

- 22 Schafer JL, Graham John W. Missing data: our view of the state of the art. *Psychol Methods* 2002;7:147–77.
- 23 van Buuren S, Groothuis-Oudshoorn K. MICE: multivariate imputation by chained equations in R. *J Stat Soft* 2011;45.
- 24 Urban D, Mayerl J, Wahl A. Regression analysis when variables have missing values: imputation or no imputation? A guide for practical regression analysis with SPSS. Working paper, ISSN: 2199-7780, University of Stuttgart, 2016.
- 25 van Buuren S. Multiple imputation of discrete and continuous data by fully conditional specification. *Stat Methods Med Res* 2007;16:219–42.
- 26 White IR, Royston P, Wood AM. Multiple imputation using chained equations: issues and guidance for practice. *Statist Med* 2011;30:377–99.
- 27 Abayomi K, Gelman A, Levy M. Diagnostics for multivariate imputations. *J R Stat Soc C* 2008;57:273–91.
- 28 Haemer M, Cluett S, Hassink SG, et al. Building capacity for childhood obesity prevention and treatment in the medical community: call to action. *Pediatrics* 2011; 128:S71–7.
- 29 Uerlich MF, Yumuk V, Finer N, et al. Obesity management in Europe: current status and objectives for the future. *Obes Facts* 2016;9:273–83.
- 30 Miller Perrin E, Flower KB, Garrett J, Ammerman AS. Preventing and treating obesity: pediatricians' self-efficacy, barriers, resources, and advocacy. *Ambul Pediatr* 2005;5:150–6.
- 31 Farpour-Lambert NJ, Baker JL, Hassapidou M, et al. Childhood obesity is a chronic disease demanding specific health care—a position statement from the Childhood Obesity Task Force (COTF) of the European Association for the Study of Obesity (EASO). *Obes Facts* 2015;8:342–9.
- 32 Bender R, Lange S. Adjusting for multiple testing—when and how? *J Clin Epidemiol* 2001;54:343–9.

.....
The European Journal of Public Health, Vol. 30, No. 3, 414–421

© The Author(s) 2020. Published by Oxford University Press on behalf of the European Public Health Association.

This is an Open Access article distributed under the terms of the Creative Commons Attribution License (<http://creativecommons.org/licenses/by/4.0/>), which permits unrestricted reuse, distribution, and reproduction in any medium, provided the original work is properly cited.

doi:10.1093/eurpub/ckaa007 Advance Access published on 7 February 2020

.....

Social inequalities in the provision of obstetric services in Norway 1967–2009: a population-based cohort study

Helene Sofie Eriksen^{1,*}, Susanne Høy^{2,*}, Lorentz M. Irgens^{3,4}, Svein Rasmussen^{5,6}, Kjell Haug³

1 Department of Internal Medicine, Ringerike Hospital, Vestre Viken Hospital Trust, Hønefoss, Norway

2 Department of Surgery, Lillehammer Hospital, Innlandet Hospital Trust, Lillehammer, Norway

3 Department of Global Public Health and Primary Care, University of Bergen, Bergen, Norway

4 Medical Birth Registry of Norway, Norwegian Institute of Public Health, Bergen, Norway

5 Department of Clinical Science, University of Bergen, Bergen, Norway

6 Department of Obstetrics and Gynecology, Haukeland University Hospital, Bergen, Norway

Correspondence: Kjell Haug, Department of Global Public Health and Primary Care, University of Bergen, Kalfarveien 31, Bergen 5020, Norway, Tel: +47 5 5586136, e-mail: kjell.haug@uib.no

*These authors contributed equally to this article.

Background: Socioeconomic (SE) inequalities have been observed in a number of adverse outcomes of pregnancy and many of the risk factors for such outcomes are associated with a low SE level. However, SE inequalities persist even after adjustment for these risk factors. Less well-off women are more vulnerable, but may also get less adequate health services. The objective of the present study was to assess possible associations between SE conditions in terms of maternal education as well as ethnic background and obstetric care. **Methods:** A population-based national cohort study from the Medical Birth Registry of Norway. The study population comprised 2 305 780 births from the observation period 1967–2009. Multilevel analysis was used because of the hierarchical structure of the data. Outcome variables included induction of labour, epidural analgesia, caesarean section, neonatal intensive care and perinatal death. **Results:** While medical interventions in the 1970s were employed less frequently in women of short education and non-western immigrants, this difference was eliminated or even reversed towards the end of the observation period. However, an excess perinatal mortality in both the short-educated [adjusted relative risk (aRR) = 2.49] and the non-western immigrant groups (aRR = 1.75) remained and may indicate increasing health problems in these groups. **Conclusion:** Even though our study suggests a fair and favourable development during the last decades in the distribution across SE groups of obstetric health services, the results suggest that the needs for obstetric care have increased in vulnerable groups, requiring a closer follow-up.

.....

Introduction

Over a long period of time, numerous studies in developed countries have reported substantial socioeconomic (SE) inequalities in morbidity and mortality, irrespective of how SE inequalities have been measured.^{1–3} In the interpretation of the results, emphasis has been put on inequalities in the exposure to the putative risk factors. The conclusion has been drawn that SE factors, such as educational level, may represent a proxy variable for material, psychosocial and behavioural risk factors.^{4–6}

However, registered morbidity and mortality are not only dependent on risk factors, but also on the health services provided.^{7,8} From an ethical point of view, substantial SE differences in the occurrence of diseases are unacceptable, but SE differences in the provision of health services in favour of well-off groups are intolerable. Medico-ethical rules have recommended that health services should be provided according to need rather than social background. Still, recent legislation in Norway, one of the world's most egalitarian countries,⁹ has found it necessary to emphasize equal provision of health services irrespective of social background.¹⁰

Also in adverse outcomes of pregnancy, substantial SE inequalities have been observed in a number of countries.^{11–16} Even with a perinatal mortality in Norway in 2011 of 4.8 per 1000 births,¹⁷ substantial differences have been reported between social groups.¹⁸ Many of the risk factors for adverse outcomes of pregnancy are associated with a low SE level.^{5,16,19,20} However, SE inequalities persist even after adjustment for these risk factors, suggesting the influence of other factors not accounted for.

