Predator traits determine food-web architecture across ecosystems

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4 Authors and Affiliations:

- 5 Ulrich Brose^{1,2}, Phillippe Archambault³, Andrew D. Barnes^{1,2,4}, Louis-Felix Bersier⁵, Thomas
- 6 Boy^{1,2}, João Canning-Clode^{6, 7, 8}, Erminia Conti⁹, Marta Dias¹⁰, Christoph Digel^{1,11}, Awantha
- 7 Dissanayake^{12, 13}, Augusto A.V. Flores¹⁴, Katarina Fussmann^{1,2}, Benoit Gauzens^{1,2}, Clare
- 8 Gray¹⁵, Johanna Häussler^{1,2}, Myriam R. Hirt^{1,2}, Ute Jacob^{1,16}, Malte Jochum¹⁷, Sonia Kéfi¹⁸,
- 9 Orla McLaughlin¹⁹, Muriel M. MacPherson²⁰, Ellen Latz^{1,2}, Katrin Layer-Dobra²¹, Pierre
- 10 Legagneux^{22,23}, Yuanheng Li^{1,2,24}, Carolina Madeira¹⁰, Neo D. Martinez²⁶, Vanessa
- 11 Mendonça¹⁰, Christian Mulder⁹, Sergio A. Navarrete²⁷, Eoin J. O'Gorman²⁸, David Ott²⁹, José
- 12 Paula¹⁰, Daniel Perkins¹⁵, Denise Piechnik³⁰, Ivan Pokrovsky^{31,32}, David Raffaelli³³, Björn C.
- 13 Rall^{1,2}, Benjamin Rosenbaum^{1,2}, Remo Ryser^{1,2}, Ana Silva³⁴, Esra H. Sohlström^{1,2}, Natalia
- 14 Sokolova³⁵, Murray S. A. Thompson³⁶, Ross M. Thompson³⁷, Fanny Vermandele³⁸, Catarina
- 15 Vinagre¹⁰, Shaopeng Wang^{1,2,39}, Jori M. Wefer^{1,2}, Richard J. Williams⁴⁰, Evie Wieters²⁷, Guy
- 16 Woodward²¹, Alison C. Iles¹
- 17
- 18 1: EcoNetLab, German Centre for Integrative Biodiversity Research (iDiv) Halle-Jena-
- 19 Leipzig, Deutscher Platz 5e, 04103, Leipzig, Germany.
- 20 2: EcoNetLab, Friedrich Schiller University Jena, Dornburger-Str. 159, 07743, Jena,
- 21 Germany
- 22 3: Département de Biologie, Takuvik, Québec-Océan, Université Laval, Québec, Canada
- 23 4: School of Science, University of Waikato, Private Bag 3105, Hamilton, New Zealand
- 24 5: Department of Biology, University of Fribourg, Chemin du Musée 10, 1700 Fribourg,
- 25 Switzerland
- 26 6: MARE Marine and Environmental Sciences Centre, Quinta do Lorde Marina, Sítio da
- 27 Piedade, 9200-044 Caniçal, Madeira Island, Portugal
- 28 7: Centre of IMAR of the University of the Azores, Department of Oceanography and
- 29 Fisheries. Rua Prof. Dr. Frederico Machado, 4, PT-9901-862 Horta, Azores, Portugal
- 30 8: Smithsonian Environmental Research Center, 647 Contees Wharf Road, Edgewater, MD
- 31 21037, USA

- 32 9: Department of Biological, Geological and Environmental Sciences, University of Catania,
- 33 Via Androne 81, 95124, Catania, Italy
- 34 10: MARE Marine and Environmental Sciences Centre, Faculdade de Ciências,
- 35 Universidade de Lisboa, Campo Grande, 1749-016 Lisboa, Portugal
- 36 11: Umweltbundesamt, Wörlitzer Platz 1, 06844 Dessau-Roßlau, Germany.
- 12: School of Biological Sciences, Plymouth University, Plymouth, PL4 8AA, United Kingdom
- 38 13: University of Gibraltar, Europa Point Campus, Gibraltar, GX11 1AA
- 39 14: Centro de Biologia Marinha, Universidade de São Paulo, Rod. Manoel Hipólito do Rego,
- 40 km 131.5, São Sebastião, SP, Brazil
- 41 15: Department of Life Sciences, Whitelands College, University of Roehampton, London
- 42 SW15 4JD, UK
- 43 16: Helmholtz Institute for Functional Marine Biodiversity (HIFMB), Ammerländer Heerstraße
- 44 231, 26129 Oldenburg, Germany
- 45 17. Institute of Plant Sciences, University of Bern, Altenbergrain 21, 3013 Bern, Switzerland
- 46 18: ISEM, CNRS, Université de Montpellier, IRD, EPHE, Montpellier, France
- 47 19: Great Barr Academy, Aldridge Road, Birmingham, B44 8NU, United Kingdom.
- 48 20: Institute of Marine Ecosystem and Fishery Science, University of Hamburg, Germany
- 49 21: Imperial College London, Silwood Park Campus, Buckhurst Road, Ascot, Berkshire, SL5
- 50 7PY, UK
- 51 22: Centre d'Etudes Biologiques de Chizé, UMR 7372 CNRS/Univ La Rochelle, 79360
- 52 Villiers en Bois, France
- 53 23: Département de Biologie, Centre d'Études Nordiques, Université Laval, Québec, Canada
- 54 24: Department of Organismic and Evolutionary Biology, Harvard University, Cambridge, MA
- 55 02138, USA
- 26: Department of Ecology and Evolutionary Biology, University of Arizona, Tucson, Arizona
 85721 USA
- 58 27: Estación Costera de Investigaciones Marinas, Las Cruces, LINCGlobal, Center for
- 59 Applied Ecology and Sustainability (CAPES), Pontificia Universidad Católica de Chile,
- 60 Santiago, Chile
- 61 28: School of Biological Sciences, University of Essex, Wivenhoe Park, Colchester CO4
- 62 3SQ, UK
- 63 29: Institute of Landscape Ecology, University of Münster, Heisenbergstr. 2, 48149, Münster,
- 64 Germany.
- 30: Division of Biological and Health Sciences, University of Pittsburgh at Bradford, Bradford,
- 66 Pennsylvania 16701 USA
- 67 31: Max-Planck Institute for Ornithology, Am Obstberg 1, 78315 Radolfzell, Germany.

