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Uncertainties in eddy covariance air–sea CO_2 flux measurements and implications for gas transfer velocity parameterisations

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Abstract. Air–sea carbon dioxide (CO_2) flux is often indirectly estimated by the bulk method using the air–sea difference in CO_2 fugacity (ΔfCO_2) and a parameterisation of the gas transfer velocity (K). Direct flux measurements by eddy covariance (EC) provide an independent reference for bulk flux estimates and are often used to study processes that drive K. However, inherent uncertainties in EC air–sea CO_2 flux measurements from ships have not been well quantified and may confound analyses of K. This paper evaluates the uncertainties in EC CO_2 fluxes from four cruises. Fluxes were measured with two state-of-the-art closed-path CO_2 analysers on two ships. The mean bias in the EC CO_2 flux is low, but the random error is relatively large over short timescales. The uncertainty (1 standard deviation) in hourly averaged EC air–sea CO_2 fluxes (cruise mean) ranges from 1.4 to 3.2 $\mathrm{mmol\,m^{-2}\,d^{-1}}$. This corresponds to a relative uncertainty of $\sim 20\,\%$ during two Arctic cruises that observed large CO_2 flux magnitude. The relative uncertainty was greater ($\sim 50\,\%$) when the CO_2 flux magnitude was small during two Atlantic cruises. Random uncertainty in the EC CO_2 flux is mostly caused by sampling error. Instrument noise is relatively unimportant. Random uncertainty in EC CO_2 fluxes can be reduced by averaging for longer. However, averaging for too long will result in the inclusion of more natural variability. Auto-covariance analysis of CO_2 fluxes suggests that the optimal timescale for averaging EC CO_2 flux measurements ranges from 1 to 3 h, which increases the mean signal-to-noise ratio of the four cruises to higher than 3. Applying an appropriate averaging timescale and suitable ΔfCO_2 threshold (20 µatm) to EC flux data enables an optimal analysis of K.

1 Introduction

Since the Industrial Revolution, atmospheric CO_2 levels have risen steeply due to human activities (Broecker and Peng, 1998). The ocean plays a key role in the global carbon cycle, having taken up roughly one-quarter of anthropogenic CO_2 emissions over the last decade (Friedlingstein et al., 2020). Accurate estimates of air–sea CO_2 flux are vital to forecast climate change and to quantify the effects of ocean CO_2 uptake on the marine biosphere.

Air–sea CO_2 flux (F, e.g. in $\mathrm{mmol\,m^{-2}\,d^{-1}}$) is typically estimated indirectly by the bulk equation:

$$F = K_{660}(Sc/660)^{-0.5}\alpha(fCO_{2w} - fCO_{2a}), \quad (1)$$

where K_{660} (in $\mathrm{cm\,h^{-1}}$) is the gas transfer velocity, usually parameterised as a function of wind speed (e.g. Nightingale et al., 2000); Sc (dimensionless) is the Schmidt number (Wanninkhof, 2014); and α ($\mathrm{mol\,L^{-1}\,atm^{-1}}$) is the solubility (Weiss, 1974). Sc is equal to 660 for CO_2 at 20 °C and 35 ‰ saltwater (Wanninkhof et al., 2009). fCO_{2w} and fCO_{2a} are the CO_2 fugacity (in µatm) at the sea surface and in the overlying atmosphere, respectively, with $fCO_{2w} - fCO_{2a}$ the air–sea CO_2 fugacity difference (ΔfCO_2). Uncertainties in the K_{660} parameterisation and limited coverage of fCO_{2w} measurements result in considerable uncertainties in global bulk flux estimates (Takahashi et al., 2009; Woolf et al., 2019).

Eddy covariance (EC) is the most direct method for measuring the air–sea CO_2 flux F:

$$F = \rho\overline{w'c'}, \quad (2)$$

where ρ is the mean mole density of dry air (e.g. in $\mathrm{mol\,m^{-3}}$). The dry CO_2 mixing ratio c (in ppm or $\mathrm{\mu mol\,mol^{-1}}$) is mea-

sured by a fast-response gas analyser, and the vertical wind velocity w (in $m\,s^{-1}$) is often measured by a sonic anemometer. The prime denotes the fluctuations from the mean, while the overbar indicates time average. Equation (2) does not rely on ΔfCO_2 measurements, nor empirical parameters and assumptions of the gas properties (Wanninkhof, 2014). EC flux measurements can therefore be considered useful as an independent reference for bulk air–sea CO_2 flux estimates. Furthermore, the typical temporal and spatial scales of EC flux measurements are ca. hourly and 1–10 km^2. These scales are much smaller than the temporal and spatial scales of alternative techniques for measuring gas transfer, e.g. by dual tracer methods (daily and 1000 km^2) (Nightingale et al., 2000; Ho et al., 2006). EC measurements are thus potentially better suited to capture variations in gas exchange due to small-scale processes at the air–sea interface (Garbe et al., 2014).

The EC CO_2 flux method has developed and improved over time. Before 1990, EC was successfully used to measure air–sea momentum and heat fluxes. EC air–sea CO_2 flux measurements made during those times were unreasonably high (Jones and Smith, 1977; Wesely et al., 1982; Smith and Jones, 1985; Broecker et al., 1986). After 1990, with the development of the infrared gas analyser, EC became routinely used for terrestrial carbon cycle research (Baldocchi et al., 2001). Development of the EC method was accompanied by improvements in the flux uncertainty analysis, which was generally based on momentum, heat and land–atmosphere gas flux measurements (Lenschow and Kristensen, 1985; Businger, 1986; Lenschow et al., 1994; Wienhold et al., 1995; Mahrt, 1998; Finkelstein and Sims, 2001; Loescher et al., 2006; Rannik et al., 2009, 2016; Billesbach, 2011; Mauder et al., 2013; Langford et al., 2015; Post et al., 2015).

In the late 1990s, the advancement in motion correction of wind measurements (Edson et al., 1998; Yelland et al., 1998) facilitated ship-based EC CO_2 flux measurements from a moving platform (McGillis et al., 2001; 2004). After 2000, a commercial open-path infrared gas analyser LI-7500 (Li-COR Inc. USA) became widely used for air–sea CO_2 flux measurements (Weiss et al., 2007; Kondo and Tsukamoto, 2007; Prytherch et al., 2010; Edson et al., 2011; Else et al., 2011; Lauvset et al., 2011). The LI-7500 generated extremely large and highly variable CO_2 fluxes in comparison to expected fluxes (Kondo and Tsukamoto, 2007; Prytherch et al., 2010; Edson et al., 2011; Else et al., 2011; Lauvset et al., 2011). This problem is generally considered to be an artefact caused by water vapour cross-sensitivity (Kohsiek, 2000; Prytherch et al., 2010; Edson et al., 2011; Landwehr et al., 2014). Mathematical corrections proposed to address this artefact (Edson et al., 2011; Prytherch et al., 2010) were later shown to be unsatisfactory (Else et al., 2011; Ikawa et al., 2013; Blomquist et al., 2014; Tsukamoto et al., 2014) or incorrect (Landwehr et al., 2014).

The most reliable method for measuring EC air–sea CO_2 fluxes involves the physical removal of water vapour fluctuations from the sampled air. The simplest approach is to combine a closed-path gas analyser with a physical dryer to eliminate most of the water vapour fluctuation (Miller et al., 2010; Blomquist et al., 2014; Landwehr et al., 2014; Yang et al., 2016; Nilsson et al., 2018). The tuneable-diode-laser-based cavity ring-down spectrometer (CRDS) made by Picarro Inc. (Santa Clara, California, USA) is the most precise closed-path analyser currently available (Blomquist et al., 2014). The closed-path infrared gas analyser LI-7200 (LI-COR Biosciences, Lincoln, Nebraska, USA) is another popular choice.

The advancements in instrumentation and in motion correction methods have significantly improved the quality of air–sea EC CO_2 flux observations, but, despite these changes, the flux uncertainties have not been well quantified. The aims of this study are to (1) analyse uncertainties in EC air–sea CO_2 flux measurements; (2) propose practical methods to reduce the systematic and random flux uncertainty; and (3) investigate how the EC flux uncertainty influences our ability to estimate and parameterise K_{660}.

2 Experiment and methods

2.1 Instrumental setup

The basic information of four cruises is summarised in Table 1. Appendix A shows the four cruise tracks (Figs. A1 and A2). Data from the Atlantic cruises (AMT28 and AMT29) are limited to 3° N–20° S in order to focus specifically on the performance of two different gas analysers in the same region with low flux signal (tropical zone).

The CO_2 flux and data logging systems installed on the *JCR* and *Discovery* were operated autonomously. The EC systems were approximately 20 m a.m.s.l. on both ships (at the top of the foremasts; Fig. 1) to minimise flow distortion and exposure to sea spray. Computational fluid dynamics (CFD) simulation indicates that the airflow distortion at the top of the *JCR* foremast is small ($\sim 1\,\%$ of the free stream wind speed when the ship is head to wind; Moat and Yelland, 2015). The hull structure of RRS *Discovery* is nearly identical to that of RRS *James Cook*. CFD simulation of the *James Cook* indicates that the airflow at the top foremast is distorted by $\sim 2\,\%$ for bow-on flows (Moat et al., 2006). The deflection of the streamline from horizontal and effects on the vertical wind component is accounted for by the double rotation (motion correction processes; see Sect. 2.2) prior to the EC flux calculation for both ships.

The EC system on the *JCR* consists of a three-dimensional sonic anemometer (Metek Inc., Sonic-3 Scientific), a motion sensor (initially Systron Donner Motionpak II, which compared favourably with and was then replaced by a Life Performance-Research LPMS-RS232AL2 in April 2019), and a Picarro G2311-f gas analyser. All instruments sampled at a frequency of 10 Hz or greater, and the data were

Table 1. Basic information for all four cruises on the RRS *James Clark Ross* (*JCR*) and RRS *Discovery* that measured air–sea EC CO_2 fluxes.

