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RESEARCH

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Machine learning guided postnatal gestational age assessment using new-born screening metabolomic data in South Asia and sub-Saharan Africa

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Abstract

Background: Babies born early and/or small for gestational age in Low and Middle-income countries (LMICs) contribute substantially to global neonatal and infant mortality. Tracking this metric is critical at a population level for informed policy, advocacy, resources allocation and program evaluation and at an individual level for targeted care. Early prenatal ultrasound examination is not available in these settings, gestational age (GA) is estimated using new-born assessment, last menstrual period (LMP) recalls and birth weight, which are unreliable. Algorithms in developed settings, using metabolic screen data, provided GA estimates within 1–2 weeks of ultrasonography-based GA. We sought to leverage machine learning algorithms to improve accuracy and applicability of this approach to LMICs settings.

Methods: This study uses data from AMANHI-ACT, a prospective pregnancy cohorts in Asia and Africa where early pregnancy ultrasonography estimated GA and birth weight are available and metabolite screening data in a subset of 1318 new-borns were also available. We utilized this opportunity to develop machine learning (ML) algorithms. Random Forest Regressor was used where data was randomly split into model-building and model-testing dataset. Mean absolute error (MAE) and root mean square error (RMSE) were used to evaluate performance. Bootstrap procedures were used to estimate confidence intervals (CI) for RMSE and MAE. For pre-term birth identification ROC analysis with bootstrap and exact estimation of CI for area under curve (AUC) were performed.

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Results: Overall model estimated GA had MAE of 5.2 days (95% CI 4.6–6.8), which was similar to performance in SGA, MAE 5.3 days (95% CI 4.6–6.2). GA was correctly estimated to within 1 week for 85.21% (95% CI 72.31–94.65). For preterm birth classification, AUC in ROC analysis was 98.1% (95% CI 96.0–99.0; $p < 0.001$). This model performed better than lowa regression, AUC Difference 14.4% (95% CI 5–23.7; $p = 0.002$).

Conclusions: Machine learning algorithms and models applied to metabolomic gestational age dating offer a ladder of opportunity for providing accurate population-level gestational age estimates in LMICs settings. These findings also point to an opportunity for investigation of region-specific models, more focused feasible analyte models, and broad untargeted metabolome investigation.

Keywords: Pre-term births, Machine learning, Gestational age, New born screening

Background

Of 15 million preterm births annually, 90% happen in Low and Middle Income Countries (LMICs) [1, 2] contributing to 1 million deaths < 5 years, 35% of deaths < 28 days [3]. Further 23.3 million infants (19.3% of live births) are born small for gestational age (SGA) in LMICs. Reduction by 10.0% in these would reduce neonatal deaths by 254,600 deaths [4]. Identifying and tracking this metric is therefore critical for advocacy, surveillance, research, evaluation of preventive strategies, and care of these high risk infants in LMICs. These in turn are essential to achieving United Nations Sustainable Development Goal 3 target 3.25 [2, 5] (elimination of preventable under-five deaths by 2030).

The estimation of accurate gestational age at birth is essential for identifying both preterm and SGA births. Early ultrasonography examination, considered as gold standard for gestational age (GA) assessment is unavailable due to high equipment cost and lack of trained manpower in most LMICs settings. Recall of last menstrual period (LMP) [6] used in these settings is unreliable in estimation of GA [7]. Postnatal methods, birth weight and standardized scoring system (Dubowitz or Ballard scales) have poor reliability and high inter user variability limiting their usage [8–11].

In global health a need for novel tools is required that could help monitor these metrics in LMIC countries on a population scale [12]. Algorithms developed in three North American settings using routine metabolic screen data to derive GA estimates, have been shown to provide accurate estimates to within 1–2 weeks of ultrasonography based GA [13–15]. Limited data for external validity of these methods in LMIC populations [16, 17] demonstrated satisfactory performance but lower accuracy for GA predictions especially among SGA new-borns in Africa and Asia. Reported publications used conventional statistical modelling approaches like linear/logistic regression and discriminant analysis. These statistical methods mainly focus on inference from fitting of a project-specific probability model [18]. Recent advances in Machine

Learning (ML) techniques and big data analysis, allows for efficient handling of large numbers of predictors while incorporating non-linear association and complex interactions. ML techniques are more robust in nature and they mainly deal with the prediction of outcomes by using general-purpose learning algorithms to find patterns in a dataset without assumptions needed for conventional modelling. Recently, ML approaches has been shown to help preterm identification in hospital setting [19].

We hypothesized that application of ML to metabolite profile datasets of Alliance for Maternal and New-born Health Improvement (AMANHI) All children thrive (ACT) cohorts (representing both South Asia and Sub-Saharan Africa) would potentially improve the prediction of GA as compared to conventional approaches previously reported. The rationale for these cohorts and associated bio-bank procedures, and cohort characteristics have been described in previous publications [20, 21]. One of the objectives of the AMANHI study was to develop and validate programmatically feasible approaches to accurately assess the gestational age of babies after they are born. Additionally this method will not be dependent on the North American population datasets for generation of equation coefficients and enable regional adjustments in the future. Among various ML classifiers, we chose to use random forest as particularly well suited for clinical predictions [22]. We report the performance of our machine learning based GA estimation algorithms.

Methods

Study population

This study was undertaken using data from the Alliance for Maternal and New-born Health Improvement (AMANHI) All children thrive (ACT), community based, prospective pregnancy and New-born cohorts from Pemba (Tanzania), Sylhet (Bangladesh) and Karachi (Pakistan). These studies received ethical approval from the local and institutional ethics committees of all the three sites: ICDDR, B and John Hopkins

University for Bangladesh, Aga Khan University for Pakistan and ZAMREC for Tanzania. The protocols were also approved by the Ethical Review Committee (ERC) at World Health Organization (WHO). Briefly, women were enrolled in early pregnancy and followed through delivery and the postpartum period (supplementary Fig. 1). All pregnant women were eligible to participate in the study. Written informed consent was obtained from all participants prior to study enrolment for collection of maternal and new-born data and samples. GA was established by ultrasonography [23] at screening using the fetal crown rump length (if < 14 weeks gestation) [24] or biparietal diameter and femur length (if ≥ 14 weeks) [23, 24]. All fetal biometry measurements were measured twice and then averaged for gestational age calculations [23, 25]. Birth weight (5 g sensitivity) was measured using standard new-born weighing scale (SECA corporation, Columbia, MD).

