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1 **Harbin: A quantitation PCR analysis tool**

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16

17 **Abstract**

18 Objectives

19 To enable analysis and comparisons of different relative quantitation experiments, a web-
20 browser application called Harbin was created that uses a quantile-based scoring system for
21 the comparison of samples at different time points and between experiments.

22 Results

23 Harbin uses the standard curve method for relative quantitation to calculate concentration
24 ratios (CRs). To evaluate if different datasets can be combined the Harbin quantile bootstrap
25 test is proposed. This test is more sensitive in detecting distributional differences between

26 data sets than the Kolmogorov-Smirnov test. The utility of the test is demonstrated in a
27 comparison of three grapevine leafroll associated virus 3 (GLRaV-3) RT-qPCR data sets.

28 Conclusions

29 The quantile-based scoring system of CRs will enable the monitoring of virus titre or gene
30 expression over different time points and be useful in other genomic applications where the
31 combining of data sets are required.

32

33 **Keywords**

34 Bootstrap; genetic variants; GLRaV-3; Location-scale-shape problem; Quantile-based
35 scoring; RT-qPCR

36

37 **Article section**

38 Plant Cell Technology

39 **Introduction**

40

41 Quantitative polymerase chain reaction (qPCR) is a widely used technique to measure
42 expression levels of nucleic acids. Absolute quantitation uses a fixed calibration curve that
43 makes comparing different experiments easier, however relative quantitation compensates for
44 differences in tissue types, environmental conditions, integrity of RNA, loading error and
45 reaction efficiency. A concentration ratio (CR) can be obtained to compare the concentration
46 of a gene of interest relative to stable reference genes. A relative quantitation model with an
47 efficiency correction is recommended since a small difference in target assay efficiency and
48 reference gene assay efficiency can result in a false expression ratio (Pfaffl 2001; Bester et al.
49 2014).

50 One of the most important viral diseases of grapevine worldwide is grapevine leafroll disease
51 (GLD) with grapevine leafroll-associated virus 3 (GLRaV-3) considered as the main
52 etiological agent contributing to the disease (Maree et al. 2013). Currently, the complete
53 genomes of 13 distinct GLRaV-3 isolates representing five of the eight major genetic variant
54 groups are available (Maree et al. 2015). Little is known about the biological characteristics
55 of the different GLRaV-3 genetic variants and it is therefore important to investigate whether
56 there is significant variation between the variant groups beyond the genome. One parameter
57 to investigate would be the CR of the different groups over time.

58 The comparison of replicate experiments over time is complicated by differences in the
59 location, scale and shape of the population distributions, i.e. data with differences in these
60 parameters are not directly comparable. The most commonly used method to determine the
61 compatibility of data is to test for shifts in shape (distribution), location (mean) and scale
62 (variance). To address this, we propose a new bootstrap test for hypothesis against the
63 location-scale-shape alternative, based on quantiles of the empirical distributions of two data

64 sets. This test is not based on any assumptions about the shapes of the distributions, and is
65 powerful in detecting differences in location, scale and/or shape simultaneously. The
66 accuracy of this novel test was compared in a Monte Carlo simulation study to the well-
67 known Kolmogorov-Smirnov test (Kolmogorov 1933).

68 A software tool called Harbin is presented, for the analysis of real-time qPCR data using a
69 relative quantitation strategy. It allows for the combining of different qPCR data
70 sets/experiments to enable comparisons of different relative quantitation experiments. Harbin
71 runs within the R statistical computing environment (R Core Team, 2013) on all major
72 platforms. It is also freely available as a graphical user interface (GUI) utilizing the Shiny
73 web-based package that requires no additional software installations. The utility of Harbin
74 was demonstrated using three GLRaV-3 RT-qPCR data sets to investigate if the data sets can
75 be combined to study variation in virus variant concentrations.

