

1 **Prognostic models for mortality after cardiac surgery in patients with infective endocarditis: a**
2 **systematic review and aggregation of prediction models**

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48 **Abstract**

49 *Background:* Several prognostic models have been developed trying to estimate the risk of
50 mortality after surgery for active infective endocarditis (IE). However, these models
51 incorporate different predictors and their performance is uncertain.

52 *Objective:* We aimed to systematically review and critically appraise all available prediction
53 models of post-operative mortality in patients with IE, and to synthesize them into a meta-
54 model.

55 *Data sources:* We searched Medline and EMBASE databases from inception to June 2020 to
56 identify post-operative prognostic models.

57 *Study eligibility criteria:* We included studies that developed or updated a prognostic model
58 of post-operative mortality in patient with IE.

59 *Methods:* We assessed the risk of bias of the models using PROBAST (Prediction model Risk
60 Of Bias ASsessment Tool) and we synthesized them into an aggregate meta-model based on
61 stacked regressions and optimized it for a nationwide registry of IE patients. The meta-model
62 performance was assessed using bootstrap validation methods and adjusted for optimism.

63 *Results:* We identified 9 studies reporting the development of 11 prognostic models for post-
64 operative mortality. Eight models were rated as high risk of bias. The meta-model included
65 weighted predictors from the remaining three models (i.e. EndoSCORE, specific ES-I and
66 specific ES-II), which were not rated as high risk of bias and provided full model equation.
67 Additionally, two variables (i.e. age and infectious agent) which had been modeled
68 differently across studies, were estimated from scratch based on the nationwide registry. The
69 meta-model performance was better than that of initial three models, with the corresponding
70 performance measures: C-statistics 0.79 (95% CI 0.76 to 0.82), calibration slope 0.98 (95%
71 CI 0.86 to 1.13) and calibration-in-the-large -0.05 (95% CI -0.20 to 0.11).

72 *Conclusions:* The meta-model outperformed published models and showed a robust predictive
73 capacity for predicting the individualized risk of post-operative mortality in patients with IE.

74 *Protocol Registration:* PROSPERO (registration number CRD42020192602)

75 *Key words:* Prognostic models, systematic review, meta-model, aggregation, validation,
76 infective endocarditis.

77

78 **Background**

79 Infective endocarditis (IE) is an uncommon but severe disease with a high mortality rate. Its
80 current estimated incidence is 3-10 episodes per 100.000 person-years, while its in-hospital
81 mortality rate ranges between 15% and 40% (1,2). Management of IE is often complex and,
82 although indications for surgery are established in current guidelines (3), the decision whether
83 to perform surgery remains a challenge because of the high mortality rate associated with the
84 procedure. For that reason, it is estimated than less than half of the patients with surgical
85 indication finally undergo cardiac surgery (4)leading to a significant decreased chance of
86 survival (5). In this context, there has been a great interest on modeling prognosis of patients
87 with IE to accurately estimate the risk of mortality and to help in the decision-making processes.

88 In the last decade, several IE prognostic models using preoperative patient´s-related and IE-
89 specific factors, have been proposed (6). Unfortunately, these models have not been
90 implemented in guidelines or applied in clinical practice. In fact, clinicians seldomly trust these
91 models because they have usually been built in relatively small cohorts and have not been
92 externally validated. Consequently, researchers carry on developing new models from their own
93 data without considering prior knowledge, which leads to an scenario with multiple prognostic
94 models of dubious validity. Therefore, we aimed to systematically review and critically
95 appraise all available prediction models for post-operative mortality in patients with IE, and to
96 synthesize them into a meta-model based on stacked regressions.

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98 **Methods**

99 The protocol for this study was registered on PROSPERO (registration number
100 CRD42020192602). We designed this systematic review according to the recent guidance by
101 Debray et al.(7,8), and reported its results following PRISMA (Preferred Reporting Items for
102 Systematic Reviews and Meta-Analyses) (9) and TRIPOD (Transparent Reporting of a
103 Multivariable Prediction Model for Individual Prognosis or Diagnosis) (10,11)
104 recommendations.

