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## ASSOCIATION STUDIES ARTICLE

## Genome-wide association analysis identifies three new susceptibility loci for childhood body mass index

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## Abstract

A large number of genetic loci are associated with adult body mass index. However, the genetics of childhood body mass index are largely unknown. We performed a meta-analysis of genome-wide association studies of childhood body mass index, using sex- and age-adjusted standard deviation scores. We included 35 668 children from 20 studies in the discovery phase and 11 873 children from 13 studies in the replication phase. In total, 15 loci reached genome-wide significance ( $P$ -value  $< 5 \times 10^{-8}$ ) in the joint discovery and replication analysis, of which 12 are previously identified loci in or close to *ADCY3*, *GNPDA2*, *TMEM18*, *SEC16B*, *FAIM2*, *FTO*, *TFAP2B*, *TNNI3K*, *MC4R*, *GPR61*, *LMX1B* and *OLFM4* associated with adult body mass index or childhood obesity. We identified three novel loci: rs13253111 near *ELP3*, rs8092503 near *RAB27B* and rs13387838 near *ADAM23*. Per additional risk allele, body mass index increased 0.04 Standard Deviation Score (SDS) [Standard Error (SE) 0.007], 0.05 SDS (SE 0.008) and 0.14 SDS (SE 0.025), for rs13253111, rs8092503 and rs13387838, respectively. A genetic risk score combining all 15 SNPs

showed that each additional average risk allele was associated with a 0.073 SDS (SE 0.011,  $P$ -value =  $3.12 \times 10^{-10}$ ) increase in childhood body mass index in a population of 1955 children. This risk score explained 2% of the variance in childhood body mass index. This study highlights the shared genetic background between childhood and adult body mass index and adds three novel loci. These loci likely represent age-related differences in strength of the associations with body mass index.

## Introduction

Childhood obesity is an important public health problem with severe consequences, including an increased risk of premature death (1–5). Body mass index (BMI) has a strong genetic component with some reported heritability estimates being over 80% (6–8). Large genome-wide association studies (GWAS) have revealed many genetic loci associated with BMI or adiposity in adults (9–13). However, the genetic loci underlying BMI in children are less well known. The biological background of BMI may differ between children and adults. In addition, it may be that the relative contributions of the same genetic loci differ depending on age, for example due to different gene–environment interactions or body fat distributions (6,14,15). A limited number of loci have been identified to associate with dichotomous definitions of childhood obesity (16–18). Also, the roles of specific known adult loci for BMI, such as *FTO* and *ADCY3*, have been described in children (13,19). The age-specific effects are illustrated by longitudinal studies on the effects of the well-known adult BMI increasing risk allele of *FTO* with BMI throughout childhood (15). It has been reported that the adult BMI increasing risk allele is associated with lower BMI in infancy, an earlier adiposity rebound and a higher BMI from the age of 5 years onwards (14,15,20). To date, studies did not present a large GWAS meta-analysis on the full spectrum of childhood BMI (13,16–19).

To identify genetic loci influencing childhood BMI, we meta-analyzed 20 GWAS with a total of 35 668 children of European ancestry, combining data for around 2.5 million single-nucleotide polymorphisms (SNPs) imputed to the HapMap imputation panel. We used as outcome sex- and age-adjusted standard deviation scores at the oldest age between 2 and 10 years.

## Results

Study characteristics are shown in Supplementary Material, Table S1. Childhood BMI was transformed into sex- and age-adjusted standard deviation scores (SDS) (LMS growth; Pan H, Cole TJ, 2012; <http://www.healthforallchildren.co.uk>).

### Meta-analysis of genome-wide association studies

Inverse-variance weighted fixed-effects meta-analysis revealed 861 SNPs with genome-wide significant or suggestive  $P$ -values ( $<5 \times 10^{-6}$ ). Two SNPs with high heterogeneity were not followed up ( $I^2$  values of 89.4 and 96.0), leaving 859 SNPs representing 43 loci. A locus was defined as a region of 500 kb to either side of the most significant SNP. The Manhattan and Quantile–Quantile plots of the discovery meta-analysis are shown in Figure 1 and Supplementary Material, Figure S1, respectively. The lambda for the discovery meta-analysis was 1.10. LD score regression analysis showed that this slight inflation was mainly due to polygenicity of the trait, rather than to population stratification, cryptic relatedness or other confounding factors (intercept 1.01). Individual study lambdas are shown in Supplementary Material, Table S2. All 43 loci were taken forward for replication in a sample of 11 873 children from 13 studies. Table 1 and Supplementary Material, Tables S3 and S4 show the results of the discovery,

replication and joint analyses for the 43 genome-wide and suggestive loci.

In total, 15 of these reached genome-wide significance in the joint analysis. Twelve out of these 15 had been reported previously for related phenotypes. SNPs in or close to *ADCY3*, *GNPDA2*, *TMEM18*, *SEC16B*, *FAIM2*, *FTO*, *TFAP2B*, *TNNI3K*, *MC4R*, *GPR61*, *LMX1B* and *OLFM4* are associated with adult BMI or childhood obesity (11,13,16). We identified three novel loci: rs13253111 near *ELP3*, rs8092503 near *RAB27B* and rs13387838 near *ADAM23*. Per additional risk allele, BMI increased 0.04 Standard Deviation Score (SDS) [Standard Error (SE) 0.007], 0.05 SDS (SE 0.008) and 0.14 SDS (SE 0.025) for rs13253111, rs8092503 and rs13387838, respectively. Figure 2 and Supplementary Material, Figure S2 show the regional plots and the forest plots, respectively, for these loci.

