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AGING AND QUENCHED LOCALIZATION FOR ONE-DIMENSIONAL RANDOM WALKS IN RANDOM ENVIRONMENT IN THE SUB-BALLISTIC REGIME

NATHANAËL ENRIQUEZ, CHRISTOPHE SABOT, AND OLIVIER ZINDY

Abstract. We consider transient one-dimensional random walks in a random environment with zero asymptotic speed. An aging phenomenon involving the generalized Arcsine law is proved using the localization of the walk at the foot of “valleys” of height $\log t$. In the quenched setting, we also sharply estimate the distribution of the walk at time t .

1. INTRODUCTION

One-dimensional random walks in random environment have been the subject of constant interest in physics and mathematics for the last thirty years since they naturally appear in a great variety of situations in physics and biology.

In 1975, Solomon gave, in a seminal work [26], a criterion of transience-recurrence for such walks moving to the nearest neighbours, and shows that three different regimes can be distinguished: the random walk may be recurrent, or transient with a positive asymptotic speed, but it may also be transient with zero asymptotic speed. This last regime, which does not exist among usual random walks, is probably the one which is the less well understood and its study is the purpose of the present paper.

Let us first recall the main existing results concerning the other regimes. In his paper, Solomon computes the asymptotic speed of transient regimes. In 1982, Sinai states, in [25], a limit theorem in the recurrent case. It turns out that the motion in this case is unusually slow. Namely, the position of the walk at time n has to be normalized by $(\log n)^2$ in order to present a non trivial limit. In 1986, the limiting law is characterized independently by Kesten [22] and Golosov [19]. Let us notice here that, beyond the interest of his result, Sinai introduces a very powerful and intuitive tool in the study of one-dimensional random walks in random environment. This tool is the potential, which is a function on \mathbb{Z} canonically associated to the random environment. The potential itself is a usual random walk when the transition probabilities at each site are independent and identically distributed (i.i.d.).

The proof by Sinai of an annealed limit law in the recurrent case is based on a quenched localization result. Namely, a notion of valley of the potential is introduced, as well as an order on the set of valleys. It is then proved that the walk is localized at time t , with a probability converging to 1, around the bottom of the smallest valley of depth bigger than $\log t$ surrounding the origin. An annealed convergence in law of this site normalized by $(\log t)^2$ implies the annealed limiting law for the walk.

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In the case of transient random walks in random environment with zero asymptotic speed, the proof of the limiting law by Kesten, Kozlov and Spitzer [23] does not follow this scheme. Therefore an analogous result to Sinai's localization in the quenched setting was missing. As we will see, the answer to this question is more complicated than in the recurrent case but still very explicit.

In the setting of sub-ballistic transient random walks, the valleys we introduce are, like in [15] and [24], related to the excursions of the potential above its past minimum. Now, the key observation is that with a probability converging to 1, the particle at time t is located at the foot of a valley having depth and width of order $\log t$. Therefore, since the walk spends a random time of order t inside a valley of depth $\log t$, it is not surprising that this random walk exhibits an aging phenomenon.

What is usually called aging is a dynamical out-of-equilibrium physical phenomenon observed in disordered systems like spin-glasses at low temperature, defined by the existence of a limit of a given two-time correlation function of the system as both times diverge keeping a fixed ratio between them; the limit should be a non-trivial function of the ratio. It has been extensively studied in the physics literature, see [9] and therein references.

More precisely, in our setting, Theorem 1 expresses that, for each given ratio $h > 1$, the probability that the particle remains confined within the same valley during the time interval $[t, th]$. This probability is expressed in terms of the generalized Arcsine law, which confirms the status of universality ascribed to this law by Ben Arous and Černý in their study of aging phenomena arising in trap models [4].

Recall that the trap model is a model of random walk that was first proposed by Bouchaud and Dean [8, 10] as a toy model for studying this aging phenomenon. In the mathematics literature, much attention has recently been given to the trap model, and many aging result were derived from it, on \mathbb{Z} in [17] and [3], on \mathbb{Z}^2 in [7], on \mathbb{Z}^d ($d \geq 3$) in [5], or on the hypercube in [1, 2]. A comprehensive approach to obtaining aging results for the trap model in various settings was later developed in [6].

Let us finally mention that Theorem 1 generalizes the aging result obtained by heuristical methods of renormalization by Le Doussal, Fisher and Monthus in [13] in the limit case when the bias of the random walk defining the potential tends to 0 (the case when this bias is 0 corresponding to the recurrent regime for the random walk in random environment). The recurrent case, which also leads to aging phenomenon, was treated in the same article and rigorous arguments were later presented by Dembo, Guionnet and Zeitouni in [12].

The second aspect of our work concerns localization properties of the walk and can be considered as the analog of Sinai's localization result in the transient setting. Unlike the recurrent case, the random walk is not localized near the bottom of a single valley. Nevertheless, if one introduces a confidence threshold α , one can say that, asymptotically, at time t , with a probability converging to 1 on the environment, the walk is localized with probability bigger than α around the bottoms of a finite number of valleys having depth of order $\log t$. This number depends on t and on the environment, but is not converging to infinity with t . Moreover, in Theorem 2 and Corollary 1 we sharply estimate the probability for the walk of being at time t in each of these valleys.

2. NOTATION AND MAIN RESULTS

Let $\omega := (\omega_i, i \in \mathbb{Z})$ be a family of i.i.d. random variables taking values in $(0, 1)$ defined on Ω , which stands for the random environment. Denote by P the distribution of ω and by E the corresponding expectation. Conditioning on ω (i.e. choosing an environment), we define the random walk in random environment $X = (X_n, n \geq 0)$ on $\mathbb{Z}^{\mathbb{N}}$ as a nearest-neighbor random walk on \mathbb{Z} with transition probabilities given by ω : $(X_n, n \geq 0)$ is the Markov chain satisfying $X_0 = 0$ and for $n \geq 0$,

$$\begin{aligned} P_\omega(X_{n+1} = x+1 | X_n = x) &= \omega_x, \\ P_\omega(X_{n+1} = x-1 | X_n = x) &= 1 - \omega_x. \end{aligned}$$

We denote by P_ω the law of $(X_n, n \geq 0)$ and E_ω the corresponding expectation. We denote by \mathbb{P} the joint law of $(\omega, (X_n)_{n \geq 0})$. We refer to Zeitouni [27] for an overview of results on random walks in random environment. Let us introduce

$$\rho_i := \frac{1 - \omega_i}{\omega_i}, \quad i \in \mathbb{Z}.$$

Our first main result is the following theorem which shows aging phenomenon in the transient sub-ballistic regime.

Theorem 1. *Let $\omega := (\omega_i, i \in \mathbb{Z})$ be a family of independent and identically distributed random variables such that*

- (a) *there exists $0 < \kappa < 1$ for which $E[\rho_0^\kappa] = 1$ and $E[\rho_0^\kappa \log^+ \rho_0] < \infty$,*
- (b) *the distribution of $\log \rho_0$ is non-lattice.*

Then, for all $h > 1$ and all $\eta > 0$, we have

$$\lim_{t \rightarrow \infty} \mathbb{P}(|X_{th} - X_t| \leq \eta \log t) = \frac{\sin(\kappa\pi)}{\pi} \int_0^{1/h} y^{\kappa-1} (1-y)^{-\kappa} dy.$$

Remark 1. *The statement of Theorem 1 could be improved in the following way: the size of the localization window $\eta \log t$ could be replaced by any positive function $a(t)$ such that $\lim_{t \rightarrow \infty} a(t) = +\infty$ and $a(t) = o(t^\kappa)$ (the authors would like to thank Yueyun Hu who raised this question). The extra constraint $a(t) = o(t^\kappa)$ comes from the fact that t^κ is the order of the distance between successive valleys where the RWRE can be localized. We did not write the proof of the theorem in this more general version since it induces several extra technicalities and makes the proof harder to read. Moreover $\eta \log t$ represents an arbitrary portion of a typical valley (which is of size of order $\log t$) where the RWRE can be localized, and is therefore a natural localization window.*

Let us now recall some basic result about X_n : under the same assumptions (a)-(b), Kesten, Kozlov and Spitzer [23] proved that X_n/n^κ converges in law to $C_\kappa (\frac{1}{\mathcal{S}_\kappa^{ca}})^\kappa$ where C_κ is a positive parameter and \mathcal{S}_κ^{ca} is the normalized positive stable law of index κ , i.e. with Laplace transform

$$E[e^{-\lambda \mathcal{S}_\kappa^{ca}}] = e^{-\lambda^\kappa}, \quad \forall \lambda > 0.$$

In [14, 15] we gave a different proof of this result and we were able to give an explicit expression for the constant C_κ .

The proof was based on a precise analysis of the potential associated with the environment, as it was defined by Sinai for its analysis of the recurrent case, see [25].

In this paper, we use the techniques developed in [14, 15] to prove Theorem 1. The potential, denoted by $V = (V(x), x \in \mathbb{Z})$, is a function of the environment ω . It is defined as follows:

$$V(x) := \begin{cases} \sum_{i=1}^x \log \rho_i & \text{if } x \geq 1, \\ 0 & \text{if } x = 0, \\ -\sum_{i=x+1}^0 \log \rho_i & \text{if } x \leq -1. \end{cases}$$

Furthermore, we consider the weak descending ladder epochs for the potential defined by $e_0 := 0$ and

$$e_i := \inf\{k > e_{i-1} : V(k) \leq V(e_{i-1})\}, \quad i \geq 1,$$

which play a crucial role in our proof. Observe that the sequence $(e_i - e_{i-1})_{i \geq 1}$ is a family of i.i.d. random variables. Moreover, classical results of fluctuation theory (see [16], p. 396), tell us that, under assumptions (a)-(b) of Theorem 1,

$$(2.1) \quad E[e_1] < \infty.$$

Now, observe that the sequence $((e_i, e_{i+1}))_{i \geq 0}$ stands for the set of excursions of the potential above its past minimum. Let us introduce H_i , the height of the excursion $[e_i, e_{i+1}]$ defined by

$$(2.2) \quad H_i := \max_{e_i \leq k \leq e_{i+1}} (V(k) - V(e_i)), \quad i \geq 0.$$

Note that the $(H_i)_{i \geq 0}$'s are i.i.d. random variables.

For $t \in \mathbb{N}$, we introduce the critical height

$$(2.3) \quad h_t := \log t - \log \log t.$$

As in [15] we define the deep valleys from the excursions which are higher than the critical height h_t . Let $(\sigma(j))_{j \geq 1}$ be the successive indexes of excursions, whose heights are greater than h_t . More precisely,

$$\begin{aligned} \sigma(1) &:= \inf\{i \geq 0 : H_i \geq h_t\}, \\ \sigma(j) &:= \inf\{i > \sigma(j-1) : H_i \geq h_t\}, \quad j \geq 2. \end{aligned}$$

We consider now some random variables depending only on the environment, which define the deep valleys.