Only few studies have addressed the question to which extent inequality in the provision of health services can account for inequality in health.^{21–23} Less well-off women may get less adequate health services.

Another vulnerable category of mothers is the increasing group of immigrants to western countries.^{24,25} Irrespective of their length of education or SE status in their country of origin, they may experience poor living conditions with a limited network and language difficulties.

Use of interventions is related to pathological conditions with different occurrence in different SE groups. By analysing subgroups with the same occurrence of pathology, valid estimates of possible effects of short education or status as immigrant on use of obstetrical interventions may be obtained. One such subgroup is preterm birth in which degree of pathology seems to be more independent of SE conditions than in term birth.^{26,27} Accordingly, need of intervention in preterm birth would be expected to be more independent of SE level.

A generation ago, long education among women was less common than today. As a group, women with short education were more heterogeneous. Today, women with short education represent a smaller and more homogeneous group with a higher burden of risk factors. Therefore, we might expect more health problems and more interventions among these women.

The objective of the present study was to assess possible associations between SE conditions as well as ethnic background and obstetric care. Based on data from the Medical Birth Registry of Norway, we wanted to assess to which extent provision of various types of obstetric intervention is dependent on the parent's SE level measured in terms of length of education. We also wanted to assess whether immigrants are provided with less adequate obstetric care.

A possible secular trend in the use of obstetric interventions across educational groups might also reflect attitudes among professionals that change over a period of time. Thus, we studied time trends in the associations of level of education and ethnic background with obstetric interventions and perinatal mortality throughout the observation period 1967–2009.

Methods

Based on compulsory notification, the Medical Birth Registry of Norway has, since 1967, registered all pregnancies in the country. Items of personal data include the national identification number as well as medical data on maternal health before and during the pregnancy, on the delivery and on the newborn. All data are notified by the attending midwife and doctor.²⁸ The national identification number provides linkage with the Statistics Norway for data on parental education and country of birth.

During the observation period 1967–2009 altogether 2 518 758 births were notified. After exclusion of multiple births and birth-weight under 500 g or gestational age under 22 weeks, the study population comprised 2 305 780 births. After further exclusion of births with lacking data on both parental education and country of birth, and births with gestational age 43+ weeks, the main analyses included 2 234 568 births.

Exposure variables included education defined as the highest level of education for the mother or the father obtained by 2009 and categorized as (i) university or college (more than 14 years), (ii) intermediate (11–14 years) and (iii) none or no more than

compulsory education (0–10 years). All births to mothers who were born in Norway, Western Europe, the USA, Australia and New Zealand were categorized according to the highest parental level of education. As we also wanted to assess whether non-western born mothers were provided with an adequate obstetric care, all these women were included in a group 4 irrespective of their educational attainment. Since this group was small until the 1980s, we do not report complete data for the first period 1967–80.

Outcome variables included induction of labour, epidural analgesia, caesarean section (CS), transfer of the newborn to a neonatal intensive care unit (NICU) and perinatal death.

Statistical analyses

For statistical analysis, we used SPSS (version 20) and the MIWin programme (version 2.30). Multilevel analysis was used due to the hierarchical structure (the first level was the pregnancy and the second level was the mother). The highest level of parental education was used as reference.

The analyses were stratified according to gestational age categorized as preterm (22–36 weeks) and term (37–42 weeks) and adjusted for maternal age categorized as <20, 20–24, 25–29, 30–34, 35–39 and 40+ years and for birth order categorized as 1, 2, 3, 4 and 5+. Further adjustment was made for size of the maternity unit categorized as 1–499, 500–1499, 1500–2999 and 3000+ births per year since they are related to exposure and outcome.

Particularly in the first part of the observation period, a considerable proportion of the births took place in rural smaller hospitals with limited possibilities to provide some of the interventions studied. An association between rural domicile and short education would confound the results. Thus, the relative risks (RRs) were also adjusted for annual number of births in the hospital where the delivery took place.

Results

The proportion of mothers with the highest level of parental education increased from 29.7% in 1967–80 to 53.7% in 2001–09 while the two other groups decreased correspondingly (table 1). The proportion of mothers born in Western world decreased continuously from 99.2% in 1967–80 to 85.0% in 2001–09.

Preterm births

In preterm births, during the first part of the observation period 1967–80, women with short education (<11 years) had less frequently induction of labour (RR 0.88), CS (RR 0.91) and epidural analgesia (RR 0.79) and had their newborns less frequently transferred to a NICU (RR 0.81) (table 2). Their offspring also had an excess perinatal mortality (RR 1.13) compared with those with long education (15+ years) (table 4).

During the last part 2001–09, women with short parental education had more frequently induction of labour (RR 1.08), CS (RR 1.08) and epidural analgesia (1.05) and had their newborns more frequently transferred to a NICU (RR 1.05) (table 2). However, the excess perinatal mortality was even higher than in the first period (RR 1.69) (table 4).

Term births

In term births during the first period, women with short parental education had their newborn transferred to a NICU, induction of labour and CS more close to the reference group, while epidural analgesia was less frequent (RR 0.69) (table 3). However, during the last period, provision of a number of services increased: transfer to NICU (RR 1.45), induction of labour (RR 1.04), CS (RR 1.36) and epidural analgesia (RR 1.19).

