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# Association of XRCC3 18067 C>T (Thr241Met) polymorphism with risk of cervical and ovarian cancers: A systematic review and meta-analysis

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**Abstract:** The 18067 C>T polymorphism of XRCC3 gene has been considered to be implicated in the development of cervical and ovarian cancers, but the results are inconsistent. Thus, we conducted a meta-analysis to assess the association of XRCC3 18067 C>T polymorphism with risk of cervical and ovarian cancers. All studies on the association of XRCC3 18067 C>T polymorphism with cervical and ovarian cancers risk were retrieved. Finally, a total of 17 studies including 10 studies with 5,637 cases and 10,057 controls on ovarian cancer and 7 studies with 1,112 cases and 1,233 controls on cervical cancer were selected. Overall, pooled results showed that the XRCC3 18067 C>T polymorphism was significantly associated with increased risk of ovarian cancer (TC vs. CC: OR = 0.904, 95% CI = 0.841–0.972,  $p = 0.006$ ; TT + TC vs. CC: OR = 0.914, 95% CI = 0.853–0.979,  $p = 0.010$ ) and cervical cancer (TC vs. CC: OR = 1.00, 95% CI = 1.066–1.585,  $p = 0.009$ ). Further subgroup analysis by ethnicity revealed an increased risk of cervical and ovarian cancer in Asians and Caucasians, respectively. The present meta-analysis inconsistent with the previous meta-analysis suggests that the XRCC3 18067 C>T polymorphism might be implicated in the pathogenesis of cervical and ovarian cancers.

**Keywords:** cervical cancer, ovarian cancer, XRCC3 gene, polymorphism, meta-analysis

## Introduction

Cervical and ovarian cancers remain two of the leading cause of cancer mortality worldwide among women and the most common site in several low-income countries [1, 2]. It is widely accepted that certain oncogenic types of human papilloma virus (HPV) are essential cause of cervical cancer development [3]. Almost 100% of women with a diagnosis of cervical cancer

have been found to have had an HPV infection [4]. 37 Ovarian cancer is characterized by few early symptoms, 38 presentation at an advanced stage, and poor survival [5, 6]. 39 The exact causes of ovarian cancer are not known<sup>1</sup>. 40 Q2 Relatively few risk factors for ovarian cancer have been 41 identified, including age, parity, oral contraceptive use, 42 lifestyle factors, and family history of breast or ovarian 43 cancer, many of these are not easily modifiable on the 44 population level [4, 7]. 45

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46 Genome-wide association studies have been extremely  
47 successful at finding susceptibility loci for cervical and  
48 ovarian cancers [8]. Molecular epidemiological studies  
49 have been conducted with the candidate gene approach  
50 to identify susceptibility genes for cervical and ovarian  
51 cancers, many of which have showed inconsistent  
52 result [9]. DNA repair plays an important role in the  
53 maintenance of genomic integrity by correcting DNA  
54 alterations caused by endogenous and exogenous  
55 genotoxic agents [10]. At present, several DNA repair  
56 genes (e.g., *XPD*, *XPF*, *ERCC1*, *XRCC1*, *XRCC3*,  
57 *XPA*, *XPB*, *XPC*, and *hOGG1*) have been reported to  
58 be associated with cervical and ovarian cancers, and the  
59 X-ray cross-complementing group 3 (*XRCC3*) gene has  
60 received an increasing attention [11, 12].

61 The human *XRCC3* gene (MIM: 600675) is localized  
62 on chromosomes 14q32.3 [13]. It is involved in the  
63 homologous recombination repair (HR) pathway,  
64 responsible for DNA double-strand breaks [14]. *XRCC3*  
65 is a polymorphic gene where many SNPs have been  
66 already described. Several polymorphisms in the *XRCC3*  
67 gene have been described to affect the enzyme function  
68 and/or its interaction with other proteins involved in  
69 DNA damage and repair [13, 14]. Of these, C18607T  
70 transition (rs861539) at exon 7 resulting in an amino  
71 acid change at codon 241 (Thr241Met) has been  
72 studied frequently [13]. This polymorphism has been  
73 reported to be associated with the development of some  
74 cancers, such as bladder, skin, breast, lung, and colorectal  
75 cancers [15].

76 Several epidemiological studies were conducted in  
77 recent years to evaluate the association of the *XRCC3*  
78 18067 C>T polymorphism with cervical and ovarian  
79 cancers [16, 17]. Some studies have shown a significant  
80 statistical correlation of this polymorphism with cervical  
81 and ovarian cancers, whereas others did not find any such  
82 association. Thus, these inconsistent results fail to clarify  
83 this complicated genetic relationship, presumably due to  
84 small sample size in each published study, various genetic  
85 backgrounds, and possible selection bias. To reliably  
86 demonstrate the effect of *XRCC3* 18067 C>T polymor-  
87 phism on cervical and ovarian cancer risks, we performed  
88 a comprehensive systematic review and meta-analysis of  
89 all eligible studies to resolve this pivotal issue.