Definition 1. For all $j \geq 1$, let us introduce

$$\begin{aligned} b_j &:= e_{\sigma(j)}, \\ a_j &:= \sup\{k \leq b_j : V(k) - V(b_j) \geq D_t\}, \\ T_j^\uparrow &:= \inf\{k \geq b_j : V(k) - V(b_j) \geq h_t\}, \\ \bar{d}_j &:= e_{\sigma(j)+1}, \\ c_j &:= \inf\{k \geq b_j : V(k) = \max_{b_j \leq x \leq \bar{d}_j} V(x)\}, \\ d_j &:= \inf\{k \geq \bar{d}_j : V(k) - V(\bar{d}_j) \leq -D_t\}. \end{aligned}$$

where $D_t := (1 + \kappa) \log t$. We call (a_j, b_j, c_j, d_j) a deep valley and denote by $H^{(j)}$ the height of the j -th deep valley.

Moreover, let us introduce the first hitting time of x , denoted by

$$\tau(x) := \inf\{n \geq 1 : X_n = x\}, \quad x \in \mathbb{Z},$$

and the index of the last visited deep valley at time t , defined by

$$\ell_t := \sup\{n \geq 0 : \tau(b_n) \leq t\}.$$

Before stating the quenched localization result, recall that X is defined on the sample probability space $\mathbb{Z}^{\mathbb{N}}$. Then, let us introduce $\mathbf{e} = (\mathbf{e}_i, i \geq 1)$ a sequence of i.i.d. exponential random variables with parameter 1, independent of X . We define \mathbf{e} on a probability space Ξ and denote its law by $P^{(\mathbf{e})}$. In order to express the independence between X and \mathbf{e} , we consider for each environment ω , the probability space $(\mathbb{Z}^{\mathbb{N}} \times \Xi, P_\omega \times P^{(\mathbf{e})})$ on which we define (X, \mathbf{e}) .

Furthermore, let us define the *weight* of the k -th deep valley by

$$W_k(\omega) := 2 \sum_{\substack{a_k \leq m \leq n \\ b_k \leq n \leq d_k}} e^{V_\omega(n) - V_\omega(m)}.$$

Moreover, let us introduce the following integer, for any $t \geq 0$,

$$\ell_{t,\omega}^{(\mathbf{e})} := \sup \left\{ i \geq 0 : \sum_{k=1}^i W_k(\omega) \mathbf{e}_k \leq t \right\}.$$

We are now able to state our second main result.

Theorem 2. *Under assumptions (a)-(b) of Theorem 1, we have,*

(i) for all $\eta > 0$,

$$\lim_{t \rightarrow \infty} \mathbb{P}(|X_t - b_{\ell_t}| \leq \eta \log t) = 1,$$

(ii) for all $\delta > 0$,

$$\lim_{t \rightarrow \infty} P\left(d_{TV}(\ell_t, \ell_{t,\omega}^{(\mathbf{e})} + 1) > \delta\right) = 0,$$

where d_{TV} denotes the distance in total variation.

Remark 2. *The statement of Theorem 2 could be improved in the following way: the choice of the critical height h_t is in some way arbitrary and we could take for h_t any positive function such that $\lim_{t \rightarrow \infty} h_t = \infty$ and $e^{h_t} = o(t)$. The meaning is that at time t the RWRE is localized at the bottom of a deep valley, deep meaning that its height H is such that e^H is of order t . Furthermore, as in Theorem 1, the size of the localization window $\eta \log t$ could be replaced by any positive function $a(t)$ such that $\lim_{t \rightarrow \infty} a(t) = \infty$.*

We remark that we can easily deduce the following quenched localization in probability result by assembling part (i) and part (ii) of Theorem 2. We precise that our quenched localization result is in probability because one should not expect an almost sure result here, since no almost sure quenched limit results are expected to hold, see [24]. For $y < x$, we denote by E_ω^x the expectation associated with the law P_ω^x of the particle in the environment ω , started at x .

Corollary 1. *Under assumptions (a)-(b) of Theorem 1, we have, for all $\delta, \eta > 0$, that*

$$P\left(\sum_{i \geq 1} \left| P_{0, \omega}(|X_t - b_i| \leq \eta \log t) - P^{(e)}\left(\sum_{k=1}^{i-1} W_k(\omega) \mathbf{e}_k \leq t < \sum_{k=1}^i W_k(\omega) \mathbf{e}_k\right) \right| > \delta\right)$$

converges to 0, when t tends to ∞ .

The content of this result is twofold. It first says that, with a probability converging to 1, the process at time t is concentrated near the bottom of a valley of depth of order $\log t$. It also determines, for each of these valleys, the probability that, at time t , the particle lies at the bottom of it. This probability is driven by a renewal Poisson process which is skewed by the weights of each of these valleys.

This result may be of interest when trying to get information on the environment on the basis of the observation of a sample of trajectories of the particle. See [11] for a recent example of this in a paper on DNA reconstruction.

3. NOTATION

A result of Iglehart [21] which will be of constant use, says that, under assumptions (a)-(b) of Theorem 1, the tail of the height H_i of an excursion above its past minimum is given by

$$(3.1) \quad P(H_1 > h) \sim C_I e^{-\kappa h}, \quad h \rightarrow \infty,$$

for a positive constant C_I (we will not need its explicit value).

The analysis done in [14, 15] shows that on the interval $[0, t]$, $t \in \mathbb{N}$, the walk X_n spends asymptotically all its time trying to climb excursions of height of order $\log t + C$ for a real C . Let us now introduce the integer

$$n_t := \lfloor t^\kappa \log \log t \rfloor.$$

The integer n_t will be used to bound the number of excursions the walk can cross before time t . The strategy will be to show that we can neglect the time spent between two excursions of size smaller than h_t , and to show that at time t the walk X_t is close to the foot of an excursion of height larger than h_t .

3.1. The deep valleys. Let us define the number of deep valleys in the n_t first excursions by

$$K_t := \sup\{j \geq 0 : \sigma(j) \leq n_t\},$$

which is the number of excursions higher than the critical height h_t in the n_t first excursions.

Remark 3. *This definition corresponds to the definition of deep valleys introduced in [15] with $n = n_t$, but with a different critical height. In [15] the critical height was $h_n = \frac{1-\varepsilon}{\kappa} \log n$, for ε such that $0 < \varepsilon < 1$. Here, we see that h_{n_t} would be equal to $(1-\varepsilon) \log t + \frac{1-\varepsilon}{\kappa} \log \log \log t$ which is smaller than our critical height $h_t = \log t - \log \log t$. This means that the deep valleys are higher and less numerous in the present paper than in [15]. We will see that this choice makes possible the control of the localization of the particle in any neighborhood of size $\eta \log t$ around the bottom of the last visited valley (recall Part (i) of Theorem 2).*

3.2. The *-valleys. Let us first define the maximal variations of the potential before site x by:

$$\begin{aligned} V^\uparrow(x) &:= \max_{0 \leq i \leq j \leq x} (V(j) - V(i)), & x \in \mathbb{N}, \\ V^\downarrow(x) &:= \min_{0 \leq i \leq j \leq x} (V(j) - V(i)), & x \in \mathbb{N}. \end{aligned}$$

By extension, we introduce

$$\begin{aligned} V^\uparrow(x, y) &:= \max_{x \leq i \leq j \leq y} (V(j) - V(i)), & x < y, \\ V^\downarrow(x, y) &:= \min_{x \leq i \leq j \leq y} (V(j) - V(i)), & x < y. \end{aligned}$$

The deep valleys defined above are not necessarily made of disjoint portions of the environment. To overcome this difficulty we defined another type of valleys, called *-valleys, which form a subsequence of the previous valleys. By construction, the *-valleys are made of disjoint portions of environment and will coincide with high probability with the previous valleys on the portion of the environment visited by the walk before time t .

$$\begin{aligned} \gamma_1^* &:= \inf\{k \geq 0 : V(k) \leq -D_t\}, \\ T_1^* &:= \inf\{k \geq \gamma_1^* : V^\uparrow(\gamma_1^*, k) \geq h_t\}, \\ b_1^* &:= \sup\{k \leq T_1^* : V(k) = \min_{0 \leq x \leq T_1^*} V(x)\}, \\ a_1^* &:= \sup\{k \leq b_1^* : V(k) - V(b_1^*) \geq D_t\}, \\ \bar{d}_1^* &:= \inf\{k \geq T_1^* : V(k) \leq V(b_1^*)\}, \\ c_1^* &:= \inf\{k \geq b_1^* : V(k) = \max_{b_1^* \leq x \leq \bar{d}_1^*} V(x)\}, \\ d_1^* &:= \inf\{k \geq \bar{d}_1^* : V(k) - V(\bar{d}_1^*) \leq -D_t\}. \end{aligned}$$

Let us define the following sextuplets of points by iteration

$$(\gamma_j^*, a_j^*, b_j^*, T_j^*, c_j^*, \bar{d}_j^*, d_j^*) := (\gamma_1^*, a_1^*, b_1^*, T_1^*, c_1^*, \bar{d}_1^*, d_1^*) \circ \theta_{d_{j-1}^*}, \quad j \geq 2,$$

where θ_i denotes the i -shift operator.

Definition 2. We call a *-valley any quadruplet $(a_j^*, b_j^*, c_j^*, d_j^*)$ for $j \geq 1$. Moreover, we shall denote by K_t^* the number of such *-valleys before e_{n_t} , i.e. $K_t^* := \sup\{j \geq 0 : T_j^* \leq e_{n_t}\}$.

The *-valleys will be made of independent and identically distributed portions of potential (up to some translation).

4. PRELIMINARY ESTIMATES

4.1. Good environments. We define in this subsection the *good environments* in the same manner as we did in [15] to give a complete characterization of the limit law. Since the critical height considered here is not the same (see Remark 3), the following results are not taken from [15] but proved with the same ideas, that we recall in this subsection. Let us introduce the following series of events, which will

occur with high probability when t tends to infinity.

$$\begin{aligned} A_1(t) &:= \{e_{n_t} \leq C' n_t\}, \\ A_2(t) &:= \left\{K_t \leq (\log t)^{\frac{1+\kappa}{2}}\right\}, \\ A_3(t) &:= \bigcap_{j=0}^{K_t} \{\sigma(j+1) - \sigma(j) \geq t^{\kappa/2}\}, \\ A_4(t) &:= \bigcap_{j=1}^{K_t+1} \{d_j - a_j \leq C'' \log t\}, \end{aligned}$$

where $\sigma(0) := 0$ (for convenience of notation) and C', C'' stand for positive constants (large enough) which will be specified below. In words, $A_1(t)$ bounds the total length of the first n_t excursions. The event $A_2(t)$ gives a control on the number of deep valleys while $A_3(t)$ ensures that they are well separated and $A_4(t)$ bounds finely the length of each of them.

Lemma 1. *Let $A(t) := A_1(t) \cap A_2(t) \cap A_3(t) \cap A_4(t)$, then*

$$\lim_{t \rightarrow \infty} P(A(t)) = 1.$$

Proof. The fact that $P(A_1(t)) \rightarrow 1$ is a consequence of the law of large numbers. Concerning $A_2(t)$ and $A_3(t)$, we know that the number of excursions higher than h_t in the first n_t excursions is a binomial random variable with parameter (n_t, q_t) where $q_t := P(H_1 \geq h_t)$, from which we can easily deduce that $P(A_2(t) \cap A_3(t)) \rightarrow 1$. For example, since (3.1) implies $q_t \sim C_I e^{-\kappa h_t}$, $t \rightarrow \infty$, we have that $E[K_t] = n_t q_t \sim C_I \log \log t (\log t)^\kappa$. Using the Markov inequality we get that $P(A_2(t))$ tends to 1, when t tends to infinity.

