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**Decentralization in Africa and the nature of local governments' competition:
evidence from Benin**

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Abstract

Without denying particular dimensions of the decentralisation in Sub-Saharan countries, this paper applies standard reasoning from the fiscal federalism literature to a developing country and tests the existence of strategic interactions among local Beninese governments, called '*communes*'. We first propose a two-jurisdiction model of public expenditure interactions, considering a constrained Nash equilibrium to capture the extreme poverty of some *communes*. We show that spillovers among jurisdictions involve strategic behaviours of local officials who have sufficient levels of fiscal resources. Second, by estimating a spatial lag model, our analysis provides evidence for the presence of strategic interactions in Benin, contingent on *communes*' fiscal autonomy. Such interactions arise among *communes* which are geographically or ethnically close. We also highlight both an opportunistic behaviour of local governments before local elections and an effect of partisan affiliations. This African democracy appears to be as concerned as developed democracies with strategic fiscal interactions.

JEL Classification: D72, H2, H7

Keywords: Fiscal interactions, Benin, decentralisation, local government, dynamic panel data.

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1 Introduction

Decentralisation has recently been embraced by a large number of developing countries, especially in Africa, since it has been presented as a powerful tool to reduce poverty and improve governance. The World Bank in particular views it as one of the major reforms on its agenda. In response to the failure of a central state to run the countries' development or to limit the risk of civil conflicts in ethnically fragmented countries, decentralisation is perceived as a way to ensure political stability, to improve accountability and responsiveness of local leaders, to increase the efficiency of public policy, and ultimately to reduce poverty.

Two principal and non-exclusive arguments might explain this infatuation with decentralisation in developing countries. The first one is what we call the 'proximity principle', given that decentralisation moves governments closer to citizens. Oates (1972) decentralisation theorem states that decentralisation improves preference matching by offering a greater diversity of public services to a heterogeneous population. Moreover, by reducing informational asymmetries between those in power and those governed, decentralisation should induce a higher accountability of governments and in fine a better efficiency in public spending. The second main argument in favour of decentralisation dates at least from Tiebout (1961) and may be called the 'competition principle'. Indeed, decentralisation is supposed to induce some interjurisdictional competition among political powers: 'voting with feet' and 'yardstick' competition (Salmon, 1987) are another ways to increase the efficiency of public spending.

Bardhan (2002) argued, however, that 'the institutional context (and therefore the structure of incentives and organisation) in developing and transition economies is quite different from those in advanced industrial economies', and he recommended us 'to go beyond the traditional fiscal federalism literature'. The reason is that most developing countries do not meet implicit or explicit assumptions posed by fiscal federalism theory. For instance, the Tiebout model cannot be applied to developing countries where the population mobility appears to be strongly limited. The existence of a yardstick competition appears at least debatable in the context of young democracies. Moreover, apart from the corruption issue emphasised by Prud'homme (1995) or Bardhan and Mookherjee (2005), these countries face

some administrative capacity constraints that the rich countries do not suffer.

These pitfalls have induced the literature on decentralisation in developing countries to focus on the effectiveness of the first argument, the ‘proximity principle’. For instance, Faguet (2004) shows that decentralisation in Bolivia has improved the responsiveness of public investment to local needs. Alderman (2002) established that local officials in Albania manage anti-poverty programmes more accurately and cost-effectively than a central government agency since they are better informed. Bardhan and Mookherjee (2005) and Galasso and Ravallion (2005) have also highlighted that decentralisation improves anti-poverty systems in particular through better intra-regional targeting. These analyses suggest that decentralisation may lead to poverty reduction from the bottom up.¹ None of these authors, however, consider the second aspect of decentralisation, the ‘competition principle’, which induces some jurisdictions’ interactions.

Do we have to believe that no interaction exists among local governments in developing countries and that decentralisation is only guided by the ‘proximity principle’ argument? The ‘departure from the classic fiscal federalism literature’ advocated by Bardhan (2002) might have been followed too far.² The aim of this paper is to come back to the ‘competition principle’ by examining interactions among local governments in a developing country. Without ignoring some of the specificities of developing countries, such as the extreme poverty of some jurisdictions, for instance, we apply some standard reasoning of the fiscal federalism literature to a developing country, Benin, which has recently experienced a decentralisation process.

Considering the degree of freedom of local governments in developing countries, we focus on the public expenditure side and not on the revenues side to shed light on jurisdictional interactions. With respect to the huge literature on tax competition the analysis of public spending interaction seems very limited: for instance, Kelejian and Prucha (1998), Redoano (2007) and Foucault, Madies, and Paty (2008) respectively study the US, the EU and France. For developing countries this kind of work is even rarer: Akin, Hutchinson, and Strumpf

¹ The relationship between decentralisation and poverty alleviation has been reviewed by Klugman (1997) and Bird and Rodriguez (1999).

² Chavis (2009) studies the effect of competition on decentralisation efficiency in Indonesia. The author considers the extent to which the cost per square metre of road project decreases in the number of villages which compete to obtain grants from the central government. The appreciation of competition is limited to the number of competitors. There is no analysis of interactions.

(2005) analyse decentralised health care in Uganda; Arze, Martinez-Vasquez, and Puwanti (2008) investigate local public spending in Indonesia.

The aim of this paper is first to establish the existence (or not) of interactions among local Beninese governments, and second to identify its nature. This last notion directly refers to the industrial organisation literature and the well-established notion of strategic complementarity or strategic substitutability provided by Bulow, Geanakoplos, and Klemperer (1985). We intend to tackle the following issues: (1) does decentralisation induce interjurisdictional interactions in Benin? (2) if yes, what is the nature of this competition (strategic complements or substitutes): does an increase in the spending of a jurisdiction induce a similar or inversed variation in neighbouring jurisdictions? (3) which factors (geographical, political or ethnic) help to boost this competition? (4) finally, what happens during election periods?

Our empirical analysis through a spatial lag model confirms the existence of strategic interactions among Beninese local governments. Moreover public spending choices are strategic complements. These interactions, which are contingent on *communes*' resources, exist not only among neighbouring *communes* but also among those close in terms of ethnic composition. Lastly, our analysis highlights an effect of partisan affiliations and opportunistic behaviour of local governments, which increase public spending in pre-electoral periods.

The remainder of the paper is divided into three sections. Section 2 develops a theoretical analysis of local public spending interactions which takes into account resource constraints of some local governments. In Section 3, after a brief overview of Benin we test the existence of interactions among Beninese local governments between 2002 and 2008; we also identify the fiscal setting behaviour of local governments during election periods. Section 4 concludes.

2 Theoretical background

In this section, we present a theoretical model in order to highlight behaviours at play in determining the levels of public spending of local governments in a developing country. Our approach is very standard regarding the literature on fiscal federalism, but we take into account some earmarks of developing countries. In particular we develop a constrained Nash equilibrium in order to capture the extreme poverty of some *communes*. Indeed, the poorest

communes in Benin, Bassila, Cibly, Kandi or Kari-Mama, for instance, respectively display an average level of annual resources of 168, 526, 734 and 861 FCFA per capita (respectively equivalent to \$0.31, \$0.97, \$1.35 and \$1.58). Beyond their realism, our assumptions have to be consistent with our empirical tools, in particular those of spatial econometrics, which we use in the next section.

2.1 The model

We consider two jurisdictions (i and j) of the same level. We do not study political issues and then adopt a normative approach. The utility function of a representative individual in jurisdiction i is given by $W^i(x_i, g_i, \theta_{ij}g_j)$, where x_i is the private consumption, g_i the public spending in jurisdiction i , and θ_{ij} is an exogenous non negative parameter, which represents the degree of spillover effect for inhabitants in jurisdiction i from the public good provided in jurisdiction j . We may have situations where spillovers are not symmetric ($\theta_{ij} \neq \theta_{ji}$).³ We define $\theta = (\theta_{ij}, \theta_{ji})$.

Since spatial empirics use weighting matrices for the strategic variables ($g_{i,j}$), the unique consistent aggregation technology of local public goods is the weighted summation. Thus, it follows:

$$W^i(x_i, g_i, \theta_{ij}g_j) = V^i(x_i, g_i + \theta_{ij}g_j),$$

where the weight is the parameter θ_{ij} .