## 90 Materials and Methods

### 91 Study identification and selection

92 This meta-analysis conformed to the Preferred Reporting  
93 Items for Systematic Reviews and Meta-analyses criteria.  
94 Two investigators independently searched the MED-  
95 LINE (PubMed), Google Scholar, Web of Science  
96 (Thomson-Reuters), Scientific Information Database  
97 (SID), Chinese National Knowledge Infrastructure

(CNKI), the Chinese Wanfang, and the Chinese VIP 98  
databases for eligible articles examined the association of 99  
*XRCC3* 18067 C>T polymorphism with cervical and 100  
ovarian cancer risks published up to January 30, 2019. 101  
The following terms were utilized: (“ovarian cancer” OR 102  
“cervical cancer”) AND (“X-ray repair cross comple- 103  
menting 3” OR “*XRCC3*”) AND (“*XRCC3* 18067 104  
C>T” OR “Thr241Met” OR “rs861539”) AND 105  
 (“polymorphism”, OR “mutation” OR “variant” OR 106  
“gene” OR “genotype” OR “SNP” OR “allele”). The 107  
search was performed without any restrictions on 108  
language and was focused on studies that had been 109  
conducted in humans. In addition, manual searching of 110  
the references of eligible studies, reviews and related 111  
meta-analyses, and the abstracts presented at relevant 112  
conferences were performed to identify potentially 113  
relevant studies. If there were multiple reports of the 114  
same study or overlapping data, only the study with the 115  
largest sample sizes or the most recent one should be in 116  
the final analysis. 117

### Data extraction 118

Information was carefully extracted from all eligible stud- 119  
ies independently by two investigators according to the 120  
inclusion criteria listed above, and potential disagree- 121  
ments were resolved by consensus. The following data 122  
were collected from each study: name of first author, 123  
publication year, country where the study was conducted, 124  
racial descent (categorized as Asian, Caucasian, or mixed 125  
descent), polymorphisms, genotypic testing method, 126  
number of cases and controls, genotype frequency of 127  
cases and controls, minor allele frequencies in control 128  
subjects, and result of Hardy–Weinberg equilibrium 129  
(HWE) test in control subjects. In this meta-analysis, 130  
ethnicity was categorized as: Caucasian, Asian, and 131  
Mixed. 132

### Inclusion and exclusion criteria 133

To be included in the meta-analysis, studies had to meet 134  
all the criteria: (1) use a case-control or cohort design; 135  
(2) assess the association of the *XRCC3* 18067 C>T 136  
polymorphism with ovarian and cervical cancers; and 137  
(3) provide sufficient data for estimating odds ratios 138  
(ORs) with 95% confidence intervals (CIs). The exclusion 139  
criteria were: (1) studies that could not offer the number 140  
of cases and controls or other essential information; 141  
(2) case only or studies without control group; (3) family 142  
based or linkage studies; (4) case reports, reviews, and 143  
studies; and (5) overlapping data. In the case of multiple 144  
studies by the same researchers involving the same or 145  
overlapping data sets, the most recent study with the 146  
largest number of participants was included in the meta- 147  
analysis. 148

149 *Statistical analyses*

150 The strength of association of the XRCC3 18067 C>T  
 151 polymorphism with ovarian and cervical cancers suscep-  
 152 tibility was assessed by OR with the corresponding 95%  
 153 CI. The  $Z$ -test was performed to determine the signifi-  
 154 cance of the pooled OR, with  $p < 0.05$  defined as the  
 155 significance threshold. The pooled ORs were calculated  
 156 for the risk associated with the XRCC3 18067 C>T  
 157 polymorphism in the allele model (T vs. C), homozygote  
 158 model (TT vs. CC), heterozygote model (TC vs. CC),  
 159 dominant model (TT + TC vs. CC), and recessive model  
 160 (TT + TC vs. CC). The between-studies heterogeneity  
 161 was tested using the  $Q$  statistic. If  $p < 0.10$ , the hetero-  
 162 geneity was considered statistically significant. Venice  
 163 criteria for the  $I^2$  test included:  $I^2 < 25\%$  represents no  
 164 heterogeneity,  $I^2 = 25\%–50\%$  represents moderate  
 165 heterogeneity,  $I^2 = 50\%–75\%$  represents large heteroge-  
 166 neity, and  $I^2 > 75\%$  represents extreme heterogeneity.  
 167 The  $p$  value of  $< 0.05$  for the  $Q$ -test indicated a lack of  
 168 heterogeneity among studies, so that the pooled OR  
 169 estimate of each study was calculated by the fixed-effects  
 170 model (the Mantel–Haenszel method), otherwise the  
 171 random effects model (the DerSimonian–Laird method)  
 172 was utilized. Furthermore, to explore the source of  
 173 between-study heterogeneity, the subgroup analyses were  
 174 performed. The one-way sensitivity analyses were  
 175 performed to survey the stability of the results, namely,  
 176 a single study in the meta-analysis was omitted each time  
 177 to reflect the influence of the individual data set to the  
 178 pooled OR. Publication bias was assessed by visually  
 179 examining the asymmetry of a funnel plot in which the  
 180 log estimates were plotted against their standard errors.  
 181 Furthermore, we also employed an Egger’s regression test  
 182 in our analysis to calculate two-tailed  $p$  values for quanti-  
 183 fying publication bias. A HWE test of the VDR gene  
 184 polymorphisms in healthy subjects was examined using  $\chi^2$   
 185 test. If  $p$  value  $> 0.05$ , the genotype distribution of the  
 186 control group conformed to HWE. All the statistical  
 187 analyses were performed by comprehensive meta-analysis  
 188 version 2.0 software (Biostat, USA). All the  $p$  values were  
 189 two sides and less than 0.05 were considered significant.