The excess perinatal mortality for term births ranged between RR 1.92 and 1.85 throughout the whole observation period 1967–2009

Table 1 Births in Norway 1967–2009 by demographic characteristics and year of birth

| | Year of birth/number of births (n) and percent (%) | | | | |
|--|--|----------------|------------------|----------------|-----------|
| | 1967–80, n (%) | 1981–90, n (%) | 1991–2000, n (%) | 2001–09, n (%) | Total (%) |
| Highest parental level of education | | | | | |
| Education <11 years | 106 111 (13.6) | 47 182 (9.8) | 44 150 (8.2) | 47 325 (9.4) | 10.6 |
| Education 11–14 years | 441 659 (56.4) | 232 168 (48.1) | 232 945 (43.5) | 173 897 (34.4) | 46.9 |
| Education ≥15 years | 232 497 (29.7) | 201 553 (41.7) | 255 475 (47.7) | 271 399 (53.7) | 41.7 |
| Education not recorded | 2146 (0.3) | 2114 (0.4) | 2713 (0.5) | 12 446 (2.5) | 0.8 |
| Gestational age | | | | | |
| Preterm (22–36 weeks) | 41 653 (5.3) | 25 696 (5.3) | 29 895 (5.6) | 27 849 (5.5) | 5.4 |
| Term (37–42 weeks) | 710 291 (90.8) | 439 009 (90.9) | 491 464 (91.8) | 475 748 (94.2) | 91.8 |
| Post term (43+ weeks) | 30 469 (3.9) | 18 312 (3.8) | 13 924 (2.6) | 1470 (0.3) | 2.8 |
| Maternal age | | | | | |
| <20 | 74 718 (9.5) | 25 998 (5.4) | 16 176 (3.0) | 11 825 (2.3) | 5.6 |
| 20–24 | 283 378 (36.2) | 134 173 (27.8) | 104 240 (19.5) | 73 964 (14.6) | 25.8 |
| 25–29 | 249 591 (31.9) | 180 213 (37.3) | 198 748 (37.1) | 162 378 (32.1) | 34.3 |
| 30–34 | 117 779 (15.1) | 104 542 (21.6) | 151 221 (28.3) | 169 204 (33.5) | 23.5 |
| 35–39 | 44 496 (5.7) | 33 058 (6.8) | 56 027 (10.5) | 74 881 (14.8) | 9.0 |
| 40+ | 12 451 (1.6) | 5033 (1.0) | 8871 (1.7) | 12 774 (2.5) | 1.7 |
| Not recorded | 0 (0.0) | 0 (0.0) | 0 (0.0) | 41 (0.0) | 0.0 |
| Birth number | | | | | |
| 1 | 319 742 (40.9) | 208 169 (43.1) | 218 791 (40.9) | 208 832 (41.3) | 41.4 |
| 2 | 261 903 (33.5) | 169 872 (35.2) | 190 439 (35.6) | 180 191 (35.7) | 34.8 |
| 3 | 125 079 (16.0) | 76 962 (15.9) | 91 257 (17.0) | 81 616 (16.2) | 16.3 |
| 4 | 47 348 (6.1) | 20 079 (4.2) | 24 840 (4.6) | 23 375 (4.6) | 5.0 |
| 5+ | 28 341 (3.6) | 7935 (1.6) | 9956 (1.9) | 11 053 (2.2) | 2.5 |
| Mother's country of birth | | | | | |
| Western world | 775 792 (99.2) | 466 496 (96.6) | 495 201 (92.5) | 429 179 (85.0) | 94.0 |
| Non-western world | 6621 (0.8) | 16 521 (3.4) | 40 082 (7.5) | 75 888 (15.0) | 6.0 |
| Maternity unit (births per year) | | | | | |
| 1–499 | 179 201 (22.9) | 75 786 (15.7) | 58 495 (10.9) | 51 538 (10.2) | 15.8 |
| 500–1499 | 307 356 (39.3) | 168 031 (34.8) | 130 753 (24.4) | 114 344 (22.6) | 31.2 |
| 1500–2999 | 215 349 (27.5) | 132 609 (27.5) | 167 808 (31.3) | 134 188 (26.6) | 28.2 |
| 3000+ | 73 489 (9.4) | 104 450 (21.6) | 174 138 (32.5) | 200 584 (39.7) | 24.0 |
| Outside institution | 7018 (0.9) | 2141 (0.4) | 4089 (0.8) | 4413 (0.9) | 0.8 |
| Total | 782 413 (100) | 483 017 (100) | 535 283 (100) | 505 067 (100) | 100 |

(table 4). In the same period the excess 'birthweight below 2500 g' increased from RR 2.19 to RR 2.80 (data not shown).

Non-western preterm group

In the non-western preterm group, transfer to NICU increased during the study period 1981–2009 (RR from 0.51 to 0.94), and so did induction of labour (RR from 0.86 to 0.97), CS (RR from 0.83 to 0.98) and epidural analgesia (RR from 0.63 to 0.87) (table 2). Perinatal mortality increased during the whole study period 1967–2009 (RR from 0.73 to 1.23) (table 4).

Non-western term group

In the non-western term group there was an increase in transfer to NICU (RR from 0.97 to 1.11), while the RR of induction of labour and epidural analgesia were close to the reference group while CS decreased (RR from 1.59 to 1.38) throughout the observation period 1981–2009 (table 3). The RR of perinatal mortality increased from 1.24 to 1.96 during the period 1967–2009 (table 4). In the same period the excess 'birthweight below 2500 g' ranged between RR 2.34 and RR 2.41 (data not shown).

Discussion

Our results indicate that in the 1970s, women with short education were provided with less obstetric services than women with long education. During the observation period, the utilization of obstetric services in the short-educated group increased both in preterm as well as term births; in preterm births to the level observed in the

longer-educated group, and in term births beyond this level. In non-western women, similar trends were observed.

In an equality of health care perspective, this result is gratifying as the distribution of obstetric services moved in a beneficial direction with more services provided to less privileged groups. To some extent these trends may reflect changing attitudes among health professionals. However, the increasing excess perinatal mortality observed both in the non-western and the short-educated groups, is disquieting and may indicate that these women as a group are more vulnerable than one generation ago. This is supported by the prevalence of low birth weight.

A strength of the present study is the large study population which provided statistical power. The variables were well defined and acquired from official registers with low occurrence of misclassification, which, to the extent it exists, would be non-differential. Education is a reliable indicator of SE status since it is clearly related to income, occupation, living conditions, social integration, lifestyle, quality of life, burden of disease and health in general.^{2,5,29} The long observation period made it possible to assess secular trends in the provision of health care to different groups of women.

