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Devaluation, competitiveness and new business formation in emerging countries

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DEVALUATION, COMPETITIVENESS AND NEW BUSINESS FORMATION IN EMERGING COUNTRIES

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In September 2010, Brazil's Finance Minister, Guido Mantega, used the term "currency war" with reference to monetary policies implemented by different countries to generate an artificial devaluation of their currency and achieve a cheaper, more competitive domestic economy that may be attractive to foreign investors. Similar cases have been documented since the 1930s Great Depression, when several countries abandoned the gold standard as backing for their currencies. More recently, a large-scale asset purchase by Japan's Central Bank in 2013 was singled out as a strategy aimed at generating devaluation of the yen. This research uses statistics of new business formation density reported by Doing Business for 30 emerging countries in the period 2004-2011 to evaluate the impact of devaluation measured by the behavior of the real effective exchange rate (REER) on the rate of new business formation (NBF). It is determined how variables associated with competitiveness affect the relationship between devaluation and business formation. Results show that devaluation has a positive effect on NBF in the short term, which gets diluted in the long term. Countries with greater competitiveness have less dependence on devaluation to increase the number of businesses.

Key words: New business formation, emerging countries, devaluation, competitiveness.

Classification codes: G18, G24, M13

1. Introduction

In September 2010, Brazil's Finance Minister, Guido Mantega, used the term "currency war" with reference to monetary policies implemented mainly by China, Japan, Thailand, South Korea, Colombia and other countries to generate an artificial devaluation¹ of their currency in order to achieve a cheaper, more competitive domestic economy that may be attractive to foreign investors, as well as to be able to withstand the 2008 economic crisis (Eichgreen, 2013).

Similar cases have been documented since the Great Depression of the 1930s, when several countries abandoned the gold standard in order to depreciate their currencies so as to achieve a faster recovery (Eichengreen and Sachs, 1985; 1986), which triggered off the Tripartite Agreement of 1936 between United States, France and Great Britain, which other countries later joined in.

The variation of REER and its close relation to the economic growth experienced by emerging countries has been the subject matter of extended research in the economic literature for several years. Positive effects such as the Foreign Direct Investment (FDI) inflow incentive, lower labor costs and the improvement of investment efficiency are transmission mechanisms of the devaluation effect to the economy (Bhalla, 2012). However, the relationship between devaluation and business formation hasn't been studied so far. Academics have focused on the relationship between devaluation and FDI or FDI and business formation, without directly analyzing the relationship between devaluation and NBF.

The study of the phenomenon of entrepreneurship and the rise of new firms has captured the attention both of academics and government agencies in the last few decades. Its impact on economic growth and its close relation to factors that mark countries' competitiveness turn it into a central topic for policy definition by national and international authorities. As part of the strategies undertaken to encourage business formation, both developed and developing countries have fostered FDI on the basis of the existence of a "spillover" effect that adds more knowledge and demand for new products to local economies (Markusen and Venables, 1999; Barrios et al., 2005; Herrera-Echeverri et al., 2014). The results of Zhang (2006) and of Yao and Wei (2007) support a positive effect of FDI in the case of China, which since the

¹ We understand that devaluation (corresponds to price fixing) and depreciation (corresponds to natural variation of currency) are different, we are keen to examine whether the increase in exchange rate has implications on business formation regardless of its origins; we will use the term devaluation for its effect.

1980s and 1990s has initiated a strategy of devaluation of the yuan to boost the growth of its economy in the following years.

Other factors that come into play in business formation have been mentioned, such as fiscal policy in terms of consumption by governments, credit access, technological developments and trade openness, which the literature has related to devaluation phenomena (Edwards, 1989; Reynolds et al., 2000; Black and Strahan, 2002; Choi and Phan, 2006; Takii, 2008); it is therefore worth wondering whether upon separating such effects devaluation continues having an impact on business formation. If such effect persists, is it short-term or long-term? And the final question is whether the relationship between business formation and devaluation is affected by the increase in competitiveness of the countries.

Answering those questions so as to contribute to public policy design that may stimulate entrepreneurship in emerging economies is the aim of this research paper. The panel data technique is employed on a sample of 30 emerging countries from 2004 to 2011 to analyze the following: the effect of devaluation on NFB and its period of duration; the variables that affect both devaluation and business formation; the residual value of devaluation on business formation after removing the effect of the variables that affect devaluation and business formation simultaneously; and, finally, how competitiveness variables affect the relationship between devaluation and business formation.

Research results reveal a positive relationship between devaluation and business formation rate in emerging countries; however, this effect is only short-term, as it gets diluted for periods longer than 2 years. Devaluation does not compensate for the importance of competitiveness in promoting business formation in emerging economies. The devaluation effect loses magnitude and significance against very low competitiveness levels in terms of credit supply for the private sector and institutional quality.

These findings show the importance of maintaining a good balance between short- and long-term strategies to boost the dynamics of business formation in developing economies. Natural or artificial devaluation can contribute short-term growth advantages, but public policy should seek to improve competitiveness to achieve higher levels of entrepreneurship for the country in the long term.

In the following section the main papers, both theoretical and empirical ones, related to devaluation and business formation are reviewed. In section 3, the premises of the models, the hypotheses raised, the methodology and the data used to test those hypotheses are presented. The results and their analyses are detailed in section 4, while in section 5 the conclusions, the discussion of main findings, the limitations encountered and some proposals for future research are presented.

2. Devaluation and business formation in different competitive contexts

Although a direct relationship is not established, the available literature is focused on mentioning how devaluation can dynamize some aspects of the economy that influence business formation, such as lower labor costs, reallocation of production factors and FDI increases. Theoretical approaches and empirical outcomes indicate that devaluation generates a reduction in labor costs in comparative terms. As labor costs are reduced, profit margins are increased and investment, savings and capital accumulation are fostered and these conditions favor NBF. (Levy-Yeyati and Sturzenegger, 2007; Bhalla, 2007 and Rodrik, 2008).