190 **Results**191 *Study selection and characteristics*

192 A flow diagram schematizing the inclusion and exclusion  
 193 process of identified articles with the inclusion criteria is  
 194 presented in *Fig. 1*. After a comprehensive search, a total  
 195 of 126 literatures were identified. Of these studies, the  
 196 first screening excluded 47 were considered as duplicates  
 197 or not relevant, leaving 79 studies for further selection.  
 198 Finally, a total of 17 case–control studies (in 14 publica-  
 199 tions) were included in this meta-analysis [18–31].

Of these, there were seven studies with 1,112 cases and 200  
 1,233 controls on cervical cancer [18–24] and 10 studies 201  
 with 5,637 cases and 9,267 controls on ovarian cancer 202  
 [25, 27–31]. The main characteristics of studies included 203  
 in the present meta-analysis are presented in *Table I*. Of 204  
 all the eligible studies, four were conducted in Asian, two 205  
 were in Caucasians, and one was in mixed for cervical 206  
 cancer; eight were conducted in Caucasians and two were 207  
 in mixed for ovarian cancer. Twelve studies were popula- 208  
 tion-based and four were hospital-based studies. One 209  
 study in the present meta-analysis did not state the source 210  
 of controls. Four genotyping methods were used, 211  
 including AS-PCR, PCR-RFLP, Pyrosequencing<sup>TM</sup>, and 212  
 TaqMan assay. The genotype distributions among the 213  
 controls in two studies were not consistent with HWE on 214  
 ovarian cancer (*Table I*). 215

216 *Quantitative synthesis*

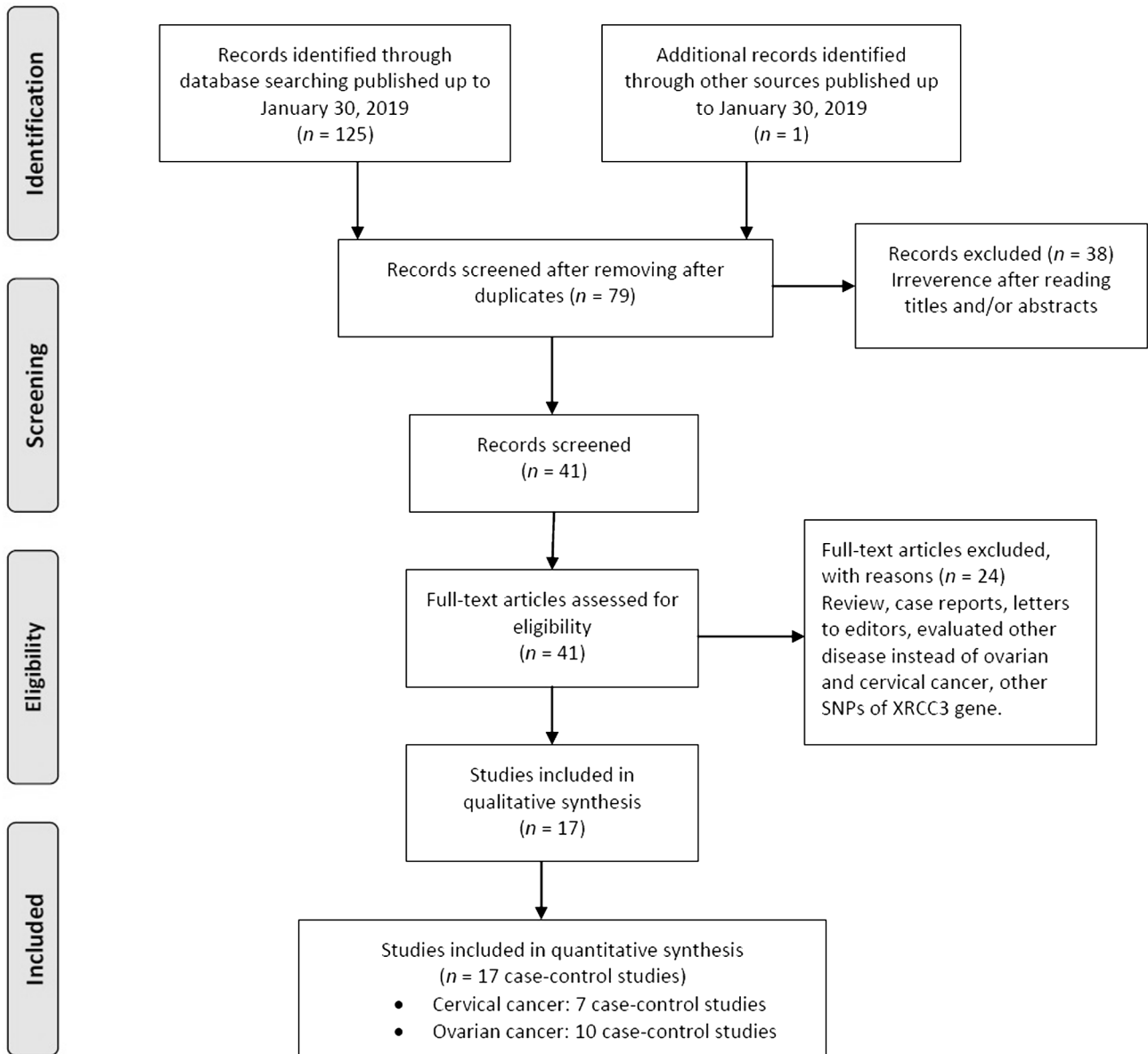
Table II listed the main results of the meta-analysis of 217  
 XRCC3 18067 C>T polymorphism with cervical and 218  
 ovarian cancers risk. When all the eligible studies were 219  
 pooled into meta-analysis, the results showed that 220  
 XRCC3 18067 C>T polymorphism was not significantly 221  
 associated with increased risk of cervical and ovarian 222  
 cancers under all genetic models genetic models, 223  
 i.e., allele (T vs. C: OR = 1.014, 95% CI = 0.930– 224  
 1.106,  $p = 0.745$ ), homozygote (TT vs. CC: OR = 225  
 1.010, 95% = CI 0.855–1.194,  $p = 0.906$ ), heterozygote 226  
 (TC vs. CC: OR = 0.967, 95% CI = 0.876–1.067, 227  
 $p = 0.530$ ), dominant (TT + TC vs. CC: OR = 0.993, 228  
 95% CI = 0.889–1.108,  $p = 0.897$ ), and recessive 229  
 (TT vs. TC + CC: OR = 1.028, 95% CI = 0.894– 230  
 1.183,  $p = 0.700$ ). 231

