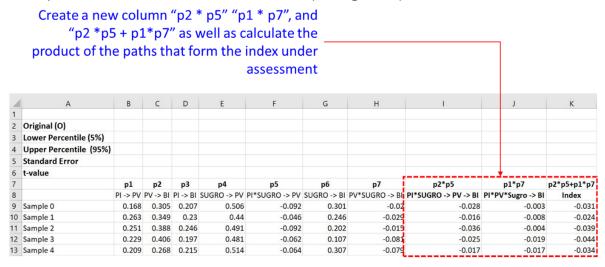


Figure D2. Step 2 Illustrative Example for Use of Continuous Moderator (CoMe Model C)

3) Calculate the original values (O) to obtain the index value of " $p_2 \cdot p_5$ ", " $p_1 \cdot p_7$ ", and " $p_2 \cdot p_5 + p_1 \cdot p_7$ " (see Figure D3).

Calculate the original values (O) to obtain the index value of "p2 * p5", "p1 * p7", and "p2 * p5 + p1 * p7".

The detail are: =AVERAGE(I9:I10008), =AVERAGE(J9:J10008), and =AVERAGE(K9:K10008)

1	A	В	С	D	E	F	G	н	1	J	K
1											
2	Original (O)								-0.026	-0.009	-0.035
3	Lower Percentile (5%)										
4	Upper Percentile (95%)										
5	Standard Error								0.012	0.006	0.013
6	t-value								-2.153	-1.397	-2.658
7		p1	p2	р3	p4	р5	р6	р7	p2*p5	p1*p7	p2*p5+p1*p7
8		PI -> PV	PV -> BI	PI -> BI	SUGRO -> PV	PI*SUGRO -> PV	SUGRO -> BI	PV*SUGRO -> BI	PI*SUGRO -> PV -> BI	PI*PV*Sugro -> BI	Index
9	Sample 0	0.168	0.305	0.207	0.506	-0.092	0.301	-0.02	-0.028	-0.003	-0.031
10	Sample 1	0.263	0.349	0.23	0.44	-0.046	0.246	-0.029	-0.016	-0.008	-0.024
11	Sample 2	0.251	0.388	0.246	0.491	-0.092	0.202	-0.015	-0.036	-0.004	-0.039
12	Sample 3	0.229	0.406	0.197	0.481	-0.062	0.107	-0.081	-0.025	-0.019	-0.044
13	Sample 4	0.209	0.268	0.215	0.514	-0.064	0.307	-0.079	-0.017	-0.017	-0.034

Figure D3. Step 3 Illustrative Example for Use of Continuous Moderator (CoMe Model C)

4) Estimate the percentile bootstrap CI, particularly for columns " $p_2 \cdot p_5 + p_1 \cdot p_7$ ", using the function PERCENTILE (range, k), where k is the percentile value between 0 and 1. In our case, given that our hypothesis postulates a direction, we use a one-sided test, and we estimate a 90% CI (Figure D4). If the CI for the product of " $p_2 \cdot p_5 + p_1 \cdot p_7$ " does not include a 0 value, this result means that the conditional mediating effect is significantly different from 0.8 In our example, the index " $p_2 \cdot p_5 + p_1 \cdot p_7$ " is significant because both extremes of the CI have negative sign, thus, there is a change in sign between them. Therefore, we do find evidence to support the hypothesis of CoMe Model C; that is, the mediated effect of a consumer's PI on BI via PV reduces as SUGRO increases, even though " $p_1 \cdot p_7$ " is insignificant.