The proof for $A_4(t)$ requires a bit more explanations. Since $K_t \leq (\log t)^{\frac{1+\kappa}{2}}$ with probability tending to one, we only have to prove, for $j \geq 1$ that $P(d_j - a_j \geq C'' \log t) = o((\log t)^{-\frac{1+\kappa}{2}})$. Furthermore, observe that we can write $d_j - a_j = (d_j - \bar{d}_j) + (\bar{d}_j - T_j^\uparrow) + (T_j^\uparrow - b_j) + (b_j - a_j)$. Therefore, the proof boils down to showing that, for each term in the previous sum, the probability that it is larger than $\frac{C''}{4} \log t$ is a $o((\log t)^{-\frac{1+\kappa}{2}})$. Here, we only prove that

$$(4.1) \quad P(T_j^\uparrow - b_j \geq \frac{C''}{4} \log t) = o((\log t)^{-\frac{1+\kappa}{2}}), \quad t \rightarrow \infty,$$

the arguments for the other terms being similar and the results more intuitive.

Let us first introduce $T_h := \inf\{x \geq 0 : V(x) \geq h\}$ for any $h > 0$. Then, recalling (3.1), we can write

$$(4.2) \quad P(T_j^\uparrow - b_j \geq \frac{C''}{4} \log t) \leq C e^{\kappa h_t} P(\frac{C''}{4} \log t \leq T_{h_t} < \infty).$$

Denoting by $I(\cdot)$ the convex rate function associated with the potential, we apply Chebychev's inequality in the same manner as is done in the proof of the upper bound in Cramer's theorem (see [20]) and obtain that the probability on the right-hand side in (4.2) is bounded above by

$$(4.3) \quad \sum_{k \geq \frac{C''}{4} \log t} P(V(k) \geq h_t) \leq \sum_{k \geq \frac{C''}{4} \log t} e^{-k I(\frac{h_t}{k})} \leq \sum_{k \geq \frac{C''}{4} \log t} e^{-k I(0)} \leq C t^{-\frac{C''}{4} I(0)}.$$

Now, let us recall that $h_t \leq \log t$ by definition. Moreover, observe that the assumption (a) implies that $E[\rho_0^\kappa] = 1$, which yields $I(0) > 0$. Then, assembling (4.2) and (4.3)

yields (4.1) by choosing C'' larger than $4\kappa/I(0)$, which concludes the proof of Lemma 1. \square

The following lemma tells us that the $*$ -valleys, which are i.i.d., coincide with the sequence of deep valleys with an overwhelming probability when t goes to infinity.

Lemma 2. *If $A^*(t) := \{K_t = K_t^*; (a_j, b_j, c_j, d_j) = (a_j^*, b_j^*, c_j^*, d_j^*), 1 \leq j \leq K_t^*\}$, then we have that the probability $P(A^*(t))$ converges to 1, when t goes to infinity.*

Proof. By definition, the $*$ -valleys constitute a subsequence of the deep valleys, and $A^*(t)$ occurs as soon as the valleys (a_j, b_j, c_j, d_j) are disjoint for $1 \leq j \leq K_t$. Hence, we see that $A_3(t) \cap A_4(t) \subset A^*(t)$. Then, Lemma 2 is a consequence of Lemma 1. \square

4.2. Directed traps. Let us first recall that it is well-known (see for example [27], formula (2.1.4)) that for $r < x < s$,

$$(4.4) \quad P_\omega^x(\tau(r) < \tau(s)) = \sum_{j=x}^{s-1} e^{V(j)} \left(\sum_{j=r}^{s-1} e^{V(j)} \right)^{-1}.$$

Moreover, we introduce here the inter-arrival times, defined, for any $x, y \in \mathbb{Z}$, by

$$\tau(x, y) := \inf\{k \geq 0 : X_{\tau(x)+k} = y\}.$$

With the two following lemmas, we prove that the particle never backtracks before a_j after reaching the bottom b_j of the j -th valley, uniformly in $1 \leq j \leq K_t$, and that it visits each of them only once.

Lemma 3. *Defining $DT(t) := A(t) \cap \bigcap_{j=1}^{K_t} \{\tau(d_j, b_{j+1}) < \tau(d_j, \bar{d}_j)\}$, we have*

$$\lim_{t \rightarrow \infty} \mathbb{P}(DT(t)) = 1.$$

Proof. Recalling that $K_t \leq (\log t)^{\frac{1+\kappa}{2}}$ with probability tending to one, we have to prove, for $j \geq 1$, that $E[\mathbf{1}_{A(t) \cap \{j \leq K_t\}} P_\omega(\tau(d_j, b_{j+1}) > \tau(d_j, \bar{d}_j))] = o((\log t)^{-\frac{1+\kappa}{2}})$, when t tends to infinity. Therefore, applying the strong Markov property at $\tau(d_j)$, we need to prove that

$$(4.5) \quad E[\mathbf{1}_{A(t) \cap \{j \leq K_t\}} P_\omega^{d_j}(\tau(b_{j+1}) > \tau(\bar{d}_j))] = o((\log t)^{-\frac{1+\kappa}{2}}), \quad t \rightarrow \infty.$$

By (4.4) we get that $P_\omega^{d_j}(\tau(b_{j+1}) > \tau(\bar{d}_j))$ is bounded by $(b_{j+1} - d_j)e^{V(d_j) - V(\bar{d}_j) + h_t}$. Observe first that $b_{j+1} - d_j \leq e_{n_t} \leq C'n_t$ on $A(t) \cap \{j \leq K_t\}$. Then, recalling that $V(d_j) - V(\bar{d}_j) \leq -D_t$ by definition (where $D_t = (1 + \kappa) \log t$) together with $h_t \leq \log t$ yields (4.5) and concludes the proof of Lemma 3. \square

Lemma 4. *Defining $DT^*(t) := \bigcap_{j=1}^{K_t^*} \{\tau(b_j^*, d_j^*) < \tau(b_j^*, a_j^*)\}$, we have*

$$\lim_{t \rightarrow \infty} \mathbb{P}(DT^*(t)) = 1.$$

Proof. We omit the details here since the arguments are very similar to the proof of Lemma 3. \square

Finally, we need to know that the time spent between the deep valleys is small. Let us first recall the following technical result proved in Lemma 7 of [15].

Lemma 5. *Let T^\uparrow be defined by $T^\uparrow(h) := \inf\{x \geq 0 : V^\uparrow(x) \geq h\}$, for any $h \geq 0$. Then, there exists $C > 0$ such that, for all h ,*

$$\mathbb{E}_{|_0} [\tau(T^\uparrow(h) - 1)] \leq Ce^{ht},$$

where $\mathbb{E}_{|_0}$ denotes the expectation under the annealed law $\mathbb{P}_{|_0}$ associated with the random walk in random environment reflected at 0.

Now, we can prove that the time spent by the particle between the K_t first deep valleys is negligible with respect to t with an overwhelming probability when t goes to infinity, which is the statement of the following lemma.

Lemma 6. *Let us introduce the following event*

$$IA(t) := A(t) \cap \left\{ \tau(b_1) + \sum_{j=1}^{K_t} \tau(d_j, b_{j+1}) < \frac{t}{\log \log t} \right\}.$$

Then, we have

$$\lim_{t \rightarrow \infty} \mathbb{P}(IA(t)) = 1.$$

Proof. Recalling Lemma 1, Lemma 3 and using the Markov inequality, we only need to prove that $\mathbb{E}[\mathbf{1}_{A(t) \cap DT(t)}(\tau(b_1) + \sum_{j=1}^{K_t} \tau(d_j, b_{j+1}))]$ is $o(\frac{t}{\log \log t})$, when t goes to infinity. For $y < x$, let us denote by $E_{\omega,|y}^x$ the expectation associated with the law $P_{\omega,|y}^x$ of the particle in the environment ω , started at x and reflected at site y . Then, applying the strong Markov property at times $\tau(d_{K_t}), \dots, \tau(d_1)$, we get that the above expectation is smaller than

$$(4.6) \quad E[\mathbf{1}_{A(t) \cap DT(t)} \tau(b_1)] + E\left[\mathbf{1}_{A(t) \cap DT(t)} \sum_{j=1}^{K_t} E_{\omega,|\bar{d}_j}^{d_j}[\tau(b_{j+1})]\right],$$

since $(X_{\tau(d_j)+n})_{n \geq 0}$ under P_ω has the same law as $(X_n)_{n \geq 0}$ under $P_{\omega,|\bar{d}_j}^{d_j}$ on $A(t) \cap DT(t)$. Concerning the second term of (4.6), we apply the strong Markov property for the potential at times $\bar{d}_{K_t}, \dots, \bar{d}_1$, such that we get

$$\begin{aligned} E\left[\mathbf{1}_{A(t) \cap DT(t)} \sum_{j=1}^{K_t} E_{\omega,|\bar{d}_j}^{d_j}[\tau(b_{j+1})]\right] &\leq (\log t)^{\frac{1+\kappa}{2}} \mathbb{E}_{|_0} [\tau(T^\uparrow(h_t) - 1)] \\ &\leq C(\log t)^{\frac{1+\kappa}{2}} e^{ht} \leq Ct(\log t)^{-\frac{1-\kappa}{2}}, \end{aligned}$$

the second inequality being a consequence of Lemma 5. Now, let us mention that the bound Ce^{ht} can be obtained in a similar way for the first term of (4.6), which yields that the expression in (4.6) is a $o(\frac{t}{\log \log t})$, when t tends to infinity and concludes the proof of Lemma 6. \square

4.3. Localization in deep traps. In a first step, we state a technical result which ensures that the potential does not have excessive fluctuations in a typical box and which will be very useful to control the localization of the particle in a valley.

Lemma 7. *If $F_\gamma(t) := \cap_{j=1}^{K_t} \{\max\{V^\uparrow(a_j, b_j); -V^\downarrow(b_j, c_j); V^\uparrow(c_j, d_j)\} \leq \gamma \log t\}$, then we have, for any $\gamma > 0$,*

$$\lim_{t \rightarrow \infty} \mathbb{P}(F_\gamma(t)) = 1.$$

Proof. Observe first that Lemma 14 in [15] implies that, for all $\varepsilon > 0$, the valleys with height larger than $(1 - \varepsilon) \log t + \frac{1-\varepsilon}{\kappa} \log \log \log t$ have fluctuations bounded by $\gamma \log t$, with a probability tending to one, for any $\gamma > \varepsilon/\kappa$. Now, since h_t is larger than $(1 - \varepsilon) \log t + \frac{1-\varepsilon}{\kappa} \log \log \log t$ for any $\varepsilon > 0$ (see Remark 3), the deep valleys considered here are included in the valleys treated by Lemma 14 in [15] for any $\varepsilon > 0$, which concludes the proof of Lemma 7. \square

For each deep valley, let us introduce the position \bar{c}_i defined by

$$\bar{c}_i := \inf\{n \geq c_i : V(n) \leq V(c_i) - h_t/3\}.$$

We first need to know that during its sojourn time inside a deep valley, the random walk spends almost all its time inside the interval (a_i, c_i) . This is a consequence of the following lemma.

Lemma 8. *Let $LT(t)$ be the event*

$$LT(t) := \bigcap_{i=1}^{K_t} \left\{ \tau(\bar{c}_i, d_i) \leq \frac{t}{\log t} \right\}.$$

Then,

$$\lim_{t \rightarrow \infty} \mathbb{P}(LT(t)) = 1.$$

This result just means that at the time scale t , if the walk reaches \bar{c}_i , then soon after it exits the deep valley (a_i, d_i) .

