

The ALHAMBRA survey: Estimation of the clustering signal encoded in the cosmic variance[★]

C. López-Sanjuan¹, A. J. Cenarro¹, C. Hernández-Monteagudo¹, P. Arnalte-Mur², J. Varela¹, K. Viironen¹, A. Fernández-Soto^{3,4}, V. J. Martínez^{2,5,4}, E. Alfaro⁶, B. Ascaso⁷, A. del Olmo⁶, L. A. Díaz-García¹, Ll. Hurtado-Gil^{2,3}, M. Moles^{1,6}, A. Molino^{8,6}, J. Perea⁶, M. Pović⁶, J. A. L. Aguerri^{9,10}, T. Aparicio-Villegas^{11,6}, N. Benítez⁶, T. Broadhurst^{12,13}, J. Cabrera-Caño¹⁴, F. J. Castander¹⁵, J. Cepa^{9,10}, M. Cerviño^{6,9,10}, D. Cristóbal-Hornillos¹, R. M. González Delgado⁶, C. Husillos⁶, L. Infante¹⁶, I. Márquez⁶, J. Masegosa⁶, F. Prada^{6,17,18}, and J. M. Quintana⁶

¹ Centro de Estudios de Física del Cosmos de Aragón, Plaza San Juan 1, 44001 Teruel, Spain
e-mail: clsj@cefca.es

² Observatori Astronòmic, Universitat de València, C/ Catedràtic José Beltrán 2, 46980 Paterna, Spain

³ Instituto de Física de Cantabria (CSIC-UC), 39005 Santander, Spain

⁴ Unidad Asociada Observatorio Astronómico (IFCA-UV), 46980 Paterna, Spain

⁵ Departament d'Astronomia i Astrofísica, Universitat de València, 46100 Burjassot, Spain

⁶ IAA-CSIC, Glorieta de la Astronomía s/n, 18008 Granada, Spain

⁷ GEPI, Observatoire de Paris, CNRS, Université Paris Diderot, 61 avenue de l'Observatoire, 75014 Paris, France

⁸ Instituto de Astronomía, Geofísica e Ciências Atmosféricas, Universidade de São Paulo, 05508-090 São Paulo, Brazil

⁹ Instituto de Astrofísica de Canarias, Vía Láctea s/n, 38200 La Laguna, Tenerife, Spain

¹⁰ Departamento de Astrofísica, Facultad de Física, Universidad de La Laguna, 38206 La Laguna, Spain

¹¹ Observatório Nacional-MCT, Rua José Cristino 77, CEP 20921-400, Rio de Janeiro-RJ, Brazil

¹² Department of Theoretical Physics, University of the Basque Country UPV/EHU, 48080 Bilbao, Spain

¹³ IKERBASQUE, Basque Foundation for Science, 48013 Bilbao, Spain

¹⁴ Departamento de Física Atómica, Molecular y Nuclear, Facultad de Física, Universidad de Sevilla, 41012 Sevilla, Spain

¹⁵ Institut de Ciències de l'Espai (IEEC-CSIC), Facultat de Ciències, Campus UAB, 08193 Bellaterra, Spain

¹⁶ Departamento de Astronomía, Pontificia Universidad Católica, 782-0436 Santiago, Chile

¹⁷ Instituto de Física Teórica, UAM/CSIC, Universidad Autónoma de Madrid, Cantoblanco, 28049 Madrid, Spain

¹⁸ Campus of International Excellence UAM+CSIC, Cantoblanco, 28049 Madrid, Spain

Received 12 June 2015 / Accepted 1 August 2015

ABSTRACT

Aims. The relative cosmic variance (σ_v) is a fundamental source of uncertainty in pencil-beam surveys and, as a particular case of count-in-cell statistics, can be used to estimate the bias between galaxies and their underlying dark-matter distribution. Our goal is to test the significance of the clustering information encoded in the σ_v measured in the ALHAMBRA survey.

Methods. We measure the cosmic variance of several galaxy populations selected with B -band luminosity at $0.35 \leq z < 1.05$ as the intrinsic dispersion in the number density distribution derived from the 48 ALHAMBRA subfields. We compare the observational σ_v with the cosmic variance of the dark matter expected from the theory, $\sigma_{v,\text{dm}}$. This provides an estimation of the galaxy bias b .

Results. The galaxy bias from the cosmic variance is in excellent agreement with the bias estimated by two-point correlation function analysis in ALHAMBRA. This holds for different redshift bins, for red and blue subsamples, and for several B -band luminosity selections. We find that b increases with the B -band luminosity and the redshift, as expected from previous work. Moreover, red galaxies have a larger bias than blue galaxies, with a relative bias of $b_{\text{rel}} = 1.4 \pm 0.2$.

Conclusions. Our results demonstrate that the cosmic variance measured in ALHAMBRA is due to the clustering of galaxies and can be used to characterise the σ_v affecting pencil-beam surveys. In addition, it can also be used to estimate the galaxy bias b from a method independent of correlation functions.

Key words. dark matter – galaxies: statistics

1. Introduction

One fundamental uncertainty in any observational measurement derived from galaxy surveys is the relative cosmic variance

[★] Based on observations collected at the German-Spanish Astronomical Center, Calar Alto, jointly operated by the Max-Planck-Institut für Astronomie (MPIA) at Heidelberg and the Instituto de Astrofísica de Andalucía (CSIC).

(σ_v , also called sample variance), arising from the underlying large-scale density fluctuations and leading to variances larger than those expected from simple Poisson statistics. The most efficient way to tackle the cosmic variance is to split the survey into several independent areas on the sky. This minimises the sampling problem and it is better than increasing the volume in a wide contiguous field (e.g. [Driver & Robotham 2010](#)). However, the uncertainties in many existing surveys are dominated by the

cosmic variance because of the observational constraints (depth vs. area). Thus, a proper estimation of σ_v is needed to fully describe the error budget in deep cosmological surveys.

The impact of the cosmic variance in a given survey and redshift range can be estimated using two basic methods: theoretically, by analysing cosmological simulations (e.g. Somerville et al. 2004; Trenti & Stiavelli 2008; Stringer et al. 2009; Moster et al. 2011), or empirically, by sampling a larger survey (e.g. Driver & Robotham 2010; López-Sanjuan et al. 2014; Keenan et al. 2014). Unfortunately, the value of previous empirical work is limited by the precise understanding of the measured cosmic variance: the estimated σ_v must encode, by definition, the clustering of the studied populations, but no previous work in the literature has tested this requirement.

The galaxy bias b is the relationship between the spatial distribution of galaxies and the underlying dark-matter density field (Kaiser 1984; Bardeen et al. 1986; Mo & White 1996). The cosmic variance σ_v is a particular case of count-in-cell statistics (Peebles 1980; Efstathiou et al. 1990; Efstathiou 1995; Andreani et al. 1994; Adelberger et al. 1998; Dekel & Lahav 1999; Marinoni et al. 2005; Robertson 2010; Yang & Saslaw 2011; Di Porto et al. 2014), and can be therefore used to estimate b by comparison with $\sigma_{v,\text{dm}}$, the cosmic variance of the dark matter expected from the theory. The galaxy bias computed from the cosmic variance, noted b_v , encodes the clustering information of σ_v and has to be the same that the bias estimated with an independent method at similar scales. We use the bias computed from the two-point correlation function, noted b_ξ , as a benchmark because it is a well-tested method widely used in the literature. The agreement between b_v and b_ξ implies that the origin of the observational σ_v is the clustering of galaxies, as desired, while the discrepancy reflects deficiencies in our methodology to measure σ_v .

We take advantage of the unique design, depth, and photometric redshift accuracy of the Advanced, Large, Homogeneous Area, Medium-Band Redshift Astronomical (ALHAMBRA) survey¹ (Moles et al. 2008) to study the clustering signal encoded in the cosmic variance. We compare the observed cosmic variance with the expectations from non-linear theory to estimate the galaxy bias b_v . Then, we compare b_v with the b_ξ reported by Arnalte-Mur et al. (2014, AM14 hereafter) from two-point correlation function analysis using the same data set, cosmological parameters, and theoretical expectations.

