- 1 Title page
- 2 Title: Associations among eating behaviour traits, diet quality and food labelling: A mediation model
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- 31 participated in data interpretation, critically reviewed the manuscript and approved the final version.
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- 34 the Laval University ethics committee. Written informed consent was obtained from all subjects.

35 Abstract

Objectives: The aims of this study were to assess the associations among eating behaviour traits, food label use and diet quality and to evaluate if the association between eating behaviour traits and diet quality is mediated by food label use.

Design: Eating behaviour traits were assessed using the Three-Factor Eating Questionnaire (TFEQ), the Restraint Scale and the Intuitive Eating Scale whereas food label use was measured with the Label Reading Survey. Diet quality (Canadian Healthy Eating Index) was assessed with a food frequency questionnaire.

43 Setting: Cross-sectional study

Subjects: Three-hundred eighty-five adults (BMI= 26.0 ± 4.9 kg/m², age= 41.1 ± 15.0 years) involved in two previous experimental studies.

Results: When controlling for potential covariates, general food label use (β =1.18±0.26, *P*<0.0001) was the main determinant of diet quality, explaining 6.7% of its variance. General food label use partly mediated the association between TFEQ-cognitive restraint and diet quality and the indirect effect was stronger in men ($\beta_{indirect}$ =0.32±0.10, 95%CI (0.15, 0.55)) than in women ($\beta_{indirect}$ =0.16±0.05, 95%CI (0.08, 0.27)). General food label use also partly mediated the negative association between unconditional permission to eat and diet quality and the indirect effect was also stronger in men ($\beta_{indirect}$ =-1.88±0.55, 95%CI (-3.11, -0.96)) compared with women ($\beta_{indirect}$ =-1.03±0.33, 95%CI (-1.81, -0.49)).

53 **Conclusions:** General food label use was the main determinant of diet quality and partly mediated the 54 association between eating behaviour traits and diet quality. The stronger mediating effect observed in 55 men suggest that they rely more on food labeling when attempting to restrained themselves, which 56 translate in a better diet quality.

57 Keywords: Eating behavior traits, Food label use, Diet Quality, Intuitive Eating, Restrained Eating

58 Introduction

Many factors are involved in the etiology of obesity, including behavioural and psychological factors. 59 Among these, eating behaviour traits that have been widely studied in association with body weight are 60 cognitive restraint, disinhibition and susceptibility to hunger^(1,2). Dietary habits are also involved in 61 weight management. Improvements in diet quality have indeed been associated with a lower weight gain 62 over a 20-year period⁽³⁾. One way that eating behaviours can influence body weight is through diet 63 quality, which can impact energy intake. Accordingly, cognitive restraint, defined as the intent to restrain 64 food intake in order to control body weight, has been associated with a higher intake of healthy foods 65 such as green vegetables⁽⁴⁾. Moreover, flexible control, a more relaxed or graduated approach towards 66 eating, dieting and weight, has been associated with a better diet quality⁽⁵⁾. Disinhibition, defined by a 67 loss of control over eating, has been associated with a higher intake of energy-dense foods⁽²⁾. 68 69 Susceptibility to hunger, which refers to the susceptibility to feel hungry triggered by internal or external cues, is strongly associated with disinhibition^(6,7) and has been positively associated with energy intake⁽⁸⁾. 70 71 Intuitive eating, an eating style that relies on hunger and satiety cues to determine when, what and how much to eat⁽⁹⁾, showed a very weak, but positive association with vegetable intake⁽¹⁰⁾ and a weak and 72 positive association with self-reported food diversity⁽¹¹⁾, although research regarding this eating 73 behaviour is more limited and no association with dietary intakes has also been reported⁽¹²⁾. Moreover, 74 75 gender differences have been observed in these eating behaviour traits. Women generally have higher levels of cognitive restraint and disinhibition than men⁽¹³⁻¹⁵⁾. Gender difference for susceptibility to 76 hunger is less clear, as studies observed either no difference between men and women^(13,15) or that women 77 present a lower⁽¹⁴⁾ level of susceptibility to hunger than men. Finally, a higher level of intuitive eating 78 has been observed in men compared to women^(11,16,17). 79

80 In addition to eating behaviour traits, food labelling, which represents a primary source of nutrition information, may be another factor influencing diet quality. Accordingly, food labelling has been 81 proposed as a tool to help individuals make better and informed food choices⁽¹⁸⁾ and it has been reported 82 that food label use was associated with a better diet quality^(19,20). Studies generally show that women 83 report using food labels more frequently than men^(20,21) and they are more likely to report that food labels 84 influenced their food choices⁽²⁰⁾. Despite this beneficial effect of food labels on food choices and diet 85 quality, several studies have also shown that food labelling may be confusing for some individuals^(18,20) 86 and their use does not always translate into healthier food choices or eating habits⁽²²⁾. These conflicting 87 88 results may be explained by different uses of food labelling among individuals presenting diverse eating

behaviour traits. For instance, restrained individuals may be more receptive to food labels, since nutrition 89 information could be viewed for them as salient cues to support dieting rules, which may not always be 90 in accordance to healthy eating patterns⁽²²⁾. Indeed, a greater use of the nutrition fact table has been 91 associated with an increased likelihood of engaging in both healthy and unhealthy weight control 92 behaviours⁽²³⁾ and individuals attempting to control their body weight have also reported a greater use of 93 food labels^(20,21). Consistent with this previous result, Christoph et al. recently showed that while nutrition 94 fact use was unrelated to intuitive eating among young women, it was associated with a lower level of 95 intuitive eating in young men⁽²³⁾. This study also observed that a greater nutrition fact use in women was 96 associated with a greater likelihood of engaging in binge eating⁽²³⁾, an eating disorder that has been 97 positively associated with disinhibition⁽²⁾. To our knowledge, Christoph et al. was the first study to 98 specifically assess the associations between one of the specific eating behaviour traits presented above, 99 i.e., intuitive eating, and the frequency of food label use⁽²³⁾ and no study has yet assessed the associations 100 between the other eating behaviour traits presented above and food label use. While the need to better 101 understand how individuals that may be at risk for disordered eating use food labels was recently 102 emphasized⁽²³⁾, no study has yet examined the associations among eating behaviour traits, food label use 103 and diet quality in a mediation model which allows to identify the indirect effects by which eating 104 behaviour traits are associated with diet quality. 105

