

On top of the game? The double-edged sword of incorporating social features into freemium products

Joost Rietveld  | Joe N. Ploog 

UCL School of Management, University College London, London, UK

Correspondence

Joost Rietveld, UCL School of Management, University College London, Level 38, One Canada Square, Canary Wharf, London E14 5AA, UK.
Email: j.rietveld@ucl.ac.uk

Abstract

Research Summary: Freemium products require widespread diffusion for their success. One way to do this is by incorporating social features (e.g., multiplayer functionality, virtual collaboration, ridesharing), which can generate network effects and result in a product becoming a *superstar*. However, social features can be a double-edged sword: When demand potential for freemium products is large, social features can significantly boost a product's appeal resulting in more adoption, more usage, and more in-app purchases; but when demand potential is constrained, network effects might fall short and users may feel they are missing out on key aspects of the product. We test this dynamic on a sample of 9,700 digital games on Steam. Findings contribute to our understanding of network effects, freemium strategies, and superstar products in platform markets.

Managerial Summary: Freemium has become a popular business model among firms competing on digital platforms. Freemium products require widespread diffusion because most consumers do not pay for premium upgrades. One way to stimulate a product's diffusion is by incorporating social features (e.g., multiplayer functionality, virtual collaboration, ridesharing). Social

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features can boost a product's appeal resulting in more adoption, more usage, and more in-app purchases. Our analysis of 9,700 digital PC games on Steam reveals that the efficacy of incorporating social features importantly depends on the number of users on the platform itself. Social features can help freemium products become a *superstar* when the platform's installed base is large, but they hinder a freemium product's success when the platform's installed base is small.

KEYWORDS

business models, freemium, network effects, platform markets, product strategy

1 | INTRODUCTION

The freemium business model—where a firm offers a base product for free and users can pay for premium content and features after they have adopted the base product—has gained widespread popularity on digital platforms (Rietveld, 2018; Tidhar & Eisenhardt, 2020). The share of freemium apps on Apple's App Store, for example, increased from 25% of all apps in 2009 to over 75% in 2018. Freemium's popularity is further signified by a handful of extremely successful products, including the telecommunications program *Skype* and the online dating application *Tinder*. Users can download and use these products free of charge and have the option to make in-app purchases, including *Skype*'s credit for making calls to mobile phones and landlines and *Tinder*'s Super Likes for signaling interest to romantic partners.

Freemium products require widespread diffusion for their success: Only a small portion of freemium users spends money on premium upgrades, and there exists substantial variation in the amount users spend. Freemium products must thus attract a disproportionately large user base compared to paid products to generate revenues and capture value.

Firms that operate the freemium business model therefore need to devise strategies to maximize their products' diffusion. One such strategy is to incorporate social product features, such as multiplayer functionality in video games, virtual collaboration tools in productivity software, and carpooling in ride-hailing applications. Social features can enhance social referral and stimulate a product's diffusion by adding network functionality to a product's standalone functionality (Cabral, Salant, & Woroch, 1999; Lee & O'Connor, 2003). Products that incorporate social features—and which also manage to attract a large user base—generate network effects (Aral & Walker, 2011; Dou, Niculescu, & Wu, 2013). Freemium's low barriers to adoption paired with strong network effects from social features can create a virtuous cycle resulting in a product's widespread diffusion, which may ultimately lead to a product becoming a “superstar” (Parakhonyak & Vikander, 2019; Shi, Zhang, & Srinivasan, 2019).

Incorporating social product features, however, could be a double-edged sword. When demand potential for freemium products is large (Lilien & Yoon, 1990), social features can indeed significantly boost a product's appeal, resulting in more adoption, more usage, and more in-app purchases. On the other hand, when demand potential for freemium products is

constrained—because, for example, a product is launched on a platform with a small installed base—network effects likely fall short and users may feel they are missing out on key aspects of the product that are integral to its value proposition. In this case, users derive more benefits from products that are less reliant on network functionality and focus more on standalone functionality. Thus, social product features raise the stakes: They might both increase and decrease the chances of a product becoming a superstar, depending on the market a product is released in. Given the uncertain results of incorporating social features, we ask: *How and when do social product features affect the likelihood of a freemium product becoming a superstar?*

We explore this question by analyzing a sample of 9,700 products released between 2011 and 2016 on Steam; the market-leading distribution platform for digital PC games. Freemium—or, free-to-play—and paid games compete side-by-side on Steam. Free-to-play games generate revenues exclusively from in-app purchases (e.g., cosmetic enhancements, additional content, etc.) and represent about 10% of all observations. Social product features include whether a game can be played jointly by multiple players simultaneously (e.g., online multiplayer, local cooperative play, cross-platform multiplayer). A game's demand potential is defined by the number of registered users on Steam—the platform's installed base—at the time of release. After controlling for a firm's decision to operate the freemium business model and the timing of a game's market launch, we find support for our arguments: When the platform's installed base is large, free-to-play games with many social features have a 49 percentage points higher probability of becoming a superstar than free-to-play games without any social features, whereas when the platform's installed base is small, free-to-play games with many social features have a 26 percentage points lower probability of becoming a superstar.

We further argue and find that the mixed effects of incorporating social features and a product's demand potential on becoming a superstar are specific to freemium products (i.e., do not apply to paid products). Freemium products enjoy stronger social referral than paid products because consumers are more inclined to recommend products that exhibit low risks to adoption, such as those that are free (Bond, He, & Wen, 2019; Lin, Zhang, & Tan, 2019). Freemium products additionally have different use dynamics than paid products. Paid products exhibit strong lock-in mechanisms given that paying consumers “want to get their money's worth”. This means that freemium products will diffuse more quickly, but also that freemium users are less engaged than paying users (Bapna, Ramaprasad, & Umyarov, 2018; Rietveld, 2018). Taken together, these considerations suggest that the effect of a product's demand potential on generating network effects from social features is stronger for freemium products.

Our study aims to make three contributions. First, we offer two important insights about network effects. One of these insights—which requires relaxing the common assumption that network effects are a market-level construct (e.g., McIntyre & Srinivasan, 2017; Schilling, 2002; Suarez, 2004; Zhu & Iansiti, 2012)—is that firms can add network functionality to their products by incorporating social features (also see: Aral & Walker, 2011; Dou et al., 2013). Competing products might thus vary in strength of network effects as a function of their user base *and* their design features (Shankar & Bayus, 2003). The other insight is that social features are not unequivocally associated with a product's superior performance. When a product's demand potential is limited, firms will, in fact, decrease their chances of becoming a superstar if they overly depend on network functionality. Combined, these insights imply that firms must think strategically about network effects: Products can be designed to have stronger or weaker network effects, but the efficacy of these choices depends on external factors.

Second, we show how and when freemium products can take advantage from network effects. We document that incorporating social features to boost a freemium product's diffusion

is especially effective on platforms with a larger installed base, while it is detrimental on platforms with a smaller installed base. In so doing, we contribute to our understanding of when the freemium business model works (Kumar, 2014), and how freemium strategies differ from strategies for paid products (Bapna et al., 2018; Bond et al., 2019; Eckhardt, 2016; Lee & Csaszar, 2020; Pauwels & Weiss, 2008; Rietveld, 2018; Shi et al., 2019; Tidhar & Eisenhardt, 2020).¹ Our research relates to work by Boudreau, Jeppesen, and Miric (2020) who study how a change in market-level network effects impacts the performance of freemium versus paid market leaders. We complement their work by identifying product design features as an important—and strategic—predictor of which freemium products become market leaders.

Third, we add to a growing literature on superstar products in platform settings. Superstars are the very best-performing products which enjoy exponentially superior performance (e.g., downloads or revenues) compared to the products they compete with (e.g., Benner & Waldfogel, 2020; Brynjolfsson, Hu, & Smith, 2010; Rosen, 1981).² Scholars are increasingly interested in superstar products given their important role in digital platforms, both as drivers of an ecosystem's overall value (Binken & Stremersch, 2009; Gretz, Malshe, Bauer, & Basuroy, 2019) and as drivers of value for products' commercializing firms (Cox, 2014; Yin, Davis, & Muzrya, 2014). We propose a novel measure for operationalizing superstar products in platform markets that accounts for variation in a product's demand potential and for the extent of competition a product faces.

2 | FREEMIUM AND SOCIAL PRODUCT FEATURES

Facilitated by the Internet, the freemium business model has gained widespread popularity among firms competing on digital platforms including mobile app stores and video game consoles.³ Freemium departs from the traditional paid model by introducing novel transaction structures and novel transaction content (Amit & Zott, 2001; Rietveld, 2018). First, instead of offering a complete product, freemium products offer users a menu of paid items in the form of in-app purchases (*product bundle decomposition*). In the freemium video game *Fortnite*, for example, players can purchase cosmetic items such as virtual clothing and accessories for their avatars as well as unlock entire game modes through in-app purchases. Second, users can download and perpetually use freemium products before making any in-app purchases (*temporal decoupling*). For example, *Fortnite* players can play the base game for as long as they want before potentially committing to any of its premium content. Firms operating the freemium model must therefore develop distinct capabilities in such areas as user engagement, data analytics, price menu design, and product life cycle management (Kumar, 2014; Lee &

¹Related literatures in marketing and information systems look at feature-limited software applications that act as a free trial or sampling instrument for paid software applications. These literatures study whether offering a free trial version benefits the paid application, and what makes this more effective or less (e.g., Arora, ter Hofstede, & Mahajan, 2017; Cheng & Liu, 2012; Gu, Kannan, & Ma, 2018; Lee, Zhang, & Wedel, 2021).