Proof. Recalling Lemma 1 and Lemma 7, we only have to prove that

$$\mathbb{P} \left(\tau(\bar{c}_j, d_j) > \frac{t}{\log t}; A_4(t); F_\gamma(t); j \leq K_t \right) = o((\log t)^{-\frac{1+\kappa}{2}}), \quad t \rightarrow \infty,$$

for any $j \geq 1$. Now, applying the strong Markov property at $\tau(\bar{c}_j)$, we get that the previous probability is bounded by

$$E \left[\mathbf{1}_{A_4(t) \cap F_\gamma(t) \cap \{j \leq K_t\}} \left(P_{\omega, |c_j}^{\bar{c}_j} (\tau(d_j) > t/\log t) + P_{\omega}^{\bar{c}_j} (\tau(c_j) < \tau(d_j)) \right) \right].$$

Concerning the first term, we use the fact that $E_{\omega, |c_j}^{\bar{c}_j} [\tau(d_j)] \leq \sum_{c_j \leq u \leq v \leq d_j} e^{V(v)-V(u)}$ (see (A1) in [18]) and Chebychev inequality, such that we obtain

$$(4.7) \quad P_{\omega, |c_j}^{\bar{c}_j} (\tau(d_j) > t/\log t) \leq \frac{\log t}{t} \sum_{c_j \leq u \leq v \leq d_j} e^{V(v)-V(u)} \leq C'' \frac{(\log t)^2}{t} e^{\gamma \log t},$$

on $A_4(t) \cap F_\gamma(t) \cap \{j \leq K_t\}$. For the second term, by (4.4) we obtain that the probability $P_{\omega}^{\bar{c}_j} (\tau(c_j) < \tau(d_j))$ is less than

$$(4.8) \quad \sum_{k=\bar{c}_j}^{d_j-1} e^{V(k)} \left(\sum_{k=c_j}^{d_j-1} e^{V(k)} \right)^{-1} \leq (d_j - c_j) e^{V(\bar{c}_j) + \gamma \log t - V(c_j)} \leq C'' (\log t) e^{\gamma \log t - \frac{h_t}{3}},$$

on $A_4(t) \cap F_\gamma(t) \cap \{j \leq K_t\}$. Then, assembling (4.7) and (4.8) yields

$$\mathbb{P} \left(\tau(\bar{c}_j, d_j) > \frac{t}{\log t}; A_4(t); F_\gamma(t) \right) \leq C (\log t) e^{\gamma \log t - \frac{h_t}{3}},$$

which concludes the proof of Lemma 8 by choosing $\gamma < 1/3$. \square

Now, we need to be sure that the bottom of the deep valleys are sharp. For $\eta > 0$, we introduce the following subsets of the deep valleys

$$O_i := [a_i + 1, \bar{c}_i - 1] \setminus (b_i - \eta \log t + 1, b_i + \eta \log t - 1), \quad i \in \mathbb{N},$$

and the event

$$A_5(t, \eta) := \bigcap_{i=1}^{K_t} \left\{ \min_{k \in O_i \cap \mathbb{Z}} (V(k) - V(b_i)) \geq C''' \eta \log t \right\},$$

for a constant C''' (small enough and independent of η) to be defined later. Then, we have the following result.

Lemma 9. *For all $\eta > 0$,*

$$\lim_{t \rightarrow \infty} P(A_5(t, \eta)) = 1.$$

Proof. Observe first that if $\eta > C''$, then the sets $(O_i, 1 \leq i \leq K_t)$ are empty on $A_4(t)$. Therefore, Lemma 9 is a consequence of Lemma 1.

Now, let us assume $\eta \leq C''$. The definition of \bar{c}_i implies that $\min_{c_i \leq k < \bar{c}_i} (V(k) - V(b_i)) \geq \frac{2}{3} h_t$. Then, choosing C''' such that $C''' C'' < 2/3$ implies that $C''' \eta \log t < \frac{2}{3} h_t$ for all large t , which yields

$$(4.9) \quad P \left(\bigcap_{i=1}^{K_t} \left\{ \min_{c_i \leq k < \bar{c}_i} (V(k) - V(b_i)) \geq C''' \eta \log t \right\} \right) = 1,$$

for all large t . Then, let us introduce the sets

$$O'_i := O_i \cap [b_i, c_i], \quad O''_i := O_i \cap [a_i, b_i], \quad i \in \mathbb{Z},$$

and the events

$$A'_5(t, \eta) := \bigcap_{i=1}^{K_t} \left\{ \min_{k \in O'_i \cap \mathbb{Z}} (V(k) - V(b_i)) \geq C''' \eta \log t \right\},$$

$$A''_5(t, \eta) := \bigcap_{i=1}^{K_t} \left\{ \min_{k \in O''_i \cap \mathbb{Z}} (V(k) - V(b_i)) \geq C''' \eta \log t \right\}.$$

Now, recalling (4.9), the proof of Lemma 9 boils down to showing that

$$(4.10) \quad \lim_{t \rightarrow \infty} P(A'_5(t, \eta)) = 1,$$

$$(4.11) \quad \lim_{t \rightarrow \infty} P(A''_5(t, \eta)) = 1.$$

Let us first prove (4.10). Recalling Lemma 1 and Lemma 7, we only need to prove that it is possible to choose C''' small enough such that for some $\gamma > 0$

$$(4.12) \quad P \left(\min_{k \in O'_1 \cap \mathbb{Z}} (V(k) - V(b_1)) < C''' \eta \log t; F_\gamma(t) \right) = o((\log t)^{-\frac{1+\kappa}{2}}),$$

when $t \rightarrow \infty$. Now recalling assumption (a) of Theorem 1 and denoting by μ the law of $\log \rho_0$, we can define the law $\tilde{\mu} = \rho_0^\kappa \mu$, and the law $\tilde{P} = \tilde{\mu}^{\otimes \mathbb{Z}}$ which is the law of a sequence of i.i.d. random variables with law $\tilde{\mu}$. The definition of κ implies that $\int \log \rho \tilde{\mu}(d\rho) > 0$. Now, let us simplify the notation by writing

$$H := H_0$$

(where H_0 is the height of the first excursion defined by $H_0 := \max_{0 \leq k \leq e_1} V(k)$) and define the hitting time of level h for the potential by

$$T_h := \min\{x \geq 0 : V(x) \geq h\}, \quad h > 0.$$

Then, introducing $\tilde{F}_\gamma(t) := \{-V^\downarrow(0, T_H) \leq \gamma \log t\}$, we can write that the probability term in (4.12) is smaller than

$$\begin{aligned} & P\left(\min_{\lfloor \eta \log t \rfloor \leq k \leq T_H} V(k) < C''' \eta \log t; \tilde{F}_\gamma(t) \mid H \geq h_t\right) \\ & \leq C e^{\kappa h_t} P\left(\min_{\lfloor \eta \log t \rfloor \leq k \leq T_H} V(k) < C''' \eta \log t; \tilde{F}_\gamma(t); H \geq h_t\right) \\ & = C \tilde{E}\left[e^{-\kappa(V(T_H) - h_t)} \mathbf{1}_{\{\min_{\lfloor \eta \log t \rfloor \leq k \leq T_H} V(k) < C''' \eta \log t; \tilde{F}_\gamma(t); H \geq h_t\}}\right] \\ (4.13) \quad & \leq C \tilde{P}\left(\min_{\lfloor \eta \log t \rfloor \leq k \leq T_H} V(k) < C''' \eta \log t; \tilde{F}_\gamma(t); H \geq h_t\right), \end{aligned}$$

the first inequality being a consequence of (3.1) and the equality deduced from Girsanov property. Now, let us introduce $\alpha = \alpha(\eta) := c\eta$ with c satisfying $0 < c < \min\{\tilde{E}[V(1)]; 1/C'''\}$ and $\gamma = \gamma(\eta) := c\eta/2$. Observe that $\alpha \log t < h_t$ for all large t , so that $T_{\alpha \log t} \leq T_{h_t} \leq T_H < \infty$ on $\{H \geq h_t\}$. Now since $c < \tilde{E}[V(1)]$, we use Chebychev's inequality in the same manner as is done in the proof of the upper bound in Cramer's theorem, see [20], and obtain that $\tilde{P}(V(\lfloor \eta \log t \rfloor) < \alpha \log t) \leq C \exp\{-\eta \tilde{I}(c) \log t\} = o((\log t)^{-\frac{1+\kappa}{2}})$, where $\tilde{I}(\cdot)$ denotes the convex rate function associated with V under \tilde{P} . This yields $\tilde{P}(T_{\alpha \log t} \leq \lfloor \eta \log t \rfloor) = 1 - o((\log t)^{-\frac{1+\kappa}{2}})$, when t tends to infinity. Therefore, we get

$$\begin{aligned} & \tilde{P}\left(\min_{\lfloor \eta \log t \rfloor \leq k \leq T_H} V(k) < C''' \eta \log t; \tilde{F}_\gamma(t); H \geq h_t\right) \\ (4.14) \quad & \leq \tilde{P}\left(\min_{T_{\alpha \log t} \leq k \leq T_H} V(k) < C''' \eta \log t; \tilde{F}_\gamma(t); H \geq h_t\right) + o((\log t)^{-\frac{1+\kappa}{2}}). \end{aligned}$$

Furthermore, observe that on $\tilde{F}_\gamma(t)$, we have $\min_{T_{\alpha \log t} \leq k \leq T_H} V(k) \geq (\alpha - \gamma) \log t$, which yields $\min_{T_{\alpha \log t} \leq k \leq T_H} V(k) \geq C''' \eta \log t$, if we choose C''' smaller than $c/2$. Therefore, for C''' small enough (independently of $\eta \leq C'''$), we get that the probability term in (4.14) is null for all large t . Now, assembling (4.13) and (4.14) concludes the proof of (4.10).

The proof of (4.11) is similar but easier. Indeed, we do not have to use Girsanov property to study the potential on $[a_i, b_i]$. \square

5. TWO VERSIONS OF A DYNKIN TYPE RENEWAL RESULT

We define the sequence of random times $(\tau_i^*)_{i \geq 1}$ as follows: conditioning on the environment ω , $(\tau_i^*)_{i \geq 1}$ is defined as an independent sequence of random variables with the law of $\tau(d_i^*)$ under $P_{\omega, |a_i^*}^{b_i^*}$, where $\tau(d_i^*)$ denotes the first hitting time of d_i^* and $P_{\omega, |a_i^*}^{b_i^*}$ is the law of the Markov chain in environment ω , starting from b_i^* and reflected at a_i^* . Hence, under the annealed law \mathbb{P} , $(\tau_i^*)_{i \geq 1}$ is an i.i.d. sequence since the $*$ -valleys are independent and identically distributed. The first step in our proof is to derive the following result.