The paper is organised as follows. In Sect. 2, we present the ALHAMBRA survey and its photometric redshifts and probability distribution functions, and in Sect. 3 our estimation of the cosmic variance and the galaxy bias. We compare our results with the bias estimated from correlation function analysis in Sect. 4. Finally, we present our summary and conclusions in Sect. 5. Throughout this paper we use a standard cosmology with $\Omega_m = 0.27$, $\Omega_\Lambda = 0.73$, $H_0 = 100 h \text{ km s}^{-1} \text{ Mpc}^{-1}$, and $\sigma_8 = 0.816$. To avoid systematics in our results, these cosmological parameters are the same as in AM14. Magnitudes are given in the AB system (Oke & Gunn 1983) and the absolute B -band magnitudes are given as $M_B - 5 \log_{10} h$, even when not explicitly indicated.

2. ALHAMBRA survey

The ALHAMBRA survey provides a photometric data set over 20 contiguous, equal-width ($\sim 300 \text{ \AA}$), non-overlapping, medium-band optical filters (3500–9700 \AA) plus three standard

broad-band near-infrared (NIR) filters (J , H , and K_s) over eight different regions of the northern sky (Moles et al. 2008). The survey has the aim of understanding the evolution of galaxies throughout cosmic time by sampling a large cosmological fraction of the Universe, for which reliable spectral energy distributions (SEDs) and precise photometric redshifts (z_p) are needed. The simulations of Benítez et al. (2009), which relate the image depth and the accuracy of the photometric redshifts to the number of filters, have demonstrated that the filter set chosen for ALHAMBRA can achieve a photometric redshift precision that is three times better than a classical 4–5 optical broad-band filter set. The final survey parameters and scientific goals, as well as the technical requirements of the filter set, were described by Moles et al. (2008). The survey has collected its data for the 20+3 optical-NIR filters in the 3.5 m telescope at the Calar Alto observatory, using the wide-field camera LAICA (Large Area Imager for Calar Alto) in the optical and the OMEGA–2000 camera in the NIR. The full characterisation, description, and performance of the ALHAMBRA optical photometric system were presented in Aparicio-Villegas et al. (2010). A summary of the optical reduction can be found in Cristóbal-Hornillos et al. (in prep.), the NIR reduction is reported in Cristóbal-Hornillos et al. (2009).

The ALHAMBRA survey has observed eight well-separated regions of the sky. The wide-field camera LAICA has four chips, each with a $15' \times 15'$ field of view ($0.22 \text{ arcsec pixel}^{-1}$). The separation between chips is $13'$. Thus, each LAICA pointing provides four distinct areas in the sky, one per chip. Six ALHAMBRA regions comprise two LAICA pointings. In these cases, the pointings define two separate strips in the sky. We assumed the four chips in each LAICA pointing to be independent subfields. We summarise the properties of the seven fields included in the first ALHAMBRA data release in Table 1. Currently, ALHAMBRA comprises 48 subfields of $\sim 183.5 \text{ arcmin}^2$, which can be assumed to be independent for cosmic variance studies as demonstrated by López-Sanjuan et al. (2014).

2.1. Bayesian photometric redshifts in ALHAMBRA

The photometric redshifts used throughout were fully presented and tested in Molino et al. (2014), and we summarise their principal characteristics below.

The photometric redshifts of ALHAMBRA were estimated with BPZ2.0, a new version of the Bayesian Photometric Redshift (BPZ, Benítez 2000) estimator. This is a SED-fitting method based on a Bayesian inference, where a maximum likelihood is weighted by a prior probability. ALHAMBRA relied on the ColorPro software (Coe et al. 2006) to perform point spread function (PSF) matched aperture-corrected photometry, which provided both total magnitudes and isophotal colours for the galaxies. In addition, a homogeneous photometric zero-point recalibration was performed using either spectroscopic redshifts (when available) or accurate photometric redshifts from emission-line galaxies (Molino et al. 2014). Sources were detected in a synthetic $F814W$ filter image, noted I in the following, defined to resemble the HST/ $F814W$ filter. The areas of the images affected by bright stars and those with lower exposure times (e.g. the edges of the images) were masked following Arnalte-Mur et al. (2014). The total area covered by the current ALHAMBRA data after masking is 2.38 deg^2 (Table 1). Finally, a statistical star/galaxy separation was encoded in the variable `Stellar_Flag` of the ALHAMBRA catalogues, and we kept ALHAMBRA sources with `Stellar_Flag` ≤ 0.5 as galaxies.

¹ <http://www.alhambrasurvey.com>

Table 1. ALHAMBRA survey fields.

Field name	Overlapping survey	RA (J2000)	Dec (J2000)	Subfields/area (no./deg ²)
ALHAMBRA-2	DEEP2 (Newman et al. 2013)	02 28 32.0	+00 47 00	8/0.377
ALHAMBRA-3	SDSS (Aihara et al. 2011)	09 16 20.0	+46 02 20	8/0.404
ALHAMBRA-4	COSMOS (Scoville et al. 2007)	10 00 00.0	+02 05 11	4/0.203
ALHAMBRA-5	GOODS-N (Giavalisco et al. 2004)	12 35 00.0	+61 57 00	4/0.216
ALHAMBRA-6	AEGIS (Davis et al. 2007)	14 16 38.0	+52 24 50	8/0.400
ALHAMBRA-7	ELAIS-N1 (Rowan-Robinson et al. 2004)	16 12 10.0	+54 30 15	8/0.406
ALHAMBRA-8	SDSS (Aihara et al. 2011)	23 45 50.0	+15 35 05	8/0.375
Total				48/2.381

The photometric redshift accuracy, as estimated by comparison with ~ 7200 spectroscopic redshifts (z_s), was encoded in the normalised median absolute deviation (NMAD) of the photometric vs. spectroscopic redshift distribution (Ilbert et al. 2006; Brammer et al. 2008),

$$\sigma_{\text{NMAD}} = 1.48 \times \left\langle \left| \frac{\delta_z - \langle \delta_z \rangle}{1 + z_s} \right| \right\rangle, \quad (1)$$

where $\delta_z = z_p - z_s$ and $\langle \cdot \rangle$ is the median operator. The fraction of catastrophic outliers η is defined as the fraction of galaxies with $|\delta_z|/(1 + z_s) > 0.2$. In the case of ALHAMBRA, $\sigma_{\text{NMAD}} = 0.011$ for $I \leq 22.5$ galaxies with a fraction of catastrophic outliers of $\eta = 2.1\%$. We refer to Molino et al. (2014) for a more detailed discussion of the ALHAMBRA photometric redshifts.

2.2. Probability distribution functions in ALHAMBRA

This section is devoted to the description of the probability distribution functions (PDFs) of the ALHAMBRA sources. When dealing with photometric redshifts, several studies show that it is better to use the full PDF than to use only the best z_p (e.g. Fernández-Soto et al. 2002; Cunha et al. 2009; Wittman 2009; Myers et al. 2009; Schmidt & Thorman 2013; Carrasco Kind & Brunner 2014). The ALHAMBRA PDFs have been successfully used to study high-redshift ($z > 2$) galaxies (Viironen et al. 2015), to detect galaxy groups and clusters (Ascaso et al. 2015), to estimate the merger fraction (López-Sanjuan et al. 2015), or to improve the estimation of stellar population parameters (Díaz-García et al. 2015).