106 The primary aim of this study was thus to assess the associations among eating behaviour traits (i.e., cognitive restraint, disinhibition, susceptibility to hunger and intuitive eating), food label use and diet 107 quality in men and women. A second aim was to evaluate if the associations between eating behaviour 108 traits and diet quality are mediated by food label use. Based on the previous but limited literature, three 109 hypotheses were stated: 1) cognitive restraint and intuitive eating are positively associated with diet 110 quality and conversely, disinhibition is negatively associated with diet quality while susceptibility to 111 hunger is not associated with diet quality; 2) cognitive restraint, disinhibition and susceptibility to hunger 112 are positively associated with food label use whereas intuitive eating showed a negative association with 113 food label use; and 3) the use of food labels may partly mediate the association between cognitive 114 restraint and diet quality. The first hypothesis is confirmatory but is a previous step for the other two 115 hypotheses which are exploratory, except for the associations between intuitive eating and disinhibition 116 with food label used since the association has been previously observed or a similar behaviour has been 117 associated with food label use, respectively. 118

119 Methods

120 **Participants**

This cross-sectional study was conducted among participants resulting from a posteriori pooling of 121 participants involved in two previous studies^(24,25). These studies aimed to assess the impact of food 122 labelling on energy intake, appetite sensations and food perceptions during either a 10-day experimental 123 period, where they received three *ad libitum* take-home meals per day⁽²⁴⁾ or a single *ad libitum* snack 124 test⁽²⁵⁾. In the 10-day experimental period, a label indicating either "low-fat" or the energy content of the 125 meal, or no label as a control differentiated the three experimental groups ⁽²⁴⁾. In the snack test, oatmeal-126 raisin cookies were described either as healthy (i.e., high-fiber oatmeal), diet (i.e., satiating effect) or 127 hedonic (less healthy ingredients i.e., brown sugar and butter), depending on the experimental groups⁽²⁵⁾. 128 Note that the experimental conditions of these two studies had no impact on measured energy intake^(24,25). 129 Participants were recruited through different media at Blinded for review University or in the Blinded for 130 review area. Inclusion criteria for the present study were as follows: age between 18 and 68 years, self-131 132 reported stable body weight $(\pm 2.5 \text{ kg})$ in the last two to three months prior to the study, no medications that could interfere with study outcomes (e.g., corticosteroids, antidepressants, antipsychotics), no 133 weight-related or chronic health diseases (e.g., eating disorders, type 1 or type 2 diabetes, uncontrolled 134 hypo- or hyperthyroidism, food allergies) and not being pregnant or lactating. Participants were blinded 135 to the objectives of each study. 136

137 Measurements

138 Anthropometric measurements

Height was measured to the nearest 0.1 cm using a standard stadiometer and body weight was measured to the nearest 0.1 kg with a digital scale. Body mass index (BMI) was calculated as body weight divided by height squared (kg/m^2). Table 1 presents a summary of measurement times of the previous studies.

142 **Diet quality assessment**

Self-reported dietary intake was measured using a validated food frequency questionnaire (FFQ), either 143 in interview $(n=269)^{(26)}$, or using a Web-based self-administered format $(n=116)^{(27)}$. The Web-based FFQ 144 contained 136 items and was developed based on the interviewer-administered FFQ which contained 91 145 146 items with a total of 33 subquestions. Both FFQs measure dietary intakes over the last month. The Webbased FFQ required approximately 45 minutes to complete and the interviewer-administered FFQ 147 required between 30 and 45 minutes. The nutritional analysis was based on the Nutrition Data System 148 for Research, version 4.03 for the interviewer-administered FFQ and on a food composition database 149 150 created based on the Nutrition Data System for Research, version 4.03 and the Canadian Nutrient File,

version 2007b for the Web-based FFQ. Servings of the 2007 Canada's Food Guide were computed using 151 an Excel File created for that purpose or electronically, depending on the FFQ. A reasonable agreement, 152 as assessed by cross-classification between quartiles of dietary intake, has been demonstrated (i.e., mean 153 of 84.3%±5.9 of participants classified within the same or adjacent quartiles of dietary intakes for all 154 nutrients, with 2.5% ±2.0 of subjects classified in non-adjacent quartiles), and significant correlations for 155 the majority of nutrients (average of Pearson correlation coefficients, r=0.59±0.15) have also been shown 156 between both FFQs⁽²⁷⁾. Diet quality was assessed using the Healthy Eating Index (HEI) adapted for the 157 Canadian nutrition recommendations⁽²⁸⁾. This index reflects the global quality of the diet on a 100-point 158 score comprising 10 components. The HEI score was calculated based on data obtained from the 159 nutritional analysis. 160

161 **Questionnaires**

Eating behaviour traits were assessed using a validated French version⁽²⁹⁾ of the Three-Factor Eating 162 Questionnaire $(TFEQ)^{(1,15)}$, the Restraint Scale⁽³⁰⁾ and the Intuitive Eating Scale⁽³¹⁾ translated in French. 163 164 The TFEQ measures cognitive restraint (21 items, Cronbach's alpha=0.81) and its two subscales, i.e., rigid and flexible control (7 items each, Cronbach's alpha=0.59 and 0.62, respectively), disinhibition (16 165 items, Cronbach's alpha=0.71), and susceptibility to hunger (14 items, Cronbach's alpha=0.72). The 166 Restraint Scale (10 items, Cronbach's alpha=0.64) also assesses restrained eating but combined with a 167 weight fluctuation factor⁽³⁰⁾. The Intuitive Eating Scale measures total intuitive eating score (21 items, 168 Cronbach's alpha=0.85) as well as three subscales, i.e., unconditional permission to eat (9 items), eating 169 for physical rather than emotional reasons (6 items) and reliance on hunger and satiety cues to determine 170 when and how much to eat (6 items) (Cronbach's alpha=0.79, 0.89 and 0.74, respectively)⁽³¹⁾. 171