Our analysis focuses exclusively on current local public spending, since it is better controlled by local governments than investment expenditures. Indeed these latter are often ordered and financed by central government. The current spending is a mix of public and merit goods. We are not able to say whether local public spending is a complement to or substitute for private consumption. Thus, without loss of generality concerning our analysis of jurisdictions' interactions, we will consider a quasi-linear utility function, that is, local public spending and private consumption are Edgeworth-independent:

$$V^i(x_i, g_i + \theta_{ij}g_j) = x_i + v_i(g_i + \theta_{ij}g_j),$$

³ This assumption is linked to our empirical work too. Since proximity matrices are normalised, their sum is equal to unity for each i . Thus, we have $\theta_{ij} = \theta_{ji}$ if and only if jurisdictions i and j have the same number of neighbours for a proximity matrix based on contiguity.

where the function $v_i(\cdot)$ is the appreciation of local public goods in jurisdiction i . This function is increasing in its argument $v_i'(\cdot) > 0$. The sign of its second derivative, however, remains indeterminate. Indeed the concavity of function $v_i(\cdot)$, which is often assumed in the literature, would restrict our theoretical analysis of jurisdictional interactions to the case of strategic substitutes.

We ignore the issue of local debt, which is the focus in important literature on soft budget constraints. Very few countries in Africa, however, allow their local governments to run into debt. Thus, private consumption is equal to net income, and the local government faces the following hard budget constraint (BC):

$$R_i = x_i + c(g_i), \quad (1)$$

where R_i is the income of jurisdiction i and $c(\cdot)$ is the cost of providing an amount g_i of local public good. This cost is assumed to be increasing and convex: $\frac{dc(g_i)}{dg_i} > 0$ and $\frac{d^2c(g_i)}{dg_i^2} > 0$. The convexity reflects the increasing marginal cost of public funds. Since we focus on current spending and not on public investments, we ignore scale economies. This assumption is not rejected by a preliminary empirical test on our data.⁴ In order to have some interior solutions when the jurisdiction is not constrained by its wealth, we have to assume that

$$\forall i, j, \quad c''(g_i) > v_i''(g_i + \theta_{ij}g_j). \quad (2)$$

The convexity of public spending must be superior to the variation of the marginal utility of public goods. This condition obviously holds as soon as the function $v_i(\cdot)$ is concave.

Substituting the expression of the private consumption given by (1) into the initial welfare function, we obtain the following objective function, denoted by U^i , which only depends on the strategic variables (g_i, g_j) :

$$U^i(g_i, \theta_{ij}g_j) = R_i - c(g_i) + v^i(g_i + \theta_{ij}g_j).$$

⁴ We show the absence of scale economies in providing current public spending according to the size of the jurisdiction (measured by the population density, *dens*). Both signs of first and second derivatives are positive and significantly different from zero: $g_i = 3.751^{**}.dens_i + 0.001^{***}.dens_i^2$. Detailed results are provided in Table 1 in the appendix.

Each local government chooses its level of public spending, considering as given the levels of public good in the other jurisdiction. The played game is static and the Nash equilibrium may be constrained. Indeed, we take into account situations where a local government is too poor to finance the minimum of public spending. This leads us to consider a constrained Nash equilibrium denoted by $g_i^*(\theta)$,

$$g_i^*(\theta) = \min \{\bar{g}_i, \tilde{g}_i(\theta)\},$$

where \bar{g}_i is given by

$$R_i - c(\bar{g}_i) = 0,$$

and $\tilde{g}_i(\theta)$ is the solution of the unconstrained Nash equilibrium:

$$\begin{cases} \tilde{g}_i(\theta) \equiv \arg \max_{g_i \geq 0} U^i(g_i, \theta_{ij} g_j^*) \\ \tilde{g}_j(\theta) \equiv \arg \max_{g_j \geq 0} U^j(g_j, \theta_{ij} g_i^*) \end{cases}$$

The set of strategies for each jurisdiction i is compact and it corresponds to $[0, \bar{g}_i]$. The First Order Condition (FOC) of the preceding programme for player i is

$$-\frac{dc(g_i)}{dg_i} + v'_i(g_i + \theta_{ij} g_j^*(\theta)) = 0. \quad (3)$$

The Second Order Condition (SOC) is respected under condition (2).

We focus on the nature of competition among jurisdictions when it exists. These strategic interactions are captured through the sign of $\frac{dg_i}{dg_j}$. Following Bulow, Geanakoplos, and Klemperer (1985), we define local public goods as strategic complements (resp. substitutes) if and only if the marginal utility of public good in jurisdiction i is increasing in the level of local public goods in the other jurisdictions, more formally if $\frac{\partial^2 U^i(g_i, \theta_{ij} g_j)}{\partial g_i \partial g_j} > 0$ (resp. < 0). If jurisdiction i is constrained by its wealth, that is, if $c(\tilde{g}_i(\theta)) > R_i$, we have $g_i^* = \bar{g}_i$ and $\frac{\partial g_i}{\partial g_j} = 0$; otherwise $g_i^* = \tilde{g}_i(\theta)$ and the application of the envelope theorem to (3) yields:

$$\frac{\partial g_i}{\partial g_j} = -\frac{\frac{\partial^2 U^i(g_i, \theta_{ij} g_j)}{\partial g_i \partial g_j}}{\frac{\partial^2 U^i(g_i, \theta_{ij} g_j)}{\partial g_i^2}}. \quad (4)$$

Since the denominator corresponds to the SOC of the maximisation programme, the sign of $\frac{\partial g_i}{\partial g_j}$ is then equivalent to the sign of $\frac{\partial^2 U^i(g_i, \theta_{ij} g_j)}{\partial g_i \partial g_j}$, which also corresponds to the sign of $v_i''(\cdot)$.⁵

2.2 Comparative statics

We will now consider a unilateral change in the degree of the spillovers experienced in jurisdiction i from jurisdiction j . By so doing we can compare the effects on the behaviour of jurisdiction i between an increase of public spending of a neighbouring jurisdiction and the same variation of a more distant jurisdiction. In other words, we estimate the consequences of geographic or ethnic proximity on local governments' public spending.

For comparative statics' analysis, we follow Caputo (1996). Indeed, unlike single-agent models, knowledge of how a parameter affects the marginal value of the i th player's decision variables in a static game is not sufficient to determine the Nash equilibrium comparative statics for the level of the i th player's decision variables. We also have to determine how the parameter's change affects the other player's best reply, and finally how these last variations impact on the marginal value of the i th player's decision variable.

Considering the unconstrained Nash equilibrium ($\forall i, g_i^*(\theta) = \tilde{g}_i(\theta)$) the differentiation of (3) with respect to θ_{ij} for both jurisdictions yields:

$$\begin{pmatrix} U_{11}^i(\cdot) & \theta_{ij} v_i''(\cdot) \\ \theta_{ji} v_j''(\cdot) & U_{11}^j(\cdot) \end{pmatrix} \begin{pmatrix} \frac{\partial \tilde{g}_i(\theta)}{\partial \theta_{ij}} \\ \frac{\partial \tilde{g}_j(\theta)}{\partial \theta_{ij}} \end{pmatrix} = \begin{pmatrix} -\theta_{ij} \tilde{g}_j(\theta) v_i''(\cdot) \\ 0 \end{pmatrix}.$$

Applying the Cramer rule we then obtain:

$$\begin{aligned} \frac{\partial \tilde{g}_i(\theta)}{\partial \theta_{ij}} &= -\frac{\theta_{ij} \tilde{g}_j(\theta)}{|J|} v_i''(\tilde{g}_i(\theta) + \theta_{ij} \tilde{g}_j(\theta)) U_{11}^j(\tilde{g}_j(\theta), \theta_{ji} \tilde{g}_i(\theta)), \\ \frac{\partial \tilde{g}_j(\theta)}{\partial \theta_{ij}} &= \frac{\theta_{ij} \theta_{ji} \tilde{g}_j(\theta)}{|J|} v_i''(\tilde{g}_i(\theta) + \theta_{ij} \tilde{g}_j(\theta)) v_j''(\tilde{g}_j(\theta) + \theta_{ji} \tilde{g}_i(\theta)). \end{aligned} \quad (5)$$

where J is the Jacoby matrix and its determinant is given by

$$|J| = \begin{vmatrix} U_{11}^i(\tilde{g}_i(\theta), \theta_{ij} \tilde{g}_j(\theta)) & \theta_{ij} v_i''(\tilde{g}_i(\theta) + \theta_{ij} \tilde{g}_j(\theta)) \\ \theta_{ji} v_j''(\tilde{g}_j(\theta) + \theta_{ji} \tilde{g}_i(\theta)) & U_{11}^j(\tilde{g}_j(\theta), \theta_{ji} \tilde{g}_i(\theta)) \end{vmatrix}. \quad (6)$$

⁵ If this last expression is positive, then the game played by each jurisdiction is supermodular as noted by Topkis (1998) and at least one equilibrium exists.