The studies were further stratified by cancer type and 232  
 ethnicity. When stratified by cancer, there was a signifi- 233  
 cant association between XRCC3 18067 C>T polymor- 234  
 phism and increased risk of cervical cancer under the 235  
 heterozygote model (TC vs. CC: OR = 1.00, 95% 236  
 CI = 1.066–1.585,  $p = 0.009$ ; *Fig. 2A*). Moreover, the 237 Q7  
 XRCC3 18067 C>T polymorphism was significantly 238  
 associated with increased risk of ovarian cancer under 239  
 two genetic models, i.e., heterozygote (TC vs. CC: 240  
 OR = 0.904, 95% CI = 0.841–0.972,  $p = 0.006$ ) and 241  
 dominant (TT + TC vs. CC: OR = 0.914, 95% 242  
 CI = 0.853–0.979,  $p = 0.010$ ; *Fig. 2B*). 243

Subgroup analysis by ethnicity showed that there was a 244  
 significant association between XRCC3 18067 C>T 245  
 polymorphism and cervical cancer in Asian under three 246  
 genetic models, i.e., model (T vs. C: OR = 1.302, 95% 247  
 CI = 1.076–1.576,  $p = 0.007$ ), heterozygote (TC vs. CC: 248  
 OR = 1.441, 95% CI = 1.113–1.867,  $p = 0.006$ ) and 249  
 dominant (TT + TC vs. CC: OR = 1.469, 95% CI = 250  
 1.148–1.880,  $p = 0.002$ ), but not in Caucasians. More- 251  
 over, subgroup analysis showed that there was a 252



PRISMA 2009 Flow Diagram



Q6 Fig. 1.

253 significant association between XRCC3 18067 C>T poly-  
 254 morphism and increased risk of ovarian cancer in Cauca-  
 255 sians under two genetic models, i.e., heterozygote (TC vs.  
 256 CC: OR = 0.898, 95% CI = 0.834–0.967,  $p = 0.004$ ) and  
 257 dominant (TT + TC vs. CC: OR = 0.905, 95% CI =  
 258 0.844–0.970,  $p = 0.005$ ). In the subgroup analyses by  
 259 ethnicity, no studies were performed for ovarian cancer  
 260 in Asians suggesting that our results might be not applica-  
 261 ble for these populations.