105 *Literature search*

106 We searched Medline through Ovid and Embase through Elsevier from inception to
107 01/06/2020. We conducted a literature search to identify all potential studies for inclusion. We
108 applied no restriction considering language or publication dates. We used the methodologic
109 filter developed by Geersing et al. for prediction models research in MEDLINE (12), which
110 was adapted for EMBASE. We added terms related to cardiac surgery and
111 endocarditis. We further searched bibliographic references of included articles to
112 identify other potential eligible studies. Complete search strings are shown in

113 **Supplementary Material: S1.**

114 *Eligibility criteria*

115 We included original studies that developed prognostic models, with or without external
116 validation, to predict the risk of post-operative mortality after cardiac surgery in patient with
117 IE, as well as studies that updated previously published models. We accepted the authors`
118 definition of post-operative mortality (either 30 days and/or in-hospital mortality), but excluded
119 models that predicted mortality as part of a composite adverse outcome. Titles, abstracts, and
120 full texts were screened for eligibility in pairs by three reviewers
121 independently (BMFF, LVB, ACP) using EPPI-Reviewer 4 (13). Discrepancies were
122 resolved by consensus.

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123 *Data extraction*

124 Data extraction of included articles was done by three reviewers independently
125 (pairs from BMFF, LVB, ACP). Reviewers used a standardized data extraction form based on
126 CHARMS (CHecklist for critical Appraisal and data extraction for systematic Reviews of
127 prediction Modelling Studies) (8), and discrepancies were solved by consensus. We extracted
128 data on the following items: general information of the study, source of data, participants'
129 characteristics, outcome definition and time of occurrence, candidate predictors, and analysis
130 methods. (**Supplementary Material: S2**). When the completed model equation or relevant data
131 were not provided, we contacted the correspondence authors to require this information.

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132 *Risk of bias assessment*

133 We used a standardized form based on PROBAST (PRediction model risk of Bias ASsessment
134 Tool) (14,15) to evaluate risk of bias (RoB) and applicability. We defined the presence of RoB
135 as the existence of deficiencies in the study design or analysis that may have led to
136 systematically distorted estimates of the model performance or its composition. Concerns
137 regarding the applicability of a primary study would arise when the population, predictors, or
138 outcomes of the study differed from those specified in our review question. RoB and
139 applicability were assessed by two independent reviewers (pairs from BMFF, LVB, ACP). We
140 evaluated the relevant items on the following domains: Participants, predictors, outcome and
141 analysis. Each domain was rated according to our review question as having a *high*, *low* or
142 *unclear* RoB, and as providing *high*, *low* or *unclear* concerns regarding applicability. Any
143 discrepancies were discussed between reviewers and resolved through discussion. The
144 supplementary material provides details on critical appraisal and applicability (**Supplementary**
145 **Material: S3**).

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146 *GAMES registry*

147 We used the infective endocarditis nationwide Spanish registry (GAMES)
148 as the validation dataset, to
149 estimate existing model weights for the meta-model development and its validation, and to
150 externally validate the previously published models. Since January 2008, all consecutive
151 episodes of IE in 34 Spanish hospitals were prospectively registered in GAMES
152 using a standardized form.
153 Regional and local ethics committees approved the study, and patients gave their informed
154 consent in each center. For the present study, we selected all the infective episodes (n=1,453)
155 registered in the GAMES cohort involving adult patients (aged ≥ 18 years) who had undergone
156 cardiac surgery with preoperative diagnosis of active IE. From these, 354 (24.4%) died after
157 surgery (273 in the first 30 days and the remaining 81 during hospitalization). **Supplementary**
158 **Material: Table S1** shows the main descriptive characteristic of patients in the validation
159 nationwide registry.

