# Extreme Rainfall Intensities and Long-term Rainfall Risk from Tropical Cyclones 

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# Extreme Rainfall Intensities and Long-term Rainfall Risk from Tropical Cyclones 

by<br>Andreas Langousis<br>Submitted to the Department of Civil and Environmental Engineering on December 16, 2008 in partial fulfillment of the requirements for the Degree of Doctor of Science in the Field of Civil and Environmental Engineering.


#### Abstract

We develop a methodology for the frequency of extreme rainfall intensities caused by tropical cyclones (TCs) in coastal areas. The mean rainfall field associated with a TC with maximum tangential wind speed $V_{\max }$, radius of maximum winds $R_{\max }$, and translation speed $V_{t}$ is obtained using a physically-based model, whereas rainfall variability at both large scales (from storm to storm) and small scales (due to rainbands and local convection) is modeled statistically.

The statistical component is estimated using precipitation radar (PR) data from the TRMM mission. Taylor's hypothesis is used to convert spatial rainfall intensity fluctuations to temporal fluctuations at a given location A. The combined physical-statistical model gives the distribution of the maximum rainfall intensity at A during a period of duration $D$ for a TC with characteristics ( $V_{\max }, R_{\max }, V_{t}$ ) that passes at a given distance from A.

To illustrate the use of the model for long-term rainfall risk analysis, we formulate a recurrence model for tropical cyclones in the Gulf of Mexico that make landfall between longitudes $85^{\circ}$ $95^{\circ} \mathrm{W}$. We then use the rainfall and recurrence models to assess the rainfall risk for New Orleans. For return periods of 100 years or more and long averaging durations ( $D$ around 12-24 hours), tropical cyclones dominate over other rainfall event types, whereas the reverse is true for shorter return periods or shorter averaging durations.


Thesis Supervisor: Daniele Veneziano<br>Title: Professor of Civil and Environmental Engineering

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## Introduction

In this thesis we develop a framework to estimate extreme rainfalls from tropical cyclones at coastal sites. Interest is in the distribution of $I_{D, \max }$ : the maximum rainfall intensity at a site A over a period $D$ during the passage of a storm.

The distribution of rainfall extremes like $I_{D, \max }$ is usually displayed in the form of intensity-duration-frequency (IDF) curves; see e.g. Chow et al. (1988) and Singh (1992). These are plots of the rainfall intensity $i$ as a function of the average duration $D$ for different return period values $T$, where $T$ is defined as the reciprocal of the exceedance rate $\lambda_{I_{D, \max }>i}=\lambda P\left[I_{D, \max }>i\right] . \lambda$ is the annual rate at which TCs "hit" site A.

Knowledge of the distribution of $I_{D, \max }$ is important when one assesses flood risk in TCprone areas as the combination of rainfall, surges and waves (Herbert et al., 1997; Rappaport, 2000). In addition, interesting insight can be produced when comparing theoretical IDF curves for TC rainfall at a site to empirically obtained IDF values for all rainstorms included (TCs and non-TCs). For example, one can conclude on the range of durations $D$ and return periods $T$ for which TCs dominate over other rainfall event types.

The episodic and spatially localized nature of TCs make the distribution of $I_{D, \max }$ hard to obtain using standard methods of probabilistic assessment of extreme rainfalls. These include fitting a probability distribution directly to the historical annual maxima observed at the site (e.g. Koutsoyiannis et al., 1998 for a review) or derive the maximum distribution from peak-over-threshold (PoT) information (e.g. Madsen et al., 1997). For example, both annual maximum TC rainfalls and annual PoT statistics at a site depend on whether in any given year the site was "hit" by one or more TCs and therefore are highly erratic.

In this case, the assessment of risk is best done parametrically, using the effects of events with various characteristics and the frequency with which those events occur. For tropical cyclones, the storm characteristics of interest include the intensity, size, location, translation speed and direction at landfall, and possibly other parameters. We list these relevant parameters as components of a vector $\theta$. This parametric approach has been used with tropical cyclones in the context of wind, surge and waves (Myers, 1975; Ho and Myers, 1975; Ho et al., 1987; Powell et al., 2005; IPET, 2006, 2008), but not rain. The objective of this thesis is to fill this gap.

The main problem in a parametric assessment of TC rainfall risk is the evaluation of the extreme rainfalls caused by a TC with given characteristics $\boldsymbol{\theta}$ : i.e. the data on TC rainfall are sparse and the potentially important TC parameters too many to directly infer a parametric rainfall model from empirical observations. For example, current parametric approaches (Lonfat et al., 2004, 2007; Tuleya et al., 2007) classify storms into three coarse categories of hurricane intensity, use microwave imager (TMI) data from TRMM (Simpson et al. 1988) to calculate the ensemble-average rainrate for each category as a function of distance from the TC center, and interpolate from these results to calculate the $24-\mathrm{hr}$ average rainfall accumulation from the passage of a TC. Limitations of these approaches are: 1) the coarse and incomplete storm parameterization: the size of the vortex $R_{\max }$ is only implicitly taken into account through its dependence on $V_{\max }$, while other factors (e.g. the radial wind velocity profile in the main vortex, the surface roughness, and the storm translation velocity) are ignored and 2) the fact that they do not account for rainfall intensifications due to rainbands and local convection.

A more complete approach is to combine physical modeling with data analysis. Here, we use a relatively simple physical model to capture the main effects of $\boldsymbol{\theta}$ on rainfall, and statistical data analysis to assess the storm-to-storm fluctuations due to features like rainbands and convective cells that the physical model does not resolve.

The thesis is organized into three main chapters. Chapter 1 describes a theoretical method to estimate the large-scale horizontal and vertical winds inside TCs (the vertical winds are largely responsible for rain). The model is an extension of Smith's (1968) formulation and is referred to here as the Modified Smith (MS) model. Characteristics of the TC that are explicitly considered by the MS model are the maximum tangential wind speed $V_{\max }$, the radius of maximum winds $R_{m a x}$, the parameter $B$ that controls the shape of the radial profile of the tangential wind speed (Holland, 1980), the storm translation velocity $V_{t}$, the surface drag coefficient $C_{D}$ and the vertical diffusion coefficient $K$. The MS model does not resolve rainbands, local convection and turbulent phenomena and therefore produces smooth wind fields.

Chapter 2 extends the MS model to predict TC rain. This is done under the assumption that water vapor condensation in the main vortex converts all the upward moisture flux at the top of the TC boundary layer to rainfall. The vertical moisture flux itself is evaluated from the vertical winds generated by the MS model and two additional parameters: the average temperature $\bar{T}$ and the average saturation ratio $\bar{Q}$ inside the TC boundary layer. The calculated rainfall field is not
simply proportional to the vertical winds at the top of the boundary layer produced by the MS model. This is because (1) the trajectory of moisted air parcels has an outward slant depending on distance from the TC center and (2) the ascending air parcels and descending rain drops are advected into a helical motion by the cyclonic circulation; therefore a parcel of air that leaves the TC boundary layer contributes rainfall to a range of azimuthal locations. We call this the modified-Smith-for-rainfall (MSR) model. The MSR model should prove useful for climatologic studies, but for hazard analysis it has the major limitation of ignoring the smallscale rainfall fluctuations associated with rainbands and local convection.

Chapter 3 adds a stochastic fluctuation component to the MSR predictions to determine the probability distribution of the maximum rainfall intensity at a site A during a period of duration $D, I_{D, \max }(y, \boldsymbol{\theta})$, for a TC with characteristics $\boldsymbol{\theta}$ whose center passes at distance $y$ from A. The statistical component is fitted and validated using precipitation radar (PR) data from the TRMM mission. Also, Taylor's hypothesis is used to convert spatial rainfall intensity fluctuations to temporal fluctuations at A. To illustrate the use of the model for long-term rainfall risk analysis, we formulate a recurrence model for tropical cyclones in the Gulf of Mexico. We then use the rainfall and recurrence models to assess the rainfall risk for New Orleans.

Chapters 1, 2 and 3 are written in a stand-alone format and are very similar to published (Langousis et al., 2008; Langousis and Veneziano, 2008) or submitted (Langousis and Veneziano, 2009) papers.

Conclusions and possible future research directions are given in Chapter 4. Supplemental material for the theoretical model in Chapter 1 and alternative methodologies for evaluating the statistical component for the rainfall fluctuations in Chapter 3 are presented in Appendices A and $B$, respectively.

## 1. Boundary Layer Model for Moving Tropical Cyclones

### 1.1 Introduction

Tropical cyclones (TCs) are a particular class of rotating low-pressure systems that develop over tropical and subtropical waters. The systems have a warm-core, a well-organized convection, and cyclonic surface wind circulation (Anthes, 1982; Landsea, 2000).

Empirical observations (La Seur and Hawkins, 1963; Hawkins and Rubsam, 1968; Holland, 1980; Willoughby, 1990, 1991; Vickery et al., 2000; among others) show that in the altitude range from $2-3 \mathrm{~km}$ to about 10 km , the tangential winds are in approximate gradient balance and the radial inflow is negligible. Based on earlier work by Schloemer (1954) and Myers (1957), Holland (1980) used a symmetric pressure distribution to derive the tangential gradient wind $V_{g r}$, as a function of distance $R$ from the TC center. His result, which we refer to here as Holland's wind profile, is

$$
\begin{equation*}
V_{g r}(R)=V_{\max } \sqrt{\left(R_{\max } / R\right)^{B} \exp \left[1-\left(R_{\max } / R\right)^{B}\right]} \tag{1.1}
\end{equation*}
$$

where $V_{\max }, R_{\max }$, and $B$ are TC-specific constants. The tangential velocity $V_{g r}$ increases with $R$ to a maximum $V_{\max }$ at $R=R_{\max }$ (usually referred to as the radius of maximum winds). For $R \gg$ $R_{\max }, V_{g r}$ has an approximately power-law decay with distance, with exponent $-B / 2$. According to Willoughby and Rahn (2004), $B$ varies in the range [1, 2] with typical values around 1.4.

Inside the TC boundary layer (BL) (within approximately 1-2km from the surface), frictional stresses are important and result in an inward net force that drives low-level convergence. Consequently, the horizontal and vertical wind fields are strongly coupled and equation (1.1) does not apply. Horizontal convergence drives the vertical winds, which are maximum at the top of the boundary layer near the radius of maximum winds $R_{\max }$ (e.g. Kepert, 2001 and Kepert and Wang, 2001).

Since the convergence of moisture inside the BL is of major importance for the maintenance, evolution and destructive potential of TCs (Emanuel, 1986, 1989; Renno and Ingersoll, 1996), a number of studies (Myers and Malkin, 1961; Chow, 1971; Shapiro, 1983; Kepert, 2001) have focused on developing theoretical models for the boundary layer of moving TCs. These models derive the radial and tangential winds inside the boundary layer from an assumed radial profile
of the tangential wind velocity under gradient balance, for example the profile in equation (1.1), and from suitable surface boundary conditions.

Section 1.2 reviews these BL models and their limitations. Section 1.3 describes our proposed model by giving the governing equations (an extension of the equations of Smith, 1968) and discussing their numerical solution. In Section 1.4, we compare model results with earlier models and with simulations using the Fifth-Generation Pennsylvania State University/NCAR Mesoscale Model (MM5). Section 1.5 shows how the calculated winds depend on various storm parameters. Conclusions are stated in Section 1.6.

### 1.2 Review of Boundary Layer Models

The focus of this review is on BL models for moving tropical cyclones, but studies of stationary TCs that are relevant to what follows are also mentioned.

Boundary layer models differ mainly in their treatment of altitude $Z$ and the surface boundary conditions. In one of the earlier studies of moving TCs, Myers and Malkin (1961) used a Lagrangian parcel trajectory approach to study the horizontal winds inside the BL. The authors assume that the frictional drag force is proportional to the square of the wind speed with equal tangential and radial components. Another (implicit) assumption is that the velocity of the background flow is zero rather than equal to the translation velocity of the TC. A finding of the study is that, when a TC in the northern (southern) hemisphere moves, the radial convergence is maximum at the right-front (left-front) quadrant of the vortex and the location of this maximum rotates anticyclonically as the translation velocity $V_{t}$ increases.

Based on the work of Chow (1971), Shapiro (1983) approximated the boundary layer of a moving TC by a slab of constant depth $H=1 \mathrm{~km}$. The horizontal momentum equations are formulated in cylindrical coordinates that translate with the vortex and then averaged in the vertical direction. This results in a system of two partial differential equations (PDEs) that are solved numerically for the vertically averaged tangential $\bar{V}(R, \theta)$ and radial $\bar{U}(R, \theta)$ wind velocity as a function of radius $R$ and the azimuth $\theta$ relative to the direction of TC motion. Contrary to Myers and Malkin (1961), Shapiro's (1983) formulation assumes that the frictional drag force is parallel to the surface-relative flow and its magnitude is proportional to the square of the composite surface-relative wind velocity. Although the two studies use different formulations for the friction-induced convergence, they both produce maximum convergence at the right-front
quadrant of vortices in the northern hemisphere. However, in Shapiro's model the location of the maximum does not depend on the translation velocity, whereas in Myers and Malkin's (1961) analysis it does.

The main limitation of Shapiro's (1983) approach is that it approximates the BL as a slab of constant depth and hence cannot resolve the variation of $U$ and $V$ with height. According to Anthes (1971), this leads to overestimation of the radial and vertical velocities close to the vortex core; see discussion of Figure 1.3 below. Another limitation of Shapiro's analysis is that the numerical stability of the system depends on TC parameters such as the depth of the boundary layer $H$, the translation velocity $V_{t}$, the vertical diffusion coefficient $K$ and the surface drag coefficient $C_{D}$. This limitation becomes important when, as for example in risk studies, one needs to calculate the wind field under a wide variety of conditions.

A third BL model for moving tropical cyclones was proposed by Kepert (2001) (see also refinements in Kepert, 2006b). Kepert's formulation neglects vertical advection and linearizes the horizontal advection. This produces a system of linear PDEs that is solved analytically for the radial, tangential and vertical wind velocities ( $U, V$ and $W$, respectively) as a function of $R, \theta$, and Z. Kepert (2001) uses a bulk formulation of the surface stresses similar to those of Rosenthal (1962), Shapiro (1983) and Smith (1968, 2003). However, the surface boundary condition is linearized to allow analytical integration. Linearization produces inaccurate results close to the TC center ( $R<2-3 R_{\max }$ ) where the horizontal gradient of the wind components is high, when the vertical gradient of the horizontal wind velocity is large (this happens for large surface drag coefficient $C_{D}$; see discussion on Figure 1.2 below), for high translation velocities ( $V_{t}>5 \mathrm{~m} / \mathrm{s}$ ), and under inertially neutral conditions ( $B>1.6-1.8$ ). Other linearizations of the horizontal momentum equations have been proposed by Haurwitz (1935), Rosenthal (1962), Miller (1965) and Elliassen and Lystad (1977), but these formulations are for stationary vortices ( $V_{t}=0$ ).

An order of magnitude analysis by Smith (1968) shows that in the near-core region the nonlinear terms are as important as the linear ones. To include the non-linear terms, Smith (1968) (see also refinements in Leslie and Smith, 1970 and Bode and Smith, 1975) used the Karman and Pohlhausen momentum integral method to calculate the radial $U(R, Z)$ and tangential $V(R, Z)$ wind velocities in a stationary vortex. In the momentum integral method (Schlichting, 1960), one avoids an explicit analysis of altitude $Z$ by assuming vertical profiles for $U$ and $V$ that satisfy the boundary conditions at the surface $(Z=0)$, and tend asymptotically to gradient
balance as $Z \rightarrow \infty$; see equation (1.3) below. Specifically, Smith (1968) used profiles of the Ekman type, with an amplitude coefficient $E$ and a dimensionless BL scale thickness $\delta$ as parameters. The horizontal momentum equations are vertically integrated to produce a system of ordinary differential equations that are solved numerically to obtain $E$ and $\delta$ as a function of $R$.

The main limitation of Smith's (1968) model is that it does not consider storm motion. Also, Smith's (1968) formulation is theoretically correct only for the case of no slip at the surface boundary (i.e. for $C_{D} \rightarrow \infty$ ); see Section 1.3. In the following section, we extend Smith's (1968) model to include storm motion and correct the formulation for the general case of stress surface boundary conditions.

### 1.3 Proposed Model

In a cylindrical coordinate system $(R, \theta, Z)$ that follows the vortex motion, the boundary layer equations are (see Smith, 1968 and Kepert 2001 for a detailed derivation),

$$
\begin{align*}
& U \frac{\partial U}{\partial R}+\frac{V}{R} \frac{\partial U}{\partial \theta}+W \frac{\partial U}{\partial Z}+\frac{V_{g r}{ }^{2}-V^{2}}{R}+f\left(V_{g r}-V\right)=K \frac{\partial^{2} U}{\partial Z^{2}}  \tag{a}\\
& U \frac{\partial V}{\partial R}+\frac{V}{R} \frac{\partial V}{\partial \theta}+W \frac{\partial V}{\partial Z}+\frac{U V}{R}+f U=K \frac{\partial^{2} V}{\partial Z^{2}}  \tag{b}\\
& \frac{\partial(R U)}{\partial R}+\frac{\partial V}{\partial \theta}+\frac{\partial(R W)}{\partial Z}=0 \tag{c}
\end{align*}
$$

where $R$ is distance from the vortex center, $\theta$ is azimuth relative to the direction of motion, $f$ is the Coriolis parameter, $U, V$, and $W$ are the storm-relative radial, tangential and vertical wind velocities, respectively, $K$ is the vertical diffusion coefficient of the horizontal momentum, and $V_{g r}$ is the tangential wind velocity under gradient wind balance; see for example equation (1.1).

As Kepert (2001) and Smith (1968), we solve equations (1.2) for the case of a semi-infinite domain, when gradient wind balance is satisfied asymptotically as $Z \rightarrow \infty$,

$$
\begin{equation*}
\frac{\partial U}{\partial Z}=\frac{\partial V}{\partial Z}=U=0, \text { and } V=V_{g r}, \text { at } Z \rightarrow \infty \tag{1.3}
\end{equation*}
$$

The translation velocity $V_{t}$ enters through the surface boundary conditions. Suppose that the vortex is translating in the positive $x$-direction with constant speed $V_{t}$. Using a viscous surface stress formulation similar to Smith $(1968,2003)$ and Kepert $(2001)$, the conditions to be satisfied at the surface boundary $(Z=0)$ are:

$$
\begin{align*}
& U+V_{t} \cos \theta=\frac{K}{C_{D} V_{g r}} \frac{\partial U}{\partial Z}  \tag{a}\\
& V-V_{t} \sin \theta=\frac{K}{C_{D} V_{g r}} \frac{\partial V}{\partial Z}  \tag{b}\\
& W=0 \tag{c}
\end{align*}
$$

where $C_{D}$ is a surface drag coefficient; see for example Rosenthal (1962), Smith (1968, 2003), Kepert (2001), and Kepert and Wang (2001). For $1 / C_{D}=0$, equation (1.4) corresponds to no slip conditions ( $\left.U=-V_{t} \cos \theta, V=V_{t} \sin \theta\right)$ at the surface boundary.

For $R \rightarrow \infty$, the system in equation (1.2) reduces to the classic Ekman BL equations under geostrophic conditions (Kundu and Cohen, 2004):

$$
\begin{align*}
& f\left(V_{g r}-V\right)=K \frac{\partial^{2} U}{\partial Z^{2}} \\
& f U=K \frac{\partial^{2} V}{\partial Z^{2}}  \tag{1.5}\\
& W=\text { const. }
\end{align*}
$$

Denote by $R_{g}$ the distance from the TC center beyond which the geostrophic model in equation (1.5) is approximately valid; say $R_{g} \approx 1000 \mathrm{~km}$ (Smith, 1968; Kundu and Cohen, 2004). Also denote by $K_{M}$ the vertical diffusion coefficient under geostrophic conditions and let $V_{g}=V_{g r}\left(R_{g}\right)$ be the gradient tangential wind for $R=R_{g}$ and $Z_{g}=\left(K_{M} / f\right)^{1 / 2}\left(Z_{\mathrm{g}}\right.$ has the meaning of vertical length scale for the depth of the boundary layer; see below). Then one can write equation (1.2) in dimensionless form, as

$$
\begin{align*}
& R_{o}\left[u \frac{\partial u}{\partial r}+\frac{v}{r} \frac{\partial u}{\partial \theta}+w \frac{\partial u}{\partial z}+\frac{v_{g r}^{2}-v^{2}}{r}\right]+v_{g r}-v=k \frac{\partial^{2} u}{\partial z^{2}}  \tag{a}\\
& R_{q}\left[u \frac{\partial v}{\partial r}+\frac{v}{r} \frac{\partial v}{\partial \theta}+w \frac{\partial v}{\partial z}+\frac{u v}{r}\right]+u=k \frac{\partial^{2} v}{\partial z^{2}}  \tag{b}\\
& \frac{\partial(r u)}{\partial r}+\frac{\partial v}{\partial \theta}+\frac{\partial(r w)}{\partial z}=0 \tag{c}
\end{align*}
$$

and equations (1.3) and (1.4) become

$$
\begin{align*}
& \frac{\partial u}{\partial z}=\frac{\partial v}{\partial z}=u=0, \text { and } v=v_{g r}, \text { at } z \rightarrow \infty  \tag{a}\\
& u+v_{t} \cos \theta=\frac{\alpha k}{v_{g r}} \frac{\partial u}{\partial z}, v-v_{t} \sin \theta=\frac{\alpha k}{v_{g r}} \frac{\partial v}{\partial z}, \text { and } w=0, \text { at } z=0 \tag{b}
\end{align*}
$$

In equations (1.6) and (1.7), $R_{o}=V_{g} /\left(R_{g} f\right)$ is the Rossby number and $r=R / R_{g}, z=Z / Z_{g}, v_{t}=$ $V_{t} / V_{g}, u=U / V_{g}, v=V / V_{g}, w=\left(W R_{g}\right) /\left(V_{g} Z_{g}\right), v_{g r}=V_{g r} / V_{g}, k=K / K_{M}$, and $\alpha=K_{M} /\left(C_{D} Z_{g} V_{g}\right)$ are dimensionless quantities. In his model for axi-symmetric vortices, Smith (1968) allows $K$ to be different from $K_{M}$ and to vary radially. In the present extension to moving vortices, we further allow $K$ to vary azimuthally.

Vertical integration of equation (1.6) under boundary conditions (1.7.a) gives

$$
\begin{align*}
& R_{o}\left[\frac{\partial^{2}}{\partial r} \int_{0}^{\infty} r u^{2} d z+\frac{\partial^{2}}{\partial \theta} \int_{0}^{\infty} u v d z+\int_{0}^{\infty} v_{g r}^{2}-v^{2} d z\right]+r \int_{0}^{\infty} v_{g r}-v d z=-\left.k r \frac{\partial u}{\partial z}\right|_{z=0}  \tag{a}\\
& R_{o}\left[\frac{\partial}{\partial r} \int_{0}^{\infty} r^{2} u v d z+r \frac{\partial}{\partial \theta} \int_{0}^{\infty} v^{2} d z+r^{2} w_{\infty} v_{g r}\right]+r^{2} \int_{0}^{\infty} u d z=-\left.k r^{2} \frac{\partial v}{\partial z}\right|_{z=0}  \tag{b}\\
& w_{\infty}=-\frac{1}{r}\left[\frac{\partial}{\partial r}\left(r \int_{0}^{\infty} u d z\right)+\int_{0}^{\infty} \frac{\partial v}{\partial \theta} d z\right] \tag{c}
\end{align*}
$$

where $w_{\infty}$ is the dimensionless vertical wind velocity at $Z \rightarrow \infty$. Next we discuss how equation (1.8) is solved under the conditions (1.7.b).

### 1.3.1 Momentum integral method

Similar to Smith (1968) we take the boundary layer thickness to be proportional to $Z_{g}=\left(K_{M} / f\right)^{1 / 2}$, with proportionality coefficient $\delta$. In Smith's (1968) axi-symmetric formulation $\delta$ exhibits only radial variation, but in the present case of moving cyclones we allow $\delta$ to vary also azimuthally.