The probability that a galaxy i is located at redshift z and has a spectral type T is $\text{PDF}_i(z, T)$. This probability density function is the posterior provided by BPZ2.0. The probability that the galaxy i is located at redshift z is then

$$\text{PDF}_i(z) = \int \text{PDF}_i(z, T) dT. \quad (2)$$

Moreover, the total probability that the galaxy i is located at $z_1 \leq z < z_2$ is

$$P_i(z_1, z_2) = \int_{z_1}^{z_2} \text{PDF}_i(z) dz. \quad (3)$$

The probability density function $\text{PDF}_i(z, T)$ is normalised to one by definition; in other words, the probability of any galaxy i being found in the whole parameter space is one. Formally,

$$1 = \int \text{PDF}_i(z) dz = \iint \text{PDF}_i(z, T) dT dz. \quad (4)$$

The B -band absolute magnitude of a galaxy with observed magnitude $I = 20$ and spectral type T that is located at redshift z is

denoted as $M_B^{20}(z, T)$, which is also provided by BPZ2.0. This function condenses the information about the luminosity distance, which depends on z , and the k correction, which depends on z and T , needed to translate observed I -band magnitudes to absolute B -band magnitudes. Thus, the M_B of the galaxy i depends on its z and T , and we have to weight $M_B^{20}(z, T)$ with the corresponding PDF $_i(z, T)$ to obtain $M_{B,i}$. Formally, we estimate M_B as a function of z as

$$M_{B,i}(z) = \frac{\int M_B^{20}(z, T) \times \text{PDF}_i(z, T) dT}{\int \text{PDF}_i(z, T) dT} + (I_i - 20), \quad (5)$$

where I_i is the observed I -band magnitude of the source.

Thanks to the probability functions defined in this section, we are able to statistically use the output of current photometric redshift codes without losing information.

2.3. Sample selection

We focus our analysis on the galaxies in the ALHAMBRA first data release². This catalogue comprises ~ 500 k sources and is complete (5σ , $3''$ aperture) for $I \leq 24.5$ galaxies (Molino et al. 2014).

We perform our study in a given redshift range $z \in [z_{\min}, z_{\max})$ with B -band luminosity selected samples. Because the M_B of a galaxy depends on its redshift as shown in Sect. 2.2, we manage the galaxy samples under study with the function \mathcal{S} , defined as

$$\mathcal{S}_i(z, Q, M_B^{\text{sel}}) = \begin{cases} 1, & \text{if } M_{B,i}(z) + Q \times z \leq M_B^{\text{sel}}, \\ 0, & \text{otherwise,} \end{cases} \quad (6)$$

where $M_{B,i}(z)$ is the B -band luminosity of the galaxy i from Eq. (5), M_B^{sel} is the selection magnitude of the sample, and the term $Q \times z$ accounts for the evolution of the luminosity function with redshift. This last term ensures that we explore similar L_B/L_B^* luminosities at every z , where L_B^* is the typical B -band luminosity at a given redshift reported by Ilbert et al. (2006).

Our goal is to compare the galaxy bias derived from the cosmic variance with the bias measured from the two-point correlation function analysis performed by AM14. To minimise any possible difference between the two studies, we define the same redshift ranges and B -band luminosity selections as AM14. The studied samples, defined with $Q = 0.6$ to mimic AM14 selection, are summarised in the first two columns of Table 2, and cover a B -band luminosity from $L_B/L_B^* = 0.2$ to 2 at $0.35 \leq z < 1.05$. The L_B/L_B^* values in the table are those reported by AM14.

² <http://cloud.iaa.es/alhambra/>

3. Estimation of the galaxy bias from the cosmic variance

In this section we present the steps we follow to estimate the galaxy bias from the cosmic variance in ALHAMBRA. First, we expose the theoretical relation between the cosmic variance and b (Sect. 3.1). Then, we detail how to measure the cosmic variance from ALHAMBRA data (Sect. 3.2). Finally, we estimate the cosmic variance of the dark matter expected from theory (Sect. 3.3).

3.1. Theoretical link between σ_v and b

To connect the cosmic variance with the galaxy bias, we follow Somerville et al. (2004) and Moster et al. (2011). The mean $\langle N \rangle$ and the variance $\langle N^2 \rangle - \langle N \rangle^2$ in the distribution of galaxies are given by the first and second moments of the probability distribution $P_N(V)$, which describes the probability of counting N objects within a volume V . The relative cosmic variance is defined as

$$\sigma_v^2 = \frac{\langle N^2 \rangle - \langle N \rangle^2}{\langle N \rangle^2} - \frac{1}{\langle N \rangle}. \quad (7)$$

The second term represents the correction for the Poisson shot noise. The second moment of the object counts is

$$\langle N^2 \rangle = \langle N \rangle^2 + \langle N \rangle + \frac{\langle N \rangle^2}{V^2} \int_V \xi(|\mathbf{r}_a - \mathbf{r}_b|) dV_a dV_b, \quad (8)$$

where ξ is the two-point correlation function of the sample under study (Peebles 1980). Combining this with Eq. (7), the relative cosmic variance can be written as

$$\sigma_v^2 = \frac{1}{V^2} \int_V \xi(|\mathbf{r}_a - \mathbf{r}_b|) dV_a dV_b. \quad (9)$$

We can approximate the galaxy correlation function in Eq. (9) using the correlation function for dark matter ξ_{dm} as $\xi = b^2 \xi_{\text{dm}}$, where b is the galaxy linear bias and we assume that b does not depend on scale. With this definition of the correlation function, we find that

$$\sigma_v^2 = \frac{b^2}{V^2} \int_V \xi_{\text{dm}}(|\mathbf{r}_a - \mathbf{r}_b|) dV_a dV_b = b^2 \sigma_{v,\text{dm}}^2, \quad (10)$$

where $\sigma_{v,\text{dm}}$ is the cosmic variance of the dark matter. Finally, the galaxy bias from the cosmic variance is

$$b_v = \frac{\sigma_v}{\sigma_{v,\text{dm}}}. \quad (11)$$

We estimate σ_v from ALHAMBRA data in Sect. 3.2, and $\sigma_{v,\text{dm}}$ from theory in Sect. 3.3.

3.2. Measuring the cosmic variance σ_v in ALHAMBRA

In this section we define the procedure used to measure the cosmic variance from ALHAMBRA data. We estimate the number density of galaxies (n) in the ALHAMBRA subfield j for a given B -band luminosity selection M_B^{sel} and redshift range $z \in [z_{\text{min}}, z_{\text{max}})$ as

$$n_j = \frac{N_j}{V_j} = \frac{1}{V_j} \sum_i \int_{z_{\text{min}}}^{z_{\text{max}}} \text{PDF}_i(z) \times \mathcal{S}_i(z, 0.6, M_B^{\text{sel}}) dz, \quad (12)$$

where N_j is the number of galaxies that fulfil the selection, and V_j is the cosmic volume probed by the subfield j at

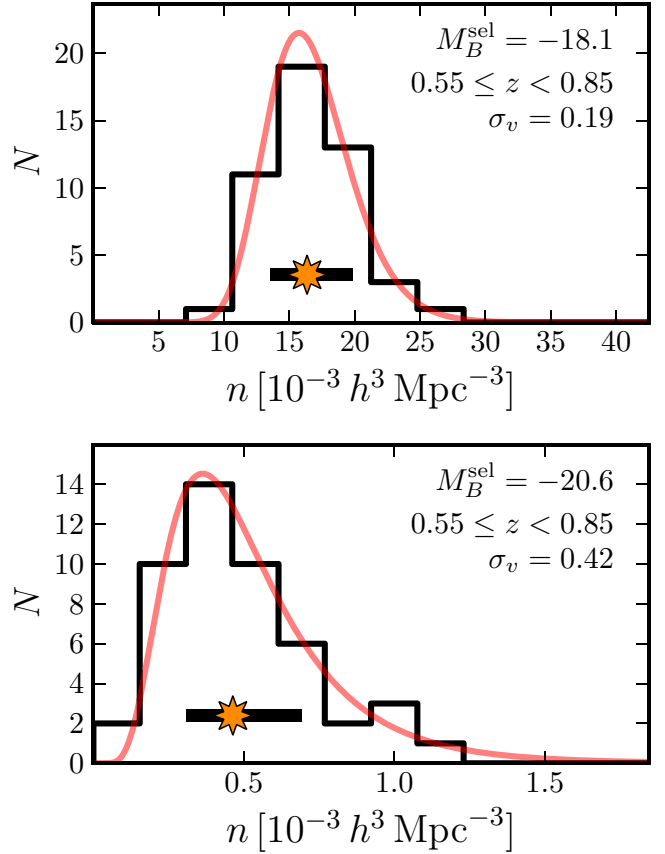


Fig. 1. Number density distribution of $M_B^{\text{sel}} = -18.1$ (top panel) and $M_B^{\text{sel}} = -20.6$ (bottom panel) galaxies at $0.55 \leq z < 0.85$ from the 48 ALHAMBRA subfields. The star and the black bar mark the median and the intrinsic dispersion retrieved by the MLE, respectively. The red solid line shows the best MLE solution convolved with the observational errors and it is independent of the histogram binning. The derived relative cosmic variance σ_v is labelled in the panels. A colour version of this plot is available in the electronic edition.