Sample collection and processing

The metabolic screening data from 1283 samples used for this analysis was generated as part of the AMANHI collaboration with Department of Epidemiology, College of Public Health, University of Iowa, for evaluating external validity of the GA estimation methods developed based on American samples [13, 17]. An overview is provided in consort flow diagram (Fig. 1) and protocol process flow (supplementary Fig. 1). Heel prick blood spots were obtained on a protein saver card (Whatman^R 903, GE healthcare, USA), within 24–72 h of birth from new-born's as per standard procedures. All 903 cards were labelled with barcoded unique identifier (ID's), were air-dried and stored in air tight Ziplock bags with desiccant at -80°C . The 903 cards were shipped in dry ice to the University of Iowa where they were examined for quality. Then samples were sent to State Hygienic Laboratory, Ankeny, Iowa, USA at regular intervals (ensuring processing before potency window). Sixty-six metabolites (Supplementary Table 4) which included amino acid, acylcarnitines, enzymes and hormones were analysed using tandem mass spectrometry. Only singleton births (women who gave birth to only one child during a delivery, excluding twins, triplets) were included in the final analysis since analyte values are associated with birth status [26]. We excluded 35 non-singleton pregnancies in the final analysis.

Selection of the machine learning algorithm

For implementing the machine learning algorithms, the dataset was divided into training and test datasets. This selection was made without replacement. Training dataset used was derived by combining stratified sampling of 80% from Africa and Asia separately. A K fold cross validation method was used as resampling procedure to

evaluate the machine learning algorithms which effectively performed a 10 fold cross validation, with three repeats [27]. Performance of the models was assessed by comparing RMSE and MAE values on the training dataset. Four ML algorithms considered appropriate for this analysis; artificial neural network (ANN), decision tree (DT), support vector machines (SVM) and random forest (RF) were evaluated and compared [28]. Amongst these RF regressor performed better and was selected for testing on test dataset (results provided in Supplementary Table 1).

Architecture of random Forest Regressor

ML models were generated using metabolite profiles along with birth weight and gender. RF regressor model for this analysis was built using 1) bootstrap sampling -with replacement and 2) Random feature $-m = 10$ in RF regressor. In this generation procedure was repeated until 10 decision trees were created to form a randomly generated "Forest". We denoted the hyper-parameters for RF regressor as "mtry" (a variable imputed in the model) at each split node for performing regression; the default value of "mtry" is $p/3$ where p is the number of predictors [29]. Random forest is an algorithm built by multiple decision trees. So it is important to choose the hyper-parameter as it defines how each tree will be built.

Implementation of random Forest Regressor

Sklearn.ensemble [30] (Random Forest Regressor package), a Python module was used for running the RF regressor. NumPy, Scipy and Pandas were used as python dependencies for running the module. R coding was used to create train and test datasets.

Models used for the analysis

Four different models were used for the prediction of Gestational Age. (Supplementary Table 3; details of variables in the models provided in Supplementary Table 4).

Model 1: only one variable per metabolite (for all 66 metabolites) from the profile of the blood metabolites along with birthweight and gender.

Model 2: was designed to replicate model published by Ryckman et al. [13] which included the linear, squared and cubic values of the metabolites as predictors.

Model 3: was designed to replicate the model published by Wilson et al. 2017 [15] which included the linear, squared and cubic values of the metabolites along with birthweight and gender as predictors.

Model 4 (selected Model): Contained all the predictors used in the above three models.

