

# Community-Onset *Escherichia coli* Infection Resistant to Expanded-Spectrum Cephalosporins in Low-Prevalence Countries

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**By global standards, the prevalence of community-onset expanded-spectrum-cephalosporin-resistant (ESC-R) *Escherichia coli* remains low in Australia and New Zealand. Of concern, our countries are in a unique position, with high extramural resistance pressure from close population and trade links to Asia-Pacific neighbors with high ESC-R *E. coli* rates. We aimed to characterize the risks and dynamics of community-onset ESC-R *E. coli* infection in our low-prevalence region. A case-control methodology was used. Patients with ESC-R *E. coli* or ESC-susceptible *E. coli* isolated from blood or urine were recruited at six geographically dispersed tertiary care hospitals in Australia and New Zealand. Epidemiological data were prospectively collected, and bacteria were retained for analysis. In total, 182 patients (91 cases and 91 controls) were recruited. Multivariate logistic regression identified risk factors for ESC-R among *E. coli* strains, including birth on the Indian subcontinent (odds ratio [OR] = 11.13, 95% confidence interval [95% CI] = 2.17 to 56.98,  $P = 0.003$ ), urinary tract infection in the past year (per-infection OR = 1.430, 95% CI = 1.13 to 1.82,  $P = 0.003$ ), travel to southeast Asia, China, the Indian subcontinent, Africa, and the Middle East (OR = 3.089, 95% CI = 1.29 to 7.38,  $P = 0.011$ ), prior exposure to trimethoprim with or without sulfamethoxazole and with or without an expanded-spectrum cephalosporin (OR = 3.665, 95% CI = 1.30 to 10.35,  $P = 0.014$ ), and health care exposure in the previous 6 months (OR = 3.16, 95% CI = 1.54 to 6.46,  $P = 0.02$ ). Among our ESC-R *E. coli* strains, the *bla*<sub>CTX-M</sub> ESBLs were dominant (83% of ESC-R *E. coli* strains), and the worldwide pandemic ST-131 clone was frequent (45% of ESC-R *E. coli* strains). In our low-prevalence setting, ESC-R among community-onset *E. coli* strains may be associated with both “export” from health care facilities into the community and direct “import” into the community from high-prevalence regions.**

Despite a dramatic global rise in the prevalence of expanded-spectrum-beta-lactamase (ESBL)-producing *Escherichia coli*, infections by expanded-spectrum-cephalosporin-resistant (ESC-R) *E. coli* in Australia, New Zealand, North America, and selected European countries remain at relatively low levels. Recent Australian national data show that 3.2% of community isolates carry such resistance. Approximately 80% of these harbor a globally dominant *bla*<sub>CTX-M</sub> ESBL gene and 12% a plasmid-borne AmpC-type mechanism (1). European surveillance data show that a significant proportion of countries have ESC resistance rates below 10% among invasive *E. coli* isolates (2). In the United States, a recent large sample of *E. coli* isolates indicated that 3.9% were ESBL-producing strains (3). Although these low rates offer reassurance in the near term, a year-on-year rise in the incidence of community-onset ESC-R *E. coli* infections in low-prevalence countries is of concern (2, 4).

Australia and New Zealand are in a globally unique position. We have low rates of use of antimicrobials traditionally identified as a risk factor for ESC-R *E. coli*. This includes very low fluoroquinolone use among humans and a ban on the use of ESC and fluoroquinolones in food production (5, 6). In contrast, we have considerable extramural pressure on antimicrobial resistance rates. Our countries are located within the Asia-Pacific region, with which we share a mobile population (7) and frequent commerce (although no land

borders). A high proportion of our regional neighbors have rates of ESC-R among *E. coli* strains in excess of 25% (8, 9).

The aim of our study was to define the risk factors for, and dynamics of, ESC-R among community-onset *E. coli* infections in the low-prevalence settings of Australia and New Zealand by using a case-control methodology. Furthermore, we characterized the resistance genes and membership of the worldwide pandemic clone ST131 in implicated isolates.

## MATERIALS AND METHODS

The COOEE Study (COmmunity Onset ESBL and AmpC *E. coli* Study) was a multisite case-control study, with prospective recruitment of patients and data collection. Six geographically dispersed tertiary centers in Australia ( $n = 5$ ) and New Zealand ( $n = 1$ ) participated. The human

Received 19 September 2013 Returned for modification 14 November 2013

Accepted 19 January 2014

Published ahead of print 27 January 2014

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Supplemental material for this article may be found at <http://dx.doi.org/10.1128/AAC.02052-13>.

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doi:10.1128/AAC.02052-13

research ethics committees at The University of Queensland and participating sites approved this study.

