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## Early-life seasonal, weather and social effects on telomere length in a wild mammal

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17

### 18 Abstract

19 Early-life environmental conditions can provide a source of individual variation in life-history 20 strategies and senescence patterns. Conditions experienced in early life can be quantified by 21 measuring telomere length, which can act as a biomarker of survival probability in some species. Here, 22 we investigate whether seasonal changes, weather conditions, and group size are associated with 23 early-life and/or early-adulthood telomere length in a wild population of European badgers (Meles 24 meles). We found substantial intra-annual changes in telomere length during the first three years of 25 life, where within-individual effects showed shorter telomere lengths in the winter following the first 26 spring and a trend for longer telomere lengths in the second spring compared to the first winter. In

27 terms of weather conditions, cubs born in warmer, wetter springs with low rainfall variability had 28 longer early-life (3–12 months old) telomere lengths. Additionally, cubs born in groups with more cubs 29 had marginally longer early-life telomeres, providing no evidence of resource constraint from cub 30 competition. We also found that the positive association between early-life telomere length and cub 31 survival probability remained when social and weather variables were included. Finally, after sexual 32 maturity, in early adulthood (i.e. 12-36 months) we found no significant association between same-33 sex adult group size and telomere length (i.e. no effect of intra-sexual competition). Overall, we show 34 that controlling for seasonal effects, which are linked to food availability, is important in telomere 35 length analyses, and that variation in telomere length in badgers reflects early-life conditions and also 36 predicts first year cub survival.

37

38 Keywords: telomere length, early-life environment, group size, weather conditions, senescence,
 39 season

40

### 41 **1. Introduction**

42 The early-life environment can affect individual fitness (Lindström, 1999), with consequences for 43 variation in life-history strategies (Metcalfe & Monaghan, 2001) and senescence patterns (Nussey, 44 Kruuk, Morris, & Clutton-Brock, 2007). For example, it has been hypothesised that senescence, the 45 decline in performance in older age, is faster in individuals that experienced adverse early-life 46 environments, due to different energy allocation trade-offs between early- and later-life in response 47 to the environment (Kirkwood & Rose, 1991; Medawar, 1952; Williams, 1957). A more stressful early-48 life environment, either through a sub-optimal mean or more variable early-life environment, during 49 this sensitive developmental period, could trigger early reproductive investment at the expense of 50 somatic maintenance, leading to faster rates of senescence (Kirkwood & Rose, 1991; Lemaitre et al., 51 2015). Empirical evidence for such detrimental effects has been found in various wild animal 52 populations (Cooper & Kruuk, 2018; Hammers, Richardson, Burke, & Komdeur, 2013; Reed et al., 53 2008).

54 Telomere length has been suggested as a non-causal biomarker of senescence in some species 55 (López-Otín, Blasco, Partridge, Serrano, & Kroemer, 2013; Monaghan & Haussmann, 2006), that 56 facilitates quantification of physiological consequences of the conditions experienced (Monaghan, 57 2014). Telomeres are highly conserved nucleoprotein structures at the end of chromosomes 58 consisting of a non-coding sequence (5'-TTAGGG-3') and shelterin proteins (Blackburn, 2000; de 59 Lange, 2005). Telomeres maintain genomic integrity by preventing chromosome degradation and 60 fusion of chromosome ends by forming T-loops (de Lange, 2004). Generally, telomeres shorten with 61 each cell replication due to the end-replication problem (Olovnikov, 1973), but telomere shortening 62 can be accelerated potentially by oxidative damage (Boonekamp, Bauch, Mulder, & Verhulst, 2017; 63 Reichert & Stier, 2017; von Zglinicki, 2002) and through stressors (Epel et al., 2004; Heidinger et al., 64 2012). Telomeres can, however, elongate via the enzyme telomerase (Blackburn et al., 1989) – which 65 shows a negative correlation with mammalian body mass (Tian et al., 2018) - and other telomere-66 elongation pathways (Cesare & Reddel, 2010; Mendez-Bermudez et al., 2012). Cells with critically 67 short telomeres ultimately enter replicative senescence, where the accumulation of senescent cells 68 can impair tissue function due to reduced renewal capacity (Campisi, 2005; Campisi & di Fagagna, 69 2007) and potentially lead to organismal senescence (Young, 2018).

70 In some species, variation in early-life telomere length has been linked to season, specifically 71 with winter effects when torpor and hibernation facilitate tolerance of winter food scarcity and 72 reduction of thermoregulatory costs. During hibernation, more frequent arousal – which increases 73 metabolic rate and potentially increases oxidative stress - is associated in arctic ground squirrels 74 (Urocitellus parryii) with shorter telomere length (Wilbur, Barnes, Kitaysky, & Williams, 2019) and in 75 edible dormice (Glis glis) with increased telomere shortening (Turbill, Ruf, Smith, & Bieber, 2013). 76 Telomere shortening is reduced when the animals' core temperature difference between hibernation 77 and arousal is smaller, in both edible and garden (Eliomys quercinus) dormice (Nowack et al., 2019).

78 Conversely, the use of spontaneous daily torpor in non-hibernating Djungarian hamsters (Phodopus 79 sungorus) is associated with telomere lengthening due to a relatively low energy investment to return 80 to euthermia along with the benefits of reduced metabolic rate in torpor compared to hibernation 81 (Turbill, Smith, Deimel, & Ruf, 2012). In contrast, non-hibernating juvenile garden dormice that more 82 frequently underwent fasting-induced torpor showed higher telomere shortening than individuals 83 undergoing torpor less frequently (Giroud et al., 2014). Species that undergo facultative winter torpor 84 may conserve energy for somatic maintenance that could potentially be invested in telomere 85 restoration/elongation. Additionally, there is evidence in non-hibernating rodents for seasonal effects 86 of food availability on telomere dynamics (Criscuolo, Pillay, Zahn, & Schradin, 2020). However, since 87 telomere length, season and body mass might be intercorrelated (Réale, Festa-Bianchet, & Jorgenson, 88 1999; Tian et al., 2018), body mass needs to be taken into account when studying seasonal effects.

89 In addition to these intra-annual changes in telomere length, extensive evidence links adverse 90 early-life conditions to shorter telomeres (McLennan et al., 2016; Mizutani, Tomita, Niizuma, & Yoda, 91 2013; Watson, Bolton, & Monaghan, 2015), where shorter telomeres are associated with reduced 92 survival probability (Wilbourn et al., 2018). Food availability, often determined by weather conditions 93 (e.g. Campbell, Nouvellet, Newman, Macdonald, & Rosell, 2012), has been positively associated with 94 early-life telomere length (e.g. Foley et al., 2020; Spurgin et al., 2017). Interestingly, early-life food 95 availability may also impact life-history strategies (Bright Ross, Newman, Buesching, & Macdonald, 96 2020). It has been hypothesised that individuals in temporally stochastic environments should 97 modulate their energy trade-offs (Erikstad, Fauchald, Tveraa, & Steen, 1998; Reid, Bignal, Bignal, 98 McCracken, & Monaghan, 2003; Weimerskirch, Zimmermann, & Prince, 2001) and adopt a bet-99 hedging strategy (Wilbur & Rudolf, 2006). Since weather variability is predicted to increase in the 100 future (IPCC, 2018), it is important to understand the implications of variable early-life conditions for 101 life-history strategies and early-life telomere length. The interplay between the mean of and variability 102 in early-life environmental conditions, such as the availability and variation in food, foraging success 103 and thermal stress for young individuals (Noonan et al., 2015; Nouvellet, Newman, Buesching, &

Macdonald, 2013; Webb & King, 1984), can thus impact developmental stress, longevity and may be
 reflected in early-life telomere length.

106 Social conditions in early-life can also shape life-history strategies and senescence due to 107 increased competition for food and social stress. For example, female red deer (Cervus elaphus) that 108 experienced high levels of resource competition in early-life showed faster rates of reproductive 109 senescence (Nussey et al., 2007). Additionally, there is evidence for conspecific resource competition 110 in early-life leading to greater telomere shortening in birds (Boonekamp, Mulder, Salomons, Dijkstra, 111 & Verhulst, 2014; Nettle et al., 2015; Stier, Massemin, Zahn, Tissier, & Criscuolo, 2015), and shorter 112 telomere lengths in wild meerkats (Cram, Monaghan, Gillespie, & Clutton-Brock, 2017). Such patterns 113 can be explained because stressors (including competition) are associated with both shorter telomere 114 lengths and higher telomere shortening (Chatelain, Drobniak, & Szulkin, 2020).

115 The effects of social conditions on senescence may also become apparent after sexual 116 maturity, when individuals compete for mating opportunities (Andersson, 1994; Beirne, Delahay, & 117 Young, 2015). In polygynous species, sex differences in senescence may be attributable to intense 118 intra-sexual competition between males (Clutton-Brock & Isvaran, 2007; Promislow, 1992; Williams, 119 1957). Male investment for mating opportunities may trade off with self-maintenance (Kirkwood & 120 Rose, 1991). Intense male-male competition drives selection for shorter lifespan and faster 121 senescence in males, compared to females (Clutton-Brock & Isvaran, 2007; Williams, 1957). While this 122 prediction has been challenged (Bonduriansky, Maklakov, Zajitschek, & Brooks, 2008; Graves, 2007; 123 Promislow, 2003), and sex-specific senescence may be trait-dependent with respect to the underlying 124 physiological processes (Nussey et al., 2009), higher rates of male-biased actuarial senescence in 125 polygynous and sexual dimorphic species exist (Clutton-Brock & Isvaran, 2007; Promislow, 1992). 126 While social effects may also contribute to senescence in females (Sharp & Clutton-Brock, 2011; 127 Woodroffe & Macdonald, 1995), such sex-specific social effects on senescence are expected to be 128 greater in males (Bonduriansky et al., 2008; Clutton-Brock & Isvaran, 2007; Maklakov & Lummaa,

129 2013). However, whether increased intra-sexual competition (e.g. higher local densities of same-sex130 individuals) is associated with shorter telomere lengths remains to be tested.

131 To test the effects of early-life social and environmental conditions on telomere length, we 132 use a long-term dataset from a wild population of European badgers (Meles meles; henceforth 133 'badgers'). Badgers show reproductive senescence with males having a later onset but faster rate of 134 senescence than females (Dugdale, Pope, Newman, Macdonald, & Burke, 2011). Additionally, early-135 life telomere length (3–12 months old) positively correlates with first-year survival and lifespan in 136 badgers (van Lieshout et al., 2019). In the UK and Ireland, badgers are natally philopatric and can form 137 large social groups (mean group size = 11.3, range = 2-29; da Silva, Macdonald, & Evans, 1994) with 138 latrine-marked borders (Buesching, Newman, Service, Macdonald, & Riordan, 2016; Delahay et al., 139 2000), although they do transgress these borders when foraging (Ellwood et al., 2017; Noonan et al., 140 2015) without any sex difference in foraging niche (Robertson, McDonald, Delahay, Kelly, & Bearhop, 141 2014).

142 Regardless of whether badgers undergo facultative winter torpor (Johansson, 1957) or true 143 hibernation (Ruf & Geiser, 2015), badgers do reduce their body temperature by up to 8.9°C (Fowler & 144 Racey, 1988), thus reducing energy expenditure (Newman, Zhou, Buesching, Kaneko, & Macdonald, 145 2011). Badgers in Britain mainly feed on earthworms (Lumbricus terrestris; Johnson, Baker, Morecroft, 146 & Macdonald, 2001; Kruuk & Parish, 1981). Earthworms are sensitive to microclimatic conditions 147 (Edwards & Bohlen, 1996; Gerard, 1967; Newman, Buesching, & Macdonald, 2017), making their 148 abundance and distribution highly dependent on weather conditions. High-density badger 149 populations occur in mild areas with damp conditions where earthworms are available (Johnson, Jetz, 150 & Macdonald, 2002; Kruuk, 1978; Macdonald, Newman, & Buesching, 2015; Newman et al., 2017). 151 Foraging efficiency is reduced in adverse weather conditions, due to reduced availability of 152 earthworms, thermal stress when foraging in cold and wet conditions, and/or the choice to remain in 153 thermally-stable underground dens, termed setts (Noonan et al., 2014; Noonan et al., 2018; Nouvellet 154 et al., 2013; Tsunoda, Newman, Buesching, Macdonald, & Kaneko, 2018). Weather conditions can

therefore impact survival probability where, for example, higher annual mean daily rainfall is positively
associated with adult survival probability in badgers, whereas high annual variability in temperature
has detrimental consequences for cub and adult survival (Nouvellet et al., 2013).