Lower Percentile (5%) for "p2 * p5" =PERCENTILE(I9:I10008; 0,05) Upper Percentile (95%) for "p2 * p5" =PERCENTILE(I9:I10008; 0,95) Lower Percentile (5%) for "p1 * p7" =PERCENTILE(J9:J10008; 0,05) Upper Percentile (95%) for "p1 * p7" =PERCENTILE(J9:J10008; 0,95) Lower Percentile (5%) for "p2 * p5 + p1 * p7" =PERCENTILE(K9:K10008; 0,05) Upper Percentile (95%) for "p1 * p5 + p1 * p7" =PERCENTILE(K9:K10008; 0,95) 2 Original (O) 3 Lower Percentile (5%) -0.046-0.020 -0.057 4 Upper Percentile (95%) -0.007 0.001 -0.013 5 Standard Error 0.012 0.006 0.013 6 t-value -1.397 -2.153 -2.658 p2*p5 p1*p7 p2*p5+p1*p7 p2 p3 p4 p5 p6 p7 p1 PI -> PV PV -> BI PI -> BI SUGRO -> PV PI*SUGRO -> PV SUGRO -> BI PV*SUGRO -> BI PI*SUGRO -> PV -> BI PI*PV*Sugro -> BI 8 Index 9 Sample 0 0.168 0.305 0.207 0.506 -0.092 0.301 -0.02 -0.028 -0.003 -0.031 0.349 0.44 10 Sample 1 0.263 0.23 -0.046 0.246 -0.029 -0.016 -0.008 -0.024 11 Sample 2 0.251 0.388 0.246 0.491 -0.092 0.202 -0.015 -0.036 -0.004 -0.039 12 Sample 3 0.229 0.406 0.197 0.481 -0.062 0.107 -0.081 -0.025 -0.019 -0.044 0.514 -0.079 -0.017 13 Sample 4 0.209 0.268 0.215 -0.0640.307

Figure D4. Step 4 Illustrative example for use of continuous moderator (CoMe Model C)

⁸ To calculate the standard error and t-value, researchers can consider Footnote 1 and Footnote 2 in this appendix.

Next, we will continue to examine the mediated effects at different levels of the moderator. To formally test this effect for CoMe Model C, we initially assess the standard deviation of the moderator (SUGRO). The standard deviation for Sugro is 1.0012. Subsequently, we continue using MS Excel to execute the following steps.

5) Create a new column for the CoMe effects at low (or low Sugro), medium (or medium Sugro), and high (or high Sugro) levels of the moderator using the formula suggested in Figure D5 to Figure D7.

Create values for columns using Equation 12 in Table 1 (see main body in paper). This enables obtaining the subsample results for CoMe effect at low moderator level (or Low Sugro). **Low Sugro** 2 Original (O) Lower Percentile (5%) 4 Upper Percentile (95%) Standard Error 6 t-value 8 **Formula** $(p1 + (p5 \cdot SD \text{ of } -1.0012)) \cdot ((p7 \cdot SD \text{ of } -1.0012) + p2) \cdot [(p1 + (p5 \cdot SD \text{ of } -1.0012))*((p7 \cdot SD \text{ of } -1.0012) + p2)]$ 9 0.260 0.325 0.085 10 0.309 0.378 0.117 11 0.343 0.403 0.138 12 0.291 0.487 0.142

Figure D5. Step 5 Illustrative Example for Use of Continuous Moderator (CoMe Model C)

Create values for columns using Equation 13 in Table 1 (see main body in paper). This enables obtaining the subsample results for CoMe effect at medium moderator level (or Medium Sugro).

13

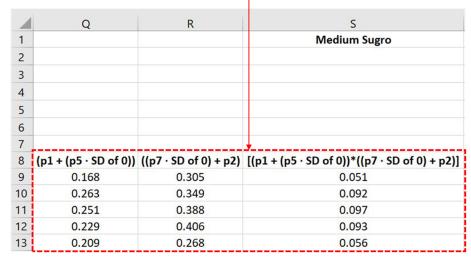


Figure D6. Step 6 Illustrative Example for Use of Continuous Moderator (CoMe Model C)

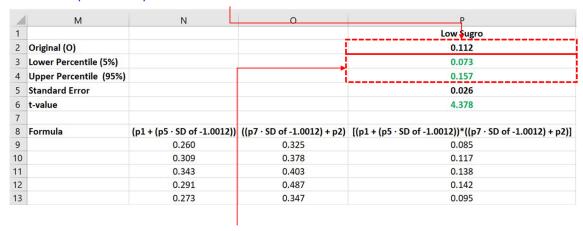
Create values for columns using Equation 14 in Table 1 (see main body in paper). This enables obtaining the subsample results for CoMe effect at high moderator level (or High Sugro).

