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Fishing-induced versus natural selection in different brown trout (*Salmo trutta*) strains

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26 **Abstract**

27 Wild, adfluvial brown trout (*Salmo trutta*) are iconic targets in recreational fisheries but also
28 endangered in many native locations. We compared how fishing and natural selection affect
29 the fitness-proxies of brown trout from two pure angling-selected strains and experimental
30 crosses between an adfluvial, hatchery-bred strain and three wild, resident strains. We exposed
31 age 1+ parr to predation risk under controlled conditions where their behaviour was monitored
32 with PIT-telemetry, and stocked age 2+ fish in two natural lakes for experimental fishing.
33 Predation mortality (16% of the fish) was negatively size-dependent, while capture probability,
34 also reflecting survival, in the lakes (38.9% of the fish) was positively length- and condition-
35 dependent. Angling-induced selection against low boldness and slow growth rates relative to
36 gillnet fishing indicated gear-dependent potential for fisheries-induced evolution in behaviours
37 and life-histories. Offspring of wild, resident fish showed slower growth rates than the
38 crossbred strains. Strain effects suggested significant heritable scope for artificial selection on
39 life-history traits and demonstrated that choices of fish supplementation by stocking may
40 override the genetic effects induced by angling.

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42 **Key words:** fisheries-induced selection, angling, gillnet fishing, predation, personality

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51 **Introduction**

52 Recreational fisheries impose significant mortality on freshwater fish populations (Lewin et al.
53 2006). To compensate for fisheries harvest and incidental mortality or to offset natural
54 recruitment bottlenecks, many populations are supported or entirely maintained by stocking
55 (Hutchings and Fraser 2008; Lorenzen et al. 2012). Both fishing and hatchery-rearing practices
56 can reduce behavioural variation with significant fitness consequences (Hutchings and Fraser
57 2008; Tsuboi et al. 2016, 2019). For example, high exploration tendency (Härkönen et al. 2014)
58 and activity phenotype (Koeck et al. 2019b) associated with unintended domestication under
59 captive breeding (Mezzera and Largiadér 2001) have been linked to increased vulnerability to
60 angling in salmonids, implying fisheries-induced selection acts on these traits. In general, any
61 individually consistent, repeatable and heritable behavioural (or other) trait may respond to
62 direct or correlated selection (Réale et al. 2007; 2010; Kortet et al. 2014; Ågren et al. 2019).
63 Adaptive behaviours may make individuals less vulnerable to fishing (Uusi-Heikkilä et al.
64 2008; Arlinghaus et al. 2017; Andersen et al. 2018), but limited experimental research exists
65 on multigenerational behavioural effects of fishing in salmonids (Alioravainen et al. 2020;
66 Prokkola et al. 2021). To alleviate the ecological and genetic risks associated with stocking, a
67 combination of management tools enhancing natural reproduction (Aas et al. 2018) and
68 hatchery techniques to restore behavioural and genetic diversity in stocked fish to improve their
69 ability to form naturally reproducing populations (Hughes et al. 2008) are increasingly adopted.

70 For the evaluation of evolutionary impacts of fishing, the strength and direction
71 of selection gradients from natural predation and fishing must be determined (Monk et al.
72 2021). Natural predation is a major survival threat to juvenile salmonids (Hyvärinen and
73 Vehanen 2004). Studies in Northern pike (*Esox lucius*) and Atlantic cod (*Gadus morhua*) have
74 suggested that negatively size-dependent natural selection is a potent counterforce for
75 positively size-selective fishing (Edeline et al. 2007; Olsen and Moland 2011; Monk et al.

76 2021). While angling can be positively size-selective, state-dependent characteristics along
77 with behavioural and physiological processes influence vulnerability to angling, particularly
78 with artificial lures or flies (Lennox et al. 2017), making the interplay between predation and
79 fishing induced selection more complex in recreational fishing than in size-selective
80 commercial fisheries. Whilst the debate over the intrinsic or learned nature of antipredator
81 behaviour continues, most studies agree that hatchery-rearing and unintended domestication
82 are deleterious for natural survival-enhancing behaviours (Johnsson et al. 1996; Álvarez and
83 Nicieza 2003; Tsuboi et al. 2019). The domestication-induced loss of intrinsic antipredator
84 behaviour may be mitigated by crossbreeding hatchery lines with wild-provenance fish
85 (Alioravainen et al. 2018) as long as the wild fish are genetically compatible with the hatchery
86 fish and produce offspring with equivalent or improved fitness compared to the hatchery fish
87 (Houde et al. 2011; Rollinson et al. 2014).

88 Brown trout (*Salmo trutta*) is one of the most widespread and economically valuable
89 salmonids and a common target for intensive recreational fisheries throughout its distribution
90 (Elliot 1989; Wills 2006). It forms resident and migratory (both anadromous and
91 potamodromous) populations (Ferguson et al. 2019), with large-growing, migratory (adfluvial)
92 stocks being the most valued targets in freshwater fisheries but also the most threatened
93 (Syrjänen and Valkeajärvi 2010; Syrjänen et al. 2017). Due to the virtual absence of wild
94 adfluvial individuals available for broodstock renewal in many areas (Syrjänen et al. 2017), it
95 has become relevant to assess the potential of wild resident populations to mitigate
96 domestication-effects in hatchery fish (Kallio-Nyberg et al. 2010). The potential for genetic
97 rescue in re-introduced brown trout populations, which typically have lower fitness than wild
98 populations (Wills 2006; Lorenzen et al. 2012), might be improved by crossbreeding captive-
99 bred adfluvial brown trout with wild resident fish (Alioravainen et al. 2018; Kelly and Phillips

100 2018). Yet, little is known about how such outbreeding would affect the growth rate (Serbezov
101 et al. 2010), survival and fishing vulnerability of the progeny.

102 Here, our primary aim was to examine whether parental vulnerability to angling or
103 outbreeding influence the vulnerability of brown trout offspring to predation under seminatural
104 conditions, or to experimental fishing in small, natural lakes. To understand the mechanisms,
105 the angling-selected and experimentally crossed, common-garden raised study strains were
106 monitored for survival and growth differences throughout the study. Additionally, we
107 examined whether exploration behaviour quantified during predator exposure or size-related
108 traits would explain capture probability by angling or gillnets in natural lakes. Bold, fast
109 growing fish with highly angling-vulnerable parents were predicted to show the highest
110 vulnerability to both predation and fishing (Biro and Post 2008; Philipp et al. 2009; Klefoth
111 2017) and that outbreeding of hatchery-strain fish with wild resident brown trout from nearby
112 sources would increase survival under predation (Alioravainen et al. 2018). Crossbred fish
113 were expected to demonstrate capture and growth rates intermediate to pure hatchery and wild
114 strains in natural lakes but superior survival.

115

116 **Materials and Methods**

117 *Overview of the study design*

118 All experiments were conducted at the Kainuu Fisheries Research Station (National Resources
119 Institute Finland) in Paltamo, Finland. Nine experimental strains of brown trout (Fig. 1.,
120 Supplementary Table S1) were created via artificial breeding in October 2015. Eggs and larval
121 fish were monitored for survival under standard hatchery rearing conditions. 1+ fish, reared in
122 common garden from September 2016, were exposed to predation by Northern pike under
123 continuous passive integrated telemetry (PIT) during summer 2017 (Alioravainen et al. 2018
124 for a similar set-up). In July 2018, predation survivors and additional fish from the same crosses

125 were stocked into two small natural lakes, which were experimentally fished in autumn 2018
126 (age 2+) and autumn 2019 (age 3+) using both angling and gillnets (Fig. 1, Supplementary
127 Table S1). Two gear types were used to determine if angling would select different phenotypes
128 than gillnetting that was considered relatively unselective.

129

130 *Source strains and breeding design*

131 Adult hatchery-bred trout (OUV) and wild-caught resident brown trout from River Vaarainjoki
132 (VAA, 64° 28' 50.510" N/27° 34' 17.340" E) were first divided into high vulnerability to angling
133 (HV, captured at least once) and low vulnerability (LV, not captured) groups by experimental
134 fly fishing in 50 m² – 75 m² outdoor ponds during summer 2015 using two size-assorted pools
135 per origin (Alioravainen et al. 2020, Supplementary Table S1). In addition, wild adult fish from
136 Rivers Tuhkajoki (TUH, 64° 2' 28.337" N/28° 7' 10.099" E) and Pohjanjoki (POH, 64° 17'
137 50.703" N/28° 3' 0.416" E) were used for the production of outbred crosses (Lemopoulos et al.
138 2019a). The OUV hatchery strain was the product of 4-5 generations of captive breeding from
139 adfluvial (migratory) brown trout originally captured in Rivers Varisjoki (~3 km downstream
140 from VAA) and Kongasjoki (parallel in < 1 km distance with VAA; Lemopoulos et al. 2018,
141 2019a). By contrast, the wild strains were assumed to be residents and non-migratory
142 (Lemopoulos et al. 2018). Wild fish were captured via electrofishing (VAA: 28–30 September
143 2010; 15 September–11 October 2011 and 2 October 2012; TUH: 17 September 2013; POH:
144 16–17 September 2015) and held in separate seminatural gravel-bottomed 50 m² riffle-pool
145 ponds prior to breeding.