²The literature has referred to such products interchangeably as “superstars,” “blockbusters,” and “killer apps”. We follow the predominant convention by using the term superstars (see Online Appendix A for a review).

³The term “freemium” was coined in 2006 by Fred Wilson, who used it to describe a novel business model where firms “Give your service away for free, possibly ad supported but maybe not, acquire a lot of customers very efficiently through word of mouth, referral networks, organic search marketing, and so forth, then offer premium priced value added services or an enhanced version of your service to your customer base.” See: https://avc.com/2006/03/my_favorite_bus/; https://avc.com/2006/03/the_freemium_bu/ (accessed February, 2021).

Csaszar, 2020; Tidhar & Eisenhardt, 2020). That is, the activities required for successfully developing and commercializing freemium products are distinct from paid products.

There are three main reasons why freemium products require more widespread diffusion than paid products to generate revenues and capture value. First, because freemium products can be adopted free of charge, the user base of freemium products is often characterized by substantial demand heterogeneity. A lack of adoption barriers entices users with a wide range of willingness-to-pay to download and use freemium products. Second, product bundle decomposition allows heterogeneous users to mix-and-match a combination of premium items that closely reflects their willingness-to-pay. This implies that the majority of freemium users will spend very little—if anything—on in-app purchases while a fraction of avid users (so-called “whales”) spends large amounts on freemium products. In the video game industry, for example, it is well-known that between two and 5% of a freemium game’s player base spends money on in-app purchases (Luton, 2013; Seufert, 2013). Third, temporally decoupling a product’s use from optional payments for premium content or features negatively impacts users’ perception of value (Datta, Foubert, & van Heerde, 2015; Gourville & Soman, 1998). Put differently, keeping all else equal, users are willing to pay more for the same product bundle when it is commercialized with a paid business model than a freemium business model.

To overcome these challenges and stimulate a product’s diffusion, firms often incorporate social features into freemium products. Social product features enable user interactions, which, when present, generate value in use. These features range from multiplayer functionality in video games to online collaborative tools in productivity software and carpooling functionality in ride-hailing apps. In *Fortnite’s* Battle Royale mode, for example, up to 100 players can play both cooperatively and competitively to accomplish a common objective (e.g., be the last players standing). If a product manages to attract a large user base, social features add network value to a product’s standalone value (Cabral et al., 1999; Lee & O’Connor, 2003). Such products generate network effects, where a user’s benefits increase with the total number of users of the same product (Farrell & Saloner, 1986; Katz & Shapiro, 1985). Products that exhibit strong network effects have increased chances of becoming a superstar (Cennamo & Santalo, 2013; Shankar & Bayus, 2003; Suarez, 2004). Indeed, the network effects literature has long asserted that a single product or technology can end up dominating an entire market in the presence of network effects.

Incorporating social product features is particularly beneficial for freemium products, because freemium products enjoy stronger social referral through word-of-mouth than paid products (Cheng & Liu, 2012; Shi et al., 2019). Consumers are more inclined to recommend products that exhibit low risks to adoption, such as those products that are free to use, and they are also more likely to reciprocate any benefits that they receive for free by endorsing a firm’s products (Bond et al., 2019; Lin et al., 2019). In sum, freemium products are advantageously positioned to generate network effects by incorporating social product features because they enjoy lower barriers to adoption and stronger social referral compared to paid products.

3 | HYPOTHESES

3.1 | The mixed effects of incorporating social features on becoming a superstar

Superstars are the top-performing products in a market, which enjoy exponentially superior performance—often expressed in terms of downloads or revenues—compared to the products

they compete with (Rosen, 1981). Superstar products are highly salient in digital platforms where the distribution of performance (at the market level) tends to be skewed (Benner & Waldfogel, 2020; Brynjolfsson et al., 2010). These contexts are characterized by marked differences between the performance of the few products ranked at the very top of the market (i.e., superstars) and of the numerous products at lower performance ranks (i.e., the long tail).

Social product features can help freemium products become a superstar by setting in motion a self-reinforcing network effect. It should be noted, however, that network effects fail to materialize in the absence of a large user base. That is, a product's network value, or the network externalities it generates, is a function of its social features *and* the size of its active user base (Shankar & Bayus, 2003). Given that products on digital platforms can only be adopted by those users who have first adopted the platform itself (i.e., the platform's installed base), we argue that freemium products will be more likely to benefit from incorporating social features when they are launched on a platform with a large (rather than a small) installed base.

Products launched on a platform with a small installed base are constrained in their demand potential (Lilien & Yoon, 1990), whereas products launched on a platform with a large installed base can potentially reach a much wider audience (Rietveld & Eggers, 2018). Even though all products launched at the same time can be offered to an equisized installed base, the extent of a product's value creation through social features will be greater the larger the platform's installed base. That is, the likelihood of a product becoming a superstar from incorporating social features will be higher when the product's demand potential is larger.

Consider the stylized example of a consumer who is deciding between two products—one that relies exclusively on network functionality for its value proposition (i.e., many social features) and another that relies exclusively on standalone functionality (i.e., no social features). If the consumer is the only one that has adopted the platform these products are launched on (i.e., the platform has an installed base of one), she will anticipate no benefits from social features. In this case, the consumer will be more inclined to adopt the product without social features, anticipating (greater) benefits. When the consumer considers the same two products but this time on a platform that has been adopted by other users (i.e., it has an installed base greater than one), she may expect to derive some benefits from the product's social features. This is true even if she does not know the exact size of the product's user base, given that consumers often *anticipate* the size of a product's user base in the presence of network effects (e.g., Farrell & Saloner, 1986; Parakhonyak & Vikander, 2019; Schilling, 2003).

Note that our stylized example holds even when we allow a product to offer any combination of standalone functionality and network functionality, so long as there is some (either perceived or actual) tradeoff between a firm's investments in standalone functionality and its investments in social features. In the aggregate, it is intuitive that a product's demand potential can hinder or help a product's market performance if it relies on network functionality for its value proposition: When a product's demand potential is large, consumers may anticipate the product to have a larger user base, which makes the inclusion of social features valuable to some extent. When a product's demand potential is constrained, however, consumers will anticipate a smaller user base, which makes the inclusion of social features less beneficial or even detrimental to the extent that it compromises the product's value proposition.

Social features therefore can be a double-edged sword: If firms are to fully exploit the benefits conferred by the freemium business model, they must carefully consider their products' demand potential. On the one hand, when a platform has a small installed base, freemium products have limited potential to create value from social features. In this case, users will anticipate greater benefits from those products that fully depend on standalone

functionality, which will offer benefits in the absence of a (large) user base. On the other hand, when a platform does boast a large installed base, freemium products will be in an opportune position to take advantage from network functionality. In this case, users will derive greater benefits from freemium products that incorporate social features than from those that do not. The combination of freemium's social referral *and* potentially strong network effects generated by social features can boost a product's diffusion to the point where it becomes a superstar⁴:

Hypothesis (H1). *Platform installed base size moderates the effect of incorporating social features on becoming a superstar, such that freemium products with many social features will be more likely to become a superstar when the platform's installed base is large.*

3.2 | Freemium versus paid superstars

There are two reasons why this dynamic is specific to freemium products (i.e., does not apply to paid products). First, the potential benefits of a large demand potential are more pronounced. Low barriers to adoption and strong social referral allow freemium products to diffuse more quickly and more widely than paid products. A product that incorporates social features, has low barriers to adoption *and* strong social referral can quickly capture a large share of the market (Cheng & Liu, 2012; Shi et al., 2019). This happened, for instance, with the free-to-play game *Dota 2*. Released in 2013 by publisher Valve, *Dota 2* quickly became the all-time most downloaded game on Steam with over 112 million downloads. *Dota 2* was highly rated among gamers and its online multiplayer and cooperative play modes generated strong network effects, which set in motion a virtuous cycle further amplifying the game's popularity.

Second, there exist qualitative differences between how freemium and paid products are used. First, given that consumers must spend money before they can use a paid product, only those users who *ex ante* anticipate sufficient benefits will adopt a paid product (Rietveld, 2018). Second, upfront payments for paid products create a sunk-cost effect wherein consumers want to “get their money's worth” (Arkes & Blumer, 1985; Staw, 1976). Paying users will thus be more committed to fully experiencing all of a product's benefits (Bapna et al., 2018). Finally, users perceive comparatively greater benefits from paid products given that paid products' payment and use are conjoined—instead of temporally decoupled (Datta et al., 2015; Gourville & Soman, 1998). Combined, these differences suggest that use rates will be significantly higher for paid products than for freemium products. It also suggests that consumers will be less inclined to adopt paid products solely on the basis of social referral.

Taken together, these differences between paying users and freemium users suggest that paid products diffuse less quickly, but also that paying users are more engaged than freemium users. The strength of paid products' network effects therefore will depend less on the size of their user base than on the (larger) amount of time paying users spend consuming these products. That is, paid products' higher average use rates will render the platform's installed base size less of a contingent factor in creating value from social features:

⁴We do not hypothesize about the main effect of incorporating social features because we do not expect a main effect given the strong contingency of a freemium product's demand potential on becoming a superstar.

Hypothesis (H2). *The interaction between platform installed base size and a product's social features on becoming a superstar is specific to freemium products; it will not apply to paid products.*

4 | DATA SAMPLE AND MEASURES

We test our hypotheses in the context of Steam, the market-leading platform for digitally distributed PC games for the Windows, Mac, and Linux operating systems. Steam was created in 2003 by game publisher Valve as a platform initially for the maintenance and distribution of its internally developed PC games *Counterstrike* and *Half-Life 2*. Shortly after Steam was launched, however, Valve recognized an opportunity as the PC gaming industry underwent a resurgence and started developing tools to facilitate third-party game developers in offering their products on the platform. The first externally developed PC games on Steam were released in 2005, and the number of games has grown exponentially since. By 2016, Steam listed over 10,000 games—the vast majority of which (>99%) were released by external developers.