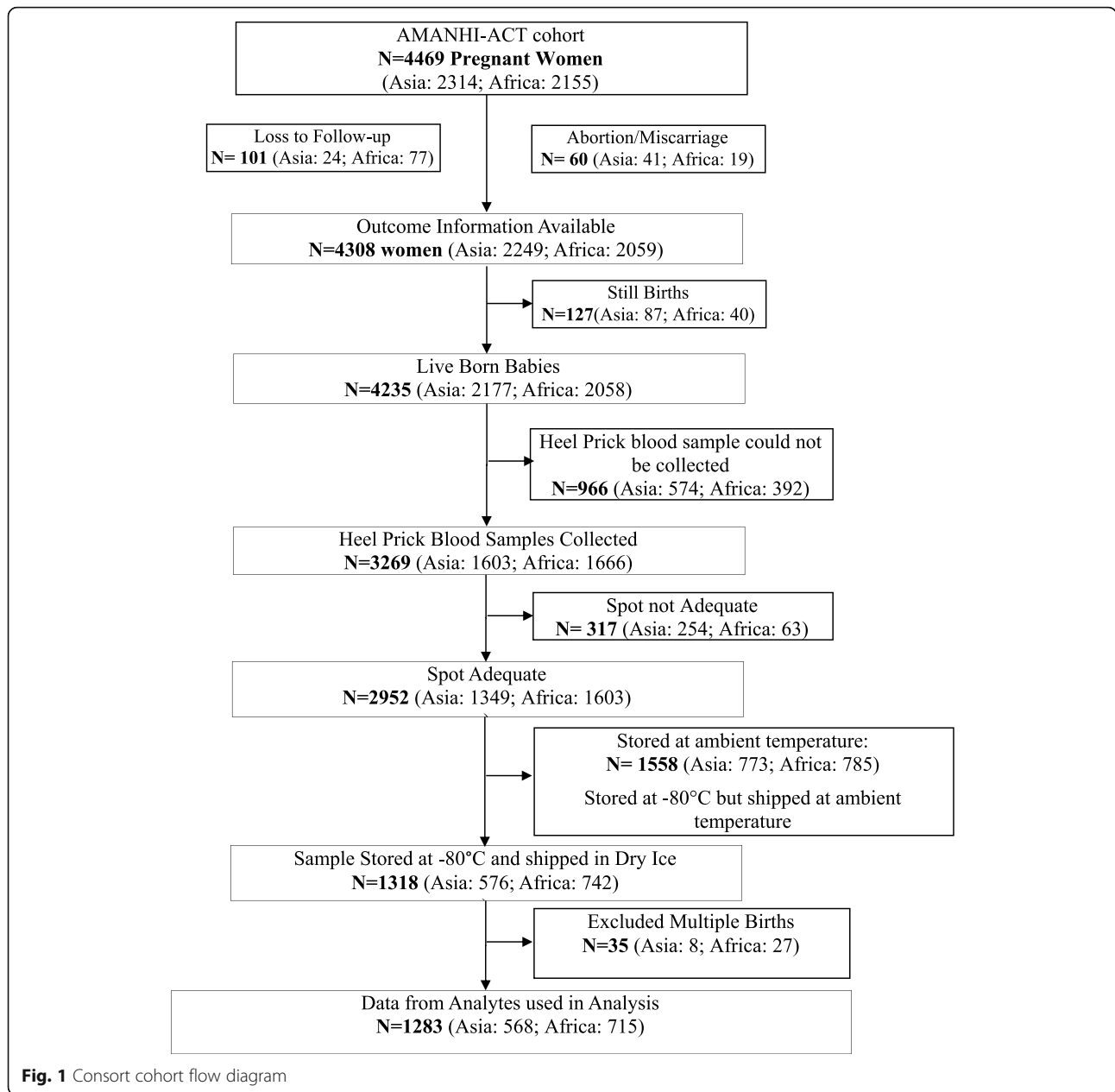


Fig. 1 Consort cohort flow diagram

Best model selection

The best model was selected on the basis of root mean square error (RMSE) and mean absolute error (MAE).

The RMSE of a predicted model with respect to the estimated variable x_{model} has been defined as the square root of the mean squared error [31].

$$RMSE = \sqrt{\frac{\sum_{i=1}^n (X_{obs,i} - X_{model,i})^2}{n}}$$

Where, x_{obs} is observed values, x_{model} is modelled values at time i.

Mean absolute error (MAE) has been calculated as

$$MAE = \frac{\sum_{i=1}^n |y_i - \chi_i|}{n} = \frac{\sum_{i=1}^n |e_i|}{n}$$

where x_i is the prediction and y_i is the true value.

Confidence intervals for RMSE and MAE

Efficient computation of RMSE, MAE values and 95% confidence interval were estimated using bootstrapped procedures [32] (Python package, (Bootstrapped0.0.2) <https://pypi.org/project/bootstrapped/>) and for t with a fixed seed number of 1 using boot [33] and metrics packages in R.

ROC analysis for evaluating discriminatory ability of the ML based GA

For ROC analysis we used Stata 16.1 (Stata Corp LLC, Texas USA) and MedCalc (MedCalc Software Ltd. Belgium). The predictions of the continuous outcome (predicted GA) were dichotomized to < 37 weeks (pre-term birth) and \geq 37 weeks (term birth) for carrying out ROC analysis. Generation of ROC curve and AUC estimation was performed and interpreted using standard methods [34, 35]. We estimated Youden index J [36]

$$J = \max\{sensitivityf[c] + specificityf[c] - 1\}$$

where c ranges over all possible criterion values. Graphically, J is the maximum vertical distance between the ROC curve and the diagonal line. Bootstrapped 95% CI for Youden index and its corresponding criterion value were estimated [37, 38]. 95% CI for sensitivities and specificities were also estimated for a range of fixed and pre-specified sensitivities/specificities [33] and 95% CI estimated using bootstrapping [37, 38]. Comparison of ROC curves estimating difference, confidence interval and p -value were also performed using bootstrap methods [39, 40]. For the Bootstrap estimation a fixed seed was used to enable replication of the analysis.

Results

General characteristic of the cohort

Data from all 1318 new-borns having new-born metabolic screen analytes were included in the current analysis. Of these 742 samples were from Africa and 576 from Asia (Pakistan and Bangladesh). Baseline characteristics of the sample are provided in Table 1. The distribution of male and female subjects in the cohort was almost in the ratio of 1:1. The mean GA as confirmed by ultrasound was 38.5 ± 1.68 weeks (mean \pm SD). Sample included 153 (11.6%) preterm births, 199 (15.1%) low birth weight and 271 (20.6%) SGA new-borns. Birth weight in African new-borns (3240.75 ± 585.88 g) tended to be higher than Asian new-borns (2774.55 ± 513.99 g).

Comparison of performance of gestational age estimation models

Initially we evaluated 4 models for performance predicting gestational age (Supplementary Table 4), the model 1 with only base terms for analytes was least accurate RMSE 1.38, model 2 using variables in final Iowa regression model [13] had RMSE of 1.29 and model 3 using variables in Ontario regression model [14] had RMSE 1.20 (Supplementary Table 2), the final all-inclusive model providing a RMSE of 1.02 (95% CI 0.91–1.14) was selected and evaluated further. For identification of

Table 1 Cohort Characteristics Of Infants Included In The Metabolic Screening Study

Heel Prick Samples	All sites Combined (Total cohort) N = 1318	Asia (Pakistan and Bangladesh) N = 576	Africa (Tanzania) N = 742
Gender			
Male	695 (52.7%)	268 (46.5%)	428 (57.6%)
Female	623 (47.3%)	308 (53.5%)	315 (42.3%)
Gestational Age Mean \pm S.D	38.53 ± 1.68	38.35 ± 1.67	38.68 ± 1.68
\geq 37 weeks	1165 (88.4%)	492 (85.4%)	673 (90.7%)
< 37 weeks	153 (11.6%)	85 (14.6%)	69 (9.3%)
34–37 weeks	126 (82.4%)	71 (83.5%)	54 (78.3%)
< 34 weeks	27 (17.6%)	14 (16.5%)	15 (21.7%)
Birthweight (Mean \pm S.D)	3037.21 ± 601.67	2774.55 ± 513.99	3240.75 ± 585.88
Birth Weight Category, n(%)			
\leq 2500 g	199 (15.1%)	153 (26.6%)	46 (6.2%)
> 2500 g	1119 (84.9%)	423 (73.4%)	696 (83.8%)
SGA Status			
Yes	272 (20.6%)	91 (15.8%)	181 (24.4%)
Multiple Birth Status	35 (2.7%)	8 (1.4%)	27 (3.6%)
Newborn Sample Collected (Hrs), Mean \pm SD	49.0 ± 16.2	52.1 ± 19.4	46.6 ± 12.7

preterm births, AUC of Model 4 was significantly better than model 1 [3.6% (95%CI - 1.2 to 8.5; $p = 0.014$)] and Model 2 [2.5% (95% CI - 2.2 to 7.3; $p = 0.03$)] (Supplementary Fig. 2).

