Fifty Years of Twin Studies: A Meta-Analysis of the Heritability of Human Traits

Tinca JC Polderman^{1†}, Beben Benyamin^{2†}, Christiaan A. de Leeuw^{1,3}, Patrick F Sullivan^{4,5,6}, Arjen van Bochoven⁷, Peter M Visscher^{2,8‡}, Danielle Posthuma^{1,9‡*}

- ¹ Department of Complex Trait Genetics, VU University, Center for Neurogenomics and Cognitive Research, Amsterdam, The Netherlands
- ² The University of Queensland, Queensland Brain Institute, Brisbane, Queensland 4072, Australia
- ³ Radboud University Nijmegen, Institute for Computing and Information Sciences, Nijmegen, The Netherlands
- ⁴ Center for Psychiatric Genomics, Department of Genetics, University of North Carolina, Chapel Hill, NC, USA
- ⁵ Department of Psychiatry, University of North Carolina, Chapel Hill, NC, USA
- ⁶ Department of Medical Epidemiology and Biostatistics, Karolinska Institutet, Stockholm, Sweden
- ⁷ Faculty of Sciences, VU University, Amsterdam, the Netherlands
- ⁸ The University of Queensland Diamantina Institute, The Translational Research Institute, Brisbane, Queensland, Australia
- ⁹ Department of Clinical Genetics, VU University Medical Center, Neuroscience Campus Amsterdam, The Netherlands

[†] These authors contributed equally to this work

[‡] These authors contributed equally to this work

*Correspondence should be addressed to: Danielle Posthuma: Department of Complex Trait Genetics, VU University, De Boelelaan 1085, 1081 HV, Amsterdam, The Netherlands. Phone: +31 20 598 2823, Fax: +31 20 5986926, <u>d.posthuma@vu.nl</u>

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1 Abstract: Despite a century of research in complex traits in humans, the relative importance and 2 specific nature of the influences of genes and environment on human traits remains controversial. 3 We report a meta-analysis of twin correlations and reported variance components for 17,804 4 traits from 2,748 publications including 14,558,903 partly dependent twin pairs, virtually all 5 published twin studies of complex traits. Estimates of heritability cluster strongly within 6 functional domains and across all traits the reported heritability is 49%. For a majority (69%) of 7 traits, the observed twin correlations are consistent with a simple and parsimonious model where 8 twin resemblance is solely due to additive genetic variation. The data are inconsistent with a 9 substantial influence of shared environment or non-additive genetic variation. This study 10 provides the most comprehensive analysis of the causes of individual differences in human traits 11 to date, and will guide future gene-mapping efforts. All results can be visualized at 12 http://match.ctglab.nl.

13 Insight into the nature of observed variation in human traits is important in medicine, 14 psychology, social sciences and evolutionary biology. It has gained new relevance with both the ability to map genes for human traits and the availability of large, collaborative datasets to do so 15 16 on an extensive and comprehensive scale. Individual differences in human traits have been 17 studied for more than a century, yet the causes of human trait variation remain uncertain and 18 controversial. Specifically, the partitioning of observed variability into underlying genetic and 19 environmental sources and the relative importance of additive and non-additive genetic variation are continually debated¹⁻⁵. Recent results from large-scale genome-wide association studies 20 21 (GWAS) show that many genetic variants contribute to complex trait variation and that effect sizes are typically small^{6,7}. However, the sum of the variance explained by the detected variants 22 is much smaller than the reported heritability of the trait^{4,6-10}. This 'missing heritability' has led 23 some investigators to conclude that non-additive variation must be important^{4,11}. Although the 24 presence of gene-gene interaction has been demonstrated empirically^{5,12-17}, little is known about 25 its relative contribution to observed variation 18 . 26

In this study, our aim is two-fold. First, we analyze empirical estimates of the relative 27 28 contributions of genes and environment for virtually all human traits investigated in the past 50 29 years. Second, we assess empirical evidence for the presence and relative importance of nonadditive genetic influences on all human traits. We rely on classical twin studies, as the twin 30 design has been used widely to disentangle the relative contributions of genes and environment, 31 32 across a variety of human traits. The classical twin design is based on contrasting trait 33 resemblance of monozygotic (MZ) and dizygotic (DZ) twin pairs. MZ twins are genetically identical and DZ twins are genetically full siblings. We show that for a majority (69%) of traits, 34 35 the observed statistics are consistent with a simple and parsimonious model where observed 36 variation is solely due to additive genetic variation. The data are inconsistent with a substantial 37 influence of shared environment or non-additive genetic variation. We also show that estimates of heritability cluster strongly within functional domains and across all traits the reported 38 39 heritability is 49%. Our results are based on a meta-analysis of twin correlations and reported 40 variance components for 17,804 traits from 2,748 publications including 14,558,903 partly 41 dependent twin pairs, virtually all published twin studies of complex traits between 1958 and 42 2012. This study provides the most comprehensive analysis of the causes of individual differences in human traits to date, and will guide future gene-mapping efforts. All results can be 43 44 visualized at http://match.ctglab.nl.

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46 **RESULTS**

47 The distribution of studied traits in twin studies is non-random across traits and countries

48 We systematically retrieved published classical twin studies in which observed variation of 49 human traits was partitioned into genetic and environmental influences. For each study we 50 collected reported twin correlations for continuous traits and contingency tables for dichotomous 51 traits, estimates from genetic model fitting, and study characteristics (sample size, population, 52 age cohort and ascertainment scheme) (Supplementary Table S1). Investigated traits were manually classified using the chapter and subchapter levels of the International Classification of 53 Functioning, Disability and Health (ICF) or the International Statistical Classification of 54 55 Diseases and Related Health Problems (ICD-10) (Online Methods). The ICD10 and ICF subchapter levels refer to actual diseases (e.g. 'Atopic Dermatitis') or traits (e.g. 'Temperament 56 and Personality Functions'). We identified 2,748 relevant twin studies, published between 1958 57 and 2012²⁰. Half of these were published after 2004, with sample sizes per study in 2012 around 58

59 1000 pairs (Supplementary Table S2). Each study could report on multiple traits measured in 60 one or several samples. These 2,748 studies reported on 17,804 traits. Twin subjects came from 61 39 different countries with a large proportion of studies (34%) based on US twin samples. The 62 continents of South-America (0.5%), Africa (0.2%) and Asia (5%) were heavily 63 underrepresented (Fig. 1a,b, Supplementary Table S3). The total number of studied twins was 64 14,558,903 partly dependent pairs, or 2,247,128 when correcting for reporting on multiple traits per study. The majority of studies (59%) were based on the adult (aged 18-64) population, 65 although the sample sizes available for the elderly (aged 65+) population studies were the largest 66 67 (Supplementary Table S4). Authorship network analyses revealed that 61 communities of 68 authors wrote the 2,748 published studies. The 11 largest authorship communities contained >65 69 authors and could be mapped back to the main international twin registries, such as the Vietnam 70 Era Twin Registry, the Finnish Twin Cohort and the Swedish Twin Registry (Supplementary 71 Figure S12).