$z_{\text{min}} \leq z < z_{\text{max}}$. The index i spans all the galaxies in the subfield; that is, no pre-selection of the sources is performed. A similar probabilistic approach is used by Viironen et al. (2015) to study the number counts of high-redshift galaxies.

The 48 number densities n_j from the ALHAMBRA subfields follow a log-normal distribution,

$$P_{\text{LN}}(n | \mu, \sigma) = \frac{1}{n \sqrt{2\pi} \sigma} \exp \left[-\frac{(\ln n - \mu)^2}{2\sigma^2} \right], \quad (13)$$

where μ and σ are the median and the dispersion of a Gaussian function in log-space,

$$P_{\text{G}}(n' | \mu, \sigma) = \frac{1}{\sqrt{2\pi} \sigma} \exp \left[-\frac{(n' - \mu)^2}{2\sigma^2} \right], \quad (14)$$

where $n' = \ln n$. This log-normal distribution of the number densities was expected because the distribution of overdense structures in the Universe is described by a log-normal function (e.g. Coles & Jones 1991; de la Torre et al. 2010; Kovač et al. 2010; Yang & Saslaw 2011). We present two representative examples in Fig. 1. We checked that the number densities of all the samples under study follow a log-normal distribution with an Anderson & Darling (1952) test. This test confirms that the distribution of the number densities is compatible with a log-normal shape and always disfavours a Gaussian distribution.

Table 2. Galaxy bias estimated from the cosmic variance in ALHAMBRA.

$M_B^{\text{sel}} - 5 \log_{10} h$	L_B/L_B^*	\bar{n} ($10^{-3} h^3 \text{ Mpc}^{-3}$)	σ_v	b_v	b_ξ
(1)	(2)	(3)	(4)	(5)	(6)
$0.35 \leq z < 0.65$ $\bar{z} = 0.52$ $\sigma_{v,\text{dm}} = 0.175$					
-16.8	0.16	32.4 ± 0.8	0.174 ± 0.019	1.00 ± 0.11	1.00 ± 0.09
-17.6	0.27	20.7 ± 0.6	0.200 ± 0.023	1.14 ± 0.13	1.16 ± 0.11
-18.1	0.37	15.0 ± 0.5	0.216 ± 0.025	1.24 ± 0.14	1.23 ± 0.13
-18.6	0.50	10.1 ± 0.4	0.228 ± 0.027	1.30 ± 0.15	1.27 ± 0.12
-19.1	0.73	6.2 ± 0.2	0.251 ± 0.031	1.43 ± 0.17	1.35 ± 0.17
-19.6	1.06	3.3 ± 0.1	0.271 ± 0.036	1.55 ± 0.20	1.50 ± 0.16
-20.1	1.52	1.4 ± 0.1	0.318 ± 0.049	1.81 ± 0.28	1.79 ± 0.17
-20.6	2.40	0.42 ± 0.04	0.387 ± 0.096	2.21 ± 0.55	2.40 ± 0.50
$0.55 \leq z < 0.85$ $\bar{z} = 0.74$ $\sigma_{v,\text{dm}} = 0.148$					
-17.6	0.26	22.4 ± 0.6	0.174 ± 0.019	1.18 ± 0.13	1.22 ± 0.17
-18.1	0.35	16.3 ± 0.5	0.194 ± 0.022	1.31 ± 0.15	1.34 ± 0.21
-18.6	0.49	11.0 ± 0.4	0.215 ± 0.025	1.46 ± 0.17	1.52 ± 0.24
-19.1	0.69	6.7 ± 0.3	0.250 ± 0.029	1.69 ± 0.30	1.74 ± 0.29
-19.6	0.97	3.5 ± 0.1	0.278 ± 0.035	1.88 ± 0.24	2.00 ± 0.30
-20.1	1.44	1.5 ± 0.1	0.325 ± 0.044	2.20 ± 0.30	2.30 ± 0.40
-20.6	2.13	0.50 ± 0.03	0.422 ± 0.072	2.85 ± 0.49	3.20 ± 0.60
$0.75 \leq z < 1.05$ $\bar{z} = 0.91$ $\sigma_{v,\text{dm}} = 0.134$					
-18.1	0.34	17.6 ± 0.6	0.190 ± 0.022	1.41 ± 0.16	1.37 ± 0.15
-18.6	0.46	11.7 ± 0.3	0.198 ± 0.023	1.48 ± 0.17	1.47 ± 0.15
-19.1	0.66	7.0 ± 0.2	0.220 ± 0.026	1.64 ± 0.19	1.62 ± 0.17
-19.6	0.93	3.5 ± 0.1	0.247 ± 0.030	1.84 ± 0.22	1.90 ± 0.30
-20.1	1.35	1.4 ± 0.1	0.330 ± 0.044	2.47 ± 0.33	2.21 ± 0.23
-20.6	1.95	0.38 ± 0.03	0.427 ± 0.070	3.18 ± 0.52	2.90 ± 0.30

Notes. Column (1): absolute B -band magnitude at $z = 0$ used to select the sample with Eq. (6). Column (2): median luminosity of the sample in units of L_B^* from AM14. Column (3): median number density of the sample. Column (4): observed cosmic variance. Column (5): galaxy bias measured from the cosmic variance. Column (6): galaxy bias measured from the two-point correlation function by AM14.

There are two origins of the observed σ : the intrinsic dispersion σ_{int} (i.e. the field-to-field variation due to the clustering of the galaxies), and the dispersion due to the Poisson shot noise σ_{p} . We estimate the Poisson shot noise in each subfield as

$$\sigma_{\text{p},j} = \sqrt{N_j}/V_j \quad (15)$$

and apply the maximum likelihood estimator (MLE) presented in López-Sanjuan et al. (2014) to measure the median number density of the distribution \bar{n} , the intrinsic dispersion σ_{int} , and their uncertainties.

Applying Eq. (7) to $P_{\text{LN}}(n)$, we conclude that the observed relative cosmic variance is

$$\sigma_v^2 = e^{\sigma_{\text{int}}^2} - 1. \quad (16)$$

We estimate the σ_v uncertainty by propagating the uncertainty in σ_{int} , as estimated with the MLE (see López-Sanjuan et al. 2014, for details). We summarise the median number density \bar{n} and the measured cosmic variance σ_v of the studied samples in Cols. 3 and 4 of Table 2. We note that the average number densities reported by AM14 differ from our values. This discrepancy is explained by the number density definition: we estimate a probabilistic number density from Eq. (12) and AM14 assumed the galaxies located at their best photometric redshift.

3.3. Estimation of the dark-matter cosmic variance $\sigma_{v,\text{dm}}$

The final ingredient needed to estimate the galaxy bias b_v from Eq. (11) is $\sigma_{v,\text{dm}}$, the cosmic variance of the dark matter expected

from the theory. We calculate $\sigma_{v,\text{dm}}$ in each volume by solving the integral in Eq. (9) for dark matter using the code QUICKCV³, which is described in Newman & Davis (2002). The code computes the cosmic variance from the dark-matter power spectrum using a window function which is 1 inside the interest volume and 0 otherwise. We obtain the dark-matter power spectrum at each redshift bin using the CAMB software (Lewis et al. 2000), including the non-linear corrections of HALOFIT (Smith et al. 2003).