Theory also holds that devaluation phenomena influence productivity improvement by the reallocation of production factors (capital and labor force) from less profitable sectors to more profitable ones (Gala, 2007; Mbaye, 2012) and this can generate a positive effect on business formation. The productivity increase of production factors promote investment development and, in turn, NBF. This idea has been supported empirically by several papers, for example Mbaye (2012), using data from 72 countries between 1970 and 2008, found that an increase in devaluation favors growth through improvement in productivity.

Another impact channel of devaluation on business formation is FDI. Both the theory of imperfect capital markets and the theory of relative labor costs point out that an increase in REER goes hand in hand with increases in FDI inflows. According to Froot and Stein (1991), information asymmetry in capital markets causes foreign sources of funding to be more costly than domestic ones, which encourages the search for lower capital costs through asset acquisition in economies with devalued currencies. Likewise, the relative labor costs theory predicts that a depreciation of the real exchange rate increases FDI inflows attracted by lower labor costs (Kosteletou and Liargovas, 2000).

Working papers by Markusen and Venables (1999) and Barrios et al. (2005) put forward analytical models regarding the positive impact of FDI on the rise of new businesses. A “*spillover*” effect comes about because FDI increases the demand for intermediate products and services as well as the supply of capital, technology and new knowledge.

The results of empirical models reported by Zhang (2006), Yao and Wei (2007) confirm the existence of a positive FDI effect on the growth of the Chinese economy from the beginning of the 1980s and 1990s, as a result of the positive interaction between FDI and human capital that speeds up growth, and the stimulus to technological progress generated by FDI. Burke et al. (2007), Barbosa and Eiriz (2007) confirm an overall predominance of the “*spillover*”

effect over the “*crowd-out*” effect in United Kingdom and Portugal. Likewise, Xu and Chang (2008) report the existence of the “*spillover*” and “*crowd-out*” effects in China nationwide and regionwide respectively. Ayyagari and Kosova (2010) also contribute evidence of the positive effect of FDI on entrepreneurship at both inter-sectorial and intra-sectorial levels in the Czech Republic.

Despite the positive effect that REER devaluation is expected to have on investment, it has been empirically shown that its effect on the growth of the economies is only short-term, and on the contrary it doesn't turn out to be significant in the long term. The Bahmani-Oskooee research (1998) in 17 developing countries confirms the inexistence of a long-term devaluation effect on their production in the period 1973-1988. Karadeloglou et al. (2001) also reveal the absence of a long-term devaluation effect on Bulgaria's economy.

Another element to consider in the analysis of the relationship between REER and NBF is the existence of phenomena that affect them jointly. Reynolds et al. (1999, 2000), Edwards (1988), Elbadawi (1994) and Montiel (1999) identify the factors affecting the real exchange rate and new business formation simultaneously. Financial resources, institutional quality, government policies, technological transfer and trade openness, together with government consumption expenditure and capital inflows form part of the factors called “*economic fundamentals*”, whose behavior may generate revaluation or devaluation while affecting entrepreneurship levels.

The theoretical approaches discussed so far indicate that devaluation might generate some conditions that can promote investment in the short term; however, it is not clear whether such effects are transferred to the increase in new businesses and how lasting those effects can turn out to be. There is also evidence of the existence of variables that may simultaneously affect NBF and devaluation. It is necessary to establish empirically whether devaluation may have a residual significance on NBF after isolating those effects. These considerations will be taken into account in the methodological development of the econometric models.

The second focus of interest in this paper lies in establishing the benefits of a devaluation policy for business formation in countries with different competitiveness conditions. The interrelation between competitiveness, devaluation and business formation has been studied by some authors; for example Drine and Rault (2003) and Joyce and Kamas (2010) mention that REER appreciation in the long term is explained by the increase in exchange terms, increase in GDP per capita, capital inflows, but also by the countries' improvement in productivity.

Likewise, an increase in a country's competitiveness has positive effects on the increase in FDI, international trade and domestic demand. Several research studies in the field of entrepreneurship contribute evidence of the benefits of these phenomena on business formation (Clercq, 2008; Ayyagari y Kosova, 2010 and Misra et al., 2012). Apart from market factors, other institutional factors are evidence of the competitiveness increase of a country and at the same time favor new business formation. A clear example is the decrease in corruption levels, the decrease in costs and in the number of formalities to register new firms (Koveos et al., 2011 and Munemo, 2012).

In view of public policy developers' confidence in using devaluation to promote the growth of the domestic industry (to name but a few, governments of China, Japan, Thailand, South Korea and Colombia have resorted to devaluation), it is necessary not only to validate empirically whether devaluation has had a significant impact on new business formation but also to establish how this impact behaves when the country has worked on improving the competitiveness of its economy.

3. Empirical modeling

The objective is to model the effect of devaluation on new business formation in emerging countries through the panel data technique, establish if such effect is long-lasting and whether it is preserved after controlling the variables that, according to the literature, may simultaneously affect the behavior of REER and NBF. The methodological strategy is the following: a first model shows NBF in terms of REER, the complete specification of the model is described in the following way:

$$LNNBF_{it} = \alpha_i + \beta_1 LNREER_{i(t-l)} + \beta_k^T X_{i(t-j_k)} + \gamma_i + e_{it} \quad (3.1)$$

$LNNBF$ represents NBF density annual growth rate, i denotes countries, t denotes years, j denotes lags in the variables, and γ_i gathers the effect of temporal dummy variables on the model². α_i is a vector that contains the fixed effects per countries as confirmed by Hausman's (1978) and Breusch and Pagan's (1980)³ tests. $LNREER$ captures annual REER movements, while X constitutes a vector with control variables identified in the economic literature as NBF determinants. In order to verify whether the devaluation effect on NBF remains in the long term, the variable $LNREER$ is evaluated in lags greater than 2 years as shown below:

$$LNNBF_{it} = \alpha_i + \beta_1 LNREER_{i(t-l>2)} + \beta_k^T X_{i(t-j_k)} + \gamma_i + e_{it} \quad (3.2)$$

² The tests carried out in the Stata software revealed the effect of a dummy variable in the year 2009.