158 Badgers have one litter per year, with a mean litter size of  $1.5 \pm 0.3$  (95% Cl; range = 1–5; 159 Annavi et al., 2014). Badger cub growth and maturation depends on the number of other cubs and 160 adults present within the social group (Sugianto, Newman, Macdonald, & Buesching, 2019), 161 potentially indicating resource competition within social groups. Adult male badgers invest substantial 162 energy into promiscuity and repeated mounting (Dugdale, Griffiths, & Macdonald, 2011) both within 163 and outside their social group, resulting in high rates (i.e. 48%) of extra-group paternity, of which 85% 164 were in neighbouring groups (Annavi et al., 2014; Dugdale, Macdonald, Pope, & Burke, 2007). Males 165 also exhibit substantial inter-individual variance in reproductive success (Dugdale et al., 2007; 166 Dugdale, Pope, et al., 2011) and evidence of reproductive skew among females within a group 167 (Dugdale, Macdonald, Pope, Johnson, & Burke, 2008; Woodroffe & Macdonald, 1995). With the 168 polygynandrous system (Dugdale, Griffiths, et al., 2011), a slight sexual dimorphism and slight male-169 biased mortality (Bright Ross et al., 2020; Johnson & Macdonald, 2001; Sugianto, Newman, 170 Macdonald, & Buesching, 2019), and evidence of downstream effects of male-male competition on 171 body mass senescence (Beirne et al., 2015), such intra-sexual competition may be reflected in 172 telomere length in early adulthood.

173 Here, we investigate the relationships between early-life conditions and relative leukocyte 174 telomere length (RLTL), by testing whether: (i) between-individual and within-individual variation in 175 RLTL in early life and early adulthood can be explained by seasonal changes; (ii) adverse early-life 176 weather, as a proxy for food availability and thermal stress, is associated with shorter early-life RLTL 177 and the social conditions that cubs are exposed to (with more cubs potentially leading to resource 178 competition and associated with shorter early-life RLTL, or more cubs reflecting more resources and 179 thus being associated with longer early-life RLTL); (iii) the strength of the association between early-180 life RLTL and first-year survival probability is dependent on early-life conditions and (iv) adverse social

181 conditions after sexual maturity (i.e. larger same-sex adult group size for females and, for males, more 182 within-group and neighbouring-group adult (>1 year old) males), are associated with shorter RLTL in 183 early post-maturity adulthood.

184

185 **2. Methods** 

### 186 (a) Study population and trapping

187 We conducted this study in a high-density population of badgers (mean  $\pm$  SE = 36.4  $\pm$  2.55 188 badgers/km2; Macdonald, Newman, Nouvellet, & Buesching, 2009) in Wytham Woods, Oxfordshire, 189 UK (51°46'24"N, 1°20'04"W); a 424 ha mixed semi-natural woodland surrounded by mixed arable and 190 permanent pasture (Macdonald et al., 2015). The population consisted of 19 ± 2 (mean ± 95% CI; range 191 = 14–26; Dugdale et al., 2008) mixed-sex social groups (Johnson, Jetz, et al., 2002; Newman et al., 192 2011) during the period that we analysed, with a 50% offspring sex ratio (Dugdale, Macdonald, & 193 Newman, 2003). The Wytham badger population is geographically discrete (Macdonald et al., 2009) 194 with only ca. 3% annual immigration/emigration per year (Macdonald & Newman, 2002).

195 We used long-term data (1987 – 2016) from a badger population that was trapped over three 196 two-week periods in May–June, August–September and November, with further trapping in January 197 in focal years (i.e. specific years when ultrasound studies were conducted to calculate implantation 198 dates, see Fig. 1). Badgers were anaesthetised using an intra-muscular injection of 0.2 ml ketamine 199 hydrochloride per kg body weight (McLaren et al., 2005). Upon first capture, badgers were assigned a 200 unique inguinal tattoo for permanent identification. Sex, age class (cub <1 year old; adult ≥1 year old), 201 capture date and social group were recorded. Age of badgers was defined as the number of days 202 elapsed since the 14<sup>th</sup> of February, reflecting the average date of synchronised parturition, in the 203 respective birth year (Yamaguchi, Dugdale, & Macdonald, 2006). Age of badgers first caught as adults 204 was inferred from tooth wear, which is commonly used and highly correlated ( $r^2 = 0.80$ ) with known 205 age in this population (Bright Ross et al., 2020; da Silva & Macdonald, 1989; Hancox, 1988; Macdonald 206 et al., 2009). Only badgers that did not have an already-known age and had a tooth wear of 2 (on a 1–

5 scale) were included since these typically indicate a 1-year old adult (Bright Ross et al., 2020). We
used data on cohorts up to and including 2010, as all cohort members were dead by the end of 2016.
Whole blood samples were collected from anaesthetised badgers through jugular venipuncture into
vacutainers with an EDTA anticoagulant, and stored immediately at -20°C. Badgers were released after
full recovery from anaesthesia. Additionally, bait-marking (Delahay et al., 2000; Macdonald &
Newman, 2002) was conducted periodically to delimit group range sizes and deduce social groups.

213

### 214 (b) Telomere analyses

215 Genomic DNA was extracted from whole blood samples (n = 814 samples; 533 badgers) using the 216 DNeasy Blood & Tissue kit (Qiagen, Manchester, UK) according to the manufacturer's protocol, with 217 changes by conducting a double elution step (2x 75  $\mu$ l AE buffer) and using 125  $\mu$ l of anticoagulated 218 blood. DNA integrity was checked by running a random selection of DNA extracts (ca. 20%) on agarose 219 gels to ensure high molecular weight. DNA concentration of all samples was quantified using the 220 Fluostar Optima fluorometer (BMG Labtech, Ortenberg, Germany) and standardized to 20 ng/ $\mu$ l, after 221 which samples were stored at -20 °C. We used monochrome multiplex quantitative PCR (MMqPCR) 222 analysis to measure RLTL (Cawthon, 2009). This is a measure that reflects the abundance of telomeric 223 sequence relative to a reference gene, which are both analysed in the same well, and although subject 224 to error represents the mean telomere length across cells in a sample. We used a sub-set of 814 samples from the full dataset of 1248 samples detailed in van Lieshout et al. (2019). In the full dataset, 225 226 Cq-values on the qPCR plates (n = 34) declined in a log-linear fashion ( $r^2 > 0.99$ ). Reaction efficiencies 227 were (mean  $\pm$  SE) 1.793  $\pm$  0.004 for IRBP and 1.909  $\pm$  0.004 for telomeres. Inter-plate repeatability 228 (intraclass correlation coefficient) calculated with rptr 0.9.2 (Stoffel, Nakagawa, & Schielzeth, 2017) -229 by comparing variance among duplicates of the reference sample within a plate, to variance of the 230 reference sample among plates – was 0.82 for RLTL measurements (95% CI = 0.76-0.87; n = 142231 samples; 34 plates). Intra-plate repeatability calculated with duplicates of the same sample on the 232 same plate, while controlling for plate effects, was 0.90 (95%CI = 0.86–0.93; n = 1,248 samples; 34

plates) for IRBP, 0.84 (95%CI = 0.79–0.90; n = 1,248 samples; 34 plates) for telomere Cq-values and
0.87 (95% CI = 0.82–0.91; n = 1,248 samples; 34 plates) for RLTL measurements. A detailed description
of the MMqPCR analysis can be found in van Lieshout et al. (2019).

236

### 237 (c) Weather conditions

238 Four weather metrics (mean daily temperature, temperature variability, mean daily rainfall, and 239 rainfall variability) were calculated for each season (Spring = end of March to end of June, Summer = 240 end of June to end of September, Autumn = end of September to end of December, Winter = end of 241 December to end of March) from 1987 to 2010 to characterise the developmental stress associated 242 with variation in earthworm food availability and thermoregulatory costs (Macdonald, Newman, 243 Buesching, & Nouvellet, 2010; Noonan et al., 2014; Nouvellet et al., 2013). Wytham Woods had a 244 mean annual temperature of 10.6 °C (± 5.5 SD) and mean annual precipitation of 684 (± 129 SD) mm, 245 1987–2010. Mean daily temperature and rainfall were calculated using mean daily temperature and 246 total daily precipitation values provided by the Radcliffe Meteorological Station, School of Geography, 247 University of Oxford (6 km from the field site). Daily temperatures followed a sinusoidal pattern, and 248 so seasonal temperature variability was calculated as the sum of daily squared residuals from a 249 sinusoidal fit to the corresponding year's temperatures (i.e. cumulative unpredictability). Rainfall did 250 not show annual trends and its seasonal variability was therefore characterised simply as the 251 coefficient of variation (SD/mean) in daily rainfall.

252

### 253 (d) Group sizes

Natal group sizes were determined by the number of individuals (cubs and adults) that were present in a social group in the year of an individual's birth. Given high lifetime natal philopatry (35.8%), low permanent dispersal rates (19.1%), and high levels of short-term inter-group movements (Macdonald, Newman, Buesching, & Johnson, 2008), individuals (*n* = 1726) were assigned as a resident of a social group each year, according to published criteria (van Lieshout, Badás, et al., 2020). The number of

individuals in a natal social group was then calculated as the sum of individuals present in the socialgroup in that year.

261 Yearly social group size measures were then separated by age class (i.e. cub/adult) and sex 262 (i.e. male/female) to determine sex- and age-specific group sizes per year. To measure intra-sexual 263 competition in females, we calculated female adult group sizes, as females compete with other within-264 group females (Woodroffe & Macdonald, 1995). However, for males, extra-group paternity is high 265 (48%) and affected by the number of within-group and neighbouring-group candidate fathers (Annavi et al., 2014), so we combined both the number of within-group males and neighbouring-group males. 266 267 The mean number of cubs in a social group for badgers in our dataset (n = 533 badgers) was 3.4 ( $\pm 2.3$ 268 SD; range 0–14), the mean number of female adults in a social group was 6.1 (± 3.4 SD; range 0–19) 269 and the mean number of male adults in focal plus neighbouring social groups was 25.2 (± 11.5 SD; 270 range 1–59).

271

### 272 (e) Statistical analyses

273 Statistical analyses were conducted in R 3.3.1 (R Development Core Team, 2020), using parametric 274 bootstrapping (n = 5000) to estimate 95% confidence intervals and determine significance of 275 predictors in Ime4 1.1-14 (Bates, Machler, Bolker, & Walker, 2015). Model fit was assessed using 276 standard residual plot techniques to ensure approximately normal distribution and constant variance, 277 and fixed effects were ensured not to be collinear (VIF < 3). Relative leukocyte telomere length (RLTL) 278 as response variable was first square-root and then Z-transformed (mean = 0, SD = 1) for comparability 279 (Verhulst, 2020). Quadratic fixed effects were included if such relationships were plausible a priori, 280 and removed if p > 0.1 to test the significance of first-order effects.

In this study, we focus on early-life (3–12 months old), but badgers typically reach sexual maturity by 2 years of age (Sugianto et al., 2019), occasionally at age 1 year (Dugdale et al., 2007). Due to delayed implantation resulting in a full year between conception and parturition, badgers thus first

produce offspring when they are 2–3 years of age, therefore we define early adulthood as 12–36
months old.

286

### 287 (i) Seasonal effects on RLTL in early-life and early adulthood

288 We first tested for an association between season and RLTL (≤36 months old) in early-life and early-289 adulthood in a Gaussian distribution model (identity link function) with RLTL as the response variable 290 (*n* = 814 samples; 533 badgers). Including threshold functions of age at 29 months, such that the slope 291 of the regression of RLTL with age differed for ≤29 months and >29 months of age best explained the 292 relationship between RLTL and age (van Lieshout et al., 2019). Threshold age, age at last capture, 293 season, weight and body length were included as fixed effects, and qPCR plate, row on qPCR plate, 294 social group, cohort (i.e. birth year; 24 levels), year and individual ID as random effects as these may 295 impact RLTL in badgers (van Lieshout, Sparks, et al., 2020).

296 As we found a significant cross-sectional difference in RLTL between spring and winter, we 297 then applied the 'within-subject centring' approach described by van de Pol and Wright (2009) to 298 distinguish within- and between-individual effects between spring and winter. Following (Schroeder, 299 Nakagawa, Cleasby, & Burke, 2012), we included two new fixed effects: 1) to estimate the within-300 individual variation component ( $\beta_w$ ) we removed between-individual variation by subtracting the 301 mean season value (coded as: spring = 0, winter = 1) for each individual across all years, from the 302 season value for each RLTL measurement. So, if an individual was measured once in spring and once 303 in winter, it was scored as -0.5 for spring and 0.5 for winter; and, 2) to estimate the between-individual 304 variation between seasons ( $\beta_B$ ), we included the mean season value for each individual (van de Pol & 305 Wright, 2009). We then ran a Gaussian distribution model (identity link function) with RLTL as the 306 response variable (n = 503 samples; 402 badgers) and threshold age (van Lieshout, 2019), age at last 307 capture, within-individual season effect ( $\beta_W$ ), between-individual season effect ( $\beta_B$ ), weight and body 308 length as fixed effects, and qPCR plate, row on qPCR plate, social group, cohort, year and individual ID 309 as random effects. Subsequently, we tested whether the within-individual ( $\beta_w$ ) and between310 individual ( $\beta_B$ ) slopes differed by including season and the between-individual effect ( $\beta_B$ ; i.e. mean 311 season value) in the same model (i.e. season now reflects the within-individual effect).

