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HADES RV Programme with HARPS-N at TNG

IV. Time resolved analysis of the Ca II H&K and H α chromospheric emission of low-activity early-type M dwarfs

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ABSTRACT

Context. M dwarfs are prime targets for current and future planet search programs, particularly those focused on the detection and characterization of rocky planets in the habitable zone. In this context, understanding their magnetic activity is important for two main reasons: it affects our ability to detect small planets and it plays a key role in the characterization of the stellar environment.

Aims. We analyze observations of the Ca II H&K and H α lines as diagnostics of chromospheric activity for low-activity early-type M dwarfs.

Methods. We analyze the time series of spectra of 71 early-type M dwarfs collected in the framework of the HADES project for planet search purposes. The HARPS-N spectra simultaneously provide the Ca II H&K doublet and the H α line. We develop a reduction scheme able to correct the HARPS-N spectra for instrumental and atmospheric effects, and also to provide flux-calibrated spectra in units of flux at the stellar surface. The Ca II H&K and H α fluxes are then compared with each other, and their time variability is analyzed.

Results. We find that the Ca II H and K flux excesses are strongly correlated with each other, while the H α flux excess is generally less correlated with the Ca II H&K doublet. We also find that H α emission does not increase monotonically with the Ca II H&K line flux, showing some absorption before being filled in by chromospheric emission when Ca II H&K activity increases. Analyzing the time variability of the emission fluxes, we derive a tentative estimate of the rotation period (on the order of a few tens of days) for some of the program stars, and the typical lifetime of chromospheric active regions (on the order of a few stellar rotations).

Conclusions. Our results are in good agreement with similar previous studies. In particular, we find evidence that the chromospheres of early-type M dwarfs could be characterized by different filament coverage, affecting the formation mechanism of the H α line. We also show that chromospheric structure is likely related to spectral type.

Key words. techniques: spectroscopic – stars: late-type – stars: low-mass – stars: activity

1. Introduction

The term “stellar activity” indicates a class of phenomena that, in stars with an external convective envelope, are triggered by the reconfiguration of the surface magnetic field. For the Sun, these phenomena are commonly classified as sunspots, plagues, flares, or coronal holes, for example, and they take place from the photosphere to the corona. The portions of the stellar surface

and atmosphere where these phenomena take place are referred to as active regions (ARs).

From the observational point of view, traditionally the most frequently investigated diagnostic of chromospheric activity for solar-type stars is the Ca II H&K doublet (3933.67 Å and 3968.47 Å; see, e.g., the milestone project of the Mount Wilson Observatory as summarized in Baliunas et al. 1998). The main advantage of this doublet is that the photospheric line profiles

have large depths with respect to the continuum, such that the chromospheric emission can be detected with high contrast against the photospheric background. Nonetheless, a particularly relevant drawback especially for M-type stars is that detailed spectroscopy is observationally more demanding because of the low continuum level at wavelengths shortward of 4000 Å. For this reason, these observations have been focusing on stars ranging from F to K spectral types.

This led to the identification of another diagnostic more suited for red dwarfs: the H α line (6562.80 Å; see, e.g., Pasquini & Pallavicini 1991; Montes et al. 1995). It is generally accepted that the Ca II H&K and H α excess fluxes are tightly correlated even though they form in different layers of the chromosphere. However, this conclusion has been drawn historically by analyzing averaged measurements of chromospheric fluxes of calcium and hydrogen taken at different epochs and rarely obtained simultaneously (Giampapa et al. 1989; Robinson et al. 1990; Strassmeier et al. 1990). Another approach is to collect simultaneous measurements for a number of stars at a given epoch, without any follow-up investigation of the observed targets at different epochs (Thatcher & Robinson 1993; Walkowicz & Hawley 2009).

Recently, planet search programs carried out with the radial velocity techniques have started to monitor samples of M dwarfs. These stars are extremely interesting targets for planet discoveries. First of all, they represent $\sim 75\%$ of the stars in the solar neighborhood (Reid et al. 2002; Henry et al. 2006). Moreover, from an observational point of view, the chances of finding an Earth-like planet in the habitable zone of a star increase as the stellar mass decreases. Still, habitability is not guaranteed simply by an assessment of the distance from the star; several other factors, such as stellar activity, may move and/or shrink the habitability zone of a star (see Vidotto et al. 2013, and references therein). Thus, it is crucial to better understand the activity of M dwarfs and how it can affect the circumstellar environment.

Intensive monitoring aimed at an accurate measurement of radial velocities has provided the community with large databases of high-resolution and high signal-to-noise ratio (S/N) optical spectra per star. These databases are valuable for the analysis of the activity of M dwarfs. For instance, Gomes da Silva et al. (2011) have analyzed the long-term variability of a number of chromospheric optical indexes for a sample of 30 M0–M5.5 stars from the HARPS at ESO M-dwarf planet search program (Bonfils et al. 2013) with a median timespan of observations of 5.2 yr.

The HARPS-N red Dwarf Exoplanet Survey (HADES, Affer et al. 2016; Perger et al. 2017) project is a collaborative program between the Global Architecture of Planetary Systems project¹ (GAPS, Covino et al. 2013), the Institut de Ciències de l’Espai (IEEC/CSIC), and the Instituto de Astrofísica de Canarias (IAC), that monitors the radial velocities of a sample of low-activity M-type dwarfs to search for planets. In this framework, we are analyzing the collected spectra to study the activity of the monitored stars from different points of view. In Maldonado et al. (2017) we analyze the Ca II H&K and Balmer line flux excesses in relation to rotation and age, while in Suárez Mascareño et al. (in prep.) we analyze the time series of spectra in order to measure the rotation period of the stars and the periodicity of their activity cycles, if any is present.

In this paper our aim is to characterize the daily to monthly variability due to chromospheric ARs based on the analysis of

simultaneous measurements of the Ca II H&K and H α lines. In Sect. 2 we present the target selection and the database of spectra collected so far; in Sect. 3 we describe the data reduction and the measurement of the chromospheric Ca II H&K and H α flux excesses; in Sect. 4 we analyze the flux-flux relationships and in Sect. 5 we discuss the amount and timescales of chromospheric variability in the selected sample of stars.

2. Target selection

The HADES sample of stars is made up of 78 red dwarfs, 71 of which have been observed so far. The observed stars have spectral types ranging between K7.5V and M3V (corresponding to the ~ 3400 – 3900 K temperature range). These stars are being monitored with the HARPS-N spectrograph (Cosentino et al. 2012) mounted at the Telescopio Nazionale Galileo in the framework of the HADES project. The spectrograph covers the 383–693 nm wavelength range with a spectral resolution of $\sim 115\,000$. HARPS-N spectra were reduced using the most recent version of the Data Reduction Software (DRS) pipeline (Lovis & Pepe 2007).

The monitored stars have been selected as targets favorable for planet search, thus the sample is biased towards low activity levels with some exceptions. The program stars are listed in Table A.1, together with the number of observations for each of them. The data analyzed in this paper have been collected over seven semesters, from September 2012 to February 2016. In Maldonado et al. (2017) we measured the stellar parameters using the methodology described in Maldonado et al. (2015). The most relevant stellar parameters for the current analysis are listed in Table A.1. We refer to Affer et al. (2016) for further details of the target selection and to Maldonado et al. (2017) for a more detailed description of the determination of the stellar parameters.

3. Post-processing analysis

During the spectral analysis, we find evidence of variability in the instrumental response, in the atmospheric reddening, and in the intensity of telluric lines, either in absorption or emission, from night to night. These effects are usually corrected by observing a spectroscopic standard star and a telluric standard star close to the observed target. The aim of the HADES and GAPS projects is to find new exoplanets around pre-selected targets or to refine the characterization of known exoplanetary systems. Thus, the instrumental setup and the observational campaign are arranged to maximize the number of targeted stars and the precision of the radial velocity measurements, while the flux calibration of the spectra is not taken into account (e.g., spectroscopic or telluric standard stars are not usually observed).

For our purposes, however, the correction of all these effects is necessary in order to compare the Ca II H&K and H α lines taken on different nights. In this section we describe in detail our post-processing analysis of the spectra reduced with the DRS pipeline.

3.1. Flux rescaling

To correct the observed fluxes for instrumental response and atmospheric reddening, for each star in our catalog we interpolate the synthetic spectral library provided by Allard et al. (2011)² to compute a model spectrum corresponding to the measured T_{eff} ,

¹ http://www.oact.inaf.it/exoit/EXO-IT/Projects/Entries/2011/12/27_GAPS.html

² We adopt the CIFIST2011 models (<https://phoenix.ens-lyon.fr/Grids/BT-Sett1/CIFIST2011bc/SPECTRA/>).