A French version of the validated Label Reading Survey⁽³²⁾ was used to measure a general (i.e., general 172 food label use) and specific behaviour (i.e., item seeking on food labels) towards food label use. The 173 Cronbach's alpha coefficient for the whole questionnaire, that also measures attitudes and knowledge 174 towards food labels, was 0.78, which was similar to the value of the original questionnaire (i.e., 0.80)⁽³²⁾. 175 Cronbach's alpha coefficients were 0.57 and 0.81 for general food label use and item seeking on food 176 177 labels, respectively. General food label use is measured as the sum of three items on a 5-point scale (never, 1 to always, 5), e.g., "When you purchase a food product for the first time, do you look at the 178 179 Nutrition Facts label on the package?". Item seeking is measured as the sum of 15 items appearing on the Nutrition Facts table (e.g., serving size, calories, sodium, etc.) and 2 items related to health and 180 181 nutrition claims on food labels. Participants were asked to indicate whether they used each item when looking at food labels (No, 1; Yes, 2). This French version of the Label Reading Survey was adapted to the Canadian food labelling context (e.g., by replacing the word "Americans" for "Canadians" and modifying examples of American Nutrition Facts labels for Canadian labels), but these adaptations did not change the nature of the questionnaire. No changes were made to the items related to general food label use and the only change to item seeking on food labels related items was modifying "Calories from fat" for "percentage of daily value from fat" as the former do not appear on the Canadian Nutrition Facts Table. Participants also completed a sociodemographic questionnaire.

189 Statistical analyses

Descriptive statistics (mean \pm SD or frequency) were computed to assess participant characteristics and 190 191 eating behaviour traits in the whole sample. T-test and chi-square analyses were performed to assess differences between men and women. To account for the possibility of under- and overreporting of 192 193 dietary intakes, participants having a ratio of self-reported energy intake to estimated basal metabolic rate (BMR), calculated with the Harris-Benedict equation, lower than 1.14 and higher than 2.4 were 194 excluded from the analyses⁽³³⁾. A ratio below 1.14 rather than 1.35 was chosen to identify underreporters 195 of energy intake as it represents the lowest energy intake to BMR ratio that may reflect actual energy 196 intake over a given period of time⁽³⁴⁾. Moreover, this ratio was chosen because restrained eaters and 197 individuals with obesity are more likely to underreport dietary intake^(33,35) and excluding these 198 199 individuals was not desired given the objectives of this study. A total of 88 and 35 participants were identified as under- and overreporters, respectively. Therefore, 385 participants were included in the 200 analyses. Analyses were adjusted for the experimental conditions of the two previous studies by creating 201 five indicator variables (i.e., experimental groups 1 to 3 were assigned to the three groups of Blinded for 202 review et al. study⁽²⁴⁾ and experimental groups 4 to 6 were assigned to the three groups of Blinded for 203 *review* et al. $study^{(25)}$. The indicator variables were created for experimental conditions 1, 2, 4, 5, 6 and 204 the control group of *Blinded for review* et al. study⁽²⁴⁾ (experimental condition 3) was used as the 205 reference). These indicator variables were added as covariates in each analysis, even though no difference 206 was observed in the main eating behaviour traits and food label use among the different experimental 207 groups of the two previous studies (data not shown, P>0.05). 208

Partial Pearson's correlations were used to assess the associations among eating behaviour traits, diet quality and food label variables. These associations were first tested in a model that was only adjusted for experimental conditions and then in a model that was further adjusted for potential confounders (i.e., experimental conditions, age, gender, BMI, education level and household income)⁽²⁰⁾. Age, BMI,

gender, education level [2 to 5 (no participant reported having no education level or not having completed 213 elementary school which was coded as 1)] and household income (1 to 6) were treated as continuous 214 variables while gender (men, 0; women, 1) was treated as a binary variable. Total scores of the main 215 eating behaviour traits and food label variables that were significantly associated with HEI score were 216 included in multiple stepwise regression analyses. These analyses were performed using an unadjusted 217 model except for experimental conditions, and a fully adjusted model considering experimental 218 conditions, BMI, age, gender, education level and household income as covariates. A second series of 219 multiple stepwise regressions was run using the subscales of eating behaviour traits and food label 220 variables that were significantly associated with HEI score, again in an unadjusted model, except for 221 experimental conditions, and a fully-adjusted model for potential confounders. 222

Moderated mediation analyses were conducted to assess whether food label use mediate the association 223 between eating behaviour traits and diet quality, and whether the mediation effect vary according to 224 gender since gender differences have been observed in eating behaviour traits, food label use and diet 225 quality^(5,20). These analyses were conducted with the use of model 58 in the Process macro version 2.16.3 226 for SAS that calculates bias-corrected 95% confidence intervals (CI) using bootstrapping with 5,000 227 samples⁽³⁶⁾. Based on the location of the gender interaction identified, the analysis was rerun using the 228 most suitable model (i.e., models 14 or 7) and if no moderated mediation was observed, the simple 229 230 mediation model was used (i.e., model 4). The mediations were only tested in the model that was fully adjusted for covariates while considering gender as a potential moderator rather than a covariate. In cases 231 where no moderating effect was observed, gender was thereafter considered as a covariate. Statistical 232 significance was set to a *P*<0.05. Bonferroni adjustments for multiple comparisons were not used because 233 of the exploratory nature of this study⁽³⁷⁾, particularly regarding mediation analyses among diet quality, 234 eating behaviour traits and food label use, since correlation analyses are generally a previous step for 235 mediation analyses. It is however possible that chance associations are presented for some findings, 236 especially for those close to a *P* value of 0.05. However, to minimize this possibility, the only mediation 237 models tested were those where significant associations between the independent and dependent 238 variables (path c), between the independent variable and the mediator (path a) and between the mediator 239 and the dependent variable (path b) were observed, according to the traditional view of interpretation of 240 mediation analysis according to Baron and Kenny⁽³⁸⁾. This rational was used because the aim of this 241 study was to better understand the observed associations between eating behaviour traits and diet quality. 242 All statistical analyses were performed using SAS software version 9.4 (SAS Institute, Cary, NC, USA). 243

