

Common-variant associations with fragile X syndrome

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Abstract

Fragile X syndrome is rare but a prominent cause of intellectual disability. It is usually caused by a de novo mutation that occurs on multiple haplotypes and thus would not be expected to be detectible using genome-wide association (GWA). We conducted GWA in 89 male FXS cases and 266 male controls, and detected multiple genome-wide significant signals near *FMR1* (odds ratio = 8.10, $P = 2.5 \times 10^{-10}$). These findings withstood robust attempts at falsification. Fine-mapping yielded a minimum $P = 1.13 \times 10^{-14}$, but did not narrow the interval. Comprehensive functional genomic integration did not provide a mechanistic hypothesis. Controls carrying a risk haplotype had significantly longer *FMR1* CGG repeats than controls with the protective haplotype ($P = 4.75 \times 10^{-5}$), which may predispose toward increases in CGG number to the premutation range over many generations. This is a salutary reminder of the complexity of even “simple” monogenetic disorders.

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Introduction

Fragile X syndrome (FXS) [1] is rare (0.25–1/1000 male births) but a prominent cause of intellectual disability [2, 3]. It is characterized by intellectual disability, autistic behavior, hyperactivity, anxiety, and a range of physical abnormalities (e.g., tall stature and macroorchidism) [4]. FXS is caused by CGG expansion in the 5′ UTR of the chromosome X gene *FMR1* in most cases [5–7]. Full FXS mutations are characterized by expansion of the *FMR1* 5′ UTR CGG repeat to ≥ 200 copies with premutations in the 55–200 copy range [8].

FMR1 5′ UTR CGG expansions generally arise de novo when mutable premutations expand to full mutations during oogenesis. Although the probability of de novo mutations can be influenced by local DNA features, detection of de novo events using linkage disequilibrium would be unexpected for high-penetrance single-gene disorders [9, 10]. This implies that genome-wide association (GWA) of FXS cases versus controls should not detect the *FMR1* region as a susceptibility locus for FXS. As part of a study of FXS and autism, we conducted a case–control GWAS for FXS and found a strong common-variant association signal near *FMR1* that withstood efforts at falsification. We determined that this association was consistent with case–control association studies using small numbers of microsatellite markers from the 1990s.

Table 1 Summary of samples and genotyping

Purpose	Genotyping	FXS cases	Controls	Control source
Establish FXS status	<i>FMRI</i> 5'UTR CGG repeat (CLIA assay)	89	0	N/A
	<i>FMRI</i> CGG repeat analysis	82 ^a	109	Sweden
GWAS	Illumina HumanOmni1-Quad	89	266	GABC
Allele frequency comparisons	Illumina Human1M & Affymetrix 6.0	–	101	HapMap3
	Illumina OmniExpress	–	3525	Sweden
Verify key genotypes and allele assignments	TaqMan rs2197706	82 ^a	0 ^b	N/A
	TaqMan rs5905149	82 ^a	94 ^b	Sweden
Common-variant fine-mapping	Sequenom (31 SNPs)	89	467	Sweden
	TaqMan rs45631657	103	467	Sweden

^aInsufficient DNA for seven FXS cases

^bThe purpose of genotyping was to verify allele assignments; as rs5905149 verified in cases and controls and as rs2197706 verified in cases, we were confident that allele misidentification did not explain case–control differences

Materials and methods

Subjects

Males with a genetically confirmed diagnosis of FXS were recruited from volunteer registries (URLs). All available medical records were reviewed, and any features suggestive of a complex or atypical presentation led to exclusion. Controls were from the Genes and Blood Clotting Study (GABC) in dbGaP (URLs, accession phs000304.v1.p1). GABC participants were male university students who volunteered for a study of the genetics of hemostasis and who had no acute or chronic illnesses. Additional male comparison subjects were from the Swedish Schizophrenia Study ($N = 3525$, 46.4% cases); briefly, these subjects were sampled from population registers and non-European ancestry outliers were removed, and we selected males genotyped with OmniExpress arrays [11]. As we found no evidence for association with schizophrenia in the *FMRI* region in this sample or in the literature [11, 12], cases and controls were combined. We also used male HapMap3 founders from northwestern Europe and Tuscany (CEU and TSI, $N = 101$) [13]. All procedures were approved by Institutional Review Boards, and written informed consent was obtained from the parents/legal guardians of FXS cases and from control subjects.