72 The investigated traits fell into 28 general trait domains. The distribution of traits evaluated in 73 twin studies was highly skewed with 62% focusing on traits classified under the *Psychiatric*, 74 Metabolic and Cognitive domains whereas traits classified under the Developmental, Connective 75 *Tissue*, and *Infection* domains together accounted for less than 1% of all investigated traits (**Fig.** 76 1c, Supplementary Tables S5-S7). The ten most investigated traits were *Temperament and* Personality Functions, Weight Maintenance Functions, General Metabolic Functions, 77 78 Depressive Episode, Higher-Level Cognitive Functions, Conduct Disorders, Mental and 79 Behavioural Disorders Due to Use of Alcohol, Anxiety Disorders, Height, and Mental and 80 Behavioural Disorders Due to Use of Tobacco. Collectively these traits accounted for 59% of all 81 investigated traits.

83 Across all traits, the reported relative contribution of genes and environment is equal

We did not find evidence of systematic publication bias as a function of sample size (e.g. where 84 85 studies based on relatively small samples were only published when larger effects were reported) (Fig. 1d, Supplementary Figure S1, Supplementary Tables S8-S11). We calculated weighted 86 87 averages of MZ and DZ twin correlations and of the reported estimates of the relative contribution of genetic and environmental influences to investigated traits using a random effects 88 89 meta-analytic model to allow for heterogeneity across different studies. The meta-analyses of all traits yielded an r_{MZ} of 0.64 and an r_{DZ} of 0.34. The reported heritability (h²) across all traits was 90 0.49, and the reported estimate of shared environmental effects (c^2) was 0.17 (**Table 1, Fig. 2a,b**, 91 Supplementary Figure S6). 92

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94 Estimates of variance components cluster in functional domains

Heritability estimates were found to cluster in functional domains, with the largest heritability 95 96 estimates for traits classified under the Ophthalmological domain, followed by the Ear-Nose-97 Throat, Dermatological and Skeletal domains. The lowest heritability estimates were for traits in the Environment, Reproduction and Social Values domains (Fig. 2d, Supplementary Table 98 **S17**). All weighted averages of h^2 across >500 distinct traits had a mean greater than zero 99 100 (Supplementary Tables S18-S20, S31, S32). The lowest reported heritability for a specific trait was for *Gene Expression* with an estimate of $h^2=0.05$ (SE=0.03), and a c² of 0.74 (SE=0.03), but 101 note these estimates are based on only 20 entries. The largest influence of c^2 was found for traits 102 in the Cell domain (0.67, SE=0.05), followed by traits in the Infection (0.35, SE=0.05), 103

Hematological (0.32, SE=0.09), Endocrine (0.32, SE=0.05), Reproduction (0.32, SE=0.06),
Social Values (0.27, SE=0.03), Environment (0.25, SE=0.02), and Skeletal (0.25, SE=0.02)
domains (Fig. 2d, Supplementary Table S17).

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108 Heterogeneity of twin correlations across sex and age

Across all traits, the weighted averages of twin correlations and reported h^2 and c^2 did not show evidence of heterogeneity across sex, although there was some evidence for a lower correlation in opposite sex twin pairs versus same sex dizygotic pairs (**Table 1**, **Supplementary Information, section 5.1**). The data showed a decrease in MZ and DZ resemblance after adolescence and an accompanying decrease in estimates of both h^2 and c^2 (**Fig. 2c, Supplementary Table S16**).

115 In the top 20 most investigated traits for twin correlations the weighted estimates did not show consistent evidence for heterogeneity across sex, with r_{MZM} and r_{MZF} as well as r_{DZM} and r_{DZF} 116 117 remarkably similar across the majority of the top 20 investigated specific traits (Fig. 3), although 118 for several traits the opposite sex twin correlations were lower than the same sex estimates, 119 mostly after age 11 (e.g. for Weight Maintenance Functions, Functions of Brain and Mental, 120 Behavioral Disorders due to the Use of Alcohol, and Mental and Behavioral Disorders due to the 121 Use of Tobacco). Heterogeneity of weighted twin correlations across age was more prominent 122 than heterogeneity across sex (Fig. 3). For example when considering the r_{MZ} we note that for 123 most of the top 20 investigated traits the estimate tends to decrease with age, especially after 124 adolescence, which was generally mirrored in the r_{DZ} (Fig. 3).

126 Per study, variance components model fitting and selection leads to an underestimation of the 127 h^2

Falconer's equations can be used to calculate \hat{h}^2 and \hat{c}^2 based on twin correlations¹⁸. The 128 equation for \hat{h}^2 is $2 \times (r_{MZ} - r_{DZ})$, and for \hat{c}^2 it is $2 \times r_{DZ} - r_{MZ}$. When these are applied to the weighted 129 averages of r_{MZ} and r_{DZ} we find an \hat{h}^2 estimate of 2×(0.64-0.34)=0.60 and a \hat{c}^2 estimate of 130 $2 \times 0.34 + 0.64 = 0.04$ (Table 1, Supplementary Figure S7). We note that the estimated \hat{h}^2 based on 131 twin correlations is larger than the weighted average of reported h^2 . As a consequence, the 132 estimated \hat{c}^2 based on twin correlations is lower than the weighted average of the reported c^2 133 134 component. To test whether this discrepancy was due to a bias in studies reporting only twin correlations or only variance components, we conducted the meta-analysis only on studies 135 136 reporting both. This yielded similar estimates with a similar discrepancy (Supplementary Table 137 **S13**), ruling out the explanation that twin correlations may have been reported on traits that 138 happened to be more heritable than traits for which the estimates of variance components were 139 reported. Through theory we show that such a discrepancy can arise when the individual studies represent a mix of traits that follow a pattern of $r_{MZ} > 2r_{DZ}$ and $r_{MZ} < 2r_{DZ}$, and where the choice 140 141 of fitting a model that includes shared environment or non-additive genetic influences is based 142 on the observed pattern of twin correlations (see Supplementary Information, section 5.2). More specifically, because c^2 and non-additive genetic influences cannot be estimated 143 144 simultaneously from twin correlations an 'ACE' model (Additive genetic, Common environmental and Error or non-shared environmental influences) is fitted to the data if $2r_{DZ}$ – 145 $r_{MZ} > 0$. On the other hand, if $2r_{DZ} - r_{MZ} < 0$, an 'ADE' model, including non-additive genetic 146 147 instead of common environmental influences, is selected. This leads to sampling bias in the estimate of h² from the full model. We show (Supplementary Table S14) that such per-study 148

choices cause bias and can lead to a 10% downward bias of reported estimates of h² compared to those based on twin correlations, consistent with the observed discrepancy between our metaanalysis of variance component estimates calculated from twin correlations and the reported variance components.