244 **Results**

245 **Participant characteristics**

Mean age of participants (women n=265; men n=120) was 42.9 ± 15.1 and 37.0 ± 14.1 years for women and men, respectively, and slightly more than two third of the sample were women (**Table 2**). Women were significantly older, had a higher BMI, HEI score and level of restrained eating, as assessed with the TFEQ or with the Restraint Scale, compared with men. Women also presented a higher level of disinhibition and a lower intuitive eating score than men (Table 2).

251 Associations of eating behaviour traits and food label use with diet quality

TFEQ-cognitive restraint and its subscales were positively associated with HEI score in the model adjusted only for experimental conditions and in the fully-adjusted model (**Table 3**). Intuitive eating was negatively, but weakly, associated with HEI score in the model adjusted for experimental conditions (P=0.03) whereas the association was no longer significant in the fully-adjusted model (P=0.052). A negative association was observed with one of the intuitive eating subscales, i.e., unconditional permission to eat, and diet quality in both models. Scores reflecting general food label use and item seeking on food labels were all positively associated with HEI score in each model (Table 3).

259 Associations among eating behaviour traits and food label use

Positive correlations were observed for the association of TFEQ-cognitive restraint, rigid control and 260 flexible control with general food label use and item seeking on food labels (Table 4). Small but positive 261 correlations were also observed for disinhibition and susceptibility to hunger with general food label use 262 in the fully-adjusted model. Restrained eating, assessed with the Restraint Scale, was positively 263 associated with general food label use but not with item seeking on food labels. Intuitive eating and its 264 subscale unconditional permission to eat were negatively associated with general food label use and with 265 266 item seeking on food labels in both statistical models while the subscale eating for physical rather than emotional reasons was negatively associated with general food label use in the fully-adjusted model 267 (Table 4). 268

269 Multiple regression analyses

270 The first multiple regression model tested for diet quality (HEI score) included TFEQ-cognitive restraint,

intuitive eating, general food label use, item seeking on food labels as well as experimental conditions.

- General food label use (β =1.21±0.26, *P*<0.0001) and TFEQ-cognitive restraint (β =0.39±0.15, *P*=0.009)
- explained 6.4% and 2.1% of the variance in the HEI score, respectively. The model explained 11.9% of
- the variance in the HEI score (P < 0.0001). Adding potential confounders (i.e., age, gender, BMI,

275 education level and household income) into the model increased the percent of variance explained in HEI score to 18.2% (P<0.0001). General food label use (β =1.18±0.26, P<0.0001) remained the only 276 277 significant variable among the main eating behaviour traits and food label variables and it explained 6.7% of the variance in HEI score, although a tendency was observed for TFEQ-cognitive restraint 278 $(\beta=0.28\pm0.15, R^2=1.2\% P=0.06)$. Gender $(\beta=4.71\pm1.30, P=0.0004)$ and BMI $(\beta=-0.32\pm0.12, P=0.01)$ 279 explained respectively 4.3 and 2.2% of the variance in HEI score respectively, indicating that women 280 and those with a lower BMI had a higher diet quality. Age and experimental conditions 1 and 2 remained 281 in the model but were not significant (P>0.05). 282

The model was also tested with the subscales that were significantly correlated with the HEI score. The 283 284 first model thus included rigid and flexible control, unconditional permission to eat, general food label use, item seeking on food labels and experimental conditions. General food label use (β =1.26±0.26, 285 P < 0.0001) and flexible control ($\beta = 0.85 \pm 0.38$, P = 0.03) respectively explained 6.9 and 1.5% of the 286 variance in HEI score. The model explained 10.3% of the variance in HEI score (P < 0.0001). In the fully-287 288 adjusted model, the percentage of variance explained in HEI score increased to 17.5% (P<0.0001). Among the eating behaviour traits and food label variables, general food label use (β =1.09±0.26, 289 290 P < 0.0001) remained again the only significant determinant of HEI score, explaining 5.7% of its variance. Gender (β=4.88±1.31, P=0.0002) and BMI (β=-0.42±0.12, P=0.0006) explained 4.6 and 4.0% in the 291 292 variance in HEI score, respectively, again suggesting that women and individuals with lower BMI had a better diet quality. Unconditional permission to eat and experimental conditions 1 and 2 remained in the 293 model but did not significantly contribute to explain the HEI score (P>0.05). 294

Additional analyses were performed to test whether food label variables could mediate the association 295 between eating behaviour traits (i.e., cognitive restraint and its two subscales, and unconditional 296 297 permission to eat) and diet quality since these eating behaviour traits were associated with food label use variables and diet quality (Figure 1, Table 5). Results showed that general food label use was a partial 298 mediator of the associations between TFEQ-cognitive restraint, flexible control and unconditional 299 permission to eat and HEI score and that general food label use mediated the association between rigid 300 301 control and HEI score. Moreover, the index of moderated mediation indicates that the mediating effects were stronger in men than in women except for the model with rigid control since the index of moderated 302 mediated did not reach significance [95% bootstrap IC (-1.46, 0.02)]. A similar pattern of association 303 was observed in the models testing the mediating effect of item seeking on food labels. Indeed, item 304 305 seeking on food labels partially mediated the association of TFEQ-cognitive restraint and rigid control with HEI score, but this was observed only in men. Item seeking on food labels partially mediated the association between flexible control and HEI score with no moderating effect of gender. Finally, item seeking on food labels partially mediated the association between unconditional permission to eat and HEI score and the mediating effect was stronger in men than in women (Table 5).