**Definitions.** *E. coli* infection was defined as “community onset” where a patient was resident in the community (including nursing homes) or had been hospitalized less than 48 h at the time of onset; “expanded-spectrum cephalosporin resistance” included all “nonsusceptible” isolates and was identified phenotypically. For ceftriaxone, MIC > 1 mg/liter was used. For ceftazidime, laboratories used MIC > 1 mg/liter or MIC > 4 mg/liter, depending on their use of EUCAST or Clinical and Laboratory Standards Institute (CLSI) criteria, respectively (10, 11); “site of infection” was determined by the researcher from available information. Guidance for urinary tract infections (UTI) was given as follows: “asymptomatic” = a positive urine culture, with no attributable symptoms; “lower tract infection” = lower-urinary-tract-only symptoms such as urgency, frequency, and dysuria; and “upper urinary tract infection” = temperature  $\geq 38^{\circ}\text{C}$ , flank pain, or costovertebral angle tenderness and/or any bacteremia from a urinary source. “Immune suppression” referred to use at the time of the sample collection of corticosteroids (>15 mg/day prednisolone or equivalent), calcineurin inhibitors, other nonbiologics (e.g., mycophenolate and methotrexate), cytotoxic agents, biological agents, or radiation therapy; the Charlson comorbidity index (12) was calculated on the basis of data available from the survey, with the exception of neurological impairment (dementia and hemiplegia), which was inadvertently omitted from the survey questioning. A McCabe score was assigned based on the investigator’s estimate of participant survival (<1 month, 1 month to 2 years, or >2 years) (13). “International travel” (excluding travel between Australia and New Zealand) was classified into geographical regions as follows: South Pacific islands, southeast Asia, Indian subcontinent, China, Japan, North America, Europe, and Africa/the Middle East. “High-risk travel” (regions of the Indian subcontinent, southeast Asia, Africa, the Middle East, and China) was defined *a priori* based on Australian data (14). “Health care exposure” was assessed by the Friedman criteria (15) with two modifications: (i) day procedures were recorded, and (ii) the criteria were assessed in three “discrete” time periods (<1 month earlier, 1 to 6 months earlier, and 7 to 12 months earlier). In addition, exact dates and details of any hospital admissions or surgical procedures were recorded and the interval (in days) from the termination of health care contact to the date of the first medical review with the enrolling *E. coli* infection subject was calculated. Further definitions are provided in the supplemental material.

**Clinical methods.** A case-control methodology was used. Case patients with community-onset ESC-R *E. coli* in a culture of blood or urine were identified in the microbiology laboratory of participating hospitals. Control patients had community-onset ESC-susceptible (ESC-S) *E. coli* isolated from the same specimen type (urine or blood) as the case. Controls were not matched by any clinical presentation, comorbidity, or demographic factors. They were selected as the next appropriate patient, after an enrolled case patient, within the same laboratory’s specimen registration system. If the next appropriate control patient could not be recruited, the process was repeated, at the same time of day and day of week, in a later week of the study. A single control was recruited for each case.

**Inclusions and exclusions.** A laboratory-specific protocol was developed by each site to identify all potentially appropriate patients aged  $\geq 16$  years with an isolate of ESC-R *E. coli* managed at the participating site. Patients cared for by external health care providers such as family doctors and external clinics (utilizing the participating laboratories as an external provider) were not considered for recruitment, due to the complex human-research ethics requirements in our jurisdiction. Initial screening to determine likely community onset and the presence of exclusion criteria was by review of available electronic laboratory data and/or contact with the clinician caring for the patient. Two exclusion criteria were applied: (i) inability of the patient to give informed consent to participate, and (ii) extra-anatomical urinary drainage such as an indwelling urinary catheter (in the community), intermittent catheterization, ileal conduit, or similar. These two groups whose members local clinicians had already identified

as a high risk factor for resistant infection appeared to have relatively distinct demographic and health profiles. Hence, they were excluded in order to focus study resources on a more generalized population group.

**Data collection.** Hospitalized patients or those attending ambulatory clinics were approached for recruitment and data collection in person, whereas the remainder were contacted by telephone. By telephone, at least three contact attempts on different days were made. After informed consent, including explanation of the aims of this study, a structured interview was conducted using a standardized data collection form completed by an investigator under non-blind conditions. Data were primarily self-reported by participants. Where the participant was uncertain of details (e.g., dates of hospitalization or antimicrobial use) or the details were not clear to the investigators on the basis of the answer(s) provided, the investigators were able to review the patient’s medical records held at their institution.

For intermittent exposures (e.g., travel, health care exposure, use of antimicrobials, etc.), participants were asked to recall 12 months before presentation. Exact dates of exposure were recorded. If the exact date was not recalled, it was estimated (“start of month” = the 1st of the month, “middle” or no date specified = 15th, “end of month” = last day).

Data were forwarded to a central coordinator where they were checked and entered into a secure database. Any omissions or discrepancies were clarified with the individual sites.

**Laboratory methods.** All phenotypic susceptibility data presented in this study have been assessed by EUCAST criteria (10). All nonsusceptible isolates were considered “resistant” for the purpose of this analysis. *E. coli* isolates from each patient were forwarded to the research laboratory, with phenotypic identification and antimicrobial susceptibility undertaken by the use of disk diffusion susceptibility testing (DST), an automated system (VITEK 2), or agar dilution, based on the criteria in use by the laboratory at the time. Where data for susceptibility to an ancillary antimicrobial (e.g., nitrofurantoin) were not available, this was assessed by DST in the research laboratory. Where an isolate was originally tested by CLSI, DST using EUCAST criteria was undertaken (in the research laboratory) for agents for which the nonsusceptibility breakpoints of these two criteria differ (ceftazidime, cefepime, amikacin, gentamicin, ciprofloxacin, and nitrofurantoin). Where stated, MICs were performed by Etest (bioMérieux, France). For each isolate, a summative antimicrobial resistance score was calculated from 11 antimicrobials tested (ampicillin, amoxicillin plus clavulanate, ceftriaxone, ceftazidime, cefepime, meropenem, trimethoprim-sulfamethoxazole [SXT], ciprofloxacin, nitrofurantoin, gentamicin, and amikacin).