Developers that wish to release their games on Steam must design their software to be compatible with Steamworks, Valve's proprietary software development kit. In 2011, Valve added the in-game purchasing application programming interface (API) to Steamworks, allowing developers to forgo charging an upfront fee for their games and monetize in-game components. The introduction of this API enabled the freemium business model on Steam. Free-to-play games quickly followed and have been among the most successful games on the platform since. Examples of popular free-to-play games include *Dota 2* (Valve, 2013); *Paladins* (Hi-Rez Studios, 2016); and *Heroes and Generals* (Reto-Moto, 2016). These games have all been downloaded more than 10 million times and generate revenues from in-game purchases.

4.1 | Data

Data were collected primarily from two sources. First, we collected game-level downloads data for every game released between 2005 and 2017 from Steam Spy, an online analytics service that uses Valve's web-based API. Every minute, Steam Spy culls from the API a random sample of user profiles and obtains lists of games these users own. Linking this information to the number of registered users, Steam Spy extrapolates ownership statistics for each game. Although Steam Spy provides estimates rather than exact downloads statistics, game developers were not allowed to publicly disclose these figures and the industry has largely relied on Steam Spy for accessing Steam performance data. Steam Spy's margin of error is less than 10%, and game developers regularly confirm the accuracy of Steam Spy's estimates for their games.⁵ In addition to downloads statistics, we also collected data on games' release dates, whether games are free-to-play or paid, and games' publishers. We further collected data on the number of registered Steam users for every year that we observe data in our sample.

⁵See, for example: <https://www.pcgamesn.com/steam/steam-spy-accuracy-developers> (accessed March, 2021). We note that since we collected our data, Valve implemented a number of important changes to its API (in April 2018), and also updated its terms for developers (in November 2018). As a result of these changes, developers may now disclose downloads data. Furthermore, citing GDPR legislation, Valve restricted its API functionality, which has significantly reduced the accuracy of Steam Spy's downloads estimates after November 2018.

Our second data source is Valve's public web-based API. Using web scraping techniques, we requested game-level data on various aspects directly from Steam. These data include the type and number of social features embedded in a game (discussed below), a game's genre(s), the type of publisher and its prior experience on Steam, and several technical elements such as system requirements and usage statistics that we use for robustness checks.

The Steam Spy data also contain information on game quality as curated by Metacritic.com. Metacritic is a publicly accessible expert review database that collects, combines, weighs, and transforms expert review scores from over 180 online and offline publications (at the time of data collection). For each game, Metacritic publishes a so-called Metascore, which reflects a weighted average of all expert review scores, ranging from 0 to 100 (100 indicating a perfect score).⁶ Metacritic assigns different colors to its Metascores to distinguish between “good” Metascores (green; ranging from 75 to 100), “average” Metascores (orange; ranging from 50 to 74), and “bad” Metascores (red; scores below 50). Metascores are a good proxy for game quality given the aggregated and independent nature of these data.

4.2 | Estimation sample

We start our estimation sample by considering all games released on Steam from 2011 to 2016. We begin the sample in 2011 when Valve introduced the in-game purchasing API, and we end the sample in 2016 to allow all games at least 1 year to accumulate downloads. We exclude 142 observations that are non-game software (mostly software development tools), 119 observations for which we do not observe any information at the publisher level, and 514 observations that are compilations, demo versions, add-on packages released as standalone products, or games that were removed from Steam after our data collection period. Our final sample for analysis includes 9,700 games, of which 771 games operate the freemium model.

Table 1 provides an overview for some of the key measures in our estimation sample. The distribution of downloads is heavily skewed as superstar games (defined below) generate on average 2.6 million cumulative downloads per game whereas non-superstar games generate 107,000 downloads per game. The asymmetry in downloads between superstars and non-superstars has widened over time.⁷ The share of free-to-play games also grew during our study time-period from 0.05 in 2011 to 0.10 in 2016. The share of superstars within the sub-sample of free-to-play games is 0.16. The number of social features per game (defined below) declined. In 2011, free-to-play games had an average of 1.58 social features, whereas in 2016 this had declined to 0.81 social features. We observe a similar trend for paid games. Further exploration of the data suggests that this decline is largely driven by a slight absolute decrease in games incorporating local multiplayer functionality and by an increase of low-quality games entering the platform without any social features. We control for these trends through our identification strategy and our various robustness checks. Steam's installed base grew exponentially from more than 38 million registered users in 2011 to nearly 223 million registered users in 2016.

⁶For more information, see: <https://www.metacritic.com/about-metascores> (accessed March, 2021).

⁷Downloads for superstar games in 2013 are distorted by the release of Valve's *Dota 2*, the all-time most downloaded (free-to-play) game on Steam with more than 112 million downloads.

TABLE 1 Game and platform statistics by year

| Year | 2011 | 2012 | 2013 | 2014 | 2015 | 2016 | All |
|--------------------------------|------------|------------|------------|-------------|-------------|-------------|-------------|
| Games released | 256 | 332 | 462 | 1,609 | 2,661 | 4,380 | 9,700 |
| Free-to-play games | 12 | 21 | 28 | 81 | 190 | 439 | 771 |
| Share of superstars in | | | | | | | |
| <i>Free-to-play</i> | 0.25 | 0.19 | 0.18 | 0.36 | 0.22 | 0.09 | 0.16 |
| <i>Paid</i> | 0.07 | 0.05 | 0.00 | 0.04 | 0.02 | 0.02 | 0.03 |
| Mean downloads | | | | | | | |
| <i>Superstars</i> | 4,598,066 | 5,859,613 | 31,996,919 | 1,735,460 | 1,633,832 | 1,470,115 | 2,633,332 |
| <i>Non-superstars</i> | 366,269 | 403,848 | 472,266 | 94,787 | 56,840 | 38,543 | 106,677 |
| Mean number of social features | | | | | | | |
| <i>Free-to-play</i> | 1.58 | 1.43 | 1.46 | 1.11 | 0.82 | 0.81 | 0.90 |
| <i>Paid</i> | 0.49 | 0.39 | 0.46 | 0.30 | 0.31 | 0.33 | 0.33 |
| Platform statistics | | | | | | | |
| <i>New platform users</i> | 10,708,000 | 13,747,000 | 25,579,000 | 36,915,000 | 47,682,000 | 60,317,000 | 32,491,333 |
| <i>Installed base</i> | 38,738,000 | 52,485,000 | 78,064,000 | 114,979,000 | 162,661,000 | 222,978,000 | 222,978,000 |

Note: Based on estimation sample.

4.3 | Variable definitions

4.3.1 | Dependent variable

We follow prior studies on superstar products in platform markets by applying a performance-based cutoff threshold to distinguish superstar products from non-superstar products (Binken & Stremersch, 2009; Cox, 2014; Ershov, 2020; Gretz et al., 2019; Lee, 2013; Sun, Rajiv, & Chu, 2016; Yin et al., 2014). In deciding on an appropriate cutoff value, however, the researcher faces at least three challenges. First, products face different levels of competition at different points in a platform's life. A product released at the start of a platform's life might face only a handful of relevant competitors, whereas a product released towards the end of a platform's life may face thousands (Boudreau, 2012). The same applies to a product's demand potential (i.e., the installed base): Users may be reluctant to adopt a new platform when it first launches, whereas once a platform establishes itself as the dominant design, it may command a large share of the overall consumer base in a market (Boudreau & Jeppesen, 2015). Second, the distribution of demand changes over time such that the gap between the top- and bottom-performing products widens as a platform matures (Rietveld & Eggers, 2018; Rietveld, Ploog, & Nieborg, 2020). Finally, there often exists variation in the amount of time products can accumulate performance—especially when analyzing cross-sectional product-level data. This implies that a product's performance as observed by the researcher will be highly contextual.⁸

We aim to address these challenges by creating a standardized measure of a game's download performance based on the subsample of all games released in the same year as a focal game. We treat each year as a separate market and calculate a game's z-score such that:

$$z_i = \frac{x_i - \bar{x}_t}{S_t}$$

where z_i represents game i 's standardized performance as a function of the difference between its own cumulative downloads (x_i) and the mean cumulative downloads of those games released in the same year t as game i (\bar{x}_t), divided by the standard deviation of all games released in the same year t as game i (S_t). A game's z-score is contingent on a subset of games that were released around the same time and thus face similar conditions in terms of installed base size and competition dynamics. Nevertheless, because z-scores are standardized, we can meaningfully compare the z-score of a game released in 2016 to one released in 2011.

Next, we determine a game's status as a superstar by applying the following cutoff:

$$Superstar_i = \begin{cases} 1, & z_i \geq 1 \\ 0, & z_i < 1 \end{cases}$$

where the variable *superstar* takes the value of 1 if game i has a z-score (z_i) equal to or greater than 1, and 0 otherwise. It should be noted that games' downloads are not normally distributed and that a z-score of 1 does not correspond with the standard normal cumulative density function (see Online Appendix B). Rather, 3.58% of all games are coded superstars ($n = 347$). This

⁸Our review of the literature suggests that not all prior studies have fully considered these challenges (see Appendix A in the Online Supplement).

proportion is consistent with prior work on superstar products in platform markets (e.g., Gretz et al., 2019; Lee, 2013; Sun et al., 2016). Furthermore, research on breakthrough innovations and blockbuster products such as patents and drugs either found or applied similar thresholds to denote outlier performance (e.g., Kaplan & Vakili, 2015; Kneeland, Schilling, & Aharonson, 2020). Our measure for superstar products thus is generally representative, while also giving us sufficient power for statistical analysis. We assess the sensitivity of our results by estimating various alternative measures in Section 6.2.