- 68 32: Laboratory of Ornithology, Institute of Biological Problems of the North FEB RAS, 18
- 69 Portovaya Str., 685000 Magadan, Russia.
- 70 33: Environment, University of York, York, UK, YO10 5DD
- 71 34: CERIS, Instituto Superior Técnico, Universidade de Lisboa, Avenida Rovisco Pais 1,
- 72 1049-001-Lisboa, Portugal
- 73 35: Arctic research station of Institute of plant and animal ecology, Ural branch, Russian
- 74 Academy of Sciences, 629400, Zelenaya Gorka Str., 21, Labytnangi, Russia
- 75 36: Centre for Environment, Fisheries and Aquaculture Science, Lowestoft Laboratory,
- 76 Pakefield Road, Lowestoft, Suffolk, NR33 0HT, UK
- 37: Institute for Applied Ecology, University of Canberra, Bruce, ACT 2617, Australia
- 78 38: Université du Québec à Rimouski, Département de Biologie, Chimie et Géographie,
- 79 Rimouski, Canada
- 39: Institute of Ecology, College of Urban and Environmental Science, and Key Laboratory
- for Earth Surface Processes of the Ministry of Education, Peking University, 100871 Beijing,
- 82 China
- 40: Rakuten Slice, 800 Concar Dr, San Mateo, CA 94402, USA
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- 85
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88 Abstract

89 Predator-prey interactions in natural ecosystems generate complex food webs that 90 have a simple universal body-size architecture where predators are systematically larger 91 than their prey. Food-web theory shows that the highest predator-prey body-mass ratios 92 found in natural food webs may be especially important as they create weak interactions with 93 slow dynamics that stabilize communities against perturbations and maintain ecosystem 94 functioning. Identifying these vital interactions in real communities typically requires arduous 95 identification of interactions in complex food webs. Here, we overcome this obstacle by 96 developing predator-trait models to predict average body-mass ratios based on a database 97 comprising 290 food webs from freshwater, marine and terrestrial ecosystems across all 98 continents. We analyzed how species traits constrain body-size architecture by changing the 99 slope of the predator-prey body-mass scaling. Across ecosystems, we found high body-100 mass ratios for predator groups with specific trait combinations including (1) small 101 vertebrates and (2) large swimming or flying predators. Including the metabolic and 102 movement types of predators increased the accuracy of predicting which species are 103 engaged in high body-mass ratio interactions. We demonstrate that species traits explain 104 striking patterns in the body-size architecture of natural food webs that underpin the stability 105 and functioning of ecosystems, paving the way for community-level management of the most 106 complex natural ecosystems.

108 Introduction

109 Prey rarely if ever give up their lives willingly to their predators. Predators overcome 110 their prey's resistance by being, on average, larger than their prey, yielding a systematic 111 pattern in communities where the non-cannibalistic average ratio of predator-to-prey body 112 mass (hereafter: body-mass ratio) is higher than unity^{1,2}. The notable exceptions to this 113 include when animals cooperate to overcome larger prey (e.g. pack hunters) and where 114 consumers are parasites or parasitoids. The variation in body-mass ratios within food webs 115 typically spans several orders of magnitude and includes some predators that are smaller 116 than their prey³, but is dominated by situations of the larger feeding on the smaller. The 117 varying body-mass ratios limit which trophic interactions are realized in a community⁴⁻⁶, and the strength of these interactions^{7–9}. Predators typically exert the strongest feeding pressure 118 119 on prey that are one to two orders of magnitude smaller¹, while weaker interaction strengths are realized with prey that are smaller or larger than this size^{10,11}. Specifically, interactions of 120 121 predators with small prey are characterized by high body-mass ratios that yield weak 122 interactions with slow dynamics, which play a central role in maintaining food-web stability¹²⁻ ¹⁶ and ecosystem functioning^{10,17,18}. Moreover, they also buffer natural communities against 123 perturbations from global warming¹⁹, eutrophication²⁰ and secondary extinction waves²¹. 124 125 Therefore, identifying these unique interactions is paramount to determining the stability of 126 natural food webs to perturbations and functioning, but applications of this concept to natural 127 communities have been hampered by the difficulty of describing the myriads of interactions 128 present in natural food webs. Using traits of predator species as proxies of body-mass ratios 129 and the resulting interaction strengths that they govern could provide the means to 130 understand which species are drivers of community stability and functioning without having 131 to perform the often logistically and economically impossible task of quantifying entire 132 interaction networks. This approach could ultimately help predict how extinctions, invasions 133 and other anthropogenic environmental changes affect community stability and functioning 134 through shifts in community trait structure. Despite the importance of understanding how species traits affect body-mass ratios, 135 there is much uncertainty about these relationships. A pioneering study²² showed that 136 137 vertebrate predators exhibit systematically higher body-mass ratios than invertebrates and 138 this has been supported by subsequent analyses also documenting higher body-mass ratios