76

77 **Methods**

78

79 **Plant material**

80

81 Three independent sample groups were selected for this study, all consisting of *Vitis vinifera*
82 cv. Cabernet Sauvignon plants. The first data set included 30 samples of which 15 samples
83 were infected with grapevine leafroll associated virus 3 (GLRaV-3) variant group II and 15
84 samples infected with GLRaV-3 variant group VI. The second data set included 12 plants
85 singly infected with either variant group I, II, III or VI (three plants each). The third data set
86 included 37 plants of which seven plants were infected with variant groups I, eight plants
87 infected with variant group II, eight plants infected with variant group III, eight plants
88 infected with variant group VI and six plants infected with variant group VII.

89 Due to GLRaV-3 being a phloem-limited virus, phloem material from each plant shoot was
90 collected and stored at $-80\text{ }^{\circ}\text{C}$. Total RNA was extracted from 2 grams of phloem material
91 using a modified CTAB extraction protocol (Carra et al. 2009; Bester et al. 2014). All plants
92 were confirmed to be infected with only GLRaV-3 after testing negative for frequently
93 occurring grapevine viruses using RT-PCRs (Jooste et al. 2015). GLRaV-3 variant group
94 status of all plants was confirmed using the previously designed real-time RT-PCR high-
95 resolution melting curve analysis assay (Bester et al. 2012).

96

97 RT-qPCRs

98

99 In order to calculate the virus CR in each plant, RT-qPCRs were performed using previously
100 designed assays targeting ORF1a of GLRaV-3 and three *V. vinifera* reference genes targeting
101 actin, alpha-tubulin and glyceraldehyde 3-phosphate dehydrogenase (GAPDH) (Bester et al.
102 2014). The stability of the reference genes was assessed using BestKeeper (Pfaffl et al. 2004).

103

104 Data analysis

105

106 The Rotor-gene Q software version 2.3.1 (Qiagen) was used to calculate primer efficiencies,
107 Cq values and gene quantitation values for all targets. For further analysis of the three data
108 sets, an R based application called Harbin was developed to ease the data handling and
109 computational aspects. Harbin runs within the R statistical computing environment (R Core
110 Team, <http://www.R-project.org/>) on all major platforms, and is available under an open
111 source licence. Harbin is dependent on base R and additional packages (psych, car,
112 beeswarm) available from the Comprehensive R Archive Network (CRAN). Harbin is also
113 available as a graphical user interface (GUI) utilizing the Shiny web-based package. The GUI

114 can be used in most web browsers and requires only an Internet connection and no
115 installation. The Harbin user manual is available for download from within the application or
116 at <https://github.com/Rbester18/Harbin>. Harbin has a direct input option for the quantitation
117 files (.csv) generated by the Rotor-Gene Q software (version 2.3.11 and above). The
118 application also allows for the upload of Cq values from any other qPCR platform, provided
119 that a standard curve equation for each gene is available. An example template is available
120 for download from within the application. Normalisation of the gene of interest
121 concentrations are performed with a reference gene index, calculated using the geometric
122 mean of up to ten reference genes. The calculation of fold changes between genes often
123 entails only limited comparisons of values across two conditions, however the Harbin
124 application allows for significance testing of two or more groups using either parametric or
125 non-parametric tests by selecting and classifying individual data points to the number of
126 groups specified. The non-parametric Wilcoxon rank-sum test can be used to assess
127 statistically significant differences between samples infected with different variant groups.
128 The Harbin application and additional information can be used and downloaded at
129 <https://rbester.shinyapps.io/Harbin/> and <https://github.com/Rbester18/Harbin>.

130

131 Harbin quantile-based bootstrap test

132

133 The Harbin application was used to perform the quantile-based bootstrap test (Harbin-test) to
134 determine if the three data sets are compatible to be combined. For each data set, the 20th,
135 40th, 60th and 80th percentiles of the CRs distribution are calculated and assigned a score (1–
136 5). A CR in the lowest quantile (0–20%) is assigned a “1”, and a CR in the highest quantile
137 (80–100%) is assigned a “5”. If data from a previous experiment is available and the option
138 to use it as a reference data set is selected, the application will compare the test data to the

139 reference data set. The Harbin-test adds the data set to the reference dataset and calculates the
140 number of CRs in the reference data set for which the “scores” (1–5) have changed. This test
141 statistic is compared to the distribution of the same statistic calculated from 1000 bootstrap
142 samples (each of the same size as the test data) drawn from the reference data set.