Genetic assays

Table 1 summarizes the samples and assays used in this study. Cases and controls were randomized where possible and analyses were performed blinded to case/control status where possible. *FMRI*: the number of CGG repeats in the 5' UTR of *FMRI* was determined in 89 FXS cases with a validated diagnostic assay (Kimball Genetics, Denver, CO

[14–16]. To understand the internal structure of *FMRI* CGG repeats and to place these on common haplotypes, we used AmpliX *FMRI* PCR kits (Asuragen, Inc.; Austin, TX; catalog #49402) to quantify *FMRI* 5'UTR repeat sizes and to count AGG interruptions. GWAS: FXS cases and GABC controls were genotyped with Illumina HumanOmni1-Quad arrays, and genotypes were called using predefined clusters using GenomeStudio. Quality control was performed using PLINK [17]. SNPs were excluded for missingness >0.03 , minor allele frequency <0.01 , deviation from Hardy–Weinberg expectations in controls ($P < 1 \times 10^{-6}$), SNP missingness differences between cases and controls ($P < 0.05$), or if a SNP probe did not map uniquely to the human genome. Subjects were excluded for missingness >0.05 , excessive autosomal homozygosity or heterozygosity, or relatedness ($\hat{\pi} > 0.2$ based on LD-pruned autosomal SNPs). One FXS case was genotyped in duplicate with 0.99998 concordance, and a CEPH sample previously assayed with the same array had 0.99981 concordance. *TaqMan*: rs2197706 and rs5905149 genotypes were verified with TaqMan Assays (catalog #4351379 and #4351379, Applied Biosystems, Carlsbad, CA). A SNP from Gerhardt et al. [18]. (rs45631657) was genotyped with a custom TaqMan assay. *Sequenom*: we designed two massARRAY iPLEX (San Diego, CA) genotyping panels for common-variant fine-mapping. SNPs were selected from GWAS results, haplotype analyses, and common variation databases, and then pruned using TAGGER [19]. All assays used genomic DNA isolated from peripheral blood. The genome reference was GRCh37/UCSC hg19.

Statistical analysis

Case–control comparisons were performed using PLINK [17] using logistic regression under an additive model with

Table 2 Genome-wide significant results of GWAS of male FXS cases and controls

SNP	chrX (hg19)	Alleles	OR (95% CI)	P	F_{case}	F_{control}	F_{Sweden}	F_{CEU}	F_{TSl}
rs5952060	146852679	C/T	5.33 (2.93–9.71)	4.57×10^{-8}	0.764	0.365	0.357	0.386	0.409
rs2197706	146895120	A/C	8.10 (4.24–15.49)	2.53×10^{-10}	0.807	0.351	–	0.357	0.386
rs5905149	146908213	A/C	5.99 (3.40–10.56)	5.76×10^{-10}	0.584	0.184	0.183	–	–
rs7876251	146913828	G/A	5.64 (3.17–10.01)	3.68×10^{-9}	0.693	0.286	0.279	0.263	0.279
rs4824253	146918268	G/A	5.35 (3.04–9.43)	6.41×10^{-9}	0.685	0.286	0.278	0.263	0.296

The first allele given is the least common in this sample and is the reference for the odds ratio (OR) and frequencies. CI is confidence interval. P is from the logistic regression including ancestry covariates. Logistic regression P values after removing nine cases with divergent ancestry were 5.86×10^{-8} , 3.36×10^{-10} , 8.52×10^{-10} , 4.79×10^{-9} , and 8.38×10^{-9} . Shown are allele frequencies for male FXS cases (F_{case}), GABC controls (F_{control}), subjects from Sweden (F_{Sweden}), and HapMap3 northwestern European (F_{CEU}), and Tuscan control samples (F_{TSl})

three ancestry principal components as covariates. Bioinformatic integration included a range of functional genomic results from fetal brain (see Supplemental methods).