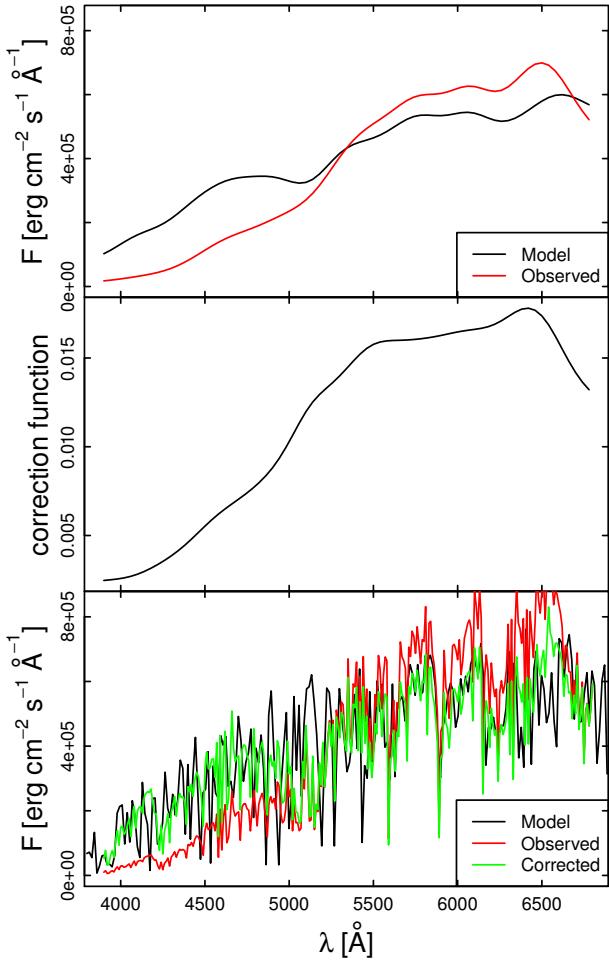


Fig. 1. Example of the correction of the observed spectra discussed in the text. *Top panel:* comparison between the BT-Settl model spectrum (in black) and the observed spectrum (in red), both degraded to low resolution. For plotting reasons, the observed spectrum, still in instrumental units, is rescaled to approximately match the flux scale of the model spectrum. *Middle panel:* observed-to-model ratio correction. The curve describes the rescaling of the observed spectrum into absolute units. *Bottom panel:* result of the rescaling: the high-resolution spectra corresponding to those shown in the top panel are plotted in black and red again. The green line shows the corrected high-resolution spectrum.

$\log(g)$, and $[\text{Fe}/\text{H}]$. We arbitrarily degrade the spectral resolution of both the model spectrum and the series of observed spectra convolving them with a Gaussian kernel with $\sigma = 120 \text{ \AA}$, corresponding to a final spectral resolution of ~ 50 over the spectral range in consideration (Fig. 1, top panel). The ratio of each low-resolution spectrum and the model is thus a continuous function of wavelength (Fig. 1, middle panel), which allows us to rescale the high-resolution instrumental fluxes to fluxes at stellar surface and, at the same time, to correct for the differential instrumental response and atmospheric reddening (Fig. 1, bottom panel).

The degradation to such a low spectral resolution was motivated by several factors:

- it allows us to safely mask out the narrow spectral windows centered on the chromospheric emission lines (Table 1) that may introduce spurious features in the correction function;
- it removes the line component of the telluric spectrum, leaving us only with the continuous component of the atmospheric absorption;

Table 1. Chromospheric emission lines analyzed in this work.

Ion	Line	Central wavelength (\AA)	Spectral width (\AA)
Ca II	K	3933.67	1.5
Ca II	H	3968.47	1.5
H I	H α	6562.80	4

- at low resolution, the discrepancies between observed and model line profiles are negligible, the low-resolution spectra being dominated by the continuum flux.

3.2. Telluric line removal

The other effect to take into account is the telluric contamination. We focus only on the 130 \AA -wide spectral range around the H α line because the Ca II H&K spectral region is free from telluric lines.

In order to have a comparison spectrum suited for the removal of telluric lines, for each star we compute the median of all the flux-rescaled high-resolution spectra. This median spectrum has a higher S/N than the single spectra, and is also free from telluric lines. Actually, telluric lines are Doppler shifted in the reference frame of the targeted star because the observations are generally performed at different hour angles and different months along the observing seasons. Since their wavelengths change in the stellar rest frame from night to night, the median algorithm filters them out as outliers.

For each star we compute the ratio between each spectrum and the corresponding median value, obtaining the normalized telluric spectrum affecting each given observation. Then we cross-correlate it with the model telluric spectrum provided in Borsa et al. (2015). While the cross-correlation allows us to align this model telluric spectrum to the ratio of the spectra, the intensity of the telluric contamination is computed by a $3\text{-}\sigma$ clipped comparison between the wavelength-shifted model telluric spectrum and the ratio. These two parameters (wavelength shift and intensity) allow us to correct the telluric lines within the noise of the observed spectra.

In the following sections, we always mean the flux-rescaled telluric-corrected spectra when referring to the observed spectra.

3.3. Selection of the reference spectra

For each star, we have a series of individual spectra and the corresponding median spectrum (Sects. 3.1 and 3.2). To make the computation faster, we extract a 70 \AA spectral interval around the H α and a 100 \AA interval around the Ca II H&K doublet comprising both lines (Table 1).

In each interval, we want to subtract the photospheric flux to measure the chromospheric flux. To this end, our first approach was to use synthetic spectra, which soon turned out to loosely fit the absorption line profiles. We then decided to use our spectral database to obtain the best “inactive” template for each star in the program.

It is reasonable to assume that those stars with minimum Ca II H&K emission are the “least active” and therefore the ones that should be used as templates for the spectral subtraction. For each star we measure the Ca II H&K flux $F_{\text{HK}}^{\text{int}}$ integrating the median spectrum in the spectral intervals around the Ca II H and K lines defined in Table 1 (see the end of Sect. 3.4 for the definition of the spectral intervals). These integrated fluxes thus contain the contribution of both the photospheric flux and chromospheric

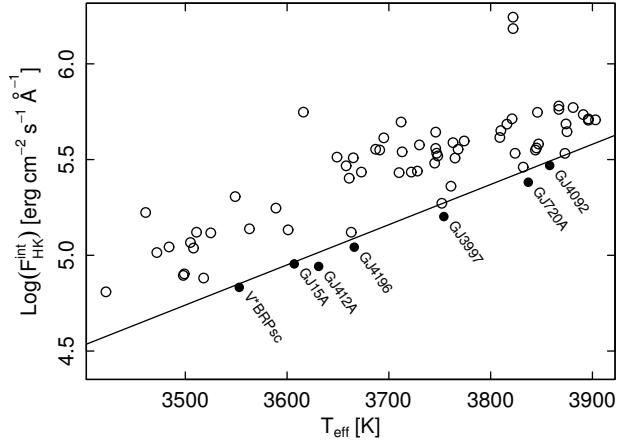


Fig. 2. $F_{\text{HK}}^{\text{int}}$ vs. T_{eff} diagram. The solid line (Eq. (1)) represent our selection threshold of reference stars. The selected stars are indicated as plain dots and their names are shown.

basal flux, plus the flux from chromospheric ARs, if present. In Fig. 2 we plot the integrated Ca II H&K flux versus the stellar effective temperature; we find that the lower boundary of the diagram is an increasing function of T_{eff} . This trend is due to the increase in the Ca II H&K photospheric flux of early-type M dwarfs with T_{eff} , as predicted by the BT-Settl models.

We select the template stars at the lower boundary (i.e., at minimum Ca II H&K flux) of the diagram in Fig. 2, arbitrarily considering the stars below the line with equation

$$\text{Log}(F_{\text{HK}}^{\text{int}}) = -2.61 + 0.0021 \cdot T_{\text{eff}}. \quad (1)$$

The selected reference stars are V* BR Psc, GJ 15A, GJ 412A, GJ 4196, GJ 3997, GJ720A, and GJ 4092. All these stars have subsolar metallicities, with the exception of GJ 4196. The selection of these stars could reflect a trend between Ca II flux and metallicity that is due to the dependence of the Ca II line core strength on the calcium elemental abundance. This correlation could therefore be linked to the line core strength curve of growth in the atmosphere discussed by Houdebine (2011).

As we discuss below, stellar metallicity cannot be neglected in the spectral subtraction. For this reason, we reject GJ 4196 as a reference star, and we keep the remaining stars as a spectral grid for the computation of the reference quiet spectrum.

3.4. Spectral subtraction and line flux measurement

For each program star, the reference spectrum is computed by interpolating the grid of reference spectra over T_{eff} . Before the interpolation takes place, we broaden the spectra such that they all have the same rotational broadening. For this purpose, we select the maximum $v \sin i$ among those of the program and reference stars as measured by Maldonado et al. (2017), and we broaden each spectrum to this maximum $v \sin i$ using the rotational profile provided by Gray (1992).