4.3.2 | Independent variables

First, *social features* measures the extent that a game can be played by more than one player simultaneously. Some games on Steam are designed exclusively for single-player experiences whereas others provide multiplayer functionality. Among multiplayer games, three further distinctions are worth mentioning. First, the potential pool of players for any multiplayer game depends on the type of connectivity a game offers, which can be either local—between players using the same computer or those connected via a local area network—or online. Second, games differ in whether multiplayer functionality is meant to be cooperative, competitive, or a combination of the two. Finally, while most multiplayer games can only be played by players on the same operating system, some can be played across operating systems, and sometimes even on different platforms altogether (e.g., PlayStation 3, Xbox 360). These three dimensions of multiplayer functionality are independent of one another; developers can vary each dimension as they see fit.

We collected data on whether a game includes any of the following social features at launch: local cooperative play, online cooperative play, local competitive play, online competitive play, and cross-platform multiplayer. Table 2 provides distributions from our sample, broken out by free-to-play and paid games. While all features are equally present across both subsamples, free-to-play games, on average, have a higher number of multiplayer modes. We measure a game's social features by counting the number of multiplayer modes it offers.⁹ The assumption is that games offer greater network functionality when they provide more ways for users to play together via different multiplayer modes. Our results are robust to various alternative measurements, including a dummy variable indicating whether a game has any social features as well as excluding local multiplayer from the count-based measure.

Second, since we observe data from one platform only, we exploit temporal variation in genre popularity to construct a measure of installed base size at the genre-year level. Specifically, for each year in our data we calculate a genre's market share by dividing the sum of downloads for all games released in a genre by the sum of downloads for all games released that year. We then multiply these market shares by the number of yearly new Steam adopters (in millions) to arrive at a measure of *genre installed base*. Thus, for each year in our data, we observe the number of new platform adopters per genre based on a genre's relative popularity.

⁹We abstain from weighting the different multiplayer modes since we do not observe the exact number of players a multiplayer game can accommodate. Moreover, prior research on network structure found that users often exhibit local bias, suggesting that network effects are stronger when users have stronger ties to each other (Afuah, 2013; Lee, Lee, & Lee, 2006; Suarez, 2005). Players that play together locally likely will have stronger ties than players that play together in an online setting. Additionally, while local multiplayer on the same computer will involve a limited number of players, the same does not necessarily hold for local multiplayer via LANs, which Steam lumps into a single category. Finally, cooperative play is not by definition restricted to two players only, and can, in fact, include coordination between large groups of players in online settings.

TABLE 2 Distribution of social features by type and count

| Social features (type) | Free-to-play games | Paid games | Net difference (%) | |
|----------------------------|--------------------|------------|--------------------|-------|
| Local cooperative play | | 115 | 628 | 7.88 |
| Online cooperative play | | 50 | 196 | 4.29 |
| Local competitive play | | 288 | 1,395 | 21.73 |
| Online competitive play | | 118 | 367 | 11.19 |
| Cross-platform multiplayer | | 120 | 391 | 11.19 |

| Social features (count) | Free-to-play games | Paid games | Net difference (%) | |
|-------------------------|--------------------|------------|--------------------|--------|
| 0 | | 428 | 7,328 | -26.56 |
| 1 | | 122 | 730 | 7.65 |
| 2 | | 139 | 539 | 11.99 |
| 3 | | 50 | 202 | 4.22 |
| 4 | | 19 | 87 | 1.49 |
| 5 | | 13 | 43 | 1.20 |

Note: Based on estimation sample. Local play includes multiplayer functionality on the same PC as well as over multiple PCs connected to the same LAN. Cross-platform multiplayer facilitates online multiplayer functionality between Steam accounts on different operating systems (i.e., Windows, Mac, and Linux), and sometimes between non-Steam accounts (e.g., video game consoles or other PC platforms).

Abbreviation: LAN, local area network.

Since there is extensive research showing that the number of (superstar) products positively impacts the diffusion of the platform itself (e.g., Binken & Stremersch, 2009; Clements & Ohashi, 2005), we lag our measure by 1 year to avoid issues of reverse causality.

We chose to calculate *genre installed base* using yearly new platform adoption and yearly genre market shares (as opposed to cumulative statistics) for two main reasons. First, there exists significant variation in terms of which genres are popular at different times (also see: Ozalp & Kretschmer, 2019). Second, cumulative installed base measures tend to overstate a product's demand potential given that users adopt most products shortly after joining a platform (Nair, Chintagunta, & Dube, 2004; Tellis, Yin, & Niraj, 2009). That said, our results are fully robust to using a measure that is based on cumulative statistics as well as to using measures based on platform-wide new adoption and Steam's cumulative installed base.

For (H1) to be supported, we expect the interaction between *social features* and *genre installed base* to be positive for the subsample of free-to-play games. For (H2) to be supported, we expect the interaction between *social features* and *genre installed base* to have a stronger positive effect on free-to-play games than paid games, for which we expect no effect.

4.3.3 | Control variables

We include several control variables at the platform, publisher, and game levels. Though the overall effect of competitive crowding in multisided platforms is ambiguous given the positive spillover effect of product variety on platform diffusion (Parker & van Alstyne, 2005), it is well-established that entry by similar products can have a negative effect on the performance of a focal product (Boudreau, 2012)—especially when rivals enter the platform around the same time (Rietveld & Eggers, 2018). We therefore control for competitive crowding by including the variable *genre competition*. For each game i , *genre competition* counts the number of newly released games within the same genre(s) as game i , from 30 days before to 30 days after the game's release, divided by the number of genres game i lists on Steam. We apply this timeframe because games typically have very short lifecycles and generate the bulk of their downloads and revenues shortly after release (Nair, 2007). This measure can be interpreted as the mean competition a game faces across the genres it competes in; we expect it to have a negative effect on the probability of becoming a superstar.

We include two measures to account for heterogeneous capabilities and resources at the publisher level. First, we control for publisher type by including the variable *indie publisher*. The industry broadly distinguishes between two types of publishers. Independent—or *indie*—publishers are smaller in size, focus their development efforts on creative and innovative output, and tend to have less (financial) resources. Incumbent publishers, on the other hand, are larger, focus on exploiting established intellectual properties, and are typically flush with resources, financial and otherwise (also see: Benner & Waldfoegel, 2016). The variable *indie publisher* takes the value of 1 for games by indie publishers and 0 for games by incumbent publishers. Second, we control for publishers' prior experience on Steam. Not all publishers embraced Steam when it was first launched, whereas others are Steam specialists. The variable *past releases publisher* counts the number of games a publisher launched on Steam over a rolling window of 5 years dating back from a focal game's release. We chose a rolling window rather than publishers' cumulative experience because prior experience may become obsolete due to the dynamic and evolving nature of the platform. We log-transform the measure to account for the skew in our data. We expect *indie publisher* to have a negative effect on the probability of becoming a superstar and *past releases publisher* to have a positive effect.

Finally, we control for several game-level factors. First, we control for game quality by including Metacritic's Metascore in our models. We distinguish between games with good Metascores, games with average Metascores, and games with bad Metascores. We include Metacritic's review classification as a vector of dummies and omit as the base category games with bad Metascores and games with missing review scores, the latter of which typically denotes very poor quality. We thus expect both included dummies to have a positive effect on the probability of becoming a superstar. Second, we control for a game's listed genres given that players often have heterogeneous preferences for different types of games. On Steam, games can list one or more of the following genres: *Action*, *Adventure*, *Casual*, *Massively Multiplayer*, *Racing*, *Role Playing Game*, *Simulation*, *Sports*, and *Strategy*. We include all of these as dummy variables in our models. Third, we control for seasonality by including 11 calendar month-of-release fixed effects and exclude *January* as the base category. Finally, we control for macrolevel and platform-level trends (e.g., changing consumer preferences, increasing technological requirements, updates, etc.) by including year-of-release dummies.

5 | METHODS

Since we rely on archival data for our study and cannot take advantage of some quasi-exogenous shock, we are confronted with a potential endogeneity problem: The existence of unobserved factors that are correlated with the *free-to-play* variable and with our outcome variable of becoming a *superstar*. Structural differences between developers of free-to-play games and developers of paid games may bias our results (Rietveld, 2018; Tidhar & Eisenhardt, 2020). Moreover, shrewd developers may refrain from releasing free-to-play games with social features when Steam's installed base is small, while being more inclined to do so when the installed base is large. Though we cannot fully rule out these potential concerns, we take several precautionary steps to account for the choice of business model as well as for the potentially endogenous timing of game launches.