Overall model estimated gestational age had a mean absolute error (MAE) of 5.2 days (95% CI 4.5–6.8), compared to gold standard ultrasound dating. Accuracy was slightly lower in Africa MAE 5.3 days (95% CI 54.8–6.2) than Asia MAE 5 days (95% CI 4.3–6.2) (Table 2). Contrary to the results from external validity of regression models [15, 17], performance in SGA new-borns was not appreciably reduced, MAE 5.3 days (95% CI 4.6–6.2 days). GA was correctly estimated to within 1 week of ultrasound-assigned values for 85.21% (95% CI 72.31–94.65) overall, 83.2% (95% CI 78.31–90.05) in African and 87.7% (95% CI 76.63–95.39) in Asian new-borns. Estimations performed as well in SGA new-borns within 1 week 83.9% (95% CI 71.21–92.32) (Table 2).

To evaluate impact of using a regionally trained algorithm (an important future prospect), we repeated the analysis with machine learning being trained by African sample for Africa estimations and Asian sample for Asia estimations. The model performance in spite of reduced samples for training improved both for Africa and Asia. The precision of MAE improved rather being reduced for Africa to 5 days (95% CI 4.1–6.0) and Asia to 4.8 days (95% CI 4.1–6.1) (Table 2). The same pattern was seen for RMSE (Table 2).

Model discrimination of preterm birth

For the ability to classify correctly preterm births (GA < 37 weeks), model in ROC analysis showed an area under curve (AUC) of 92.6% (BC 95% CI 87.5–96.1; $p < 0.001$). Criterion of ≥ 37 providing a sensitivity of 100% and specificity of 92.61% (Fig. 2). This model provided a significant improvement (difference in AUC 14.4% (95% CI 5.0–23.7; $p = 0.002$) over predicting GA by regression

models in the same dataset, AUC 84% (95 CI 78.6–88.0), (Fig. 3b). There was no significant difference, 1.3% (95% CI - 1.5–4.3; $p = 0.333$) in AUC between Africa and Asia (Fig. 3a). The AUC between SGA and non-SGA new-borns also did not differ (AUC difference 1.7%;(95% CI - 1.2–4.8; $p = 0.891$) (Fig. 3c).

Performance across gestational age categories

Estimation of RMSE and MAE as well as cross tabulation of actual and predicted GA by 2 weekly categories (Table 3), indicated that the accuracy of the current set of analytes was diminishing only at < 35.

Discussion

This study has highlighted promising application of ML methodology to birth weight and new-born metabolomic screening data for improving postnatal prediction of gestation age at birth and discriminating between preterm and term new-borns. It also demonstrated ability of using LMICs data for training ML models and not needing external estimators from developed country datasets. In LMICs setting of South Asia and Sub-Saharan Africa, GA estimates from ML model were within an average of 5.2 days of ultrasonography based GA. The ML estimated GA enabled discrimination between pre-term and term births, AUC 98% was significantly better than regression estimated GA AUC 84%. The optimal criterion of ≤ 37 weeks providing a sensitivity of 100% and specificity of 92.616%.

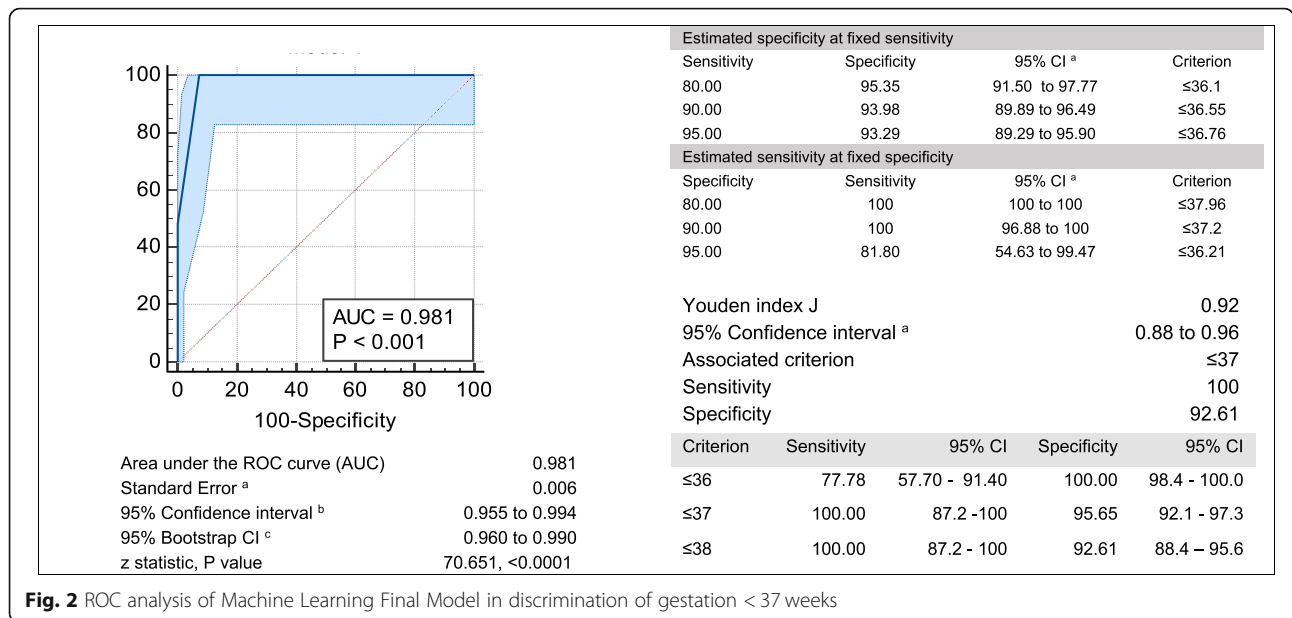
As against lower performance of previous approaches [13, 15–17], in estimating GA in SGA sub population, our ML model estimates were within 5.3 days of ultrasonography based GA. This also reflected in the finding of a similar proportion with estimated gestation being within 1 week of the ultrasound confirmed gestation, 85.2% overall vs 83.9% in SGA subgroup. Use of data with 80% each of Asian and African data for training the