Cruise	JR18006	JR18007	AMT28	AMT29
Data period	30 Jun–1 Aug 2019	5 Aug–29 Sep 2019	9–16 Oct 2018	4–11 Nov 2019
Visited region	Arctic Ocean (Barents Sea)	Arctic Ocean (Fram Strait)	Tropical Atlantic Ocean	Tropical Atlantic Ocean
Research vessel	*JCR*	*JCR*	*JCR*	*Discovery*
Gas analyser	Picarro G2311-f	Picarro G2311-f	Picarro G2311-f	LI-7200

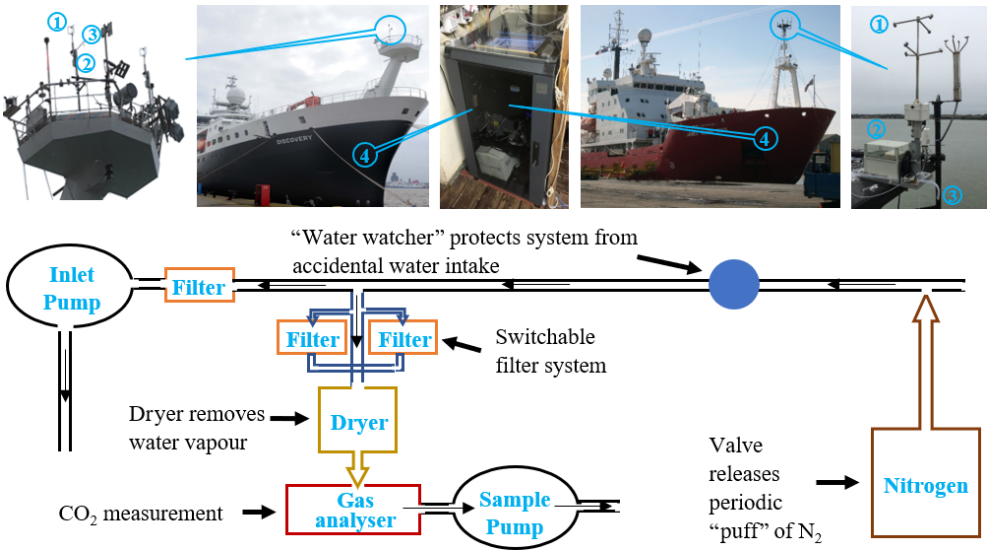

Figure 1. EC system (upper panel) and a diagram of the system setup (bottom panel). EC instruments: (1) sonic anemometer, (2) motion sensor, (3) air sample inlet for gas analyser, (4) data logger/gas analyser. Arctic and Atlantic data from 2018 were collected on the RRS *James Clark Ross* (*JCR*; upper right) using a Picarro G2311-f, and Atlantic data from 2019 were collected using a LI-7200 on the RRS *Discovery* (upper left).

logged at 10 Hz with a data logger (CR6, Campbell Scientific, Inc.), similar to the setup by Butterworth and Miller (2016). Air is pulled through a long tube (30 m, inner diameter 0.95 cm, Reynolds number 5957) with a dry vane pump at a flow rate of $\sim 40\,\mathrm{L\,min^{-1}}$ (Gast 1023 series). The Picarro gas analyser subsamples from this tube through a particle filter (Swagelok 2 µm) and a dryer (Nafion PD-200T-24M) at a flow of $\sim 5\,\mathrm{L\,min^{-1}}$ (Fig. 1). The dryer is set up in the "re-flux" configuration and uses the lower pressure Picarro exhaust to dry the sample air. This method removes $\sim 80\,\%$ of the water vapour and essentially all of the humidity fluctuations (Yang et al., 2016). The Picarro internal calculation accounts for the detected residual water vapour and yields a dry CO_2 mixing ratio that is used in the flux calculations. A valve controlled by the Picarro instrument injects a "puff" of nitrogen (N_2) into the tip of the inlet tube for 30 s every 6 h. This enables estimates of the time delay and high-frequency signal attenuation (Sect. 2.2).

The EC system on RRS *Discovery* consists of a Gill R3-50 sonic anemometer, a LPMS motion sensor package, and a LI-7200 gas analyser. The LI-7200 gas analyser was mounted within the enclosed staircase, directly underneath the meteorological platform and close to the inlet (inlet length 7.5 m, inner diameter 0.95 cm, Reynolds number 1042). A single pump (Gast 1023) was sufficient to pull air through a particle filter (Swagelok 2 µm), a dryer (Nafion PD-200T-24M), and the LI-7200 at a flow of $\sim 7\,\mathrm{L\,min^{-1}}$. There was no N_2 puff system setup on *Discovery*, but equivalent lab tests confirmed that the delay time was less than on the *JCR* because of the shorter inlet line. The dryer on the *Discovery* is set up in the same re-flux configuration as the *JCR* and uses the lower pressure at the LI-7200 exhaust (limited by an additional 0.08 cm diameter critical orifice) to dry the sample air. This setup removes $\sim 60\,\%$–$70\,\%$ of the water vapour and essentially all of the humidity fluctuations. The dry CO_2 mixing ratio, computed by accounting for the LI-7200 temperature, pressure, and residual water vapour measurements, is used in the flux calculations.

2.2 Flux processing

The EC air–sea CO_2 flux calculation steps using the raw data are outlined with a flow chart (Fig. 2) and detailed below. The

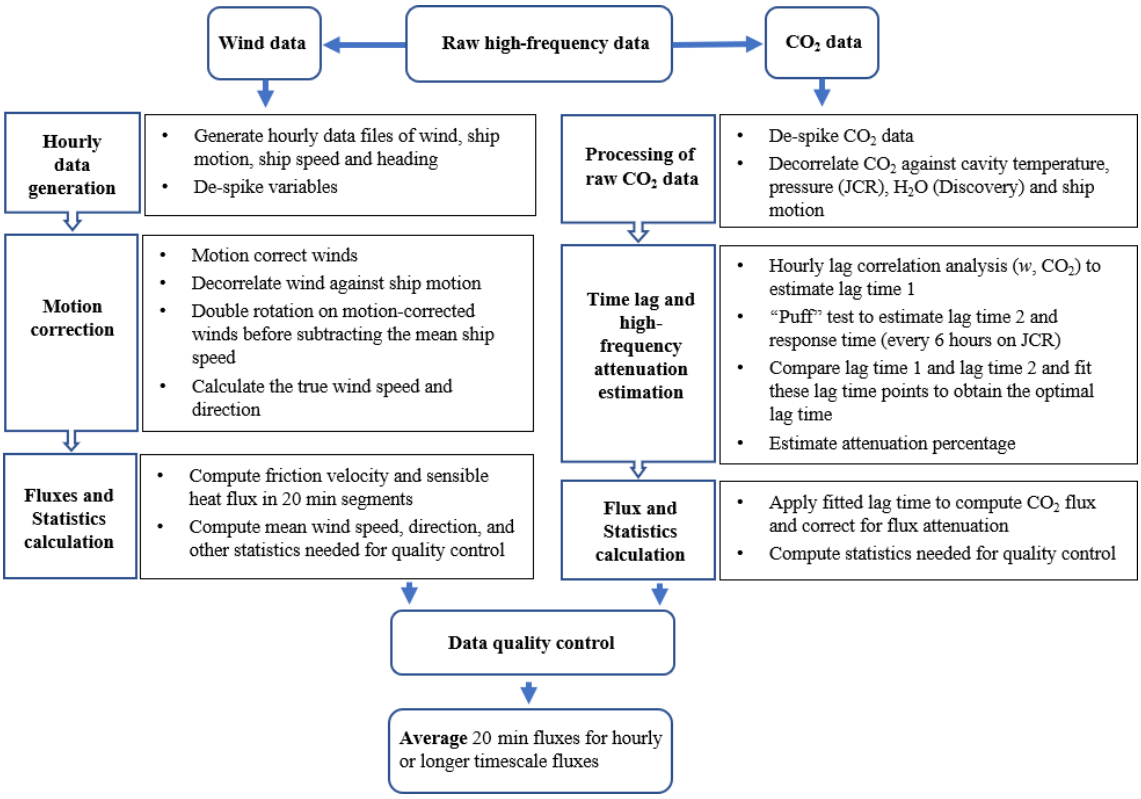

Figure 2. Flow chart of EC data processing. The raw high-frequency (10 Hz) wind and CO_2 data were initially processed separately and then combined to calculate fluxes. CO_2 fluxes were filtered by a series of data quality control criteria. The 20 min flux intervals were averaged to longer timescales (hourly or more). The data processing is detailed in the text.

raw high-frequency wind and CO_2 data are processed first, yielding fluxes in 20 min averaging time interval and related statistics. These statistics are then used for quality control of the fluxes. Further averaging of the quality-controlled 20 min fluxes to hourly or longer timescales is then used to reduce random error (Sect. 4.1). Linear detrending was used to identify the turbulent fluctuations (i.e. w' and c') throughout the analyses.

To correct the wind data for ship motion, we first generated hourly data files containing the measurements from the sonic anemometer (three-dimensional wind speed components: u, v, and w and sonic temperature T_s), motion sensor (three axis accelerations: accel_x, accel_y, accel_z; and rotation angles: rot_x, rot_y, rot_z), ship heading over ground (HDG; from the gyro compass), and ship speed over ground (SOG; from Global Position System). Spikes larger than 4 standard deviations (SDs) from the median were removed. Secondly, a complementary filtering method using Euler angles (see Edson et al., 1998) was applied to the hourly data files to remove apparent winds generated by the ship movements. The motion-corrected winds were further decorrelated against ship motion to remove any residual motion-sensitivity (Miller et al., 2010; Yang et al., 2013). The motion-corrected winds were double-rotated to account for the wind streamline over the ship, yielding the vertical wind velocity (w) required in Eq. (2). Inspection of frequency spectra showed that the spectral peak at the ship motion frequencies (approximately 0.1–0.3 Hz) had disappeared after the motion correction (Fig. S1 in the Supplement). This indicates that the majority of ship motion had been removed from the measured wind speed. The last step in the wind data processing was the calculation of 20 min average friction velocity, sensible heat flux, and other key variables used for data quality control (Table S1 in the Supplement).

The CO_2 data were de-spiked (by removing values > 4 SDs from the median). The Picarro CO_2 mixing ratio was further decorrelated against analyser cell pressure and temperature to remove CO_2 variations due to the ship's motion. The LI-7200 CO_2 mixing ratio was further decorrelated against the LI-7200 H_2O mixing ratio and temperature to remove residual air density fluctuations, following Landwehr et al. (2018). CO_2 data were also decorrelated against the ship's heave and accelerations because these can produce spurious CO_2 variability (Miller et al., 2010; Blomquist et al., 2014).

A lag between CO_2 data acquisition and the wind data is created because of the time taken for sample air to travel through the inlet tube. On the *JCR*, we use the puff system

where the lag time is the time difference between the N_2 puff start (when the on/off valve is switched) and the time when the diluted signal is sensed by the gas analyser. The lag time can also be estimated by the maximum covariance method, calculated by shifting the time base of the CO_2 signal and finding the shift that achieves maximum covariance between the vertical wind velocity (w) signal and the shifted CO_2 signal. The lag times estimated by the maximum covariance method agree well with the estimates of the puff procedure (Fig. S2 in the Supplement). These estimates indicate a lag time of 3.3–3.4 s for the Arctic cruises and 3.3 s for cruise AMT28 on the *JCR*. The lag time on *Discovery* (AMT29) estimated by the maximum covariance method was 2.6 s, consistent with laboratory test results prior to the cruise.