³ The presence of fixed effects in the model coincides with the findings of Sutaria and Hicks (2004).

In order to verify the behavior of REER in terms of the variables simultaneously affecting NBF, Edwards' (1988, 1989) and Hussain's (2011) approaches are followed. The authors agree on stating that REER movements in emerging economies are determined by short- and long-term factors, both domestic and foreign, that affect the relative price between tradable and non-tradable goods:

$$LNREER_{it} = \varphi_i + \phi_1 \psi_{i(t-n_k)} + \lambda_t + \mu_{it} \quad (3.3)$$

Where *LNREER* represents annual REER growth, *i* denotes countries, *t* denotes years, $k = 1, 2, 3, \dots, 6$, *n* represents the lags in the variables and φ_i is a vector that contains the fixed effects per countries as revealed by Hausman's (1978) and Breusch and Pagan's (1980)⁴ tests. λ is a dummy variable of temporal effects⁵, whereas ψ is the vector composed of control variables.

From the models above, the common variables are established that show significance to explain NBF and REER. The methodology used by Gwartney et al. (2004) is followed, substituting REER for the residual after extracting the effects of the variables that prove significant in both models. The econometric approach for this solution is the following:

$$LNNBF_{it} = \alpha_i + \pi_1 LNREER_{i(t-1)}^* + \pi_k^T X_{i(t-j_k)} + \gamma_t + \varepsilon_{it} \quad (3.4)$$

With

$$LNREER_{it}^* = LNREER_{it} - \phi_k^T Z_{i(t-n_k)} \quad (3.5)$$

Where *LNREER** is the residuals that result from extracting the effect of the variables causing both NBF and REER, which are represented in the vector *Z*. Please note that in the equation 3.5 common variables maintain the coefficients and lags defined in equation 3.3 and that $K \leq 3$. With the results of this model, it is sought to establish whether devaluation maintains its positive effect on NBF, even after subtracting the effect the common variables cause to REER and NBF.

Three final models are used to check whether the competitiveness level (measured specifically in terms of credit supply and institutional quality) of the countries significantly influences the devaluation effect on NBF. The effect of REER on NBF is evaluated when credit supply and institutional quality are below their mean in the country sample. The econometric approach is the following:

⁴ Tests carried out in the Stata software confirm the presence of fixed effects in the model.

⁵ The tests carried out in the Stata software revealed the effect of dummy variables in the years 2010 and 2011.

$$LNNBF_{it} = \alpha_i + \pi_1 LNREER_{i(t-1)}^* + \pi_2 LNCCI_{i(t-j_2)}^{\leq Q2} + \pi_k^T X_{i(t-j_k)}^* + \gamma_t + \varepsilon_{it} \quad (3.6)$$

$$LNNBF_{it} = \alpha_i + \pi_1 LNREER_{i(t-1)}^* + \pi_2 LNDPCS_{i(t-j_2)}^{\leq Q2} + \pi_k^T X_{i(t-j_k)}^* + \gamma_t + \varepsilon_{it} \quad (3.7)$$

$$LNNBF_{it} = \alpha_i + \pi_1 LNREER_{i(t-1)}^* + \pi_2 LNDPCS_{i(t-j_2)}^{\leq Q2} + \pi_3 LNCCI_{i(t-j_3)}^{\leq Q2} + \pi_k^T X_{i(t-j_k)}^* + \gamma_t + \varepsilon_{it} \quad (3.8)$$

In each of the equations, the variables $LNDPCS^{\leq Q2}$ and $LNCCI^{\leq Q2}$ represent the quartiles one and two of domestic credit supply to the private sector as a percentage of GDP and institutional quality, whereas the vector X contains the control variables. The results of these models will make it possible to determine whether the devaluation impact on NBF in developing economies is only significant when countries reach better competitiveness levels in terms of financial resource supply to entrepreneurs and the quality of their public institutions.

4. Description of variables

The dependent variable (NBF) is measured with the new firm formation density, defined as the number of new firms registered per 1,000 economically active people reported by Doing Business for 30 emerging countries in the period 2004-2011 (Klapper et al., 2007; Klapper and Love, 2010; Munemo, 2012; Herrera et al., 2014). Unlike other entrepreneurship measures such as those of Global Entrepreneurship Monitor (GEM) that reflect the informality of the economies by capturing individuals' intention of starting a business, the new firm formation density shows greater adjustment and consistency in measuring NBF across countries by taking into account only the number of firms that become part of the formal economy annually; however, it has the limitation of not considering other forms of organization of entrepreneurial activity.

Some control variables are used to verify the robustness of the results. First, government expenditure (LNGCE). Greater government spending on acquiring consumer goods, which later reach households through social assistance programs or subsidies, can have a "crowd-out" effect on private consumption and may become a barrier to the entry of new firms. In order to measure this effect, recent research studies have used government final consumption expenditure as a percentage of GDP (Aidis, 2012).