Proposition 1. Let ℓ_t^* be the random integer defined by

$$\ell_t^* := \sup\{n \geq 0 : \tau_1^* + \cdots + \tau_n^* \leq t\}.$$

For all $0 \leq x_1 < x_2 \leq 1$, we have

$$\lim_{t \rightarrow \infty} \mathbb{P}(t(1 - x_2) \leq \tau_1^* + \cdots + \tau_{\ell_t^*}^* \leq t(1 - x_1)) = \frac{\sin(\kappa\pi)}{\pi} \int_{x_1}^{x_2} (1 - x)^{\kappa-1} x^{-\kappa} dx.$$

For all $0 \leq x_1 < x_2$, we have

$$\lim_{t \rightarrow \infty} \mathbb{P}(t(1 + x_1) \leq \tau_1^* + \cdots + \tau_{\ell_t^*+1}^* \leq t(1 + x_2)) = \frac{\sin(\kappa\pi)}{\pi} \int_{x_1}^{x_2} \frac{dx}{x^\kappa(1+x)}.$$

Observe that the result would exactly be Dynkin's theorem (cf e.g. Feller, vol II, [16], p. 472) if the sequence $(\tau_i^*)_{i \geq 1}$ was an independent sequence of random variables in the domain of attraction of a stable law with index κ . Here, the sequence $(\tau_i^*)_{i \geq 1}$ implicitly depends on the time t , since the $*$ -valleys are defined from the critical height h_t . We will use the main intermediate result of [15] which gives an estimate of the Laplace transform of τ_1^* at 0. We deduce from Corollary 2 and Remark 7 of [15] the following lemma.

Lemma 10. We have

$$\mathbb{E} \left[1 - e^{-\lambda \frac{\tau_1^*}{t}} \right] \sim 2^\kappa \frac{\pi \kappa}{\sin(\pi \kappa)} \frac{C_U}{t^\kappa P(H \geq h_t)} \lambda^\kappa, \quad t \rightarrow \infty,$$

for all $\lambda > 0$.

Proof. We apply Corollary 2 of [15] to $n = \lfloor t^\kappa \rfloor$ and $h_n = h_t = \log t - \log \log t$ which satisfies the condition of Remark 7 of [15]. The constant C_U was made explicit in [14] but we will not need this value here.

For the convenience of the reader, we give a brief idea of the arguments of the proof of this formula. Let us simply write (a, b, c, d) for $(a_1^*, b_1^*, c_1^*, d_1^*)$. The time it takes to cross the valley can be decomposed in a geometric number of unsuccessful attempts and a successful attempt, hence we can write

$$\tau_1^* = \tau(b, d) = F_1 + \cdots + F_N + S,$$

where N is a geometric random variable with parameter

$$1 - p(\omega) := P_\omega^b(\tau(d) < \tau^+(b)) = \omega_b \frac{e^{V(b)}}{\sum_{x=b}^{d-1} e^{V(x)}},$$

where $\tau^+(b) := \inf\{n > 0 : X_n = b\}$. The random variables $(F_i)_{i \geq 1}$ are i.i.d. and distributed as $\tau^+(b)$ under $P_\omega^b(\cdot | \tau^+(b) < \tau(d))$ and S is distributed as $\tau(d)$ under $P_\omega^b(\cdot | \tau(d) < \tau^+(b))$. The first step is to prove that the successful attempt S can be neglected (this is done in [15] using some estimates on h-processes). Thus, we can write

$$\mathbb{E} \left[e^{-\lambda \frac{\tau_1^*}{t}} \right] \sim E \left[\frac{1 - p(\omega)}{1 - p(\omega) E_\omega \left[e^{-\frac{\lambda}{t} F_1} \right]} \right], \quad t \rightarrow \infty.$$

The second step is to linearize $E_\omega \left[e^{-\frac{\lambda}{t} F_1} \right]$, i.e. to show that it can be replaced by $(1 - \frac{\lambda}{t} E_\omega[F_1])$ (using again estimates on h-processes). This leads to

$$\mathbb{E} \left[e^{-\lambda \frac{\tau_1^*}{t}} \right] \sim E \left[\frac{1}{1 + \frac{\lambda}{t} \frac{p(\omega)}{1-p(\omega)} E_\omega[F_1]} \right], \quad t \rightarrow \infty.$$

Then we prove that $\frac{p(\omega)}{1-p(\omega)}E_\omega[F_1]$ is of order $Z = 2e^H M_1 M_2$, where M_1 and M_2 are defined by $M_1 := \sum_{k=a}^c e^{-(V(k)-V(b))}$ and $M_2 := \sum_{k=b}^d e^{V(k)-V(c)}$. Then, we use the main result of [14], where the tail estimate of Z is obtained (see Theorem 2.2). \square

Proof of Proposition 1. The arguments are essentially the same as in [16]. Let us introduce $S_0^* = 0$ and $S_n^* := \sum_{i=1}^n \tau_i^*$, for $n \geq 1$. Then, the inequality $t(1-x_2) \leq \tau_1^* + \dots + \tau_{\ell_t^*}^* \leq t(1-x_1)$ occurs iff $S_n^* = ty$ and $\tau_{n+1}^* > t(1-y)$ for some combination n, y such that $1-x_2 < y < 1-x_1$. Summing over all n and possible y we get

$$(5.1) \quad \mathbb{P}(t(1-x_2) \leq S_{\ell_t^*}^* \leq t(1-x_1)) = \int_{1-x_2}^{1-x_1} \frac{G_t(1-y)}{P(H \geq h_t)} U_t\{dy\},$$

where $G_t(x) := P(H \geq h_t) \mathbb{P}(t^{-1}\tau_1^* \geq x)$, and $U_t\{dx\}$ denotes the measure associated with $U_t(x) := \sum_{n \geq 0} \mathbb{P}(t^{-1}S_n^* \leq x)$. We introduce the measure $dH_t(u)$ such that $\int_x^\infty dH_t(u) = G_t(x)$, for all $x \geq 0$.

Lemma 11. *For any $x > 0$, we have*

$$(5.2) \quad \lim_{t \rightarrow \infty} x^\kappa t^\kappa G_t(x) = 2^\kappa \Gamma(1 + \kappa) C_U.$$

Moreover, the convergence is uniform on any compact set.

Proof. In a first step, observe that $\mathbb{E}[1 - e^{-\lambda \frac{\tau_1^*}{t}}] = P(H \geq h_t)^{-1} \int_0^\infty (1 - e^{-\lambda u}) dH_t(u)$. Recalling Lemma 10, we obtain

$$\lim_{t \rightarrow \infty} t^\kappa \int_0^\infty (1 - e^{-\lambda u}) dH_t(u) = 2^\kappa \Gamma(1 + \kappa) C_U \Gamma(1 - \kappa) \lambda^\kappa.$$

Since $\Gamma(1 - \kappa) \lambda^\kappa = \lambda \int_0^\infty e^{-\lambda u} u^{-\kappa} du$, this implies

$$(5.3) \quad \lim_{t \rightarrow \infty} t^\kappa \int_0^\infty (1 - e^{-\lambda u}) dH_t(u) = 2^\kappa \Gamma(1 + \kappa) C_U \lambda \int_0^\infty e^{-\lambda u} u^{-\kappa} du.$$

On the other hand, integrating by parts, we get, for any $t \geq 0$,

$$(5.4) \quad \int_0^\infty (1 - e^{-\lambda u}) dH_t(u) = \lambda \int_0^\infty e^{-\lambda u} G_t(u) du.$$

Combining (5.3) and (5.4) implies that the measure $t^\kappa G_t(u) du$ tends to the measure with density $2^\kappa \Gamma(1 + \kappa) C_U u^{-\kappa}$. Therefore, we have for all $x \geq 0$,

$$(5.5) \quad \lim_{t \rightarrow \infty} t^\kappa \int_0^x G_t(u) du = 2^\kappa \Gamma(1 + \kappa) C_U \frac{x^{1-\kappa}}{1-\kappa},$$

which yields

$$(5.6) \quad \lim_{\varepsilon \rightarrow 0} \lim_{t \rightarrow \infty} \frac{\int_x^{(1+\varepsilon)x} G_t(u) du}{\varepsilon \int_0^x G_t(u) du} = 1 - \kappa.$$

Moreover, observe that the monotonicity of $G_t(\cdot)$ implies

$$(5.7) \quad \frac{x G_t((1+\varepsilon)x)}{\int_0^x G_t(u) du} \leq \frac{\int_x^{(1+\varepsilon)x} G_t(u) du}{\varepsilon \int_0^x G_t(u) du} \leq \frac{x G_t(x)}{\int_0^x G_t(u) du}.$$

Now, combining (5.6) and (5.7), we obtain

$$\liminf_{t \rightarrow \infty} \frac{x G_t(x)}{\int_0^x G_t(u) du} \geq 1 - \kappa.$$

Recalling (5.5), this yields

$$(5.8) \quad \liminf_{t \rightarrow \infty} x^\kappa t^\kappa G_t(x) \geq 2^\kappa \Gamma(1 + \kappa) C_U.$$

Similarly, we obtain, for any $\varepsilon > 0$,

$$(5.9) \quad \limsup_{t \rightarrow \infty} x^\kappa t^\kappa G_t((1 + \varepsilon)x) \leq 2^\kappa \Gamma(1 + \kappa) C_U.$$

Assembling (5.8) and (5.9) and letting $\varepsilon \rightarrow 0$ conclude the proof of (5.2).

Furthermore, observe that the uniform convergence on any compact set is a consequence of the monotonicity of $x \mapsto G_t(x)$, the continuity of the limit and Dini's theorem. \square

Lemma 12. *The measure $\frac{P(H \geq h_t)^{-1}}{t^\kappa} U_t\{dx\}$ converges vaguely to the measure*

$$\frac{1}{\Gamma(\kappa)\Gamma(1 + \kappa)\Gamma(1 - \kappa)2^\kappa C_U} x^{\kappa-1} dx.$$

Proof. Observe first that the Laplace transform $\widehat{U}_t(\lambda) := \int_0^\infty e^{-\lambda u} U_t\{du\}$ satisfies $\widehat{U}_t(\lambda) = \sum_{n \geq 0} \mathbb{E}[e^{-\lambda \frac{S_n^*}{t}}] = (1 - \mathbb{E}[e^{-\lambda \frac{\tau_1^*}{t}}])^{-1}$. Therefore, Lemma 10 yields

$$\lim_{t \rightarrow \infty} \frac{P(H \geq h_t)^{-1}}{t^\kappa} \widehat{U}_t(\lambda) = \frac{\lambda^{-\kappa}}{\Gamma(1 + \kappa)\Gamma(1 - \kappa)2^\kappa C_U}.$$

Furthermore, since $\Gamma(\kappa)\lambda^{-\kappa} = \int_0^\infty e^{-\lambda u} u^{\kappa-1} du$, we deduce the vague convergence of the measure from the pointwise convergence of the Laplace transforms. \square

Now, recalling (5.1), we observe that Lemma 11 together with Lemma 12 imply

$$\begin{aligned} \lim_{t \rightarrow \infty} \mathbb{P}(t(1 - x_2) \leq S_{\ell_t^*}^* \leq t(1 - x_1)) &= \frac{1}{\Gamma(\kappa)\Gamma(1 - \kappa)} \int_{1-x_2}^{1-x_1} (1 - y)^{-\kappa} y^{\kappa-1} dy, \\ &= \frac{\sin(\kappa\pi)}{\pi} \int_{x_1}^{x_2} (1 - y)^{\kappa-1} y^{-\kappa} dy. \end{aligned}$$

This concludes the proof of the first part of Proposition 1. The second part of Proposition 1 is obtained using similar arguments.