The inlet tube, particle filter, and dryer cause high-frequency CO_2 flux signal attenuation. The N_2 puff was also used to assess the response time by considering the e-folding time in the CO_2 signal change (similar approaches have been used by Bariteau et al., 2010; Blomquist et al., 2014, Bell et al., 2015). The response time is 0.35 s for the EC system on *JCR* and 0.25 s for the EC system on *Discovery* (estimated in the laboratory prior to cruise). These response times were combined with the relative wind speed-dependent, theoretical shapes of the cospectra (Kaimal et al., 1972) to estimate the percentage flux loss due to the inlet attenuation (Yang et al., 2013). The mean attenuation percentage is less than 10 %, with a relative wind speed dependence (Fig. S3 in the Supplement). The attenuation percentage value was applied to the computed flux to compensate for the flux loss due to the high-frequency signal attenuation. Finally, horizontal CO_2 fluxes and other statistics such as CO_2 range and CO_2 trend were computed for quality control purposes (Table S1, Supplement).

The computed 20 min fluxes were filtered for non-ideal ship manoeuvres or violations of the homogeneity/stationary requirement of EC (see Supplement for the quality control criteria).

2.3 Uncertainty analysis methods

2.3.1 Uncertainty components

Uncertainty contains two components: systematic error (δF_S) and random error (δF_R). According to propagation of uncertainty theory (JCGM, 2008), the total uncertainty in EC CO_2 fluxes (from random and systematic errors) can be expressed as

$$\delta F = \sqrt{\delta F_R^2 + \delta F_S^2}. \tag{3}$$

Systematic errors (Sect. 2.3.2) will cause bias in the flux. They thus should be eliminated/minimised with the appropriate system setup and, if needed, effective numerical corrections. Random error results in imprecision (but not bias) and can be reduced by averaging repeated measurements (Sect. 2.3.3). Errors due to insufficient sampling and instrument noise are generally considered most important in EC flux measurements (Lenschow and Kristensen, 1985; Businger, 1986; Mauder et al., 2013; Rannik et al., 2016).

Sampling error is an inherent issue for EC flux measurements and is typically the main source of the CO_2 flux uncertainty (Mauder et al., 2013). The sampling error is caused by the difference between the ensemble average and the time average. The calculation of EC flux (Eq. 2) requires the separation between the mean and fluctuating components, which can be represented fully for CO_2 mixing ratio c as

$$c(x,t) = \overline{c}(x,t) + c'(x,t). \tag{4}$$

The mean component $\overline{c}$ represents ensemble average over time (t) and space (x) and does not contribute to the flux. The time average of a stationary turbulent signal and space average of a homogenous turbulent signal theoretically converge on the ensemble average when the averaging time approaches infinity, i.e. $T \to \infty$ (Wyngaard, 2010). In practice, Reynolds averaging over a much shorter time interval (10 min to an hour) is typically used for EC flux measurements from a fixed point or from a slow-moving platform such as a ship. This is because the atmospheric boundary layer is only quasi-stationary for a few hours. Non-stationarity (e.g. diurnal variability and synoptic conditions) is an inherent property of the atmospheric boundary layer (Wyngaard, 2010). EC flux observations thus inevitably contain some random error due to insufficient sampling time, and this error is greater at shorter averaging times.

Random error due to instrument noise comes mainly from the white noise of the gas analyser, as the noise from the sonic anemometer is relatively unimportant (Blomquist et al., 2010; Fairall et al., 2000; Mauder et al., 2013). Blomquist et al. (2014) show "pink" noise with a weak spectral slope for their CRDS gas analyser (G1301-f), but the gas analysers on *JCR* (G2311-f) and *Discovery* (LI-7200) demonstrate white noise with a constant variance at high frequency (Fig. B2, Appendix B).

2.3.2 Systematic error

Table 2 details the measures taken during instrument setup and data processing that help eliminate most sources of systematic error in EC CO_2 fluxes.

In addition to bias sources related to the instrument setup (Table 2), insufficient sampling time (an inherent issue of EC fluxes) may also generate a systematic error. We use a theoretical method to estimate this systematic error in EC CO_2 flux (Lenschow et al., 1994):

$$|\delta F_S| \le 2\sigma_w \sigma_{c_a} \frac{\sqrt{\tau_w \tau_c}}{T}, \tag{5}$$

where σ_w (m s^{-1}) and σ_{c_a} (ppm) are the standard deviations of the vertical wind velocity and the CO_2 mixing ratio due to

Table 2. Potential sources of bias in our EC air–sea CO_2 flux measurements and the methods used to minimise them.

Potential source of bias	Methods used to minimise the bias	Flux uncertainty
$\delta F_{S,1}$ Water vapour cross-sensitivity	Closed-path gas analyser with a dryer removes essentially all of the water vapour fluctuation (Blomquist et al., 2014; Yang et al., 2016). The residual H_2O signal is measured by the gas analyser and used in the calculation of the dry CO_2 mixing ratio, which removes water cross-sensitivity.	Negligible
$\delta F_{S,2}$ Ship motion	Flux uncertainty from an earlier version of the motion correction procedure (less rigorous than the one used by ourselves) is estimated to be 10 %–20 % (Edson et al., 1998). The more recently adopted decorrelation of vertical winds and CO_2 against platform motion (Miller et al., 2010; Yang et al., 2013) reduces this uncertainty. Flügge et al. (2016) compare EC momentum fluxes measured from a moving platform (buoy) with fluxes measured from a nearby fixed tower. Flux estimates from these two platforms agree well (relative flux bias due to the motion correction ≤ 6 %).	≤ 6 %
$\delta F_{S,3}$ Airflow distortion	The EC flux system is deployed as far forward and as high as possible on the ship (top of the foremast), which minimises the impacts of flow distortion. Subsequent distortion correction using the CFD simulation (Moat et al., 2006; Moat and Yelland, 2015) along with a relative wind direction restriction further reduces the impact of flow distortion on the fluxes. Measured EC friction velocities and friction velocities from the COARE3.5 model (Edson et al., 2013) agree well (e.g. $R^2 = 0.95$, slope $= 0.97$) for data collected during cruise JR18006. Good comparison between observed and COARE3.5 friction velocity estimates indicates that we have fully accounted for flow distortion effects.	Negligible
$\delta F_{S,4}$ Inlet effects (high-frequency flux attenuation and CO_2 sampling delay)	High-frequency flux signal attenuation (in the inlet tube, particle filter, and dryer) is evaluated by the CO_2 signal response to a puff of N_2 gas. Flux attenuation is calculated from the "inlet puff" response and applied as a correction (< 10 %; see Sect. 2.2). The uncertainty in the attenuation correction is about 1 % for unstable/neutral atmospheric conditions, which is generally the case over the ocean (e.g. 93 % of the time for the Atlantic cruises, 80 % of the time for the Arctic cruises). During stable conditions, the attenuation correction is larger (Landwehr et al., 2018), and the uncertainty is also greater (~ 20 %). The lag time adjustment prior to the flux calculation aligns the CO_2 and wind signals. Two methods are used to estimate the optimal lag time: puff injection and maximum covariance. The two lag estimates are in good agreement (Sect. 2.2). Random adjustment of ± 0.2 s (1σ of the puff test result) to the optimal lag time impacts the CO_2 flux by < 1 %.	< 2 % for vast majority of the cruises
$\delta F_{S,5}$ Spatial separation between the sonic anemometer and the gas inlet	The CO_2 inlet is ~ 70 cm directly below the centre volume of the sonic anemometer. This distance is small relative to the size of the dominant flux-carrying eddies encountered by the EC measurement system height above sea level. The excellent agreement between the lag time determined by the puff system and by the optimal covariance method further confirms that the distance between the CO_2 inlet and anemometer is sufficiently small.	Negligible
$\delta F_{S,6}$ Imperfect calibration of the sensors	The potential flux bias resulting from instrument calibration (gas analyser, anemometer and meteorological sensors required to calculate air density: air temperature, relative humidity and pressure) is up to 4 % for the *JCR* setup. The largest instrument calibration uncertainty derives from the wind sensor accuracy ($\pm 0.15\,m\,s^{-1}$ at $4\,m\,s^{-1}$ winds according to the Metek uSonic instrument specification). This bias is even lower (< 2 %) for the *Discovery* setup because the Gill R3 sonic anemometer is more accurate.	≤ 4 %
Propagated bias	Estimated from the individual bias estimates above ($\delta F_{S,1}$, $\delta F_{S,2}$, etc.) using $\delta F_S = \sqrt{\sum_1^n \delta F_{S,n}^2}$	< 7.5 %

atmospheric processes, respectively. T is the averaging time interval (s), and τ_w and τ_c are integral timescales (s) for vertical wind velocity and CO_2 signal, respectively. The definition and estimation of the integral timescale are shown in Appendix B. The sign of δF_S could be positive or negative (i.e. under- or overestimation) because of the poor statistics in capturing low-frequency eddies within the flux averaging period (Lenschow et al., 1993). The mean hourly relative systematic error due to insufficient sampling time for four cruises estimated by Eq. (5) is $< 5\,\%$. According to propagation of uncertainty theory (JCGM, 2008), the total systematic error is less than $9\,\%$ ($= \sqrt{7.5\,\%^2 + 5\,\%^2}$).

2.3.3 Random error

Five approaches used to estimate the total random error (A–C) and the random error component due to instrument noise (C–E) in EC CO_2 fluxes are discussed below. The random error assessments are empirical (A and D) or theoretical (B, C, and E).

A. An empirical approach to estimate total random error involves shifting the w data relative to the CO_2 data (or vice versa) by a large, unrealistic time shift and then computing the "null fluxes" from the time-desynchronized CO_2 and w time series (Rannik et al., 2016). The shift removes any real correlation between CO_2 and w due to vertical exchange. The standard deviation of the resultant null fluxes represents the random flux uncertainty (Wienhold et al., 1995). We applied a series of time shifts of ~ 20–$60 \cdot \tau_w$ (i.e. using time shifts ranging from -300 to -100 and 100 to 300 s; Rannik et al., 2016). This empirical estimation of total random flux uncertainty will hereafter be referred to as $\delta F_{R,Wienhold}$.

