



**Manchester
Metropolitan
University**

Middlebrook, N and Heneghan, NR and Evans, DW and Rushton, A and Falla, D (2020) Reliability of temporal summation, thermal and pressure pain thresholds in a healthy cohort and musculoskeletal trauma population. PLoS One, 15 (5). ISSN 1932-6203

Downloaded from: <http://e-space.mmu.ac.uk/627740/>

Version: Published Version

Publisher: Public Library of Science (PLoS)

DOI: <https://doi.org/10.1371/journal.pone.0233521>

Usage rights: Creative Commons: Attribution 4.0

Please cite the published version

<https://e-space.mmu.ac.uk>

RESEARCH ARTICLE

Reliability of temporal summation, thermal and pressure pain thresholds in a healthy cohort and musculoskeletal trauma population

Nicola Middlebrook^{1,2}, Nicola R. Heneghan¹, David W. Evans¹, Alison Rushton^{1,2}, Deborah Falla^{1,2*}

1 Centre of Precision Rehabilitation for Spinal Pain, School of Sport, Exercise and Rehabilitation Sciences, College of Life and Environmental Sciences, University of Birmingham, Edgbaston, Birmingham, United Kingdom, **2** NIHR Surgical Reconstruction & Microbiology Research Centre, University of Birmingham, Edgbaston, Birmingham, United Kingdom

* d.falla@bham.ac.uk



OPEN ACCESS

Citation: Middlebrook N, Heneghan NR, Evans DW, Rushton A, Falla D (2020) Reliability of temporal summation, thermal and pressure pain thresholds in a healthy cohort and musculoskeletal trauma population. *PLoS ONE* 15(5): e0233521. <https://doi.org/10.1371/journal.pone.0233521>

Editor: César Calvo-Lobo, Universidad de Leon, SPAIN

Received: January 20, 2020

Accepted: May 6, 2020

Published: May 29, 2020

Copyright: © 2020 Middlebrook et al. This is an open access article distributed under the terms of the [Creative Commons Attribution License](https://creativecommons.org/licenses/by/4.0/), which permits unrestricted use, distribution, and reproduction in any medium, provided the original author and source are credited.

Data Availability Statement: All data are available at DOI: [10.6084/m9.figshare.12102495](https://doi.org/10.6084/m9.figshare.12102495).

Funding: NM is a doctoral researcher, DF is the lead supervisor, AR and NH are supervisors. This study/project is funded by the National Institute for Health Research (NIHR) Surgical Reconstruction and Microbiology Research Centre (SRMRC). The views expressed are those of the author(s) and not necessarily those of the NIHR or the Department of Health and Social Care. The funders had no role in

Abstract

Traumatic injuries affect approximately 978 million people worldwide with 56.2 million requiring inpatient care. Quantitative sensory testing (QST) can be useful in predicting outcome following trauma, however the reliability of multiple QST including temporal summation (TS), heat and cold pain thresholds (HPT, CPT) and pressure pain thresholds (PPT) is unknown. We assessed intra (between day) and inter-rater (within day) reliability of QST in asymptomatic participants ($n = 21$), and inter-rater (within day) reliability in participants presenting with acute musculoskeletal trauma ($n = 25$). Intra-class correlations with 95% confidence intervals (ICC 3,2), standard error of measurement (SEM) and Bland Altman Plots for limits of agreement were calculated. For asymptomatic participants, reliability was good to excellent for HPT (ICC range 0.76–0.95), moderate to good for PPT (ICC range 0.52–0.93), with one site rated poor (ICC 0.41), and poor to excellent for TS scores (ICC range 0.20–0.91). For musculoskeletal trauma participants reliability was good to excellent for HPT and PPT (ICC range 0.76–0.86), and moderate to good reliability for TS (ICC range 0.69–0.91). SEM for HPT for both sets of participants was $\sim 1^\circ\text{C}$ and an average of 7N for asymptomatic participants and less than 8N for acute musculoskeletal trauma participants for PPT. This study demonstrates moderate to excellent intra and inter-rater reliability for HPT and PPT in asymptomatic participants and good to excellent inter-rater reliability for acute musculoskeletal trauma participants, with TS showing more variability for both sets of participants. This study provides foundations for future work evaluating the sensory function over time following acute musculoskeletal trauma.

Introduction

Approximately 978 million people worldwide sustain injuries, of which 56.2 million require inpatient care. A high proportion of these injuries are musculoskeletal in nature with 21.7 million sustaining fractures [1]. Acute pain is common, if not expected, following trauma [2] and persistent pain is associated with a poor rate of return to work and reduced activities of daily living [3–5]. Pain assessment in a musculoskeletal trauma population is often limited, with the

study design, data collection and analysis, decision to publish, or preparation of the manuscript.

Competing interests: The authors have declared that no competing interests exist.

numerical rating scale (NRS) or visual analogue scale (VAS) most commonly utilised in combination with patient reported questionnaires [6]. Quantitative Sensory Testing (QST), a psychophysical method used to assess sensation and pain perception, can also be used to evaluate the presence of peripheral and central sensitisation; a method commonly used to evaluate pain perception in patients with spinal cord injuries and musculoskeletal conditions such as low back pain, whiplash and osteoarthritis [7–10]. QST includes multiple tests such as assessing pain thresholds and temporal summation [11]. This can be achieved with different modalities such as thermal or mechanical stimuli [11]. Pain thresholds are defined as the moment that a sensation (e.g. pressure) changes to pain [11], whereas temporal summation can be defined as a gradual increase in the pain response following a number of repetitive stimuli such as heat or pressure [12]. Collectively, cold pain thresholds (CPT), heat pain thresholds (HPT), pressure pain thresholds (PPT) and temporal summation (TS), when used at both local and remote sites of injured tissue, can provide insight into the functioning of A and C fibres together with their central pathways, and can detect the presence of both local (peripheral sensitisation) and more centrally driven symptoms (e.g. secondary hyperalgesia and temporal summation of pain) [11, 13, 14]. QST can therefore be useful to understand which pain mechanisms are at play for an individual patient, knowledge which can inform a more personalised approach to rehabilitation and pain management [14].

Previous studies have found QST to be useful in the assessment of sensory function, and that the results of QST can be used to predict outcome such as the development of chronic pain following injury. For example, secondary hyperalgesia is common in patients with whiplash associated disorders [15–17], and the presence of widespread hyperalgesia including both mechanical and thermal changes at one month following injury can predict poorer outcome at 6 months post injury [18]. QST therefore has the potential to be utilised in a similar manner to predict outcome in a broader musculoskeletal trauma population.

Reliability of QST is well established in asymptomatic populations for PPT, thermal pain thresholds, TS and using the German Research Network on Neuropathic Pain (DNFS) protocol [19–26]. Within symptomatic populations, reliability studies have been conducted in participants with knee osteoarthritis [27], neuropathic pain [28], acute neck pain [29], chronic whiplash [30], spinal cord injury [31, 32] and fractured wrists [33]. However, no reliability study has been performed within an acute hospital environment, evaluating the reliability of QST in people with acute pain following major musculoskeletal trauma. Furthermore, there are limited systematic reviews evaluating the quality of published QST reliability studies. Of those which do exist, they have evaluated thermal sensory testing, conditioned pain modulation and QST test procedures for peripheral joint pain [34–36]. Variable approaches in reporting, methods used and statistical analysis were highlighted in these reviews with a call for higher quality reliability studies to be conducted. There is the need for a rigorous reliability study of multiple QST measures within an asymptomatic population and the need to establish the reliability of these measures for people with acute pain following musculoskeletal trauma.

The aims of this study were to 1) establish both intra and inter-rater reliability of HPT, CPT, PPT and TS measures in healthy asymptomatic volunteers and 2) establish inter-rater reliability of the same measures for people with acute pain following musculoskeletal trauma whilst based in an acute hospital setting.

Materials and methods

Ethical approval was obtained from the University of Birmingham Ethics Committee (ERN_17–0893) for asymptomatic participants and NHS ethical approval, HRA approval and individual site confirmation (17/WA/0421/IRAS 229790) was obtained for acute

musculoskeletal trauma participants. The study was conducted according to the Declaration of Helsinki. All data are available at DOI: [10.6084/m9.figshare.12102495](https://doi.org/10.6084/m9.figshare.12102495).