Generally, the sign of $|J|$ remains indeterminate, since it does not rely on the sign of the Hessian matrix of a single optimisation problem as Caputo (1996) emphasises it. Thus, without additional assumptions about the stability or uniqueness of the Nash equilibrium, for instance, we cannot sign $|J|$. We then obtain the following PROPOSITION:

Proposition 1 *Under our assumptions, we have*

- (i) *If the jurisdiction i is constrained by its wealth ($\tilde{g}_i(\theta) > c^{-1}(R_i)$), a change in θ_{ij} has no effect on the level of provided public good in both jurisdictions;*
- (ii) *If the jurisdiction j is constrained by its wealth, a change in θ_{ij} has no effect on the level of provided public good in jurisdiction j but increases (decreases) the level of public good in jurisdiction i if public goods are strategic complements (substitutes);*
- (iii) *If no jurisdiction is constrained, an increase in the degree of spillover from jurisdiction j to i (θ_{ij}) involves a variation in the same (opposite) sense in both jurisdictions if local public goods are strategic complements (substitutes).*

Proof. (i) If $g_i^*(\theta) = \bar{g}_i$, it is then obvious that $\frac{\partial g_i^*(\theta)}{\partial \theta_{ij}} = 0$ and $\frac{\partial g_j^*(\theta)}{\partial \theta_{ij}} = 0$ from differentiation of (3) with respect to θ_{ij} .

(ii) If $g_j^*(\theta) = \bar{g}_j$ and $g_i^*(\theta) = \tilde{g}_i(\theta)$, then we have $\frac{\partial g_j^*(\theta)}{\partial \theta_{ij}} = 0$ which yields

$$\frac{\partial g_i^*(\theta)}{\partial \theta_{ij}} = -\frac{\theta_{ij} \bar{g}_j v_i''(g_i + \theta_{ij} \bar{g}_j)}{\frac{\partial^2 U^i(g_i, \theta_{ij} \bar{g}_j)}{\partial g_i^2}}.$$

(iii) If $g_i^*(\theta) = \tilde{g}_i(\theta)$ and $g_j^*(\theta) = \tilde{g}_j(\theta)$, we obtain from (5)

$$\frac{\partial \tilde{g}_i(\theta)}{\partial \theta_{ij}} \frac{\partial \tilde{g}_j(\theta)}{\partial \theta_{ij}} = -\theta_{ji} \left(\frac{\theta_{ij} \tilde{g}_j(\theta) v_i''(\tilde{g}_i(\theta) + \theta_{ij} \tilde{g}_j(\theta))}{|J|} \right)^2 v_j''(\tilde{g}_j(\theta) + \theta_{ji} \tilde{g}_i(\theta)) U_{11}^j(\tilde{g}_j(\theta), \theta_{ji} \tilde{g}_i(\theta)).$$

■

The parameter θ_{ij} may represent, for instance, the degree of ‘proximity’ that jurisdiction i experiences from the local public good provided by jurisdiction j . This ‘proximity’ will be expressed in geographic or ethnic terms in our empirical works. An increase in θ_{ij} would induce two effects on g_i : one direct and one indirect (strategic) effect through the level of public good provided by the neighbour (g_j). If jurisdiction i is constrained by its wealth, any change in θ_{ij} does not affect the equilibrium value. Indeed, neither the direct effect, nor the strategic effect would come into play, since the level of public spending in this jurisdiction is at the corner. If it is the other jurisdiction, namely j , which is constrained, then only the direct effect of θ_{ij} would play on g_i . An increase of θ_{ij} induces an increase (decrease) in

g_i local public expenditures are strategic complements (substitutes). Remember that in the presence of strategic complements, the marginal utility of local public good increases in the level of the other local public good or in the spillover effect.

Finally, if no jurisdiction is constrained, then both effects are at play. Without additional assumptions, however, particularly on the sign of $|J|$, we can only conclude that an increase in θ_{ij} would induce an increase or a decrease of levels of local public goods in both jurisdictions in the presence of strategic complements. Otherwise, that is in the presence of strategic substitutes, an exogenous change of θ_{ij} would involve opposite variations among jurisdictions.

Following Dixit (1986) or Kolstad and Mathiesen (1987), we specify the uniqueness and the stability of the Nash equilibrium through the following assumption:⁶

$$|J| > 0. \tag{7}$$

This relationship allows us to pinpoint the sense of variations resulting from the two kinds of parameter changes. We obtain the following PROPOSITION:

Proposition 2 *Under our assumptions and (7), an increase in the degree of spillover from jurisdiction j to i (θ_{ij}) involves an increase (decrease) of the level of public goods in both jurisdictions if local public goods are strategic complements (substitutes).*

Proof. Immediate from (5). ■

Assuming the uniqueness of the Nash equilibrium allows us to specify the sense of deviation of public spending when the degree of spillovers varies. To sum up our theoretical results, we showed that spillovers among jurisdictions involve strategic behaviours, which in turn lead to a competition process. Without restricting the nature of such a competition, we estimated to what extent the level of provided public good is affected by a deviation in the degree of spillovers.

Our theoretical framework yields the following implications: (1) the provision of local public goods with spillovers may induce two cases: (a) strategic interactions in terms of

⁶ If we adopt the contraction approach (see Vives, 1999), the condition of equilibrium uniqueness involves

$$U_{11}^i(g_i, \theta_{ij}g_j) + \left| v_{12}^i(g_i, \theta_{ij}g_j) \right| < 0,$$

which yields that $|J|$ is positive.

complements or substitutes (classical result), (b) no strategic interactions owing to the insufficient level of fiscal resources and despite positive externalities (largely ignored by the relevant literature); (2) under the presence of strategic complements, the expected quantity of public goods in jurisdiction i will positively depend on the level of public good allocated by jurisdiction j ; (3) in the presence of strategic substitutes, an opposite relationship is expected; (4) the sign of such a strategic interaction is not determined *a priori*, since different measures of contiguity may be put forward (geographical or ethnic contiguity).

3 Empirical evidence of public spending interactions in a less developed country: the case of Benin

Our empirical analysis focuses on Benin, a young democracy, which is quite representative of the Sub-Saharan region. After a brief overview of this country, we test the existence of strategic interactions among local governments' spending with the spatial econometric method. We then examine if a change in these interactions occurs during election years.

3.1 Benin overview

With a per capita income of US\$ 570 in 2007 and a ranking of 163 out of 177 countries,⁷ Benin remains one of the poorest countries of the world. It is ethnically fragmented.⁸ Since its independence on 1 August, 1960, the history of Benin has been chaotic. A succession of military governments ended in 1972 with the last military coup led by Mathieu Kerekou and the establishment of a government based on Marxist-Leninist principles. A move to democracy began in 1989. Two years later, free elections ushered in former Prime Minister Nicephore Soglo (a former World Bank official) as President. Kerekou regained power in 1996 in elections fraught with irregularities and won subsequent elections in 2001. Having served two terms and being over 70, he was ineligible to run in the presidential elections of 2006. He was succeeded by Thomas Boni Yayi, an independent political outsider who had previously headed the West African Development Bank. In March 2007, President Yayi

⁷ Human Development Report (2007).

⁸ Among the 42 ethnic groups, the most prominent are the Fon and the Adjias in the south, the Baribas and the Sombas in the north and the Yorubas in the south-east.