312 Lastly, to test whether telomere length decreases or increases from spring to winter we used 313 a subset of individuals measured either in their first spring or first winter, plus 11 individuals measured 314 in both their first spring and first winter (n = 214 samples; 203 badgers). For the direction of the effect 315 from winter to spring we used a subset of individuals measured either in their first winter or second 316 spring, plus 6 individuals measured in both their first winter and second spring (n = 84 samples; 78 317 badgers). In the two models (spring to winter and winter to spring) with a Gaussian distribution and 318 RLTL as the response variable, we included age, age at last capture, season, weight and body length 319 as fixed effects, and qPCR plate, row on qPCR plate, social group, cohort, year (not in winter to spring 320 model due to singularity) and individual ID as random effects. Subsequently, we used the within-321 subject centring approach again to separate within- and between-individual effects and test whether 322 these slopes differ (van de Pol & Wright, 2009).

323

### 324 (ii) Weather and natal group size effects on early-life RLTL

325 We tested whether weather and social conditions experienced as a cub (3-12 months old) were 326 associated with early-life RLTL. We first used a GLMM to confirm the previous observation (van 327 Lieshout et al., 2019) that early-life RLTL did not vary with age (in months), controlling for season, 328 weight and body length (n = 406,  $\beta = 0.154$ , 95% CI = -0.158–0.464), and excluded age from subsequent 329 analyses. The effects of first-year conditions on early-life RLTL were then modelled with early-life RLTL 330 as the response variable in a Gaussian-distributed model (identity link function; *n* = 406, samples; 406 331 badgers). First, we determined the season in which the weather conditions (i.e. mean temperature, 332 mean rainfall, temperature variability and rainfall variability) best explained the variation in early-life 333 RLTL (AICc spring = 1133.1 was lowest, versus summer  $\Delta$ AICc = 11.3, autumn  $\Delta$ AICc = 10.3, winter 334  $\Delta$ AlCc = 11.0), with models with  $\Delta$ AlCc <7 from the top model being plausible (Burnham, Anderson, & 335 Huyvaert, 2011). The weather window of spring (end of March to end of June) is the season in which 336 cubs grow the most and thus encounter the strongest developmental stress. This period includes when 337 cubs first emerge above ground from the end of February, are weaned around mid-May, and reach 338 independence at the start of June (Dugdale, Ellwood, & Macdonald, 2010) during which time cubs 339 exhibit high growth rates depending on food availability and social conditions (Sugianto et al., 2019). 340 Secondly, we determined whether the number of cubs, adults or the total number of individuals in 341 the natal group best predicted early-life RLTL using AICc (the lowest AICc = 1133.1 was for number of 342 cubs, versus number of adults  $\Delta$ AICc = 3.8, total number of individuals  $\Delta$ AICc = 4.0, number of cubs 343 plus number of adults  $\Delta$ AICc = 5.8, number of cubs plus total number of individuals  $\Delta$ AICc = 5.6). Since 344  $\Delta$ AlCc <7, and VIF>3 for the other combinations in the same model, we ran five separate models with 345 either the number of cubs, number of adults, the total number of individuals, number of cubs plus 346 adults or number of cubs plus total number of individuals in the natal group as a fixed effect along 347 with season, weight, body length, and mean daily temperature, temperature variability, mean daily 348 rainfall and rainfall variability in spring. qPCR plate, row on qPCR plate, social group and cohort were 349 included as random effects.

350

### 351 (iii) Covariation between early-life RLTL and weather conditions on cub survival probability

352 To understand whether the association between early-life RLTL and cub survival probability (van 353 Lieshout et al., 2019) is due to or independent of weather effects, we tested whether the association 354 between early-life RLTL and cub survival probability was still detected when social and weather 355 conditions were included in the model. We first modelled survival to adulthood (≥1 year old) as a 356 binary term in a binomially distributed model (logit link function; *n* = 406 samples; 406 badgers), where 357 cubs only caught in their first year of life were coded as 0 and cubs that were caught when older than 358 1 year of age were coded as 1, with early-life RLTL, weight and body length as fixed effects and qPCR 359 plate, row on qPCR plate, social group and cohort were included as random effects. We then also 360 included as fixed effects: number of cubs in the natal group, mean daily temperature, temperature 361 variability, mean daily rainfall and rainfall variability in a given season. We determined the season in

362 which weather conditions best explained the variation in cub survival probability, using AICc (the 363 lowest AICc = 408.9 was in winter, versus spring  $\Delta$ AICc = 21.6, summer  $\Delta$ AICc = 16.3 and autumn  $\Delta$ AICc 364 = 22.5) where models with  $\Delta$ AICc <7 from the top model are plausible (Burnham et al., 2011). We 365 checked whether the model was overdispersed. While cub survival is negatively impacted by 366 endoparasitic coccidia infection (Newman, Macdonald, & Anwar, 2001), we did not have data to 367 control for coccidia infection. We then applied model selection to test whether including weather and 368 social variables knocked early-life RLTL out of the plausible models. This would indicate that the early-369 life RLTL and survival probability relationship is driven by covariation between the environment and 370 physiological state (early-life RLTL). As early-life RLTL was retained, we estimated the RLTL model-371 averaged parameter and 95% confidence interval using the natural averaged method (where the 372 parameter was averaged over models in which it was present; Burnham & Anderson, 2002). This 373 avoids the parameter estimate shrinking towards zero, from inclusion of the relatively less important 374 models where the parameter was not retained (Nakagawa & Freckleton, 2011).

375

### 376 (iv) Same-sex group size effects on RLTL in early adulthood

377 We examined whether same-sex adult group sizes were reflected in RLTL in early adulthood (i.e. 12-378 36 months old). In a GLMM with RLTL in early adulthood as the response variable with one age 379 threshold separating two periods of 12 to ≤29 months and >29 and ≤36 months (see van Lieshout et 380 al., 2019) and season, weight and body length as fixed effects, we determined that RLTL did not vary 381 with age (n = 376, 12 to  $\leq 29$  months,  $\beta = -0.064$ , 95%Cl = -0.175-0.050; >29 and  $\leq 36$  months,  $\beta = -0.040$ , 382 95%CI = -0.184–0.110), and excluded age from the subsequent analysis. The effects of same-sex adult 383 group sizes on RLTL in early adulthood were then modelled with RLTL in early adulthood as the 384 response variable (*n* = 376 samples; 308 badgers). Same-sex adult group size (within-group for females 385 and within- plus neighbouring-group for males), sex and its interaction with group size (to model 386 differential strength in intra-sexual competition among the sexes), age at last capture (to control for 387 selective disappearance), season, weight and body length were included as fixed effects, and qPCR

388 plate, row on qPCR plate, social group, cohort, year and individual ID as random effects.

389

**390 3. Results** 

### 391 (i) Seasonal effects on RLTL in early-life and early adulthood

392 When controlling for age, weight and body length, we found a significant effect of season on RLTL with 393 badgers having shorter RLTL in winter compared to spring (Figure 1; Table S1). After partitioning the 394 within- and between-individual effects we found that there was a within-individual effect of shorter 395 RLTL in winter than in spring (Table S2). There was no significant difference between the within- and 396 between- individual slopes (Table S3), and thus, the significant between-individual effect (Table S2) 397 was driven by the within-individual effect. Using a subset of individuals measured only at consecutive 398 seasons, combined with individuals measured once, we found that from spring to winter there was a 399 within-individual decline in RLTL (Table S4 & Figure S1), whereas from winter to the following spring 400 there was a marginally non-significant within-individual increase in RLTL (Table S5 & Figure S1). For 401 both spring to winter and winter to spring the slopes for within- and between-individual effects did 402 not differ (Table S6).

403

### 404 (ii) Weather and natal group size effects on early-life RLTL

405 We found a positive association between spring temperature and early-life RLTL (Figure 2; Table 1 & 406 S7–S10), with cubs experiencing cooler-than-average first springs having shorter early-life RLTL. We 407 also found that cubs experiencing intermediate-to-high mean daily rainfall had longer early-life RLTL 408 (Figure 3; Table 1 & S7–S10) than cubs developing during drier years. Cubs experiencing low rainfall 409 variability also had longer early-life RLTL (Figure 4; Table 1 & S7–S10). We found, while controlling for 410 weather effects, a marginal effect where more cubs in the natal group leads to longer early-life RLTL. 411 In contrast, we found no evidence for an association between the number of adults or total number 412 of individuals in the natal group and early-life RLTL (Table 1 & S7–S10).

413

### 414 (iii) Covariation between early-life RLTL and weather conditions on cub survival probability

415 We first replicated our published finding (van Lieshout et al., 2019) of a positive association between 416 early-life RLTL and survival to adulthood, not controlling for social and weather effects (Table S11). 417 Then we included social and weather conditions in the model: cub survival probability exhibited a negative quadratic relationship with mean daily temperature (Figure S2; Table S12), a negative 418 419 quadratic association with winter temperature variability (Figure S3; Table S12), a marginal non-420 significant positive effect of mean daily rainfall (Table S12), a negative association with winter rainfall 421 variability (Figure S4; Table S12) but no significant effect of the number of cubs in a group (Table S12). 422 Using model selection, early-life RLTL was present in the top 39 models and retained in 82/100 423 plausible models (Table S13). The naturally averaged estimate for RLTL in the plausible models was 424 0.366 (95% CI = 0.064 - 0.666; Table S14) and thus the 95% CIs of early-life RLTL overlapped between 425 the models with and without ( $\beta$  = 0.386, 95% CI = 0.095 to 0.713, Table S11) early-life social and 426 weather variables.

427

### 428 (iv) Same-sex group size effects on RLTL in early adulthood

We found no evidence of same-sex adult group size effects on RLTL in early adulthood for females ormales (Table S15).

431

### 432 **4. Discussion**

Our results show both between-individual variation and within-individual changes in RLTL across seasons, where a cub's RLTL in their first spring was longer than in the following winter, and an indication that RLTL was longer again in the following spring compared to the preceding winter. Since the between- and within-individual slopes did not differ, the between-individual effect is driven by within-individual change and not selective (dis)appearance. We also found that cubs born in conditions that were warmer and wetter, with little variation in rainfall, had longer early-life RLTL. Sociologically, the number of cubs had a positive effect on early-life RLTL, in contrast to the number of adults or total number of individuals. Our results also suggest that the link between early-life RLTL and cub survival probability is driven by conditions experienced in addition to the early-life social and weather conditions modelled. Additionally, we found no effect of the number of within-group adult females, or both within-group and extra-group adult males (i.e. intra-sexual competition) on RLTL in early adulthood.