160 *Statistical analyses*

161 The validation dataset was depurated for the outcomes and predictors included in the prognostic
162 models included in the systematic review.

163 Model aggregation was based on stacked regressions (16), which allows the synthesis of
164 literature models in a meta-model using the prior evidence optimized for the validation dataset
165 (17,18). Only the models that reported the full model equation and were not flagged as high
166 risk of bias were considered for aggregation. Stacked regressions used the linear predictor of
167 each model as a co-variable in the meta-model, and subsequently created a linear combination
168 of model predictions. That is, the original coefficients of each model are weighted by an
169 independent parameter estimated in the meta-model, so that the models with worse performance
170 in the validation dataset are penalized more. When aggregation of the coefficients was not

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171 possible, either because the definition of the predictor from primary studies was too
172 heterogeneous or because predictors had been modeled differently in the published models (for
173 instance, a numerical variable treated as a continuous predictor in one model and being
174 categorized at different cut-points in the others), these predictors were dropped, and were
175 included in the meta-model as independent covariables to re-estimate their coefficients entirely
176 from scratch based on the validation dataset. Non-linear relationships for continuous predictors
177 were tested using fractional polynomials (19).

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178 Predictors with missing data in the validation dataset were imputed under the missing at random
179 assumption using multiple imputation with chained equations (20). We included all predictors
180 and the outcome in the imputation models to ensure compatibility. (**Supplementary Material:**
181 **S4**). Imputation checks were completed by looking at the distributions of imputed values to
182 ensure plausibility. We generated 10 multiple imputed datasets and all primary analyses were
183 performed in each imputed dataset. Pooled parameters were estimated both in the aggregation
184 and validation processes using Rubin's rules (21).

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185 The model validation was assessed in terms of discrimination (i.e. through the use of the C-
186 statistic, with values from 1 indicating perfect discrimination to 0.5 no discrimination) and
187 calibration (i.e. through the calibration slope and calibration-in-the-large [CITL], with 1 and 0
188 as ideal values, respectively; as well as with calibration plots). Calibration plots represent the
189 average predicted probability for risk groups categorized using deciles of predicted probability
190 against observed proportion in each group, and fitting a less smoother to show calibration
191 across the entire range of predicted probabilities at the individual-level (22,23). For the
192 calibration plots we used the average predicted probabilities for individuals by pooling the
193 imputed datasets using Rubin's rules (21). Because the meta-model was optimized to the

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194 validation dataset, we assessed its optimism-corrected performance measures by applying
195 bootstrap validation with 500 replicates. As sensitivity analyses, we tested all model

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196 performance regardless of their critical appraisal. In addition, the meta-model performance was
197 assessed only for 30-days mortality to investigate the meta-model robustness.

198 All analyses were performed using Stata software version 16 (24).
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200 **Results**

201 *Literature search and study selection*

202 We retrieved 4,862 titles through our systematic search combining Medline and Embase. From
203 these, 684 duplicate references were identified. Of 4,178 titles assessed by title and abstract, 34
204 studies were retained for full text screening, and 2 additional studies were detected in the
205 bibliographic references of these articles. Nine studies describing 11 prediction models met the
206 inclusion criteria (**Figure 1 and Supplementary Table S2**).

207 *Source of data and participants*

208 All prognostic model development studies were carried out in the last decade. Six used data
209 from a study cohort (three of them from a single center (25–27) and three from multiple centers
210 (28–30)); two studies used data from multicenter registries (6,31); and one study used data
211 from both a multicenter cohort and a local clinical registry (32). Eight studies used data from
212 patients in Europe (Spain, Italy, France or Portugal) and one from patients in North America.
213 Participants were recruited between 1980 and 2015. (**Supplementary Table S3**).

214 *Outcomes*

215 Three models were developed to predict any death occurring before discharge or within 30 days
216 of surgery (6,25,27), five models **were built** to predict any death occurring before discharge
217 (26,30–32), and the remaining three **models predicted** death within 30 days of surgery (28,
218 29). The incidence of deaths varied between 8.2% and 29.2% (**Table 1**).

