

1 **The role of C-reactive protein (CRP) as a prognostic marker in COVID-19**

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1 **Abstract**

2 **Background**

3 C-reactive protein (CRP) is a non-specific acute phase reactant elevated in infection
4 or inflammation. Higher levels indicate more severe infection and has been used as
5 an indicator of COVID-19 disease severity. However, the evidence for CRP as a
6 prognostic marker is yet to be determined. The aim of this study is to examine the
7 CRP response in patients hospitalised with COVID-19 and determine the utility of
8 CRP on admission for predicting inpatient mortality.

9 **Methods**

10 Data were collected between February 27th and 10th June incorporating two cohorts:
11 the COPE (COVID-19 in Older People) study of 1564 adult patients with a diagnosis
12 of COVID-19 admitted to 11 hospital sites (test cohort) and a later validation cohort
13 of 271 patients. Admission CRP was investigated and finite mixture models were fit
14 to assess the likely underlying distribution. Further, different prognostic thresholds of
15 CRP were analysed in a time-to-mortality Cox regression to determine a cut-off.
16 Bootstrapping was used to compare model performance (Harrell's C and AIC).

17 **Results**

18 The test and validation cohort distribution of CRP was not affected by age and
19 mixture models indicated a bimodal distribution. A threshold cut-off of CRP ≥ 40 mg/L
20 performed well to predict mortality (and performed similarly to treating CRP as a
21 linear variable).

22 **Conclusions**

23 The distributional characteristics of CRP indicated an optimal cut-off of ≥ 40 mg/L
24 was found associated with mortality. This threshold may assist clinicians in using
25 CRP as an early trigger for enhanced observation, treatment decisions, and
26 advanced care planning.

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1 **Key Messages**

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- CRP has been used inconsistently in both patient management and as a prognostic marker during COVID-19.

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- Admission elevated CRP for patients with COVID-19 was associated with increased inpatient mortality and indicative of disease severity at admission.

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- The distribution of CRP at admission was found to be bimodally distributed, and a CRP ≥ 40 mg/L was the optimal threshold of increased risk of mortality.

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- Admission CRP ≥ 40 mg/L may be used by treating clinicians as an early warning of for enhanced care and patient centred decision making.

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11

1 **Introduction**

2 Elevated levels of serum C-reactive protein (CRP) have been observed in patients
3 with COVID-19, and used to assist with triage, diagnostics, and prognostication (1,2).

4 CRP is a non-specific acute phase protein that is produced by hepatocytes and
5 elevated in acute infection or inflammation (3). Secretion begins 4-10 hours after an
6 inflammatory insult and peaks at 48 hours with a short half-life of 19 hours. Crucially
7 it may be elevated before a patients' vital signs are affected or leucocytes are raised
8 (3).The profile of this biomarker has made CRP useful and routinely available in
9 clinical medicine for diagnostics.

10 CRP can be used to assist with differentiation between viral and bacterial infections,
11 for example, influenza produces a mean CRP of 25.65 mg/L (CI 18.88 mg/L -
12 32.41mg/L) versus bacterial pneumonia a CRP of 135.96 mg/L (CI 99.38 mg/L -
13 172.54 mg/L) (4). In COVID-19 a CRP of ≥ 4 mg/L has been shown to be useful for
14 triaging suspected cases when comparing PCR positive patients versus negative
15 controls who have presented to a fever clinic with respiratory symptoms or a high
16 temperature (OR 4.75 CI 3.28 – 6.88) (5).

17 However, debate remains over the utility of CRP as a prognostic marker for patients
18 admitted to hospital with COVID-19. In a recent systematic review, 10 of the 22
19 included COVID-19 prognostic models treated CRP either as a factor or covariate
20 (6). Most these studies used CRP with a binary threshold, proposed values to predict
21 inpatient mortality varied from ≥ 10 mg/L to ≥ 76 mg/L. In addition to a binary
22 threshold, CRP has been examined in a trichotomized model with the two thresholds
23 at ≥ 40 mg/L and ≥ 100 mg/L (9). A lower cut-off of ≥ 20.44 mg/L was used as a
24 threshold related lung injury (7), and >32.5 mg/L was found to offer 80% predictive
25 power for a person needing mechanical ventilation (8). The studies adjusted for
26 admission CRP as a covariate to account for baseline disease severity have
27 assumed a linear or natural logarithm transformation ($\text{Ln}(\text{CRP})$) relationship with
28 outcome (10,11) . Although using CRP in a continuous manner may offer an
29 improved understanding of the contribution of CRP within each analysis, it does not
30 allow CRP to be used by clinical teams to guide management of patients with
31 COVID-19.