We control for the choice of business model by fitting a treatment effects model in which both the treatment and the outcome are binary, also known as a recursive bivariate probit model (Greene, 2018). This model is akin to the Heckman two-step control function, such that:

$$\text{Outcome equation : } y_i = x_i\beta + w_i\delta + \varepsilon_i,$$

$$\text{Treatment equation: } w_i^* = z_i\gamma + u_i, w_i = 1 \text{ if } w_i^* > 0, \text{ and } w_i = 0 \text{ otherwise}$$

$$\text{Prob}(w_i = 1 | z_i) = \Phi(z_i\gamma)$$

and

$$\text{Prob}(w_i = 0 | z_i) = 1 - \Phi(z_i\gamma)$$

where x_i is a vector of exogenous variables determining a binary outcome y_i (i.e., *superstar*), and w_i is an endogenous dummy variable indicating the treatment condition (i.e., *free-to-play*). Contrary to the Heckman two-step control function model, however, the outcome y_i is observed for both $w_i = 1$ (i.e., *free-to-play*) and $w_i = 0$ (i.e., *paid*). This allows us to conduct a Chow test on the interactions between *social features* and *genre installed base* to assess their equality (and test (H2)). Notably, there need not be an exclusion restriction for recursive bivariate probit models to be identified—granted the exogenous variables provide sufficient variation.¹⁰

We include several covariates in our treatment equation. Since we expect some genres to be a better fit for the freemium business model than others, we include the vector of genre dummies. We additionally include the *indie publisher* and *past releases publisher* variables to control for variation at the publisher level. At the platform level, we control for *genre installed base* at the time of a game's release. We further include a variable that counts the number of *past freemium superstars* measured over a rolling window of 3 years before a game's release up to 1 year before a game's release, to reflect a typical video game development cycle. We expect that, at the time a game goes into production, publishers will be guided by the extant success of

¹⁰Formal proofs for why exclusion restrictions are not strictly required in recursive bivariate probit models are beyond the scope of this paper. The general intuition is that the identification of such models relies on the nonlinearity of the function and variation in the derivatives of the probability that $y_i = 1$ with respect to the covariates in x_i and z_i . The treatment correction hazard (λ_i) is linearly independent of x_i , even if $z_i = 1$; all that is required is that there is variation in x_i and in z_i (William Greene, personal communication, March 2019).

the freemium model on Steam in deciding whether their games should be free-to-play or paid. Finally, we add year-of-release dummies to account for platform-level time trends. First-stage results are reported in Appendix C in the Online Supplement.

The treatment effects model accounts for the nonrandom assignment of games into the free-to-play and paid conditions. The treatment correction hazard (λ_i), which is included in the outcome equation, balances against differences in genre and year of release, the type of publisher, the size of the installed base, and the success of the freemium business model.

We take an additional step to reduce imbalance in the empirical distribution of our covariates (x_i) by applying a coarsened exact matching (CEM) algorithm (Iacus, King, & Porro, 2012). Based on a set of matching covariates, CEM prunes observations so that the remaining data have a better balance between the treatment and the control conditions. For each free-to-play game i , the algorithm finds at least one paid game that is similar on the following covariates: *social features*, *genre installed base*, *indie publisher*, *past releases publisher*, and *genre*. Thus, for each free-to-play game, the algorithm finds at least one same-genre paid game with an equivalent number of social features that is released around the same time by a comparable publisher. After pruning 4,641 observations, the imbalance (L_1) in our estimation sample reduces from 0.66 to 0.06. The CEM algorithm assigns weights based on the number of control group observations for each treated observation, which we use in our models to further improve the quality of our inferences (Blackwell, Iacus, King, & Porro, 2009).

In sum, while our results are correlational, we attempt to mitigate endogeneity concerns that are both structural (i.e., choice of business model) and time varying (i.e., release timing of games on Steam) by fitting a treatment effects model on a matched and rebalanced sample.

6 | RESULTS

6.1 | Main results

Table 3 lists descriptive statistics and pairwise correlations for our covariates. Our main results are reported in Table 4. Model 1 includes control variables only. Model 2 adds independent variables (*social features* and *genre installed base*). Model 3 adds the interaction between *social features* and *genre installed base*, testing H1. Model 4 controls for the nonrandom (treatment) assignment into *free-to-play*. Model 5 prunes and rebalances the sample by applying the CEM

TABLE 3 Descriptive statistics and pairwise correlations

| Variable | Mean | SD | Min | Max | 1 | 2 | 3 | 4 | 5 | 6 |
|---|--------|--------|------|--------|-------|-------|-------|-------|-------|-------|
| 1 <i>Superstar</i> | 0.04 | 0.19 | 0.00 | 1.00 | | | | | | |
| 2 <i>Free-to-play</i> | 0.08 | 0.27 | 0.00 | 1.00 | 0.19 | | | | | |
| 3 <i>Social features</i> | 0.38 | 0.88 | 0.00 | 5.00 | 0.17 | 0.17 | | | | |
| 4 <i>Genre installed base_{t-1}</i> | 4.57 | 3.23 | 0.00 | 10.66 | 0.02 | 0.07 | 0.05 | | | |
| 5 <i>Genre competition</i> | 159.90 | 100.59 | 1.00 | 456.00 | -0.08 | -0.11 | -0.06 | 0.52 | | |
| 6 <i>Indie publisher</i> | 0.67 | 0.47 | 0.00 | 1.00 | -0.10 | -0.07 | -0.03 | 0.26 | 0.21 | |
| 7 <i>ln(Past releases publisher)</i> | 1.05 | 1.32 | 0.00 | 4.82 | 0.05 | -0.13 | -0.06 | -0.12 | -0.08 | -0.37 |

Note: Based on estimation sample ($n = 9,700$). Mean VIF = 2.54.
Abbreviation: VIF, variance inflation factor.

TABLE 4 Regression estimates of games' likelihood of becoming a superstar

| Variable | Superstar | | | | | | Test of 5 ≠ 6 |
|---|----------------|----------------|----------------|----------------|----------------|----------------|---------------|
| | 1 | 2 | 3 | 4 | 5 | 6 | |
| <i>Social features</i> | | | | | | | |
| <i>Genre installed base_{t-1}</i> | | 0.21 [0.05] | 0.07 [0.09] | 0.06 [0.09] | -0.22 [0.15] | 0.006 [0.13] | 1.28 |
| <i>Social features × genre installed base_{t-1}</i> | | 0.002 [0.019] | -0.03 [0.03] | -0.08 [0.06] | -0.38 [0.13] | 0.07 [0.04] | 10.24 |
| <i>Genre competition</i> | -0.004 [0.001] | -0.004 [0.001] | 0.02 [0.01] | 0.03 [0.01] | 0.10 [0.03] | 0.02 [0.02] | 6.32 |
| <i>Indie publisher</i> | -0.24 [0.14] | -0.23 [0.14] | -0.004 [0.001] | -0.004 [0.001] | -0.006 [0.002] | -0.001 [0.001] | 3.40 |
| <i>ln(Past releases publisher)</i> | 0.13 [0.07] | 0.15 [0.07] | -0.23 [0.14] | 0.11 [0.37] | 1.70 [0.87] | -0.22 [0.32] | 4.31 |
| Quality dummies (2) | Yes | Yes | Yes | Yes | Yes | -0.16 [0.14] | 7.59 |
| Genre dummies (9) | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Month of release dummies (11) | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Year of release dummies (5) | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Endogenous treatment correction | No | No | No | Yes | Yes | Yes | Yes |
| Matched and rebalanced sample | No | No | No | No | Yes | Yes | Yes |
| Constant | -1.04 [0.29] | -1.14 [0.31] | -0.99 [0.32] | 0.47 [1.48] | 5.95 [3.28] | -2.10 [0.70] | 6.85 |
| McFadden's <i>pseudo R</i> ² | .22 | .23 | .24 | .24 | | | |
| Observations | 771 | 771 | 771 | 771 | 456 | 4,603 | |

Note: Heteroscedasticity robust SEs in parentheses. Models 1–3 report stepwise results from probit regressions on the subsample of free-to-play games. Model 4 estimates an endogenous treatment effects model on the subsample of free-to-play games. First-stage results are reported in Appendix C in the Online Supplement. Models 5 and 6 estimate results on a pruned and rebalanced sample via coarsened exact matching, on the subsamples of free-to-play and paid games, respectively. The error terms in Models 5 and 6 are jointly estimated. Model 7 reports Chow test statistics estimating equality of coefficients between coefficients in Models 5 and 6.

algorithm. Model 6 estimates the fully-specified model on the matched subsample of *paid* games. (Error terms in Models 5 and 6 are jointly estimated to facilitate cross-model hypothesis tests.) Model 7 tests (H2) by comparing regression coefficients across both models.

Results lend support to our hypotheses. Consistent with H1, the interaction between *social features* and *installed base* in Model 5 is positive and significant ($\beta = .102, p = .000$). This suggests that free-to-play games with many social features have a better chance of becoming a superstar when the platform's installed base is large. Given that the interpretation of interaction effects in nonlinear models is complicated, we obtain the marginal effects on *superstar* at different values of *social features* and *genre installed base* (Hoetker, 2007; Zelner, 2009). Figure 1 depicts the margins slopes for games with five social features and games without any social features at different values of the installed base. The figure shows that when the installed base is large, free-to-play games with many social features have a 49 percentage points higher probability of becoming a superstar than free-to-play games without any social features, whereas when the platform's installed base is small, free-to-play games with many social features have a 26 percentage points lower probability of becoming a superstar.

To test (H2), we take the coefficient on the interaction between *social features* and *genre installed base* for the subsample of free-to-play games and compare it to the same coefficient for the subsample of paid games. To do so, we conduct a Chow test to assess the equality of coefficients (Chow, 1960). The interaction between *social features* and *genre installed base* for the subsample of paid games is positive but not statistically significant ($\beta = .021, p = .226$). Moreover, the Chi-square test statistic is significant ($\chi^2 = 6.32, p = .012$), confirming that the coefficients on the interaction effects are statistically different across both subsamples. The right-hand margins plot in Figure 1 further confirms the absence of an interaction effect for paid games. Notably, free-to-play games with many social features that are released on a platform with a large installed base are 37 percentage points more likely to become a superstar than paid games with many social features released on a platform with a similar installed base.