Define $\eta=Z J\left[Z_{g} \delta(r, \theta)\right]=z / \delta(r, \theta)$ and notice that the geostrophic model in equation (1.5) is satisfied for $\delta=\sqrt{2}$; see e.g. Kundu and Cohen (2004). Following the derivations of Schlichting (1960), Mack (1962) and Smith (1968), but allowing azimuthal dependence of $u$ and $v$, the solution of equation (1.8) can be approximated as

$$
\begin{align*}
& u(r, \theta, \eta)=E(r, \theta) \Psi(r, \theta, \eta)  \tag{1.9}\\
& v(r, \theta, \eta)=\Omega(r, \theta, \eta) \tag{b}
\end{align*}
$$

where $E(r, \theta)$ is an unknown function, usually referred to as the amplitude coefficient, and $\Psi$ and $\Omega$ are such that $(u, v)=(\Psi, \Omega)$ satisfy equations (1.6) and (1.7) under geostrophic conditions ( $R>R_{g}$ ). After some algebra one obtains

$$
\begin{align*}
& \Psi(r, \theta, \eta)=g(r, \theta, \eta) v_{t} \cos \theta+f(r, \theta, \eta)\left(v_{g r}-v_{t} \sin \theta\right)-v_{t} \cos \theta \\
& \Omega(r, \theta, \eta)=g(r, \theta, \eta)\left(v_{g r}-v_{t} \sin \theta\right)-f(r, \theta, \eta) v_{t} \cos \theta+v_{t} \sin \theta \tag{1.10}
\end{align*}
$$

where

$$
\begin{align*}
& f(r, \theta, \eta)=-e^{-\eta}\left[a_{1}(r, \theta) \sin \eta+a_{2}(r, \theta) \cos \eta\right] \\
& \mathrm{g}(r, \theta, \eta)=1-e^{-\eta}\left[a_{1}(r, \theta) \cos \eta+a_{2}(r, \theta) \sin \eta\right] \tag{1.11}
\end{align*}
$$

The parameters $a_{1}(r, \theta)$ and $a_{2}(r, \theta)$ in equation (1.11) are calculated so that ( $u, v$ ) in equation (1.9) satisfy condition (1.7.b). This gives

$$
\begin{equation*}
a_{2}(r, \theta)=\frac{L_{3} L_{4}-L_{1} L_{6}}{L_{2} L_{4}-L_{1} L_{5}}, \quad a_{1}(r, \theta)=\frac{L_{3}-L_{2} a_{2}}{L_{1}} \tag{1.12}
\end{equation*}
$$

where

$$
\begin{align*}
& L_{1}=v_{g r}-v_{t} \sin \theta+\left(1+\frac{2 k \alpha}{v_{g r} \delta}\right) v_{t} \cos \theta, \quad L_{2}=v_{g r}-v_{t} \sin \theta-\left(1+\frac{2 k \alpha}{v_{g r} \delta}\right) v_{t} \cos \theta \\
& L_{3}=v_{g r}-v_{t} \sin \theta+\frac{v_{t} \cos \theta}{E}, \quad L_{4}=\left(v_{g r}-v_{t} \sin \theta\right)\left(1+\frac{2 k \alpha}{v_{g r} \delta}\right)-v_{t} \cos \theta  \tag{1.13}\\
& L_{5}=\left(v_{t} \sin \theta-v_{g r}\right)\left(1+\frac{2 k \alpha}{v_{g r} \delta}\right)-v_{t} \cos \theta, \quad L_{6}=v_{g r}-v_{t} \sin \theta-\frac{v_{t} \cos \theta}{E}
\end{align*}
$$

The parameters $a_{1}$ and $a_{2}$ are constants independent of $r$ and $\theta\left(a_{1}=1\right.$ and $\left.a_{2}=0\right)$ only for a stationary TC and $\alpha=0$ (i.e. $C_{D} \rightarrow \infty$, no slip conditions). Hence, Smith's (1968) axi-symmetric formulation, where $a_{1}$ and $a_{2}$ are assumed constant independent of $r$, is theoretically correct only in the case of no slip conditions at the surface boundary.

By combining equations (1.8)-(1.11) and after some algebra, one obtains the following system of differential equations in $E$ and $\delta$ :

$$
\begin{align*}
& B_{1}(r, \theta) \frac{\partial E}{\partial r}+B_{2}(r, \theta) \frac{\partial E}{\partial \theta}+B_{3}(r, \theta) \frac{\partial \delta}{\partial r}+B_{4}(r, \theta) \frac{\partial \delta}{\partial \theta}=-B_{5}(r, \theta)  \tag{a}\\
& C_{1}(r, \theta) \frac{\partial E}{\partial r}+C_{3}(r, \theta) \frac{\partial \delta}{\partial r}+C_{4}(r, \theta) \frac{\partial \delta}{\partial \theta}=C_{5}(r, \theta) \tag{b}
\end{align*}
$$

with coefficients

$$
\begin{align*}
& B_{1}=r \delta R_{o} A_{10}, B_{2}=\delta R_{o} A_{12}, B_{3}=r R_{o} A_{1}, B_{4}=R_{o} A_{2} \\
& B_{5}=\frac{k r}{\delta} A_{6}+\delta R_{o}\left(r A_{5} / R_{o}+A_{3}+A_{13}+A_{1}+r A_{11}\right), C_{1}=r^{2} R_{o} \delta\left(A_{12}-v_{g r} A_{16}\right) \\
& C_{3}=r^{2} R_{o}\left(A_{2}-v_{g r} A_{7}\right), C_{4}=r R_{o}\left(A_{4}+A_{19}-v_{g r} A_{8}\right)  \tag{1.15}\\
& C_{5}=-\frac{k r^{2}}{\delta} A_{9}+r R_{o} \delta\left[v_{g r}\left(A_{18}+r A_{17}+A_{7}-A_{20}\right)-r A_{7} / R_{0}-A_{15}-r A_{14}-2 A_{2}\right]
\end{align*}
$$

Analytical expressions for the parameters $A_{1}-A_{20}$ in equation (1.15) are derived in Appendix A.
The nonlinear system in equation (1.14) can be integrated numerically to obtain $E(r, \theta)$ and $\delta(r, \theta)$. For this purpose we use a scheme that is implicit in $\theta$ and explicit in $r$. Integration starts at $r=1$ where we set $E=1$ and $\delta=\sqrt{2}$ (these are the values under geostrophic conditions) and moves inward using a stepwise integration procedure. At each step $i$ in $r$, integration with respect to $\theta$ is performed simultaneously for all azimuthal locations using a central difference approach; see e.g. Chapra and Canale (2002).

Since the parameters in equation (1.14) depend on $E$ and $\delta$, a first approximation to the solution at step $i$ is obtained by evaluating all parameters using the values of $E$ and $\delta$ from step $i$ 1. This procedure is iterated until $E$ and $\delta$ at step $i$ converge.

After $E$ and $\delta$ (and hence the horizontal wind components $U$ and $V$ ) are obtained, the vertical wind velocity $W$ is calculated using mass conservation, as

$$
\begin{equation*}
W(R, \theta, Z)=-\frac{1}{R}\left[\int_{0}^{Z} \frac{\partial(R U)}{\partial R} d Z+\int_{0}^{Z} \frac{\partial V}{\partial \theta} d Z\right] \tag{1.16}
\end{equation*}
$$

In what follows, we refer to equations (1.6) and (1.7) and their solution by the momentum integral method as the Modified Smith (MS) model. In the next section we compare results for a specific storm with wind estimates from the Shapiro (1983) and Kepert (2001) formulations and with MM5 simulations, and in Section 1.5 we examine how the winds generated by the MS model vary with a number of storm parameters.

### 1.4 Model Comparison

First we illustrate the results from the MS model and then compare the performance of that model with others for stationary and moving tropical cyclones.

Figure 1.1 shows the storm-relative radial and tangential wind velocity fields $U$ and $V$ and the vertical wind velocity $W$ at elevations $Z=0,0.5,1$ and 2 km obtained using the MS model. The tropical cyclone translates eastward in the northern hemisphere, with velocity $V_{t}=5 \mathrm{~m} / \mathrm{s}$. Asymptotically as $Z \rightarrow \infty$, the tangential winds satisfy Holland's profile in equation (1.1) with parameters $V_{\max }=50 \mathrm{~m} / \mathrm{s}, R_{\max }=40 \mathrm{~km}$ and $B=1.6$. Results are for non-slip conditions at the surface boundary $\left(1 / C_{D}=0\right)$ and constant vertical diffusion coefficient $K=K_{M}=50 \mathrm{~m}^{2} / \mathrm{s}$. This value of $K$ is often quoted in the literature (e.g. Smith, 1968; Shapiro, 1983; Kepert, 2001;

Kepert, 2006b) and is close to values extracted from MM5 simulations (M. Desflots 2006, personal communication). The Coriolis parameter $f$ is set to $510^{-5} \mathrm{sec}^{-1}$, which corresponds to a latitude $\varphi$ of $20^{\circ}$ North.


Figure 1.1: MS boundary layer solution for the tangential $V$, radial $U$ and vertical $W$ wind velocity fields at altitudes $\mathrm{Z}=0,0.5,1$, and $2 \mathrm{~km} . V$ and $U$ are velocities relative to the moving vortex. The location of the maximum of $U$ is denoted by an x symbol. The tropical cyclone translates eastwards (to the right) with velocity $V_{t}=5 \mathrm{~m} / \mathrm{s}$. All figures are generated under non-slip conditions at the surface boundary and using a constant vertical diffusion coefficient $K=50 \mathrm{~m}^{2} / \mathrm{s}$. Other parameters are $V_{\max }=50 \mathrm{~m} / \mathrm{s}, R_{\max }=40 \mathrm{~km}$ and $B=1.6$.

The model reproduces the conditions $U l_{\mathrm{Z}=0}=-V_{t} \cos \theta, V l_{\mathrm{Z}=0}=V_{t} \sin \theta$, and $W_{\mathrm{Z}=0}=0$ at the surface boundary; see also discussion of Figure 1.2 below. Translation of the tropical cyclone causes intensification of the radial fluxes at the right and right-front of the vortex. Specifically, the maximum of $U$ (x symbol in Figure 1.1) is located at the right/right-front of motion and close to the vortex center for $Z=500 \mathrm{~m}$, and moves outward while rotating clockwise as $Z$ increases. Intensification of $V$ is at the left-front of the vortex, with an asymmetry that decreases as one approaches gradient balance. Due to radial convergence at different altitudes, the vertical velocity $W$ increases monotonically with $Z$. In addition, storm translation causes $W$ to intensify at the right and right-front of the vortex. Similar qualitative findings on $U, V$ and $W$ have been reported by Kepert (2001) and Kepert and Wang (2001).

For the same storm, Figure 1.2 compares the vertical profiles of the azimuthally averaged radial and tangential winds generated by the MS and Kepert (2001) models. Azimuthal averaging produces results that correspond to a stationary, and hence, axi-symmetric TC. The profiles shown in Figure 1.2 are at radial distances $R=R_{\max }=40 \mathrm{~km}$ and $R=2.5 R_{\max }=100 \mathrm{~km}$ from the vortex center.


Figure 1.2: Comparison of the axi-symmetric component of the wind field from Kepert's (2001) model and the MS model. Vertical profiles of (a) the radial wind velocity $U$, and (b) the tangential wind velocity $V$ at 40 km and 100 km from the vortex center. Same storm as in Figure 1.1.

As was pointed out in Section 1.2, the linearized stress formulation of Kepert (2001) is accurate only for small drag coefficients $C_{D}$. In the present case $C_{D} \rightarrow \infty$ and Kepert's (2001) model correctly reproduces the non-slip condition $\left.U\right|_{\mathrm{Z}=0}=0$ in the radial direction (see Figure 1.2.a) but fails to reproduce the condition $V_{Z=0}=0$ in the tangential direction (Figure 1.2.b). This inconsistency results in underestimation of the frictional stresses at the surface boundary and hence of the radial fluxes relative to the MS formulation; see Figure 1.2.a and discussion of Figure 1.3 below.

Kepert's (2001) model also fails to reproduce the variation of the radial wind $U$ with $R$. Observations using GPS dropsonde data (Kepert 2006a,b) show that the maximum of $U$ increases as one approaches the center of the TC. While the MS model is consistent with this observation, the maximum values of $U$ at 40 and 100 km from Kepert's (2001) approach are about the same.

In both the MS and Kepert (2001) models, the depth of the boundary layer $H$, defined as the height $Z$ where $U \approx 0$, increases with increasing $R$. In the MS solution, $H$ is about 2.2 km at $R=$ 100 km and about 1.5 km at $R=40 \mathrm{~km}$, whereas the corresponding values from Kepert's (2001) analysis are 1.5 and 0.8 km . Both sets of estimates are order-of-magnitude correct, with the MS values been closer to observations (e.g. Frank, 1977 and Kepert, 2006a).

Figures 1.3 and 1.4 compare the vertically averaged radial and tangential winds and the vertical winds obtained using the Shapiro (1983), Kepert (2001) and MS models as well as MM5 simulations. Figure 1.3 shows the radial variation of the axi-symmetric component, whereas Figure 1.4 includes the asymmetry due to motion through contour plots. The reasons for vertically averaging $U$ and $V$ (over a depth of 1 km ) are to ease model comparison, since Shapiro (1983) treats the BL as a slab of constant depth $H=1 \mathrm{~km}$, and to reduce the effect of vertical fluctuations in the MM5 solution. All storm parameters are the same as in Figure 1.1, except for the drag coefficient, which is set to 0.003 . This value of $C_{D}$ is close to values extracted from MM5 simulations for oversea conditions (M. Desflots 2006, personal communication), does not introduce significant distortions in Kepert's (2001) linear formulation, and does not cause numerical oscillations in Shapiro's (1983) approach. Also, values of $C_{D}$ close to 0.003 are often quoted in the literature for wind speeds in the range encountered in hurricanes (e.g. Large and Pond, 1981; Powell et al., 2003; Donelan et al., 2004).


Figure 1.3: Comparison of the Shapiro (1983), Kepert (2001) and MS boundary layer solutions with MM5 simulations for an axi-symmetric (stationary) TC. (a) Vertically averaged radial wind velocity $\bar{U}$, (b) vertically averaged tangential wind velocity $\bar{V}$, and (c) vertical wind velocity $W$ at an altitude of 1 km . All figures are generated using a constant vertical diffusion coefficient $K=50 \mathrm{~m}^{2} / \mathrm{s}$ and a surface drag coefficient $C_{D}=0.003$ and averaging is over a depth of 1 km . Other parameters are $V_{\max }=50 \mathrm{~m} / \mathrm{s}, R_{\max }=40 \mathrm{~km}$ and $B=1.6$.

In general, the MS model predictions are close to the MM5 simulations, except that in MM5 the vertical velocities are more peaked near $R_{\max }$ (mainly due to rainband effects, which the BL models do not resolve) and the contour plots are more erratic due to local fluctuations. Also the Kepert (2001) model reproduces well the vertically averaged tangential winds of MM5, but that model severely underestimates the radial and vertical flows, especially in the near-core region. This is due to inaccuracy of Kepert's linearization at radial distances smaller than 2-3 $R_{\max }$; see Section 1.2. For example, flight observations show that, in the vicinity of $R_{\max }, W$ is in the range $0.5-3 \mathrm{~m} / \mathrm{s}$ or higher (Willoughby et al., 1982; Jorgensen, 1984a, b; Black et al., 2002), whereas

Kepert's (2001) estimates are around 0.1-0.4 m/s. In the far field ( $R>2.5 R_{m a x}$ ), Kepert's (2001) model and MM5 produce vertical velocities that are similar and in good agreement with observations (Jorgensen 1984a, b; Black et al. 2002).


Distance from storm center (km)
Figure 1.4: Comparison of results form the Kepert (2001), Shapiro (1983) and MS models with MM5 simulations for a TC that translates eastwards (to the right) with velocity $V_{t}=5 \mathrm{~m} / \mathrm{s}$. Vertically averaged radial wind velocity $\bar{U}$ (left column), vertically averaged tangential wind velocity $\bar{V}$ (middle column) and vertical wind velocity $W$ at an altitude of 1 km (right column). $\bar{V}$ and $\bar{U}$ are relative to the moving vortex and averaging is over a depth of 1 km . All other parameters are the same as in Figure 1.3.

Shapiro's (1983) model has an opposite behavior in the near-core region, where it overpredicts the radial and vertical flows by a factor of about 2 . This is consistent with the finding by Anthes (1971) for slab-layer models; see Section 1.2.

Overall, we find that the MS model produces realistic estimates of the tangential, radial and vertical wind velocities and is numerically more stable and more accurate than the boundary layer models of Kepert (2001) and Shapiro (1983). Next we use the MS model to study the sensitivity of the velocity fields to various storm parameters.

### 1.5 Sensitivity Analysis

Figures $1.5,1.6$ and 1.7 show the sensitivity of MS model results to tropical cyclone characteristics: the tangential wind speed under gradient balance [parameterized here in terms of $V_{\max }, R_{\max }$ and $B$; see equation (1.1)], the vertical diffusion coefficient $K$, the surface drag coefficient $C_{D}$, and the translation velocity $V_{t}$.

Figure 1.5 shows the sensitivity of the azimuthally averaged (axi-symmetric) velocities to $V_{\max }, R_{\max }, B, K$ and $C_{D}$. The base case (solid lines in Figure 1.5) corresponds to the storm used in Figure 1.1. Sensitivity is evaluated by varying the parameters one at a time around their basecase values. $V_{\max }$ affects only the amplitude of the radial profiles, whereas other parameters affect mainly the shape of the profile (e.g. rate of decay with distance) or both. The effects on the three velocity components are however not uniform. For example, $R_{\max }$ has negligible influence on the peak value of the horizontal winds $\bar{U}$ and $\bar{V}$, but affects significantly the peak vertical velocity.

Figure 1.6 shows the effect of $C_{D}$ on the wind fields for a TC that moves eastward with velocity $V_{l}=5 \mathrm{~m} / \mathrm{s}$. Higher values of $C_{D}$ correspond to increased friction at the surface boundary and stronger asymmetry of the three velocity components.

Finally, Figure 1.7 shows sensitivity to the translation velocity $V_{t}$ when the other parameters are fixed at $V_{\max }=50 \mathrm{~m} / \mathrm{s}, R_{\max }=40 \mathrm{~km}, B=1.6, K=50 \mathrm{~m}^{2} / \mathrm{s}$ and $C_{D}=0.003$. One sees that $\bar{U}$ and $W$ intensify in the NE quadrant relative to the direction of motion, whereas $\bar{V}$ intensifies to the left of the velocity vector. The asymmetry increases with increasing $V_{t}$.

Qualitatively similar sensitivity results are obtained when using Kepert's (2001) model (not shown here). This is expected, since Kepert's (2001) model is a linearized variant of the present

MS formulation. However, Kepert's linearization has approximate validity only for large $R, C_{D}$ $\ll 1$ and small $V_{t}$; see Section 1.2.


Figure 1.5: Sensitivity analysis of MS model results to storm parameters for a stationary TC. Vertically averaged (over a 1 km depth) radial wind velocity $\bar{U}$ (left column) and tangential wind velocity $\bar{V}$ (middle column) and vertical wind velocity $W$ at $Z=1 \mathrm{~km}$ (right column). Solid lines correspond to $V_{\max }=50 \mathrm{~m} / \mathrm{s}, R_{\text {max }}=40 \mathrm{~km}, B=1.6$, $K=50 \mathrm{~m}^{2} / \mathrm{s}$ and non-slip conditions at the surface boundary. Each row shows results under perturbation of one parameter.


Distance from storm center (km)
Figure 1.6: Sensitivity of MS results to the drag coefficient $C_{D}$ for a tropical cyclone that translates eastwards (to the right) with velocity $V_{t}=5 \mathrm{~m} / \mathrm{s}$. Vertically averaged radial wind velocity $\bar{U}$ (left column), vertically averaged tangential wind velocity $\bar{V}$ (middle column), and vertical wind velocity $W$ at an altitude of 1 km (right column). $\bar{V}$ and $\bar{U}$ are relative to the moving vortex and averaging is over a depth of 1 km . Other parameters are the same as in Figure 1.4.


Distance from storm center (km)
Figure 1.7: Sensitivity of MS model results to the translation velocity $V_{t}$. Notation and parameters are the same as in Figure 1.6, except for $C_{D}=0.003$.

### 1.6 Conclusions

We developed a simple theoretical model for the boundary layer (BL) of translating tropical cyclones (TCs). The model estimates the tangential, radial and vertical wind fields from a given radial profile of the tangential wind speed under geostrophic conditions [e.g. Holland's profile in equation (1.1)] and surface stress boundary conditions.

The model is based on Smith's (1968) formulation, which is corrected for the general case of a stress surface boundary (see Section 1.3) and modified to account for motion-induced asymmetries. The governing equations are solved using the momentum integral method, resulting in a very efficient computational scheme (less than a minute to run on a conventional computer). Contrary to the Shapiro (1983) and Kepert (2001) BL formulations, the present modified Smith (MS) model is stable over a wide range of storm parameters and produces results that are in good agreement with MM5 simulations and observations; see Figures 1.5, 1.6 and 1.7.

In a parametric analysis, we have examined how the symmetric wind components and motion-induced asymmetries generated by the MS model depend on various storm characteristics. For TCs that translate in the northern hemisphere, the vertical and storm-relative radial wind velocity fields intensify at the right-front quadrant of the vortex, whereas the stormrelative tangential winds intensify at the left-front of the storm. The asymmetry is higher for faster-moving TCs. These findings are in accordance with field observations; see for example Shapiro (1983).

As the intensity of the TC (expressed through the maximum tangential velocity at gradient balance $V_{\max }$ ) increases, the horizontal and vertical velocity components also increase. Larger values of Holland's $B$ parameter and lower values of the radius of maximum winds $R_{\max }$ produce horizontal and vertical wind profiles that are more picked and concentrated closer to the TC center.

The effect of the frictional drag coefficient $C_{D}$ is more complex. As $C_{D}$ increases, the frictional drag force also increases resulting in lower tangential winds and higher radial convergence (stronger vertical winds). Another effect of increasing $C_{D}$ is higher asymmetry of the velocity field.

## 2. Theoretical Model of Rainfall in Tropical Cyclones for the Assessment of Long-Term Risk

### 2.1 Introduction

Tropical cyclones (TCs) are atmospheric disturbances capable of producing extreme rainfall with devastating social and economic impact (Landsea, 2000; Rappaport, 2000). Consequently, there is much interest in the assessment of the rainfall hazards posed by TCs, either in real time (with leads of hours or days) or in the long run; see e.g. Marks et al. (1998). For the latter purpose, when interest is in the rate at which different rainfall intensity levels are exceeded, one needs to parameterize the storms and for each set of parameters evaluate rainfall at the site or over the region of interest as a random process in time or a random field in space-time. In principle, the stochastic rainfall model could be directly fitted to data from historical events, but the large number of parameters and the relative lack of historical data make an empirical model identification and fitting approach unfeasible. Moreover, it would be difficult in such an approach to incorporate knowledge of the physics of the phenomenon. A better approach, which we follow here, is to formulate a physically-based rainfall model. The model should be simple enough that it can be run under a very large set of scenario conditions; hence detailed numerical TC models would not be suited for this purpose.

Neither simple nor sophisticated TC models can produce accurate statistical estimates of space-time rainfall for a given set of global TC parameters. Therefore, any deterministic rainfall model must be complemented by a statistical representation of the rainfall "residuals", defined as the difference between observed rainfall and model prediction. For example, the model developed here ignores the rainfall fluctuations due to rainbands and local convection. The statistical characterization of these fluctuations (residuals) is discussed in Chapter 3.

The third and final component of a long-term TC rainfall risk analysis method is the recurrence model, which specifies the frequency with which different TC parameter combinations occur in the region of interest. This component has been the subject of numerous studies, as the recurrence relation is common to the assessment of any TC-related risk, such as wind, waves and surges; see for example Vickery and Twisdale (1995), Vickery et al. (2000), Willoughby and Rahn (2004) and Powell et al. (2005).

In the late 1950s, R.H. Kraft (as referenced by Pfost, 2000, and Kidder et al., 2005) used raingauge rainfall depths to estimate the maximum $24-\mathrm{hr}$ rainfall accumulation due to the passage of a TC. According to Kraft, this maximum is 100 inches $(254 \mathrm{~cm})$ divided by the storm translation speed in knots ( $1 \mathrm{knot}=0.514 \mathrm{~m} / \mathrm{s}$ ). Limitations of Kraft's analysis are that it does not provide information on the spatial distribution of rainfall and does not account for TC characteristics such as size and intensity.