Since the reference spectrum is computed using the spectra of stars with minimum metallicities in our sample, we find a systematic offset (see Fig. 3) correlated with the metallicity of the star. In particular, this occurs because the continuum flux predicted by the BT-Settl models decreases with increasing [Fe/H], consistently with the increase in the absorption by the multitude of atomic and molecular lines in M star spectra. In Fig. 3 we show the example of the H α spectrum of GJ 9793 (see also Fig. 4), the highest metallicity star in our sample: its spectrum

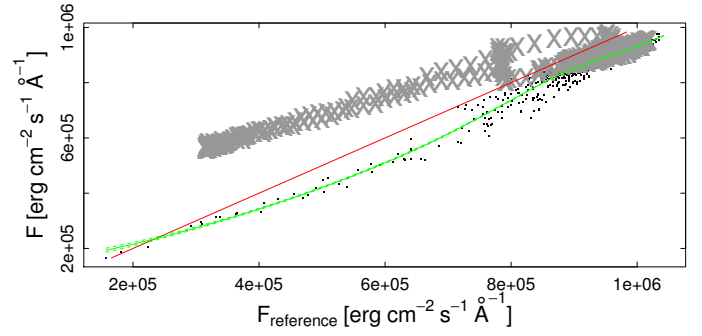


Fig. 3. Comparison between fluxes in the H α region of the spectrum of GJ 9793 and the corresponding interpolated reference spectrum (see text). Gray crosses represent fluxes in the H α core, and are not taken into consideration in the flux rescale. The red solid line is the 1:1 line. The green solid line is the fitted function used for the flux rescaling.

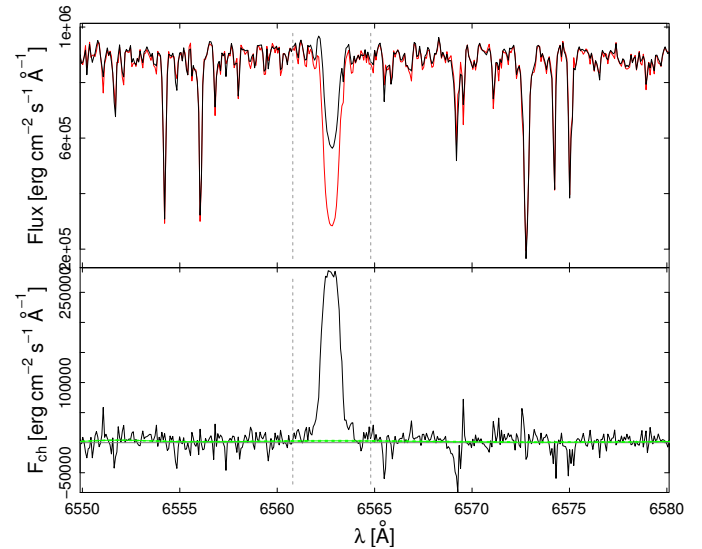


Fig. 4. Example of the extraction of the H α line flux. *Top panel:* zoom-in of the spectral comparison between the observed spectrum (in black) and the rescaled template (in red). The vertical gray dashed lines bracket the line core. *Bottom panel:* chromospheric spectrum computed as the difference between the observed and template spectrum shown in the top panel. The green line is the polynomial fit for the removal of low-order trends in the differential spectrum.

shows lower fluxes than those of the corresponding reference spectrum. To correct this offset, we fit a low-order polynomial curve to the flux points. In the fit, we exclude the core of the line (gray crosses in the figure) because it may be affected by chromospheric emission. Finally, the reference spectrum is rescaled to the observed spectrum using the best-fit curve.

Once the reference spectrum is rescaled to the observed spectrum, the chromospheric spectrum is computed by subtraction of the template (see also Herbig 1985; Frasca & Catalano 1994, for previous application of the spectral subtraction technique). Outside of the line cores we sometimes find a residual wavelength-dependent offset significantly different from zero. This residual offset arises from several factors, such as inaccuracies in the measured stellar parameters, the interpolation of the reference spectrum, and spectral peculiarities of the analyzed spectrum. To remove these systematics, we subtract a low-order polynomial function from the chromospheric spectrum, fitted over the spectral range excluding the line cores. We note that the 1σ confidence band on the polynomial fit is narrow enough to introduce

a negligible uncertainty on the line flux compared to the noise (Fig. 4, bottom panel).

Finally, the chromospheric emission is computed by integrating the chromospheric spectrum, i.e., the template-subtracted systematics-corrected spectrum, over the wavelength ranges listed in Table 1. The widths of the spectral windows used to compute the emission flux are conveniently set after visual inspection of the differential spectra, representing a compromise between the bracketing of the whole flux radiated by the chromosphere and the exclusion of the nearby spectral regions that introduce noise in the flux measurement. To compute the uncertainty on the integrated flux, we measure the S/N of the differential spectrum outside of the line cores, and we propagate it in the integration of the line flux. This error thus includes both the noise in the observed spectra and the uncertainties introduced by our data reduction.

We caution that the spectral subtraction removes the flux of both the photosphere and the “quiet” chromosphere from the observed spectra. For the Ca II H&K doublet, this means that the measured emission fluxes do not include the basal chromospheric emission, i.e., they are representative of the excess chromospheric flux radiated by ARs. The case of the H α line is more subtle. Following the chromospheric models of Cram & Giampapa (1987), the absorption in the H α line seen in the spectra of early-type M dwarfs originates in the chromosphere. In particular, in these models even the unperturbed chromospheres are hot enough to populate the $N = 2$ levels of the hydrogen ions, making the absorption of the H α line possible. Thus, in the least active M stars in our sample we see a “basal” H α absorption depth, and the line fluxes measured with the spectral subtraction technique have this basal absorption as zero point.

4. Flux-flux relationships

In the following we analyze the relationships between the chromospheric fluxes in the Ca II H&K and H α lines resolved in time. To improve the statistical reliability of our analysis, we arbitrarily discard all the stars with fewer than 20 observations, which leaves us with 41 stars. Moreover, to clean up our samples of measurements, for each star and each line we reject the outliers with a 5σ -clipping criterion. We apply the same selection to the measurement uncertainties, thus discarding the data with anomalously large error bars.

4.1. Ca II H vs. Ca II K

In Fig. 5 we compare the individual measurements of the Ca II H and K excess fluxes (F_{CaIIH} and F_{CaIIK} , respectively) of the stars in the analyzed sample, omitting the measurements less than 3σ above zero because they are not representative of flux in excess. We find that the fluxes are well aligned despite differences in stellar parameters and activity levels. Since the uncertainties on the quantities are comparable, we compute the best-fit line by means of the ranged major axis (RMA) regression (Legendre & Legendre 1983). The result of the fit is

$$\log F_{\text{CaIIK}} = (0.11 \pm 0.04) + (0.990 \pm 0.009) \log F_{\text{CaIIH}}, \quad (2)$$

which is consistent within 3σ with similar previous studies performed on samples of M dwarfs (Martínez-Arnáiz et al. 2011; Stelzer et al. 2012). From Eq. (2) we derive that the K/H flux ratio is close to 1, and following Houdebine & Stempels (1997) we argue that the emission of the Ca II H&K lines is optically thick.

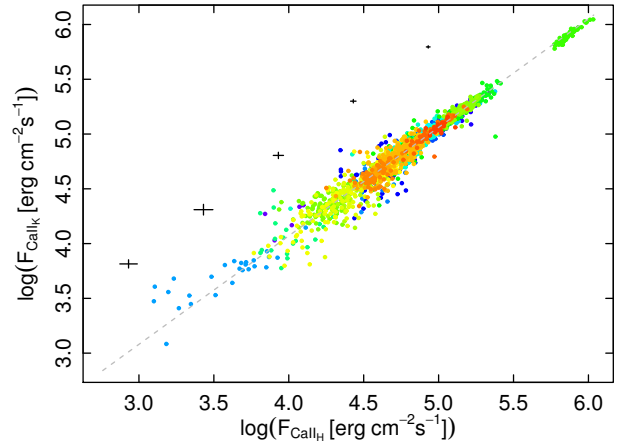


Fig. 5. $\log(F_{\text{CaIIH}})$ vs. $\log(F_{\text{CaIIK}})$. Different colors represent different stars. The gray dashes show the RMA best fit in Eq. (2). Black crosses represent the typical measurement uncertainties at different activity levels.

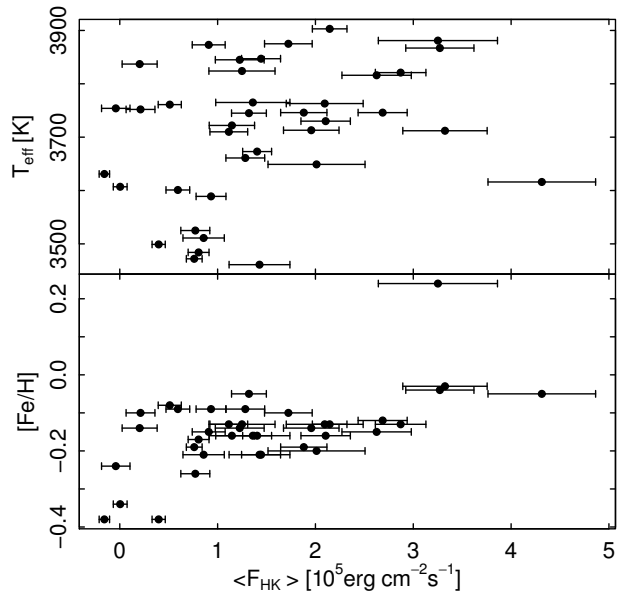


Fig. 6. T_{eff} and $[\text{Fe}/\text{H}]$ versus the median total flux excess $\langle F_{\text{HK}} \rangle$. Error bars show the median absolute deviation of the sample of F_{HK} measurements for each star.