146 We used a replicated, fully factorial 3 males × 3 females matrix breeding design
147 to create the F₁-generation of the HV and LV selection lines and the three outcrosses with wild
148 VAA, TUH and POH strains (Supplementary Table S1). Parental fish were randomized *a priori*
149 from the HV and LV groups, but due to logistic constraints and non-ripe status of the

150 randomized fish, parents were occasionally replaced with other haphazardly dip-netted fish
 151 from the same group. To minimize the risk of inbreeding, females and males within the OUV
 152 group were taken from different year classes (2008 or 2012). Within the VAA strain, males
 153 were mostly taken from the small size group (mean \pm S.D. body mass 531.9 ± 249.3 g on 25
 154 September 2015; 15 out of 18) and females from the large size group (body mass $1268.5 \pm$
 155 487.6 g on 25 September 2015; 14 out of 18). Due to limited availability of mature females
 156 among the HV fish, two females that had not been captured in experimental angling had to be
 157 included as broodstock in the HV line. This resulted in 88.9% genetic contribution of strictly
 158 HV fish in the VAA HV line. Nine of the crossbred parent fish were HV fish and 15 were LV
 159 fish.

160 Adfluvial OUV and resident VAA strains differed genetically by F_{ST} 0.109 -
 161 0.112, with greater diversity in the OUV versus VAA strain (Lemopoulos et al. 2019a, b). One
 162 VAA male in one replicate matrix of OUV ♀ \times VAA ♂ was later genetically identified as an
 163 OUV fish despite its capture in River Vaarainjoki (i.e. VAA). This was taken into account in
 164 egg mortality calculations by strain but could not be controlled for in the other analyses. TUH
 165 and OUV strains differed by F_{ST} 0.064 - 0.074 (Lemopoulos et al. 2019a). The stocking history
 166 and migratory status of the River Tuhkajoki (TUH) population could not be confirmed, but the
 167 strain was assumed to represent wild resident fish (Lemopoulos et al. 2019a). Due to the lack
 168 of mature females in the original pool of wild-caught fish ($N = 65$), the OUV \times POH strain was
 169 created by crossing hatchery OUV females with wild-caught POH males (Supplementary Table
 170 S1). The original F_{ST} -value between the populations was 0.119 - 0.120 (Lemopoulos et al.
 171 2019a). No known stocking history of the POH or VAA populations suggests that they were
 172 original (Lemopoulos et al. 2019a).

173

174

175 *Rearing of eggs and larvae with monitoring for survival*

176 Fertilized eggs (100 per full-sib family) were incubated in family-specific mesh-bottomed
177 floating circular tubes (diameter 10 cm) placed in 32 0.4 m² flow-through fibreglass tanks (5-
178 6 tubes per tank) with daily observations of egg mortality (total average 25.5%) from 12
179 October 2015 to 17 May 2016. In addition, all the live eggs were counted on 15 February 2016.
180 Fry were pooled by breeding-matrix on 17 May 2016 and reared in replicated (N = 2) 0.4 m²
181 tanks (42 tanks, 225 fish in each, 9-10 fish per tank used for Alioravainen et al. 2020) until
182 PIT-tagging over 18-21 September 2016 (50 haphazardly chosen fish per tank, N = 2105 in
183 total). 12 mm HDX pit-tags (Oregon RFID, Oregon, USA) were inserted into the body cavity
184 through a small scalpel incision under benzocaine (40 mg L⁻¹) anaesthesia. In the pooling of
185 the fry, the aim was to take 25 fry from each family, but due to variability in egg mortality, the
186 effective number of fish per family in each of the two replicates varied from 0 to 75
187 (Supplementary Table S1). PIT-tagged fish were further reared in four 3.2 m² fibreglass tanks
188 under standard feeding regimes with commercial dry feeds (Veronesi Vita 0.5 mm – 1.5 mm,
189 Veronesi, Italy & Hercules 1.7 mm – 2.5 mm, Raisioaqua, Finland) until transfer to two 15 m²
190 indoor tanks on 14 July 2017 and finally to the predation experiment (see below) or a 50 m²
191 outdoor concrete pond (21 and 28 November 2017).

192

193 *Behaviour-tracked predator exposure experiment*

194 To quantify individual boldness, behavioural responses to predator presence and vulnerability
195 to actual predation, we conducted two batches of a 16-day experiment in replicated (N = 8) 50
196 m² seminatural riffle-pool ponds (described in Alioravainen et al. 2018). Each pond was
197 stocked with N = 40 fish (N = 320 fish per round, N = 640 total) consisting of N = 72 fish per
198 strain except for the OUV × POH cross (N = 64, Supplementary Table S2 for details). Fish
199 from each strain were equally divided between the eight ponds. The ponds were equipped with

200 plastic grids (45 mm squares) placed in half-tube chutes between the gravel-bottomed stream-
201 like riffle area (10.9 m², 1.5 m wide, water depth 0.2-0.5 m) and the deeper circular inner
202 concrete-bottomed pool area with water outlet (water depth ~0.8 m). Randomised half of the
203 ponds were stocked with Northern pike that were released to the deep pool section. The grids
204 were permeable to brown trout but not to pike, ensuring the safety of the trout and ethicality of
205 the experiment. Looped PIT-antennae were installed on both sides of the grid to record (9 times
206 per second) each fish passage and the direction of movement (see Alioravainen et al. 2018 for
207 details). The plain text format PIT detection data were recorded using the TIRIS program
208 (Citius solutions Oy, Kajaani, Finland). The recording computers had to be restarted twice
209 during the first replicate batch and once during the second batch. During the offline time (10-
210 15 min), the fish were assumed to have maintained their previous positions.

211 Two pike per predator pond (686.3 ± 44.0 mm, 2048 ± 300 g, mean \pm SD) were
212 transferred to the ponds on 13 July 2017 at 15:00. The original pike were returned to their
213 rearing ponds and replaced with new pike (639.3 ± 80.2 mm, 1905 ± 686 g, mean \pm SD) for
214 the second batch on 7 August 2017. The pike were fed once per week with dead brown trout.
215 The experimental trout were not fed during the experiment but were supplied by natural drift
216 and benthic macroinvertebrate fauna occurring in the riffle section of the ponds (Rodewald et
217 al. 2011). The trout were released to the riffle section of the ponds so that the grids dividing
218 the safe (outer riffle) and risky (deep center) areas were blocked by metal gates on 13 July 2017
219 between 16:00 and 22:00 (escapement recorded by PIT-data). Behavioural observations started
220 on 17 July 2017 at 10:13:00 for each pond, and the gates were removed within the next 9
221 minutes with the exact time recorded for each pond. PIT-data were collected continuously until
222 2 August 2017 at 08:00:00 after which the water level was lowered, remaining fish were dip-
223 netted, measured for size and mass under benzocaine anaesthesia and transferred back to a 50
224 m² rearing pond. New trout were introduced to the ponds on 3 August 2017, and the gates were

225 lifted on 7 August 2017. PIT-data were again collected for 16 days from 7 August 2017 at
226 10:13:00 to 23 August 2017 at 08:00:00. The barrier to prevent movement of fish between the
227 predator and safe sides failed in one of the control ponds before the start of the trial, and though
228 it was repaired, this pond was assigned a third category of predator treatment to account for the
229 disturbances.

230 All ponds were thoroughly scanned with portable PIT-readers on 24 August
231 2017 and any retrieved PIT-tags were recorded as consumed by pike when found within the
232 predator areas. No accidental tag loss was observed during the project. Each pike was also
233 carefully PIT-scanned at removal. All missing fish were assumed to be eaten by pike, but it is
234 possible that some fish were lost to birds or mammalian predators as the shades set on the riffle
235 sections did not prevent terrestrial predators from accessing the ponds. Three fish were not
236 recovered for final length and mass measurements but were later found alive and recorded as
237 having survived. After the experiments, all fish were placed in a 50 m² concrete rearing pond
238 together with other trout from the same crosses until stocking (Supplementary Table S2).