Study design

Two test-retest reliability studies were conducted. A within (inter-rater) and between day (intra-rater) reliability study was conducted for asymptomatic participants, and within (inter-rater) day reliability study was conducted for participants with acute musculoskeletal trauma. For both intra and inter-rater reliability all measures of HPT, CPT, PPT and TS were evaluated. For reliability studies using 2 raters, a sample size of $n \geq 19$ was required, based on a power calculation of 5% significance with true reliability exceeding 0.7 [37, 38]. Three raters were recruited with varying levels of QST experience (range 6 months to 10 years). Rater 1 and 2 completed the asymptomatic intra and inter-rater study and rater 1 and 3 completed the acute musculoskeletal trauma inter-rater study. Raters underwent training prior to data collection in the use of the equipment and testing procedure. This included training on the testing procedure of all modalities to ensure the same wording was used by each rater to the participants, as well as training on the use of equipment to ensure that a similar technique was used for all raters. For both sets of participants, room temperature and background noise level were not strictly controlled.

Participants

Asymptomatic participants. Participants were recruited from the staff and student population of the University of Birmingham, United Kingdom via poster advertisement. All willing participants were screened to ensure they met the inclusion criteria, and written consent was gained if eligible. Inclusion criteria were adults (aged ≥ 18) with the capacity to understand both written and verbal English language. Exclusion criteria included any pain which had affected activities of daily living within the last month, or neurological and rheumatological conditions (e.g. Parkinson's disease, multiple sclerosis, rheumatoid arthritis etc) [39].

Acute musculoskeletal trauma participants. Participants were recruited from a major trauma centre in the United Kingdom. All potential participants were screened by a team of research nurses. If eligible, participants were approached by one of the research nurses and were given a participant information sheet. All participants who were interested in the study were then approached by one of the research team to obtain consent. Musculoskeletal trauma was defined as any trauma (e.g. from a road traffic accident) which involved the musculoskeletal system. This included fractures, stab and gunshot wounds. The broad definition is in keeping with studies and systematic reviews evaluating musculoskeletal trauma [40–42]. Inclusion criteria included being admitted to the major trauma centre within 14 days of injury, adult (aged ≥ 16), and the capacity to understand both written and verbal English language. The difference in inclusion criteria age between the two sets of participants was due the University population where the asymptomatic participants were recruited were aged ≥ 18 , whereas within the hospital setting the adult age was defined as ≥ 16 [40]. Therefore, to ensure we recruited all potential eligible participants the age range was kept at ≥ 16 for the acute musculoskeletal trauma participants. Exclusion criteria included an acute intra-cranial lesion with a Glasgow Coma Scale score of ≤ 14 , any neurological or rheumatological disorders, ongoing terminal illness (e.g. comorbid cancer) with short life expectancy or prolonged use of corticosteroids [40]. The inclusion and exclusion criteria for the acute musculoskeletal trauma participants is intentionally diverse in order to evaluate reliability of these measures within this environment and population. By evaluating reliability in a lab-based environment on asymptomatic participants and then within a real-life environment allows applicability to a real-life clinical situation where these measures can be used.

Equipment

For both sets of participants the same equipment was used. For the thermal testing (HPT, CPT), the TSA-II NeuroSensory Analyzer thermal stimulator (Medoc Ltd, Israel) and accompanying software with a 30 x 30mm peltier thermode was used. For both PPT and TS testing, a handheld digital pressure algometer (Series 7 force gauge, Mark-10 corporation, USA) with custom made software (LabVIEW, National Instruments, USA) was used.

Procedure

Asymptomatic intra and inter-rater study. A total of four sessions took place over two days. Two sessions took place on each day with a minimum of two hours between sessions to allow a washout period to take place and avoid potential learning effects such as participants remembering the pain ratings of previous sessions for TS [37]. There was a minimum of 48 hours and a maximum of seven days between the two testing days. Six sites in total were tested in each participant including upper limb (extensor carpi radialis muscle belly), lower limb (tibialis anterior muscle belly) and spinal (lumbar erector spinae); all tested bilaterally. A combination of upper and lower limb and spinal sites were chosen to test within the asymptomatic population in an attempt to reflect the range of potential sites which maybe encountered for the acute musculoskeletal trauma participants [5]. Each site was marked with a Schuco Surgical Skin Marking Pen at the first session to allow the same sites to be tested in each session. The same six sites were used at each session for every participant.

Acute musculoskeletal trauma inter-rater study. A total of two sessions on the same day with a minimum of two hours between sessions to allow a washout period to take place and avoid potential learning effects was observed [37]. Due to the changeable nature of symptoms following musculoskeletal trauma, together with known sensory changes i.e. peripheral sensitisation following injury to tissue/inflammation, [43], intra-rater reliability testing on a different day was deemed unsuitable due to the likelihood of capturing change of symptoms and sensory changes over time rather than assessing reliability of testers. Due to patient burden, it was deemed unnecessary to test six sites as per the asymptomatic participants. Additionally, these sites would not have been applicable to the injuries in which the participants sustained. Therefore a total of two sites were tested for each participant—a local site within or close to the same dermatome of the injury and a remote site which was on the opposite side to the injury where possible. The choice of both the local and remote site depended on injuries the participant had sustained resulting from trauma and therefore could not be standardised across all participants [40]. Each site was marked with a Schuco Surgical Skin Marking Pen at the first session, and the same sites were used for both sessions but were individual to each participant.

Testing procedure. The testing procedure for both sets of participants was standardised. For both HPT, CPT and PPT, a method of limits design [44] was used. For each modality, two measurements were taken per session and the mean calculated. A minimum of 30 seconds was observed between each measurement. Within the first session a period of familiarisation was conducted prior to testing before each modality [40].

Thermal testing. Thermal testing was tested using skin contact stimulation. For HPT and CPT temperature either increased or decreased from a baseline of 32°C at a rate of 1°C/second until the maximum temperature of 50.5°C for HPT and 0°C for CPT had been reached in which the temperature then automatically returned to baseline. The participants were instructed to press a button when the sensation changes from a warm/cold sensation to one of pain. Once the button was pressed, the temperature returned to baseline at the same rate of 1°C/Second [44].

Pressure pain threshold and temporal summation testing. For the PPT testing, the tip of the force gauge was applied perpendicular to the skin prior to testing and then pressure was applied at a constant rate of 5 newtons/second [19, 44, 45]. The participant was instructed to press a button when the pressure sensation changed to pain. Upon pressing the button, the pressure was released immediately.

For TS testing, the mean PPT score from that session was used. Ten consecutive pulses were applied at this threshold level. The pressure was increased over five seconds, held for one second before being released immediately, with a one second interval before the next pulse [46, 47]. At each peak pulse, the participant was asked to rate their pain on an NRS of 0–10 with 0 being no pain and 10 being pain as worse as it could be.

Order of testing

To minimise learning effects and potential bias, the order of modalities, site and raters were randomised. Due to the nature of TS testing, and the requirement of session PPT scores, this was always the last test performed. For consistency, one rater performed all of the site markings for each participant prior to testing commencing.

Data analysis

For measures of reliability, intraclass correlation coefficients, two-way mixed effects, absolute agreement (ICC 3,2) for PPT, CPT and HPT and ICC [3,10] for TS, with 95% confidence intervals (CIs) were calculated. The ICC model was chosen *a priori* with a two-way mixed effects model (model 3) chosen as the raters in the study were only raters of interest, and an average measurement as a mean of more than one measurement was calculated [48]. Bland Altman Plots were calculated for limits of agreement (LOA) as well as standard error of measurement (SEM) calculated. For interpretation of the Bland Altman Plots in terms of good agreement, the majority of the points were required to be within the 95% limits of agreement, with an even distribution of points on both sides of the mean difference to indicate no systematic bias [49]. The descriptive data for CPT, HPT and PPT are reported as the mean of two measurements (SD). CPT and HPT are reported in degrees Celsius, and PPT in Newtons. For TS analyses and to capture the variation in the pain ratings over the 10 pulses, the mean NRS of pulses 1 to 4 (M1), 5 to 7 (M2), and 8 to 10 (M3) was calculated as well as the ratio between M3 and M1 [20]. In order to calculate the ratio where a number greater than 0 was required, 0.1 was added to all temporal summation scores. SPSS (version 25) was used for analyses. For interpretation of ICC results the following criteria were used: <0.5 = poor, 0.5–0.74 = moderate, 0.75–0.9 = Good and >0.91 = Excellent [50, 51].

Results

Participant demographics

Asymptomatic participants. A total of 21 participants were recruited. One participant was excluded from data analyses due to extra-ordinarily high ratings on all measures, the majority of which reached the safety limits of the testing equipment. The sample size was reduced to 20 participants (50% male) for analysis. Mean age was 27.55 (Standard Deviation (SD) 8.06).

Acute musculoskeletal trauma participants. A total of 25 participants were recruited. Five participants were excluded from data analysis for the following reasons: slow reaction times with an inability to press the button when pain thresholds were reached ($n = 2$), second data set not collected due to reduction in mental capacity ($n = 1$), or discharged from hospital

Table 1. Patient characteristics for both asymptomatic and musculoskeletal trauma participants.