Our results might have been influenced by a number of confounders. Obviously, the association between SE group and pathology will influence the needs of care. The stratification into term and preterm births to some extent reduced the confounding. Maternal age, parity and year of birth were adjusted for as potential confounders. The size of the maternity unit would have different effects in the short-educated and the non-western groups. Women living in remote areas generally have shorter education and deliver in smaller maternity units in which transfer to an NICU or other services are more complicated. This would cause a spurious association between short education and low

Table 2 Obstetric services provided in preterm birth (22–36 weeks) according to parental educational level and mothers' country of origin by year of birth, Norway 1967–2009

| | Year of birth/number (n), per 1000 births and aRR and 95% confidence interval (CI) | | | | | | | | | | | |
|-------------------------|--|----------|------------------|---------|----------|------------------|-----------|----------|------------------|---------|----------|------------------|
| | 1967–80 | | | 1981–90 | | | 1991–2000 | | | 2001–09 | | |
| | n | Per 1000 | aRR (95% CI) | n | Per 1000 | aRR (95% CI) | n | Per 1000 | aRR (95% CI) | n | Per 1000 | aRR (95% CI) |
| Induction of labour | 892 | 118.2 | 0.88 (0.81–0.97) | 653 | 209.6 | 1.02 (0.94–1.12) | 864 | 328.6 | 1.04 (0.96–1.12) | 908 | 422.1 | 1.08 (1.00–1.16) |
| Education <11 years | 2852 | 121.4 | 0.92 (0.87–0.99) | 2626 | 214.1 | 0.99 (0.94–1.06) | 4174 | 331.1 | 1.03 (0.98–1.07) | 3894 | 427.0 | 1.05 (1.00–1.09) |
| Education 11–14 years | 1457 | 145.5 | Ref | 2054 | 231.0 | Ref | 3915 | 337.8 | Ref | 5076 | 431.0 | Ref |
| Education ≥15 years | | | | 275 | 205.8 | 0.86 (0.76–0.98) | 932 | 308.8 | 0.92 (0.86–0.99) | 1903 | 409.2 | 0.97 (0.91–1.02) |
| Non-western | | | | | | | | | | | | |
| Caesarean section | 628 | 83.2 | 0.91 (0.83–0.99) | 794 | 254.9 | 0.94 (0.87–1.02) | 838 | 318.8 | 1.03 (0.96–1.12) | 787 | 365.9 | 1.08 (1.00–1.17) |
| Education <11 years | 2155 | 91.7 | 1.05 (0.97–1.13) | 3494 | 284.9 | 1.00 (0.95–1.05) | 4261 | 338.0 | 1.07 (1.03–1.12) | 3481 | 381.7 | 1.07 (1.02–1.12) |
| Education 11–14 years | 1123 | 112.1 | Ref | 2762 | 310.7 | Ref | 3881 | 334.8 | Ref | 4496 | 381.8 | Ref |
| Education ≥15 years | | | | 350 | 262.0 | 0.83 (0.74–0.93) | 919 | 304.5 | 0.92 (0.85–0.99) | 1674 | 359.9 | 0.98 (0.93–1.04) |
| Non-western | | | | | | | | | | | | |
| Epidural analgesia | 68 | 9.0 | 0.79 (0.60–1.05) | 443 | 142.2 | 0.99 (0.88–1.10) | 433 | 164.7 | 0.99 (0.89–1.10) | 496 | 230.6 | 1.05 (0.95–1.17) |
| Education <11 years | 267 | 11.4 | 0.81 (0.67–0.97) | 1675 | 136.6 | 0.92 (0.85–0.98) | 2054 | 162.9 | 0.96 (0.90–1.02) | 1992 | 218.4 | 1.06 (0.99–1.12) |
| Education 11–14 years | 212 | 21.2 | Ref | 1465 | 164.8 | Ref | 2134 | 184.1 | Ref | 2477 | 210.3 | Ref |
| Education ≥15 years | | | | 145 | 108.5 | 0.63 (0.53–0.74) | 437 | 144.8 | 0.79 (0.71–0.88) | 878 | 188.8 | 0.87 (0.80–0.94) |
| Non-western | | | | | | | | | | | | |
| Transfer intensive care | 590 | 78.2 | 0.81 (0.74–0.90) | 352 | 113.0 | 0.94 (0.83–1.05) | 469 | 178.4 | 0.90 (0.83–0.99) | 1194 | 555.1 | 1.05 (1.00–1.10) |
| Education <11 years | 2104 | 89.5 | 0.84 (0.79–0.91) | 1480 | 120.7 | 0.98 (0.90–1.05) | 2458 | 195.0 | 0.99 (0.94–1.05) | 5101 | 559.4 | 1.06 (1.03–1.09) |
| Education 11–14 years | 1280 | 127.8 | Ref | 1140 | 128.2 | Ref | 2468 | 212.9 | Ref | 6318 | 536.5 | Ref |
| Education ≥15 years | | | | 96 | 71.9 | 0.51 (0.42–0.61) | 576 | 190.9 | 0.89 (0.82–0.96) | 2340 | 503.1 | 0.94 (0.91–0.97) |
| Non-western | | | | | | | | | | | | |