239 The PIT-data were analysed using custom codes in AV Bio-Statistics 5.2
240 software (by A.V.) for the response variables used in the final statistical analyses: 1) Entering
241 the predator zone (or not) as binary variable, 2) Time to enter the predator zone for the first
242 time (in minutes, maximum time if not entering), 3) Total time (min) spent in the predator
243 section of the pond (defined only for the fish recovered alive), and 4) Number of antenna
244 switches (requiring a minimum of 300 seconds between detections, defined only for the fish
245 recovered alive). All variables were assumed to indicate boldness in the predator ponds and
246 exploration tendency more generally in all ponds.

247

248 **Stocking experiment**

249 All stocked fish (N = 885 with all data, Supplementary Table S3, Supplementary Fig. S1)
 250 originated from the same breeding set-up but had varying history in behavioural tests. In
 251 addition to fish reared only for stocking (N = 325) and taken from the original 50 m² concrete
 252 rearing pond, 560 fish held in another similar rearing pond had previous history in behavioural
 253 experiments: N = 21 fish had been used in a behavioural study following group-behaviour in a
 254 short-term flume experiment and in a long-term seminatural migration experiment (April 2017
 255 – June 2018, Alioravainen et al. 2021), N = 64 fish had been exposed to measurements of
 256 metabolic rate and individual behaviour under olfactory predator cues (June-July 2017,
 257 Prokkola et al. 2021) and N = 475 fish (of which 211 had been exposed to predators) were
 258 included in the predation exposure experiment described above.

259 The 2+ fish were dip-netted from the rearing ponds, anaesthetized with
 260 benzocaine, measured for total length (to the nearest mm) and weighed (to 0.1 g) on 11 July
 261 2018. The availability of OUV × POH hybrids was limited compared to other strains (27 vs.
 262 33-38 fish per strain in Lake Kylmälampi, 44 vs. 56-78 fish per strain in Lake Koukkulampi;
 263 Supplementary Table S3). The fish were transported ~100 km in two 1.5 m³ oxygenated tanks
 264 on a truck to Kuhmo and released at early morning on 17 July 2018 to Lake Kylmälampi
 265 (63°58.55' N/30°13.29' E, 2.66 ha, 243 m above sea level) and Lake Koukkulampi (64°07.54'
 266 N/30°16.00' 10.099" E, 5.71 ha, 247 m above sea level) by carrying them in batches of 30-50
 267 fish in 50 l buckets for the last 100 – 500 m. Lake Koukkulampi received 586 trout (length 233
 268 ± 27.2 mm, mass 136 ± 52.2 g, 101 fish ha⁻¹) and Lake Kylmälampi 310 trout (length 246 ±
 269 28.5 mm, mass 169 ± 78.3 g, 116 fish ha⁻¹). Lake Kylmälampi naturally supported only
 270 ninespine stickleback (*Pungitius pungitius*), and Lake Koukkulampi naturally supported
 271 Eurasian perch (*Perca fluviatilis*) as well as stocked whitefish (*Coregonus lavaretus*).

272 Experimental fishing was conducted in three sessions, 1) August 2018, 2)
 273 October 2018 and 3) August 2019 (Supplementary Table S4). The main fishing methods were

274 fly fishing from a float tube and spincasting from shore with artificial lures. In sessions 2 and
 275 3, gillnets (30 m × 1.5-1.8 m, mesh sizes 20 mm, 25 mm, 27 mm, 30 mm, 35 mm, 40 mm and
 276 Nordic survey multimesh gillnets) were also used. Bycatch of other fishes occurred only in
 277 Lake Koukkulampi and consisted of ~250 Eurasian perch and one whitefish. For fly fishing,
 278 gear in AFTM classes 6-7 were used with sinking lines and barbless flies (woolly bugger and
 279 nymph patterns) in hook sizes 16-6 (two flies in 0.5-1 m intervals). For spincasting, 6-8' light-
 280 action rods with multifilament lines (line nominal diameter 0.10-0.15 mm) and small inline
 281 spinners and spoons with barbless hooks were used. All fish were killed immediately after
 282 landing with a sharp blow to the head. Each fishing day angling efforts were divided into 1 hr
 283 periods after which all captured fish were identified by PIT-tag, measured and weighed. A
 284 subset of fish showing gonad development were assessed visually for sexual maturity and sex,
 285 but the data did not allow proper analysis (77.6% of the dissected 85 fish were assigned as
 286 spawning in the same autumn, 60.4% of mature fish were males, no apparent differences among
 287 the strains).

288

289 *Growth and condition*

290 Instantaneous (specific) growth rate was calculated for each fish in the predation experiment
 291 and for each fish captured during the stocking experiment as:

292

$$293 \quad IGR_{100} = 100 \frac{\ln(M_{stocking}) - \ln(M_{capture})}{days\ between}, \quad (1)$$

294

295 where M is body mass (g). This common metric (Lugert et al. 2016) was chosen despite its
 296 complex unit as it was normally distributed and efficiently controls for different capture dates
 297 among the fish (Pearson's correlation between capture date and IGR, $R = 0.057$, $N = 340$, $P =$
 298 0.297). Condition factor (CF) was calculated for the stocked fish as:

299

$$CF = 100\,000 \frac{M_{stocking}}{L_{stocking}^{3.2752}}, \quad (2)$$

301

302 where L is the fish total length (mm) and the exponent 3.2752 was empirically derived from
303 the pooled size (total length and mass) measurements at stocking using non-linear regression.

304 Mean-standardized selection gradients β_μ were calculated as instructed by Matsumura et al.

305 (2012).

306

307 *Statistical analyses for egg and juvenile survival*

308 To account for potential failure to fertilize all eggs, egg batches with zero survival (N = 6) were
309 excluded from the analysis, and the egg and alevin survival was analysed for two time periods:

310 I: from fertilizations to eyed-egg stage (14 October 2015 to 15 February 2016) and II from
311 eyed-egg stage to the counting of live fry (15 February 2016 to 17 May 2016). Survival

312 proportions were arcsine square root (egg, alevin and fry survival) and logit-transformed
313 (fingerling survival) to meet normality and analysed for strain differences using linear mixed

314 effect models (LME, eggs) and linear models (GLM, larvae). In the LME, male and female
315 parent ID's and their interaction (dropped from the final models due to being non-significant)

316 were included as random factors, and strain, male body mass and female body mass as fixed
317 factors and variables. The GLM was based on the tanks as independent statistical units, and

318 strain was the only fixed factor.

319

320 *Statistical analyses for the predation experiment*

321 Survival likelihood in predator experiment was determined through logistic regression with
322 backward model selection based on log-likelihood ratio tests. Fish strain, predator presence,

323 original length, batch and strain \times original length were first entered as explanatory terms.

324 Behavioural measures were not included as they were confounded with the other factors. The
325 raw survival frequencies were supplementarily tested using a χ^2 test (Preacher 2001). Latency
326 time to enter the predator section was analysed for all the successfully tested fish (N = 600)
327 using Cox regression in the *coxme* package (Therneau et al. 2003) in R 3.5.2 (The R Foundation
328 for Statistical Computing). The initial model included the random effect of pond (ponds of the
329 second batch treated as new ponds), factors strain, predator presence and predator presence \times
330 strain -interaction, and covariates initial fish length and the interaction between length and
331 strain (Alioravainen et al. 2018), but the interaction terms were dropped from the final models
332 (AICs 5844.8, 5834.4 and 5824.9 in the order of simplification). Among the fish that survived
333 the predation experiment, the three response variables (time to enter the predator section, time
334 spent in the predator section and the total count of section switches) were subject to principal
335 component analysis (PCA) to extract the common boldness/exploration-related variation. The
336 extracted PC scores (by regression) were further analysed using LME models with pond \times
337 batch as a random factor, predator presence and strain as fixed effects and initial total length
338 as covariate. The predator \times strain interaction was omitted as it was non-significant, $P = 0.119$.
339 A model without the strain term was used to derive the model residuals (named boldness scores
340 from here on) used to represent the boldness of fish later stocked for fishing experiments. LME
341 was also used to determine if strain explained specific growth rate during the predation
342 experiment among surviving fish.