143 The purpose of the Harbin-test function is to determine whether the samples in a new data set
144 are compatible with those in a well-defined reference data set. The combining of different
145 data sets is performed under the assumption that the samples originate from populations that
146 can be described by the same probability distribution function. Suppose that $\mathbf{x}' =$
147 $[x_1, \dots, x_n]$ and $\mathbf{y}' = [y_1, \dots, y_m]$ are representative data sets from two continuous
148 univariate populations, G_{ref} and F , respectively. It is of interest to determine whether the two
149 population distributions are homogeneous, or in particular, whether the new data set \mathbf{y} is
150 compatible with the reference data set, \mathbf{x} . The hypothesis of interest is

151

$$152 \quad H_0 : F(x) = G_{ref}(x), \text{ for all } x \in (-\infty, \infty),$$

$$153 \quad (1)$$

154

155 against the general location-scale-shape alternative,

156

$$157 \quad H_1 : F(x) \neq G_{ref}(x), \text{ for some } x \in (-\infty, \infty),$$

$$158 \quad (2)$$

159

160 where F and G_{ref} are continuous univariate probability distribution functions describing the
161 two populations. Hypothesis (2) implies a difference at any point on the two distributions:
162 The medians, variances and/or shapes of the two distributions differ.

163 The Harbin-test is a quantile-based bootstrap test for hypothesis (1) against the general
 164 alternative in (2). The test works as follows: Calculate the 20th, 40th, 60th and 80th percentiles
 165 of \mathbf{x} , indicating these percentiles with Q_{20} , Q_{40} , Q_{60} and Q_{80} , respectively. Let $g_i, i =$
 166 $1, \dots, n$ be a variable taking the values,

$$168 \quad g_i = \begin{cases} 1 & \text{if } x_i \leq Q_{20}, \\ 2 & \text{if } Q_{20} < x_i \leq Q_{40}, \\ 3 & \text{if } Q_{40} < x_i \leq Q_{60}, \\ 4 & \text{if } Q_{60} < x_i \leq Q_{80}, \\ 5 & \text{if } x_i > Q_{80}. \end{cases}$$

169 (3)

170

171 Combine the reference and new data sets in a vector, $\mathbf{z}' = [\mathbf{x}' \mathbf{y}']$ and construct a variable,
 172 $h_i, i = 1, \dots, n$, taking the values,

$$174 \quad h_i = \begin{cases} 1 & \text{if } x_i \leq Q_{20}^*, \\ 2 & \text{if } Q_{20}^* < x_i \leq Q_{40}^*, \\ 3 & \text{if } Q_{40}^* < x_i \leq Q_{60}^*, \\ 4 & \text{if } Q_{60}^* < x_i \leq Q_{80}^*, \\ 5 & \text{if } x_i > Q_{80}^*. \end{cases}$$

175 (4)

176

177 where Q_p^* indicates the p^{th} percentile of \mathbf{z} . Let

178

$$179 \quad c_i = \begin{cases} 0 & \text{if } g_i = h_i, \\ 1 & \text{if } g_i \neq h_i. \end{cases} \quad (5)$$

180

181 The quantity $\sum_{i=1}^n c_i$ is thus the number of elements in \mathbf{x} for which the “scores” (1–5) have
 182 changed in the combined data set, \mathbf{z} . The test statistic for hypothesis (1) is

183

$$184 \quad u = \frac{1}{n} \sum_{i=1}^n c_i, \tag{6}$$

185

186 which is the proportion of the elements in x for which the scores have changed in the
187 combined data set. To find the distribution of u under the null hypothesis, $r = 1000$
188 bootstrap samples (Efron and Tibshirani 1994) of size m are drawn from x . Let

189

$$190 \quad z_0^{(j)} = \begin{bmatrix} x \\ y_0^{(j)} \end{bmatrix}, j = 1, \dots, r, \tag{7}$$

191

192 where $y_0^{(j)}$ indicates the j^{th} bootstrap sample. Using x and $z_0^{(j)}$, the j^{th} bootstrap replication
193 of the test statistic, $u_0^{(j)}$, is calculated as in (6). The null hypothesis in (1) is rejected at a
194 significance level of α if the test statistic in (6) exceeds the $100(1 - \alpha)^{th}$ percentile of $u'_0 =$
195 $[u_0^{(1)}, \dots, u_0^{(r)}]$.