- 139 in aquatic versus terrestrial communities¹. These studies also found that predator-prey body-
- 140 mass scaling is superlinear with slopes higher than unity, meaning that body-mass ratios
- 141 increase with body mass (see Supplementary Figure 1, red line). However, other studies

142 have proposed that the scaling relationship is either sublinear (decreasing body-mass ratios with body mass; Supplementary Figure 1, yellow line)^{2,22,23} or superlinear^{1,2,23,24}, depending 143 144 on the ecosystem type^{2,23}, predator metabolic group^{23–25} or resource supply²⁶. Additionally, 145 besides body mass and metabolic type, little is known about how body-mass ratios vary with 146 other species traits and across different ecosystem types. Predator and prey movement 147 types and feeding behavior are likely to influence scaling relationships by limiting maximum 148 achievable attack speeds^{24,27}. Interaction dimensionality, which describes whether predators 149 forage in three dimensions (e.g. the water column of lakes and oceans) or on two-150 dimensional surfaces (e.g. epigeic terrestrial or benthic aquatic predators), also influences predator-prey attack rates^{24,28}. As these variables affect the likelihood and strength of 151 152 predator attack rates and scale with individual body mass, we expected that they should also 153 modify the scaling relationship between predator and prey body masses. 154

155 Insert Fig. 1 here

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157 Here, we provide a comprehensive assessment of how species traits modulate 158 predator-prey body-mass scaling relationships and body-mass ratios in natural, complex 159 food webs across an unprecedented range of ecosystems. To do so, we collated the most 160 extensive global food-web database (GlobAl daTabasE of traits and food Web Architecture, 161 GATEWAy version 1.0) to date, comprising 290 food webs (with 222,151 feeding links 162 between 5736 species; see Supplementary Table 1) distributed across the globe (Fig. 1), 163 and including information on four different species traits and five ecosystem types (see 164 Supplementary Table 2). First, we analyzed the scaling of predator and prey body masses 165 over 17 orders of magnitude (fresh masses ranging from the 2 10⁻⁹ g protozoan *Bodo* 166 saltans, to the 275 10⁶ g sperm whale *Physeter microcephalus*). Some prior studies 167 advocated the use of major axis regressions to account for the bidirectional causality 168 between predator and prey body mass¹, whereas others used mixed-effects models to 169 include random effects of the study². As these two types of analyses are mutually exclusive 170 in traditional statistics, we used Bayesian modelling to implement a combination of major 171 axis regressions with mixed effects. Second, we tested for the importance of co-factors in 172 this scaling relationship (ecosystem type, predator and prev metabolic types, interaction 173 dimensionality, predator and prey movement types). These analyses address relationships 174 between species traits and food-web architecture across ecosystems. Third, we developed 175 predictions of average body-mass ratios of predators by their traits, which identifies 176 ecological attributes that broadly predict ecological perturbation stability and functioning in 177 natural communities without requiring detailed knowledge of complex food-web structure.

178 Results

179 In our first analysis, we addressed the scaling of predator and prey body masses. 180 Ordinary least squares (OLS) regressions relating these two variables generate different 181 slopes depending on which variable is chosen as the independent variable (Fig. 2, magenta 182 and blue lines). This discrepancy arises because there is no unidirectional causal 183 relationship between the two variables and because both have measurement errors of the 184 same magnitude, which renders major axis regression the appropriate tool for analyzing 185 these data³¹. Thus, we used Bayesian modelling to fit a major axis regression, which makes 186 no assumptions about a causal relationship between the variables. The major axis 187 regression showed that the overall allometric scaling relationship between predator and prey 188 body mass was superlinear, exhibiting a slope higher than unity (Fig. 2, black solid line, 189 Bayesian major axis regression, slope = 1.315, 95% CI: 1.307-1.323). This suggests that the 190 body-mass ratios between predators and their prey increase with the masses of prey and 191 predators (i.e. the distance between the regression line and the dashed diagonal indicating 192 equally sized predator-prey pairs). On average, interactions between relatively large 193 predator and prey species are characterized by higher body-mass ratios than interactions 194 between smaller species. The Bayesian approach also allowed us to fit mixed-effects 195 models to the data, particularly the random effects on the intercept of the different studies 196 (database variable: link.citation; see Supplementary Table GATEWAy metadata). This 197 yielded a very similar scaling relationship as the non-mixed major axis regression (Fig. 2, 198 green line). Based on the similarity of the results and statistical arguments (see Methods), 199 we have based the following analyses on Bayesian major axis regressions, while results of 200 the analyses with Bayesian mixed-effects major axis regressions are shown in the 201 supplement.