Riehl and Malkus (1961), Goodyear (1968) and more recently Simpson and Riehl (1981) have addressed some of these limitations. From the examination of 46 TCs making landfall along the Gulf Coast of the United States, Goodyear (1968) concluded that the $48-\mathrm{hr}$ maximum rainfall depth is about 150 mm and occurs $40-80 \mathrm{~km}$ inland and $40-80 \mathrm{~km}$ to the right of the storm. Using a similar approach, Riehl and Malkus (1961) and Simpson and Riehl (1981) found that for hurricane-strength cyclones rainfall intensity averages about $33 \mathrm{~mm} / \mathrm{h}$ within 37 km from the cyclone center and for larger distances decays almost exponentially. While these studies extend and improve upon Kraft's rule, they too fail to resolve the dependence of rainfall on storm characteristics.

NASA's Tropical Rainfall Measuring Mission (TRMM) (Simpson et al., 1988) produced vast amounts of TC rainfall data, making it possible to conduct more systematic statistical analyses. Lonfat et al. (2004) extracted 2121 tropical cyclone microwave images from the TMI TRMM data set to find how the azimuthally averaged rainfall intensity varies with distance $R$ from the TC center in three storm intensity ranges: tropical storms (TSs) with maximum tangential wind speed $V_{\max }$ in the range $18-33 \mathrm{~m} / \mathrm{s}$; CAT12 cyclones with $V_{\max }=34-48 \mathrm{~m} / \mathrm{s}$ and CAT 35 cyclones with $V_{\max }>49 \mathrm{~m} / \mathrm{s}$. The study concluded that TC rainfall intensifies with increasing $V_{\max }$ and the symmetric component of the rainfall intensity reaches its maximum at a distance from the hurricane center close to the radius of maximum winds $R_{\max }$. For larger distances, rainfall intensity decays approximately as a power law; see their Figure 11. Due mainly to storm translation and vertical wind shear, rainfall intensity lacks circular symmetry and varies also with the azimuth relative to the directions of shear and motion.

Chen et al. (2006) used the same TRMM storms to further assess the dependence of rainfall on vertical wind shear $S$, defined as the difference between the 200 and $850-\mathrm{hPa}$ horizontal wind velocities in the annular region between 200 and 800 km from the TC center. The study calculated the average rainfall asymmetry, defined as the ratio of the wavenumber-1 Fourier
amplitude to the azimuthal average of the rainfall intensity, for the nine combinations of the 3 intensity categories in Lonfat et al. (2004) and three shear magnitude ranges ( $S<5 \mathrm{~m} / \mathrm{s}$, $5 \leq S \leq 7.5 \mathrm{~m} / \mathrm{s}$, and $S>7.5 \mathrm{~m} / \mathrm{s}$ ). Chen et al. (2006) found that, in storms in the Northern (Southern) hemisphere with high wind shear ( $S>5 \mathrm{~m} / \mathrm{s}$ ), rainfall intensifies downshear and downshear-left (-right) of the storm.

Parametric rainfall models have also been developed. Using the radial rainfall profiles of Lonfat et al. (2004), Tuleya et al. (2007) suggested one such model for 24-hr rainfall totals (RCLIPER) based on climatological and persistence information. The model assumes that storms are symmetric and therefore ignores vertical wind shear and storm motion. Lonfat et al. (2007) built on the R-CLIPER algorithm to construct a parametric rainfall model (PHRaM) that includes shear-related asymmetries according to the results of Chen et al. (2006).

Due to data limitations, R-CLIPER and PHRaM use a coarse and incomplete storm parameterization: the effects of storm intensity and vertical wind shear are modeled by interpolating from 3 classes of each variable, the size of the vortex $R_{\max }$ is only implicitly taken into account by allowing the location of the maximum rainrate depend on the intensity of the storm according to the results of Lonfat et al. (2004), while other factors (e.g. the radial wind velocity profile in the main vortex, the surface roughness, and the storm translation velocity) are ignored. Another limitation is that the Lonfat et al. (2004) profiles on which R-CLIPER and PHRaM are based use ensemble averages of storms with significantly different $R_{\max }$ values. Since rainfall intensity has a sharp peak near $R_{\max }$, this averaging operation depresses the maximum rainfall estimate. For example, for CAT35 storms Lonfat et al. (2004) find maximum rainfall intensities around $12 \mathrm{~mm} / \mathrm{h}$, which is $2.5-3$ times lower than the values most often reported in the literature; see for example Riehl and Malkus (1961), Jiang et al. (2006), Trenberth et al. (2007) and the rainfall intensities implied by the radar reflectivities in Marks (1985) and Kepert (2006a,b). Finally, the Lonfat et al. (2004) profiles are based on TMI rainfall products, which are known to be biased towards low values for high rainfall intensities and towards high values for low rainfall intensities (Viltard et al., 2006).

Here we develop a simple theoretical model of TC rainfall based on the vertical outflow of water vapor from the TC boundary layer (BL). This water vapor flux originates from the lowlevel convergence of the horizontal flow. The analysis combines a user-specified tangential wind profile at gradient level, an Ekman-type solution for the horizontal and vertical winds inside the
boundary layer (BL), and basic thermodynamics. Evaluation of the BL winds is based on Smith's (1968) axi-symmetric formulation, modified in Chapter 1 to account for storm motion. The resulting models of wind and rainfall are referred to as the modified-Smith (MS) BL model and the modified-Smith-for-rainfall (MSR) model, respectively.

The MSR model produces asymmetric rainfall fields that explicitly depend on: the maximum tangential wind velocity at gradient level $V_{\max }$, the radius of maximum winds $R_{\max }$, Holland's $B$ parameter (Holland, 1980), the surface drag coefficient $C_{D}$, the storm translation velocity $V_{t}$, the vertical diffusion coefficient of the horizontal momentum $K$, and the average temperature $\bar{T}$ and saturation ratio $\bar{Q}$ inside the TC boundary layer.

An important departure from previous studies is that we parameterize asymmetries in terms of storm motion not vertical wind shear. The degree to which TC motion and shear contribute to wind, lightning, and rainfall asymmetries has been intensely discussed in the literature; see for example Black et al. (2002), Corbosiero and Molinari (2002, 2003), Rogers et al. (2003), Lonfat et al. (2004) and Chen et al. (2006). Separation of the two effects through data analysis is made difficult by the high correlation between the directions and magnitudes of motion and shear in any given geographical region (Corbosiero and Molinari, 2003; Lonfat et al., 2004; Chen et al., 2006). As a consequence, the calculated rainfall asymmetry is almost the same when storms are aligned in the direction of motion or shear, except for a region-specific rotation; see e.g. Corbosiero and Molinari (2003) and Section 2.5 below. Another consequence is that, in risk analysis, one may equivalently use shear or motion as conditioning parameter. Since it is easier to include motion than shear when modeling rainfall and the historical records readily provide storm motion information (e.g. Vickery and Twisdale, 1995, and Vickery et al., 2000), we have chosen to develop a motion-based rather than shear-based parameterization of rainfall asymmetry.

Section 2.2 summarizes the boundary layer model developed In Chapter 1 and Section 2.3 uses the vertical fluxes from that model to estimate surface rainrates in the case of stationary (i.e. symmetric) cyclones. Model predictions are compared to MM5 simulations and R-CLIPER rainrate estimates. The choice of MM5 is based on the fact that this code has been successfully used to simulate a number of TCs, including Hurricanes Bonnie (1998) (Rogers et al. 2003, 2007), Floyd (1998) (Tenerelli and Chen, 2001, Rogers et al. 2007) and Frances (2004) (Chen et al., 2007). Section 2.4 validates the symmetric MSR predictions using precipitation radar (PR)
rainfall products from 38 TRMM frames. The PR rainfall products are less biased than the microwave imager (TMI) data used in previous studies, especially in the core region where rainfall intensities are high (Viltard et al., 2006). Section 2.5 extends the analysis to translating TCs, which generate asymmetric rainfall fields, assesses the effect of motion on the spatial variation of TC rainfall, and suggests a motion-based parameterization of rainfall asymmetry. Section 2.6 assesses the sensitivity of the symmetric and asymmetric rainfall components to various TC parameters and Section 2.7 summarizes the main conclusions.

### 2.2 Modified Smith Boundary Layer Model for Moving Tropical Cyclones

A number of studies (Myers and Malkin 1961; Shapiro 1983; Kepert 2001) have developed theoretical boundary layer (BL) models for moving tropical cyclones. These models derive the radial, tangential and vertical winds inside the boundary layer from an assumed radial profile of the tangential wind velocity under gradient balance, $V_{g r}(R)$, and suitable surface boundary conditions. For example, a widely used gradient wind profile is (Holland, 1980)

$$
\begin{equation*}
V_{g r}(R)=V_{\max } \sqrt{\left(R_{\max } / R\right)^{B} \exp \left[1-\left(R_{\max } / R\right)^{B}\right]} \tag{2.1}
\end{equation*}
$$

where $V_{\max }, R_{\max }$, and $B$ are TC-specific parameters. According to equation (2.1), the tangential velocity $V_{g r}$ increases radially to a maximum $V_{\max }$ at $R=R_{\max }$ and for $R \gg R_{\max }$ decays approximately as a power-law of distance with exponent $-B / 2$. The shape parameter $B$ varies in the range [1, 2], with typical values around 1.4 (Willoughby and Rahn, 2004). Next we briefly describe the boundary layer model developed in Chapter 1 and in Sections 2.3-2.5 use this model to calculate water vapor fluxes that are responsible for rainfall.

The model in Chapter 1 corrects Smith's (1968) BL formulation for the case of stress surface boundary conditions and accounts for storm motion. Like in Smith (1968), vertical diffusion of the horizontal momentum is parameterized through a vertical diffusion coefficient $K$. The horizontal momentum equations are written in cylindrical coordinates that move with the storm and solved using the Karman and Pohlhausen momentum integral method. In this method, one specifies vertical profiles for the radial $U$ and tangential $V$ wind velocity components, which satisfy the boundary conditions at the surface (elevation $Z=0$ ) and for $Z \rightarrow \infty$ tend to the gradient winds, for example the profile in equation (2.1). The boundary conditions are modeled using a surface stress formulation with drag coefficient $C_{D}$.

For $U$ and $V$, we use functions of the Ekman type with parameters $E$ (amplitude coefficient) and $\delta$ (dimensionless BL scale thickness) that vary both radially and azimuthally. The horizontal momentum equations are vertically integrated through the BL to produce a system of two partial differential equations, which are solved numerically to obtain $E$ and $\delta$ as functions of radius $R$ and azimuth $\theta$ relative to the direction of storm motion. Once the horizontal wind components $U$ and $V$ are obtained, the vertical wind velocity $W$ is calculated using mass conservation, as

$$
\begin{equation*}
W(R, \theta, Z)=-\frac{1}{R}\left[\int_{0}^{Z} \frac{\partial(R U)}{\partial R} d Z+\int_{0}^{Z} \frac{\partial V}{\partial \theta} d Z\right] \tag{2.2}
\end{equation*}
$$

For stationary cyclones $\left(V_{t}=0\right)$, there is no azimuthal variation of $V$ and $U$ and equation (2.2) reduces to

$$
\begin{equation*}
W(R, Z)=-\frac{1}{R} \frac{d}{d R}\left(R \int_{0}^{Z} U d Z\right) \tag{2.3}
\end{equation*}
$$

$W(R, Z)$ in equation (2.3) is also the symmetric component of the vertical wind speed for a storm that translates with velocity $V_{t} \neq 0$.

The above modified Smith (MS) scheme is computationally very efficient and stable over a wide range of parameter values. Model predictions are close to MM5 simulations and to observed wind speeds; see Chapter 1 for details.

### 2.3 Estimation of the Symmetric Component of Rainfall

Estimates of rainfall intensity are obtained assuming that, with corrections to be made later, the surface rain rate $i$ is proportional to the water vapor up-flux at a reference height $H$. Similar approaches have been used in the past to evaluate the rainfall potential of extra-tropical storms (Palmen, 1958), orographic precipitation (Alpert, 1986) and latent heat (Magaki and Barros, 2004), as well as to predict rainfall extremes (Abbs, 1999; Wilson and Toumi, 2005).

To verify how strongly rainfall intensity is related to the vertical velocity $W_{H}(R, \theta)=$ $W(R, \theta, Z=H)$ from equation (2.2) at different elevations $H$, we used MM5 simulations. Figure 2.1 shows the correlation between the two quantities using 12 frames of Hurricane Frances, simulated at 6 hr intervals for the period Aug. 29-Sep. 01, 2004. The correlation is maximum around 0.85 at an elevation of $2-3 \mathrm{~km}$, which can be taken as the reference height $H$. The inset of

Figure 2.1 compares the MM5 radial profiles of the simulated rainfall intensity and vertical wind velocity at 3 km elevation for the 06:00UTC Aug. 29, 2004 frame. Both profiles are normalized to have unit maximum value. This detailed comparison shows that the correlation coefficient is below 1 due mainly to fluctuations of the rainfall intensity caused by rainbands and other local convective phenomena. If these fluctuations in the MM5 profiles are smoothed out, which is what the present MSR model effectively does, the surface rainfall intensity and vertical wind speed are in even better agreement.


Figure 2.1: (a) Ensemble correlation function of the vertical wind velocity at different elevations and the surface rainfall intensity from MM5 simulations of Hurricane Frances. Ensemble averaging is over 12 frames (at 6 hr intervals) during the period Aug. 29-Sep. 01, 2004. (b) Normalized radial profiles of surface rainfall intensity and vertical wind velocity on Aug. 29, 2004 at 06:00UTC at 3 km elevation.

To complete the symmetric rainfall model one needs the proportionality constant between rainfall intensity and vertical wind speed. From simple calculations using a lapse-rate of about 6$7^{\circ} \mathrm{C} / \mathrm{km}$ (Rogers and Yau, 1996), one obtains that at elevations in excess of $6-8 \mathrm{~km}$ the water vapor mixing ratio is close to zero. Consequently, one may accurately assume that the upward water vapor flux from the TC boundary layer equals the downward flux of rainwater. To keep the rainfall model simple, we assume that below the reference height $H$ the temperature $T$ and saturation ratio $Q$ are constant and equal to the depth-averaged values $\bar{T}$ and $\bar{Q}$. For cyclones over tropical and sub-tropical waters, $\bar{T}$ ranges between $20-24^{\circ} \mathrm{C}$ and $\bar{Q}$ is between $75-85 \%$; see Gray et al. (1975), Frank (1977) and Smith (2003). Under these conditions, the symmetric rainfall intensity $i_{\text {sym }}$ is given by

$$
i_{s y m}(R)= \begin{cases}\alpha(\bar{T}) \bar{Q} W_{H}(R) & , W_{H}(R)>0  \tag{2.4}\\ 0 & , W_{H}(R) \leq 0\end{cases}
$$

where $\alpha(\bar{T})$ is the volume of liquid water per unit volume of saturated air after complete condensation (see below), and $W_{H}(R)=W(R, Z=H)$ is the vertical wind velocity in equation (2.3) for $Z=H$. The function $\alpha(\bar{T})$ is obtained by combining the ideal gas law with the ClausiusClapeyron equation. Using a liquid water density $\rho_{w}=1000 \mathrm{kgr} / \mathrm{m}^{3}$, this gives

$$
\begin{equation*}
\alpha(\bar{T})=\frac{1.32410^{-3}}{\bar{T}+273} \exp \left(\frac{17.67 \bar{T}}{\bar{T}+243.5}\right) \tag{2.5}
\end{equation*}
$$

where $\bar{T}$ is in ${ }^{\circ} \mathrm{C}$. Notice that in downdraft regions where $W_{H}$ is negative, equation (2.4) sets the rainfall intensity to zero. This means that rainfall generation is limited to regions where moist air updrafts. However, due to the slant of the wall updrafts and the cyclonic advection, rainfall may be nonzero also in downdraft regions. This effect is modeled below through a rainfall redistribution scheme.

### 2.3.1 Correction for the sloping angle of the wall

Flight observations (e.g. Jorgensen, 1984b; Marks and Houze, 1984) show that the wall updraft of a tropical cyclone slopes outward to altitudes $H_{0} \approx 5-7 \mathrm{~km}$, with an angle $\psi_{0}$ from the vertical in the $45^{\circ}-60^{\circ}$ range. The MS model in Chapter 1 assumes fixed vertical profiles of the radial and tangential wind velocities and therefore does not account for such sloping angle. Consequently, equation (2.4) tends to underpredict the radius of maximum rainfall.

To include radial advection of the rainwater by the wall updraft while avoiding discontinuities in the radial distribution of rainfall, we assume that the angle of the updrafts decreases exponentially with distance $R$ from the storm center, as

$$
\begin{equation*}
\psi(R)=\psi_{0} \exp \left(-\frac{\left|R-R_{m}\right|}{R_{m}}\right) \tag{2.6}
\end{equation*}
$$

where $R_{m}$ is the location where $i_{s y m}$ and $W_{H}$ in equation (2.4) are maximum. The outward radial displacement $\Delta R$ of the rainwater due to the sloping updrafs is then

$$
\begin{equation*}
\Delta R=H_{0} \tan \psi \tag{2.7}
\end{equation*}
$$

Notice that estimating rainfall intensities at distance $R$ from the cyclone center as $i_{\text {sym }}(R-\Delta R)$ is technically incorrect because the model does not satisfy mass conservation. However, we have verified that the error is very small and negligible in practice.

### 2.3.2 Comparison with MM5 and R-CLIPER

Figure 2.2 compares the azimuthally averaged rainfall intensities $i_{\text {sym }}$ for Hurricane Frances (2004) estimated by MM5, R-CLIPER (see Introduction), and the present modified-Smith-forrainfall (MSR) model. The MM5 and MSR curves are the ensemble averages of 12 rainfields simulated at 6 hr intervals during the period Aug. 29-Sep. 01, 2004, using the two models.


Figure 2.2: Comparison of the ensemble average rainrates for Hurricane Frances 2004 during the period Aug. 29Sep. 01, produced by the MSR, MM5 and R-CLIPER rainfall models.

The MM5 simulations where conducted at 1.67 km resolution using the nested grid capability at the University of Miami (Houze et al., 2006; 2007), whereas the MSR estimates were obtained as follows:

1) For each frame, the parameters $V_{\max }$ and $R_{\max }$ in equation (2.1) were extracted from the azimuthally averaged tangential winds simulated by MM5 at 5 km elevation;
2) Holland's (1980) gradient wind profile with $B=1$ was used in the MS model in Chapter 1 to calculate the vertical wind profile $W_{H}(R)$ at elevation $H=3 \mathrm{~km}$;
3) Equations (2.4) and (2.5) were used to estimate how the azimuthally averaged rainfall intensity $i_{\text {sym }}$ varies with distance $R$ from the TC center;
4) Finally, the results were corrected for sloping-updrafts using equations (2.6) and (2.7) and averaged over the 12 frames.

Setting Holland's $B$ to 1 reproduces well the MM5 rainfall fields, as well as the PR rainfall estimates from TRMM; see Section 2.4.

The MS model in Chapter 1 requires also specification of the Coriolis parameter $f$, the vertical diffusion coefficient $K$, and the surface drag coefficient $C_{D}$. In our simulations we have set $f=$ $4.710^{-5} \mathrm{sec}^{-1}$, which corresponds to latitude $19^{\circ}$ North (the approximate latitude of TC Frances during the period considered), $K=50 \mathrm{~m}^{2} / \mathrm{s}$, and $C_{D}=0.002$. Values of $K$ near $50 \mathrm{~m}^{2} / \mathrm{s}$ are often quoted in the literature (e.g. Smith, 1968; Shapiro, 1983; Kepert, 2001; Kepert 2006b) and are consistent with back-calculations from MM5 simulations (Melicie Desflots, 2007, personal communication). The value 0.002 is representative of drag coefficients extracted from oversea MM5 simulations and to values in the literature for winds in the hurricane range (e.g. Kepert, 2001; Powell et al., 2003; Donelan et al., 2004). The vertically averaged temperature $\bar{T}$ (over a depth of 3 km ) and saturation ratio $\bar{Q}$ in equation (2.4) have been set to $22^{\circ} \mathrm{C}$ and $80 \%$, respectively. These values correspond to a depth-averaged mixing ratio of approximately $13 \mathrm{gr} / \mathrm{kg}$, which is slightly lower than the ensemble average value of $15 \mathrm{gr} / \mathrm{Kg}$ extracted from MM5 simulations for Hurricane Frances (Melicie Desflots, 2007, personal communication). For the wall updraft correction in equations (2.6) and (2.7), we have assumed an outwards slope of $\psi_{0}=$ $50^{\circ}$ from the vertical to an altitude $H_{0}=6 \mathrm{~km}$.

The solid lines in Figure 2.2 are the profiles of $i_{\text {sym }}$ before the correction for sloping updrafts (thin lines) and after that correction (thick lines). The rainfall estimates from the MSR model are close in shape and magnitude to the MM5 profiles. This is especially true after the correction for out-sloping updrafts. Differences are mostly due to local rainfall intensifications in MM5 caused by rainbands. By contrast, the rain rates of Lonfat et al. (2004), which form the basis of the RCLIPER algorithm, agree with MM5 in the far field but severely underestimate rainfall in the near-core region. As discussed in the Introduction, reasons for the much-reduced rain rate maximum in R-CLIPER are the smoothing effect of ensemble averaging and the bias of the TMI rainfall retrievals used by Lonfat et al. (2004).

### 2.4 Validation of Symmetric MSR Predictions

Figure 2.3 compares PR and MM5 rainfall estimates with rainfall intensities generated by the present MSR model using the procedure described in Section 2.3. Figure 2.3.a shows a scatterplot of the ratio between the PR and MSR rainfall estimates as a function of the
normalized distance $R / R_{\max }$ from the storm center, using a $5 \mathrm{~km} \times 5 \mathrm{~km}$ grid of spatial locations and the 38 TRMM frames in Table 2.1 (a total number of 48483 points). The number of points in different ranges of $R / R_{\max }$ is shown in Table 2.2.


Figure 2.3: Comparison of PR, MM5 and MSR point rainfall intensity estimates. (a) Scatterplot of the ratio between PR and MSR rainfall estimates as a function of the normalized distance $R / R_{\max }$ from the storm center, for 38 TRMM frames; see Table 2.1. The number of data points in different ranges of $R / R_{\max }$ is shown in Table 2.2. (b) Scatterplot of the ratio between MM5 and MSR rainfall estimates as a function of $R / R_{\max }$, for hurricane Frances 2004 during the period Aug. 29-Sep. 01. (c) Local averages and standard deviation of the ratios in (a) using a moving window of 2000 points. (d) Same as (c) but for the ratios in (b).