After overnight culture, bacterial DNA was extracted using an Ultra-Clean microbial DNA isolation kit (Mo Bio Laboratories). ESC resistance genes were investigated by PCR using previously published primers and conditions (16–18). A stepwise approach based on local epidemiology of resistance mechanisms was employed. All isolates were investigated for *bla*<sub>CTX-M-1</sub> group and *bla*<sub>CTX-M-9</sub> group genes. Isolates negative for these were investigated for *bla*<sub>CTX-M</sub> (consensus sequence), *bla*<sub>CMY</sub>, *bla*<sub>DHA</sub>, *bla*<sub>TEM</sub>, *bla*<sub>SHV</sub>, and *bla*<sub>VEB</sub>. All isolates were screened for carbapenemase genes (*bla*<sub>NDM</sub>, *bla*<sub>KPC</sub>, and *bla*<sub>IMP</sub>) using an in-house multiplex PCR (19) and a singleplex PCR for *bla*<sub>OXA-48</sub>-like enzymes (20). All PCR amplicons were sequenced in the forward and reverse directions using an ABI3730XL (Life Technologies) capillary sequencer and compared to published sequences in GenBank ([www.ncbi.nlm.nih.gov/GenBank](http://www.ncbi.nlm.nih.gov/GenBank)).

Presumptive ST131 *E. coli* isolates were determined by use of semiautomated repetitive sequence-based PCR (rep-PCR) (DiversiLab, bioMérieux, France). Isolates clustering within 95% similarity to multilocus sequence type (MLST)-confirmed ST131 reference clones, using a Pearson correlation coefficient, were considered members of this clone (21). A selection of isolates ( $n = 4$ ) were confirmed as ST131 by formal MLST analysis (22).

**Statistical methods.** Sample sizes with overseas travel as a risk factor for resistant infection were calculated. With an estimated annual rate of overseas travel of 250/1,000 population (7), a sample size of 95 cases with

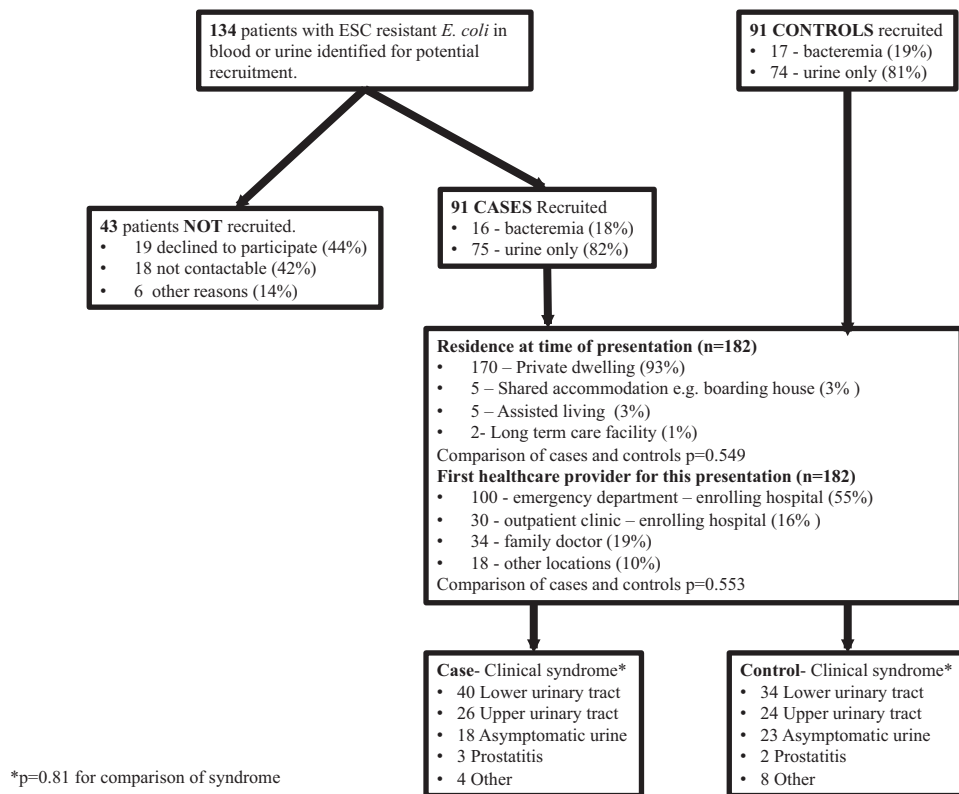


FIG 1 Participant identification and recruitment and characteristics of presentation and clinical syndrome.

matched controls was required to detect this risk with odds ratio  $\geq 2.5$  (power of 0.8 and two-sided alpha of 0.05).

Continuous data on health care exposure was right-censored at 365 days. Univariate comparison was undertaken by a  $\chi^2$  test, Fischer's exact test, Wilcoxon rank sum test, and logistic regression as indicated. Interactions were examined. A multivariate logistic regression model with variables significant in univariate analysis at a  $P = 0.2$  level was constructed. Using backward selection, variables were retained in the final logistic regression model if their significance remained below  $P = 0.2$ . Models were assessed by calculation of a receiver operating characteristic (ROC) curve and Hosmer-Lemeshow goodness of fit. All statistical tests were two-tailed, and  $P < 0.05$  was considered significant. STATA version 12.1 (Statacorp) was used.

## RESULTS

In total, 182 patients (91 cases and 91 controls) were recruited between March 2011 and October 2012 (Fig. 1). Patients were recruited over 12 continuous months at five sites and over 9 months at one site. Sites contributed between 8 and 58 patients.

Bacteremia was detected in 33 patients (18%), and positive cultures were grown from urine samples collected from the remaining 149 (82%). Uneven numbers of bacteremias occurred as one control patient recruited with a positive urine culture subsequently manifested a positive blood culture. The residences of the patients before presentation, clinical syndromes of presentation, and characteristics of hospital presentation did not differ significantly between case and control patients (Fig. 1).