1 Whilst CRP has been argued as an important marker of disease progression in
2 COVID-19 (6), its distribution has never been explored to understand whether
3 distinct patterns exist in a heterogeneous population. The use of CRP as a
4 biomarker in COVID-19 may present a quick and accessible tool in clinical
5 management, and trigger longer periods of enhanced observation, and may provide
6 information around likely disease progression, and assist with early therapeutic,
7 ventilation, and palliative care discussions.

8 The aim of this study is to examine the distribution of CRP at hospital admission, and
9 objectives are to: 1) assess CRP as a prognostic bimodal, or trimodal distribution; 2)
10 propose and compare the categorisation of CRP as a prognostic marker to either a
11 linear or log-linear measure of CRP.

12 **Methods**

13 **Study design**

14 This observational study used two cohorts at different time points to examine the
15 contribution of CRP to clinical outcomes. Permission to conduct this study was
16 granted in the UK by the Health Research Authority (20/HRA/1898), and in Italy by
17 the ethics committee of University Hospital of Modena Policlinico
18 (369/2020/OSS/AOUMO). Written consent was not required from participants as per
19 ethical review. This study has been written in accordance with the STROBE
20 statement (12).

21 **Settings**

22 Thirteen hospital sites participated, twelve from the UK and one from Italy. All were
23 acute hospitals directly admitting patients with suspected or confirmed COVID-19.

24 **Participants**

25 *Original Cohort (Cohort 1)*

26 Participants in Cohort 1 were included as part of the COPE study (COVID in Older
27 People study) as reported in the paper by Hewitt, Carter et al (13)(14). Briefly, this
28 was a European multicentre observational study recruiting 1564 hospitalised adults
29 between February 27th to April 28th 2020 with either SARS-CoV-2 viral PCR
30 confirmed disease (95.9%) or clinically diagnosed (4.1%) COVID-19. Any patients

1 aged 18 years or older admitted to the participating hospitals with a diagnosis of
2 COVID-19 were included. The study found frailty was associated with longer hospital
3 stay, and a better predictor of mortality as an inpatient, and at day 7, than age or
4 comorbidity alone.

5 *Validation Cohort (Cohort 2)*

6 Cohort 2 consisted of an additional 271 patients recruited between 29th April and 10th
7 June 2020 from a combination of six of Cohort 1's hospitals plus two additional
8 recruiting hospitals. All patients were SARS-CoV-2 viral PCR positive.

9 **Variables**

10 A prognostic threshold for CRP was needed within the COPE protocol (March 2020).
11 The limited literature available early in the pandemic included a case series of 73
12 patients with COVID-19 presenting with a mean CRP of 51.4 mg/L (SD 41.8) (1).
13 Based on this paper, and proposed by the clinical experience of the authors who
14 delivered acute care a dichotomous threshold was chosen with <40 mg/L (lower
15 admission CRP), and ≥ 40 mg/L (CRP-elevated, indicating increased disease severity
16 (14)).

17

18 **Data sources**

19 CRP was measured at hospital admission and transcribed from patients' medical
20 records. There was no attempt to standardise the CRP assay between sites.
21 A standardised case reporting form was used for all hospital sites. Data were
22 transferred to King's College London in anonymous format for statistical analysis.