219 *Predictors*

220 The number of candidate predictors considered in the models varied from 15 to 57 and included
221 patient-, clinical-, surgery- and IE-related factors. The number of parameters retained in the
222 final models ranged from 2 to 15 (**Table 1**): The most common factors were critical
223 preoperative state (n=9), renal failure (n = 7), age (n = 6), New York Heart Association (NYHA)

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224 classification (n = 6), paravalvular complications (n = 6) and infection etiology (n = 5). The
225 predictor definitions and the models' composition are shown in the **Supplementary Table S4**
226 **and Table S5**.

227 *Model development and presentation*

228 Sample sizes for models' development varied between 128 and 13,617 patients, and the
229 number of events ranged from 21 and 1,117. Only two models from the same study adequately
230 informed the handling of missing data (29), and these used complete data analyses. Logistic
231 regression analysis was the most common modelling technique (n = 9), while logistic mixed
232 effects (28) and logistic GEE (Generalized Estimating Equation) models (6) were only used in
233 one model development each. Nine models used univariable analyses to select the candidate
234 predictors. In nine out of eleven models the number of events per parameter (EPP) assessed for
235 inclusion in the final model were lower than the minimum required for development of a
236 new prediction model, based on the sample size estimation proposed by Riley et al.(33,34)
237 (**Supplementary Table S6**). The method of predictors selection during multivariable modelling
238 was backward selection in three models (26,32), stepwise selection in two models (30,31), and
239 an automatic algorithm based on Akaike information criteria in multiple bootstrap samples in
240 the other two models, with predictors selected in at least 70% of the bootstrapped samples being
241 included in the final model (29). Four models did not inform about the method used to select
242 predictors. (**Table 1**)

243 In seven out of 11 models the authors did not inform the complete model equation, and five of
244 them did not respond when were asked for further details
245 (**Supplementary Table S7**). Nine models were presented as a scoring system, and two of them
246 included nomograms.

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247 *Model performance*

248 The model performance was assessed in terms of discrimination in all models through the C-
249 statistic. Nevertheless calibration was often wrongly assessed using the Hosmer-Lemeshow test
250 (37) in six models. Only three models (27,29) used calibration slopes and CITL. Eight models
251 were internally validated: three models were evaluated by bootstrapping with correction for
252 optimism (28,29), one was assessed through the 0.632 bootstrap method (26), two used
253 temporal split samples (32) and two used random split samples (6,30). Three models only
254 estimated the apparent performance (25,27,31). Three models were externally validated in the
255 same development study using very small sample sizes, with only 18 events in the Olmos
256 model (30) and 21 in the Gattis model (32). Clinical utility of the models was never assessed.

257 *Risk of bias*

258 The RoB was high in eight models, unclear in one (28) and low in the remaining two (29)
259 (Table 1, Supplementary Table S8 and Figure S1). Two of the eight models with high RoB
260 scored at “high risk” in the participants domain. Eight models scored at “high risk” in the
261 analysis domain. Most of the models had small sample sizes and the number of EPP was
262 close to 1 in several models, increasing the risk of overfitting (34). Many studies decided model
263 predictors based on univariable analysis, three reported only the apparent performance and two
264 used random splitting validation. The calibration was sub-optimally assessed in all models
265 classified as high risk of bias, with most of them using the Hosmer-Lemeshow test.

266 *Derivation of the Meta-model*

267 The eight models with high RoB were excluded from the statistical synthesis so that only the
268 EndoScore, Specifics EuroSCORE-I (Specific ES-I and EuroSCORE-II (Specific ES-II)
269 models were aggregated in the meta-model. The model developed by Di Mauro (EndoSCORE)
270 (28) included 15 parameters, while the other two (Specific ES-I and Specific ES-II) developed
271 by Fernández-Hidalgo (29), presented 10 and 9 parameters respectively, from the EuroSCORE

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272 models predictors (35, 36) and IE-related factors (**Table 2 and Supplementary Table S7**). The
273 dependent variable for the meta-model was mortality (either 30-days or in-hospital).