The MSR estimates where generated using the $V_{\max }, R_{\max }$ and latitude information in the extended best track record (Demuth et al., 2006; M. DeMaria, 2008; personal communication). Figure 2.3.b shows a similar scatterplot of the ratio between the MM5 and MSR rainfall estimates. In this case the comparison is based on the 12 simulated rainfields of Hurricane Frances, for a total of 43919 points. All MSR simulations were performed using $B=1$, $K=50 \mathrm{~m}^{2} / \mathrm{s}$ and $C_{D}=0.002$. Both Figures 2.3.a and 2.3.b show a large dispersion, which reflects
the significant small-scale variability of rainfall intensity due to rainbands and local convection.
Those fluctuations are not resolved by the MSR model.
Table 2.1: Storm characteristics for the PR-TRMM rainfields used in Figure 2.3. The estimates of $V_{\max }$ and $R_{\max }$ are obtained from the extended best track record (M. DeMaria, 2008; personal communication).

|  | Storm center |  | $\begin{gathered} V_{\text {max }} \\ (\mathrm{m} / \mathrm{s}) \end{gathered}$ | $\begin{aligned} & \boldsymbol{R}_{\text {max }} \\ & (\mathbf{k m}) \end{aligned}$ | Intensity category |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Lat. (deg) | $\begin{gathered} \hline \text { Lon. } \\ \text { (deg) } \end{gathered}$ |  |  |  |
| 㦯 | 21.7 | -61.6 | 48.8 | 41 | CAT2 |
|  | 23.5 | -68.7 | 64.0 | 37 | CAT4 |
|  | 23.7 | -70.6 | 69.3 | 37 | CAT4 |
| $\begin{aligned} & \pm \\ & 0 \\ & 0 \\ & 0 \\ & 0 \\ & \text { Hy } \\ & \text { in } \end{aligned}$ | 12.6 | -43.7 | 23.1 | 37 | TS |
|  | 15.7 | -49.8 | 51.4 | 19 | CAT3 |
|  | 17 | -51.3 | 54.0 | 28 | CAT3 |
|  | 17.9 | -52.6 | 59.1 | 28 | CAT4 |
|  | 19 | -57.3 | 51.4 | 28 | CAT3 |
|  | 21.2 | -68.5 | 61.7 | 28 | CAT4 |
| $\begin{aligned} & \underset{Z}{\Xi} \\ & \vdots \\ & \tilde{Z} \end{aligned}$ | 8.9 | -38.9 | 25.7 | 37 | TS |
|  | 10.7 | -50.6 | 57.5 | 28 | CAT4 |
|  | 11.2 | -53.4 | 51.4 | 28 | CAT3 |
|  | 12.3 | -64.1 | 61.7 | 19 | CAT4 |
|  | 12.7 | -66.2 | 61.7 | 20 | CAT4 |
|  | 17.4 | -77.3 | 66.8 | 28 | CAT4 |
|  | 17.7 | -78.4 | 64.3 | 28 | CAT4 |
|  | 25.6 | -87.4 | 61.7 | 46 | CAT4 |
|  | 27.4 | -70.6 | 38.6 | 42 | CAT1 |
|  | 25.5 | -69.5 | 41.1 | 37 | CAT2 |
|  | 26.5 | -74.3 | 43.7 | 60 | CAT2 |
|  | 26.5 | -75.6 | 46.3 | 46 | CAT2 |
| U | 11.5 | -35.3 | 26.7 | 37 | TS |
|  | 17.3 | -45.5 | 57.8 | 32 | CAT3 |
|  | 19.1 | -47.4 | 64.0 | 32 | CAT4 |
|  | 22.9 | -48.6 | 54.0 | 28 | CAT3 |
|  | 25.7 | -49.5 | 48.8 | 28 | CAT3 |
| 感 | 24.6 | -85.6 | 51.5 | 56 | CAT3 |
|  | 25 | -86.2 | 56.5 | 50 | CAT3 |
|  | 26.9 | -89 | 75.0 | 38 | CAT5 |
|  | 23.6 | -87.2 | 51.5 | 20 | CAT2 |
|  | 24.4 | -88.4 | 56.5 | 20 | CAT2 |
|  | 28.4 | -91.4 | 54.0 | 20 | CAT4 |
|  | 29 | -91.9 | 41.1 | 20 | CAT2 |
|  | 24.3 | -85.9 | 61.7 | 28 | CAT4 |
|  | 24.9 | -88 | 77.1 | 19 | CAT5 |
|  | 25.4 | -88.7 | 72.0 | 19 | CAT5 |
|  | 26.8 | -91 | 59.1 | 37 | CAT4 |
|  | 27.4 | -91.9 | 59.1 | 37 | CAT4 |

Table 2.2.: Number of data shown in Figure 2.3.a that fall into different ranges of $R / R_{\max }$.

| $R / R_{\max }$ <br> range | No. of data <br> points |
| :---: | :---: |
| $0-1.5$ | 3586 |
| $1.5-3$ | 8772 |
| $3-4.5$ | 11025 |
| $4.5-6$ | 9250 |
| $6-7.5$ | 6626 |
| $7.5-9$ | 4027 |
| $9-10.5$ | 2272 |
| $10.5-12$ | 1273 |
| $12-19$ | 1652 |

Figures 2.3.c and 2.3.d show the moving average and standard deviation of the ratios in Figures 2.3.a and 2.3.b, using a window of 2000 points. Except for a small region close to the core ( $R<$ $1.5 R_{\max }$ ), the local average in Figure 2.3.c fluctuates around 1. This means that on average the MSR model generates unbiased rainfall profiles for radial distances up to $15 R_{\max }$ from the TC center. For distances $R<1.5 R_{\max }$ the MSR model tends to overpredict the PR rainrates.

As noted above, the large local standard deviations in Figure 2.3.c reflect the significant small-scale variability of TC rainfall. It is interesting that the standard deviation tends to increase as the distance from the TC center increases. This is in accordance with the findings of other studies (Jorgensen, 1984a; Powell, 1990, and Molinari et al., 1994) that the outer TC environment exhibits more cellular structure and higher small-scale variability relative to the inner region.

Figure 2.3.d shows that for radial distances up to $8 R_{\max }$ the MSR model tends to underpredict the MM5 rainfall intensities by about $50 \%$, whereas for larger distances the opposite is true. Since the MSR model displays good skills in reproducing the PR rain rates, it is possible that these differences reflect MM5 biases. This is consistent with what other studies have found when comparing MM5 rainfall estimates to empirical and radar observations; see e.g. Fall et al. (2007), Juneng et al. (2007), Chen et al. (2007) and Rogers et al. (2007). The higher standard deviations in Figure 2.3.d compared to Figure 2.3.c further suggest that MM5 may enhance local convective activity. One should however caution that these observations are based on just one simulated hurricane and should be validated through a more extensive comparison.

### 2.5 Asymmetry of the Rainfall Field

In the case of a moving TC, equation (2.4) becomes

$$
i(R, \theta)= \begin{cases}\alpha(\bar{T}) \bar{Q} W_{H}(R, \theta) & , W_{H}(R, \theta)>0  \tag{2.8}\\ 0 & , W_{H}(R, \theta) \leq 0\end{cases}
$$

where the vertical wind speed $W_{H}$ depends on both $R$ and $\theta$ and is given by equation (2.2) for $Z=H$. In this asymmetric case the rainfall intensities from equation (2.8) must be corrected both radially using equations (2.6) and (2.7) and azimuthally to account for the redistribution of rainwater due to cyclonic circulation; on the latter, see Corbosiero and Molinari (2002), Black et al. (2002) and Rogers et al. (2003).

To keep the correction simple, we perform the azimuthal redistribution uniformly within an angular interval $[\theta, \theta+\Delta \theta]$ where $\Delta \theta$ is given by

$$
\begin{equation*}
\Delta \theta=\frac{V_{g r}(R)}{R}\left(t_{f}+t_{r}\right) \tag{2.9}
\end{equation*}
$$

The angle $\Delta \theta$ is in radians (positive clockwise in the Northern hemisphere), $V_{g r}$ is the tangential wind velocity at gradient level [equation (2.1)], $t_{f} \approx 30 \mathrm{~min}$ is the time needed for rain generating features like convective cells to develop (Weisman and Klemp, 1986; Rogers and Yau, 1996) and $t_{r}$ is the time needed for a raindrop at height $H$ to reach the ground. A rough estimate of $t_{r}$ comes from assuming an average raindrop velocity of $2-3 \mathrm{~m} / \mathrm{s}$ and a boundary layer depth $H \approx$ $2.5-3 \mathrm{~km}$. This gives $t_{r} \approx 25 \mathrm{~min}$.

Next we use equations (2.8) and (2.9) for $t_{f}+t_{r}=60 \mathrm{~min}$ to assess the effect of motion on the spatial variation of TC rainfall and propose a motion-based, rather than shear-based, parameterization of rainfall asymmetry.

### 2.5.1 Motion-based versus shear-based parameterization of rainfall asymmetry

MSR is a boundary layer model that generates spatial rainfall without explicitly considering vertical shear $S$. Rather, rainfall asymmetries are linked to storm motion. Since most of the rainfall originates at low altitudes relative to those that define wind shear, one may expect this to be a suitable approach.

To verify this assertion, Figure 2.4 compares the shear-aligned rainfall asymmetry from TRMM with the motion-aligned rainfall asymmetry from MSR. In both cases, asymmetry is
defined as

$$
\begin{equation*}
A(R, \theta)=\frac{i(R, \theta)-i_{\text {sym }}(R)}{i_{s y m}(R)} \tag{2.10}
\end{equation*}
$$

where $i(R, \theta)$ is rainfall intensity at $(R, \theta)$ and $i_{s y m}(R)$ is the azimuthal average. More specifically, Figure 2.4.a shows the average of the rainfall asymmetries in Figure 7 of Chen et al. (2006) over all TC-intensities and shear magnitudes after aligning the shear vector to point North. For shear we have used the distribution in Figure 6 of the same study, whereas for TC intensity we have used the discrete distribution in Table 1 of Lonfat et al. (2004).


Figure 2.4: Comparison of rainfall asymmetry from TRMM and the MSR model. (a) Ensemble average of rainfall asymmetries in Figure 7 of Chen et al. (2006) over all TC intensities and shear magnitudes. (b) Ensemble average of rainfall asymmetries from MSR over all TC intensities and translation velocities. In (b), the TC moves in the Northern hemisphere at an angle $75^{\circ}$ to the west of the shear vector in (a).

Similarly, Figure 2.4.b was generated by averaging rainfall asymmetries from the MSR model over a range of TC intensities and translation velocities. Storms are assumed to move in the Northern hemisphere at an angle of $75^{\circ}$ west of the shear-direction in Figure 2.4.a. This is the average angle between shear and motion from Figures 3 and 12 of Chen et al. (2006) and is in the range reported by Corbosiero and Molinari (2003). For storm intensity we have used the same discrete distribution as in Figure 2.4.a, setting $V_{\max }=30 \mathrm{~m} / \mathrm{s}$ for tropical storms, $V_{\max }=$ $42 \mathrm{~m} / \mathrm{s}$ for CAT12 and $V_{\max }=60 \mathrm{~m} / \mathrm{s}$ for CAT35 systems. The distribution of the translation velocity was taken from Figure 11 of Chen et al. (2006). All other storm parameters have been
kept constant, with values $f=4.710^{-5} \mathrm{sec}^{-1}, R_{\max }=40 \mathrm{~km}, B=1, \bar{T}=22^{\circ} \mathrm{C}, \bar{Q}=0.8, K=50 \mathrm{~m}^{2} / \mathrm{s}$, and $C_{D}=0.002$.

One sees that the two asymmetries are very similar in both pattern and magnitude, validating the contention that for rainfall risk one can use the MSR model with motion as the driver of asymmetry. Differences between Figures 2.4.a and 2.4.b occur mainly far away from the core ( $R>250 \mathrm{~km}$ ), but these differences are statistically not significant and inconsequential for risk analysis.

### 2.6 Sensitivity Analysis

Figures 2.5 and 2.6 show the sensitivity of the MSR model results to various tropical cyclone characteristics: the tangential wind speed under gradient balance [parameterized by $V_{\max }, R_{\max }$ and $B$; see equation (2.1)], the vertical diffusion coefficient $K$, the surface drag coefficient $C_{D}$, the depth-averaged temperature $\bar{T}$ inside the BL and the translation velocity $V_{t}$ of the storm. Since rainfall intensity is proportional to the depth-averaged saturation ratio $\bar{Q}$ [see equations (2.4) and (2.8)], dependence on $\bar{Q}$ is not illustrated.

Figure 2.5 shows the sensitivity of the azimuthally averaged rainfall intensity $i_{s y m}$ to $V_{\max }, R_{\max }$, $B, K, C_{D}$ and $\bar{T}$. Parameters are varied one at a time around the base-case values $V_{\max }=50 \mathrm{~m} / \mathrm{s}$, $R_{\max }=40 \mathrm{~km}, B=1, K=50 \mathrm{~m}^{2} / \mathrm{s}, C_{D}=0.002, \bar{T}=22^{\circ} \mathrm{C}$ and $\bar{Q}=0.8$ (solid lines). The figure shows that the maximum tangential velocity $V_{\max }$ and the roughness of the surface boundary (expressed through $C_{D}$ ) have significant effects on rainfall intensity and that lower values of $R_{\max }$ produce rain rates that are more peaked and more concentrated near the TC center.

Dependence of the azimuthally averaged rainrate $i_{s y m}$ on $V_{\max }$ of the type produced by the model has been observed in TC rainfall data (Lonfat et al., 2004, Tuleya et al., 2007; see Introduction). For example, the expressions used by the R-CLIPER parameterization (Tuleya et al., 2007) indicate that when $V_{\max }$ increases from 50 to $70 \mathrm{~m} / \mathrm{s}$, the maximum rainrate increases by a factor of about 1.5. This is also what the MSR model predicts. However, to our knowledge the effect of $C_{D}$ and $R_{\max }$ on $i_{s y m}$ have not been isolated from data. The effect of surface roughness can be qualitatively assessed using the finding in Trenberth et al. (2007) that low-level horizontal wind convergence is by far the dominant factor for TC rainfall. Hence, if one considers that lowlevel convergence increases with increasing surface drag (see Chapter 1), one concludes that higher surface drag coefficients should cause TC rainfall to intensify.

The $B$ parameter has a small effect on the peak rainfall intensity, but influences significantly the rate at which rainfall decays with radial distance (higher values of $B$ resulting in faster decay). The azimuthally averaged rainfall intensity $i_{s y m}$ has small sensitivity to temperature $\bar{T}$ and the vertical diffusion coefficient $K$. Consequently, setting those parameters to constant values (e.g. to $\bar{T}=22^{\circ} \mathrm{C}$ and $K=50 \mathrm{~m}^{2} / \mathrm{s}$, as was done in Sections 2.3-2.5) does not induce large errors.


Figure 2.5: Sensitivity of the azimuthally averaged MSR rainfall profiles. Solid lines correspond to $V_{\max }=50 \mathrm{~m} / \mathrm{s}$, $R_{\text {max }}=40 \mathrm{~km}, B=1, C_{D}=0.002, K=50 \mathrm{~m}^{2} / \mathrm{s}, \bar{T}=22^{\circ} \mathrm{C}$ and $\bar{Q}=0.8$. Each panel shows results under perturbation of one parameter.

Figure 2.6 shows the effect of the drag coefficient $C_{D}$ and translation velocity $V_{t}$ on rainfall asymmetry for a TC that translates northward in the Northern hemisphere. All other parameters
are the same as for the base case in Figure 2.5. As expected and in accordance with findings in Lonfat et al. (2004), the asymmetry increases as $V_{t}$ increases. The effect of $C_{D}$ is more complex: at the front of the storm, rainfall asymmetry is insensitive to $C_{D}$, whereas at the rear-right the rainfall asymmetry increases with decreasing $C_{D}$.


Figure 2.6: Sensitivity of MSR rainfall asymmetry to the drag coefficient $C_{D}$ and the storm translation velocity $V_{t}$ for a tropical cyclone that moves northward. All other parameters are the same as for the base case in Figure 2.5.

### 2.7 Conclusions

We have developed a simple theoretical model for the large-scale rainfall intensity field generated by translating tropical cyclones (TCs). The model assumes that, with corrections for sloping updrafts and azimuthal redistribution, the upward water vapor flux originated from the boundary layer is a good predictor of rainfall intensity. Vertical moisture fluxes are calculated
using elementary thermodynamic principles in combination with a boundary layer model that extends Smith's (1968) analysis to moving storms.

The proposed modified-Smith-for-rainfall (MSR) model estimates the rainfall field from a given radial profile of the tangential wind speed at gradient level, the translation velocity $V_{t}$ of the storm, the surface drag coefficient $C_{D}$, and the average temperature $\bar{T}$ and saturation ratio $\bar{Q}$ inside the TC boundary layer. Model predictions are compared to MM5 simulations and RCLIPER estimates and validated through precipitation radar (PR) rainfall products from TRMM. The MSR model displays good skills in reproducing the shape and magnitude of PR rainfall fields. We have also verified that the asymmetries produced by storm motion are close to those observed and often parameterized in terms of vertical wind shear. In a parametric analysis, we have studied how the model predictions depend on various storm characteristics.

The combination of a rich parameterization and computational efficiency makes the present model an attractive instrument for risk applications, where one must assess tropical cyclone rainfall under many storm and environmental scenarios. For the latter purpose one needs tools with computational times on the order of minutes. This constraint effectively rules out the use of full-physics high-resolution numerical weather prediction models. An important limitation of the MSR model relative to high-resolution schemes is that it does not account for local rainfall intensifications due to rainbands and local convection. In Chapter 3 we model those intensifications by adding a stochastic component to the MSR model predictions. Another limitation of the MSR model is that it does not account for after-landfall conditions and therefore is applicable only to open-water or near-water sites. Extension of the model to inland conditions should be pursued in the future.

## 3. Long-Term Rainfall Risk from Tropical Cyclones in Coastal Areas

### 3.1 Introduction

The quantification of long-term rainfall risk is a basic problem of stochastic hydrology (e.g. Chow et al., 1988; Singh, 1992). Our specific interest is in the risk of extreme rainfall posed at coastal sites by tropical cyclones (TCs). These events are relatively rare, but in combination with wind, surge and waves, high rainfall intensities may have devastating consequences (Herbert et al., 1997; Rappaport, 2000).

For ordinary rainfall, standard risk analysis techniques use historical annual-maximum data (e.g. Koutsoyiannis et al., 1998) or peak-over-threshold (PoT) information (e.g. Madsen et al., 1997). The episodic and spatially localized nature of tropical cyclones prevents one from using these standard techniques. For example, the annual maximum and PoT rainfall statistics due to tropical cyclones depend on whether the site is "hit" by one or more TCs in any one year and therefore are highly erratic. For this reason, the risk is best assessed parametrically, by combining a probabilistic model of the maximum rainfall due to a TC with given characteristics $\boldsymbol{\theta}=\left[\theta_{1}, \ldots, \theta_{r}\right]$ with the rate at which those events occur. For coastal sites, the vector $\boldsymbol{\theta}$ might include the intensity and size of the storm, the location and translational velocity at landfall, and possibly other parameters related to atmospheric conditions, the radial profile of the tangential winds, etc. Parametric approaches of this type have been used to assess the risk posed by tropical cyclones for wind, surge and waves (Myers, 1975; Ho and Myers, 1975; Ho et al., 1987; Powell et al., 2005; IPET, 2006, 2008), but not rain. Our objective is to fill this gap.

The main problem for rainfall is to evaluate the extreme precipitation intensities caused by a TC with given characteristics $\boldsymbol{\theta}$. The historical data are too sparse and the potentially important TC parameters are too many to infer such extreme rainfalls from empirical observations alone. For example, current empirical approaches (Lonfat et al., 2004, 2007; Tuleya et al., 2007) classify storms into three coarse intensity categories and use microwave imager (TMI) data from TRMM (Simpson et al. 1988) to calculate the ensemble-average rainrate for each category as a function of distance from the TC center.

The alternative we pursue here is to use a physical model to assess the dependence of the mean rainfall field on $\boldsymbol{\theta}$ and statistical analysis to quantify the fluctuations of rainfall intensity
around this mean field. The physical model is that developed in Chapters 1 and 2. In Chapter 1 we proposed a theoretical method to estimate the large-scale horizontal and vertical winds inside TCs (the vertical winds are largely responsible for rain). The model is an extension of Smith's (1968) formulation and is referred to here as the Modified Smith (MS) model. Characteristics of the TC that are explicitly considered by the model are the maximum tangential wind speed $V_{\max }$, the radius of maximum winds $R_{\max }$, the parameter $B$ that controls the shape of the radial profile of the tangential wind speed (Holland, 1980), the storm translation velocity $V_{t}$, the surface drag coefficient $C_{D}$, and the vertical diffusion coefficient $K$. When $V_{t}=0$, the wind field is symmetric around the storm center, whereas when the TC translates in the Northern (Southern) hemisphere the field is asymmetric, with stronger horizontal and vertical winds right-front (left-front) of the storm. The model does not resolve rainbands, local convection and turbulent phenomena and therefore produces smooth wind fields.

In Chapter 2 we extended the MS model to predict TC rain, assuming that the upward moisture flux at the top of the TC boundary layer is all converted into rainfall. The vertical moisture flux is evaluated from the vertical winds generated by the MS model and two additional parameters: the average temperature $\bar{T}$ and average saturation ratio $\bar{Q}$ inside the TC boundary layer. We call this the modified-Smith-for-rainfall (MSR) model. The MSR model should prove useful for climatologic studies, but for hazard analysis it has the limitation of ignoring the interstorm and intra-storm variations of rainfall intensity. These variations are highly significant for the assessment of risk. For example, Lonfat et al. (2004) found that, also within a given TC strength category, the average of the positive rainfall intensity inside annular regions of 10 km width may deviate from the median value by more than one order of magnitude.

Our main objectives are: (1) Extend the MSR model to obtain the probability distribution of the maximum rainfall intensity in a time interval of given duration $D$ at a fixed geographical location during the passage of a tropical cyclone with given characteristics $\boldsymbol{\theta}$ and (2) Combine this maximum rainfall model with a TC recurrence model to quantify rainfall risk in the form of intensity-duration-frequency (IDF) curves. For the first objective, we consider a site A at some distance $y$ to the right $(y<0)$ or left $(y>0)$ of the TC center, as shown in Figure 3.1. As the storm passes, the rainfall intensity at A fluctuates as a random process $I(t)$. Our interest is in $I_{D}(t)$, the moving average of $I(t)$ for an averaging duration $D$, and more specifically in the distribution of $I_{D, \max }(y, \theta)$, the maximum of $I_{D}(t)$ during the storm.


Figure 3.1: Schematic representation of a moving storm. Point O translates with the storm at speed $V_{r}$. Point A is the geographical location of interest.

Section 3.2 presents our general approach to calculate the distribution of $I_{D, \max }(y, \boldsymbol{\theta})$. This distribution is obtained in Section 3.3 and validated in Section 3.4. Section 3.4 also shows how the distribution depends on various storm characteristics, the normalized distance $y / R_{\max }$ from the center of the storm, and the averaging duration $D$. Section 3.5 uses the model of $I_{D, \max }(y, \boldsymbol{\theta})$ and a recurrence relation for hurricanes in the Gulf of Mexico to obtain IDF curves for New Orleans and compares these curves with published IDF values for all rainstorms (TCs and non-TCs) combined. Conclusions are stated in Section 3.6.

### 3.2 Framework for the Estimation of Extreme TC Rainfall

Our first objective is to relate the distribution of the maximum rainfall intensity $I_{D, \max }(y, \theta)$ to the smooth rainfall intensities produced by the MSR model in Chapter 2. The storm parameters are $\boldsymbol{\theta}$ $=\left[V_{\max }, R_{\max }, V_{t}\right]$. The analysis uses a Cartesian reference frame $(x, y)$, translated and rotated such that the center of the storm O moves to the right along the $x$ axis; see Figure 3.1. In this reference, the ordinate $y$ of A is also the closest (signed) distance of A from the storm center.

To estimate this relationship, we use precipitation radar (PR) data from the TRMM mission (Simpson et al. 1988; Kummerow et al., 1998; Lee et al., 2002). These data are in the form of swaths about 200 km wide with a spatial resolution of approximately 5 km and have been validated against ground-based radar and rain gauge measurements (Bolen and Chandrasekar, 2000; Liao et al., 2001; Wolff et al., 2005). Due to their long inter-frame time (about 12 hours), the PR snapshots cannot be interpolated to produce the rainfall intensities in continuous time that are needed to estimate rainfall maxima. A common way to overcome this limitation is to use Taylor's frozen turbulence hypothesis (Taylor, 1921, 1938). Under this hypothesis, the temporal
variability of rainfall at a fixed location A is statistically the same as the variability that results from translating the frozen-in-time rainfield over A with the storm velocity $V_{t}$. For example, Vicente et al. (1998), Scofield and Kuligowski (2003), Kidder et al. (2005) and Ferraro et al. (2005) used Taylor's hypothesis to obtain rainfall totals at fixed locations from satellite and radar rainfall snapshots.

It follows from Taylor's hypothesis that $I_{D, \max }(y, \boldsymbol{\theta})$ has the same distribution as $I_{l, \max }(y, \boldsymbol{\theta})$, the maximum of the rainfall intensity averaged in a spatial window of length $l$ along cross-section C in Figure 3.1, for $l=D V_{t}$. As an example, Figure 3.2 shows moving-average rainfall intensities from Hurricane Katrina (2005) along a cross-section at distance $y=100 \mathrm{~km}$ from the storm center, for averaging lengths $l=6 \mathrm{~km}$ (dashed line) and $l=24 \mathrm{~km}$ (solid line). The cross-section extends over $L=384$ kilometers and is symmetrical relative to the storm center.