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The overall pattern of twin correlations is consistent with a model where all genetic variance is additive and all environmental variance is non-shared

156 There may be many causes of similarities and differences within MZ and DZ pairs and typically 157 these are interpreted in terms of (non-) additive genetic and (non-) shared environmental influences¹⁹. Yet, there are essentially only two estimable and testable variance components of 158 159 interest in the twin design. Therefore, inference from classical twin studies on all underlying 160 unobserved sources of variation that lead to the resemblance between relatives is limited. 161 However, there are two simple and parsimonious hypotheses that can be tested across traits from 162 estimated correlation coefficients of MZ pairs (r_{MZ}) and DZ pairs (r_{DZ}). The first is that the MZ 163 and DZ population correlations (ρ_{MZ} and ρ_{DZ}) are the same, consistent with twin resemblance 164 being solely due to non-genetic factors. The second is a two-fold ratio of ρ_{MZ} and ρ_{DZ} , consistent 165 with twin resemblance being solely due to additive genetic variation. An across-trait consistency 166 with either of these hypotheses is not a proof of these simple models but would provide an 167 extremely parsimonious model against which other experimental designs (e.g. DNA-sequence 168 based) should be tested. For the vast majority (84%) of traits, we find that MZ correlations are 169 larger than DZ correlations. Using the weighted estimates of r_{MZ} and r_{DZ} across all traits, we 170 show that the average $2r_{DZ}-r_{MZ} = .04$ (SE=.01) (**Table 1**), which is very close to a two-fold 171 difference in the MZ to DZ correlation. The proportion of single studies in which the pattern of twin correlations was consistent with the null hypothesis that $2r_{DZ}=r_{MZ}$ was 69%. This observed pattern of twin correlations is consistent with a simple and parsimonious underlying model of the absence of environmental effects shared by twin pairs and the presence of genetic effects that are entirely due to additive genetic variation (**Table 2**). This remarkable fitting of the data with a simple mode of family resemblance is inconsistent with the hypothesis that a substantial part of human trait variation is due to shared environmental variation or to substantial non-additive genetic variation.

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180 For most specific traits the empirical data is consistent with a model where all genetic 181 variance is additive

182 Although across all traits 69% of studies show a pattern of MZ and DZ twin correlations 183 consistent with an r_{MZ} that is exactly twice the r_{DZ} , this finding is not necessarily representative 184 of the majority of studies in functional domains or for every specific trait (i.e., at ICD10/ICF 185 subchapter level). We thus calculated the proportion of studies consistent with $2r_{DZ}=r_{MZ}$ within functional domains and for each specific trait, and found that traits consistent with this 186 187 hypothesis tend to cluster in specific functional domains. A pattern of twin correlations 188 consistent with 2r_{DZ}=r_{MZ} was most prominent for traits included in the Neurological, Ear-Nose-189 Throat, Cardiovascular and Ophthalmological domains, with 99.5%, 97%, 95% and 87% of 190 studies being consistent with a model where all resemblance was entirely due to additive genetic 191 variance. In only three of 28 general trait domains most studies were inconsistent with this model. These domains were Activities (35%), Reproduction (44%), and Dermatological (45%) 192 193 (Table 2 and Supplementary Table S24). Of 59 specific traits (i.e. ICD10/ICF subchapter 194 classification) for which we had sufficient information to calculate the proportion of studies

195 consistent with $2r_{DZ}=r_{MZ}$, 21 traits showed a proportion less than 0.50, whereas for the remaining 196 38 traits the majority of individual studies were consistent with $2r_{DZ}=r_{MZ}$ (Supplementary Table 197 **S26**). Of the top 20 most investigated specific traits, we found that for 12 traits the majority of 198 individual studies were consistent with a model where variance was solely due to additive 199 genetic variance and non-shared environmental variance, while the pattern of MZ and DZ twin 200 correlations was inconsistent with this model for eight traits, suggesting that apart from additive 201 genetic influences and non-shared environmental influences, either or both non-additive genetic influences and shared environmental influences are needed to explain the observed pattern of 202 203 twin correlations (Table 2 and Supplementary Table S26). These eight traits were Conduct 204 disorders, Height, Higher Level Cognitive Functions, Hyperkinetic disorder, Mental and 205 Behavioral Disorders due to the use of Alcohol, Mental and Behavioral Disorders due to the use 206 of Tobacco, Other Anxiety Disorders, and Weight Maintenance Functions. For all eight traits, 207 meta-analyses on reported variance components resulted in a weighted estimate of reported 208 shared environmental influences that was statistically different from zero (Supplementary 209 Table S32). Comparison of weighted twin correlations for these specific traits resulted in 210 positive estimates of 2r_{DZ}-r_{MZ}, except for Hyperkinetic Disorders where 2r_{DZ}-r_{MZ} was -0.13 211 (SE=0.03), based on 144 individual reports and 207,589 twin pairs), which suggests the 212 influence of non-additive genetic variation for this trait, or any other source of variation that 213 leads to a disproportionate similarity among monozygotic twin pairs.

214

215 **DISCUSSION**

We have conducted a meta-analysis of virtually all twin studies published in the past 50 years, ona wide range of traits and reporting on more than 14 million twin pairs across 39 different

218 countries. Our results provide compelling evidence that all human traits are heritable: not one 219 trait had a weighted heritability estimate of zero. The relative influence of genes and 220 environment is not randomly distributed across all traits but clusters in functional domains. In 221 general, we showed that reported estimates of variance components from model fitting can 222 underestimate the true trait heritability, when compared with the heritability based on the twin 223 correlations. Roughly two-thirds of traits show a pattern of MZ and DZ twin correlations that is 224 consistent with a simple model whereby trait resemblance is solely due to additive genetic 225 variation. This implies that for the majority of complex traits causal genetic variants can be 226 detected using a simple additive genetic model.

227 Approximately one-third of traits did not follow the simple pattern of a two-fold ratio in MZ and 228 DZ correlations. For these traits, a simple additive genetic model does not sufficiently describe 229 the population variance. An incorrect assumption about narrow sense heritability (the proportion 230 of total phenotypic variation due to additive genetic variation) can lead to a mismatch between 231 results from gene finding studies and prior expectations. If the pattern of twin correlations is 232 consistent with a substantial contribution from shared environmental factors, like we find for 233 Conduct Disorders, Religion and Spirituality, and Education, then gene-mapping studies may 234 yield disappointing results. If the cause of a departure from a simple additive genetic model is 235 the existence of non-additive genetic variation, as is for example suggested by the average twin 236 correlations for Recurrent Depressive Disorder, Hyperkinetic Disorders, and Atopic Dermatitis, 237 then it may be tempting to fit non-additive models in gene mapping studies (e.g. GWAS or 238 sequencing studies). However, the statistical power of such scans is extremely low due to the 239 many non-additive models that can be fitted (e.g. within-locus dominance versus between locus 240 additive by additive effects) and the penalty incurred by multiple testing. Our current results

241 signal traits for which an additive model cannot be assumed. In most of these traits DZ twin 242 correlations are higher than half the MZ correlations, suggesting that shared environmental 243 effects are causing the deviation from a simple additive genetic model. Yet, data from twin pairs 244 only does not provide sufficient information to resolve the actual causes of the deviation from a 245 simple additive genetic model. More detailed studies may reveal the likely causes of such 246 deviation, and may as such uncover epidemiological or biological factors that drive family 247 resemblance. To make stronger inferences about the causes underlying the resemblance between 248 relatives for traits that deviate from the additive genetic model, additional data are required, for 249 example from large population samples with extensive phenotypic and DNA sequence 250 information, detailed measures of environmental exposures, and larger pedigrees including non-251 twin relationships.