23

24 **Graphical Data Analysis**

25 Using the test cohort the distribution of CRP was examined graphically and stratified
26 by age. Finite bivariate and trivariate Gaussian mixture models were fit to CRP,
27 representing two and three latent classes respectively. The theoretical distribution
28 from these models was compared to the empirical data and the threshold between
29 the two and three classes examined. The normality assumptions were assessed
30 visually

31

32 **Statistical Analysis**

1 *Primary Analysis: Mixture modelling analysis*

2 The empirical data from the test cohort were fit to a Gaussian mixture model with
3 one, two or three components using an Expectation-Maximisation algorithm (to refine
4 the starting values) then maximum likelihood estimation (Stata routine “*fmm*”). The
5 models were compared using the AIC and the thresholds determined by the
6 posterior probability of belonging to the two or three class models.

7 *Secondary Analysis: Prognostic modelling analysis*

8 To assess differing thresholds for CRP as a prognostic factor of outcome, a series of
9 mixed-effects multivariable Cox proportional hazards models for time to mortality, in
10 a method consistent with the COPE study primary analysis (13). The model was
11 adjusted for elevated CRP using a level of ≥ 40 mg/L, in addition to: patient age
12 group (<65, 65-79, ≥ 80 years old), sex, diabetes (yes/no), hypertension (yes/no),
13 coronary artery disease (yes/no), and kidney disease (eGFR <60 ml/min/1.73m²).
14 Dichotomized thresholds of CRP were compared within a range of 10mg/L to
15 100mg/L in 5mg/L intervals (≥ 10 mg/L, ≥ 15 mg/L, etc). Model performance was
16 evaluated and compared using Harrell’s C and the Akaike information criterion (AIC)
17 (15). We compared the dichotomised thresholds against linear CRP and Ln(CRP) (as
18 CRP is known to be skewed) as benchmarks of performance. This method was
19 chosen as dichotomising results can lead to a loss of information resulting in a lower
20 predictive power compared to using a continuous measure (16). Bootstrapping was
21 used to construct 95% percentile confidence intervals for differences in model
22 performance between the best-fitting models. Bootstrapping was stratified by site
23 with 1000 replications for each comparison. A complete case analysis was used in
24 all cases due to negligible missing data (<4%).

25

26 **Validation Cohort (Cohort 2)**

27 To provide an indication of whether the original results from Cohort one were likely to
28 be replicable to a wider group of patients with COVID-19, the analysis was repeated
29 on an independent validation sample (Cohort 2). Using the validation cohort, two-
30 class and three-class mixture models were estimated using the empirical data
31 without restriction. On evidence of overfitting, to assess the additional benefit of a
32 very elevated category for CRP, the validation cohort was fitted using a three-class

1 mixture model, with the class-two mean fixed using the validation cohort two-class
2 mixture model mean.

3

4 **Comparison of the prognostic effect of CRP**

5 Using a mixed-effect multivariable Cox regression, the effect of elevated CRP will be
6 reported using an adjusted hazards ratio (aHR), alongside the respective 95%
7 confidence interval (95%CI), for a linear CRP, $\text{Ln}(\text{CRP})$.

8

9 **Role of the funding source**

10 There was no funding source for this study. The corresponding author had full
11 access to all the data in the study and had final responsibility for the decision to
12 submit for publication.

13 **Results**

14 The study included 1835 patients across Cohorts 1 and 2, which were drawn from 12
15 hospitals in the UK and one from Italy. Of the total study participants, 26.4% (n=484)
16 died in-hospital, varying between sites from 13.3% to 42.9%. A comparison for those
17 who died in hospital was carried out in Table 1 split into Cohort 1 (n=1564) and
18 Cohort 2 (n=271). In Cohort 1, 27.2% died and the median CRP for those that died
19 was 115 mg/L (63 mg/L – 191 mg/L [IQR]) compared to 69 mg/L (29 mg/L – 140
20 mg/L) among those who survived. For patients with $\text{CRP} \geq 40$ mg/L mortality was
21 31.9% compared to 15.0% for patients with $\text{CRP} < 40$ mg/L. Median follow up time
22 (time to mortality or discharge) was 13 days (6 – 22 days).

23

24 Cohort 2 experienced 21.8% mortality. Among those who died, median CRP was 86
25 mg/L (48 mg/L – 173.5 mg/L) compared to 53 mg/L (16 mg/L – 109 mg/L) among
26 those who survived. For patients with $\text{CRP} \geq 40$ mg/L mortality was 28.6% compared
27 to 10.4% for patients with $\text{CRP} < 40$ mg/L. The median follow up time (time to death
28 or discharge) was 10 days (5 – 18 days).