Figure 3.2: Rainfall intensities from Hurricane Katrina (Aug. 28, 2005, at 03:00UTC; TRMM frame 44361) along a cross-section C at distance $y=100 \mathrm{~km}$ from the storm center, for spatial averaging scales $l=6$ and 24 km . The maximum values $I_{l, \max }$ are indicated by circles. $I_{L}$ is the average value for the entire cross-section and $I_{L, M S R}$ is the estimate of $I_{L}$ produced by the MSR model.

The intensity labeled $I_{L}$ in Figure 3.2 is the average PR rainrate in $L$, whereas $I_{L, M S R}$ is the estimate of that average rainrate produced by the MSR model. These average intensities play an important role in our analysis. For any given $(y, \boldsymbol{\theta})$ combination, $I_{L, M S R}$ is a deterministic quantity, whereas $I_{L}$ is regarded as a random variable with different values for different tropical cyclones. We model this storm-to-storm variability by expressing $I_{L}(y, \boldsymbol{\theta})$ as

$$
\begin{equation*}
I_{L}(y, \boldsymbol{\theta})=I_{L, M S R}(y, \boldsymbol{\theta}) \beta_{L}(y, \boldsymbol{\theta}) \tag{3.1}
\end{equation*}
$$

where $\beta_{L}$ is a random variable.
Figure 3.2 also shows significant amplification of the rainfall intensity when one considers the maximum over lengths $l<L$. One may express the maximum in $l, I_{l, \max }$, as

$$
\begin{equation*}
I_{l, \max }(y, \boldsymbol{\theta})=I_{L, M S R}(y, \boldsymbol{\theta}) \beta_{l, \max }(y, \boldsymbol{\theta})=I_{L, M S R}(y, \boldsymbol{\theta}) \beta_{L}(y, \boldsymbol{\theta}) \gamma_{l, \max }(y, \boldsymbol{\theta}) \tag{3.2}
\end{equation*}
$$

where the total factor relative to $I_{L, M S R}, \beta_{l, \max }$, is the product of $\beta_{L}$ in equation (3.1) and a random amplification factor $\gamma_{l, \max }$ for the change of scale from $L$ to $l$. The next section uses PR/TRMM data from 8 tropical cyclones (a total of 38 frames) to derive the distributions of $\beta_{L}$ and $\gamma_{l, m a x}$. The selected frames (see Table 3.1) cover a wide range of TC intensities, from tropical storms to CAT5 systems, under pre-landfall conditions. This makes our model best suited for use in coastal areas. Due to the limited lateral coverage of the PR instrument, an additional requirement for selecting the frames was to cover regions close to the hurricane core (with radial distance less than 300 km from the storm center), as these are the regions that are most critical for rainfall.

### 3.3 Distribution of $\boldsymbol{\beta}_{L}$ and $\boldsymbol{\gamma}_{l \text { max }}$

Equation (3.2) relates the maximum rainfall intensity in $l$ to the average intensity in $L$ produced by the MSR model using two random correction factors: a factor $\beta_{L}$ to obtain the average rainfall in $L$, and a factor $\gamma_{l, \max }$ to obtain the maximum average intensity at a smaller scale $l$. Sections 3.3.1 and 3.3.2 obtain the distribution of these factors using the rainfall information in Table 3.1 and MSR model simulations.

### 3.3.1 Distribution of $\boldsymbol{\beta}_{\boldsymbol{L}}$

The factor $\beta_{L}$ is given by

$$
\begin{equation*}
\beta_{L}(y, \boldsymbol{\theta})=\frac{I_{L}(y, \boldsymbol{\theta})}{I_{L, M S R}(y, \boldsymbol{\theta})} \tag{3.3}
\end{equation*}
$$

Its distribution generally depends on the distance $y$ from the TC center and the vector $\boldsymbol{\theta}=\left[V_{\max }\right.$, $\left.R_{\max }, V_{t}\right]$ of storm characteristics, but as we show next, a simple parameterization in terms of the normalized distance $y^{\prime}=\left|y / R_{\max }\right|$ and the large-scale MSR rainfall intensity $I_{L, M S R}$ suffices. Of course, $I_{L, M S R}$ is itself a function of $\boldsymbol{\theta}$.

Table 3.1: Characteristics of the PR/TRMM rain frames used in the analysis. The direction of storm translation is relative to the East and is positive counter-clockwise. The estimates of $V_{\max }$ and $R_{\max }$ are from the extended best track record (M. DeMaria, 2008; personal communication).

|  | Storm center |  | Storm speed ( $\mathrm{m} / \mathrm{s}$ ) | storm direction (deg) | $\begin{gathered} V_{\max } \\ (\mathrm{m} / \mathrm{s}) \end{gathered}$ | $\begin{aligned} & \boldsymbol{R}_{\max } \\ & (\mathbf{k m}) \end{aligned}$ | TRMM <br> frame number | Storm intensity |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Lat. (deg) | Lon. (deg) |  |  |  |  |  |  |
|  | 21.7 | -61.6 | 4.9 | 143 | 48.8 | 41 | 10290 | CAT2 |
|  | 23.5 | -68.7 | 4.8 | 169 | 64.0 | 37 | 10317 | CAT4 |
|  | 23.7 | -70.6 | 5.8 | 171 | 69.3 | 37 | 10321 | CAT4 |
|  | 12.6 | -43.7 | 10.9 | 158 | 23.1 | 37 | 38646 | TS |
|  | 15.7 | -49.8 | 5.4 | 139 | 51.4 | 19 | 38667 | CAT3 |
|  | 17 | -51.3 | 5.3 | 139 | 54.0 | 28 | 38677 | CAT3 |
|  | 17.9 | -52.6 | 4.3 | 144 | 59.1 | 28 | 38682 | CAT4 |
|  | 19 | -57.3 | 4.9 | 180 | 51.4 | 28 | 38708 | CAT3 |
|  | 21.2 | -68.5 | 6.1 | 162 | 61.7 | 28 | 38739 | CAT4 |
| $\begin{aligned} & \underset{\Xi}{Z} \\ & \underset{\sim}{\pi} \end{aligned}$ | 8.9 | -38.9 | 7.6 | 184 | 25.7 | 37 | 38789 | TS |
|  | 10.7 | -50.6 | 12.2 | 185 | 57.5 | 28 | 38814 | CAT4 |
|  | 11.2 | -53.4 | 8.1 | 173 | 51.4 | 28 | 38820 | CAT3 |
|  | 12.3 | -64.1 | 8.3 | 166 | 61.7 | 19 | 38845 | CAT4 |
|  | 12.7 | -66.2 | 7.3 | 164 | 61.7 | 20 | 38851 | CAT4 |
|  | 17.4 | -77.3 | 4.1 | 194 | 66.8 | 28 | 38892 | CAT4 |
|  | 17.7 | -78.4 | 4.4 | 153 | 64.3 | 28 | 38897 | CAT4 |
|  | 25.6 | -87.4 | 5.5 | 112 | 61.7 | 46 | 38954 | CAT4 |
|  | 27.4 | -70.6 | 5.5 | 0 | 38.6 | 42 | 39045 | CAT1 |
|  | 25.5 | -69.5 | 1.1 | 207 | 41.1 | 37 | 39079 | CAT2 |
|  | 26.5 | -74.3 | 7.4 | 173 | 43.7 | 60 | 39106 | CAT2 |
|  | 26.5 | -75.6 | 6.5 | 180 | 46.3 | 46 | 39110 | CAT2 |
|  | 11.5 | -35.3 | 7.1 | 176 | 26.7 | 37 | 38987 | TS |
|  | 17.3 | -45.5 | 2.0 | 166 | 57.8 | 32 | 39033 | CAT3 |
|  | 19.1 | -47.4 | 5.9 | 121 | 64.0 | 32 | 39048 | CAT4 |
|  | 22.9 | -48.6 | 8.2 | 112 | 54.0 | 28 | 39059 | CAT3 |
|  | 25.7 | -49.5 | 6.8 | 117 | 48.8 | 28 | 39063 | CAT3 |
|  | 24.6 | -85.6 | 2.1 | 153 | 51.5 | 56 | 44357 | CAT3 |
|  | 25 | -86.2 | 3.5 | 146 | 56.5 | 50 | 44361 | CAT3 |
|  | 26.9 | -89 | 5.5 | 135 | 75.0 | 38 | 44373 | CAT5 |
| $\begin{aligned} & \text { N } \\ & \text { N } \\ & \end{aligned}$ | 23.6 | -87.2 | 9.0 | 162 | 51.5 | 20 | 27826 | CAT2 |
|  | 24.4 | -88.4 | 6.2 | 141 | 56.5 | 20 | 27830 | CAT2 |
|  | 28.4 | -91.4 | 10.1 | 117 | 54.0 | 20 | 27842 | CAT4 |
|  | 29 | -91.9 | 5.4 | 124 | 41.1 | 20 | 27845 | CAT2 |
| $\begin{aligned} & \stackrel{0}{0} \\ & \text { Fix } \end{aligned}$ | 24.3 | -85.9 | 5.7 | 189 | 61.7 | 28 | 44743 | CAT4 |
|  | 24.9 | -88 | 3.9 | 166 | 77.1 | 19 | 44754 | CAT5 |
|  | 25.4 | -88.7 | 4.3 | 153 | 72.0 | 19 | 44758 | CAT5 |
|  | 26.8 | -91 | 5.5 | 135 | 59.1 | 37 | 44770 | CAT4 |
|  | 27.4 | -91.9 | 4.8 | 143 | 59.1 | 37 | 44773 | CAT4 |

Figure 3.3 shows statistics of $\beta_{L}$ as a function of $y^{\prime}$ and $I_{L, M S R}$ for the TRMM frames in Table 3.1. The MSR intensities $I_{L, M S R}$ where obtained using the values of $V_{\max }, R_{\max }$, and $V_{t}$ in the
extended best track record (M. DeMaria, 2008; personal communication; Demuth et al., 2006). In addition, the MSR model requires the vertical diffusion coefficient $K$, the surface drag coefficient $C_{D}$, the vertically averaged temperature $\bar{T}$ and saturation ratio $\bar{Q}$ inside the boundary layer (BL), Holland's $B$ parameter for the profile of gradient winds, the sloping angle $\psi_{0}$ and height $H_{0}$ of the wall updraft, and the temporal scale $t_{r}$ for azimuthal re-distribution of rainfall by the cyclonic circulation; see Chapter 2 for details. In our simulations we have set $K=50 \mathrm{~m}^{2} / \mathrm{s}, C_{D}$ $=0.002, \bar{T}=22^{\circ} \mathrm{C}, \bar{Q}=0.8, B=1, \psi_{0}=50^{\circ}, H_{0}=6 \mathrm{~km}$ and $t_{r}=60 \mathrm{~min}$. In Chapter 2 we recommend these settings as representative of tropical cyclones in the North Atlantic and as values that reproduce well the TRMM/PR rainfall fields in an ensemble-average sense.

Figures 3.3.a and 3.3.b show smooth contour plots of the log-mean $m_{\ln \beta_{L}}$ and $\log$-standard deviation $\sigma_{\ln \beta_{L}}$ of $\beta_{L}$ as a function of the normalized distance $y^{\prime}=\left|y / R_{\max }\right|$ and the rainfall intensity $I_{L, M S R}$ for the 38 frames in Table 3.1. For each frame, a regular spacing $\Delta y=10 \mathrm{~km}$ was maintained between adjacent transects, producing a total of 789 points; see Figure 3.3.a. In all cases, averaging is over segments of length $L=384 \mathrm{~km}$, symmetric relative to the storm center. This value of $L$ encompasses more than $95 \%$ of the total rainfall volume along each transect; see for example Figure 3.2. Smooth estimates of the mean value and variance of $\ln \beta_{L}$ were obtained using an isotropic Gaussian kernel with standard deviation 0.5 in the $\left[\ln \left(I_{L, M S R}\right), \ln \left(y^{\prime}\right)\right]$-plane. Hence, if $g(\mathbf{x})$ denotes this kernel, local estimates of $m_{\ln \beta_{L}}\left(\mathbf{x}_{0}\right)$ and $\sigma_{\ln \beta_{L}}^{2}\left(\mathbf{x}_{0}\right)$ around a given point $\mathbf{x}_{0}=\left[\ln \left(I_{L, M S R}\right), \ln \left(y^{\prime}\right)\right]$ are given by

$$
\begin{equation*}
m_{\ln \beta_{L}}\left(\mathbf{x}_{0}\right)=\frac{\sum_{i} \ln \beta_{L}\left(\mathbf{x}_{i}\right) \mathrm{g}\left(\mathbf{x}_{i}-\mathbf{x}_{0}\right)}{\sum_{i} \mathrm{~g}\left(\mathbf{x}_{i}-\mathbf{x}_{0}\right)}, \quad \sigma_{\ln \beta_{L}}^{2}\left(\mathbf{x}_{0}\right)=\frac{\sum_{i}\left[\ln \beta_{L}\left(\mathbf{x}_{i}\right)-m_{\ln \beta_{L}}\left(\mathbf{x}_{i}\right)\right]^{2} \mathrm{~g}\left(\mathbf{x}_{i}-\mathbf{x}_{0}\right)}{\sum_{i} \mathrm{~g}\left(\mathbf{x}_{i}-\mathbf{x}_{0}\right)} \tag{3.4}
\end{equation*}
$$

where $\mathbf{x}_{i}$ is the generic $\left[\ln \left(I_{L, M S R}\right), \ln \left(y^{\prime}\right)\right]$ combination for which a value of $\beta_{L}$ is available. To use values of $\beta_{L}$ at locations close to the center of the storm where $\ln \left(y^{\prime}\right)$ diverges, 59 points with $|y|<0.5 R_{\max }$ where moved to $y=0.5 R_{\max }$.


Figure 3.3: (a,b) Mean value and standard deviation of $\ln \beta_{L}$ as a function of the model rainfall intensity $I_{L, M S R}$ and the normalized distance $y^{\prime}=\left|y / R_{\max }\right|$ from the TC center using 789 cross-sections of the 38 frames in Table 3.1. The contour plots are obtained using a smoothing Gaussian kernel with standard deviation 0.5 . The dashed lines delimit the region of high data density along the direction of the gradient of $\sigma_{\ln \beta_{L}}$ (white arrow). (c) Plots of $m_{\ln \beta_{L}}$ and $\sigma_{\ln \beta_{L}}$ as a function of $\omega=\ln \left(y^{\prime}\right)-0.4 \ln \left(I_{L, M S R}\right)$ along cross-section A. (d) Comparison between the standard normal density and the empirical PDF of $\ln \left(\beta_{L}\right)$, normalized to have zero mean and unit variance.

The overall mean value of $\beta_{L}$ is 1.02 , indicating that on average the MSR model produces unbiased large-scale estimates of the PR rainrates. The dashed lines in Figure 3.3.b delimit the region of high data density and are generally oriented along the gradient of $\sigma_{\ln \beta_{\iota}}$. Figure 3.3.c shows plots of $m_{\ln \beta_{L}}$ and $\sigma_{\ln \beta_{L}}$ as a function of the transformed variable $\omega=\ln \left(y^{\prime}\right)-0.4 \ln \left(I_{L, M S R}\right)$ along the dashed-dotted line in Figure 3.3.b. The log-mean $m_{\ln \beta_{L}}$ is approximately constant and equal to -0.5 , whereas $\sigma_{\ln \beta_{L}}$ increases as the normalized distance $y^{\prime}$ increases or the large-scale mean rainfall intensity $I_{L, M S R}$ decreases. This higher $\log$ variability in regions of lower intensity is
expected due to the more episodic nature of rainfall in those regions. This is also in qualitative agreement with the findings of Lonfat et al. (2004) and Molinari et al. (1994). The solid lines in Figure 3.3.c are least-squares fits for the mean and standard deviation of $\ln \beta_{L}$. For $y$ close to zero, the fitted standard deviation becomes very small or negative. To avoid this inconsistency, we have imposed a lower bound of 0.5 to the fitted standard deviation.

To investigate the distribution type, we standardize the empirical values of $\ln \beta_{L}$ by removing the parametrically fitted mean -0.5 and dividing by the parametrically fitted standard deviation $0.25 \omega+0.87$. Figure 3.3.d shows a histogram of these normalized quantities and suggests that $\ln \beta_{L}$ has near-normal distribution. To check for possible lack of fit and possible dependence of $\ln \beta_{L}$ on other parameters, we generated histograms of the type in Figure 3.3.d separately for different ranges of $y, y^{\prime}, I_{L, M S R}, R_{\max }$ and $V_{\max }$; see Figures 3.4-3.6.


Figure 3.4: Comparison between a standard normal density (solid lines) and histograms of the normalized quantities in Figure 3.3.d when classified into two categories of positive and negative distances $y$ from the center of the storm.


Figure 3.5: Comparison between a standard normal density (solid lines) and histograms of the quantities in Figure 3.3.d when classified into three equally sized bins of the normalized distance $y^{\prime}=\left|y / R_{\max }\right|$ (left column) and the theoretical large-scale rainfall intensity $I_{L, M S R}$ (right column).

As none of these analyses reveals significant dependence, we use the fits in Figure 3.3.c and model $\ln \beta_{L}$ as a normal variable with parameters

$$
\begin{align*}
& m_{\ln \beta_{L}}(\omega)=-0.5 \\
& \sigma_{\ln \beta_{L}}(\omega)=\max \{0.5,0.25 \omega+0.87\} \tag{3.5}
\end{align*}
$$

where $\omega=\ln \left(y^{\prime}\right)-0.4 \ln \left(I_{L, M S R}\right)$.


Figure 3.6: Same as Figure 3.5 but for different ranges of the radius of maximum winds $\boldsymbol{R}_{\max }$ (left column) and the intensity $V_{\max }$ of the storm (right column).

### 3.3.2 Distribution of $\gamma_{l, \max }$

Next we consider the amplification factor $\gamma_{l, \max }$ in equation (3.2). The distribution of this factor can be found by a variety of methods, from the direct use of data on $\gamma_{l, \max }$ from the frames in Table 3.1 to theoretical analysis of the maximum of the moving-average processes $I_{l}(x)$ illustrated in Figure 3.2. In Appendix B we compare several such approaches and find similar results. Here we follow the empirical approach, which is the simpler and more transparent method. We start by calculating the empirical ratio

$$
\begin{equation*}
\gamma_{l, \max }=\frac{I_{l, \max }}{I_{L}}, l \leq L \tag{3.6}
\end{equation*}
$$

where $I_{L}$ is the average PR rainrate along a cross section C of fixed length $L=384 \mathrm{~km}$ and $I_{l, \max }$ is the maximum rainfall intensity when the same cross section is continuously scanned using an averaging window of length $l$; see Figures 3.1 and 3.2 and Section 3.2. Ideally, the cross section

C should be in the direction of the storm motion, but since the TRMM swaths are not always aligned with that direction, we calculate the factor $\gamma_{l, \max }$ using cross-sections parallel to the swath track. Hence, the resulting factor $\gamma_{l, \max }$ does not depend on the orientation of C relative to the storm motion. We verified that $\gamma_{l, \max }$ is insensitive to this orientation by dividing the swaths into two groups: those that are generally aligned with the storm trajectory and those that are not. The distribution of $\gamma_{l, \max }$ is similar in the two cases.

We also studied the dependence of the distribution of $\gamma_{l, \max }$ on $R_{\max }$. Dependence is expected because smaller values of $R_{\max }$ produce more picked radial rainfall profiles and hence higher rainfall maxima. Figure 3.7 shows histograms of $\gamma_{l, \max }$ as a function of the scale of spatial averaging $l$ when classified into three equally sized bins of the radius of maximum winds $R_{\max }$. One sees that for small spatial scales ( $l \leq 12 \mathrm{~km}$ ) the mean value and standard deviation of $\gamma_{l, \max }$ increase somewhat with decreasing $R_{\max }$, whereas at larger spatial scales the increase is modest. Based on these results, we ignore the dependence of $\gamma_{l, \max }$ on $R_{\max }$ and use a simple parameterization in terms of the averaging length $l$ and the large-scale average intensity $I_{L}$. The latter quantity depends significantly on both the storm intensity $V_{\max }$ and the distance $y$ from the storm center; see Chapter 2.

Figure 3.8 shows $\log -\log$ plots of $E\left[\gamma_{l, \max }\right]$ and $\operatorname{Var}\left[\gamma_{l, \max }\right]$ against $l$ after classifying the 789 cross-sections in Figure 3.3.a into 12 equally-sized $I_{L}$ bins. As expected, $\operatorname{Var}\left[\gamma_{l, \max }\right]$ increases with decreasing spatial scale $l$. A less obvious finding is that the variability of $\gamma_{l, \max }$ increases as the large-scale intensity $I_{L}$ decreases. Considering that lower values of $I_{L}$ are generally found at larger distances $y$ from the storm center, Figure 3.8 shows that the outer TC environment exhibits higher (multiplicative) variability relative to the inner region. The higher variability inside low- $I_{L}$ regions is due for the most part to an increase in the dry area fraction (see Appendix B) and has been noted also in other studies (Molinari et al., 1994; Lonfat et al., 2004). This feature is also commonly observed in extra-tropical rainfall (e.g. Over and Gupta, 1996; Deidda et al., 2006; Veneziano et al., 2006a; Gebremichael et al., 2006).


Figure 3.7: Empirical PDFs of $\gamma_{l, m a x}$ as a function of the scale of spatial averaging $l$ when classified into three equally sized bins of the radius of maximum winds $R_{m a x} . m$ and $\sigma$ are the mean value and standard deviation of the empirical distributions.

For each intensity category $I_{L}$, we use least squares to fit linear and quadratic expressions for the log-mean and log-variance of $\gamma_{l, \max }$,

$$
\begin{align*}
& \ln E\left[\gamma_{l, \max }\right]=a_{1} \ln l+a_{2}  \tag{a}\\
& \ln \operatorname{Var}\left[\gamma_{l, \max }\right]=a_{3}(\ln l)^{2}+a_{4} \ln l+a_{5} \tag{b}
\end{align*}
$$

where $l \leq L$ is in km and $a_{1}-a_{5}$ are parameters. Figure 3.9 shows how the parameters $a_{1}-a_{5}$ in equation (3.7) vary with the large-scale rainfall intensity $I_{L}$. The solid lines in Figure 3.9 are smooth least-squares estimates of $a_{i}(i=1, \ldots 5)$. Use of the smooth estimates reproduces well the empirical moments of $\gamma_{l, \max }$; see solid lines in Figure 3.8.


Figure 3.8: $\log -\log$ plots of $E\left[\gamma_{l m a x}\right]$ and $\operatorname{Var}\left[\gamma_{l_{\text {max }}}\right]$ against $l$ for different ranges of $I_{L}$. Triangles and circles indicate empirical values. The solid lines are from equation (3.7).


Figure 3.9: Dependence of the parameters $a_{1}-a_{5}$ in equation (3.7) on $I_{L}$. The solid lines are least squares fits.
The amplification factor $\gamma_{l, \max }$ has values between 1 and $L / l$. The lower bound corresponds to a uniform distribution of rainfall inside $L$, whereas the upper bound is attained when all the rainfall in $L$ is concentrated in a single $l$ interval. We model $\gamma_{l, \max }$ using a beta distribution with moments in equation (3.7). One may write this cumulative distribution as

$$
\begin{equation*}
F_{\gamma_{l, m a x}}(\gamma)=F_{X}\left(\frac{\gamma-1}{L / l-1}\right), \gamma \geq 1 \tag{3.8}
\end{equation*}
$$

where $F_{X}$ is the beta distribution in $[0,1]$ with parameters

$$
\begin{equation*}
E[X]=\frac{E\left[\gamma_{l, \max }\right]-1}{L / l-1}, \operatorname{Var}[X]=\frac{\operatorname{Var}\left[\gamma_{l, \max }\right]}{(L / l-1)^{2}} \tag{3.9}
\end{equation*}
$$

Figure 3.10 compares the empirical distribution of $\gamma_{l, \max }$ at spatial scales $l=96$ and 6 km for different large-scale average intensities $I_{L}$ with theoretical distributions from equations (3.8) and (3.9). The moments $E\left[\gamma_{l, \max }\right]$ and $\operatorname{Var}\left[\gamma_{l, \max }\right]$ in equation (3.9) are calculated using equation (3.7) with parameters $a_{1}-a_{5}$ in Figure 3.9. Equally good fits are obtained for other window sizes $l$; see Appendix B.