Table 2 Mean Abs Error and RMSE in weeks in final machine learning model

STATISTICS	Cohort		Africa		Asia	
	Overall	SGA	Overall	SGA	Overall	SGA
Training Dataset	Test Data Set- Pooled remaining		Test Data Set - remaining		Test Data Set- remaining	
80% Pemba Samples + 80% Asian samples	20% PembaSamples +	20% Asiansamples	20% Pemba Samples		20% Asian Samples	
MAE (95% CI)*	0.74 (0.65–0.98)	0.76 (0.65–0.88)	0.75 (0.61–0.89)	0.88 (0.75–1.16)	0.72 (0.62–0.88)	0.73 (0.61–0.95)
RMSE(95% CI)*	1.02 (0.91–1.14)	1.05 (0.91–1.19)	1.04 (0.89–1.16)	1.20 (1.10–1.31)	1.00 (0.89–1.16)	1.01 (0.93–1.19)
1 week difference (%)*	85.21 (72.31–94.65)	83.9 (71.21–92.32)	83.21 (78.31–90.05)	72 (65.67–79.34)	87.71 (76.63–95.39)	87.09 (77.67–94.21)
2 weeks difference (%)*	99.61 (91.42–100)	98.31 (89.74–100)	100 (93.32–100)	100 (92–79–100)	99.12 (91.56–100)	99.15 (90.45–100)
Training Dataset 80% Africa samples for Africa and 80% Asia samples for Asia			Test Dataset (20% Africa samples)		Test Dataset (20% Asia samples)	
MAE (95% CI)*			0.71 (0.58–0.85)	0.83 (0.71–1.10)	0.68 (0.58–0.87)	0.71 (0.62–0.83)
RMSE (95% CI)*			0.96 (0.82–1.07)	1.13 (1.01–1.27)	0.93 (0.82–01.05)	0.97 (0.84–1.08)

*Bootstrapped,

*Detailed description of the analytes used in the models have been given in supplementary information



models, was associated with some variation in predictive accuracy for Asia (average of 5 days) compared to Africa (average of 5.2 days). Using region specific data for training reduced the variation to 4.8 and 4.9 days respectively and improved the precision in spite of reduced sample size of training dataset (Table 2). Being preliminary proof of principle, these findings provide a vision for future implementations, where in region specific training

datasets may improve global application of metabolomics based data for gestational age assessment.

Our study had a number of important strengths and also some limitations which need consideration while interpreting the results. The strengths included 1) a sampling frame which utilized samples from both South Asia and East Africa, home to most of the global mortality associated with preterm and SGA births, 2) the study

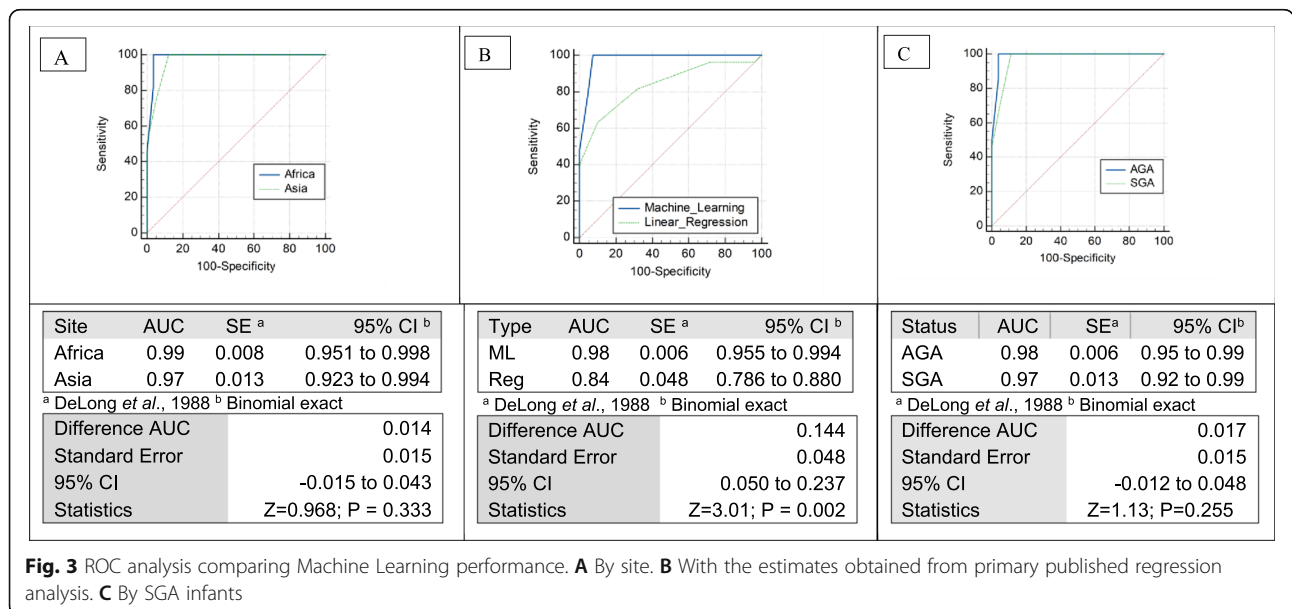
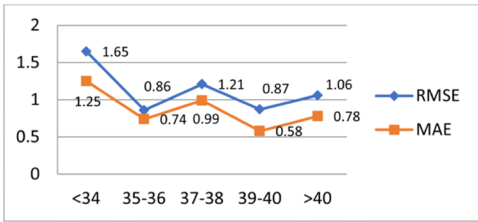


Table 3 Performance and Concordance of predicted gestational age by actual gestation age group

Performance and Concordance of predicted gestational age by actual gestation age group						
						
Table 3a: Model Performance, RMSE and MAE by GA group						
	AGE GROUPS					
Measures	<=34	35-36	37-38	39-40	>40	
RMSE*	1.65	0.86	1.21	0.87	1.06	
95% CI*	1.42-1.79	0.71-0.98	1.13-1.35	0.72-0.94	0.91-1.16	
MAE*	1.25	0.74	0.99	0.58	0.78	
95% CI*	1.12-1.39	0.62-0.88	0.87-1.06	0.41-0.74	0.61-0.89	
*Bootstrapped						
Table 3b: Concordance between actual and predicted gestational age (50% cohort)						
Actual Gestational Age, wks	n	Predicted Gestational Age, wks				
		<=34	35-36	37-38	39-40	>40
<=34	8	4 (50%)	3	1		
35-36	19		14 (73.68%)	5		
37-38	74		10	23 (31.08%)	41	
39-40	133			14	111 (83.45%)	8
>40	23				4	19 (82.60%)
Asia	Total	<=34	35-36	37-38	39-40	>40
<=34	5	2 (40%)	2	1		
35-36	11		11 (100%)			
37-38	29			21 (72.41%)	8	
39-40	59			14	45 (76.27%)	
>40	10					100 (100%)
Africa	Total	<=34	35-36	37-38	39-40	>40
<=34	3	2 (66.67%)	1			
35-36	8		7 (87.5%)	1		
37-38	45			8 (17.8%)	37	
39-40	74				66 (89.19%)	8
>40	13				4	9 (69.23%)