29

30 **Results of Cohort 1 (n=1564)**

31 **Distribution of CRP**

32 On graphical examination of the distribution of $\text{Ln}(\text{CRP})$, it exhibited negative skew,
33 with two “peaks” suggestive of a bimodal distribution, see Figure 1: Plot (i), and

1 Supplementary Figure 1: Plot(i, ii). The distribution of $\text{Ln}(\text{CRP})$ was observed in age
2 stratified groups of <65, 65-79, and ≥ 80 years old. On inspection there was no
3 difference between the distribution age stratified, or the complete data set.

4 5 **Primary Analysis: Mixture Modelling Analysis**

6 Following the two suggested peaks in the examination of the $\text{Ln}(\text{CRP})$ distribution a
7 two-latent class finite mixture model was fitted. It appeared to graphically fit the data
8 when examined against the empirical distribution in Figure 1: Plot (i). This was
9 supported by a comparison with the one-class (or null) model, which displayed a
10 higher AIC (4739 compared to 4524). The simple threshold at which the predicted
11 probability of belonging to a two-class model being greater than one-class was 38
12 mg/L. This will be implemented as ≥ 40 mg/L herein to account for the imprecision of
13 the measurement of CRP, and also for ease of recall in a busy clinical setting.

14 The three-class finite mixture model fit slightly better than the two-class finite mixture
15 model (AIC of 4484) with probability of class-one membership highest between
16 range 0-14 mg/L, class-two between 15-120 mg/L, and class-three for values of CRP
17 ≥ 120 mg/L, see Figure 1: Plot (iii).

18 The primary analysis proposed a single optimal threshold of CRP ≥ 40 mg/L to
19 indicate elevated CRP.

20 **Secondary Analysis: Prognostic modelling**

21 The time to mortality analysis included 1502 participants (96%) in the complete case
22 population. A cut-off of ≥ 65 mg/L appeared to fit best in the sample on all measures
23 (Harrell's C of 0.7068, AIC of 5124) (Table 2) after fitting different binary
24 categorisations of CRP in a Cox model for time to mortality. Differences in measures
25 of goodness of fit were small especially between cut-offs in the range of ≥ 40 mg/L to
26 ≥ 90 mg/L. CRP as a continuous $\text{Ln}(\text{CRP})$ measure performed considerably better
27 (Harrell's C of 0.7157, AIC of 5001), and with little improvement on this using a linear
28 scale (Harrell's C of 0.7040, AIC of 5024). Bootstrapped differences in the measures
29 of goodness of fit between a cut-off of ≥ 40 mg/L and the marginally better performing
30 cut-off of ≥ 65 mg/L; no difference in performance was seen with 95%CI for all
31 measures (Table 3). There was evidence that both a cut-off of ≥ 40 mg/L and of ≥ 65

1 mg/L outperformed a cut-off of ≥ 10 mg/L, the upper limit of the normal range for CRP
2 (17). It should be noted that that $\text{Ln}(\text{CRP})$ was the optimal parameterisation compared
3 to either ≥ 40 mg/L (-135.1 AIC, bootstrapped 95% CI -210.4 to -65.1) or ≥ 65 mg/L (-
4 123.5 AIC, bootstrapped 95% CI -197.6, to -55.8)

5

6 **Results of Cohort 2 (n=271)**

7 **Distribution of CRP**

8 Cohort 2 included 271 new patients from eight hospital sites: 85 (31.4%) were fully
9 independent, recruited from two new hospital sites; 186 were pseudo-independent,
10 being newly recruited patients from original hospital sites in Cohort 1. There was no
11 difference in the demographic, comorbidities and distribution of CRP was seen in
12 Cohort 2 and Cohort 1 (Table 1).