B. Lenschow and Kristensen (1985) derived a rigorous theoretical equation for total random error estimation, which contains both the auto-covariance and cross-covariance functions. The theoretical equation has been numerically approximated by Finkelstein and Sims (2001):

$$\delta F_{R,Finkelstein} = \left\{ \frac{1}{n} \left[\sum_{p=-m}^{m} r_{ww}(p) r_{cc}(p) + \sum_{p=-m}^{m} r_{wc}(p) r_{cw}(p) \right] \right\}^{1/2}, \quad (6)$$

where n is the number of data points within an averaging time interval, and p is the number of shifting points. The maximum shifting point m can be chosen subjectively ($< n$). We found that the random error for m between 1000 and 2000 data points was similar, so for this study we use $m = 1500$ (150 s shift time). The first term in the brackets represents the auto-covariance component, and the second term is the cross-covariance component. r_{ww} and r_{cc} are the auto-covariance functions for vertical wind velocity (w) and CO_2 mixing ratio (c), respectively. r_{wc} and r_{cw} are the cross-covariance functions for w and c. Here r_{wc} represents shifting w data relative to CO_2 data, while r_{cw} represents shifting CO_2 data relative to w data.

C. Blomquist et al. (2010) attributed the sources of CO_2 variance σ_c^2 to atmospheric processes ($\sigma_{c_a}^2$) and white noise ($\sigma_{c_n}^2$). The sources of variance are considered to be independent of each other, and the sonic anemometer is assumed to be relatively noise-free. According to propagation of uncertainty theory (JCGM, 2008), the total random flux error can be defined as

$$\delta F_{R,Blomquist} \leq \frac{a\sigma_w}{\sqrt{T}} \left(\sigma_{c_a}^2 \tau_{wc} + \sigma_{c_n}^2 \tau_{c_n} \right)^{1/2}, \quad (7)$$

where the constant a varies from $\sqrt{2}$ to 2, depending on the relationship between the covariance of the two variables (w and CO_2) and the product of their autocorrelations (Lenschow and Kristensen, 1985). Here, τ_{wc} is equal to the shorter of τ_w and τ_c, which is typically τ_w (Blomquist et al., 2010), and τ_{c_n} is the integral timescale of white noise in the CO_2 signal. The CO_2 variance due to atmospheric processes ($\sigma_{c_a}^2$) includes two components: variance due to vertical flux (i.e. air–sea CO_2 flux), $\sigma_{c_{av}}^2$, and variance due to other atmospheric processes, $\sigma_{c_{ao}}^2$ (Fairall et al., 2000). The variance in CO_2 due to vertical flux ($\sigma_{c_{av}}^2$) depends on atmospheric stability. $\sigma_{c_{av}}^2$ can be estimated with Monin–Obukhov similarity theory (Blomquist et al., 2010, 2014; Fairall et al., 2000):

$$\sigma_{c_{av}}^2 = \left[3 \frac{\overline{w'c'}}{u_*} f_c(z/L) \right]^2, \quad (8)$$

where u_* is the friction velocity ($\mathrm{m\,s^{-1}}$), and the similarity function (f_c) depends on the stability parameter z/L, where z is the observational height (m), and L is the Obukhov length (m). The expression of f_c can be found in Blomquist et al. (2010).

Equation (7) can be used to assess the random error due to instrument noise by setting $\sigma_{c_a}^2 = 0$, referred to hereafter as $\delta F_{RN,Blomquist}$. We use the CO_2 variance spectra to directly estimate the white noise term $\sigma_{c_n}^2 \tau_{c_n}$ in Eq. (7). The variance is fairly constant at high frequency (1–5 Hz; Fig. B2, Appendix B), which is often referred to as band-limited white noise. The relationship between $\sigma_{c_n}^2 \tau_{c_n}$ and the band-limited noise spectral value φ_{c_n} is expressed in Blomquist et al. (2010) as

$$\sigma_{c_n}^2 \tau_{c_n} = \frac{\varphi_{c_n}}{4}. \quad (9)$$

D. Billesbach (2011) developed an empirical method to estimate the random error due to instrument noise alone (referred to as $\Delta F_{\mathrm{RN,Billesbach}}$). This involves random shuffling of the CO_2 time series within an averaging interval and then calculating the covariance of w and CO_2. The correlation between w and CO_2 is minimised by the shuffling, and any remaining correlation between w and CO_2 is due to the unintentional correlations contributed by instrument noise.

E. Mauder et al. (2013) describe another theoretical approach to estimate the random flux error due to instrument noise:

$$\delta F_{\mathrm{RN,Mauder}} = \frac{\sigma_{\mathrm{w}}\sigma_{\mathrm{c_n}}}{\sqrt{n}}. \tag{10}$$

White noise correlates with itself but is uncorrelated with atmospheric turbulence. Thus, the white-noise-induced CO_2 variance ($\sigma_{\mathrm{c_n}}$) only contributes to the total variance. The value of $\sigma_{\mathrm{c_n}}$ can be estimated from the difference between the zero-shift auto-covariance value (CO_2 variance σ_{c}^2) and the noise-free variance extrapolated to a time shift of zero (Lenschow et al., 2000):

$$\sigma_{\mathrm{c_n}}^2 = \sigma_{\mathrm{c}}^2 - \sigma^2(t \to 0), \tag{11}$$

where $\sigma^2(t \to 0)$ represents the extrapolation of auto-covariance to a zero shift, which is considered equal to variance due to atmospheric processes ($\sigma_{\mathrm{c_a}}^2$). Figure 3 shows the normalised auto-covariance function curves of w and CO_2 as measured by the Picarro G2311-f and the LI-7200. There is a sharp decrease in the CO_2 auto-covariance when shifting from 0 s shift to 0.1 s shift for both the Picarro G2311-f and LI-7200 gas analyser. The same sharp decrease is not seen in the vertical wind velocity (w) auto-covariance. The relative difference in the change in normalised auto-covariance shows that white noise makes a much larger relative contribution to the CO_2 variance than to the vertical wind velocity variance.

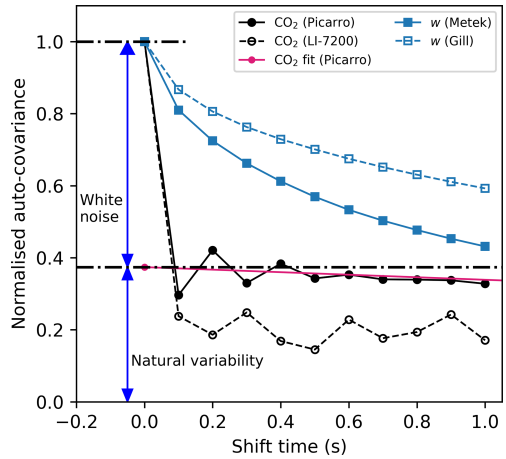

Figure 3. Mean normalised auto-covariance functions of CO_2 and vertical wind velocity (w) by four different instruments. The magenta line represents a fit to the noise-free auto-covariance function of CO_2 (measured by Picarro) extrapolated back to a zero time shift. An example of the white noise and natural variability contributions to the total CO_2 (measured by Picarro) variance is indicated by two blue arrows. The sharp decrease of the CO_2 auto-covariance between the zero shift and the initial 0.1 s shift corresponds to the large contribution of white noise from the gas analysers. The LI-7200 is the noisier instrument. The noise contributions from the anemometer are relatively small (< 10 %).

3 Results

Measurements from AMT28 and AMT29 set the scene for our uncertainty analysis. These two Atlantic cruises transited across the same tropical region (Fig. A2, Appendix A) in October 2018 and September 2019 with different eddy covariance systems (Sect. 2.1). AMT28 and AMT29 show broadly similar latitudinal patterns (Fig. 4a). An obvious question of interest is whether the measured fluxes were the same for the 2 years. To answer this question, the measurement uncertainties must be quantified. The total random uncertainties in CO_2 flux ($\delta F_{\mathrm{R,Finkelstein}}$) are comparable for the two cruises, even though the random error component due to instrument noise ($\delta F_{\mathrm{RN,Mauder}}$) is about 3 times higher during AMT29 using LI-7200 than during AMT28 using Picarro G2311-f (Fig. 4b; Fig. D1, Appendix D). The similar total random uncertainty in the AMT28 and AMT29 fluxes shows that both gas analysers are equally suitable for air–sea EC CO_2 flux measurements. The variance budgets of atmospheric CO_2 mixing ratio (used to estimate random flux uncertainty; see Sect. 3.1) are shown in Fig. 4c. Total variance in CO_2 mixing ratio is dominated by instrument noise on both cruises. CO_2 mixing ratio variance (total and instrument noise) was substantially higher during AMT29.

3.1 Random uncertainty

Theoretical derivation of flux uncertainty ($\delta F_{\mathrm{RN,Blomquist}}$, Eq. 7) requires knowledge of the contributions to CO_2 mixing ratio variance. Total CO_2 variance is made up of instrument noise ($\sigma_{\mathrm{c_n}}^2$) and atmospheric processes ($\sigma_{\mathrm{c_a}}^2$). Atmospheric processes include vertical flux ($\sigma_{\mathrm{c_{av}}}^2$) and other atmospheric processes ($\sigma_{\mathrm{c_{ao}}}^2$). The variance budgets of CO_2 mixing ratio for the four cruises are listed in Table 3. Atmospheric processes contribute a larger CO_2 variance in the Arctic (where flux magnitudes are greater) compared to the Atlantic. Vertical flux accounts for ∼ 10 % of the variance in CO_2 mixing ratio in the Arctic and ∼ 1 % of the CO_2 variance in the Atlantic. Previous results demonstrate that horizontal transport is a major source of $\sigma_{\mathrm{c_{ao}}}^2$ for long-lived greenhouse gases (Blomquist et al., 2012). Small changes in CO_2