NBF is expected to be influenced by the behavior of demand in the markets (Evans and Jovanovic, 1989). The greater demand for goods and services opens up opportunities for the entry of new firms in expansion times (Reynolds et al., 1999). GDP per capita (LNGDPP) has

been widely used as proxy for demand behavior (Spencer and Gómez, 2004). Since some empirical studies suggest the existence of a shelter effect, whereby individuals undertake new businesses as a way of generating income in a context of increased unemployment (Reynolds et al., 2000), the unemployment rate is also included (LNUE).

Flexibility in the procedures required for new business formation is an institutional factor of great impact on NBF dynamics. Bruton et al. (2010) point out that the excess of procedures required for registering a business (LNSPRB) is a factor that can discourage the rise of new companies. Research studies in emerging countries reveal a negative association between the increase in the number of procedures and the entry rate of new firms (Klapper et al., 2007; Bruhn, 2011).

The variables used to measure specific aspects of the competitiveness level of an economy are described and justified below. First, there's the availability of credit for the private sector as a percentage of GDP (LNDCPS), since availability of credit access for entrepreneurs constitutes a decisive factor for the leverage of nascent firms (Reynolds et al., 1999, 2000). A large number of empirical studies contribute evidence for the positive relationship between credit supply and entrepreneurship rate (Black and Strahan, 2002; Klapper et al., 2007; Robb and Robinson, 2012).

The degree of business transparency is another institutional factor that affects the entry of new firms in emerging countries. Lower corruption levels generate greater confidence in individuals towards public institutions, reduce inefficiencies and eliminate additional costs for entrepreneurial activity that end up facilitating entrepreneurship. In order to assess the institutional quality of countries, the corruption index (LNCI) reported by Transparency International is used, which was also applied in recent research (Wu and Liang, 2012; Dreher and Gassebner, 2013). Empirically, authors such as Wei (2000) and Quéré et al. (2005) contribute evidence on the negative impact that corruption in public institutions has on investment in developing countries. High corruption levels deprive countries of competitiveness and turn them less attractive for investors.

It is known that the competitiveness level reached by countries positively influences the NBF rate. The World Economic Forum, from the 12 pillars that make up the global competitiveness index, highlights the importance that government transparency has to avoid extra costs for companies and to facilitate economic development, as well as the credit access facility for entrepreneurs that allows them to leverage their investment projects (Schwab, 2011, 2012). The Global Entrepreneurship Monitor (2011) in its conceptual model also remarks how the funding of ventures and the strength of institutions accompanied by public policy help in promoting entrepreneurial activity.

The control variables for the devaluation model are mentioned below. Technological progress: According to the productivity differential theory (Balassa, 1964 and Samuelson, 1964), a productivity increase in tradable sectors boosts an increase in salaries in the overall economy, which in turn generates a price increase in non-tradable goods and, as a result, a decrease in REER. Several studies have considered the growth of GDP per capita (LNGDPP) as an appropriate proxy that gathers the effect of productivity derived from technological progress on REER (Drine and Rault, 2003; Kim and Korhonen, 2005).

Government consumption: According to the dynamic model put forward by Edwards (1988), the increase in government consumption of non-tradable goods has the effect of appreciating the real exchange rate in the short term. In order to measure this variable, government final consumption expenditure as a percentage of GDP (LNGCE) is used. Authors like Hussain (2011) confirm appreciation of REER in six emerging countries of the Southeast Asia derived from a greater share in government final expenditure on GDP.

Domestic credit: An expansive policy in terms of credit growth can result in a greater demand for tradable goods, non-tradable goods and financial assets (Edwards, 1989). It is expected that the increase in domestic credit can generate a greater demand for non-tradable goods thus devaluing REER. In order to measure the effect of domestic credit on REER, domestic credit to private sector (LNDPCPS) as a percentage of GDP is used as a proxy variable; this variable has been employed in other studies for example in Daboh (2010).

The nominal exchange rate: Nominal exchange rate movements are expected to influence REER behavior positively in the short term. Several studies in emerging economies back up the positive effect of the nominal exchange rate (LNNEER) on REER (Edwards, 1988; Wahid and Shahbaz, 2009).

Openness degree of the economy: Following Elbadawi (1994), the sum of exports and imports of goods and services (LNTDE) as a percentage of GDP is used in this study to measure the openness degree of the economies. Greater openness of the emerging economies is expected to boost the non-tradable goods sector and to generate revaluation of REER.

Foreign direct investment inflows: FDI increments increase capital stock in local economies and generate a “spillover” effect of technological transfer that increases production and reduces prices of non-tradable goods, which in turn causes a devaluation of REER (Rehman et al., 2010). A higher FDI also translates itself into greater investment in tradable sectors, which generates a devaluation effect of the real exchange rate, as reported by Elbadawi (1994). Net foreign direct investment inflows (LNFDI) as a percentage of GDP are used to measure the FDI effect.

5. Data and sample

The information on NBF density analyzed in this study corresponds to 209 data reported by Doing Business in the period 2004-2011 regarding the number of firms with limited liability registered per every 1,000 people of working age (15 to 64 years old) for a sample of 30 emerging countries or frontier economies according to the Dow Jones and The Economist. The sample is composed by six Latin American countries, ten European countries, ten countries from the Asian continent and four African ones, as shown on Table 1. The sample heterogeneity makes it possible to obtain general results on the NBF pattern according to the specific characteristics of this type of economies, regardless of the geographical variable.

<Insert Table 1 here>

The 300 data analyzed about the REER to evaluate its impact on NBF are obtained from databases made available by Bruegel for the period 2002-2011. The REER of each country is calculated against a currency basket of its trading partners (Darvas, 2012). Bruegel is an organization specialized in economic studies founded in 2004 thanks to the initiative of 12 governments of the European Union (EU), and its databases have also been used in recent research (Darvas, 2013). The additional information on the control variables used in the models was obtained from the World Bank database, as described in Appendix A. The difference in time periods of the two variables of interest is due to the limitations in data availability for NBF, which reduces the period under analysis. Descriptive statistics of the variables included in the study are found in Appendix C.