Recall Lemma 6 which tells that the inter-arrival times are negligible. Now, we will prove that the results of Proposition 1 are still true if we consider, in addition, these inter-arrival times. Let $\delta_1 := \tau(b_1)$, $\tau_1 := \tau(b_1, d_1)$ and

$$\delta_k := \tau(d_{k-1}, b_k), \quad \tau_k := \tau(b_k, d_k), \quad k \geq 2.$$

Moreover, we set

$$T_k := \delta_1 + \tau_1 + \cdots + \tau_{k-1} + \delta_k, \quad k \geq 1,$$

the entering time in the k -th deep valley.

Proposition 2. *Recall $\ell_t = \sup\{n \geq 0 : \tau(b_n) \leq t\}$. Then, we have*

$$\mathbb{P}(T_{\ell_t} \leq t < T_{\ell_t} + \tau_{\ell_t}) \rightarrow 1, \quad t \rightarrow \infty.$$

For all $0 \leq x_1 < x_2 \leq 1$, we have

$$\lim_{t \rightarrow \infty} \mathbb{P}(t(1 - x_2) \leq T_{\ell_t} \leq t(1 - x_1)) = \frac{\sin(\kappa\pi)}{\pi} \int_{x_1}^{x_2} (1 - x)^{\kappa-1} x^{-\kappa} dx.$$

For all $0 \leq x_1 < x_2$, we have

$$\lim_{t \rightarrow \infty} \mathbb{P}(t(1+x_1) \leq T_{\ell_t+1} \leq t(1+x_2)) = \frac{\sin(\kappa\pi)}{\pi} \int_{x_1}^{x_2} \frac{dx}{x^\kappa(1+x)}.$$

Proof. On the event $A(t) \cap DT^*(t)$, we know that the random times $(\tau_i)_{1 \leq i \leq K_t^*}$ have the same law as the random times $(\tau_i^*)_{1 \leq i \leq K_t^*}$ defined in Section 5. If we define $\tilde{\ell}_t := \sup\{n \geq 0 : \tau_1 + \dots + \tau_n \leq t\}$, then, using Proposition 1 and Lemma 3, we get that the result of Proposition 1 is true with τ and $\tilde{\ell}_t$ in place of τ^* and ℓ_t^* . Now, using Lemma 6 we see that

$$\begin{aligned} & \liminf_{t \rightarrow \infty} \mathbb{P}(\tilde{\ell}_t = \ell_t - 1; T_{\ell_t} \leq t < T_{\ell_t} + \tau_{\ell_t}) \\ & \geq \liminf_{t \rightarrow \infty} \mathbb{P}(IA(t); |t - (\tau_1 + \dots + \tau_{\tilde{\ell}_t})| \geq \xi t), \end{aligned}$$

for all $\xi > 0$. Thus, using Proposition 1 (for $\tilde{\ell}_t$ and τ_i) and letting ξ tends to 0, we get that

$$\lim_{t \rightarrow \infty} \mathbb{P}(\tilde{\ell}_t = \ell_t - 1; T_{\ell_t} \leq t < T_{\ell_t} + \tau_{\ell_t}) = 1.$$

We conclude the proof by the same type of arguments. \square

6. PROOF OF PART (i) OF THEOREM 2: A LOCALIZATION RESULT

We follow the strategy developed by Sinai for the recurrent case. For each valley we denote by π_i the invariant measure of the random walk on $[a_i, \bar{c}_i]$ in environment ω , reflected at a_i and \bar{c}_i and normalized so that $\pi_i(b_i) = 1$. Clearly, π_i is the reversible measure given, for $k \in [b_i + 1, \bar{c}_i - 1]$, by

$$\begin{aligned} \pi_i(k) &= \frac{\omega_{b_i}}{1 - \omega_{b_i+1}} \dots \frac{\omega_{k-1}}{1 - \omega_k} \\ &= \omega_{b_i} \rho_{b_i+1}^{-1} \dots \rho_{k-1}^{-1} (\rho_k^{-1} + 1) \\ &\leq e^{-(V(k)-V(b_i))} + e^{-(V(k-1)-V(b_i))}. \end{aligned}$$

Similarly, $\pi_i(k) \leq e^{-(V(k)-V(b_i))} + e^{-(V(k+1)-V(b_i))}$ for $k \in [a_i + 1, b_i - 1]$. Since the walk is reflected at a_i and \bar{c}_i , we have $\pi_i(a_i) = e^{-(V(a_i+1)-V(b_i))}$ and $\pi_i(\bar{c}_i) = e^{-(V(\bar{c}_i-1)-V(b_i))}$. Hence on the event $A_5(t, \eta)$ we have

$$\sup\{\pi_i(k); k \in [a_i, \bar{c}_i] \setminus (b_i - \eta \log t, b_i + \eta \log t)\} \leq C e^{-C''' \eta \log t} = C t^{-C''' \eta}.$$

Moreover, since π_i is an invariant measure and since $\pi_i(b_i) = 1$, we have, for all $k \geq 0$,

$$P_{\omega, |a_i, \bar{c}_i|}^{b_i}(X_k = x) \leq \pi_i(x).$$

Hence, on the event $A(t) \cap A_5(t, \eta)$ we have, for all $k \geq 0$,

$$(6.1) \quad P_{\omega, |a_i, \bar{c}_i|}^{b_i}(|X_k - b_i| > \eta \log t) \leq C(\log t) t^{-C''' \eta}.$$

Let ξ be a positive real, $0 < \xi < 1$. Then, let us write

$$\begin{aligned} & \liminf_{t \rightarrow \infty} \mathbb{P}(|X_t - b_{\ell_t}| \leq \eta \log t) \\ & \geq \liminf_{t \rightarrow \infty} \mathbb{P}(|X_t - b_{\ell_t}| \leq \eta \log t; \ell_t = \ell_{t(1+\xi)}) \\ & \geq \liminf_{t \rightarrow \infty} \mathbb{P}(\ell_t = \ell_{t(1+\xi)}) - \limsup_{t \rightarrow \infty} \mathbb{P}(|X_t - b_{\ell_t}| > \eta \log t; \ell_t = \ell_{t(1+\xi)}). \end{aligned}$$

Considering the first term, we get by using Proposition 2,

$$(6.2) \quad \begin{aligned} \liminf_{t \rightarrow \infty} \mathbb{P}(\ell_t = \ell_{t(1+\xi)}) &= \liminf_{t \rightarrow \infty} \mathbb{P}(T_{\ell_{t+1}} > t(1+\xi)) \\ &= \frac{\sin(\kappa\pi)}{\pi} \int_{\xi}^{\infty} \frac{dx}{x^{\kappa}(1+x)}. \end{aligned}$$

In order to estimate the second term, let us introduce the event

$$TT(t) := A(t) \cap A_5(t, \eta) \cap DT(t) \cap DT^*(t) \cap A^*(t) \cap IA(t) \cap LT(t) \cap IT(t),$$

where $IT(t) := \{T_{\ell_t} \leq t < T_{\ell_t} + \tau_{\ell_t}\}$. Observe that the preliminary results obtained in Section 4 together with Proposition 2 imply that $\mathbb{P}(TT(t)) \rightarrow 1$, when $t \rightarrow \infty$. Then, we have

$$\begin{aligned} &\limsup_{t \rightarrow \infty} \mathbb{P}(|X_t - b_{\ell_t}| > \eta \log t; \ell_t = \ell_{t(1+\xi)}) \\ &\leq \limsup_{t \rightarrow \infty} \mathbb{P}(TT(t); |X_t - b_{\ell_t}| > \eta \log t; \ell_t = \ell_{t(1+\xi)}) \\ &\leq \limsup_{t \rightarrow \infty} \mathbb{E} \left[\mathbf{1}_{TT(t)} \sum_{i=1}^{K_t} \mathbf{1}_{\{|X_t - b_i| > \eta \log t; \ell_t = \ell_{t(1+\xi)} = i\}} \right]. \end{aligned}$$

But on the event $TT(t) \cap \{\ell_t = \ell_{t(1+\xi)} = i\}$ we know that for all $k \in [T_i, t]$ the walk X_k is in the interval $[a_i, \bar{c}_i - 1]$. Indeed, on the event $LT(t) \cap DT(t) \cap IA(t)$ we know that once the position \bar{c}_i is reached then within a time $t/\log t$ the position b_{i+1} is reached which would contradict the fact that $\ell_{t(1+\xi)} = i$. Hence, we obtain, for all $i \in \mathbb{N}$,

$$\begin{aligned} &\mathbb{P}(TT(t); i \leq K_t; |X_t - b_i| > \eta \log t; \ell_t = \ell_{t(1+\xi)} = i) \\ &\leq \mathbb{E} \left[\mathbf{1}_{\{i \leq K_t\}} \mathbf{1}_{A(t) \cap A_5(t, \eta)} \sup_{k \in [0, t]} P_{\omega, |a_i, \bar{c}_i|}^{b_i} (|X_k - b_i| > \eta \log t) \right] \\ &\leq C(\log t) t^{-C'''\eta}, \end{aligned}$$

where we used the estimate (6.1) on the event $A(t) \cap A_5(t, \eta)$. Considering now that, on the event $A(t)$, the number $K(t)$ of deep valleys is smaller than $(\log t)^{\frac{\kappa+1}{2}}$ we get

$$\begin{aligned} \limsup_{t \rightarrow \infty} \mathbb{P}(|X_t - b_{\ell_t}| > \eta \log t; \ell_t = \ell_{t(1+\xi)}) &\leq \limsup_{t \rightarrow \infty} C(\log t)^{\frac{3+\kappa}{2}} t^{-C'''\eta} \\ &= 0. \end{aligned}$$

Then, letting ξ tends to 0 in (6.2) concludes the proof of part (i) of Theorem 2. \square

7. PART (ii) OF THEOREM 2: THE QUENCHED LAW OF THE LAST VISITED VALLEY

In order to prove the proximity of the distributions of ℓ_t and $\ell_{t, \omega}^{(e)}$, we go through $\ell_t^* = \sup\{n \geq 0, \tau_1^* + \dots + \tau_n^* \leq t\}$ whose advantage is to involve independent random variables whose laws are clearly identified.

Proposition 3. *Under assumptions (a)-(b) of Theorem 1, we have, for all $\delta > 0$,*

$$\lim_{t \rightarrow \infty} P\left(d_{TV}(\ell_t^*, \ell_{t, \omega}^{(e)}) > \delta\right) = 0,$$

where d_{TV} denotes the distance in total variation.

Proof. The strategy is to build a coupling between ℓ_t^* and $\ell_{t,\omega}^{(e)}$ such that

$$\lim_{t \rightarrow \infty} P(P_{0,\omega}(\ell_t^* \neq \ell_{t,\omega}^{(e)}) > \delta) = 0.$$

Let us first associate to the exponential variable \mathbf{e}_i the following geometric random variable

$$N_i := \left\lfloor \left(-\frac{1}{\log(p_i(\omega))} \right) \mathbf{e}_i \right\rfloor,$$

where $1 - p_i(\omega)$ denotes the probability for the random walk starting at b_i to go to d_i before returning to b_i , which is equal to $\omega_b \frac{e^{V(b_i)}}{\sum_{x=b_i}^{d_i-1} e^{V(x)}}$. The parameter of this geometric law is now clearly equal to $1 - p_i(\omega)$.

