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Hanene Mohamed, Philippe Robert

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DYNAMIC TREE ALGORITHMS

HANÈNE MOHAMED AND PHILIPPE ROBERT

ABSTRACT. In this paper, a general tree algorithm processing a random flow of arrivals is analyzed. Capetanakis-Tsybakov-Mikhailov's protocol in the context of communication networks with random access is an example of such an algorithm. In computer science, this corresponds to a trie structure with a dynamic input. Mathematically, it is related to a stopped branching process with exogeneous arrivals (immigration). Under quite general assumptions on the distribution of the number of arrivals and on the branching procedure, it is shown that there exists a *positive* constant λ_c so that if the arrival rate is smaller than λ_c , then the algorithm is stable under the flow of requests, i.e. that the total size of an associated tree is integrable. At the same time a gap in the earlier proofs of stability of the literature is fixed. When the arrivals are Poisson, an explicit characterization of λ_c is given. Under the stability condition, the asymptotic behavior of the average size of a tree starting with a large number of individuals is analyzed. The results are obtained with the help of a probabilistic rewriting of the functional equations describing the dynamic of the system. The proofs use extensively this stochastic background throughout the paper. In this analysis, two basic limit theorems play a key role: the renewal theorem and the convergence to equilibrium of an auto-regressive process with moving average.

This paper is dedicated to Philippe Flajolet on the occasion of his 60th birthday.

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1. INTRODUCTION

This paper investigates probabilistic algorithms which decompose recursively a given set of elements (also referred to as items) into random subsets until all subsets have a cardinality less than some fixed number D . The dynamic aspect of the algorithms analyzed here is that new elements are added to the subsets created during each decomposition.

The general procedure is as follows: If the cardinality of the set is strictly less than $D > 0$, the process is stopped. Otherwise the set is split into several subsets and each subset receives a random number of new elements. The algorithm is then

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recursively applied to each of these subsets. A tree is naturally associated to this algorithm: the root node having the initial items, the subsequent nodes contain the corresponding subsets and so on. . . At the end of this process, one ends up with a tree whose leaves contain less than D items and all internal nodes contain more than D items. Such an algorithm can be seen as a state dependent branching process which dies out whenever a termination condition is satisfied. When there are no arrivals of new elements, the algorithm is called *static*. Static tree algorithms are of fundamental importance as a generic class and also because of their use in computer science (where the corresponding data structure is called a *trie*) and in many other areas such as communication protocols or distributed systems. See Mohamed and Robert [14] for a general overview of static tree algorithms.

The extension analyzed here, with the introduction of new elements, consists in introducing immigration in the language of branching processes. This situation is quite natural in the context of communication protocols for which new requests (immigration) arrive continuously at the communication node. On the use of these algorithms in the context of communication networks, see the surveys by Berger [2] and Massey [12]. On the mathematical side, as it will be seen, the analysis of these algorithms turns out to be more delicate.

TREE ALGORITHM WITH IMMIGRATION $\mathcal{S}(n)$.

- TERMINATION CONDITION.
If $n < D$ \longrightarrow STOP.
 - TREE STRUCTURE.
If $n \geq D$, randomly divide n into n_1, \dots, n_G , with $n_1 + \dots + n_G = n$.
 \longrightarrow APPLY $\mathcal{S}(n_1 + A_1)$, $\mathcal{S}(n_2 + A_2)$, \dots , $\mathcal{S}(n_G + A_G)$.
where (A_i) are i.i.d. random variables.
-

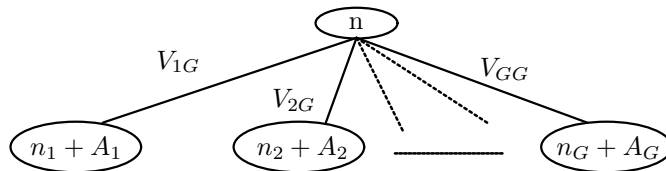


FIGURE 1. First level of the tree algorithm to decompose $n \geq D$, $n = n_1 + \dots + n_G$ and with arrivals (A_i)

For the static algorithm, when there are no new arrivals, provided that the decomposition mechanism is not degenerated, it is easily seen that the associated tree is almost surely finite, in fact that the total number of its nodes is integrable. Mohamed and Robert [14] investigates this case.

Finiteness of the Associated Tree and Law of Large numbers. When there is a set of new items arriving every time unit, it may happen that the algorithm does not terminate with probability 1, i.e. that the associated tree is infinite. In this case, the algorithm cannot cope with the flow of arriving requests. This non-trivial phenomenon is analyzed in this paper. Furthermore, in the case where the process terminates almost surely and that there are n items at the root, another

problem is to describe the asymptotic behavior of the average size of the tree as n gets large. As it will be seen this is a quite challenging question.

Tree Recurrences. In this setting the main quantity of interest R_n^A is the total number of nodes of the associated tree when the algorithm starts with n items. The superscript “A” of R_n^A refers to the common distribution of the i.i.d. sequence (A_i) . In this case, one gets naturally the recursive relation

$$(1) \quad R_n^A \stackrel{\text{dist.}}{=} 1 + (R_{1,n_1+A_1}^A + R_{2,n_2+A_2}^A + \cdots + R_{G,n_G+A_G}^A) \mathbb{1}_{\{n \geq D\}},$$

where, for $i \geq 1$, $(R_{i,n}^A, n \geq 0)$ are independent random variables with the same distribution as $(R_n^A, n \geq 0)$. The total size of the tree is 1 plus the size of all of sub-trees of level 1.

A Markovian Representation. This algorithm can also be represented by a Markov chain (\underline{L}_n) on the set $\mathcal{S} = \cup_{n \geq 0} \mathbb{N}^n$ of finite sequences on \mathbb{N} , its transitions are described as follows: if $(\underline{L}_0) = (l_0, l_1, l_2, \dots)$

$$(2) \quad (\underline{L}_1) = \begin{cases} (l_1 + A_1, l_2, \dots, l_n, \dots) & \text{if } l_0 < D, & \text{/SHIFT/} \\ (n_1 + A_1, n_2, \dots, n_G, l_1, l_2, l_3, \dots) & \text{if } l_0 \geq D, & \text{/SPLIT/} \end{cases}$$

if the integer l_0 is decomposed into $l_0 = n_1 + n_2 + \cdots + n_G$ by the splitting procedure (see the precise description below). For $n \geq 1$, if the initial state is $(\underline{L}_0 = (n, 0, \dots, 0, \dots))$ then it is not difficult to check that

$$R_n^A \stackrel{\text{dist.}}{=} \inf\{t \geq 1 : (\underline{L}_t = (0, \dots, 0, \dots))\},$$

the variable R_n^A can also be viewed as the hitting time of the empty state by the Markov chain (\underline{L}_n) . The ergodicity of the Markov chain (\underline{L}_n) is therefore equivalent to the fact that the variable $R_{A_1}^A$ is integrable. Because of the description by a sequence (a *stack* in the language of computer science), the tree algorithm is, sometimes, also called *stack algorithm*.

Provided that the variables (R_n^A) are integrable, the Poisson transform $\phi(x)$ of the sequence $(\mathbb{E}(R_n^A))$ is defined as

$$(3) \quad \phi(x) \stackrel{\text{def.}}{=} \sum_{n \geq 0} \mathbb{E}(R_n^A) \frac{x^n}{n!} e^{-x}.$$

In the case where the arrivals have a Poisson distribution with parameter λ , the ergodicity of (\underline{L}_n) implies that the Poisson transform of the sequence $(\mathbb{E}(R_n^A))$ is well defined at $x = \lambda$.

Iteration of non-commutative functional operators. Mathematically, these tree algorithms are quite challenging. By using an iterating scheme, a family of simple functional operators $(P_v^\lambda, v \in (0, 1))$ play a central role in the analysis, they are defined by, if $f: \mathbb{R}_+ \rightarrow \mathbb{R}_+$ is a continuous function,

$$P_v^\lambda(f)(x) = \frac{1}{v} f(xv + \lambda),$$

the static algorithm corresponds to the case when $\lambda = 0$.

By taking the expected value of Equation (1) and by iterating the functional obtained, it turns out that the Poisson transform can be expressed by the following

equation

$$(4) \quad \sum_{n=1}^{+\infty} \int_{[0,1]^n} P_{v_1}^\lambda \circ P_{v_2}^\lambda \circ \dots \circ P_{v_n}^\lambda (f)(x) \bigotimes_{i=1}^n \mathcal{W}(dv_i), \quad x \geq 0,$$

for a convenient function f depending on some unknown constants and where \mathcal{W} is some probability distribution on $(0, 1)$.

Note that the operators $(P_v^\lambda, v \in [0, 1])$ commute only when $\lambda = 0$. This complicates significantly the analysis of the asymptotic behavior of Expression (4) as x goes to infinity. In the static case, one has that

$$P_{v_1}^0 \circ P_{v_2}^0 \circ \dots \circ P_{v_n}^0 = P_{v_1 v_2 \dots v_n}^0$$

which gives a multiplicative representation of Expression (4) which is analyzed by using Mellin transform methods in an analytical context (See Flajolet *et al.* [10]) or by using random walks methods with a probabilistic approach (See Mohamed and Robert [14]). When $\lambda > 0$, such a multiplicative formulation is not available, these methods have therefore to be adapted. It turns out that such a generalization is not straightforward.

