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# Articles

# Non-invasive risk scores for prediction of type 2 diabetes (EPIC-InterAct): a validation of existing models



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#### Summary

**Background** The comparative performance of existing models for prediction of type 2 diabetes across populations has not been investigated. We validated existing non-laboratory-based models and assessed variability in predictive performance in European populations.

Methods We selected non-invasive prediction models for incident diabetes developed in populations of European ancestry and validated them using data from the EPIC-InterAct case-cohort sample (27779 individuals from eight European countries, of whom 12 403 had incident diabetes). We assessed model discrimination and calibration for the first 10 years of follow-up. The models were first adjusted to the country-specific diabetes incidence. We did the main analyses for each country and for subgroups defined by sex, age (<60 years  $vs \ge 60$  years), BMI (<25 kg/m<sup>2</sup>  $vs \ge 25$  kg/m<sup>2</sup>), and waist circumference (men <102 cm  $vs \ge 102$  cm; women <88 cm  $vs \ge 88$  cm).

**Findings** We validated 12 prediction models. Discrimination was acceptable to good: C statistics ranged from 0.76 (95% CI 0.72-0.80) to 0.81 (0.77-0.84) overall, from 0.73 (0.70-0.76) to 0.79 (0.74-0.83) in men, and from 0.78 (0.74-0.82) to 0.81 (0.80-0.82) in women. We noted significant heterogeneity in discrimination ( $p_{heterogeneity} < 0.0001$ ) in all but one model. Calibration was good for most models, and consistent across countries ( $p_{heterogeneity} > 0.05$ ) except for three models. However, two models overestimated risk, DPoRT by 34% (95% CI 29–39%) and Cambridge by 40% (28–52%). Discrimination was always better in individuals younger than 60 years or with a low waist circumference than in those aged at least 60 years or with a large waist circumference. Patterns were inconsistent for BMI. All models overestimated risks for individuals with a BMI of <25 kg/m<sup>2</sup>. Calibration patterns were inconsistent for age and waist-circumference subgroups.

Interpretation Existing diabetes prediction models can be used to identify individuals at high risk of type 2 diabetes in the general population. However, the performance of each model varies with country, age, sex, and adiposity.

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## Introduction

The number of individuals with type 2 diabetes is high and increasing rapidly: 366 million people worldwide were estimated to have type 2 diabetes in 2011, and prevalence is expected to rise by 51% by 2030.<sup>1</sup> Diabetes is associated with increased morbidity and mortality, and accounts for a substantial proportion of use of health-care resources worldwide.<sup>1</sup> Several studies have convincingly shown that early interventions can prevent or postpone type 2 diabetes.<sup>23</sup> The cost of these interventions and their constraints for individuals are arguments against their provision, particularly for people at low risk of diabetes. Therefore, appropriate identification of individuals at high risk is important.

Several risk scores to predict type 2 diabetes have been developed.<sup>4-7</sup> They vary in many ways, such as the time horizon for prediction and number and nature of predictors. Some models are based on non-laboratory clinical variables (non-invasive risk scores); others have also incorporated biological variables (invasive risk scores). Non-invasive risk scores have been shown to identify a high risk of type 2 diabetes (*C* statistics  $\geq 0.8$ ), although invasive risk scores are more successful (*C* statistics  $\sim 0.9$ ).<sup>8</sup> The use of non-invasive risk scores is more likely to be cost effective and feasible for large-scale screening than is use of invasive risk scores. Generally, the use of risk scores has been widely incorporated into strategies for diabetes prevention.<sup>9,10</sup> However, risk scores might not be generalisable from one population to another, and the validity of existing diabetes risk scores in different settings and subgroups on the basis of sex or BMI has not been established.<sup>10</sup>

We aimed to validate and compare existing noninvasive prediction models for type 2 diabetes in European populations. We assessed variability in predictive performance between countries and by sex, BMI, waist circumference, and age.

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See Online for appendix

# Methods

# Study design and population

We used the EPIC-InterAct case-cohort sample (27779 participants, of whom 12403 had incident diabetes), which was selected from the European Investigation into Cancer and Nutrition (EPIC) cohort study (455569 participants; appendix p 4).<sup>11</sup> The design and methods of the InterAct study have been described elsewhere.<sup>11</sup> Briefly, the InterAct study was designed to investigate how genes and lifestyle factors interact to influence risk of type 2 diabetes. The InterAct consortium partners ascertained and verified incident cases of type 2 diabetes in EPIC cohorts between 1991 and 2007, from eight of the ten EPIC countries (26 centres).

Participants gave written informed consent for participation in the EPIC study. The study was approved by the local ethics committees in the participating countries and the internal review board of the International Agency for Research on Cancer (Lyon, France).

# Procedures

Self-administered questionnaires provided baseline information about lifestyle and medical history in the case-cohort sample. Validated instruments were used to assess nutritional variables<sup>12</sup> and physical activity.<sup>13</sup> Blood pressure and anthropometric measurements followed standard approaches (appendix p 5). We handled missing data in three ways: proxy variables (appendix p 6) or exclusion of the predictor from the model; exclusion of countries from specific analyses (appendix p 11); and imputation of missing values by country with R's AregImpute Function, which takes all aspects of uncertainty in the imputations into account by using the bootstrap to estimate the process of drawing predicted values from a full Bayesian predictive distribution.

