

1 Sexual selection can partly explain low frequencies of

2 *Segregation Distorter* alleles

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23 **Abstract**

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25 The *Segregation Distorter* (*SD*) allele found in *Drosophila melanogaster* distorts Mendelian
26 inheritance in heterozygous males by causing developmental failure of non-*SD* spermatids,
27 such that >90% of the surviving sperm carry *SD*. This within-individual advantage should
28 cause *SD* to fix, and yet *SD* is typically rare in wild populations. Here, we explore whether this
29 paradox can be resolved by sexual selection, by testing if males carrying three different
30 variants of *SD* suffer reduced pre or postcopulatory reproductive success. We find that males
31 carrying the *SD* allele are just as successful at securing matings as control males, but that one
32 *SD* variant (*SD-5*) reduces sperm competitive ability and increases the likelihood of female
33 remating. We then used these results to inform a theoretical model; we found that sexual
34 selection could limit *SD* to natural frequencies when sperm competitive ability and female
35 remating rate equalled the values observed for *SD-5*. However, sexual selection was unable
36 to explain natural frequencies of the *SD* allele when the model was parameterised with the
37 values found for two other *SD* variants, indicating that sexual selection alone is unlikely to
38 explain the rarity of *SD*.

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40 **Keywords:** Meiotic drive, gene drive, genomic conflict, sperm competition, mate choice

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46 Introduction

47 In sexually-reproducing organisms, meiosis ensures that autosomal alleles are divided evenly
48 between the haploid gametes. However, this equitable transmission can be subverted by
49 'selfish genetic elements' which encode phenotypes that are selected to increase their own
50 propagation, at the expense of other alleles in the genome [1]. These selfish alleles have
51 manifold ecological and evolutionary consequences [2], and given their potential to spread
52 even when they lower the fitness of individuals carrying them, efforts are underway to
53 develop synthetic selfish alleles that mimic their effects, with the aim to modify or suppress
54 populations [3]. This highlights a need to understand the evolutionary dynamics of naturally
55 occurring selfish alleles.

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57 One well-studied selfish allele is *Segregation Distorter (SD)*, a male gamete killer found in
58 *Drosophila melanogaster* [4]. *SD* is a large multigenic locus making up ~ 40% of the second
59 chromosome, a large autosome which itself comprises over a third of the genome. It contains
60 a distorter locus, multiple loci that enhance distortion, and a target site that is insensitive to
61 distortion [5]. In heterozygous *SD/+* males (that carry one *SD* allele and one homologous non-
62 distorting allele), *SD* causes spermatids that carry the non-distorting, sensitive allele to die
63 before completing development [5]. The result is that >90% of the male's functional sperm
64 carry *SD*, rather than the 50% expected for a typical heterozygous locus [6].

65

66 This large advantage in within-individual sperm competition should cause the *SD* allele to
67 reach fixation [7]. Contrary to this prediction, *SD* was only found on 0-8% of second
68 chromosomes in a sample of wild *D. melanogaster* populations [6]. A possible explanation for
69 this is that some variants of the *SD* allele accumulate harmful, recessive mutations causing

70 lethality, sterility, or greatly reduced fitness in *SD/SD* homozygotes [8, 9]. These recessive
71 mutations impose negative frequency-dependent selection on *SD*: as *SD* becomes more
72 common, the within-individual benefits of distortion are increasingly offset by the costs to *SD*
73 alleles in homozygotes, creating a balanced polymorphism of *SD* and non-distorting alleles.
74 However, population genetic models that consider recessive lethality [e.g. 7, 10] still
75 overestimate the equilibrium frequency of *SD*. For example, Bruck (1957) found that the
76 equilibrium frequency for a homozygous lethal segregation distorter is $\frac{1}{2} - \frac{\sqrt{k(1-k)}}{2k}$, where k
77 is the proportion of a heterozygous male's functional sperm that carry the distorting allele.
78 When $k = 0.9$, the predicted equilibrium frequency is 33%, suggesting there are unconsidered
79 fitness consequences associated with *SD* alleles.

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81 Here, we test whether sexual selection acting on males might partly explain why *SD* is rare in
82 natural populations. The population genetic effects of sexual selection have been well-
83 explored in other species harbouring segregation distorters [reviewed in 2, 11]. Moreover, a
84 recent study of *SD* showed that *SD/+* males were sometimes weak competitors in sexual
85 selection, but did not determine whether *SD/+* males have reduced success in pre- or post-
86 copulatory competition [or both; 9]. Theoretically, precopulatory sexual selection might help
87 to explain the rarity of *SD* if females tend to avoid mating with *SD/+* males if, for example,
88 females have been selected to avoid males that produce non-viable or *SD*-carrying offspring
89 [12]. *SD/+* males may also have reduced overall condition relative to *+/+* males, because the
90 large *SD* gene complex experiences little to no recombination, and is thus predicted to
91 accumulate deleterious mutations [13]. If either or both of these hold and because male
92 mating success often relies on condition-dependent traits [14], we predict females to mate
93 preferentially with non-*SD* males.