Now one can introduce like in [15] two random variables $F^{(i)}$ (resp. $S^{(i)}$) whose law are given by the time it takes for the random walk reflected at a_i , starting at b_i , to return to b_i (resp. to hit d_i) conditional on the event that d_i (resp. b_i) is not reached in between.

We introduce now a sequence of independent copies of $F^{(i)}$ we denote by $(F_n^{(i)})_{n \geq 0}$. The law of τ_i^* is clearly the same as $F_1^{(i)} + \dots + F_{N_i}^{(i)} + S^{(i)}$ which is going now to be compared with $E_\omega[\tau_i^*] \mathbf{e}_i$.

Let us now estimate, for a given $\xi > 0$ (small enough),

$$\begin{aligned} & \mathbb{P} \left(\forall i \leq K_t, \right. \\ & \quad \left. (1 - \xi)(F_1^{(i)} + \dots + F_{N_i}^{(i)} + S^{(i)}) \leq E_\omega[\tau_i^*] \mathbf{e}_i < (1 + \xi)(F_1^{(i)} + \dots + F_{N_i}^{(i)} + S^{(i)}) \right) \\ & \geq \mathbb{P} \left(\forall i \leq K_t, \quad (1 - \frac{\xi}{2})(F_1^{(i)} + \dots + F_{N_i}^{(i)}) \leq E_\omega[\tau_i^*] \mathbf{e}_i < (1 + \frac{\xi}{2})(F_1^{(i)} + \dots + F_{N_i}^{(i)}) \right) \\ (7.1) \quad & - \mathbb{P} \left(\exists i \leq K_t, \quad S^{(i)} > \frac{\xi}{3}(F_1^{(i)} + \dots + F_{N_i}^{(i)}) \right). \end{aligned}$$

Let us first treat the second quantity of the rhs of (7.1). For this purpose, we need an upper bound for $E_\omega[S^{(i)}]$ which is obtained exactly like in Lemma 13 of [15] and can be estimated by controlling the size of the falls (resp. rises) of the potential during its rises from $V(b_i)$ to $V(c_i)$ (resp. falls from $V(c_i)$ to $V(d_i)$), see Lemma 7. Indeed, the random variable $S^{(i)}$ concerns actually the random walk which is *conditioned* to hit d_i before b_i . Therefore, this involves an h -process which can be viewed as a random walk in a modified potential between b_i and d_i . This modified potential has a decreasing trend (which encourages the particle to go to the right), and the main contribution to $S^{(i)}$ comes from the small risings of this modified potential along its global fall.

More precisely, the particle starting at b_i which is conditioned to hit d_i before returning to b_i moves like a particle in the modified random potential $\bar{V}^{(i)}$ defined as follows: for all $b_i \leq x < y \leq d_i$,

$$(7.2) \quad \bar{V}^{(i)}(y) - \bar{V}^{(i)}(x) = (V(y) - V(x)) + \log \left(\frac{g^{(i)}(x)g^{(i)}(x+1)}{g^{(i)}(y)g^{(i)}(y+1)} \right),$$

where $g^{(i)}(x) := P_\omega^x(\tau(d_i) < \tau(b_i))$. The expectation of $S^{(i)}$ is given by the usual formula (see [27]), so that

$$E_\omega[S^{(i)}] \leq 1 + \sum_{k=b_i+1}^{d_i} \sum_{l=k}^{d_i} e^{\bar{V}^{(i)}(l) - \bar{V}^{(i)}(k)}.$$

We are therefore concerned by the largest rise of $\bar{V}^{(i)}$ inside the interval $[b_i, d_i]$. We first notice that, by standard arguments, for any $b_i \leq x < y \leq d_i$,

$$(7.3) \quad \frac{g^{(i)}(x)g^{(i)}(x+1)}{g^{(i)}(y)g^{(i)}(y+1)} = \frac{\sum_{j=b_i}^{x-1} e^{V(j)} \sum_{j=b_i}^x e^{V(j)}}{\sum_{j=b_i}^{y-1} e^{V(j)} \sum_{j=b_i}^y e^{V(j)}} \leq 1.$$

Therefore, we obtain for any $b_i \leq x < y \leq d_i$

$$(7.4) \quad \bar{V}^{(i)}(y) - \bar{V}^{(i)}(x) \leq V(y) - V(x).$$

This allows to bound the largest rise of $\bar{V}^{(i)}$ on the interval $[c_i, d_i]$ by the largest rise of V on this interval.

Concerning the largest rise of $\bar{V}^{(i)}$ on the interval $[b_i, c_i]$, we notice, taking into account the small size of the fluctuations of V described in Lemma 7, (7.3) and (7.4), that for all $\eta > 0$, for all $\omega \in A_4(t) \cap F_\eta(t)$, and for all $i \leq K_t$, the difference $\bar{V}^{(i)}(y) - \bar{V}^{(i)}(x)$ is less than or equal to

$$\begin{aligned} & [V(y) - \max_{b_i \leq j \leq y} V(j)] - [V(x) - \max_{b_i \leq j \leq x} V(j)] + O(\log \log t) \\ & \leq \eta \log t + O(\log \log t). \end{aligned}$$

This reasoning yields for all $\eta > 0$ that, for all $\omega \in A_4(t) \cap F_\eta(t)$,

$$\forall i \leq K_t, \quad E_\omega[S^{(i)}] \leq t^\eta.$$

This implies, by the Markov inequality, that, for all $\eta > 0$ and all $\omega \in A_4(t) \cap F_\eta(t)$,

$$\forall i \leq K_t, \quad P_\omega(S^{(i)} > t^{2\eta}) < \frac{1}{t^\eta}.$$

On the other hand, we have

$$P_\omega(F_1^{(i)} + \dots + F_{N_i}^{(i)} < t^{2\eta}) \leq P_\omega(N_i < t^{2\eta}) = 1 - p_i(\omega)^{\lfloor t^{2\eta} \rfloor} = O\left(\frac{t^{2\eta} \log t}{t}\right),$$

the last equality coming from the definition of $h_t := \log t - \log(\log t)$, which implies that $1 - p_i(\omega)$ is smaller than $\frac{\log t}{t}$. Hence, since $A_2(t) = \{K_t \leq (\log t)^{\frac{1+\kappa}{2}}\}$ satisfies $P(A_2(t)) \rightarrow 1$ (see Lemma 1), we obtain

$$\lim_{t \rightarrow +\infty} P\left(\forall i \leq K_t, \quad P_\omega(F_1^{(i)} + \dots + F_{N_i}^{(i)} < t^{2\eta}) \leq \frac{1}{t^{\frac{1}{2}-2\eta}}\right) = 1.$$

Gathering these two informations on $S^{(i)}$ and $F_1^{(i)} + \dots + F_{N_i}^{(i)}$, we obtain

$$\lim_{t \rightarrow +\infty} \mathbb{P}\left(\forall i \leq K_t, \quad S^{(i)} < \frac{\xi}{3}(F_1^{(i)} + \dots + F_{N_i}^{(i)})\right) = 1,$$

for all $\xi > 0$, which treats the second quantity of the rhs of (7.1).

The first quantity of the rhs of (7.1) is treated by going through

$$\mathbb{P}\left(\left(1 - \frac{\xi}{4}\right)N_i E_\omega[F^{(i)}] \leq F_1^{(i)} + \dots + F_{N_i}^{(i)} \leq \left(1 + \frac{\xi}{4}\right)N_i E_\omega[F^{(i)}]\right),$$

which, for all $\eta > 0$, is larger than

$$\begin{aligned} 1 - \mathbb{P}\left(\left\{\left|\frac{F_1^{(i)} + \dots + F_{N_i}^{(i)}}{N_i} - E_\omega[F^{(i)}]\right| > \frac{\xi}{4} E_\omega[F^{(i)}]\right\} \cap \{N_i \neq 0\} \cap \{E_\omega[(F^{(i)})^2] \leq t^\eta\}\right) \\ - P(E_\omega[(F^{(i)})^2] \geq t^\eta), \end{aligned}$$

which is in turn, using the Bienaimé-Chebychev's inequality, larger than

$$\begin{aligned} 1 - E \left[E \left(\frac{t^\eta}{N_i} \mathbf{1}_{\{N_i \neq 0\}} \mid N_i \right) \frac{16}{\xi^2 E_\omega [F]^2} \right] - P(E_\omega [(F^{(i)})^2] \geq t^\eta) \\ \geq 1 - \frac{16t^\eta}{\xi^2} \mathbb{E} \left[\frac{1}{N_i} \mathbf{1}_{\{N_i \neq 0\}} \right] - P(E_\omega [(F^{(i)})^2] \geq t^\eta). \end{aligned}$$

Now, we use again the reasoning based on h -processes to get an upper bound for $E_\omega [(F^{(i)})^2]$. Like in the success case, the particle starting at b_i which is conditioned to hit b_i before returning to d_i moves like a particle in the modified random potential $\widehat{V}^{(i)}$ defined as follows: for all $a_i \leq x < y \leq d_i$,

$$(7.5) \quad \widehat{V}^{(i)}(y) - \widehat{V}^{(i)}(x) = (V(y) - V(x)) + \log \left(\frac{h^{(i)}(x) h^{(i)}(x+1)}{h^{(i)}(y) h^{(i)}(y+1)} \right),$$

where $h^{(i)}(x) := P_\omega^x(\tau(b_i) < \tau(d_i))$ (notice that V and $\widehat{V}^{(i)}$ coincide on the interval $[a_i, b_i]$).

It happens now that $E_\omega [(F^{(i)})^2]$ can be computed explicitly in terms of $\widehat{V}^{(i)}$ (see Lemma 12 in [15]), and is bounded by a constant times $(d^{(i)} - a^{(i)})^2$ times the exponential of the maximum of the largest rise of V on $[a_i, b_i]$ and the largest fall of $\widehat{V}^{(i)}$ on $[b_i, d_i]$, which are treated in a similar way as the fluctuations of $\bar{V}^{(i)}$, above. So, we get

$$\forall \eta > 0, \quad P(E_\omega [(F^{(i)})^2] \geq t^\eta) = o\left(\frac{1}{(\log t)^2}\right).$$

Moreover, we have

$$E \left[\frac{1}{N_i} \mathbf{1}_{\{N_i \neq 0\}} \right] = E \left[-\frac{1 - p_i(\omega)}{p_i(\omega)} \log(1 - p_i(\omega)) \right] = O\left(\frac{(\log t)^2}{t}\right).$$

As a result, we obtain

$$\mathbb{P} \left(\left(1 - \frac{\xi}{4}\right) N_i E_\omega [F^{(i)}] \leq F_1^{(i)} + \dots + F_{N_i}^{(i)} \leq \left(1 + \frac{\xi}{4}\right) N_i E_\omega [F^{(i)}] \right) = 1 - o\left(\frac{1}{(\log t)^2}\right).$$

Now, the second step in the estimation of the first quantity of the rhs of (7.1) is the examination, for $\xi > 0$, of

$$\mathbb{P} \left(\left(1 - \frac{\xi}{4}\right) N_i E_\omega [F^{(i)}] \leq E_\omega [\tau_i] \mathbf{e}_i \leq \left(1 + \frac{\xi}{4}\right) N_i E_\omega [F^{(i)}] \right),$$

i.e.

$$\mathbb{P} \left(\left(1 - \frac{\xi}{4}\right) N_i E_\omega [F^{(i)}] \leq (E_\omega [N_i] E_\omega [F^{(i)}] + E_\omega [S^{(i)}]) \mathbf{e}_i \leq \left(1 + \frac{\xi}{4}\right) N_i E_\omega [F^{(i)}] \right).$$