274 To construct the meta-model, we first calculated the linear predictors (LP) from EndoSCORE,
275 Specific ES-I and Specific ES-II for each observation in the validation dataset, after dropping
276 the parameters for age and infection etiology because these variables were modeled
277 differently in the different studies. Subsequently, we adjusted the meta-model using a logistic
278 regression model, which incorporated the LPs as co-variables, to estimate the models' weights
279 for aggregation, as well as the predictors for age (treated as continuous) and infection etiology
280 (categorized into three groups: *Staphylococcus* spp, fungi and other microorganisms) to re-
281 estimate the coefficients from scratch. The meta-model included 18 parameters from the
282 predictors included in at least one of the three original models (**Table 2**).

283 *Validation of the models*

284 The three prediction models considered for aggregation and the meta-model were validated in
285 the GAMES registry. The C-statistics and their 95% confidence intervals (95%CI) for the
286 published models were: 0.76 (95% CI 0.73 to 0.79) for EndoSCORE, 0.76 (95% CI 0.73 to
287 0.79) for Specific ES-I, and 0.73 (95% CI 0.73 to 0.79) for Specific ES-II. The optimism
288 adjusted C-statistic for the meta-model was 0.79 (95% CI 0.76 to 0.82) (**Figure 2**). Calibration
289 slopes were < 1 for all published models: 0.80 (95% CI 0.69 to 0.92) for EndoScore, 0.82 (95%
290 CI 0.70 to 0.94) for Specific ES-I, and 0.76 (95% CI 0.65 to 0.87) for Specific ES-II. CITL was
291 0.58 (95% CI 0.44 to 0.71) for EndoSCORE and 0.62 (95% CI 0.48 to 0.76) for Specific ES-
292 II, and -0.02 (95% CI -0.16 to 0.11) for Specific ES-I. Optimism adjusted calibration measures
293 for the meta-model were 0.98 (95% CI 0.86 to 1.13) for the slope and -0.05 (95% CI -0.20 to
294 0.11) for CITL (**Figure 2**). The calibration plots for the three previously published models and
295 the meta-model are shown in **Figure 3**.

296 Sensitivity analysis showed that the meta-model had better overall performance than all
297 published models regardless of their quality assessment (**Supplementary Figure S2**).
298 Moreover, even though the meta-model was not fitted for the 30-days mortality outcome, it
299 outperformed the three models used for model aggregation. (**Supplementary Figure S3**)

300

301 **Discussion**

302 *Summary of findings*

303 In this systematic review of prediction models for post-operative mortality in patients with
304 infective endocarditis, we identified and critically appraised 11 models developed in 9 studies.
305 The predicted outcome varied between studies (in-hospital, 30-days or both in-hospital or 30-
306 days mortality). Of the eleven prognostic models, only two had low RoB and one
307 unclear, the remained eight models had high RoB mainly owing to poor
308 statistical methods used, which suggests that their predictive performance when used in practice
309 is probably lower than that reported. The sample sizes used to develop the models were limited
310 and this is a well-known problem that leads to inaccurate predictions and consequently incorrect
311 healthcare decisions in practice (34).

312 Four out of the 11 published models reported the full model equation required for a models'
313 aggregation and a complete independent external validation as recommended by reporting
314 guidelines (10,11). Two models' equations were recovered asking correspondence authors.
315 Three models that were not flagged as high RoB could be used to create the meta-model.

316 Our meta-model showed better performance than the existing models. We investigated the
317 internal validity of the meta-model using bootstrap validation, and the results indicate there was
318 no substantial over-optimism and that the validation sample was sufficiently large to combine
319 and update the published models. Therefore, the meta-model is likely less prone to over-
320 optimism and more generalizable to new patient populations or settings, because it was built
321 from the evidence of several patient cohorts and optimized to a nationwide registry.