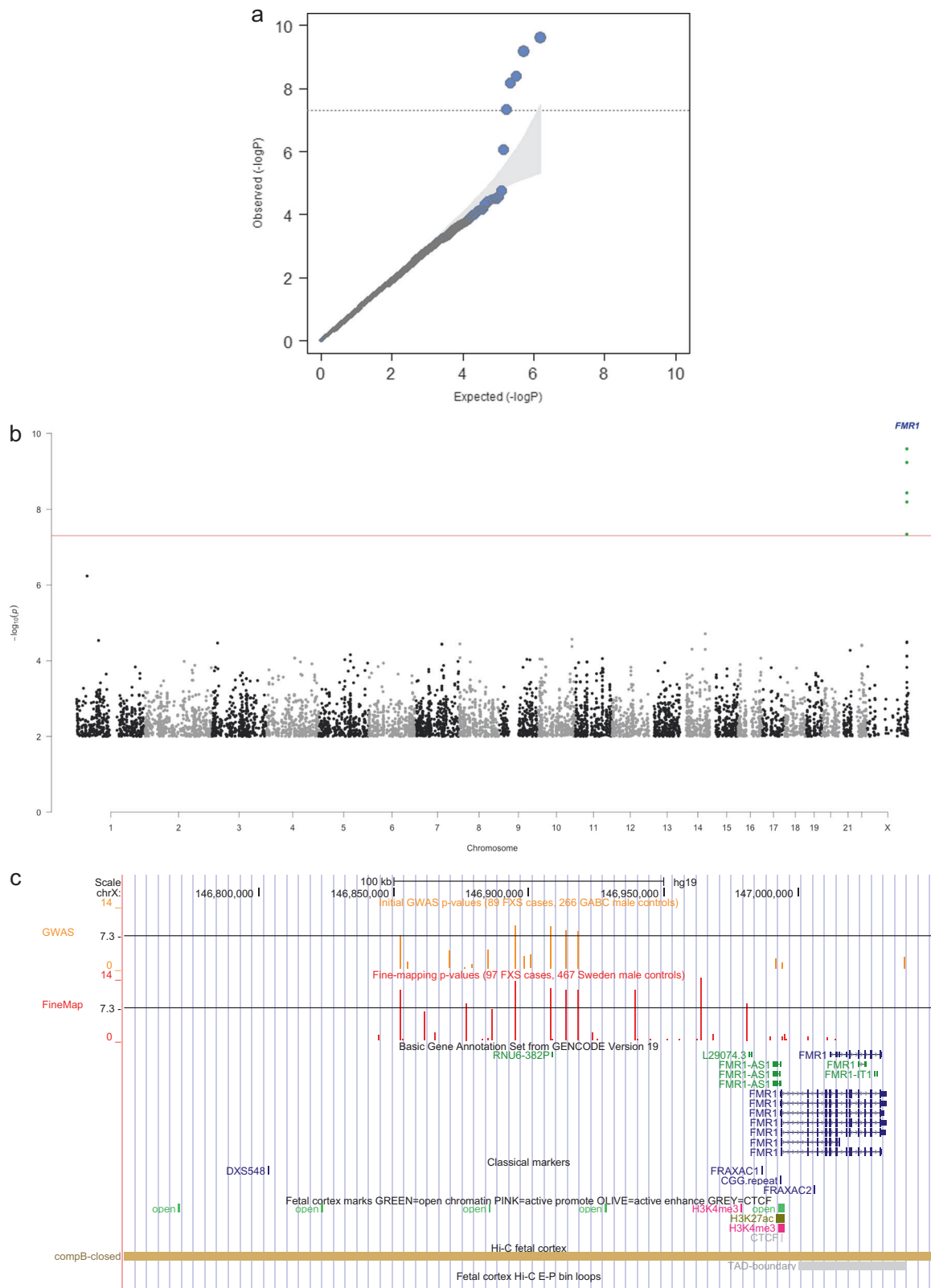
Results

We conducted GWA analyses for 750K SNPs in 89 male FXS cases and 266 male controls (Table S1). All FXS cases had full *FMRI* mutations (>200 CGG repeats) as determined using a validated diagnostic assay meeting US governmental standards for accuracy and reliability (we verified these findings using a second assay in 82 FXS cases with sufficient DNA). We assessed ancestry using principal components analysis [20] on LD-pruned autosomal SNPs (Figure S1). All controls and 90% of cases were of predominant European ancestry (we retained nine cases of mixed ancestry given the small number of cases). Logistic regression analyses identified five SNPs that met genome-wide significance with odds ratios >5 (Table 2, Fig. 1a, b). These SNPs were in a 66 kb interval from chrX:146.85–146.92 Mb located 75 kb 5' of the nearest gene, *FMRI*. Repeating the logistic regression conditioning on the most strongly associated SNP (rs2197706) markedly attenuated significance in the *FMRI* region suggesting the presence of a single association signal.