343

344 *Statistical analyses for the stocking experiment*

345 Fully factorial GLMs with rearing pond, stocking lake and strain as fixed factors were used to
346 examine if the stocked fish differed in size or condition. Levene's tests were used to assess
347 homoscedasticity of variances, but violations were accepted as biologically indicative of
348 different variances among groups. Independent samples *t*-tests were used to compare group

349 means between stocking lakes. Cox regression with lake as stratum (and to confirm results,
350 separately by lake) was used to examine which factors contributed to vulnerability to capture
351 defined by time until capture (days) and final status (captured or not) by 1) all gear, and 2)
352 angling gear only among 1) fish included in the predation experiment, 2) all fish and 3) fish
353 captured by angling or by gillnets, and thus confirmed to be alive. The initial model including
354 strain (with deviance contrasts, i.e. factor level specific differences from the mean effect),
355 rearing pond (capturing use in previous experiments), length at stocking, condition factor at
356 stocking and boldness score was simplified with backward likelihood-ratio model selection.
357 IGR_{100} was included in the model for the fish confirmed to be alive. Logistic regression of
358 capture status was used to determine probability of capture by gillnet in data excluding the fish
359 captured and removed from the lakes by angling before gillnetting. Captures in angling were
360 assumed to primarily reflect differences in vulnerability to angling, while captures in gillnetting
361 were assumed to primarily reflect differences in survival due to the small size of the lakes and
362 large number of gillnets used.

363 Histograms of variables and residuals were visually examined for normality. All
364 analyses were conducted using IBM SPSS statistics 25.0.0.1 (IBM Corporation) unless
365 otherwise stated.

366

367 *Ethics approval*

368 All animal experimentation was conducted under licenses (ESAVI/3443/04.10.07/2015 and
369 ESAVI/3385/2018) from the Animal Experiment Board (ELLA) of Finland. All animal rearing
370 and experimentation complied with the *Canadian Council on Animal Care guidelines*, and
371 Finnish laws and regulations. Introducing a new species to the used natural lakes lacking
372 migration corridors to other waterbodies was approved by ELY-centre of Lapland (permit no.
373 LAPELY 508 / 5711-2018). All other fishing was forbidden in the study lakes by decision of

374 ELY-centre of Lapland (LAPELY 507 / 5710-2018), and all fishing rights to the study lakes
 375 were rented from Metsähallitus, a state-owned organization responsible for the management of
 376 state-owned land and waterbodies in Finland (contract #43245). Fishing against the prevailing
 377 regulations was conducted under permission (LAPELY 166 / 5713-2017) of the ELY-centre of
 378 Lapland. The experiments did not threaten the endangered wild populations of brown trout in
 379 Finland. All collection of wild brown trout for the establishment of the study strains was
 380 conducted under licences from local shareholder associations managing the respective waters
 381 and ELY-centre of Kainuu (dnro 1013 / 5713-2012).

382

383 **Results**

384 *Egg-stage and first-summer survival*

385 Strains did not differ in egg survival until eyed-egg stage ($78.7\% \pm 27.3\%$, mean \pm S.D., LME,
 386 $F_{8,73.89} = 1.509$, $P = 0.169$), but sire mass had a negative effect (LME, estimate -0.00025 , $F_{1,60.44}$
 387 $= 8.538$, $P = 0.005$) and maternal mass a positive effect (LME, estimate 0.00013 , $F_{1,76.52} =$
 388 4.238 , $P = 0.043$) on egg survival. Random sire effect and maternal effect were both statistically
 389 significant (sire variance estimate 0.0581 , Wald's $Z = 4.143$, $P < 0.001$; maternal effect,
 390 variance estimate 0.0358 , Wald's $Z = 3.862$, $P < 0.001$). Eggs hatched during the two first
 391 weeks of April 2016. The survival from eyed-egg stage to fry stage was not explained by any
 392 of the studied factors ($97.3\% \pm 0.075\%$, mean \pm S.D, LME, strain $F_{8,64.61} = 1.102$, $P = 0.374$,
 393 sire mass $F_{1,57.09} = 0.324$, $P = 0.572$, maternal mass $F_{1,60.17} = 3.209$, $P = 0.078$), but maternal
 394 random effect was statistically significant (variance estimate 0.0101 , Wald's $Z = 4.078$, $P <$
 395 0.001). There were no strain-specific differences in the survival of the fingerlings during the
 396 first summer prior to pit-tagging ($92.6\% \pm 0.06\%$, mean \pm S.D., GLM, $F_{8, 757.348} = 1.375$, $P =$
 397 0.243).

398

399 *Survival in predator exposure experiment*

400 Two fish (0.63%) died in the control treatment and 51 fish (15.94%) in the predator treatment
 401 with no frequency differences among the strains ($\chi^2 = 4.64$, $df = 8$, $P = 0.795$, Supplementary
 402 Table S2). Pooled across origins, mortality was 16.7% in the hatchery group, 14.8% in the
 403 outbred group and 18.1% in the wild group. The strains showed statistically significant but
 404 biologically small differences in length at the beginning of the predation exposure experiment
 405 (July-August 2017, ANOVA, $F_{8, 578} = 4.97$, $P < 0.001$, Supplementary Fig. S1a). Pike presence
 406 and original length of the fish were the only statistically significant predictors of survival
 407 (strain: Wald's $\chi^2 = 5.88$, $df = 8$, $P = 0.661$; batch: $\chi^2 = 0.176$, $df = 1$, $P = 0.675$; strain \times length:
 408 $\chi^2 = 7.09$, $df = 8$, $P = 0.527$). Predator presence tripled the risk of mortality ($\exp(B) = 2.901$, χ^2
 409 $= 18.78$, $df = 1$; $P < 0.001$) and large size decreased the risk of mortality 3% per each mm
 410 ($\exp(B) = 0.970$, $\chi^2 = 10.77$, $df = 1$, $P = 0.001$).

411

412 *Predator-induced selection and behavioural change*

413 Pooled over the strains, predation-induced mean-standardized selection gradient for length was
 414 $\beta_\mu = 0.677$ ($S = 1.33$ mm, $\mu = 143.28$ mm, $\sigma_p = 16.78$ mm, Matsumura et al. 2012). Predator
 415 presence did not affect likelihood to enter the deep pool sections ($\chi^2 = 1.503$, $df = 1$, $P = 0.220$;
 416 average entrance with predators at 105.9 h (84.4 % of the fish) and without predators at 78.6 h
 417 (90.0 % of the fish) from the beginning of the experiment). Larger fish were less likely to enter
 418 the predator sections than smaller fish ($\exp(B) = 0.988$, $z = -4.23$, $P < 0.001$), and the length
 419 effect was independent of strain (interaction, $\chi^2 = 6.317$, $df = 8$, $P = 0.612$). The statistically
 420 significant strain effect ($\chi^2 = 23.731$, $df = 8$, $P = 0.003$) indicated that wild-strain fish had
 421 increased tendency to enter the pool section (compared to OUV HV reference, VAA HV:
 422 $\exp(B) = 1.636$, VAA LV: $\exp(B) = 1.987$, $P < 0.01$). This effect was observed both with and
 423 without predators because the predator presence \times strain interaction was non-significant ($\chi^2 =$

424 5.234, $df = 8$, $P = 0.732$). Brown trout spent 45.2 ± 75.4 h in the deep sections without predators
 425 and 55.7 ± 73.5 h with predators.

426 Principal component analysis yielded only one component (from here on
 427 boldness score) explaining 45.9% of variance in time to enter the pike sections (loading -0.572),
 428 total time spent in the pike sections (loading 0.683) and total activity (loading 0.763). Predator
 429 presence decreased boldness score ($F_{2, 12.95} = 5.214$, $P = 0.022$), and the control ponds did not
 430 differ from the control pond with the broken gate ($t_{12.696} = -0.057$, $P = 0.955$). Fish length was
 431 negatively related to boldness ($F_{1, 568.44} = 7.037$, $P = 0.008$, $R = -0.096$, $N = 640$, $P = 0.020$) but
 432 strain did not have a statistically significant effect ($F_{8, 563.56} = 0.364$, $P = 0.939$, Fig. 2a). The
 433 origin effect was not significant even when simplified to hatchery, wild and hybrid groups ($F_{2,$
 434 $572.87} = 1.241$, $P = 0.290$).

435

436 *Growth in predator exposure experiment*

437 The original LME suggested that none of the included explanatory factors or variables were
 438 statistically significant in explaining IGR_{100} , so the model was simplified by dropping off terms
 439 one by one (1. predator presence \times origin, $F_{16, 541.86} = 0.556$, $P = 0.916$; 2. predator presence,
 440 $F_{2, 12.64} = 0.051$, $P = 0.951$; 3. fish length, $F_{1, 567.35} = 0.169$, $P = 0.681$). The final model
 441 including only strain had marginal statistical significance ($F_{8, 558.95} = 1.946$, $P = 0.051$, Fig. 2a).
 442 When the origin effect was compared only between the hatchery, wild and crossbred fish, the
 443 effect was significant ($F_{2, 565.00} = 5.680$, $P = 0.004$) with offspring of wild parents having slower
 444 growth rate ($0.029 \pm 0.084 \times 100 \ln g d^{-1}$, mean + S.E.) than hatchery parents ($0.169 \pm 0.084 \times$
 445 $100 \ln g d^{-1}$) or the crossbred fish ($0.175 \pm 0.079 \times 100 \ln g d^{-1}$).