322 *Strengths and limitations*

323 To our knowledge, this is the first systematic review of prediction models of post-operative
324 mortality in patients with infective endocarditis with a complete external validation. We only

325 combined the published prediction models with low or unclear RoB and adjusted them to a new
326 patient population. We used multiple imputation of predictors to avoid loss of useful
327 information. The resulting meta-model incorporated prior knowledge optimally and
328 outperformed previously published models.

329 Our study has some limitations. The outcome definition in the validation dataset was either 30-
330 days or in-hospital post-operative mortality, and the outcome definition in the three models
331 used for aggregation was 30-days mortality. Despite this difference a sensitivity analysis
332 showed that the meta-model outperformed all published models when we explored its
333 performance for the 30-days mortality. The meta-model did not include some predictors that
334 were associated with post-operative mortality from the models with high RoB.
335 Nevertheless, except type-of-valve which was included in several models (27,30,31), the
336 remaining predictors were each only included in one model. Unfortunately, although we
337 identified 11 prediction models in our systematic review, we were only able to validate the
338 models that published the complete model equation. Although the definition of predictors in
339 GAMES registry was standardized, these could differ from definitions of published studies.

340 *Comparison to existing studies*

341 Most studies to develop new prediction models are based on small sample sizes and the
342 modelling strategies are excessively driven by available data without considering the previous
343 knowledge, leading to inefficient models. Other authors carried out external validation studies
344 but none of them made a critical appraisal (38–41). In a previous study, Varela et. al. developed
345 a prognostic meta-model based on a systematic review of pre-operative factors related to in-
346 hospital mortality, however, it was built using multiple univariate meta-analyses of the crude
347 associations, without considering possible covariable correlations (42,43).

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348 *Implications for practice*

349 The decision whether to perform surgery in IE remains a challenge in clinical practice and it
350 should come after a careful balance between the procedural risk and its estimated benefit.

351 Although risk scores in predicting mortality do not offer help in terms of establishing the
352 burdens of surgical futility, they apport a great value helping endocarditis teams to manage that
353 complex disease.

354 Although in the 2015 IE guidelines the score created by De Feo-Cotrufo et al for native IE is
355 the only one recommended, it would be expected to change with the creation of several new
356 IE specific scores and the generation of a meta-model that outperformed existing models.

357 *Challenges and opportunities*

358 Further external validation studies are necessary to confirm the improvement in predictive
359 ability of the meta-model. We will develop an online calculator to allow a simple and effective
360 use of the meta-model. Given the low incidence of infective endocarditis, available sufficient
361 sample sizes for the adequate development of new predictive models are difficult to come by.

362 We encourage authors to make their data available in order to allow building models s based on
363 available data (44).

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364 **Conclusions**

365 The meta-model we built is a robust prognostic model to calculate the individualized risk of
366 post-operative mortality in patients with infective endocarditis. It was developed based on the
367 previous evidence using aggregation methods of the existing models identified from a
368 systematic review and after critical being appraised. This meta-model outperformed existing
369 models; therefore, this preoperative tool can help guide individually tailored choices made by
370 patients and clinicians.

371

372 **Authors contributions**

373 Conceptualization: BMFF, LVB, JLA, AM, JIP, MR, JRM, EGE, JZ; Search strategies:

374 BMFF, NAD, JLA; Data extraction and Critical appraisal: BMFF, LVB, ACP; Methodology:

375 BMFF, EGE, AM, JZ; Software, Formal analysis: BMFF; Validation: AM, JZ; Data

376 acquisition/curation: BMFF, ENE, PM, MCF, MAG; Writing - Original draft: BMFF, EGE,

377 JZ; Visualization: BMFF, LVB, NFH; Supervision: EGE, JZ; Writing – Review & Editing:

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