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203 Insert Fig. 2 here

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205 In our second analysis, we used major axis regressions to fit six models of predator-206 prey body mass scaling that each contained one co-variable (ecosystem type, predator or 207 prey metabolic type, predator or prey movement type, interaction dimensionality; see 208 Supplementary Table 2 for variable description). Model comparisons demonstrated that 209 adding any of these co-variables improves the fit substantially over the simple scaling model 210 (Table 1, lower WAIC scores indicate higher model adequacy). According to these WAIC 211 ranks, the best-performing models included predator metabolic type (rank 1) or predator 212 movement type (rank 2) as co-variables, whereas models including the same trait variables

of the prey led to lower ranks (Table 1, ranks 4 and 6). Hence, WAIC values suggested that
predator traits were more important for determining body-mass scaling than prey traits.

- 216 Insert Table 1 here
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218 While the overall relationship was superlinear (Fig. 3, black solid lines), the 219 relationships for ectotherm and endotherm vertebrate predators exhibited strong sublinear 220 scaling, implying that the body-mass ratios of vertebrate predators decrease with their body 221 mass (Fig. 3a). As vertebrate prey often have vertebrate predators, a similar pattern might 222 be expected for the scaling relationship within the prey metabolic groups. Surprisingly, we 223 found superlinear scaling for all vertebrate prey groups (Fig. 3b). Together, these results 224 suggest that the sublinear scaling characterizes vertebrate predators irrespective of whether 225 their prey are vertebrate or invertebrate species.

226 The second most important co-variable in our analyses was predator movement type. 227 Interestingly, we found that swimming, flying and sessile predators exhibit superlinear 228 scaling relationships that are similar to the overall model, whereas walking predators exhibit 229 sublinear scaling (Fig. 3c). Although many walking predators feed on walking prey, our 230 analyses of the prey movement type show superlinear scaling across groups (Fig. 3d). 231 Similar to the metabolic groups, this implies that changes in predator-prey body-mass ratios 232 are mainly driven by predator movement type, irrespective of prey movement type. Together, 233 our analyses of species' traits suggest that the traits of predators have stronger implications 234 for scaling relationships and body-mass ratios than the traits of their prey.

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238 Comparing the two environmental characteristics showed that ecosystem type (WAIC 239 rank 3) improved the model substantially more than interaction dimensionality (2D vs. 3D; 240 rank 5, Table 1). Among ecosystem types, marine, stream and terrestrial aboveground 241 ecosystems follow superlinear scaling similarly to the overall relationship (although streams 242 followed steeper scaling relationships), whereas lake and terrestrial belowground 243 ecosystems exhibit sublinear scaling, parallel to each other (Fig. 4a). Both 2D and 3D 244 interaction dimensionality demonstrate superlinear scaling with a slope similar to the overall 245 pattern. However, 3D interactions tend to involve predator-prey pairs with greater body-mass 246 ratios compared to 2D interactions (Fig. 4b). Although many marine or lake interactions 247 occur in the pelagic 3D part of the ecosystem, ecosystem type does not completely overlap 248 with interaction dimensionality as these aquatic ecosystems also include benthic 2D 249 interactions.

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253 Finally, we addressed how well we can predict which predators in a food web have the 254 highest average body-mass ratios compared with their prey in the absence of information on 255 food-web structure and traits of the prey species. The statistical models thus included the 256 predator traits (body mass, metabolic and movement type) and ecosystem type as 257 independent variables ("predator-trait model", see Methods for details and Supplementary 258 Table 4 for parameters). The overall predator-trait model fitted the data well (Fig. 5a). We 259 found that predictive accuracy varied across ecosystem types and with the fraction of target 260 predators (Fig. 5b). We anticipated that typically a low fraction of predators will be chosen for 261 applied population management and used a fraction of target predators of 25% as an 262 arbitrary example to illustrate our results (Fig. 5b, grey area; note that qualitatively similar 263 results could be obtained for any fraction of 30% or lower). At this fraction of target 264 predators, the accuracy of the predator-trait model predictions is almost always higher than 265 the prediction accuracy when the same fraction of predators is chosen at random (Fig. 5b, 266 diagonal line). An exception to this pattern were the terrestrial belowground systems (Fig. 267 5b), potentially as a consequence of the substantially higher degree of omnivory in soil communities²⁹ or the widespread use of poison by soil predators³². In contrast, the predator-268 269 trait model had high accuracy in streams (89%), marine (61%), terrestrial aboveground 270 (64%) and lake ecosystems (61%), exceeding the 25% accuracy of random predictions (Fig. 271 5b, diagonal line in the grey area). This implies that for these ecosystems the predator-trait 272 model improves the predictions by a factor between 2.44 (marine and lake ecosystems) and 273 3.56 (streams), which is close to the maximum improvement factor of 4 (occurring with 100% 274 prediction accuracy relative to the 25% random prediction accuracy at a fraction of target 275 predators of 25%, grey shaded area in Fig. 5b).