A further 43 patients with presumed community-onset ESC-R *E. coli* infection and no overt exclusion criteria were not recruited (declined to participate,  $n = 19$ ; not contactable,  $n = 18$ ; other,  $n = 6$ ). On comparison with recruited study participants, the me-

dian age (56 years,  $P = 0.81$ ) and gender (11/43, 26% male,  $P = 0.39$ ) did not differ significantly from those of the recruited patients and they were not analyzed further.

Close temporal matching of cases and controls was not frequent. Samples from 9 controls originated from the same calendar day as those from the matched case. For the entire cohort, there was a median interval of 22 days between the dates of collection of the case and control samples.

**Phenotype, resistance genes, and ST131.** All case patients' *E. coli* isolates demonstrated phenotypic ESC resistance (ceftriaxone plus ceftazidime = 60 [68%], ceftriaxone only = 28 [32%], ceftazidime only = 3 [3%]). For the three *E. coli* isolates with ceftazidime resistance, the MICs of ceftazidime in the study laboratory were  $>256$  mg/liter, 2 mg/liter, and 0.25 mg/liter, respectively. All control patient isolates were susceptible to ceftriaxone and ceftazidime. For all antimicrobials studied, with the exception of meropenem (100% susceptible) and amikacin (case = 4 resistant/91 [4%], control = 0 resistant/91 [0%],  $P = 0.121$ ), resistance was significantly more likely in the ESC-R isolates than in the ESC-susceptible (ESC-S) isolates. For ESC-R *E. coli*, there was significant resistance to the oral therapeutic options investigated, including amoxicillin plus clavulanate (ESC-R = 59/91 [65%] versus ESC-S = 15/91 [16%],  $P = <0.001$ ), ciprofloxacin (57/91 [63%] versus 6/91 [7%],  $P < 0.001$ ) and SXT (64/91 [70%] versus 20/91 [22%],  $P < 0.01$ ).

*E. coli* isolates from 89 cases (98%) and 90 (99%) controls were available for further analysis. Carbapenemases were not detected in any isolates. Expanded-spectrum cephalosporinase genes were detected in 87 of 89 (98%) ESC-R *E. coli* isolates as follows: for

ESBLs, the *bla*<sub>CTX-M-1</sub> group (36/89, 40%), *bla*<sub>CTX-M-9</sub> group (35/89, 39%), *bla*<sub>CTX-M-1</sub> and *bla*<sub>CTX-M-9</sub> group (3/89, 3%), and *bla*<sub>SHV-5</sub> group ( $n = 1$ ; 1%); and for non-ESBLs, the *bla*<sub>CMY-2</sub> group ( $n = 11$ ; 12%) and *bla*<sub>DHA-1</sub> group ( $n = 1$ ; 1%). The two remaining samples included two of the three *E. coli* isolates with ceftazidime resistance (MICs, 2 mg/liter and 0.25 mg/liter) and contained only *bla*<sub>TEM-1</sub>, a non-expanded-spectrum beta-lactamase. ESC nonsusceptibility most likely originated from hyperproduction of this enzyme, with loss of this trait during passage and storage in the case of the isolate with the lower drug MIC.

The worldwide pandemic ST131 clone was presumptively identified in 46 patients (24%), who were significantly more likely to be case patients than controls (40/89 [45%] versus 6/90 [7%],  $P < 0.001$ ). Among ESC-R *E. coli* isolates, ST131 was not associated with any non-CTX-M enzymes. They constituted 54% of the entire group of CTX-M isolates. In total, 24 (60%) harbored a CTX-M-1 group enzyme and 19 (48%) a CTX-M-9 group enzyme ( $P = 0.173$  for the comparison). This includes three isolates (8%) harboring both enzymes. There was no significant difference in the proportions of ST131 by sample type (blood versus urine,  $P = 0.514$ ) or hospital site ( $P = 0.574$ ). With the exception of the smallest site (where 0 of 8 were ST131), the clone constituted 19% to 32% of the isolates from each site.

**Demographics, comorbidities, and antimicrobial use.** Age data were compared by visual inspection of histograms. Cases and controls had similar bimodal distributions, with peaks at approximately 25 and 65 years. Median and 25th to 75th percentiles for ages of cases and controls, respectively, were 61 years (21 to 82) and 59 years (19 to 87) ( $P = 0.769$ ). Results of univariate comparisons of demographic factors and medical comorbidities between cases and controls are shown in Table 1. Male sex was the only variable with a significant difference (odds ratio [OR] = 2.3, 95% confidence interval [CI] = 1.5 to 4.6,  $P = 0.018$ ).

Risk from previous urinary tract infection, renal allograft transplant, and anatomical abnormality of the renal tract was investigated (Table 1). The number of urinary tract infections in the previous year was significantly associated with ESC-R *E. coli*, with an odds ratio of 1.32 (95% CI = 1.08 to 1.63,  $P = 0.008$ ) per infection.

Results of univariate analysis of antimicrobial use in the previous year are shown in Table 1. Where the patient could not recall the antimicrobial taken, it was recorded as “unknown.” Exposure to trimethoprim or trimethoprim with sulfamethoxazole (SXT) (OR = 3.02, 95% CI = 1.13 to 8.12,  $P = 0.028$ ) was a significant risk factor for ESC-R *E. coli*. In addition, 7 of 7 patients who had been exposed to an expanded-spectrum cephalosporin (ceftriaxone, ceftazidime, or cefepime) had ESC-R *E. coli* isolated.