### Genetic risk score

We combined the 15 identified genome-wide significant SNPs into a genetic risk score that summed the number of BMI-increasing alleles weighted by their betas from the discovery analysis and rescaled to a range of 0 to 30, which is the maximum number of risk alleles. The risk score was associated with childhood BMI ( $P$ -value =  $3.12 \times 10^{-10}$ ) in 1955 children from the PIAMA Study, one of our largest replication cohorts. For each additional average risk allele in the score, childhood BMI increased by 0.073 SDS (SE 0.011) (Fig. 3). This risk score explained 2.0% of the variance in childhood BMI.

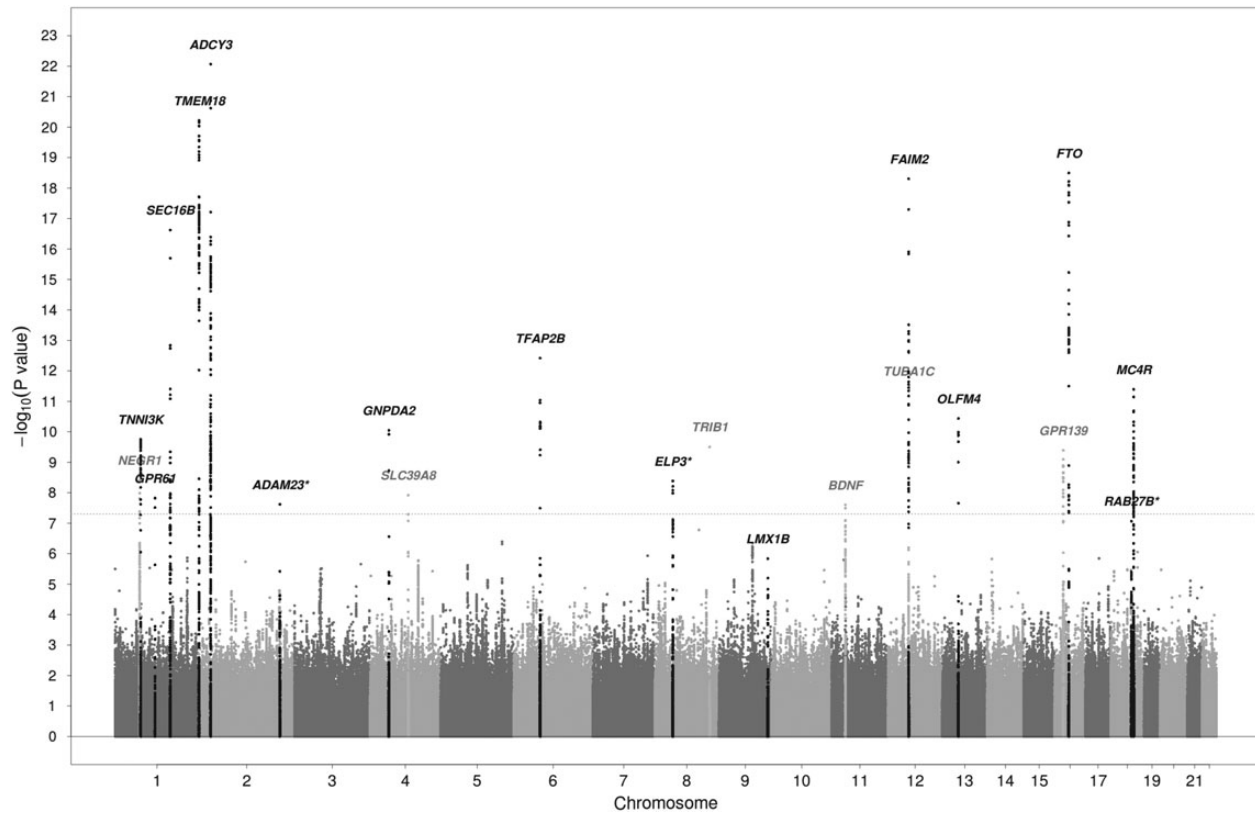
### Associations with adult body mass index and childhood obesity

The genetic correlation between childhood BMI and adult BMI was 0.73. A lookup of the 15 SNPs associated with childhood BMI in a recently published GWAS meta-analysis on adult BMI in >300 000 participants revealed that all SNPs showed evidence for association with adult BMI, with  $P$ -values of 0.005,  $5.76 \times 10^{-5}$  and 0.003 for the novel SNPs rs13253111, rs8092503 and rs13387838, respectively. Also, the direction of the effect estimates for all 15 SNPs was the same in children and adults (Supplementary Material, Table S5) (11). The 15 SNPs found in this study explained 0.94% of the variance in adult BMI in the GIANT consortium (11).

A reverse lookup in our dataset of the 97 known genome-wide significant loci previously reported to be associated with adult BMI showed that 22 out of the 97 loci were significantly associated with childhood BMI, using a Bonferroni-adjusted  $P$ -value cutoff of  $5.2 \times 10^{-4}$  for 97 SNPs. A total of 50 out of the 97 known adult BMI SNPs were nominally associated with childhood BMI ( $P$ -value  $<0.05$ ). The direction of the effect estimates was the same in adults and children for 86 SNPs ( $P$ -value binomial sign test  $<1.0 \times 10^{-4}$ ; Supplementary Material, Table S6).

We looked up the association of the three novel loci in a GWAS meta-analysis of childhood obesity. In this study, childhood obesity cases were defined as having a BMI  $\geq$  95th percentile, whereas childhood normal weight controls were defined as having a BMI  $<$ 50th percentile. This meta-analysis included 22





**Figure 1.** Manhattan plot of results of the discovery meta-analysis of 20 studies. Chromosomes are shown on the x-axis, the  $-\log_{10}$  of the P-value on the y-axis. The gray dotted line represents the genome-wide significance cutoff of  $5 \times 10^{-8}$ . Genes shown in black are the known loci that were significantly associated with childhood BMI in the joint discovery and replication analysis. Genes shown in gray were significant in the discovery, but not in the joint discovery and replication analysis. \* indicates novel loci that were significantly associated with childhood BMI in the joint discovery and replication analysis. See also Table 1.

studies, of which 16 were also included in our current meta-analysis. All three SNPs were associated with childhood obesity (P-values 0.01, 0.005 and  $6.0 \times 10^{-4}$  for rs13253111, rs8092503 and rs13387838, respectively) (16).