Example: The Binary Tree. The splitting mechanism is binary, the branching number G is deterministic and equal to 2, $G \equiv 2$, and with deterministic weights, $V_{11} \equiv p$ and $V_{12} \equiv q = 1 - p$ with $p \in (0, 1)$. The variables (A_i) are assumed to be Poisson with parameter λ . The upper index A in (R_n^A) is replaced by λ in this case. Provided that the variables are integrable, and if $\alpha_n^\lambda = \mathbb{E}(R_n^\lambda)$, the integration of Equation (1) gives the identities $\alpha_0^\lambda = \alpha_1^\lambda = 1$ and, for $n \geq 2$,

$$(5) \quad \alpha_n^\lambda = 1 + \sum_{i=0}^n \binom{n}{i} p^i q^{n-i} \sum_{k, \ell \geq 0} \frac{\lambda^k}{k!} e^{-\lambda} \frac{\lambda^\ell}{\ell!} e^{-\lambda} (\alpha_{i+k}^\lambda + \alpha_{n-i+\ell}^\lambda).$$

In this case it is not difficult to check that the corresponding Poisson transform ϕ satisfies the equation

$$(6) \quad \phi(x) = \phi(px + \lambda) + \phi(qx + \lambda) + h(x),$$

where $x \rightarrow h(x)$ is some fixed function with a specific form and unknown coefficients. This functional relation can be rewritten as

$$\phi(x) = \int_0^1 P_w^\lambda(\phi)(x) \mathcal{W}(dw) + h(x),$$

where $\mathcal{W} = p\delta_p + q\delta_q$, where δ_x is the Dirac measure at $x \in \mathbb{R}$. A (formal) iteration of this function gives Representation (4) for ϕ .

Literature. These problems have been analyzed in several ways in the past. Motivated by the design of stable communication protocols, Tsybakov and Mikhaïlov [15] and Capetanakis [6] did the early studies in this domain (and, at the same time, designed the algorithms in the context of distributed systems), see also Tsybakov and Vvedenskaya [16].

At the end of the 1980's, Flajolet and his co-authors in a series of interesting papers [8, 9, 13] have obtained rigorous asymptotics for solutions of Equations of the type (5) in several cases. In the first of them [8], Recurrence (5) for the binary tree is investigated. It is shown that for λ smaller than some threshold, there is a unique sequence (α_n^λ) of real numbers which is solution of this recurrence. In this paper, a sophisticated asymptotic analysis of the sequence (α_n^λ) is presented.

Basically, it is shown that the sequence (α_n^λ) grows linearly with respect to n and its rate is, in some cases, a function with small fluctuations. It is conducted in three steps:

- (1) By iteration of this equation, express the Poisson transform ϕ of (α_n^λ) under the form (4),

$$(7) \quad \phi(x) = \sum_{n \geq 0} \sum_{i \in \{1,2\}^n} h(\sigma_{i_1} \circ \sigma_{i_2} \circ \cdots \circ \sigma_{i_n}(x)),$$

where $\sigma_1(x) = px + \lambda$ and $\sigma_2(x) = qx + \lambda$.

- (2) Obtain the asymptotics of the Poisson transform $\phi(x)$ as x goes to infinity.
 (3) Prove that $(\phi(x))$ and (α_n^λ) have the same behavior as x [resp n] goes to infinity.

The main part of the analysis is devoted to step (2) where, via several estimates of series and contour integrals with arguments from complex analysis, the authors can identify in the series (7) the main contributing terms when x goes to infinity. Given the complexity of this analysis for this example of the binary tree, an extension of these methods to a more general branching mechanism seems to be more than challenging. The case when \mathcal{W} is the uniform distribution on Q points is discussed in Section IV of Mathys and Flajolet [13]. Note that if \mathcal{W} has some Lebesgue component, a representation of ϕ as a series similar to (7) is no longer available, one has to go back to the general representation (4).

In addition to extend these results to a quite general branching scheme, the purpose of this paper is also of developing probabilistic methods to analyze additive functionals of various tree structures. This program has been initiated in Mohamed and Robert [14] in the context of static tree algorithms. We nevertheless believe that, because of the intricacies of its associated equations, the tree algorithm with arrivals provides a real significant test for this approach. It turns out that the method proposed in this paper has some advantage in that it can handle more easily and in a more general setting the complexity of the underlying non-commutative iterating scheme of this algorithm.

Recent works deal with some aspects of these fundamental algorithms, see Boxma *et al.* [5], Janssen and de De Jong [11] and Velthoven *et al.* [17] for example. For surveys on the communication protocols in random access networks, see Berger [2], Massey [12] and Ephremides and Hajek [7].

Contributions of the paper. The main objective of the paper is to present a probabilistic approach to these problems that can tackle, with a limited technical complexity, quite general models of tree algorithms with immigration.

For the model considered in this paper, Equation (6) for the Poisson transform becomes

$$(8) \quad \phi(x) = \int_0^1 \frac{\phi(wx + \lambda)}{w} \mathcal{W}(dw) + h(x),$$

for some probability distribution \mathcal{W} on $(0, 1)$ describing the branching mechanism of the splitting algorithm and some function h . The example of the binary tree corresponds to the case where \mathcal{W} is carried by two points p and q as noted before. When the measure \mathcal{W} is carried by Q points of $(0, 1)$, the equivalent expression for the series (7) involves the various products of Q functions $(\sigma_m, 1 \leq m \leq Q)$.

A. Stability of Tree Algorithms

The stability property of the tree algorithm, i.e. the fact that the tree is almost surely finite, is a crucial issue for communication networks. Assuming Poisson arrivals with parameter λ , it amounts to the existence of some threshold $\lambda_0 > 0$ such that if the arrival rate λ is strictly less than λ_0 then the associated Markov process describing the tree algorithm (see above) is ergodic. During the 70's and 80's the design of stable protocols and the mathematical proof of their stability has been a very active research domain. Recently, because of the emergence of wireless and mobile networks, there is a renewed interest in these models. The first protocols, Aloha and Ethernet, turned out to be unstable when there are an infinite number of possible sources, i.e. for these algorithms the number of requests waiting for transmission goes to infinity in distribution for any arrival rate. See Aldous [1]. The tree algorithm corresponding to the example of the binary tree with $p = q = 1/2$ is the first such protocol in a really distributed system where the stability region is non-empty, see Massey [12] and Bertsekas and Gallager [4]. Later the tree algorithm has been improved in order to have a larger stability region.

A gap in the proof of previous stability results. The stability results obtained up to now are under the assumption of Poisson arrivals. The proofs known to us rely on the analysis of Equations of the type (5) for the sequence $(\mathbb{E}(R_n^\lambda))$: it is shown that there exists some $\lambda_0 > 0$ such that for $\lambda < \lambda_0$, there is a unique finite solution (α_n) and from there it is concluded that the corresponding Markov chain is ergodic. The problem here is that this analysis shows only that, for $\lambda < \lambda_0$, there exists a unique sequence (α_n) of finite real numbers satisfying Relation (5). The sequence $(\beta_n) = (1, 1, +\infty, \dots, +\infty, \dots)$ also satisfies Relation (5). At this point, without an additional argument it cannot be concluded that, for $\lambda < \lambda_0$, $(\mathbb{E}(R_n^\lambda))$ is indeed (α_n) and not (β_n) . To make the identification with (α_n) , it has to be proved that the random variables R_n^λ , $n \in \mathbb{N}$ are indeed integrable but this is precisely the final result. Strictly speaking, the previous results have only shown that the system is unstable whenever $\lambda \geq \lambda_0$.

This gap is fixed in this paper. The key ingredient to relate recurrence relation (5) and the sequence (R_n^λ) is a perturbation result of the static case, i.e. when $\lambda = 0$. It is also shown that, under general assumptions on the distribution of the inputs (A_i) and on the branching mechanism, the tree algorithm is stable for a sufficiently small arrival rate. As far as we know, this is the first stability result for tree algorithms with non-Poissonian arrivals.

B. Analysis of General Tree Recurrences.

The second part of the paper investigates the asymptotic behavior of the sequence $(\mathbb{E}(R_n^A)/n)$ where (R_n^A) is a solution of the tree recurrence (1) under the condition that A is a Poisson random variable with a parameter λ less than some constant.

Some of the ingredients of the analysis of static algorithms ($\lambda = 0$) of Mohamed and Robert [14] are used. The situation is nevertheless completely different when $\lambda > 0$. As mentioned above, the non-commutativity of the operators P_v is a major issue. An additional important difficulty is the fact that, contrary to the case $\lambda = 0$, the function h of Equation (8) is unknown, D coefficients have to be determined.

To study these recurrences, the approach of the paper consists in expressing Series (7) for the binary tree or Equation (4) in the general case as the expected value of some random variable depending on some auto-regressive process with

moving average (X_n) defined by, $X_0 = x$ and

$$(9) \quad X_{n+1} = W_n X_n + 1, \quad n \geq 0,$$

where (W_n) is an i.i.d. sequence whose common distribution is \mathcal{W} .

Key limit theorems are used to derive the asymptotic behavior of the sequence $(\mathbb{E}(R_n^\lambda))$: The renewal theorem and the convergence of (X_n) to its stationary distribution. In the particular case of the binary tree, they avoid the use of estimations of Fayolle *et al.* [8] necessary to get the significant terms of the series (7) in the asymptotic expansion. Roughly speaking, with these two limit theorems, the probabilist “knows” what are the most likely trajectories of the compositions of σ_1 and σ_2 . This approach simplifies *significantly* the asymptotic analysis of the algorithm. Another key step is to identify the D unknown coefficients of function h , they are expressed as a functional of the auto-regressive process, of its invariant distribution in particular.