445 Our finding that badgers had shorter early-life RLTL (both between and within individuals) in 446 winter, compared to the preceding spring could be linked to the end-replication problem and stressful 447 effects such as disease (Newman et al., 2001), sub-optimal foraging conditions and food availability 448 (Macdonald & Newman, 2002; Newman et al., 2017). The within-individual effect means that between 449 seasons there is an increase or decrease in telomere length for the same individual. Since there is no 450 difference in the slopes for the within-individual and between-individual effect there is no selective 451 disappearance of individuals and the between-individual effect is driven by within-individual changes. 452 We then found a non-significant trend for positive within-individual changes in RLTL from the first 453 winter to the following spring. Body temperatures in badgers fall from November to December (by a 454 maximum of 8.9 °C compared to late-spring) and steadily rise until euthermic levels are reached by 455 late April (Fowler & Racey, 1988; Geiser & Ruf, 1995). During harsh winter conditions, badgers use 456 facultative torpor to reduce their core temperature and metabolic rate, conserving energy (Newman 457 et al., 2011). This reduction of basal metabolic rates (Geiser, 2004) can reduce mitosis (Kruman, 458 Ilyasova, Rudchenko, & Khurkhulu, 1988) and therefore potentially reduce telomere shortening. 459 Similarly, daily torpor cycles in Djungarian hamsters had a positive effect on telomere length (Turbill 460 et al., 2012). However, for species using torpor as a seasonal energy conservation strategy (e.g. edible 461 dormice, garden dormice, and arctic ground squirrels; as do badgers), arousal and return to euthermia 462 has been linked to telomere shortening; although this appears to be in proportion to the extent that 463 body temperature must be re-warmed (Giroud et al., 2014; Hoelzl, Cornils, Smith, Moodley, & Ruf, 464 2016; Turbill et al., 2013; Turbill et al., 2012; Wilbur et al., 2019). We postulate that badgers use torpor

465 and their ability to remain within thermally stable setts (Tsunoda et al., 2018) to try to mitigate RLTL 466 shortening that would otherwise be incurred by the stresses of maintaining activity during winter, 467 when food is scarce and thermal losses are high. More detailed analyses are needed to explore this 468 further, for example, comparing badgers in different regions that experience different degrees of 469 winter severity, with a large longitudinal sample size to disentangle within- and between-individual 470 effects. Importantly, we would need to track which badgers go into torpor, for how long and how 471 often, and then calculate how much energy is conserved. We also do not yet know to what extent 472 torpor-arousal cycles may affect telomere shortening, and where there is likely an optimal balance. In 473 this regard, predicted increases in weather variability (IPCC 2018) that may cause more frequent 474 warm-cold winter episodes, could add to the allostatic load of badgers, causing accelerated RLTL 475 shortening. Since positive within-individual changes in badger telomere length occur, that are greater 476 than measurement error (van Lieshout et al., 2019), such seasonal patterns may explain some of the 477 variability in telomere length patterns across life in badgers. Indeed, there is also evidence of seasonal 478 telomere dynamics in non-hibernating rodents (Criscuolo et al., 2020). Even though we accounted for 479 body weight and length, other factors such as seasonal changes in leukocyte cell composition can also 480 lead to apparent changes in telomere length (Beaulieu, Benoit, Abaga, Kappeler, & Charpentier, 2017), 481 which would require further investigation. For example, there is a greater proportion of neutrophils 482 and lymphocytes that were lymphocytes in spring compared to autumn in badgers (van Lieshout, 483 Badás, et al., 2020), and lymphocytes have shorter telomere lengths than neutrophils in humans and 484 baboons (Baerlocher, Rice, Vulto, & Lansdorp, 2007; Kimura et al., 2010). Nonetheless, our findings 485 also highlight the importance of controlling for seasonal effects when analysing telomere dynamics.

Cubs born into more energetically favourable springs (warm, rainy, and low rainfall variability) had longer early-life RLTL. These weather conditions present optimal soil conditions for earthworm surfacing, enhancing food supply (Kruuk, 1978; Newman et al., 2017). Dry conditions in spring have negative consequences for badger foraging success (Macdonald & Newman, 2002). However, while we found no effect of spring temperature variability on early-life RLTL, cubs experiencing lower daily 491 rainfall variability in spring had longer early-life RLTL. Greater rainfall variability can reduce the 492 predictability of food availability and impact foraging activity (Noonan et al., 2014), and may require 493 individuals to modulate their energy trade-offs (Erikstad et al., 1998; Reid et al., 2003; Weimerskirch 494 et al., 2001) and adopt a bet-hedging strategy (Wilbur & Rudolf, 2006). The variability in spring rainfall 495 and thus early-life conditions experienced shape life-history trade-offs, and since variability is likely to 496 increase under current climate change (IPCC, 2018), this can impact ecological and individual resilience 497 (Bright Ross et al., 2020).

498 Our estimate of post-dependence social effects was positive. An explanation for this positive 499 effect may be that in badgers, variation in maternal capacity to lactate may exceed the low variation 500 that is observed in litter size (Dugdale et al., 2007), causing the per-offspring suckling rate to increase 501 with litter size. In contrast, in other species or experimental brood size enlargements in birds, variation 502 in clutch size can exceed variation in parental resource acquisition, causing the per-offspring feeding 503 rate to decrease with litter size (van Noordwijk & de Jong, 1986; Vedder, Verhulst, Bauch, & Bouwhuis, 504 2017; Wilson & Nussey, 2010). An increase in the per-offspring suckling rate with litter size could result 505 in more available resources for cubs and thus longer early-life telomere length. Secondly, groups with 506 more independent cubs may also potentially have more food available per capita which permits faster 507 growth and cell replication without inducing stress, hence facilitating longer early-life telomere length. 508 This result is in contrast with studies reporting that competition for food within litters and juvenile 509 cohorts can cause telomere shortening (Boonekamp et al., 2014; Cram et al., 2017; Nettle et al., 2015). 510 However, these studies were able to measure telomere length within the first month of life. In 511 contrast, we were unable to sample individuals until at least 3 months of age, due to welfare 512 legislation (Protection of Badgers Act, 1992), when the weakest cubs could have already succumbed, 513 reducing group sizes. We therefore do not have a measure of the number of dependent cubs in a 514 group and could only measure RLTL in the first year from 3–12 months of age; thus, we cannot test 515 for social effects during the dependent period, including selective disappearance which may also lead 516 to similar positive associations between the number of cubs and early-life RLTL.

517 We found that the association between early-life RLTL and cub survival probability was 518 retained in the top 39 most plausible models and 82/100 plausible models when including early-life 519 weather and social variables. This indicates that, in badgers, the association between early-life RLTL 520 and survival is not solely driven by covariation between the early-life environment and early-life RLTL 521 (i.e. physiological state). While early-life RLTL in badgers appears to reflect the physiological 522 consequences of conditions experienced, independent of the weather and social variables included in 523 the models, there could still be a genetic component to telomere length or telomere length may 524 genetically covary with survival as seen in other species (Froy et al., 2021; Vedder et al., 2021). 525 Nonetheless, in badgers telomere length can be used as a comprehensive measure of the 526 environmental consequences for physiology and first-year survival probability.

527 There was no significant association between same-sex adult group size and RLTL in early 528 adulthood. While female–female reproductive competition occurs in badgers (Sharp & Clutton-Brock, 529 2011; Woodroffe & Macdonald, 1995), in polygynous species, theory predicts intra-sexual competition 530 for mating opportunities to be stronger among males than females. In Wytham badgers, there is slight 531 sexual dimorphism (Johnson & Macdonald, 2001) and slight male-biased mortality (Bright Ross et al., 532 2020). Reproductive skew is higher in sexually-mature males than females (Dugdale et al., 2008) and 533 males with a higher body-condition index attain more reproductive success (Dugdale, Griffiths, et al., 534 2011). High levels of polygynandrous and repeated mounting behaviour may however reduce male-535 male aggression and infanticide from males (Dugdale, Griffiths, et al., 2011; Wolff & Macdonald, 536 2004). Secondly, cryptic female choice (i.e. delayed implantation, superfecundation and 537 superfetation) may promote sperm competition and mask paternity, and reduce pre-copulatory 538 male-male competition (Birkhead & Pizzari, 2002). Finally, group size and/or density could be a poor 539 metric for competition due to foraging niche variation or variation in sex-ratio; additionally, although 540 the resource dispersion hypothesis predicts that groups approximate territorial carrying capacity, 541 results are mixed (Revilla, 2003). In fact, in our study population results vary with year such that only 542 in some situations larger groups may have proportionally more resources available (Johnson et al.,

543 2001; Johnson, Kays, Blackwell, & Macdonald, 2002). In line with this, we found no evidence that 544 variation in telomere length is due to intra-sexual competition in early adulthood. Badger early-life 545 telomere length may reflect the consequences of the weather conditions experienced, with little 546 impact of early-adulthood social conditions. However, in bad quality years only females in good 547 condition breed, whereas in good quality years breeding success is related to status (Woodroffe & 548 Macdonald, 1995). We can therefore not exclude that there may only be female–female competition 549 in good years. Additionally, early-adulthood male-male competition impacts on body mass 550 senescence in a badger population at the Woodchester Park study population (Beirne et al., 2015). 551 While we detected no significant evidence of direct effects of early-adulthood intra-sexual 552 competition on telomere length, there may be downstream effects on senescence.

In conclusion, we demonstrate the importance of accounting for seasonal variation when analysing telomere dynamics because of potential decreases as well as increases in telomere length across seasons. We also evidence that early-life adversity is reflected in shorter early-life telomere lengths in badgers, where the physical (weather) and social environment predict early-life telomere length. When accounting for these environmental effects, the positive association between early-life telomere length and survival probability remains. We conclude that variation in telomere length in badgers reflects early-life conditions, and in addition to this predicts first year cub survival.

560

561 **Ethics** 

All work was approved by the University of Oxford's Animal Welfare and Ethical Review Board, ratified
by the University of Leeds, and carried out under Natural England Licenses, currently 2017-27589-SCISCI and Home Office Licence (Animals, Scientific Procedures, Act, 1986) PPL: 30/3379.

565

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576

### 577 Author contributions

578 This study was conceived by S.H.J.v.L, A.B., H.L.D; Samples were collected by S.H.J.v.L., C.N., C.D.B.,

579 D.W.M. and H.L.D.; S.H.J.v.L. conducted laboratory work with input from T.B., environmental metrics

580 were calculated by S.H.J.v.L, E.P.B, J.G.B. and statistical analyses were conducted by S.H.J.v.L with

581 input from E.P.B and H.L.D; The paper was written by S.H.J.v.L and H.L.D. with extensive input from all

authors. All authors gave final approval for publication.

583

## 584 Data accessibility

- 585 Data are available from the Dryad Digital Repository (https://doi.org/10.5061/dryad.3r2280gf5) (van
- 586 Lieshout et al., 2021)
- 587

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## 990 **Tables and figures**

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**Table 1:** Parameter estimates and 95% confidence intervals of fixed effects from a mixed model and parametric bootstrap tests of the number of cubs in the natal group, season and weather effects in spring on early-life (3–12 months old) relative leukocyte telomere length (Z-score) in European badgers (full model and with p > 0.10 2<sup>nd</sup> order effects removed).  $\beta$  = parameter estimate, S.E. = standard error, 95% CI = 95% confidence intervals; reference terms in brackets = reference level for factors. Significant parameter estimates (95% CI does not overlap zero) are in bold.

Parameter (reference level)	β	S.E.	95% CI
Intercept <sup>+</sup>	-0.009	0.118	-0.228 to 0.218
Number of cubs in natal group	0.106	0.052	0.008 to 0.206
Season (Spring)			
Summer	0.196	0.137	-0.072 to 0.464
Autumn	0.131	0.277	-0.409 to 0.656
Winter	-1.001	0.383	-1.741 to -0.232
Mean temperature	-4.036	3.767	-11.38 to 3.367
Mean temperature <sup>2</sup>	4.519	3.830	-3.089 to 11.94
Daily temperature variability	0.588	1.709	-2.849 to 3.850
Daily temperature variability <sup>2</sup>	-0.457	1.733	-3.780 to 3.044
Mean daily rainfall	-1.894	0.810	-3.473 to -0.267
Mean daily rainfall <sup>2</sup>	2.074	0.836	0.368 to 3.692
Daily rainfall variability	-3.911	2.019	-7.818 to 0.014
Daily rainfall variability <sup>2</sup>	3.790	2.041	-0.211 to 7.698
Weight	0.075	0.101	-0.124 to 0.275
Body length	-0.089	0.097	-0.274 to 0.100
Without non-significant 2 <sup>nd</sup> order e	effects		
Intercept <sup>++</sup>	0.009	0.110	-0.197 to 0.216
Number of cubs in natal group	0.100	0.051	0.001 to 0.200
Season (Spring)			
Summer	0.178	0.136	-0.088 to 0.446
Autumn	0.097	0.274	-0.447 to 0.614
Winter	-0.995	0.380	-1.726 to -0.237
Mean temperature	0.403	0.087	0.227 to 0.577
Daily temperature variability	0.135	0.095	-0.053 to 0.321
Mean daily rainfall	-1.225	0.559	-2.333 to -0.124
Mean daily rainfall <sup>2</sup>	1.356	0.555	0.255 to 2.467
Daily rainfall variability	-2.843	1.414	-5.563 to -0.112
Daily rainfall variability <sup>2</sup>	2.745	1.405	0.039 to 5.464
Weight	0.072	0.100	-0.124 to 0.272

Random effect estimates (variance): <sup>†</sup>qPCR plate (4.955\*10<sup>-2</sup>), Row on qPCR plate (1.861\*10<sup>-3</sup>), Social
 group (2.798\*10<sup>-2</sup>), Cohort (7.745\*10<sup>-2</sup>), Residual (7.537\*10<sup>-1</sup>)

1000 <sup>++</sup>qPCR plate (4.911\*10<sup>-2</sup>), Row on qPCR plate (2.895\*10<sup>-3</sup>), Social group (2.642\*10<sup>-2</sup>), Cohort 1001 (5.287\*10<sup>-2</sup>), Residual (7.572\*10<sup>-1</sup>)



 $\begin{array}{c} 1003\\ 1004 \end{array}$ 

**Figure 1:** Variation in early-life relative leukocyte telomere length (RLTL) among seasons in European

1005 badgers. The data distributions and probability densities are shown (n = 814 samples; 533 badgers -

1006 the sum of badgers in the plot is >533 due to repeated measures). Data were collected in 19 years, 1007 across 59 trapping periods. The line in the boxplot represents the median, with first and third quartiles,

1008 and whiskers represent 1.57 times the inter-quartile range.