We note that our inference is based on twin studies published between 1950 and 2012 and 252 253 generally applies to complex traits and not necessarily generalizes to Mendelian subtypes of 254 traits. Most Mendelian traits are rare in the population and are therefore not studied by twin 255 researchers because they cannot ascertain enough affected twin pairs to estimate genetic 256 parameters reliably. In the rare case that they were available, the Mendelian subtypes were 257 analyzed together with the subtypes of the same trait that are due to common causes. In that case 258 our inference would be biased away from our main result because Mendelian diseases tend to be 259 dominant or recessive, not additive. In addition, there may be heterogeneity in measurement 260 errors between studies for the same trait and between traits. A test-retest correlation would 261 quantify measurement error when contrasted with a correlation between MZ twins but few twin 262 studies report such correlations in the same papers that estimate heritability.

Our results provide the most comprehensive empirical overview of the relative contributions of genes and environment to all human traits that have been studied in twins to date, which can guide and serve as a reference for future gene-mapping efforts.

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308	Supplementary Table S33
309	
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311	
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329 competing financial interests. Correspondence and requests for materials should be addressed to330 d.posthuma@vu.nl.

	Est.	SE	NTraits	Npairs
r _{MZ}	0.64	0.0024	9568	2,563,627
r _{MZM}	0.62	0.0037	4518	1,070,962
r _{MZF}	0.63	0.0035	4360	1,171,841
r _{DZ}	0.34	0.0031	5220	2,606,252
r _{DZSS}	0.35	0.0030	6108	1,752,952
r _{DZM}	0.32	0.0034	4412	1,039,238
r _{DZF}	0.34	0.0037	4255	1,068,562
r _{DOS}	0.30	0.0047	2342	898,610
h ²	0.49	0.0042	2929	4,341,721
h ² ss	0.47	0.0047	1795	1,187,837
h ² _M	0.47	0.0054	2095	1,732,622
h ² _F	0.47	0.0050	1957	1,539,582
c ²	0.17	0.0040	2771	4,272,318
c ² _{ss}	0.19	0.0055	1769	1,185,116
c ² _M	0.16	0.0042	1988	1,519,148
C ² _F	0.17	0.0047	1925	1,516,192
2(r _{MZ} -r _{DZ})	0.59	0.0078	9568	5,169,879
2(r _{MZ} -r _{DZ}) SS	0.58	0.0077	9568	4,316,578
2(r _{MZ} -r _{DZ}) M	0.59	0.0102	4518	2,110,200
2(r _{MZ} -r _{DZ}) F	0.57	0.0102	4360	2,240,403
$2r_{DZ}-r_{MZ}$	0.04	0.0066	9568	5,169,879
2r _{DZ} -r _{MZ} SS	0.05	0.0065	9568	4,316,578
$2r_{DZ}-r_{MZ}M$	0.02	0.0078	4518	2,110,200
$2r_{\text{DZ}}-r_{\text{MZ}}$ F	0.06	0.0082	4360	2,240,403

Table 1. Weighted means of twin correlations and variance components across all human
 traits investigated in a classical twin study and published between 1958-2012

334 r=correlation; MZ=monozygotic twins; DZ=dizygotic twins; M=males; F=females; SS=same sex

pairs only; DOS=DZ opposite sex; h^2 =heritability; c^2 =proportion of variance due to share

environmental variation; Est.=estimate based on random effects meta-analysis; SE=standard

337 error; NTraits=number of investigated traits; Npairs=number of dependent twin pairs. The pairs

are not independent as the same or an overlapping sample of twins may have been used for

339 multiple traits and across multiple studies.

Table 2. Proportion (π_0) of studies that are consistent with a model where trait resemblance is

solely due to additive genetic variation, and weighted means of twin correlations, for the main

342 trait domains and the top 20 investigated traits

	π_0			r _{MZ}			r _{DZ}			
	NTraits	Est.	Est.	SE	NTraits	Npairs	Est.	SE	NTraits	Npairs
All traits	5185	0.69	0.64	0.0024	9568	2,563,628	0.34	0.0031	5220	2,606,252
General trait domains							-			
Activities	62	0.35	0.57	0.0185	118	58,227	0.34	0.0217	63	55,864
Cardiovascular	267	0.95	0.56	0.0080	380	41,669	0.29	0.0096	268	25,544
Cell	54	0.59	0.72	0.0219	72	3,188	0.52	0.0428	54	1,667
Cognitive	450	0.57	0.65	0.0072	931	288,867	0.37	0.0103	454	304,720
Dermatological	74	0.45	0.73	0.0251	109	19,509	0.40	0.0175	75	23,245
Ear, Nose, Throat	165	0.97	0.76	0.0134	200	27,882	0.33	0.0155	172	14,222
Endocrine	108	0.69	0.56	0.0171	162	10,112	0.39	0.0218	110	9,140
Environment	145	0.50	0.55	0.0136	295	120,606	0.40	0.0167	145	99,137
Gastrointestinal	32	0.59	0.55	0.0243	64	10,982	0.27	0.0281	39	28,431
Hematological	19	0.65	0.76	0.0229	50	5,541	0.56	0.0316	19	3,218
Immunological	230	0.67	0.61	0.0116	280	18,051	0.36	0.0128	231	36,075
Metabolic	464	0.60	0.75	0.0049	912	210,189	0.41	0.0078	464	197,921
Neurological	702	1.00	0.68	0.0048	1751	129,076	0.29	0.0058	705	89,103
Nutritional	110	0.72	0.48	0.0157	205	75,751	0.22	0.0146	110	79,188
Ophthalmological	106	0.87	0.73	0.0165	199	26,139	0.39	0.0171	106	16,189
Psychiatric	1778	0.62	0.55	0.0044	2865	1,232,382	0.31	0.0050	1781	1,374,817
Reproduction	16	0.44	0.77	0.0336	34	12,130	0.33	0.0633	16	27,879
Respiratory	125	0.74	0.70	0.0183	184	34,443	0.33	0.0190	127	51,150
Skeletal	190	0.51	0.83	0.0077	395	111,282	0.50	0.0121	191	113,080
Social Interactions	24	0.63	0.34	0.0168	146	43,501	0.27	0.0411	24	22,764
Social Values	45	0.69	0.49	0.0297	120	52,492	0.41	0.0619	45	28,071
Top 20 investigated traits	for r _{MZ} a	nd r _{DZ}					_			
Blood Pressure Funct.	110	0.93	0.58	0.0095	179	20,621	0.31	0.0125	110	11,620
Conduct Dis.	216	0.41	0.66	0.0091	289	147,974	0.41	0.0096	216	192,651
Depressive Episode	115	0.60	0.45	0.0136	173	98,315	0.25	0.0153	115	121,936
Endocrine Gland Funct.	92	0.72	0.54	0.0170	139	8,533	0.38	0.0251	92	7,295
Food	110	0.72	0.48	0.0157	205	75,751	0.22	0.0146	110	79,188
Funct. of Brain	594	0.99	0.68	0.0062	1010	69,722	0.29	0.0064	594	58,951
General Metab. Funct.	219	0.69	0.68	0.0074	462	62,108	0.37	0.0101	219	58,338
Heart Funct.	140	1.00	0.53	0.0095	174	15,070	0.27	0.0111	140	11,109
Height	87	0.29	0.91	0.0045	128	53,076	0.54	0.0079	87	68,358
High-L. Cognitive Funct.	188	0.44	0.71	0.0087	419	152,197	0.44	0.0156	188	158,626
Hyperkinetic Dis.	100	0.37	0.65	0.0130	144	86,450	0.26	0.0159	100	121,139
Imm. System Funct.	223	0.67	0.61	0.0118	276	16,703	0.36	0.0131	223	32,964
Ment. Beh. Dis. Alc.	100	0.36	0.63	0.0152	158	94,477	0.41	0.0199	101	94,196
Ment. Beh. Dis. Tob.	70	0.47	0.72	0.0164	110	51,102	0.47	0.0220	72	34,186
Other Anxiety Dis.	145	0.29	0.55	0.0131	191	105,902	0.33	0.0157	145	153,730
Spec. Personal. Dis.	140	0.93	0.45	0.0092	162	41,460	0.23	0.0073	140	33,681
Structure of the Eyeball	86	0.91	0.73	0.0221	121	19,276	0.37	0.0190	86	13,580
Structure of Mouth	117	0.89	0.82	0.0103	127	7,769	0.40	0.0118	119	8,493
Temp. Pers. Funct.	568	0.84	0.47	0.0079	1134	334,190	0.23	0.0105	568	296,114
Weight Maint. Funct.	215	0.48	0.81	0.0051	391	141,152	0.44	0.0102	215	134,867