Existing non-invasive prediction models for incident diabetes had been identified through a systematic review, as previously described.8 We included in this study only models developed in general populations of European ancestry. We first validated the original models by computing the predicted probability of diabetes for each participant in the EPIC-InterAct casecohort with baseline values (appendix p 7). We expressed models' performance in terms of discrimination (whether the model can distinguish between people who do and do not develop diabetes) and calibration (to what extent the predicted probabilities agree with the reported risk across groups of individuals). We assessed discrimination with the C statistic (which is comparable to the area under the receiver operator characteristic curve) adapted for time-to-event data.14 C statistics vary from 0.5 (no discrimination) to 1 (perfect discrimination), with values of 0.7-0.8 deemed acceptable and 0.8–0.9 good.15 We assessed calibration graphically with calibration plots, and by computing the ratio of expected to recorded probabilities and the accompanying 95% CIs by assuming a Poisson variance.<sup>16</sup> We also calculated the Yates slope (difference between mean predicted probability of type 2 diabetes for participants with and without incident diabetes, with higher values indicating better performance), and Brier score (squared difference between predicted probability and actual outcome for each participant; these values vary from 0 for a perfect prediction model to 0.25 for a non-informative model with 50% incident outcome).<sup>17</sup> We did not use tests of calibration such as the Hosmer-Lemeshow test or *U* statistic,<sup>18</sup> because they are sensitive to study sample size.

Because we were comparing the performance of several models in a case-cohort design across many countries, we took some extra steps before model validation. First, the models have different so-called time horizons. In the main analysis, we compared all models for prediction of type 2 diabetes within 10 years. Therefore, participants who developed diabetes after 10 years of follow-up were included as non-cases. Second, calibration is strongly affected by the outcome's incidence in the population; we therefore recalibrated models to the country-specific incidence through intercept adjustment (appendix pp 9, 12).18,19 These approaches eliminate differences between the models due to differences in incidence between the populations for which the scores were developed.<sup>19</sup> Third, the incidence of diabetes is artificially increased in the casecohort design. To arrive at the true incidences from the original cohort, we reconstituted the EPIC cohort within countries by applying a so-called blow-up approach to extrapolate the case-cohort data to a full cohort (appendix p 10).8 Finally, to account for any heterogeneity between countries, we estimated the C statistic and ratio of expected to recorded probabilities for the entire cohort by pooling country-level estimates with random effects models. We then used the Cochran's Q and I<sup>2</sup> statistics to assess heterogeneity across countries.

We did the main analyses for each country and for subgroups defined by sex, age (<60 years  $vs \ge 60$  years), BMI (<25 kg/m<sup>2</sup>  $vs \ge 25$  kg/m<sup>2</sup>), and waist circumference (men <102 cm  $vs \ge 102$  cm; women <88 cm  $vs \ge 88$  cm). In four of the eight countries (Denmark, Italy, Spain, and Sweden), not all models had been validated (appendix p 11). We did a sensitivity analysis by restricting the assessment of the heterogeneity between countries to the four countries in which all models had been validated (France, Germany, the Netherlands, and the UK) to investigate if assessments of heterogeneity of different models on the basis of a different number of countries affected the comparisons. We also compared all models for prediction of diabetes within 5 years, which corresponds to the time horizon for predicted probability for six models.

We used the Hmisc, Design, Survival, and meta packages of R (version 2.13.0) for data analysis.

#### Role of the funding source

The sponsors of the study had no role in study design, data collection, data analysis, data interpretation, or writing of the report. APK and JWJB had full access to all the data in the study; JWJB, YTvdS, and NJW had final responsibility for the decision to submit for publication.

# Results

We validated 12 prediction models (table 1, appendix p 7). Age was a predictor of type 2 diabetes in all models except DESIR.<sup>24</sup> Most models were based on logistic regressions (table 1). Varying definitions for incident diabetes were applied (table 1). Data for baseline age and sex were available for all participants, but data were missing for some participants for some predictors (table 1).

More than half the overall EPIC-InterAct study population were women, but the proportion varied by country (table 2). French participants were all women, whereas the proportion of women among participants in other countries ranged from 44% to 81%. Mean baseline age was  $54 \cdot 1$  years (SD 8.4) in men and  $53 \cdot 5$  years (9.0) in women.

During the first 10 years of follow-up, almost 10000 cases of incident diabetes were recorded in the EPIC-InterAct population (table 2). Discrimination of the 12 prediction models was acceptable to good, with *C* statistics ranging from 0.76 to 0.81 (figure 1, table 3). For nine models, the recalibrated calibration curves mostly followed the ideal calibration line (figure 2; see appendix p 21 for calibration curves before recalibration). For the ARIC 2009 model,<sup>21</sup> the curves were mostly steeper than the ideal calibration line, with increasing risk (figure 2), indicating risk underestimation in high-

risk participants. For the Cambridge<sup>23</sup> and DPoRT<sup>25</sup> models, calibration curves were fairly flat, with increasing risk (figure 2), indicating risk overestimation in high-risk participants. Four models accurately estimated the overall rate of incident diabetes, six marginally overestimated or underestimated, and two largely overestimated (figure 2, table 3).

We noted significant heterogeneity in discrimination across countries for all models except for the ARIC 2009 model (figure 1). The highest *C* statistics were almost always recorded in France, and the lowest mostly in Denmark (figure 1). The worst and best discriminating models varied by country (figure 1). Patterns of calibration were largely similar in each country to those recorded for the overall cohort (appendix pp 22–33). As expected, because of recalibration by country, we recorded almost no heterogeneity in calibration.

Discrimination was generally higher in women than men (table 3). *C* statistics in men ranged from 0.73 to 0.79 in the overall cohort (table 3), and from 0.68 to 0.86 across countries (appendix p 13). In men, the best discrimination was recorded in the Netherlands (appendix p 13). In women, *C* statistics ranged from 0.78 to 0.81 in the overall cohort (table 3), and from 0.70 to 0.85 across countries (appendix p 13). The highest *C* statistics in women were mostly recorded in Germany, and the lowest in Denmark (appendix p 13). Calibration differed greatly between men and women per country, which could be masked by the overall ratio of expected to recorded probabilities (appendix p 13).