Figure 3.10: Comparison of histograms of $\gamma_{l, \max }$ for $l=96$ and 6 km and different large-scale intensities with theoretical distributions from equations (3.7) and (3.8). The intensity categories are the same as in the left column of Figure 3.8.

### 3.4 Validation of Maximum Rainfall Model and Sensitivity Analysis

For a tropical cyclone with parameters $\boldsymbol{\theta}=\left[V_{\max }, R_{\max }, V_{t}\right]$ and a given distance $y$ from the storm center, one may use equation (3.2) and the distributions of $\beta_{L}$ and $\gamma_{l, \max }$ in Section 3.3 to obtain the distribution of the maximum rainfall intensity $I_{l, \max }$ as

$$
\begin{equation*}
P\left[I_{l, \max }(y, \boldsymbol{\theta}) \leq i\right]=\int_{0}^{\infty} f_{I_{L} \mid y, \boldsymbol{\theta}}(u) F_{\gamma_{l, \max } I I_{L}=u}(i / u) d u \tag{3.10}
\end{equation*}
$$

where $f_{I_{L} t y, \boldsymbol{\theta}}$ is the probability density function of $I_{L}=I_{L, M S R} \beta_{L}$ given $(y, \boldsymbol{\theta})$ and $F_{\gamma_{l, m a x} I I_{L}}$ is the cumulative distribution function of $\gamma_{l, \max }$ given $I_{L}$. To assess the validity of the theoretical probabilities generated by equation (3.10), we compare them with in-sample observed relative frequencies, as follows. For each of the 789 transects extracted from the PR data in Table 3.1,

1. We calculate the maximum intensity $I_{l, \max }$ over segments of different length $l$;
2. We use $\left(V_{\max }, R_{\max }, V_{t}\right)$ from Table 3.1 and the distance $y$ of the transect from the TC center to obtain model estimates of the large-scale mean rainfall intensity $I_{L, M S R}(y, \boldsymbol{\theta})$ for $L=384 \mathrm{~km}$. All other MSR model parameters are fixed to the values in Section 3.3.1.
3. We use equation (3.10) and the parametric expressions in equations (3.5) and (3.7) and Figure 3.9 to find the distribution of $I_{l, \max }$ and the probability $P$ with which the observed value from step (1) is not exceeded.

If the model is correct, the probabilities $P$ from step (3) should have uniform distribution between 0 and 1 . Figure 3.11 shows histograms of $P$ for different $l$. The fact that the histograms differ little from a uniform density validates the model in the sense that calculated exceedance probabilities are close to observed relative frequencies.

The distribution of $I_{l, \max }$ in equation (3.10) depends critically on the amplification factor $\beta_{l, \max }$ in equation (3.2). Figure 3.12 shows how the distribution of $\beta_{l, \max }$ depends on $l, V_{\max }$, and $y^{\prime}=\left|y / R_{\max }\right|$. The effect of the translation velocity $V_{t}$ is modest and is not displayed. Also, for given $V_{\max }$ and $y^{\prime}=\mid y / R_{\max }, \beta_{l, \max }$ is insensitive to $R_{\max }$. The dispersion of $\beta_{l, \max }$ increases as $l$ decreases. It also increases for smaller $V_{\max }$ and larger $y^{\prime}$. The latter effects are related to the increased spatial variability of the rainfall intensity in regions of lower average precipitation.


Figure 3.11: Histogram of the non-exceedance probability $P$ in equation (3.10) for different spatial scales $l$. Each histogram is based on a sample of size 789.

### 3.5 Long-term Rainfall Risk for New Orleans

To assess rainfall risk at a given location $A$, one must find the rate $\lambda_{I_{D, \max }>i}$ of tropical cyclones for which $I_{D, \max }$, the maximum rainfall intensity at A for a given averaging duration $D$, exceeds different threshold levels $i$. This rate is given by

$$
\begin{equation*}
\lambda_{I_{D, \max }>i}=\lambda \int_{\text {all }(y, \boldsymbol{\theta})} P\left[I_{D, \max }(y, \boldsymbol{\theta})>i\right] f_{y, \boldsymbol{\theta}}(y, \boldsymbol{\theta}) d y d \boldsymbol{\theta} \tag{3.11}
\end{equation*}
$$

where $\lambda$ is the rate of TCs in the region, $P\left[I_{D, \max }(y, \theta)>i\right]$ is the probability that, for a storm with characteristics $\boldsymbol{\theta}, I_{D, \max }$ at distance $y$ from the storm center exceeds $i$, and $f_{y, \boldsymbol{\theta}}$ is the joint density of $(y, \theta)$. The joint density $f_{y, \boldsymbol{\theta}}$ and the rate $\lambda$ are region-specific and define the TC recurrence model. Under Taylor's hypothesis, $P\left[I_{D, \max }(y, \theta)>i\right]$ is obtained by setting $l=D V_{t}$ in equation (3.10).

To exemplify, we use equation (3.11) and a recurrence model for an appropriate coastal region of the Gulf of Mexico to obtain intensity-duration-frequency (IDF) relationships for New

Orleans. We select this location because: 1) the site is close to the coast; hence our pre-landfall model should produce accurate results, 2) a number of studies have developed TC recurrence models for the Louisiana coast, and 3) one can compare the TC rainfall results with available IDF curves from continuous rainfall records in the region.


Figure 3.12: Comparison of the PDF of $\beta_{l \text { max }}=I_{l, \text { max }} / I_{L, M S R}$ for different $V_{\max }, y^{\prime}=\mid y / R_{\max } I$, and $l$.

### 3.5.1 TC recurrence model for the northern Gulf of Mexico

We start by specifying the distribution of the distance $y$ between the center of the storm and the city of New Orleans (point A), which is located at approximately $\left(90^{\circ} \mathrm{W}, 30^{\circ} \mathrm{N}\right)$. Then we
consider the distribution of $\boldsymbol{\theta}=\left[V_{\max }, R_{\max }, V_{t}\right]$. The joint model for $V_{\max }$ and $R_{\max }$ is specified through the distribution of the maximum pressure deficit $\Delta P_{\max }$ and the conditional distributions of $\left[V_{\max } \mid \Delta P_{\max }\right]$ and $\left[R_{\max } \mid \Delta P_{\max }\right]$. Finally we specify the TC rate $\lambda$. To keep the model simple, we approximate the coastline by a line segment with constant latitude $30^{\circ} \mathrm{N}$ and longitudinal range $85^{\circ}-95^{\circ} \mathrm{W}(\approx 960 \mathrm{~km})$, centered at A .

Let $z$ be the location (positive eastward) of landfall relative to A. Assuming a straight storm path, the closest distance of the storm center from the site is

$$
\begin{equation*}
y=-z \cos (\alpha) \tag{3.12}
\end{equation*}
$$

where $\alpha$ is the azimuth of the storm track at landfall, positive clockwise. The distribution of $y$ can be obtained numerically from equation (3.12) and the distributions of $\alpha$ and $z$, assumed here to be independent. For $z$ we use a uniform distribution in the interval $\left[85^{\circ} \mathrm{W}, 95^{\circ} \mathrm{W}\right]$. The distribution of the angle $\alpha$ in the region is usually found to be normal or the mixture of two normal distributions, one for easterly storms and the other for westerly storms (Vickery and Twisdale, 1995; IPET, 2006, 2008). Here we model $\alpha$ using a single normal distribution with mean value $m_{\alpha}=-5.4^{\circ}$ and standard deviation $\sigma_{\alpha}=34.9^{\circ}$. This distribution was obtained by IPET (2006) using NOAA's HURDAT data set (Jarvinen et al., 1984) and found to describe well storms with central pressure deficit $\Delta P_{\max }>34 \mathrm{mb}$ that make landfall in the longitudinal range $85^{\circ}-95^{\circ} \mathrm{W}$.

Several studies (Holland, 1980; Atkinson and Holiday, 1977; Willoughby and Rahn, 2004) have used theoretical arguments and pressure-wind observations to relate $V_{\max }$ to $\Delta P_{\max }$. The relationships are typically of the power-law type

$$
\begin{equation*}
V_{\max }=c\left(\Delta P_{\max }\right)^{g} \tag{3.13}
\end{equation*}
$$

where $c$ and $g$ are positive constants. Using flight level data from 23 hurricane seasons, Willoughby and Rahn (2004) found $c=4.8$ and $g=0.559$ for $V_{\max }$ in $\mathrm{m} / \mathrm{s}$ and $\Delta P_{\max }$ in mb . Based on these and other findings of Willoughby and Rahn (2004), we model [ $V_{\max } \Delta P_{\max }$ ] as a lognormal variable with mean value $4.8\left(\Delta P_{\max }\right)^{0.559}$ and coefficient of variation 0.15 .

Empirical evidence (Vickery and Twisdale, 1995; Vickery et al., 2000; Willoughby and Rahn, 2004; Powell et al., 2005; IPET, 2008) and theoretical arguments (Shen, 2006) show that $R_{\max }$ increases when the hurricane intensity $\Delta P_{\max }$ decreases or the latitude $\varphi$ increases. Here we
assume that $\left(\ln R_{\max } \Delta P_{\max }\right)$ has the normal distribution proposed by Vickery et al. (2000), which for the region of New Orleans ( $\varphi \approx 30^{\circ} \mathrm{N}$ ) has parameters

$$
\begin{align*}
m_{\ln R_{\max } \Delta P_{\max }} & =3.962-0.00567 \Delta P_{\max } \\
\sigma_{\ln R_{\max }} \Delta P_{\max } & =0.313 \tag{3.14}
\end{align*}
$$

where $R_{\max }$ is in km and $\Delta P_{\max }$ is in mb .
The translational speed $V_{t}$ has weak dependence on the intensity of the TC (Chen et al., 2006; IPET, 2008) and is usually modeled as a lognormal variable with mean value around $6 \mathrm{~m} / \mathrm{s}$ and standard deviation around $2.5 \mathrm{~m} / \mathrm{s}$; see Vickery and Twisdale (1995), Vickery et al. (2000), and Chen et al. (2006). The former two studies report a slight dependence of $V_{t}$ on the approach angle $\alpha$. To keep the TC recurrence model simple, we use for $V_{t}$ a lognormal distribution with the above mean value and standard deviation and assume that $V_{t}$ and $\alpha$ are independent.

Different studies have concluded that the pressure deficit $\Delta P_{\max }$ has lognormal, Weibull or Gumbel distribution. The Weibull distribution gives better fits when all tropical cyclones are considered, whereas the lognormal distribution is more appropriate for storms in the hurricane intensity range; see Vickery and Twisdale (1995), Chouinard et al. (1997) and IPET (2006). The Gumbel distribution has been suggested by IPET (2008) for storms in the CAT35 range ( $\Delta P_{\max }>$ 58 mb ). While the Gumbel distribution is appropriate for the analysis of surges, winds and waves (for which the long-term risk is dominated by intense storms), significant rainfall is contributed by less intense slow-moving systems; see Section 3.5 . 2 below. For this reason we model $\Delta P_{\text {max }}$ using the lognormal distribution suggested by IPET (2006). This study shows that for TCs with $\Delta P_{\max }>34 \mathrm{mb}$ that made landfall in the longitudinal range $85^{\circ}-95^{\circ} \mathrm{W}, \Delta P_{\max }$ is accurately described by a shifted lognormal distribution with shift parameter 18 mb , log-mean 3.15 and logstandard deviation 0.68.

Finally, we set $\lambda=0.57$ events/yr, which is the rate found by IPET (2006) for TCs with $\Delta P_{\max }>34 \mathrm{mb}$ making landfall between $85^{\circ}-95^{\circ} \mathrm{W}$ along the Gulf of Mexico coast.

### 3.5.2 IDF curves for TC-rainfall and comparison with other storms

Next we use equation (3.11) with the recurrence model in Section 3.5.1 to estimate the intensity-duration-frequency (IDF) curves for New Orleans associated with tropical cyclones. The model explicitly accounts for variability in $y, V_{\max }, R_{\max }$ and $V_{t}$. All other input parameters to the MSR model are fixed to the values used in Sections 3.3 and 3.4. The joint density of $\left\{y, V_{\max }, R_{\max }, V_{t}\right\}$
for a TC that makes landfall between longitudes $85^{\circ}-95^{\circ} \mathrm{W}, f_{y, \boldsymbol{\theta}}$, is obtained by first calculating the joint density conditional on the pressure deficit $\Delta P_{\text {max }}$ under the assumption that the variables $y,\left[V_{\max } \Delta P_{\max }\right],\left[R_{\max } \Delta P_{\max }\right]$ and $V_{t}$ are independent and then averaging the conditional density with respect to $\Delta P_{\max }$.

Figure 3.13.a shows the calculated IDF curves as plots of rainfall intensity $i$ against duration $D$ for different return periods $T$.


Figure 3.13: Theoretical IDF curves for New Orleans obtained from equation (3.11). (a) Maximum rainfall intensity $i$ as a function of averaging duration $D$ for different return periods $T$. (b) Comparison of the IDF values in (a) for different averaging durations $D$ (solid lines) with intensities obtained from continuous rainfall records.

For durations below about 12 hours, the decay of $i$ with $D$ follows a power law $D^{-\gamma_{D}}$ where $\gamma_{D} \approx$ 0.55 . This exponent is slightly smaller than the values around $0.6-0.7$ that are typical of extratropical rainfall (because the rainfall intensities associated with long durations in TCs tend to be high relative to extra-tropical events); see for example Langousis et al. (2007). For longer averaging durations, the exponent $\gamma_{D}$ rapidly increases and is effectively 1 for $D>24$ hours; see dashed lines in Figure 3.13.a. The reason is that the passage of a hurricane usually lasts less than 24 hours; hence for $D>24$ the total rainfall depth is approximately constant and the average rainfall intensity depends on $D$ like $D^{-1}$.

Figure 3.13.b shows the same results as plots of $T$ against $i$ for different averaging durations D. The curves are compared with values from TP-40 (Hersfield, 1961); Babak et al. (1991) and Singh and Zhang (2007) for return periods $T=5,10,25,50$ and 100 yr . The latter values refer to generic rainfall in the New Orleans area and therefore include both TC and non-TC events. The rainfall values reported in TP-40 cover the whole duration range from $0.5-24 \mathrm{hrs}$, whereas Babak et al. (1991) and Singh and Zhang (2007) give rainfall values only for $D=6,12$ and 24 hr . It is clear from Figure 3.13.b that for $T>100 \mathrm{yr}$ also the dependence of the rainfall intensity on $T$ is of the power-law type, say $T^{\gamma_{T}}$ with $\gamma_{T} \approx 0.32$. This exponent is higher than the values around 0.20 0.25 that are typical of ordinary rainfall (Langousis et al., 2007; Veneziano et al., 2006b). The higher exponent in tropical cyclones is related to the large dispersion of the amplification factor $\beta_{l, \max }$ (see example plots in Figure 3.12).

Another feature of the TC curves in Figure 3.13.b is the lower asymptote at $T=1 / \lambda=1.75 \mathrm{yr}$. This lower bound is a consequence of the fact that the return period of any TC-induced rainfall intensity cannot be lower than the return period of the TCs themselves. The effect of this lower bound is that for short return periods, say $T<10 \mathrm{yr}$, the precipitation intensities from tropical cyclones are far below those from ordinary rainfall (frontal events, mesoscale convective systems etc.), for which the recurrence rate is much higher. By contrast, for long durations ( $D>$ 12 hr ) and long return periods ( $T=100 \mathrm{yr}$ ), the calculated TC intensities are close to the empirical intensities, indicating that tropical cyclones have a dominant effect on those extreme values. Given that the TC curves in Figure 3.13.b are flatter than those for overall rain, it is expected that tropical cyclones become even more dominant for longer return periods. For short durations (e.g. $D$ on the order of 1 hr ), the contribution of tropical rainfall to the risk is negligible, irrespective of the return period.

It is also of interest to determine which tropical cyclones contribute the most to rainfalls in excess of the IDF values $i(D, T)$. Such TCs might for example be used as scenario events when designing for return period $T$. The main parameters to be considered are $\theta=\left[V_{\max }, R_{\max }, V_{t}\right]$ and the distance $y$ to the cyclone center. Their modal (most likely) values are obtained by maximizing the conditional probability density of $(y, \boldsymbol{\theta})$ given $I_{D, \max }>i(D, T)$. This conditional density is given by

$$
\begin{equation*}
f_{y, \theta D, T}(y, \boldsymbol{\theta}) \propto f_{y, \boldsymbol{\theta}}(y, \boldsymbol{\theta}) P\left[I_{D, \max }(y, \boldsymbol{\theta})>i(D, T)\right] \tag{3.15}
\end{equation*}
$$

Figure 3.14 shows the modal values of $V_{\max }, R_{\max }$ and $V_{t}$ for different $D$ and $T$. The most likely distance $y$ always satisfies $y \approx R_{\max }$. This makes sense because $R_{\max }$ is the distance at which the MSR model predicts maximum large-scale rainfall intensities.

Figure 3.14.a shows that the mode of $V_{\max }$ increases when either $D$ or $T$ increase. This makes physical sense since for any given $D$, higher rainfall intensities require more intense storms, and for any given $T$, intense precipitation over longer durations is associated with more intense systems. Figure 3.14.b shows that the mode of $V_{t}$ decreases as $T$ increases, meaning that more intense rainfall is generally produced by slower-moving systems. For durations smaller than 12 hours, the modal value of $V_{t}$ is insensitive to $D$, whereas for longer durations $V_{t}$ decreases faster with $T$. This faster decay is related to the fact that, for durations $D$ on the order of one day or longer, extremely high rainfall intensities are produced by storms that take a time close to $D$ to pass over the site. Therefore, for $T$ large the translation speed $V_{t}$ tends to be inversely proportional to $D$. Finally, Figure 3.14.c shows that the mode of $R_{\max }$ decreases when either $D$ or $T$ increase. This makes sense, since more intense storms tend to have smaller values of $R_{\max }$; see Section 3.5.1.


Figure 3.14: Modal values of ( $V_{\max }, V_{t}, R_{\max }$ ) conditioned on exceeding the $T$-yr rainfall intensity for duration $D=$ $0.5,1,3,6,12$ and 24 hr .

### 3.6 Conclusions

We have developed a methodology to assess the frequency of extreme rainfall intensities from tropical cyclones (TCs) in coastal areas. The mean rainfall field associated with a TC with maximum tangential wind speed $V_{\max }$, radius of maximum winds $R_{\max }$, and translation speed $V_{t}$ is
obtained using a physically-based ("MSR") model (Chapter 2), whereas rainfall variability at both large scales (from storm to storm) and small scales (due to rainbands and local convection within a single storm) is modeled statistically. The statistical component of the model is estimated using 38 precipitation radar (PR) frames from the TRMM mission; see Table 3.1. These frames cover a wide range of TC intensities $V_{\max }$ and vortex sizes $R_{\max }$. To make the model easier to use in risk analysis, we developed approximate analytical expressions for the statistical parameters. We use Taylor's hypothesis to convert spatial rainfall intensity fluctuations to temporal fluctuations as the storm passes over a given geographical location A . The combined physical-statistical model predicts the maximum rainfall intensity at A during a period of duration $D$ for a TC with characteristics $\left(V_{\max }, R_{\max }, V_{t}\right)$ whose center passes at distance $y$ from A. To illustrate the use of the model for long-term rainfall risk analysis, we formulated a recurrence model for tropical cyclones in the Gulf of Mexico that make landfall between longitudes $85^{\circ}-95^{\circ} \mathrm{W}$ and used the rainfall and recurrence models to assess the rainfall risk for New Orleans. Our main findings are as follows.

The maximum rainfall $I_{l, \text { max }}$ in a spatial interval of length $l$ depends on $l$, the distance $y$ from the center of the TC, and the intensity $V_{\max }$ and size $R_{\max }$ of the vortex. We expressed $I_{l, \max }$ as the product of the large-scale ( $L \approx 400 \mathrm{~km}$ ) average rainfall intensity produced by the MSR model, $I_{L, M S R}$, and an amplification factor $\beta_{l, \max }$ that includes both storm-to-storm variability and spatial fluctuations of rainfall intensity within a storm. The distribution of $\beta_{l, \max }$ depends of course on $l$, but in addition depends significantly on the large-scale intensity $I_{L, M S R}$ and the normalized distance from the storm center, $y^{\prime}=\left|y / R_{\max }\right|$. Specifically, the dispersion of $\beta_{l, \max }$ increases as $l$ and $I_{L, M S R}$ decrease or $y^{\prime}=\left|y / R_{\max }\right|$ increases. These trends with $I_{L, M S R}$ and $y^{\prime}$ are linked to the fact that lower intensity storms and larger distances $y^{\prime}$ are associated with higher dry area fractions, more intermittent rainfall, and therefore an increased dispersion of the rainfall maxima.

Application of the model to TC rainfall risk for New Orleans has produced interesting insight into the importance of tropical cyclones relative to other rainfall-producing events. For short return periods $T$, the TC intensities are significantly below those from other storms, which have a much higher rate of occurrence. However, as the return period $T$ increases, the TC estimates for long averaging durations ( $D$ around 12-24 hours) approach the values found from continuous rainfall records. This means that for long return periods, the long-duration TC rainfalls tend to dominate. In New Orleans, this happens for $T$ around 100 years.

To determine how the most likely TC scenario varies with the averaging duration $D$ and the return period $T$, we calculated the joint distribution of $\left\{V_{\max }, R_{\max }, V_{t}, y\right\}$ conditioned on exceeding the $T$-yr rainfall intensity for duration $D$. Then we plotted the modal values of $V_{\max }$, $R_{\max }$, and $V_{t}$ against $D$ and $T$; see Figure 3.14 (for $y$, the modal value is always close to $R_{\max }$ ). The modal value of $V_{\max }$ increases when $D$ or $T$ increase, whereas the opposite is true for $R_{\max }$. The mode of the translation velocity $V_{t}$ is insensitive to $D$ for $D<24 \mathrm{hr}$, but decreases with increasing $T$ and with increasing $D$ for $D>24 \mathrm{hr}$.

## 4. Conclusions

In this thesis we combined data analysis and physical modeling to develop a methodology for the frequency of extreme rainfall intensities caused by tropical cyclones (TCs) in coastal areas. The mean rainfall field associated with a TC with maximum tangential wind speed $V_{\max }$, radius of maximum winds $R_{m a x}$, and translation speed $V_{t}$ is obtained using a physically-based model (Chapters 1 and 2), whereas rainfall variability at both large scales (from storm to storm) and small scales (due to rainbands and local convection) is modeled statistically (Chapter 3).

The theoretical model of TC rainfall is based on the assumption that the vertical outflow of water vapor from the TC boundary layer (BL) is all converted into rainfall. This water vapor flux originates from the low-level convergence of the horizontal flow. Evaluation of the BL winds is based on Smith's (1968) axi-symmetric formulation, modified in Chapter 1 to account for storm motion. The resulting models of wind and rainfall are referred to as the modifiedSmith (MS) BL model and the modified-Smith-for-rainfall (MSR) model, respectively. Our main findings from Chapters 1-3 are outlined below.

## MS model for evaluation of the boundary layer winds

The MS scheme estimates the horizontal and vertical wind velocity fields inside the BL from a few TC characteristics: the maximum tangential wind speed $V_{\max }$, the radius of maximum winds $R_{m a x}$, and Holland's $B$ parameter, in addition to the storm translation velocity $V_{t}$, the surface drag coefficient $C_{D}$, and the vertical diffusion coefficient of the horizontal momentum $K$. The scheme is computationally very efficient and is stable also for large $B$ values and fast-moving storms. Model results were compared to those from other studies (Shapiro 1983; Kepert 2001) and validated using the Fifth-Generation Pennsylvania State University/NCAR Mesoscale Model (MM5).

We find that Kepert's (2001) BL model significantly underestimates the radial and vertical fluxes, whereas Shapiro's (1983) slab-layer formulation produces radial and vertical winds that are a factor of about 2 higher than those produced by MM5. The velocity fields generated by the present model are consistent with MM5 and with tropical cyclone observations.

In a sensitivity analysis we used the MS model to study how the symmetric and asymmetric components of the wind field vary with the storm parameters mentioned above. In accordance with observations, we find that larger values of $B$ and lower values of $R_{\max }$ produce horizontal
and vertical wind profiles that are more picked near the radius of maximum winds. We also find that, when cyclones in the Northern (Southern) hemisphere move, the vertical and storm-relative radial winds intensify at the right-front (left-front) quadrant of the vortex, whereas the stormrelative tangential winds are more intense in the left-front (right-front) region. The asymmetry is higher for faster moving TCs and for higher surface drag coefficients $C_{D}$.