We analyze the Ca II H&K flux excess as a function of temperature and metallicity. For this purpose, in Fig. 6 we plot the median total flux excess $\langle F_{\text{HK}} \rangle = \langle F_{\text{H}} + F_{\text{K}} \rangle$ of the stars, together with the corresponding median absolute deviation³ (MAD) of the sample of measurement, versus T_{eff} and $[\text{Fe}/\text{H}]$.

We find that the flux excess drops for stars cooler than $T_{\text{eff}} = 3600$ K (spectral type M1). Similar results have been found by West et al. (2004), who analyzed the H α emission, and by West et al. (2011), who included the Ca II K line, but for stars later than M4. We cannot claim any stronger conclusion because we are investigating a narrow spectral range. Moreover, we note that our subsample of 41 stars is biased towards low activity levels, thus our findings are strongly affected by selection effects.

³ The mean absolute deviation (MAD) is a measure of the scatter of a sample of data. It is defined as the median of the absolute residuals with respect to the sample median, and is thus robust against the presence of outliers.

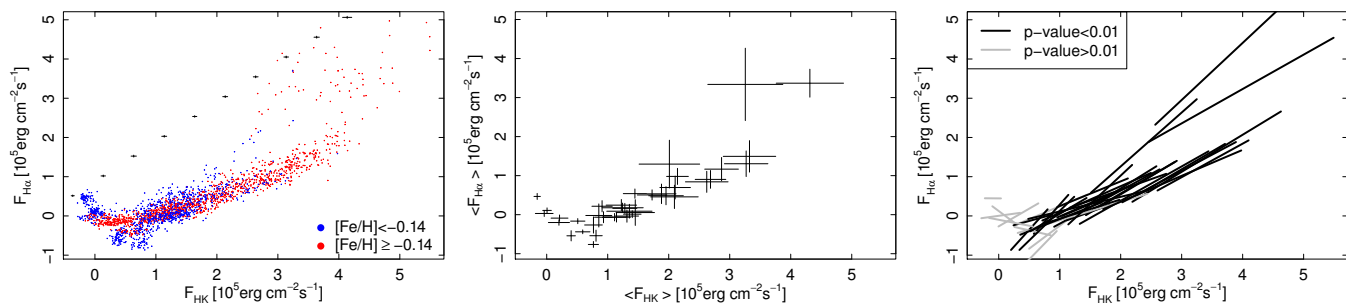


Fig. 7. *Left panel:* close-up view of the F_{HK} vs. $F_{\text{H}\alpha}$ diagram. Colors code the metallicity of the stars as shown in the legend. Black crosses represent the typical measurement uncertainties at different activity levels. *Middle panel:* same as in the left panel for the median of the flux measurements. For each star, the error bars represent the MAD of the measurements shown in the left panel. *Right panel:* the same as in the left panel, representing each star with the corresponding linear fit to its flux-flux values. Gray lines mark best fits with low statistical significance, while black lines represent p -values lower than 1% based on Spearman's correlation test.

The dependence on the stellar metallicity is slightly more convincing: more metallic stars tending towards larger $\langle F_{\text{HK}} \rangle$ excess. One possible explanation is that the Ca II H&K emission increases with the abundance of Ca II ions in the chromosphere, thus with metallicity at first approximation, as also discussed by Houdebine (2011). Another possibility is that this could be an age effect, as higher metallicities generally correspond to younger ages. This is consistent with the fact that younger stars tend to be more active, as we discuss in Sect. 5.1.

4.2. Ca II H&K vs. $H\alpha$

In the left panel of Fig. 7 we show the emission flux $F_{\text{H}\alpha}$ radiated in the $H\alpha$ line plotted against the flux F_{HK} radiated by the Ca II H&K doublet. The comparison between $H\alpha$ and Ca II H&K fluxes is more scattered than that of Ca II H and K lines. To avoid cluttering, in the middle panel we show the same diagram for the median of the flux measurements $\langle F_{\text{HK}} \rangle$ and $\langle F_{\text{H}\alpha} \rangle$, where the plotted uncertainties represent the MAD of the individual flux measurements.

One possible explanation of the scatter in Fig. 7 is that, because they have a larger number of Ca II ions in the atmosphere, stars with higher metallicities have more efficient Ca II H&K radiative cooling (i.e., higher fluxes in emission), as found by Houdebine (2011). To test this hypothesis, we split the sample of stars into two halves at $[\text{Fe}/\text{H}] = -0.14$, and plot them with different colors in the left panel of Fig. 7; we find that the two subsamples follow the same locus. In particular, we do not find that the F_{HK} fluxes tend to increase with metallicity at fixed $F_{\text{H}\alpha}$. We thus conclude that we do not have enough statistical evidence to state that metallicity affects the $H\alpha$ vs. Ca II H&K excess relationship, maybe because the metallicity range spanned by our program stars is not large enough (see Table A.1).

We perform the same analysis splitting the sample of stars in two halves on the basis of their T_{eff} (Fig. 8). Also in this case we do not find any evidence of correlation, maybe because we are investigating a narrow range in T_{eff} .

Another interesting feature, visible in the middle panel of Fig. 7, is that the flux-flux relationship is not monotonic. A visual analysis of the plot shows that the flux radiated in the $H\alpha$ line seems to initially decrease with increasing Ca II H&K excess flux, then goes below zero at $\langle F_{\text{HK}} \rangle \simeq 10^{-5} \text{ erg cm}^{-2} \text{ s}^{-1}$ (i.e., the $H\alpha$ line deepens), and finally increases with the Ca II H&K flux. Similar results were obtained by Robinson et al. (1990). The absorption of the $H\alpha$ line towards negative $\langle F_{\text{H}\alpha} \rangle$ values is the

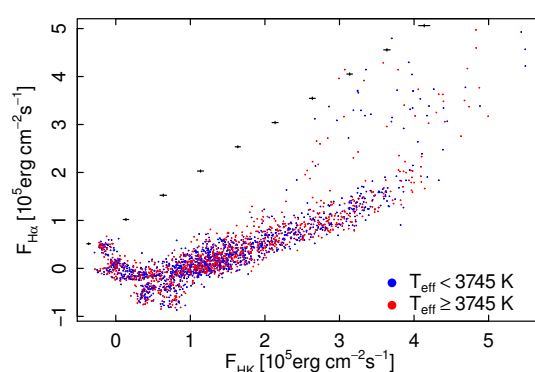


Fig. 8. Same as in the left panel of Fig. 7, but splitting the analyzed sample into two halves at $T_{\text{eff}} = 3745 \text{ K}$.

reason why we plot the Ca II H&K vs. $H\alpha$ relationship in linear rather than logarithmic scale.

The non-monotonic trend is supported by statistical tests. Kendall's correlation test⁴ significantly supports the idea that the two datasets are correlated. Moreover, we naively fit a linear and a quadratic function to the data, and we find that the corrected Akaike Information Criterion⁵ significantly rejects the former compared to the latter.

From the theoretical point of view, this result has been predicted by chromospheric models (Cram & Mullan 1979; Cram & Giampapa 1987; Rutten et al. 1989; Houdebine et al. 1995; Houdebine & Stempels 1997). According to these models, the most feasible scenario is that the chromospheric Ca II H&K emission lines are collisionally dominated, thus the radiated flux steadily increases with pressure. Conversely, the $H\alpha$ line is radiation-dominated and the increase of the optical depth initially leads to a deeper absorption profile, until the electron density is high enough to take the $H\alpha$ line into the collisionally dominated formation regime, leading to the fill-in of the line.

Previous observational studies performed on similar samples of stars (K- and M-type main sequence stars, e.g., Stauffer & Hartmann 1986; Giampapa et al. 1989; Rauscher & Marcy 2006; Walkowicz & Hawley 2009) have

⁴ Kendall's rank coefficient is often used as a non-parametric test statistic to establish whether two variables are statistically dependent (Kendall 1938).

⁵ Given a collection of models for the data, the Akaike Information Criterion (Burnham & Anderson 2002) estimates the quality of the models relative to each other. Hence, it provides a means for model selection.