Boni strengthened his position following the legislative elections in which his coalition, ‘Force Cauris pour un Bénin Emergent (FCBE)’ won the largest number of seats (35 out of 83) and negotiated a pro-government majority in Parliament with seven minor parties and coalitions joining the FCBE.

This democratic process was accompanied by a huge transformation of the political and administrative organisation. Since 1998, Benin has undergone a decentralisation process that became effective with local elections in 2002. The second local elections took place in 2008.⁹ As depicted in Figure 1, Benin is divided into twelve *départements* which are subdivided into 77 *communes*, which are further divided into 546 districts. *Départements* are managed by a representative of the central government. In contrast, *communes* are ruled by a local government directly elected by inhabitants. The average size of *communes*, presented in the following map, numbers about 90,000 inhabitants.

Insert Figure 1

In January 1999, law 97-029 defined the transferred competencies from the centre to the 77 *communes*. Theoretically, competencies of Beninese *communes* range from elementary school to economic development and include transport infrastructure, environment (hygiene), health and social goods, tourism, security or market-place management. As in most of African countries, however, this competences’ transfer was not accompanied by an adequate transfer of resources. Beninese *communes* are characterised by a very low level of resources (only about 4.5% of country tax revenues or equivalently 0.7% of GDP).¹⁰ Moreover, important inequalities appear between *communes*: the resources of the ten poorest *communes* represent 5 per cent of the resources of the five richest ones.

⁹ The first round of municipal elections held on 15 December, 2002 and the second round on 19 January, 2003 with an average rate of turnout estimated at 70 per cent.

¹⁰ Local resources are mainly *communes*’ own resources (about 70%). Property taxes and licences to exercise a trade or profession (‘patente’) represent 90% of local tax revenues (see Chambas, Brun, and Rota Graziosi, 2008 for a detailed analysis of local fiscal resources in Sub-Saharan Africa, particularly in Benin). Retroceded taxes, which come from transfers of state tax revenue to local governments, account for about 10% of local resources.

3.2 Econometric framework

Horizontal interactions entail a fiscal reaction function that depicts how the decision variable for a given jurisdiction depends on the decisions of other jurisdictions. To test the existence of such functions, we have to test spatial dependence in a panel data framework. We consider a specification in the most general form in which *commune* i public expenditure in year t , defined by G_{it} , is a function of its neighbours' same public choice, G_{jt} . Moreover, we allow G_{it} to depend on a vector of specific controls X_{it} and we include a *commune*-specific effect, α_i .¹¹ This gives the following specification:

$$G_{it} = \sum_{ij} \rho_{ij} \cdot G_{jt} + \beta \cdot X_{it} + \alpha_i + \varepsilon_{it}, \quad (8)$$

where $i = 1, \dots, n$ denotes a *commune* and $t = 1, \dots, T$ a time period, α, β and ρ are unknown parameters and ε_{it} a random error. Since there are too many parameters ρ_{ij} to be estimated, the usual procedure is to consider:

$$G_{it} = \rho \cdot A_{jt} + \beta \cdot X_{it} + \alpha_i + \varepsilon_{it}, \quad (9)$$

where $A_{jt} = \sum \theta_{ij} \cdot G_{jt}$. G_{it} , the vector of public spending in a local government i at time t depends on A_{jt} , the weighted average vector of public spending in the set of the other local governments j at time t and a set of specific controls X_{it} .

We explore a variety of weighting schemes to allow different patterns of spatial interactions. We are then able to examine the nature of the neighbouring effects and to discuss the definition of distance. First, we have chosen a common geographical definition of neighbouring communities based on a contiguity matrix where the value one is assigned if two jurisdictions share the same border and zero otherwise. This scheme is given by the weight matrix θ^{neigh} . Contiguity is commonly used in the relevant empirical literature. Second, we define an ethnic weight matrix θ^{ethn} based on the ethnic proximity of *communes*' inhabitants. Ethnic proximity is defined as the probability that two individuals randomly drawn from two distinct *communes* are from the same ethnic group. In doing so, we test

¹¹ All time-invariant community characteristics, observed or unobserved can be represented by community-specific intercepts.

the existence of spending interactions between *communes* which are similar with respect to ethnicity. The first two weighting schemes are based on the idea that *communes* follows *communes* close to them either geographically or with similar socio-economic structure.¹²

Then, we consider a possible weighting scheme in which weights are assumed to be identical for all *communes* j (θ^{uni}). This uniform weighting scheme will give us a useful benchmark to ascertain whether the potential observed spatial auto-correlation can be attributed to a substantive strategic interaction process and not to a ‘common intellectual trend’.¹³ Moreover, following Lockwood and Migali (2009), we compare ethnic and contiguity weights with ‘placebo’ weights, θ^{plac} , which are chosen in a random way. We generate a random number distributed between zero and one for each *commune*. Then, the weight assigned between two *communes* is the difference between its random numbers. Evidence of strategic interactions with this placebo matrix would indicate some general positive correlations between all public spending owing to omitted common shocks.¹⁴

Finally, we present a dynamic version of the model. We introduce the lagged dependent variable, G_{it-1} , as a right-hand side in order to take into account the persistency in public expenditure (see Veiga and Veiga, 2007). The dynamic model can be written as follows:

$$G_{it} = \lambda.G_{it-1} + \rho.A_{jt} + \beta.X_{it} + \alpha_i + \varepsilon_{it}. \quad (10)$$

By estimating this reaction function we are confronted with important econometric issues as described by Brueckner (2003). First, because of strategic interactions, public expenditure in different jurisdictions is jointly determined. Therefore, neighbours’ decisions are endogenous and correlated with the error term ε_{it} and ordinary least squares estimation of the parameters is inconsistent, requiring alternative estimation methods based on the instrumental variables

¹² In the welfare competition literature, jurisdictions are likely to take into account inhabitants’ mobility in the neighbouring communities induced by an increase in its own welfare. In the yardstick competition literature, residents consider neighbouring jurisdictions—on which they are likely to get better information—as a yardstick to compare the performance of their incumbent. It is easy to understand that ethnic as well as geographic proximity could explain such interactions.

¹³ Indeed, our theory relies on the common assumption that local governments behave strategically with each other. Alternatively, Manski (1993) suggests that fiscal choices appear to be interdependent not because jurisdictions behave strategically but because they actually follow a ‘common intellectual trend’ that drives fiscal choices in the same directions. Thus, we extend our analysis to determine whether these interdependencies are owed to strategic interactions or only to a common trend.

¹⁴ Weights are normalised so that their sum equals unity for each i for all weight matrices. This assumes that spatial interactions are homogeneous: each neighbour has the same impact on the *commune*.

method (IV) or on maximum likelihood (ML).¹⁵

Second, the omission of explanatory variables that are spatially dependent may generate spatial dependence in the error term, which is given by: $\varepsilon_{it} = \tau\theta\varepsilon_{it} + v_{it}$.¹⁶ When spatial error dependence is ignored, estimation can provide false evidence of strategic interaction. To deal with this problem, one possible approach is to use the ML estimator, taking into account the error structure (see Case, Rosen, and Hines, 1993) or the IV method which yields consistent estimations even with spatial error dependence (see Kelejian and Prucha, 1998). Brueckner (1998), Brueckner and Saavedra (2000), Saavedra (2000) and Foucault, Madies, and Paty (2008) use the tests of Anselin, Bera, Florax, and Yoon (1996) to verify the hypothesis of error independence.¹⁷

Lastly, since we introduced the lagged dependent variable as a right-hand side to consider the autoregressive component of the time series, the previous estimators are inconsistent (Nickell, 1981). We propose to use the GMM System estimator after verifying the hypothesis of error independence and estimating the static model with the ML estimator. The GMM estimators allow us to control for both unobserved country-specific effects and potential endogeneity of the explanatory variables. The GMM System estimator combines in one system the regressions in difference and the regressions in level. Basically, Blundell and Bond (1998) show that this extended GMM estimator is preferable to that of Arellano and Bond (1991) when the dependent variable, the independent variables, or both are persistent.¹⁸

The first challenge of our empirical work is to determine that the observed spatial autocorrelation can be attributed to a substantive strategic interaction process and not to exogenous correlations in omitted jurisdictional characteristics or common shocks to local fiscal policy. To go beyond this issue, we introduce time dummies to capture shocks in each period

¹⁵ With the IV approach, a typical procedure is to use the weighted average of neighbours' control variables as instruments (see Kelejian and Prucha, 1998). The ML method consists in using a non-linear optimisation routine to estimate the spatial coefficient ρ (see Brueckner, 2003).