446

447

448

449 *Capture rates by strain and individual traits*

450 For the length and condition information of the stocked fish, see the statistics and Fig. S1 in
451 Supplementary Materials. In total, 345 trout were captured during the experimental fishing
452 sessions (297 in 2018 and 48 in 2019) resulting in a 38.8% recovery rate (77.9% in Lake
453 Kymänlampi and 18.2% in Lake Koukkulampi, Fig. 2. For lake-specific results see
454 Supplementary Figs. S2 and S3). Fish captured in Lake Kymänlampi were, on average,
455 significantly longer ($t_{342} = -4.901$, $P < 0.001$) and heavier ($t = -6.250$, $df = 206.58$, $P < 0.001$)
456 than fish captured in Lake Koukkulampi (278.6 mm and 228.3 g vs 261.1 mm and 171.4 g,
457 respectively). For strain-specific capture sizes see Supplementary Table S3. The catch per unit
458 effort (CPUE, number of fish per hour of fishing) decreased rapidly over time and was lower
459 in Lake Koukkulampi than in Lake Kymänlampi (Supplementary Table S4). Of the captured
460 fish, 119 were caught with artificial lures, 117 by fly fishing and 109 with gillnets (including
461 12 with Nordic survey gillnets).

462 Condition factor and length at stocking but not strain or boldness score explained
463 vulnerability to capture among the fish included in the predation exposure experiment ($N =$
464 475, of which 163 were captured; Table 1). The model explaining the vulnerability to angling
465 (115 fish out of 475) did not differ qualitatively and produced very similar results (Table 1).
466 When only captured fish were included, boldness score was the only retained variable but was
467 not statistically significant in explaining angling vulnerability (Table 1).

468 Cox regression indicated that strain or previous use in experiments did not
469 explain vulnerability to capture (340 captured fish) among all fish (882 in total; Supplementary
470 Table S5). Strain was dropped from the model even when simplified to three categories of
471 origin ($\chi^2 = 4.266$, $df = 2$, $P = 0.112$). The strain effect was close to significant ($P = 0.096$) in
472 model explaining vulnerability to angling among all stocked fish (Supplementary Table S5),
473 and strain-specific odd ratios revealed that the OUV HV line had the highest probability for

474 becoming captured by angling (Supplementary Table S3). When the strain effect was
475 simplified to three groups, the hatchery strain (OUV) showed increased vulnerability to angling
476 ($1.354 \times$ the probability in crossbred group, 39.3% capture rate) and the wild strain (VAA)
477 decreased vulnerability ($0.749 \times$ the probability in crossbred group, 33.8% capture rate;
478 Supplementary Table S5). The crossed groups had the highest recapture rate (40.7%) but, due
479 to the influence of covariates, intermediate risk to become captured. When the model was
480 limited to fish confirmed to be alive (captured), only length at stocking was retained in the
481 model although its effect was not statistically significant (Supplementary Table S5). When the
482 analysis for the fish confirmed to be alive was conducted separately for the lakes, length
483 appeared to be the main determinant of vulnerability to angling in Lake Koukkulampi and
484 condition factor the main determinant in Lake Kylmälampi (Supplementary Table S5).

485

486 *Fishing-induced selection*

487 Comparison of boldness score, length, condition and IGR_{100} between the angled and non-
488 angled fish among the fish confirmed to be alive revealed statistically significant angling
489 selection on boldness and IGR_{100} (Table 2). Logistic regression fitted on the non-angled fish to
490 predict vulnerability to gillnet fishing (survival from angling and natural mortality), the model
491 selection procedure retained only condition factor ($\exp(B) = 18596.49$, $\chi^2 = 4.73$, $df = 1$, $P =$
492 0.030) and lake effects ($\exp(B) = 0.054$ for Lake Koukkulampi, $\chi^2 = 132.40$, $df = 1$, $P < 0.001$).
493 The final model predicted 86.9% of the fish to the correct capture category. Pooled over the
494 strains, mean-standardized selection gradient for length-at-stocking was $\beta_{\mu} = -1.556$ assuming
495 that all non-captured fish survived, and $\beta_{\mu} = 2.480$ assuming that all non-captured fish died for
496 natural reasons.

497

498

499 *Growth in natural lakes*

500 IGR₁₀₀ among captured fish was negatively influenced by length at stocking (GLM, $F_{1, 327} =$
501 56.81, $P < 0.001$, estimate $-0.0068 \times 100 \ln \text{g day}^{-1} \text{mm}^{-1}$). The strain \times length at stocking
502 interaction was non-significant (GLM, $F_{8, 319} = 1.86$, $P = 0.066$). The final model indicated that
503 there were differences among strains ($F_{8, 327} = 4.48$, $P < 0.001$; Fig. 3c), that fish captured by
504 angling had slower growth rates ($0.016 \pm 0.026 \times 100 \ln \text{g day}^{-1}$, mean \pm S.E.) than fish captured
505 with gillnets ($0.115 \pm 0.037 \times 100 \ln \text{g day}^{-1}$), that fish used in other experiments had faster
506 growth rates ($0.125 \pm 0.028 \times 100 \ln \text{g day}^{-1}$) than untouched fish ($0.005 \pm 0.035 \times 100 \ln \text{g}$
507 day^{-1}) and that fish in Lake Kylmälampi grew faster ($0.371 \pm 0.025 \times 100 \ln \text{g day}^{-1}$) than fish
508 in Lake Koukkulampi ($-0.241 \pm 0.038 \times 100 \ln \text{g day}^{-1}$). Growth rate was not correlated with
509 boldness scores (Pearson's $r = 0.022$, $N = 340$, $P = 0.783$) or time until capture ($r = 0.057$, $N =$
510 340, $P = 0.297$) indicating that differences between fish captured by angling vs gillnets were
511 not explained by the inherently longer times until capture by gillnet (confirmed by including
512 capture date as a covariate in the GLM).

513 When the comparison of growth rates was simplified to hatchery, hybrid and wild
514 fish, the patterns remained qualitatively similar (simple strain effect, $F_{2, 333} = 12.40$, $P < 0.001$;
515 captured via angling, $F_{1, 333} = 4.31$, $P = 0.039$; lake, $F_{1, 333} = 198.16$, $P < 0.001$; background
516 pond, $F_{1, 333} = 9.64$, $P = 0.002$). Crossbred fish had the fastest ($0.116 \pm 0.028 \times 100 \ln \text{g day}^{-1}$,
517 mean \pm S.E.) growth rates but these were not significantly different from the hatchery fish with
518 average IGR₁₀₀ ($0.111 \pm 0.044 \times 100 \ln \text{g day}^{-1}$). Wild fish had the slowest IGR₁₀₀ ($-0.140 \pm$
519 $0.048 \times 100 \ln \text{g day}^{-1}$).

520

521

522

523

524 **Discussion**

525 Crossbreeding did not improve survival during predation exposure, but crossbred strains
526 showed good growth rates with the hatchery female \times wild male crosses demonstrating fastest
527 growth (Fig. 3c). Hatchery origin of brown trout appeared to increase post-release capture rate
528 as both pure hatchery strain fish and crossbred fish showed higher angling capture rates than
529 the pure wild-strain fish in small natural lakes. However, offspring of angling-selected parents
530 did not show any detectable transgenerational responses to fishing. Individual vulnerability to
531 angling was best yet statistically non-significantly explained by boldness score, but compared
532 to fish captured by gillnets, angled fish showed lower boldness scores and slower growth rate
533 suggesting strong fishing-induced selection on these traits. Overall capture rates were
534 explained by length (especially in Lake Koukkulampi with poor recapture rate) and condition
535 factor (especially in Lake Kylmälampi with excellent recapture rate) at stocking suggesting
536 that post-release survival, potentially related to better starvation-resistance of large fish
537 (Anderson 1988) played a role in explaining capture rates while vulnerability to angling could
538 also have been condition- and state-dependent, and potentially mediated by individual
539 behavioural differences (Lennox et al. 2017).