**Health care exposure.** Health care exposure was analyzed using two distinct sets of data. First, health care exposure, classified using Friedman criteria for health care-associated (HA) infection, was analyzed in three time windows, with and without the inclusion of day procedures. Exclusion of day procedures performed marginally better at predicting ESC-R; exposure 0 to 1 month earlier (OR = 3.56, 95% CI = 1.14 to 11.14,  $P = 0.029$ ) and 2 to 6 months earlier (OR = 2.99, 95% CI = 1.50 to 5.98,  $P = 0.002$ ) was associated with ESC-R *E. coli* whereas exposure 7 to 12 months earlier ( $P = 0.705$ ) was not (full details are provided in the supplemental material).

Second, a continuous model of the temporal risk of ESC-R *E. coli* infection after health care exposure was generated using the

exact time interval since last hospital admission. Day procedures were excluded based on the results of the first analysis. This smoothed curve of the odds ratios shows the lower bound of the 95% CI approaching an odds ratio of 1.0 at approximately 4 to 5 months (Fig. 2).

**Travel, community, and occupational exposure.** Travel in the previous year was analyzed by region. Travel to the Indian subcontinent approached but did not achieve significance ( $P = 0.09$ ). Birth on the Indian subcontinent was a significant risk factor (OR = 6.119, 95% CI = 1.32 to 28.44,  $P = 0.021$ ) (Table 1).

Occupational exposure to animals, medical patients, and potential household risks was assessed, as was consumption of a variety of meats. No factors were significant (Table 1). Probable household transmission of ESC-R *E. coli* was suggested in one case where the partner of an enrolled patient had an infection with a highly similar isolate (99% identical by rep-PCR using Diversilab) 3 months prior.

**Multivariate analysis.** For the multivariate model, health care exposure in the previous 6 months, excluding day procedures, was selected as a pragmatic option (univariate OR = 2.95, 95% CI = 1.59 to 5.46,  $P = 0.001$ ). This dichotomous measure was nonsignificantly different from the four-category measurement used earlier (likelihood ratio test  $P = 0.821$ ). Travel to high-risk regions was selected from the travel group (OR = 1.97, 95% CI = 0.94 to 4.11,  $P = 0.071$ ). Use of an expanded-spectrum cephalosporin was combined with use of trimethoprim and SXT, in order to enter the former into the model, given its accepted prominence as a risk factor for ESC-R *E. coli* infection.

Significant variables in multivariate analysis were health care exposure, excluding day procedures in the previous 6 months ( $P = 0.002$ ), birth on the Indian subcontinent ( $P = 0.004$ ), travel to high-risk regions ( $P = 0.011$ ), SXT/ESC use ( $P = 0.014$ ), and number of UTIs in the previous year ( $P = 0.003$ ) (Table 2). Assessment of the final model demonstrated an area under the ROC curve of 0.77 and a nonsignificant Hosmer-Lemeshow goodness of fit ( $P = 0.289$ ).

**Interactions and alternative models.** A significant correlation between travel to high-risk regions and region of birth occurred. Those born in high-risk regions were more likely to undertake high-risk travel than those born elsewhere (17/28 [61%] versus 21/154 [14%],  $P < 0.001$ ). This was particularly noted for birth and travel to the Indian subcontinent (7/13 [54%] versus 31/169 [18%],  $P = 0.002$ ). This correlation, and the use of differing parameters for health care contact and antimicrobial exposure, were explored in alternative multivariate models (see the supplemental material). Specific population subgroups (symptomatic patients only, ESC-R *bla*<sub>CTX-M</sub> patients only, and ESC-R ST131 patients only) were also tested in the model. None of the alternative models performed better than the final model, although the levels of significance of health care exposure, male sex, and region of birth/travel differed depending on the model parameters selected.

**HA and non-HA ESC-R *E. coli*.** A difference in the levels of risk of health care-associated (HA) ESC-R *E. coli* and non-HA ESC-R *E. coli* was separately investigated by analysis of risks within the HA ( $n = 73$ ) and non-HA ( $n = 109$ ) cohorts (full details are provided in the supplemental material). Several of the identified risks for ESC-R *E. coli* appeared to be most concentrated in one cohort. Data corresponding to travel to high-risk regions ( $P = 0.001$ ), birth on the Indian subcontinent ( $P = 0.006$ ), and male sex ( $P = 0.018$ ) were statistically significant only among the mem-

TABLE 1 Univariate analysis of demographics, comorbidities, antimicrobial use, region of travel and birth, occupational and household exposure