The cosmic variance is a particular case of count-in-cell statistics, and the procedure used to estimate the galaxy linear bias is similar if either spherical or cubic volumes are used instead (see references in Sect. 1). In the next section, we compare the galaxy bias b_v estimated from the cosmic variance with the values reported in ALHAMBRA by AM14 from two-point correlation function analysis. They estimate the galaxy bias as

$$b_\xi = \sqrt{\frac{w_{\text{p}}}{w_{\text{p},\text{dm}}}}, \quad (17)$$

where w_{p} is the measured projected correlation function and $w_{\text{p},\text{dm}}$ is the projected dark-matter correlation function expected from the theory.

³ QUICKCV is available at www.phyast.pitt.edu/~janewman/quickcv

4. Clustering signal encoded in the cosmic variance

We summarise the studied samples, the measured cosmic variance σ_v , the inferred galaxy bias from the cosmic variance b_v , and the galaxy bias b_ξ reported by AM14 from correlation function analysis in Table 2. We show the comparison between b_v and b_ξ in three redshift bins and for several B -band luminosity selections in Fig. 2.

We find that the galaxy bias from the cosmic variance nicely agrees with the values from AM14, with all the estimations consistent at the 1σ level. This result has two important implications. First, it confirms that the cosmic variance measured in ALHAMBRA is due to the clustering of galaxies and the methodology applied to measure σ_v is therefore correct. Thus, the ALHAMBRA σ_v can be used to characterise the cosmic variance affecting pencil-beam surveys, reinforcing the results and the parametrisation of σ_v for close-pair studies presented by López-Sanjuan et al. (2014). Second, we can estimate the galaxy bias b from the cosmic variance, complementing the bias measurements from correlation function analysis with an independent method. The σ_v is measured in large volumes defined by the observational strategy of the survey, and it is unaffected by the photometric redshift precision and peculiar motions as the counts-in-cells performed with spherical or cubic volumes (Papageorgiou et al. 2012).

Regarding the observed trends, we find that b_v increases with the B -band luminosity of the sample and, for a fixed selection, increases with redshift. These trends have already been found in several studies (e.g. Benoist et al. 1996; Norberg et al. 2001; Pollo et al. 2006; Coil et al. 2006; Coupon et al. 2012; Skibba et al. 2014; AM14). The quantitative comparison with the galaxy bias measured by other surveys is beyond the scope of this paper. Because of the excellent agreement with the AM14 values, we refer the reader to AM14 for a detailed discussion of the galaxy bias measurements in ALHAMBRA and a comparison with previous work in the literature.

In the next sections we explore the impact on the b_v vs. b_ξ comparison of the ALHAMBRA fields explored (Sect. 4.1), the power-spectrum model used (Sect. 4.2), the probed scale (Sect. 4.3), and the colour of the galaxies (Sect. 4.4).

4.1. Dependence on the explored fields

Arnalte-Mur et al. (2014) conclude that the fields ALHAMBRA-4 (COSMOS) and ALHAMBRA-7 (ELAIS-N1) have significantly different clustering properties from the other ALHAMBRA fields. Thus, in addition to the global measurements presented in the previous section, AM14 also reported the b_ξ estimated without using the fields ALHAMBRA-4 and ALHAMBRA-7. To further check the ALHAMBRA cosmic variance, we also estimated the galaxy bias b_v without the 12 subfields corresponding to the ALHAMBRA-4 and ALHAMBRA-7 fields. The results are summarised in Fig. 3. As in the global case, the agreement between both estimations of the galaxy bias is remarkable and further support that the clustering measured by the correlation function is encoded in the estimated cosmic variance, as desired.

4.2. Dependence on the power-spectrum model

The power spectrum that we used to estimate $\sigma_{v, \text{dm}}$ accounts for the non-linear evolution of the dark matter. The popular work of Moster et al. (2011) also uses QUICKCV but with a linear power spectrum. We checked that the values from Moster et al. (2011)

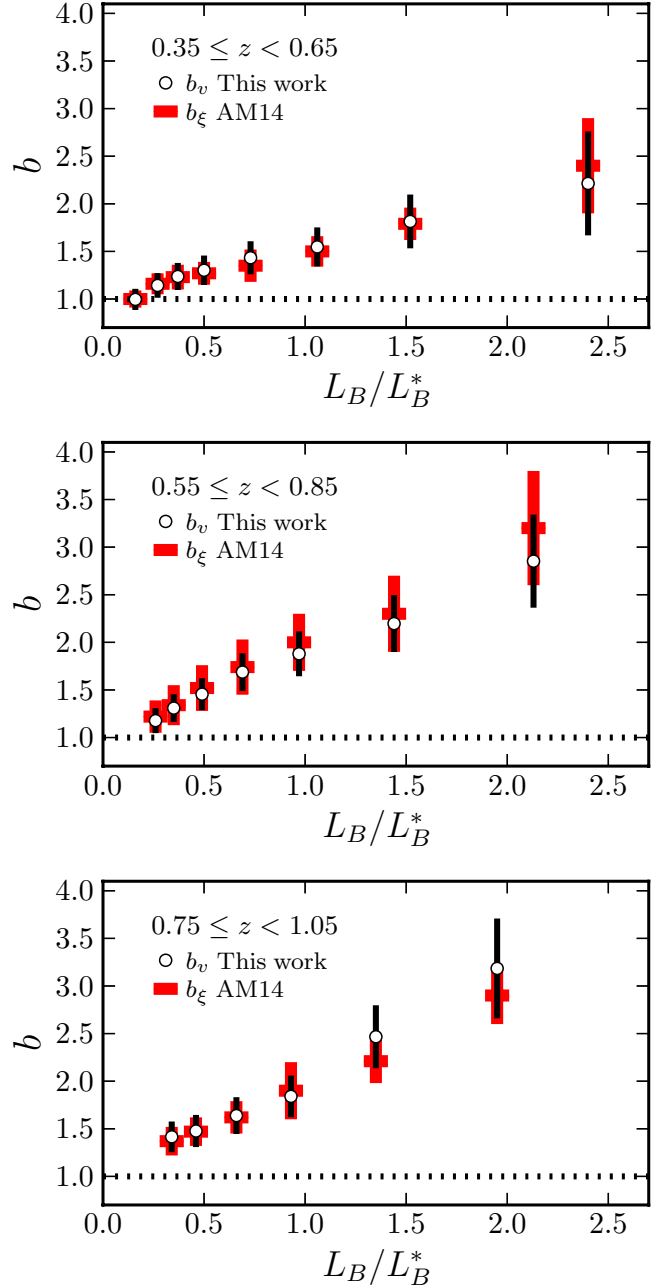


Fig. 2. Galaxy bias b as a function of the B -band luminosity in three redshift bins (labelled in the panels). The dots are from this work by cosmic variance analysis, b_v . The red crosses are from Arnalte-Mur et al. (2014) from the two-point correlation function analysis, b_ξ . The dotted line marks unity. A colour version of this plot is available in the electronic edition.

are recovered if linear theory is used, and we estimated the impact of non-linear evolution in the theoretical models. We find that the $\sigma_{v, \text{dm}}$ from non-linear theory is larger by 15%, 9%, and 6% than the cosmic variance from linear theory reported by Moster et al. (2011) at $\bar{z} = 0.52, 0.74,$ and 0.91 , respectively. These differences are below the 20% accuracy targeted by Moster et al. (2011) and decrease with redshift for two main reasons: first, the impact of the non-linear evolution increases with the growth of the structures, and therefore decreases with z . Second, the non-linear effects are important at scales smaller than $\sim 1 h^{-1}$ Mpc. The scale probed by each ALHAMBRA

subfield increases with z , decreasing the impact of the non-linear evolution.