### Functional analysis

To explore functionality, we first analyzed whether the 15 identified SNPs affect messenger RNA expression (eQTLs). We analyzed eQTLs from peripheral blood samples from 5311 individuals, which revealed two *cis*-eQTLs [false discovery rate (FDR) P-value  $< 0.05$ ] for rs11676272, the top SNP in one of the previously identified loci (ADCY3). One of these eQTLs was for ADCY3, and one was for DNAJC27 (21). Also, we found a *cis*-eQTL for FAM125B for rs3829849, which is located in LMX1B (Supplementary Material, Table S7). eQTL analysis in adipose tissue, a more specific target tissue in relation to BMI, from 856 healthy female twins in the MuTHER resource in Genevar revealed two significant *cis*-eQTLs (distance to SNP  $< 1$  Mb) for rs11676272, for transcripts of ADCY3 and POMC, with a Bonferroni-corrected P-value of  $< 0.003$  (22,23). The association of rs11676272 with expression of ADCY3 was also validated in a second eQTL analysis in a smaller set of 206 lymphoblastoid cell lines (24). We did not identify eQTLs related to our three novel loci.

Second, we performed functional analyses with the tool Data-Driven Expression Prioritized Integration for Complex Traits (DEPICT) using all SNPs with a P-value  $< 1 \times 10^{-5}$  in the discovery analysis (see Materials and Methods for details) (25). Gene prioritization analysis did not show prioritized genes, nor did the gene

set enrichment analysis reveal evidence for enriched reconstituted gene sets and genes near the associated SNPs were not found to enrich for expression in a panel of 2009 tissue and cell types (FDR  $< 0.05$ ; Supplementary Material, Tables S8a, b and c).

### Discussion

In this GWAS meta-analysis of childhood BMI among  $> 47$  000 children, we identified 15 genome-wide significant loci, of which three loci, rs13253111 near ELP3, rs8092503 near RAB27B and rs13387838 near ADAM23, have not been associated with adiposity-related phenotypes before.

Large GWAS have revealed many genetic loci associated with BMI or adiposity in adults (9–13). A recent meta-analysis in up to 339 224 individuals identified 97 BMI-associated loci, explaining 2.7% of the adult BMI variation. Pathway analyses showed that the central nervous system may play a large role in obesity susceptibility. The number of identified loci associated with BMI or obesity in childhood is scarce. Of the total of 15 loci associated with childhood BMI in the current study, 12 have previously been associated with adiposity outcomes in adults or children. All 12 loci are known to be associated with adult BMI (11). Also, eight loci, including those in or near ADCY3 (annotated to the nearby gene POMC in the previous paper), TMEM18, SEC16B, FAIM2, FTO, TNNI3K, MC4R and OLFM4, have previously been associated with childhood obesity (16). All three novel loci were nominally associated with the more extreme outcome of childhood obesity in a largely overlapping population of child cohorts (16).

**Table 1.** Results of the discovery, replication and joint analyses for 43 loci with P-values <math> < 5 \times 10^{-6}</math> in the discovery phase