Given that FXS usually results from de novo mutations, strong associations with common SNPs could be considered unexpected. Indeed, the strongest association (rs2197706, odds ratio = 8.10, $P = 2.5 \times 10^{-10}$) is among the top dichotomous trait associations in the GWAS catalog [21] (URLs). We, therefore, evaluated alternative explanations for these findings. First, given the marked allele frequency differences in cases and controls, resequencing rs5905149 and rs2197706 with TaqMan assays showed perfect agreement with Illumina array genotypes, and served to exclude allele assignment errors. Second, the allele frequencies in cases and controls were similar genome-wide except for SNPs 5' of *FMRI* (Figure S2). Given our use of controls genotyped

independently from cases, it is important to note that the control allele frequencies for the significant SNPs in the *FMRI* region were similar to those from two external samples (Table 2). Third, genome-wide P values conformed closely to the null expectation (mean P value = 0.500 over 750K SNPs, Fig. 1a), findings inconsistent with uncontrolled bias. Fourth, exclusion of nine cases with mixed ancestry had little impact on the results (Table 2). Fifth, a trivial explanation for these findings is if cases were cryptically related via a recent shared ancestor; however, case–case pairs were slightly *less* related on average than control–control pairs (Figure S3). Cases and controls had similar proportions of autosomal homozygous SNPs as well as the number and size of autosomal runs of homozygosity (no comparisons were significantly different and cases had lower means in each instance). Sixth, asymptotic P values can be inaccurate in small samples, but Fisher's exact test and permutation procedures yielded similar significance levels. Thus, we could identify no plausible alternative explanation for our findings.

In a fine-mapping experiment, we genotyped 32 SNPs (chrX:146844358–147013704, the association region extending into *FMRI*) in an expanded but overlapping set of 89 FXS cases (27 new male FXS male cases plus 62 of the 89 original FXS cases) along with 467 male controls from a different study than for the initial GWAS (see Table 2). We included rs45631657 which was reported to inactivate an important replication origin [18]. Variable numbers of SNPs were genotyped twice on the same subjects, and we observed 100% agreement (data not shown). Nine SNPs exceeded genome-wide significance (Table S2). Analyses of all five SNPs in Table 2 yielded consistent odd ratios and greater significance (P values ranging from 4×10^{-12} to 7×10^{-14}), and four other SNPs reached genome-wide significance (rs4824231, $P = 1.13 \times 10^{-14}$; rs25705, $P = 5.74 \times 10^{-9}$; rs45631657, $P = 5.20 \times 10^{-12}$; and rs112146098, $P = 6.60 \times 10^{-9}$). Repeating the logistic regression conditioning on rs2197706 or rs4824231 markedly attenuated significance in the *FMRI*



region again suggesting the presence of a single association signal. Thus, we continued to observe a broad region of significance.

Figure 1c depicts the association region. The association region includes the *FMR1* promoter CGG repeat. Table 3 shows haplotypes from the genome-wide significant SNPs,

◀ **Fig. 1** FXS case–control GWAS. **a** Quantile–quantile plot for logistic regression of male FXS cases and GABC controls (including ancestry principle components). The observed P values conform closely to the null except for five SNPs in the *FMRI* region. The shaded region indicates the expected 95% probability interval for ordered P values. **b** Manhattan plot for the GWAS of male FXS cases and GABC controls (logistic regression including ancestry principle components). The X -axis is chromosomal position from 1ptel to Xqtel. The Y -axis is $-\log_{10}(P)$. Genome-wide significant SNPs near *FMRI* are indicated. **c** Detail of *FMRI* region (hg19, chrX:146850000-147040000). Tracks are: GENCODE gene annotations; positions of FRAXAC1, FRAXAC2, and promoter CGG repeat; selected ChIP-seq marks; SNP positions and $-\log_{10}(P)$ for SNPs in the fine-mapping study and in the GWAS; DNA–DNA chromosomal looping from 5C based on the *FMRI* promoter; and open chromatin in prefrontal cortex of nine adult schizophrenia (SCZ) cases and nine fetal samples