¹⁶ We note that the use of a panel helps to eliminate spatial error dependence which arises through spatial autocorrelation of omitted variable, since the influence of such variables is partly captured in community-specific intercept terms.

¹⁷ These tests are not contaminated by uncorrected spatial error dependence and may detect the presence of spatial lag dependence.

¹⁸ Arellano and Bond (1991) present a first-difference GMM estimator. There are, however, conceptual and statistical shortcomings with this estimator as the first difference estimator exacerbates the bias owed to errors in variables (Hausman, Hall, and Griliches, 1984). Thus, we use an alternative system estimator that reduces the potential biases and imprecision associated with the usual difference estimators (Arellano and Bover, 1995 and Blundell and Bond, 1998) and also greatly reduces the finite sample bias (Blundell, Bond, and Windmeijer, 2000).

which are common to all local governments and other specific controls. We then have

$$G_{it} = \lambda.G_{it-1} + \rho.A_{jt} + \beta_1.D_{it} + \beta_2.N_{dt} + \alpha_i + \zeta_t + \varepsilon_{it}, \quad (11)$$

where D_{it} is the population density of jurisdiction i on year t , which captures the possibility of scale economies in public spending,¹⁹ N_{dt} is the percentage of men who have a job in *département* d on year t . The employment rate is an indicator of the economic conjuncture and allows the partial control of common shocks which would be spatially correlated.

We extend our analysis to test the effect of partisan affiliation and the fiscal behaviour of local policymakers in an election period. Until he stepped down in March 2006, Mathieu Kérékou enjoyed strong support in the north of the country (Alibori, Atacora, Borgou and Donga) which is considered as his fief.²⁰ When Boni Yayi was elected, he affirmed his desire for political openness. Municipal elections took place in April 2008. Parties allied with the President won a majority of local council seats, but the most of municipalities in the south were won by opposition parties. Departments that can be considered as fiefs are concentrated in the south of the country, in particular, Atlantic, Collines and Mono. Finally, over the whole time period, about 40% of the departments have shared the same partisan affiliation as the President in office. To capture this effect (having the same partisan affiliation as the president in office), we extend our analysis by including dummy variables for political affiliation. We also introduce dummy variables for election years to test opportunistic behaviour of local policymakers.²¹ It follows that:

$$G_{it} = \lambda.G_{it-1} + \rho.A_{jt} + \beta_1.D_{it} + \beta_2.N_{dt} + \beta_3.O_{ct} + \beta_4.T + \beta_5.PR_{it} + \beta_6.E_{t-1} + \beta_7.E_t + \beta_8.E_{t+1} + \alpha_i + \varepsilon_{it}, \quad (12)$$

where T is a trend variable, which accounts for the common trend in local governments, O_{ct} is a trade openness measure at country level which controls for macroeconomic common shocks, since developing countries are vulnerable to foreign trade.²² Moreover, it could have

¹⁹ Population density is the number of inhabitants per square kilometre. Per capita expenditures and population density are in log. Per capita expenditures are corrected for inflation.

²⁰ Kérékou was born in 1933 in Kouarfa, in the north-west of the country.

²¹ We can note that we have a small number of observations for election years, so, results will only give us an indication of the existence of opportunistic behaviour and it will be difficult to come to a firm conclusion.

²² Since we introduce dummy variables for election years we can no longer introduce time dummies.

many effects on public finances.²³ PR_{it} is a dummy variable which takes the value one if the local government i has the same political affiliation as the President in office and zero otherwise, E_{t-1} is a dummy variable, which takes the value one the year before the election and zero otherwise, E_t is a dummy variable, which takes the value 1 the year of the election and zero otherwise, E_{t+1} is a dummy variable, which takes the value one the year after the election and zero otherwise.²⁴

With respect to our theoretical model, we see that $\rho \neq 0$ is equivalent to the existence of some strategic interactions and that local governments are not constrained by their wealth, on average. Moreover, if $\rho > 0$ ($\rho < 0$), that means an increase in the degree of spillover between jurisdictions involves a variation in the same (opposite) sense of the level of local public goods, we can conclude that local public goods are strategic complements (substitutes). To refine this result, we will also empirically test the effect of wealth constraints on the existence of public spending interactions. Indeed, our theoretical model highlights that strategic interactions may be restricted by the extreme poverty of some local governments. To test the hypothesis that interactions are stronger when local governments are less constrained by their wealth, we use an indicator of fiscal autonomy,²⁵ denoted by F_{it} . We have:

$$G_{it} = \lambda \cdot G_{it-1} + \rho \cdot A_{jt} + \varphi \cdot AF_{it} + \beta_1 \cdot D_{it} + \beta_2 \cdot N_{dt} + \beta_3 \cdot O_{ct} + \beta_4 \cdot T_{it} + \beta_5 \cdot PR_{it} + \beta_6 \cdot E_{t-1} + \beta_7 \cdot E_t + \beta_8 \cdot E_{t+1} + \beta_9 \cdot F_{it} + \alpha_i + \varepsilon_{it}, \quad (13)$$

where $AF_{it} = A_{jt} \cdot F_{it}$. If fiscal autonomy actually reinforces strategic interactions, we should observe the coefficient of AF_{it} being more significant and higher than the coefficient of A_{jt} ($\varphi > \rho$). Moreover, as policymakers should interact with their neighbours only if they can choose the level of their local resources, ρ may no longer be significant. If that is the case, we will conclude that strategic interactions are contingent on a certain degree of fiscal autonomy and that strategic interactions exist only among unconstrained local governments.

²³ Rodrik (1998) shows that there is a positive correlation between an economy's exposure to international trade and the size of its government because government spending plays a risk-reducing role in economies exposed to a significant amount of external risk.

²⁴ Descriptive statistics are presented in Table 2.

²⁵ A usual approximation of fiscal autonomy is the ratio of jurisdictions' own resources to their total resources.

3.3 Results

Our dataset covers the two local elections (2002 and 2008) and the 77 *communes* of Benin. Data for *communes*' current expenditure come from *Beninese Ministry of Finances and Economy*. The other control variables are drawn from *WDI* (World Development Indicators), *Afrobarometers*, *Demographic and Health Surveys* provided by the *National Institute of Statistic and Economic Analysis of Benin* and 77 monographs realised for the *Mission of Decentralisation*.

First, as we have noted before, it is important to investigate whether the policy of a local jurisdiction is actually correlated with the policies of other jurisdictions and whether spatial lag or spatial error dependence are the more likely sources of correlation. Anselin, Le Gallo, and Jayet (2006) have proposed two in-depth tests based on the Lagrange Multiplier principle for panel data that indicate the most likely source of spatial dependence (spatial lag or spatial error dependence). We compute the same robust tests for spatial lag dependence and for spatial error dependence which require only the OLS residuals from a non-spatial model. Therefore, we first estimate (12) using OLS for both contiguity and ethnic matrix without taking into account the influence of public spending in other jurisdictions ($\rho = 0$) and the lagged value of our dependent variable ($\lambda = 0$). The estimation results are shown in Table 3. Spatial tests indicate the presence of spatial lag dependence for public spending but not the existence of spatial error dependence for both matrices.