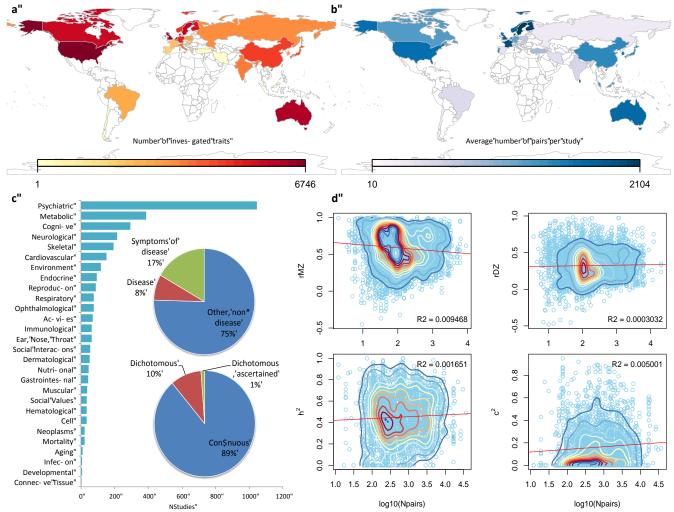
General trait domain categories with < 10 entries for π_0 were excluded. Alc. = Alcohol; Dis.=

344 Disorders; High-L. = Higher Level; Imm. = Immunological; Funct. = Functions; Maint.=

345 Maintenance; Metab.= Metabolic; Ment. Beh. = Mental and Behavioural; Spec. Personal.= Specific

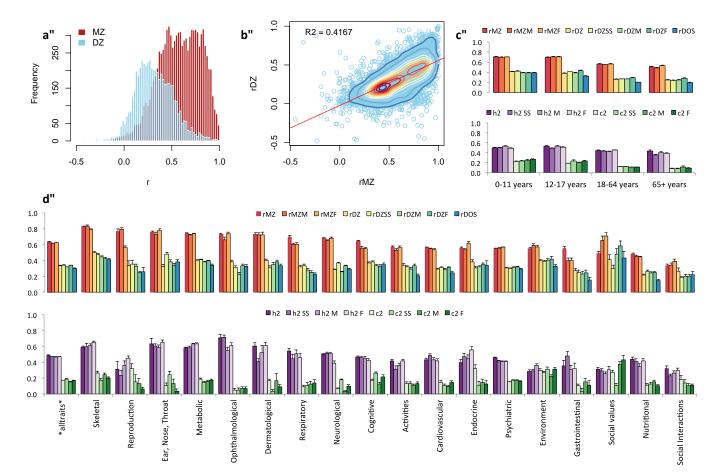
- Personality; Temp. Pers. =Temperament and Personality; Tob. = Tobacco; other abbreviations, see Table 1. The top 20 investigated traits are conditional on the reporting of r_{MZ} and r_{DZ} . 346
- 347

- 348 **Figures and legends**
- Fig. 1. Distribution of investigated traits in virtually all twin studies published between 349
- 350 1958 and 2012





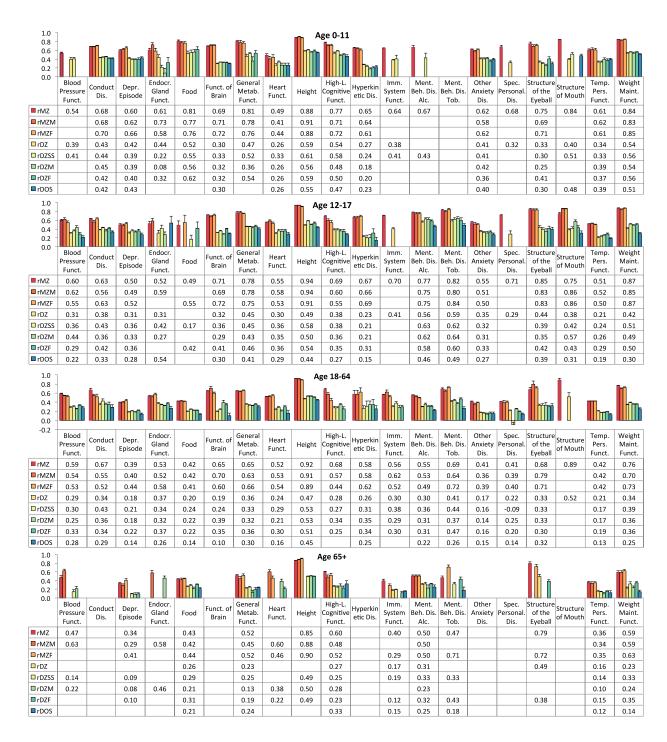
(a) The number of investigated traits in classical twins studies across all countries. (b) The 353 average number of included pairs per study across countries. (c) The number of investigated traits according to functional trait domain and trait characteristic (inset). (d) MZ and DZ twin 354 correlations and reported estimates of h^2 and c^2 as a function of sample size. Contour lines 355 indicate the density of the data in that region. The lines are 'heat' colored from blue to red, 356 357 indicating increasing data density.