4.3. Dependence on the probed scale

Throughout this paper, we assume that the galaxy bias b does not depend on scale, but b is a scale-dependent parameter (e.g. Heavens et al. 1998; Mann et al. 1998; Cresswell & Percival 2009). The observational studies find that the bias is nearly constant for scales larger than $r \sim 1 h^{-1}$ Mpc (the linear bias regime), with larger b values at smaller scales (e.g. White et al. 2011; Parejko et al. 2013). The values of b_ξ reported by AM14 are computed in the linear bias regime, from $1 h^{-1}$ Mpc to $10 h^{-1}$ Mpc. In the cosmic variance case, we integrate all the scales inside the volume of interest, and the measurements of b_v are affected by scales smaller than $r = 1 h^{-1}$ Mpc. The integration in Eq. (9) weights each scale by r^2 , so the impact of the small scales decreases with the area subtended by the ALHAMBRA subfields. This subtended area increases from $\sim 3.6 \times 3.6 h^{-2}$ Mpc² at $\bar{z} = 0.52$ to $\sim 4.5 \times 4.5 h^{-2}$ Mpc² at $\bar{z} = 0.91$. We find that b_v and b_ξ are similar, implying that b_v is dominated by the linear bias regime and that the ALHAMBRA subfields are large enough to reduce the scale dependence of the bias.

We further test the impact of small scales with a toy model. In our model, we assumed a scale-dependent bias of the form

$$b(r) = \begin{cases} 4 - 2r, & \text{if } r < 1 h^{-1} \text{ Mpc,} \\ 2, & \text{if } r \geq 1 h^{-1} \text{ Mpc,} \end{cases} \quad (18)$$

as suggested by observations (White et al. 2011; Parejko et al. 2013). We applied this bias model to the non-linear power spectrum and computed the cosmic variance of this biased power spectrum, noted σ_v^{model} . The comparison between σ_v^{model} and $\sigma_{v,\text{dm}}$ provides an effective bias b_{eff} that should be close to $b_{\text{eff}} = 2$ if the signal is dominated by the linear bias regime. We find $b_{\text{eff}} = 2.2, 2.1$, and 2.1 at $\bar{z} = 0.52, 0.74$, and 0.91 , respectively. This confirms that small scales should affect our measurements by less than 10%, supporting that b_v is comparable to the linear bias measured with b_ξ given the geometry of the ALHAMBRA survey.

4.4. Dependence on the colour

Following the work of AM14, Hurtado-Gil et al. (2015) study the two-point correlation function and the galaxy bias of red and blue galaxies with $M_B^{\text{sel}} \leq -18.6$ in the ALHAMBRA survey. To complete the comparison between b_v and b_ξ , we also estimated the galaxy bias of red and blue galaxies with $M_B^{\text{sel}} \leq -18.6$ from the cosmic variance.

To define red and blue galaxies, we take advantage of the profuse information encoded in the PDFs. Instead of selecting galaxies according to their observed colour or their best spectral template, we split each PDF into red templates ($T = \text{E/S0}$), denoted PDF^{red}, and blue templates ($T = \text{S/SB}$), denoted PDF^{blue}. Formally,

$$\begin{aligned} \text{PDF}(z) &= \text{PDF}^{\text{red}}(z) + \text{PDF}^{\text{blue}}(z) \\ &= \int \text{PDF}(z, \text{E/S0}) dT + \int \text{PDF}(z, \text{S/SB}) dT. \end{aligned} \quad (19)$$

In practice, the red templates have $T \in [1, 5.5]$ and the blue templates have $T \in (5.5, 11]$ in the ALHAMBRA catalogues. Thus, we can reliably work with colour segregations without any pre-selection of the sources (e.g. López-Sanjuan et al. 2015).

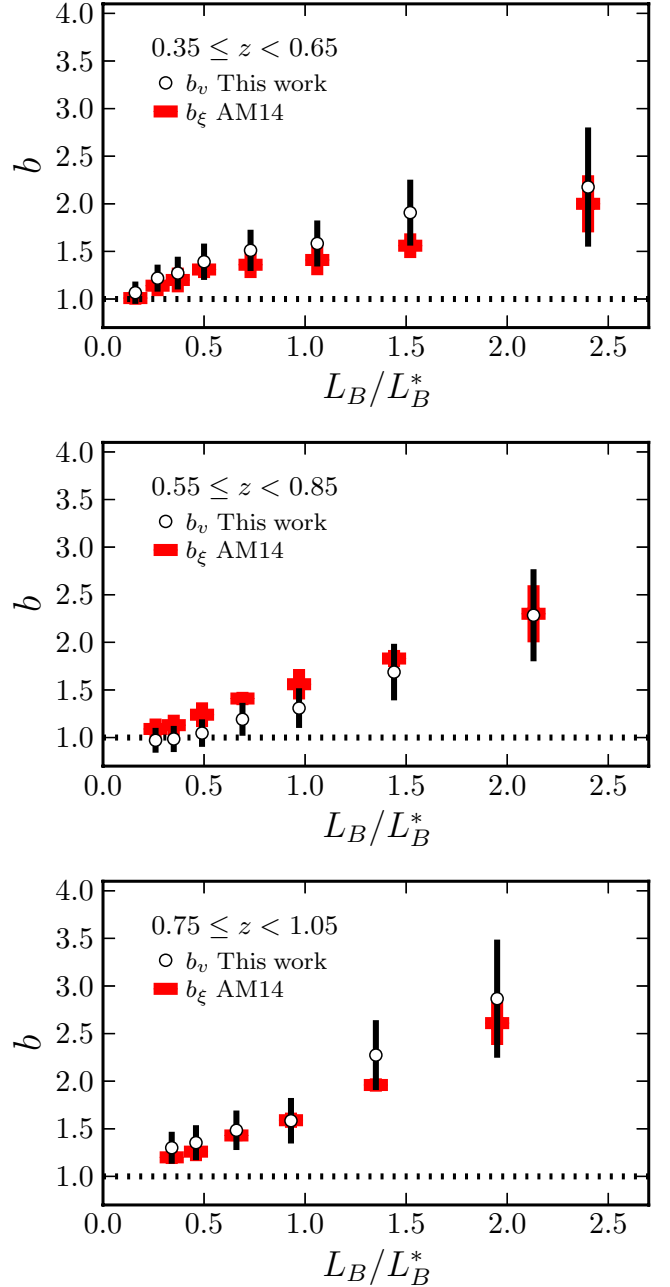


Fig. 3. Galaxy bias b as a function of the B -band luminosity in three redshift bins (labelled in the panels) with the 12 subfields from ALHAMBRA-4 and ALHAMBRA-7 removed. The dots are from this work by cosmic variance analysis, b_v . The red crosses are from Arnalte-Mur et al. (2014) by two-point correlation function analysis, b_ξ . The dotted line marks unity. A colour version of this plot is available in the electronic edition.

We summarise the basic properties of the samples under study, the measured σ_v and b_v , and the b_ξ reported by Hurtado-Gil et al. (2015) in Table 3. We note that the subfields from ALHAMBRA-4 are not included in this analysis because Hurtado-Gil et al. (2015) discard this field owing to its peculiar clustering properties. As in the general case, b_v agrees with b_ξ at the 1σ level both for red and blue galaxies, reinforcing our results and conclusions (Fig. 4). We find that red galaxies are more biased than blue galaxies, with b_v being compatible with a constant bias in the redshift range under study. The error-weighted averages at $0.35 \leq z < 1.0$ are $b_v^{\text{red}} = 1.69 \pm 0.14$

Table 3. Galaxy bias of galaxies with $M_B^{\text{sel}} \leq -18.6$ as a function of colour estimated from the cosmic variance in ALHAMBRA.

Redshift range (1)	\bar{n} ($10^{-3} h^3 \text{ Mpc}^{-3}$) (2)	σ_v (3)	$\sigma_{v,\text{dm}}$ (4)	b_v (5)	b_ξ (6)
Red templates (E/S0)					
$0.35 \leq z < 0.6$	3.4 ± 0.2	0.349 ± 0.046	0.195	1.84 ± 0.26	1.69 ± 0.08
$0.6 \leq z < 0.8$	3.0 ± 0.2	0.317 ± 0.045	0.183	1.78 ± 0.26	1.55 ± 0.13
$0.8 \leq z < 1.0$	2.7 ± 0.2	0.244 ± 0.033	0.163	1.52 ± 0.21	1.83 ± 0.11
Blue templates (S/SB)					
$0.35 \leq z < 0.6$	7.2 ± 0.3	0.229 ± 0.032	0.195	1.19 ± 0.17	1.18 ± 0.06
$0.6 \leq z < 0.8$	7.4 ± 0.3	0.224 ± 0.031	0.183	1.24 ± 0.18	1.13 ± 0.07
$0.8 \leq z < 1.0$	8.4 ± 0.3	0.190 ± 0.024	0.163	1.18 ± 0.15	1.10 ± 0.20

Notes. Column (1): redshift range under study. Column (2): median number density of the sample. Column (3): observed cosmic variance. Column (4): dark-matter cosmic variance from the theory. Column (5): galaxy bias measured from the cosmic variance. Column (6): galaxy bias measured from the two-point correlation function by [Hurtado-Gil et al. \(2015\)](#).