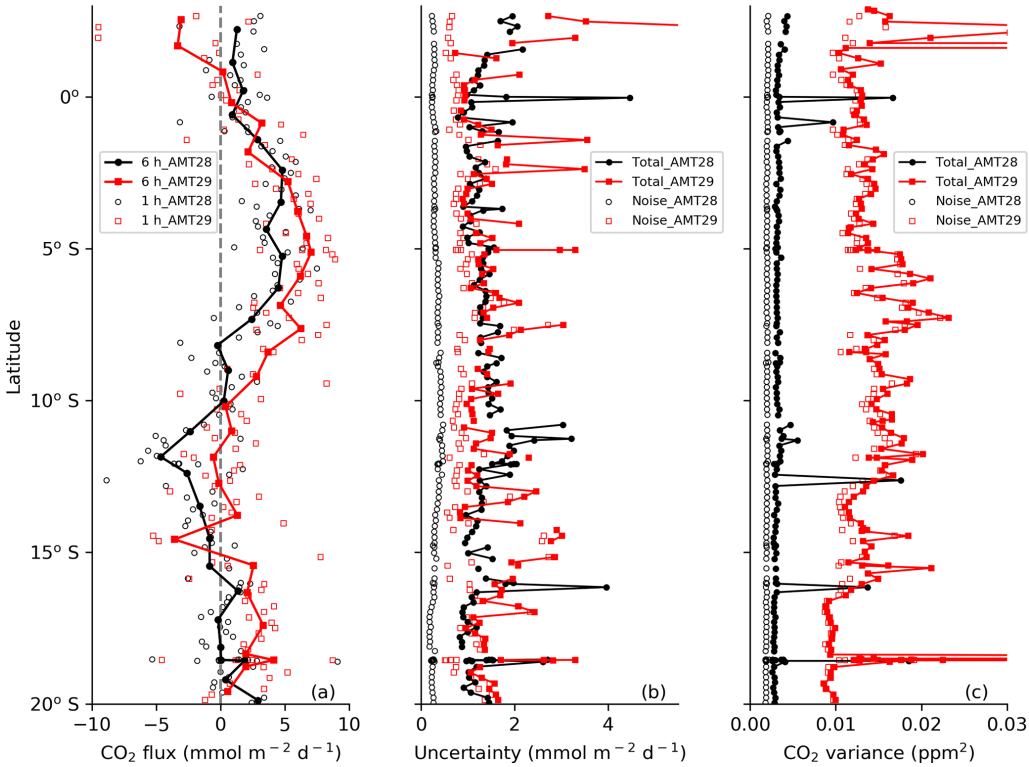

Figure 4. **(a)** Air–sea CO_2 fluxes (hourly and 6 h averages), **(b)** random uncertainty in flux (total and due to instrument noise only), and **(c)** variance in CO_2 mixing ratio (total and due to instrument noise only) for two Atlantic cruises.

Table 3. Variance in the CO_2 mixing ratio estimated using Eqs. (8) and (11) for the Arctic (JR18006/7, Picarro G2311-f) and Atlantic cruises (AMT28, Picarro G2311-f; AMT29, LI-7200). Total CO_2 variance (σ_c^2) consists of white noise ($\sigma_{c_n}^2$) and atmospheric processes ($\sigma_{c_a}^2$). The latter can be further broken down to the CO_2 variance due to vertical flux ($\sigma_{c_{av}}^2$) and due to other processes ($\sigma_{c_{ao}}^2$).

CO_2 variance ($\times 10^{-3}$ ppm^2)	JR18006	JR18007	AMT28	AMT29
Total, σ_c^2	9.9	8.6	3.6	13.9
Due to instrument white noise, $\sigma_{c_n}^2$	5.8	5.4	2.0	12.6
Due to atmospheric processes, $\sigma_{c_a}^2$	4.1	3.3	1.6	1.3
– Due to vertical flux, $\sigma_{c_{av}}^2$	1.3	0.8	0.03	0.08
– Due to other atmospheric processes, $\sigma_{c_{ao}}^2$	2.8	2.5	1.6	1.2

mixing ratio transported horizontally can yield variance that greatly exceeds the variance from vertical flux.

Three quasi-independent methods were used to estimate random uncertainty in EC air–sea CO_2 fluxes caused by instrument noise (δF_{RN}; Methods C–E, Sect. 2.3.3). Good agreement was found between all three estimates (Fig. C2, Appendix C) when $\sqrt{2}$ is used as the constant in Eq. (7) (a). The $\Delta F_{RN,Billesbach}$ estimates have more scatter and are slightly higher than the theoretical results, possibly because the random shuffling of data fails to fully exclude the contribution from atmospheric turbulence (Rannik et al., 2016). For the remainder of this study, we use the $\delta F_{RN,Mauder}$ method to estimate δF_{RN}.

We used three methods to estimate the total random uncertainty (δF_R; Methods A–C, Sect. 2.3.3) in the hourly averaged air–sea CO_2 fluxes. There is good agreement among the three estimates ($r > 0.88$; Fig. C1, Appendix C). Again, the constant in Eq. (7) (a) is set to $\sqrt{2}$, as informed by the instrument noise uncertainty analysis above. We use $\delta F_{R,Finkelstein}$ (Eq. 6) to estimate the total random flux uncertainty hereafter. Our decision is based on $\delta F_{R,Finkelstein}$ not requiring the integral timescale (unlike $\delta F_{R,Blomquist}$) and showing less scatter than $\delta F_{R,Wienhold}$.

Figure 5 shows the different relative contributions to the random flux uncertainty for the Arctic cruises (hourly average). Here the uncertainty is normalised by the flux magnitude and then averaged into flux magnitude bins. When the

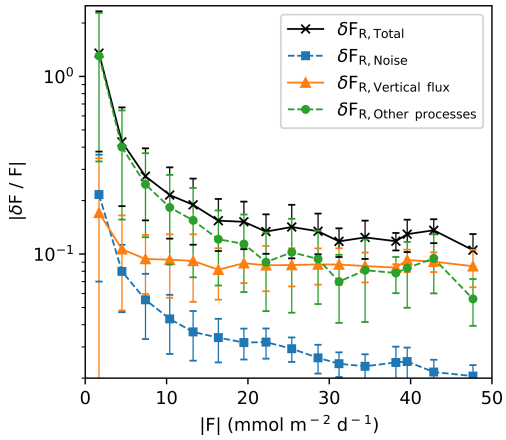

Figure 5. Relative random uncertainty in hourly CO_2 flux and its contribution from noise, vertical flux, and other processes during two Arctic cruises. Relative random uncertainty data are binned into 3 mmol m^{-2} d^{-1} flux magnitude bins (error bars represent 1 standard deviation).

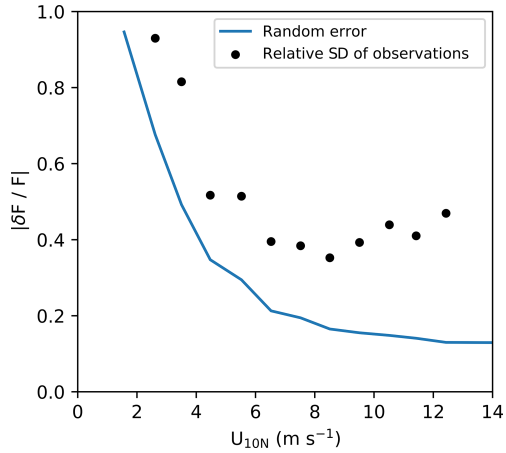

Figure 6. Comparison of relative random uncertainty in hourly CO_2 flux and relative standard deviation (RSTD; standard deviation/|flux mean|) of the EC CO_2 flux from two Arctic cruises. These results are binned in 1 m s^{-1} wind speed bins.

flux magnitude is sufficiently large (> 20 mmol m^{-2} d^{-1}), the total relative random uncertainty in flux asymptotes to about 15 % and is driven by variance associated with both vertical flux and other atmospheric processes. This estimate is similar to uncertainties in air–sea fluxes of other well resolved (i.e. high signal-to-noise ratio) variables (Fairall et al., 2000). At a lower flux magnitude, uncertainty due to atmospheric processes other than vertical flux dominates the total random uncertainty. Uncertainty due to the white noise from the Picarro G2311-f gas analyser is small.

3.2 Summary of systematic and random uncertainties

The total uncertainty δF in the hourly average EC CO_2 flux (estimated using Eq. 3) ranges from 1.4 to 3.2 mmol m^{-2} d^{-1} in the mean for the four cruises (Table 4). Our EC flux system setup was optimal, and subsequent corrections have minimised any bias to < 9 % (Sect. 2.3.2). Systematic error is on average much lower than random error (Table 4). This means the accuracy of the EC CO_2 flux measurements is very high, but the precision of hourly averaged EC CO_2 air–sea flux measurements is relatively low. In Sect. 4.1, we discuss how the precision can be improved by averaging the observed fluxes for longer.

The theoretical uncertainty estimates above can be compared with a portion of the AMT28 cruise data (15–20° S, ∼ 25° W; Fig. 4), when the ship encountered sea surface CO_2 fugacity close to equilibrium with the atmosphere (i.e. $\Delta fCO_2 \sim 0$; Fig. A2, Appendix A). The data from this region are useful for assessing the random and systematic flux uncertainties. The standard deviation of the EC CO_2 flux during cruise AMT28 when $\Delta fCO_2 \sim 0$ is 1.6 mmol m^{-2} d^{-1}, which compares well with the theoretical random flux uncertainty in this region (1.4 mmol m^{-2} d^{-1}). The mean EC CO_2 flux from this region was 0.5 mmol m^{-2} d^{-1}, which is indistinguishable from zero considering the random uncertainty. This further confirms the minimal bias in our flux observations.

Figure 6 shows a comparison between the relative uncertainty and the relative standard deviation (RSTD) in the hourly CO_2 flux for the two Arctic cruises. Results have been binned into 1 m s^{-1} wind speed bins. Wind speed was converted to 10 m neutral wind speed (U_{10N}) using the COARE3.5 model (Edson et al., 2013). The relative random error decreases with increasing wind speed. This is partly because the fluxes tend to be larger at higher wind speeds, and so the signal-to-noise ratio in the flux is greater. In addition, at higher wind speeds, a greater number of high-frequency turbulent eddies are sampled by the EC system, providing better statistics of turbulent eddies and lower sampling error.

The RSTD of the flux is greater in magnitude than the estimated flux uncertainty because it also contains environmental variability. The CO_2 flux auto-covariance analysis (Sect. 4.1) shows that random error in hourly flux explains ∼ 20 % of the flux variance on average for the two Arctic cruises. This implies that the remaining variability in the EC flux (∼ 80 %) is due to natural phenomena (e.g. changes in ΔfCO_2 or wind speed). Similarly, substantial variability is typical in EC-derived CO_2 gas transfer velocity at a given wind speed (e.g. Edson et al., 2011; Butterworth and Miller, 2016). K_{660} is derived from (EC CO_2 flux)/ΔfCO_2, and thus an understanding of EC flux uncertainty can help understand and explain the variability in EC-derived gas transfer velocity estimates (Sect. 4.2).

Table 4. Summary of hourly average EC CO_2 fluxes and associated uncertainties in the mean for the four cruises (mmol m^{-2} d^{-1}). Shown are the mean CO_2 flux magnitude ($\overline{|F|}$, mmol m^{-2} d^{-1}), upper limitation of the total uncertainty (δF; Eq. 3), upper limitation of the absolute systematic error ($|\delta F_S|$; propagated from Table 2 and Eq. 5), and random error (δF_R; Eq. 6). The random error components are white noise (δF_{RN}; Eq. 10), vertical flux (δF_{RV}; Eqs. 7 and 8), and other atmospheric processes ($\delta F_{RO} = \sqrt{\delta F_R^2 - \delta F_{RN}^2 - \delta F_{RV}^2}$). The total uncertainty is also expressed as a percentage of the mean flux magnitude ($\delta F/|F| \cdot 100\,\%$).