1009 1010

Figure 2: The association between mean spring temperature and early-life relative leukocyte telomere 1011 length (RLTL). Raw data points (n = 406 samples; 406 badgers) are shown, and jittered for clarity on 1012 the amount of data. The fitted line represents the regression from the mixed model, and the 95% 1013 confidence intervals as shaded areas.



 $\begin{array}{c} 1014\\ 1015 \end{array}$ 

Mean daily rainfall (mm)

Figure 3: The association between mean daily rainfall in spring and early-life relative leukocyte 1016 telomere length (RLTL). Raw data points (n = 406 samples; 406 badgers) are shown, and jittered for 1017 clarity on the amount of data. The fitted line represents the quadratic regression from the mixed 1018 model, and the 95% confidence intervals as shaded areas.



 $\begin{array}{c} 1019\\ 1020 \end{array}$ 

Rainfall variability in spring  $(\sigma/\mu)$ 

Figure 4: The association between the rainfall variability in spring and early-life relative leukocyte 1021 telomere length (RLTL). Raw data points (n = 406 samples; 406 badgers) are shown, and jittered for 1022 clarity on the amount of data. The fitted line represents the quadratic regression from the mixed 1023 model, and the 95% confidence intervals as shaded areas.



## **Supplemental Information for:**

## Early-life seasonal, weather and social effects on telomere lengths in a wild mammal

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This supplementary materials document includes supplementary tables and figures.



**Table S1:** Parameter estimates and 95% confidence intervals (from parametric bootstrapping) of fixed effects from a mixed model of age and seasonal effects on relative leukocyte telomere length (Z-score) in European badgers.  $\beta$  = parameter estimate, S.E. = standard error, 95% CI = 95% confidence intervals; reference terms in brackets = reference level for factors. A threshold age of 29 months was modelled, such that two parameter estimates are generated for <29 months and >29 to <36 months of age. Significant parameter estimates (95% CI does not overlap zero) are in bold.

Parameter (reference level)	β	S.E.	95% CI
Intercept	-0.159	0.123	-0.400 to 0.085
Age ≤29 months	-0.049	0.064	-0.178 to 0.079
Age >29 and ≤ 36 months	0.018	0.038	-0.057 to 0.094
Age at last capture	0.040	0.036	-0.032 to 0.110
Season (Spring)			
Summer	-0.001	0.076	-0.157 to 0.153
Autumn	0.057	0.135	-0.198 to 0.315
Winter	-0.518	0.183	-0.875 to -0.172
Weight	0.042	0.035	-0.027 to 0.113
Body length	0.058	0.058	-0.058 to 0.170

Random effect estimates (variance): qPCR plate (5.301\*10<sup>-2</sup>), Row on qPCR plate (4.723\*10<sup>-3</sup>), Social group (7.617\*10<sup>-3</sup>), Cohort (6.470\*10<sup>-2</sup>), Year (1.175\*10<sup>-1</sup>), Individual ID (<1.000\*10<sup>-12</sup>), Residual (7.841\*10<sup>-1</sup>)

**Table S2:** Parameter estimates and 95% confidence intervals (from parametric bootstrapping) of fixed effects from a mixed model of age effects and both within-individual ( $\beta_W$ ; the mean season value for each individual subtracted from the season value for each RLTL measurement) and between-individual ( $\beta_B$ ; the mean season value for each individual) seasonal effects on relative leukocyte telomere length (Z-score) in European badgers.  $\beta$  = parameter estimate, S.E. = standard error, 95% CI = 95% confidence intervals. A threshold age of 29 months was modelled, such that two parameter estimates are generated for ≤29 months and >29 to ≤36 months of age. Significant parameter estimates (95% CI does not overlap zero) are in bold.

Parameter (reference level)	β	S.E.	95% CI
Intercept	-0.091	0.087	-0.260 to 0.078
Age ≤29 months	0.045	0.084	-0.112 to 0.210
Age >29 and ≤36 months	0.041	0.047	-0.054 to 0.136
Age at last capture	0.063	0.045	-0.026 to 0.152
Season (β <sub>w</sub> )	-0.109	0.042	-0.192 to -0.026
Season (β <sub>B</sub> )	-0.119	0.048	-0.212 to -0.026
Weight	0.026	0.049	-0.070 to 0.121
Body length	0.006	0.079	-0.150 to 0.161

Random effect estimates (variance): qPCR plate (2.404\*10<sup>-2</sup>), Row on qPCR plate (2.684\*10<sup>-3</sup>), Social group 3.325\*10<sup>-4</sup>), Cohort (6.338\*10<sup>-2</sup>), Year (1.480\*10<sup>-2</sup>), Individual ID (<1.000\*10<sup>-12</sup>), Residual (7.955\*10<sup>-1</sup>)

**Table S3:** Parameter estimates and 95% confidence intervals of fixed effects from a mixed model and parametric bootstrap tests of age effects and within-individual ( $\beta_{w}$ ; the mean season value for each individual subtracted from the season value for each RLTL measurement) and the difference in betweenand within-individual ( $\beta_{B} - \beta_{w}$ ) seasonal effects on relative leukocyte telomere length (Z-score) in European badgers.  $\beta$  = parameter estimate, S.E. = standard error, 95% CI = 95% confidence intervals. A threshold age of 29 months was modelled, such that two parameter estimates are generated for  $\leq$  29 months and >29 to  $\leq$ 36 months of age. Significant parameter estimates (95% CI does not overlap zero) are in bold.

Parameter (reference level)	β	S.E.	95% CI
Intercept	-0.036	0.089	-0.209 to 0.141
Age ≤29 months	0.045	0.084	-0.112 to 0.210
Age >29 and ≤36 months	0.041	0.047	-0.054 to 0.136
Age at last capture	0.063	0.045	-0.026 to 0.152
Season (β <sub>w</sub> )	-0.786	0.307	-1.384 to -0.187
Season (β <sub>B</sub> - β <sub>W</sub> )	0.055	0.078	-0.093 to 0.206
Weight	0.026	0.049	-0.070 to 0.121
Body length	0.006	0.079	-0.150 to 0.161

Random effect estimates (variance): qPCR plate (2.404\*10<sup>-2</sup>), Row on qPCR plate (2.684\*10<sup>-3</sup>), Social group (3.325\*10<sup>-4</sup>), Cohort (6.338\*10<sup>-2</sup>), Year (1.480\*10<sup>-2</sup>), Individual ID (<1.000\*10<sup>-12</sup>), Residual (7.955\*10<sup>-1</sup>)

**Table S4:** Parameter estimates and 95% confidence intervals of fixed effects from a mixed model and parametric bootstrap tests of the direction of the effect on relative leukocyte telomere length (Z-score) from a cross-sectional model and when separating within- from between-individual effects from spring to winter in European badgers.  $\beta$  = parameter estimate, S.E. = standard error, 95% CI = 95% confidence intervals. Significant parameter estimates (95% CI does not overlap zero) are in bold.

Parameter (reference level)	β	S.E.	95% CI
Intercept <sup>+</sup>	-0.001	0.123	-0.236 to 0.249
Age ≤29 months	0.289	0.144	0.004 to 0.579
Age at last capture	0.111	0.068	-0.022 to 0.249
Season (Spring)			
Winter	-1.431	0.488	-2.392 to -0.461
Weight	0.176	0.172	-0.181 to 0.519
Body length	-0.156	0.127	-0.400 to 0.092
Intercept <sup>++</sup>	-0.165	0.112	-0.380 to 0.062
Age ≤29 months	0.290	0.149	-0.007 to 0.586
Age at last capture	0.102	0.068	-0.030 to 0.238
Season (β <sub>w</sub> )	-0.287	0.096	-0.470 to -0.094
Season (β <sub>B</sub> )	-0.353	0.146	-0.643 to -0.068
Weight	0.158	0.177	-0.209 to 0.500
Body length	-0.153	0.128	-0.402 to 0.098

Random effect estimates (variance):

<sup>+</sup>qPCR plate (<1.000\*10<sup>-12</sup>), Row on qPCR plate (4.175\*10<sup>-3</sup>), Social group (3.878\*10<sup>-2</sup>), Cohort (6.284\*10<sup>-4</sup>), Year (8.181\*10<sup>-1</sup>), Individual ID (<1.000\*10<sup>-12</sup>), Residual (7.569\*10<sup>-1</sup>)

<sup>++</sup>qPCR plate (<1.000\*10<sup>-12</sup>), Row on qPCR plate (3.965\*10<sup>-3</sup>), Social group (3.830\*10<sup>-2</sup>), Cohort (5.964\*10<sup>-3</sup>), Year (7.920\*10<sup>-2</sup>), Individual ID (<1.000\*10<sup>-12</sup>), Residual (7.566\*10<sup>-1</sup>)

**Table S5:** Parameter estimates and 95% confidence intervals of fixed effects from a mixed model and parametric bootstrap tests of the direction of the effect on relative leukocyte telomere length (Z-score) from a cross-sectional model and when separating within- from between-individual effects from winter to spring in European badgers.  $\beta$  = parameter estimate, S.E. = standard error, 95% CI = 95% confidence intervals. Significant parameter estimates (95% CI does not overlap zero) are in bold.

Parameter (reference level)	β	S.E.	95% CI
Intercept <sup>+</sup>	-0.041	0.183	-0.399 to 0.308
Age ≤29 months	0.160	0.168	-0.172 to 0.498
Age at last capture	-0.086	0.123	-0.333 to 0.157
Season (Spring)			
Winter	-0.356	0.416	-1.173 to 0.486
Weight	0.055	0.185	-0.309 to 0.423
Body length	0.014	0.186	-0.356 to 0.369
Intercept <sup>++</sup>	-0.110	0.132	-0.380 to 0.143
Age ≤29 months	-0.124	0.238	-0.560 to 0.358
Age at last capture	-0.095	0.112	-0.323 to 0.132
Season (β <sub>w</sub> )	-0.408	0.217	-0.807 to 0.034
Season (β <sub>B</sub> )	-0.173	0.142	-0.454 to 0.113
Weight	-0.004	0.176	-0.334 to 0.344
Body length	0.042	0.188	-0.329 to 0.410

Random effect estimates (variance):

<sup>†</sup>qPCR plate (6.170<sup>\*10<sup>-2</sup></sup>), Row on qPCR plate (<1.000<sup>\*10<sup>-12</sup></sup>), Social group (1.599<sup>\*10<sup>-2</sup></sup>), Cohort (7.467<sup>\*10<sup>-2</sup></sup>), Individual ID (<1.000<sup>\*10<sup>-12</sup></sup>), Residual (7.587<sup>\*10<sup>-1</sup></sup>)

<sup>++</sup>qPCR plate (5.188\*10<sup>-2</sup>), Row on qPCR plate (<1.000\*10<sup>-12</sup>), Social group (<1.000\*10<sup>-12</sup>), Cohort (4.522\*10<sup>-2</sup>), Individual ID (<1.000\*10<sup>-12</sup>), Residual (7.555\*10<sup>-1</sup>)

**Table S6:** Parameter estimates and 95% confidence intervals of fixed effects from a mixed model and parametric bootstrap tests of the direction of the effect on relative leukocyte telomere length (Z-score) to test whether slopes differ between within- and between-individual slopes from spring to winter and winter to spring in European badgers.  $\beta$  = parameter estimate, S.E. = standard error, 95% CI = 95% confidence intervals. Significant parameter estimates (95% CI does not overlap zero) are in bold.