Outline of the paper. The paper is organized as follows: Section 2 shows that, under a quite general assumptions on arrivals, the stability region is not empty. In Section 3, by denoting R_n^λ the size of the tree when $n \geq 0$ elements are at the root, under the assumption of Poisson arrivals, a probabilistic representation of the Poisson transform of the sequence $(\mathbb{E}(R_n^\lambda))$ is established and an auto-regressive process with moving average is introduced. Section 4 established the main stability result (Theorem 9) for a general tree algorithm Poisson arrivals. Section 5 investigates the delicate asymptotics of the sequence $(\mathbb{E}(R_n^\lambda)/n)$, Theorem 12 summarizes the results obtained.

2. EXISTENCE OF A NON-EMPTY STABLE REGION

This section it is proved that, if the arrival rate is sufficiently small, then the tree obtained with the algorithm is almost surely finite, its size is in fact integrable.

Formulation of the problem. The algorithm starts with a set of n items. If $n < D$ then it stops. Otherwise, this set is randomly split into G subsets, where G is some random variable. Now, conditionally on the event $\{G = \ell\}$, for $1 \leq i \leq \ell$, each of the n items is sent into the i th subset with probability $V_{i,\ell}$, where $V_\ell = (V_{i,\ell}; 1 \leq i \leq \ell)$ is a random probability vector on $\{1, \dots, \ell\}$. The quantity $V_{i,\ell}$ is the weight on the i th edge of the splitting structure. Additionally, a vector (A_1, \dots, A_ℓ) of independent random variables with the same distribution as some random variable A is given, A_i new items are added to the i th subset. If n_i is the cardinality of the i th subset then, conditionally on the event $\{G = \ell\}$ and on the random variables $V_{1,\ell}, V_{2,\ell}, \dots, V_{\ell,\ell}$, the distribution of the vector (n_1, \dots, n_ℓ) is multinomial with parameter n and $(V_{1,\ell}, V_{2,\ell}, \dots, V_{\ell,\ell})$. If the i th subset, $1 \leq i \leq n$, is such that $n_i + A_i < D$, the algorithm stops for this subset. Otherwise, it is applied to the i th subset: a variable G_i , with the same distribution as G , is drawn and this i th subset is split into G_i subsets, and so on ... See Figure 1.

The Q -ary algorithm considered by Mathys and Flajolet [13] corresponds to the case where G is constant and equal to Q and the vector of weights $(V_{i,Q}, 1 \leq i \leq Q)$ is deterministic.

As in Mohamed and Robert [14] for the static case, the key characteristic of this algorithm is a probability distribution \mathcal{W} on $[0, 1]$ defined with the branching distribution (the variable G) and the weights on each arc (the vector $(V_{1,G}, \dots, V_{G,G})$).

As it will be seen, the asymptotic behavior of the algorithm can be described only in terms of the distribution \mathcal{W} .

Definition 1. *The splitting measure is the probability distribution \mathcal{W} on $[0, 1]$ defined by, for a non-negative Borelian function f ,*

$$(10) \quad \int f(x) \mathcal{W}(dx) = \mathbb{E} \left(\sum_{i=1}^G V_{i,G} f(V_{i,G}) \right) = \sum_{\ell=2}^{+\infty} \sum_{i=1}^{\ell} \mathbb{P}(G = \ell) \mathbb{E}(V_{i,\ell} f(V_{i,\ell})).$$

In the context of fragmentation processes, the measure is related to the *dislocation measure*, see Bertoin [3].

The following conditions will be assumed throughout the paper.

Assumptions (H)

- **(H₁)** *There exists some $\delta > 0$ such that the relation $\mathcal{W}([0, \delta]) = 1$ holds.*
- **(H₂)**

$$\int_0^1 \frac{|\log x|}{x} \mathcal{W}(dx) < +\infty.$$

Condition **(H₁)** implies in particular the non-degeneracy of the splitting mechanism:

$$\sup_{\ell \geq 2} \sup_{1 \leq i \leq \ell} V_{i,\ell} \leq \delta < 1.$$

Definition 2. *For $n \geq 0$ and $p \geq 1$, define by $R_n^{A,p}$, the number of nodes with level (generation) less than p th when n items are at the root node. The variable A has the same distribution as the common distribution of the i.i.d. sequence (A_i) of the arrivals. By convention, $R_n^{A,\infty} = R_n^A$ and $(R_n^{0,p})$ refers to the static case, i.e. when $A \equiv 0$. The variable G_A is defined as*

$$G_A = \sum_{i=0}^{+\infty} \sum_{k=1}^{A_i} B_k^i,$$

where, for $i \geq 0$, $(B_k^i, k \geq 0)$ is an i.i.d. Bernoulli sequence independent of (A_i) such that $\mathbb{P}(B_k^i = 1) = \delta^i$.

Note that if $E(A) < +\infty$, the variable G_A is well defined and integrable.

The following lemma establishes a useful relation for tree sizes, the symbol \leq_{st} is, as usual, for the classical stochastic ordering.

Lemma 1 (Stochastic Inequality). *For $n \in \mathbb{N}$, under Assumption **(H)**, the relation*

$$(11) \quad R_n^{A,p} \leq_{\text{st}} R_n^{0,p} + \sum_{k=1}^{R_n^{0,p}} \tilde{R}_{D-1+G_{A,i}}^{A,p} \mathbb{1}_{\{G_{A,i} > 0\}}$$

holds, where $(\tilde{R}_n^{A,p})$ [resp $(G_{A,i})$] is a sequence of random variables with the same distribution as $(R_n^{A,p})$ [resp G_A]. The variables $(\tilde{R}_n^{A,p})$, $(R_n^{0,p})$ and $(G_{A,i})$ are independent.

Proof. First note that, for $n \leq m$, one has clearly that $R_n^{A,p} \leq_{\text{st}} R_m^{A,p}$. A coupling is used to prove the stochastic ordering (11). The splitting algorithm is first played only for the initial n items. The total number of nodes up to level p for the associated tree T^0 is $R_n^{0,p}$. The leaves of the tree have at most D items and the internal nodes more than D items.

Now external arrivals are added to the internal nodes and the splitting algorithm played on these new items with the branching structure associated with \mathcal{T}^0 until they reach one of the leaves of \mathcal{T}^0 . From there, for all the leaves which have more than D items, the dynamic algorithm is played starting from this node. The number of external items at the leaves has thus to be estimated.

Because of Assumption **(H₁)**, an item in a node containing more than D items is sent to a given children of this node with probability at most δ . Hence, a given leaf $L = I_i$ of \mathcal{T}^0 with depth $i \leq p$ contains at most $D-1$ initial items and a fraction of the number of new items $A_{L,k}$, $0 \leq k \leq p$ arrived successively at the internal nodes connecting this leaf to the root. Each of the external items arrived at some node goes to some fixed node below with probability at most δ so that, in distribution, there are at most

$$B_1^1 + B_2^1 + \cdots + B_{A_{L,i-1}}^1$$

such items at this node. Similarly, for $2 \leq k \leq i-1$, external items arrived at a node of level k will reach some fixed node of level i with probability at most δ^{i-k} . Consequently, the total number N_L of items at leaf L is, in distribution, at most

$$D-1 + \sum_{i=1}^{p-2} \sum_{k=1}^{A_{L,i}} B_k^i.$$

If N_L is greater than D , the dynamic algorithm continues at that leaf starting with $N_L + A_0 \leq_{\text{st.}} D-1 + G_A$ items. Note that if this event happens, there must be new items at this leaf and thus $G_A > 0$. Otherwise, there are only the initial items at L and therefore the algorithm stops.

By noting that the number of leaves of \mathcal{T}^0 whose depth $\leq p$ is less than $R_n^{0,p}$, one thus get the desired relation

$$R_n^{A,p} \leq_{\text{st.}} R_n^{0,p} + \sum_{k=1}^{R_n^{0,p}} \tilde{R}_{D-1+G_{A,i}}^{A,p} \mathbb{1}_{\{G_{A,i} > 0\}}.$$

□

Theorem 3 (Existence of a stable system). *Under Conditions **(H)** for the splitting algorithm and if A^ε is a family of integrable integer valued random variables such that*

$$\lim_{\varepsilon \rightarrow 0} A^\varepsilon = 0 \text{ in distribution and, } \limsup_{\varepsilon \rightarrow 0} \mathbb{E}(A^\varepsilon \mid A^\varepsilon > 0) < +\infty,$$

then there exists some $\varepsilon_0 > 0$ and a finite constant $C_{\mathcal{W}}^1$ such that for any $\varepsilon \leq \varepsilon_0$, $R_n^{A^\varepsilon}$ is integrable and

$$(12) \quad \mathbb{E} \left(R_n^{A^\varepsilon} \right) \leq n C_{\mathcal{W}}^1, \quad \forall n \geq 0.$$

In particular, for such an ε , the Poisson transform of the sequence $(\mathbb{E}(R_n^{A^\varepsilon}))$ is defined and differentiable on \mathbb{R} .

Note that the conditions on the family (A^ε) are quite weak to assert the stability of the algorithm for ε sufficiently small.

Proof. If F_{A^ε} is a random variable with the same distribution as $(G_{A^\varepsilon} | G_{A^\varepsilon} > 0)$,

$$\begin{aligned} \mathbb{E}(F_{A^\varepsilon}) &= \frac{1}{\mathbb{P}(G_{A^\varepsilon} > 0)} \mathbb{E}(G_{A^\varepsilon} \mathbb{1}_{\{G_{A^\varepsilon} > 0\}}) \\ &\leq \frac{\mathbb{E}(A^\varepsilon)}{\mathbb{P}(A^\varepsilon > 0)(1 - \delta)} = \mathbb{E}(A^\varepsilon | A^\varepsilon > 0) \frac{1}{1 - \delta}, \end{aligned}$$

the assumptions of the theorem imply that $\mathbb{E}(F_{A^\varepsilon})$ is bounded by some constant K as ε goes to 0.