310 Discussion

311 This study aims to assess the associations among eating behaviour traits, food label use and diet quality and to examine whether the association between eating behaviour traits and diet quality was mediated by 312 food label use. Among the variables examined in this study, general food label use appears to be the main 313 determinant of diet quality although correlation analyses also showed positive associations between 314 TFEQ-cognitive restraint, and its subscales, and diet quality (HEI score) and negative associations for 315 intuitive eating (in the model that was only adjusted of experimental condition), and its subscale 316 unconditional permission to eat, with diet quality. Results also revealed that the associations between 317 most of these eating behaviour traits and diet quality were partially mediated by general food label use 318 and item seeking on food labels and the mediating effect was stronger in men than in women in most 319 320 models.

The pattern of associations between eating behaviour traits and diet quality is consistent with the 321 literature. Indeed, the positive association between cognitive restraint and diet quality has been 322 previously reported by studies showing that cognitive restraint and flexible control were associated with 323 324 higher intakes of foods that are components of healthy eating such as green vegetables, fish and yogurts or with a higher diet quality score based on fruit and vegetables, whole-grain products and fish intakes, 325 respectively^(4,5). The positive association between rigid control and diet quality must although be 326 interpreted with caution since restrained eaters may be more prone to social desirability bias when 327 reporting eating habits⁽³⁵⁾. Because rigid control is characterized by a dichotomous (all or nothing) 328 approach towards eating and has been positively associated with disinhibition⁽¹³⁻¹⁵⁾, it may not be a 329 positive determinant of diet quality in the longer term. Accordingly, it has been suggested that a high 330 level of cognitive restraint in women may be difficult to sustain over time⁽¹⁴⁾. In contrast to rigid control, 331 flexible control represents an approach towards eating that is characterized by a higher probability of 332 successful weight reduction or weight management and by a negative association with disinhibition^(14,15). 333 Such literature suggests that flexible control may be easier to maintain over time, so individuals may be 334 less likely to show important deviations from their usual dietary habits and thus, they may eat less 335 unhealthy foods as supported by a positive association with diet quality⁽⁵⁾. Although the negative 336

association between unconditional permission to eat and diet quality has not been previously reported,
this intuitive eating subscale has been associated with lower fruit and vegetables and whole-grain
intakes⁽³⁹⁾, which represent important components of diet quality.

Food label use, as measured by both general food label use and item seeking on food labels, was 340 associated with a better diet quality. This is in line with the majority of studies on this topic demonstrating 341 that food label use benefit eating habits⁽²⁰⁾. Considering the cross-sectional nature of the present study, it 342 is also possible that individuals having a better diet quality pay more attention to food labels, as it has 343 been reported that individuals having better eating habits report a greater use of food labels⁽²⁰⁾. To our 344 knowledge, this study is one of the first to document the associations between eating behaviour traits and 345 346 food label use. The positive associations among restrained eating and food label use were expected since restrained eating requires cognitive effort to adhere to a diet in order to lose or to maintain body weight, 347 and nutritional information found on food labels could support that effort. Similarly, the negative 348 association between intuitive eating and food label use suggests that because intuitive eating relies more 349 350 on internal sensations of hunger and satiety, it could be less related to cognitive processes towards eating such as using food labels when purchasing or consuming food. Disinhibition and susceptibility to hunger 351 352 were both positively, but weakly, associated with general food label use in the fully-adjusted model, suggesting that individuals presenting a higher level of disinhibition or susceptibility to hunger may use 353 354 food labels to select food products that seem healthier or lower in fat or in calories to compensate for their overeating tendencies. Likewise, individuals with susceptibility to hunger may also use food labels 355 to choose foods that seem more satiating, but is it also possible that choosing low-fat or low-calorie foods 356 triggers hunger sensations. 357

Although eating behaviour traits and food label use only explained a small proportion of the variance in 358 359 diet quality, this study suggests that food label use greater explains diet quality than eating behaviour traits. Given the myriad of factors influencing dietary habits⁽⁴⁰⁾, the percentage of variance explained in 360 diet quality by the different models and, mainly by food label use, is nonetheless considerable. Moreover, 361 using food labels seems to represent a tool that explained a greater part of the association between 362 363 cognitive restraint and diet quality in men than in women. In contrast, not using food labels seems to greater explain the negative association between unconditional permission to eat and diet quality in men 364 compared with women. Considering that men are more prone to give themselves an unconditional 365 permission to eat compared with women⁽³⁹⁾, this latter result suggests that when men allow themselves 366 367 to eat unconditionally, they use food labels less often, which contribute to a lower diet quality. This result

is in line with Christoph et al. study showing that a higher level of intuitive eating was associated with a 368 lower level of food label use in men⁽²³⁾. The fact that women usually eat less intuitively ^(11,16,17), are more 369 restrained $^{(13-15)}$, and use food labels more often than men $^{(20,21)}$ may explain why the mediating effect of 370 food label use was weaker in women. Similarly, restrained eating and dieting represent a cultural norm 371 for women in western countries^(41,42), suggesting that women may more importantly internalize these 372 behaviours. This context might explain why the association between cognitive restraint and diet quality 373 is less mediated by food label use in women as opposed to restrained men who more essentially need to 374 375 rely on such tool to achieve a better diet quality.