To interpret our control variables we estimate a model on the matched and rebalanced subsample of free-to-play games that does not include any interactions or variable transformations (Wiersema & Bowen, 2009). We find that one additional social feature is associated with a 4.36 percentage points increase in the probability of becoming a superstar ($p = .001$). *Genre installed base* has no significant effect on becoming a superstar. One hundred additional same-genre

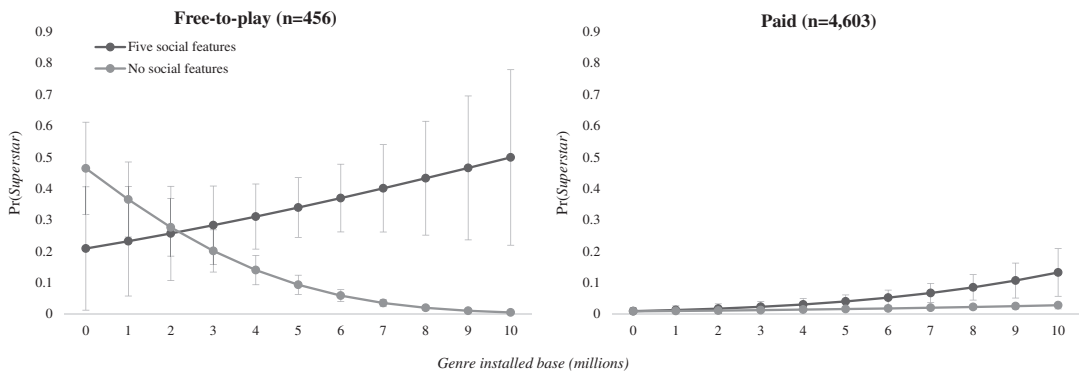


FIGURE 1 Predicted probability plots based on subsample analyses. Predicted probability plots for free-to-play and paid video games with many (5) and without any social features based on split-sample probit regressions reported in Models 5 and 6 in Table 4

games (i.e., *genre competition*) reduces the probability of becoming a superstar by 7.79 percentage points ($p = .018$). Publisher type (i.e., *indie publisher*) has no effect on becoming a superstar. Adding 10 games to a publisher's past releases is associated with a 3.75 percentage points higher probability of becoming a superstar ($p = .081$). Free-to-play games with good Metascores have a 29 percentage points higher probability of becoming a superstar than games with bad or missing Metascores ($p = .082$). *Action* games are 13 percentage points more likely to become a superstar than any of the other genres ($p = .086$).

6.2 | Robustness tests and mechanism checks

We took several steps to mitigate potential endogeneity concerns: First, we constructed a novel measure for superstar products that accounts for supply- and demand-side variation as well as for time trends by standardizing performance for products released in different circumstances. Second, to control for structural differences between free-to-play and paid games we first estimated an endogenous treatment effects model before estimating the equation of interest. Third, we pruned and rebalanced the estimation sample via CEM to further reduce imbalance in our data. Our hypotheses remain fully supported after these steps. Nevertheless, we conduct several additional robustness checks to further rule out potential alternative explanations.¹¹

First, one might argue that variation in installed base size cannot be fully separated from other potentially relevant time-varying factors. Consequently, the argument goes, what determines which games become superstars likely changes over time. We run an additional set of tests to further check for this. First, because fluctuations in the average number of social features may affect the relative value of incorporating these features, we control for either the average number of social features per game at the genre-year level, or for the share of games with any social features. Second, to account for time-varying demand heterogeneity, we include the median number of games owned by different user cohorts based on the year of joining Steam. Consistent with Rietveld and Eggers (2018), we note that the median number of games owned decreases over time. Third, we add games' system requirements in the form of random-access memory and hard disk space (measured in gigabytes) to account for increasing technological demands, which could affect the type of games users download (Ghose & Han, 2014). In all these alternative specifications, our arguments remain supported.

Second, we construct three alternative measures for determining whether a game is a superstar. First, instead of basing a game's z-score on the subsample of all games released in the same year, we create a more granular measure at the level of a game's genre(s). Computing separate z-scores for each genre a game competes in, we denote a game as a superstar when the statistical average of these genre-based z-scores is equal to or greater than 1. Second, to rule out the possibility that our findings are specific to our z-score-based performance measure, we estimate our results using measures that indicate whether a game is among the top 5% most downloaded games in a year or genre. Results are consistent with our reported main results.

Third, we run several supplementary analyses for our social features variable. The first alternative measure operationalization takes the value of 1 if a game has any social features and 0 if a game has no social features. The second measure pools online and local multiplayer modes, restricting the count to three. The third drops local multiplayer functionality from the count, focusing exclusively on the various online multiplayer modes. The fourth takes the

¹¹Fully tabulated results from these robustness checks are reported in Appendix C in the Online Supplement.

log-transformation of the number of social features to account for the skewness in the data. Furthermore, we add control variables indicating whether games include leaderboards and achievements as examples of social features that do not materially alter a game's value proposition. Results from all these specifications are fully consistent with our main results.

Fourth, we run several supplementary analyses for our installed base measure. First, instead of operationalizing the genre installed base measure using yearly statistics, we recreate the measure using cumulative data (from the start of the platform up to a focal year). Second, instead of using a genre-specific measure, we estimate platform-wide measures. Relatedly, we create a measure of the platform's yearly growth rate and control for platform age (in months). Year-of-release fixed effects are absorbed in these robustness checks, and multicollinearity is an issue in some of the models. That said, our arguments remain supported.

Fifth, one of the mechanisms driving (H2) is consumers spending less time on freemium products than paid products. To generate network effects, freemium products therefore require a disproportionately larger user base (i.e., to offset lower engagement rates). To check for this, we estimate results on the log of games' median playing time, or the median number of minutes per user playing a game, as dependent variable. Results from a linear regression with endogenous treatment effects suggests that free-to-play games indeed have disproportionately lower use rates ($\beta = -4.769$; $p = .000$; also see: Rietveld, 2018). In fact, we note that the median playing time for free-to-play games in our estimation sample is 54 min while it is 270 min for paid games. Exponentiating the regression coefficient suggests that free-to-play games have a 99 percentage points lower median playing time (per user) than paid games. These findings on playing time are consistent with the suggested mechanisms for our hypotheses.

Finally, we estimate the fully specified model on the pooled sample of free-to-play and paid games and test the three-way interaction between *free-to-play*, *social features*, and *genre installed base*. Our findings are fully robust; the three-way interaction effect is positive and statistically significant ($\beta = .069$, $p = .030$). We additionally find support for our hypotheses using a rare events logit estimator (King & Zeng, 2001). We conclude that our results are robust to various alternative model specifications and measurement operationalizations.

7 | DISCUSSION AND CONCLUSIONS

Recognizing the need for widespread diffusion and the potentially mixed effects of incorporating social features into freemium products, we asked: *How and when do social product features affect the likelihood of a freemium product becoming a superstar?* To answer this question, we analyzed a sample of 9,700 digital PC games launched on Steam between 2011 and 2016. Results from an endogenous treatment effects model estimated on a matched and rebalanced sample show that, when the platform's installed base is large, free-to-play games with many social features have a 49 percentage points higher probability of becoming a superstar than free-to-play games without any social features. Conversely, when the platform's installed base is small, free-to-play games with many social features have a 26 percentage points lower probability of becoming a superstar. Furthermore, we find that the contingent effect of incorporating social features is specific to free-to-play games; the effect does not apply to paid games. Our study makes several contributions and holds implications for future research on network effects, freemium strategies, and superstar products in platform markets.

First, our study contributes to the literature on network effects. By relaxing the common assumption in the strategy literature that network effects are a market-level construct, we

explored how firms can increase the strength of network effects through their product design choices. By incorporating multiplayer modes into freemium games, the publishers in our sample added network functionality to products' standalone functionality, thereby increasing their chances of attaining market-leading performance. The notion that firms can alter the strength of their products' network effects has previously been examined in the fields of marketing and information systems (e.g., Aral & Walker, 2011; Dou et al., 2013), but remains largely unexplored in strategic management (for an exception, see: McIntyre & Subramaniam, 2009). This lack of attention is surprising given the myriad implications *tractable network effects* can have on such issues as market entry timing, product innovation management, and competitive dynamics. Allowing for product-level variation in the strength of network effects, for example, may shed new light on how new entrants can successfully compete with dominant incumbents or when winner-take-all dynamics are likely to occur.

However, designing products around network effects is no panacea. The efficacy of this strategy importantly depends on external factors, including a product's demand potential. Our results suggest that firms can increase the chances of their products becoming a superstar if they incorporate social features when demand potential is large, whereas they decrease the chances of their products becoming a superstar when demand potential is constrained. This finding is consistent with Srinivasan, Lilien, and Rangaswamy (2004), who found that network effects have a negative impact on the survival rate of pioneering products—those products that are first to enter a nascent market when demand is still limited. We conclude that there exist contingencies as to when firms should rely more on network functionality versus standalone functionality for their products' value proposition (Cabral et al., 1999; Lee & O'Connor, 2003). Future research should explore how other factors, such as demand heterogeneity or competitive dynamics, impact the efficacy of incorporating social product features.