By Theorem 3 of Mohamed and Robert [14], there exists a finite constant $C_{\mathcal{W}}^0$ such that

$$\mathbb{E}(R_n^0) \leq nC_{\mathcal{W}}^0, \quad \forall n \geq 0,$$

in particular

$$\mathbb{E}(R_{D-1+F_{A^\varepsilon}}^0) \leq (D + K)C_{\mathcal{W}}^0.$$

The variable $R_n^{A^\varepsilon, p}$ being integrable, Relation (11) gives the inequality

$$(13) \quad \mathbb{E}(R_n^{A^\varepsilon, p}) \leq \mathbb{E}(R_n^{0, p}) + \mathbb{E}(R_n^{0, p}) \mathbb{P}(G_{A^\varepsilon} > 0) \mathbb{E}(R_{D-1+F_{A^\varepsilon}}^{A^\varepsilon, p}), \quad n \geq 0,$$

and therefore the relation

$$\begin{aligned} \mathbb{E}(R_{D-1+F_{A^\varepsilon}}^{A^\varepsilon, p}) &\leq \mathbb{E}(R_{D-1+F_{A^\varepsilon}}^{0, p}) \left(1 + \mathbb{P}(G_{A^\varepsilon} > 0) \mathbb{E}(R_{D-1+F_{A^\varepsilon}}^{A^\varepsilon, p})\right) \\ &\leq (D + K)C_{\mathcal{W}}^0 \left(1 + \mathbb{P}(G_{A^\varepsilon} > 0) \mathbb{E}(R_{D-1+F_{A^\varepsilon}}^{A^\varepsilon, p})\right) \end{aligned}$$

holds. It is easy to check that the variables (G_{A^ε}) converge to 0 in distribution as ε goes to 0. Consequently, there exists some $\varepsilon_0 > 0$ such that if $\varepsilon \leq \varepsilon_0$, then $\mathbb{P}(G_{A^\varepsilon} > 0)(D + K)C_{\mathcal{W}}^0 < 1/2$, so that

$$\mathbb{E}(R_{D-1+F_{A^\varepsilon}}^{A^\varepsilon, p}) \leq 2(D + K)C_{\mathcal{W}}^0,$$

by letting p go to infinity, one gets $\mathbb{E}(R_{D-1+F_{A^\varepsilon}}^{A^\varepsilon}) \leq 2(D + K)C_{\mathcal{W}}^0$. By using again Equation (13) and Theorem, this last inequality gives the relation

$$\mathbb{E}(R_n^{A^\varepsilon, p}) \leq nC_{\mathcal{W}}^0 (1 + 2(D + K)C_{\mathcal{W}}^0),$$

the theorem is proved. \square

The above theorem shows that for ε sufficiently small the random variables $(R_n^{A^\varepsilon})$ are integrable but also that the variable $R_{A_1^\varepsilon}^{A^\varepsilon}$ is integrable, in particular the Markov chain defined by the transitions (2) is ergodic.

Corollary 4 (Stability Region for Tree Algorithms with Poisson Arrivals). *When arrivals are Poisson with parameter λ and the branching mechanism defined by \mathcal{W} satisfies Assumptions **(H)**, there exists $\lambda_{\mathcal{W}} > 0$ such that*

- (1) *If $\lambda < \lambda_{\mathcal{W}}$, the random variables (R_n^λ) are integrable.*
- (2) *If $\lambda > \lambda_{\mathcal{W}}$ then $\mathbb{E}(R_n^\lambda) = +\infty$ for all $n \geq D$.*

Proof. If A_1^λ is random variable with a Poisson distribution with parameter λ , the family (A_1^λ) clearly satisfies the assumptions of the above theorem. Hence, there exists $\lambda_0 > 0$ and a constant C such that $\mathbb{E}(R_n^\lambda) < nC$ for all $n \geq 0$.

The sum of the components of the Markov chain defined by the transitions (2) decreases at most of $D - 1$ during a time unit and new arrivals have mean λ , therefore, if $\lambda > D - 1$, then $\mathbb{E}(R_n^\lambda) = +\infty$ for all $n \geq D$. The quantity

$$\lambda_{\mathcal{W}} = \sup\{\lambda \geq 0 : \mathbb{E}(R_n^\lambda) < +\infty\}$$

is thus positive and finite. Moreover it does not depend on $n \geq D$: indeed, if $\mathbb{E}(R_n^\lambda) < +\infty$ and for $m \geq D$, if $m \leq n$, clearly $\mathbb{E}(R_m^\lambda) < \mathbb{E}(R_n^\lambda)$. If $m \geq n$, from Equation (1) one obtains that $R_n^\lambda \geq R_m^\lambda \mathbb{1}_{\{n_1+A_1=m\}}$, so that R_m^λ is integrable. Since the function $\lambda \rightarrow \mathbb{E}(R_n^\lambda)$ is non-decreasing one obtains that if $\lambda < \lambda_{\mathcal{W}}$ then, for all $n \geq 0$, the variable R_n^λ is integrable. \square

3. POISSON TRANSFORM

From now on and for the rest of the paper, it is assumed that the arrivals are Poisson with parameter λ and, as before, one writes R_n^λ instead of R_n^A . The sequence $\mathcal{N} = (t_n)$, $0 \leq t_1 \leq \dots \leq t_n \leq \dots$ is assumed to be a Poisson process with intensity 1 and, for $x \geq 0$, $\mathcal{N}([0, x])$ denotes the number of t_n 's in the interval $[0, x]$:

$$\mathcal{N}([0, x]) = \inf\{n : t_{n+1} \geq x\}.$$

The Poisson transform ϕ_r of a sequence (r_n) is given by

$$\phi_r(x) = \sum_{n \geq 0} r_n \frac{x^n}{n!} e^{-x} = \mathbb{E}(r_{\mathcal{N}([0, x])}).$$

Provided that this function is well defined on \mathbb{R} , formally

$$\phi_r'(x) = \sum_{n \geq 0} (r_{n+1} - r_n) \frac{x^n}{n!} e^{-x} = \phi_{\Delta r}(x),$$

where $\Delta r = (r_{n+1} - r_n, n \geq 0)$. In other words, the Poisson transform commutes with the differentiation: the derivative of the Poisson transform of (r_n) is the Poisson transform of the (discrete) derivative of (r_n) . The following relation is easily checked by induction, for $n \geq 0$ and $x \geq 0$,

$$(14) \quad \mathbb{E}(r_{n+\mathcal{N}([0, x])}) = \sum_{k=0}^n \binom{n}{k} \phi_r^{(k)}(x).$$

The next proposition establishes an important functional equation for the Poisson transform. For convenience, the Poisson transform of $(\mathbb{E}(R_n^\lambda))$ is denoted by ϕ_λ instead of ϕ_{R^λ} .

Proposition 5 (Poisson transform). *Provided that λ is small enough, the Poisson transform $\phi_\lambda(x)$ of the sequence (R_n^λ) satisfies the relation*

$$(15) \quad \phi_\lambda(x) = \mathbb{E}\left(\frac{1}{W_1} \phi_\lambda(\lambda + W_1 x)\right) + 1 - \phi_C(x),$$

where W_1 is a random variable with distribution \mathcal{W} and $C = (C_m)$ is the sequence defined by $C_m = 0$ for $m \geq D$ and, for $0 \leq m < D$,

$$C_m = \sum_{k=0}^m \binom{m}{k} \mathbb{E}(W_1^{k-1}) \phi_\lambda^{(k)}(\lambda)$$

Proof. From Theorem 3 and Relations (12), one gets that there exists some λ_0 such that ϕ_λ is defined on \mathbb{R} when $\lambda < \lambda_0$. By using the splitting property of Poisson processes and by including the boundary cases of Equation (1), one gets the relation

$$(16) \quad R_{\mathcal{N}([0,x])}^\lambda \stackrel{\text{dist.}}{=} 1 + \sum_{i=1}^G R_{i,\mathcal{N}([xS_{i-1,G},xS_{i,G}])+Z_i}^\lambda - \mathbb{1}_{\{\mathcal{N}([0,x]) < D\}} \sum_{i=1}^G R_{i,\mathcal{N}([xS_{i-1,G},xS_{i,G}])+Z_i}^\lambda$$

where

- For $1 \leq i \leq G$, $S_{i,G}$ is the i th partial sum of the weights

$$S_{i,G} = V_{1,G} + V_{2,G} + \cdots + V_{i,G},$$

- in particular $S_{G,G} = 1$;
- the variables $R_{i,j}$, $i \geq 1$, $j \geq 0$ are independent and $R_{i,n}$ has the same distribution as R_n for any $n \geq 0$;
- (Z_i) is an i.i.d sequence of Poisson random variables with parameter λ .