	Asymptomatic	Acute Musculoskeletal Trauma
Age, mean (SD)	27.55 (8.06)	44.8 (19.32)
Sex, male n (%)	10 (50%)	14 (70%)
Height, cm (SD)	172.72 (9.50)	170.4 (12.96)
Weight, kg (SD)	72.89 (15.62)	77.45 (14.67)
Body Mass Index (SD)	24.32 (4.71)	26.61 (3.99)

<https://doi.org/10.1371/journal.pone.0233521.t001>

before a full data set had been collected ($n = 2$). The mean age of participants (70% male) included in data analyses was 44.8 (SD 19.32). Participant demographics are summarised in [Table 1](#) with injury characteristics for the acute musculoskeletal trauma participants summarised in [Table 2](#).

Intra-class correlation coefficients and standard error of measurement

Asymptomatic participants. Heat pain thresholds. Intra-rater reliability at all sites was good or excellent (ICC range 0.76, 0.91). Inter-rater reliability for both days of testing at all sites was good or excellent (ICC range 0.83, 0.95). The average SEM for rater 1 across all sites was 0.97°C for day 1 and 0.84°C for day 2. For rater 2, the average SEM was similar at 0.89°C for day 1 and 0.81°C for day 2. All ICCs with 95% CIs, SEM and means (SD) for all sites for HPT are reported in [Table 3](#).

Cold pain thresholds. Ten participants were excluded from cold data analyses as pain was not evoked within the safety limit of the thermal testing equipment in three or more sessions for each site. Of the remaining 10 participants, only two participants had a full data set for all sessions and sites. Eight participants had a partial data set where either intra or inter-rater reliability could be calculated. Due to low number of participants included in data analysis, it was deemed insufficient to perform ICCs and SEM, therefore descriptive data only is reported. A summary of means and SDs can be found in [Table 4](#).

Pressure pain thresholds. For intra-rater reliability, rater 1 demonstrated a wide range of ICC scores which ranged from poor to good reliability (ICC range 0.41, 0.83), with the lower limb sites and one upper limb site demonstrating lower ICCs compared to spinal sites. Rater 2 demonstrated higher intra-rater reliability (ICC range 0.70, 0.88), although demonstrated wide CIs. Inter-rater reliability for day 1 of testing was rated as moderate to excellent (ICC range 0.66, 0.92) depending on site, with day 2 demonstrated improved inter-rater reliability rated as good to excellent (ICC range 0.80, 0.93). The average SEM across all sites for rater 1 was 7.12 newtons (N) for day one and 5.96 for day 2. For rater 2, the average SEM was slightly lower at 6.93N for day one and 5.98N for day two. ICCs with 95% CIs, SEM and means for all sites for PPT are reported in [Table 5](#).

Table 2. Injury characteristics for acute musculoskeletal participants.

Number of Injuries	Single Injury	10
	Multiple Injury	10
Location of Injury	Upper Extremity	1
	Lower Extremity	16
	Trunk	1
	Upper and Lower Extremity	2
Type of Injury	Fractures	19
	Soft Tissue	1

<https://doi.org/10.1371/journal.pone.0233521.t002>

Table 3. Intra-class correlations, means and standard error of measurement for heat pain thresholds for the asymptomatic participants.

Site	ICC (95% CI) Intra-Rater		ICC (95% CI) Inter-Rater		SEM Day 1 (°C)		SEM Day 2 (°C)		Mean (SD) Day 1 (°C)		Mean (SD) Day 2 (°C)	
	Rater 1	Rater 2	Day 1	Day 2	Rater 1	Rater 2	Rater 1	Rater 2	Rater 1	Rater 2	Rater 1	Rater 2
ULR	0.76 (0.41, 0.90)	0.88 (0.71, 0.95)	0.88 (0.70, 0.95)	0.94 (0.85, 0.98)	1.13	1.10	0.76	0.74	44.75 (3.27)	45.19 (3.18)	45.70 (3.08)	45.48 (3.01)
ULL	0.82 (0.56, 0.93)	0.91 (0.77, 0.97)	0.90 (0.74, 0.96)	0.91 (0.76, 0.96)	1.34	1.04	0.99	1.09	44.10 (4.18)	44.58 (3.23)	45.00 (3.20)	45.35 (3.54)
LLR	0.78 (0.43, 0.91)	0.85 (0.42, 0.95)	0.95 (0.88, 0.98)	0.88 (0.61, 0.96)	0.59	0.54	0.94	0.72	44.97 (2.67)	44.97 (2.43)	45.18 (2.67)	46.05 (2.05)
LLL	0.76 (0.41, 0.91)	0.90 (0.74, 0.96)	0.83 (0.57, 0.93)	0.94 (0.86, 0.98)	0.97	0.95	0.52	0.59	45.00 (2.32)	45.51 (2.28)	45.84 (2.19)	45.71 (2.47)
SR	0.83 (0.57, 0.93)	0.91 (0.78, 0.97)	0.92 (0.81, 0.97)	0.91 (0.76, 0.96)	0.71	0.67	0.89	0.88	43.13 (2.54)	43.51 (2.41)	44.03 (2.92)	43.37 (2.90)
SL	0.77 (0.42, 0.91)	0.94 (0.85, 0.98)	0.84 (0.61, 0.94)	0.92 (0.80, 0.97)	1.07	1.04	0.91	0.83	42.96 (2.70)	43.35 (2.63)	44.07 (3.20)	43.63 (2.93)

ULR, Upper Limb Right; ULL, Upper Limb Left; LLR, Lower Limb Right; LLL, Lower Limb Left; SR, Spine Right; SL, Spine Left; ICC, Intraclass correlation; CI, Confidence Interval; SD, Standard Deviation; SEM, Standard Error of Measurement

<https://doi.org/10.1371/journal.pone.0233521.t003>

Temporal summation. Table 6 summarises all ICCs for M1, M2, M3 and ratios (M3/M1). A number of participants asked for temporal summation testing to stop due to high pain levels therefore a full set of NRS were not collected. Where a full set of 10 NRS were not collected, these were excluded from the analyses with Table 6 summarising number of participants included in analyses. For intra-rater reliability, ICCs for the three means (M1, M2, M3) ranged

Table 4. Asymptomatic participant’s descriptive data results for cold pain threshold.

Site	Means (SD) (°C) Intra-Rater		Site	Means (SD) (°C) Inter-Rater		Site	Mean (SD) all participants Day 1 (°C)		Mean (SD) all participants Day 2 (°C)	
	Rater 1	Rater 2		Day 1	Day 2		Rater 1	Rater 2	Rater 1	Rater 2
ULR	N = 7	N = 6	ULR	N = 9	N = 7	ULR	N = 9	N = 9	N = 8	N = 7
Session 1	18.55 (10.71)	18.06 (9.09)	Rater 1	15.49 (11.10)	16.50 (8.52)		15.49 (11.10)	14.16 (9.44)	14.81 (9.23)	14.23 (8.95)
Session 2	6.03 (9.26)	15.60 (8.97)	Rater 2	14.16 (9.44)	14.23 (8.95)					
ULL	N = 8	N = 6	ULL	N = 8	N = 6	ULL	N = 9	N = 8	N = 8	N = 6
Session 1	18.87 (10.50)	20.43 (8.54)	Rater 1	19.24 (9.93)	20.60 (7.77)		17.54 (10.59)	16.81 (9.85)	16.30 (10.34)	20.23 (8.76)
Session 2	16.30 (10.34)	20.23 (8.76)	Rater 2	16.81 (9.85)	20.23 (8.76)					
LLR	N = 4	N = 4	LLR	N = 5	N = 4	LLR	N = 6	N = 5	N = 5	N = 5
Session 1	21.53 (7.45)	23.23 (1.97)	Rater 1	20.19 (9.57)	18.53 (10.79)		18.68 (9.32)	19.47 (8.58)	15.03 (12.18)	18.71 (9.44)
Session 2	18.04 (11.73)	22.78 (2.90)	Rater 2	19.46 (8.58)	17.78 (10.63)					
LLL	N = 7	N = 6	LLL	N = 7	N = 6	LLL	N = 7	N = 7	N = 7	N = 6
Session1	17.71 (7.89)	17.90 (7.35)	Rater 1	17.71 (7.89)	18.16 (6.44)		17.71 (7.89)	15.55 (9.14)	16.14 (7.94)	18.75 (6.18)
Session 2	16.14 (7.94)	18.75 (6.18)	Rater 2	15.55 (9.14)	18.75 (6.18)					
SR	N = 7	N = 6	SR	N = 6	N = 6	SR	N = 7	N = 6	N = 7	N = 6
Session 1	22.78 (8.44)	23.40 (4.76)	Rater 1	25.62 (4.20)	21.65 (6.75)		22.78 (8.44)	23.40 (4.76)	19.78 (7.90)	20.50 (6.71)
Session 2	19.78 (7.90)	20.50 (6.71)	Rater 2	23.40 (4.76)	20.50 (6.71)					
SL	N = 6	N = 5	SL	N = 6	N = 5	SL	N = 6	N = 6	N = 6	N = 5
Session 1	22.48 (6.92)	20.67 (9.08)	Rater 1	22.48 (6.92)	19.45 (7.35)		22.48 (6.92)	19.83 (8.38)	17.36 (8.33)	18.06 (8.51)
Session 2	17.36 (8.33)	18.06 (8.51)	Rater 2	19.83 (8.38)	18.06 (8.51)					

ULR, Upper Limb Right; ULL, Upper Limb Left; LLR, Lower Limb Right; LLL, Lower Limb Left; SR, Spine Right; SL, Spine Left; N, number of participants; SD, Standard Deviation

<https://doi.org/10.1371/journal.pone.0233521.t004>

Table 5. Intraclass correlations, means and standard error of measurement for pressure pain thresholds for the asymptomatic participants.