Second, we contribute to our understanding of how and when the freemium business model works (Kumar, 2014). There exist qualitative differences between freemium and paid products—including usage patterns and diffusion dynamics—that have implications for firms operating the freemium business model. On the one hand, freemium products diffuse more quickly and more widely than paid products, due to their lower barriers to adoption and stronger word-of-mouth dynamics (Bond et al., 2019; Lin et al., 2019). On the other hand, freemium consumers have a lower willingness-to-pay and they are less engaged than paying consumers (Bapna et al., 2018; Rietveld, 2018). Given that the strength of any product's network effects is a function of the product's network functionality *and* its active user base (Shankar & Bayus, 2003), incorporating social features will therefore only benefit those freemium products that manage to attract a disproportionately large user base. Thus, while prior literature has addressed *why* firms should adopt a freemium business model in the presence of network effects (Cheng & Liu, 2012; Shi et al., 2019), we add by documenting *how* firms operating the freemium business model can design their products to optimally benefit from network effects.

Third, we contribute a novel measure for operationalizing superstars in platform settings. Many digital platforms are characterized by markedly skewed demand distributions (e.g., Benner & Waldfogel, 2020). When this is the case, there exists a strong demarcation between superstar products and (less successful) products that reside in the long tail of these platforms (Brynjolfsson et al., 2010; Elberse, 2008). What exactly constitutes a superstar, however, is a moving target. Considering platforms' two-sidedness and products' embedded nature within these markets, the size of the market for such products—both on the demand side and on the supply side—is subject to constant variation. As a result, applying a fixed cutoff threshold based on a product's downloads or its sales ranking may yield inconsistent results. To

account for this, we proposed a measure based on a product's standardized performance relative to similar products released around the same time on the platform. We hope our z-score-based measure for superstar products will benefit scholars conducting research in similar settings.

Our study also has some limitations that must be acknowledged. Most notably, we only have data on one platform. While this enabled us to develop a deep understanding of games on Steam, as well as to construct measures and conduct analyses that are highly contextual, we cannot unequivocally ascertain that our results are caused by variation in Steam's installed base. Neither can we assure the generalizability of our findings to other digital platforms or different types of social features. We encourage colleagues to duplicate and extend our findings in related settings such as mobile app stores. The share of freemium products on Apple's App Store and Google's Play Store, for example, is much higher. Developers on these platforms also often operate hybrid business models, where income is generated from both consumers and advertisers (Casadesus-Masanell & Zhu, 2010, 2013). Non-gaming apps further incorporate different types of social features such as ride-sharing functionality and virtual collaborative tools.

How do our findings replicate to these different platforms and social features? A study by Boudreau et al. (2020) documents dynamics similar to ours in the context of Apple's App Store: An increase in market-level network effects benefitted freemium market leaders more than paid market leaders. Thus, in line with our arguments, they document that network effects are more beneficial to freemium than paid superstars. Their study differs, however, in its focus on market-level network effects, and it does not explore the antecedents of market leadership.

A potential boundary condition of our work is the extent that a platform's life cycle is finite. Platforms tied to a specialized technology product, such as a video game console or a smartphone, are often replaced by a next-generation platform after some time (e.g., Kretschmer & Claussen, 2016). Consumers may be tempted to abandon the focal platform in lieu of a next-generation platform (e.g., Kim & Srinivasan, 2009), causing a contraction of the installed base—and with it products' demand potential. Thus, in the presence of generational breaks, a platform's installed base size positively impacts the inclusion of social product features *to the point that consumers won't anticipate the introduction of a next-generation platform*.

We also believe it will be important to gain better understanding about the life cycle dynamics of freemium products. At what point do freemium superstars fail to retain and engage their users, why? These questions are pertinent given that consumers seem to lose interest in these products virtually overnight—as witnessed by *H1Z1*, a freemium superstar game on Steam that was abruptly dethroned by *PlayerUnknown's Battleground*, another freemium game in the same genre (also see: Lee et al., 2021). Similarly, it should be interesting to investigate differences between social features. For example, in our data, we observed a decline in the number of games incorporating local multiplayer modes. While this trend can be explained by increased connectivity among users and by firms targeting larger user bases via online multiplayer, firms may exercise some caution. Prior research found that users often exhibit local bias, meaning that network effects are stronger when users have stronger ties to each other (Afuah, 2013; Lee et al., 2006; Suarez, 2005). Local multiplayer functionality may therefore not accommodate the same scale of connectedness, but it will likely result in stronger engagement, which could offset the smaller user base. More work is needed on the tradeoff between user base growth and engagement dynamics for network effects (also see: Claussen, Kretschmer, & Mayrhofer, 2013). Finally, scholars may want to study value capture strategies for freemium products. Little is known still about the effective design of freemium price menus (for exceptions, see: Meng, Hao, & Tan, 2021; Tidhar & Eisenhardt, 2020). This is a complex issue at the intersection of product design and consumer psychology. Archival data are unlikely to generate conclusive evidence, which may be better studied using experiments or machine learning techniques.

The freemium business model has become the leading business model on many digital platforms. Some of the most popular products today are commercialized with a freemium model. These products often incorporate social features such as multiplayer functionality in video games, virtual collaboration tools in productivity software, and carpooling in ride-hailing applications. Social features add network functionality to a product's standalone functionality, and they can increase the strength of network effects. Incorporating social features, however, is a double-edged sword: On the one hand, when a freemium product's demand potential is constrained, network effects will not materialize and users may feel they are missing out. On the other hand, when a freemium product's demand potential is large, social product features can set in motion a virtuous cycle of more adoption, more usage, and more in-app purchases.

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DATA AVAILABILITY STATEMENT

The data that support the findings of this study are available from the corresponding author upon reasonable request.

ORCID

Joost Rietveld  <https://orcid.org/0000-0001-8722-4442>

Joe N. Ploog  <https://orcid.org/0000-0002-8895-5228>

REFERENCES

- Afuah, A. (2013). Are network effects really all about size? The role of structure and conduct. *Strategic Management Journal*, 34(3), 257–273.
- Amit, R., & Zott, C. (2001). Value creation in e-business. *Strategic Management Journal*, 22(6–7), 493–520.
- Aral, S., & Walker, D. (2011). Creating social contagion through viral product design: A randomized trial of peer influence in networks. *Management Science*, 57(9), 1623–1639.
- Arkes, H. R., & Blumer, C. (1985). The psychology of sunk cost. *Organizational Behavior and Human Decision Processes*, 35(1), 124–140.
- Arora, S., ter Hofstede, F., & Mahajan, V. (2017). The implications of offering free versions for the performance of paid mobile apps. *Journal of Marketing*, 81(6), 62–78.
- Bapna, R., Ramaprasad, J., & Umyarov, A. (2018). Monetizing freemium communities: Does paying for premium increase social engagement? *MIS Quarterly*, 42(3), 719–736.
- Benner, M. J., & Waldfoegel, J. (2016). The song remains the same? Technological change and positioning in the recorded music industry. *Strategy Science*, 1(3), 129–147.

- Benner, M. J., & Waldfogel, J. (2020). Changing the channel: Digitization and the rise of “middle tail” strategies. *Strategic Management Journal*.
- Binken, J. L., & Stremersch, S. (2009). The effect of superstar software on hardware sales in system markets. *Journal of Marketing*, 73(2), 88–104.
- Blackwell, M., Iacus, S., King, G., & Porro, G. (2009). Cem: Coarsened exact matching in stata. *The Stata Journal*, 9(4), 524–546.
- Bond, S. D., He, S. X., & Wen, W. (2019). Speaking for “free”: Word of mouth in free- and paid-product settings. *Journal of Marketing Research*, 33(1), 1–15.
- Boudreau, K. J. (2012). Let a thousand flowers bloom? An early look at large numbers of software app developers and patterns of innovation. *Organization Science*, 23(5), 1409–1427.
- Boudreau, K. J., & Jeppesen, L. B. (2015). Unpaid crowd complementors: The platform network effect mirage. *Strategic Management Journal*, 36(12), 1761–1777.
- Boudreau, K. J., Jeppesen, L. B., & Miric, M. (2020). *Competing on freemium: Digital competition with network effects* (Working Paper). Available at SSRN: <https://ssrn.com/abstract=2984546>
- Brynjolfsson, E., Hu, Y., & Smith, M. D. (2010). Research commentary—Long tails vs. superstars: The effect of information technology on product variety and sales concentration patterns. *Information Systems Research*, 21(4), 736–747.
- Cabral, L. M., Salant, D. J., & Wroch, G. A. (1999). Monopoly pricing with network externalities. *International Journal of Industrial Organization*, 17, 199–214.
- Casadesus-Masanell, R., & Zhu, F. (2010). Strategies to fight ad-sponsored rivals. *Management Science*, 56(9), 1484–1499.
- Casadesus-Masanell, R., & Zhu, F. (2013). Business model innovation and competitive imitation: The case of sponsor-based business models. *Strategic Management Journal*, 34(4), 464–482.
- Cennamo, C., & Santalo, J. (2013). Platform competition: Strategic trade-offs in platform markets. *Strategic Management Journal*, 34(11), 1331–1350.
- Cheng, H. K., & Liu, Y. (2012). Optimal software free trial strategy: The impact of network externalities and consumer uncertainty. *Information Systems Research*, 23(2), 488–504.
- Chow, G. C. (1960). Tests of equality between sets of coefficients in two linear regressions. *Econometrica*, 28(3), 591–605.
- Claussen, J., Kretschmer, T., & Mayrhofer, P. (2013). The effects of rewarding user engagement: The case of facebook apps. *Information Systems Research*, 24(1), 186–200.
- Clements, M. T., & Ohashi, H. (2005). Indirect network effects and the product cycle: Video games in the U.S., 1994–2002. *The Journal of Industrial Economics*, 53(4), 515–542.
- Cox, J. (2014). What makes a blockbuster video game? An empirical analysis of US sales data. *Managerial and Decision Economics*, 35(3), 189–198.
- Datta, H., Foubert, B., & van Heerde, H. J. (2015). The challenge of retaining customers acquired with free trials. *Journal of Marketing Research*, 52(2), 217–234.
- Dou, Y., Niculescu, M. F., & Wu, D. J. (2013). Engineering optimal network effects via social media features and seeding in markets for digital goods and services. *Information Systems Research*, 24(1), 164–185.
- Eckhardt, J. T. (2016). Welcome contributor or no price competitor? The competitive interaction of free and priced technologies. *Strategic Management Journal*, 37(4), 742–762.
- Elberse, A. (2008). Should you invest in the long tail? *Harvard Business Review*, 86(7/8), 88.
- Ershov, D. (2020). *Competing with superstars in the mobile app market* (Working Paper).
- Farrell, J., & Saloner, G. (1986). Installed base and compatibility: Innovation, product preannouncements, and predation. *The American Economic Review*, 76(5), 940–955.
- Ghose, A., & Han, S. P. (2014). Estimating demand for mobile applications in the new economy. *Management Science*, 60(6), 1470–1488.
- Gourville, J. T., & Soman, D. (1998). Payment depreciation: The behavioral effects of temporally separating payments from consumption. *Journal of Consumer Research*, 25(2), 160–174.
- Greene, W. H. (2018). *Econometric analysis* (8th ed.). New York, NY: Pearson.
- Gretz, R. T., Malshe, A., Bauer, C., & Basuroy, S. (2019). The impact of superstar and non-superstar software on hardware sales: The moderating role of hardware lifecycle. *Journal of the Academy of Marketing Science*, 47(3), 394–416.