Discrimination was acceptable to good across age strata, and was always better in the younger age group

	Incident diabetes risk models												Variable or proxy in InterAct	Participants with missing data for InterAct
	ARIC 2005 <sup>20</sup>	ARIC 2009 <sup>21</sup>	AUS- DRISK <sup>22</sup>	Cam- bridge <sup>23</sup>	DESIR <sup>24</sup>	DPoRT <sup>25</sup>	FINDRISK concise <sup>6</sup>	FINDRISK full⁰	Framingham personal <sup>26</sup>	KORA S4/F4 <sup>27</sup>	EPIC- Potsdam <sup>28</sup>	QD Score <sup>29</sup>	-	
Year	2005	2009	2010	2008	2008	2010	2003	2003	2007	2010	2007	2009		
Cohort size	7915	12729	6060	24495	3817	19861	4595	4435	3140	873	25167	2540753		
Follow-up (years)	9	10	5	4.6	9	9	5	5	7	5	5	10		
Definition of incident diabetes	Any incident	Any incident	Treated with drugs, diag- nosed with OGTT	Self-report, registries	Treated with drugs, diagnosed with fasting glucose measure- ments	Any incident	Treated with drugs	Treated with drugs	Treated with drugs, diagnosed with fasting glucose measure- ments	Diag- nosed with OGTT	Diag- nosed with OGTT	Any incident		
Number of cases of incident diabetes	1292	2407	362	323	203	1410	194	182	160	91	849	78081		
Country	USA	USA	Australia	UK	France	Canada	Finland	Finland	USA	Germany	Germany	UK		
Age group included (years)	≥25	45-64	≥25	40-79	30-64	>20	35-64	35-64	54*	55-74	35-65	25-79	20–79	
Statistical model	Logistic	Weibull	Logistic	Logistic	Logistic	Weibull	Logistic	Logistic	Logistic	Logistic	Cox	Cox	NA	NA
Validation	External	External	External	External	External	External	External	External	External	External	External	External	NA (Continue	NA es on next page)

	Incident diabetes risk models											Variable or proxy in InterAct	Participants with missing data for InterAct	
	ARIC 2005 <sup>20</sup>	ARIC 2009 <sup>21</sup>	AUS- DRISK <sup>22</sup>	Cam- bridge²³	DESIR <sup>24</sup>	DPoRT <sup>25</sup>	FINDRISK concise <sup>6</sup>	FINDRISK full⁰	Framingham personal <sup>26</sup>	KORA S4/F4 <sup>27</sup>	EPIC- Potsdam <sup>28</sup>	QD Score <sup>29</sup>		
(Continued from previous page)														
Predictors														
Age	Yes	Yes	Yes	Yes	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	0
Sex	No	No	Yes	Yes	Sex specific†	Sex specific†	No	No	Yes	Yes	No	Sex specific†	Yes	0
Smoking	No	Yes	Yes	Yes	Yes	Yes	No	No	No	Yes	Yes	Yes	Yes	339
Family history	No	No	No	Yes	Yes	No	No	No	No	No	No	Yes	Yes	13795‡
Parental diabetes	Yes	Yes	Yes	No	Sex specific†	No	No	No	Yes	Yes	No	No	Yes	15761‡
Maternal diabetes	No	Yes	No	No	No	No	No	No	No	No	No	No	Yes	15761‡
Hypertension	No	Yes	No	Yes	Yes	Yes	No	No	No	Yes	Yes	Yes	Yes	556
Pulse	No	Yes	No	No	No	No	No	No	No	No	No	No	Yes	16311§
Ethnic origin	Yes	Yes	Yes	No	No	Yes	No	No	No	No	No	Yes	Yes	0
BMI	No	No	No	Yes	No	Yes	Yes	Yes	Yes	Yes	No	Yes	Yes	197
Body weight	No	Yes	No	No	No	No	No	No	No	No	No	No	Yes	177
Waist circumference	Yes	Yes	Yes	No	Yes	No	Yes	Yes	No	No	Yes	No	Yes	2018
Height	Yes	Yes	No	No	No	No	No	No	No	No	Yes	No	Yes	113
Steroid use	No	No	No	Yes	No	No	No	No	No	No	No	Yes	NA	All¶
High glucose concentration	No	No	Yes	No	No	No	Yes	Yes	No	No	No	No	NA	All¶
Physical activity	No	No	Yes	No	No	No	No	Yes	No	No	Yes	No	Yes	2217
Townsend score	No	No	No	No	No	No	No	No	No	No	No	Yes	NA	All¶
Education	No	Yes	No	No	No	Yes	No	No	No	No	No	No	Yes	330
Consumption of red meat	No	No	No	No	No	No	No	No	No	No	Yes	No	Yes	736
Consumption of whole-grain bread	No	No	No	No	No	No	No	No	No	No	Yes	No	Yes	736
Coffee intake	No	No	No	No	No	No	No	No	No	No	Yes	No	Yes	736
Alcohol intake	No	No	No	No	No	No	No	No	No	No	Yes	No	Yes	117
Immigration into country where study done	No	No	Yes	No	No	No	No	No	No	No	No	No	NA	All¶
Systolic blood pressure	Yes	No	No	No	No	No	No	No	No	No	No	No	Yes	6343
Heart disease	No	No	No	No	No	Yes	No	No	No	No	No	No	Yes	55

ARIC=Atherosclerosis Risk in Communities. AUSDRISK=Australian Type 2 Diabetes Risk Assessment Tool. DESIR=Epidemiological Study on the Insulin Resistance Syndrome. DPoRT=Diabetes Population Risk Tool. FINDRISK=Finnish Diabetes Risk Score. KORA S4/F4=Cooperative Health Research in the Region of Augsburg (KORA), Survey 4. EPIC=European Investigation into Cancer and Nutrition. OGTT=oral glucose tolerance test. NA=not applicable. \*Mean. †Model built separately for women and men. ‡Predictor completely missing in centres in Italy and Spain. \$Predictor completely missing in centres in Denmark and Sweden. ¶All 27779 participants. ||Index of social and economic deprivation of a location.

Table 1: Models of incident diabetes risk prediction included

than in the older (table 3). Discrimination was below the acceptable range (*C* statistic <0.70) in the lower stratum of waist circumference but was acceptable in the upper one. Discrimination was similar across BMI groups, although *C* statistics were higher in people with increased BMI than in those with lower BMIs in seven of the 12 models (table 3).