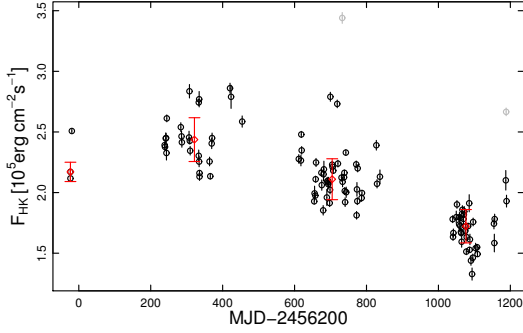


Fig. 9. Time series of the H&K emission flux of GJ 4306. For each season, red circles and red error bars represent the median and the MAD of the corresponding sample of measurements. Gray symbols show the rejected outliers.

found comparable results. In particular, while the deepening of the $H\alpha$ line is still somehow controversial, it is clear that the $H\alpha$ is not filled in until a certain Ca II H&K emission flux is reached. Thus, the $H\alpha$ line is not a good activity indicator for low to intermediate activity levels. Moreover, Houdebine & Stempels (1997) noted that observational data do not strictly match the models likely owing to surface inhomogeneities, which may also explain the scatter in Fig. 7.

A careful analysis of the left panel of Fig. 7 also shows that the stars follow individual loci with different slopes in the diagram. To better show this effect, in the right panel of Fig. 7 we plot the locus occupied by each single star defined as the RMA best fit (see Sect. 4.1) of its data points.

First of all, we check whether the scatter of each star in the diagram is due to an ongoing long-term activity cycle or to variability on shorter timescales. In this regard, for each star we visually inspect the Ca II H&K and $H\alpha$ time series, and in some cases we find hints of increasing or decreasing activity over year-long timescales or even turnarounds (see Fig. 9 for an example). Robertson et al. (2013) found similar evidence analyzing the ~ 11 -yr series of $H\alpha$ observations of a sample of 93 M-type dwarfs, and concluded that $\sim 10\%$ of M dwarfs show either long-term trends or activity cycles that are longer than one year.

To remove the above-mentioned trends, for each season we reject outliers (generally positive outliers likely due to flares) and we compute the median of the remaining measurements. Then, for each season we subtract the corresponding median flux, thus filtering out any variability longer than one year. We find that long-term variability is generally able to explain up to 40% of the scatter in the emission line fluxes. We thus conclude that $\geq 60\%$ of the variance in the data is due to daily to weekly to monthly variability.

In support of this conclusion, we find that slopes, correlation coefficients and significances are generally preserved by the correction for long-term trends, which confirms that the displacement of the stars in Fig. 7 is mainly due to correlated shorter term variability. In the following, we always refer to the long-term corrected flux excess.

In Fig. 10 we plot the distribution function of the slopes of the residual F_{HK} vs. $F_{H\alpha}$ relationship for those stars that have slopes with high statistical significance (p -value $< 1\%$). This distribution shows that the slope is always positive, i.e., $F_{H\alpha}$ is always an increasing function of F_{HK} .

Moreover, the distribution function is not symmetric: it has a broad right tail. The Anderson-Darling normality test (Anderson & Darling 1952) supports the non-normality of the

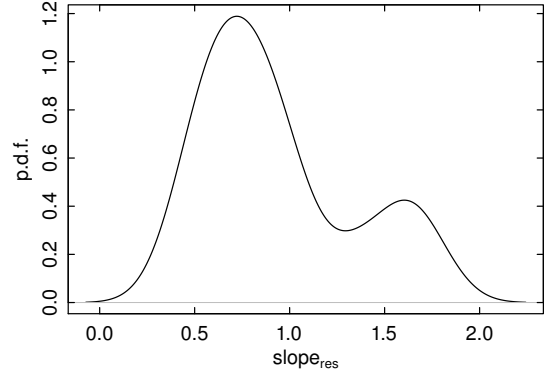


Fig. 10. Distribution function of the slopes of the F_{HK} vs. $F_{H\alpha}$ relationship with p -value $< 1\%$.

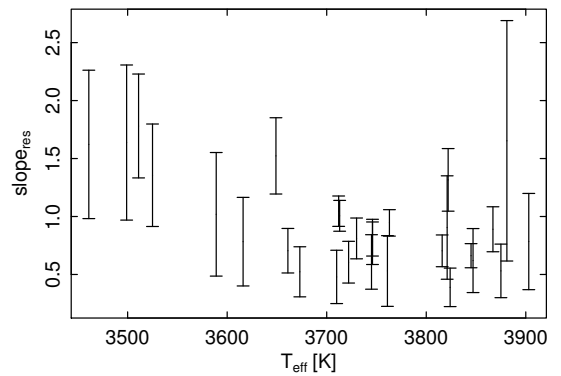


Fig. 11. Slope of the F_{CaIIHK} vs. $F_{H\alpha}$ relationship as a function of T_{eff} . Only the stars with significant correlation (p -value $< 1\%$) have been taken into account.

sample distribution at the level of $\sim 4 \times 10^{-3}$. We thus infer that the distribution of the data is not dominated by measurement errors (which would lead to a Gaussian distribution function), but reflects an intrinsic characteristic of the analyzed sample.

We also test the dependence of the slopes of the F_{HK} vs. $F_{H\alpha}$ relationship on $\langle F_{\text{HK}} \rangle$, T_{eff} , and $[\text{Fe}/\text{H}]$, respectively, and we find evidence of decreasing steepness of the flux-flux relationship with increasing T_{eff} with a confidence level of $\sim 4\%$ (Fig. 11). Stelzer et al. (2013) found a similar trend by analyzing a sample of pre-MS low-mass stars in the 2500–4500 K temperature range. Conversely, Kendall's correlation test cannot reject the hypothesis that the slope of the flux-flux relationship is independent of the average level of activity or the metallicity of the star. On the other hand, we remark that we cannot claim the independence from the aforementioned parameters as a lack of statistical dependence may be a consequence of the narrow ranges of the investigated stellar parameters.

We compare these results with the model provided by Meunier & Delfosse (2009), who explain the F_{HK} vs. $F_{H\alpha}$ relationship for the Sun. In their work, the authors use spatially resolved observations of the solar surface covering 1.5 activity cycles. They find that the presence of dark filaments affects the correlation between F_{HK} and $F_{H\alpha}$ on short timescales (daily or weekly), sometimes leading to negative correlations depending on the surface coverage of plages and filaments. Analogous phenomena may affect the distribution in Fig. 10 leading to a non-Gaussian shape, and may also explain why we obtain negative slopes in some cases, even though with low statistical significance (see Fig. 7, gray lines in the right panel). In particular, in

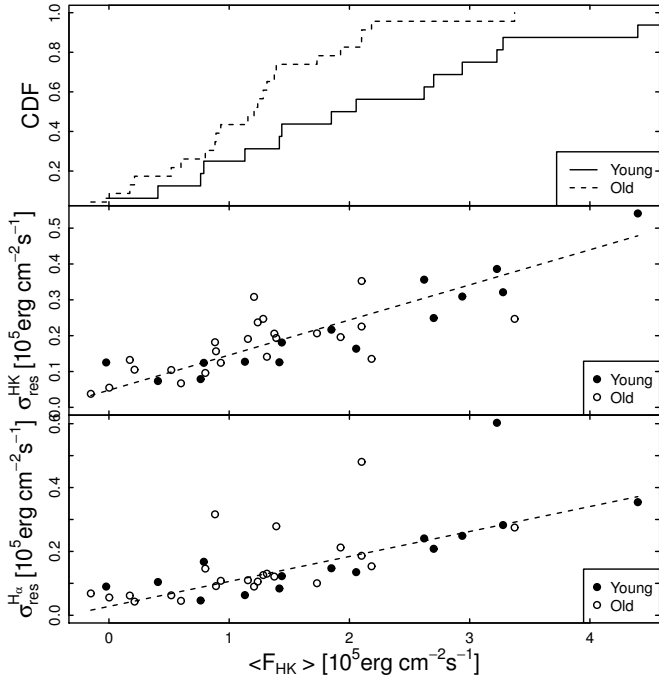


Fig. 12. *Top panel:* empirical cumulative distribution function of the $\langle F_{\text{HK}} \rangle$ measurement for young (solid line) and old (dashed line) stars, following the classification of Maldonado et al. (2017) based on kinematics. *Central panel:* standard deviation $\sigma_{\text{res}}^{\text{HK}}$ of the H&K measurements corrected for the quadratic trend (Sect. 4.2) as a function of the average level of activity $\langle F_{\text{HK}} \rangle$. The black dashes represent the best-fit line; the best-fit coefficients are given in Table 2. Filled and open circles mark young and old stars, respectively. *Bottom panel:* same as in the middle panel, but for the residual H α line.

their model, larger values for the slopes correspond to a scenario in which the chromospheric H α emission is less affected by the absorption of cold filaments than the rest of the sample is. For a fixed configuration of plages seen in the Ca II H&K lines on the Sun, the slope of the flux-flux relationship depends on the contrast of the H α emission in the plages and the absorption by filaments, and their respective filling factors. For a fixed contrast of H α plages on the solar disk, the slope of the F_{HK} vs. $F_{\text{H}\alpha}$ relationship tends to decrease as the contrast of the filaments increases. This is explained by the fact that solar filaments are optically thicker in the H α than in the Ca II H&K lines. Thus, the increase in slope with decreasing T_{eff} (Fig. 11) may indicate that the H α absorption in equivalent active regions decreases towards later M types. Analyzing the Balmer decrements of the same sample of stars we independently draw the same conclusion (Maldonado et al. 2017), consistently with similar previous studies (Bochanski et al. 2007; Stelzer et al. 2013).