Second, since the hypothesis of error independence is verified, we estimate (12) using ML with specific-effects for both contiguity and ethnic matrices without taking into account the lagged value of our dependent variable ($\lambda = 0$), but we introduce the influence of the expenditure set by other jurisdictions ($\rho \neq 0$). The estimation results are shown in Table 4. The coefficient of the weighted average vector of public expenditure in the set of the other local governments is always significant and positive for both matrices.²⁶

Finally, we estimate with the GMM System the dynamic model (12) for all weighting

²⁶ In these first estimations, departments' employment rate and population density coefficients are not significant. There are no economies of scale in the studied public spending. The trade openness indicator coefficient is significantly positive. Dummies associated with election years indicate, *a priori*, an opportunistic use of public spending during the year before the election. Indeed, current expenditure seems to increase during the year before the election and to decrease after. Lastly, a *commune* governed by a local government which has the same political affiliation as the president in office has higher public expenditure.

schemes, taking into account the lagged value of our dependent variable ($\lambda \neq 0$). We adopt the assumption of weak exogeneity of employment rates and trade openness, in the sense that they are assumed to be uncorrelated with future realisations of error terms. The weighted average vector of per capita public spending of other local governments is, as noted before, suspected of endogeneity. Other explanatory variables²⁷ are assumed to be strictly exogenous. The lagged levels of variables are used as instruments in the regressions in level as well as in the regressions in difference. We collapse instruments and limit the number since too many instruments lead to inaccurate estimation of the optimal weight matrix, biased standard errors and, therefore, incorrect inference of overidentification tests (see Roodman, 2009).²⁸ Table 5 displays estimation results.

Insert Table 5

We focus our attention on (1) (2) (3) (4) and (5), that is, the GMM System estimations for contiguity, ethnic, uniform and placebo matrices. We can first note that orthogonality conditions are correct and that the coefficient on the lagged dependent variable is always significant and positive.²⁹ As the coefficient on lagged public spending provides an estimate λ varying between 0.411 and 0.629, the result indicates some level of persistency in public expenditure which is likely to change slowly over time. Moreover, it confirms the consistency of the autoregressive specification. After correction for endogeneity, the coefficient of the weighted average vector of public expenditure in the set of the other local governments is significant at least at 1% level and positive for ethnic and contiguity matrices.

²⁷ Population density, time dummies, election dummies, partisan affiliation, trends.

²⁸ The lags of at least two earlier periods for weak exogenous variables and three earlier periods for endogenous variables are used as instruments. The lagged dependent variable is instrumented by lags of the dependent variable from at least two earlier periods. We use two lags for endogenous and weak exogenous variables.

²⁹ The consistency of the estimator depends on whether lagged values of explanatory variables are valid instruments. The criteria for the selection of instruments are two specification tests (Arellano and Bond, 1991). With the Hansen test, we test the null hypothesis of the overall validity of instruments' orthogonality conditions (over-identifying restrictions). The second test is about the serial correlation of residuals. It examines the hypothesis that the residuals from the first-differentiated estimating equation are not second-order correlated. First, we test the null hypothesis of no first-order serial correlation of differentiated residuals (AR (1) test) and second, the null hypothesis of no second-order serial correlation of differentiated residuals (AR (2) test). If we reject the null hypothesis of no first-order serial correlation and do not reject the null hypothesis of no second-order serial correlation of differentiated residuals, the residuals are serially uncorrelated and we conclude that orthogonality conditions are correct. In our case, both statistics confirm the validity of the instruments used.

At this stage, we cannot conclude that there are strategic spending interactions as Manski (1993) suggested. If this is only a common trend that drives local governments in the same direction, we should expect a positive sign of the interaction coefficient but not a specific pattern in the type of *communes* with which to interact. As the coefficient of interaction with the uniform matrix is significant (column (3)), we can say that local governments follow, in part, a common trend. When, however, we estimate the coefficient for the contiguity matrix after checking for common trends in column (4), the neighbouring interaction coefficient remains significantly positive. Beyond the common trend, local governments interact with each other. Moreover, the placebo matrix (column (5)) does not show any evidence of positive strategic interactions. This shows that the phenomenon of fiscal interactions detected with geographical and ethnical matrices is not an artefact of the estimation procedure. Note that we have also ascertained that before 1998, the beginning of the decentralisation process, there were no strategic interactions (see Table 6).³⁰

Hence, we can conclude that there are strategic interactions between neighbouring jurisdictions and that these interactions also exist between *communes* that are close with respect to ethnic composition. Public expenditure seems to be a strategic complement. An average public spending increase of 10% in the neighbouring municipalities induces an increase of around 6.2% in local primary expenditure. We find a smaller coefficient (5.1%) for the ethnic matrix, suggesting the existence of stronger interactions among neighbouring *communes* than among ethnically close ones. Since different ethnic groups are located in close geographical areas, we can assume that the geographic matrix overlies the ethnic matrix. We estimate the coefficient for the ethnic matrix after checking for geographical interactions in column (6), which remains significant and stable. Therefore, even if geographical distance remains more relevant for explaining interjurisdictional competition, interactions also exist among ethnically close *communes*.

Finally, regarding our theoretical results, we can conclude that current expenditure is a strategic complement, since an increase in the degree of spillover involves a variation in the same sense in different jurisdictions. Moreover, jurisdictions seem not to be constrained by their wealth, on average.

³⁰ We run the same regressions as previously for the period 1994 to 1998. The coefficients of interaction with all matrices are not significant.

In columns (7) and (8), we test the robustness of these results by estimating the same econometric model with alternative matrices. The θ^{neigh2} matrix, in which the value of one is assigned if two *communes* belong to the same *département* and zero otherwise, is a proxy of the contiguity matrix. Indeed, two *communes* that belong to the same *département* are generally close. The θ^{ethn2} matrix is defined as follows: the value one is assigned if two *communes* have the same dominant ethnic group and zero otherwise. The coefficient of the weighted average vector of public expenditure in the set of the other local governments remains positive and significant at the 5% level for the θ^{neigh2} matrix but only at 10% for the ethnic matrix θ^{ethn2} . Additionally, these estimations tend to confirm the existence of opportunistic behaviour of local governments and the effects of partisan affiliation.

We find a positive and significant sign for the parameter associated with the employment rate, which indicates the effect of economic conjuncture. The trend variable remains, as expected, significant and negative. Indeed, per capita public expenditure has decreased by 75% over the period despite little growth between 2004 and 2006. There seems to be opportunistic behaviour of local jurisdictions since dummies associated with the pre-election year indicate an increase in public spending.³¹ Thus, we find some evidence of a political budget cycle for current expenditure in this young democracy.³² As we noted before, however, it is relatively difficult to determine since we have a small number of observations. In addition, it appears that having a local government which has the same political affiliation as the President in office is likely to increase public expenditure. Indeed, the coefficient of the dummy variable that indicates whether the local government has the same political affiliation as the President in office is always significant.³³

In columns (9) and (10), we test the effect of wealth constraints on the existence of public spending interactions by estimating equation (13) for both matrices. As expected, the

³¹ To understand the sign of the coefficient associated with the election year dummy, one must refer to the election calendar and budget votes. Local elections took place at the beginning of March and the definitive budget must be adopted before 31 March. Therefore, in the year before the elections, decision-makers increase current expenditures and decrease them the year after, since the definitive budget is approved.

³² Shi and Svensson (2006) have shown that political budget cycles are much larger in young democracies than in developed countries, particularly in countries with weak institutional constraints on incumbents' rent-seeking ability.

³³ These municipalities benefit from higher state subsidies.

Note that the parameter associated with population density is positive but not always significant. This is also the case for the coefficient of trade openness, which is negative but not always significant. As some of these effects offset each other, it is often difficult to predict the net effect of trade openness on public expenditure which is not significant.

coefficient of the interaction variable between the neighbours' spending decisions and the indicator of fiscal autonomy (φ) is positive and significant. We can thus conclude that interactions are stronger when local governments have a more significant share of their own local resources. Furthermore, since the coefficient for strategic interaction alone (ρ) is no longer significant, strategic interactions are contingent on fiscal autonomy: strategic interactions exist only among local governments which have a certain degree of autonomy.

At this stage, our results suggest that decentralisation has induced interjurisdictional competition. Indeed, there are strategic interactions between Beninese local governments with regard to current expenditure that appears to be strategic complement. These interdependences exist between neighbouring *communes* but also between those which are close in terms of ethnic composition. Finally, we have found a variation of public expenditure in the same sense between close local jurisdictions: local public goods are strategic complements and, on average, local jurisdictions are not constrained by their wealth. Moreover, our results confirm that strategic interactions are contingent on fiscal autonomy. Estimations tend to show that local governments adopt opportunistic behaviour before elections. Lastly, *communes* where local government has the same political affiliation as the President enjoy higher public spending.