Cruises	JR18006	JR18007	AMT28	AMT29
$\overline{\lvert CO_2\ \text{flux}\rvert}$, $\overline{\lvert F\rvert}$	10.1	16.3	2.5	3.5
Total uncertainty, δF ($\delta F/\lvert F\rvert \cdot 100\,\%$)	2.3 (23 %)	3.2 (20 %)	1.4 (58 %)	1.7 (49 %)
Systematic error, $\lvert\delta F_S\rvert$	0.8	1.2	0.3	0.3
Total random error, δF_R	2.2	2.9	1.4	1.7
Random error due to white noise, δF_{RN}	0.5	0.6	0.3	1.0
Random error due to vertical flux, δF_{RV}	1.1	1.4	0.2	0.4
Random error due to other atmospheric processes, δF_{RO}	1.5	2.4	1.4	1.5

4 Discussion

4.1 Impact of averaging timescale on flux uncertainty

The random error in flux decreases with increasing averaging time interval T or the number of sampling points n (Eqs. 6, 7, and 10). This is because a longer averaging time interval results in better statistics of the turbulent eddies. However, averaging for too long is also not ideal since the atmosphere is less likely to maintain stationarity. The typical averaging time interval is thus typically between 10 min and 60 min for air–sea flux measurements (20 min intervals were used in this study). The time series of quality controlled 20 min flux intervals can be further averaged over a longer timescale to reduce the random uncertainty. Averaging the 20 min flux intervals assumes that the flux interval data are essentially repeat measurements within a chosen averaging timescale. If the 20 min flux intervals are averaged, one can ask the following question: what is the optimal averaging timescale for interpreting air–sea EC CO_2 fluxes?

We use an auto-covariance method to determine the optimal averaging timescale. The observed variance in CO_2 flux consists of random uncertainty (random noise) as well as natural variability. The random noise component should only contribute to the CO_2 flux variance when the data are zero-shifted. After the CO_2 flux data are shifted, the noise will not contribute to the auto-covariance function. Figure 7 shows the auto-covariance function of the air–sea CO_2 flux with different averaging timescales for Arctic cruise JR18007. For the 20 min fluxes (Fig. 7a), the auto-covariance decreases rapidly between the zero shift and the initial time shift, which indicates that a large fraction of the 20 min flux variance is due to random noise.

The random noise in the CO_2 fluxes decreases with a longer averaging timescale, with the greatest effect observed from 20 min to 1 h (Fig. 7b). A fit to the noise-free auto-covariance function extrapolated back to a zero time shift gives us an estimate of the non-noise variability in the natural CO_2 flux. Subtracting the extrapolated natural flux variability from the total variance in CO_2 flux provides an estimate of the random noise in the flux for each averaging timescale (Fig. 7a). All four cruises consistently demonstrate a non-linear reduction in the noise contribution to the flux measurements when the averaging timescale increases (Fig. 8). The random noise in flux can be expressed relative to the natural variance in flux representing the inverse of the signal-to-noise ratio (i.e. random noise in flux/natural flux variability, hereafter referred to as noise : signal).

The noise : signal also facilitates comparison of all four cruises (Fig. 8) and demonstrates the consistent effect that increasing the averaging timescale has on noise : signal. Consistent with Table 4, the Arctic cruises show much lower noise : signal because the flux magnitudes are much larger. Typical detection limits in analytical science are often defined by a 1 : 3 noise : signal ratio. A 1 : 3 noise : signal is achieved with a 1 h averaging timescale for the Arctic cruises. The Atlantic cruises encountered much lower air–sea CO_2 fluxes, and an averaging timescale of at least 3 h is required to achieve the same 1 : 3 noise : signal ratio.

The flux measurement uncertainty at a 6 h averaging timescale for the AMT cruises is ~ 0.6 mmol m^{-2} d^{-1}. The analysis presented above permits an answer to the question posed at the beginning of the Results section. The mean difference between the 6 h averaged EC CO_2 flux observations on AMT29 and AMT28 (1.3 mmol m^{-2} d^{-1}; Fig. 4a) is much greater than the measurement uncertainty. This significant difference was likely because of the interannual variability in AMT CO_2 flux due to changes in the natural environment (e.g. ΔfCO_2, sea surface temperature, and physical drivers of interfacial turbulence such as wind speed) during the two cruises.

At a typical research ship speed of ~ 10 knots, the AMT cruises cover ~ 110 km in 6 h, which is equivalent to $\sim 1°$ latitude. Averaging for longer than 6 h is likely to cause substantial loss of real information about the natural variations

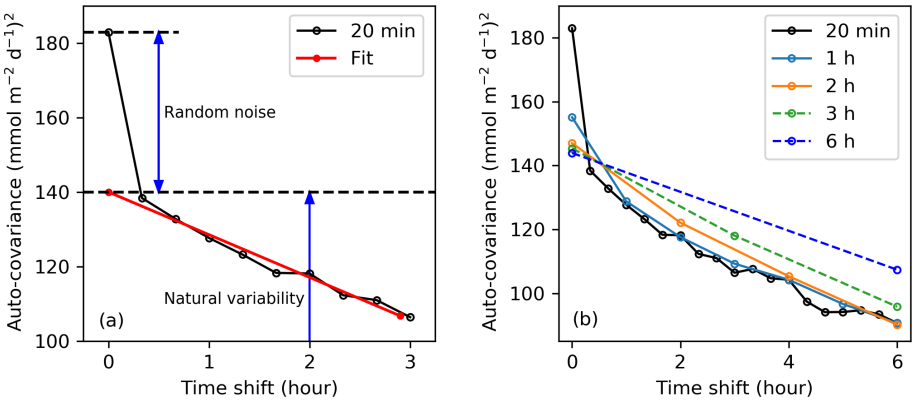

Figure 7. **(a)** Auto-covariance of the original 20 min fluxes (cruise JR18007) and a fit to the noise-free auto-covariance function extrapolated back to a zero time shift. **(b)** CO_2 flux auto-covariance functions with different averaging timescales. The black line represents the auto-covariance of the original 20 min fluxes. The 20 min fluxes are further averaged at different timescales (1, 2, 3, and 6 h), and the corresponding auto-covariance functions are shown with different colours (dark blue, orange, green, and light blue).

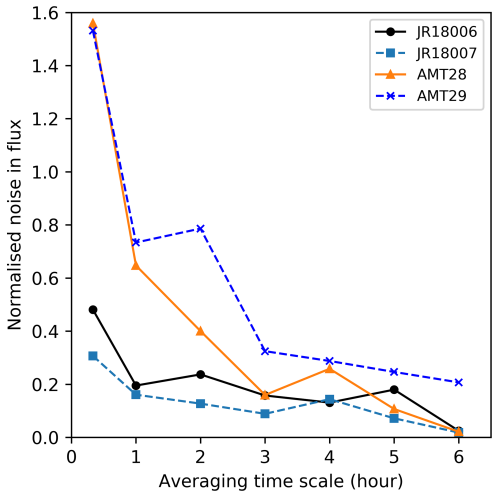

Figure 8. Effect of the averaging timescale on the noise : signal (random noise in flux/natural flux variability) for EC air–sea CO_2 flux measurements during four cruises.

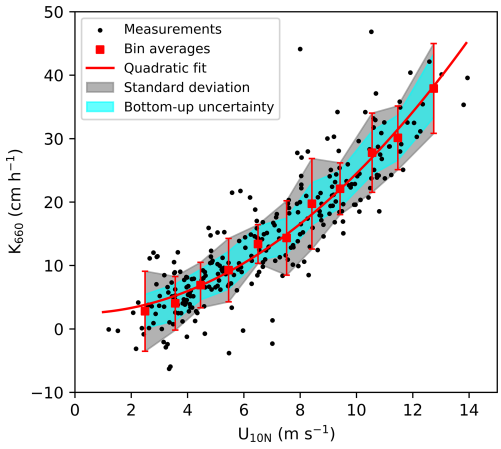

Figure 9. Gas transfer velocity (K_{660}) measured on Arctic cruise JR18007 (hourly average, signal : noise ~ 5) vs. 10 m neutral wind speed (U_{10N}). Red squares represent 1 m s^{-1} bin averages, with error bars representing 1 standard deviation (SD). The red curve represents a quadratic fit using the bin averages: $K_{660} = 0.22\,U_{10N}^2 + 2.46$ ($R^2 = 0.76$). The grey shaded area represents the standard deviation calculated for each wind speed bin ($K_{660} \pm 1$ SD). The cyan region represents the upper and lower bounds in K_{660} uncertainty computed from the EC flux uncertainty ($K_{660} \pm \delta K_{660}$; see text for detail).

in air–sea CO_2 flux and the drivers of flux variability. For example, the mean flux between 0–20° S during cruise AMT28 is 0.9 mmol m^{-2} d^{-1}. However, the 6 h average EC measurements show that the flux varied between +5 mmol m^{-2} d^{-1} ($\sim$ 2–6° S) and −5 mmol m^{-2} d^{-1} ($\sim$ 11–13° S; Fig. 4a).

4.2 Effect of CO_2 flux uncertainty on the gas transfer velocity K

The uncertainties in the EC CO_2 air–sea flux measurement will influence the uncertainty that translates to EC-based estimates of the gas transfer velocity, K. For illustration, K is computed for Arctic cruise JR18007, which had a high flux signal : noise ratio of ~ 5 (Fig. 8). Any data potentially influenced by ice and sea ice melt were excluded using a sea surface salinity filter (data excluded when salinity < 32). Equation (1) is rearranged and used with concurrent measurements of CO_2 flux (F), ΔfCO_2, and sea surface temperature (SST) to obtain K adjusted for the effect of temperature (K_{660}).

The determination coefficient (R^2) of the quadratic fit between wind speed (U_{10N}) and EC-derived K_{660} (Fig. 9) demonstrates that wind speed explains 76 % of the K_{660} variance during Arctic cruise JR18007. How much of the remaining 24 % can be attributed to uncertainties in EC CO_2 fluxes?