1	T	U	V
1			High Sugro
2			
3			
4			
5			
6			
7		<u> </u>	
8	(p1 + (p5 · SD of 1.0012))	((p7 · SD of 1.0012) + p2)	[(p1 + (p5 · SD of 1.0012))*((p7 · SD of 1.0012) + p2)]
9	0.076	0.285	0.022
10	0.217	0.320	0.069
11	0.159	0.373	0.059
12	0.167	0.325	0.054
13	0.145	0.189	0.027

Figure D7. Step 7 Illustrative Example for Use of Continuous Moderator (CoMe Model C)

6) Calculate the original values (O) provided by the PLS-SEM path coefficients of each subsample (see Figure D8 to D10). The result shows that the path coefficient effects of the CoMe effects become weaker when the Sugro effect is one standard deviation above the mean. Subsequently, estimate the percentile bootstrap confidence interval (CI) using the function PERCENTILE (range, k), where k is the percentile value between 0 and 1 (see Figure D8 to D10). If the CI for the product "does not include a 0 value, this result means that the mediated effects at different levels of the moderator are significantly different from 0." In our example, the CoMe effects at low, medium, and high levels of the moderator are significant because both extremes of the CI are positive. Therefore, we find some evidence to support the hypothesis of CoMe Model C; that is, the mediated effect of the consumer's PI on BI via PV reduces as SUGRO increases.

Calculate the original value (O) provided by PLS for path coefficients for CoMe effect at low moderator level (or Low Sugro). The detail is: =AVERAGE(P9:P10008)



Calculate both the lower percentile (5%) and the upper percentile (95%).

- The details of the lower percentile for CoMe effect at low moderator level (or Low Sugro) is : =PERCENTILE(P9:P10008, 0.05)
- The details of the upper percentile for CoMe effect at low moderator level (or Low Sugro) is: =PERCENTILE(P9:P10008, 0.95)

Figure D8: Step 8 Illustrative Example for Use of Continuous Moderator (CoMe Model C)

Calculate the original value (O) provided by PLS for path coefficients for CoMe effect at medium moderator level (or Medium Sugro). The detail is: =AVERAGE(S9:S10008)

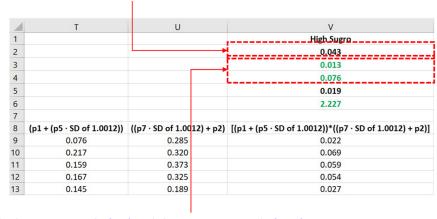
1	Q	R	S
1			Medium Sugro
2		•	0.075
3			0.047
4			0.107
5			0.019
6			4.035
7			
8	(p1 + (p5 · SD of 0))	((p7 · SD of 0) + p2)	$[(p1 + (p5 \cdot SD \text{ of } 0))*((p7 \cdot SD \text{ of } 0) + p2)]$
9	0.168	0.305	0.051
10	0.263	0.349	0.092
11	0.251	0.388	0.097
12	0.229	0.406	0.093
13	0.209	0.268	0.056

Calculate both the lower percentile (5%) and the upper percentile (95%).

- The details of the lower percentile for CoMe effect at medium moderator level (or Medium Sugro) is: =PERCENTILE(S9:S10008, 0.05)
- The details of the upper percentile for CoMe effect at medium moderator level (or Medium Sugro) is: =PERCENTILE(S9:S10008, 0.95)

Figure D9: Step 9 Illustrative Example for Use of Continuous Moderator (CoMe Model C)

Calculate the original value (O) provided by PLS for path coefficients for CoMe effect at high moderator level (or High Sugro). The detail is: =AVERAGE(V9:V10008)



Calculate both the lower percentile (5%) and the upper percentile (95%).