- Gu, X., Kannan, P. K., & Ma, L. (2018). Selling the premium in freemium. *Journal of Marketing*, 82(6), 10–27.
- Hoetker, G. (2007). The use of logit and probit models in strategic management research: Critical issues. *Strategic Management Journal*, 28(4), 331–343.
- Iacus, S. M., King, G., & Porro, G. (2012). Causal inference without balance checking: Coarsened exact matching. *Political Analysis*, 20(1), 1–24.
- Kaplan, S., & Vakili, K. (2015). The double-edged sword of recombination in breakthrough innovation. *Strategic Management Journal*, 36(10), 1435–1457.
- Katz, M. L., & Shapiro, C. (1985). Network externalities, competition, and compatibility. *The American Economic Review*, 75(3), 424–440.
- Kim, S. H., & Srinivasan, V. (2009). A conjoint-hazard model of the timing of buyers' upgrading to improved versions of high-technology products. *Journal of Product Innovation Management*, 26(3), 278–290.
- King, G., & Zeng, L. (2001). Logistic regression in rare events data. *Political Analysis*, 9(2), 137–163.
- Kneeland, M. K., Schilling, M. A., & Aharonson, B. S. (2020). Exploring uncharted territory: Knowledge search processes in the origination of outlier innovation. *Organization Science*, 31(3), 535–557.
- Kretschmer, T., & Claussen, J. (2016). Generational transitions in platform markets—The role of backward compatibility. *Strategy Science*, 1(2), 90–104.
- Kumar, V. (2014). Making "freemium" work: Many start-ups fail to recognize the challenges of this popular business model. *Harvard Business Review*, 92(5), 27–30.
- Lee, E., Lee, J., & Lee, J. (2006). Reconsideration of the winner-take-all hypothesis: Complex networks and local bias. *Management Science*, 52(12), 1838–1848.
- Lee, R. S. (2013). Vertical integration and exclusivity in platform and two-sided markets. *The American Economic Review*, 103(7), 2960–3000.
- Lee, S., & Csaszar, F. A. (2020). Cognitive and structural antecedents of innovation: A large-sample study. *Strategy Science*, 5(2), 71–97.
- Lee, S., Zhang, J., & Wedel, M. (2021). Managing the versioning decision over an app's lifetime. *Journal of Marketing*, 85(6), 44–62.
- Lee, Y., & O'Connor, G. C. (2003). New product launch strategy for network effects products. *Journal of the Academy of Marketing Science*, 31(3), 241–255.
- Lilien, G. L., & Yoon, E. (1990). The timing of competitive market entry: An exploratory study of new industrial products. *Management Science*, 36(5), 568–585.
- Lin, Z., Zhang, Y., & Tan, Y. (2019). An empirical study of free product sampling and rating bias. *Information Systems Research*, 30(1), 260–275.
- Luton, W. (2013). *Free-to-play: Making money from games you give away*. London: New Riders.
- McIntyre, D. P., & Srinivasan, A. (2017). Networks, platforms, and strategy: Emerging views and next steps. *Strategic Management Journal*, 38(1), 141–160.
- McIntyre, D. P., & Subramaniam, M. (2009). Strategy in network industries: A review and research agenda. *Journal of Management*, 35(6), 1494–1517.
- Meng, Z., Hao, L., & Tan, Y. (2021). Freemium pricing in digital games with virtual currency. *Information Systems Research*, 32(2), 481–496.
- Nair, H. (2007). Intertemporal price discrimination with forward-looking consumers: Application to the US market for console video-games. *Quantitative Marketing and Economics*, 5(3), 239–292.
- Nair, H., Chintagunta, P. K., & Dube, J.-P. H. (2004). Empirical analysis of indirect network effects in the market for personal digital assistants. *Quantitative Marketing and Economics*, 2(1), 23–58.
- Ozalp, H., & Kretschmer, T. (2019). Follow the crowd or follow the trailblazer? The differential role of firm experience in product entry decisions in the US video game industry. *Journal of Management Studies*, 56(7), 1452–1481.
- Parakhonyak, A., & Vikander, N. (2019). Optimal sales schemes for network goods. *Management Science*, 65(2), 819–841.
- Parker, G. G., & van Alstyne, M. W. (2005). Two-sided network effects: A theory of information product design. *Management Science*, 51(10), 1494–1504.
- Pauwels, K., & Weiss, A. (2008). Moving from free to fee: How online firms market to change their business model successfully. *Journal of Marketing*, 73(3), 14–31.

- Rietveld, J. (2018). Creating and capturing value from freemium business models: A demand-side perspective. *Strategic Entrepreneurship Journal*, 12(2), 171–193.
- Rietveld, J., & Eggers, J. P. (2018). Demand heterogeneity in platform markets: Implications for complementors. *Organization Science*, 29(2), 304–322.
- Rietveld, J., Ploog, J. N., & Nieborg, D. B. (2020). Coevolution of platform dominance and governance strategies: Effects on complementor performance outcomes. *Academy of Management Discoveries*, 6(3), 488–513.
- Rosen, S. (1981). The economics of superstars. *The American Economic Review*, 71(5), 845–858.
- Schilling, M. A. (2002). Technology success and failure in winner-take-all markets: The impact of learning orientation, timing, and network externalities. *Academy of Management Journal*, 45(2), 387–398.
- Schilling, M. A. (2003). Technological leapfrogging: Lessons from the U.S. video game console industry. *California Management Review*, 45(3), 6–32.
- Seufert, E. B. (2013). *Freemium economics: Leveraging analytics and user segmentation to drive revenue*. Amsterdam: Elsevier Science.
- Shankar, V., & Bayus, B. L. (2003). Network effects and competition: An empirical analysis of the home video game industry. *Strategic Management Journal*, 24(4), 375–384.
- Shi, Z., Zhang, K., & Srinivasan, K. (2019). Freemium as an optimal strategy for market dominant firms. *Marketing Science*, 38(1), 150–169.
- Srinivasan, R., Lilien, G. L., & Rangaswamy, A. (2004). First in, first out? The effects of network externalities on pioneer survival. *Journal of Marketing*, 68(1), 41–58.
- Staw, B. M. (1976). Knee-deep in the big muddy: A study of escalating commitment to a chosen course of action. *Organizational Behavior and Human Performance*, 16(1), 27–44.
- Suarez, F. F. (2004). Battles for technological dominance: An integrative framework. *Research Policy*, 33(2), 271–286.
- Suarez, F. F. (2005). Network effects revisited: The role of strong ties in technology selection. *Academy of Management Journal*, 48(4), 710–720.
- Sun, L., Rajiv, S., & Chu, J. (2016). Beyond the more the merrier: The variety effect and consumer heterogeneity in system markets. *International Journal of Research in Marketing*, 33(2), 261–275.
- Tellis, G. J., Yin, E., & Niraj, R. (2009). Does quality win? Network effects versus quality in high-tech markets. *Journal of Marketing Research*, 46(2), 135–149.
- Tidhar, R., & Eisenhardt, K. M. (2020). Get rich or die trying... Finding revenue model fit using machine learning and multiple cases. *Strategic Management Journal*, 41(7), 1245–1273.
- Wiersema, M. F., & Bowen, H. P. (2009). The use of limited dependent variable techniques in strategy research: Issues and methods. *Strategic Management Journal*, 30(6), 679–692.
- Yin, P. L., Davis, J. P., & Muzyrya, Y. (2014). Entrepreneurial innovation: Killer apps in the iPhone ecosystem. *The American Economic Review*, 104(5), 255–259.
- Zelner, B. A. (2009). Using simulation to interpret results from logit, probit, and other nonlinear models. *Strategic Management Journal*, 30(12), 1335–1348.
- Zhu, F., & Iansiti, M. (2012). Entry into platform-based markets. *Strategic Management Journal*, 33(1), 88–106.

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