Variability in K_{660} within each $1\,\mathrm{m\,s^{-1}}$ wind speed bin can be considered to have minimal wind speed influence. It is thus useful to compare the variability within each wind speed bin ($K_{660} \pm 1$ SD) with the upper and lower uncertainty bounds derived from the EC flux measurements. Uncertainty in EC flux-derived K_{660} (δK_{660}) is calculated from the uncertainty in hourly EC flux (δF) by rearranging Eq. (1) (bulk flux equation) and replacing F with δF. The resultant δK_{660} is then averaged in wind speed bins. The shaded cyan band in Fig. 9 ($K_{660} \pm \delta K_{660}$) is consistently narrower than the grey shaded band ($K_{660} \pm 1$ SD). On average, EC flux-derived uncertainty in K_{660} can only account for a quarter of the K_{660} variance within each wind speed bin, and the remaining variance is most likely due to the non-wind speed factors that influence gas exchange (e.g. breaking waves, surfactants).

The analysis above can be extended to assess how EC flux-derived uncertainty affects our ability to parameterise K_{660} (e.g. as a function of wind speed). To do so, a set of synthetic K_{660} data is generated (same U_{10N} as the K_{660} measurements in Fig. 9). The synthetic K_{660} data are initialised using a quadratic wind speed dependence that matches JR18007 (i.e. $K_{660} = 0.22\,U_{10N}^2 + 2.46$). Random Gaussian noise is then added to the synthetic K_{660} data, with relative noise level corresponding to the relative flux uncertainty values taken from JR18007 (mean of 20 %; Table 4). The relative uncertainty in K_{660} due to EC flux uncertainty ($\delta K_{660}/K_{660}$) shows a wind speed dependence (Fig. S4a in the Supplement), and the artificially generated Gaussian noise incorporates this wind speed dependence (Fig. S4b, Supplement). The R^2 of the quadratic fit to the synthetic data as a function of U_{10N} is 0.90 (the rest of the variance is due to uncertainty in K_{660}). Since wind speed explains 76 % of variance in the observed K_{660}, it can be inferred that non-wind speed factors can account for 14 % (i.e. (100–76) %–(100–90) %) of the total variance in K_{660} from this Arctic cruise. If the synthetic K_{660} data are assigned a relative flux uncertainty of 50 % (reflective of a region with low fluxes, e.g. AMT28/29), the R^2 of the wind speed dependence in the synthetic data decreases to 0.60.

The relative uncertainty in EC flux-derived K_{660} ($\delta K_{660}/K_{660}$) is large when $|\Delta fCO_2|$ is small (Fig. 10). Previous EC studies have filtered EC flux data to remove fluxes when the $|\Delta fCO_2|$ falls below a specified threshold (e.g. 20 µatm, Blomquist et al. (2017); 40 µatm, Miller et al. (2010), Landwehr et al. (2014), Butterworth and Miller (2016), Prytherch et al. (2017); 50 µatm, Landwehr et al. (2018)). Analysis of the data presented here suggests that a $|\Delta fCO_2|$ threshold of at least 20 µatm is reasonable for hourly K_{660} measurements, leading to δK_{660} of $\sim 10\,\mathrm{cm\,h^{-1}}$ ($\delta K_{660}/K_{660} \sim 1/3$) or less on average. At very large $|\Delta fCO_2|$ of over 100 µatm, δK_{660} is reduced to only a few centimetres per hour ($\mathrm{cm\,h^{-1}}$) ($\delta K_{660}/K_{660} \sim 1/5$). At longer flux averaging timescales, it may be possible to relax the minimal $|\Delta fCO_2|$ threshold.

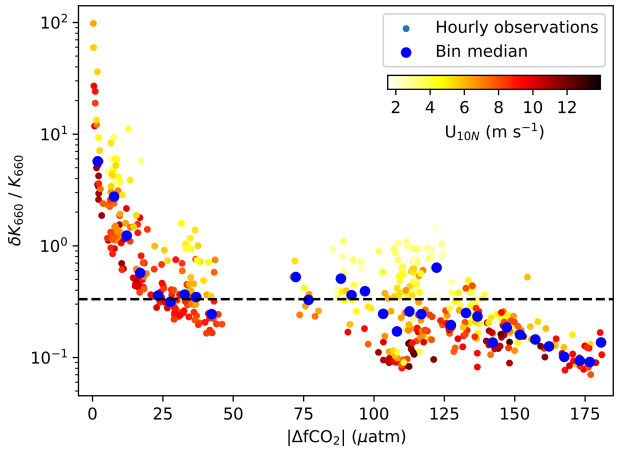

Figure 10. Relative uncertainty in EC-estimated hourly K_{660} ($\delta K_{660}/K_{660}$) vs. the magnitude of the air–sea CO_2 fugacity difference ($|\Delta fCO_2|$) during Arctic cruise JR18007 and Atlantic cruises AMT28 and AMT29 (no ΔfCO_2 data were collected on JR18006). The data points are colour-coded by wind speed. Blue points are medians of $\delta K_{660}/K_{660}$ in 5 µatm bins. Here we use the parameterised K_{660} ($= 0.22\,U_{10N}^2 + 2.46$) to normalise the uncertainty in K_{660}. The dashed line represents the 3 : 1 signal : noise ratio ($\delta K_{660}/K_{660} = 1/3$).

5 Conclusions

This study uses data from four cruises with a range in air–sea CO_2 flux magnitude to comprehensively assess the sources of uncertainty in EC air–sea CO_2 flux measurements. Data from two ships and two different state-of-the-art CO_2 analysers (Picarro G2311-f and LI-7200, both fitted with a dryer) are analysed using multiple methods (Sect. 2.3). Random error accounts for the majority of the flux uncertainty, while the systematic error (bias) is small (Table 4). Random flux uncertainty is primarily caused by variance in CO_2 mixing ratio due to atmospheric processes. The random error due to instrument noise for the Picarro G2311-f is 3-fold smaller than for LI-7200 (Table 4 and Fig. D1, Appendix D). However, the contribution of the instrument noise to the total random uncertainty is much smaller than the contribution of atmospheric processes such that both gas analysers are well suited for air–sea CO_2 flux measurements.

The mean uncertainty in hourly EC flux is estimated to be 1.4–3.2 $\mathrm{mmol\,m^{-2}\,d^{-1}}$, which equates to the relative uncertainty of ~ 20 % in high CO_2 flux regions and ~ 50 % in low CO_2 flux regions. Lengthening the averaging timescale can improve the signal : noise ratio in EC CO_2 flux through the reduction of random uncertainty. Auto-covariance analysis of CO_2 flux is used to quantify the optimal averaging timescale (Figs. 7 and 8, Sect. 4.1). The optimal averaging timescale varies between 1 h for regions of large CO_2 flux (Arctic in our analysis) and at least 3 h for regions of low CO_2 flux (tropical/subtropical Atlantic in our analysis).

The measurement uncertainty in EC CO_2 flux contributes directly to scatter in the derived gas transfer velocity, K_{660}. Flux uncertainties determined in this paper are applied to a synthetic K_{660} dataset. This enables a partitioning of the variance in measured K_{660} that is due to EC CO_2 flux uncertainty, wind speed, and other processes (10 %, 76 %, 14 % for Arctic cruise JR18007). At a given averaging timescale, a $|\Delta fCO_2|$ threshold helps to reduce the scatter in K_{660}. A minimum $|\Delta fCO_2|$ filter of 20 µatm is needed for interpreting hourly K_{660} data, with the signal : noise ratio in K_{660} improving further at higher $|\Delta fCO_2|$.

Appendix A: Cruise tracks

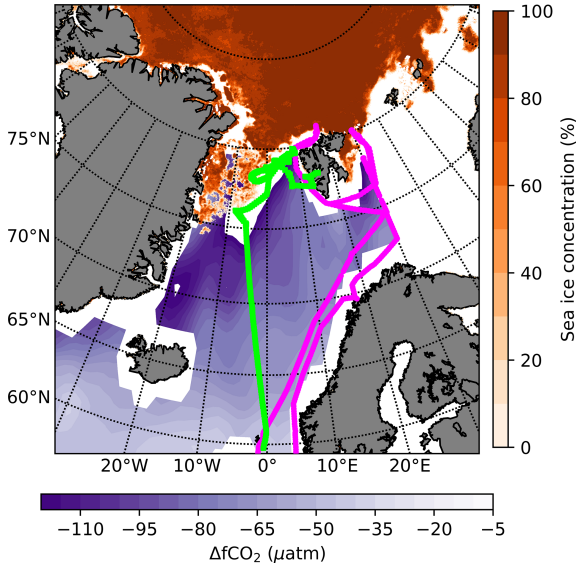

Figure A1. Cruise tracks of JR18006 (magenta) and JR18007 (green). The bottom colour bar indicates the CO_2 fugacity difference (ΔfCO_2) of August 2019 (Bakker et al., 2016; Landschützer et al., 2020), while the right colour bar shows the Arctic sea ice concentrations of 1 August 2019 measured by Advanced Microwave Scanning Radiometer – Earth Observing System Sensor (AMSR-E; Spreen et al., 2008).

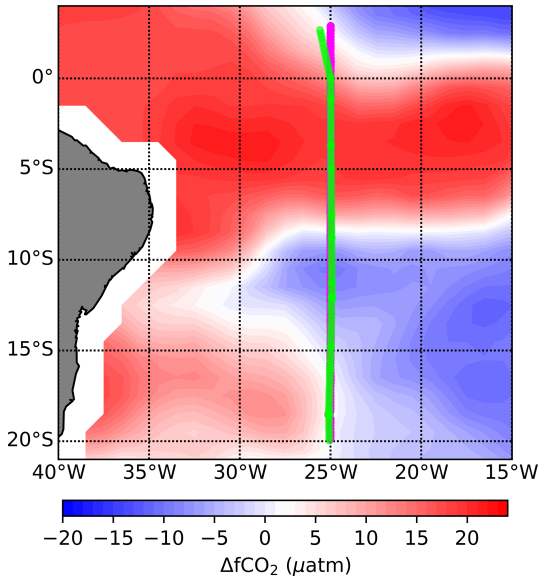

Figure A2. Cruise tracks of AMT28 (magenta) and AMT29 (green). The ocean is coloured with the ΔfCO_2 for October 2018 (Bakker et al., 2016; Landschützer et al., 2020).

Appendix B: Integral timescale and variance spectra of CO_2 and vertical wind velocity

Integral timescale is used in the flux uncertainty calculation (Eqs. 5 and 7). The definition of integral timescale τ_x of variable x is

$$\tau_x = \frac{1}{\sigma_x^2}\int_0^\infty r_{xx}(t)\mathrm{d}t, \tag{B1}$$

where σ_x^2 is the variance of x, and r_{xx} is the auto-covariance function of x. t is the shifting time of auto-covariance (which is different from the lag time between w and CO_2 in the EC flux calculation). We can use Eq. (B1) to estimate the integral timescale of w and CO_2 directly. However, integration up to infinity is not practical. Instead we can numerically estimate the timescale by determining the time corresponding to the auto-covariance coefficient function (r_{xx}/σ_x^2) value decaying to $1/e$ ($1/e$ decaying method) or by integrating the auto-covariance function up to the first zero crossing of the function (zero crossing method) (Rannik et al., 2009).