Table 3 Haplotype analyses of *FMRI* region

Haplotype	Subjects	Controls	FXS cases	Freq control	Freq case
TGACGGTCC	17	14	3	0.030	0.031
CGACGGTCC	20	17	3	0.036	0.031
CGAAGGTTT	21	20	1	0.043	0.010
CGACAATTC	26	17	9	0.036	0.093
CGCCAATCC	31	29	2	0.062	0.021
CGAAGGTTC	65	42	23	0.090	0.237
CCAAGGCTT	77	44	33	0.094	0.340
TGCCAATCC	261	250	11	0.535	0.113

Observed haplotypes from fine-mapping data (32 SNPs in 89 FXS cases and 467 male controls). Haplotypes were created using nine genome-wide significant SNPs (rs5952060-rs112146098-rs2197706-rs5905149-rs7876251-rs4824253-rs45631657-rs4824231-rs25705). Haplotypes counts <10 ($N=46$) were removed. Logistic regression highlighted a strongly protective haplotypes (green) and two risk haplotypes (red)

and the LD matrix of pairwise r^2 values is given in Table S3. The most common haplotype was strongly protective, and there were two risk haplotypes. *FMRI* CGG analysis was available on 109 controls (selected to represent the most common haplotypes). Controls with the risk haplotype had significantly longer CGG repeats (mean = 33.7, SD = 6.9, $N=53$) than controls with the protective haplotype (mean = 28.7, SD = 5.3, $N=56$; $F_{1,107} = 18.7$, $P = 3.52 \times 10^{-5}$). Around 40% of cases had additional phenotype measures (e.g., vineland adaptive behavior scale and the social responsiveness scale) and there were no significant differences between cases with the risk or protective haplotypes (data not shown).

We next evaluated possible functions of the association region using functional genomic data. Using RNA-seq data from human dorsolateral prefrontal cortex (DLPFC) in nine schizophrenia cases and nine controls along with PFC from nine fetuses and three neural progenitor cell lines, we saw that *FMRI* (but not the antisense transcript, *FMRI-ASI*) was robustly expressed (Figure S4). There was no evidence

of substantial gene expression or an unannotated feature in the association region 5' to *FMRI*. The expression of *FMRI* in DLPFC is associated with a common genetic variant but the associated SNP is far outside the region [22]. We evaluated functional genomic data from fetal brain [23]. As expected, the *FMRI* transcriptional start site had a pattern of open chromatin and H3K27ac and H3K4me3 histone marks consistent with an actively transcribed gene. The 5' association region had several additional open chromatin regions, but these are common in the genome and there were no suggestions of a distal enhancer-promoter regulatory mechanism (i.e., the open chromatin regions did not have H3K4me marks or evidence of distal chromatin loop interactions using Hi-C; Fig. 1c). The SNPs associated with FXS were not notable for open chromatin or key ChIP-seq marks.

Discussion

GWA of FXS cases versus controls identified an unusually strong association with the *FMRI* region. The largest association (odds ratio = 8.10, $P = 2.5 \times 10^{-10}$) is among the top dichotomous trait associations in the EBI GWAS catalog [21] (*URLs*), and generally exceeded only by rare adverse drug reactions. Given the small sample size (89 FXS cases and 266 male controls), it is notable that the association survived robust attempts at falsification, and became more significant in a fine-mapping experiment with some new cases and independent controls.

In some respects, it was unexpected to identify the causal locus for FXS (a rare, single-gene disorder) in an outbred population using linkage disequilibrium-based GWA. GWA in case–control samples can detect rare causal genes in special circumstances that do not apply here (e.g., if sample size is very large, if cases inherit a causal mutation from a relatively recent common ancestor [24, 25], or if multiple rare mutations yield an aggregate signal detectable by GWA [26]). De novo mutations in particular may be invisible to GWA: although de novo mutational processes can be influenced by local genomic context, replication timing, and genotypes at other loci [9, 10, 27–29], these effects are generally not deterministic, and most de novo mutations occur on different haplotypes.