Site	ICC (95% CI) Intra-Rater		ICC (95% CI) Inter-Rater		SEM Day 1 (N)		SEM Day 2 (N)		Mean (SD) Day 1 (N)		Mean (SD) Day 2 (N)	
	Rater 1	Rater 2	Day 1	Day 2	Rater 1	Rater 2	Rater 1	Rater 2	Rater 1	Rater 2	Rater 1	Rater 2
ULR	0.78 (0.44, 0.91)	0.75 (0.35, 0.90)	0.66 (0.15, 0.86)	0.93 (0.83, 0.97)	7.14	6.64	3.77	4.15	33.70 (12.17)	36.69 (11.32)	38.62 (14.68)	37.80 (16.17)
ULL	0.65 (0.11, 0.86)	0.70 (0.27, 0.88)	0.66 (0.14, 0.87)	0.80 (0.49, 0.92)	7.35	4.20	6.45	5.65	33.07 (12.61)	31.30 (7.20)	35.53 (14.35)	33.71 (12.57)
LLR	0.41 (-0.43, 0.77)	0.77 (0.43, 0.91)	0.79 (0.48, 0.92)	0.81 (0.51, 0.92)	9.91	9.11	9.00	7.82	44.97 (21.53)	50.77 (19.78)	51.80 (20.48)	54.13 (17.80)
LLL	0.52 (-0.11, 0.80)	0.88 (0.68, 0.95)	0.75 (0.38, 0.90)	0.90 (0.75, 0.96)	7.11	9.14	5.26	5.56	43.64 (14.17)	49.49 (18.22)	50.59 (16.80)	50.29 (17.76)
SR	0.83 (0.57, 0.93)	0.86 (0.64, 0.94)	0.92 (0.77, 0.97)	0.92 (0.74, 0.97)	6.03	6.77	5.42	6.87	48.11 (21.32)	54.12 (23.93)	54.33 (19.04)	60.72 (24.14)
SL	0.76 (0.41, 0.91)	0.87 (0.32, 0.96)*	0.90 (0.76, 0.96)	0.93 (0.74, 0.98)*	5.15	5.71	5.87	5.80	48.62 (16.63)	51.67 (18.42)	55.72 (23.25)	60.92 (22.57)

ULR, Upper Limb Right; ULL, Upper Limb Left; LLR, Lower Limb Right; LLL, Lower Limb Left; SR, Spine Right; SL, Spine Left; ICC, Intraclass correlation; CI, Confidence Interval; SD, Standard Deviation; SEM, Standard Error of Measurement; * number of participants = 19 N; Newtons

<https://doi.org/10.1371/journal.pone.0233521.t005>

from poor to excellent (ICC range 0.20, 0.91) with the lower limb sites overall showing a wider range in ICCs compared to other sites. For inter-rater reliability, day one and day two ICCs for the means were all rated poor to excellent (ICC range 0.30, 0.92). For the ratio ICCs both intra and inter-rater were rated as poor or moderate (ICC range -0.31, 0.66) with the exception of right upper limb site which was rated as good (ICC 0.86).

Acute musculoskeletal trauma participants. *Heat pain thresholds.* A total of three sets of data for the remote site were excluded from analyses due to pain not being evoked within the safety limit of the thermal testing device (50.5⁰). No sets of data were excluded for the local site analyses. Both local (ICC 0.86 CI 0.65, 0.95) and remote (ICC 0.81 CI 0.49, 0.93) sites were rated as good for HPT. For SEM, rater 1 for the local site was 1.02°C and less for the remote site at 0.85°C. For rater 3, both sites were similar to rater 1 with the local site SEM being 1.11°C, and remote site 1.24°C. ICCs with 95% CIs, SEM and means for both local and remote sites for HPT are reported in [Table 7](#).

Cold pain thresholds. Thirteen participants were excluded from cold data analyses as pain was not evoked within the safety limit of the thermal testing equipment for the local site and 17 participants excluded for the remote site. Therefore, a total of 7 participants for the local site and 3 for the remote site were included in analyses. Due to low numbers of participants included in data analyses, descriptive data only are reported. A summary of the descriptive data for CPT can be found in [Table 7](#).

Pressure pain thresholds. Both local (ICC 0.78 CI 0.43, 0.91) and remote (ICC 0.76 CI 0.43, 0.91) were rated as good. The SEM for rater 1 was 6.04N for the local site and 6.95N for the remote site. For rater 3, SEM was similar to rater 1 with the 6.07N for the local site and 7.45N for the remote site. ICCs with 95% CIs, SEM and means for both local and remote sites for PPT are reported in [Table 7](#).

Temporal summation. Three sets of data were excluded from analyses for the local site and two sets of data for the remote site due to the participant asking for testing to stop due to high pain levels therefore 10 pulses were not achieved. When calculating the ratio both local (ICC 0.57 CI -0.07, 0.84) and remote (ICC 0.62 CI -0.04, 0.86) were rated as moderate. However, when calculating the 3 means over the 10 pulses, for the local site M1 (ICC 0.87 CI 0.63, 0.95) and M3 (ICC 0.79 CI 0.45, 0.92) were rated as good whereas M2 (ICC 0.69 CI 0.13, 0.89) was rated as moderate. For the remote site all three means were rated as excellent (M1 ICC 0.91 CI

The Bland Altman plots for CPT (Fig 2) show more variation in limits of agreement with Rater 1 intra-rater average mean difference at 2.95°C whereas rater 2 mean difference at 1.30°C. Inter-rater agreement between raters was improved on day 2 compared to 1. Although more variation in agreement, no systematic or proportional bias was demonstrated.

The Bland Altman plots for PPT (Fig 3) show similar limits of agreement for both raters and both days. For both intra and inter rater agreement, less agreement appears to be consistent in higher pain thresholds compared to lower pain thresholds, however no systematic or proportional bias is evident.

Fig 4 depicts the Bland Altman plots and limits of agreement for HPT, CPT and PPT for the acute musculoskeletal trauma participants. All plots show good agreement between raters with the average mean difference no more than 1°C or 1N. For HPT, higher temperature scores are associated with rater 3 compared to rater 1, and further investigation into the dataset showed this to be more for the remote site than local site. No other proportional or systematic bias was shown in for PPT and CPT.

Discussion

This is the first methodologically rigorous reliability study evaluating intra and inter-rater reliability within asymptomatic participants and inter-rater reliability for acute musculoskeletal

Table 7. Intraclass correlations, means and standard error of measurement for heat and cold pain thresholds, pressure pain thresholds and temporal summation for the acute musculoskeletal trauma participants.