## MSR model for evaluation of the mean rainfall field in TCs

The MSR model calculates the upward moisture flux that converts into rainfall from the vertical winds generated by the MS model and two additional parameters: the average temperature $\bar{T}$ and average saturation ratio $\bar{Q}$ inside the TC boundary layer. The rainfall field generated by the MSR model is not simply proportional to the vertical wind field at the top of the boundary layer produced by the MS model. This is because (1) the trajectory of moisted air parcels has an outward slant depending on distance from the TC center and (2) the ascending air parcels and descending rain drops are advected into a helical motion by the cyclonic circulation; therefore a parcel of air that leaves the TC boundary layer contributes rainfall to a range of azimuthal locations. The azimuthal redistribution is important only for radially asymmetric flows; i.e. when $V_{t} \neq 0$. The relation between vertical winds and surface rainfall has been verified through MM5 simulations. The MSR model accurately reproduces the rainrate estimates extracted from the precipitation radar (PR) data of the TRMM mission.

In a sensitivity analysis we found that the temperature $\bar{T}$, the saturation ratio $\bar{Q}$ and the vertical diffusion coefficient $K$ have small effect on the calculated rainfall intensities. Consequently, setting those parameters to constant values (e.g. to $\bar{T}=22^{\circ} \mathrm{C}, \bar{Q}=0.8$ and $K=50 \mathrm{~m}^{2} / \mathrm{s}$, as was done in Sections 2.3-2.5) does not induce large errors. By contrast, the maximum tangential velocity $V_{\max }$ and the roughness of the surface boundary (expressed through $C_{D}$ ) have significant effects on rainfall intensity. The latter increases when either $V_{\max }$ or $C_{D}$ increase. Lower values of $R_{\max }$ produce rain rates that are more picked and more concentrated near the TC center. The $B$ parameter has a small effect on the peak rainfall intensity, but influences the rate at which rainfall decays with radial distance (higher values of $B$ result in faster decay).

For a TC that moves in the Northern (Southern) hemisphere the effect of the translation velocity $V_{t}$ is to intensify rainfall at the right (left) and right-front (left-front) of the storm.

Larger asymmetries are associated with higher velocities $V_{t}$. These findings are in accordance with empirical observations (Lonfat et al., 2004). The effect of $C_{D}$ is more complex: at the front of the storm, rainfall asymmetry is insensitive to $C_{D}$, whereas at the rear-right (rear-left) the rainfall asymmetry increases with decreasing $C_{D}$.

An important departure from previous studies is that we parameterize rainfall asymmetries in TCs in terms of storm motion not vertical wind shear; see Chapter 2. We have used MSR model simulations and empirical data on rainfall asymmetry (Chen et al., 2006) to verify that the asymmetries produced by storm motion are close to those observed and often parameterized in terms of vertical wind shear; see Section 2.5 . Since, in risk analysis, one may equivalently use shear or motion as conditioning parameter and, in addition, it is easier to include motion than shear when modeling rainfall (see Section 3.5), we have chosen to develop a motion-based rather than shear-based parameterization of rainfall asymmetry.

## Stochastic fluctuation component and an application to rainfall risk for New Orleans

The MSR model should prove useful for climatologic studies, but for hazard analysis it has the major limitation of ignoring the small-scale rainfall variations associated with rainbands and local convection. These variations have an important effect on the distribution of $I_{D, \max }$ : the maximum of the average rainfall intensity over duration $D$ at location A during the passage of a storm. We obtained the distribution of $I_{D, \max }$ by adding a stochastic component to the MSR model predictions.

The combined physical-statistical model predicts the maximum rainfall intensity at A during a period of duration $D$ for a TC with characteristics ( $V_{\max }, R_{\max }, V_{t}$ ) whose center passes at distance $y$ from A (the other parameters of the MSR model are set to $K=50 \mathrm{~m}^{2} / \mathrm{s}, C_{D}=0.002, \bar{T}$ $=22^{\circ} \mathrm{C}, \bar{Q}=0.8$ and $B=1$ ). To illustrate the use of the model for long-term rainfall risk analysis, we formulated a recurrence model for tropical cyclones in the Gulf of Mexico. We then used the rainfall and recurrence models to assess the rainfall risk for New Orleans and compared our results with empirical IDF estimates for all rainstorms (TCs and non-TCs) in the region.

We found that for short return periods $T$, the TC intensities are significantly below those from other storms, which have a much higher rate of occurrence. However, as the return period $T$ increases, the TC estimates for long averaging durations ( $D$ around 12-24 hours) approach the
values found from continuous rainfall records. This means that for long return periods, the longduration TC rainfalls tend to dominate. In New Orleans, this happens for $T$ around 100 years.

We also determined how the most likely TC scenario varies with the averaging duration $D$ and the return period $T$. We did so by plotting the modal values of $V_{\max }, R_{\max }$, and $V_{t}$ conditioned on exceeding the $T$-yr rainfall intensity for duration $D$ (for $y$, the modal value is always close to $R_{\max }$, the location where the MSR model predicts maximum large-scale rainfall intensities). These modal values can for example be used as scenario events when designing for return period $T$.

The mode of $V_{t}$ decreases as $T$ increases. This means that more intense rainfall is generally produced by slower-moving systems. For durations smaller than 12 hours, the modal value of $V_{t}$ is insensitive to $D$, whereas for longer durations $V_{t}$ decreases faster with $T$. This faster decay is related to the fact that, for durations $D$ on the order of one day or longer, extremely high rainfall intensities are produced by storms that take a time close to $D$ to pass over the site.

The modal value of $R_{\max }$ decreases when $D$ or $T$ increase, whereas the opposite is true for $V_{\max }$. Both findings make sense since intense storms tend to have smaller values of $R_{\max }$ (see Section 3.5.1) and for any given $D$, higher rainfall intensities require more intense storms, and for any given $T$, intense precipitation over longer durations is associated with more intense systems.

## Future research directions

A rich parameterization and high computational efficiency make the proposed model attractive for rainfall risk applications in TC-prone areas, where one must assess tropical cyclone rainfall under many storm and environmental scenarios. This requirement effectively rules out the use of full-physics high-resolution numerical weather prediction models.

Although computationally efficient, the current version of the model requires numerical calculation of the large-scale rainfall intensity $I_{L, M S R}$ in equation (3.2) using the MS and MSR models. In Section 3.4 we found that when $V_{\max }$ and $y^{\prime}=\left|y / R_{\max }\right|$ are kept fixed, the sensitivity of $I_{L, M S R}$ to both $V_{t}$ and $R_{m a x}$ is low. Hence, one could significantly reduce the computational effort and make the model more user-friendly by parameterizing $I_{L, M S R}$ in terms of $V_{\max }$ and $y^{\prime}$. This parameterization would be worth pursuing.

A limitation of the current framework is that it does not account for post-landfall conditions and therefore is applicable only to open-water or coastal sites. When extending the model to
inland locations, one should investigate whether the statistics of the small scale rainfall variations in Chapter 3 remain valid over land. Based on findings of other studies on the effects of orography on the spatial variability of extra-tropical rainfall (Pathirana and Herath, 2002; Badas et al., 2005; Deidda et al., 2006), we anticipate that for sites not very far inland the statistical model for the rainfall fluctuations would not change. In this case, one can use a correction factor to multiply the theoretical $I_{L, M S R}$ rainrates to account for the effects of topography at inland locations. For example, Lonfat et al. (2007) use a simple formulation where this factor is proportional to the product of surface winds and the gradient of ground elevation.

It is possible to extend the current framework to deal with short term rainfall prediction. For this purpose one needs to start with real-time information on the storm and use some type of a TC-track model to calculate the distribution of the location $y$ and the characteristics $\boldsymbol{\theta}$ of the storm at landfall. The joint distribution of $(y, \boldsymbol{\theta})$ can be used as a short-term recurrence model in equation 3.11 (for $\lambda=1$ ) to assess rainfall risk at the site. In this case one should check the accuracy of the rainfall estimates produced by the suggested model relative to those calculated by simple translation and possibly intensification/deintensification of real-time radar rainfall estimates over the site. The latter method is usually referred to as tropical rainfall potential (TRaP) technique (e.g. Kidder et al., 2005). We anticipate that for leads shorter than a day the TRaP technique should produce more accurate rainfall estimates.

Another possible research direction is to extend the current framework to maximum TC wind speeds. For example, one could use the MS model to predict the mean wind field in a TC with given characteristics $\boldsymbol{\theta}$ and analysis of surface wind observations to evaluate statistically the wind gust factor.

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## Appendix A: Analytical Expressions for the Parameters in Equation (1.15)

Let $I_{1}=\int_{0}^{\infty} f^{2} d \eta, I_{2}=\int_{0}^{\infty}\left(1-g^{2}\right) d \eta, I_{3}=\int_{0}^{\infty}(1-g) d \eta, I_{4}=\int_{0}^{\infty} f g d \eta, I_{5}=\int_{0}^{\infty} f d \eta I_{6}=\int_{0}^{\infty}\left(g^{2}-g\right) d \eta$ and $I_{7}=\int_{0}^{\infty}(1-g)^{2} d \eta$ where $f$ and $g$ are the functions given in equation (1.11). Then

$$
\begin{align*}
& \left.\frac{\partial f}{\partial \eta}\right|_{\eta=0}=a_{2}-a_{1},\left.\quad \frac{\partial g}{\partial \eta}\right|_{\eta=0}=a_{1}-a_{2} \\
& I_{1}=\frac{1}{8}\left(a_{1}^{2}+2 a_{1} a_{2}+3 a_{2}^{2}\right), \quad I_{2}=\left(a_{1}+a_{2}\right)-\frac{1}{8}\left(3 a_{1}^{2}+2 a_{1} a_{2}+a_{2}^{2}\right) \\
& I_{3}=\frac{a_{1}+a_{2}}{2}, \quad I_{4}=\frac{1}{8}\left(a_{1}^{2}+4 a_{1} a_{2}+a_{2}^{2}\right)-\frac{a_{1}+a_{2}}{2}  \tag{A.1}\\
& I_{5}=-\frac{a_{1}+a_{2}}{2}, \quad I_{6}=\frac{1}{8}\left(3 a_{1}^{2}+2 a_{1} a_{2}+a_{2}^{2}\right)-\frac{a_{1}+a_{2}}{2} \\
& I_{7}=\frac{1}{8}\left(3 a_{1}^{2}+2 a_{1} a_{2}+a_{2}^{2}\right)
\end{align*}
$$

where $a_{1}$ and $a_{2}$ are calculated from equation (1.12). Under equation (A.1), the parameters $A_{1}-A_{20}$ in equation (1.15) are given by

$$
\begin{gather*}
A_{1}=E^{2}\left[v_{t}^{2} \cos ^{2} \theta I_{7}+\left(v_{t} \sin \theta-v_{\mathrm{gr}}\right)^{2} I_{1}+\left(I_{5}-I_{4}\right)\left(v_{t}^{2} \sin 2 \theta-2 v_{t} v_{g r} \cos \theta\right)\right]  \tag{A.2}\\
A_{2}=E\left[\left(v_{t}^{2} \cos 2 \theta+v_{t} v_{g r} \sin \theta\right) I_{5}+\left(1 / 2 v_{t}^{2} \sin 2 \theta-v_{t} v_{g r} \cos \theta\right) I_{1}+\left(v_{g r}^{2}-2 v_{t} v_{g r} \sin \theta-v_{t}^{2} \cos 2 \theta\right) I_{4}+\right. \\
\left.v_{t} v_{g r} I_{6} \cos \theta-1 / 2 I_{7} v_{t}^{2} \sin 2 \theta\right] \\
A_{3}=-v_{t}^{2} I_{1} \cos ^{2} \theta+v_{t}^{2} I_{5} \sin 2 \theta+\left(2 v_{t} v_{g r} \cos \theta-v_{t}^{2} \sin 2 \theta\right) I_{4}-v_{t}^{2} I_{7} \sin ^{2} \theta+v_{g r}^{2} I_{2}+2 v_{t} v_{g r} I_{6} \sin \theta \\
A_{4}=v_{t}^{2} I_{1} \cos ^{2} \theta-v_{t}^{2} I_{5} \sin 2 \theta+\left(v_{t}^{2} \sin 2 \theta-2 v_{t} v_{g r} \cos \theta\right) I_{4}+v_{t}^{2} I_{7} \sin ^{2} \theta-2 v_{t} v_{g r} I_{6} \sin \theta \\
A_{5}=v_{t} I_{5} \cos \theta+\left(v_{g r}-v_{t} \sin \theta\right) I_{3} \\
A_{7}=E\left[\left(v_{g r}-v_{t} \sin \theta\right) I_{5}-v_{t} I_{3} \cos \theta\right], A_{8}=-v_{t} \cos \theta I_{5}+v_{t} I_{3} \sin \theta \\
A_{6}=E\left[\left.v_{t} \cos \theta \frac{\partial g}{\partial \eta}\right|_{\eta=0}+\left.\left(v_{g r}-v_{t} \sin \theta\right) \frac{\partial f}{\partial \eta}\right|_{\eta=0}\right]  \tag{A.8}\\
A_{9}=-\left.v_{t} \cos \theta \frac{\partial f}{\partial \eta}\right|_{\eta=0}+\left.\left(v_{g r}-v_{t} \sin \theta\right) \frac{\partial g}{\partial \eta}\right|_{\eta=0} \tag{A.9}
\end{gather*}
$$

$$
\begin{align*}
& A_{10}=2 E\left[v_{t}^{2} I_{7} \cos ^{2} \theta+I_{1}\left(v_{g r}-v_{t} \sin \theta\right)^{2}+\left(v_{t}^{2} \sin 2 \theta-2 v_{t} v_{g r} \cos \theta\right)\left(I_{5}-I_{4}\right)\right]  \tag{A.10}\\
& A_{11}= E^{2}\left[v_{t}^{2} \cos ^{2} \theta \frac{\partial I_{7}}{\partial r}-2 I_{1}\left(v_{t} \sin \theta-v_{g r}\right) \frac{\partial v_{g r}}{\partial r}+\left(v_{g r}-v_{t} \sin \theta\right)^{2} \frac{\partial I_{1}}{\partial r}-2 v_{t}\left(I_{5}-I_{4}\right) \frac{\partial v_{g r}}{\partial r} \cos \theta+\right. \\
&\left.\left(v_{t}^{2} \sin 2 \theta-2 v_{t} v_{g r} \cos \theta\right) \frac{\partial\left(I_{5}-I_{4}\right)}{\partial r}\right]  \tag{A.11}\\
& A_{12}=\left(v_{t}^{2} \cos 2 \theta+v_{t} v_{g r} \sin \theta\right) I_{5}+\left(1 / 2 v_{t}^{2} \sin 2 \theta-v_{t} v_{g r} \cos \theta\right) I_{1}+\left(v_{g r}^{2}-2 v_{t} v_{g r} \sin \theta-v_{t}^{2} \cos 2 \theta\right) I_{4}+ \\
& v_{t} v_{g r} I_{6} \cos \theta-1 / 2 v_{t}^{2} I_{7} \sin 2 \theta  \tag{A.12}\\
& A_{13}= E\left[\left(-2 v_{t}^{2} \sin 2 \theta+v_{t} v_{g r} \cos \theta\right) I_{5}+\left(v_{t}^{2} \cos 2 \theta+v_{t} v_{g r} \sin \theta\right)\left(\frac{\partial I_{5}}{\partial \theta}+I_{1}\right)+\left(1 / 2 v_{t}^{2} \sin 2 \theta-v_{t} v_{g r} \cos \theta\right) \frac{\partial I_{1}}{\partial \theta}\right. \\
&+\left(2 v_{t}^{2} \sin 2 \theta-2 v_{t} v_{g r} \cos \theta\right) I_{4}+\left(v_{g r}^{2}-2 v_{t} v_{g r} \sin \theta-v_{t}^{2} \cos 2 \theta\right) \frac{\partial I_{4}}{\partial \theta}-v_{t} v_{g r} I_{6} \sin \theta+v_{t} v_{g r} \cos \theta \frac{\partial I_{6}}{\partial \theta} \\
&\left.-v_{t}^{2} I_{7} \cos 2 \theta-1 / 2 v_{t}^{2} \sin 2 \theta \frac{\partial I_{7}}{\partial \theta}\right]  \tag{A.13}\\
& A_{14}= E\left[v_{t} \sin \theta I_{5} \frac{\partial v_{g r}}{\partial r}+\left(v_{t}^{2} \cos 2 \theta+v_{t} v_{g r} \sin \theta\right) \frac{\partial I_{5}}{\partial r}-v_{t} \frac{\partial v_{g r}}{\partial r} I_{1} \cos \theta+\left(1 / 2 v_{t}^{2} \sin 2 \theta-v_{t} v_{g r} \cos \theta\right) \frac{\partial I_{1}}{\partial r}\right. \\
&+2 I_{4}\left(v_{g r}-v_{t} \sin \theta\right) \frac{\partial v_{g r}}{\partial r}+I_{6} v_{t} \cos \theta \frac{\partial v_{g r}}{\partial r}+\left(v_{g r}^{2}-2 v_{t} v_{g r} \sin \theta-v_{t}^{2} \cos 2 \theta\right) \frac{\partial I_{4}}{\partial r}+ \\
&\left.v_{t} v_{g r} \cos \theta \frac{\partial I_{6}}{\partial r}-1 / 2 v_{t}^{2} \sin 2 \theta \frac{\partial I_{7}}{\partial r}\right]
\end{align*}
$$

$$
\begin{align*}
A_{15}= & -v_{t}^{2} I_{1} \sin 2 \theta+v_{t}^{2} \cos ^{2} \theta \frac{\partial I_{1}}{\partial \theta}-2 v_{t}^{2} I_{5} \cos 2 \theta-v_{t}^{2} \sin 2 \theta \frac{\partial I_{5}}{\partial \theta}+\left(2 v_{t}^{2} \cos 2 \theta+2 v_{t} v_{g r} \sin \theta\right) I_{4}+ \\
& \left(v_{t}^{2} \sin 2 \theta-2 v_{t} v_{g r} \cos \theta\right) \frac{\partial I_{4}}{\partial \theta}+v_{t}^{2} I_{7} \sin 2 \theta+v_{t}^{2} \sin ^{2} \theta \frac{\partial I_{7}}{\partial \theta}-2 v_{t} v_{g r} I_{6} \cos \theta-2 v_{t} v_{g r} \sin \theta \frac{\partial I_{6}}{\partial \theta} \tag{A.15}
\end{align*}
$$

$$
\begin{equation*}
A_{16}=\left(v_{g r}-v_{t} \sin \theta\right) I_{5}-v_{t} I_{3} \cos \theta \tag{A.16}
\end{equation*}
$$

$$
\begin{equation*}
A_{17}=E\left[I_{5} \frac{\partial v_{g r}}{\partial r}+\left(v_{g r}-v_{t} \sin \theta\right) \frac{\partial I_{5}}{\partial r}-v_{t} \cos \theta \frac{\partial I_{3}}{\partial r}\right] \tag{A.17}
\end{equation*}
$$

$$
A_{18}=v_{t} I_{5} \sin \theta-v_{t} \cos \theta \frac{\partial I_{5}}{\partial \theta}+v_{t} I_{3} \cos \theta+v_{t} \sin \theta \frac{\partial I_{3}}{\partial \theta}
$$

$$
\begin{equation*}
A_{19}=v_{g r}^{2} I_{6}, A_{20}=v_{g r} \frac{\partial I_{6}}{\partial \theta} \tag{A.19}
\end{equation*}
$$

In equations (A.2)-(A.19), the derivatives $\frac{\partial I_{i}}{\partial r}$ and $\frac{\partial I_{j}}{\partial \theta}$ for $j=1, \ldots, 7$ can be calculated analytically using the chain rule

$$
\begin{equation*}
\frac{\partial I_{i}}{\partial s}=\frac{\partial I_{i}}{\partial a_{1}} \frac{\partial a_{1}}{\partial s}+\frac{\partial I_{i}}{\partial a_{2}} \frac{\partial a_{2}}{\partial s}, \quad j=1, . ., 7 \tag{A.20}
\end{equation*}
$$

where $s$ is either $r$ or $\theta$ and $\frac{\partial a_{1}}{\partial s}$ and $\frac{\partial a_{2}}{\partial s}$ can be analytically derived from equations (1.12) and (1.13).

## Appendix B: Alternative Methods to Calculate the Distribution of the Amplification Factor $\gamma_{l \text { max }}$

In Section 3.3.2 we used data on $\gamma_{l, \max }$ to fit a theoretical distribution model for the amplification factor. Alternatively, one can evaluate the distribution of $\gamma_{l, \max }$ by studying the maximum of the moving average process $I_{l}(x)$ illustrated in Figure 3.2.

Here we develop two alternative approaches for the distribution of $\gamma_{l, \max }$ starting from the ratio $\gamma_{l}$, defined as

$$
\begin{equation*}
\gamma_{l}=\frac{I_{l}}{I_{L}}, l \leq L \tag{B.1}
\end{equation*}
$$

where $I_{l}$ is obtained by continuously averaging the PR rainrate along a cross section C (see Figure 3.1) using a window of length $l$ and $I_{L}=I_{l=L}$.

The distribution of $\gamma_{l, \max }$ is calculated from the distribution of $\gamma_{l}$ under the assumption of independence and then corrected for dependence and continuous sampling; see Sections B. 1 and B.2. Since we calculate the distribution of the maximum amplification factor $\gamma_{l, \max }$ starting from a marginal distribution (i.e. the distribution of $\gamma_{l}$ ), we refer to those approaches as the "marginal approaches". The approach in Section 3.3.2 is referred to as the "maxima approach". Section B. 3 compares results on the total amplification factor $\beta_{l, \max }$ in equation (3.2), as well as calculated IDF values for New Orleans using the three approaches for $\gamma_{l, \max }$.

## B. 1 Marginal Approach Using a Modified-beta Distribution Model

We start by calculating the empirical moments of the ratio $\gamma_{l}$ in equation (B.1) using the frames in Table 3.1.

Figure B. 1 shows plots of $\ln E\left[\left(\gamma_{l}\right)^{q}\right]$ against $\ln l$ for $q=0,1,2,3$, for the same categories of the large-scale intensity $I_{L}$ as in Figure 3.8. One sees that the variability of the rainfall field, expressed here through $\ln E\left[\left(\gamma_{l}\right)^{q}\right]$ for $q>1$, increases with decreasing spatial scale $l$ and largescale rainfall intensity $I_{L}$. This is in accordance to what we found in Section 3.3.2 when we studied the moments of the amplification factor $\gamma_{l, \max }$. An additional observation one makes from Figure B. 1 is that an important contributor to the increased variability at low $I_{L}$ regions is the intermittency of the rainfall process.


Figure B.1: Plots of $\ln E\left[(\gamma)^{\left.)^{q}\right]}\right.$ against $\ln l$ for $q=0,1,2,3$, for the same categories of $I_{L}$ as in Figure 3.8. The markers indicate empirical moment values, whereas the solid lines correspond to the parametric fits in equation (B.2). The solid dashed lines correspond to the theoretical third moments calculated from the distribution in equation (B.9); see text for details.

The latter is expressed through the probability $P_{0}=1-E\left[\left(y_{l}\right)^{0}\right]$ that an interval of size $l$ is dry. One sees that $P_{0}$ increases with decreasing $I_{L}$. Since lower values of $I_{L}$ are associated with larger distances $y$ from the storm center, one concludes that the outer TC environment is more lacunar and exhibits more cellular structure and higher small-scale variability relative to the inner region. This is in accordance with the findings of other studies on the spatial variability of TC rainfall; see e.g. Molinari et al. (1994) and Lonfat et al. (2004).