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Post-copulatory sexual selection may also explain the discrepancy between predicted and observed *SD* frequencies. Segregation distorters increase their relative within-individual frequency by destroying or incapacitating sperm carrying non-distorting homologous alleles. This means that *SD/+* males should produce half as many sperm as *+/+* males [5], assuming no compensatory increase in sperm production by the male [see 15]. The deleterious mutations carried by *SD*, or off-target effects of the sperm incapacitation mechanism, might reduce the number of sperm still further, and/or reduce their average competitive ability [16]. Sperm number and quality are key determinants of post-copulatory mating success [17, 18], such that *SD* alleles might have reduced fitness in populations where females mate multiply [as hypothesised for other distorters e.g. 19, 20]. In support of this hypothesis, segregation distorters reduce sperm competitive ability in other fly species and mice [21-25]. Building upon earlier models [7, 10], evolutionary simulations accounting for sperm competition costs paired with homozygous viability costs have produced distorter frequency estimates that match observations from wild populations [26, 27]. However, the effect of *SD* on sperm competitive ability has never been measured.

Here we examined pre- and post-copulatory success for *SD/+* males, and also measured whether females preferentially re-mate after mating with *SD/+* males. *D. melanogaster* has strong last-male sperm precedence [28], and so effects of male genotype on female remating latency could strongly affect the fitness of the *SD* allele. In *Drosophila*, females tend to remate faster when their sperm storage organs are comparatively empty, e.g. because stored sperm steadily release chemicals such as sex peptide that suppress remating [29]. One might therefore expect *SD/+* males, which probably transfer fewer sperm [as found for a

118 segregation distorter in *D. simulans*; 30], to create a shorter post-mating refractory period in
119 their mates. Female remating is also strongly affected by seminal fluid proteins from the male
120 ejaculate [31], and it is also possible that the deleterious mutations linked to *SD* affect
121 seminal fluid quantity or quality.

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123 Finally, we present a population genetic model incorporating these effects in conjunction
124 with segregation distortion and homozygote lethality, which we parameterised with our
125 empirical results. We use the model to explore the effects that precopulatory mating success,
126 sperm competitive ability and female remating propensity have on the allele frequency of *SD*,
127 and to test whether the fitness costs we identified are sufficient to explain the observed
128 rarity of *SD* in nature [6].

129

130 **Methods**

131 **Fly stocks**

132 We maintained all stocks at 25°C under a 16:8h photoperiod in *Drosophila* vials (95 mm x
133 25mm) on food medium (recipe in Table S1; ~ 8cm³ in each vial), supplemented with dry
134 yeast. We used four genotypes in this study: three of these were heterozygous for three
135 different variants of *segregation distorter* (*SD*), all of which were originally collected in
136 Madison, Wisconsin [4]. The *SD* variants are named *SD-5* (Bloomington stock number:
137 64322), *SD-72* (64323), and *SD-Mad* (64324). Each variant is characterised by the inversions it
138 carries and/or its viability effects [5]; *SD-5* and *SD-72* are homozygous lethal, while *SD-Mad* is
139 not [though its homozygotes have low fitness; 9]. To minimise extraneous genetic differences
140 between the three *SD* genotypes, we first standardised the genotype of both of the sex
141 chromosomes, the non-*SD* copy of chromosome 2, and both copies of chromosome 3 using a

142 crossing scheme involving balancers (Figure S1). This scheme produced experimental lines
143 (hereafter *SD/+* lines) that carried one copy of a *SD*-variant chromosome and one copy of the
144 *w¹¹¹⁸* chromosome 2, and were otherwise genetically uniform, with the possible exception of
145 the tiny fourth chromosome. We confirmed that each of the *SD/+* lines exhibited segregation
146 distortion in a pilot experiment (see supplementary methods and Figure S2). The fourth
147 genotype (hereafter *+/+*) was a non-*SD* control, which we generated in identical fashion,
148 except that the flies carried a copy of chromosome 2 from the isogenic *w¹¹¹⁸* line (and were
149 therefore homozygous for both major autosomes), instead of an *SD*-bearing chromosome.
150 The *SD-5* line was not included in Experiment 1 because it went extinct when access to the
151 laboratory was restricted due to Covid-19 (Experiment 1 was the last to be completed).

152

153 We also used three other fly stocks to compete or mate with the *SD/+* and *+/+* lines. In our
154 experiments, we used males from two outbred strains to provide a standardised source of
155 competition against the *SD/+* and *+/+* males. For Experiment 1, we sourced males from a *LH_m*
156 population that is homozygous for the *bw* mutation and therefore expresses a brown eye
157 phenotype (hereafter *Lbw*). For Experiment 2 we used males from another *LH_m* population,
158 that is homozygous for the transgenic construct *Ubi-GFP* (hereafter *LH_m^{Ubi}*). The *Ubi-GFP*
159 construct is attached to chromosome three and causes ubiquitous expression of green
160 fluorescence in *D. melanogaster* when viewed under fluorescent light. Females that mated
161 with experimental and competitor males were sourced from a large, outbred population of
162 the *LH_m* line that does not harbour the *Ubi-GFP* construct.