Subject variable	Frequency in ESC-R cases (%) (n = 91)	Frequency in ESC-S controls (%) (n = 91)	Odds ratio	95% CI	P value
<b>Demographics + comorbidities</b>					
Male sex	30 (33)	16 (18)	2.31	1.51–4.62	0.018 <sup>b</sup>
Age < 30 or > 59 yrs	66 (73)	60 (66)	1.36	0.72–2.57	0.336
Immune suppression	19 (20)	10 (11)	1.99	0.87–4.60	0.105 <sup>b</sup>
Charlson score ≥ 1	44 (48)	34 (37)	1.57	0.87–2.83	0.135 <sup>b</sup>
Active malignancy	11 (13)	9 (8)	1.43	0.55–3.73	0.469
Renal failure	11 (13)	9 (10)	1.25	0.49–3.19	0.636
McCabe score ≥ 2+	78 (86)	76 (84)	1.18	0.53–2.65	0.681
Indigenous	7 (8)	6 (7)	1.18	0.38–3.66	0.774
Heart disease	7 (8)	7 (8)	1		
Long-term-care-facility resident	1 (1)	1 (1)	1		
Smoker	12 (13)	14 (15)	0.83	0.36–1.92	0.672
Liver disease	3 (3)	4 (4)	0.74	0.16–3.41	0.701
Lung disease	5 (5)	7 (8)	0.70	0.21–2.85	0.552
Pregnant or postpartum	3 (3)	7 (8)	0.41	0.10–1.63	0.206
<b>Renal tract background</b>					
Renal transplant	8 (9)	4 (4)	2.1	0.61–7.22	0.241
Anatomical or structural abnormality	23 (25)	15 (16)	1.71	0.83–3.55	0.147 <sup>b</sup>
UTIs in past 12 mos (per UTI) <sup>c</sup>	Median = 1; IQR = 0–3	Median = 0; IQR = 0–1	1.32	1.08–1.63	0.008 <sup>b</sup>
UTIs in lifetime (per UTI) <sup>c</sup>	Median = 2; IQR = 0–5	Median = 2; IQR = 0–5	1.03	0.90–1.18	0.657
<b>Antimicrobial use</b>					
Any antimicrobials in past 12 mos	69 (76)	62 (68)	1.47	0.76–2.81	0.249
Trimethoprim ± sulfamethoxazole	16 (17.58)	6 (6.59)	3.022	1.13–8.12	0.028 <sup>b</sup>
Expanded-spectrum cephalosporin(s)	7 (8)	0	NA		0.014 <sup>b</sup>
Fluoroquinolone(s)	7(8)	3 (3)	2.44	0.61–9.77	0.206
β-Lactam + β-lactamase inhibitor	16 (17.58)	11 (12.09)	1.552	0.68–3.56	0.300
Carbapenem(s)	3 (3.3)	2 (2.2)	1.517	0.25–9.30	0.652
Aminoglycoside(s)	5 (5)	4 (4)	1.26	0.33–4.87	0.733
Macrolide	6 (6.59)	5 (5.49)	1.214	0.36–4.13	0.756
“Unknown” antimicrobial(s)	35 (38)	33 (36)	1.1	0.60–2.00	0.759
Narrow-spectrum cephalosporin(s)	16 (17.58)	15 (16.48)	1.081	0.50–2.34	0.844
Narrow-spectrum penicillin(s)	10 (10.99)	14 (15.38)	0.679	0.28–1.62	0.383
<b>Travel by region<sup>d</sup></b>					
Any overseas travel	28 (30.8)	22 (24.18)	1.39	0.72–2.68	0.32
High-risk regions	24 (26)	14 (15)	1.97	0.94–4.11	0.071 <sup>b</sup>
Indian subcontinent	6 (6.59)	1 (1.1)	6.928	0.75–53.87	0.09
North America	5 (5.49)	2 (2.20)	2.199	0.49–13.69	0.264
Africa + the Middle East	3 (3.3)	2 (2.2)	1.517	0.25–9.30	0.652
Southeast Asia	15 (16)	13 (14)	1.18	0.53–2.65	0.681
South Pacific	3 (3.30)	3 (3.30)	1		
Europe	3 (3.30)	5 (5.49)	0.586	0.14–2.53	0.474
China	4 (4.4)	0			0.121
Japan	1 (1.1)	0			0.500
<b>Birth by region</b>					
High-risk region	18 (20)	10 (11)	2.0	0.87–4.60	0.105
Indian subcontinent	11 (13)	2 (2)	6.12	1.32–28.45	0.021 <sup>b</sup>
Australia + New Zealand	58 (64)	59 (65)	0.95	0.52–1.75	0.877
Europe	15 (16)	18 (20)	0.80	0.38–1.71	0.564
Southeast Asia	3 (3)	4 (4)	0.74	0.16–3.41	0.701
Africa + Middle East	2 (2)	4 (4)	0.49	0.09–2.74	0.415
China	2 (2)	0			0.497
South Pacific	0	3 (3)			0.246
Latin America	0	1 (1)			1.0
<b>Occupation and household exposure</b>					
Partner with recent ESC-R <i>E. coli</i> infection	2 (2)	Not assessed			
Occupational health care exposure	10 (11)	7 (8)	1.48	0.54–4.08	0.447

(Continued on following page)

TABLE 1 (Continued)

Subject variable	Frequency in ESC-R cases (%) (n = 91)	Frequency in ESC-S controls (%) (n = 91)	Odds ratio	95% CI	P value
Pet cat or dog or both at home	32 (35)	33 (36)	0.95	0.52–1.75	0.877
Occupational animal exposure	4 (4)	5 (5)	0.79	0.21–3.05	0.733
Preschoolers at home (<5 yrs of age)	7 (8)	9 (10)	0.76	0.27–2.13	0.601
<b>Food consumption</b>					
Any meat in past 12 mos	89 (98)	87 (98)	2.05	0.37–11.46	0.415
Poultry	88 (97)	83 (92)	2.47	0.62–9.89	0.206
Processed/preserved meats	51 (56)	52 (58)	0.93	0.52–1.68	0.814
Pork	60 (66)	63 (70)	0.83	0.44–1.55	0.558
Red meat	76 (84)	78 (88)	0.72	0.31–1.66	0.433

<sup>a</sup> Destinations of travel by region were as follows: for the Indian subcontinent, India, Pakistan, Nepal, and Bangladesh; for North America, the United States and Canada; for Africa and the Middle East, Zimbabwe, Kenya, Sudan, Liberia, Turkey, and Afghanistan; for Southeast Asia, Malaysia, Singapore, Thailand, Laos, Cambodia, Vietnam, Burma, Indonesia, and The Philippines; for South Pacific, New Caledonia, Papua New Guinea (PNG), Fiji, Samoa, Cook Islands, and boat cruises through the South Pacific; for Europe, the United Kingdom, Italy, Holland, Portugal, and Poland; for China, China, Hong Kong, and Macau; for Japan, Japan. High-risk regions include the Indian subcontinent, Africa, the Middle East, Southeast Asia, and China regions.