Nonetheless, and irrespective of gender, the mediating effect of food label use in the association between 376 377 cognitive restraint and diet quality is in line with the use of dietary restraint as a self-regulation strategy as proposed by Schaumberg et al.⁽⁴³⁾. Accordingly, our results suggest that food labelling could support 378 379 self-monitoring among restrained eaters, helping them to implement their dieting rules and reach a better diet quality, which is in line with a previous study showing that the association between attitude toward 380 381 healthy meal preparation and diet quality was mediated be a greater use of food labels among college students⁽⁴⁴⁾. Christoph et al. also showed that food label use was associated with a greater likelihood of 382 383 engaging in healthy weight control behaviours, but also, and to a lesser extent, to a greater likelihood of engaging in unhealthy weight control behaviours⁽²³⁾, suggesting that some individuals may use food 384 385 labels to implement unfavorable eating behaviours. It is important to note that the associations observed in the present study are not causal due to its cross-sectional nature. 386

This study has several strengths and limitations that need to be outlined. First, it is the first study to assess 387 the associations of many eating behaviour traits with food label use and global diet quality in the same 388 sample of men and women and to our knowledge, the mediating effect of food label use in the association 389 390 between these eating behaviour traits and diet quality has never been explored. As previously discussed, the main limitation of this study is its cross-sectional nature that does not allow to determine causality 391 among variables. Therefore, it is not possible to confirm if restrained individuals use food labels because 392 of their dieting behaviours or if using food labels when making food choices can lead to restrained eating. 393 394 Moreover, because of the rather conservative strategy used regarding mediation analyses, it is possible that other mediating effects could have been observed for eating behaviour traits that were not directly 395 associated with diet quality (HEI score). However, this was beyond the scope of this study and therefore 396 remain to be investigated. Dietary intake was self-reported, implying that potential social desirability 397 398 bias could have influenced the results. The use of a FFQ may be implicated in the small proportion of

the variance in diet quality explained by eating behaviour traits and food label use variables. While the 399 Cronbach's alpha coefficient of one of the TFEQ subscales is rather low (i.e., <0.60) and could be 400 considered as a limitation, the TFEO remains an established questionnaire used to measure eating 401 behaviour traits and the Cronbach's alpha coefficients of its three main components were adequate in this 402 sample. With regards to the low Cronbach alpha coefficient for general food label use (i.e., 0.57), the 403 validity of the Label Reading Survey has been previously reported⁽³²⁾, and as mentioned earlier, the 404 Cronbach's alpha coefficient for our adapted questionnaire was similar to the value of the original one. 405 It is also likely that the different contexts specified in the questions (i.e., using food labels when 406 purchasing a food for the first time or when eating food) are implicated in this relatively low internal 407 consistency for the general food label use variable since one could only use food labels at the point of 408 purchase. The Label Reading Survey covered the main food label components (i.e., nutrition fact table, 409 health and nutrient-related allegations), but did not cover all information included on food labels (e.g., 410 411 ingredients list and other types of allegations), and this may reduce the accuracy of the food label use measure. The high education level of participants could limit the generalization of the results to other 412 populations. Finally, it is important to mention that this is a cross-sectional study that was conducted 413 among participants of two previous studies. One could argue that this design implicated priming which 414 can impact the results of the present study. However, as previously mentioned, no experimental 415 conditions effect was observed for the main eating behaviour traits, food label use variables, measured 416 energy intake in the main studies (i.e., ad libitum snack test or 10-d energy intake). Moreover, this 417 potential priming effect was considered in all analyses by adding the experimental conditions of the main 418 studies as a covariate, so it is likely that the results observed in the present study are not explained by the 419 priming effect of the main studies or that this bias is therefore greatly reduced. 420

421 Conclusions

This study showed that food label use was a better determinant of diet quality than eating behaviour traits. Moreover, food label use partially mediated the association between cognitive restraint or unconditional permission to eat and diet quality and the mediating effects were stronger in men than in women. While food labels could be helpful to adopt a healthy diet, the psychobehavioural profile of individuals seen in a clinical context should be assessed to individualize strategies used to facilitate healthy eating. It is indeed important to support appropriate self-regulation strategies and not favour the adoption of unhealthy eating behaviour traits that may result in conterregulatory eating⁽⁴³⁾. Future studies 429 should assess the impact of food label use on eating behaviour traits and diet quality in an intervention

430 context.

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537		

538 Figure legend.

- 539 Figure 1. Mediating effect of general food label use on the association between cognitive
- restraint (A) or unconditional permission to eat (B) and diet quality.
- $a = \beta$ coefficient for the association between cognitive restraint (A), or unconditional permission to
- 542 eat (B) and general food label use.
- $b = \beta$ coefficient for the association between general food label use and diet quality (HEI score).
- 544 $c'=\beta$ coefficient for the association between cognitive restraint (A) or unconditional permission to
- 545 eat (B), and diet quality (HEI score) when the mediator (general food label use) is in the model.
- 546 Data obtained from Process model 7 for A and B. Adjusted for experimental conditions, age, BMI,
- 547 household income and education level. Prefer not to answer for education level and household
- 548 income were recoded as missing data. Education level and household income were treated as
- 549 continuous variables. Five indicator variables were created for experimental conditions and the
- 550 control group of *Blinded for review* et al. study was used as reference). Cognitive restraint was
- assessed with the TFEQ. A) n=329, B) n=336.

Measures	Blinded for rev	<i>iew</i> et al. 2012	Blinded for review et al. 2015				
Measures	Before	After	Before	After			
Height		Х	х				
Weight		Х	х				
FFQ		Х	х				
Questionnaires							
TFEQ		Х		Х			
Restraint Scale		Х		Х			
Intuitive Eating Scale		Х		Х			
Food label use		Х	х				
Sociodemographic		Х		Х			

Table 1. Summary of measurement times (before or after experimentation) of the two previousstudies.