design was nested in a well-described population-based cohort of pregnancy with WHO coordinated and harmonized protocols and SOP, 3) Active surveillance for early pregnancy identification with added measures (menstrual calendar, pregnancy), culminating in harmonized ultrasonography based gestation assessment between 8 and 19 weeks of gestation and 4) Sample collection, storage and shipment SOP based on pilot QC, resulting in high quality of samples. The primary limitation of this study is the participation bias against early preterm and early deaths before sample collection window. Relatively small proportion of actual births in this sub-sample limits our ability to comment on model performance in these sub-groups. Our finding of lower accuracy in pre-term ≤ 34 weeks may either reflect lack of association of the metabolites in that sub-group, a function of lack of sample in that group and/or bias introduced by selective exclusion of early deaths. Additionally, we were working with the limitation of a smaller sample size as compared to the usual sample sizes in machine learning universe. We did try to use methods appropriate to accommodating smaller sample sizes, however would not have been protected against extreme chance affecting the sample. The ability to train the model and precision of estimates is somewhat reassuring but would need confirmation.

Preterm births and SGA account for a substantial burden of mortality in first 5 years [3, 4]. Tracking these metrics is therefore critical for advocacy, allocation of resources, for surveillance, research, evaluation of preventive strategies, and care of these high-risk infants in low- and middle-income countries [3, 41]. At the core of this is the estimation of gestational age at birth and

being able to discriminate pre-term births accurately. Difference in GA at birth of a week impacts neonatal morbidity, mortality, and long-term outcomes significantly [42, 43]. Our findings provide evidence that ML based gestational dating models improve upon the currently-used postnatal gestational age estimation methods [6, 8, 10, 11, 13–17]. However while considering implementation of metabolic gestational dating approaches for robust population-level estimates, current challenges and future opportunities that machine learning brings to this domain need consideration. Heel prick samples for new-born screening are typically collected at least 24 h after birth to accommodate postpartum fluctuations in analyte levels. This introduces a bias due to early deaths selectively occurring in pre-term births, further in LMICs settings as most mother-infant pairs do not stay in hospital beyond 24 after delivery [44]. In most LMICs new-born screening is not a standard practice and will entail challenges in sample collection and processing for metabolic screening, therefore scale up needs to include rethinking about development of cord-blood-specific models restricted to analytes less susceptible to fluctuations in the postnatal environment, establishing a profile of fewer selected metabolites that are measurable using less sophisticated equipment. While rethinking and investigating low-tech variations suitable to LMICs settings, also to consider are, newer high throughput trans proteome/metabolome platforms which are now becoming affordable (i.e. Seers Nano peptide technology [45]). An untargeted metabolomic approach may improve our ability to estimate GA postnatally while also identifying infants at risk of a variety

of conditions. Use of a broader spectrum of analytes may also help select a restrictive model for cord blood. Building on this study, use of ML methodology would positively influence development of all the above approaches, due to flexibility, ability to use regional data for ML and not requiring circling back to accumulating large datasets with new intended analyte profiles.

Summary

Towards implementing preterm birth surveillance initiatives [46] ML algorithms and models applied to metabolomic gestational age dating offer an opportunity ladder to provide accurate population-level gestational age estimates in LMICs settings. Further research should focus on application of ML enabling investigation and incorporation of region-specific models, evaluating broad untargeted metabolome or more focused feasible analyte pool with ML approaches. Derivation and optimization of cord blood metabolic profiles models predicting gestational age accurately would usher a new feasibility for use of this approach in LMICs settings.

Abbreviations

LMICs: Low And Middle Income Countries; SGA: Small For Gestational Age; GA: Gestational Age; LMP: Last Menstrual Period; ML: Machine Learning; AMANHI: Alliance For Maternal And Newborn Health Improvement; ACT: All Children Thrive; ANN: Artificial Neural Network; DT: Decision Tree; RF: Random Forest; RMSE: Root Mean Square Error; MAE: Mean Absolute Error; CI: Confidence Interval; SD: Standard Deviation; WHO: World Health Organisation; QC: Quality Control

Supplementary Information

The online version contains supplementary material available at <https://doi.org/10.1186/s12884-021-04067-y>.

Additional file 1: Fig. S1. Pregnancy cohort and Study design. **Table S1.** Performance metrics of the different machine learning algorithms. **Table S2.** RMSE and Mean Abs Error in weeks obtained in three comparative models. **Fig. S2.** ROC analysis comparing discriminatory ability of four models evaluated. **Table S3.** Newborn screening metabolomic analytes estimated. **Table S4A.** Description of the model. **Table S4B.** Description of the model.

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Authors' contributions

a. SS, KR, AB, FJ, RB contributed to conceptualization, design and implementation, quality control, participated in analysis and interpretation of data. b. SS, S. Das, AD, RB Conceptualized Machine learning exercise and contributed to pooled data analysis and machine learning analysis. c. RK, IN, EJ, UD, contributed to the implementation, analysis and quality control in the field. d. KKR, EJ, BB, contributed to laboratory analysis of samples. e. UM, AD, NHC, AB, SR, S. Deb, SA, FK, RR, AM, SY contributed to field implementation, data collection and quality control. f. RB, AM, SY

coordinated the study. g. SS, SDAs, HM, AD additionally contributed to first draft write-up. h. All authors reviewed and contributed to manuscript writing. The author(s) read and approved the final manuscript.

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Availability of data and materials

The datasets used and/or analysed during the current study are available from the corresponding author on reasonable request.

Declarations

Ethics approval and consent to participate

The study was performed in accordance with the ethical standards of the Declaration of Helsinki (1964) and its subsequent amendments. All study protocols for AMANHI cohorts were approved by appropriate institutional review board in each of the participating sites. ICDDR, B and John Hopkins University for Bangladesh, Aga Khan University for Pakistan and ZAMREC for Tanzania. The protocols were also approved by the Ethical Review Committee (ERC) at World Health Organization (WHO). Additionally, institutional/Local sample utilization committees approved shipment of samples to Iowa for metabolic screening assay. Written informed consent for ultrasonography and additional procedures was obtained from all participating mothers in their local or preferred language by study supervisors during enrolment. Mothers were asked for additional informed consent before obtaining a heel prick from the baby.