13

14 **Fitting finite mixture models**

15 The empirical distribution of the Cohort 2 $\text{Ln}(\text{CRP})$ appeared to graphically reasonable
16 a similar pattern to Cohort 1, see Figure 1: Plot (ii). The two-class finite mixture
17 model gave a consistent threshold (CRP ≥ 41 mg/L). The unrestricted three-class
18 finite mixture model exhibited likely overfitting to the data on examination of the
19 distributions. Inconclusive evidence for the additional second cut-off was found with
20 the class three distribution entirely contained within class-two, with a large variance.
21 There was no additional benefit for fixing the central distribution mean and allowing
22 the mixture proportion to vary, but this can be seen graphically in Figure 1: Plot (iv).
23 The simple threshold between class-one and class-two was ≥ 41 mg/L.

24

25 The time to mortality analysis included 208 of the participants (77%) with complete
26 data. Fitting different binary categorisations of CRP in a Cox model for time to
27 mortality gave a CRP cut-off of ≥ 40 mg/L as the best fitting model (Harrell's C of
28 0.7187, AIC of 424), outperforming the $\text{Ln}(\text{CRP})$ model (Harrell's C of 0.7014, AIC of
29 427), see Table 2. There was no evidence of difference in performance between a
30 cut-off ≥ 65 mg/L and ≥ 40 mg/L, or between ≥ 40 mg/L and $\text{Ln}(\text{CRP})$ on examination of
31 bootstrapped 95%CI Supplementary Table 1.

32

33 **The prognostic effect of elevated CRP with prognostic properties**

1 The aHR for CRP ≥ 40 mg/L were 2.58 (95%CI 1.95 to 3.41) and 2.61 (95%CI 0.54 to
2 4.63) for Cohorts 1 and 2 and the estimate of CRP appeared stable (Supplementary
3 Table 2). For comparison CRP ≥ 65 mg/L, the aHR was consistent in Cohort 1
4 (aHR=2.48; 95%CI 1.96 to 3.14) but appeared unstable in Cohort 2 (aHR=1.61;
5 95%CI 0.84 to 3.09). Using a cut-off of ≥ 40 the sensitivity, specificity, positive
6 predictive value and negative predictive value was 0.84; 0.33; 0.32; 0.85 for Cohort 1
7 and 0.82, 0.43, 0.29, and 0.90 for Cohort 2.

8

9 **Discussion**

10 **Key results**

11 CRP reasonably followed a bimodal distribution using data from two independent
12 cohorts. There was inconclusive evidence of a trimodal distribution; whilst the AIC
13 metric suggested it fit better, on graphical examination there appeared to be
14 overfitting.

15

16 In an analysis of 1835 patients across 13 hospital sites using a binary cut-off for CRP
17 as a prognostic factor of COVID-19 inpatient death appeared to have similar
18 predictive power compared to treating it as a linear or $\text{Ln}(\text{CRP})$. In addition, a cut-off
19 value to indicate disease severity is simpler to use in a clinical setting than a linear
20 predictor. These findings support the use of a simple binary threshold for CRP in
21 daily clinical medicine. These results are well aligned with many published analyses
22 in COVID-19 that have already employed a binary cut-off (4,18–20).

23

24 The bimodal distribution of CRP may reflect the presence of a latent class influence.
25 Candidate variables for this latent class may include confounders that were not fully
26 controlled for: chronic inflammatory conditions, genomic variation of the virus,
27 genetic susceptibility of populations, or other binary exposures such as BCG
28 vaccination status (21–23).

29

30 The association of higher CRP with worse outcomes may be due to the severity of
31 the disease consistent with the “cytokine storm” theory of COVID-19 where the
32 innate immune system is activated releasing TNF-alpha, IL-6 and IL-1. Elshazli et al
33 found CRP to be a valid biomarker of death from COVID-19 when examining a range

1 of haematological and immunological markers. IL-6 was found to be most predictive
2 (OR=13.87) of death, and CRP the next best marker (OR=7.09)(24). However, IL-6
3 is not routinely available to clinicians, but being linked to CRP as a trigger for its
4 transcription makes CRP a better candidate tool for front line hospital usage (25). In
5 the same Elshazli paper a threshold level of 38.2 mg/L was demonstrated to have
6 the best sensitivity and specificity, which fits well with our findings, this was also
7 found within a recent Cochrane Diagnostic Test Accuracy review (26). In addition an
8 elevated CRP may not be attributable to COVID-19 alone, and may represent
9 concomitant pathology such as secondary bacterial pneumonia. Although co-
10 infection is well known in other viral respiratory illnesses, the rate in COVID-19 has
11 been found to be far less, being present in around 5.9% of the general COVID-19
12 hospital population, and 8.1% of those with critical illness (18).