Parameter (reference level)	β	S.E.	95% CI
Intercept <sup>+</sup>	0.022	0.128	-0.225 to 0.279
Age ≤29 months	0.289	0.149	-0.007 to 0.589
Age at last capture	0.102	0.068	-0.030 to 0.239
Season (β <sub>w</sub> )	-1.676	0.563	-2.750 to -0.545
Season (β <sub>B</sub> - β <sub>W</sub> )	0.110	0.123	-0.129 to 0.349
Weight	0.158	0.177	-0.209 to 0.501
Body length	-0.153	0.128	-0.401 to 0.100
Intercept <sup>++</sup>	0.420	0.302	-0.213 to 0.963
Age ≤29 months	-0.143	0.228	-0.561 to 0.336
Age at last capture	-0.095	0.113	-0.322 to 0.132
Season (β <sub>w</sub> )	-2.229	1.121	-4.288 to 0.105
Season (β <sub>B</sub> - β <sub>W</sub> )	0.679	0.406	-0.153 to 1.436
Weight	-0.006	0.175	-0.337 to 0.337
Body length	0.043	0.187	-0.319 to 0.410

Random effect estimates (variance):

<sup>+</sup>qPCR plate (<1.000\*10<sup>-12</sup>), Row on qPCR plate (3.965\*10<sup>-3</sup>), Social group (3.830\*10<sup>-2</sup>), Cohort (5.964\*10<sup>-3</sup>), Year (7.920\*10<sup>-2</sup>), Individual ID (<1.000\*10<sup>-12</sup>), Residual (7.566\*10<sup>-1</sup>)

<sup>++</sup>qPCR plate (5.188\*10<sup>-2</sup>), Row on qPCR plate (<1.000\*10<sup>-12</sup>), Social group (<1.000\*10<sup>-12</sup>), Cohort (4.522\*10<sup>-2</sup>), Individual ID (<1.000\*10<sup>-12</sup>), Residual (7.554\*10<sup>-1</sup>)

**Table S7:** Parameter estimates and 95% confidence intervals of fixed effects from a mixed model and parametric bootstrap tests of the number of adults in the natal group, season and weather effects in spring on early-life (3–12 months old) relative leukocyte telomere length (Z-score) in European badgers (full model and with p > 0.10 2<sup>nd</sup> order effects removed).  $\beta$  = parameter estimate, S.E. = standard error, 95% CI = 95% confidence intervals; reference terms in brackets = reference level for factors. Significant parameter estimates (95% CI does not overlap zero) are in bold.

Parameter (reference level)	β	S.E.	95% CI
Intercept <sup>*</sup>	-0.037	0.117	-0.254 to 0.194
Number of adults in natal group	-0.028	0.053	-0.130 to 0.080
Season (Spring)			
Summer	0.238	0.136	-0.030 to 0.507
Autumn	0.197	0.278	-0.341 to 0.729
Winter	-1.003	0.386	-1.742 to -0.231
Mean temperature	-3.373	3.820	-10.81 to 4.218
Mean temperature <sup>2</sup>	3.827	3.883	-3.869 to 11.40
Daily temperature variability	0.044	1.714	-3.413 to 3.382
Daily temperature variability <sup>2</sup>	0.093	1.739	-3.317 to 3.580
Mean daily rainfall	-1.711	0.818	-3.309 to -0.060
Mean daily rainfall <sup>2</sup>	1.878	0.843	0.154 to 3.513
Daily rainfall variability	-3.671	2.048	-7.680 to 0.273
Daily rainfall variability <sup>2</sup>	3.528	2.070	-0.463 to 7.585
Weight	0.064	0.102	-0.134 to 0.267
Body length	-0.099	0.097	-0.284 to 0.091
Without non-significant 2 <sup>nd</sup> order ef	fects		
Intercept <sup>++</sup>	-0.022	0.107	-0.222 to 0.186
Number of adults in natal group	-0.032	0.053	-0.132 to 0.074
Season (Spring)			
Summer	0.220	0.135	-0.046 to 0.485
Autumn	0.162	0.276	-0.377 to 0.690
Winter	-0.996	0.382	-1.725 to -0.236
Mean temperature	0.400	0.087	0.223 to 0.574
Daily temperature variability	0.133	0.094	-0.054 to 0.321
Mean daily rainfall	-1.141	0.560	-2.250 to -0.046
Mean daily rainfall <sup>2</sup>	1.272	0.555	0.170 to 2.392
Daily rainfall variability	-3.038	1.424	-5.804 to -0.317
Daily rainfall variability <sup>2</sup>	2.938	1.414	0.216 to 5.670
Weight	0.058	0.102	-0.138 to 0.260
Body length	-0 090	0 097	-0 273 to 0 098

Random effect estimates (variance):

<sup>+</sup>qPCR plate (4.802\*10<sup>-2</sup>), Row on qPCR plate (2.073\*10<sup>-3</sup>), Social group (1.727\*10<sup>-2</sup>), Cohort (8.195\*10<sup>-2</sup>), Residual (7.687\*10<sup>-1</sup>)

<sup>++</sup>qPCR plate (4.721\*10<sup>-2</sup>), Row on qPCR plate (2.853\*10<sup>-3</sup>), Social group (1.562\*10<sup>-2</sup>), Cohort (5.253\*10<sup>-2</sup>), Residual (7.727\*10<sup>-1</sup>)

**Table S8:** Parameter estimates and 95% confidence intervals of fixed effects from a mixed model and parametric bootstrap tests of the total number of individuals in the natal group, season and weather effects in spring on early-life (3–12 months old) relative leukocyte telomere length (Z-score) in European badgers (full model and with p > 0.10 2<sup>nd</sup> order effects removed).  $\beta$  = parameter estimate, S.E. = standard error, 95% CI = 95% confidence intervals; reference terms in brackets = reference level for factors. Significant parameter estimates (95% CI does not overlap zero) are in bold.

Parameter (reference level)	β	S.E.	95% CI
Intercept <sup>+</sup>	-0.030	0.117	-0.249 to 0.202
Number of individuals in natal group	0.016	0.054	-0.086 to 0.122
Season (Spring)			
Summer	0.234	0.136	-0.032 to 0.501
Autumn	0.176	0.279	-0.366 to 0.707
Winter	-1.024	0.385	-1.762 to -0.253
Mean temperature	-3.558	3.818	-11.03 to 4.027
Mean temperature <sup>2</sup>	4.018	3.882	-3.672 to 11.59
Daily temperature variability	0.064	1.713	-3.419 to 3.379
Daily temperature variability <sup>2</sup>	0.072	1.738	-3.301 to 3.554
Mean daily rainfall	-1.719	0.816	-3.296 to -0.084
Mean daily rainfall <sup>2</sup>	1.893	0.842	0.170 to 3.527
Daily rainfall variability	-3.593	2.040	-7.596 to 0.339
Daily rainfall variability <sup>2</sup>	3.452	2.062	-0.506 to 7.490
Weight	0.076	0.102	-0.122 to 0.280
Body length	-0.105	0.097	-0.290 to 0.084
Without non-significant 2 <sup>nd</sup> order effe	ects		
Intercept <sup>**</sup>	-0.015	0.108	-0.218 to 0.193
Number of individuals in natal group	0.011	0.054	-0.091 to 0.115
Season (Spring)			
Summer	0.216	0.135	-0.051 to 0.482
Autumn	0.140	0.276	-0.404 to 0.668
Winter	-1.019	0.382	-1.750 to -0.260
Mean temperature	0.403	0.088	0.225 to 0.578
Daily temperature variability	0.133	0.095	-0.056 to 0.322
Mean daily rainfall	-1.127	0.562	-2.247 to -0.022
Mean daily rainfall <sup>2</sup>	1.264	0.558	0.170 to 2.374
Daily rainfall variability	-2.929	1.434	-5.687 to -0.186
Daily rainfall variability <sup>2</sup>	2.830	1.425	0.094 to 5.582
Weight	0.071	0.101	-0.124 to 0.272
Body length	-0 096	0 097	-0 277 to 0 093

Random effect estimates (variance):

<sup>+</sup>qPCR plate (4.782\*10<sup>-2</sup>), Row on qPCR plate (1.108\*10<sup>-3</sup>), Social group (2.372\*10<sup>-2</sup>), Cohort (8.127\*10<sup>-2</sup>), Residual (7.658\*10<sup>-1</sup>)

<sup>++</sup>qPCR plate (4.718\*10<sup>-2</sup>), Row on qPCR plate (1.910\*10<sup>-3</sup>), Social group (2.200\*10<sup>-2</sup>), Cohort (5.394\*10<sup>-2</sup>), Residual (7.695\*10<sup>-1</sup>)

**Table S9:** Parameter estimates and 95% confidence intervals of fixed effects from a mixed model and parametric bootstrap tests of the number of cubs and number of adults in the natal group, season and weather effects in spring on early-life (3–12 months old) relative leukocyte telomere length (Z-score) in European badgers (full model and with p > 0.10 2<sup>nd</sup> order effects removed).  $\beta$  = parameter estimate, S.E. = standard error, 95% CI = 95% confidence intervals; reference terms in brackets = reference level for factors. Significant parameter estimates (95% CI does not overlap zero) are in bold.

Parameter (reference level)	β	S.E.	95% CI
Intercept <sup>+</sup>	-0.014	0.118	-0.235 to 0.215
Number of cubs in natal group	0.109	0.052	0.008 to 0.209
Number of adults in natal group	-0.039	0.055	-0.142 to 0.071
Season (Spring)			
Summer	0.195	0.137	-0.073 to 0.463
Autumn	0.147	0.278	-0.394 to 0.668
Winter	-0.981	0.384	-1.715 to -0.214
Mean temperature	-3.942	3.782	-11.36 to 3.503
Mean temperature <sup>2</sup>	4.421	3.845	-3.180 to 11.96
Daily temperature variability	0.617	1.715	-2.789 to 3.909
Daily temperature variability <sup>2</sup>	-0.486	1.740	-3.815 to 2.970
Mean daily rainfall	-1.908	0.814	-3.493 to -0.282
Mean daily rainfall <sup>2</sup>	2.082	0.839	0.371 to 3.708
Daily rainfall variability	-4.007	2.033	-7.941 to -0.074
Daily rainfall variability <sup>2</sup>	3.887	2.055	-0.107 to 7.875
Weight	0.064	0.101	-0.133 to 0.267
Body length	-0.082	0.097	-0.267 to 0.106
Without non-significant 2 <sup>nd</sup> order effe	ects		
Intercept <sup>++</sup>	0.003	0.109	-0.204 to 0.211
Number of cubs in natal group	0.104	0.051	0.005 to 0.204
Number of adults in natal group	-0.042	0.054	-0.144 to 0.067
Season (Spring)			
Summer	0.178	0.136	-0.088 to 0.445
Autumn	0.116	0.275	-0.429 to 0.642
Winter	-0.972	0.381	-1.698 to -0.212
Mean temperature	0.400	0.087	0.224 to 0.573
Daily temperature variability	0.136	0.094	-0.051 to 0.322
Mean daily rainfall	-1.248	0.559	-2.358 to -0.144
Mean daily rainfall <sup>2</sup>	1.375	0.555	0.276 to 2.488
Daily rainfall variability	-2.942	1.418	-5.675 to -0.195
Daily rainfall variability <sup>2</sup>	2.843	1.409	0.110 to 5.576
Weight	0.060	0.101	-0.136 to 0.261
Body length	-0.076	0.097	-0.257 to 0.110

Random effect estimates (variance):

<sup>+</sup>qPCR plate (4.987\*10<sup>-2</sup>), Row on qPCR plate (2.843\*10<sup>-3</sup>), Social group (2.409\*10<sup>-2</sup>), Cohort (7.829\*10<sup>-2</sup>), Residual (7.561\*10<sup>-1</sup>)



<sup>++</sup>qPCR plate (4.929\*10<sup>-2</sup>), Row on qPCR plate (3.963\*10<sup>-3</sup>), Social group (2.182\*10<sup>-2</sup>), Cohort (5.223\*10<sup>-2</sup>), Residual (7.602\*10<sup>-1</sup>)

**Table S10:** Parameter estimates and 95% confidence intervals of fixed effects from a mixed model and parametric bootstrap tests of the number of cubs and total number of individuals in the natal group, season and weather effects in spring on early-life (3–12 months old) relative leukocyte telomere length (Z-score) in European badgers (full model and with p > 0.10 2<sup>nd</sup> order effects removed).  $\beta$  = parameter estimate, S.E. = standard error, 95% CI = 95% confidence intervals; reference terms in brackets = reference level for factors. Significant parameter estimates (95% CI does not overlap zero) are in bold.