Consent for publication

Not applicable.

Competing interests

No competing interests.

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References

- Lawn JE, Kinney M. Preterm birth: now the leading cause of child death worldwide. *Sci Transl Med.* 2014;6:263ed221.
- Liu L, Oza S, Hogan D, Chu Y, Perin J, Zhu J, et al. Global, regional, and national causes of under-5 mortality in 2000–15: an updated systematic analysis with implications for the Sustainable Development Goals. *Lancet.* 2016;388(10063):3027–35. [https://doi.org/10.1016/S0140-6736\(16\)31593-8](https://doi.org/10.1016/S0140-6736(16)31593-8) Epub 2016 Nov 11. Erratum in: *Lancet.* 2017 May 13;389(10082):1884. PMID: 27839855; PMCID: PMC5161777.
- Walani SR. Global burden of preterm birth. *Int J Gynecol Obstet.* 2020;150(1): 31–3. <https://doi.org/10.1002/ijgo.13195>.
- AnneCC L, Naoko K, Simon C, Stevens Gretchen A, Hannah B, Silveira Mariangela F, et al. Estimates of burden and consequences of infants born small for gestational age in low and middle income countries with INTERGROWTH-21st standard: analysis of CHERG datasets. *BMJ.* 2017;358:j3677.

5. United Nations. Sustainable development goals. New York: United Nations; 2015. (accessed Sept 11, 2015). <http://www.un.org.proxy1.library.jhu.edu/sustainabledevelopment/summit/>
6. Alexander GR, de Caunes F, Hulsey TC, Tompkins ME, Allen M. Validity of postnatal assessments of gestational age: a comparison of the method of Ballard et al. and early ultrasonography. *Am J Obstet Gynecol.* 1992;166:891–5. [https://doi.org/10.1016/0002-9378\(92\)91357-G](https://doi.org/10.1016/0002-9378(92)91357-G) PMID: 1550159.
7. Lynch CD, Zhang J. The research implication for the estimation of gestational age dating method. *Paediatr Perinat Epidemiol.* 2007;21(Suppl 2):86–96. <https://doi.org/10.1111/j.1365-3016.2007.00865.x>.
8. Taylor R, Beyai S, Owens S, Denison F. The external ballard examination does not assess gestational age accurately in a rural field setting in the Gambia. *Arch Dis Child Fetal Neonatal Ed.* 2010;95:Fa103.
9. Spinnato JA, Sibai BM, Shaver DC, Anderson GD. Inaccuracy of Dubowitz gestational age in low birth weight infants. *Obstet Gynecol.* 1984;63(4):491–5. 6700894.
10. Sanders M, Allen M, Alexander GR, Yankowitz J, Graeber J, Johnson TR, et al. Gestational age assessment in preterm neonates weighing less than 1500 grams. *Pediatrics.* 1991;88(3):542–6.
11. Robillard PY, De Caunes F, Alexander GR, Sergent MP. Validity of postnatal assessments of gestational age in low birthweight infants from a Caribbean community. *J Perinatol.* 1992;12(2):115–9.
12. Wyber R, Vaillancourt S, Perry W, Mannava P, Folaranmi T, Celi LA. Big data in global health: improving health in low- and middle-income countries. *Bull World Health Organ.* 2015;93(3):203–8. <https://doi.org/10.2471/BLT.14.139022>.
13. Ryckman KK, Berberich SL, Dagle JM. Predicting gestational age using neonatal metabolic markers. *Am J Obstet Gynecol.* 2016;214(4):515.e1–515.e13. <https://doi.org/10.1016/j.ajog.2015.11.028>.
14. Wilson LA, Murphy MS, Ducharme R, Denize K, Jadavji NM, Potter B, et al. Postnatal gestational age estimation via newborn screening analysis: application and potential. *Expert Rev Proteomics.* 2019;16(9):727–31. <https://doi.org/10.1080/14789450.2019.1654863> Epub 2019 Aug 17. PMID: 31422714; PMCID: PMC6816481.
15. Jelliffe-Pawlowski LL, Norton ME, Baer RJ, Santos N, Rutherford GW. Gestational dating by metabolic profile at birth: a California cohort study. *Am J Obstet Gynecol.* 2016;214(4):511.e1–511.e13. <https://doi.org/10.1016/j.ajog.2015.11.029>.
16. Murphy MSQ, Hawken S, Cheng W, Wilson LA, Lamoureaux M, Henderson M, et al. External validation of postnatal gestational age estimation using newborn metabolic profiles in Matlab, Bangladesh. *eLife.* 2019;8:e42627.
17. Sazawal et al., 2021. Using AMANHI-ACT cohorts for external validation of lowa new-born metabolic profiles based models for postnatal Gestational age estimation. In submission.
18. Bzdok D, Altman N, Krzywinski M. Statistics versus machine learning. *Nat Methods.* 2018;15(4):233–4. <https://doi.org/10.1038/nmeth.4642>.
19. Rittenhouse KJ, Vwalika B, Keil A, et al. Improving preterm newborn identification in low-resource settings with machine learning. *PLoS One.* 2019;14(2):e0198919. Published 2019 Feb 27. <https://doi.org/10.1371/journal.pone.0198919>.
20. AMANHI Study Group 2017. Development and validation of a simplified algorithm for neonatal gestational age assessment– protocol for the Alliance for Maternal Newborn Health Improvement (AMANHI) prospective cohort study. *J Glob Health.* 2017;7(2):021201. <https://doi.org/10.7189/jogh.07.021201>.
21. AMANHI Study Group 2017. Understanding biological mechanisms underlying adverse birth outcomes in developing countries: protocol for a prospective cohort (AMANHI bio-banking) study. *J Glob Health.* 2017;7(2):021202. <https://doi.org/10.7189/jogh.07.021202>.
22. Fawagreh K, Gaber MM, Elyan E. Random forests: from early developments to recent advancements. *Systems Sci Control Eng.* 2014;2(1):602–9. <https://doi.org/10.1080/21642583.2014.956265>.
23. Butt K, Lim KI. Guideline no. 388-determination of gestational age by ultrasound. *J Obstet Gynaecol Can.* 2019;41(10):1497–507. 3154803. <https://doi.org/10.1016/j.jogc.2019.04.010>.