With the exception of unusual exonic mutations, FXS is caused by a de novo mutational event in the expansion of a premutation to a full mutation during oogenesis [5–7]. However, the local genomic context of *FMRI* de novo promoter mutations is influential [30–32]. In fact, our finding could be anticipated given the literature from the 1990s: there is substantial evidence that this region is detectable via linkage disequilibrium in case–control studies using a few microsatellite markers [32–37]. Indeed, a 1992

paper [34] reported a FXS case versus control haplotype difference as “ $P < 0.001$ ” but the P value actually reached genome-wide significance ($P \sim 9 \times 10^{-9}$). The association of common variation upstream of *FMRI* with FXS has strong replication evidence in the literature: this is unquestionably a replicated association.

Fine-mapping of the interval and integration with a number of types of functional genomic data did not narrow the region or yield a mechanistic hypothesis. A lack of early fetal data limits this conclusion. It is possible that a population genetic mechanism is at work: the risk haplotype is present in ~18% of European-ancestry controls, and tends to carry a greater number of CGG repeats which may predispose toward increases in CGG number to the premutation range over many generations. Similar mechanisms may apply for repeat expansions in Huntington’s disease [38] and amyotrophic lateral sclerosis [39]. These findings also suggest a potential clinical use of *FMRI* region haplotype data. For example, if SNP array or whole genome sequencing data are consistent with the FXS-risk haplotype, subjects with a risk haplotype and features of FXS could reasonably be prioritized for testing for the presence of the causal *FMRI* CGG repeat (a specialized test not readily obtainable otherwise).

This is a salutary reminder of the complexity of even “simple” monogenetic disorders, and of the continuing importance of robust findings from older papers before the advent of high-throughput/high-resolution genomic methods. Publication of our finding will allow inclusion in the EBI GWAS catalog, and maximize the chances that these older but highly relevant papers remain known to the community.

URLs

dbGaP, <http://www.ncbi.nlm.nih.gov/gap>

Fragile X Research Registry, <https://www.fragilexregistry.org>
EBI/NHGRI GWAS catalog, <https://www.ebi.ac.uk/gwas>

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Compliance with ethical standards

Conflict of interest PFS is a scientific advisor for Lundbeck.

References

1. Bhakar AL, Dolen G, Bear MF. The pathophysiology of fragile X (and what it teaches us about synapses). *Annu Rev Neurosci*. 2012;35:417–43.
2. Turner G, Webb T, Wake S, Robinson H. Prevalence of fragile X syndrome. *Am J Med Genet*. 1996;64:196–7.
3. Sherman SL. Epidemiology. In: Hagerman R, Hagerman P, editors. *Fragile X Syndrome: Diagnosis, Treatment and Research*. Baltimore: The Johns Hopkins University Press; 2002. p. 136–68.
4. Terracciano A, Chiurazzi P, Neri G. Fragile X syndrome. *Am J Med Genet C Semin Med Genet*. 2005;137:32–7.
5. Oberle I, et al. Instability of a 550-base pair DNA segment and abnormal methylation in fragile X syndrome. *Science*. 1991; 252:1097–102.
6. Verkerk AJ, et al. Identification of a gene (FMR-1) containing a CGG repeat coincident with a breakpoint cluster region exhibiting length variation in fragile X syndrome. *Cell*. 1991;65:905–14.
7. Yu S, et al. Fragile X genotype characterized by an unstable region of DNA. *Science*. 1991;252:1179–81.
8. Jin P, Alisch RS, Warren ST. RNA and microRNAs in fragile X mental retardation. *Nat Cell Biol*. 2004;6:1048–53.
9. Veltman JA, Brunner HG. De novo mutations in human genetic disease. *Nat Rev Genet*. 2012;13:565–75.
10. Lek M, et al. Analysis of protein-coding genetic variation in 60,706 humans. *Nature*. 2016;536:285–91.
11. Ripke S, et al. Genome-wide association analysis identifies 13 new risk loci for schizophrenia. *Nat Genet*. 2013;45:1150–9.
12. Ripke S, et al. Biological insights from 108 schizophrenia-associated genetic loci. *Nature*. 2014;511:421–7.
13. Altshuler DM, et al. Integrating common and rare genetic variation in diverse human populations. *Nature*. 2010;467:52–8.
14. Taylor AK, et al. Molecular predictors of cognitive involvement in female carriers of fragile X syndrome. *J Am Med Assoc*. 1994; 271:507–14.
15. Merenstein SA, et al. Molecular-clinical correlations in males with an expanded FMR1 mutation. *Am J Med Genet*. 1996;64:388–94.
16. Loesch DZ, Huggins RM, Bui QM, Taylor AK, Hagerman RJ. Relationship of deficits of FMR1 gene specific protein with physical phenotype of fragile X males and females in pedigrees: a new perspective. *Am J Med Genet A*. 2003;118:127–34.
17. Chang CC, et al. Second-generation PLINK: rising to the challenge of larger and richer datasets. *Gigascience*. 2015;4:7.
18. Gerhardt J, et al. Cis-acting DNA sequence at a replication origin promotes repeat expansion to fragile X full mutation. *J Cell Biol*. 2014;206:599–607.
19. Barrett JC, Fry B, Maller J, Daly MJ. Haploview: analysis and visualization of LD and haplotype maps. *Bioinformatics*. 2005; 21:263–5.
20. Price AL, et al. Principal components analysis corrects for stratification in genome-wide association studies. *Nat Genet*. 2006; 38:904–9.
21. Hindorf LA, et al. Potential etiologic and functional implications of genome-wide association loci for human diseases and traits. *Proc Natl Acad Sci USA*. 2009;106:9362–7.
22. Fromer M, et al. Gene expression elucidates functional impact of polygenic risk for schizophrenia. *Nat Neurosci*. 2016;19:1442–53.
23. Giusti-Rodríguez, P. et al. Schizophrenia and a high-resolution map of the three-dimensional chromatin interactome of adult and fetal cortex. 2018 [submitted for publication].
24. Peltonen L, Jalanko A, Varilo T. Molecular genetics of the Finnish disease heritage. *Hum Mol Genet*. 1999;8:1913–23.
25. Struwing JP, et al. The risk of cancer associated with specific mutations of BRCA1 and BRCA2 among Ashkenazi Jews. *N Engl J Med*. 1997;336:1401–8.