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279 Discussion

Using a global database of 290 food webs we show that (1) the overall allometric scaling relationship between predator and prey body mass is superlinear, implying that the largest species have the highest body-mass ratios and that (2) predator traits (metabolic and movement type) are more important than prey traits in determining these scaling

284 relationships. Subsequently, we developed a predator-trait model that successfully predicted 285 the predators with the highest average body-mass ratio. Food-web theory has shown that 286 these high body-mass ratios yield weak interactions with slow dynamics that are critically 287 important for buffering communities against external perturbations and maintaining ecosystem functioning^{10,12–15,17,18}. Historically, these theoretical results have had little real-288 289 world application, as they require the logistically challenging task of assessing all or at least 290 a large fraction of the food-web links. By focusing on predator traits and ecosystem type 291 while discarding prey traits and the specific links of the food-webs, our predator-trait model 292 provides a generalizable and feasible solution that can bridge the gap between food-web 293 theory and applied ecosystem conservation. For instance, our results suggest that 294 population protection of small vertebrates (e.g. mustelids) and large swimming (e.g. sharks) 295 or flying predators (e.g. birds of prey) might be most effective at buffering natural 296 communities against external perturbations such as extinctions, invasions, pollution, 297 eutrophication and warming. This trait-based approach enables the management of 298 perturbation vulnerability in natural communities without detailed knowledge of the food-web 299 structure.

300 Within the debate over the allometric scaling relationships of predator and prey body-301 masses in natural food webs, the superlinear relationship presented here is consistent with 302 some prior studies^{1,2,23}, while deviating from others that demonstrate sublinear scaling^{2,22,23}. 303 Our comparison of regression methods suggests that this discrepancy could be partially 304 attributed to the alternative use of major axis regressions¹ (consistently yielding superlinear scaling) or ordinary least square (OLS) regressions^{2,22,23} (suggesting superlinear or sublinear 305 306 scaling depending on which is the independent variable). Our comparison of the two OLS 307 regressions with either predator or prey mass as the independent variable reveals 308 substantial uncertainty as they make opposite predictions on how body-mass ratios scale 309 with predator and prey mass, and there is no a priori argument over which OLS regression 310 should be preferred. Hence, major axis regressions are the most appropriate statistical 311 method because: (1) there is no a priori expectation for a causal relationship between 312 predator and prey mass, and (2) both body masses are quantified with the same measurement error³¹. Our results show that major axis regression is not only statistically 313 314 more appropriate but also that the choice of statistical approach has important biological 315 implications for interpretation of the allometric scaling relationship. This approach, combined 316 with our newly compiled food-web database, has enabled refining our understanding of how 317 the scaling relationship between predator and prey body mass varies across ecosystems 318 and between predator-prey combinations of different movement type and metabolic group. 319 Despite the overall superlinear relationship between predator and prey body mass, our 320 analyses identified several species' traits and ecosystem characteristics that are associated

321 with a sublinear scaling relationship. Most notably, both ectotherm and endotherm vertebrate 322 predators demonstrate strong sublinear scaling, making predator metabolic type the most 323 important factor among those we considered for predicting predator-prey body-mass scaling relationships. Consistent with previous research^{1,22,23}, we found that large vertebrate 324 325 predators tend, on average, to feed on prey that are more equally sized (e.g. orcas feeding 326 on minke whales), whereas small vertebrate predators consume relatively smaller prey (e.g. 327 arctic foxes preying on lemmings). This result suggests that large and small vertebrate 328 predators may be constrained by different factors, such as the limitations of maximum attack 329 speed which are only experienced by the largest species²⁷. Interestingly, some (often large) 330 vertebrate predators hunt in groups to attack larger prey to improve their attacking success 331 and overcome the body mass and speed constraints. Indeed, the next most important factor 332 in our analysis was predator movement type, which separates species categories of different 333 speeds (e.g. flying predators are faster than walking predators). In our analysis, walking 334 predators demonstrate sublinear scaling in contrast to all other movement types. The highest 335 body-mass ratios were observed for the largest swimming and flying predators. Further 336 investigations of the physiological constraints related to predator movement type, metabolic 337 type and relative predator-prey body masses on predator feeding rates would help illuminate 338 the processes behind these observed patterns.

339 Generally, our model selection results suggest that predator metabolic and movement 340 traits had much stronger effects on the scaling relationship than the equivalent prey traits. 341 This is partially supported by the greater similarity between the major axis regression 342 (accounting for bi-directional causalities) and the OLS regression with prey body mass as 343 the dependent variable. We therefore conclude that top-down prey selection by predators 344 has a stronger effect on prey mass than does the bottom-up influence of prey mass on 345 predator masses. It is likely that both top-down and bottom-up influences are important, but 346 our results indicate the dominance of the former, which stimulated the development of the 347 predator-trait models of our third analyses predicting which predators have the highest 348 average body-mass ratios across food webs.