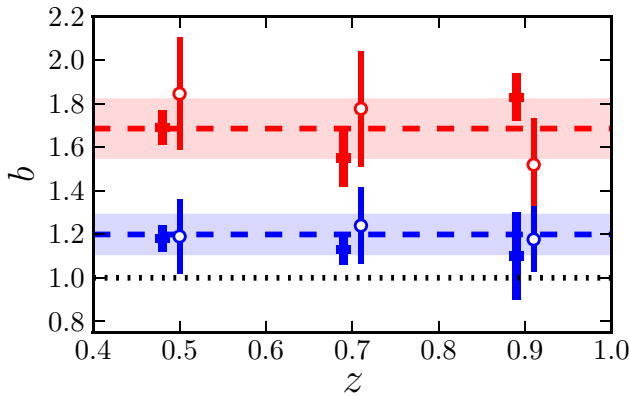


Fig. 4. Galaxy bias b of E/S0 (red symbols) and S/SB (blue symbols) galaxies with $M_B^{\text{sel}} \leq -18.6$ as a function of redshift. The dots are from this work by cosmic variance analysis, b_v . The crosses are from [Hurtado-Gil et al. \(2015\)](#) by two-point correlation function analysis, b_ξ . The dashed lines and the coloured areas mark the error-weighted average of b_v and its error, respectively, for red ($b_v^{\text{red}} = 1.69 \pm 0.14$) and blue ($b_v^{\text{blue}} = 1.20 \pm 0.10$) galaxies. The dotted line marks unity. A colour version of this plot is available in the electronic edition.

and $b_v^{\text{blue}} = 1.20 \pm 0.10$. This implies a relative bias of $b_{\text{rel}} = b_v^{\text{red}}/b_v^{\text{blue}} = 1.4 \pm 0.2$, in agreement with previous work that find $b_{\text{rel}} \sim 1.3 - 1.7$ at $z < 1$ (e.g. [Madgwick et al. 2003](#); [Meneux et al. 2006](#); [Coil et al. 2008](#); [de la Torre et al. 2011](#); [Skibba et al. 2014](#); [Hurtado-Gil et al. 2015](#)).

5. Summary and conclusions

We estimate the significance of the clustering signal encoded in the cosmic variance σ_v measured in the ALHAMBRA survey. With this aim, we measured the galaxy bias b from the ALHAMBRA cosmic variance. The cosmic variance is defined as the intrinsic dispersion (i.e. due to the field-to-field variations) in the number density distribution of the 48 ALHAMBRA sub-fields. The number densities were computed using the full probability distribution functions (PDFs) of the ALHAMBRA photometric redshifts. We compared the observed σ_v with the cosmic variance of the dark matter expected from non-linear theory to estimate the galaxy bias b .

The galaxy bias from the cosmic variance nicely agrees with the bias estimated by AM14 in ALHAMBRA from two-point

correlation function analysis. This confirms that the cosmic variance measured in ALHAMBRA is due to the clustering of galaxies and the methodology applied to measure σ_v is therefore correct. Thus, the ALHAMBRA σ_v can be used to characterise the cosmic variance affecting pencil-beam surveys, reinforcing the results and the parametrisation of the cosmic variance for close-pair studies presented by [López-Sanjuan et al. \(2014\)](#). We find that the bias increases with the B -band luminosity of the sample and, for a fixed selection, increases with redshift, in agreement with previous studies (e.g. [Norberg et al. 2001](#); [Coupon et al. 2012](#); AM14). Moreover, red galaxies have a larger bias than blue galaxies, with a relative bias of $b_{\text{rel}} = 1.4 \pm 0.2$ (e.g. [Madgwick et al. 2003](#); [Meneux et al. 2006](#); [Skibba et al. 2014](#)).

The technique outlined in this paper can be used to estimate the galaxy bias b of high-redshift galaxies and of different galaxy populations at $z < 1$ (Sect. 4.4) simply from the dispersion of the observed number densities, complementing the bias measurements from correlation function analysis with an independent method. These number densities are measured in large volumes defined by the observational strategy of the survey rather than in the common spherical or cubic volumes, which are affected by the photometric redshift precision and peculiar motions ([Papageorgiou et al. 2012](#)). In addition to its cosmological and observational value, the study of the number density distribution can be useful in order to detect systematics and photometric inhomogeneities in the next generation of large photometric surveys, such as the Dark Energy Survey (DES, [Flaugher 2012](#)), the Javalambre-Physics of the accelerating Universe Astrophysical Survey (J-PAS, [Benítez et al. 2014](#)), and the Large Synoptic Survey Telescope (LSST, [Ivezic et al. 2008](#)).

Acknowledgements. We dedicate this paper to the memory of our six IAC colleagues and friends who met with a fatal accident in Piedra de los Cochinos, Tenerife, in February 2007, with special thanks to Maurizio Panniello, whose teachings of python were so important for this paper. We thank R. Angulo, S. Bonoli, A. Ederoclite, A. Orsi, and all the CEFA staff for useful and productive discussions. We thank the anonymous referee for his/her suggestions, and K. Xu for the pertinent comments that motivated this work. This work has been mainly funded by the FITE (Fondos de Inversiones de Teruel) and the projects AYA2012-30789, AYA2006-14056, and CSD2007-00060. We also acknowledge support from the Spanish Ministry for Economy and Competitiveness and FEDER funds through grants AYA2010-15081, AYA2010-15169, AYA2010-22111-C03-01, AYA2010-22111-C03-02, AYA2011-29517-C03-01, AYA2012-39620, AYA2013-40611-P, AYA2013-42227-P, AYA2013-43188-P, AYA2013-48623-C2-1, AYA2013-48623-C2-2, ESP2013-48274, AYA2014-58861-C3-1, Aragón Government

Research Group E103, Generalitat Valenciana projects Prometeo 2009/064 and PROMETEOII/2014/060, Junta de Andalucía grants TIC114, JA2828, P10-FQM-6444, and Generalitat de Catalunya project SGR-1398. A.J.C. and C.H.-M. are Ramón y Cajal fellows of the Spanish government. A. M. acknowledges the financial support of the Brazilian funding agency FAPESP (Post-doc fellowship - process number 2014/11806-9). M.P. acknowledges financial support from JAE-Doc program of the Spanish National Research Council (CSIC), co-funded by the European Social Fund. This research made use of *Astropy*, a community-developed core Python package for Astronomy (*Astropy Collaboration et al. 2013*), and *matplotlib*, a 2D graphics package used for Python for publication-quality image generation across user interfaces and operating systems (*Hunter 2007*).