5. Chromospheric variability

5.1. Variability vs. activity vs. age

In Fig. 12 we plot the standard deviation of the residual Ca II H&K and H α flux measurements (see Sect. 4.2) of each program star ($\sigma_{\text{res}}^{\text{HK}}$ and $\sigma_{\text{res}}^{\text{H}\alpha}$ respectively) as a function of the median level of Ca II H&K flux, assumed to monotonically increase with the level of magnetic activity (see Sect. 4). In these diagrams we find a tight correlation (supported by the Spearman's correlation test) between variability and median excess flux. The linear best fits suggest that the fractional variability with respect to the average Ca II H&K flux excess (which is represented by the slope of the

Table 2. Best-fit models of the form $y = q + m \cdot x$ shown in Fig. 12.

y [$\text{erg cm}^{-2} \text{s}^{-1}$]	x [$\text{erg cm}^{-2} \text{s}^{-1}$]	$q \pm \sigma_q$ [$\text{erg cm}^{-2} \text{s}^{-1}$]	$m \pm \sigma_m$
$\sigma_{\text{res}}^{\text{HK}}$	$\langle F_{\text{HK}} \rangle$	0.05 ± 0.03	0.10 ± 0.02
$\sigma_{\text{res}}^{\text{H}\alpha}$	$\langle F_{\text{HK}} \rangle$	0.03 ± 0.03	0.08 ± 0.02

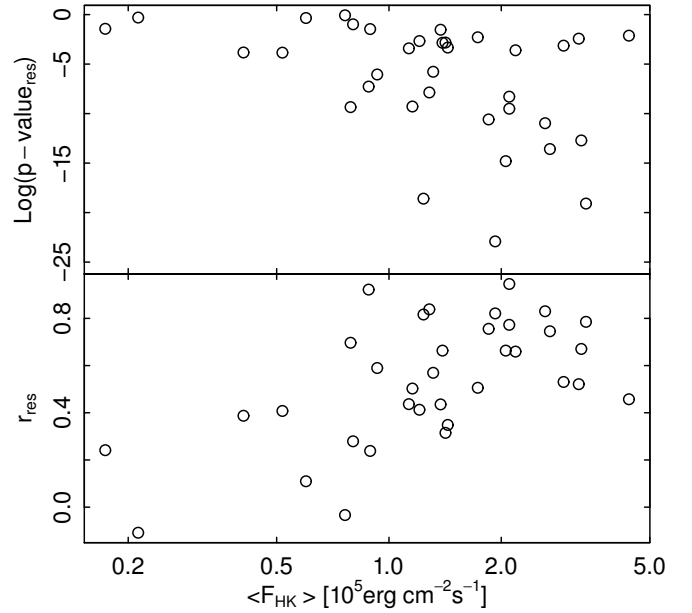


Fig. 13. p -values (top panel) and Pearson's correlation coefficient (bottom panel) of the best fits between H α and Ca II H&K fluxes shown in Fig. 7 vs. the average level of activity $\langle F_{\text{HK}} \rangle$.

linear fit) is roughly 10% and that even at the minimum level of activity there is a residual variability of $\sim 0.5 \times 10^{-5} \text{ erg cm}^{-2} \text{ s}^{-1}$, which is different from zero at the $\sim 2\sigma$ level for the Ca II H&K lines (see Table 2). If this is a real effect rather than a bias introduced by our analysis, it is evidence that even at minimum activity there is some variability.

We also separate the stars younger and older than ~ 650 Myr as classified in Maldonado et al. (2017) based on their kinematics (Table A.1). Although kinematics provide only lower limits to stellar ages (Maldonado et al. 2010, and references therein), with this classification we find evidence that the two populations of stars show significantly different distributions in terms of Ca II H&K emission flux. The Kolmogorov-Smirnov test on the $\langle F_{\text{HK}} \rangle$ we measure confirms this result (top panel in Fig. 12), indicating that younger stars have higher Ca II H&K fluxes. The same test on $\langle F_{\text{H}\alpha} \rangle$ is less conclusive probably because the H α is seen in absorption for a range of $\langle F_{\text{HK}} \rangle$, i.e., the H α is not a good diagnostic at low and intermediate activity levels (see Sect. 4.2).

The Kolmogorov-Smirnov test on the $\sigma_{\text{res}}^{\text{HK}}$ and $\sigma_{\text{res}}^{\text{H}\alpha}$ does not provide any firm indication that the amount of variability is age-dependent, i.e., variability is not a good age estimator.

It can be argued that the increase of variability in the Ca II H&K and H α residual emission fluxes may lower the degree of correlation in the flux-flux relationship. On the contrary, we find that larger variability generally corresponds to higher levels of Ca II H&K activity (Fig. 12, middle panel), which in turn are associated with stronger correlation (Pearson's coefficient r_{res} closer to 1, smaller p -value $_{\text{res}}$) of the flux-flux relationship (Fig. 13). Conversely, the quietest stars tend to show high p -value $_{\text{res}}$, and in general have lower or negative

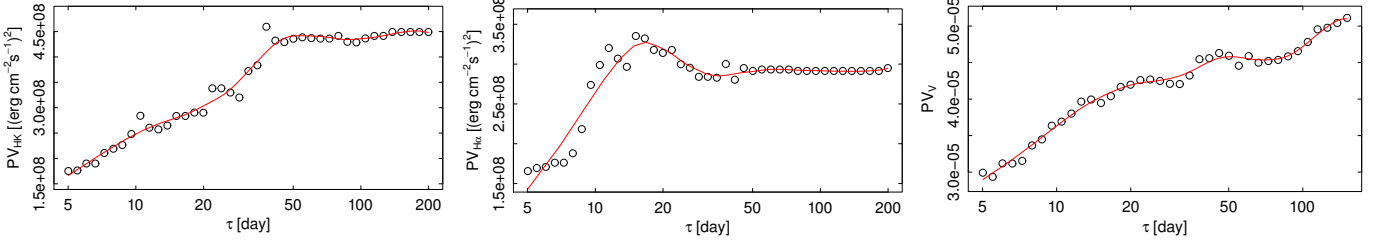


Fig. 14. PV diagrams of GJ 2. From *left to right*, the PV diagrams of F_{HK} , $F_{H\alpha}$, and the V photometric band are shown. The red line is a smoothing function for ease of reading the graphs.

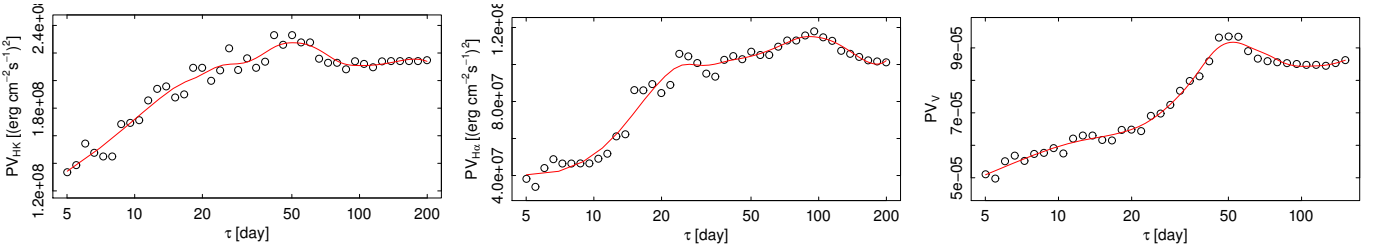


Fig. 15. Same as in Fig. 14, but for GJ 16.

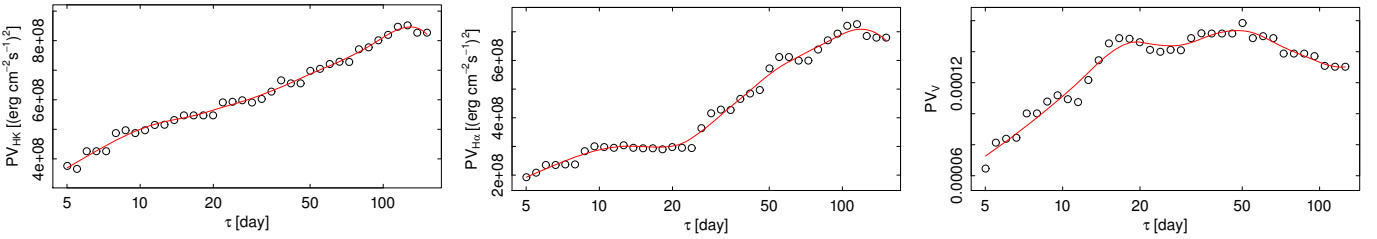


Fig. 16. Same as in Fig. 14, but for GJ 3942.