*Test of heterogeneity and sensitivity analyses*

262

For cervical cancer, statistical significant heterogeneity 263  
 among studies under four genetic models was observed 264  
 when all eligible studies were pooled into the meta- 265  
 analysis. However, the heterogeneity test showed that 266  
 there was no significant heterogeneity in terms of 267  
 the XRCC3 18067 C>T polymorphism association 268  
 with ovarian cancer. Therefore, to explore the potential 269



Table I. Characteristics of studies included in the meta-analysis

First author (year)	Country (ethnicity)	SOC	Genotyping technique	Case/control	Cases				Controls							
					CC	CT	TT	C	T	CC	CT	TT	C	T		
<i>Cervical cancer</i>																
He (2008)	China (Asian)	PB	AS-PCR	200/200	177	19	4	373	27	182	17	1	381	19	0.047	0.391
Xiao (2010)	China (Asian)	PB	PCR-RFLP	158/164	82	59	17	223	93	115	41	8	271	57	0.173	0.097
Sertheetham-Ishida (2011)	Thailand (Asian)	PB	PCR-RFLP	111/118	101	10	0	212	10	106	12	0	224	12	0.050	0.560
Pérez (2013)	Argentina (Caucasian)	PB	Sequencing	117/205	50	56	11	156	78	78	95	32	251	159	0.387	0.730
Djansugurova (2013)	Kazakhstan (Caucasian)	PB	AS-PCR	217/160	140	57	20	337	97	124	32	4	280	40	0.125	0.278
Colacino-Silva (2017)	Brazil (Mixed)	HB	PCR-RFLP	77/73	43	28	6	114	40	36	30	7	102	44	0.301	0.837
Al-Harbi (2017)	Saudi Arabia (Asian)	NS	PCR-RFLP	232/313	79	126	27	284	180	126	145	42	397	229	0.365	0.977
<i>Ovarian cancer</i>																
Auranen (2005a)	UK (Caucasian)	PB	TaqMan	1039/2614	427	468	144	1322	756	1046	1231	337	3336	1892	0.361	0.394
Auranen (2005b)	USA (Caucasian)	PB	TaqMan	270/344	125	114	31	364	176	130	174	40	434	254	0.369	0.110
Auranen (2005c)	Danish (Caucasian)	PB	TaqMan	361/891	144	168	49	456	266	358	394	139	1110	672	0.377	0.079
Webb (2005)	Australia (Caucasian)	HB	PCR	543/1125	229	238	76	696	390	438	538	149	1416	834	0.371	0.420
Beesley (2007a)	Australia (Caucasian)	PB	PCR-RFLP	504/972	207	223	74	637	371	370	471	131	1211	733	0.377	0.326
Beesley (2007b)	Australia (Caucasian)	PB	PCR-RFLP	731/747	291	339	101	921	541	288	351	108	927	567	0.379	0.949
Quaye (2009)	UK-USA-DK (Caucasian)	PB	Sequencing	1332/2024	545	612	175	1702	962	784	958	282	2526	1522	0.376	0.695
Gonzalez-Hormazaba (2012)	Chile (Mixed)	PB	TaqMan	87/570	45	32	10	122	52	335	209	23	879	261	0.224	0.171
Monteiro (2014)	Brazil (Mixed)	HB	PCR-RFLP	70/70	32	33	5	97	43	32	33	5	97	43	0.307	0.023
Michalska (2016)	Poland (Caucasian)	HB	PCR-RFLP	700/700	180	340	180	700	700	150	350	200	650	750	0.535	0.892

SOC: source of control; PB: population-based; HB: hospital-based; NA: not stated; PCR: polymerase chain reaction; AS-PCR: allele-specific PCR; RFLP: restriction fragment length polymorphism; MAF: minor allele frequency; HWE: Hardy-Weinberg equilibrium

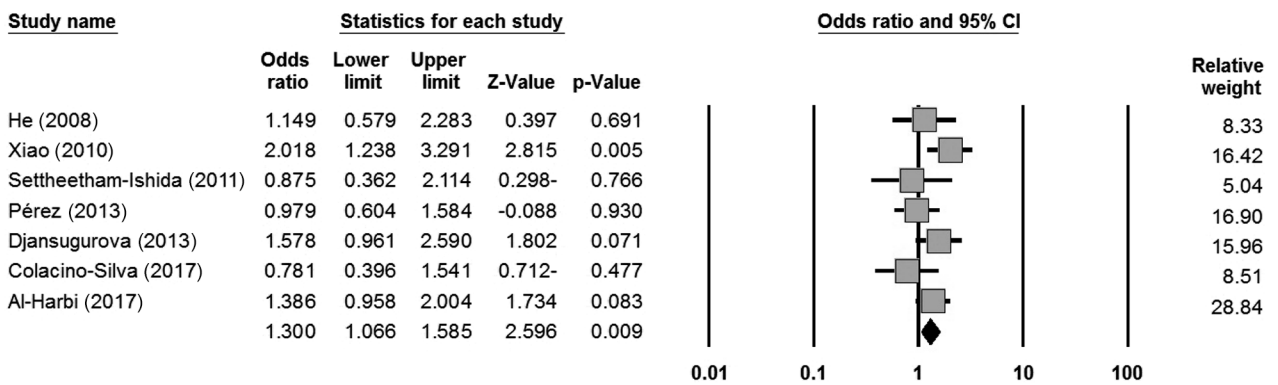
**Table II** Results of the association of XRCC3 18067 C>T polymorphism with cervical and ovarian cancers risk