Table 2. Participant characteristics, eating behaviour traits, diet quality and food label use of the 555

whole sample (n=385) and of women and men 556

	Total (n=	385)	Women (n	=265)	Men (n=		
Variables	mean or frequency	SD	Mean or frequency	SD	Mean or frequency	SD	- P
Gender (%)	-		68.8		31.2		< 0.0001
Age (years)	41.1	15.0	42.9	15.1	37.0	14.1	0.0004
BMI (k g/m^2)	26.0	4.9	26.4	5.5	25.1	3.3	0.006
Overweight/obese (%)	50.4		49.8		51.7		0.74
Education level (%) *							
Elementary school	0.8		1.2		0.0		
High school	12.5		11.7		14.4		
College	30.7		28.4		35.6		0.34 [¶]
University	55.7		58.4		50.0		
Prefer not to answer	0.3		0.4		0.0		
Household income (\$CA) (%) [†]							
<20 000	21.3		19.7		24.8		
20 000-39 999	17.8		17.3		18.8		
40 000-59 999	20.0		20.5		18.8		
60 000-79 999	14.6		13.8		16.2		0.71
80 000-99 9999	7.3		8.3		5.1		
≥100 000	11.9		12.2		11.1		
Prefer not to answer	7.3		8.3		5.1		
HEI score (scale 0 to 100)	79.3	11.0	80.5	10.5	76.6	11.5	0.001
Eating behaviours							
Cognitive restraint (scale 1 to 21) ^{\ddagger}	7.6	4.3	8.2	7.7	6.2	3.8	< 0.0001
Rigid control (scale 1 to 7)	2.1	1.6	2.4	1.7	1.6	1.3	< 0.0001
Flexible control (scale 1 to 7)	2.8	1.7	3.0	1.7	2.4	1.5	0.002
Disinhibition (scale 1 to 16)	5.5	2.9	5.8	3.0	4.8	2.6	0.002
Susceptibility to hunger (scale 1 to 14)	4.5	3.0	4.7	3.0	4.3	3.1	0.27
Restraint (scale 0 to 35) §	13.2	4.7	13.9	4.6	11.6	4.6	< 0.0001
Intuitive eating (scale 1 to 5)	3.4	0.5	3.3	0.5	3.6	0.5	< 0.0001
Unconditional permission to eat (scale 1 to 5)	3.2	0.7	3.1	0.6	3.4	0.7	< 0.0001
Eating for physical rather than emotional reasons (scale 1 to 5)	3.4	0.9	3.2	0.9	3.8	0.9	< 0.0001
Reliance on internal hunger and satiety cues (scale 1 to 5)	3.6	0.6	3.6	0.6	3.6	0.5	0.42
General food label use (scale 3 to 15)	9.7	2.3	9.8	2.2	9.4	2.5	0.07
Item seeking on food labels (scale 17 to 34)	26.7	3.8	27.0	3.5	26.2	4.3	0.08

BMI, body mass index; HEI score, Healthy Eating Index score. 557

* Missing values n=10 (women n=8, men n=2); [†]Missing values n=14 (women n=11, men n=3) [‡]Assessed with the TFEQ; [§] Assessed with the Restraint scale 558

- 560
- [¶]*P* value from Fisher exact test 561
- Values are presented as mean (SD) or as frequency 562

563 Table 3. Associations of eating behaviour traits and food label use with diet quality (HEI score)

Variables	Unadj	usted model	Fully-adjusted model			
	r	Р	r	Р		
Cognitive restraint [*]	0.26	< 0.0001	0.20	0.0004		
Rigid control	0.19	0.0003	0.13	0.02		
Flexible control	0.24	< 0.0001	0.17	0.003		
Disinhibition	-0.05	0.31	-0.03	0.59		
Susceptibility to hunger	0.01	0.90	0.06	0.32		
Restraint [†]	0.03	0.62	0.00	0.98		
Intuitive eating	-0.11	0.03	-0.11	0.052		
Unconditional permission to eat	-0.20	0.0001	-0.16	0.003		
Eating for physical rather than emotional reasons	0.02	0.72	0.03	0.55		
Reliance on internal hunger and satiety cues	-0.03	0.57	-0.10	0.07		
General food label use	0.30	< 0.0001	0.27	< 0.0001		
Item seeking on food labels	0.20	0.0002	0.19	0.0009		

564 HEI score, Healthy eating index score.

565 * Assessed with the TFEQ;

566 † Assessed with the Restraint scale

567 Values are partial Pearson's correlation coefficients (r).

568 Unadjusted model: adjusted only for experimental conditions; Fully-adjusted model: adjusted for

569 experimental conditions, age, gender, BMI, education level and household income. Prefer not to answer

570 for education level and household income were recoded as missing data. Education level and household

571 income were treated as continuous variables. Five indicator variables were created for experimental

572 conditions and the control group of *Blinded for review* et al. study was used as reference).

573 Unadjusted model: n=349 to 384; Fully-adjusted model: n=317 to 342.

575 Table 4. Associations between eating behaviour traits and food label use.

		General for	od label u	ise	Item seeking on food labels					
Variables		adjusted nodel	-	y-adjusted nodel		adjusted nodel	2	adjusted odel		
	r	р	r	р	r	р	Fully-ac mod r 0.17 0.13 0.14 0.00 0.01 0.01 -0.17	р		
Cognitive restraint [*]	0.37	< 0.0001	0.32	< 0.0001	0.26	< 0.0001	0.17	0.004		
Rigid control	0.35	< 0.0001	0.30	< 0.0001	0.22	< 0.0001	0.13	0.03		
Flexible control	0.33	< 0.0001	0.27	< 0.0001	0.22	< 0.0001	0.14	0.02		
Disinhibition	0.06	0.27	0.11	0.04	0.01	0.87	0.00	0.96		
Susceptibility to hunger	0.03	0.53	0.12	0.03	-0.05	0.34	0.01	0.84		
Restraint [†]	0.19	0.0002	0.22	< 0.0001	0.06	0.24	0.01	0.84		
Intuitive eating	-0.23	< 0.0001	-0.24	< 0.0001	-0.19	0.0004	-0.17	0.003		
Unconditional permission to eat	-0.35	< 0.0001	-0.33	< 0.0001	-0.33	< 0.0001	-0.27	< 0.0001		
Eating for physical rather than emotional reasons	-0.08	0.12	-0.12	0.04	-0.06	0.30	-0.08	0.18		
Reliance on internal hunger and satiety cues	0.04	0.45	0.00	0.98	0.06	0.28	0.05	0.36		