- The details of the lower percentile for CoMe effect at high moderator level (or High Sugro) is : =PERCENTILE(V9:V10008, 0.05)
- The details of the upper percentile for CoMe effect at high moderator level (or High Sugro) is: =PERCENTILE(V9:V10008, 0.95)

Figure D10. Step 10 Illustrative Example for Use of Continuous Moderator (CoMe Model C)

7) Finally, the researcher can create a graphical plot of the mediated effects at different levels of the moderator by inserting the estimated values in Figure D8 to D10 to an Excel spreadsheet, similar to the earlier suggestion in Figure B6 (see Figure 14 in the main body of paper).

Appendix D(ii). Step-by-Step Procedure for Conditional Mediation Analysis (CoMe Model C) When Moderator Is Converted from Continuous to Categorical

Given that the index of " $p_2 \cdot p_5$ " can be obtained except for " $p_1 \cdot p_7$ ", we decided to examine the results of the path coefficients using the output of the bootstrapping sample. Consequently, once we estimate the model, we next perform a bootstrapping procedure with 10,000 subsamples. Then, using a spreadsheet and multiplying the bootstrapping outputs, we calculate the percentile bootstrap CI for the index of " $p_2 \cdot p_5$ " and " $p_1 \cdot p_7$ ". Subsequently, we estimate the index of CoMe Model C by estimating the sum of both index of " $p_2 \cdot p_5$ " and " $p_1 \cdot p_7$ ". Thus, using an Excel spreadsheet, we suggest carrying out the following steps:

1) Take the 10,000 path coefficients from all direct effects created by the bootstrap procedure and copy and paste them into a spreadsheet (Figure D11).

1	А	В	C	D	E	F	G	Н
1	Original (O)							
2	Lower Percentile (5%)							
3	Upper Percentile (95%)							
4	Standard Error							
5	t-value							
6								
7		p1	p2	p3	p4	p5	p6	р7
8		PI -> PV	PV -> BI	PI -> BI	Sugro -> PV	PI*Sugro -> PV	Sugro -> BI	PV*Sugro -> BI
9	Sample 0	0.425	0.356	0.316	0.297	-0.094	0.157	-0.139
10	Sample 1	0.358	0.448	0.276	0.388	-0.229	0.137	-0.117
11	Sample 2	0.323	0.4	0.177	0.337	-0.127	0.219	-0.125
12	Sample 3	0.377	0.43	0.283	0.391	-0.119	0.124	-0.1
13	Sample 4	0.335	0.45	0.257	0.372	-0.094	0.118	-0.1

Figure D11. Step 1 Illustrative Example for Use of Categorical Moderator (CoMe Model C)

2) Create a new column for both the index of " $p_2 \cdot p_5$ ", " $p_1 \cdot p_7$ ", and " $p_2 \cdot p_5 + p_1 \cdot p_7$ " as well as calculate the product of the paths that form the index under assessment (see Figure D12).

Calculate the original values (O) to obtain the index value of "p2 * p5", "p1 * p7", and "p2 * p5 + p1 * p7".

The detail are: =AVERAGE(I9:I10008), =AVERAGE(J9:J10008), and =AVERAGE(K9:K10008)

A B C D E F G H J J Original (O) 2 Lower Percentile (5%) 3 Upper Percentile (95%) 4 Standard Error 9 0.023 0.017 5 t-value -1.761 -3.069

p5

-0.094

-0.229

-0.127

-0.119

-0.094

p4

0.297

0.388

0.337

0.391

0.372

Lower Percentile (5%) for "p2 * p5" =PERCENTILE(I9:I10008; 0,05)

p1

0.425

0.358

0.323

0.377

0.335

p2

0.356

0.448 0.276

0.4 0.177

0.43

0.45

p3

0.316

0.283

0.257

Figure D12. Step 2 Illustrative Example for Use of Categorical Moderator (CoMe Model C)