349 Our results also identify ecosystem type as an important co-factor of the predator-prey body-mass scaling relationship, which is generally consistent with prior studies^{23,24,33}. We 350 351 expected this effect to be partially explained by the habitat dimensionality of the interaction 352 (2D or 3D), which has important consequences for the strength of predator attack rates^{24,28}. 353 Although we found an effect of interaction dimensionality with overall higher body-mass 354 ratios in 3D than in 2D habitats, surprisingly it did not explain the different scaling 355 relationships between different ecosystem types. This may be explained by the fact that the 356 ecosystem type varies across food webs, whereas variance in interaction dimensionality 357 plays an important role across the different predator-prey pairs with food webs. The superior

358 explanatory power of the model including ecosystem type compared to that including 359 interaction habitat dimensionality suggests that there are ecosystem characteristics not 360 related to dimensionality, such as laminar viscosity, that may have a stronger effect on 361 predator-prey interactions. We found relatively high body-mass ratios and a very steep body-362 mass scaling relationship in stream ecosystems. The streams exhibit some differences to 363 the other ecosystem types of our database: (1) the higher physical drag force of the water, 364 (2) the higher dependence on allochthonous resources, (3) the dendritic environmental 365 structure, and (4) the relatively narrower range of body masses included in our data. While 366 each of these points could be responsible for the difference in scaling relationships, the last 367 point calls for additional data on stream interactions between larger species such as fish to 368 see if the steep increase in the scaling relationship holds. While terrestrial aboveground and 369 marine interactions exhibited superlinear scaling relationships as the overall relationship, 370 those of lake and terrestrial belowground systems were sublinear. Furthermore, the lack of 371 vertebrate predators with high body masses and high body-mass ratios may at least partially 372 explain this for soil communities, but this surprising result requires more mechanistic 373 investigation of the so far untested similarity between lake and belowground interactions.

374 Our approach to characterize predator-prey body-mass ratios in natural food webs has 375 some limitations. First, in order to encompass a wide range of body masses, taxonomy and 376 ecosystem types, we assume that interacting individuals have population-average body masses³⁴. As in prior studies^{1,22,23,33}, we rely on population-averaged body masses, since we 377 378 do rarely have measurements for the actual body masses of the interacting individuals. 379 Thus, for many predator species, particularly those with ontogenetic diet shifts, actual body-380 mass ratios are likely to have a lower variation than body-mass ratios calculated from 381 population averages. Unfortunately, the lack of individual data for entire food webs across 382 ecosystems hampers any alternative approach. As prior comparisons of individual-based versus population-based food webs have shown^{34,35}, our population-based approach likely 383 384 underestimates the intercepts of the scaling relationships. Second, the study sampling 385 design, environmental factors such as temperature and the species' phylogeny may also affect the scaling relationship^{25,36–38}, and these would ideally be included as co-variables in 386 387 the analyses. As these data were not systematically available for the data sets included, we 388 accounted for them by random effects in mixed models², which leaves the need for more 389 detailed analyses for future studies. As major axis regressions with random effects are not 390 generally available, we addressed this issue by using Bayesian models throughout the 391 study, which allowed comparisons with hierarchical models including random effects (i.e. 392 mixed-effects models). Although the mixed effects model results do not change our findings 393 substantially (see Supplement for a comparison between mixed and non-mixed Bayesian 394 models) and, due to potentially confounding clustering effects (see Methods), we have

395 focused our analysis on the model without mixed effects. Third, our analyses were restricted 396 to predator-prey interactions, whereas interactions of other consumer types such as 397 parasites, parasitoids or herbivores were excluded. As these interaction types are typically characterized by different body-mass ratios^{1,39,40}, future studies should address their scaling 398 399 relationships in our GATEWAy database. Fourth, we employed simple scaling relationships 400 with up to one single co-variable to gain an in-depth mechanistic understanding, whereas 401 models with interactions between multiple co-variables were omitted from our analyses of 402 the predator-prey body-mass scaling. These more complex relationships with higher order 403 interactive effects, however, could be addressed by black box approaches such as machine 404 learning algorithms, which could provide accurate predictions of food-web structures³². Fifth, 405 our study illustrates systematic differences in body-mass ratios across ecosystem types and 406 species' traits, whereas explanations for these differences remain to be revealed by studies 407 integrating mechanistic models with our data.

408 Our analyses provide insights into how predator and prey body masses scale with 409 each other in natural food webs. The discovery that predator traits are more important than 410 prey traits in predicting body-mass scaling and that ecosystem type has a greater effect than 411 interaction dimensionality offers new possibilities for understanding and predicting 412 differences in food-web structure, community stability and ecosystem functioning across 413 community and ecosystem types. Specifically, our results highlight that critically important 414 high body-mass ratios occur in interactions with predators that are (1) small vertebrates or 415 (2) large swimming or flying species. With only three species traits (body mass, metabolic 416 and movement type), our models were able to predict which 25% of the predators possess 417 the highest average body-mass ratios with surprisingly high accuracy in most ecosystem 418 types (58-89%). We anticipate that this accuracy will be increased by additional species traits (e.g. predation strategy, use of poison, sub-habitat association) that compose the 419 multiple dimensions of natural food webs^{6,32}. Our trait-based food-web analyses enable 420 421 generalizations of food-web theory from the food webs studied to the vast majority of 422 communities for which only species and trait information is available. Updated with additional 423 traits, this approach has great potential for managing ecosystem functioning and stability 424 against external perturbations such as pollution, eutrophication and warming without full 425 knowledge of food-web structure. The trait-based body-mass ratio approach therefore 426 presents an important integration of food-web theory with applied ecosystem management 427 that provides a theoretical foundation for the community-level conservation of the most 428 complex natural ecosystems.