6. Results

The results are estimated with Panel-Corrected Standard Errors (PCSE) (Beck and Katz, 1995) in order to solve problems of heteroskedasticity, autocorrelation and potential cross-correlation bias identified in the data. Despite argued limitations regarding the PCSE method associated with the difficulty of achieving a non-singular estimate of the covariance matrix when $T < N$ (Hoechle, 2007), empirical applications in this type of samples carried out by Friedland and Sanders (1985) and Scholz (1986) report efficient estimates in such cases.

The Hausman test (1978) reveals the presence of fixed effects, which matches the findings of Sutaria and Hicks (2004). The influence of non-observed characteristics in some years and countries on the results is controlled by dummy variables. The initial results of the NBF model are presented in Table 2.

<Insert Table 2 here>

General statistical parameters show that both models prove to be well-defined, with R-squared above 90% indicating that the exogenous variables used account for a great deal of NBF behavior. The significance of LNREER results in $t=2$, while LNGCE, LNDCPS and LNGDPP variables prove to be significant in $t=1$, which diminishes potential problems of endogeneity. Correlation matrix outputs contained in Appendix B rule out collinearity among the variables. The unit root tests in Appendix D reject the presence of unit roots and ensure stationarity.

The result of model 1 indicates that REER devaluation episodes generate a positive impact on NBF rate in emerging countries in the short term. In model 2, the devaluation effect is seen as not lasting over time and it disappears within a three-year window.

Another part of the results shows the positive impact of competitiveness variables such as LNCI and LNDCPS on NBF. LNCI with a coefficient of 0.906 generates the greatest positive impact on business formation, which reaffirms the claim that higher transparency levels favor entrepreneurial climate by creating confidence in business people, as remarked by other authors (Koveos et al., 2011 and Munemo, 2012). LNDCPS also has a positive impact on NBF; international evidence about the importance of credit availability facilitating entrepreneurship processes is confirmed. While the positive sign of LNGDPP reveals the effect of the increased demand for goods and services that encourages the entry of new firms as expected.

On the contrary, other control variables such as LNGCE and LNSPRB become barriers for NBF as pointed out by the theory. The negative sign of LNGCE confirms the “*crowd-out*” effect on private demand that limits opportunities for the entry of new firms, as remarked by Startz (1989). Likewise, the excess of procedures required for organizing new firms ends up discouraging entrepreneurship initiatives.

The results of the devaluation model confirm that LNGCE, LNDCPS and LNGDPP variables also determine REER behavior as shown in Table 3.

<Insert Table 3 here>

The impact of LNGCE, LNNEER, LNTDE and FDI variables on REER behavior matches what is stated by the literature. The positive sign of LNGDPP contrasts with what the Balassa-Samuelson theory predicts, indicating that the increase in productivity in tradable sectors ends up decreasing the price of non-tradable goods instead of increasing it, thus devaluing REER. However, this theory has been structured based on evidence from developing countries, which does not necessarily match the behavior of emerging economies. For emerging countries evidence has been found that matches our results.

For example, some Asian countries like China, Hong Kong, Malaysia and Singapore experienced fast economic growth before the financial crisis in 1997, but their currencies didn't become revalued, they even tended to get devalued. Such a seemingly contradictory phenomenon has been explained by He (2010) as a result of the increase in the variety of non-tradable products on the markets which, together with the monopolistic effect that companies may have, tends to maintain or reduce the prices of these goods, thus generating REER devaluation.

In the fourth model, it can be observed how LNREER* residuals continue being significant for NBF, showing that despite the effect of LNGCE, LNDCPS and LNGDPP variables on REER devaluation generates additional effects on NBF. The sign and significance of the remaining control variables continue behaving properly, though their coefficients are slightly increased. The adjusted results are shown in the second column of Table 4.

<Insert Table 4 here>

Finally, it is also being tested how the devaluation effect varies when countries experience low competitiveness levels in LNCI and LNDCPS variables. The results of models five and six show that the devaluation impact diminishes in cases where the financial supply and institutional quality are low. Model seven, in turn, reveals how the devaluation effect loses significance when both variables present levels below their mean or quartile two. These findings support the conclusion that when institutional quality and credit supply to the private sector in emerging economies are low, devaluation loses its effect on NBF. Therefore, minimum competitiveness conditions are required in the economies for devaluation phenomena to influence business formation in emerging countries.

The negative sign in the $LNCI^{=Q2}$ variable in models six and seven shows that the relationship between institutional quality and NBF does not follow a linear behavior; on the contrary, it charts a "U" course. This indicates that countries with high corruption levels or high institutional quality levels experience higher NBF rates, as pointed out by Wu and Liang (2012). Individuals make use of corruption through their influence on public power as a mechanism to gain advantages in resource appropriation and market access that benefit their businesses. As Dreher and Gassebner (2013) argue, when countries have inefficient administrative systems with overregulation (excess in formalities) or high capital requirements, corruption becomes an alternate mechanism that favors NBF. High corruption levels, the low quality of public goods and the State's weakness in guaranteeing law enforcement encourage informal entrepreneurship in the economies (Johnson et al., 1997). Low institutional quality is also reflected in a weak protection of property rights that triggers a predatory attitude in entrepreneurs, who seek to obtain higher returns seizing on technological developments created by others (González, 2005).

7. Conclusions

Research results yield evidence regarding the initial question about the impact of devaluation on increased business formation in emerging countries. By means of a system of four models under a panel data structure for 30 emerging countries in the period 2004-2011, it is shown that REER devaluation stimulates NBF in the short term. However, this effect is not lasting over time and loses significance in a time horizon greater than two years, which suggests that other effects such as the increase in import costs brought about by devaluation can dilute its benefits over time (Krugman, 1978 and Edwards, 1986).