360 Fig 2. Twin correlations and heritabilities for all human traits

361

(a) Distribution of r_{MZ} and r_{DZ} across traits investigated in 2,748 twin studies published between 362 363 1958 and 2012. The r_{MZ} is based on 9,568 traits and 2,563,628 partly dependent pairs; r_{DZ} is 364 based on 5,220 traits and 2,606,252 partly dependent pairs (see Table 1). (b) Relation between 365 the r_{MZ} and r_{DZ} , using all 5,185 traits for which both were reported. (c) Random effect meta-366 analytic estimates of twin correlations (top) and reported variance components (bottom) across all traits separate for four age cohorts. Error bars denote standard errors. (d) Random effect 367 368 meta-analytic estimates of twin correlations (top) and reported variance components (bottom) 369 across all traits and within functional domains for which data on all correlations and variance 370 components was available. Error bars denote standard errors. For data Tables see Supplementary 371 Information. 372



373 Fig. 3 Twin correlations for the top 20 most investigated specific traits across age and sex

374

375 Alc. = Alcohol; Dis.= Disorders; High-L. = Higher Level; Imm. = Immunological; Funct. =

376 Functions; Maint.= Maintenance; Metab.= Metabolic; Ment. Beh. = Mental and Behavioural;

377 Spec. Personal.= Specific Personality; Temp. Pers. =Temperament and Personality; Tob. =

- 378 Tobacco; other abbreviations, see Table 1. The top 20 investigated traits are conditional on the
- 379 reporting of r_{MZ} and r_{DZ} . Empty cells denote insufficient information available to calculate
- 380 weighted estimates; error bars denoted standard errors. We note that estimates and Figures for *all*
- 381 specific traits are available in the online webtool MaTCH.

382 Online Methods

383 Identifying relevant studies

We searched PubMed for all studies published between Jan 1st, 1900 and December 31st, 2012

- that provided twin correlations, concordance rates, a heritability (h^2) , or environmentability (c^2)
- 386 estimate, based on monozygotic (MZ) and dizygotic (DZ) twin data. The following search term
- 387 was used in PubMed:
- 388 ("english"[Language] AND ("1900/01/01"[Date Publication] : "2012/12/31"[Date Publication]) AND twin AND

389 "journal article"[Publication Type] AND "humans"[Filter] AND (heritability[Title/Abstract] OR "genetic

390 influence"[Title/Abstract] OR "environmental influence"[Title/Abstract] OR "genetic factors"[Title/Abstract] OR

- 391 "environmental factors"[Title/Abstract]) AND "journal article"[Publication Type]) NOT review[Title] NOT
 392 review[Publication Type].
- The search was run on 31/01/2013 and again on 29/04/2013, which yielded an additional 44 publications, with the difference likely due to keywords or tags that had been added to publications in PubMed in the intermittent period.
- The last PubMed search yielded 4,388 unique studies. From these we deleted studies that werenot relevant for the current purpose using the following exclusion criteria:
- Studies in which only MZ twins were included (including discordant twin case studies)
- 399 Studies in which no heritability estimates, twin correlations or concordances were
 400 provided
- 401 Review studies
- 402 Meta-analysis of other studies
- 403 Multivariate studies that provided information on completely overlapping traits and
 404 samples with previously published univariate studies

Some studies investigating h² of the brain (e.g. voxel-based brain measures) were not
included for practical purposes. These studies typically present their results in graphs
with color-coded point estimates of heritability mapped onto the brain. Such estimates
could not be quantified and thus were not entered into the database.

From the remaining 2,748 studies we were able to retrieve full texts from all but five (99.8%). Of

the studies without full-text availability, we included available relevant information based on theabstract.

412

413 Primary information obtained from each study

From every study we retrieved basic information on PubMed ID, authors, the trait as named in
the study, and the year of publication. All information was entered into a database. In addition
the following information was retrieved:

- 417
- The country of origin of the study population: We used standard ISO country names and
 where possible data entry was done separately for each country investigated in the study.
 Note that in studies where pooled estimates across different countries were provided, yet
 sample sizes were available for each country separately, we entered information
 separately for each country.
- *The age group of the study population:* The study population was classified into four age cohorts based on the average age of the population included in the study, as: age > 0 and
 (12; age >= 12 and < 18; age >= 18 and < 65; and age >= 65.
- *The MZ and DZ twin correlations:* Twin correlations were entered as provided in the
 study, and could be calculated as intra-class, Pearson, polychoric or tetrachoric

428 correlations, or based on least squares or maximum likelihood estimates. When available

429 we entered the twin correlations separated for males and females (i.e., MZ males

430 (MZM), MZ females (MZF), DZ males (DZM), DZ females (DZF), DZ opposite sex

- 431 (DOS)). If correlations were not available for males and females separately, we entered
 432 the MZall and DZall correlations, i.e. based on both sexes. In cases where it was clear
- that the DZ correlation was based on same-sex twins only, we entered the DZ same-sex
- 434 (DZSS) correlation.
- The estimates of the heritability (h^2) and shared environmental component (c^2) , under 435 the full ACE (or ADE) model: We entered ' h^2 FULL' and ' c^2 FULL', based on 436 437 estimates under the full ACE (including additive genetic, shared and non-shared 438 environmental influences) or ADE (including additive and non-additive genetic and non-439 shared environmental influences) model. When an ACE model was fitted, the estimate for A was entered in 'h²_FULL' and the estimated for C in 'c²_FULL'. When an ADE 440 model was fitted, the estimates of A and D were summed and entered for ' h^2 FULL' 441 and zero was entered for ' c^2 FULL'. When estimates were provided separately for 442 443 males and females, they were entered separately; otherwise they were entered for the 444 total sample of males and females. In the case of multivariate analyses, univariate estimates were always preferred to allow comparison across studies. 445
- 446 The estimates of the heritability (h^2) and shared environmental component (c^2) , under 447 the best fitting ACE (or ADE) model: We entered 'h²_BEST' and 'c²_BEST', based on 448 estimates under the best fitting ACE or ADE model as provided in the study. When an 449 ACE model was the best fitting model, the estimate for A was entered in 'h²_BEST' and 450 the estimate for C in 'c²_BEST'. When an ADE model was the best fitting model, the