13

14 There data presented here support a single threshold, and whilst there was
15 argument for competing cut-offs of ≥ 40 , ≥ 65 , or greater, the single cut-off is
16 consistent with other studies (8,24). In addition, it would be clearer and safer to offer
17 a conservative approach taking for the lower value of CRP, as a higher threshold
18 may falsely reassure clinicians

19

20 There is a need for simple tests to aid clinical management, as the behaviour of CRP
21 in COVID-19 may provide useful immediate risk stratification as to whom may have a
22 poor outcome. The threshold of CRP ≥ 40 offered a high negative predictive value, so
23 patient presenting with a low CRP are unlikely to exhibit disease progression, and
24 high sensitivity analysis which might lead to opening discussions with patients and
25 their carers about the possible course of the disease. This may assist with early
26 resource planning around the potential for critical care support, and may help guide
27 rapid safe discharge from acute hospitals (5). Although the results within this paper
28 give a population based cut-off, any interpretation and management plan must be
29 made on an individual patient basis, with clinicians using CRP in context of clinical
30 history, examination, investigation, noting the threshold offered a low positive
31 predictive value. Beyond clinical predictive value, this model may be useful for
32 monitoring the outcomes of treatments, for example in a trial of Tocilizumab, CRP
33 monitoring was used as a marker of efficacy (27).

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Strengths and Limitations

This was a large study that included participants admitted to 13 hospital sites. The demographics, case-mix and mortality are similar to other larger studies reported within the UK, increasing the findings generalisability (20). We have also shown good replication between the two UK wide cohorts. However, caution should be given to the threshold reported for CRP as studies identifying optimal cut-offs may be subject to selection bias, and may not be replicable (28). Using a threshold of ≥ 40 offered a high sensitivity and negative predictive value, but low positive predictive value.

A limitation of this study is that due to the urgent nature of research data collection in a pandemic, disease severity on admission was only assessed using CRP without collection of circulating lymphocytes, interleukin-6, procalcitonin, serum lactate, and viral load, all of which may also contribute to disease severity (29).

Interpretation

A simple threshold ≥ 40 mg/L should be used within clinical practice to guide disease severity and likely disease progression. Future studies should analyse using this simple threshold.

Generalisability

The impact of these findings support the routine assessment of serum CRP as an adjunct in the early diagnosis and assessment of illness severity of hospitalised patients with COVID-19. We recommend CRP ≥ 40 mg/L on admission may indicate an increased risk of disease progression and death and warrants an enhanced level of discussion and clinical support.

Conclusions

We have demonstrated that CRP follows a bimodal distribution in hospitalised patients with COVID-19. This requires further exploration to discover the latent class effect of unobserved factors influencing the distribution of CRP. A CRP of ≥ 40 mg/L on admission to hospital should be seen as a reliable indicator of disease severity

1 and increased risk of death. We recommend clinicians to use this cut-off as a
2 prognostic indicator only in conjunction with an individualised clinical assessment,
3 frailty assessment, incorporating a person's wishes and values, to make early
4 decisions about enhanced observation, critical care support, and advanced care
5 planning.

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16 **Declarations**

17 **Contributions of authors**

18 Conceived the study (BC, PB), developed the protocol (BC, DS, PB), collected the data (PB,
19 PM, LE, JC, VA, TQ, AV, MS, LP, JH, SM, KMc), analysed the data (DS, BC), interpreted
20 the findings (DS, BC, PB), drafted the initial manuscript (DS, PB, BC), all authors approved
21 the final manuscript.

22 BC is the guarantor of the study findings

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27 **Data Sharing Agreement**

28 Data is available on request from the corresponding author after submission of a statistical
29 analysis plan, after approval from the COPE Study Investigators.

30 **Competing Interests Statement**

31 No author has a competing interest