SNP	CHR	Position	Nearest gene	EA/Non-EA	EAF <sup>a</sup>	Beta <sup>a</sup>	SE <sup>a</sup>	P-value discovery	P-value replication	P-value joint
rs13130484 <sup>b</sup>	4	44870448	GNPDA2	T/C	0.44	0.067	0.007	$8.94 \times 10^{-11}$	$4.29 \times 10^{-18}$	$1.58 \times 10^{-23}$
rs11676272 <sup>b</sup>	2	24995042	ADCY3	G/A	0.46	0.068	0.007	$8.55 \times 10^{-23}$	0.020	$7.12 \times 10^{-23}$
rs4854349 <sup>b</sup>	2	637861	TMEM18	C/T	0.83	0.090	0.009	$6.00 \times 10^{-21}$	0.005	$5.41 \times 10^{-22}$
rs543874 <sup>b</sup>	1	176156103	SEC16B	G/A	0.20	0.077	0.009	$2.38 \times 10^{-17}$	$8.77 \times 10^{-4}$	$2.20 \times 10^{-19}$
rs7132908 <sup>b</sup>	12	48549415	FAIM2	A/G	0.39	0.066	0.008	$4.99 \times 10^{-19}$	0.043	$1.57 \times 10^{-18}$
rs1421085 <sup>b</sup>	16	52358455	FTO	C/T	0.41	0.059	0.007	$3.20 \times 10^{-19}$	0.654	$4.53 \times 10^{-16}$
rs12429545 <sup>c</sup>	13	53000207	OLFM4	A/G	0.13	0.076	0.010	$3.66 \times 10^{-11}$	$1.01 \times 10^{-4}$	$2.08 \times 10^{-14}$
rs987237 <sup>b</sup>	6	50911009	TFAP2B	G/A	0.19	0.062	0.009	$3.81 \times 10^{-13}$	0.224	$1.80 \times 10^{-12}$
rs12041852 <sup>b</sup>	1	74776088	TNNI3K	G/A	0.46	0.046	0.007	$1.77 \times 10^{-10}$	0.142	$2.28 \times 10^{-10}$
rs6567160 <sup>b</sup>	18	55980115	MC4R	C/T	0.23	0.050	0.008	$4.06 \times 10^{-12}$	0.996	$1.21 \times 10^{-9}$
rs13253111	8	28117893	ELP3	A/G	0.57	0.042	0.007	$4.13 \times 10^{-9}$	0.114	$4.89 \times 10^{-9}$
rs8092503	18	50630485	RAB27B	G/A	0.27	0.045	0.008	$8.55 \times 10^{-8}$	0.034	$8.17 \times 10^{-9}$
rs3829849 <sup>b</sup>	9	128430621	LMX1B	T/C	0.36	0.041	0.007	$1.46 \times 10^{-6}$	0.001	$8.81 \times 10^{-9}$
rs13387838	2	206989692	ADAM23	A/G	0.04	0.139	0.025	$2.40 \times 10^{-8}$	0.306	$2.84 \times 10^{-8}$
rs7550711 <sup>d</sup>	1	109884409	GPR61	T/C	0.04	0.105	0.019	$1.50 \times 10^{-8}$	0.401	$4.52 \times 10^{-8}$
rs17309930 <sup>b</sup>	11	27705069	BDNF	A/C	0.21	0.045	0.009	$2.47 \times 10^{-8}$	0.540	$1.41 \times 10^{-7}$
rs2590942 <sup>b</sup>	1	72657869	NEGR1	T/G	0.82	0.047	0.009	$3.88 \times 10^{-9}$	0.966	$1.91 \times 10^{-7}$
rs13107325 <sup>b</sup>	4	103407732	SLC39A8	T/C	0.07	0.081	0.016	$1.19 \times 10^{-8}$	0.970	$3.79 \times 10^{-7}$
rs10151686 <sup>b</sup>	14	29536217	PRKD1	A/G	0.04	0.096	0.019	$1.50 \times 10^{-6}$	0.109	$6.99 \times 10^{-7}$
rs25832	5	66211438	LOC375449	A/G	0.71	0.039	0.008	$2.41 \times 10^{-6}$	0.177	$1.62 \times 10^{-6}$
rs7869969	9	95257268	FAM120A	G/A	0.33	0.036	0.008	$4.43 \times 10^{-7}$	0.425	$1.68 \times 10^{-6}$
rs11079830 <sup>c</sup>	17	44037629	HOXB6	A/G	0.58	0.034	0.007	$1.43 \times 10^{-6}$	0.254	$1.98 \times 10^{-6}$
rs4569924	5	153520218	GALNT10	T/C	0.43	0.032	0.007	$4.06 \times 10^{-7}$	0.823	$3.48 \times 10^{-6}$
rs8046312 <sup>b</sup>	16	19886835	GPR139	A/C	0.81	0.042	0.009	$4.06 \times 10^{-10}$	0.185	$3.97 \times 10^{-6}$
rs1838856	2	113822060	PAX8	A/C	0.46	0.034	0.008	$1.85 \times 10^{-6}$	0.588	$1.47 \times 10^{-5}$
rs633143	1	179716108	CACNA1E	T/C	0.14	0.044	0.011	$3.40 \times 10^{-6}$	0.648	$2.44 \times 10^{-5}$
rs4923207	11	24713901	LUZP2	T/G	0.79	0.039	0.010	$1.60 \times 10^{-6}$	0.834	$3.52 \times 10^{-5}$
rs6971577	7	140350204	MRPS33	C/G	0.78	0.036	0.009	$1.17 \times 10^{-6}$	0.690	$6.80 \times 10^{-5}$
rs10866069	3	64366964	ADAMTS9	T/C	0.17	0.041	0.011	$3.05 \times 10^{-6}$	0.687	$8.43 \times 10^{-5}$
rs12457682	18	7216505	LAMA1	C/A	0.23	0.035	0.009	$3.81 \times 10^{-6}$	0.942	$9.01 \times 10^{-5}$
rs11165675 <sup>b</sup>	1	96812556	PTBP2	A/G	0.27	0.031	0.008	$2.93 \times 10^{-6}$	0.520	$1.01 \times 10^{-4}$
rs12096993	1	217931859	SLC30A10	T/C	0.27	0.031	0.008	$1.38 \times 10^{-6}$	0.583	$1.02 \times 10^{-4}$
rs760931	1	1637388	CDC2L1	C/G	0.93	0.103	0.027	$3.15 \times 10^{-6}$	0.160	$1.27 \times 10^{-4}$
rs2968990	4	131098524	C4orf33	C/T	0.37	0.028	0.007	$1.67 \times 10^{-6}$	0.487	$1.42 \times 10^{-4}$
rs1247117	10	120418792	C10orf46	G/A	0.11	0.040	0.011	$3.47 \times 10^{-6}$	0.339	$2.62 \times 10^{-4}$
rs6580706	12	47959818	TUBA1C	C/G	0.34	0.031	0.009	$1.83 \times 10^{-8}$	0.047	$3.68 \times 10^{-4}$
rs8092620	18	41433991	SLC14A2	G/T	0.48	0.024	0.007	$3.31 \times 10^{-6}$	0.199	$7.32 \times 10^{-4}$
rs188584	3	62675007	CADPS	C/A	0.77	0.028	0.008	$3.23 \times 10^{-6}$	0.139	0.001
rs4870949	8	126704776	TRIB1	T/C	0.07	0.164	0.054	$3.13 \times 10^{-10}$	0.061	0.002
rs1573972	4	171559399	AADAT	C/T	0.19	0.030	0.013	$3.78 \times 10^{-6}$	0.232	0.020
rs214821	20	2258291	TGM3	T/C	0.02	0.479	0.208	$3.38 \times 10^{-6}$	0.575	0.021
rs8084077	18	49532928	DCC	T/C	0.73	0.014	0.008	$3.47 \times 10^{-6}$	$2.72 \times 10^{-5}$	0.062
rs3845265	18	63690108	DSEL	G/A	0.71	0.013	0.008	$8.86 \times 10^{-7}$	$1.44 \times 10^{-6}$	0.087

CHR, chromosome; EA, effect allele; EAF, effect allele frequency; SE, standard error.

Bolded P-values indicate genome-wide significance in the joint analysis.