3.4 What happens in electoral years?

This section is really an empirical extension of our theoretical model, in which perfect information is assumed. Electoral years implicitly refer to a yardstick competition model, which is based on the informational asymmetry among local elected decision-makers and jurisdictions' populations. Indeed, in this model, voters can use the performance cues of other governments as a benchmark to judge whether their representative wastes resources and deserves to remain in office. During an electoral year, we may expect that political campaigns increase interactions among *communes*, since more information is available on the fiscal policies of local decision-makers, reinforcing the effect of yardstick competition.

Here, we evaluate the effects of electoral pressure on inter-jurisdictional interactions. The challenge consists in evaluating the real electoral pressure induced by decentralisation by identifying the effect of elections on interjurisdictional competition. We consider the election

cycle variables to test the hypothesis that interactions may be stronger in election periods. A straightforward way to test this is to interact the neighbours' spending decisions (A_{jt}) with the election years dummy, EY (E_{t-1} and E_t) and estimate two different interaction coefficients, one for years of election ($A_{jt} \times EY$) and one for all the other periods, NEY ($= 1 - EY$), ($A_{jt} \times NEY$).

$$G_{it} = \lambda S_{it-1} + \rho' \cdot (A_{jt} \times EY) + \rho'' \cdot (A_{jt} \times NEY) + \beta_1 \cdot D_{it} + \beta_2 \cdot N_{dt} + \beta_3 \cdot O_{ct} + \beta_4 \cdot T + \beta_5 \cdot PR_{it} + \beta_6 \cdot EY + \beta_7 \cdot NEY + \alpha_i + \varepsilon_{it}, \quad (14)$$

where $EY = E_{t-1} + E_t$ and $NEY = (1 - (E_{t-1} + E_t))$. If elections actually reinforce the exposure of jurisdictions, we should observe the coefficient of ($A_{jt} \times EY$) being more significant and higher than the coefficient ($A_{jt} \times NEY$) as policymakers should be particularly concerned about their neighbours' decisions during election periods.

Insert Table 7

We note that significant coefficients have the same sign as in previous estimations and that statistics confirm the validity of used instruments. Therefore, we focus our analysis on the comparison of the interaction term coefficients. As expected, the coefficient is slightly higher and more significant in election periods than in other periods with contiguity and ethnic matrices. Therefore, current expenditure settings appear to be a little more dependent on neighbours during election periods. We complete our analysis with Wald tests to test how coefficients during election periods are significantly different from those in other periods. These tests do not indicate that the election year interaction coefficients are significantly different at the 10% level from the non-election ones. Thus, it is not a definite fact that yardstick competition is the main channel of *communes*' interactions.

4 Conclusion

Decentralisation has been advocated to improve the performance of the public sector by stimulating interjurisdictional competition. The originality of our paper consists in reasoning from an African case in which decentralisation has been recently implemented. One may

expect that fiscal interactions in a developing country where local public resources are scarce would remain modest. Our paper shows that this is not the case. Indeed, for the first time, we provide a new contribution in the fiscal federalism literature by ascertaining that decentralisation entails interjurisdictional interactions in an African country. These interactions are not a common trend. They exist not only among neighbouring local jurisdictions but also among *communes* which are close in terms of ethnic composition. We also emphasise both the influence of partisan affiliation and the opportunistic behaviour of local governments before elections. Finally, Benin which has recently experienced a decentralisation process seems to be as concerned with fiscal strategic interactions as developed democracies.

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A Appendix

A.1 Figure and Tables



Figure 1: Administrative map of Benin

Table 1: Estimation results for the presence of scale economies - Specific effects

Dependent variable: current expenditure of ‘commune’ i ($G_{i,t}$)	
Population density	2.540** (1.41)
Squared Population density	0.001*** (0.00)
Haussman test: p-value	0.34
Observations	429

Robust standard errors are in brackets.***: coefficient significant at 1 % level, **: at 5 % level, *: at 10 % level.

Table 2: Descriptive statistics

Variables	Obs.	Mean	Std. dev.	Min	Max
Current expenditure	448	2915	10934	15.06	123779
				(Basilla)	(Cotonou)
Population density	378	334.94	1019.25	7.61	7684
				(Tangui.)	(Cotonou)
Employment rate	462	26.67	10.68	3.46	59.40
				(Kandi)	(Cotonou)
Trade openness	462	39.14	0.81	37.95	40.20
				(2003)	(2008)
Partisan affiliation	462	0.38	0.48	0	1

Table 3: Estimation results for LM tests - Specific effects

Weighting scheme	Dependent variable: current expenditure of 'commune' i ($G_{i,t}$)	
	(1) θ^{neigh}	(2) θ^{ethn}
Population density	0.216 (0.19)	0.216 (0.19)
Employment rate	0.176*** (0.05)	0.176*** (0.05)
Trade openness	-0.026*** (0.00)	-0.026*** (0.00)
Partisan Affiliation	0.074 (0.27)	0.074 (0.27)
Trend	0.059* (0.03)	0.059* (0.03)
Election year t-1	0.114* (0.06)	0.114* (0.06)
Election year t	-1.467*** (0.17)	-1.467*** (0.17)
Election year t+1	-0.497*** (0.07)	-0.497*** (0.07)
LMlag (p-value)	13.33 (0.001)	11.97 (0.005)
LMerr (p-value)	1.35 (0.25)	0.60 (0.43)
Observations	462	462

Robust standard errors are in brackets.***: coefficient significant at 1 % level, **: at 5 % level, *: at 10 % level.

Table 4: Estimation results with spatial lag dependence - ML estimator

Weighting scheme	Dependent variable: current expenditure of 'commune' i ($G_{i,t}$)	
	(1) θ^{neigh}	(2) θ^{ethn}
Spending in city j	0.255*** (0.07)	0.443** (0.19)
Population density	0.025 (0.06)	0.022 (0.06)
Employment rate	-0.003 (0.01)	-0.003 (0.01)
Trade openness	0.115** (0.05)	0.167*** (0.05)
Partisan Affiliation	0.288** (0.11)	0.244** (0.11)
Trend	-0.124** (0.05)	-0.065** (0.01)
Election year t-1	0.214** (0.09)	0.169* (0.10)
Election year t	-0.666*** (0.19)	-0.361 (0.30)
Election year t+1	-0.568*** (0.09)	-0.549*** (0.10)
Log-likelihood	-206.54	-207.57
N	462	462

Robust standard errors are in brackets.***: coefficient significant at 1 % level, **: at 5 % level, *: at 10 % level.