In terms of calibration, the ARIC 2005,<sup>20</sup> AUSDRISK,<sup>22</sup> FINDRISK concise,<sup>6</sup> and FINDRISK full models<sup>6</sup> acceptably predicted the overall rate of incident diabetes in age subgroups (table 3). The Cambridge model<sup>23</sup> overestimated risk in all participants (table 3). Most other models overestimated risk in younger participants and underestimated risk in the older people, but we recorded

	Overall	Denmark	France	Germany	Italy	Netherlands	Spain	Sweden	ик	
	26	2	6	2		2	-	2	2	
Participating centres	20	2	0	2	5	2	5	2	2	
Total eligible EPIC cohort	326805	20410	45 877	36755	35471	35 0 9 3	46104	52146	54949	
Total InterAct case-cohort	27779	4037	867	3578	3393	2290	5889	5401	2324	
Women	15887 (57-2%)	1787 (44·3%)	867 (100%)	1795 (50·2%)	2177 (64·2%)	1845 (80.6%)	3323 (56·4%)	2824 (52.3%)	1269 (54.6%)	
Age (years)	53.7 (8.7)	57.0 (4.4)	56.9 (6.5)	52.4 (8.3)	51.3 (7.7)	53.7 (10.2)	50.3 (7.8)	54.9 (9.7)	58.5 (10.2)	
Smoking	Smoking									
Never	12335 (44·4%)	1264 (31·3%)	588 (67.8%)	1529 (42.7%)	1509 (44·5%)	908 (39.7%)	3149 (53·5%)	2309 (42.8%)	1079 (46·4%)	
Former	8026 (28.9%)	1281 (31.7%)	200 (23.1%)	1286 (35-9%)	925 (27·3%)	762 (33·3%)	1028 (17.5%)	1618 (30.0%)	926 (39.8%)	
Current	7418 (26.7%)	1492 (37.0%)	79 (9·1%)	763 (21·3%)	959 (28·3%)	620 (27.1%)	1712 (29·1%)	1474 (27·3%)	319 (13·7%)	
Hypertension*	7461 (26·9%)	1020 (25.3%)	172 (19.8%)	1506 (42·1%)	912 (26·9%)	626 (27.3%)	1469 (24·9%)	1332 (24.7%)	424 (18·2%)	
Systolic blood pressure (mm Hg)	136 (21)	144 (21)	130 (20)	134 (19)	134 (19)	133 (22)	132 (20)	140 (20)	139 (20)	
Pulse rate (beats per min)	73 (12)	NA	70 (11)	74 (12)	71 (9)	74 (12)	74 (12)	NA	71 (12)	
BMI (kg/m²)	27.6 (4.8)	27.4 (4.6)	24.6 (4.8)	27.7 (4.9)	27.5 (4.9)	26.6 (4.5)	29.4 (4.6)	26.8 (4.6)	27.1 (4.7)	
Waist circumference (cm)	91 (14)	93 (14)	80 (12)	92 (14)	88 (13)	87 (13)	95 (12)	89 (14)	89 (14)	
Hip circumference (cm)	103 (9)	103 (9)	101 (10)	103 (9)	102 (9)	104 (9)	107 (9)	101 (10)	103 (9)	
Family history of diabetes	4907 (26·5%)†	955 (23·7%)	147 (17.0%)	1114 (31·1%)	NA	771 (33.7%)	NA	1534 (28·4%)	386 (16-6%)	
Follow-up (days)	3968 (2603–4634)	3745 (2299-4242)	3391 (2603–3835)	3472 (1766–4099)	3967 (2480–4638)	4065 (2389–4655)	4557 (3262–4975)	4334 (3335-4952)	3797 (2294–4401)	
Cases of incident diabetes duri	Cases of incident diabetes during follow-up									
Overall	12 403	2055	288	1584	1437	828	2564	2622	1025	
Within 5 years	3698	633	132	916	470	406	418	418	305	
Within 10 years	9951	1845	278	1554	1292	778	1759	1520	925	

Data are n, n (%), mean (SD), or median (IQR). Includes imputed data for those with missing values. NA=not available. \*Self-reported diagnosis or use of drugs, or both. †Percentage calculated with number of individuals for whom data available (ie, not Italy and Spain; n=18 497).

Table 2: Baseline characteristics of the InterAct case-cohort and cases of incident diabetes during follow-up, overall and by country

the opposite finding for KORA<sup>27</sup> and Potsdam<sup>28</sup> models. Absolute risk of diabetes was variably overestimated or underestimated by models across weight circumference strata, while all models systematically overestimated the risk within the lower BMI stratum (table 3).

The pattern of overall event rate prediction in the four countries with valid data for all models (France, Germany, the Netherlands, and the UK) was very similar to that across all countries (data not shown), again with significant heterogeneity for the Cambridge<sup>23</sup> and DPoRT<sup>25</sup> models (both p≤0.0002). Heterogeneity in discrimination across countries decreased when we examined data for countries with full data, with the ARIC 2009,<sup>21</sup> DPoRT,<sup>25</sup> and FINDRISK<sup>6</sup> models showing no heterogeneity (p>0.11).

During the first 5 years of follow-up, 3698 cases of incident diabetes were recorded (table 2). Discrimination was slightly better and less heterogeneous at 5 years than at 10 years; otherwise, the pattern of predictions was broadly similar to that recorded at 10 years (appendix pp 16–20, 34–36).

## Discussion

We have shown that non-invasive models for prediction of incident type 2 diabetes have acceptable to good discrimination over 10 years, both overall and across countries. After recalibration, most models showed good calibration, which was consistent across countries, although discrimination varied significantly. We showed that the models' performance is worse in men than in women. Discrimination is better in people younger than 60 years, but risk can be overestimated in this age group. Discrimination is generally lower in participants with a BMI of less than 25 kg/m<sup>2</sup> than in those with a BMI of at least 25 kg/m<sup>2</sup>. Risk is systematically overestimated in participants with a BMI of less than 25 kg/m<sup>2</sup>. No model significantly outperforms others enough to be uniquely recommended for routine risk stratification.