One can also use similarity theory to estimate the integral timescale theoretically (Blomquist et al., 2010):

$$\tau_\mathrm{w} = 2.8\frac{z}{\overline{u_\mathrm{r}}}f_\tau(z/L). \tag{B2}$$

Here, $\overline{u_\mathrm{r}}$ is the relative wind speed. The similarity function $f_\tau(z/L)$ is described by the stability parameter z/L, where z is the observation height (m), and L is the Obukhov length (m) (Blomquist et al., 2010).

Yet another method to estimate the integral timescale is from the peak frequency (f_max) in the w variance spectrum (Kaimal and Finnigan, 1994):

$$\tau_\mathrm{w} = \frac{1}{2\pi f_\mathrm{max}}. \tag{B3}$$

The integral timescales of w estimated by these four methods for cruise JR18007 are shown in Fig. B1. The integral timescale estimated by the zero crossing method agrees well with the peak frequency estimates using Eq. (B3). The 1/e decaying method tends to underestimate the integral timescale, which is generally observed for turbulent signals (Rannik et al., 2009), whereas the similarity method (Eq. B2) considerably overestimates the integral timescale. Based on the recent analysis (as yet unpublished) of the entire NOAA PSL flux database, the Eq. (B2) formulation is now thought to be an overestimate (review comment for this paper from Blomquist, 2021). In this study we use the integral timescale of w from the zero crossing method to estimate the theoretical flux uncertainty (Eqs. 5 and 7). The theoretical systematic error estimates (Eq. 8) also require the integral timescale of CO_2. The integral timescale of CO_2 is difficult to evaluate from the above four methods due to instrument noise. Instead, we estimate it by directly integrating the auto-covariance function (Eq. B1) to a shift time of 200 s (we found no significant difference of the integral timescale when integrating the CO_2 auto-covariance function for shift times ranging from 150 to 250 s).

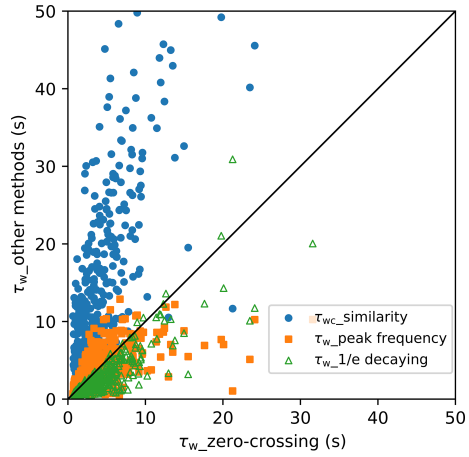

Figure B1. Comparison of integral timescales of w estimated by four different methods. Estimated integral timescales from the zero crossing method (integrating the auto-covariance function up to first zero crossing the function) agree well with the estimation of peak frequency method (Eq. B3). However, the similarity method (Eq. B2) overestimates the integral timescale, whereas the $1/e$ decaying method (determining the time needed for the auto-covariance coefficient function value to decay to $1/e$) tends to underestimate the integral timescale.

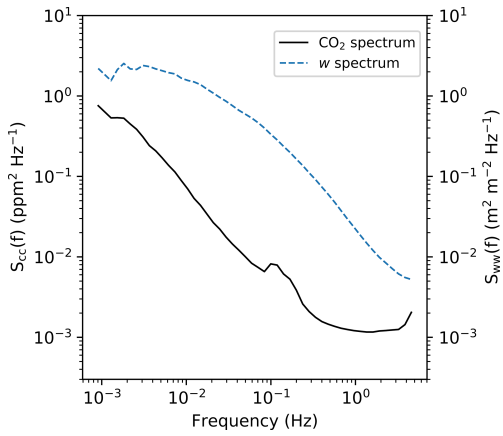

Figure B2. Mean variance spectra for CO_2 and w for one Arctic cruise JR18007. The near-constant CO_2 variance at high frequency (1–5 Hz) indicates the band-limited noise in the CO_2 signal. In contrast, the w spectrum does not show a similar band-limited noise at < 10 Hz.

Appendix C: Comparison of the uncertainty estimates by different methods

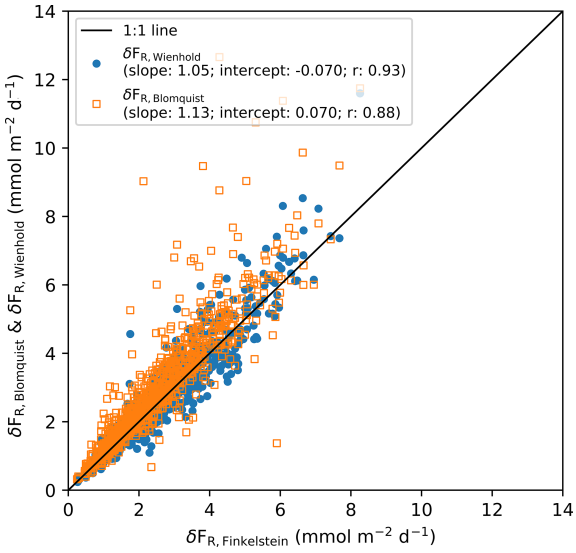

Figure C1. Comparison of total random uncertainties in hourly flux estimated by three different methods for the Arctic cruises. The empirical estimates $F_{\mathrm{R,Wienhold}}$ agree well with one of the theoretical estimates $\Delta F_{\mathrm{R,Finkelstein}}$ ($r = 0.93$). The other theoretical estimate $\Delta F_{\mathrm{R,Blomquist}}$ is slightly higher than the random uncertainties $\Delta F_{\mathrm{R,Finkelstein}}$ (slope $= 1.13$) if the constant in Eq. (8) is set equal to $\sqrt{2}$.

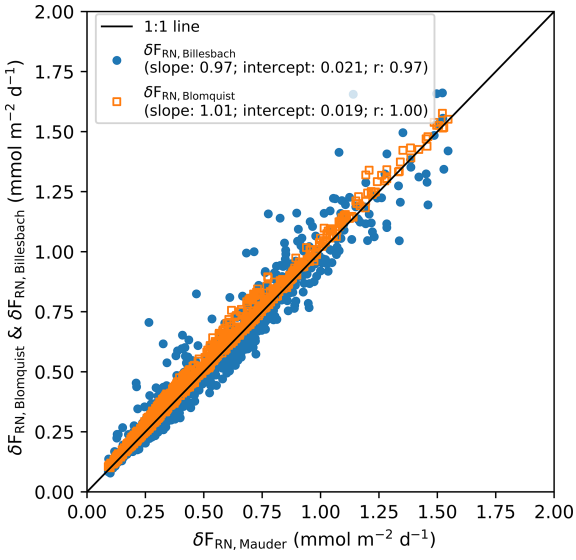

Figure C2. Comparison of random error in hourly flux due to instrument white noise, estimated by three different methods for the Arctic cruises. The three uncertainty estimations agree well. The correlation coefficient (r) between $\delta F_{\mathrm{RN,Mauder}}$ and $\delta F_{\mathrm{RN,Blomquist}}$ is 1 if the constant in Eq. (7) (a) is set to $\sqrt{2}$.

Appendix D: Performance of two gas analysers

Figure D1 shows a comparison between the performance of the Picarro 2311-f and the LI-7200 gas analysers. We estimated that the noise of the LI-7200 is on average 3 times higher than that of the Picarro 2311-f (Table 3). Indeed, random error in the CO_2 flux due to the white noise is much higher for the LI-7200 than for the Picarro 2311-f, but the total flux uncertainty of the EC system with the LI-7200 on AMT29 is only slightly higher than that of the EC system with the Picarro 2311-f on AMT28 (Table 4). Again, this is because for both EC systems, sampling error dominates the total random uncertainty, while the contribution of instrument noise ($< 30\,\%$) to the total uncertainty is relatively small (Billesbach, 2011; Langford et al., 2015; Mauder et al., 2013; Rannik et al., 2016). Another often used CRDS gas analyser in EC measurements is the Los Gatos Research (LGR) Fast Greenhouse Gas Analyser (FGGA) (Prytherch et al., 2017). Yang et al. (2016) showed that LGR FGGA is ca. 10 times noisier than the Picarro G2311-f, and as a result, the total CO_2 flux uncertainty measured by the LGR is 4 times higher than that by the Picarro. From the perspective of measurement noise, Picarro and LI-7200 gas analysers are better suited for air–sea CO_2 flux measurements than the LGR FGGA.

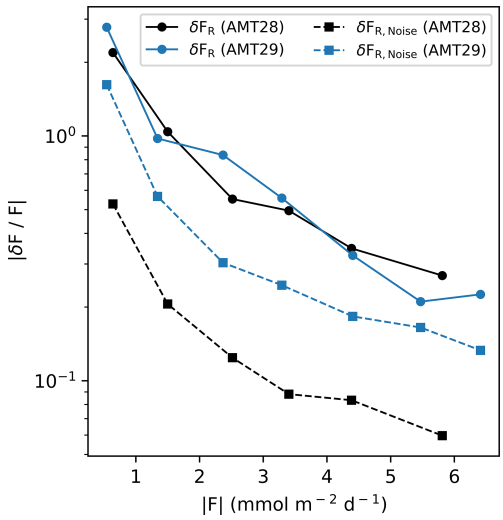

Figure D1. Comparison of the relative total random uncertainty and the relative random error component due to white noise for different gas analysers. A Picarro G2311-f gas analyser was used on AMT28 and a LI-7200 infrared gas analyser on AMT29.

Data availability. The processed hourly EC CO_2 fluxes and uncertainties can be found in the Supplement of this paper. Raw, high-frequency (10 Hz) data are large (tens of gigabytes) and are archived at PML. Please contact the authors directly if you are interested in the raw data.

Supplement. The supplement related to this article is available online at: https://doi.org/10.5194/acp-21-8089-2021-supplement.

Author contributions. TGB and MY designed and installed the eddy covariance systems on ships and managed the collections of measurements. VK collected and processed the CO_2 fugacity data. YD processed and analysed the data with the help of MY and TGB. YD wrote the paper with input from DCEB, TGB, and MY. All authors contributed to and approved the final paper.

Competing interests. The authors declare that they have no conflict of interest.

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