For each intensity category $I_{L}$, we use the method of least squares to fit linear expressions to the logs of the $0^{\text {th }}, 1^{\text {st }}$ and $2^{\text {nd }}$ initial moments of $\gamma_{l}$ :

$$
\ln E\left[\left(\gamma_{l}\right)^{q}\right]= \begin{cases}b_{q} \ln \left(\frac{c_{q}}{l}\right) & , l \leq c_{q}  \tag{B.2}\\ 0 & , l>c_{q}\end{cases}
$$

where $l$ is in km and $b_{q}$ and $c_{q}(q=0,1,2)$ are parameters. The intercept of the log-linear expressions is equation (B.2), $c_{q}$, is set to 220 km for $q=0$ and to 384 km for $q=1$ and 2 . The former value of $c_{q}$ is the average of the $0^{\text {th }}$ moment intercepts from all $I_{L}$ categories, whereas the latter corresponds to the maximum scale $L$ used in the analysis.

Figure B. 2 shows how the parameters $b_{q}(q=0,1,2)$ in equation (B.2) vary with the largescale intensity $I_{L}$. The solid lines in Figure B. 2 are smooth least-squares estimates of $b_{q}(q=0,1$, $2)$.


Figure B.2: Plots of the parameters $b_{0}-b_{2}$ in equation (B.2) against $I_{L}$. Solid lines are smooth least-squares estimates.

## B.1.1 Distribution of $\boldsymbol{\gamma}_{l}$

The distribution of $\gamma_{l}$ has unit mean (see Figure B.1), it is bounded between 0 and $L / l$ and has concentrated masses at zero: $P\left[\gamma_{l}=0\right]=P_{0}=1-E\left[\left(\gamma_{l}\right)^{0}\right]$ and at $L / l: P\left[\gamma_{l}=L / l\right] . \gamma_{l}=0$ corresponds to the case of no rainfall inside an interval of size $l$, whereas $\gamma_{l}=L / l$ to the case when all rainfall volume is concentrated inside one $l$ interval. By assuming independence between different $l$ intervals, one obtains an estimate of the probability mass at $L / l: P\left[\gamma_{l}=L / l\right]=\left(P_{0}\right)^{L /-1}$. In this case, the mean value $m_{c}$ and variance $\sigma_{c}{ }^{2}$ of the conditional variable ( $\gamma_{l} \mid \gamma_{l} \neq 0, L / l$ ) are calculated to be:

$$
\begin{align*}
& m_{c}=1+\frac{P_{0}-(L / l-1)\left(P_{0}\right)^{L l-1}}{1-P_{0}-\left(P_{0}\right)^{L l-1}} \\
& \sigma_{c}{ }^{2}=\frac{E\left[\left(\gamma_{l}\right)^{2}\right]-1+\left(m_{c}-1\right)^{2}-P_{0} m_{c}{ }^{2}-\left(L / l-m_{c}\right)^{2}\left(P_{0}\right)^{L l-1}}{1-P_{0}-\left(P_{0}\right)^{L / l-1}} \tag{B.3}
\end{align*}
$$

where $P_{0}$ and $E\left[\left(\gamma_{l}\right)^{2}\right]$ can be obtained from equation (B.2) and the parametric expressions in Figure B.2, as functions of the spatial scale $l$ and the large-scale rainfall intensity $I_{L}$.

We approximate the conditional distribution of $\left(\gamma_{l} \gamma_{l} \neq 0, L / l\right)$ by using a beta distribution. In this case the cumulative distribution function (CDF) of $\gamma_{l}$ is given by

$$
F_{\gamma_{1}}(\gamma)= \begin{cases}P_{0} & , \gamma=0  \tag{B.4}\\ P_{0}+\left(1-P_{0}-P_{0}{ }^{L l-1}\right) F_{X}\left(\frac{l y}{L}\right) & , 0<\gamma<L / l \\ 1 & , \gamma \geq L / l\end{cases}
$$

where $F_{X}$ is a standard beta distribution in $[0,1]$ with mean value $m_{X}=m_{c} l / L$ and variance $\sigma_{X}{ }^{2}=$ $\left(\sigma_{c} l / L\right)^{2}$.

Figure B. 3 compares the empirical CDF of $\gamma$ at different spatial scales $l$ with those obtained theoretically by fitting the distribution model in equation (B.4) to the $0^{\text {th }}, 1^{\text {st }}$ and $2^{\text {nd }}$ initial moments obtained from equation (B.2). Comparison is done for the lowest ( $I_{L}=0-0.8 \mathrm{~mm} / \mathrm{h}$ ) and highest ( $I_{L}>10 \mathrm{~mm} / \mathrm{h}$ ) intensity categories in Figure B.1.

The distribution model in equation (B.4) produces good analytical fits to the empirical distribution of $\gamma$. Equally good fits have been obtained for all other intensity categories and spatial scales $l$. This is evident from Figure B.1, where we show the theoretical third moment of the distribution in equation (B.4) (thick dashed lines) when fitted to the $0^{\text {th }}, 1^{\text {st }}$ and $2^{\text {nd }}$ initial moments obtained from equation (B.2). The fact that the theoretical and empirical $3^{\text {rd }}$ initial
moments are close shows that the distribution in equation (B.4) produces good analytical fits to the distribution of $\gamma_{l}$.


Figure B.3: Comparison between the empirical and theoretical CDFs of the ratios $\gamma_{l}$ at different spatial scales $l$. Comparison is done for the lowest ( $I_{L}=0-0.8 \mathrm{~mm} / \mathrm{h}$ ) and highest ( $I_{L}>10 \mathrm{~mm} / \mathrm{h}$ ) intensity categories in Figure B.1.

## B.1.2 Distribution of $\gamma_{l, \max }$

With corrections to be made latter, we approximate $\gamma_{l, \max }$ in equation (3.6) by $\gamma_{l, \text { max,ind }}$, the maximum of $L / l$ independent copies of $\gamma_{l}$. In this case

$$
\begin{equation*}
F_{\gamma_{l, m a x, i n d}}(\gamma)=\left[F_{\gamma_{l}}(\gamma)\right]^{I l l}, \quad \gamma \geq 0 \tag{B.5}
\end{equation*}
$$

$\gamma_{l, \text { max, ind }}$ is biased relative to $\gamma_{l, \max }$ in two respects. The first is that, by definition [equation (B.1)], the variables $\gamma_{l, i}$ for $i=1, \ldots, L / l$ are dependent and average to 1 :

$$
\begin{equation*}
\frac{l}{L} \sum_{i=1}^{L / l} \gamma_{l, i}=\frac{l}{L} \sum_{i=1}^{L / l} \frac{I_{L, i}}{I_{L}}=1 \tag{B.6}
\end{equation*}
$$

The second source of bias is that $\gamma_{l, \max }$ in equation (3.6) is the maximum amplification factor obtained by continuously scanning each cross section of length $L$ using an averaging window of size $l$, whereas equation (B.5) considers non-overlapping intervals.

To quantify the two sources of bias, Figures B. 4 and B. 5 show the empirical distribution of $\gamma_{l, \max }$ at spatial scale $l=192 \mathrm{~km}$ obtained: as the maximum of two non-overlapping intervals of size $l=192 \mathrm{~km}$ (Figure B.4), and by continuously scanning each cross section of length $L=384 \mathrm{~km}$ using an averaging window of size $l=192 \mathrm{~km}$ (Figure B.5). Results are shown for the 12 intensity categories in Figure B.1.

The dotted lines in Figures B. 4 and B. 5 denote the theoretical PDF of $\gamma_{l, \text { max,ind }}$ under independence [equation (B.5)], whereas the solid blue lines show the exact PDF of the maximum of two dependent random variables $\gamma_{L / 2}$ with marginal distribution given by equation (B.4) that satisfy $\gamma_{L / 2}(1)+\gamma_{L 2}(2)=2$. For the latter, one obtains

$$
F_{\max \left\{\gamma_{L 2}(1), \gamma_{L 2}(2)\right\}}(z)= \begin{cases}0 & , z<1  \tag{B.7}\\ 2 F_{\gamma_{L 2}}(z)-1 & , z \geq 1\end{cases}
$$

One sees that $\gamma_{l, \max , \text { ind }}$ fails to reproduce the distribution of $\gamma_{l, \max }$ in both cases of non-overlaping intervals (Figure B.4) and continuous sampling (Figure B.5). Note in particular that the effect of dependence causes the distribution of $\gamma_{l, \max }$ to have a lower bound at 1 (this corresponds to a uniform rainfall inside $L$ ), whereas $\gamma_{l, m a x, \text { ind }}$ is bounded at 0 . The exact theoretical distribution in equation (B.7) resembles well the empirical distribution $\gamma_{l, \max }$ for the case of non-overlaping intervals (see Figure B.4), but underestimates the maxima under continuous sampling (see Figure B.5).


Figure B.4: Comparison of the empirical distribution of $\gamma_{l, \max }$ obtained for the case of two non-overlaping intervals of size $l=192 \mathrm{~km}$, and those calculated theoretically under independence [dotted lines; equation (B.5)] and including dependence [solid blue lines; equation (B.7)]. Comparison is done for the intensity categories in Figure B.1.


Figure B.5: Comparison of the empirical distribution of $\gamma_{l, \max }$ obtained for the case of continuous sampling using a sliding window of size $l=192 \mathrm{~km}$, and those calculated theoretically under independence [dotted lines; equation (B.5)], including dependence [solid blue lines; equation (B.7)], using the linear correction in equation (B.9) (dashed lines), using the maxima approach described in Section 3.3 .2 [solid black lines; equation (3.8)] and using the marginal approach with a modified lognormal (LN) distribution described in Section B. 2 (dashed-dotted lines). Comparison is done for the intensity categories in Figure B.1.

To correct the distribution of $\gamma_{l, m a x, \text { ind }}$ for the effects of dependence and continuous sampling, we use a linear transformation to rescale the CDF in equation (B.5) so it is bounded between 1 and $L / l$. Specifically we take

$$
\begin{equation*}
\gamma_{l, \text { max }, \text { corr }}=1+\frac{L / l-1}{L / l} \gamma_{l, \text { max,ind }} \tag{B.8}
\end{equation*}
$$

The distribution of $\gamma_{l, \max , \text { corr }}$ is given by

$$
F_{\gamma_{l, \text { mauc,carr }}}(\gamma) \approx \begin{cases}{\left[F_{\gamma_{l}}\left(\frac{(\gamma-1) L}{L-l}\right)\right]^{L / l}} & , \gamma \geq 1 \text { and } l<L  \tag{B.9}\\ F_{\gamma_{l}}(\gamma) & , \gamma \geq 0 \text { and } l=L\end{cases}
$$

For $\gamma \gg 1$ and $L / l$ large (say $>10$ ) the models in equations (B.5) and (B.9) produce similar results.
The dashed lines in Figure B. 5 show the probability density function of $\gamma_{l, m a x, c o r r}$ calculated from equation (B.9) for the case of $l=192 \mathrm{~km}$, whereas the solid black lines show the distribution of $\gamma_{l, \max }$ using the maxima approach with a beta distribution model described in Section 3.3.2. Figures B. 6 -B. 8 make similar comparisons for $l=96,48$ and 6 km . One sees that the theoretical model of equation (B.9) reproduces well the empirical PDF of $\gamma_{l, \max }$ for different spatial scales $l$ and large-scale average intensities $I_{L}$ and produces results close to those obtained using the maxima approach in Section 3.3.2.

Figure B. 9 compares theoretically calculated exceedance probabilities of rainfall maxima at different spatial scales $l$ to observed relative frequencies. It was generated using the procedure in Section 3.4 with $\gamma_{l, \max }$ calculated from equations (B.9) and (B.4) and the parametric expressions in equation (B.2) and Figure B.2. The fact that all histograms are nearly uniform validates the model in the sense that observed relative frequencies are close to theoretically calculated probabilities.


Figure B.6: Comparison of the empirical distribution of $\gamma_{l, \max }$ obtained for the case of continuous sampling using a sliding window of size $l=96 \mathrm{~km}$, and those calculated theoretically using the linear correction in equation (B.9) (dashed lines), using the maxima approach described in Section 3.3.2 [solid lines; equation (3.8)] and using the marginal approach with a modified lognormal (LN) distribution described in Section B. 2 (dashed-dotted lines). Comparison is done for the intensity categories in Figure B.1.


Figure B.7: Same as Figure B. 6 but for $l=48 \mathrm{~km}$


Figure B.8: Same as Figure B. 6 but for $l=6 \mathrm{~km}$.


Figure B.9: Histogram of the non-exceedance probability $P$ in equation (3.10) for different spatial scales $l . \gamma_{l \text { max }}$ is calculated from equations (B.9) and (B.4) and the parametric expressions in equation (B.2) and Figure B.2. Each histogram is based on a sample of size 789 .

## B. 2 A Simpler Marginal Approach Using a Modified-lognormal Distribution Model

Figure B. 1 shows that, for all categories of the large-scale intensity $I_{L}$, the moments of $\gamma_{l}$ vary with scale $l$ in an approximately log-linear way. This observation allows one to simplify the marginal model developed in Section B. 1 using concepts from the theory of multifractal cascades. To do so, we employ standard probability techniques with a minimum of multifractal formalism. This makes the approach more transparent than using multifractal concepts like dressed and bare densities, singularities and their fractal dimensions; although the latter tools are more elegant and general. For more details on multifractal measures and models, the reader is referred to Schertzer and Lovejoy (1987), Gupta and Waymire (1993), Lovejoy and Schertzer (1995), Veneziano and Langousis (2005a,b, 2009), Langousis and Veneziano (2007) and Langousis et al. (2007).

Define random variable $S_{l}$ such that

$$
\begin{equation*}
E\left[\left(S_{l}\right)^{q}\right]=\left(\frac{L}{l}\right)^{K(q)} \tag{B.10}
\end{equation*}
$$

Up to a maximum scale $L$ (in our case $L=384 \mathrm{~km}$ ), the moments of $S_{l}$ vary with $l$ in a log-linear way with slope $K(q)$ that depends on the moment order $q$.

In what follows we model $S_{l}$ using a modified lognormal (LN) distribution with unit mean. In particular, $S_{l}$ has a probability mass $P_{0}$ at $0: P\left[S_{l}=0\right]=P_{0}=1-(L / l)^{-C_{b}}$ and $\left(S_{l} \mid S_{l}>0\right)$ has lognormal distribution with log-mean $m_{L N}=\left(C_{b}-C_{L N}\right) \ln (L / l)$ and log-variance $\sigma_{L N}{ }^{2}=2 C_{L N} \ln (L / l)$ where $C_{b}$ and $C_{L N}$ are parameters. $C_{b}$ controls the lacunarity of the rainfall field (i.e. duration of wet and dry periods), whereas $C_{L N}$ controls the amplitude of the fluctuations when it rains. In this case the log-log slope of the moments, $K(q)$, is given by

$$
\begin{equation*}
K(q)=C_{b}(q-1)+C_{L N}\left(q^{2}-q\right) \tag{B.11}
\end{equation*}
$$

We call this distribution of $S_{l}$ a modified lognormal (LN) distribution with cumulative distribution function (CDF)

$$
F_{S_{l}}(s)= \begin{cases}P_{0}+\left(1-P_{0}\right) \Phi\left(\frac{\ln s-m_{L N}}{\sigma_{L N}}\right) & , s>0  \tag{B.12}\\ P_{0} & , s=0\end{cases}
$$

where $\Phi$ is the standard normal CDF. The distribution in equation (B.12) has been extensively used in the past for developing scaling representations of rainfall using the notion of multiplicative cascades; see e.g. Over and Gupta (1996); Schmitt et al. (1998); Langousis and Veneziano (2007), Langousis et al. (2007) and Veneziano et al. (2007).

An important difference between the distribution of $S_{l}$ in equation (B.12) and that of $\gamma_{l}$ in equation (B.4) is that the former does not have an upper bound, whereas the latter is upper bounded at $L / l$. In the multifractal literature this difference is referred to as the effect of "dressing" (see e.g. Schertzer and Lovejoy, 1987; Veneziano and Furcolo, 2003; Langousis and Veneziano, 2007), which is an alternative way of saying that random variable $\gamma_{l}$ accounts for rainfall variability at scales smaller than $l$ whereas $S_{l}$ does not. Since random variable $S_{l}$ does not have an upper bound, one expects equation (B.12) to produce better fits to the empirical ratios $\gamma_{l}$ as the resolution $L / l$ increases; see also discussion on Figures B.5-B. 8 below.

We allow the parameters $C_{b}$ and $C_{L N}$ of the modified lognormal distribution in equation (B.12) to depend on the large-scale rainfall intensity $I_{L}$. In particular, for each intensity category
$I_{L}$ in Figure B.1, we use the method of least squares to fit equation (B.10) to the $2^{\text {nd }}$ and $3^{\text {rd }}$ initial moments of $y_{l}$. The obtained slopes $K(2)$ and $K(3)$ and equation (B.11) are then used to calculate the parameters $C_{b}$ and $C_{L N}$ :

$$
\begin{equation*}
C_{b}=3 K(2)-K(3), C_{L N}=\frac{K(3)-2 K(2)}{2} \tag{B.13}
\end{equation*}
$$

We select to fit the $2^{\text {nd }}$ and $3^{\text {rd }}$ empirical moments for two reasons. The first reason is that a number of studies (Veneziano et al., 2006b, 2007; Langousis and Veneziano, 2007; Langousis et al., 2007) have shown that when modeling extremes of point rainfall, it is important to focus on moment orders larger than 2 . The other reason is that the empirical moments of order larger than 3.5-4 tend to underestimate the true moments due to sample limitations (Ossiander and Waymire, 2000, 2002; Veneziano et al., 2006b). Note, however, that similar findings on the variation of $C_{b}$ and $C_{L N}$ with $I_{L}$ (see below) where obtained also when fitting the $0^{\text {th }}$ and $2^{\text {nd }}$ moments of the empirical data.

Figure B. 10 shows how the parameters of the modified lognormal model vary with the largescale rainfall intensity $I_{L}$. The solid lines are smooth least-squares estimates. $C_{b}$ decreases exponentially with $I_{L}$. This is in accordance with the observed increase of the fraction of dry regions with increasing distance from the center of the storm; see Section B.1. $C_{L N}$ is approximately constant equal to 0.08 . Values of $C_{L N}$ around 0.1 have also been reported by studies on extra-tropical rainfall; see e.g. Langousis and Veneziano (2007) and Veneziano et al. (2007).


Figure B.10: Plots of the parameters $C_{b}$ and $C_{L N}$ against $I_{L}$. The solid lines are least-squares fits.

## B.2.1 Distribution of $\gamma_{l, \max }$

From equation (B.12) one can derive the CDF of $\gamma_{l, m a x, i n d}$, the maximum of $L / l$ independent copies of $S_{l}$. One obtains

$$
F_{\gamma_{l / \text { maxi.ind }}}(\gamma)= \begin{cases}{\left[P_{0}+\left(1-P_{0}\right) \Phi\left(\frac{\ln \gamma-m_{L N}}{\sigma_{L N}}\right)\right]^{L l l}} & , \gamma>0  \tag{B.14}\\ P_{0}^{L l l} & , \gamma=0\end{cases}
$$

where $m_{L N}=\left(C_{b}-C_{L N}\right) \ln (L / l)$ and $\sigma_{L N}{ }^{2}=2 C_{L N} \ln (L / l)$ can be obtained from Figure B. 10 as functions of $I_{L}$.

As discussed in the previous section, $\gamma_{l, \text { max,ind }}$ is biased relative to $\gamma_{l, \max }$ by the assumption that maximum rainfall occurs in one of the non-overlapping intervals of size $l$. In the case considered here, the maxima distribution in equation (B.14) is obtained from a modified lognormal marginal distribution, for which the classical way to eliminate this source of bias is to multiply $\gamma_{l, \text { max }, \text { ind }}$ by a continuity correction factor $\delta \approx 1.13-1.15$; see Hershfield (1961), Weiss (1964), Young and McEnroe (2003) and Veneziano et al. (2007). After corrected for continuity, equation (B.14) becomes

$$
F_{\gamma_{l, \text { max, corr }}}(\gamma)= \begin{cases}{\left[P_{0}+\left(1-P_{0}\right) \Phi\left(\frac{\ln (\gamma / \delta)-m_{L N}}{\sigma_{L N}}\right)\right]^{L l}} & , \gamma>0  \tag{B.15}\\ P_{0}{ }^{L l l} & , \gamma=0\end{cases}
$$

The dashed dotted lines in Figures B.5-B.8 show the probability density function of $\gamma_{l, \text { max }, \text { corr }}$ calculated from equation (B.15) for $l=192,96,48$ and 6 km . One sees that the probability model in equation (B.15) produces results close to those obtained from equation (B.9) only for spatial scales $l \leq 48 \mathrm{~km}$ or, equivalently, resolutions $L / l \geq 4$. This makes sense since the variable $\gamma_{l, \text { max,corr }}$ in equation (B.15) does not include dependencies between rainfall amounts inside different $l$ intervals and, in addition, it does not account for rainfall variability at scales smaller than $l$. The first source of bias causes the distribution in equation (B.15) to have a lower bound at zero, whereas the second source of bias results in a distribution that does not have an upper bound. Note that the distribution model in equation (B.9) accounts for these two effects and, hence, it is bounded in the range $[1, L / l]$. For high resolutions $L / l$ (say $L / l>4$ ) the models in equations (B.9) and (B.15) should produce similar results; see Figures B. 7 and B.8.

Similar to Figure B.9, Figure B. 11 compares theoretically calculated exceedance probabilities of rainfall maxima to observed relative frequencies. Here $\gamma_{l, \max }$ is calculated from equation (B.15) and the parametric expressions in Figure B.10. One sees that for spatial scales $l \geq 24 \mathrm{~km}$ the modified log-normal model for $\gamma_{l, \max }$ performs slightly worse relative to the modified beta model in Section B. 1 (see Figure B.9) and the maxima approach with a beta distribution in Section 3.3.2; see Figure 3.11. This reduced performance of the model is due to the two sources of bias discussed above.


Figure B.11: Histogram of the non-exceedance probability $P$ in equation (3.10) for different spatial scales $l . \gamma_{l m a x}$ is calculated from equation (B.15) and the parametric expressions in Figure B.10. Each histogram is based on a sample of size 789.

## B. 3 Results from Different Approaches

Similar to Figure 3.12, Figure B. 12 shows how the distribution of the total amplification factor $\beta_{l, \max }=I_{l, \max } / I_{L, M S R}$ depends on $l, V_{\max }$, and $y^{\prime}=\mid y / R_{\max } l$. Results are shown for the three approaches developed in Sections 3.3.2, B.1 and B. 2 to evaluate the distribution of the amplification factor $\gamma_{l, \max }$. One sees that the maxima approach using a beta distribution (MB, dashed dotted lines), the marginal approach using a modified beta distribution (MMB, solid lines) and the marginal
approach using a modified lognormal distribution (MML, dashed lines) produce similar results. Small differences can be noted, but this is due to the fact that the approaches are different and not derived one from another. However, these differences are small and negligible for practical applications; see Figure B. 13 below.


Figure B.12: Comparison of the theoretical probability density functions (PDFs) of the total amplification factor $\beta_{l, \max }=I_{l, \max } / I_{L, M S R}$ using the maxima approach with a beta distribution [MB (Section 3.3.2), dashed dotted lines], the marginal approach with a modified beta distribution [MMB (Section B.1), solid lines] and the marginal approach with a modified lognormal distribution [MML (Section B.2), dashed lines]. Results are shown for two different length scales $l=10,50 \mathrm{~km}$, different storm intensities $V_{\max }=30,45$ and $60 \mathrm{~m} / \mathrm{s}$ and different normalized distances $y^{\prime}=$ $\left|y / R_{\max }\right|$ from the center of the storm: $y^{\prime}=1$ (thin lines) and $y^{\prime}=5$ (thick lines).

Figure B. 13 shows calculated IDF curves for New Orleans as plots of rainfall intensity $i$ against return period $T$ for different durations $D$. Results are shown for the three different approaches to evaluate the distribution of the amplification factor $\gamma_{l, \max }$. All other model components (i.e. recurrence model and distribution of $\beta_{L}$ ) are the same as those used in Section 3.5. One sees that the IDF curves produced by the three approaches are similar.


Figure B.13: Theoretical IDF curves for the region New Orleans obtained from equation (3.11) for $\lambda=0.57$ events/yr using the maxima approach with a beta distribution [MB (Section 3.3.2), dashed dotted lines], the marginal approach with a modified beta distribution [MMB (Section B.1), solid lines] and the marginal approach with a modified lognormal distribution [MML (Section B.2), dashed lines]. Results are shown for averaging durations $D=$ $0.5,1,6,12$ and 24 hours.