p6

0.157

0.137

0.219

0.124

0.118

PI -> PV PV -> BI PI -> BI Sugro -> PV PI*Sugro -> PV Sugro -> BI PV*Sugro -> BI PI*Sugro -> PV -> BI PI*PV*Sugro -> BI

p7

-0.139

-0.117

-0.125

-0.1

-0.1

p2*p5

-0.033

-0.102

-0.051

-0.051

-0.042

3) Estimate the percentile bootstrap CI, particularly for columns " $p_2 \cdot p_5 + p_1 \cdot p_7$ ", using the function PERCENTILE (range, k), where k is the percentile value between 0 and 1. In our case, given that our hypothesis postulates a direction, we use a one-sided test, and we estimate a 90% CI (Figure D13). If the CI for the product of " $p_2 \cdot p_5 + p_1 \cdot p_7$ " does not include a 0 value, this result means that the conditional mediating effect is significantly different from 0.9 In our example, the index " $p_2 \cdot p_5 + p_1 \cdot p_7$ " is significant because both extremes of the CI have negative sign. Therefore, we do find evidence to support the hypothesis of CoMe Model C; that is, the mediated effect of a consumer's PI on BI via PV reduces as SUGRO increases.

Upper Percentile (95%) for "p2 * p5" =PERCENTILE(I9:I10009; 0,95) Lower Percentile (5%) for "p1 * p7" =PERCENTILE(J9:J10008; 0,05) Upper Percentile (95%) for "p1 * p7" =PERCENTILE(J9:J10008; 0,95) Lower Percentile (5%) for "p2 * p5 + p1 * p7" =PERCENTILE(K9:K10008; 0,05) Upper Percentile (95%) for "p1 * p5 + p1 * p7" =PERCENTILE(K9:K10008; 0,95) 1 Original (O) -0.040 -0.053 -0.093Lower Percentile (5%) -0.078 -0.084 -0.137 Upper Percentile (95%) -0.048 -0.004 -0.026 0.027 Standard Error 0.023 0.017 5 t-value -1.761-3.069-3.4436 p2*p5 p1*p7 p2*p5 + p1*p7 p2 p3 p7 PI*PV*Sugro 8 PI -> PV PV -> BI PI -> BI S ugro -> PV PI*Sugro -> PV igro -> BI PV*Sugro -> BI PI*Sugro -> PV -> Index 0.157 9 Sample 0 0.297 -0.094 -0.139 -0.033 -0.059 -0.092 0.425 0.356 0.316 10 Sample 1 0.388 -0.229 0.137 -0.117 0.358 0.448 0.276 -0.102 -0.042 -0.144 11 Sample 2 0.323 0.4 0.177 0.337 -0.1270.219 -0.125 -0.051 -0.040 -0.091 0.377 0.43 0.391 -0.119 0.124 -0.051 -0.038 -0.089 0.335 -0.076

Figure D13. Step 3 Illustrative Example for Use of Categorical Moderator (CoMe Model C)

5

9 Sample 0

10 Sample 1

11 Sample 2

12 Sample 3

13 Sample 4

-0.093

0.027

-3.443

p2*p5 + p1*p7

Index

-0.092

-0.144

-0.091

-0.089

-0.076

p1*p7

-0.059

-0.042

-0.040

-0.038

-0.034

⁹ To calculate the standard error and t-value, researchers can consider comments in Footnote 1 and Footnote 2 in this appendix.

4) Create a new column for CoMe effects at low (or Low Sugro) and CoMe effects at high (or High Sugro) levels of moderator using the illustrated formula suggested in Figure D14.