^{*} Assessed with the TFEQ; [†] Assessed with the Restraint scale

577 Values are partial Pearson's correlation coefficients (r).

578 Unadjusted model: Adjusted only for experimental conditions; Fully-adjusted model: Adjusted for

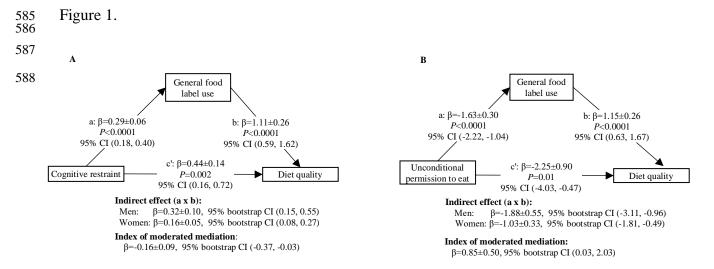
579 experimental conditions, age, gender, BMI, education level and household income. Prefer not to answer

for education level and household income were recoded as missing data. Education level and household income were treated as continuous variables. Five indicator variables were created for experimental

581 income were treated as continuous variables. Five indicator variables were created for 6 582 conditions and the control group of *Blinded for review* et al. study was used as reference).

582 General food label use: Unadjusted model: n=359 to 379; Fully-adjusted model: n=323 to 339

584 Item seeking on food labels: Unadjusted model: n=333 to 346; Fully-adjusted model: n=302 to 314.



	a *				b †				Direct effect (c') \ddagger			Indirect effect (a x b)				Index of moderated mediation			Process model used	
-	β	SE	р	95% CI	β	SE	р	95% CI	β	SE	р	95% CI		β	SE	95% Bootstrap CI	β	SE	95% Bootstrap CI	_
Mediator: General food label use																				
Rigid control §	0.45	0.08	< 0.0001	(0.29, 0.62)	1.14	0.26	< 0.0001	(0.63, 1.65)	0.37	0.39	0.35	(-0.41, 1.15)		0.52	0.14	(0.28, 0.83)	-0.63	0.38	(-1.46, 0.02)	4∥
Flexible control	0.61	0.14	< 0.0001	(0.33, 0.89)	1.21	0.26	< 0.0001	(0.70, 1.73)	0.90	0.37	0.01	(0.18, 1.62)	Men	0.74	0.23	(0.37, 1.30)	-0.39	0.23	(-0.93, -0.03)	7
													Women	0.36	0.13	(0.15, 0.64)				
Mediator: Item seeking on food labels																				
Cognitive restraint [¶]	0.35	0.09	0.0002	(0.16, 0.53)	0.59	0.16	0.0004	(0.27, 0.92)	0.53	0.14	0.0003	(0.24, 0.81)	Men	0.21	0.10	(0.06, 0.44)	-0.17	0.09	(-0.40, -0.02)	7
													Women	0.04	0.03	(-0.02, 0.12)				
Rigid control	0.90	0.27	0.001	(0.37, 1.43)	0.61	0.17	0.0004	(0.27, 0.94)	1.07	0.39	0.007	(0.30, 1.85)	Men	0.54	0.26	(0.14, 1.18)	-0.48	0.27	(-1.11, -0.06)	7
													Women	0.07	0.09	(-0.08, 0.28)				
Flexible control §	0.31	0.13	0.02	(0.05, 0.58)	0.52	0.16	0.002	(0.20, 0.85)	0.75	0.38	0.047	(0.01, 1.49)		0.16	0.09	(0.03, 0.40)	-0.29	0.23	(-0.86, 0.08)	4**
Unconditional permission to eat	-2.54	0.50	< 0.0001	(-3.52, -1.56)	0.50	0.17	0.003	(0.17, 0.83)	-2.81	0.94	0.003	(-4.67, -0.96)	Men	-1.27	0.50	(-2.51, -0.46)	0.82	0.44	(0.14, 1.96)	7
													Women	-0.45	0.26	(-1.16, -0.08)				

Table 5. Mediation models between eating behaviour traits, food label use and diet quality (HEI score)

 $* a = \beta$ coefficient for the association between eating behaviour traits and general food label use/item seeking on food labels.

[†] $b = \beta$ coefficient for the association between general food label use/item seeking on food labels and diet quality (HEI score).

 $\frac{1}{c} = \beta$ coefficient for the association between eating behaviour traits and diet quality (HEI score) when the mediator (general food label use/item

seeking on food labels) is in the model. Adjusted for experimental conditions, age, BMI, household income and education level. Prefer not to answer

594 for education level and household income were recoded as missing data. Education level and household income were treated as continuous variables.

595 Five indicator variables were created for experimental conditions and the control group of *Blinded for review* et al. study was used as reference).

[§] This model was also adjusted for gender as gender did not moderate the mediation effect.

¹ Index of moderated mediation was obtained from model 58, but it was also non-significant in other moderated mediation models (i.e., models 14 and 7).

⁵⁹⁹ [¶]Assessed with the TFEQ.

^{**} Index of moderated mediation was obtained from model 7, but it was also non-significant in other moderated mediation models (i.e., models 58 and 14).

Models using general food label use as a mediator: n=323 to 325; Models using item seeking on food labels as a mediator: n=302 to 312.