451		estimates of A and D were summed and entered for 'h ² _BEST' and zero was entered for
452		'c ² _BEST'. When estimates were provided separately for males and females, they were
453		entered separately; otherwise they were entered for the total sample of males and
454		females. In the case of multivariate analyses, univariate estimates were always preferred
455		to allow comparison across studies. In cases where estimates for the best fitting model
456		were not directly provided but information available in the paper indicated that the best
457		fitting model was AE (or CE, or E), we entered zero for ' c^2 _BEST' and missing for ' h^2 _
458		BEST' (when the best fitting model was an AE model), missing for ' c^2 _BEST' and zero
459		for 'h ² _BEST' (when the best fitting model was a CE model), zero for 'c ² _BEST' and
460		zero for ' h^2 _BEST' (when the best fitting model was described to be an E model).
461	-	The total number of twin pairs as used for each entered correlation
462	-	Whether or not the study was a classical twin study: All 2,748 studies in the database
463		include MZ and DZ twins. However a classical twin study was defined as a study that
464		involved only reared-together MZ and DZ twins. From studies that included siblings,
465		extended families, adoptees, or reared apart twins only estimates based on the reared-
466		together twin sample were used for the meta-analyses. Most of these non-classical twin
467		studies did provide twin correlations for the classical twin design, and were thus
468		included in the meta-analysis for twin correlations. When A and C estimates were based
469		on extended twin designs they were excluded from the meta-analyses.
470	-	The method used for estimating the variance components: We entered the statistical
471		method used for estimating the variance components, which included e.g. ANOVA,
472		Bayesian, Maximum Likelihood (ML), DeFries-Fulker regression, least squares (LS), or
473		intra-pair differences. We also listed a dichotomized version of this indicating whether

474		the method used was 'ML or LS' or 'not ML or LS' for all other methods. In the meta-
475		analyses for h^2 and c^2 estimates, only those based on ML or LS were included.
476	-	Whether the trait was dichotomous or continuous: Traits measured as 0/1 as well as
477		traits measured on a quantitative scale but dichotomized before analysis, were listed as
478		dichotomous. All other traits, including ordinal traits were listed as 'not dichotomous'
479		and treated as continuous.
480	-	Whether the study involved ascertainment for the trait: When the trait under
481		investigation was the same trait that was used to select probands, the study was listed as
482		'ascertained'.
483	-	Number of concordant and discordant pairs: In cases of dichotomous traits, the total
484		numbers of pairs for discordant and concordant affected were entered separately for each
485		zygosity. In cases of dichotomous traits that were not ascertained the number of
486		concordant unaffected pairs was also entered.
487	-	Prevalence: In cases of dichotomous traits, the population prevalence, separately for MZ
488		and DZ twins when available was entered. Prevalence was based on what was provided
489		in the study or was calculated using $(2C+D)/2N$, where C is the number of concordant
490		affected pairs, D the number of discordant pairs and N is the total number of pairs in
491		non-ascertained traits.
492		
493	Thus,	provided availability, the statistics in Supplementary Table S1 were obtained for each

Thus, provided availability, the statistics in **Supplementary Table S1** were obtained for each trait reported on in every study. When the five basic twin correlations were available (r_{MZM} , r_{MZF} , r_{DZM} , r_{DZF} , r_{DOS}) we calculated the r_{MZall} , r_{DZSS} and r_{DZall} , using the weighted average via Fisher Z transformation, and using sample size as weights. In situations where r_{MZM} (or r_{MZF}) was exactly 497 1 (or -1), we subtracted (or added in case of -1) .00001 to the correlation to ensure non-498 problematic Fisher Z transformation. Sample sizes of MZall, DZSS and DZall were obtained by 499 summing the sample sizes of MZM and MZF, of DZM and DZF, and of DZM, DZF, and DOS respectively. Estimates of h^2 and c^2 were calculated across sex as the N-weighted average across 500 501 the separate male and female estimates when available. For the number of concordant and 502 discordant pairs, the MZall, DZall and DZSS were calculated based on the numbers available for 503 MZM, MZF, DZM, DZF and DOS. Prevalences for pooled entries were calculated as an N-504 weighted average.

505 Data entry checks

506 Studies were entered and cross-checked for obvious typos by TP and DP. After initial data-entry 507 and initial cross-checking, all data-points were manually checked a second time (DP) by looking 508 up the entered values in the original paper. In addition, automatic checks were run (DP, BB) to 509 identify typos, strange outliers, or obvious mistakes. These checks included:

- Identifying highly unlikely values (clear typos, e.g. correlation of 120)

511 - Testing whether the sum of
$$h^2$$
 and $c^2 < 100$

512 - Testing for strange discrepancies between estimates from the full and best fitting model

513 - Checking for outliers based on extreme sample size and extreme Chi-square values for 514 rejecting the null hypothesis that either $2 \times (r_{MZ} - r_{DZ})$ or $2 \times r_{DZ} - r_{MZ}$ is equal to zero.

515 Classification of traits

516 After data-entry, all traits were manually classified using the International Classification of 517 Functioning, Disability and Health (ICF). The ICF is the WHO's framework for health and 518 disability, and provides the conceptual basis for the definition, measurement and policy formulations for health and disability. It is a universal classification of disability and health for use in health and health-related sectors. ICF belongs to the World Health Organization (WHO) family of international classifications, the best-known member of which is the ICD-10 (International Statistical Classification of Diseases and Related Health Problems). ICD-10 provides an etiological framework for the classification of diseases, disorders and other health conditions, whereas ICF classifies functioning and disability associated with health conditions. The ICD-10 and ICF are therefore complementary.

526

527 Most traits investigated in twin studies concern healthy functioning (e.g. cognitive function, 528 social attitudes, body height, personality) and could be classified according to ICF. In cases 529 where the studied traits were diseases or symptoms of disease, ICD-10 was used. Traits were 530 given two hierarchical classifications corresponding to the ICF/ICD-10 hierarchical structure, 531 using the chapter-structure (e.g. b1), and the level directly under the chapter (e.g. b110), which 532 corresponds to the actual disease (ICD10) or trait (ICF) code.

533 Six new classes at the chapter level and 17 new classes at the sub-chapter level were created to 534 accommodate traits that could not be classified under either the ICF or ICD. For the chapter level 535 the created classes were Cell, Function of DNA, Functions of the nervous system, Medication 536 effects, Mortality, and Structure of DNA. For the sub-chapter level these were All-cause 537 mortality, Cell cycle, Cell growth, Diazepam effects, Expression, Function of brain, Gene 538 expression, Height, Methylation, Mortality from heart disease, mtDNA, Physical appearance, 539 Receptor binding, Sister chromatid exchange, Structure of DNA, Telomeres, and X-inactivation.