196 Two example data sets are available on github (<https://github.com/Rbester18/Harbin>) and
197 will be able to serve as independent reference data sets if the same qPCR protocol and
198 reagents are used as described in this study.

199

200 Monte Carlo simulation study

201

202 A Monte Carlo simulation study was performed to compare the size and power of the Harbin-
203 test to the Kolmogorov-Smirnov test (Hollander et al. 2013). Compared to the number of
204 available tests for common location and/or homogeneity of variances for two groups,
205 relatively few tests have been proposed to test for equality of the population distributions. A

206 well-known non-parametric test for the two-sample hypothesis in (1) against the location-
207 scale-shape alternative in (2) is the Kolmogorov-Smirnov test.

208 The Kolmogorov-Smirnov test compares the empirical distribution functions of two data sets.
209 If differences in the locations, scales or shapes of the empirical distribution functions are
210 sufficiently large, the conclusion is made that the two population distribution functions differ.

211 For the first (“reference database”) group, data sets of sizes $n_1 = 10, 30$ or 50 were
212 simulated from populations with one of the following four distributions:

- 213 • 1a. Normal: $N(3, 1)$;
- 214 • 1b. Chi-squared with three degrees of freedom: χ_3^2 ;
- 215 • 1c. Uniform distribution on the $[0, 6]$ interval;
- 216 • 1d. Bimodal: Half of observations from a $N(1.5, 0.75^2)$ distribution, with the other
217 half from a $N(4.5, 0.75^2)$ distribution.

218 For all four of the distribution types, the majority of the observations will thus lie on the
219 $[0, 6]$ interval, as can be seen in Fig. 1. For the second (“new data”) group, data sets of size
220 n_2 for ratios $\frac{n_2}{n_1} = 0.5, 1$ or 2 , were simulated from populations with one of the following four

221 distribution types:

- 222 • 2a. Normal: $N(3 + \delta, 1\gamma)$;
- 223 • 2b. Chi-squared with 3γ degrees of freedom, shifted to the right by addition of the
224 value δ ; i.e. $\chi_{3\gamma}^2 + \delta$
- 225 • 2c. Uniform distribution on the $[0, 6\gamma]$ interval, shifted by addition of the quantity
226 $(-3\gamma + 3 + \delta)$;
- 227 • 2d. Bimodal: Half of observations from a $N(1.5 + \delta, (0.75\gamma)^2)$ distribution, with the
228 other half from a $N(1.5 + 3\gamma + \delta, (0.75\gamma)^2)$ distribution.

229 The mean shift values, $\delta = 0, 0.2, 0.5, 1, 1.5, 2$, and standard deviation shift values, $\gamma =$
230 $1, 1.5, 2$, were varied to determine the power of the two tests to detect shifts in location and
231 scale, respectively. For each ($n_1 : n_2 : \text{Distribution 1 type} : \text{Distribution 2 type} : \delta : \gamma$) factorial
232 treatment combination, a total of $r = 1000$ simulation runs were performed. The simulation
233 study was performed on the Rhasatsha high-performance computer (HPC) at Stellenbosch
234 University (<http://www.sun.ac.za/hpc>), using R (R Core Team, 2013). For each test per
235 simulation run, a significance level of 5 % was used to decide whether to reject the null
236 hypothesis.

237 The Harbin application has the option to apply either the Harbin-test or the Kolmogorov-
238 Smirnov test to test the two-sample hypothesis. If the hypothesis that the two data sets
239 originated from populations with the same probability distribution function seems plausible,
240 the Harbin application allows for the option to add the new data set to the reference data set.
241 The quantile scores of the data in the reference data set will be adjusted according to the new
242 combined data distribution.