For $k \geq 0$, the homogeneity properties of Poisson processes give

$$\begin{aligned} & \mathbb{E} \left(\mathbb{1}_{\{\mathcal{N}([0,x])=k\}} \sum_{i=1}^G R_{i,\mathcal{N}([xS_{i-1,G},xS_{i,G}])+Z_i}^\lambda \right) \\ &= \mathbb{E} \left(\mathbb{1}_{\{\mathcal{N}([0,x])=k\}} \sum_{i=1}^G R_{\mathcal{N}([0,xV_{i,G}])+Z_i}^\lambda \right) = \mathbb{E} \left(\mathbb{1}_{\{\mathcal{N}([0,x])=k\}} \frac{R_{\mathcal{N}([0,xW_1])+Z_1}^\lambda}{W_1} \right) \\ &= \mathbb{E} \left(\frac{R_{\mathcal{N}([0,W_1])+Z_1}^\lambda}{W_1} \middle| \mathcal{N}([0,1]) = k \right) \frac{x^k}{k!} e^{-x}, \end{aligned}$$

where W_1 is a random variable whose distribution is \mathcal{W} . By taking the expected value of Equation (16), one gets the relation

$$\begin{aligned} \phi_\lambda(x) &= \mathbb{E} \left(\sum_{i=1}^G \phi_\lambda(\lambda + xV_{i,G}) \right) \\ &\quad - \sum_{k=0}^{D-1} \mathbb{E} \left(\frac{R_{\mathcal{N}([0,W_1])+Z_1}^\lambda}{W_1} \middle| \mathcal{N}([0,1]) = k \right) \frac{x^k}{k!} e^{-x}, \end{aligned}$$

consequently,

$$\phi_\lambda(x) = \mathbb{E} \left(\frac{1}{W_1} \phi_\lambda(\lambda + xW_1) \right) - \sum_{k=0}^{D-1} \mathbb{E} \left(\frac{R_{\mathcal{N}([0,W_1])+Z_1}^\lambda}{W_1} \middle| \mathcal{N}([0,1]) = k \right) \frac{x^k}{k!} e^{-x}.$$

For $k \geq 0$, from Equation (14)

$$\begin{aligned} \mathbb{E} \left(\frac{R_{\mathcal{N}([0,1]) + Z_1}^\lambda}{W_1} \middle| \mathcal{N}([0,1]) = k, W_1 \right) &= \sum_{\ell=0}^k \binom{k}{\ell} W_1^{\ell-1} (1-W_1)^{k-\ell} \mathbb{E} (R_{\ell+Z_1}^\lambda) \\ &= \sum_{\ell=0}^k \sum_{m=0}^{\ell} \binom{k}{\ell} \binom{\ell}{m} W_1^{\ell-1} (1-W_1)^{k-\ell} \phi_\lambda^{(m)}(\lambda) \\ &= \sum_{m=0}^k \binom{k}{m} W_1^{m-1} \phi_\lambda^{(m)}(\lambda) \end{aligned}$$

The proposition is proved. \square

An auto-regressive process with moving average. At this point, it is natural to introduce the sequence of following sequence of random variables.

Definition 6. *The process (X_n^x) is defined by $X_0^x = x$ and,*

$$(17) \quad X_n^x = W_n X_{n-1}^x + 1, \quad n \geq 1,$$

where (W_n) is an i.i.d. sequence with the same distribution as W_1 .

The sequence (X_n^x) is an auto-regressive process with moving average. These processes have interesting theoretical properties and play an important role in many areas. In the following the upper index x may be omitted when $x=0$.

The terms of the sequence (X_n^x) can be expressed as, for $n \geq 0$,

$$X_n^x = x \prod_{i=1}^n W_i + \sum_{p=1}^n \prod_{i=p+1}^n W_i = \pi_n x + X_n^0,$$

with, for $1 \leq k$, $\pi_k = \prod_{i=1}^k W_i$. The distribution of the sequence (W_i) being exchangeable, i.e. invariant under permutations, one has

$$(18) \quad X_n \stackrel{\text{dist.}}{=} X_n^* \stackrel{\text{def.}}{=} \sum_{p=0}^{n-1} \pi_p.$$

The sequence (X_n^*) converges almost surely to X_∞^* , and therefore (X_n^x) converges in distribution to the random variable X_∞ such that

$$X_\infty \stackrel{\text{dist.}}{=} W_1 X_\infty + 1 \text{ or } X_\infty \stackrel{\text{dist.}}{=} X_\infty^* = \sum_{p=0}^{+\infty} \pi_p,$$

the distribution of X_∞ is not explicitly known in general. With this notation, Equation (15) can be rewritten as

$$\phi_\lambda(x) = \mathbb{E} \left(\frac{1}{W_1} \phi \left(\lambda X_1^{x/\lambda} \right) \right) + 1 - \phi_C(x),$$

by differentiating with respect to x , one gets

$$(19) \quad \phi'_\lambda(x) = \mathbb{E} \left(\phi'_\lambda(\lambda X_1^{x/\lambda}) \right) - \phi_{\Delta C}(x).$$

Equation (19) expresses ϕ'_λ as the solution of Poisson equation associated to the Markov chain $(X_n^{x/\lambda})$ and the function $\phi_{\Delta C}$. Note that, nevertheless, the function

$\phi_{\Delta C}$ is depending on ϕ_λ through its successive derivatives at λ . By taking $x = X_\infty$ in Equation (19) and integrating, one gets

$$(20) \quad \mathbb{E}(\phi_{\Delta C}(\lambda X_\infty)) = 0.$$

The iteration of Equation (19) shows that, for $n \geq 1$,

$$\phi'_\lambda(x) = \mathbb{E} \left(\phi'_\lambda(\lambda X_n^{x/\lambda}) \right) - \sum_{k=0}^{n-1} \mathbb{E}(\phi_{\Delta C}(\pi_k x + \lambda X_k)),$$

and consequently

$$(21) \quad \phi'_\lambda(x) = C_\infty - \sum_{k=0}^{+\infty} \mathbb{E}(\phi_{\Delta C}(\pi_k x + \lambda X_k)),$$

with C_∞ defined as $\mathbb{E}(\phi'_\lambda(\lambda X_\infty))$. For $k \geq 0$, by using Relation (20) and the exchangeability property,

$$\begin{aligned} & \mathbb{E} \left[\frac{1}{\pi_k} \left(\phi_C(\pi_k x + \lambda X_k) - \phi_C(\lambda X_k) \right) \right] \\ &= \mathbb{E} \left[\frac{1}{\pi_k} \left(\phi_C(\pi_k x + \lambda X_k^*) - \phi_C(\lambda X_k^*) - \pi_k x \phi_{\Delta C}(\lambda X_k^*) \right) \right], \end{aligned}$$

and since $|X_\infty^* - X_k^*| \leq \pi_k/(1-\delta)$ and $X_k^* \leq 1/(1-\delta)$ by Assumption **(H₁)**,

$$\begin{aligned} & |\phi_C(\pi_k x + \lambda X_k^*) - \phi_C(\lambda X_k^*) - \pi_k x \phi_{\Delta C}(\lambda X_k^*)| \\ & \leq \frac{1}{2} (\pi_k x)^2 \|\phi_{\Delta^2 C}\|_\infty + \pi_k x |\phi_{\Delta C}(\lambda X_k^*) - \phi_{\Delta C}(\lambda X_\infty^*)| \\ & \leq \pi_k^2 \left(\frac{x^2}{2} + \frac{x}{1-\delta} \right) \|\phi_{\Delta^2 C}\|_\infty. \end{aligned}$$

The integration of Equation (21) term by term is therefore valid, this gives finally the following representation.

Proposition 7 (Representation of Poisson Transform). *Provided that λ is small enough, the Poisson transform $\phi_\lambda(x)$ of the sequence (R_n^λ)*

$$(22) \quad \phi_\lambda(x) = 1 + x C_\infty + \mathbb{E} \left(\sum_{k=0}^{+\infty} \frac{1}{\pi_k} \left[\phi_C(\lambda X_k) - \phi_C(\pi_k x + \lambda X_k) \right] \right),$$

where $C = (C_n)$ is the sequence defined in Proposition 5 and $C_\infty = \mathbb{E}(\phi'_\lambda(\lambda X_\infty))$,

- (1) (W_n) is a sequence of i.i.d. random variables whose distribution is \mathcal{W} .
- (2) (X_n) is the auto-regressive process defined by $X_{n+1} = W_n X_n + 1$ for $n \geq 1$ with $X_0 = 0$ and X_∞ is its limit in distribution.

4. STABILITY CONDITION

In order to get the condition to get the existence of a first moment for the sequence (R_n^λ) , one has to establish an appropriate representation of this sequence by inverting probabilistically its Poisson transform and to get an expression for the unknown constants C_0, C_1, \dots, C_{D-1} and C_∞ .

The notations of Proposition 7 are used. Let \mathcal{F}_k be the σ -field generated by the random variables W_1, \dots, W_k and \mathcal{N}_1 is another Poisson process with rate 1 independent of \mathcal{N} and (W_n) then, for $k \geq 1$,

$$\begin{aligned} \phi_C(\pi_k x + \lambda X_k) &= \mathbb{E} \left(C_{\mathcal{N}([0, x\pi_k]) + \mathcal{N}_1([0, \lambda X_k])} \mid \mathcal{F}_k \right) \\ &= \sum_{m \geq 0} \mathbb{E} \left(C_{\mathcal{N}([0, x\pi_k]) + \mathcal{N}_1([0, \lambda X_k])} \mid \mathcal{F}_k, \mathcal{N}([0, x]) = m \right) \frac{x^m}{m!} e^{-x} \\ &= \sum_{m \geq 0} \mathbb{E} \left(C_{\mathcal{N}([0, \pi_k]) + \mathcal{N}_1([0, \lambda X_k])} \mid \mathcal{F}_k, \mathcal{N}([0, 1]) = m \right) \frac{x^m}{m!} e^{-x}. \end{aligned}$$

With Equation (22), one obtains the relation

$$\begin{aligned} \phi_\lambda(x) &= 1 + xC_\infty \\ &\quad + \sum_{m \geq 0} \frac{x^m}{m!} e^{-x} \sum_{k \geq 0} \mathbb{E} \left(\frac{1}{\pi_k} \left[C_{\mathcal{N}([0, \lambda X_k])} - C_{\mathcal{U}_m([0, \pi_k]) + \mathcal{N}([0, \lambda X_k])} \right] \right), \end{aligned}$$

where, if U_1, \dots, U_n are n i.i.d. random variables uniformly distributed on $[0, 1]$, for $0 \leq x \leq 1$, $\mathcal{U}_n([0, x])$ denotes the number of U_k 's in the interval $[0, x]$. These variables are ordered as $U_{(1)}^n \leq U_{(2)}^n \leq \dots \leq U_{(n)}^n$, in particular, for $m \geq 1$, $\{\mathcal{U}_n([0, x]) \geq m\} = \{U_{(m)}^n \leq x\}$. By identifying the coefficients of the above expression, one gets the following proposition.