429 Methods

430 We compiled a global database of traits and food-web architecture (GATEWAy version 431 1.0, see Supplement), where each link is characterized by the taxonomy and trait variables 432 of both the consumer and the resource (see Supplement metadata for variables). We 433 included food webs with (1) a sufficient quality in terms of taxonomic resolution, which 434 prevents nodes aggregating species with very different trophic interactions; (2) a reasonable 435 completeness integrating all trophic levels and community compartments; (3) trait 436 information for the trophic species including at least their population-averaged body mass, 437 their metabolic type and their movement type (see Supplementary metadata table for 438 definitions); (4) information for each trophic link such as the type (e.g. predacious), the 439 dimensionality (2D and 3D) and the classification (individual-based and non-individual-440 based) (see Supplement metadata table for definitions); (5) descriptors for the ecosystems 441 such as the ecosystem type and the geographic location. 442 In our analyses, we focused on predatory (variable: interaction.type) and individual-443 based (variable: interaction.classification) interactions. The former excludes interactions of 444 other types (e.g. herbivorous, detritivorous, parasitic, parasitoid), whereas the latter discards 445 interactions of consumers attacking groups, swarms or films of resources (e.g. filter feeding, 446 grazing). Some of the studies included in our database sampled the same ecosystem at 447 different locations, resulting in replicated predator-prey species pairs²⁹. To avoid 448 pseudoreplication, each unique combination of taxonomy, life stage, and individual body 449 mass for predator and prey species was included only once. After exclusion of interactions 450 with missing variables, the resulting data included 88,197 unique predator-prey interactions 451 among the original 222,151 feeding links.

452 First, we analyzed the reduced data for the relationship between the base-10 453 logarithms (log₁₀) of predator and prey body masses [gram fresh mass]. We compared the fit 454 of two ordinary least squares (OLS) regressions (either predator mass or prey mass as the 455 dependent variable) to that of a major axis regression and a mixed-effects major axis 456 regression including random effects on the intercept of the different studies (variable: 457 link.citation). Traditional methods only allow to fit either major axis regressions or mixed 458 models with random effects. Hence, our aim of comparing major axis regressions with and 459 without random effects (i.e. random intercepts for each study) could only be achieved by 460 realizing models that were fitted by Bayesian methods using the RStan package³⁰ (see 461 supplementary statistical methods for details). Consistent with traditional major axis 462 regressions, we minimized the sum of squared orthogonal distances of the observations

463 (x,y) to the regression line³¹ instead of the vertical distance (y) as in OLS (model I) 464 regressions.

465 Second, we used Bayesian major axis models to compare the fit of the simple scaling 466 model to six models, whereby each included one co-variable: ecosystem type, predator or 467 prey metabolic type, predator or prey movement type or interaction dimensionality (see 468 Supplementary Table 2 for variables). Overall, the results were mostly consistent between 469 the mixed-effects and non-mixed models. The mixed-effects models fit the relationships 470 separately for each study. As the body-mass ranges within studies do not cover the entire 471 body-mass gradient and the number of data points within studies is much lower than in the 472 entire database, some of the fitted scaling relationships can become arbitrary as single 473 points can strongly affect the slope. Averaging across all slopes and all intercepts using 474 hierarchical approaches can lead to clusters of such arbitrary slopes, which can exert 475 substantial leverage on the average relationship across all studies. In our data, the clustering 476 remained even when using random intercepts and a fixed slope across all studies. 477 Therefore, the mixed-effects modelling of our data suffered from two limitations: (1) it loses 478 information about the overall trend across the whole database (i.e. none of the study-specific 479 scaling relationships spans the entire body-mass gradient), and (2) the joint mean slope and 480 intercept are affected by partially arbitrary slopes (data sets with few points). As both 481 regressions also yielded qualitatively similar results, we report the results of the non-mixed 482 major axis regressions in the manuscript (Figs. 3, 4) with comparisons to the fits of the mixed 483 major axis regressions in the supplement (Supplementary Figures 2-7). Model comparison 484 (based on their WAIC values, Watanabe-Akaike Information Criterion) of these seven 485 models (the simple model without co-variable and the six models with one co-variable each) 486 provided a ranking of their performance, and we used the model parameters to gain an 487 understanding of how they modify the relationship. In the analysis of predator-prey body-488 mass scaling, we refrained from analyzing more complex models with interactions between 489 these co-variales for three reasons: (1) they imply impossible combinations (e.g. swimming 490 predators in terrestrial ecosystems), (2) their higher order interactions hamper the 491 mechanistic understanding of individual effects, and (3) their strong collinearity causes 492 interference between factors. 493 Third, we analyzed our database for the dependence of the predators' average 494 predator-prey body-mass ratios on predator traits (body mass, metabolic and movement

type) and ecosystem type. To avoid circularity in the statistical model (predator body mass in
both the dependent and independent variables), we fitted Bayesian major axis regressions
with log₁₀ prey mass as the dependent and log₁₀ predator mass as the independent variable

498 with the co-variables predator metabolic type, predator movement type, and ecosystem type.

499 We restructured the resulting predator-trait model equation to calculate the effect of the

independent and co-variables on predator-prey body-mass ratios. By discarding prey
species traits, these analyses allow prediction of which predators in a community have the
highest average body-mass ratios without knowledge of the predator-prey links.