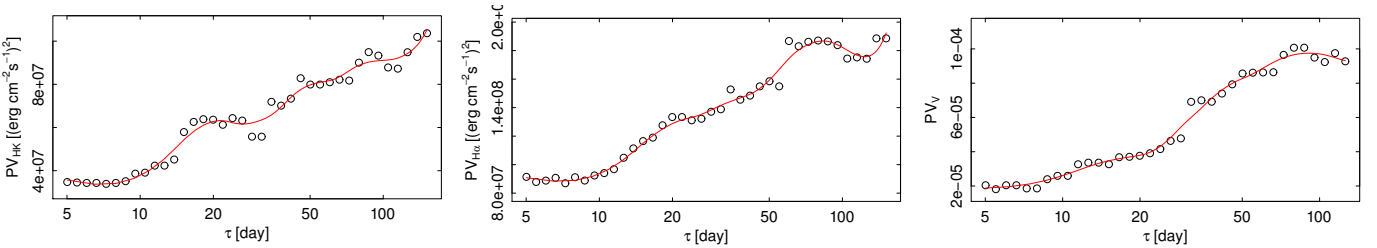


Fig. 17. Same as in Fig. 14, but for GJ 625.

correlation coefficients r_{res} . The same results were also found by [Gomes da Silva et al. \(2011\)](#) by analyzing the long-term activity of a sample of M dwarfs. This supports the assumption that line emission in the Ca II H&K doublet and the $H\alpha$ line are triggered by spatially correlated ARs.

5.2. Variability timescales

To analyze the typical timescales of the variability, we perform a time series analysis using the pooled variance (PV) approach described by [Donahue et al. \(1997a,b\)](#).

The PV is a measure of the average variance in the data over the timescale τ . By construction, the PV computed at τ is expected to be the combination of contributions from a variety of sources, instrumental and/or astrophysical, with timescales

shorter than τ . When τ increases, the PV remains constant until the effects from processes with longer timescales become more noticeable, in which case the PV increases with τ . The PV approach is thus suited for time series containing multiple periodic signals with different amplitudes and phases, and is indicated to study the timescales of variability induced by stellar rotation and the life cycle of ARs.

The PV computation needs a large number of data points to be robust ([Lanza et al. 2004](#)). In 3.5 yr of observations we have collected up to 100 spectra for a single star. Since this is not a particularly abundant data set, we apply the PV analysis only to the eight stars with more than 90 observations, namely GJ 2, GJ 16, GJ 3942, GJ 625, GJ 4306, GJ 694.2, GJ 3998, and GJ 49, and show their PVs in Figs. 14 to 21.

In the computation of the PV diagrams, we consider the Ca II H&K and $H\alpha$ flux excesses corrected for the long-term trend

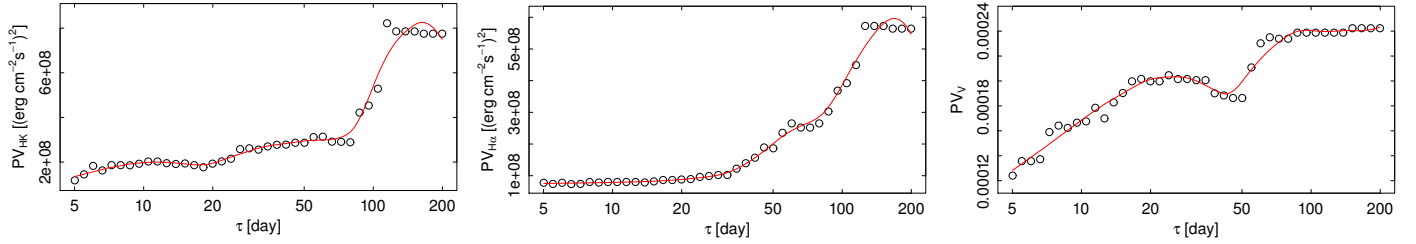


Fig. 18. Same as in Fig. 14, but for GJ 4306.

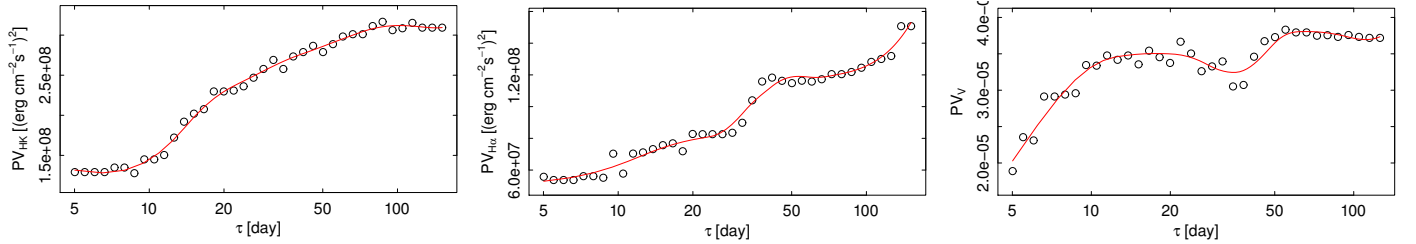


Fig. 19. Same as in Fig. 14, but for GJ 694.

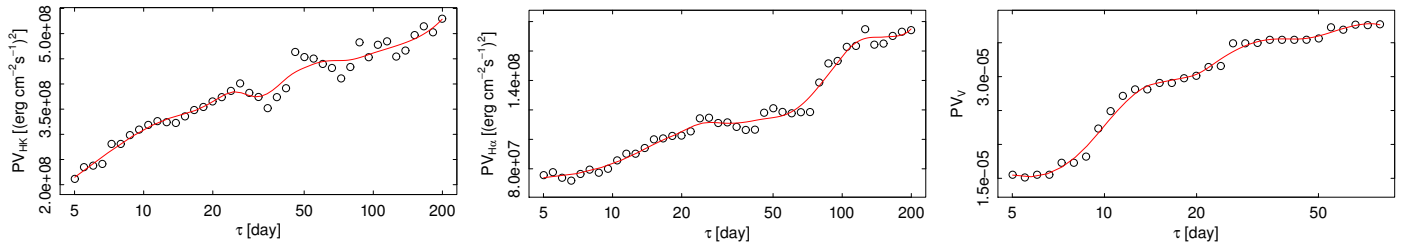


Fig. 20. Same as in Fig. 14, but for GJ 3998.

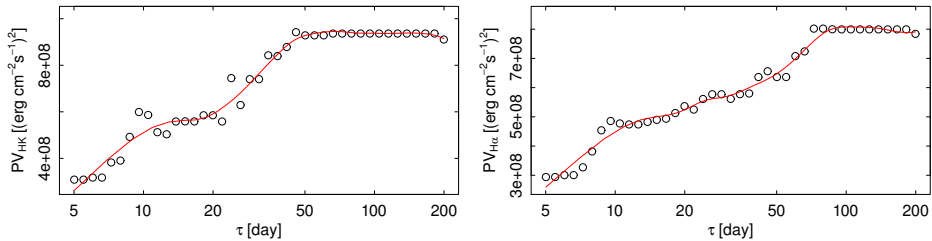


Fig. 21. Same as in Fig. 14, but for GJ 49. The V -band monitoring of this star is not long enough to allow the PV analysis.

(Sect. 4.2), hence we are confident that any contribution from long-term activity cycles is reduced to a minimum. We use the algorithm described in Donahue et al. (1997a); the only difference is that we use the median and the MAD instead of the mean and the standard deviation, respectively, the former being more robust to the presence of outliers especially for small sample sizes.

We complement the spectroscopic data with the V -band photometric monitoring of the same sample of stars, carried out at the INAF-Catania Astrophysical Observatory in the same epochs of the spectroscopic monitoring using an 80 cm telescope. The complete and detailed analysis of the photometric monitoring will be discussed in a future paper.

In general, for each star we find that the Ca II H&K, $H\alpha$, and V -band PV diagrams look alike. The PV diagrams generally increase at small timescales τ ; they reach a plateau at

roughly $\tau \approx 10$ – 40 days, and then increase to level off again at $\tau \gtrsim 50$ days.

Our interpretation of these diagrams is that the first plateau corresponds to the stellar rotation period. This is confirmed by a complementary Lomb-Scargle based analysis⁶ of the Ca II H&K and $H\alpha$ fluxes we are carrying out to measure the rotation period of the stars (see Suárez Mascareño et al., in prep.). The increase at longer τ indicates that there is an additional source of variance in the data with timescales longer than the rotation period. The PV then levels off at $\tau \sim 50$ – 60 days, suggesting that the new source of variance typically has timescales of a few stellar rotations. We exclude that the second flattening is due to cyclic activity for two reasons: dynamo cycles of M dwarfs act on year-long

⁶ The full details of the methodology are described in Suárez Mascareño et al. (2015) where the authors analyze the rotation of a sample of 48 late F-type to mid-M dwarf stars.

timescales, as indicated by Robertson et al. (2013), and their effects should be greatly reduced by the long-term correction we applied in Sect. 4.2. We thus infer that this source of variance is likely related to the growth and decay of chromospheric ARs, as also suggested by Donahue et al. (1997a,b). Other studies have found that M dwarfs may show variability on timescales on the order of 100 days (Davenport et al. 2015; Robertson et al. 2015; Newton et al. 2016), proposing that it is related to the lifecycle of ARs.