24. Ohuma EO, Papageorgiou AT, Villar J, Altman DG. Estimation of gestational age in early pregnancy from crown-rump length when gestational age range is truncated: the case study of the INTERGROWTH-21st Project. *BMC Med Res Methodol.* 2013;13(1):151. <https://doi.org/10.1186/1471-2288-13-151>.
25. Papageorgiou AT, Ohuma EO, Altman DG, et al. International standards for fetal growth based on serial ultrasound measurements: the fetal growth longitudinal study of the INTERGROWTH-21st project. *Lancet.* 2014; 384(9946):869–79. [https://doi.org/10.1016/S0140-6736\(14\)61490-2](https://doi.org/10.1016/S0140-6736(14)61490-2).
26. Ochiai M, Matsushita Y, Inoue H, et al. Blood Reference Intervals for Preterm Low-Birth-Weight Infants: A Multicenter Cohort Study in Japan. *PLoS One.* 2016;11(8):e0161439. Published 2016 Aug 23. <https://doi.org/10.1371/journal.pone.0161439>.
27. Burman P. A Comparative Study of Ordinary Cross-Validation, v-Fold Cross-Validation and the Repeated Learning-Testing Methods. *Biometrika.* 1989; 76(3):503–14. <https://doi.org/10.2307/2336116>.
28. Landset S, Kshofoftaar TM, Richter AN, Hasanin T. A survey of open source tools for machine learning with big data in the Hadoop ecosystem. *J Big Data.* 2015;2(1):24. <https://doi.org/10.1186/s40537-015-0032-1>.
29. Breiman, L. Random Forests *Machine Learning* 45, 5–32 (2001). doi: <https://doi.org/10.1023/A:1010933404324>.
30. Pedregosa, et al. Scikit-learn: Machine Learning in Python. *JMLR.* 2011;12: 2825–30.
31. Chai T, Draxler RR. Root mean square error (RMSE) or mean absolute error (MAE)? –arguments against avoiding RMSE in the literature. *Geosci Model Dev.* 2014;7(3):1247–50. <https://doi.org/10.5194/gmd-7-1247-2014>.
32. DiCiccio TJ, Efron, Bradley. Bootstrap confidence intervals. *Stat Sci.* 1996; 11(3):189–228. <https://doi.org/10.1214/ss/1032280214>.
33. Canty A, Ripley BD. Boot: bootstrap R (S-plus) functions. *R Package Version.* 2020;1:3–25.
34. Zhou XH, Obuchowski NA, McClish DK. *Statistical methods in diagnostic medicine*: Wiley-Interscience; 2002. <https://doi.org/10.1002/9780470317082>.
35. Hanley JA, McNeil BJ. The meaning and use of the area under a receiver operating characteristic (ROC) curve. *Radiology.* 1982;143(1):29–36. <https://doi.org/10.1148/radiology.143.1.7063747>.
36. Shan G. Improved Confidence Intervals for the Youden Index. *PLoS One.* 2015;10(7):e0127272. <https://doi.org/10.1371/journal.pone.0127272> PMID: 26132806; PMCID: PMC4488538.
37. Platt RW, Hanley JA, Yang H. Bootstrap confidence intervals for the sensitivity of a quantitative diagnostic test. *Stat Med.* 2000;19(3):313–22. [https://doi.org/10.1002/\(SICI\)1097-0258\(20000215\)19:3<313::AID-SIM370>3.0.CO;2-K](https://doi.org/10.1002/(SICI)1097-0258(20000215)19:3<313::AID-SIM370>3.0.CO;2-K).
38. Efron and Tibshirani. *An introduction to the bootstrap*. London: Chapman & Hall; 1993. p. 436.
39. De Long et al. 1988. Comparing the Areas under Two or More Correlated Receiver Operating Characteristic Curves: A Nonparametric Approach. *Biometrics.* Vol. 44, No. 3 (Sep., 1988).
40. Hanley JA, McNeil BJ. A method of comparing the areas under receiver operating characteristic curves derived from the same cases. *Radiology.* 1983;148(3):839–43. <https://doi.org/10.1148/radiology.148.3.6878708>.
41. Farag TH, Koplan JP, Breiman RF, Madhi SA, Heaton PM, Mundel T, et al. Precisely Tracking Childhood Death. *Am J Trop Med Hyg.* 2017;97(1):3–5. <https://doi.org/10.4269/ajtmh.16-0302> PMID: 28719334; PMCID: PMC5508885.
42. Parikh LI, Grantz KL, Iqbal SN, Huang CC, Landy HJ, Fries MH, et al. Neonatal outcomes in fetuses with cardiac anomalies and the impact of delivery route. *Am J ObstetGynecol.* 2017;217(4):469.e1–469.e12. <https://doi.org/10.1016/j.ajog.2017.05.049> Epub 2017 May 31. PMID: 28578168; PMCID: PMC5793863.
43. Boyle EM, Poulsen G, Field DJ, Kurinczuk JJ, Wolke D, Alfirevic Z, et al. Effects of gestational age at birth on health outcomes at 3 and 5 years of age: population based cohort study. *BMJ.* 2012;344:e896. <https://doi.org/10.1136/bmj.e896> PMID: 22381676; PMCID: PMC3291750.
44. Campbell OM, Cegolon L, Macleod D, Benova L. Length of Stay After Childbirth in 92 Countries and Associated Factors in 30 Low- and Middle-Income Countries: Compilation of Reported Data and a Cross-sectional Analysis from Nationally Representative Surveys. *PLoS Med.* 2016;13(3): e1001972. <https://doi.org/10.1371/journal.pmed.1001972> PMID: 26954561; PMCID: PMC478307.
45. Blume JE, Manning WC, Troiano G, Hornburg D, Figa M, Hesterberg L, et al. Rapid, deep and precise profiling of the plasma proteome with multi-nanoparticle protein corona. *Nat Commun.* 2020;11(1):3662. <https://doi.org/10.1038/s41467-020-17033-7>.

46. WHO. Every Newborn Action Plan: WHO; 2014. p. 978924150744843. [http://www.un.org.proxy1.library.jhu.edu/sustainabledevelopment/summit/United Nations . Sustainable Development Goals. United Nations; New York: 2015. \(accessed Sept 11, 2015\)](http://www.un.org.proxy1.library.jhu.edu/sustainabledevelopment/summit/United Nations . Sustainable Development Goals. United Nations; New York: 2015. (accessed Sept 11, 2015))

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