540 Body size was the most important predictor of survival in the predation exposure
541 and capture probability in the natural lakes. Strain-specific variation in growth also
542 demonstrated considerable heritable scope for selection on growth-related traits. Pike predation
543 was negatively size-dependent partly due to size-dependent behaviour (Alioravainen et al.
544 2018), while capture probability in lakes was positively size-dependent and thus supporting the
545 pattern of conflicting natural and fishing-induced selection (Edeline et al. 2007; Olsen and
546 Moland 2011; Monk et al. 2021; but see Edeline et al. 2009; Klefoth 2017). Comparison of the
547 selection pressures suggests that negatively-size-dependent predation mortality is a potent
548 counterforce ($\beta_{\mu} = 0.677$) for positively size-selective fishing ($\beta_{\mu} = -1.017$). However, the total

549 selection gradient for length-at-stocking was -1.556 - 2.480 depending on the unknown survival
550 rate of the uncaptured fish. While fish that displayed high tendency to take risks likely suffered
551 from predation in the predator exposure experiment, angling selected against fish that avoided
552 risks in the predator exposure experiment suggesting that behaviour may play a role in
553 mediating selection but not as simply as often predicted.

554

555 *Potential for assisted genetic rescue*

556 Crossbreeding of hatchery fish with wild fish was expected to improve the survival of the
557 offspring at the potential cost of genetic incompatibility issues (Houde et al. 2011; Neff et al.
558 2011; Frankham 2015). We did not observe mortality differences among the study strains at
559 egg, alevin, fry or fingerling stages suggesting that any potential issues with genetic
560 compatibility at the first generation were marginal at most (see also Ågren et al. 2019).
561 Interestingly, we detected a negatively size-dependent paternal effect on early egg survival,
562 and a less surprising positively size-dependent maternal effect on early egg survival. Brown
563 trout are known to express some parental effects in offspring (Petersson and Järvi 2007), but
564 male size has not previously been found to explain egg survival (Jacob et al. 2007). Because
565 the body size of the sire did not explain survival of the eyed-eggs and alevins, it is likely that
566 the size-dependent sire effect arose via effect on fertilization success in this study and should
567 thus be interpreted with caution. Maternal effects mediated via egg size are known to diminish
568 quickly in hatchery environments (Einum and Fleming 1999), and the maternal size effect was
569 statistically non-significant already on the survival of alevins and fry. However, our
570 experimental design did not allow formal analysis to separate phenotypic maternal and paternal
571 effects from purely genetic effects on offspring performance.

572 Unlike Alioravainen et al. (2018) but supporting the results of Petersson and Järvi
573 (2006), we did not observe any survival benefits from crossbreeding during predator exposure.

574 Surprisingly, the fish did not appear to demonstrate intrinsic predator avoidance as nearly all
575 fish entered the risky pool sections both with and without predators. The offspring of wild fish
576 were the most likely to enter the risky predation section, potentially due to size-dependent
577 within-group dominance behaviours. Although this effect was somewhat bound to the faster
578 pool side entrance of the wild offspring in the control ponds, the data did not support earlier
579 findings of wild or hatchery \times wild crosses being less bold than pure hatchery strain fish
580 (Alioravainen et al. 2018; Ågren et al. 2019; Prokkola et al. 2021). Instead, fish reduced their
581 demonstrated boldness, most likely through learning in response to predator presence
582 independent of strain. Due to the large size of the pike, negatively size-dependent mortality
583 cannot be explained by gape-limitation but may at least partly relate to faster entrances to the
584 risky areas by smaller brown trout (Alioravainen et al. 2018).

585 Post-release survival as estimated by gillnet capture rate in the natural lakes did
586 not differ between crossbred and purebred groups, but vulnerability to angling appeared to be
587 greater in the hatchery strain compared to the wild strain while the crossbred group showed
588 intermediate vulnerability (sensu Mezzera and Largiadér 2001). Domestication has been shown
589 to increase vulnerability to angling in common carp (*Cyprinus carpio*) (Beukema 1969; Klefoth
590 et al. 2012) but decrease it in ayu (*Plecoglossus altivelis*) (Tsuboi et al. 2019). The changes in
591 vulnerability to fishing were likely mediated through strain-specific variation in heritable
592 personality traits (Kortet et al. 2014; Ågren et al. 2019; Alioravainen et al. 2020).

593 The crossbred strains demonstrated good growth rates ($\sim 7\text{-}11\text{ cm y}^{-1}$ in length,
594 comparable to most Finnish adfluvial brown trout stocks, Huusko et al. 2017) both in the
595 predation experiment and in the natural lakes, particularly when the female parent was from
596 the hatchery strain and the male was from a wild strain (Fig. 3). The strong effect of maternal
597 strain could be explained by the maternal inheritance of mitochondria and their important role
598 in explaining growth rates in brown trout (Salin et al. 2019). Also earlier experiments have

599 shown that migratory × resident crosses of brown trout display growth rates comparable to
600 those of pure migratory strains (Kallio-Nyberg et al. 2010; Alioravainen et al. 2018; Ågren et
601 al. 2019). Crossbreeding of migratory strains with resident strains might lead to early
602 maturation, but the data collected here did not indicate any major effects as fish from the pure
603 hatchery strain also showed signs of sexual maturation at smaller size than typical for migratory
604 brown trout (Huusko et al. 2017).

605 Overall, our results are conditionally positive for the use of slow-growing
606 resident strains of brown trout to help the recovery of wild-type behavioural trait value
607 distributions in hatchery broodstocks (McClelland and Naish 2007, Kelly and Phillips 2018).
608 As such these results support earlier studies in anadromous brown trout (Kallio-Nyberg et al.
609 2010). However, as our study was limited to F1-generation hybrids, there may be potential
610 outbreeding depression effects that would only manifest in subsequent generations
611 (McClelland and Naish 2007; Neff et al. 2011). In addition, the impact of crossbreeding on
612 maturation schedules of brown trout should be systematically evaluated under natural
613 conditions before adopting crossbreeding in conservational hatchery programs. Nevertheless,
614 stocking should not be continued beyond the first introduction of crossbred, genetically
615 heterogenous individuals to allow the introduced fish to adapt to their new environment
616 (Boulding and Hay 2001; Aprahamian et al. 2003; Aas et al. 2018).

617

618 *Fishing-induced selection versus natural selection*

619 The groups selected for high or low vulnerability to angling within the hatchery and wild strains
620 (Alioravainen et al. 2020) did not differ in any examined trait thus demonstrating a lack of a
621 measurable evolutionary response to fly fishing. As predicted (Philipp et al. 2009), the hatchery
622 origin high vulnerability line (OUV HV) showed the highest angling capture rate in the natural
623 lakes but the statistical difference to the other strains was small, and the wild strain high

624 vulnerability line (VAA HV) had the lowest vulnerability to angling (Supplementary Table
625 S3). Strong environmental effects over two summers of hatchery rearing and behavioural
626 conformity (Webster and Ward 2010) within groups may have partially diminished any
627 inherited differences between the HV and LV lines (see also Prokkola et al. 2021). Thus, our
628 results do not provide strong support for the absence of any response to fly fishing or falsify
629 behavioural results obtained in first-summer (described in Alioravainen et al. 2020) or second-
630 summer juveniles (Prokkola et al. 2021).

631 Even though we did not find a transgenerational response to fly-fishing in the F1
632 generation, fish captured in the natural lakes were generally larger than the non-captured fish.
633 Because we could not confirm that the non-captured fish were alive, this observation does not
634 directly imply positively size-selective fishing mortality. When the effect of size and condition
635 on the survival of the fish was controlled by focusing the vulnerability analysis only on the
636 captured fish (confirmed to be alive), vulnerability to angling was (marginally) explained only
637 by boldness score (in survivors of the predation experiment, $P = 0.099$) and length (all fish, P
638 $= 0.150$). Comparison of trait means in angled and gillnetted fish (Table 2) further revealed
639 that angling-imposed selection against slow growth and low boldness relative to gillnetted fish
640 (c.f. Klefoth et al. 2017). These differences could not be explained by the intrinsically later
641 captures via gillnets.

642 Slow growth rate could be an indication of high level of hunger, and thus rather
643 indicate a state-dependent vulnerability to angling rather than strong selection on life-history
644 traits. Largemouth bass (*Micropterus salmoides*) with high vulnerability to angling are known
645 to have higher metabolic rates (but see Prokkola et al. 2021 for brown trout) and food
646 consumption than bass with low vulnerability to angling (Cooke et al. 2007; Redpath et al.
647 2010). Thus, in resource-limited natural lakes fish with high vulnerability to be captured by
648 angling should show increased energy demands and thus greater feeding motivation, but lower

649 realized growth rates due to the lack of suitable food (Lennox et al. 2017). Another potential
650 explanation for our results lies in gillnet selection against active behaviour (Alós et al. 2012),
651 as gillnet fishing is highly selective for bold behaviour and fast growth (Biro and Post 2008).
652 However, these explanations are not necessarily mutually exclusive, as angling could have
653 targeted the fish that had the highest intrinsic growth capacity and thus the highest foraging
654 motivation while the gillnet-captured fish were more successful at foraging and had lower
655 motivation to strike the flies and lures.