Subgroup	Genetic model	Type of model	I <sup>2</sup> (%)	Heterogeneity		OR		Publication bias		
				P <sub>H</sub>	OR	95% CI	Z <sub>test</sub>	P <sub>OR</sub>	P <sub>Begg's</sub>	P <sub>Figgen's</sub>
Overall	T vs. C	Random	61.03	0.001	1.014	0.930-1.106	0.326	0.745	0.091	0.112
	TT vs. CC	Random	51.70	0.009	1.010	0.855-1.194	0.118	0.906	0.162	0.079
	TC vs. CC	Random	39.28	0.049	0.967	0.876-1.067	-0.670	0.530	0.232	0.151
	TT + TC vs. CC	Random	54.91	0.003	0.993	0.889-1.108	-0.129	0.897	0.232	0.140
	TT vs. TC+CC	Random	43.09	0.034	1.028	0.894-1.183	0.385	0.700	0.224	0.099
Cervical cancer	T vs. C	Random	73.82	0.001	1.223	0.897-1.669	1.272	0.203	1.000	0.901
	TT vs. CC	Random	68.44	0.007	1.456	0.723-2.932	1.053	0.292	0.707	0.376
	TC vs. CC	Fixed	26.84	0.224	1.300	1.066-1.585	2.596	0.009	0.548	0.242
	TT + TC vs. CC	Random	58.55	0.025	1.270	0.935-1.726	1.530	0.126	0.763	0.452
	TT vs. TC+CC	Random	64.54	0.015	1.309	0.693-2.470	0.829	0.407	0.452	0.225
Asian	T vs. C	Fixed	60.39	0.056	1.302	1.076-1.576	2.716	0.007	1.000	0.862
	TT vs. CC	Fixed	58.94	0.088	1.457	0.918-2.314	1.595	0.111	1.000	0.446
	TC vs. CC	Fixed	14.62	0.319	1.441	1.113-1.867	2.768	0.006	0.308	0.474
	TT + TC vs. CC	Fixed	36.18	0.195	1.469	1.148-1.880	3.055	0.002	0.734	0.666
	TT vs. TC+CC	Fixed	61.31	0.075	1.165	0.754-1.801	0.689	0.491	1.000	0.375
Caucasians	T vs. C	Random	91.87	0.00	1.253	0.500-3.138	0.481	0.630	NA	NA
	TT vs. CC	Random	89.45	0.002	1.484	0.188-11.730	0.374	0.708	NA	NA
	TC vs. CC	Fixed	45.46	0.176	1.234	0.873-1.743	1.193	0.233	NA	NA
	TT + TC vs. CC	Random	83.94	0.013	1.248	0.551-2.826	0.532	0.595	NA	NA
	TT vs. TC+CC	Random	88.24	0.004	1.425	0.210-9.655	0.363	0.716	NA	NA
Ovarian cancer	T vs. C	Fixed	4.19	0.402	0.956	0.910-1.003	-1.830	0.067	0.210	0.554
	TT vs. CC	Fixed	31.26	0.158	0.942	0.850-1.045	-1.130	0.259	0.591	0.313
	TC vs. CC	Fixed	0.00	0.662	0.904	0.841-0.972	-2.725	0.006	1.000	0.929
	TT + TC vs. CC	Fixed	0.00	0.504	0.914	0.853-0.979	-2.569	0.010	1.000	0.849
	TT vs. TC+CC	Fixed	25.21	0.211	1.010	0.994-1.092	-0.133	0.894	0.371	0.209
Caucasians	T vs. C	Fixed	0.00	0.763	0.948	0.902-0.996	-2.137	0.033	1.000	0.171
	TT vs. CC	Fixed	0.00	0.791	0.922	0.831-1.024	-1.514	0.130	0.901	0.445
	TC vs. CC	Fixed	0.00	0.567	0.898	0.834-0.967	-2.853	0.004	0.386	0.221
	TT + TC vs. CC	Fixed	0.00	0.638	0.905	0.844-0.970	-2.818	0.005	0.386	0.133
	TT vs. TC+CC	Fixed	0.00	0.804	0.977	0.888-1.074	-0.480	0.631	0.901	0.963