Site	Modality	Number of participants included in analyses	Rater 1 Mean (SD)	Rater 3 Mean (SD)	ICC (95% CI)	Standard Error of Measurement	
						Rater 1	Rater 3
Local Site	HPT °C	N = 20	46.92 (2.72)	46.59 (2.97)	0.86 (0.65, 0.95)	1.02	1.11
	CPT °C	N = 7	13.53 (8.87)	12.90 (6.80)	N/A	N/A	N/A
	PPT Newtons	N = 20	37.60 (12.87)	36.26 (12.94)	0.78 (0.43, 0.91)	6.04	6.07
	Mean 1	N = 17	5.25 (1.88)	5.35 (1.79)	0.87 (0.63, 0.95)	N/A	N/A
	Mean 2	N = 17	5.16 (1.49)	5.10 (2.03)	0.69 (0.13, 0.89)	N/A	N/A
	Mean 3	N = 17	5.59 (1.75)	5.00 (2.14)	0.79 (0.45, 0.92)	N/A	N/A
	Ratio between Mean 3 and Mean 1 (NRS)	N = 17	1.15 (0.46)	0.96 (0.31)	0.57 (-0.07, 0.84)	N/A	N/A
Remote Site	HPT °C	N = 17	47.75 (1.95)	47.06 (2.85)	0.81 (0.49, 0.93)	0.85	1.24
	CPT °C	N = 3	11.84 (5.57)	7.37 (1.46)	N/A	N/A	N/A
	PPT Newtons	N = 20	43.36 (14.18)	42.22 (15.20)	0.76 (0.43, 0.91)	6.95	7.45
	Mean 1 (NRS)	N = 18	4.96 (2.28)	5.35 (2.47)	0.91 (0.77, 0.97)	N/A	N/A
	Mean 2 (NRS)	N = 18	5.38 (2.46)	5.75 (2.54)	0.94 (0.84, 0.98)	N/A	N/A
	Mean 3 (NRS)	N = 18	5.47 (2.62)	5.90 (2.51)	0.93 (0.82, 0.97)	N/A	N/A
	Ratio between Mean 3 and Mean 1 (NRS)	N = 18	1.11 (0.31)	1.13 (0.32)	0.62 (-0.04, 0.86)	N/A	N/A

HPT, heat pain threshold; CPT, cold pain threshold; PPT, pressure pain threshold; NRS, numerical rating scale; SD, standard deviation; ICC, Intraclass correlation coefficient

<https://doi.org/10.1371/journal.pone.0233521.t007>

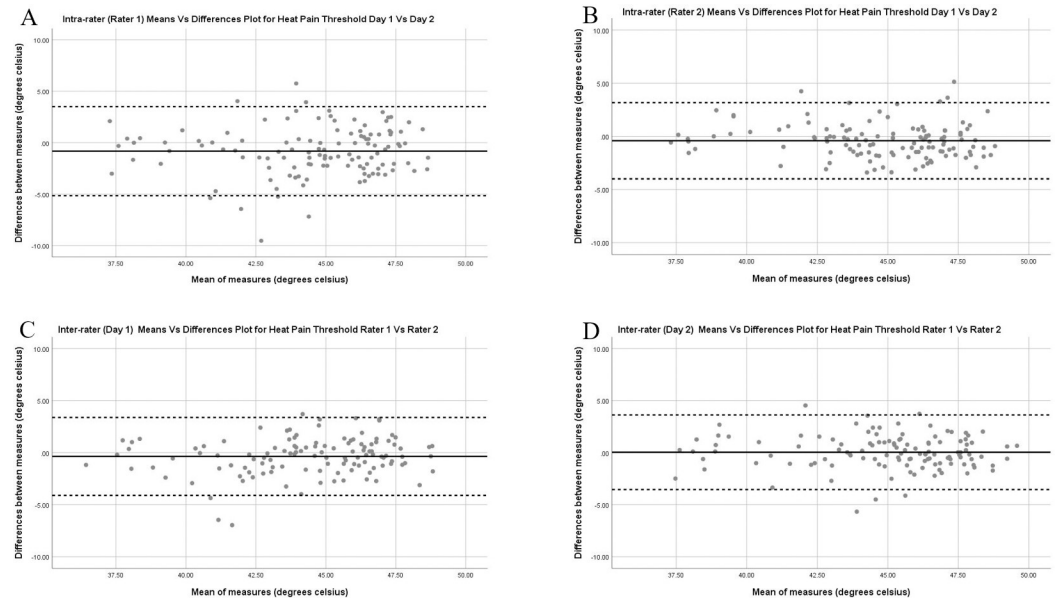


Fig 1. Intra-rater (A & B) and Inter-rater (C & D) Bland Altman Plots with limits of agreement for heat pain thresholds for the asymptomatic participants. Limits of agreement are presented as the dotted lines with the mean difference illustrated by the black line.

<https://doi.org/10.1371/journal.pone.0233521.g001>

trauma participants. In asymptomatic participants, reliability was established as moderate to good intra-rater reliability for HPT and PPT, with the exception of one lower limb site for PPT which was rated poor, and moderate to excellent inter-rater reliability for PPT and HPT. Good to excellent inter-rater reliability of HPT and PPT was demonstrated for acute musculoskeletal

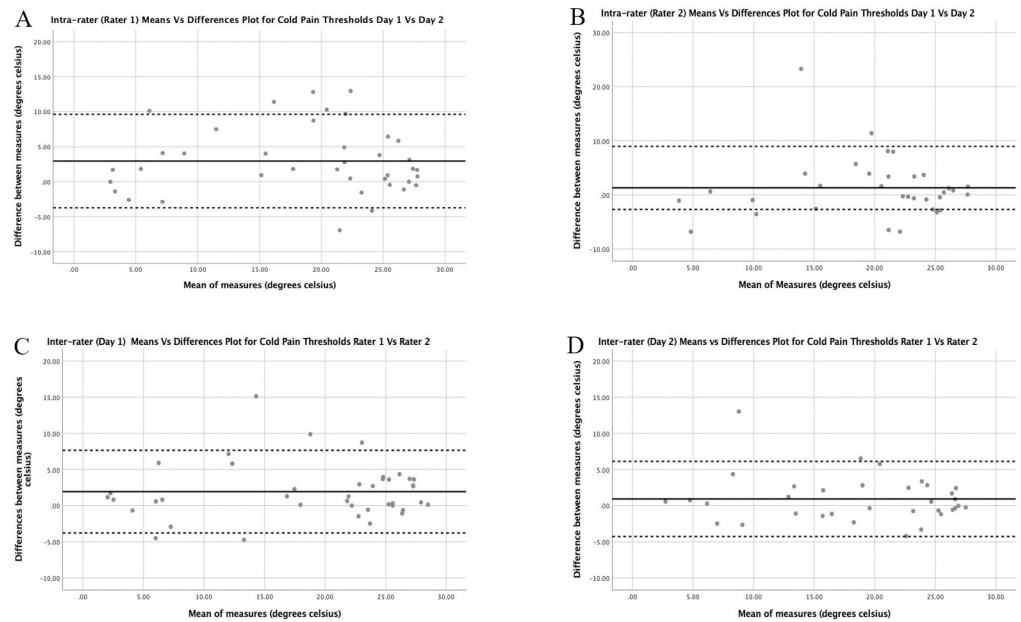


Fig 2. Intra-rater (A & B) and Inter-rater (C & D) Bland Altman Plots with limits of agreement for cold pain thresholds for the asymptomatic participants. Limits of agreement are presented as the dotted lines with the mean difference illustrated by the black line.

<https://doi.org/10.1371/journal.pone.0233521.g002>

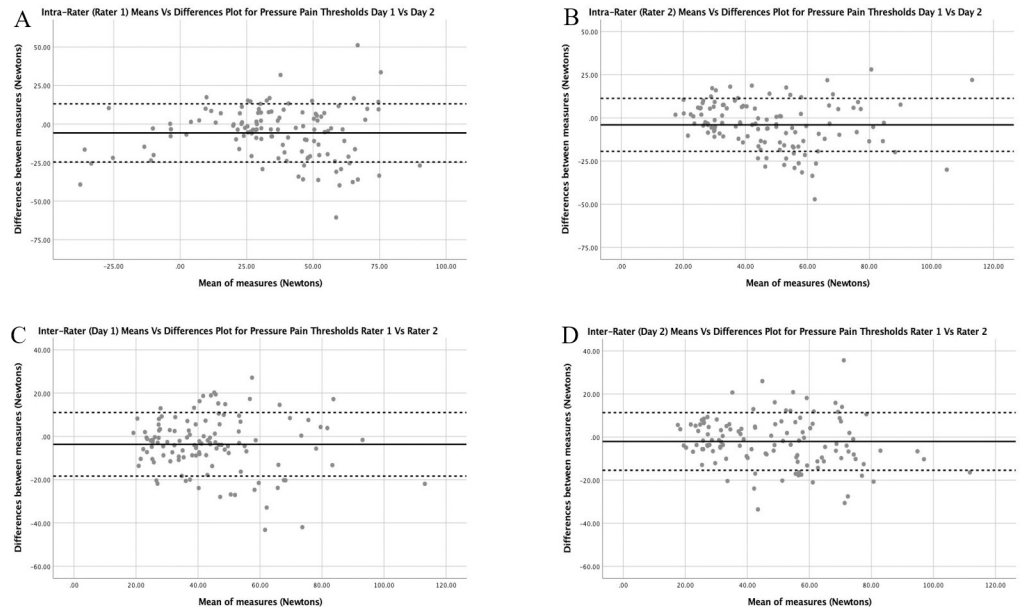


Fig 3. Intra-rater (A & B) and Inter-rater (C & D) Bland Altman Plots with limits of agreement for pressure pain thresholds for the asymptomatic participants. Limits of agreement are presented as the dotted lines with the mean difference illustrated by the black line.