A few previous validation studies of incident diabetes models have been done.<sup>4,5,8,21,24,30-34</sup> Most of those—mainly cross-sectional studies—compared several incident or prevalent risk scores in one population,<sup>21,24,32-34</sup> and another study compared three risk scores in a multiethnic population in the same country.<sup>30</sup> In one study,<sup>8</sup> the performance of both invasive and non-invasive prediction models for incident diabetes in a population was assessed. Overall, the previous studies showed that the models had modest to good discrimination and poor calibration. However, intercept adjustment to correct for differences in diabetes incidence between development and validation populations was done in only one study.

Models' performance (mostly discrimination) differed across countries in our study. This variation could be a result of differences in how predictors and outcomes

	C statistic (95% CI)		C statistic (95% CI)
ARIC 2005		FINDRISK concise	
Denmark	0.73 (0.72-0.73)	Denmark	0.72 (0.71-0.73)
France -	0.84 (0.81-0.86)	France	0.80 (0.78-0.83)
Germany	- 0.81 (0.80-0.82)	Germany	0.81 (0.80-0.82)
Italy		Italy –	0.78 (0.76–0.79)
Netherlands	- 0.83 (0.82-0.84)	Netherlands —	0.81 (0.79-0.82)
Spain		Spain	0.74 (0.73-0.75)
Sweden	0.77 (0.76-0.78)	Sweden -	0.76 (0.75-0.77)
	0.79 (0.76-0.81)		0.79(0.76-0.80)
Tests for heterogeneity: $l^2=98.6\%$ , $\tau^2=0.0021$ , p<0.0001	0./9(0./0=0.83)	Tests for heterogeneity: $l^2=97.8\%$ , $\tau^2=0.0014$ , p<0.0001	0.78 (0.75-0.80)
······································			
ARIC 2009		FINDRISK full	
Denmark France	0.91 (0.70, 0.94)	França	0.72(0.71-0.73)
Garmany	0.81 (0.79-0.84)	France	0.80 (0.76-0.83)
Italy	0.00 (0.79=0.01)	Italy -	0.77 (0.76-0.78)
Netherlands	- 0.81 (0.80-0.83)	Netherlands	0.80 (0.79-0.82)
Spain	(	Spain -	0.74 (0.73-0.75)
Sweden		Sweden -	0.76 (0.75–0.77)
ик — —	0.79 (0.78-0.81)	UK – –	0.79 (0.77-0.80)
Overall 🔶	0.80 (0.79–0.81)	Overall	0.77 (0.75-0.80)
Tests for heterogeneity: l <sup>2</sup> =18·5%, $\tau^2$ <0·0001, p=0·298		Tests for heterogeneity: I²=97·4%, $\tau^{2}$ =0·0013, p<0·0001	
ALISDRISK		Framingham personal	
Denmark	0.71 (0.70-0.72)	Denmark	0.69 (0.68-0.70)
France	0.80 (0.77–0.83)	France	- 0.82 (0.80-0.85)
Germany -	0.79 (0.78–080)	Germany -	0.78 (0.77–0.79)
Italy		Italy	
Netherlands -	- 0.81 (0.80-0.82)	Netherlands -	0.80 (0.79-0.82)
Spain	``````	Spain	
Sweden – – –	0.74 (0.72–0.75)	Sweden –	0.74 (0.73–0.76)
UK — — —	0.76 (0.74–0.77)	UK –	0.76 (0.74–0.77)
Overall	0.77 (0.73-0.80)	Overall	0.76 (0.72–0.80)
Tests for heterogeneity: $I^2$ =98%, $\tau^2$ =0.002, p<0.0001		Tests for heterogeneity: $l^2$ =98·5%, $\tau^2$ =0·0024, p<0·0001	
Cambridge		KORA S4/F4	
Denmark	0.72 (0.71–0.73)	Denmark	0.72 (0.71–0.72)
France	0.82 (0.80-0.84)	France	0.82 (0.80-0.84)
Germany -	- 0.81 (0.80–0.82)	Germany -	0.80 (0.79–0.81)
Italy		Italy	
Netherlands -	- 0.81 (0.80-0.82)	Netherlands	0.82 (0.81–0.83)
Spain Swadan	0.78 (0.77, 0.70)	Spain	0.76 (0.75, 0.77)
	0.78 (0.77 0.80)		0.70(0.75-0.77)
Overall	0.79 (0.75-0.82)	Overall	0.78 (0.74-0.82)
Tests for heterogeneity: I <sup>2</sup> =98·2%, τ <sup>2</sup> =0·0019, p<0·0001	0.79(0.75-0.02)	Tests for heterogeneity: I <sup>2</sup> =98·3%, τ <sup>2</sup> =0·0021, p<0·0001	0.70 (0.74-0.02)
DESIR		Potsdam	
Denmark	0.72 (0.72–0.73)	Denmark ·	0.72 (0.72-0.73)
France		France	- 0.84 (0.81-0.86)
Germany -	0.81 (0.80-0.82)	Germany	0.82 (0.81-0.83)
Nothorlands	0 82 (0 81 0 82)	Notherlands	0.78(0.77-0.80)
Spain	0.82 (0.81-0.83)	Spain	0.76 (0.75-0.77)
Sweden -	0.76 (0.75-0.78)	Sweden	0.77 (0.76-0.78)
	0.80 (0.79-0.81)		0.80 (0.78-0.81)
Overall	0.79 (0.76-0.83)	Overall	0.79 (0.76-0.82)
Tests for heterogeneity: I <sup>2</sup> =98·5%, τ <sup>2</sup> =0·0021, p<0·0001		Tests for heterogeneity: $l^2=98\%$ , $\tau^2=0.0015$ , p<0.0001	
DD-DT		00	
Denmark	0.71 (0.60.0.71)	Denmark	0.72 (0.72 0.74)
France	0.81 (0.78–0.84)	France	0.85 (0.82-0.87)
Germany	0.79 (0.78-0.80)	Germany	0.82 (0.81-0.83)
Italy	0.76 (0.75–0.77)	Italy	(0 01 0 05)
Netherlands —	0.79 (0.78–0.81)	Netherlands -	- 0.84 (0.83-0.85)
Spain –	0.73 (0.72-0.74)	Spain	
Sweden -	0.76 (0.75-0.77)	Sweden -	0.80 (0.79-0.81)
ик — —	0.77 (0.76–0.79)	UK –	0.81 (0.79-0.82)
Overall	0.76 (0.74–0.79)	Overall	0.81 (0.77–0.84)
Tests for heterogeneity: J2=97·4%, $\tau^2$ =0·0014, p<0·0001		Tests for heterogeneity: I²=98.6%, $\tau^{2}$ =0.0021, p<0.0001	
0.70 0.75 0.80	0.85 0.90	0.70 0.75 0.80	0.85 0.90
C statist	ic	C statistic	