<sup>a</sup>From joint analysis.

<sup>b</sup>Locus previously reported in Ref. (11).

<sup>c</sup>Locus previously reported in Refs. (11,16).

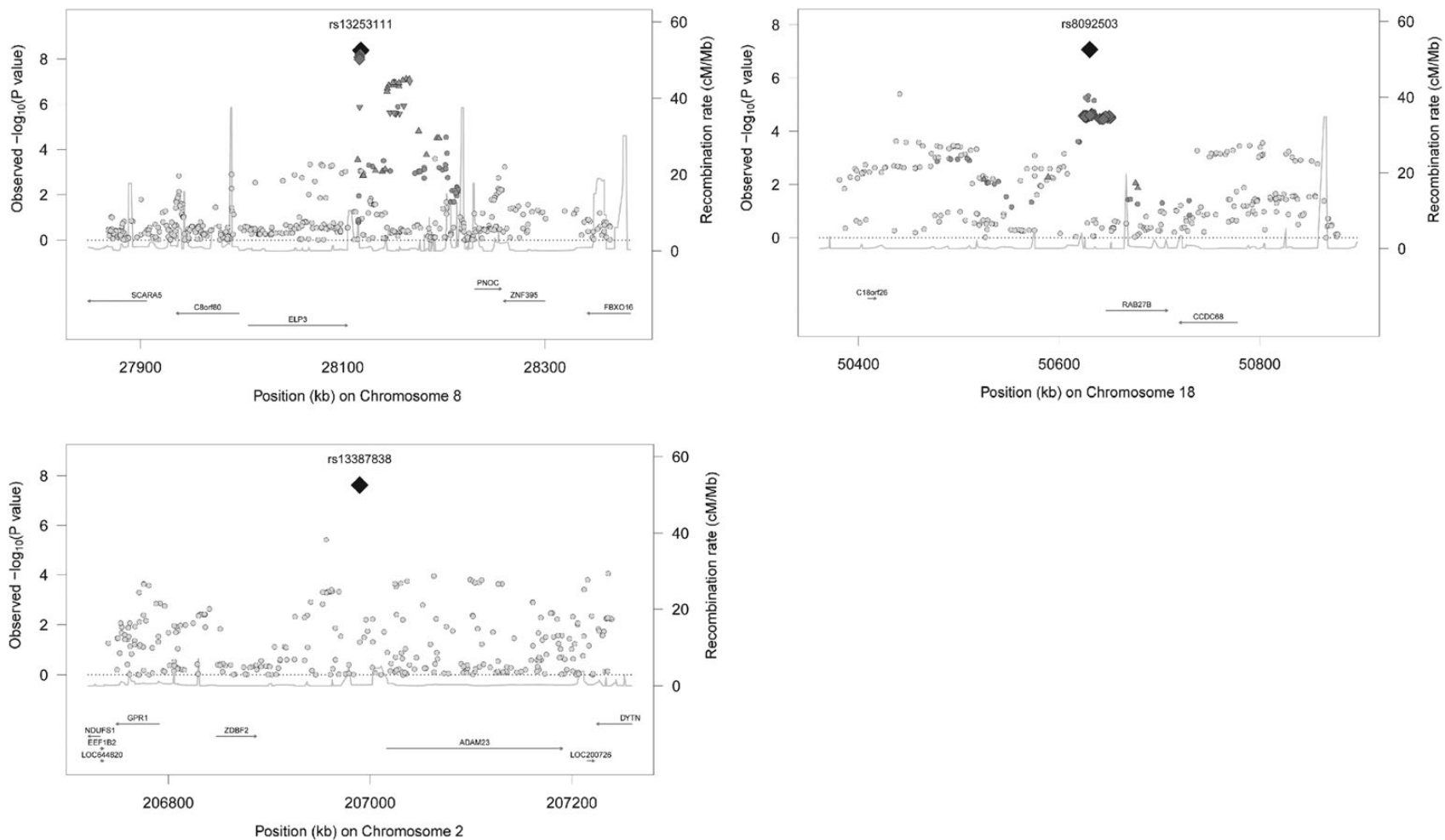
<sup>d</sup>Locus previously reported in Refs. (9,11).

A recent meta-analysis of two studies showed that the known loci *FTO*, *MC4R*, *ADCY3*, *OLFM4* are associated with BMI trajectories in childhood (26). Their findings also suggested that a locus annotated to *FAM120A* influences childhood BMI, which could not be replicated in the current study. The lead SNP in this locus, rs944990, had a P-value of  $1.61 \times 10^{-5}$  in the current analysis. These findings suggest that the overlap between the genetic background of childhood and adult BMI is relatively large, but not complete.

rs7550711 represents one of the 12 identified loci known to be associated with BMI or obesity in adults and children. rs7550711 is a proxy for rs17024258 and rs17024393 ( $R^2$  0.8 with both SNPs),

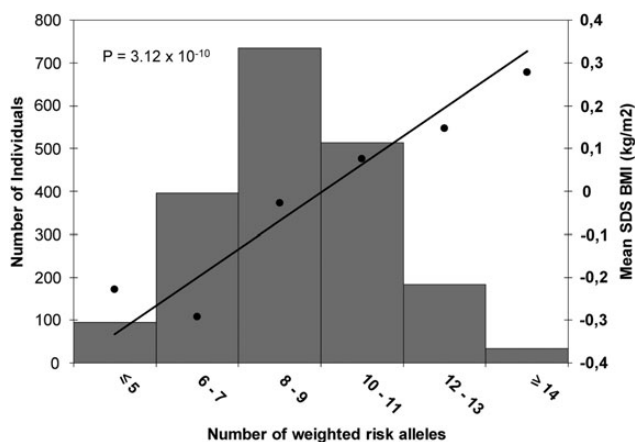
which have previously been associated with adult obesity and BMI, respectively, and annotated with the *GNAT2* gene. However, our proxy resides in *GPR61*, G protein-coupled receptor 61, the biology of which may be more relevant to BMI. *Gpr61*-deficient mice are obese and have hyperphagia, suggesting the role of *Gpr61* in food intake regulation (27). Further studies, including expression studies in relevant human tissues, are needed to establish the causal genes underlying this association.