OR: odds ratio; CI: confidence interval; NA: not applicable



A



B

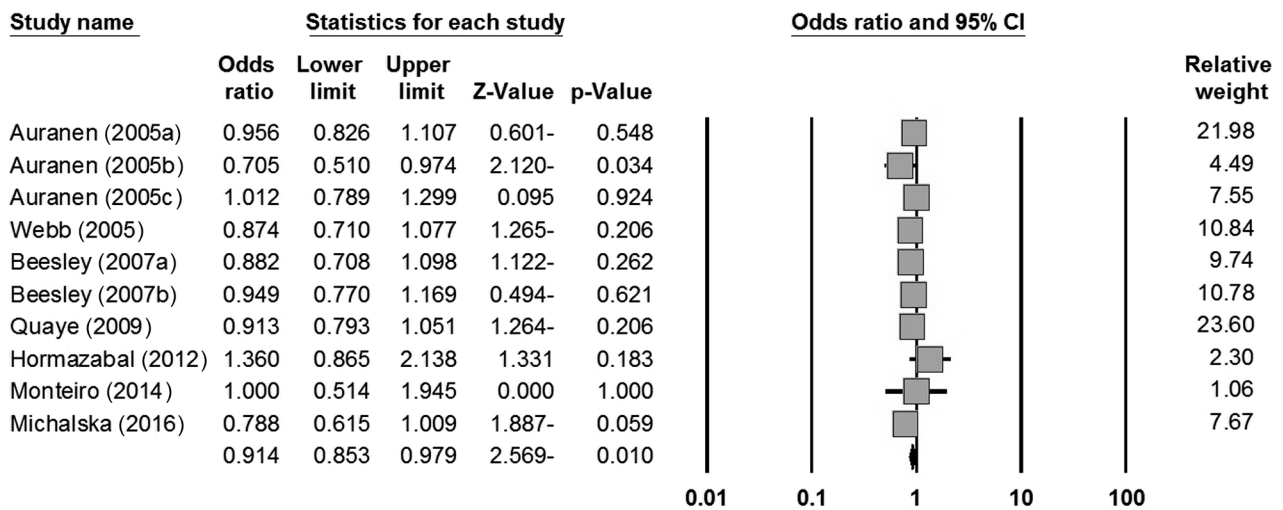


Fig. 2.

270 sources of heterogeneity across studies, we performed  
 271 subgroup analysis under all models. To explore the  
 272 sources of heterogeneity, we conducted subgroup analy-  
 273 ses by ethnicity, genotyping methods, and source of  
 274 controls. Subgroup analyses by ethnicity showed that the  
 275 heterogeneity was still significant in Caucasians popula-  
 276 tions, indicating that ethnicity was the major source that  
 277 contributed to heterogeneity for cervical cancer. In  
 278 addition, we have performed sensitivity analyses to assess  
 279 the influence of each individual study on the pooled ORs  
 280 by sequential omission of individual studies. The results  
 281 suggested that the sequential omission of individual  
 282 studies did not significantly affect the pooled ORs for  
 283 the XRCC3 18067 C>T polymorphism, the stability of  
 284 the current meta-analysis results. For ovarian cancer,  
 285 sensitivity analysis was further performed by excluded  
 286 one HWE-violating study. However, the XRCC3  
 287 18067 C>T polymorphism association with ovarian can-  
 288 cer risk was not influenced by omitting the study.

### Publication bias

289

Both Begg's funnel plot and Egger's test were performed  
 290 to assess the publication bias of literatures in all genetic  
 291 models and by ethnicity. The shape of the funnel plot did  
 292 not reveal any evidence of obvious asymmetry in overall  
 293 and by cancer type (Fig. 3). Then, we used the Egger's  
 294 test to provide statistical evidence of funnel plot symme-  
 295 try. The results still did not suggest any evidence of  
 296 publication bias in overall, by cancer type and ethnicity  
 297 (Table II).  
 298

### Discussion

299

The XRCC3 gene is one of the major genes involved in  
 300 the restoration phase of DNA damage [14]. More than  
 301 300 validated single nucleotide polymorphisms in the  
 302 XRCC3 gene were reported in the dbSNP database  
 303

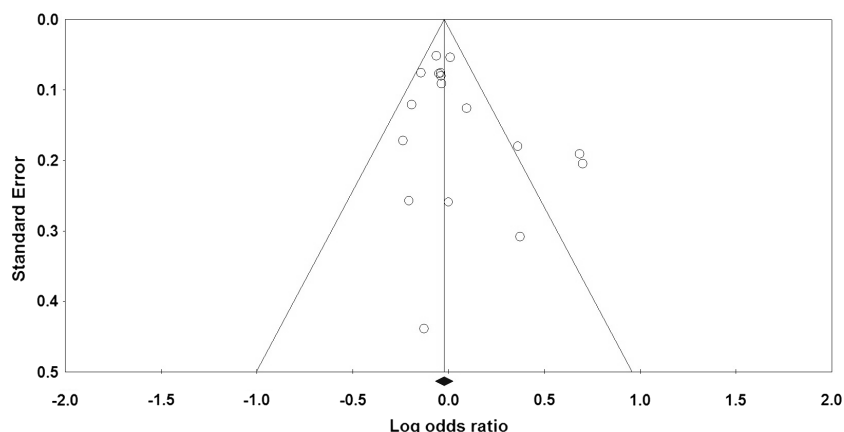


Fig. 3.

among them, 18067 C>T (rs861539) in *XRCC3* codon 241 (Thr241Met) was the most extensively studied in different malignancies [16, 17]. There is evidence that *XRCC3* 18067 C>T polymorphism is a functional variant with potential to affect the capacity of DNA repair activity [15]. The association of this polymorphism with cervical and ovarian cancer risk has been assessed in several studies, which showed inconclusive results.