Parameter (reference level)	β	S.E.	95% CI
Intercept <sup>†</sup>	-0.013	0.118	-0.234 to 0.215
Number of cubs in natal group	0.122	0.058	0.012 to 0.233
Number of individuals in natal group	-0.040	0.060	-0.156 to 0.080
Season (Spring)	010.0	01000	0.200 10 0.000
Summer	0.195	0.137	-0.073 to 0.463
Autumn	0.146	0.278	-0.395 to 0.669
Winter	-0.982	0.384	-1.714 to -0.215
Mean temperature	-3.920	3.779	-11.33 to 3.492
Mean temperature <sup>2</sup>	4.399	3.843	-3.172 to 11.95
Daily temperature variability	0.614	1.713	-2.788 to 3.903
Daily temperature variability <sup>2</sup>	-0.482	1.738	-3.821 to 2.981
Mean daily rainfall	-1.899	0.812	-3.482 to -0.273
Mean daily rainfall <sup>2</sup>	2.072	0.838	0.363 to 3.694
Daily rainfall variability	-4.000	2.031	-7.929 to -0.078
Daily rainfall variability <sup>2</sup>	3.879	2.053	-0.116 to 7.860
Weight	0.065	0.101	-0.132 to 0.267
Body length	-0.082	0.097	-0.266 to 0.106
Without non-significant 2 <sup>nd</sup> order effe	ects		
Intercept <sup>**</sup>	0.003	0.109	-0.203 to 0.211
Number of cubs in natal group	0.118	0.057	0.008 to 0.227
Number of individuals in natal group	-0.045	0.060	-0.158 to 0.075
Season (Spring)			
Summer	0.178	0.136	-0.088 to 0.445
Autumn	0.116	0.275	-0.428 to 0.640
Winter	-0.972	0.381	-1.699 to -0.215
Mean temperature	0.400	0.087	0.225 to 0.573
Daily temperature variability	0.136	0.094	-0.051 to 0.321
Mean daily rainfall	-1.242	0.558	-2.346 to -0.142
Mean daily rainfall <sup>2</sup>	1.367	0.553	0.270 to 2.474
Daily rainfall variability	-2.943	1.416	-5.674 to -0.209
Daily rainfall variability <sup>2</sup>	2.843	1.407	0.122 to 5.570
Weight	0.060	0.101	-0.136 to 0.261
Body length	-0.075	0.097	-0.256 to 0.110

Random effect estimates (variance):

<sup>+</sup>qPCR plate (4.980\*10<sup>-2</sup>), Row on qPCR plate (2.772\*10<sup>-3</sup>), Social group (2.460\*10<sup>-2</sup>), Cohort (7.799\*10<sup>-2</sup>), Residual (7.560\*10<sup>-1</sup>)



<sup>++</sup>qPCR plate (4.921\*10<sup>-2</sup>), Row on qPCR plate (3.905\*10<sup>-3</sup>), Social group (2.226\*10<sup>-2</sup>), Cohort (5.183\*10<sup>-2</sup>), Residual (7.601\*10<sup>-1</sup>)

**Table S11:** Parameter estimates and 95% confidence intervals of fixed effects from a mixed model and parametric bootstrap tests of the association between relative leukocyte telomere length and survival to adulthood (>1 year old) in European badgers.  $\beta$  = parameter estimate, S.E. = standard error, 95% CI = 95% confidence interval. Significant parameter estimates (95% CI does not overlap zero) are in bold.

Parameter (reference level)	β	S.E.	95% CI
Intercept	1.504	0.332	0.872 to 2.191
Relative leukocyte telomere length	0.386	0.153	0.095 to 0.713
Weight	0.023	0.244	-0.460 to 0.504
Body length	0.081	0.244	-0.392 to 0.580

Random effect estimates (variance): qPCR plate (<1.000\*10<sup>-12</sup>), Row on qPCR plate (1.033\*10<sup>-1</sup>), Social group (3.798\*10<sup>-1</sup>), Cohort (7.319\*10<sup>-1</sup>)

**Table S12:** Parameter estimates and 95% confidence intervals of fixed effects from a mixed model and parametric bootstrap tests of the association between relative leukocyte telomere length, winter weather conditions and survival to adulthood (>1 year old) in European badgers (full model and with p > 0.10 2<sup>nd</sup> order effects removed).  $\beta$  = parameter estimate, S.E. = standard error, 95% CI = 95% confidence interval. Significant parameter estimates (95% CI does not overlap zero) are in bold.

Parameter (reference level)	ß	S F	95% CI
	P 1 701	J.L.	
Intercept	1./81	0.301	1.239 to 2.399
Relative leukocyte telomere length	0.441	0.172	0.113 to 0.792
Number of cubs in natal group	-0.053	0.167	-0.386 to 0.292
Mean temperature	6.821	1.525	4.031 to 9.920
Mean temperature <sup>2</sup>	-6.619	1.466	-9.618 to -3.953
Daily temperature variability	-3.699	2.270	-8.346 to 0.680
Daily temperature variability <sup>2</sup>	4.236	2.381	-0.345 to 9.180
Mean daily rainfall	-0.671	1.133	-2.873 to 1.511
Mean daily rainfall <sup>2</sup>	1.636	1.319	-0.841 to 4.252
Daily rainfall variability	2.414	2.500	-2.563 to 7.346
Daily rainfall variability <sup>2</sup>	-3.365	2.407	-8.107 to 1.409
Weight	-0.090	0.297	-0.656 to 0.528
Body length	0.177	0.281	-0.369 to 0.736
Without non-significant 2 <sup>nd</sup> order effe	ects		
Intercept <sup>++</sup>	1.752	0.287	1.240 to 2.342
Relative leukocyte telomere length	0.404	0.168	0.090 to 0.735
Number of cubs in natal group	-0.036	0.164	-0.362 to 0.291
Mean temperature	6.125	1.354	3.492 to 8.980
Mean temperature <sup>2</sup>	-6.106	1.355	-8.978 to -3.470
Daily temperature variability	-3.877	2.126	-8.253 to 0.211
Daily temperature variability <sup>2</sup>	4.388	2.223	0.162 to 8.991
Mean daily rainfall	0.307	0.166	-0.003 to 0.647
Daily rainfall variability	-1.048	0.207	-1.495 to -0.664
Weight	-0.046	0.293	-0.614 to 0.512
Body length	0.142	0.276	-0.384 to 0.712

Random effect estimates (variance):

<sup>†</sup>qPCR plate (<1.000\*10<sup>-12</sup>), Row on qPCR plate (8.172\*10<sup>-2</sup>), Social group (3.824\*10<sup>-1</sup>), Cohort (2.789\*10<sup>-1</sup>) <sup>††</sup>qPCR plate (<1.000\*10<sup>-12</sup>), Row on qPCR plate (1.012\*10<sup>-1</sup>), Social group (3.952\*10<sup>-1</sup>), Cohort (6.453\*10<sup>-1</sup>)

**Table S13:** Comparison of models describing the relationship between early-life relative leukocyte telomere length (RLTL) and cub survival probability with qPCR plate, cohort, social group, and row on qPCR plate as random effects. Models were ordered and numbered by AICc, and the difference from the top model (lowest AICc) is stated in the column termed ' $\Delta$ AICc', with only models within  $\Delta$ AICc <7 (omitting 1196 models). CV<sub>R</sub> = Daily rainfall variability,  $\mu_R$  = mean daily rainfall,  $\mu_T$  = Mean temperature,  $\sigma_T$  = Daily temperature variability.

Model	Intercept	μт	(μ <sub>τ</sub> ) <sup>2</sup>	στ	(σ <sub>τ</sub> ) <sup>2</sup>	μ <sub>R</sub>	(μ <sub>R</sub> ) <sup>2</sup>	CV <sub>R</sub>	(CV <sub>R</sub> ) <sup>2</sup>	N cubs in	Body	Body	RLTL	Degrees of	AICc	ΔAICc
	-	<b>P</b> -1	16-17	-1	(-1)	Part	(F-it)			natal group	weight	length		freedom		
1	1.653	5.771	-5.749	-3.485	3.972	0.279		-0.996					0.368	12	400.6	0.000
2	1.644	6.285	-6.155	-2.894	3.358	0.604		0.872	-1.762				0.381	13	401.8	1.193
3	1.661	5.565	-5.562	-3.894	4.406	-0.514	0.802	-1.031					0.373	13	401.9	1.293
4	1.660	6.296	-6.118	-3.306	3.809	-0.613	1.464	2.079	-2.974				0.398	14	401.9	1.333
5	1.662	5.812	-5.792	-3.503	3.978	0.285		-0.982			0.093		0.374	13	402.2	1.565
6	1.639	5.200	-5.250	-3.797	4.306			-2.101	1.046				0.368	12	402.2	1.588
7	1.590	4.918	-4.915	-3.377	3.889			-0.984					0.390	11	402.3	1.650
8	1.657	5.797	-5.780	-3.472	3.951	0.282		-0.986				0.067	0.371	13	402.4	1.803
9	1.606	6.468	-6.342	0.342		0.803		2.152	-2.869				0.336	12	402.5	1.861
10	1.648	5.795	-5.779	-3.435	3.918	0.283		-0.996		-0.036			0.373	13	402.6	1.940
11	1.602	5.541	-5.617	0.337		0.272		-0.850					0.296	11	402.9	2.285
12	1.654	6.361	-6.225	-2.887	3.338	0.630		1.006	-1.875		0.104		0.388	14	403.3	2.660
13	1.613	6.448	-6.293	0.360		-0.109	1.099	3.090	-3.798				0.346	13	403.3	2.732
14	1.669	6.361	-6.179	-3.296	3.788	-0.560	1.427	2.168	-3.042		0.094		0.404	15	403.5	2.903
15	1.669	5.612	-5.610	-3.886	4.387	-0.464	0.756	-1.016			0.084		0.377	14	403.6	2.942
16	1.649	6.332	-6.202	-2.868	3.324	0.617		0.942	-1.818			0.074	0.384	14	403.6	2.950
17	1.640	6.312	-6.187	-2.839	3.301	0.608		0.877	-1.767	-0.036			0.385	14	403.7	3.133
18	1.664	5.595	-5.595	-3.862	4.367	-0.472	0.761	-1.021				0.052	0.375	14	403.8	3.177
19	1.656	5.590	-5.593	-3.841	4.349	-0.522	0.814	-1.031		-0.042			0.378	14	403.8	3.212
20	1.663	6.330	-6.153	-3.277	3.773	-0.571	1.425	2.100	-2.983			0.053	0.400	15	403.8	3.215
21	1.654	6.333	-6.160	-3.235	3.734	-0.621	1.483	2.105	-2.999	-0.047			0.404	15	403.8	3.233
22	1.647	5.223	-5.276	-3.810	4.310			-2.100	1.057		0.082		0.373	13	403.9	3.246
23	1.616	6.531	-6.401	0.330		0.827		2.269	-2.966		0.104		0.344	13	403.9	3.314
24	1.596	4.934	-4.933	-3.390	3.894			-0.972			0.077		0.394	12	403.9	3.339
25	1.643	5.217	-5.272	-3.783	4.285			-2.096	1.050			0.059	0.370	13	404.0	3.432
26	1.593	4.934	-4.936	-3.365	3.870			-0.975				0.056	0.392	12	404.1	3.502
27	1.658	5.836	-5.822	-3.451	3.924	0.288		-0.982		-0.036	0.093		0.378	14	404.1	3.505
28	1.663	5.811	-5.790	-3.515	3.992	0.285		-0.984			0.127	-0.044	0.374	14	404.1	3.537
29	1.636	5.215	-5.270	-3.763	4.270			-2.116	1.061	-0.029			0.372	13	404.2	3.550
30	1.612	6.509	-6.385	0.334		0.814		2.207	-2.911			0.080	0.340	13	404.2	3.567
31	1.568	4.844	-4.952	0.368				-0.846					0.338	10	404.2	3.569
32	1.590	4.920	-4.918	-3.369	3.881			-0.984		-0.006			0.391	12	404.3	3.649
33	1.600	6.516	-6.398	0.339		0.809		2.141	-2.861	-0.063			0.345	13	404.3	3.673
34	1.653	5.818	-5.806	-3.426	3.903	0.285		-0.986		-0.033		0.065	0.375	14	404.4	3.755
35	1.611	5.571	-5.650	0.328		0.278		-0.837			0.089		0.301	12	404.5	3.877
36	1.607	5.563	-5.644	0.331		0.275		-0.839				0.073	0.299	12	404.7	4.044
37	1.596	5.594	-5.677	0.335		0.280		-0.853		-0.068			0.306	12	404.7	4.066