Proposition 8 (Representation of the Average Size of the Tree). *Under Assumptions (H) and for λ sufficiently small*

$$(23) \quad \mathbb{E}(R_n^\lambda) = 1 + nC_\infty + \sum_{k \geq 0} \mathbb{E} \left(\frac{1}{\pi_k} \left[C_{\mathcal{N}([0, \lambda X_k])} - C_{\mathcal{U}_n([0, \pi_k]) + \mathcal{N}([0, \lambda X_k])} \right] \right), \quad n \geq 0.$$

where \mathcal{N} is a Poisson process with rate 1 and

- (1) $C = (C_n)$ is the sequence defined in Proposition 5 and $C_\infty = \mathbb{E}(\phi'_\lambda(\lambda X_\infty))$.
- (2) For uniformly distributed random variables $(U_i, 1 \leq i \leq n)$ on $[0, 1]$ and $0 \leq x \leq 1$, the quantity $\mathcal{U}_n([0, x])$ denotes the number of U_i 's in the interval $[0, x]$.

Determination of the constants. In order to get an explicit representation of the sequence $(\mathbb{E}(R_n^\lambda))$, the D unknown coefficients C_0, \dots, C_{D-1} (recall that the other C_k 's are null) and the constant $C_\infty = \mathbb{E}(\phi'_\lambda(\lambda X_\infty))$ have to be determined. The method used by Fayolle *et al.* [8] to determine these coefficients in the binary case cannot be, apparently, extended to other tree structures.

- i) The boundary conditions $\mathbb{E}(R_m^\lambda) = 1$ for $1 \leq m \leq D-1$ translate into $D-1$ linear equations involving these $D+1$ unknown constants,

$$(24) \quad \sum_{\ell=0}^{D-1} M_{m,\ell}^\lambda C_\ell + M_{m,D}^\lambda C_\infty = 0, \quad 1 \leq m \leq D-1,$$

with, for $1 \leq m \leq D-1, 0 \leq \ell \leq D-1$,

$$M_{m,\ell}^\lambda = \sum_{k \geq 0} \mathbb{E} \left(\frac{1}{\pi_k} \left[\mathbb{1}_{\{\mathcal{N}([0, \lambda X_k]) = \ell\}} - \mathbb{1}_{\{\mathcal{U}_m([0, \pi_k]) + \mathcal{N}([0, \lambda X_k]) = \ell\}} \right] \right)$$

and $M_{m,D}^\lambda = m$.

ii) Equation (20) gives the additional relation

$$(25) \quad M_{D,0}^\lambda C_0 + M_{D,1}^\lambda C_1 + \cdots + M_{D,D-1}^\lambda C_{D-1} = 0,$$

with

$$M_{D,\ell}^\lambda = \mathbb{E} \left(\frac{(\lambda X_\infty)^\ell}{\ell!} (\lambda \ell X_\infty^{-1} - 1) e^{-\lambda X_\infty} \right), \quad 0 \leq \ell \leq D-1,$$

and $M_{D,D}^\lambda = 0$.

iii) The final equation is obtained by plugging $x = \lambda$ in Equation (22) so that, since $C_0 = \mathbb{E}(G)\phi(\lambda)$ by Proposition 5,

$$(26) \quad -1 = \lambda C_\infty - \frac{1}{\mathbb{E}(G)} C_0 + \sum_{m=0}^{D-1} C_m \mathbb{E} \left(\sum_{k=0}^{+\infty} \frac{1}{\pi_k} \left[\frac{(\lambda X_k)^m}{m!} e^{-\lambda X_k} - \frac{(\pi_k \lambda + \lambda X_k)^m}{m!} e^{-(\pi_k \lambda + \lambda X_k)} \right] \right).$$

The Matrix M_λ . The square matrix $M_\lambda = (M_{m,\ell}^\lambda, 1 \leq m \leq D+1, 0 \leq \ell \leq D)$ is defined as follows:

$$\left\{ \begin{array}{l} M_{m,\ell}^\lambda = \sum_{k \geq 0} \mathbb{E} \left(\frac{1}{\pi_k} \left[\mathbb{1}_{\{\mathcal{N}([0, \lambda X_k]) = \ell\}} - \mathbb{1}_{\{\mathcal{U}_m([0, \pi_k]) + \mathcal{N}([0, \lambda X_k]) = \ell\}} \right] \right), \quad m < D, \ell \neq D, \\ M_{D,\ell}^\lambda = \mathbb{E} \left(\frac{(\lambda X_\infty)^\ell}{\ell!} [\lambda \ell X_\infty - 1] e^{-\lambda X_\infty} \right), \quad 0 \leq \ell \leq D-1, \\ M_{D+1,\ell}^\lambda = \mathbb{E} \left(\sum_{k=0}^{+\infty} \frac{1}{\pi_k} \left[X_k^\ell - (\pi_k + X_k)^\ell e^{-\lambda \pi_k} \right] \frac{\lambda^\ell}{\ell!} e^{-\lambda X_k} \right), \quad 1 \leq \ell \leq D-1, \\ M_{D+1,0}^\lambda = \mathbb{E} \left(\sum_{k=0}^{+\infty} \frac{1}{\pi_k} [1 - e^{-\lambda \pi_k}] e^{-\lambda X_k} \right) - \frac{1}{\mathbb{E}(G)}, \\ M_{D,D}^\lambda = 0, M_{D+1,D}^\lambda = \lambda. \end{array} \right.$$

By gathering Equations (24), (25) and (26) and denoting $e_{D+1} = (0, 0, \dots, 0, 1)$ and by $\underline{C} = (C_0, C_1, \dots, C_{D-1}, C_\infty)$ the vector of constants, one gets the linear relation

$$(27) \quad M_\lambda \cdot \underline{C} = -e_{D+1}.$$

The following theorem is the main result concerning the ergodicity of the tree algorithm.

Theorem 9 (Stability of Tree Algorithm with Poisson Arrivals). *Under Assumptions (H) for the splitting distribution \mathcal{W} , if M_λ is the matrix defined above and*

$$\lambda_c = \inf\{\lambda > 0 : \det M_\lambda = 0\},$$

then $\lambda_c > 0$ and for any $\lambda < \lambda_c$, the size of the tree associated to the tree algorithm is integrable.

Proof. With the same notations as before, for $\lambda = 0$, then $X_k = 0$ for $0 \leq k \leq +\infty$ and the D th and $(D + 1)$ th rows of the matrix \mathcal{M}_λ are given by

$$M_D = (-1, 1, 0, \dots, 0) \text{ and } M_{D+1} = (-1/\mathbb{E}(G), 0, \dots, 0, 0)$$

by expanding with respect to these rows, one gets

$$\det M_0 = -\frac{1}{\mathbb{E}(G)} \begin{vmatrix} M_{1,2}^0 & M_{1,3}^0 & \dots & M_{1,D-1}^0 & 1 \\ M_{2,2}^0 & M_{2,3}^0 & \dots & M_{2,D-1}^0 & 2 \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ M_{D-1,2}^0 & M_{D-1,3}^0 & \dots & M_{D-1,D-1}^0 & D-1 \end{vmatrix}.$$

Since, for $1 \leq m, \ell \leq D - 1$,

$$M_{m,\ell}^0 = -\sum_{k \geq 0} \mathbb{E} \left(\frac{1}{\pi_k} \mathbb{1}_{\{U_m([0, \pi_k]) = \ell\}} \right),$$

then, $M_{m,\ell}^0 = 0$ for $\ell > m$ hence,

$$\det \mathcal{M}_0 = \frac{-1}{\mathbb{E}(G)} M_{2,2}^0 \dots M_{D-1,D-1}^0 \neq 0.$$

Due to the explicit expression of the matrix M_λ , the function $\lambda \rightarrow \det M_\lambda$ is clearly continuous so that $\lambda_c > 0$.

Corollary 4 shows the existence of some constant $\lambda_{\mathcal{W}}$ such that, for $\lambda < \lambda_{\mathcal{W}}$, random variables (R_n^λ) are integrable and their expected values are given by Equation (23) and the constant vector C in this expression satisfies Equation (27). Hence, for $\lambda < \lambda_{\mathcal{W}} \wedge \lambda_c$, there exists a unique C such that Equation (23) holds for $\mathbb{E}(R_n^\lambda)$ for $n \geq 0$. Since the function $\lambda \rightarrow \mathbb{E}(R_n^\lambda)$ is non-decreasing and because of the existence of the a solution to Equation (27) $\lambda < \lambda_c$, the expression given by Equation (23) is finite for any $\lambda < \lambda_c$, one concludes necessarily by Corollary 4 that $\lambda_c \leq \lambda_{\mathcal{W}}$. The theorem is proved. \square

Remarks.