26. Anderson CA, Soranzo N, Zeggini E, Barrett JC. Synthetic associations are unlikely to account for many common disease genome-wide association signals. *PLoS Biol.* 2011;9:e1000580.
27. Berg IL, et al. PRDM9 variation strongly influences recombination hot-spot activity and meiotic instability in humans. *Nat Genet.* 2010;42:859–63.
28. Hastings PJ, Lupski JR, Rosenberg SM, Ira G. Mechanisms of change in gene copy number. *Nat Rev Genet.* 2009;10:551–64.
29. Koren A, et al. Differential relationship of DNA replication timing to different forms of human mutation and variation. *Am J Hum Genet.* 2012;91:1033–40.
30. Libby RT, et al. CTCF cis-regulates trinucleotide repeat instability in an epigenetic manner: a novel basis for mutational hot spot determination. *PLoS Genet.* 2008;4:e1000257.
31. Brock GJ, Anderson NH, Monckton DG. Cis-acting modifiers of expanded CAG/CTG triplet repeat expandability: associations with flanking GC content and proximity to CpG islands. *Hum Mol Genet.* 1999;8:1061–7.
32. Eichler EE, et al. Length of uninterrupted CGG repeats determines instability in the FMR1 gene. *Nat Genet.* 1994;8:88–94.
33. Fu YH, et al. Variation of the CGG repeat at the fragile X site results in genetic instability: resolution of the Sherman paradox. *Cell.* 1991;67:1047–58.
34. Richards RI, et al. Evidence of founder chromosomes in fragile X syndrome. *Nat Genet.* 1992;1:257–60.
35. Haataja R, Vaisanen ML, Li M, Ryyanen M, Leisti J. The fragile X syndrome in Finland: demonstration of a founder effect by analysis of microsatellite haplotypes. *Hum Genet.* 1994;94:479–83.
36. Chiurazzi P, Macpherson J, Sherman S, Neri G. Significance of linkage disequilibrium between the fragile X locus and its flanking markers. *Am J Med Genet.* 1996;64:203–8.
37. Gunter C, et al. Re-examination of factors associated with expansion of CGG repeats using a single nucleotide polymorphism in FMR1. *Hum Mol Genet.* 1998;7:1935–46.
38. Warby SC, et al. CAG expansion in the Huntington disease gene is associated with a specific and targetable predisposing haplogroup. *Am J Hum Genet.* 2009;84:351–66.
39. Mok K, et al. Chromosome 9 ALS and FTD locus is probably derived from a single founder. *Neurobiol Aging.* 2012;33:209 e3–8.