163

164 For our experiments, we reared the four experimental genotypes at a density of 100 larvae
165 per vial. Each genotype was sired by parents two to four days old that had also developed

166 under density-controlled conditions. We collected virgin males from the *SD/+*, *+/+* and
167 competitor male *LH_m^{Ubi}* and *Lbw* populations, and virgin females from the *LH_m* population.
168 All virgins were collected within 8h of eclosion and housed in same-sex environments until
169 they were themselves two to four days old, to ensure sexual maturity at the onset of the
170 experiments. To minimise differences in male mating investment caused by the social
171 environment during the days preceding the experiment, we standardised the number of
172 adult experimental virgin males (and *Lbw* males, for Experiment 1) to approximately 10 per
173 vial. In Experiment 2 we housed adult *LH_m^{Ubi}* competitor males at 80 per vial, due to the
174 larger number of males required.

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176 **Experiment 1: Testing whether *SD/+* males exhibit reduced mating success**

177 To assess whether *SD/+* males suffer reduced mating success when competing with other
178 males, we employed a two-choice test design. We aspirated two males into a vial containing
179 food medium; first a brown-eyed *Lbw* male, followed by a white-eyed male carrying one of
180 the experimental genotypes (either *SD-72*, *SD-MAD* or the control). We then introduced a
181 single virgin *LH_m* female and noted the time. Once the female mated with one of the males,
182 we recorded the genotype of the successful male and the time at which mating started. After
183 the mating pair separated, we immediately ended the trial, recorded the time mating
184 finished and discarded the three flies. We recorded the mating outcomes from 124 triads and
185 conducted the experiment blind to male genotype, to prevent observer bias affecting the
186 results [32].

187

188 We note that eye colour may affect mating success, and as such we expect >50% of females
189 to mate with the brown-eyed competitor male over the white-eyed focal male [33].

190 However, the purpose of this experiment is to compare the relative mating successes of the
191 four types of experimental males, and this comparison is not confounded by differences in
192 eye colour.

193

194 **Experiment 2: Testing sperm competitive success and female remating propensity**

195 The aims of this experiment were to 1) measure sperm competitive success of *SD/+* males
196 and 2) test whether female remating propensity is affected by male genotype. We ran the
197 experiment across three blocks made up of flies from three consecutive generations and
198 again conducted the experiment blind to male genotype.

199

200 To mimic natural conditions and accentuate any effects of *SD* on sperm production, we
201 mated all *SD/+* and control males once, shortly before starting the experiment. To do this, we
202 paired individual virgin *SD/+* or control males with a virgin *LH_m* female, allowed the pair to
203 interact for two hours, and recorded that mating occurred. Males that did not mate were
204 discarded, and the mated males were used in the experiment two to three hours after
205 mating.

206

207 To measure P1 (the proportion of offspring sired by *SD/+* males when mating first), as well as
208 female remating propensity, we first paired a single *SD/+* or control male with a virgin *LH_m*
209 female and allowed them three hours to mate. We confirmed mating and discarded the male
210 once they disengaged from copula. After four days, we allowed females a single opportunity
211 to remate - we aspirated a single 6- to 8-day old *LH_m^{UBI}* male and the previously mated
212 female into a new food vial and observed the pair for a maximum of three hours. For both
213 mating interaction periods we recorded whether mating occurred, the time taken for mating

214 to begin (hereafter 'mating latency'), and the copulation duration. 94/196 females remated,
215 and we collected no further mating data on females that did not remate. Throughout the
216 experiment we observed 11 females mating after three hours had passed, before we could
217 discard them from their vial. We recorded these females as failing to re-mate, but we did
218 include them in the subsequent sperm competitive ability measurements in order to
219 maximise sample size. Upon completion of the female's second mating, we discarded males
220 and transferred females into a vial containing grape juice agar and a small amount of yeast
221 paste, and left them to oviposit for 72 hours.

222

223 We recorded the number of offspring sired by the *SD/+* (or *+/+*) male and the *LH_m^{UBI}*
224 competitor to estimate P1. We determined paternity by first counting the number of
225 offspring produced by each female using a light microscope, then counting the number of
226 these offspring expressing GFP fluorescence (using a Leica M165 FC Fluorescence
227 microscope): the offspring of *SD* males did not express GFP, while offspring of *LH_m^{UBI}*
228 competitor males exhibited strong fluorescence. We measured P2 (the proportion of
229 offspring produced by the *SD/+* male when the *SD/+* male mated second) for *SD/+* males in
230 identical fashion, except that the order of matings was reversed, with *LH_m^{UBI}* males mated to
231 females first and *SD/+* or control males mated to females second. This time, 119/246 females
232 remated within the three-hour observation period (and were scored as having remated), and
233 16 females were observed remating after this time (and were scored as not having remated,
234 but were included in subsequent the sperm competition progeny counts).