Table 3 Obstetric services provided in term birth (37–42 weeks) according to parental educational level and mothers' country of origin by year of birth, Norway 1967–2009

| | Year of birth/number (n), per 1000 births and aRR with 95% confidence interval (CI) | | | | | | | | | | | |
|-------------------------|---|----------|------------------|---------|----------|------------------|-----------|----------|------------------|---------|----------|------------------|
| | 1967–80 | | | 1981–90 | | | 1991–2000 | | | 2001–09 | | |
| | n | Per 1000 | aRR (95% CI) | n | Per 1000 | aRR (95% CI) | n | Per 1000 | aRR (95% CI) | n | Per 1000 | aRR (95% CI) |
| Induction of labour | 11 152 | 120.6 | 0.90 (0.88–0.92) | 5424 | 144.2 | 1.06 (1.03–1.09) | 4966 | 160.7 | 1.16 (1.12–1.19) | 5021 | 205.3 | 1.04 (1.02–1.05) |
| Education <11 years | 49 680 | 124.7 | 0.93 (0.92–0.94) | 29 673 | 144.5 | 1.03 (1.01–1.04) | 30 946 | 154.9 | 1.08 (1.06–1.10) | 29 181 | 207.2 | 1.02 (1.02–1.03) |
| Education 11–14 years | 30 460 | 143.7 | Ref | 27 132 | 150.6 | Ref | 34 702 | 155.2 | Ref | 47 084 | 198.9 | Ref |
| Education ≥15 years | 739 | 125.0 | 0.87 (0.80–0.93) | 2129 | 145.4 | 0.97 (0.93–1.02) | 5582 | 154.1 | 1.05 (1.02–1.08) | 13 437 | 189.5 | 1.00 (1.00–1.01) |
| Caesarean section | 2889 | 31.2 | 1.04 (0.99–1.09) | 3626 | 96.4 | 1.24 (1.19–1.29) | 3383 | 109.5 | 1.32 (1.27–1.37) | 3296 | 134.8 | 1.36 (1.31–1.41) |
| Education <11 years | 12 855 | 32.3 | 1.01 (0.98–1.05) | 18 913 | 92.1 | 1.09 (1.07–1.12) | 20 574 | 103.0 | 1.15 (1.13–1.17) | 19 888 | 141.2 | 1.24 (1.22–1.26) |
| Education 11–14 years | 8326 | 39.3 | Ref | 17 272 | 95.9 | Ref | 22 759 | 101.8 | Ref | 31 386 | 132.6 | Ref |
| Education ≥15 years | | | | 2031 | 138.7 | 1.59 (1.51–1.67) | 4762 | 131.5 | 1.47 (1.42–1.52) | 11 189 | 157.8 | 1.38 (1.35–1.42) |
| Non-western | | | | | | | | | | | | |
| Epidural analgesia | 465 | 5.0 | 0.69 (0.62–0.76) | 3473 | 92.3 | 1.04 (1.00–1.08) | 4126 | 133.6 | 1.10 (1.06–1.13) | 7674 | 313.8 | 1.19 (1.16–1.22) |
| Education <11 years | 2392 | 6.0 | 0.64 (0.60–0.67) | 16 465 | 80.2 | 0.91 (0.89–0.93) | 24 848 | 124.4 | 0.99 (0.98–1.01) | 37 329 | 265.1 | 1.08 (1.07–1.09) |
| Education 11–14 years | 3128 | 14.8 | Ref | 19 002 | 105.5 | Ref | 31 883 | 142.6 | Ref | 62 617 | 264.6 | Ref |
| Education ≥15 years | | | | 1612 | 110.1 | 0.95 (0.90–1.00) | 5043 | 139.2 | 0.97 (0.94–1.00) | 17 787 | 250.9 | 0.95 (0.94–0.97) |
| Non-western | | | | | | | | | | | | |
| Transfer intensive care | 703 | 7.6 | 0.93 (0.85–1.02) | 371 | 9.9 | 1.27 (0.13–1.43) | 616 | 19.9 | 1.11 (1.02–1.21) | 1633 | 66.8 | 1.45 (1.37–1.53) |
| Education <11 years | 3360 | 8.4 | 0.95 (0.95–1.01) | 1868 | 9.1 | 1.13 (1.06–1.22) | 3627 | 18.2 | 1.02 (0.97–1.07) | 8246 | 58.6 | 1.25 (1.22–1.29) |
| Education 11–14 years | 2253 | 10.6 | Ref | 1570 | 8.7 | Ref | 4416 | 19.7 | Ref | 12 137 | 51.3 | Ref |
| Education ≥15 years | | | | 125 | 8.5 | 0.97 (0.81–1.16) | 910 | 25.1 | 1.30 (1.21–1.41) | 3945 | 55.6 | 1.11 (1.07–1.16) |
| Non-western | | | | | | | | | | | | |

Table 4 Perinatal mortality in preterm and term births according to parental educational level and mothers' country of origin by year of birth, Norway 1967–2009