656 We defined boldness score as tendency to spend time in the predator sections of
657 the study ponds but whether these scores actually reflected boldness, exploration or tendency
658 to avoid competition in the riffle section is not clear (Alioravainen et al. 2018). Thus, angling
659 targeted low-activity individuals that preferably occupied the predator-free riffle areas in the
660 predation experiment. A feasible alternative explanation is that angling selected against
661 qualities like dominance that made fish successful in obtaining a feeding territory in the riffle
662 sections (Alioravainen et al. 2018). No matter what the actual behavioural mechanism is, the
663 link between behaviours under predation risk and vulnerability to angling adds to the evidence
664 that fishing can induce selection on ecologically relevant behavioural traits (Olsen et al. 2012;
665 Klefoth et al. 2017; Arlinghaus et al. 2017).

666 Resolving how behavioural and size-dependent traits in general affect
667 vulnerability to fishing in varying environments remains a challenge (Uusi-Heikkilä et al. 2008;
668 Lennox et al. 2017). Obviously, the net result of conflicting size-selection from natural and
669 human-induced sources, as acting on released fish, is affected both the post-release survival
670 rate and the intensity of harvesting (Edeline et al. 2007; Olsen & Moland 2011). Changes in
671 growth may not be likely due to angling-selection, while natural selection might not be
672 producing a similar counterforce for angling-induced behavioural changes as for size-selection
673 induced by positively size-selective fishing gear (Monk et al. 2021).

674 The relationships between growth rate, personality, body size, survival and
675 physiology are generally environment-dependent rendering the interpretation of observational
676 results challenging. Gillnet capture was assumed to reflect survival more than vulnerability to
677 this type of fishing as gillnet fishing for brown trout in small lakes is extremely efficient
678 (Borgstrøm 1992). The findings of large size and high condition factor at stocking being the
679 major determinants of capture suggest that the stocking lakes were challenging environments
680 for the fish (Carlson et al. 2008). In the shallow, clearwater Lake Koukkulampi, length at
681 stocking was the main predictor of angling vulnerability. In the deeper and more resource-rich
682 Lake Kylmälampi, fish with high condition factor at stocking were most vulnerable to angling.
683 Because the study lakes did not support natural piscivorous fishes, size- and condition-
684 dependent survival must have arisen from other reasons such as starvation resistance and/or
685 very high water temperatures that occurred in summer 2018 right after stocking. Catch rates by
686 angling decreased rapidly with increases in cumulative fishing effort (similar to Koeck et al.
687 2019a), while overall, we recorded high capture rates in the smaller Lake Kylmälampi
688 compared to significantly lower capture rates in the larger Lake Koukkulampi. Despite the
689 better capture rates and acutely better growth rates in Lake Kylmälampi, the Lake
690 Koukkulampi fish showed better growth in length from the stocking year to the next year. Any
691 direct gear selectivity effects are unlikely to explain the results on growth because identical
692 angling (lures and flies) gear and gillnets with multiple mesh sizes were used in both lakes and
693 even the smallest trout stocked were catchable with the gears used.

694

695 *Conclusions*

696 This study compared the effects of angling-induced selection to strain-specific variation in
697 brown trout. Overall, our results show that the strain differences (e.g., domesticated vs. wild
698 fish) are much stronger than those induced by angling over a single generation. Thus, hatchery

699 impacts on brown trout stocks are expected to override fishing-induced effects (Hutchings and
700 Fraser 2008) while gillnet fishing may induce stronger undesired selection on production-
701 related traits than angling (Handford et al. 1977). Fishing-induced selection may be at least
702 partly contradictory to hatchery-induced selection as the pure hatchery group was the most
703 vulnerable to angling, and hatchery-reared brown trout typically show increased boldness over
704 wild brown trout (Ågren et al. 2019; Alioravainen et al. 2020). Fishing may select against
705 hatchery fish and wild \times hatchery hybrids (Mezzera and Largiadér 2001), further reducing risk
706 of introgressive hybridization in wild brown trout populations receiving hatchery-reared fish
707 (Wills 2006; Koeck et al. 2019b). The offspring of wild resident fish clearly experienced slower
708 growth both in the predation experiment and in the natural lakes, while crossbred groups
709 performed relatively well. Our results suggest that migratory hatchery female \times resident wild
710 male crosses could be used in stockings intended to create new, naturally reproducing
711 populations or be further selectively bred for the restoration of pre-domestication fitness in
712 hatchery broodstocks.

713

714 **Conflict of Interest**

715 The authors declare no conflicts of interest.

716

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726

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964 **Table 1.** Cox regression results on vulnerability to capture by I: all fishing gear, II: by angling
 965 among the subset of fish that were included in the predation experiment and III: by angling
 966 only among the fish confirmed to be alive after stocking (captured by either method and tested
 967 for boldness).

Subset	Factor / variable	Loss of χ^2	d.f.	Sig.	Wald	d.f.	Sig.	Exp(B)
I	Strain	6.162	8	0.629				
	Boldness score	0.941	1	0.332				
	Length at stocking	3.107	1	0.078	3.11	1	0.078	1.006
	Condition factor	7.416	1	0.006	7.28	1	0.007	1311.22
II	Strain	9.112	8	0.333				
	Boldness score	0.178	1	0.674				
	Length at stocking	4.464	1	0.035	4.493	1	0.034	1.008
	Condition factor	5.582	1	0.018	5.506	1	0.019	1683.11
III	Strain	9.623	8	0.293				
	Length at stocking	0.151	1	0.698				
	Condition factor	0.237	1	0.626				
	SGR	0.095	1	0.758				
	Boldness score	2.874	1	0.090	2.729	1	0.099	0.819

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980 **Table 2.** Selection imposed by angling on fish confirmed to be alive after stocking via capture
 981 by angling or by gillnets. Mean-standardized selection gradient β_μ is calculated according to
 982 Matsumura et al. (2012). *Not corrected for average, because of arbitrary scale and negative
 983 average boldness score (-0.0287). S.D. refers to standard deviation.

Trait	Captured			Non-captured			<i>t</i>	Sig.	β_μ
	Average	S.D.	N	Average	S.D.	N			
Boldness score	-0.143	0.744	117	0.245	1.004	49	2.75	0.007	0.383*
Stocking length (mm)	245.2	23.2	232	242.0	22.6	108	-1.207	0.228	-1.017
Condition factor	0.248	0.026	232	0.246	0.026	108	-0.585	0.559	-0.446
IGR ₁₀₀ 100 ln g d ⁻¹	0.135	0.534	232	0.250	0.364	108	2.326	0.021	0.039

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1002 **Figure captions**

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1004 **Fig. 1.** Schematic outline of the experiments. Breeding design was based on 3 males/females
1005 \times 3 males/females matrix except for OUV \times POH due to lack of mature POH females. The
1006 initial rearing of strains occurred in replicated tanks but was fully common-garden since PIT-
1007 tagging in autumn 2016. Experimental fishing in the two natural lakes was started by angling
1008 and finished with gillnets in both years.

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1010 **Fig. 2.** Realized total frequencies of different outcomings by strain in the stocking experiment
1011 pooled over the two natural lakes. See Supplementary Figures S2-S3 for lake-specific results.

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1013 **Fig. 3.** Estimated marginal mean (model-predicted average) strain differences (\pm 95%
1014 confidence intervals) in (a) boldness scores in the predation exposure experiment, (b) growth
1015 rates in the predation exposure experiment (IGR_{100} , see the main text for the calculation), and
1016 (c) realized post-stocking growth rates (IGR_{100}) of the captured fish in the natural lakes. The
1017 first strain acronym indicates the maternal origin and the latter the sire origin.