In the present meta-analysis, we examined the association of *XRCC3* 18067 C>T polymorphism with cervical and ovarian cancers risk. We found that the *XRCC3* 18067 C>T polymorphism was significantly associated with ovarian cancer risk. We also observed a significant relationship between the *XRCC3* 18067 C>T polymorphism and ovarian cancer in Caucasians. However, our results were inconsistent with previous meta-analysis. Yan et al. [16] in a meta-analysis of seven studies with 3,635 cases and 5,473 controls suggested that the *XRCC3* 18067 C>T polymorphism may not be associated with ovarian cancer in all five genetic models in overall and Caucasians population. In 2013, Qin et al. [17] in a meta-analysis of five case-control studies with a total of 806 cervical cancer cases and 850 controls estimated the association between *XRCC3* 18067 C>T polymorphism and cervical cancer risk. The results showed a significant association that *XRCC3* 18067 C>T polymorphism may contribute to the susceptibility of cervical cancer only under heterozygote model. The association was further confirmed by our meta-analysis, which involved seven studies with 1,112 cases and 1,233 controls only in the heterozygote model. Moreover, the previous [16] and the current meta-analyses findings confirmed that *XRCC3* 18067 C>T polymorphism is associated with the risk of cervical cancer among Asians, but not among Caucasians, suggesting that this polymorphism may modify the risk of cervical cancer in different ethnicities. Compared to the previous meta-analyses, the included studies to the current meta-analysis are most precise and comprehensive attributing to the largest sample size and

accumulative meta-analysis method. Hence, our results are more precise and comprehensive on the association of *XRCC3* 18067 C>T polymorphism with cervical and ovarian cancers.

The heterogeneity plays an important role when performing meta-analysis and finding the source of heterogeneity is very important for the final result of meta-analysis. There were several sources bringing in heterogeneity, such as study design, age, sex distribution, sample size, genotyping methods, and ethnicity. Obviously, there was potential to moderate level heterogeneity in the current meta-analysis. Thus, we have performed meta-regression analysis to find source of heterogeneity. The heterogeneity between our studies was significantly reduced in the analysis of the cancer type and by ethnicity subgroups, indicating that the effect of *XRCC3* 18067 C>T polymorphism may be modified by cancer etiology and ethnicity backgrounds.

The main advantage of our meta-analysis that publication bias was not observed, which indicates that the whole pooled results, may be unbiased. However, several limitations in this meta-analysis should be addressed. First, the included studies only provided data toward Asians and Caucasians. The data regarding other ethnicities such as Africans were not found. Therefore, we cannot generalize these findings to every ethnic group. Second, there were only seven studies with a total of 1,112 cases and 1,233 controls that were finally included into the meta-analysis for cervical cancer. The number of included studies was relatively limited, which may increase the risk of bias in the meta-analysis, especially in the subgroup analysis by ethnicity. Thus, more studies with a larger sample size from different ethnicities should be performed in the future. Third, we have included only published studies in the meta-analysis, and non-significant or negative findings may be unpublished. Hence, any preexisting publication bias will be reflected in the findings; however, the statistical data may not show it. Fifth, the summary ORs were based on individual

382 unadjusted estimates, while a more precise analysis might  
383 be performed if detailed individual data were available,  
384 which could allow for an adjusted estimation by age,  
385 obesity, hormone replacement therapy, reproductive  
386 history and infertility, gynecologic surgery, and environ-  
387 ment factors. Lack of information for data analysis may  
388 cause serious confounding bias. Finally, gene-gene and  
389 gene-environment interactions may have influenced our  
390 findings, as ovarian and cervical cancers are mainly caused  
391 by genetic and environmental factors. However, these  
392 interactions were not tested in the current meta-analysis  
393 because of the lack of sufficient data.

394 In summary, our meta-analysis demonstrated that the  
395 XRCC3 18067 C>T polymorphism may be associated  
396 with increased risk of cervical and ovarian cancers.  
397 Moreover, the XRCC3 18067 C>T polymorphism might  
398 be a potential risk factor for cervical cancer among Asians  
399 and for ovarian cancer among Caucasians. However, to  
400 validate this association and our findings further, large  
401 and well-designed epidemiological studies are warranted.

402 \* \* \*

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405 the study, writing and editing the paper, and validation of the final  
406 version. HA, AH, and R-ST searched literature, selected study, and  
407 drafted the article. HN analyzed the data. HN and AJ contributed to  
408 interpretation of data, writing and editing of the paper, and validation of  
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