<https://doi.org/10.1371/journal.pone.0233521.g003>

trauma participants. Bland Altman plots supported these findings with good agreement between days and raters for HPT and PPT, however agreement for PPT was less with higher PPT scores (i.e. more pressure exerted by the rater) compared to lower scores (less pressure exerted by the rater) more so in the asymptomatic participants. SEM for both HPT and PPT

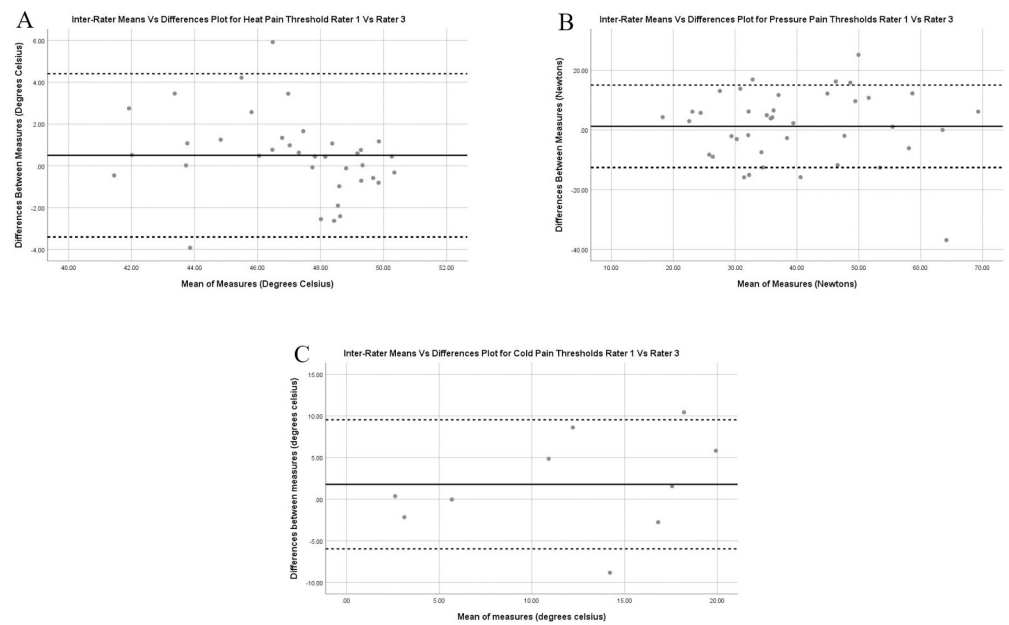


Fig 4. Inter-rater Bland Altman Plots with limits of agreement for heat pain thresholds (A), pressure pain thresholds (B) and cold pain thresholds (C) for the acute musculoskeletal trauma participants. Limits of agreement are presented as the dotted lines with mean difference the black line.

<https://doi.org/10.1371/journal.pone.0233521.g004>

was low, with PPT SEM in asymptomatic participants having more variability depending on the testing site. TS demonstrated variable results ranging from poor to excellent reliability in asymptomatic participants for intra and inter-rater reliability, but moderate to excellent reliability for acute musculoskeletal trauma participants. However, some participants in both groups asked for testing to stop due to high pain intensity during TS. Limited conclusions can be made for CPT since no statistical analyses could be performed yet, Bland Altman plots show adequate agreement in both sets of participants.

No previous study has evaluated the reliability of these specific QST tests in combination, however, results in this study are consistent with studies evaluating thermal and pressure testing as stand-alone modalities, [19, 22, 26, 52, 53] as well as recent studies combining multiple QST tests [27] and studies using the established DNFS protocol for QST [24, 25].

Although the current study demonstrated good reliability for PPT in both sets of participants, musculoskeletal trauma participants showed higher inter-rater reliability compared to asymptomatic participants. This was also observed for TS testing, where results for asymptomatic participants were more variable compared to the musculoskeletal trauma participants. A possible reason for the variation compared to previous studies could be due to the different algometers used across studies including a digital algometer [29, 53] or a mechanical algometer [19]. The algometer used in the current study was a digital handheld algometer, with an additional software programme used to guide rate of pressure and to accurately calculate the temporal summation rate for each participant. Because of this, raters were required to focus on the computer screen and not on the site during testing, which could have possibly led to variation in the technique applied, although both raters had substantial training on the method prior to commencement of the study. This could also explain why TS reliability results were variable given that same method of measuring PPT was used for TS testing.

This algometer however, was used in both sets of participants and does not explain differences in results between the asymptomatic and acute musculoskeletal trauma participants. For the asymptomatic participants, the testing protocol was extensive using six sites with four modalities which totals to 24 tests all focusing on pain thresholds. It is possible the quantity of testing caused habituation or sensitisation to testing thus having the potential for their pain thresholds to increase or decrease between sessions [54–56]. This could have the potential to explain the variability in TS results which was not observed for musculoskeletal trauma participants. Research in this area is limited, with a handful of studies evaluating the habituation/sensitisation effect [54–56]. These studies however focus on thermal testing only, over a number of days with testing every day, which differs from the multiple modalities and testing schedule used in this study. However, other protocols such as the DNFS protocol have extensive QST testing [57]. Although the DNFS protocol is extensive, all tests are not pain threshold tests but combine detection thresholds differing to the protocol for this study. Future research is required in this area to evaluate the effects of QST and the effects of habituation and sensitisation. Despite this, future studies should also ensure training of raters prior to testing, as this study has highlighted good inter-rater reliability can be achieved both in asymptomatic and acute musculoskeletal trauma participants.

Despite the differences in PPT and TS, HPT was consistently rated good or excellent for both asymptomatic and acute musculoskeletal trauma participants, with SEM for HPT being approximately 1°C for both sets of participants. However, across all modalities a trend was observed whereby both SEM and ICCs on day 2 of testing of the asymptomatic participants increased. This could indicate a potential learning effect has taken place. The raters received training before undertaking the study and since no systematic or proportional bias is demonstrated in the Bland Altman plots for the asymptomatic participants, it is unlikely there is a learning effect from the raters which could contribute to the variation in results, however

despite a period of test familiarisation for the participants prior to testing, a learning effect could have taken place which could explain increased reliability on day 2. Furthermore, a specific testing protocol was adopted for this study with a minimum of 2 hours between sessions observed to allow a period of rest for the participant and to minimise any potential bias and learning effect, with 48 hours given between testing days for the asymptomatic participants. Previous studies have used a range of times from five minutes to ten minutes for within day testing between different raters [19, 29] and up to four months for between day testing for intra-rater reliability [24]. With no consensus on the amount of time between sessions, the larger time period between sessions and days of testing in this study may have contributed to differences between days of testing. Future studies should take into consideration a potential learning effect and consider a familiarisation session on a different day rather than immediately before the testing to account for this. No short-term learning effect was observed or indicated for within day testing for the asymptomatic or acute musculoskeletal trauma participants.

Taking into account the results of the reliability studies conducted, and the confounding factors which can influence the reliability of measures such as time between testing sessions, number of sites tested, adequate training of raters and the position and methods of the algometer for PPT and TS testing, this study sets the foundations for future studies to utilise QST measures within clinical research. This can include investigating if early evidence of sensitisation relates to how well people recover from injury. However, it is important for studies to acknowledge the potential factors which can influence results of QST testing and factor this into study design to ensure a change in pain thresholds is captured rather than measurement error.

Strengths and limitations

This is the first study to evaluate intra and inter-reliability of CPT, HPT, PPT and TS combined, and the first study to evaluate inter-rater reliability in an acute hospital setting for a musculoskeletal trauma population. Establishing reliability for these measures allows future studies to evaluate sensory changes from an acute stage of injury over time as people recover from traumatic musculoskeletal injuries and further demonstrate the value of QST in predicting outcome [40]. Additionally, previous reliability studies have controlled the environment in which testing was conducted, often within a lab setting. Although the asymptomatic sample was conducted within the lab, the environment was not controlled as strictly as previous studies where temperature and background noise were controlled. This was deliberate, as there is a considerable amount of noise, distraction and variance in temperature within a hospital setting, and reliability of these measures needed to be established in these real-life clinical conditions in order to be transferable to future studies.