243

244 **Results and discussion**

245

246 RT-qPCRs

247

248 The utility of the Harbin application is demonstrated in a comparison of three GLRaV-3 RT-
249 qPCR data sets. The requisite control reactions were included in all data sets, and as expected
250 no virus CRs were generated for GLRaV-3 negative plant samples. The statistics of the
251 standard curves generated for each assay per data set can be seen in Table 1. The PCR
252 efficiencies and linearity calculated from all assays' standard curves were high and no
253 evidence of inhibition was seen from the Cq values of the dilution series. These assays

254 complied with the Minimum Information for publication of Quantitative real-time PCR
255 Experiments (MIQE) guidelines to ensure the integrity of the experiments and facilitate
256 reproducibility (Bustin et al. 2009).

257

258 Monte Carlo simulation study

259

260 The overall performance of the Harbin-test and the Kolmogorov-Smirnov test was assessed
261 by the percentage of simulation runs for which the null hypothesis was correctly rejected (or
262 not rejected) for the specific test. The Kolmogorov-Smirnov test had the smaller size (2.4%)
263 and power (54%), indicating that it is conservative compared to the Harbin-test, failing to
264 reject an incorrect null hypothesis in a larger proportion of cases. The Harbin-test was found
265 to be consistently more accurate and powerful than the Kolmogorov-Smirnov test, but had a
266 higher false positive rate (10.6%). Therefore the Harbin-test offers a good alternative in
267 situations where the purpose is to avoid considering samples from two different distributions
268 as originating from populations with the same distribution. The power of both tests increases
269 with an increase in the size of the sample from the first (“reference data set”) population. For
270 the smallest sample size considered ($n_I = 10$), the Harbin test outperformed the Kolmogorov-
271 Smirnov test. This advantage disappeared in the larger sample size scenarios ($n_I = 30, 50$),
272 where both tests have nearly equal power.

273 For populations with the same distribution types (for example, 1a vs. 2a, 1b vs. 2b, etc.), it is
274 of interest to compare the power of the two tests to detect differences in location and/or scale
275 only. The Harbin test showed greater power (68.8%) on the simulated data compared to the
276 Kolmogorov-Smirnov test (56.7%). The Harbin-test was the most powerful in detecting
277 location shifts.

278 One important purpose of the Harbin test is to detect differences in the shapes of two
279 population distributions, when the locations and scales of the populations are approximately
280 equal. To assess the performance of the two tests in this regard, the power of the tests for
281 detecting only differences in distribution types were calculated. The Harbin test is more
282 powerful (11.4%) than the Kolmogorov-Smirnov test (2.2%) in this regard.

283 Considering the detection of location and/or scale shifts for two populations with different
284 distribution types, the Harbin test is also more powerful (65.4%) than the Kolmogorov-
285 Smirnov test (53.2%). Location shifts, scale shifts and changes in sample size from two
286 populations with different distribution types showed the same effects on the tests as was
287 observed overall.

288 Both the Kolmogorov-Smirnov test and the Harbin-test are able to compare data sets
289 irrespective of the relationship between the data sets.

290

291 Harbin-test

292

293 When comparing the three qPCR data sets generated from the greenhouse samples, it seemed
294 possible that the data sets originated from populations which can be described by the same
295 probability distribution function. The Harbin-test shows that this assumption is likely, as only
296 6.67 % of the scores assigned to values in the first data set changed when the second data set
297 was added (p-value = 0.906). Only 13.04 % of scores assigned to values in the newly
298 combined data set changed with the addition of the third data set (p-value = 0.187).

299 Therefore, it was concluded that the three data sets are compatible and can be combined for
300 further analyses. The decision to combine data sets remains the user's responsibility. It is
301 important to ensure that all qPCR data were generated using the same protocol and reagents.

302 Neither the Harbin-test nor the Kolmogorov-Smirnov test takes into account any biological

303 factors and therefore careful consideration should be given to the experimental setup before
304 combining data sets. The combining of data sets is beneficial when sample numbers are large,
305 experiments need to be extended over a long period of time or when different time points
306 need to be compared. The cumulative addition of subsequent samples to a reference data set
307 will ensure an increase in the confidence with which each quantile score represents a true
308 distribution of CRs unique to the specific quantile.