503 The accuracy of this approach was determined in a five-step cross-validation process. 504 First, we chose one of the food webs ("test data") and ranked its predators according to their 505 empirical average body-mass ratios. Second, we ran the predator-trait regression model 506 described above in the remaining database containing the other 289 food webs ("training 507 data") to predict the predators' average body-mass ratios depending on their traits. Third, we 508 calculated the proportion of predators that were correctly predicted by this "predator-trait 509 model" (hereafter: accuracy) for a fraction x of the highest ranked predators of the test-data 510 food web (hereafter: fraction of target predators). For example, a fraction of target predators 511 of 0.1 implies that the 10% highest ranked predators (i.e. those with the highest average 512 body-mass ratios) of the empirical "test data" are compared to the 10% highest ranked 513 predators as predicted by the predator-trait model of the "training data". An exemplary 514 accuracy of 0.8 would indicate an 80% overlap between the two species lists. Fourth, this 515 assessment of prediction accuracy was systematically replicated across a gradient in the 516 fraction of target predators x between 5% and 95% (steps of 5%). Finally, these four steps 517 were repeated for each of the 290 food webs independently to calculate the average 518 accuracy across food webs depending on the fraction of target predators.

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527 Data availability

528 The data supporting the findings of this study (GATEWAy 1.0) are available at the iDiv data 529 repository⁴¹.

530 Code availability

531 The R code of the statistical analyses is available as a supplement.

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619

620 Author contributions

- 621 U.B. developed the study design. All authors gathered, contributed or organized data. U.B.
- and B.R. carried out statistical analyses. M.R.H. made the figures. U.B. and A.C.I. wrote the
- 623 first draft of the manuscript. All authors discussed the results and commented on the
- 624 manuscript.

625 Competing interests

626 The authors declare no competing financial interests.

- 628 **Table 1**: Comparison (Watanabe-Akaike Information Criterion, WAIC) of six predator-prey
- 629 body-mass scaling models with one co-variable. Bayesian major axis models (ma) as in Fig.
- 630 3 and Fig. 4 of the main manuscript and mixed Bayesian major axis models as in
- 631 Supplementary Figures 2-7.

Co-variable	Bayesian ma		Bayesian ma mixed	
	WAIC	Rank	WAIC	rank
Predator metabolic group	2.414 x 10 ⁵	1	2.229 x 10 ⁵	1
Predator movement type	2.720 x 10 ⁵	2	2.520 x 10⁵	2
Ecosystem type	2.722 x 10⁵	3	2.566 x 10⁵	4
Prey metabolic group	2.807 x 10 ⁵	4	2.563 x 10⁵	3
Interaction dimensionality	2.818 x 10⁵	5	2.616 x 10⁵	6
Prey movement type	2.830 x 10 ⁵	6	2.605 x 10 ⁵	5
None	2.859 x 10⁵	7	2.657 x 10⁵	7

⁶³⁴ Figure legends

Fig. 1: The global distribution of food webs in GATEWAy (GlobAl daTabasE of traits andfood Web Architecture, GATEWAy version 1.0, see Supplement).

Fig. 2: Overall scaling of predator and prey body mass assessed by four regression

638 methods (n=88,197). Ordinary least squares regression (OLS) of prey mass depending on

639 predator mass (blue line), ordinary least squares regression of predator mass depending on

640 prey mass (magenta line), Bayesian major axis regression (black line), mixed Bayesian

641 major axis (ma) regression with random intercepts (green line). Bayesian regression

642 parameters are the means of the posterior distributions. The dashed line indicates equal

body masses of predator and prey for comparison. All body masses are gram fresh masses.

644 See Supplementary Table 3 for model parameters. Our extensive statistical considerations

645 hold that the Bayesian major axis regression (black line) is the most appropriate model.

Fig. 3: Species' traits constrain the scaling of log₁₀ predator body mass with log₁₀ prey body

647 mass (n=88,197): (a) predator metabolic type, (b) prey metabolic type, (c) predator

648 movement type, (d) prey movement type. Solid black lines represent the overall scaling

relationship, and the colored lines show the relationships for subgroups. Bayesian

650 regression parameters are the means of the posterior distributions. Dashed lines indicate

equal body masses of predator and prey for comparisons. See Supplementary Table 3 formodel parameters.

Fig. 4: Ecosystem characteristics constrain the scaling of log₁₀ predator body mass with log₁₀ prey body mass (n=88,197): **(a)** ecosystem type, **(b)** interaction dimensionality. Solid black lines represent the overall scaling relationship, and the coloured lines show the relationships for the subgroups. Bayesian regression parameters are the means of the posterior distributions. Dashed lines indicate equal body masses of predator and prey for

658 comparisons. See Supplementary Table 3 for model parameters.

659

Fig. 5: The predator-trait model predicts the target predators with the highest body-mass
 ratios across different ecosystem types (color code) (n=7296). (a) Observed versus

662 predicted average body-mass ratios characterize the goodness of fit (R²=0.633,

663 RMSE=0.914). Dashed diagonal line shows where observations and predictions are

664 identical. **(b)** Accuracy (proportion of correct predictions) in an out-of-sample food web

665 depending on the fraction of target predators to be predicted. The diagonal line characterizes

666 predictions when predators are chosen at random. Grey area corresponds to an exemplary

- 667 fraction of the 25% of the predators with the highest body-mass ratios. See Supplementary
- 668 Figure 8 for variation in accuracy across the individual food webs.