Nevertheless, there are some limitations to this study. One limitation to this study is we had low participant numbers for CPT due to the safety limit of the thermal analyzer, making overall conclusions to the reliability of CPT difficult. One study has acknowledged similar issues with participants reaching the safety limit data [22] and excluded this data from analysis. The decision was made to exclude this data in the current study, as if included, the ICCs would likely portray excellent reliability and agreement, which could be misleading. Although conclusions of the reliability of CPT cannot be made from this study, the reporting of the issues with the safety limit of the thermal analyzer is important to document. Even though the device was calibrated correctly and the application followed protocols described in previous studies, issues were still observed. This has implications for future studies using CPT as well as future studies required to establish reliability of CPT with adequate sample sizes.

Secondly, due to high pain levels experienced in some participants for TS testing, sample sizes for some sites were lower than 19. Therefore, some ICC's were slightly underpowered but had adequate numbers to perform statistical analysis [37, 38]. This should be taken into consideration when interpreting the results of the TS testing. Although we are less confident in the ICC reported, it does not change the overall conclusions of this study that the TS protocol requires further testing and evaluation.

Another limitation of this study was inter-rater reliability but not intra-rater reliability was evaluated within the musculoskeletal trauma population. Due to the acute nature of musculoskeletal trauma, assessing reliability over a number of days was difficult, as symptoms and pain are often changeable making it problematic to assess reliability rather than change of symptoms. Although intra-rater could have been tested on the same day, the decision was made to evaluate inter-rater reliability, as intra-rater reliability in the asymptomatic participants was moderate to good depending on modality, and inter-rater reliability is clinically relevant for future studies planned [40].

Finally, although this study has demonstrated good reliability for the battery of QST tests used, the ICC statistical model (two-way mixed effects, absolute agreement 3,2) was chosen as the particular raters used in this study were of interest, therefore the more common ICC statistical model of two-way random effects utilised when the raters of interest are selected from random and therefore can be generalised to a wider population of raters with similar characteristics was not appropriate for this study [50]. Therefore, generalising the study to larger populations and other raters should be interpreted with caution and further research is required to confirm these results.

Conclusion

This study has demonstrated moderate to good intra-rater reliability and moderate to excellent inter-rater reliability for HPT and PPT in a sample of asymptomatic individuals with good to excellent inter-rater reliability demonstrated for HPT and PPT in an acute musculoskeletal trauma population. Good reliability was demonstrated for TS within the acute musculoskeletal trauma population. Limited conclusions can be made of CPT data, however Bland Altman plots did show good agreement between days and raters. This study has demonstrated HPT, CPT, PPT and TS can be used in an acute clinical environment and forms the foundation for future work in acute musculoskeletal trauma population evaluating sensory function over time as well as the value of QST in predicting outcome post trauma.

Acknowledgments

With thanks to Dr Alessandro De Nunzio for development of the software used with the digital algometer, and to Pauline Kuithan and Andrew Sanderson for assistance with data collection.

Author Contributions

Conceptualization: Nicola Middlebrook, Nicola R. Heneghan, Alison Rushton, Deborah Falla.

Data curation: Nicola Middlebrook.

Formal analysis: Nicola Middlebrook, Deborah Falla.

Funding acquisition: Alison Rushton, Deborah Falla.

Investigation: Nicola Middlebrook, Nicola R. Heneghan, David W. Evans.

Methodology: Nicola Middlebrook, Nicola R. Heneghan, David W. Evans, Alison Rushton, Deborah Falla.

Project administration: Nicola Middlebrook, Deborah Falla.

Supervision: Nicola R. Heneghan, Alison Rushton, Deborah Falla.

Validation: Deborah Falla.

Visualization: Nicola Middlebrook, Alison Rushton, Deborah Falla.

Writing – original draft: Nicola Middlebrook, Deborah Falla.

Writing – review & editing: Nicola Middlebrook, Nicola R. Heneghan, David W. Evans, Alison Rushton, Deborah Falla.

References

1. Haagsma JA, Graetz N, Bolliger I, Naghavi M, Higashi H, Mullany EC, et al. The global burden of injury: incidence, mortality, disability-adjusted life years and time trends from the Global Burden of Disease study 2013. *Injury Prevention*. 2016; 22(1):3–18.
2. Berben SAA, Schoonhoven L, Meijs THJM, van Vugt AB, van Grunsven PM. Prevalence and Relief of Pain in Trauma Patients in Emergency Medical Services. *The Clinical Journal of Pain*. 2011; 27(7):587–92.
3. Gabbe BJ, Simpson PM, Harrison JE, Lyons RA, Ameratunga S, Ponsford J, et al. Return to work and functional outcomes after major trauma: who recovers, when, and how well? *Annals of surgery*. 2016; 263(4):623–32.
4. Holbrook TL, Anderson JP, Sieber WJ, Browner D, Hoyt DB. Outcome after Major Trauma: 12-Month and 18-Month Follow-Up Results from the Trauma Recovery Project. *Journal of Trauma and Acute Care Surgery*. 1999; 46(5):765–73.
5. Holbrook TL, Anderson JP, Sieber WJ, Browner D, Hoyt DB. Outcome after Major Trauma: Discharge and 6-Month Follow-Up Results from the Trauma Recovery Project. *Journal of Trauma and Acute Care Surgery*. 1998; 45(2):315–24.
6. Gabbe BJ, Sutherland AM, Hart MJ, Cameron PA. Population-based capture of long-term functional and quality of life outcomes after major trauma: the experiences of the Victorian State Trauma Registry. *Journal of Trauma and Acute Care Surgery*. 2010; 69(3):532–6.
7. Klyne DM, Moseley GL, Sterling M, Barbe MF, Hodges PW. Are signs of central sensitisation in acute low back pain a precursor to poor outcome? *The Journal of Pain*. 2019.
8. Boakye M, Harkema S, Ellaway PH, Skelly AC. Quantitative testing in spinal cord injury: overview of reliability and predictive validity. *Journal of neurosurgery: Spine*. 2012; 17(Suppl1):141–50.
9. Goldsmith R, Wright C, Bell SF, Rushton A. Cold hyperalgesia as a prognostic factor in whiplash associated disorders: a systematic review. *Manual therapy*. 2012; 17(5):402–10.
10. Fingleton C, Smart K, Moloney N, Fullen BM, Doody C. Pain sensitization in people with knee osteoarthritis: a systematic review and meta-analysis. *Osteoarthritis and Cartilage*. 2015; 23(7):1043–56.
11. Cruz-Almeida Y, Fillingim RB. Can Quantitative Sensory Testing Move Us Closer to Mechanism-Based Pain Management? *Pain Medicine*. 2014; 15(1):61–72.
12. Nie H, Graven-Nielsen T, Arendt-Nielsen L. Spatial and temporal summation of pain evoked by mechanical pressure stimulation. *European journal of pain*. 2009; 13(6):592–9.
13. Woolf CJ. Central sensitization: Implications for the diagnosis and treatment of pain. *PAIN*. 2011; 152(3):S2–S15.
14. Vardeh D, Mannion RJ, Woolf CJ. Toward a Mechanism-Based Approach to Pain Diagnosis. *The Journal of Pain*. 2016; 17(9):T50–T69.
15. Curatolo M, Petersen-Felix S, Arendt-Nielsen L, Giani C, Zbinden AM, Radanov BP. Central Hypersensitivity in Chronic Pain After Whiplash Injury. *The Clinical Journal of Pain*. 2001; 17(4):306–15.
16. Lemming D, Graven-Nielsen T, Sørensen J, Arendt-Nielsen L, Gerdle B. Widespread Pain Hypersensitivity and Facilitated Temporal Summation of Deep Tissue Pain in Whiplash Associated Disorder: an Explorative Study of Women. *Journal of Rehabilitation Medicine*. 2012; 44(8):648–57.
17. Scott D, Jull G, Sterling M. Widespread Sensory Hypersensitivity Is a Feature of Chronic Whiplash-Associated Disorder but not Chronic Idiopathic Neck Pain. *The Clinical Journal of Pain*. 2005; 21(2):175–81.