We identified three loci, rs13253111 near *ELP3*, rs8092503 near *RAB27B* and rs13387838 near *ADAM23*, which have not been associated with adiposity-related phenotypes before in adulthood or childhood. The nearest genes to the novel loci have varying



**Figure 2.** Regional plots of the three novel loci for childhood BMI. On the x-axis, the position of SNPs on the chromosome is shown. On the left y-axis is the  $-\log_{10}$  of the P-values from the discovery analysis, on the right y-axis is the estimated recombination rate (from HapMap), shown by the light blue line in the figure. The named SNP is the most significant SNP in the locus from the discovery meta-analysis. The linkage disequilibrium of all SNPs with the most significant SNP is shown by the symbols, with dark gray diamonds indicating an  $R^2$  of  $\geq 0.8$ , inverted dark gray triangles indicating an  $R^2$  of 0.6–0.8, dark gray circles indicating an  $R^2$  of 0.2–0.4 and light gray circles indicating an  $R^2$  of 0–0.2. Genes (from HapMap release 22) are plotted below the x-axis.





**Figure 3.** Association of the weighted risk score with BMI. Along the x-axis, categories of the weighted risk score are presented, the mean standard deviation score (SDS)-BMI per group is shown on the right y-axis, with the line representing the regression of the mean SDS-BMI values on the categories of the weighted risk score. The left y-axis represents the number of children in each risk score category, shown in the histogram. The P-value is derived from the analysis of the continuous risk score. Analysis was performed in the PIAMA Study ( $n = 1955$ ).

functions. *ELP3*, Elongator Acetyltransferase Complex, subunit 3, has a potential role in the migration of cortical projection neurons and in paternal demethylation after fertilization in mice (28–30). *RAB27B*, RAS-associated protein *RAB27B*, encodes a membrane-bound protein with a role in secretory vesicle fusion and trafficking. It has been associated with pituitary hormone secretion, regulation of exocytosis of digestive enzyme containing granules from pancreatic acinar cells and with gastric acid secretion (31–33). Expression of *ADAM23*, A Disintegrin And Metalloproteinase Domain 23, may influence tumor progression and brain development (34,35). It has also been described to be expressed in mouse adipose tissue and to have a potential role in adipogenesis *in vitro* (36).

Two of our novel loci, rs13253111 near *ELP3* and rs13387838 near *ADAM23*, are close to rs4319045 and rs972540, respectively. Both these SNPs were reported as subthreshold results in the GWAS meta-analysis on adult BMI (11). However, the linkage disequilibrium between the SNPs in both pairs is very low ( $R^2 \leq 0.1$  for both) suggesting that these SNPs may represent different signals. It is important to note that, although both SNPs reached genome-wide significance in the joint discovery and replication analysis, the P-values in the replication stage were non-significant. This lack of significance may be due to the smaller sample size and lower power. Also, the joint P-values were slightly higher than the discovery P-values. Heterogeneity between the discovery and the replication stages was low to moderate, with  $I^2$  values of 61.1 and 27.8 for rs13253111 and rs13387838, respectively (P-values > 0.1 for both). These two signals need to be interpreted with some caution and further studies with larger sample sizes are needed to fully clarify the role of variants in these regions in the physiology of BMI.

Functional analysis showed *cis*-eQTLs for the lead SNPs in two of the known loci. rs11676272 was associated with eQTLs in *ADCY3* and *DNAJC27*, also known as *RBJ*. Both these genes have been associated with adult BMI before and the association of rs11676272 with expression of *ADCY3* has been previously described in childhood BMI (11,13,37). rs3829849 was associated with an eQTL in *FAM125B*, or *MVB12B*, multivesicular body subunit 12B. This gene encodes a component of ESCRT-I (endosomal

sorting complex required for transport 1), a plasma membrane complex with a role in vesicular trafficking was recently described to be associated with intra-ocular pressure (38). However, the LD of our SNPs with the peak markers for the *DNAJC27* ( $R^2 0.11$ ) and the *FAM125B* ( $R^2 0.03$ ) transcripts was low. Our analysis using DEPICT did not show enriched gene sets. This may reflect the relatively limited sample size in our analysis. Further studies are needed to determine the potential functional impact of all SNPs associated with childhood BMI.

Using LD score regression analysis with our meta-analysis results and the results from the recently published GWAS meta-analysis on adult BMI as input, we found that the genetic correlation between childhood and adult BMI was high (11,39). The variance in adult BMI explained by the 15 SNPs identified in this study was lower than in children. The novel SNPs reported in this study may represent loci that specifically influence childhood BMI, but not adult BMI. An alternative explanation is that the effect sizes of these loci may be larger in children than in adults, which may explain the discovery in childhood studies but not in adult studies (11). The large overlap between childhood and adult BMI loci suggests that many of these loci may not represent childhood-specific effects, but rather involvement of the same loci with differential effect sizes at different ages. Age-specific effects of genetic variants associated with BMI in children have been described for the *FTO* locus (15). However, longitudinal studies with multiple measurements of BMI are needed to confirm and quantify such varying effects with age. In discussing the genetic overlap between childhood and adult BMI, it needs to be noted that, because of the differences in body proportions and body fat distribution, childhood BMI may be a different phenotype compared with adult BMI. Our outcome was the conventional measure of BMI calculated as weight/height<sup>2</sup>. Especially in early childhood, higher orders of magnitude for height may be more appropriate. Results from a previously published GWAS study on childhood BMI in two of the cohorts included in the current meta-analysis suggest that the results for SNPs close to *ADCY3* are different when higher orders of magnitude for height are being used (37). Further studies are needed to identify loci related to more specific and directly assessed measures of adiposity and body fat distribution in young children.