| | Year of birth/number (n), per 1000 births and aRR with 95% confidence interval (CI) | | | | | | | | | | | | | |
|------------------------------|---|----------|------------------|---------|----------|------------------|-----------|----------|------------------|---------|----------|------------------|--------|----------|
| | 1967–80 | | | 1981–90 | | | 1991–2000 | | | 2001–09 | | | Total | |
| | n | Per 1000 | aRR (CI) | n | Per 1000 | aRR (CI) | n | Per 1000 | aRR (CI) | n | Per 1000 | RR (CI) | n | Per 1000 |
| Preterm (22–36 weeks) | | | | | | | | | | | | | | |
| Education <11 years | 1525 | 202.0 | 1.13 (1.05–1.21) | 458 | 147.0 | 1.72 (1.52–1.94) | 221 | 84.1 | 1.43 (1.22–1.67) | 146 | 67.9 | 1.69 (1.38–2.06) | 2350 | 152.2 |
| Education 11–14 years | 4154 | 176.8 | 1.02 (0.96–1.08) | 1228 | 100.1 | 1.18 (1.07–1.29) | 756 | 60.0 | 1.05 (0.95–1.17) | 389 | 42.7 | 1.05 (0.92–1.21) | 6527 | 113.5 |
| Education ≥15 years | 1745 | 174.3 | Ref | 761 | 85.6 | Ref | 661 | 57.0 | Ref | 497 | 42.2 | Ref | 3664 | 86.7 |
| Non-western | 64 | 132.2 | 0.73 (0.57–0.94) | 114 | 85.3 | 0.95 (0.78–1.16) | 203 | 67.3 | 1.12 (0.95–1.31) | 251 | 54.0 | 1.23 (1.06–1.44) | 632 | 66.6 |
| Term (37–42 weeks) | | | | | | | | | | | | | | |
| Education <11 years | 869 | 9.4 | 1.92 (1.74–2.11) | 220 | 5.8 | 2.31 (1.95–2.73) | 127 | 4.1 | 1.84 (1.50–2.24) | 64 | 2.6 | 1.85 (1.40–2.45) | 1280 | 6.9 |
| Education 11–14 years | 2696 | 6.8 | 1.46 (1.36–1.58) | 693 | 3.4 | 1.33 (1.18–1.50) | 531 | 2.7 | 1.18 (1.04–1.34) | 323 | 2.3 | 1.53 (1.31–1.78) | 4243 | 4.5 |
| Education ≥15 years | 1033 | 4.9 | Ref | 468 | 2.6 | Ref | 539 | 2.4 | Ref | 399 | 1.7 | Ref | 2439 | 2.9 |
| Non-western | 36 | 6.1 | 1.24 (0.89–1.73) | 59 | 4.0 | 1.52 (1.16–2.00) | 136 | 3.8 | 1.61 (1.33–1.95) | 220 | 3.1 | 1.96 (1.65–2.32) | 451 | 3.5 |
| Preterm and term | | | | | | | | | | | | | | |
| Education <11 years | 2394 | 23.5 | 1.84 (1.74–1.95) | 678 | 16.2 | 2.60 (2.35–2.87) | 348 | 10.2 | 2.10 (1.85–2.37) | 210 | 7.9 | 2.49 (2.12–2.93) | 3630 | 17.8 |
| Education 11–14 years | 6850 | 16.0 | 1.33 (1.27–1.39) | 1921 | 8.6 | 1.40 (1.30–1.51) | 1287 | 6.0 | 1.25 (1.16–1.36) | 712 | 4.7 | 1.47 (1.33–1.63) | 10 770 | 10.6 |
| Education ≥15 years | 2778 | 12.4 | Ref | 1229 | 6.4 | Ref | 1200 | 5.1 | Ref | 896 | 3.6 | Ref | 6103 | 6.8 |
| Non-western | 100 | 15.1 | 1.15 (0.94–1.41) | 173 | 10.6 | 1.54 (1.31–1.82) | 339 | 8.5 | 1.65 (1.46–1.87) | 471 | 6.3 | 1.75 (1.56–1.96) | 1083 | 7.9 |

provision of care. Non-western women would live in urban areas and deliver at larger maternity units well equipped with health services. In this group, a reported low provision of care would represent a conservative estimate. To avoid such confounding, we adjusted for size of maternity unit.

Whether our findings apply in other countries is unclear, but as SE gradients in the distribution of health services have been observed in Norway known to have comparatively low SE gradients in general, the situation may be more worrying in other countries.

In term births, we assume more pathological conditions in the short-educated women which would necessitate more frequent obstetric services. In the 1970s, this group was not provided with excess services, unlike today when this is the case. In most preterm births, need of special obstetric services is involved and more or less independent of social background. In these births, the short-educated women were initially also provided with less frequent services, as opposite to what was observed in more recent years when equal services were provided across educational groups. It is interesting to note that this secular trend in the distribution of obstetric health care did not result in a reduction of care in the women with longer education.

In a number of studies, a higher perinatal mortality has been documented among children of women with short education.^{13,30} A similar gradient was observed in the present study. To the extent that mortality is dependent on medical care, one would expect that both the improvement and the redistribution of obstetric care observed in Norway during the observation period would reduce the gradient in perinatal mortality. This was not the case. In preterm births, excess perinatal mortality increased both in the short-educated and the non-western groups. In term births, it remained high in the short-educated group and increased in the non-western group. In the interpretation of these results, we have to consider that the short education group has become much smaller and more marginalized during the study period. Consequently, the group will be more exposed to risk factors and thereby suffer an increasing perinatal mortality.

The situation is different for non-western women. They represent an increasing group of births. The group is heterogeneous, comprising different cultural, religious and social backgrounds, as well as individual factors related to life style and living conditions. This makes it difficult to identify specific risk factors causing the excess perinatal mortality. Still, it seems reasonable to suggest that the needs for more care in these births are not met by the beneficial redistribution of the obstetric services observed in Norway.

Technological development represents a factor contributing to the change in provision of health care observed since 1970s. Obstetric pain relief is an example. Epidural analgesia was implemented in obstetric care in the 1970s. In the beginning, the use was restricted and thereby limited to a smaller group; women with long education were overrepresented. Throughout the observation period, the use of epidural analgesia increased to involve more than 20% of all births, and eventually the use was quite equally distributed between the groups. The increase in availability seemed to even out the inequalities, with less discrimination as a result.

During the observation period, the CS-rate increased from 3.8% in 1967–70 to 15.5% in 2001–09. This may be accounted for by technological development as well as an increasing maternal age, but the number of CSs performed on maternal request demand has also increased.^{31,32} In these pregnancies, medical indications for a CS are rare, which involves a possibility of an SE gradient. However, we observed a 'higher' CS rate in the short-educated group throughout the observation period limited to term births. A higher occurrence of pathology in short-educated groups may explain the high CS rate, consistent with a more frequent transfer to an NICU. In non-western women, the excess CS-rate decreased through the years. Many other countries have a higher CS-rate than Norway,³³ which may account for the observed difference and trend.

In addition to occurrence of pathology and attitudes among health professionals, several other mechanisms may be involved in the processes by which health needs are met by the health services. Adequate treatment depends on the patient acknowledging a problem and seeking medical attention, communication, mutual understanding and finally the type of intervention involved.

Our results suggest a challenge to antenatal care, which needs to be addressed, most likely both by educating the parents and in the training of health professionals. Still, the fact that the RRs involved approached unity and equality towards the end of the observation period was gratifying.