541 The standardized classification schemes of the ICF and ICD-10 were queried online, via:
542 http://apps.who.int/classifications/icfbrowser/

- 543 and <u>http://apps.who.int/classifications/icd10/browse/2010/en</u>.
- 544

545 In addition to the two standard ICF/ICD-10 classification levels we added a general classification 546 of functional trait domains. We thus classified all traits using a three level scheme that included 547 28 broad, functional domains, 54 ICF/ICD-10 chapter level classes, and 313 subchapter level 548 classes. A small proportion of studied traits (<0.1%) could not be classified meaningfully on the 549 chapter level (2 traits) or the subchapter level (3 traits). There were 326 unique combinations 550 across the three levels of trait categorizations (Supplementary Table S29). All analyses were 551 conducted on all entries of each of the three levels of classification. In addition we analyzed all 552 traits together. Although this is unspecific in terms of diseases or traits, it provides a general 553 overview of the relationship between MZ and DZ twin correlations, and reveals general patterns 554 of for example sex and age differences. The most specific level is the subchapter level, which is 555 the actual ICD10 diagnosis or a similar ICF classification for normal functioning, reflecting 556 specific traits such as cleft lip, hyperkinetic disorders, or higher-level cognitive function. As 557 researchers do not necessary adhere to the ICD10 or ICF trait nomenclature, traits with the same 558 subchapter classification could have different trait names in the original study, for example for 559 Higher Level Cognitive Function the original studies included trait names total IQ score, 560 cognitive ability, intelligence, or 'g'.

562 Tests for publication bias

Meta-analysis relies on published results. Publication bias can occur when studies that report relatively large heritability estimates or high twin correlations are more likely to be submitted and/or accepted for publication than studies which report more modest effects. Such a publication bias would lead to an overestimation of the true twin correlations or the true heritability and environmental estimates. We used several standard statistical tools to aid in identifying and quantifying possible publication bias, including inspection of funnel plots, Begg and Mazumdar's test²¹, Eggers regression test²² and Rosenthal's Fail Safe N²³.

570

571 Meta-analysis methods of twin correlations and variance components

572 We used the DerSimonian-Laird (DSL) random-effect meta-analytical approach with correlation coefficients as effect sizes, as described by Schulze²⁴ and implemented in the R package 573 574 'metacor'. This function transforms a correlation to its FisherZ value with corresponding 575 standard error prior to the meta-analysis. This method is preferred over conducting a metaanalysis directly on the correlations because the standard error of a twin correlation is a function 576 577 of not only sample size but also of the correlation itself with larger correlations having a smaller 578 standard error. This can cause problems in a meta-analysis since it would lead the larger 579 correlations to appear more precise and be assigned more weight in the analysis, irrespective of 580 sample size. To avoid this problem, the DSL methods transforms correlations to the Fisher's Z 581 metric, whose standard error is determined solely by sample size. All N-weighted computations 582 were thus performed using Fisher's Z, and the results were converted back to correlations for 583 interpretation.

585 The random effects approach allows for heterogeneity of the true twin correlations across 586 different studies. That is, rather than assuming that there is one true level of the twin correlation, the random effects model allows a distribution of true correlations. The combined effect of the 587 588 random effects model represents the mean of the population of true correlations. For 589 computational reasons, correlations of -1 and 1 were converted to -.99999 and .99999 prior to 590 meta-analysis. We set a threshold of at least 5 pairs of twins available per estimate and at least 591 two studies available per category. Meta-analyses were conducted for each category of all three 592 levels of classification.

593

We note that twin samples used in different publications were not independent. For example, 594 595 studies using Australian twins are predominantly based on twins from the Australian Twin Registry. These studies sometimes include different subsamples, but may also include 596 597 completely overlapping samples used to investigate different traits. As participants are 598 anonymous, it is not possible to determine the extent of overlap in the studies included in our 599 analyses. We thus assumed independency of samples in the meta-analyses. This leads to an 600 underestimation of the variance of weighted estimates and an overestimation of their precision. 601 We expect that the dependency of study samples is lowest at the specific level of the ICD10/ICF 602 subchapter level and highest at the general functional domains.

603

604 Meta-analysis for dichotomous, non-ascertained traits

In the DSL random effects model, the standard error of a correlation is calculated based on the provided N (number of twin pairs). The estimated standard errors for continuous traits are correct, but for dichotomous traits the resulting standard error is incorrect. That is because twin 608 correlations for non-ascertained dichotomous traits are typically based on three categories of 609 pairs; concordant unaffected, discordant and concordant affected pairs. While the total number of participating twin pairs is the sum of these three, the information that determines the twin 610 611 correlation and its significance is mostly derived from the latter two. For non-ascertained 612 dichotomous traits we calculated the study-specific tetrachoric twin correlation based on the 613 contingency table (i.e. the number of concordant unaffected, discordant and concordant affected 614 pairs), under the assumption that the dichotomous traits represent latent variables that follow a bivariate normal distribution²⁵. We used a Maximum Likelihood estimator described by 615 Olsson²⁶, which is implemented in the R function 'polycor' to calculate the study specific twin 616 617 correlation and its standard error. As our meta-analysis required the twin correlation and sample 618 size (not standard error) as input and we wanted to be able to pool across continuous and 619 dichotomous traits, we calculated the 'effective' number of twin pairs based on the obtained 620 standard error. The effective number of twin pairs is defined as the number of twin pairs that 621 produces the exact same standard error within the DSL meta-analyses as the standard error 622 obtained from the contingency table.

623

624 Meta-analysis for dichotomous, ascertained traits

For ascertained traits it was not possible to calculate the twin correlations and standard errors based on the contingency table as the traits include only pairs with at least one proband. Without information on the number of concordant unaffected pairs, the prevalence of the affection status is required to calculate a twin correlation. We used the algorithms derived from Falconer²⁷ and Smith²⁸. Again, for practical purposes, calculated standard errors were transformed to an effective number of twin pairs for use in the DSL meta-analysis. 631

632 **Proportion of studies consistent with specific hypotheses**

633 We estimated the proportion of studies that are consistent with H₀:2×(r_{MZ} - r_{DZ})=0 ($\pi_0(h)$) and the 634 proportion of observations consistent with H₀:2× r_{DZ} - r_{MZ} =0 ($\pi_0(c)$), using the Jiang and Doerge 635 method²⁹, as well as the *q*-value method³⁰.

636 Authorship network analysis

We used the approach more fully described in Bulik-Sullivan and Sullivan³¹. Briefly, we retrieved from PubMed the full Medline listing for all twin studies included in this meta-analysis using NCBI "eutils". The output was parsed to capture the names of all authors or collaborators. The twin study author list was manually reviewed to resolve clear inconsistencies in the spelling of the names of authors between publications. This was important so that the same author was not incorrectly considered to be two people, and all were changed to the same character string as to uniquely identify this researcher.

644

645 We then used GEPHI³² to construct a network in order to understand twin study publication 646 patterns. For clarity, we removed individuals who had published only one paper (i.e., we 647 required authorship on \geq 2 papers). The substructure of the network was investigated by 648 estimating community membership modules using the Louvain method implemented in 649 GEPHI³².

651	Web application
652	The data used for this manuscript has been integrated in a web application, where user-specified
653	selections of traits can be made to apply the analyses presented in this work. The web application
654	is called MaTCH: Meta-analysis of Twin Correlations and Heritability and is accessible via
655	http://match.ctglab.nl.
656	
657	
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659	publication bias. <i>Biometrics</i> 50 , 1088–1101 (1994).
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