Figure 1: Discrimination of the different models for the prediction of incident type 2 diabetes at 10 years of follow-up overall and by country

Discrimination refers to the ability of the model to distinguish between participants who did and did not develop diabetes during 10 years of follow-up. ARIC=Atherosclerosis Risk in Communities. AUSDRISK=Australian Type 2 Diabetes Risk Assessment Tool. DESIR=Epidemiological Study on the Insulin Resistance Syndrome. DPoRT=Diabetes Population Risk Tool. FINDRISK=Finnish Diabetes Risk Score. KORA S4/F4=Cooperative Health Research in the Region of Augsburg (KORA), Survey 4.

	Overall	Men	Women	Age <60 years	Age ≥60 years	BMI <25 kg/m²	BMI ≥25 kg/m²	Waist circumference ≥102 cm in men or ≥88 cm in women	Waist circumference <102 cm in men or <88 cm in women
C statistics									
ARIC 2005	0·79	0·75	0·81	0·81	0·73	0·73	0·71	0·66	0·75
	(0·76–0·83)*	(0·73–0·78)*	(0·77–0·84)*	(0·77–0·85)*	(0·69–0·76)*	(0·66–0·80)*	(0·69–0·73)*	(0·65–0·67)†	(0·70–0·79)*
ARIC 2009	0·80	0·79	0·81	0·82	0·74	0·71	0·73	0·66	0·75
	(0·79–0·81)	(0·74–0·83)*	(0·80–0·82)	(0·81–0·83)	(0·71–0·77)‡	(0·68–0·73)	(0·71–0·74)	(0·63–0·69)‡	(0·73-0·76)
AUSDRISK	0·77	0·73	0·78	0·78	0·70	0·69	0·71	0·64	0·69
	(0·73–0·80)*	(0·70–0·76)*	(0·74–0·82)*	(0·74–0·82)*	(0·68–0·72)‡	(0·64–0·74)*	(0·68–0·73)*	(0·63–0·66)†	(0·65–0·73)*
Cambridge model	0·79	0·76	0·80	0·80	0·72	0·70	0·70	0·66	0·74
	(0·75–0·82)*	(0·73–0·79)*	(0·76–0·83)*	(0·76–0·84)*	(0·70–0·74)*	(0·65–0·75)*	(0·68–0·72)*	(0·64–0·68)*	(0·69–0·79)*
DESIR model	0·79	0·73	0·81	0·81	0·74	0·71	0·72	0·65	0·71
	(0·76–0·83)*	(0·71–0·75)*	(0·77–0·85)*	(0·77–0·85)*	(0·71–0·77)*	(0·65–0·76)*	(0·69–0·74)*	(0·63–0·67)*	(0·67–0·76)*
DPoRT model	0·76	0·73	0·78	0·78	0·70	0·63	0·68	0·64	0·72
	(0·74–0·79)*	(0·71–0·75)*	(0·75–0·81)*	(0·75–0·81)*	(0·68–0·73)*	(0·59–0·67)*	(0·66–0·69)*	(0·63–0·65)‡	(0·68–0·75) *
FINDRISK concise	0·78	0·75	0·79	0·79	0·71	0·68	0·70	0·63	0·72
	(0·75–0·80)*	(0·72–0·78)*	(0·76–0·81)*	(0·76–0·82) *	(0·69–0·73)*	(0·64–0·71)*	(0·69–0·72)*	(0·62–0·65)*	(0·68–0·75)*
FINDRISK full model	0·77	0·74	0·78	0·78	0·71	0·68	0·70	0·63	0·71
	(0·75–0·80)*	(0·72–0·77)*	(0·76–0·81)*	(0·75–0·81)*	(0·68–0·73)*	(0·65–0·72)*	(0·69–0·72)*	(0·62–0·65)§	(0·67–0·74)*
Framingham personal	0·76	0·74	0·78	0·78	0·69	0·64	0·65	0·61	0·70
	(0·72–0·80)*	(0·70–0·77)*	(0·74–0·82)*	(0·74–0·83)*	(0·67–0·73)*	(0·57–0·70)*	(0·63–0·60)*	(0·60–0·63)†	(0·65–0·75)*
KORA model	0·78	0·75	0·79	0·80	0·72	0·73	0·70	0·66	0·75
	(0·74–0·82)*	(0·71–0·79)*	(0·75–0·82)*	(0·76–0·84)*	(0·69–0·74)*	(0·67–0·78)*	(0·68–0·73)*	(0·65–)·68)‡	(0·70–0·79)*
Potsdam model	0·79	0·74	0·81	0·80	0·72	0·74	0·72	0·66	0·75
	(0·76–0·82)*	(0·72–0·77)*	(0·78–0·83)*	(0·77–0·83)*	(0·70–0·75)*	(0·68–0·79)*	(0·70–0·73)*	(0·65–0·68)§	(0·71–0·78)*
QDScore	0·81	0·78	0·81	0·82	0·75	0·75	0·72	0·68	0·77
	(0·77–0·84)*	(0·74–0·82)*	(0·77–0·85)*	(0·78–0·86)*	(0·72–0·77)*	(0·69–0·80)*	(0·70–0·75)*	(0·66–0·70)*	(0·72–0·82)*
Ratios of expected to o	bserved rates								
ARIC 2005	1.06	1·08	1·06	1·02	1·11	2·22	0·83	0·74	1·49
	(1.04–1.08)	(1·02–1·14)‡	(0·97–1·15)*	(0·98–1·06)†	(1·03−1·19)§	(1·74–2·68)*	(0·74–0·93)*	(0·65–0·82)*	(1·34–1·64)*
ARIC 2009	1·04	1·01	1·13	1·20	0·79	2·90	0·69	0·55	1·72
	(1·01–1·08)	(0·84–1·17)*	(0·96–1·30)*	(1·14–1·27)	(0·69–0·88)‡	(2·60–3·21)	(0·56–0·81)*	(0·56–0·65)*	(1·56–1·88)*
AUSDRISK	1·00	1·11	0·93	1·03	0·97	2·29	0·75	0·77	1·31
	(0·98–1·02)	(1·02−1·18)§	(0·86–1·00)*	(0·99–1·08)	(0·87–1·02)§	(1·70–2·81)*	(0·67–0·83)*	(0·68–0·86)*	(1·17–1·45)*
Cambridge model	1·40	1·76	1·14	1·29	1·53	1·27	1·42	1·30	1·54
	(1·28–1·52)*	(1·59–1·94)*	(0·98–1·30)*	(1·17–1·42)*	(1·30–1·77)*	(0·99–1·55)*	(1·30–1·53)*	(1·20–1·41)*	(1·34–1·73)*
DESIR model	1·06	1·06	1·06	1·18	0·88	1·35	1·02	1·13	0·99
	(1·02–1·11)‡	(1·03–1·08)	(0·99–1·13)§	(1·12–1·24)‡	(0·83–0·92)	(1·07–1·64)*	(0·98–1·06)†	(1·05–1·22)*	(0·89–1·10)*
DPoRT model	1·34	1·58	1·14	1·44	1·07	1·58	1·30	1·00	1·80
	(1·29–1·39)‡	(1·51–1·64)†	(1·11–1·17)	(1·39–1·49)†	(0·95–1·18)*	(1·33–1·83)*	(1·23–1·36)*	(0·92–1·07)*	(1·67–1·93)*
FINDRISK concise	0·98	0·85	1·11	1·00	0·94	1·28	0·93	1·08	0·84
	(0·96–0·99)	(0·79–0·91)*	(1·02–1·19)*	(0·97–1·03)	(0·88–1·01)§	(1·04–1·51)*	(0·90–0·97)§	(1·03–1·13)*	(0·79–0·90)*
FINDRISK full model	0·98	0·89	1·07	1·00	0·94	1·37	0·92	1·06	0·88
	(0·96–0·99)	(0·84–0·95)*	(1·00–1·15)*	(0·97–1·03)	(0·88–1·00)§	(1·12–1·62)*	(0·88–0·96)*	(1·01–1·11)*	(0·82–0·93)*
Framingham personal	0·99	0·78	1·22	1·12	0·81	1·67	0·88	0·73	1·36
	(0·98–1·02	(0·70–0·87)*	(1·09–1·35)*	(1·07–1·16)	(0·76–0·86)†	(1·30–2·05)*	(0·82–0·94)*	(0·66–0·79)*	(1·23–1·50)*
KORA model	1·00	1·09	0·91	0·82	1·26	1·70	0·87	0·77	1·31
	(0·97–1·02)	(1·06–1·12)	(0·84–0·99)*	(0·74–0·91)*	(1·15–1·37)*	(1·37–2·02)*	(0·82–0·92)§	(0·73–0·80)†	(1·21–1·42)*
Potsdam model	0·98	1·06	0·93	0·90	1·17	1·38	0·93	0·96	1·04
	(0·97–1·00)	(1·01–1·11)†	(0·85–1·02)*	(0·85–0·94)†	(1·08–1·25)*	(1·12–1·64)*	(0·88–0·97)*	(0·90–1·03)*	(0·94–1·14)*
QDScore	1·06	1·08	1·04	1·04	1·22	1·40	0·99	0·87	1·29
	(1·04–1·08)	(1·05–1·11)*	(1·01–1·07)	(0·99–1·10)‡	(1·07–1·36)*	(1·15–1·65)*	(0·93–1·05)*	(0·80–0·94)*	(1·19–1·39)*