235

236 **Statistical analysis**

237 We analysed the results using Bayesian generalised linear mixed models implemented in the
238 *brms* package for R [34]. For all models, we specified a prior distribution of $N(\mu = 0, \sigma = 3)$ for
239 fixed effect estimates and $N(\mu = 0, \sigma = 5)$ for intercept estimates. We ran four chains per
240 model, each with 8000 iterations (2000 discarded as warmup), and confirmed model
241 convergence and fit with \hat{R} statistics and posterior predictive checks. To make inferences
242 about our models, we calculated posterior differences between the means of the *SD*-variant
243 treatment groups and the control treatment group. We interpret differences between the *SD*
244 lines and the control line for which the 95% uncertainty intervals exclude zero as noteworthy.

245

246 For Experiment 1, we modelled whether or not each male mated using a binomial model. We
247 fit *SD*-variant as a fixed effect, and rearing vial as a random effect (to model and control for
248 similarities between individuals that developed in the same vial). We also modelled the
249 mating latency and copulation duration for the subset of trials in which the *SD/+* or control
250 male mated, in two separate models, both using the Weibull distribution and with the same
251 fixed and random effects as the mating success model.

252

253 For Experiment 2, we modelled P1 and P2 separately using binomial models, with proportion
254 of offspring sired as the response variable. We fit the P1 model using the progeny count data
255 for females that mated with an *SD/+* or *+/+* male first, and the P2 model using data from
256 females that mated with these males second. We fit *SD*-variant as a fixed effect, as well as
257 Block (which models the variance produced by the replication of the experiment across three
258 generations). We also included rearing vial and individual ID as random effects. Secondly, we
259 used another binomial model to estimate the likelihood of female remating after mating with
260 each type of male. Thirdly, we modelled remating latency to further explore the effects of

261 male genotype on female remating. These data were modelled using a Weibull distribution
262 with right censoring, where females that did not re-mate within three hours were censored.
263 Both models of remating contained the same fixed effects as the sperm competition models
264 and rearing vial as a random effect. Finally, we modelled copulation duration using two
265 separate models, where the duration of the first and second matings were used as response
266 variables. We specified a Weibull distribution for each, and used the same fixed and random
267 effects as the remating models.

268

269 The raw data and R code used to run all analyses are presented at
270 https://tomkeaney.github.io/SD_sexual_selection/.

271

272 **Population genetic model**

273 The effect that *SD* has on a male's sperm competitive ability and its capacity to limit female
274 remating is likely to affect the frequency of *SD* in natural populations. We therefore built a
275 one-locus, two-allele population genetic model – parameterised with our estimates of
276 segregation distortion, mating success, sperm competitive ability and female remating
277 probability – to assess how these variables affect the evolutionary trajectory of the *SD* allele.

278

279 The model considers an infinite, panmictic population composed of two sexes with non-
280 overlapping generations. The population contains distorting *SD* alleles and non-distorting
281 wildtype alleles. Beginning with the fertilised zygotes, all genotypes survive to breeding age
282 with equal probability, except for *SD* homozygotes, which we assume to be inviable (Table S3
283 shows that our model returns the same equilibrium frequencies as earlier analytical models
284 [e.g. 7] if we only include segregation distortion and homozygote lethality). This assumption

285 simplifies the model considerably, and reflects reality for at least two of the *SD* variants [the
286 third has low but non-zero fitness in homozygotes; 9]. Removing this assumption would
287 result in elevated allele frequencies for *SD*, while modelling a viability cost to *SD/+* individuals
288 would lower the frequency of *SD* [see 9, 26].

289

290 After removing non-viable genotypes, the population matures to adulthood and breeds. We
291 implement precopulatory sexual selection on males via a parameter S_{precop} . When $S_{precop} = 1$,
292 the two male genotypes are selected as mates randomly, i.e. with probabilities equal to their
293 frequencies in the population. Values of S_{precop} below 1 indicate that *SD/+* males are poor
294 precopulatory competitors, while values above indicate they are superior competitors. S_{precop}
295 includes the short range sexual selection we measured in Experiment 2, as well as longer
296 range processes like mate searching. We explored the evolution of *SD* for parameter spaces
297 where $0.8 \leq S_{precop} \leq 1.2$.

298

299 With S_{precop} defined and the genotype frequencies among the surviving adults known, we
300 next calculated the frequencies of each possible mating type. We make the simplifying
301 assumption that females mate with a maximum of two males, which is likely reasonable given
302 that *D. melanogaster* has a long post-mating refractory period and thrice-mated females
303 produce very few offspring sired by the first-mated male [35]. The proportion of females that
304 mate twice is $p_{+/+}$ among females whose first mate was *+/+*, or $p_{SD/+}$ for females whose first
305 mate was *SD/+*. We focus on parameter spaces where $p_{SD/+} \geq p_{+/+}$ i.e. where females are
306 equally or more likely to remate after mating with *SD/+* males. The mating types therefore
307 consist either of a male-female pair, or triads containing a female, her first mate, and her
308 second mate. We began by multiplying the population frequency of *SD/+* males by S_{precop}