18. Sterling M, Jull G, Vicenzino B, Kenardy J. Sensory hypersensitivity occurs soon after whiplash injury and is associated with poor recovery. *Pain*. 2003; 104(3):509–17.
19. Chesterton LS, Sim J, Wright CC, Foster NE. Interrater reliability of algometry in measuring pressure pain thresholds in healthy humans, using multiple raters. *The Clinical journal of pain*. 2007; 23(9):760–6.
20. Graven-Nielsen T, Vaegter HB, Finocchietti S, Handberg G, Arendt-Nielsen L. Assessment of musculoskeletal pain sensitivity and temporal summation by cuff pressure algometry: a reliability study. *Pain*. 2015; 156(11):2193–202.
21. Cathcart S, Winefield AH, Rolan P, Lushington K. Reliability of temporal summation and diffuse noxious inhibitory control. *Pain Research and Management*. 2009; 14(6):433–8.
22. Moloney NA, Hall TM, O'Sullivan TC, Doody CM. Reliability of thermal quantitative sensory testing of the hand in a cohort of young, healthy adults. *Muscle & Nerve*. 2011; 44(4):547–52.
23. Cathcart S, Pritchard D. Reliability of pain threshold measurement in young adults. *The journal of headache and pain*. 2006; 7(1):21–6.
24. Marcuzzi A, Wrigley PJ, Dean CM, Adams R, Hush JM. The long-term reliability of static and dynamic quantitative sensory testing in healthy individuals. *Pain*. 2017; 158(7):1217–23.
25. Nothnagel H, Puta C, Lehmann T, Baumbach P, Menard MB, Gabriel B, et al. How stable are quantitative sensory testing measurements over time? Report on 10-week reliability and agreement of results in healthy volunteers. *Journal of pain research*. 2017; 10:2067.
26. Knutti I, Suter M, Opsommer E. Test–retest reliability of thermal quantitative sensory testing on two sites within the L5 dermatome of the lumbar spine and lower extremity. *Neuroscience letters*. 2014; 579:157–62.
27. Wyld V, Palmer S, Learmonth ID, Dieppe P. Test–retest reliability of Quantitative Sensory Testing in knee osteoarthritis and healthy participants. *Osteoarthritis and Cartilage*. 2011; 19(6):655–8.
28. Geber C, Klein T, Azad S, Birklein F, Gierthmühlen J, Hüge V, et al. Test-retest and interobserver reliability of quantitative sensory testing according to the protocol of the German Research Network on Neuropathic Pain (DFNS): a multi-centre study. *Pain*. 2011; 152(3):548–56.
29. Walton D, MacDermid J, Nielson W, Teasell R, Chiasson M, Brown L. Reliability, standard error, and minimum detectable change of clinical pressure pain threshold testing in people with and without acute neck pain. *Journal of orthopaedic & sports physical therapy*. 2011; 41(9):644–50.
30. Prushansky T, Handelzalts S, Pevzner E. Reproducibility of pressure pain threshold and visual analog scale findings in chronic whiplash patients. *The Clinical journal of pain*. 2007; 23(4):339–45.
31. Felix ER, Widerstrom-Noga EG. Reliability and validity of quantitative sensory testing in persons with spinal cord injury and neuropathic pain. *J Rehabil Res Dev*. 2009; 46(1):69–83.
32. Widerström-Noga EG. Reliability and validity of quantitative sensory testing in persons with spinal cord injury and neuropathic pain. *Journal of rehabilitation research and development*. 2009; 46(1):69.
33. Saebo H, Naterstad IF, Stausholm MB, Bjordal JM, Joensen J. Reliability of pain pressure threshold algometry in persons with conservatively managed wrist fractures. *Physiotherapy research international: the journal for researchers and clinicians in physical therapy*. 2019:e1797–e.
34. Moloney NA, Hall TM, Doody CM. Reliability of thermal quantitative sensory testing: a systematic review. *Journal of rehabilitation research and development*. 2012; 49(2):191.
35. Kennedy DL, Kemp HI, Ridout D, Yarnitsky D, Rice ASC. Reliability of conditioned pain modulation: a systematic review. *PAIN*. 2016; 157(11):2410–9.
36. Alqarni AM, Manlapaz D, Baxter D, Tumilty S, Mani R. Test Procedures to Assess Somatosensory Abnormalities in Individuals with Peripheral Joint Pain: A Systematic Review of Psychometric Properties. *Pain Practice*. 2018; 18(7):895–924.
37. Sim J, Wright C. *Research in health care: concepts, designs and methods*. Hampshire, UK: Nelson Thornes; 2000.
38. Walter S, Eliasziw M, Donner A. Sample size and optimal designs for reliability studies. *Statistics in medicine*. 1998; 17(1):101–10.
39. Gierthmühlen J, Enax-Krumova EK, Attal N, Bouhassira D, Cruccu G, Finnerup NB, et al. Who is healthy? Aspects to consider when including healthy volunteers in QST-based studies—a consensus statement by the EUROPAIN and NEUROPAIN consortia. *Pain*. 2015; 156(11):2203–11.
40. Rushton AB, Evans DW, Middlebrook N, Heneghan NR, Small C, Lord J, et al. Development of a screening tool to predict the risk of chronic pain and disability following musculoskeletal trauma: protocol for a prospective observational study in the United Kingdom. *BMJ Open*. 2018; 8(4).

41. Middlebrook N, Rushton AB, Heneghan NR, Falla D. Measures of central sensitisation and their measurement properties in the adult musculoskeletal trauma population: a protocol for a systematic review and data synthesis. *BMJ Open*. 2019; 9(3):e023204.
42. Clay FJ, Watson WL, Newstead SV, McClure RJ. A systematic review of early prognostic factors for persistent pain following acute orthopedic trauma. *Pain Research & Management: The Journal of the Canadian Pain Society*. 2012; 17(1):35–44.
43. Woolf CJ. Pain: moving from symptom control toward mechanism-specific pharmacologic management. *Annals of internal medicine*. 2004; 140(6):441–51.
44. Rolke R, Magerl W, Campbell KA, Schalber C, Caspari S, Birklein F, et al. Quantitative sensory testing: a comprehensive protocol for clinical trials. *Eur J Pain*. 2006; 10(1):77–88.
45. Bisset LM, Evans K, Tuttle N. Reliability of 2 Protocols for Assessing Pressure Pain Threshold in Healthy Young Adults. *Journal of Manipulative and Physiological Therapeutics*. 2015; 38(4):282–7.
46. Nie H, Arendt-Nielsen L, Andersen H, Graven-Nielsen T. Temporal Summation of Pain Evoked by Mechanical Stimulation in Deep and Superficial Tissue. *The Journal of Pain*. 2005; 6(6):348–55.
47. Nie H, Arendt-Nielsen L, Madeleine P, Graven-Nielsen T. Enhanced temporal summation of pressure pain in the trapezius muscle after delayed onset muscle soreness. *Experimental Brain Research*. 2006; 170(2):182–90.
48. Rankin G, Stokes M. Reliability of assessment tools in rehabilitation: an illustration of appropriate statistical analyses. *Clinical rehabilitation*. 1998; 12(3):187–99.
49. Martin Bland J, Altman D. STATISTICAL METHODS FOR ASSESSING AGREEMENT BETWEEN TWO METHODS OF CLINICAL MEASUREMENT. *The Lancet*. 1986; 327(8476):307–10.
50. Koo TK, Li MY. A guideline of selecting and reporting intraclass correlation coefficients for reliability research. *Journal of chiropractic medicine*. 2016; 15(2):155–63.
51. Watkins MP, Portney L. *Foundations of clinical research: applications to practice*: Pearson/Prentice Hall Upper Saddle River, NJ; 2009.
52. Walton D, MacDermid J, Nielson W, Teasell R, Reese H, Levesque L. Pressure pain threshold testing demonstrates predictive ability in people with acute whiplash. *journal of orthopaedic & sports physical therapy*. 2011; 41(9):658–65.
53. Pelfort X, Torres-Claramunt R, Sánchez-Soler JF, Hinarejos P, Leal-Blanquet J, Valverde D, et al. Pressure algometry is a useful tool to quantify pain in the medial part of the knee: An intra- and inter-reliability study in healthy subjects. *Orthopaedics & Traumatology: Surgery & Research*. 2015; 101(5):559–63.
54. May A, Rodriguez-Raecke R, Schulte A, Ihle K, Breimhorst M, Birklein F, et al. Within-session sensitization and between-session habituation: A robust physiological response to repetitive painful heat stimulation. *European Journal of Pain*. 2012; 16(3):401–9.
55. Jurgens TP, Sawatzki A, Henrich F, Magerl W, May A. An Improved Model of Heat-Induced Hyperalgesia-Repetitive Phasic Heat Pain Causing Primary Hyperalgesia to Heat and Secondary Hyperalgesia to Pinprick and Light Touch. *Plos One*. 2014; 9(6).
56. Breimhorst M, Hondrich M, Rebhorn C, May A, Birklein F. Sensory and sympathetic correlates of heat pain sensitization and habituation in men and women. *European Journal of Pain*. 2012; 16(9):1281–92.
57. Rolke R, Baron R, Maier C, Tölle TR, Treede RD, Beyer A, et al. Quantitative sensory testing in the German Research Network on Neuropathic Pain (DFNS): Standardized protocol and reference values. *PAIN*. 2006; 123(3):231–43.