309 The distribution of the quantile scores and the change in quantile score distribution with the
310 addition of data sets can be seen in Fig. 2. The addition of the second data set lowered one
311 and raised one of the quantile scores of the first data set by one score. With the addition of the
312 third data set to the combined data of data set 1 and 2, six quantiles scores were raised with
313 one score of which four were in data set 1 and two in data set 2.

314

315 **Conclusions**

316

317 Harbin simplifies the analysis of high-density qPCR assays, either for individual experiments
318 or across sets of replicates and biological conditions. The Harbin-test for the combining of
319 data sets was shown to be less conservative than the Kolmogorov-Smirnov test, and therefore
320 more sensitive in detecting distributional differences between data sets. Both tests are able to
321 compare data sets irrespective of the relationship between the data sets. The quantile-based
322 scoring system of CRs will allow for comparison of samples between experiments and
323 different time points, aiding the monitoring of virus titre or gene expression over a season or
324 longer period of time. The Harbin application and the Harbin-test will ease the data analysis
325 associated with virus quantitation to monitor disease spread in vineyards. In this study a
326 quantile score was assigned to each virus concentration ratio of GLRaV-3 single variant
327 infections in three independent data sets. The addition of more data to the reference database

328 will increase the confidence of the quantile boundaries as they will eventually stabilise and
329 provide a scoring system for virus concentrations. This enables the simplified comparison of
330 virus concentrations between different variants of GLRaV-3. The addition of mixed variant
331 infections and more time points to study variation over time will aid the investigation into the
332 biological characteristics of the different variant groups and their individual contribution to
333 GLD. It is envisioned that the Harbin-test will also be useful in other genomic applications
334 where the combining of data sets can be beneficial. The application runs in any web-browser,
335 and requires no programming experience from the user. This increases the accessibility of the
336 Harbin quantitation framework for analysis of qPCR data.

337

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339

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344

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346

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386

387 **Table 1** RT-qPCR standard curve statistics per data set
 388

Assay	Efficiency	r^2	Slope	y-intercept (b)
Data set 1				
GLRaV-3 ORF1a	1.02	0.996	-3.286	20.623
actin	0.99	0.997	-3.349	17.746
GAPDH	1.01	0.997	-3.305	20.615
alpha-Tubulin	1.00	0.996	-3.317	20.090
Data set 2				
GLRaV-3 ORF1a	0.96	0.996	-3.413	27.193
actin	1.06	0.998	-3.180	25.579
GAPDH	0.97	0.995	-3.393	24.546
alpha-Tubulin	0.94	0.995	-3.469	23.998
Data set 3				
GLRaV-3 ORF1a	0.87	0.993	-3.667	17.29
actin	0.91	0.991	-3.559	18.309
GAPDH	0.72	0.99	-4.243	19.503
alpha-Tubulin	0.92	0.995	-3.524	21.153

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395 **Fig. 1** Empirical examples of the four distribution types for the first group (1a, 1b, 1c and 1d)
396 used in the simulation study

397

398 **Fig. 2** Concentration ratio (CR) distribution per data set. Dotted lines indicate the quantile
399 boundaries. The change in distribution and the quantile boundary shifts can be seen in the
400 combined data sets.

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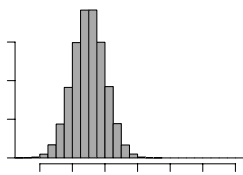
420 **Fig. 1**

421

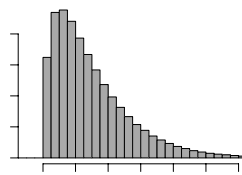
Simulation study: Distribution types

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Normal



Chi-squared



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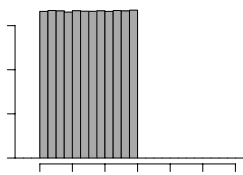
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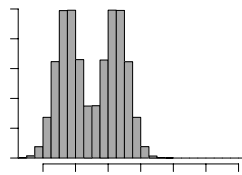
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Uniform



Bimodal



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440 Fig. 2

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