Conclusion

Our study suggests a favourable development during the last decades in the distribution of obstetric health services in Norway between different SE groups. While medical interventions in the 1970s were employed less in women of short than long education, the employment in women with short education increased to equal or even higher levels towards the end of the period. The same trend was observed for non-western women. Still, the excess perinatal mortality in both the short-educated and the non-western groups remained. The results suggest increasing perinatal health problems in vulnerable groups and indicate need for a closer obstetric follow-up and surveillance. Midwives and general practitioners should be aware of risk factors and should offer support, education and more frequent contacts. Further studies, addressing distribution of antenatal care would provide a basis for developing preventive measures aiming at SE equality.

Ethical approval

The study was approved by the internal review board of the Medical Birth Registry of Norway and by the regional ethics committee, REK Vest, Norway (2009/1868).

Conflicts of interest: All authors have completed the Unified Competing Interest form at www.icmje.org/coi_disclosure.pdf (available from the corresponding author).

Key points

- During the study period 1967–2009 the distribution of obstetric services moved in a beneficial direction with more services provided to less privileged groups.
- Perinatal mortality remained almost twice as high among children of short-educated women and non-western immigrants throughout the study period.
- Even though the study suggests a better distribution of obstetric health services across socioeconomic groups, the need for obstetric follow-up and surveillance seems to increase among vulnerable groups.

References

- Marmot MG, Smith GD, Stansfeld S, et al. Health inequalities among British civil servants: the Whitehall II Study. *Lancet* 1991;337:1387–93.
- Mackenbach JP, Kunst AE, Cavelaars A, et al.; EU Working Group on Socioeconomic Inequalities in Health. Socioeconomic inequalities in morbidity and mortality in Western Europe. *Lancet* 1997;349:1655–9.
- Mackenbach JP, Stirbu I, Roskam AJR, et al. Socioeconomic inequalities in health in 22 European countries. *N Engl J Med* 2008;358:2468–81.
- Galobardes B, Shaw M, Lawlor DA, et al. Indicators of socioeconomic position (part 1). *J Epidemiol Community Health* 2006;60:7–12.
- van Oort FVA, van Lenthe FJ, Mackenbach JP. Material, psychosocial, and behavioral factors in the explanation of educational inequalities in mortality in the Netherlands. *J Epidemiol Community Health* 2005;59:214–20.
- Kunst AE, Mackenbach JP. The size of mortality differences associated with educational level in nine industrialized countries. *Am J Public Health* 1994;84:932–7.
- Nilssen Y, Strand TE, Fjellbirkeland L, et al. Lung cancer treatment is influenced by income, education, age and place of residence in a country with universal health coverage. *Int J Cancer* 2016;138:1350–60.
- Berglund A, Holmberg L, Tishelman C, et al. Social inequalities in non-small cell lung cancer management and survival: a population-based study in central Sweden. *Thorax* 2010;65:327–33.
- Norway-Gini Index (World Bank Estimate). Available at: <http://www.indexmundi.com/facts/indicators/SI.POV.GINI/rankings> (26 January 2020, date last accessed).
- Law on Patient and User Rights (LOV-2015-12-11-97). Oslo: Ministry of Health and Care Services (in Norwegian). Available at: <https://lovdata.no/dokument/NL/lov/1999-07-02-63> (26 January 2020, date last accessed).
- Jørgensen T, Mortensen LH, Andersen A. Social inequality in fetal and perinatal mortality in the Nordic countries. *Scand J Public Health* 2008;36:635–49.
- Blumenshine P, Egarter S, Barclay CJ, et al. Socioeconomic disparities in adverse birth outcomes. A systematic review. *Am J Prev Med* 2010;39:263–72.
- Rom AL, Mortensen LH, Cnattingius S, et al. A comparative study of educational inequality in the risk of stillbirth in Denmark, Finland, Norway and Sweden 1981–2000. *J Epidemiol Community Health* 2012;66:240–6.
- Auger N, Park AL, Harper S, et al. Educational inequalities in preterm and term small-for-gestational-age birth over time. *Ann Epidemiol* 2012;22:160–7.
- Poulsen G, Strandberg-Larsen K, Morensen L, et al. Exploring educational disparities in risk of preterm delivery: a comparative study of 12 European cohorts. *Paediatr Perinat Epidemiol* 2015;29:172–83.
- Ruiz M, Goldblatt P, Morrison J, et al. Mother's education and the risk of preterm and small for gestational age birth: a DRIVERS meta-analysis of 12 European cohorts. *J Epidemiol Community Health* 2015;69:826–33.
- Medical Birth Registry of Norway. Annual Report. Bergen: National Institute of Health, 2011.
- Arntzen A, Samuelsen SO, Bakkeiteig LS, Stoltenberg C. Socioeconomic status and risk of infant death. A population-based study of trends in Norway, 1967–1998. *Int J Epidemiol* 2004;33:279–88.
- Flenady V, Koopmans L, Middleton P, et al. Major risk factors for stillbirth in high-income countries: a systematic review and meta-analysis. *Lancet* 2011;377:1331–40.
- Agardh E, Allebeck P, Hallqvist J, et al. Type 2 diabetes incidence and socioeconomic position: a systematic review and meta-analysis. *Int J Epidemiol* 2011;40:804–18.
- Morris S, Sutton M, Gravelle H. Inequity and inequality in the use of health care in England: an empirical investigation. *Soc Sci Med* 2005;60:1251–66.
- Hagen TP, Häkkinen U, Iversen T, et al.; on behalf of the EuroHOPE Study Group. Socio-economic inequality in the use of procedures and mortality among AMI patients: quantifying the effects along different paths. *Health Econ* 2015;24:102–15.
- Asaria M, Ali S, Doran T, et al. How a universal health system reduces inequalities: lessons from England. *J Epidemiol Community Health* 2016;70:637–43.
- Naimy Z, Grytten J, Monkerud L, Eskild A. Perinatal mortality in non-western migrants in Norway as compared to their countries of birth and to Norwegian women. *BMC Public Health* 2013;13:37.
- Racape J, Schoenborn C, Sow M, et al. Are all immigrant mothers really at risk of low birth weight and perinatal mortality? The crucial role of socio-economic status. *BMC Pregnancy Childbirth* 2016;16:75.
- Blencowe H, Cousens S, Chou D, et al.; the Born Too Soon Preterm Birth Action Group. Born too soon: the global epidemiology of 15 million preterm births. *Reprod Health* 2013;10:S2.
- Liu L, Oza S, Hogan D, et al. Global, regional, and national causes of under-5 mortality in 2000–15: an updated systematic analysis with implications for the Sustainable Development Goals. *Lancet* 2016;388:3027–35.
- Irgens LM. The Medical Birth registry of Norway. Epidemiological research and surveillance throughout 30 years. *Acta Obstet Gynecol Scand* 2000;79:435–9.
- Winkleby MA, Jatulis DE, Frank E, Fortmann SP. Socioeconomic status and health: how education, income, and occupation contribute to risk factors for cardiovascular disease. *Am J Public Health* 1992;82:816–20.
- Zeitlin J, Mortensen L, Prunet C, et al.; Euro-Peristat Scientific Committee. Socioeconomic inequalities in stillbirth rates in Europe: measuring the gap using routine data from the Euro-Peristat Project. *BMC Pregnancy Childbirth* 2016;16:15.