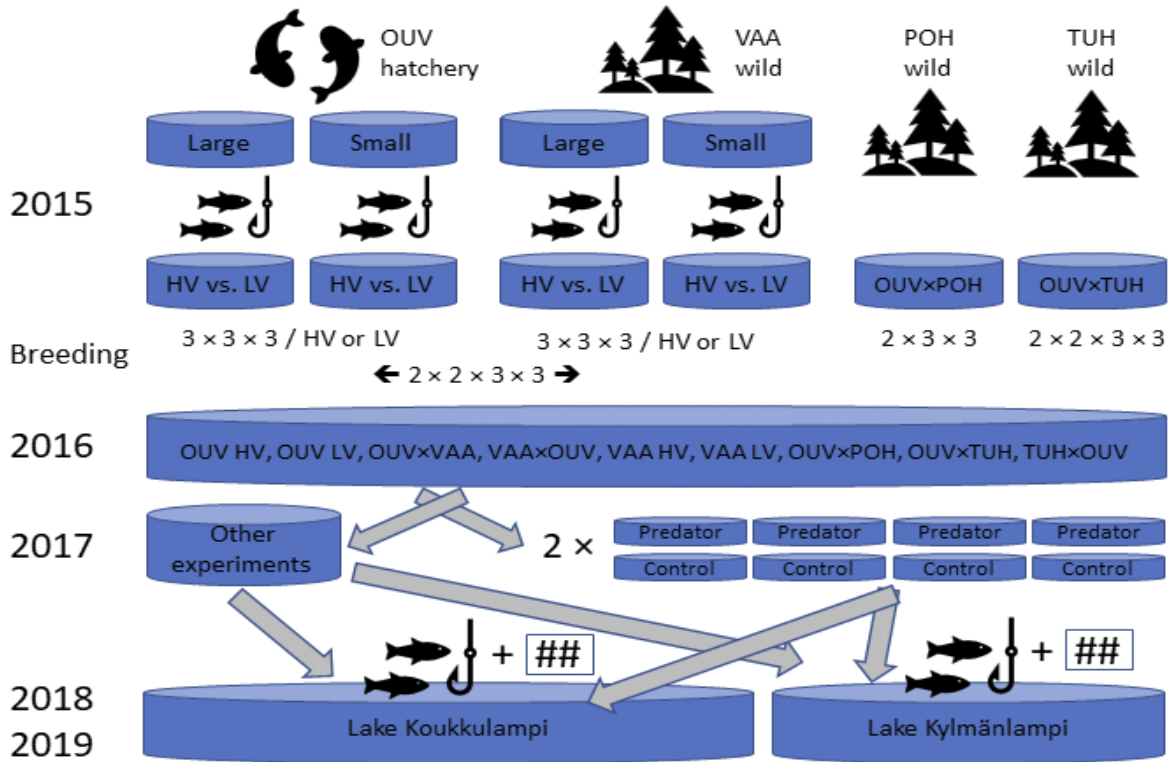
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1023 **Figure 1.**

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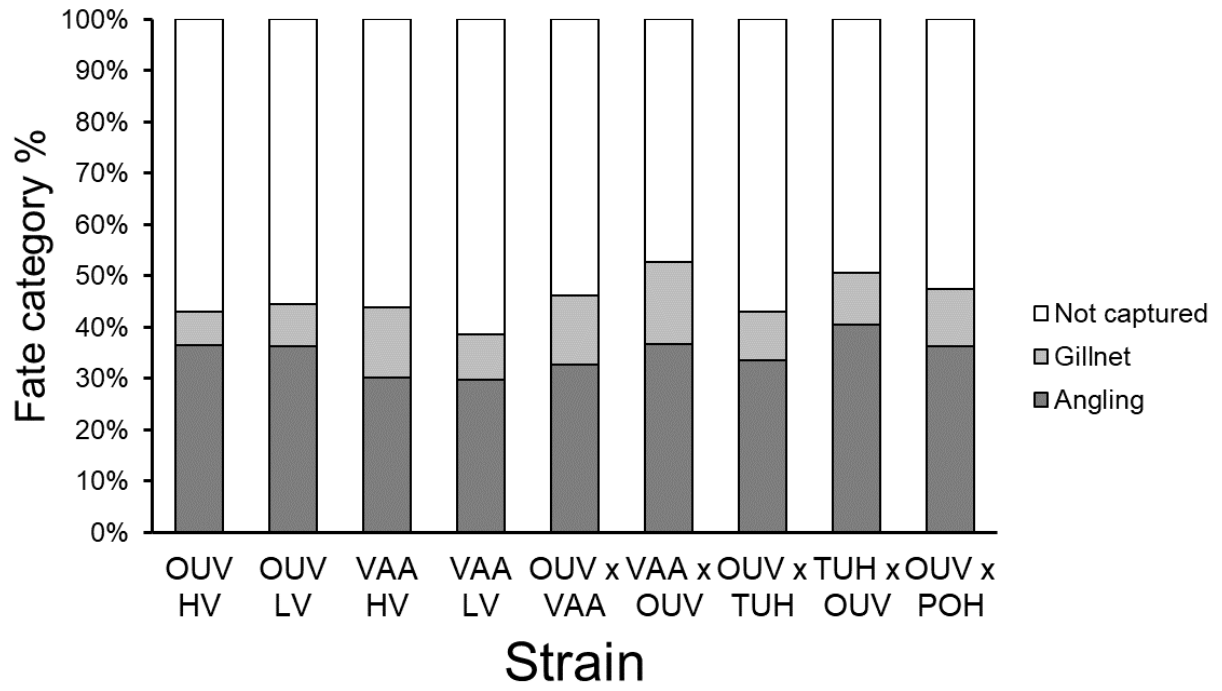
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1037 **Figure 2.**

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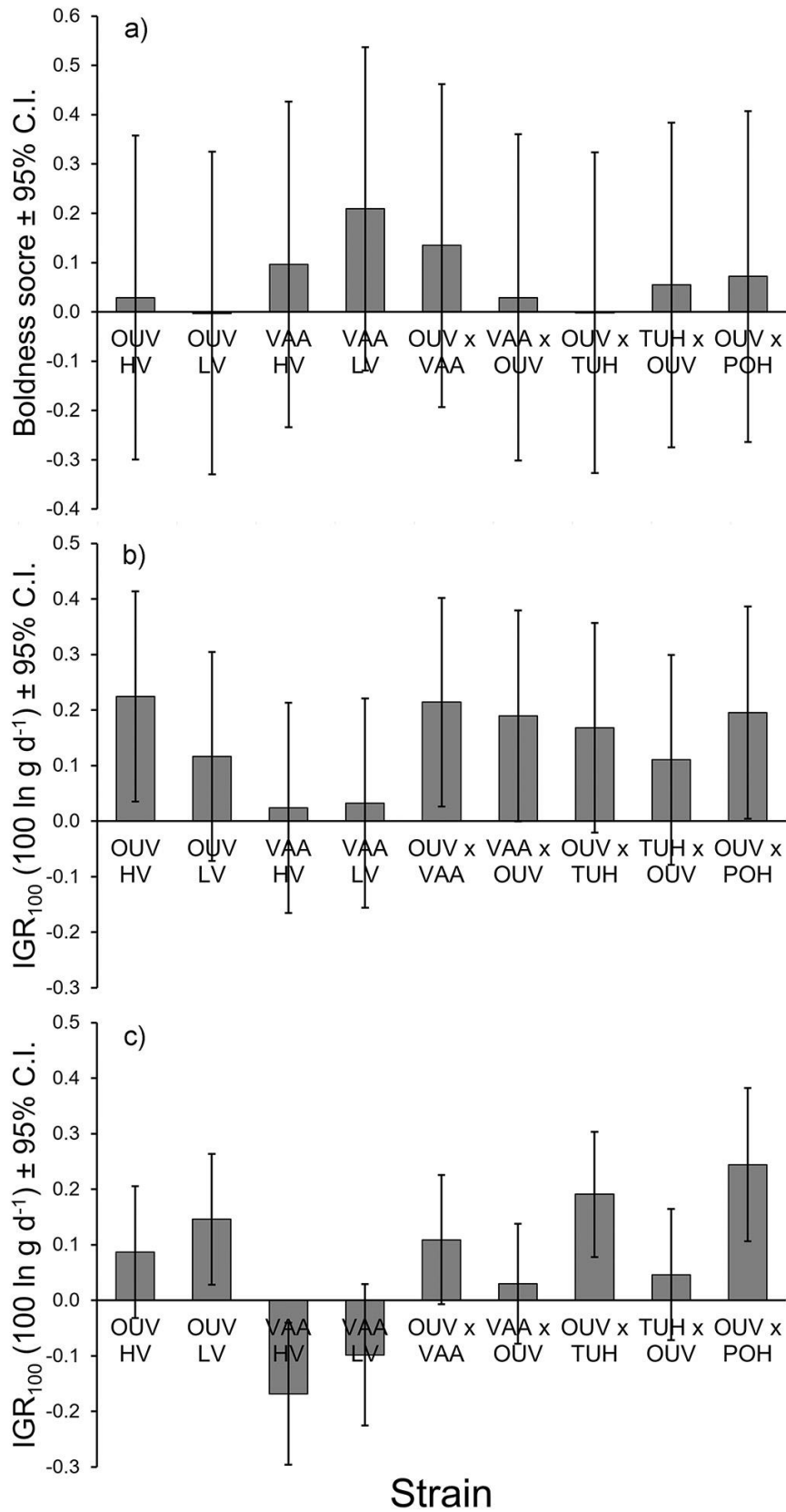
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1053 **Figure 3.**