Yet, the devaluation impact on NBF ceases to exist when countries are not competitive. Results allow us to conclude that there are minimum levels in competitiveness factors from which devaluation may have significance on NBF. That is to say that, although devaluation generates opportunities for entrepreneurship in the economies, it is necessary to create conditions that can make it easy for new entrepreneurs to develop their business ideas, such as funding programs and improvement in public institutions which may guarantee property rights and free competition. The findings reported here show how devaluation ends up being an unsustainable mechanism to foster entrepreneurship in emerging countries. Any positive effect that may be derived from devaluation can only be maintained in the long term when countries manage to improve their competitiveness.

Finally, the development of this paper does not escape information limitations on variables of interest that might contribute greater robustness to its results. In particular, the use of a wider analysis period and the need to include other competitiveness variables such as quality in infrastructure, quality in education and tax rates, among others. However, the results reported, apart from being consistent with the related literature, throw light for future research on, for example, the assessment of the devaluation effect on net NBF at an inter- and intra-sectorial level, the knowledge of NBF patterns per regions or per development level of the countries, the knowledge of the devaluation impact on NBF in the presence of other variables such as the number of free trade treaties signed by countries or the determination of the threshold of competitiveness variables from which devaluation proves to be significant for NBF.

Appendix A. Variables used

The variables included in the models were selected after a close review of the literature related to REER and NBF determinants.

Variable	Acronym	Definition	Expected effect		Sources:
			NBF	REER	
New business formation	NBF	Number of limited liability firms registered per 1,000 people of working age (15 to 64 years old).	n.a.	n.a.	Doing Business
Real effective exchange rate	REER	Evolution of the real value of a currency against a currency basket of its trading partners.	+	n.a.	Bruegel
Government final consumption expenditure	GCE	It includes all current expenditure for the purchase of goods and services (including salary payment). It also includes most defense and security expenses. It is expressed as a GDP percentage.	-	-	World Bank
Domestic credit to private sector	DCPS	Financial resources lent to the private sector through loans, purchases of non-equity securities, trade credits and other accounts receivable that establish a claim for repayment. It is expressed as a GDP percentage.	+	+	World Bank
GDP per capita	GDPP	Gross domestic product per capita converted to international dollars using purchasing power parity rates.	+	-	World Bank
Unemployment rate	UE	Share of the labor force that is without work but available for and seeking employment.	+/-	n.a.	World Bank
Start-up procedures required to register a business	SPRB	It includes permits and licenses necessary to complete all inscriptions, verifications and notifications to start operations.	-	n.a.	World Bank
Institutional quality	CI	Public sector corruption level perception, rated from 1 to 10. Where 1 is highly corrupt and 10 is highly clean.	+	n.a.	Transparency International
Nominal effective exchange rate	NEER	Cost or relative price of a country's currency in relation to a currency basket of its trading partners.	n.a.	+	Bruegel
Foreign direct investment net inflows	FDI	Net inflows of investment to acquire an interest (10 per cent or more of voting stock) in an enterprise operating in an economy other than that of the investor. It is expressed as a GDP percentage.	n.a.	+	World Bank
Openness degree of the economy	TDE	It is the sum of exports and imports of goods and services as a share of GDP.	n.a.	-	World Bank

Appendix B. Correlation matrix among variables

The correlation matrix permits the identification of possible causality relations among variables when coefficients turn out to be high and are accompanied by marked significance. Based on this, no significant problems of multicollinearity are detected among the variables of the models.

	LNNBF	LNREER	LNNEER	LNFDI	LNDCPS	LNSPRB	LNGCE	LNGDPP	LNUE	LNCI	LNTDE
LNNBF	1										
LNREER	0.081 0.243	1									
LNNEER	0.227 0.001	0.582 0.000	1								
LNFDI	0.496 0.000	0.041 0.485	-0.057 0.328	1							
LNDCPS	0.488 0.000	0.300 0.000	0.133 0.022	0.340 0.000	1						
LNSPRB	-0.603 0.000	-0.307 0.000	-0.172 0.005	-0.395 0.000	-0.482 0.000	1					
LNGCE	0.285 0.000	-0.001 0.983	-0.039 0.507	0.086 0.140	0.205 0.000	-0.121 0.048	1				
LNGDPP	0.831 0.000	0.080 0.165	-0.012 0.832	0.519 0.000	0.377 0.000	-0.505 0.000	0.239 0.000	1			
LNUE	-0.007 0.925	-0.078 0.224	0.014 0.832	-0.165 0.010	-0.173 0.007	0.104 0.127	0.386 0.000	-0.117 0.068	1		
LNCI	0.729 0.000	0.173 0.003	0.062 0.288	0.576 0.000	0.679 0.000	-0.512 0.000	0.228 0.000	0.777 0.000	-0.079 0.218	1	
LNTDE	0.636 0.000	0.104 0.072	0.086 0.137	0.541 0.000	0.622 0.000	-0.554 0.000	0.070 0.229	0.643 0.000	-0.206 0.000	0.6972 0.000	1

Appendix C. Descriptive statistics of variables

The following is a summary of the descriptive statistics of the variables analyzed.