309 then renormalising all of the genotype frequencies to again sum to 1 (this step lowers or
310 raises the frequencies of mating types involving $SD/+$ males). Then, for singly-mated females,
311 the frequency of each mating type was calculated as $F_i M_j (1 - p_j)$, where F_i and M_j are the
312 female and male parental genotype frequencies, and p_j is the probability of female remating
313 following a first mating with a male of genotype j . Similarly, we found the expected
314 frequencies of each possible mating type for females that mated with two males via the
315 formula $F_i M_j N_k p_j$, where N_k represents the genotype frequency of the second male to mate.
316
317 We next model (order-specific) sperm competition, which is only necessary for females that
318 mated with one $SD/+$ and one $+/+$ male. We set the normal P1 value for the population,
319 $P1_{normal}$, to 0.1 (i.e males mating first sire 10% of the offspring produced by a twice-mated
320 female), which is broadly consistent with our empirical estimates and those from other
321 studies of *D. melanogaster* [e.g. 28, 36]. We also explored the parameter space where
322 $P1_{normal} = 0.5$, which represents a scenario where first-mating males sire half the offspring
323 produced by twice mating females. We assume that first-mating $SD/+$ males suffer a cost to
324 their sperm competitive ability when the female mates second with a $+/+$ male, such that the
325 $SD/+$ male sires a proportion $P1_{normal} - (P1_{normal} \times P1_{cost})$ of the offspring. When they occupy
326 the second mating role and a $+/+$ male mates first, $SD/+$ males suffer a cost to P2 and sire a
327 proportion $1 - (P1_{normal} + (1 - P1_{normal}) \times P2_{cost})$ of the offspring. We investigated the full range
328 of possible values for $P1_{cost}$ and $P2_{cost}$, i.e. 0-1, where 0 indicates that $SD/+$ males are equally
329 effective in sperm competition, and 1 indicates a complete loss of paternity for the $SD/+$ male
330 when females mate twice.
331

332 After determining the mating type frequencies and the outcome of sperm competition,
333 zygotes are produced and the adults are removed, starting the next generation. We assume
334 standard Mendelian inheritance except for zygotes fertilised by *SD/+* males, where 86.8%,
335 90.9% or 94.4% of zygotes inherit their father's *SD* allele (these values correspond to the k_c
336 estimates found in our pilot experiment; see supplementary methods and Table S2), instead
337 of the typical 50%.

338

339 We calculated the genotype frequencies each generation immediately after removing the
340 inviable *SD/SD* genotype. We found the equilibrium allele frequencies numerically, by setting
341 the initial frequency of *SD* to 0.01 and iterating for multiple generations until *SD* approached
342 extinction (freq < 0.0001), fixation (freq > 0.99), or until 1,000 generations had elapsed. We
343 wrote the model in *R*; the code and a detailed explanation of it can be found at
344 https://tomkeaney.github.io/SD_sexual_selection/.

345

346 Results

347

348 Experiment 1: No evidence for an effect of *SD* on male mating success

349 There was no difference between the proportion of successfully mating males carrying either
350 of the *SD*-variants and the *+/+* male control (Fig 1a and b). Moreover, we found weak
351 evidence that males carrying either *SD-Mad* or *SD-72* had *shorter* mating latencies than the
352 control males (*SD-Mad* odds difference from *+/+* males = -0.65, 95% CIs: -1.36 to 0.09, *SD-72*
353 odds difference from *+/+* males = -0.49, 95% CIs: -1.22 to 0.24; Figure S3), the opposite of
354 predicted if *SD* reduces male attractiveness to females. There was no difference in mating
355 duration between males carrying *SD-72*, *SD-Mad* or the control allele (Figure S4).

356

357 **Experiment 2: *SD* reduces sperm competitive success and female remating**
358 **propensity**

359 We found strong mating order effects on fertilisation success: males of all genotypes (both
360 experimental and competitor males) that mated second sired 6,556 of the 7,158 offspring
361 (92%) produced by the 227 females. *SD/+* males exhibited reduced P1 values compared to
362 experimental control males (Figure 1c and d). *+/+* control males sired 8.2% (95% CIs: 1-
363 44.4%) of offspring when their mates subsequently mated with an LH_m^{UBI} male. The negative
364 effect of *SD* on fitness was greatest in males carrying a copy of *SD-5* (log-odds mean
365 difference from *+/+* males = -2.47, 95% CIs: -4.46 to -0.57) who only sired 0.8% (CIs: 0.1-
366 5.8%) of offspring when mating first. Males heterozygous for *SD-72* and *SD-Mad* appeared to
367 suffer an intermediate reduction in P1, siring 2.2% (CIs: 0.2-17%) and 1.8% of offspring (CIs:
368 0.2%-16.3%). Their P1 estimates did not differ significantly from *+/+* males (*SD-72* log-odds
369 mean difference: -1.42, CIs: -3.45 to 0.59; *SD-Mad*: -1.57, CIs: -3.67 to 0.55; Figure 1d),
370 though we note that detecting a significant difference between two small proportions
371 requires a very large sample size.