ARIC=Atherosclerosis Risk in Communities. AUSDRISK=Australian Type 2 Diabetes Risk Assessment Tool. DESIR=Epidemiological Study on the Insulin Resistance Syndrome. DPoRT=Diabetes Population Risk Tool. FINDRISK=Finnish Diabetes Risk Score. KORA S4/F4=Cooperative Health Research in the Region of Augsburg (KORA), Survey 4. \*p<sub>heterogenetity</sub><0-0001. †p<sub>heterogenetity</sub><0-001. \$p<sub>heterogenetity</sub><0-001.

Table 3: Discrimination and calibration of models for the prediction of incident type 2 diabetes at 10 years of follow-up overall and by subgroup

were measured (for discrimination) and in baseline diabetes risk (for calibration) across countries. The FINDRISK concise model<sup>6</sup> was previously tested in five

cohorts from Europe, North America, and Australia.<sup>35</sup> Discrimination varied by country, with *C* statistics between 0.63 and 0.78,<sup>35</sup> which is consistent with our



results. Therefore, risk scores might not always be generalisable to populations other than the development population.

We noted some consistent trends across countries. Discrimination was generally low for the DPoRT model,<sup>25</sup> potentially because this model was developed to estimate diabetes risk in a population rather than for an individual, which could affect quality of the predictors and outcome data. The QDScore<sup>29</sup> was one of the best discriminatory models across countries. It is based on more predictors than other models are and so uses more information, and was developed from a large sample, providing precise estimates of the associations between predictors and outcomes. Calibration of the Cambridge model<sup>23</sup> was generally poor. Although validated for incident diabetes prediction,<sup>36</sup> this model was developed as a diagnostic model, which could have inflated estimates of the association between diabetes and predictors.