Variable	Obs	Mean	Std. Dev.	Min	Max
LNNBF	209	0.2818863	1.55042	-3.618687	3.320515
LNREER	300	4.588745	0.1147184	4.160587	4.913376
LNGCE	299	-1.959631	0.3061049	-2.685486	-1.450081
LNDPCS	297	-0.7565576	0.6505255	-2.288395	0.7041503
LNGDPP	300	9.129745	0.7442018	7.505405	10.88914
LNUE	246	-2.517127	0.5627875	-4.422849	-1.164752
LNSPRB	270	2.054048	0.4024881	1.098612	2.944439
LNCI	300	1.399485	0.3391012	0.6418539	2.24071
LNNEER	300	4.59344	0.1109252	4.223138	5.045853
LNTDE	299	4.441335	0.6627107	3.096405	6.100424
LNFDI	294	-3.421338	1.033756	-7.475258	-0.6559314

Appendix D. Unit root tests

In the following table, the results of the unit root tests applied to each variable using the Phillips-Perron method are reported. The tests do not present arguments in favor of the presence of unit roots, making it possible to point out that the series analyzed prove to be stationary.

Variables	Inverse chi-square(P)		Inverse normal (Z)		Inverse logit (L*)		Modified inverse chi-square (Pm)	
	Statistical	P-value	Statistical	P-value	Statistical	P-value	Statistical	P-value
LNNBF	375.548	0.000	-9.143	0.000	-18.900	0.000	28.806	0.000
LNREER	265.639	0.000	-4.714	0.000	-9.215	0.000	18.772	0.000
LNGCE	189.156	0.000	-1.857	0.032	-5.085	0.000	11.790	0.000
LNDPCS	210.169	0.000	-1.772	0.038	-7.419	0.000	13.709	0.000
LNGDPP	115.073	0.000	-1.764	0.039	-2.065	0.020	5.028	0.000
LNUE	79.448	0.000	1.406	0.920	0.536	0.704	1.991	0.023
LNSPRB	120.295	0.000	-0.563	0.287	-3.431	0.000	5.504	0.000
LNCI	122.428	0.000	0.846	0.081	-0.201	0.421	5.699	0.000
LNNEER	326.17	0.000	-7.544	0.000	-13.61	0.000	24.298	0.000
LNFDI	291.292	0.000	-5.66	0.000	-11.165	0.000	21.114	0.000
LNTDE	115.545	0.000	-1.519	0.064	-2.433	0.008	5.071	0.000

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Table 1. List of countries that make up the study sample.

Country	Region	Country	Region
Argentina ^a	Latin America	Malaysia ^a	Asia
Brazil ^a	Latin America	Mauritius ^a	Africa
Bulgaria ^a	Europe	Mexico ^a	Latin America
Chile ^a	Latin America	Morocco ^a	Africa
Colombia ^a	Latin America	Pakistan ^a	Asia
Czech Republic ^a	Europe	Peru ^a	Latin America
Egypt ^a	Africa	Philippines ^a	Asia
Estonia ^a	Europe	Poland ^a	Europe
Hong Kong ^b	Asia	Romania ^a	Europe
Hungary ^a	Europe	Singapore ^b	Asia
India ^a	Asia	Slovakia ^a	Europe
Indonesia ^a	Asia	South Africa ^a	Africa
Jordan ^a	Asia	Sri Lanka ^a	Asia
Latvia ^a	Europe	Thailand ^a	Asia
Lithuania ^a	Europe	Turkey ^a	Europe

^a Emerging countries according to the Dow Jones and The Economist.

^b Emerging countries according to The Economist.

Table 2

<i>Business formation in emerging countries</i>				
<i>Independent variables</i>	<i>Dependent variable</i>			
	<i>New business formation</i>			
	(1)		(2)	
<i>LNREER</i> _{<i>t-2</i>}	0.388 (0.050)**		-	
<i>LNREER</i> _{<i>t-3</i>}	-		0.241 (0.269)	
<i>LNGCE</i> _{<i>t-1</i>}	-1.308 (0.000) ***		-1.298 (0.000) ***	
<i>LNDGPS</i> _{<i>t-1</i>}	0.278 (0.000) ***		0.285 (0.000) ***	
<i>LNGDPP</i> _{<i>t-2</i>}	0.611 (0.000) ***		0.572 (0.000) ***	
<i>LNUE</i> _{<i>t-1</i>}	0.256 (0.000) ***		0.245 (0.000) ***	
<i>LNSPRB</i> _{<i>t-1</i>}	-0.430 (0.000) ***		-0.507 (0.000) ***	
<i>LNCI</i> _{<i>t-1</i>}	0.906 (0.000) ***		0.966 (0.000) ***	
<i>Constant</i>	-9.549 (0.000) ***		-8.464 (0.000) ***	
<i>Temporal dummies</i>	Yes		Yes	
<i>Dummies per countries</i>	Yes		Yes	
<i>R</i> ²	0.954		0.956	
<i>Observations</i>	190		165	
<i>Number of countries</i>	30		30	
Tests	Stat.	P-value	Stat.	P-value
Wald Chi(26)	5317.50	0.000	4860.09	0.000
Hausman	35.35	0.000	25.93	0.001
Breusch-Pagan	351.09	0.000	264.82	0.000
Modified Wald het.	2160.46	0.000	56072.14	0.000
Serial Correlation LM	18.012	0.000	18.012	0.000
Model	PCSE c(ar1)		PCSE c(ar1)	

Depended variable i at year t is the number of firms registered per 1,000 economically active population. Independent variables are: Natural log of annual REER in $t=2$ ($LNREER_{t-2}$), natural log of the government expenditures as a percentage of GDP in $t=1$ ($LNGCE_{t-1}$), natural log of the availability of credit for the private sector as a percentage of GDP in $t=1$ ($LNDGPS_{t-1}$), natural log of the growth of GDP per capita movement in $t=2$ ($LNGDPP_{t-2}$), natural log of the unemployment rate in $t=1$ ($LNUE_{t-1}$), natural log of the excess of procedures required for registering a business in $t=1$ ($LNSPRB_{t-1}$) and natural log of the corruption index in $t=1$ ($LNCI_{t-1}$). The coefficient of the regression are based on Panel-Corrected Standard Errors (PCSE). The statistical T is specified in parentheses. *** and ** which means significance at 1% and 5% respectively.