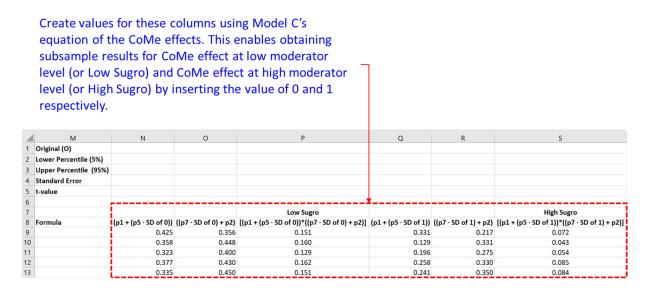
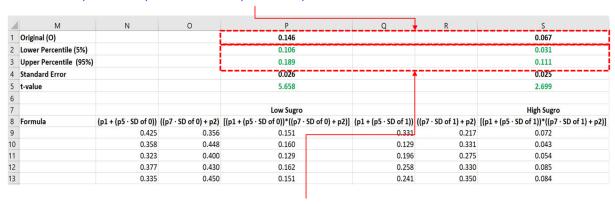


Figure D14. Step 4 Illustrative Example for Use of Categorical Moderator (CoMe Model C)

5) Calculate the original values (O) provided by the PLS-SEM path coefficients for each subsample (see Figure D15). The results show that the path coefficient effects of the CoMe effects become weaker for high Sugro than for low Sugro. Subsequently, estimate the percentile bootstrap confidence interval (CI) using the function PERCENTILE (range, k), where k is the percentile value between 0 and 1 (see Figure D15). If the CI for the product "does not include a 0 value, this result means that the mediated effects at different levels of the moderator are significantly different from 0." In our example, the CoMe effects for low Sugro and high Sugro are significant because both extremes of the CI are positive. Therefore, we find some evidence to support the hypothesis of Model C; that is, the mediated effect of the consumer's PI on BI via PV reduces as SUGRO increases.

Calculate the original value (O) provided by PLS for path coefficients for CoMe effect at low moderator level (or Low Sugro) and CoMe effect at high moderator level (or High Sugro). The detail are:

=AVERAGE(P9:P10008) and = AVERAGE(S9:S10008)



Calculate both the lower percentile (5%) and the upper percentile (95%).

- The details of the lower percentile for CoMe effect at low moderator level (or Low Sugro) and CoMe effect at high moderator level (or High Sugro) are: =PERCENTILE(P9:P10008, 0.05) and =PERCENTILE(S9:S10008, 0.05)
- The details of the upper percentile for CoMe effect at low moderator level (or Low Sugro) and CoMe effect at high moderator level (or High Sugro) are: =PERCENTILE(P9:P10008, 0.95) and =PERCENTILE(S9:S10008, 0.95)

Figure D15. Step 5 Illustrative Example for Use of Categorical Moderator (CoMe Model C)

6) Finally, the researcher can create a graphical plot of the mediated effects at different levels of the moderator via an Excel spreadsheet (see Figure 14 in the main body of paper).

Appendix E. Result of Measurement Invariance Test using MICOM

Gender	Compositional er Invariance			Equal Mean		Equal Variance			
Construc t	Configural Invariance Establishe d	c = 1	Confidence Interval (Cls) - 95%	Partial Measure ment Invarianc e Establis hed	Difference s	Confidence Interval (Cls) - 95%	Difference s	Confidence Interval (Cls) - 95%	Full measureme nt invariance established
BI	Yes	0.99 4	[0.974; 1.000]	Yes	0.831	[-0.176; 0.156]	0.551	[-0.304; 0.299]	No
PV	Yes	1.00 0	[0.999; 1.000]	Yes	0.863	[-0.164; 0.163]	0.491	[-0.289; 0.285]	No
PI	Yes	0.99 8	[0.990; 1.000]	Yes	0.671	[-0.161; 0.161]	0.592	[-0.277; 0.281]	No

Note: BI= Behavioural Intention, PV= Perceived Value; and PI= Price Image

Based on Appendix E, the results show that all composite constructs have significant differences in terms of the composite mean values and variances ratio because the results fall outside the upper and lower bounds of 95% confidence interval. Partial measurement invariance is thus established for the female and male groups (see last column in above table). It can be surmised that the different model estimations for the female and male groups are not distinct in terms of content or meaning of the constructs.