372

373 The proportion of offspring sired by a *SD/+* male when mating second (P2) depended on the
374 variant of *SD* he carried (Figure 1e and f). Males heterozygous for *SD-5* sired 93.2% (CIs: 74.5-
375 98.9%) of the offspring produced by a female that had previously mated with an LH_m^{UBI} male.
376 This was significantly lower P2 than we recorded for *+/+* males (CIs: 97.9%, 91.6-99.7%; log-
377 odds mean difference: -1.25, CIs: -2.38 to -0.12). However, males heterozygous for the *SD-*
378 *Mad* allele sired 99.5% (CIs: 97.6-99.9%) of offspring when mating second, which was
379 significantly higher than the P2 estimated for *+/+* males (log-odds mean difference: 1.5; CIs:

380 0.29 to 2.76). There was no difference between the percentage of offspring sired by males
381 carrying the *SD-72* and the w^{1118} allele when mating second (log-odds mean difference: -0.13;
382 CIs: -1.2 to 0.92; Figure 1f).

383

384 A total of 94 of 196 (48%) females mated a second time, four days after initially mating with a
385 *SD/+* male. The genotype of the female's first mate significantly affected the probability of
386 remating (Figure 1g and h). Specifically, 75.5% (CIs: 55.5-89.2%) of females that originally
387 mated with a *SD-5/+* male mated again, while only 30.4% (CIs: 15%-51.1%) of females that
388 had originally mated with *+/+* males mated again (odds mean difference: 1.97, CIs: 1.03 to
389 2.98). There was no difference in the proportion of females remating that had originally
390 mated with males carrying a copy of the *SD-72* (42.5% remating, CIs: 23.1-64.1%), *SD-Mad*
391 (42.9% remating, CIs: 23-65.1%) or control alleles (Figure 1h). Additionally, females that
392 originally mated with *SD-5/+* males remated more quickly than females that had mated with
393 *+/+* males when presented with an opportunity to remate. The estimated mean remating
394 latency of these females was 58 minutes (CIs: 37-95 mins), about half the estimated mean for
395 those females that originally remated with *+/+* males (115 mins, CIs: 65-213 mins). We found
396 no difference between the remating latencies of females that originally mated with males
397 possessing a copy of the *SD-72*, *SD-Mad* or control allele (Figure S5).

398

399 There was no variation in mating duration between *SD/+* and *+/+* males when in the first
400 mating role (Figure S6). However, males carrying the *SD-72* allele mated for significantly
401 longer than did *+/+* males, when occupying the second mating role (odds mean difference:
402 0.29, CIs: 0.01 to 0.57; Figure S7). We found no difference between the mating durations of
403 males carrying the *SD-5*, *SD-Mad* or control allele when in the second mating role.

404

405 Population genetic model

406 We found many parameter spaces in which *SD* and wildtype alleles coexisted in a balanced
407 polymorphism (Figure 2). As in earlier models [e.g. 7, 10], *SD* was unable to drive to fixation
408 because we assumed that it is lethal in homozygous form, which creates negative frequency-
409 dependent selection. At low frequencies, *SD* alleles rarely pay the cost of homozygous
410 lethality, so they increase in frequency due to their within-individual distortion advantage.
411 However, as *SD* becomes more common, *SD/SD* zygotes are formed more commonly, which
412 removes *SD* from the population. This opposes the effects of segregation distortion, creating
413 a balanced polymorphism.

414

415 Furthermore, we found that both pre- and postcopulatory sexual selection affect the
416 equilibrium frequency of *SD*. Varying the mating success of *SD/+* males (controlled by the
417 parameter S_{precop}) within the parameter space that equates with our empirical data simply
418 shifts the equilibrium frequency of *SD* (Figure 2; the mating success of *SD/+* males increases
419 as panels move left to right). Put simply, detrimental effects of *SD* from precopulatory sexual
420 selection reduce its equilibrium frequency, while benefits increase it. In combination with our
421 empirical findings, the model suggests that precopulatory sexual selection against *SD* is not
422 strong enough to explain the rarity of *SD* in natural populations.

423

424 Figure 2 shows that postcopulatory sexual selection can stop the *SD* allele from invading
425 when it is also homozygous lethal. When there is strong second male sperm precedence
426 ($P1_{normal} = 0.1$), as in *Drosophila*, a proportional reduction in $P2$ for *SD* males matters more to
427 the equilibrium allele frequency of *SD* than a correspondingly large proportional reduction in

428 P1, as shown by Figure 2's relatively horizontal isobars (as compared to Figure S8). When
429 there is no second male sperm precedence ($P1_{normal} = 0.5$), costs to P1 and P2 are of equal
430 importance for the equilibrium allele frequency of *SD* (Figure S8; note the diagonal isobars).
431 However, when the mates of *SD/+* males remate more often than the population mean
432 ($p_{SD/+} > p_{+/+}$), *SD/+* males become increasingly likely to occupy the first mating role. This has
433 two general effects on the evolutionary outcome. First, with strong second male sperm
434 precedence, the first-mating male sires few offspring, and so *SD* becomes rarer when females
435 mated to *SD/+* males are more likely to remate; this is true even if we assume that *SD* does
436 not affect a male's success in sperm competition. If there is no second male sperm
437 precedence, the effect of remating likelihood becomes less pronounced (*c.f.* Figures 2 and
438 S8). Secondly, as $p_{SD/+}$ increases, the effect of $P1_{cost}$ on *SD* frequencies becomes increasingly
439 influential, because *SD/+* males occupy the first mating role more often (Figure 2; compare
440 the three rows).