- 31 Habiba M, Kaminski M, Da Fre M, et al. Caesarean section on request: a comparison of obstetricians' attitudes in eight European countries. *BJOG* 2006;113:647–56.
- 32 Selinger H. Maternal request for caesarean section: an ethical consideration. *J Med Ethics* 2014;40:857–60.
- 33 Delnord M, Blondel B, Drewniak N, et al.; Euro-Peristat Preterm Group. Varying gestational age patterns in cesarean delivery: an international comparison. *BMC Pregnancy Childbirth* 2014;14:321.

The European Journal of Public Health, Vol. 30, No. 3, 421–426

© The Author(s) 2020. Published by Oxford University Press on behalf of the European Public Health Association. All rights reserved.
doi:10.1093/eurpub/ckaa024 Advance Access published on 10 March 2020

Social inequalities in medical rehabilitation outcomes—a registry-based study on 219 584 insured persons in Germany

Simon Götz¹, Morten Wahrendorf¹, Johannes Siegrist², Nico Dragano¹

1 Institute of Medical Sociology, Centre for Health and Society, Medical Faculty, University of Duesseldorf, Duesseldorf, Germany

2 Senior Professorship on Work Stress Research, Centre for Health and Society, Medical Faculty, University of Duesseldorf, Duesseldorf, Germany

Correspondence: Simon Götz, Institute for Medical Sociology, Centre for Health and Society, Medical Faculty, University of Duesseldorf, Moorenstr. 5, 40225 Duesseldorf, Germany, Tel: +49 (0) 211 81 06544, e-mail: simon.goetz@uni-duesseldorf.de

Background: Given limited knowledge on the extent of social inequalities in longer-term work ability of people with a chronic disease, this study analyzes social inequalities of three consecutive indicators of work ability following medical rehabilitation in a large sample of insured employees. **Methods:** Based on data from the German statutory pension insurance, a representative 20% random sample of all employed persons undergoing medical rehabilitation between 2006 and 2008 was included in a longitudinal analysis ($n=219\ 584$ persons). Three measures of consecutive work-related outcomes (physicians' assessment of work ability at discharge; return to work in the year thereafter; disability pension during follow-up) and socioeconomic position (SEP) (education, occupational position and income) were assessed. Adjusted relative risks (RRs) for each outcome were calculated according to SEP, applying Poisson regression analysis. **Results:** The measures of SEP were associated with all three outcomes of work ability in the fully adjusted models. Relatively strongest relationships were observed for education as SEP measure, and they were particularly pronounced for 'low work ability' (RR=2.38 for lower secondary education compared to tertiary education; 95% CI: 2.26–2.51). Based on average marginal effects, absolute differences of work ability by SEP indicate a socially graded pattern, with only few exceptions. **Conclusions:** Despite Germany's universal access to medical and vocational rehabilitation social inequalities in longer-term work ability following chronic disease persist, thus calling for targeted programmes of prevention and occupational health promotion.

Introduction

Scientific evidence indicates that social gradients of morbidity and mortality persist in many countries with developed social and health policies, leaving people in lower socioeconomic positions (SEPs) at higher risk to develop a disease.¹ Less knowledge, though, is available on social inequalities of the longer-term consequences for living and working conditions among those who already developed or survived an incident chronic disease, such as cardiovascular disease, cancer or depression. Recent data from the German micro census (2013), e.g. suggests that 12.7 million people with impairments are living in Germany. They are 30 percentage points less likely to participate on the labour market and 7 percentage points more likely to experience poverty compared to the population without impairments.² While several investigations documented social gradients of success in return to work following hospitalization due to a stroke,³ coronary heart disease,^{4,5} cancer^{6,7} or other long-standing illnesses,⁸ few studies only analyzed longer-term outcomes, such as disability pensions, for these patient groups. A recent report from Finland provides an exception as it explored whether social inequalities of disability pension differed between hospitalized persons with severe disease and persons without severe disease.⁹

Although this study observed social gradients in disability pensions in both groups, the association of SEP with disability pension was somewhat stronger among persons with hospitalization, in particular in case of injuries and musculoskeletal disorders.⁸ Yet, a comprehensive study on socioeconomic differences of trajectories of work ability from hospitalization to differential opportunities of returning to work and to longer-term risks of early exit from labour market is still missing.

To fill this gap, administrative data derived from the German health and social security system are of interest. There are almost no social differences in access to, and treatment obtained in medical rehabilitation clinics for persons insured under the German pension insurance scheme. The pension funds have established a detailed registry of patients' trajectories following medical rehabilitation, including data on sociodemographic characteristics, employment histories (return to work and early exit from paid work) and survival.¹⁰

With this retrospective observational study, we set out to analyze social inequalities of trajectories following medical rehabilitation from chronic disease in a large sample of insured men and women in Germany, using this administrative dataset. More specifically, three steps of these trajectories are studied, (i) the patients' work