We recorded substantial variation in models' performance by major subgroups, with consistently better performance in women than in men. For the DESIR<sup>24</sup> and DPoRT<sup>25</sup> models and the QDScore,<sup>29</sup> this difference could be a result of the use of sex-specific coefficients that could efficiently capture the predictive information from risk factors. BMI and waist circumference are generally much stronger diabetes predictors in women than in men and could also explain these results.<sup>37,38</sup> In our study, the differences in discrimination linked to BMI were small. BMI and waist circumference might not be linearly related to diabetes risk.37 By including linear or categorical terms of these predictors, most models might not completely capture the nature of the associations. Our finding that discrimination is mostly better in younger than in older participants is consistent with a previous validation of the KORA model.<sup>39</sup> However, stratification for a variable with discriminatory value by itself reduces discrimination.

The European guideline for the prevention of type 2 diabetes<sup>10</sup> and the International Diabetes Federation<sup>9</sup> recommend the use of reliable, straightforward, and practical prediction models to identify people at high risk of diabetes. Overall, our study shows that prediction models for type 2 diabetes are valid instruments for the identification of individuals at high risk. However, this finding cannot directly be translated into a

recommendation to use a specific model, because the decision to adopt a prediction model in any specific setting is driven by many factors, of which performance is only one.

For individuals, highly accurate risk scores including biochemical testing are probably more relevant than non-invasive risk models are.8 However, non-invasive risk scores could be used as part of the public health approach to diabetes prevention to identify individuals who should receive biochemical testing. Alternatively, a risk score could be used to identify a subgroup of the population for lifestyle intervention without additional testing. Mortality risk is increased in the large group of people who have positive risk scores, justifying direct action in this group.40 Risk scores like those in the Cambridge model<sup>23</sup> and QDScore<sup>29</sup> in the UK, or the DPoRT model in Canada,<sup>25</sup> use data available from records in the countries where models were developed, and new data are not needed. Because such scores are likely to be applicable to anyone coming into contact with the health system, their performance could be better overall than that of another model with a slightly better discrimination. Nevertheless, a model for which a questionnaire is designed for self-screening at the population level (eg, the two FINDRISK models) might not be based on many responses, potentially overriding any performance advantage recorded in our analysis.<sup>41</sup>

Therefore, the decision to use a particular model could be country specific and depends on factors other than model performance, such as availability of measurements in the setting where the model is used. Additionally, prediction of diabetes is useful only when used to implement interventions to affect the outcomes of high-risk people.4 Diabetes prevention trials have been mostly based on a high-risk status defined by blood tests. However, implementation studies are necessary to investigate effectiveness of such risk scores. A few implementation studies have used risk scores to select participants for interventions and have shown favourable effects on risk factor levels42-eg, a moderate weight loss effectively reduced risk of type 2 diabetes after 1 year of intervention in participants selected with the FINDRISK score.43

Our study has some limitations. Some of the approaches used to account for predictors that were completely missing (ie, proxy variables, predictor omission, and country's exclusion) could marginally decrease discrimination of the models. It is now widely accepted that deletion of participants with missing values (who are frequent in large studies) yields biased results.<sup>39</sup> Therefore, we applied robust methods to deal with missing data.<sup>44</sup> The large sample and the multicountry nature of our study has allowed us to assess and compare 12 models across countries for the first time (panel). However, we did not include non-European participants. Therefore, our results might not be generalisable to non-European countries or other

Figure 2: Calibration curves for each model for the prediction of incident type 2 diabetes at 10 years of follow-up for the total cohort

Calibration of a model describes the extent to which the expected probability of diabetes matches the recorded probability of diabetes during follow-up. The ideal calibration (perfect agreement) is graphically represented by the dotted diagonal line at 45°. The vertical lines at the bottom of the graph depict the frequency distribution of the calibrated probabilities of diabetes. Grouped observations are for groups of participants across increasing deciles of predicted risk. ARIC=Atherosclerosis Risk in Communities. AUSDRISK=Australian Type 2 Diabetes Risk Assessment Tool. DESIR=Epidemiological Study on the Insulin Resistance Syndrome. DPORT=Diabetes Population Risk Tool. FINDRISK=Finnish Diabetes Risk Score. KORA 54/F4=Cooperative Health Research in the Region of Augsburg (KORA), Survey 4.

#### Panel: Research in context

#### Systematic review

We searched PubMed for studies published in English before July 31, 2013, with the search terms "diabetes", "risk score", "prediction model", "prediction", "risk assessment", "algorithm", "validation", "calibration", and "discrimination". Previous validation studies of diabetes prediction models have tested the performance of one model in one population or compared the performance of a few models within the same population or country,<sup>45,821,24,30-34</sup> but none has assessed and compared the performance of several models within and across several countries.

#### Interpretation

Our findings support existing recommendations that risk scores for type 2 diabetes can be used to identify individuals at high risk and start interventions to lower their future risk of type 2 diabetes. Available non-laboratory-based diabetes risk scores have acceptable to good discriminatory power in the ranking of individuals on the basis of risk across diverse settings in Europe. However, model performance differed across countries and no model outperformed the others enough to be uniquely recommended. Our results provide information to identify an appropriate risk score for each country.

ethnic groups. Finally, case definition was based on a clinical diagnosis of diabetes. Because diabetes can remain undetected, false-negative cases could have been present. Such false negatives could lead to an underestimation of the *C* statistic and affect the incidence of diabetes, affecting calibration of the risk scores. However, we adjusted for such effects of diabetes incidence by recalibrating the models.

#### Contributors

All authors contributed to study conception and design, data interpretation, and critical revision of the report. APK, JWJB, and LMP analysed the data and wrote the first draft of the report.

#### Conflicts of interest

SG received an honorarium and reimbursement of travel expenses from Eli Lilly, associated with membership of an independent data monitoring committee for a randomised trail of a drug to lower blood glucose concentrations. All other authors declare that they have no conflicts of interest.

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