Neglecting again, like above, the contribution of $S^{(i)}$ we are back to prove that

$$\mathbb{P} \left(\left(1 - \frac{\xi}{4}\right) \left[\left(-\frac{1}{\log(p_i(\omega))}\right) \mathbf{e}_i \right] \leq \frac{p_i(\omega)}{1 - p_i(\omega)} \mathbf{e}_i \leq \left(1 + \frac{\xi}{4}\right) \left[\left(-\frac{1}{\log(p_i(\omega))}\right) \mathbf{e}_i \right] \right) = 1 - o\left(\frac{1}{(\log t)^2}\right),$$

which is a direct consequence of $1 - p_i(\omega) \leq \frac{\log t}{t}$ allied with

$$P^{(\mathbf{e})} \left(\mathbf{e}_i > \frac{\log t}{t} \right) = 1 - o\left(\frac{1}{(\log t)^2}\right).$$

Now, since $P(K_t \leq (\log t)^{\frac{1+\kappa}{2}}) \rightarrow 1$, when $t \rightarrow \infty$, this concludes the proof that the rhs (and therefore the lhs) of (7.1) tends to 1 when t tends to infinity. Indeed, we obtain

$$\mathbb{P}\left(\forall i \leq K_t, (1-\xi)(F_1^{(i)} + \dots + F_{N_i}^{(i)} + S^{(i)}) \leq E_\omega[\tau_i^*] \mathbf{e}_i < (1+\xi)(F_1^{(i)} + \dots + F_{N_i}^{(i)} + S^{(i)})\right) \rightarrow 1,$$

from which we deduce

$$\mathbb{P}\left(\forall i \leq K_t, (1-\xi)(\tau_1^* + \dots + \tau_i^*) \leq \sum_{k=1}^i E_\omega[\tau_k^*] \mathbf{e}_k < (1+\xi)(\tau_1^* + \dots + \tau_i^*)\right) \rightarrow 1.$$

Moreover, we use the fact that

$$E_\omega[\tau_k^*] = W_k^* - (d_k^* - b_k^*),$$

where (cf for example [27], formula (2.1.14))

$$W_k^* = 2 \sum_{\substack{a_k^* \leq m \leq n \\ b_k^* \leq n \leq d_k^*}} e^{V_\omega(n) - V_\omega(m)}.$$

Since on the event $A_4(t)$ we have for all $k \leq K_t$, $d_k^* - b_k^* \leq C'' \log t$, we see that on $A_4(t)$ we have $(1 - C'' \frac{(\log t)^2}{t}) W_k^* \leq E_\omega[\tau_k^*] \leq W_k^*$, since $W_k^* \geq e^{ht}$ by definition. Hence, it implies that

$$\mathbb{P}\left(\forall i \leq K_t, (1-\xi)(\tau_1^* + \dots + \tau_i^*) \leq \sum_{k=1}^i W_k^* \mathbf{e}_k < (1+\xi)(\tau_1^* + \dots + \tau_i^*)\right) \rightarrow 1.$$

Applying this, for $i = \ell_t^*$ and $i = \ell_{t,\omega}^{(\mathbf{e})}$ we get respectively that, for all $\xi > 0$,

$$\mathbb{P}\left(\ell_t^* \leq \ell_{\frac{t}{1-\xi}, \omega}^{(\mathbf{e})}\right) \rightarrow 1 \quad \text{and} \quad \mathbb{P}(\ell_{t,\omega}^{(\mathbf{e})} \leq \ell_{t(1+\xi)}^*) \rightarrow 1.$$

We conclude now the proof by reminding that $\lim_{\xi \rightarrow 0} \mathbb{P}(\ell_t^* = \ell_{t(1+\xi)}^*) = 1$ as well as $\lim_{\xi \rightarrow 0} \mathbb{P}(\ell_{t,\omega}^{(\mathbf{e})} = \ell_{\frac{t}{1-\xi}, \omega}^{(\mathbf{e})}) = 1$. \square

Proof of Part (ii) of Theorem 2. The passage from Proposition 3 to Part (ii) of Theorem 2 is of the same kind as the passage from Proposition 1 to Proposition 2. \square

8. PROOF OF THEOREM 1

We fix $h > 1$ and $\eta > 0$ (η was used to define the event $A_5(t, \eta)$ before Lemma 9). Let us introduce the event

$$TT(t, h) := TT(t) \cap \{X_t - b_{\ell_t} \leq \frac{\eta}{2} \log t\} \cap \{X_{t_h} - b_{\ell_{t_h}} \leq \frac{\eta}{2} \log t\},$$

whose probability tends to 1, when t tends to infinity (it is a consequence of Section 4 together with part (i) of Theorem 2). Then, we easily have

$$(\{\ell_{t_h} = \ell_t\} \cap TT(t, h)) \subset (\{|X_{t_h} - X_t| \leq \eta \log t\} \cap TT(t, h)).$$

Moreover, observe that on $TT(t)$, $\ell_{t_h} > \ell_t$ implies that $|b_{\ell_{t_h}} - b_{\ell_t}| \geq t^{\kappa/2}$ (by definition of $A_3(t)$). Therefore, we get

$$(\{|X_{t_h} - X_t| \leq \eta \log t\} \cap TT(t, h)) \subset (\{\ell_{t_h} = \ell_t\} \cap TT(t, h)),$$

for all large t . Thus, since Proposition 2 implies that $\lim_{t \rightarrow \infty} \mathbb{P}(\ell_{th} = \ell_t)$ exists, we obtain

$$\begin{aligned} \lim_{t \rightarrow \infty} \mathbb{P}(|X_{th} - X_t| \leq \eta \log t) &= \lim_{t \rightarrow \infty} \mathbb{P}(\ell_{th} = \ell_t) \\ &= \lim_{t \rightarrow \infty} \mathbb{P}(T_{\ell_{t+1}} \geq th) \\ &= \frac{\sin(\kappa\pi)}{\pi} \int_{h-1}^{\infty} \frac{dx}{x^\kappa(1+x)} \\ &= \frac{\sin(\kappa\pi)}{\pi} \int_0^{1/h} y^{\kappa-1}(1-y)^{-\kappa} dy, \end{aligned}$$

which concludes the proof of Theorem 1. \square

REFERENCES

- [1] Ben Arous, G., Bovier, A. and Gaynard, V. (2003). Glauber dynamics of the random energy model. I. Metastable motion on the extreme states. *Comm. Math. Phys.* **235**, 379–425.
- [2] Ben Arous, G., Bovier, A. and Gaynard, V. (2003). Glauber dynamics of the random energy model. II. Aging below the critical temperature. *Comm. Math. Phys.* **236**, 1–54.
- [3] Ben Arous, G. and Černý, J. (2005). Bouchaud’s model exhibits two aging regimes in dimension one. *Ann. Appl. Probab.* **15**, 1161–1192.
- [4] Ben Arous, G. and Černý, J. (2006). *Dynamics of trap models*, Ecole d’Été de Physique des Houches, Session LXXXIII “Mathematical Statistical Physics”, pp. 331–394. Elsevier.
- [5] Ben Arous, G. and Černý, J. (2007). Scaling limit for trap models on \mathbb{Z}^d . *Ann. Probab.* **35**, 2356–2384.
- [6] Ben Arous, G. and Černý, J. (2008). The arcsine law as a universal aging scheme for trap models. *Comm. Pure and Appl. Math.* **61**, 289–329.
- [7] Ben Arous, G., Černý, J. and Mountford, T. (2006). Aging for Bouchaud’s model in dimension two. *Probab. Theory Related Fields* **134**, 1–43.
- [8] Bouchaud, J.-P. (1992). Weak ergodicity breaking and aging in disordered systems. *J. Phys. I (France)* **2**, 1705–1713.
- [9] Bouchaud, J.-P., Cugliandolo, L., Kurchan, J. and Mézard, M. (1998). Out of equilibrium dynamics in spin-glasses and other glassy systems. *Spin-glasses and Random Fields*, 161–224. (A.P. Young, Ed.), World Scientific
- [10] Bouchaud, J.-P. and Dean, D. S. (1995). Aging on Parisi’s tree. *J. Phys. I (France)* **5**, 265–286.
- [11] Cocco, S. and Monasson, R. (2008). Reconstructing a random potential from its random walks. *Europhysics Letters* **81**, 20002.
- [12] Dembo, A., Guionnet, A. and Zeitouni, O. (2004). Aging properties of Sinai’s model of random walk in random environment. In *St. Flour summer school 2001 lecture notes by O. Zeitouni*. [ArXiv:math/0105215](https://arxiv.org/abs/math/0105215).
- [13] Le Doussal, P., Fisher, D.S. and Monthus, C. (1999). Random walkers in one-dimensional random environments: Exact renormalization group analysis. *Phys. Rev.* **E 59**, 4795–4840.
- [14] Enriquez, N., Sabot, C. and Zindy, O. (2009). A probabilistic representation of constants in Kesten’s renewal theorem. To appear in *Probability Theory and Related Fields*. [ArXiv:math/0703648](https://arxiv.org/abs/math/0703648).
- [15] Enriquez, N., Sabot, C. and Zindy, O. (2009). Limit laws for transient random walks in random environment on \mathbb{Z} . To appear in *Annales de l’Institut Fourier*. [ArXiv:math/0703660](https://arxiv.org/abs/math/0703660).
- [16] Feller, W. (1971). *An Introduction to Probability Theory and its Applications*, Vol. II. (2nd ed.). Wiley, New York.

- [17] Fontes, L. R., Isopi, M. and Newman, C. M. (2002). Random walks with strongly inhomogeneous rates and singular diffusions: convergence, localization and aging in one dimension. *Ann. Probab.* **30**, 579–604.
- [18] Golosov, A.O. (1984). Localization of random walks in one-dimensional random environments. *Comm. Math. Phys.* **92**, 491–506.
- [19] Golosov, A. O. (1986). Limit distributions for random walks in random environments. *Soviet Math. Dokl.* **28**, 18–22.
- [20] den Hollander, F. (2000). *Large Deviations*. Fields Institute Monographs **14**. AMS, Providence, RI.
- [21] Iglehart, D.L. (1972). Extreme values in the GI/G/1 queue. *Ann. Math. Statist.* **43**, 627–635.
- [22] Kesten, H. (1986). The limit distribution of Sinai’s random walk in random environment. *Physica A* **138**, 299–309.
- [23] Kesten, H., Kozlov, M.V. and Spitzer, F. (1975). A limit law for random walk in a random environment. *Compositio Math.* **30**, 145–168.
- [24] Peterson, J. and Zeitouni, O. (2009). Quenched limits for transient, zero speed one-dimensional random walk in random environment. *Ann. Probab.* **37**, 143–188.
- [25] Sinai, Ya.G. (1982). The limiting behavior of a one-dimensional random walk in a random medium. *Th. Probab. Appl.* **27**, 256–268.
- [26] Solomon, F. (1975). Random walks in a random environment. *Ann. Probab.* **3**, 1–31.
- [27] Zeitouni, O. (2004). *Random Walks in Random Environment*, XXXI summer school in probability, St Flour (2001), *Lecture Notes in Math.* **1837**, pp. 193–312. Springer, Berlin.

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