Table 3

<i>Devaluation in emerging countries</i>		
<i>Independent variables</i>	<i>Dependent variable</i>	
	Real effective exchange rate	
	(3)	
<i>LNGCE_{t-1}</i>	-0.069 (0.009)***	
<i>LNDCPS_{t-1}</i>	0.053 (0.000)***	
<i>LNGDPP_{t-2}</i>	0.139 (0.000)***	
<i>LNNEER_t</i>	0.646 (0.000)***	
<i>LNTDE_{t-1}</i>	-0.068 (0.000)***	
<i>LNFDI_{t-1}</i>	0.017 (0.003)***	
<i>Constant</i>	0.366 (0.152)	
<i>Temporal dummies</i>	Yes	
<i>Dummies per countries</i>	Yes	
<i>R²</i>	0.994	
<i>Observations</i>	233	
<i>Number of countries</i>	30	
Tests	Stat.	P-value
Wald Chi(18)	520.89	0.000
Hausman	143.55	0.000
Breusch-Pagan	0.74	0.389
Modified Wald het.	4386.52	0.000
Serial Correlation LM	466.221	0.000
Model	PSCE c(ar1)	

Depended variable i at year t is the number of firms registered per 1,000 economically active population. Independent variables are: Natural log of the government expenditures as a percentage of GDP ($LNGCE_{t-1}$), natural log of the availability of credit for the private sector as a percentage of GDP ($LNDCPS_{t-1}$), natural log of the growth of GDP per capita movement ($LNGDPP_{t-2}$), natural log of the effect of nominal exchange rate ($LNNEER_t$), natural log of the sum of exports and imports of goods and services as a percentage of GDP ($LNTDE_{t-1}$) and natural log of the net foreign direct investment as a percentage of GDP ($LNFDI_{t-1}$). The coefficient of the regression are based on Panel-Corrected Standard Errors (PCSE). The statistical T is specified in parentheses. *** means significance at 1%.

Table 4

Devaluation, competitiveness and business formation in emerging countries								
Independent variables	Dependent variable							
	New business formation							
	(4)	(5)	(6)	(7)				
$LNREER^*_{t-2}$	0.539 (0.020)**	0.454(0.061)*	0.468(0.067)*	0.273(0.296)				
$LNGCE_{t-1}$	-1.366 (0.000)***	-1.351(0.00)***	-1.367(0.000)***	-1.354(0.000)***				
$LNDPCS_{t-1}$	0.296 (0.000) ***	-	0.468(0.000)***	-				
$LNDPCS^{<=Q2}_{t-1}$	-	0.191(0.000)***	-	0.364(0.000)***				
$LNGDPP_{t-2}$	0.630 (0.000) ***	0.596(0.000)***	0.877(0.000)***	0.861(0.000)***				
$LNUE_{t-1}$	0.245 (0.000) ***	0.255(0.000)***	0.220(0.000)***	0.243(0.000)***				
$LNSPRB_{t-1}$	-0.491 (0.000)***	-0.568(0.000)***	-0.557(0.000)***	-0.674(0.000)***				
$LNCI_{t-1}$	1.010 (0.000) ***	1.132(0.000)***	-	-				
$LNCI^{<=Q2}_{t-1}$	-	-	-0.214(0.002)***	-0.243(0.001)***				
Constant	-9.836 (0.000)***	-9.286(0.000)***	-10.082(0.000)***	-9.038(0.000)***				
Temporal dummies	Yes	Yes	Yes	Yes				
Dummies per countries	Yes	Yes	Yes	Yes				
R^2	0.957	0.954	0.954	0.947				
Observations	165	165	165	165				
Number of countries	30	30	30	30				
Tests	Stat.	P-value	Stat.	P-value	Stat.	P-value	Stat.	P-value
Wald Chi(26)	5238.16	0.000	4352.43	0.000	4580.59	0.000	3925.07	0.000
Hausman	32.38	0.000	33.30	0.000	24.41	0.002	29.47	0.000
Breusch-Pagan	269.35	0.000	269.25	0.000	268.75	0.000	268.84	0.000
Modified Wald het.	4980.05	0.000	6721.93	0.000	4720.40	0.000	6758.41	0.000
Serial Correlation	18.288	0.000	18.861	0.000	18.225	0.000	18.848	0.000
LM								
Model	PCSE c(ar1)	PCSE c(ar1)	PCSE c(ar1)	PCSE c(ar1)				

Depended variable i at year t is the number of firms registered per 1,000 economically active population. Independent variables are: Natural log of the residuals that result from extracting the effect of the variables causing both NBF and REER ($LNREER^*_{t-2}$), natural log of the government expenditures as a percentage of GDP ($LNGCE_{t-1}$), natural log of the availability of credit for the private sector as a percentage of GDP ($LNDPCS_{t-1}$), natural log of the quartile one of domestic credit supply to the private sector as a percentage of GDP ($LNDPCS^{<=Q2}_{t-1}$), natural log of the growth of GDP per capita movement ($LNGDPP_{t-2}$), natural log of the unemployment rate ($LNUE_{t-1}$), natural log of the excess of procedures required for registering a business ($LNSPRB_{t-1}$), natural log of the corruption index ($LNCI_{t-1}$) and natural log of the quartile two of domestic credit supply to the private sector as a percentage of GDP ($LNCI^{<=Q2}_{t-1}$). The coefficient of the regression are based on Panel-Corrected Standard Errors (PCSE). The statistical T is specified in parentheses. ***, ** and * which means significance at 1%, 5% and 10% respectively.