38	1.602	5.463	-5.548	0.340		0.040	0.233	-0.855					0.296	12	404.8	4.220
39	1.621	6.503	-6.345	0.350		-0.054	1.056	3.161	-3.850		0.096		0.352	14	404.9	4.274
40	1.604	5.670	-5.670	-2.426	2.846	0.308		-0.932						11	404.9	4.280
41	1.607	6.509	-6.361	0.357		-0.138	1.144	3.121	-3.831	-0.076			0.357	14	405.1	4.460
42	1.579	4.920	-5.062	0.351				-1.667	0.784				0.302	11	405.1	4.535
43	1.617	6.483	-6.330	0.353		-0.062	1.053	3.099	-3.796			0.065	0.348	14	405.2	4.544
44	1.650	6.387	-6.256	-2.830	3.279	0.635		1.011	-1.879	-0.036	0.104		0.393	15	405.2	4,601
45	1.655	6.360	-6.222	-2.898	3.351	0.631		1.009	-1.879		0.142	-0.049	0.389	15	405.2	4.626
46	1.575	5.517	-5.588	0.318		0.302		-0.834						10	405.3	4,690
47	1 669	6 354	-6 168	-3 328	3 824	-0 588	1 460	2 193	-3 071		0 159	-0.085	0 405	16	405.4	4 800
48	1 663	6 395	-6 219	-3 225	3 713	-0 567	1 443	2 190	-3.063	-0.045	0.093	0.000	0.409	16	405.4	4 810
40	1.664	5 636	-5 640	-3 832	4 330	-0.471	0.767	-1 017	5.005	-0.041	0.055		0.382	15	405.4	4.865
50	1.004	5 604	-5 600	-3 917	4.330	-0.485	0.707	-1 020		0.041	0.004	-0.064	0.378	15	405.5	4.005
51	1.6/5	6 353	-6 229	-2.820	3 273	0.400	0.770	0.944	-1 820	-0.032	0.155	0.004	0.388	15	405.5	4.005
52	1.650	5 616	-5 622	-2 912	1 215	-0.481	0 774	-1 022	1.020	-0.032		0.072	0.300	15	405.5	5 100
52	1.033	5.010	-5.022	-3.813	4.515	-0.481	0.774	-1.022	2 057	-0.039	0.105	0.049	0.375	14	405.7	5.105 E 101
55	1.010	6.379	-0.430	0.520	2 706	0.652	1 442	2.237	-2.957	-0.004	0.105	0.040	0.334	14	405.7	5.121
54	1.038	0.300 E 227	-0.190	-3.213	3.700	-0.561	1.445	2.120	-3.004	-0.045	0.092	0.049	0.403	10	405.7	5.150
55	1.045	5.257	-5.295	-3.773	4.275			-2.115	1.071	-0.029	0.082	0.026	0.370	14	405.8	5.209
50	1.047	5.221	-5.272	-3.822	4.323	0.605		-2.104	1.059		0.110	-0.036	0.373	14	405.8	5.220
57	1.575	6.201	-6.123	0.320		0.695		1.381	-2.115		0.070		0.240	11	405.9	5.284
58	1.572	4.844	-4.952	0.361		0.020		-0.834	2 072		0.072	0.022	0.340	11	405.9	5.298
59	1.616	6.531	-6.399	0.331	2 0 0 2	0.828		2.275	-2.972		0.130	-0.033	0.345	14	405.9	5.298
60	1.596	4.931	-4.928	-3.399	3.903			-0.973			0.099	-0.029	0.394	13	405.9	5.326
61	1.595	4.936	-4.936	-3.382	3.885			-0.972		-0.006	0.077		0.395	13	405.9	5.337
62	1.570	4.850	-4.961	0.363				-0.837				0.060	0.339	11	406.0	5.396
63	1.640	5.230	-5.289	-3.752	4.253			-2.110	1.063	-0.026		0.057	0.373	14	406.0	5.402
64	1.606	6.551	-6.434	0.332		0.819		2.191	-2.899	-0.059		0.077	0.349	14	406.0	5.404
65	1.658	5.837	-5.820	-3.461	3.935	0.289		-0.985		-0.039	0.133	-0.051	0.379	15	406.1	5.468
66	1.593	4.935	-4.937	-3.361	3.866			-0.975		-0.003		0.056	0.393	13	406.1	5.502
67	1.563	4.858	-4.969	0.367				-0.847		-0.035			0.343	11	406.1	5.510
68	1.604	5.625	-5.711	0.325		0.285		-0.840		-0.069	0.089		0.312	13	406.3	5.651
69	1.613	5.482	-5.497	-2.810	3.255	-0.409	0.725	-0.964						12	406.3	5.699
70	1.597	6.049	-5.970	-1.978	2.383	0.546		0.431	-1.285					12	406.4	5.827
71	1.610	5.507	-5.593	0.330		0.092	0.187	-0.841			0.086		0.301	13	406.4	5.836
72	1.601	5.613	-5.700	0.328		0.282		-0.843		-0.065		0.069	0.309	13	406.5	5.844
73	1.611	5.570	-5.649	0.328		0.278		-0.838			0.102	-0.017	0.301	13	406.5	5.873
74	1.593	5.060	-5.142	-2.815	3.260			-2.239	1.238					11	406.5	5.874
75	1.595	5.507	-5.601	0.338		0.013	0.268	-0.859		-0.071			0.306	13	406.6	5.980
76	1.606	5.500	-5.588	0.333		0.090	0.186	-0.844				0.069	0.299	13	406.6	6.003
77	1.612	5.701	-5.703	-2.429	2.842	0.312		-0.921			0.074			12	406.6	6.003
78	1.615	6.564	-6.413	0.347		-0.081	1.098	3.187	-3.879	-0.075	0.095		0.363	15	406.6	6.008
79	1.597	6.636	-6.500			0.838		2.230	-2.830				0.309	11	406.7	6.101
80	1.609	5.695	-5.698	-2.407	2.821	0.310		-0.924				0.054		12	406.8	6.152
81	1.619	5.734	-5.823	-		0.295		-0.732					0.299	10	406.8	6.163
82	1.584	4.926	-5.070	0.343				-1.670	0.799		0.077		0.304	12	406.8	6.221
83	1.622	6.501	-6.340	0.351		-0.068	1.075	3.184	-3.874		0.140	-0.056	0.353	15	406.8	6.229
84	1.606	5.664	-5,662	-2.444	2,866	0.307		-0.932		0.011				12	406.9	6 275
85	1 610	6 5 3 7	-6 390	0 351	2.000	-0.093	1 098	3 123	-3 873	-0.072		0.059	0 359	15	406.9	6 304
	1.010	0.557	0.555	0.331		0.000	1.050	5.125	5.625	0.072		0.055	0.555	13	400.5	0.004

86	1.581	4.929	-5.075	0.345				-1.666	0.794			0.065	0.303	12	406.9	6.333
87	1.572	4.944	-5.092	0.348				-1.710	0.823	-0.056			0.307	12	407.0	6.384
88	1.583	6.189	-6.083	0.338		-0.104	0.964	2.215	-2.940					12	407.0	6.410
89	1.583	5.545	-5.617	0.311		0.306		-0.823			0.074			11	407.0	6.411
90	1.613	6.049	-5.930	-2.330	2.770	-0.490	1.239	1.433	-2.291					13	407.1	6.474
91	1.580	5.543	-5.617	0.313		0.305		-0.825				0.062		11	407.1	6.519
92	1.650	6.389	-6.256	-2.838	3.287	0.637		1.016	-1.886	-0.039	0.148	-0.056	0.394	16	407.2	6.556
93	1.576	5.431	-5.511	0.322		0.039	0.264	-0.840						11	407.2	6.606
94	1.600	5.109	-5.247					-0.719					0.348	9	407.2	6.621
95	1.573	5.532	-5.604	0.317		0.305		-0.835		-0.021				11	407.3	6.669
96	1.663	6.393	-6.213	-3.252	3.743	-0.598	1.483	2.224	-3.101	-0.052	0.166	-0.095	0.411	17	407.3	6.681
97	1.664	5.630	-5.632	-3.861	4.361	-0.495	0.793	-1.021		-0.045	0.138	-0.072	0.383	16	407.4	6.791
98	1.532	4.734	-4.755	-2.297	2.747			-0.916						10	407.5	6.845
99	1.583	6.244	-6.164	0.313		0.707		1.442	-2.162		0.081			12	407.6	6.950
100	1.551	4.846	-4.997	0.330				-1.898	1.022					10	407.6	6.970

**Table S14:** Parameter estimates and 95% confidence intervals of fixed effects from a model averaging approach (natural average) between relative leukocyte telomere length, social and weather variables and survival to adulthood (>1 year old) in European badgers.  $\beta$  = parameter estimate, S.E. = standard error, 95% CI = 95% confidence interval. Significant parameter estimates (95% CI does not overlap zero) are in bold.

Parameter (reference level)	β	S.E.	95% CI
Intercept	1.631	0.252	1.134 to 2.124
Relative leukocyte telomere length	0.366	0.153	0.064 to 0.666
Number of cubs in natal group	-0.041	0.147	-0.329 to 0.249
Mean temperature	5.791	1.401	3.024 to 8.540
Mean temperature <sup>2</sup>	-5.754	1.361	-8.428 to -3.069
Daily temperature variability	-2.325	2.375	-6.913 to 2.409
Daily temperature variability <sup>2</sup>	3.874	2.061	-0.208 to 7.906
Mean daily rainfall	0.140	0.772	-1.369 to 1.663
Mean daily rainfall <sup>2</sup>	1.039	1.080	-1.115 to 3.138
Daily rainfall variability	-0.052	2.009	-3.988 to 3.874
Daily rainfall variability <sup>2</sup>	-1.840	2.436	-6.590 to 2.972

**Table S15:** Parameter estimates and 95% confidence intervals of fixed effects from a mixed model and parametric bootstrap tests of same-sex group size effects on early-adulthood relative leukocyte telomere length (Z-score) in European badgers (full model and with  $p > 0.10 2^{nd}$  order effects removed).  $\beta$  = parameter estimate, S.E. = standard error, 95% CI = 95% confidence interval; reference terms in brackets = reference level for factors; \* = interaction. Significant parameter estimates (95% CI does not overlap zero) are in bold.

Parameter (reference level)	β	S.E.	95% CI			
Intercept <sup>+</sup>	-0.127	0.283	-0.670 to 0.436			
Number of same-sex adults in group	-0.054	0.302	-0.632 to 0.528			
Sex (Female)	0.109	0.292	-0.465 to 0.676			
Season (Spring)						
Summer	-0.128	0.113	-0.353 to 0.096			
Autumn	0.150	0.205	-0.261 to 0.563			
Winter	-0.164	0.335	-0.821 to 0.508			
Age at last capture	0.003	0.053	-0.105 to 0.109			
Weight	0.012	0.080	-0.144 to 0.164			
Body length	-0.031	-0.031 0.083 -0.195 t				
Number of same-sex adults in group	0.072	0.072 0.311 -0.540 to 0.67				
* Sex (Female)						
Without non-significant 2 <sup>nd</sup> order effe	ects					
Intercept <sup>++</sup>	-0.072	0.152	-0.362 to 0.231			
Number of same-sex adults in group	0.013	0.093	-0.169 to 0.201			
Sex (Female)	0.057	0.189	-0.328 to 0.430			
Season (Spring)						
Summer	-0.128	0.113	-0.355 to 0.090			
Autumn	0.149	0.205	-0.260 to 0.549			
Winter	-0.164	0.334	-0.814 to 0.458			
Age at last capture	0.002	0.053	-0.102 to 0.106			
Weight	0.014	0.079	-0.139 to 0.169			
Body length	-0.029	0.083	-0.189 to 0.133			

Random effect estimates (variance):

<sup>+</sup>qPCR plate (3.839\*10<sup>-2</sup>), Row on qPCR plate (<1.000\*10<sup>-12</sup>), Social group (<1.000\*10<sup>-12</sup>), Cohort (5.750\*10<sup>-2</sup>), Year (4.093\*10<sup>-2</sup>), Individual ID (7.779\*10<sup>-8</sup>), Residual (8.358\*10<sup>-1</sup>)

<sup>++</sup>qPCR plate (3.870\*10<sup>-2</sup>), Row on qPCR plate (<1.000\*10<sup>-12</sup>), Social group (<1.000\*10<sup>-12</sup>), Cohort (5.923\*10<sup>-2</sup>), Year (4.064\*10<sup>-2</sup>), Individual ID (1.280\*10<sup>-7</sup>), Residual (8.329\*10<sup>-1</sup>)



**Figure S1:** The effect of longitudinal measures from spring to winter (a) and winter to spring (b) on relative leukocyte telomere length (RLTL). Raw data for spring to winter (a; n = 214 samples; 203 badgers) and winter to spring (b; n = 84 samples; 78 badgers) are shown.



**Figure S2:** The association between mean winter temperature and survival to adulthood. Raw data points (n = 406 samples; 406 badgers) are shown, and jittered for clarity on the amount of data. The fitted line represents the quadratic regression from the mixed model, and the 95% confidence intervals as shaded areas.



**Figure S3:** The association between the temperature variability in winter and survival to adulthood. Raw data points (n = 406 samples; 406 badgers) are shown, and jittered for clarity on the amount of data. The fitted line represents the regression from the mixed model, and the 95% confidence intervals as shaded areas.



**Figure S4:** The association between the rainfall variability in winter and survival to adulthood. Raw data points (*n* = 406 samples; 406 badgers) are shown, and jittered for clarity on the amount of data. The fitted line represents the regression from the mixed model, and the 95% confidence intervals as shaded areas.