Table 5: Estimation results for dynamic model - GMM-System³⁴

Dependent variable: current expenditure of commune i ($G_{i,t}$)										
Weighting scheme	(1) θ^{neigh}	(2) θ^{ethn}	(3) θ^{uni}	(4) θ^{uni}	(5) θ^{plac}	(6) θ^{ethn}	(7) θ^{neigh2}	(8) θ^{ethn2}	(9) θ^{neigh}	(10) θ^{ethn}
Lagged dep. var.	0.569*** (0.22)	0.527*** (0.21)	0.580** (0.22)	0.411** (0.20)	0.629*** (0.21)	0.403** (0.19)	0.410* (0.27)	0.768*** (0.14)	0.678*** (0.21)	0.652*** (0.21)
Spending in communes j	0.623** (0.28)	0.513*** (0.19)	0.384* (0.20)	0.472** (0.19)	-0.202 (0.20)	0.468*** (0.18)	0.653*** (0.18)	0.769** (0.32)	0.155 (0.40)	0.130 (0.33)
Population density	0.104 (0.11)	0.252** (0.11)	0.222* (0.12)	0.179 (0.12)	0.173 (0.10)	0.210* (0.12)	0.275 (0.12)	0.101 (0.08)	0.088 (0.11)	0.158 (0.10)
Employment rate	0.052*** (0.02)	0.020* (0.01)	0.011* (0.008)	0.061*** (0.01)	0.015 (0.01)	0.060*** (0.01)	0.059*** (0.01)	0.037** (0.01)	0.032** (0.01)	0.017* (0.01)
Trade openness	-0.080 (0.06)	-0.073 (0.07)	-0.094 (0.08)	-0.001 (0.07)	-0.148** (0.07)	-0.025 (0.07)	-0.054 (0.08)	-0.106 (0.07)	-0.117* (0.06)	-0.135* (0.07)
Partisan Affiliation	0.395** (0.15)	0.722** (0.31)	0.239 (0.18)	0.612*** (0.21)	0.143 (0.15)	0.953*** (0.31)	0.528** (0.23)	0.813** (0.33)	0.131 (0.38)	0.453* (0.35)
Trend	-0.469*** (0.11)	-0.285*** (0.09)	-0.297*** (0.10)	-0.347*** (0.11)	-0.419*** (0.05)	-0.345*** (0.10)	-0.443*** (0.09)	-0.419*** (0.06)	-0.512*** (0.11)	-0.463* (0.09)
Election year t-1	0.347*** (0.11)	0.294** (0.13)	0.348** (0.14)	0.207** (0.11)	0.434*** (0.10)	0.190* (0.11)	0.305*** (0.10)	0.343*** (0.11)	0.494*** (0.17)	0.584*** (0.16)
Election year t	-0.055 (0.02)	-0.482** (0.24)	-0.502** (0.25)	0.672 (0.42)	-1.077*** (0.39)	-0.244 (0.38)	-0.215 (0.49)	-0.241 (0.53)	-0.044 (0.63)	-0.112 (0.29)
Election year t+1	-0.307** (0.12)	-0.357*** (0.09)	-0.391*** (0.09)	-0.090* (0.11)	-0.569*** (0.10)	-0.318 (0.10)	-0.318*** (0.11)	-0.257* (0.14)	-0.497** (0.19)	-0.567*** (0.11)
Spending in neighbours j				0.794*** (0.20)		0.617** (0.20)				
Interact term AF_{it}									0.592** (0.25)	0.673** (0.32)
Fiscal autonomy									-4.405** (1.78)	-4.784* (2.56)
AR(1) test: p-value	0.004	0.001	0.001	0.005	0.000	0.002	0.030	0.000	0.001	0.000
AR(2) test: p-value	0.240	0.138	0.101	0.300	0.102	0.315	0.425	0.209	0.117	0.152
Hansen test: p-value	0.176	0.201	0.126	0.568	0.007	0.502	0.403	0.130	0.242	0.584
Nb of instruments	19	19	19	28	19	28	19	19	25	25
Nb of units	63	63	63	63	63	63	62	62	63	63
N	324	324	324	324	324	324	319	318	324	324

³⁴ Robust standard errors. are in brackets. ***, coefficient significant at 1 % level, **, at 5 % level, *, at 10 % level. We adopt the assumption of weak exogeneity of employment rates and trade openness. The weighted average vector of per capita public spending of other local governments is, as noted before, suspected of endogeneity. Other explanatory variables (Population density, partisan affiliation, trends) are assumed to be strictly exogenous. The lagged levels of variables are used as instruments in the regressions in level as well as in the regressions in difference. We collapse instruments and limit the number. The lags of at least two earlier periods for weak exogenous variables and three earlier periods for endogenous variables are used as instruments. The lagged dependent variable is instrumented by lags of the dependent variable from at least two earlier periods. We use two lags for endogenous and weak exogenous variables.

Table 6: Estimation results for dynamic model 1994-1998 - GMM-System³⁵

Dependent variable: current expenditure of <i>commune i</i> ($G_{i,t}$)	(1) θ^{neigh}	(2) θ^{ethn}	(3) θ^{neigh2}	(4) θ^{ethn2}
Weighting scheme				
Lagged dep. var.	0.835*** (0.14)	0.872*** (0.12)	0.887*** (0.11)	0.965*** (0.10)
Spending in <i>communes j</i>	-0.093 (0.13)	0.205 (0.25)	0.002 (0.35)	0.926 (0.85)
Population density	0.082 (0.05)	0.061 (0.04)	0.041 (0.02)	0.001 (0.05)
Employment rate	0.012 (0.01)	0.008 (0.01)	0.001 (0.01)	0.005 (0.01)
Trade openness	-0.001 (0.005)	-0.001 (0.005)	-0.001 (0.005)	-0.006 (0.005)
Partisan Affiliation	0.095 (0.12)	0.225 (0.23)	0.022 (0.07)	0.001 (0.08)
Trend	-0.001 (0.11)	-0.038 (0.04)	-0.001 (0.03)	-0.168 (0.14)
AR(1) test: p-value	0.001	0.001	0.001	0.001
AR(2) test: p-value	0.840	0.726	0.881	0.751
Hansen test: p-value	0.262	0.467	0.494	0.553
Nb of instruments	16	16	16	16
Nb of units	63	63	62	63
N	241	241	237	241

³⁵ Robust standard errors. are in brackets.***: coefficient significant at 1 % level, **: at 5 % level, *: at 10 % level. We adopt the assumption of weak exogeneity of employment rates and trade openness. The weighted average vector of per capita public spending of other local governments is, as noted before, suspected of endogeneity. Other explanatory variables (Population density, time dummies, election dummies, partisan affiliation, trends) are assumed to be strictly exogenous. The lagged levels of variables are used as instruments in the regressions in level as well as in the regressions in difference. We collapse instruments and limit the number. The lags of at least two earlier periods for weak exogenous variables and three earlier periods for endogenous variables are used as instruments. The lagged dependent variable is instrumented by lags of the dependent variable from at least two earlier periods. We use two lags for endogenous and weak exogenous variables.

Table 7: Testing for yardstick competition - GMM-System³⁶

Dependent variable: current expenditure of 'commune' i ($G_{i,t}$)	(2) θ^{neigh}	(2) θ^{ethn}	(3) θ^{neigh2}	(4) θ^{ethn2}
Weighting scheme				
Spending in non election years	0.915*** (0.11)	1.239*** (0.12)	0.897*** (0.14)	1.013*** (0.11)
Spending in election years	0.989*** (0.09)	1.289*** (0.13)	1.002*** (0.09)	1.449*** (0.26)
Lagged dep. var.	0.569*** (0.22)	0.434** (0.24)	0.695*** (0.21)	0.521* (0.29)
Population density	0.052 (0.11)	0.333*** (0.12)	0.190* (0.10)	0.222 (0.16)
Employment rate	0.068*** (0.01)	0.031** (0.01)	0.070*** (0.01)	0.068*** (0.02)
Trade openness	-0.138* (0.07)	-0.020 (0.08)	-0.212*** (0.07)	-0.016 (0.11)
Partisan Affiliation	0.476** (0.24)	1.510*** (0.28)	0.507** (0.22)	1.462*** (0.36)
Trend	-0.430*** (0.08)	-0.097 (0.08)	-0.482*** (0.07)	-0.365*** (0.12)
Election years	-0.387 (0.40)	0.353 (0.76)	-0.570 (0.72)	3.577* (2.17)
AR(1) test: p-value	0.003	0.001	0.001	0.201
AR(2) test: p-value	0.193	0.186	0.517	0.106
Hansen test: p-value	0.153	0.123	0.492	0.125
Wald test: p-value	0.157	0.438	0.264	0.112
Nb of instruments	20	20	20	20
Nb of units	62	63	62	62
N	324	324	319	318

³⁶ Robust standard errors. are in brackets. ***, coefficient significant at 1 % level, **, at 5 % level, *, at 10 % level. We adopt the assumption of weak exogeneity of employment rates and trade openness. The weighted average vector of per capita public spending of other local governments is, as noted before, suspected of endogeneity. Other explanatory variables (Population density, time dummies, election dummies, partisan affiliation, trends) are assumed to be strictly exogenous. The lagged levels of variables are used as instruments in the regressions in level as well as in the regressions in difference. We collapse instruments and limit the number. The lags of at least two earlier periods for weak exogenous variables and three earlier periods for endogenous variables are used as instruments. The lagged dependent variable is instrumented by lags of the dependent variable from at least two earlier periods. We use two lags for endogenous and weak exogenous variables.