References

- Adelberger, K. L., Steidel, C. C., Giavalisco, M., et al. 1998, *ApJ*, **505**, 18
- Aihara, H., Allende Prieto, C., An, D., et al. 2011, *ApJS*, **193**, 29
- Anderson, T. W., & Darling, D. A. 1952, *Ann. Math. Statist.*, **23**, 193
- Andreani, P., Cristiani, S., Lucchin, F., Matarrese, S., & Moscardini, L. 1994, *ApJ*, **430**, 458
- Aparicio-Villegas, T., Alfaro, E. J., Cabrera-Caño, J., et al. 2010, *AJ*, **139**, 1242
- Arnalte-Mur, P., Martínez, V. J., Norberg, P., et al. 2014, *MNRAS*, **441**, 1783
- Ascaso, B., Benítez, N., Fernández-Soto, A., et al. 2015, *MNRAS*, **452**, 549
- Astropy Collaboration, Robitaille, T. P., Tollerud, E. J., et al. 2013, *A&A*, **558**, A33
- Bardeen, J. M., Bond, J. R., Kaiser, N., & Szalay, A. S. 1986, *ApJ*, **304**, 15
- Benítez, N. 2000, *ApJ*, **536**, 571
- Benítez, N., Moles, M., Aguerri, J. A. L., et al. 2009, *ApJ*, **692**, L5
- Benítez, N., Dupke, R., Moles, M., et al. 2014, in *Highlights of Spanish Astrophysics VIII, Proc. XI Scientific Meeting of the Spanish Astronomical Society*, eds. A. J. Cenarro, F. Figueras, C. Hernández-Monteagudo, J. Trujillo Bueno, & L. Valdivielso 148
- Benoist, C., Maurogordato, S., da Costa, L. N., Cappi, A., & Schaeffer, R. 1996, *ApJ*, **472**, 452
- Brammer, G. B., van Dokkum, P. G., & Coppi, P. 2008, *ApJ*, **686**, 1503
- Carrasco Kind, M., & Brunner, R. J. 2014, *MNRAS*, **442**, 3380
- Coe, D., Benítez, N., Sánchez, S. F., et al. 2006, *AJ*, **132**, 926
- Coil, A. L., Newman, J. A., Cooper, M. C., et al. 2006, *ApJ*, **644**, 671
- Coil, A. L., Newman, J. A., Croton, D., et al. 2008, *ApJ*, **672**, 153
- Coles, P., & Jones, B. 1991, *MNRAS*, **248**, 1
- Coupon, J., Kilbinger, M., McCracken, H. J., et al. 2012, *A&A*, **542**, A5
- Cresswell, J. G., & Percival, W. J. 2009, *MNRAS*, **392**, 682
- Cristóbal-Hornillos, D., Aguerri, J. A. L., Moles, M., et al. 2009, *ApJ*, **696**, 1554
- Cunha, C. E., Lima, M., Oyaizu, H., Frieman, J., & Lin, H. 2009, *MNRAS*, **396**, 2379
- Davis, M., Guhathakurta, P., Konidaris, N. P., et al. 2007, *ApJ*, **660**, L1
- de la Torre, S., Guzzo, L., Kovač, K., et al. 2010, *MNRAS*, **409**, 867
- de la Torre, S., Le Fèvre, O., Porciani, C., et al. 2011, *MNRAS*, **412**, 825
- Dekel, A., & Lahav, O. 1999, *ApJ*, **520**, 24
- Díaz-García, L. A., Cenarro, A. J., López-Sanjuan, C., et al. 2015, *A&A*, **582**, A14
- Di Porto, C., Branchini, E., Bel, J., et al. 2014, *A&A*, submitted [[arXiv:1406.6692](https://arxiv.org/abs/1406.6692)]
- Driver, S. P., & Robotham, A. S. G. 2010, *MNRAS*, **407**, 2131
- Efstathiou, G. 1995, *MNRAS*, **276**, 1425
- Efstathiou, G., Kaiser, N., Saunders, W., et al. 1990, *MNRAS*, **247**, 10P
- Fernández-Soto, A., Lanzetta, K. M., Chen, H.-W., Levine, B., & Yahata, N. 2002, *MNRAS*, **330**, 889
- Flaugher, B. 2012, in *APS April Meet. Abstr.*, D7007
- Giavalisco, M., Ferguson, H. C., Koekemoer, A. M., et al. 2004, *ApJ*, **600**, L93
- Heavens, A. F., Matarrese, S., & Verde, L. 1998, *MNRAS*, **301**, 797
- Hunter, J. D. 2007, *Comput. Sci. Eng.*, **9**, 90
- Hurtado-Gil, L., Arnalte-Mur, P., Martínez, V. J., et al. 2015, *ApJ*, submitted
- Ilbert, O., Lauger, S., Tresse, L., et al. 2006, *A&A*, **453**, 809
- Ivezic, Z., Tyson, J. A., Acosta, E., et al. 2008, *ArXiv e-prints* [[arXiv:0805.2366](https://arxiv.org/abs/0805.2366)]
- Kaiser, N. 1984, *ApJ*, **284**, L9
- Keenan, R. C., Foucaud, S., De Propris, R., et al. 2014, *ApJ*, **795**, 157
- Kovač, K., Lilly, S. J., Cucciati, O., et al. 2010, *ApJ*, **708**, 505
- Lewis, A., Challinor, A., & Lasenby, A. 2000, *ApJ*, **538**, 473
- López-Sanjuan, C., Cenarro, A. J., Hernández-Monteagudo, C., et al. 2014, *A&A*, **564**, A127
- López-Sanjuan, C., Cenarro, A. J., Varela, J., et al. 2015, *A&A*, **576**, A53
- Madgwick, D. S., Hawkins, E., Lahav, O., et al. 2003, *MNRAS*, **344**, 847
- Mann, R. G., Peacock, J. A., & Heavens, A. F. 1998, *MNRAS*, **293**, 209
- Marinoni, C., Le Fèvre, O., Meneux, B., et al. 2005, *A&A*, **442**, 801
- Meneux, B., Le Fèvre, O., Guzzo, L., et al. 2006, *A&A*, **452**, 387
- Mo, H. J., & White, S. D. M. 1996, *MNRAS*, **282**, 347
- Moles, M., Benítez, N., Aguerri, J. A. L., et al. 2008, *AJ*, **136**, 1325
- Molino, A., Benítez, N., Moles, M., et al. 2014, *MNRAS*, **441**, 2891
- Moster, B. P., Somerville, R. S., Newman, J. A., & Rix, H.-W. 2011, *ApJ*, **731**, 113
- Myers, A. D., White, M., & Ball, N. M. 2009, *MNRAS*, **399**, 2279
- Newman, J. A., & Davis, M. 2002, *ApJ*, **564**, 567
- Newman, J. A., Cooper, M. C., Davis, M., et al. 2013, *ApJS*, **208**, 5
- Norberg, P., Baugh, C. M., Hawkins, E., et al. 2001, *MNRAS*, **328**, 64
- Oke, J. B., & Gunn, J. E. 1983, *ApJ*, **266**, 713
- Papageorgiou, A., Plionis, M., Basilakos, S., & Ragone-Figueroa, C. 2012, *MNRAS*, **422**, 106
- Parejko, J. K., Sunayama, T., Padmanabhan, N., et al. 2013, *MNRAS*, **429**, 98
- Peebles, P. J. E. 1980, *The large-scale structure of the universe* (Princeton University Press)
- Pollo, A., Guzzo, L., Le Fèvre, O., et al. 2006, *A&A*, **451**, 409
- Robertson, B. E. 2010, *ApJ*, **716**, L229
- Rowan-Robinson, M., Lari, C., Perez-Fournon, I., et al. 2004, *MNRAS*, **351**, 1290
- Schmidt, S. J., & Thorman, P. 2013, *MNRAS*, **431**, 2766
- Scoville, N., Aussel, H., Brusa, M., et al. 2007, *ApJS*, **172**, 1
- Skibba, R. A., Smith, M. S. M., Coil, A. L., et al. 2014, *ApJ*, **784**, 128
- Smith, R. E., Peacock, J. A., Jenkins, A., et al. 2003, *MNRAS*, **341**, 1311
- Somerville, R. S., Lee, K., Ferguson, H. C., et al. 2004, *ApJ*, **600**, L171
- Stringer, M. J., Benson, A. J., Bundy, K., Ellis, R. S., & Quetin, E. L. 2009, *MNRAS*, **393**, 1127
- Trenti, M., & Stiavelli, M. 2008, *ApJ*, **676**, 767
- Viiroinen, K., Marin-Franch, A., López-Sanjuan, C., et al. 2015, *A&A*, **576**, A25
- White, M., Blanton, M., Bolton, A., et al. 2011, *ApJ*, **728**, 126
- Wittman, D. 2009, *ApJ*, **700**, L174
- Yang, A., & Saslaw, W. C. 2011, *ApJ*, **729**, 123