<sup>b</sup> Entered into multivariate model.

<sup>c</sup> Infections were recorded numerically on a scale of 0 to 5+, with all 5+ results considered 5 for analysis. Summaries are presented as a median value and an interquartile range (IQR).

bers of the non-HA group. Conversely, data corresponding to a risk from SXT or ESC use were statistically significant only among members of the HA group ( $P = 0.026$ ). The numbers of UTIs in the previous year were nonsignificantly different among the members of either group assessed separately.

**Correlates of the classes of ESC resistance enzymes.** Correlates of ESC resistance enzyme classes were investigated by a comparison of patients harboring *E. coli* with CTX-M group enzyme to those harboring other enzymes (“non-CTX-M” = CMY, DHA, SHV, TEM). Full details are provided in the supplemental material.

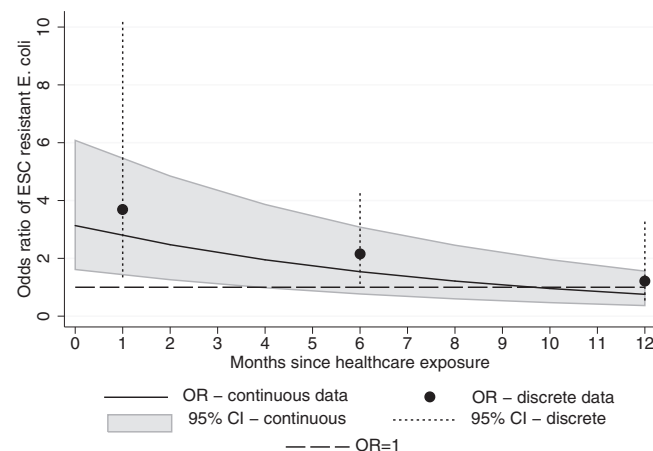
There was no statistically significant difference with respect to the site of infection between CTX-M-harboring and non-CTX-M-harboring participants ( $P = 0.473$ ), and although bacteremia was more frequent in the CTX-M group, this did not reach statis-

tical significance (15/73 [21%] versus 0/15 [0%],  $P = 0.063$ ). A significantly higher median resistance score was present in CTX-M isolates than in non-CTX-M isolates (median = 6 [interquartile range = 5 to 7] versus median = 4 [interquartile range = 4 to 5],  $P = 0.001$ ). Notable differences included higher rates of resistance to the non-beta-lactam oral agents ciprofloxacin (56/74 [76%] versus 1/15 [7%],  $P < 0.001$ ) and SXT (60/74 [81%] versus 4/15 [27%],  $P < 0.001$ ) among the members of CTX-M group.

In regard to potential risk factors, the members of the CTX-M group were significantly more likely to have had health care exposure in the previous 6 months than the members of the non-CTX-M group (45/74 [61%] versus 3/15 [20%],  $P = 0.005$ ) although the same was not true with respect to health care exposure in the previous 12 months ( $P = 0.72$ ). Other factors used in the multivariate model that trended toward significance among the members of the CTX-M group included more high-risk travel ( $P = 0.052$ ) and fewer reported UTIs in the previous 12 months ( $P = 0.054$ ). Comparisons of factors not included in the multivariate model showed that “any overseas travel” was more likely in the CTX-M group (27/74 [36%] versus 1/15 [7%],  $P = 0.033$ ).

## DISCUSSION

This multicenter prospective case-control study of community-onset ESC-R *E. coli* infection has several key findings that have implications for risk-based empirical antibiotic prescription and infection control practices and for control of ESC-R *E. coli* infections within communities.



**FIG 2** The risk of ESC-R *E. coli* infection over a 12-month period after the most-recent episode of health care exposure, excluding day procedures, estimated with two data sets. The smoothed curve was calculated using continuous data corresponding to the months since hospital admission (black line; 95% CI in gray). Discrete intervals determined using Friedman criteria are indicated (black dots; 95% CI as vertical dashes). The dashed line represents no increased risk (odds ratio = 1.0).

**TABLE 2** Multivariate logistic regression

Subject variable	Odds ratio	95% CI	P value
Health care exposure in the previous 6 mos	3.16	1.54–6.46	0.002
UTIs in previous yr (per UTI)	1.43	1.16–1.82	0.003
Birth on the Indian subcontinent	11.13	2.17–56.96	0.004
Travel to high-risk region(s)	3.09	1.29–7.38	0.011
Trimethoprim ± sulfamethoxazole ± ESC use	3.67	1.30–10.35	0.014
Male sex	2.17	0.97–4.84	0.060

First, we established that 6 months is a practical, evidence-based definition for the duration of increased risk of a community-onset *E. coli* isolate harboring ESC-R after health care exposure. The time-dependent relationship of health care exposure to resistance seems intuitive in nature; however, previously there has been little supporting data. Hence, authors have used a variety of definitions from 1 to 6 months (23–25).

Overall, the significant contribution of health care exposure (OR = 3.15) as an ongoing “exporter” of resistant infection in a low-prevalence setting highlights the importance of controlling ESC resistance in the health care system. Supporting this hypothesis, United Kingdom data have recently demonstrated a broad-based decrease in the rate of ESC resistance among invasive *Enterobacteriaceae* strains following a reduction in the use of ESC and fluoroquinolones within the hospital system (26).