441

442 To estimate how sexual selection might affect the frequencies of the three *SD*-variants we
443 studied, we plotted the points in the sperm competition parameter space where *SD-5*, *SD-72*
444 and *SD-Mad* occupy, based on our estimates from Experiment 2. Figure 2h best represents
445 the parameter space relevant to *SD-5*, as $p_{SD/+} = 0.75$ (meaning that females are ~2.5 times
446 more likely to remate relative to females that mated with a standard male), and $S_{precop} = 1$,
447 matching our empirical estimates. Here the equilibrium frequency for *SD-5* falls below 5%,
448 which is within the range of frequencies that *SD* is found to occur in real-world populations.
449 However, the predicted allele frequencies for *SD-72* and *SD-Mad* fell between 25-35% when
450 we observed the parameter space informed by our estimates of $p_{SD/+}$ and mating success for
451 these two genotypes (Figure 2e and f); this frequency is higher than observed in natural

452 populations. This likely reflects the simplifications made by of our model, especially our
453 assumption that *SD/+* males are equally fit as *+/+* males in all other contexts besides
454 precopulatory sexual selection and sperm competition, which is likely not correct [see 9].

455

456 Discussion

457 We evaluated whether sexual selection might explain the observed low allele frequencies of
458 the *SD* selfish allele, using experiments and a model. In Experiment 1, we found no evidence
459 that a single copy of *SD* reduces male mating success, suggesting that *SD* is not held at low
460 frequencies by pre-copulatory sexual selection. However, Experiment 2 revealed that males
461 carrying *SD-5* are poor sperm competitors, and that their mates are subsequently more likely
462 to mate again. Using a population genetic model, we found that if these effects on remating
463 and sperm competition are sufficiently large, they can fully explain the rarity in natural
464 populations. However, males carrying the *SD-72* or *SD-Mad* allele do not suffer sexually-
465 selected costs of the same sufficient magnitude, and so these costs seem unlikely to fully
466 explain the rarity of *SD* in nature. Overall, our results provide limited empirical support for
467 the hypothesis that post-copulatory sexual selection constrains the spread of *SD* .

468

469 We found no support for the hypothesis that male precopulatory competitive ability is
470 adversely affected by the distorting genes of *SD* or deleterious mutations found in the *SD*
471 locus. Furthermore, given that mating success is determined both by male-male competition
472 and female choice, our data suggest that females are unable to identify and/or discriminate
473 against *SD*-carrying males, as might be expected given the fitness costs of selecting *SD*-
474 carrying mates [12]. However, as with all other laboratory studies that have measured the

475 effects of segregation distorters on mating success, our experimental design removes the
476 need for males to locate females within a larger landscape. If the mutations hitchhiking
477 within the *SD* complex affect condition, this may reduce the mate-searching ability of males,
478 in which case we may underestimate precopulatory sexual selection against *SD* alleles.
479 Nevertheless, our findings align with explicit investigations of male mating success conducted
480 on the other well-known segregation distorters: *SR* elements in other *Drosophila* species [37,
481 38] and the *t* haplotype in mice [39], with one notable exception. Female *Teleopsis dalmanni*
482 stalk-eyed flies have been found to avoid mating with *SR* males [40, 41]. In these systems, *SR*
483 is genetically linked to a locus that affects eye-stalk width, a trait that is under sexual
484 selection due to female choice [42]. Here, it appears there are mutations hitchhiking within
485 the *SR* complex that reduce eye-stalk width, causing *SR* males to be disfavoured by females
486 [43]. It is unclear whether this female preference has been strengthened by the indirect
487 fitness benefits of mating with non-*SR* males, or if the female preference has evolved entirely
488 through conventional ‘good genes’ or ‘sexy sons’ processes, and *SR* males are coincidentally
489 affected because they carry deleterious mutations.

490

491 In Experiment 2, we found some evidence that *SD/+* males suffer reduced sperm competitive
492 ability. Males carrying *SD-5* sired significantly fewer offspring than *+/+* males when
493 competing against the sperm of a rival male, both in the P1 and P2 role. When paired with
494 homozygote lethality and an increased risk of sperm competition (resulting from elevated
495 rates of female remating), our model suggests that the observed sperm competition costs for
496 *SD-5* can explain the low *SD* frequencies found in wild populations. The poor sperm
497 competitive ability of *SD-5* males is consistent with previous work on other segregation
498 distorters [21-25]. Together, these studies suggest that a reduction in sperm number caused