6. Summary

In this study we analyze the spectroscopic database collected in the framework of the HADES project, focusing on the characterization of the variability of Ca II H&K and H α lines as chromospheric diagnostics. The database consists of the intensive spectroscopic monitoring performed with the HARPS-N spectrograph at TNG for a sample of 71 low-activity early-type M dwarfs, to search for planets.

We focus on the simultaneous analysis of the flux excess emitted in the Ca II H&K and H α lines. For this purpose, we developed a technique used to correct the spectra for instrumental and atmospheric effects by means of synthetic spectra. This leads to the calibration of the spectra to a common absolute flux scale, which enables the measurement of line flux excesses in units of flux at the stellar surface.

Our measurements show that the Ca II H and K flux excesses are strongly linearly correlated, consistent with previous results (see Martínez-Arnáiz et al. 2011, and references therein). When comparing the Ca II H&K with the H α chromospheric line flux we find significantly more scatter, larger than the measurement uncertainties. In the F_{HK} vs. $F_{\text{H}\alpha}$ diagram we also find statistical evidence of a non-linear flux-flux relationship. In particular, supported by the results of Houdebine et al. (1995) and Houdebine & Stempels (1997), we argue that the Ca II H&K emission flux increases monotonically with the stellar activity level, while the H α line initially goes into absorption and then is filled in by radiative emission processes. This suggests that at very low activity levels the H α absorption by filaments is more evident, while at higher activity levels the emission by plages dominates.

Searching for the physical origin of the scatter and trends in the F_{HK} vs. $F_{\text{H}\alpha}$ diagram, we find that long-term activity cycles on year-long timescales play a minor role in the overall variability. Conversely, short-term activity explains $\geq 60\%$ of the variance of the flux excess measurements.

After correcting the collected time series for year-long variability (whose analysis is not the scope of the present study), we find marginal statistical evidence that the slope of the F_{HK} vs. $F_{\text{H}\alpha}$ relationship increases with decreasing stellar T_{eff} . This is consistent with a scenario in which cooler stars tend to be less affected by chromospheric filaments as suggested by Meunier & Delfosse (2009), and as we found in Maldonado et al. (2017) with an independent analysis of the Balmer decrements of the same sample of stars.

We also find that the variance of the flux excess is an increasing function of stellar activity, and that even the quietest stars show some degree of variability. We attempt a rough time series analysis using the pooled variance approach suggested by Donahue et al. (1997a,b). We find evidence for rotation periods on the order of 10–40 days and active region lifetime cycles longer than ≈ 50 days. These findings are in agreement with

previous results (Donahue et al. 1997a,b; Reiners et al. 2012; Robertson et al. 2013) and are supported by a more detailed analysis, which we are currently preparing (Suárez Mascareño et al., in prep.).

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Appendix A: Additional table**Table A.1.** Relevant stellar parameters of the stellar sample, and number of HARPS-N spectra per target. See [Maldonado et al. \(2017\)](#) for a full description.

Star	T_{eff} (K)	Sp-Type	[Fe/H] (dex)	$\log g$ (cgs)	$v \sin i$ (km s^{-1})	Age	No. of obs.
GJ 2	3713	M1	-0.14	4.76	0.98	Old	95
GJ 3014	3695	M1.5	-0.19	4.79	<1.08	Old	1
GJ 16	3673	M1.5	-0.16	4.78	1.02	Young	107
GJ 15A	3607	M1	-0.34	4.87	1.09	Old	88
GJ 21	3746	M1	-0.12	4.74	1.46	Young	81
GJ 26	3484	M2.5	-0.17	4.88	<0.94	Old	39
GJ 47	3525	M2	-0.26	4.88	1.81	Young	66
GJ 49	3712	M1.5	-0.03	4.73	1.32	Old	95
GJ 1030	3658	M2	-0.08	4.76	<0.93	Young	4
NLTT 4188	3810	M0.5	-0.06	4.70	1.11	Young	1
GJ 70	3511	M2.5	-0.21	4.87	1.02	Old	20
GJ 3117A	3549	M2.5	-0.13	4.82	<0.91	Old	8
GJ 3126	3505	M3	0.01	4.80	<0.83	Old	2
GJ 3186	3768	M1	-0.14	4.74	<1.02	Young	1
GJ 119A	3761	M1	-0.08	4.72	0.98	Old	89
GJ 119B	3508	M3	0.05	4.79	<0.81	Old	3
TYC 1795-941-1	3774	M0	0.01	4.65	3.30	Young	1
NLTT 10614	3728	M1.5	-0.06	4.73	2.07	Young	1
TYC 3720-426-1	3822	M0	0.12	4.64	4.13	Young	2
GJ 150.1B	3730	M1	-0.16	4.76	0.87	Old	44
GJ 156.1A	3745	M1.5	-0.05	4.72	2.85	Old	65
GJ 162	3746	M1	-0.19	4.77	0.93	Young	64
GJ 1074	3765	M0.5	-0.16	4.75	1.13	Old	26
GJ 184	3752	M0.5	-0.10	4.73	1.45	Old	44
GJ 3352	3809	M0.5	-0.13	4.72	1.47	Old	10
TYC 3379-1077-1	3896	M0	0.04	4.61	1.85	Old	8
TYC 743-1836-1	3846	M0	-0.03	4.67	1.73	Young	3
GJ 272	3747	M1	-0.19	4.77	1.09	Young	9
StKM 1-650	3874	M0.5	-0.11	4.67	1.12	Old	10
NLTT 21156	3616	M2	-0.05	4.77	0.70	Young	36
GJ 399	3563	M2.5	0.15	4.72	<0.88	Young	30
GJ 408	3472	M2.5	-0.19	4.89	0.97	Young	36
GJ 412A	3631	M0.5	-0.38	4.87	1.20	Old	72
GJ 414B	3661	M2	-0.09	4.76	<0.94	Old	30
GJ 3649	3691	M1.5	-0.14	4.77	1.55	Young	17
GJ 450	3649	M1.5	-0.20	4.80	1.15	Old	22
GJ 9404	3875	M0.5	-0.10	4.67	1.25	Old	32
GJ 476	3498	M3	-0.16	4.86	<0.93	Old	12
GJ 9440	3710	M1.5	-0.13	4.76	<0.99	Young	69
GJ 521A	3601	M1.5	-0.09	4.79	<0.90	Old	53
GJ 3822	3821	M0.5	-0.13	4.71	0.98	Young	41
GJ 548A	3903	M0	-0.13	4.66	1.11	Old	29
GJ 552	3589	M2	-0.09	4.79	<0.90	Old	61
GJ 606	3665	M1.5	-0.21	4.80	1.57	Old	27
GJ 3942	3867	M0	-0.04	4.65	1.67	Young	98
GJ 625	3499	M2	-0.38	4.94	1.32	Young	99
GJ 3997	3754	M0	-0.24	4.78	0.94	Young	63
GJ 3998	3722	M1	-0.16	4.77	1.56	Old	140
GJ 2128	3518	M2.5	-0.30	4.90	1.19	Young	16
GJ 671	3422	M2.5	-0.17	4.93	0.91	Old	7
GJ 685	3816	M0.5	-0.15	4.72	1.33	Young	44
GJ 686	3663	M1	-0.30	4.83	1.01	Old	2
GJ 694.2	3847	M0.5	-0.21	4.72	1.13	Young	105
GJ 4057	3873	M0	-0.15	4.69	0.81	Old	86
GJ 720A	3837	M0.5	-0.14	4.71	1.49	Old	78
GJ 731	3844	M0	-0.16	4.71	1.59	Old	15

Table A.1. continued.

Star	T_{eff} (K)	Sp-Type	[Fe/H] (dex)	$\log g$ (cgs)	$v \sin i$ (km s^{-1})	Age	No. of obs.
GJ 740	3845	M0.5	-0.14	4.70	0.92	Old	80
GJ 4092	3858	M0.5	-0.06	4.67	1.20	Old	13
GJ 9689	3824	M0.5	-0.13	4.71	1.47	Old	58
GJ 793	3461	M3	-0.21	4.91	<1.00	Old	24
BPM 96441	3896	M0	-0.03	4.63	2.05	Young	13
TYC 2710-691-1	3867	K7.5	0.02	4.63	2.41	Young	1
TYC 2703-706-1	3822	M0.5	0.06	4.65	3.32	Young	68
GJ 4196	3666	M1	0.07	4.71	2.40	Old	1
NLTT 52021	3687	M2	-0.12	4.77	<0.97	Old	1
NLTT 53166	3832	M0	-0.11	4.70	1.45	Old	11
GJ 9793	3881	M0	0.24	4.58	2.77	Young	30
GJ 4306	3763	M1	-0.13	4.74	1.01	Young	123
GJ 895	3748	M1.5	-0.09	4.73	1.70	Old	4
V* BR Psc	3553	M1.5	-0.29	4.88	0.88	Old	10
2MASS J22353504+3712131	3891	K7.5	-0.13	4.67	1.92	Old	4