In conclusion, we identified 15 loci associated with childhood BMI, of which three are novel. Our results highlight a considerable shared genetic background between childhood and adult BMI. The novel BMI-related loci may reflect childhood-specific genetic associations or differences in strength of associations between age groups.

## Materials and Methods

### Study populations

Characteristics of each discovery and replication study population can be found in Supplementary Material, Table S1 and Methods. The discovery analysis included 20 studies with an age range from 3 to 10 years: the Avon Longitudinal Study of Parents and Children (ALSPAC, 6887 children), the Children's Hospital of Philadelphia (CHOP, 2456 children), the Copenhagen Studies on Asthma in Childhood 2000 birth cohort (COPSAC2000, 309 children), the Danish National Birth Cohort (DNBC, 1020 children), the Generation R Study (GenerationR, 2226 children), the GOYA Study (GOYA, 199 children), the Helsinki Birth Cohort Study (HBSC, 1674 children), the Infancia y Medio Ambiente Project (INMA, 756 children), the Leipzig study (Leipzig, 555 children), the Lifestyle—Immune System—Allergy Study plus German

Infant Study on the influence of Nutrition Intervention (LISA + GINI, 1147 children), the Manchester Asthma and Allergy Study (MAAS, 801 children), the Norwegian Mother and Child Cohort Study (MoBa, 126 children), the Northern Finland Birth Cohort 1966 (NFBC 1966, 3948 children), the Northern Finland Birth Cohort 1986 (NFBC 1986, 4000 children), the Netherlands Twin Register (NTR, 1810 children), the Physical Activity and Nutrition in Children Study (PANIC, 423 children), the Western Australian Pregnancy Cohort (Raine) Study (Raine, 1458 children), the Special Turku coronary Risk factor Intervention Project (STRIP, 569 children), the Young Finns Study (YFS, 1134 children), the British 1958 Birth Cohort Study, with two subcohorts that were entered into the meta-analysis separately (1958BC-T1DGC, 1974 children, and 1958BC-WTCCC2, 2196 children).

We included 13 replication studies. Eleven of these were cohort studies: 574 children from the Copenhagen Studies on Asthma in Childhood 2010 birth cohort (COPSAC2010), 676 additional children from the DNBC, 386 additional children from LISA + GINI, 3152 children from the TEDS Study, 1955 children from the Prevention and Incidence of Asthma and Mite Allergy birth cohort study (PIAMA), 1665 children from the BREATHE Study, 447 children from the Bone Mineral Density in Childhood Study (BMDCS), 200 children from the TEENs of Attica: Genes and Environment (TEENAGE) study, additional imputed data on 857 children from the Leipzig Study, 480 additional children from PANIC and additional imputed data for 569 children from STRIP. We also included two obesity case-control studies in the replication: the Danish Childhood Obesity Biobank (306 cases, 158 controls) and the French Young Study (304 cases, 144 controls). In the BREATHE Study, information was available about six SNPs only (rs8046312, rs12429545, rs13130484, rs3845265, rs543874, rs8084077).

All included children were of European ethnic origin. Sex- and age-adjusted standard deviation scores were created for BMI at the latest time point (oldest age, if multiple measurements existed) between 2 and 10 years using the same software across all studies (LMS growth; Pan H, Cole TJ, 2012; <http://www.healthforallchildren.co.uk>). Syndromic cases of obesity and children of non-European ethnic origin were excluded. In the case of twin pairs, only one twin was included, either randomly or based on genotyping or imputation quality.

### Statistical approach

Cohort-specific genome-wide association analyses were first run in the discovery cohorts, using high-density Illumina or Affymetrix SNP arrays, followed by imputation to the HapMap CEU release 22 imputation panel. The MAAS study imputed to the combined 1000 Genomes (1000G) Pilot + HapMap 3 (release June 2010/February 2009) panel. Before imputation, studies applied study-specific quality filters on samples and SNP call rate, minor allele frequency and Hardy-Weinberg disequilibrium (see Supplementary Material, Table S1 for details). Leipzig (discovery sample), NFBC1986, STRIP (discovery sample) and PANIC (discovery sample) contributed unimputed data from the Metachip. Linear regression models assuming an additive genetic model were run in each study, to assess the association of each SNP with SDS-BMI, adjusting for principal components if this was deemed needed in the individual studies. As SDS-BMI is age and sex specific, no further adjustments were made. Before the meta-analysis, we applied quality filters to each study, filtering out SNPs with a minor allele frequency below 1% and SNPs with poor imputation quality (MACH  $r^2_{\text{hat}} \leq 0.3$ , IMPUTE  $\text{proper\_info} \leq 0.4$  or  $\text{info} \leq 0.4$ ). For studies contributing unimputed metabochip data to the discovery analysis, we excluded SNPs with a SNP call rate of <0.95

or with a Hardy-Weinberg Equilibrium  $P$ -value of  $\leq 0.00001$ . We performed fixed-effects inverse-variance weighted meta-analysis of all discovery samples using Metal (40). Genomic control was applied to every study before the meta-analysis. Individual study lambdas ranged from 0.985 to 1.077 (Supplementary Material, Table S2). The lambda of the discovery meta-analysis was 1.10. After the meta-analysis, SNPs for which information was available in only one study were removed.