499 by the targeted gamete killing of a segregation distorter has direct individual-level costs to
500 male fitness in polyandrous mating systems [2, 19]. Interestingly, while we observe mild
501 reductions in P1 for the *SD-72* and *SD-Mad* male carriers, we observe no costs to P2, and
502 even a small increase in P2 for *SD-Mad/+* males. Unlike for *SD-5*, our model suggests that the
503 P1 and P2 values observed for these variants are not sufficient to explain the low frequency
504 at which they are found in natural populations. There are several potential explanations for
505 the competitive P1 and P2 values observed for males carrying *SD-72* and *SD-Mad*. First, it is
506 unknown how many sperm are inseminated by *SD/+* males, and how much variation there is
507 between variants. Males might compensate for the sperm lost to distortion by investing
508 more in spermatogenesis, as demonstrated for stalk eyed fly populations harbouring *SR* [15].
509 Under this scenario, *SD* would incur a direct material cost to the male, but not to his sperm
510 competitive ability. It is also possible that while *SD/+* males suffer a reduced absolute sperm
511 number, they ‘strategically allocate’ their sperm towards early matings [44]. If true, we might
512 not observe large deficits in sperm competition, as the maximum number of matings for a
513 male in our experiments was two. In Experiment 2, we found that males carrying the *SD-72*
514 allele, but not the *SD-5* or *SD-Mad* alleles, mated for significantly longer in the second mating
515 role than did males carrying non-distorting alleles. This may suggest variation between males
516 carrying different *SD* alleles in ejaculate investment, however, while mating duration is
517 positively correlated with the transfer of accessory seminal proteins in *D. melanogaster* [45],
518 there is no clear relationship between mating duration and sperm transfer [46]. Finally, it is
519 also possible that our control males, which were homozygous at most loci for the w^{1118}
520 genotype, have much lower sperm competitive ability than wild-type males, which would
521 lead to underestimation of the costs of *SD*.

522

523 In our model, we show that P1 becomes increasingly important for the evolutionary
524 trajectory of *SD* when *SD/+* males disproportionately occupy the first mating role. We also
525 show that this is a particularly plausible scenario, because we observed cryptic female choice
526 [as defined in 47] against *SD/+* males: the mates of *SD/+* males were more likely to remate
527 than females first mated to control males when given the opportunity. Interestingly, even in
528 the absence of sperm competition costs, the ability of males to reduce the risk of subsequent
529 sperm competition remains an important determinant of the *SD* equilibrium frequency. This
530 is likely because by inhibiting a female from remating, a male can avoid losing the majority of
531 any subsequently-produced offspring to the second male [approximately 90% in *D.*
532 *melanogaster*; 28, 36]. Accordingly, our model confirms that female remating behaviour may
533 be a more important determinant of *SD* frequencies than sperm competitive ability.

534

535 In sum, we show for the first time that post-copulatory sexual selection, combined with
536 homozygote lethality, is sufficient to explain the rarity of a particularly costly variant of *SD* in
537 wild populations. However, sexual selection alone seems unable to explain the rarity of the
538 two other *SD* variants studied here, implying that other evolutionary or ecological factors are
539 involved. For example, there may be alleles that confer resistance to segregation distortion
540 [5]. Other sources of selection against *SD* are likely important too, such as costs of *SD* to
541 survival, longevity, or mate-searching in heterozygotes. Higher order levels of selection may
542 also play a role, for example if *SD* reduces the size of a population relative to populations that
543 do not harbour the selfish allele [48]. Future empirical studies could manipulate the strength
544 of sexual selection acting on laboratory populations and test whether this affects the invasion
545 success of the *SD* allele, for example by manipulating female remating frequency [as in 49]
546 and/or the opportunity for pre-copulatory sexual selection. There is also scope to further our

547 understanding of how segregation distorters affect population dynamics [2], which
548 incidentally might inform the development of synthetic selfish genetic elements for
549 population control [3].

550

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554

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677

678

679 Figure captions

680

681 **Figure 1.** The effect of *SD* on male mating success, fertilisation success and female remating

682 propensity. Black points indicate the estimated mean, with associated 66 and 95% uncertainty

683 intervals, while coloured area shows the posterior distribution. Panels a, c, e and g show results on
684 the response scale, while panels b, d, f, and h show log-odds differences between the *SD* variants and
685 the control allele; 95% uncertainty intervals that do not overlap zero indicate a significant effect.

686

687 **Figure 2.** Predicted equilibrium frequency of the *SD* allele, calculated from the population genetic
688 model. The plot depicts the interaction between the P1 and P2 costs suffered by *SD/+* males in their
689 effects on the equilibrium frequency of *SD* (shown by the colour scale and 10% contour lines). The
690 dashed line shows an equilibrium frequency of 8%, the upper estimate for *SD* in natural populations.
691 *SD/+* male mating success (S_{precop}) increases across the columns and the risk of sperm competition
692 caused by a female remating after first mating to an *SD/+* male, $p_{SD/+}$, increases down the rows (values
693 correspond to the risk of sperm competition we estimated in Experiment 2). The three points (with
694 associated 95% credible intervals) in each panel show where males carrying each *SD* variant fall in the
695 figure's parameter space. In the parameter space presented, $k = 0.944$, $P1_{normal} = 0.1$, and
696 *SD* homozygotes are non-viable.