- (1) It is very likely that $\lambda_{\mathcal{W}}$ defined in Proposition 4 and λ_c are equal, but we have not been able to prove it. For $\lambda = \lambda_c$, at least one of the determinants of the Cramer Formula should be non zero which would imply that at least one of the (C_k) is infinite and therefore that the random variables (R_n^λ) are not integrable.
- (2) The coefficients of the matrix M_λ are expressed in terms of the distribution of the auto-regressive process (X_n) . An explicit, usable, representation of this distribution is available mostly through Laplace transform functionals, the invariant distribution included.

Despite it is not easy to handle, the introduction of the auto-regressive process is, in our opinion, the key ingredient in our analysis. It plays a major role in Representation (23) of the sequence $(\mathbb{E}(R_n^\lambda))$. It should also be kept in mind that one relation used to determine the constants (C_k) is Equation (20) which comes directly from the fact that (X_n) has an equilibrium distribution. In a purely analytic setting (i.e. without this probabilistic representation), an analogous equation would probably require some spectral analysis in a functional space.

Examples.

a) *Static Case.*

In this case $\lambda = 0$ and the components of the vector $C = (C_i, 0 \leq i \leq D - 1)$ are constant and equal to the average branching degree $\mathbb{E}(G) = \mathbb{E}(W_1)$. For $n \geq D$,

$$\mathbb{E}(R_n^0) = 1 + \mathbb{E}(G) \sum_{k \geq 0} \mathbb{E} \left(\frac{1}{\pi_k} \mathbb{1}_{\{U_{(D)}^n \leq \pi_k\}} \right),$$

which is the expression established for the static splitting algorithm in Mohamed and Robert [14].

b) *Binary Tree Algorithm.*

In the binary case, $G \equiv 2$, $D = 2$ and the splitting measure is $\mathcal{W} = p\delta_p + q\delta_q$. In this case the average cost of the algorithm is expressed as follows

$$\begin{aligned} \mathbb{E}(R_n^\lambda) = & 1 + nC_\infty - \phi'_\lambda(\lambda) \sum_{k \geq 0} \mathbb{E} \left(\frac{e^{-\lambda X_k}}{\pi_k} \mathbb{1}_{\{U_{(1)}^n \leq \pi_k\}} \right) \\ & + (2\phi_\lambda(\lambda) + \phi'_\lambda(\lambda)) \left(\sum_{k \geq 0} \mathbb{E} \left(\frac{\lambda X_k e^{-\lambda X_k}}{\pi_k} \mathbb{1}_{\{U_{(1)}^n \leq \pi_k\}} \right) \right. \\ & \left. + \sum_{k \geq 0} \mathbb{E} \left(\frac{e^{-\lambda X_k}}{\pi_k} \mathbb{1}_{\{U_{(2)}^n \leq \pi_k\}} \right) \right). \end{aligned}$$

The two coefficients C_0 and C_1 satisfy

$$C_0 = 2\phi_\lambda(\lambda), \quad C_1 = 2\phi_\lambda(\lambda) + \phi'_\lambda(\lambda).$$

Equation (20) implies that

$$[\mathbb{E}(e^{-\lambda X_\infty}) - \lambda \mathbb{E}(X_\infty e^{-\lambda X_\infty})] C_1 = \mathbb{E}(e^{-\lambda X_\infty}) C_0,$$

which gives the relation

$$(28) \quad \phi'_\lambda(\lambda) = 2 \left(\frac{\mathbb{E}(e^{-\lambda X_\infty})}{\mathbb{E}(e^{-\lambda X_\infty}) - \lambda \mathbb{E}(X_\infty e^{-\lambda X_\infty})} - 1 \right) \phi_\lambda(\lambda).$$

Note that in the case of the symmetric binary algorithm, the limit X_∞ is constant and equal to 2.

A identity similar to Equation (28) has been established in Fayolle *et al.* [9]. By taking advantage of the fact that if one plugs successively $x = \lambda/p$ and $x = \lambda/q$ successively in Equation (15) (Equation (6) in this case), one gets the relation

$$\phi'_\lambda(\lambda) = 2(K - 1) \phi_\lambda(\lambda),$$

where

$$K = \begin{cases} -\frac{e^{-\lambda/p} - e^{-\lambda/q}}{\frac{\lambda}{p}e^{-\lambda/p} - \frac{\lambda}{q}e^{-\lambda/q}} & \text{if } p \neq 1/2, \\ 1/(1 - 2\lambda) & \text{otherwise.} \end{cases}$$

This trick turns out to be specific of binary trees and does not seem to have a generalization for other random trees. Interestingly, when $p \neq 1/2$, the representation of

the constant K by Equation (28) gives the following relation for $g(\lambda)$, the Laplace transform of X_∞ at λ ,

$$-\frac{g'(\lambda)}{g(\lambda)} = \frac{(1 + \lambda/p)e^{-\lambda/p} - (1 + \lambda/q)e^{-\lambda/q}}{\lambda(e^{-\lambda/p} - e^{-\lambda/q})},$$

which can be solved as

$$(29) \quad \mathbb{E}(e^{-\lambda X_\infty}) = \frac{1}{1/(1-p) - 1/p} \frac{e^{-\lambda/p} - e^{-\lambda/(1-p)}}{\lambda},$$

which gives an explicit representation of the Laplace transform of the invariant measure of the auto-regressive process in this case.

5. ASYMPTOTIC ANALYSIS OF THE AVERAGE SIZE OF THE TREE

In this section, the asymptotic behavior of the sequence $(\mathbb{E}(R_n^\lambda))$ is investigated for $\lambda < \lambda_c$ where λ_c is defined in Theorem 9. The goal is to establish an analogue of the law of large numbers for these expected values. As noted before, Fayolle *et al.* [8] (for the binary tree) is the only rigorous result we know in this domain.

Equation (23) gives the representation, for $n \geq D$,

$$(30) \quad \mathbb{E}(R_n^\lambda) = 1 + nC_\infty - \sum_{i=0}^{D-1} \sum_{k \geq 0} \mathbb{E} \left(\frac{1}{\pi_k} \Delta C_{i+\mathcal{N}([0, \lambda X_k])} \mathbb{1}_{\{U_{(i+1)}^n \leq \pi_k\}} \right),$$

where (ΔC_i) is the sequence $(C_{i+1} - C_i)$ and, for $0 \leq i \leq D-1$ and $n \geq 1$. Recall that $U_{(i)}^n$ is the i th smallest term of n independent uniform random variables on $[0, 1]$.

In a first step, it is shown that the series associated to $i = 0$ in Equation (30) is vanishing for the asymptotic behavior of the sequence $(\mathbb{E}(R_n^\lambda)/n)$. This is a crucial result since the arguments to derive a law of large numbers rely on an integrability property which is not satisfied for this series.

Lemma 2. *Under Assumptions (H), the relation*

$$\lim_{n \rightarrow +\infty} \frac{1}{n} \sum_{k \geq 0} \mathbb{E} \left(\frac{1}{\pi_k} \Delta C_{\mathcal{N}([0, \lambda X_k])} \mathbb{1}_{\{U_{(1)}^n \leq \pi_k\}} \right) = 0$$

holds.

Proof. Equation (19) gives the relation

$$\begin{aligned} A_n &\stackrel{\text{def.}}{=} \sum_{k \geq 0} \mathbb{E} \left(\frac{1}{\pi_k} \Delta C_{\mathcal{N}([0, \lambda X_k])} \mathbb{1}_{\{U_{(1)}^n \leq \pi_k\}} \right) = \sum_{k \geq 0} \mathbb{E} \left(\frac{1}{\pi_k} \phi_{\Delta C}(\lambda X_k) \mathbb{1}_{\{U_{(1)}^n \leq \pi_k\}} \right) \\ &= \sum_{k \geq 0} \mathbb{E} \left(\frac{1}{\pi_k} \left[\Delta R_{\mathcal{N}([0, \lambda X_{k+1}])}^\lambda - \Delta R_{\mathcal{N}([0, \lambda X_k])}^\lambda \right] \mathbb{1}_{\{U_{(1)}^n \leq \pi_k\}} \right) \\ &\leq \sup_{0 \leq x \leq \lambda/(1-\delta)} \phi_{\Delta R^\lambda}(x) \cdot \sum_{k \geq 0} \mathbb{E} \left(\frac{1}{\pi_k} \mathbb{1}_{\{\mathcal{N}([\lambda X_k, \lambda X_{k+1}]) \neq 0, U_{(1)}^n \leq \pi_k\}} \right), \end{aligned}$$

by Assumption **(H₁)**. For $k \geq 0$, by exchangeability of the sequence (W_i) and Definition (18), one gets

$$\begin{aligned} \mathbb{E} \left(\frac{1}{\pi_k} \mathbb{1}_{\{\mathcal{N}([\lambda X_k, \lambda X_{k+1}]) \neq 0, U_{(1)}^n \leq \pi_k\}} \right) &= \mathbb{E} \left(\frac{\pi_1}{\pi_{k+1}} \mathbb{1}_{\{\mathcal{N}([\lambda X_k^*, \lambda X_{k+1}^*]) \neq 0, U_{(1)}^n \leq \pi_{k+1}/\pi_1\}} \right) \\ &\leq \delta \mathbb{E} \left(\frac{1}{W_{k+1}} \mathbb{1}_{\{W_1 U_{(1)}^n \leq \delta^{k+1}\}} \right) = \delta \mathbb{E}(G) \mathbb{P} \left(W_1 U_{(1)}^n \leq \delta^{k+1} \right). \end{aligned}$$

By summing up this terms, this gives the following upper bound for A_n , for some constant C ,

$$A_n \leq C \sum_{k \geq 0} \mathbb{P} \left(W_1 U_{(1)}^n \leq \delta^{k+1} \right) \leq C \mathbb{E} \left(\left[-\log_{1/\delta} \left(W_1 U_{(1)}^n \right) \right] \right),$$

Since this term is of the order of $\log n$, the sequence (A_n/n) converges to 0. \square

Before analyzing the asymptotic behavior of $(\mathbb{E}(R_n^\lambda)/n)$, Propositions 9 and 11 from Mohamed and Robert [14] obtained in the static case are summarized in the following proposition.