The final dataset consisted of 2 499 691 autosomal SNPs. The most significant SNP for each of 43 genome-wide significant or suggestive loci ( $P$ -value  $< 5 \times 10^{-6}$ ) was taken forward for replication in 13 replication cohorts. A locus was defined as a region 500 kb to either side of the most significant SNP. All replication cohorts had *in silico* data available. One of them only had non-imputed data (BREATHE), two (TEENAGE and TEDS) had data imputed to HapMap release 22, one cohort (PANIC) used exome chip data and the other nine performed imputation to 1000G. The replication samples of the STRIP and Leipzig studies only contributed 20 and 21 imputed SNPs, respectively, as the unimputed SNPs were part of the discovery analysis. Fixed-effects inverse-variance meta-analysis was performed for these 43 SNPs combining the discovery samples and all replication samples, giving a joint analysis beta, standard error and  $P$ -value (Table 1 and Supplementary Material, Table S2).

### Sensitivity analyses

Allele frequency differences between the discovery and the replication samples were small and stayed within a range of seven percentage points for all SNPs, except for rs1573972, which had a minor allele frequency of 9% in the discovery analysis and 28% in the replication analysis. This was likely due to the inclusion of one study (MAAS) that had imputed to the combined HapMap + 1000G panel, whereas all other studies with imputed data had imputed to HapMap. To increase homogeneity, we performed several sensitivity analyses. First, we reran the discovery meta-analysis excluding the MAAS study. This analysis did not materially change our findings, with one additional SNP (rs10055577) reaching the subthreshold level of significance ( $P$ -value =  $1.10 \times 10^{-6}$ ) and five SNPs (rs4870949, rs1838856, rs633143, rs10866069 and rs1573972) losing significance. None of these five SNPs had replicated in the primary analysis. Second, we reran the replication and joint meta-analysis including only those cohorts that imputed to 1000G. Results of this analysis were very similar to the primary analysis, with two additional replicated SNPs, rs17309930 near *BDNF* and rs13107325 in *SLC39A8*. Both of these are known loci for adult BMI (11,13). Third, we reran the replication including only the HapMap-imputed and unimputed studies (TEDS, TEENAGE and BREATHE). The results were very similar to those using all studies, with rs4870949 and rs2590942 now passing the significance threshold and rs8092503 and rs3829849 now just above it (results not shown). rs1573972 was not replicated in any of the analyses. As results of the third and fourth sensitivity analyses were very similar to those including all replication cohorts, we used the latter as our main analysis for reasons of power.

### Genetic risk score and percentage of variance explained

A weighted risk score was computed as the sum of the number of SDS-BMI-increasing alleles (dosage) weighted by the effect sizes from the discovery meta-analysis. Then, the score was rescaled to range from zero to the maximum number of SDS-BMI-increasing alleles (30 alleles for 15 SNPs) and rounded to the nearest integer. The association of the risk score with

SDS-BMI was assessed in one of the largest replication cohorts (PIAMA,  $N = 1955$ ) by running a linear regression model. The variance in SDS-BMI explained by the risk score was estimated by the unadjusted  $R^2$  of this model. The percentage of variance in adult BMI explained by the 15 SNPs was calculated using the published data from the recently published large meta-analysis of GWAS studies on adult BMI (11). For each SNP, the variance explained was calculated as:  $2 \times (\text{adult effect size}^2) \times \text{MAF} \times (1 - \text{MAF})$ , and these variances were then summed to give the total percentage of variance in adult BMI explained by the 15 SNPs (11,41).

### LD score regression

LD score regression was used with the standard settings (39). Changing the minor allele frequency filter from 0 to 0.05 did not change the results. Therefore, we report the results of the unfiltered analysis only.

### eQTL analysis

eQTL analysis was conducted using the most significant SNP from each of the 15 genome-wide significant loci from the joint analysis. There was no linkage disequilibrium between these SNPs. First, we assessed whether the top SNPs or their proxies, identified on the basis of  $R^2 > 0.7$ , were associated with gene expression in whole-blood cells in a sample of 5311 individuals (21). Expression in this dataset was assessed using Illumina Whole-Genome Expression BeadChips (HumanHT-12). eQTLs were deemed *cis* when the distance between the SNP chromosomal position and the probe midpoint was  $< 250$  kb. eQTLs were mapped using Spearman's rank correlation, using imputation dosage values as genotypes. An FDR  $P$ -value of  $< 0.05$  was considered significant. Second, the 15 SNPs were introduced to the online eQTL database Genevar ([www.sanger.ac.uk/resources/software/genevar](http://www.sanger.ac.uk/resources/software/genevar)) to explore their associations with expression transcripts of genes in proximity ( $< 1$  Mb distance) to the SNP in adipose tissue from 856 healthy female twins of the MuTHER resource (22,23). We used Bonferroni correction for the significance threshold ( $P$ -value  $< 0.003$ ).

### Data-driven Expression Prioritized Integration for Complex Traits

DEPICT was run using SNPs with a  $P$ -value of  $< 10^{-5}$  yielding 56 independent DEPICT loci comprising 100 genes (42). DEPICT was run using default settings, that is using 500 permutations for bias adjustment, 20 replications for FDR estimation, normalized expression data from 77 840 Affymetrix microarrays for gene set reconstitution [see Ref. (43) for details], 14 461 reconstituted gene sets for gene set enrichment analysis and testing 209 tissue/cell types assembled from 37 427 Affymetrix U133 Plus 2.0 Array samples for enrichment in tissue/cell type expression (42).

### Supplementary Material

Supplementary material is available at HMG online.

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#### ALSPAC Study

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scientists, volunteers, managers, receptionists and nurses. The UK Medical Research Council and the Wellcome Trust (Grant ref: 102215/2/13/2) and the University of Bristol provide core support for ALSPAC. ALSPAC GWAS data were generated by Sample Logistics and Genotyping Facilities at the Wellcome Trust Sanger Institute and LabCorp (Laboratory Corporation of America) supported by 23 and Me. The MRC IEU is supported by the Medical Research Council and the University of Bristol (grant code MC\_UU\_12013/1-9).

#### 1958BC-T1DGC and 1958CB-WTCCC

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#### CHOP

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