The “importation” of ESC-R *E. coli* after travel to countries with a high community incidence of ESBLs is starting to be defined (27), although fewer studies have identified infection rather than carriage (25, 28, 29). While the pathophysiology seems clear, the temporality of this remains to be confirmed. In our study, analysis of temporality, as presented for health care exposure, was precluded by the imprecision of data from the composite “high-risk” group and the low numbers involved. However, in absolute terms, 21 of 24 (87.5%) case participants with travel to high-risk regions departed those regions within the 6 months before presentation of infection. This fits with our previous research demonstrating mostly short-lived carriage of ESBL *E. coli* following travel overseas and with other studies demonstrating a decrease in the risk of resistant infection beyond 6 weeks after return from travel (29, 30).

Investigation of risks for community acquisition in the low-prevalence countries of Australia and New Zealand showed that one-quarter ( $n = 23$ ) of ESC-R *E. coli* patients reported neither health care exposure nor high-risk travel, suggesting there are as-yet-undefined risk factors for transmission within the community (25, 31).

While there was some correlation between birth and travel region data, the identification of birth on the Indian subcontinent (OR = 11.12) as a risk factor for ESC-R *E. coli* infection in our cohort appears genuine. The etiology of this risk could stem from prolonged carriage of ESC-R *E. coli* after travel more than 1 year previous, leading to delayed community-onset ESC-R *E. coli* infection. Alternatively, our observation of a mostly short interval between travel and infection supports the possibility of domestic (within Australia and New Zealand) transmission of this resistance. Transmission of ESC-R *E. coli* from others within the household or community who have had recent travel to the Indian subcontinent may occur. Although the true magnitude of risk and the broader applicability require further study, this observation is consistent with a previously published study from one of our participating sites and with other descriptions of household transmission (32, 33). Recently, “birth outside Europe” was identified as a risk factor for infection by CTX-M-producing *E. coli* in another study, although comparison with our data is complicated, as the European study did not fully account for recent travel (31).

Our molecular epidemiology data serve to confirm a number of key observations made in other regions. The first is a distinct difference between the epidemiology of CTX-M ESBLs and that of other expanded-spectrum cephalosporinase enzymes, which may be mediated by the differing modes of acquisition, phenotypes,

and characteristics of the *E. coli* strains harboring them (16, 34). Second, the high proportion of the ST131 clone among ESBLs is no surprise given its global prevalence (35). More surprising is its predominance without significant fluoroquinolone use (<6% of all participants in this study), one of the likely drivers in other regions (36). Exposure to this class of antimicrobials within Australia and New Zealand is very low (5).

Male sex has been defined by other researchers as a risk factor for community-onset ESC-R *E. coli* infection (23, 37–40) and became significant in some of our alternative models. The patient population of studies with this finding gives a clue to the etiology of this risk. On the whole, they are of older age with frequent health care exposure. This contrasts with studies conducted with a more traditional UTI population of young females that did not identify male sex as a risk factor (28, 29). In addition to males experiencing an age-dependent rise in the overall rates of *E. coli* infection (41), a limitation associated with case-control studies may also contribute to this finding. Aging patients certainly experience changes in the nuances and dynamics of health care exposure and other potential risk factors for ESC-R *E. coli* infection not identified with the data corresponding to dichotomous measures such as hospitalization and antimicrobial use that are most often collected.

The strengths of our study include its prospective collaborative nature, a geographically broad sample range, and the case-control methodology used. The low background rates of ESC-R *E. coli* infection in Australia and New Zealand have likely led to more discrete exposures and easier delineation of temporal risks than in communities where participants are frequently exposed to this form of resistance.

Limitations of our study include the moderate sample size, rate of nonrecruitment, and risk of bias due to an absence of blind investigator or patient procedures and reliance on patient recall for many exposures. Recruiting a higher ratio of controls (1:2 or 1:3) would have increased our study power and might have delineated further unidentified risks. The use of a third group of uninfected patients (a case-case-control design) would have allowed for delineation of risk factors associated with de novo acquisition of ESC-R *E. coli*, as opposed to delineation of risk factors for ESC-R within those that have *E. coli* (42). However, pragmatic limitations precluded these options.

Some unique features of Australia and New Zealand may limit extrapolation of our findings to other regions. The exclusion of day procedures in this study’s definition of health care exposure correlated with our local epidemiology and would need to be reconsidered elsewhere. Furthermore, in cases in which *bla*<sub>CTX-M</sub> was not the predominant ESC resistance mechanism in a local population, risk data might differ.

The use of only hospital patients for recruitment allowed consistent access to participants and samples, although it might limit the applicability of some risks to the wider community. The exclusion of patients unable to consent meant that we could not define risks for patients in long-term-care facilities, a known reservoir of ESBL *E. coli* in Australia and overseas (43–45).

In conclusion, we have defined a critical ESC-R risk period after health care exposure among community-onset *E. coli* infections and demonstrated that ESC-R *E. coli* infection in a low-prevalence settings may be driven by “export” from health care exposure in the previous 6 months and importation after travel to regions with a high incidence of community ESBLs.

## ACKNOWLEDGMENTS

We thank the clinical and microbiology laboratory staff at all sites. They greatly contributed to this study. In particular, we thank Tiffany Brown, Anna Sartor, and Wan Keat Yam (UQ), Julian Williams and Narelle George (RBWH), Cheryl McCullough (PathWest, RPH), and James Pollard and Owen Harris (St John of God Pathology, Geelong).

The laboratory component of this work was supported by a grant from The Royal Brisbane and Womens Hospital Foundation. The use of the Diversilab for this study was supported by bioMérieux (Australia), who supplied consumables for the Diversilab at a discounted cost. B.A.R. is funded by an Australian Postgraduate Award Scholarship. D.L.P. has received honoraria from AstraZeneca, Merck, and Pfizer. P.R.I., N.R., M.C.P., J.T.F., E.A., S.M.H., H.E.S., M.J., E.G., M.D.A., and K.S. received no funding.

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