Proposition 10. *Under Assumption **(H)**, for $i \geq 1$, if*

$$E_{i,n} \stackrel{\text{def.}}{=} \sum_{k \geq 0} \mathbb{E} \left(\frac{1}{\pi_k} \mathbb{1}_{\{U_{(i+1)}^n \leq \pi_k\}} \right)$$

(1) *If the random variable $-\log W_1$ is non-arithmetic, then*

$$\lim_{n \rightarrow +\infty} \frac{E_{i,n}}{n} = \frac{\mathbb{E}(G)}{i \mathbb{E}(|\log W_1|)}.$$

(2) *If the random variable $-\log W_1$ is arithmetic and $\xi > 0$ is its span, then, as n goes to infinity,*

$$\lim_{n \rightarrow +\infty} \frac{E_{i,n}}{n} - F_i \left(\frac{\log n}{\xi} \right) = 0,$$

where F_i is the periodic function defined by

$$(31) \quad F_i(x) = \frac{\mathbb{E}(G)}{\mathbb{E}(|\log(W_1)|)} \frac{\xi}{1 - e^{-\xi}} \int_0^{+\infty} \exp \left(-\xi \left\{ x - \frac{\log y}{\xi} \right\} \right) \frac{y^{i-1}}{i!} e^{-y} dy,$$

and $\{x\} = x - \lfloor x \rfloor$.

The next proposition “decouples” the process (X_n) and the counting process associated to the sequence (π_k) .

Proposition 11. *Under Assumption **(H)**, for $1 \leq i \leq D - 1$,*

$$\lim_{n \rightarrow +\infty} \frac{1}{n} \sum_{k \geq 0} \mathbb{E} \left(\frac{1}{\pi_k} \mathbb{1}_{\{U_{(i+1)}^n \leq \pi_k\}} \left[\Delta C_{i+\mathcal{N}([0, \lambda X_k])} - \mathbb{E} \left(\Delta C_{i+\mathcal{N}([0, \lambda X_\infty])} \right) \right] \right) = 0.$$

Proof. By using Definition (18) and the exchangeability property, one has for $p \geq 1$,

$$\begin{aligned}
& \frac{1}{n} \sum_{k \geq p} \mathbb{E} \left(\frac{1}{\pi_k} \left| \Delta C_{i+\mathcal{N}([0, \lambda X_k])} - \Delta C_{i+\mathcal{N}([0, \lambda X_p])} \right| \mathbb{1}_{\{U_{(i+1)}^n \leq \pi_k\}} \right) \\
&= \frac{1}{n} \sum_{k \geq p} \mathbb{E} \left(\frac{1}{\pi_k} \left| \Delta C_{i+\mathcal{N}([0, \lambda X_k^*])} - \Delta C_{i+\mathcal{N}([0, \lambda X_p^*])} \right| \mathbb{1}_{\{U_{(i+1)}^n \leq \pi_k\}} \right) \\
&\leq \frac{1}{n} \|\Delta C\|_\infty \sum_{k \geq 0} \mathbb{E} \left(\frac{1}{\pi_k} \mathbb{P}(\mathcal{N}([\lambda X_k^*, \lambda X_p^*]) \neq 0 \mid \mathcal{F}_k) \mathbb{1}_{\{U_{(i+1)}^n \leq \pi_k\}} \right) \\
&= \frac{1}{n} \|\Delta C\|_\infty \mathbb{E} \left(\sum_{k \geq 0} \frac{1}{\pi_k} \left(1 - e^{-\lambda |X_k^* - X_p^*|} \right) \mathbb{1}_{\{U_{(i+1)}^n \leq \pi_k\}} \right) \\
&\leq \|\Delta C\|_\infty (1 - \exp(-\lambda \delta^p / (1 - \delta))) E_{i,n},
\end{aligned}$$

by Assumption **(H₁)**, where $E_{i,n}$ is defined in Proposition 10. One can therefore choose a p sufficiently large so that the above difference is arbitrarily small, uniformly in $n \geq 1$.

One has thus to investigate the asymptotic behavior of

$$\frac{1}{n} \mathbb{E} \left(\Delta C_{i+\mathcal{N}([0, \lambda X_p])} \sum_{k \geq p} \frac{1}{\pi_k} \mathbb{1}_{\{U_{(i+1)}^n \leq \pi_k\}} \right) = \mathbb{E} (\Delta C_{i+\mathcal{N}([0, \lambda X_p])} G_p(n)),$$

with

$$G_p(n) = \mathbb{E} \left(\frac{1}{n} \sum_{k \geq p} \frac{1}{\pi_k} \mathbb{1}_{\{U_{(i+1)}^n \leq \pi_k\}} \middle| \mathcal{F}_p \right).$$

When $-\log W_1$ is non-arithmetic, Proposition 10 shows that $\mathbb{E}(G_p(n))$ converges. With the same argument as in Mohamed and Robert, the conditioning being on the first p elements of the sequence (W_k) , almost surely, $G_p(n)$ converges to the same limit as the sequence $(\mathbb{E}(G_p(n)))$. Consequently, by denoting x^+ the non-negative part of $x \in \mathbb{R}$,

$$\mathbb{E}(|\mathbb{E}(G_p(n)) - G_p(n)|) = 2\mathbb{E}([\mathbb{E}(G_p(n)) - G_p(n)]^+)$$

since the quantity $[\mathbb{E}(G_p(n)) - G_p(n)]^+$ is bounded by $\sup_n \mathbb{E}(G_p(n))$, Lebesgue's Theorem gives that the sequence $(\mathbb{E}(G_p(n)) - G_p(n))$ converges to 0 in L_1 . Therefore, the quantity

$$\left| \mathbb{E} \left(\Delta C_{i+\mathcal{N}([0, \lambda X_p])} [G_p(n) - \mathbb{E}(G_p(n))] \right) \right| \leq \|\Delta C\|_\infty \mathbb{E}(|G_p(n) - \mathbb{E}(G_p(n))|)$$

converges to 0 as n goes to infinity. The proposition is therefore proved in this case. When $-\log W_1$ is arithmetic with range $\xi\mathbb{N}$, the argument is similar by using the fact that the convergence to 0 of $(G_p(n) - F_i(\log n/\xi))$ holds almost surely and for the expected value. \square

The main result of this section can now be stated. It is a direct consequence of Representation (30), Lemma 2, Proposition 10 and Proposition 11.

Theorem 12. *If $\lambda < \lambda_c$ defined in Theorem 9 and under Assumption **(H)**,*

(1) if the random variable $-\log W_1$ is non-arithmetic, then

$$\lim_{n \rightarrow +\infty} \frac{\mathbb{E}(R_n^\lambda)}{n} = C_\infty - \frac{\mathbb{E}(G)}{\mathbb{E}(|\log W_1|)} \sum_{i=1}^{D-1} \frac{1}{i} \mathbb{E}(\Delta C_{i+\mathcal{N}([0, \lambda X_\infty])}).$$

(2) If the random variable $-\log W_1$ is arithmetic and $\xi > 0$ is its span, then

$$\lim_{n \rightarrow +\infty} \frac{\mathbb{E}(R_n^\lambda)}{n} - C_\infty - \sum_{i=1}^{D-1} \frac{1}{i} \mathbb{E}(\Delta C_{i+\mathcal{N}([0, \lambda X_\infty])}) F_i\left(\frac{\log n}{\xi}\right) = 0,$$

($F_i, 1 \leq i \leq D-1$) being the periodic functions defined by Equation (31),

where

- for $i \geq 1$, $\Delta C_i = C_{i+1} - C_i$ with $\underline{C} = (C_0, C_1, \dots, C_{D-1}, C_\infty)$ being the vector solution of the equation

$$M_\lambda \cdot \underline{C} = -e_{D+1},$$

with $C_k = 0$ for $k \geq D$ and M_λ is the matrix above Equation (27).

- \mathcal{N} is a Poisson point process with rate 1.
- The variable X_∞ has the invariant distribution of the auto-regressive process (X_n) defined by $X_{n+1} = W_n X_n + 1$, $n \geq 0$.

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E-mail address: Hanene.Mohamed@math.uvsq.fr

URL: <http://www.math.uvsq.fr/~mohamed/>

(H. Mohamed) LMV, UNIVERSITE DE VERSAILLES-SAINT-QUENTIN-EN-YVELINES, 78035 VERSAILLES CEDEX, FRANCE

E-mail address: Philippe.Robert@inria.fr

URL: <http://www-rocq.inria.fr/~robert>

(Ph. Robert) INRIA, DOMAINE DE VOLUCEAU, B.P. 105, 78153 LE CHESNAY CEDEX, FRANCE