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     Assessing the impact of a cattle risk-based trading scheme on the movement of bovine
     tuberculosis infected animals in England and Wales
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     Running header: Impact of bTB risk-based trading scheme
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39 The adoption of bovine tuberculosis (bTB) risk-based trading (RBT) schemes has the potential to reduce the risk of bTB spread. However, any scheme will have cost 40 implications that need to be balanced against its likely success in reducing bTB. This 41 paper describes the first stochastic quantitative model assessing the impact of the 42 43 implementation of a cattle risk-based trading scheme to inform policy makers and contribute to cost-benefit analyses. A risk assessment for England and Wales was 44 45 developed to estimate the number of infected cattle traded using historic movement data recorded between July 2010 and June 2011. Three scenarios were implemented: 46 cattle traded with no RBT scheme in place, voluntary provision of the score and a 47 compulsory, statutory scheme applying a bTB risk score to each farm. For each 48 49 scenario, changes in trade were estimated due to provision of the risk score to potential purchasers. An estimated mean of 3,981 bTB infected animals were sold to 50 51 purchasers with no RBT scheme in place in one year, with 90% confidence the true value was between 2,775 and 5,288. This result is dependent on the estimated 52 between herd prevalence used in the risk assessment which is uncertain. With the 53 voluntary provision of the risk score by farmers, on average, 17% of movements were 54 affected (purchaser did not wish to buy once the risk score was available), with a 55 reduction of 23% in infected animals being purchased initially. The compulsory 56 provision of the risk score in a statutory scheme resulted in an estimated mean change 57 to 26% of movements, with a reduction of 37% in infected animals being purchased 58 initially, increasing to a 53% reduction in infected movements from higher risk sellers 59 (score 4 and 5). The estimated mean reduction in infected animals being purchased 60 could be improved to 45% given a 10% reduction in risky purchase behaviour by 61

62	farmers which may be achieved through education programmes, or to an estimated
63	mean of 49% if a rule was implemented preventing farmers from the purchase of
64	animals of higher risk than their own herd.

66	Given voluntary trials currently taking place of a trading scheme, recommendations
67	for future work include the monitoring of initial uptake and changes in the purchase
68	patterns of farmers. Such data could be used to update the risk assessment to reduce
69	uncertainty associated with model estimates.

- 70
- 71 Keywords: risk factors, risk-based trading, bovine tuberculosis, risk scores

72 INTRODUCTION

73

Bovine tuberculosis (bTB) is an infectious disease of cattle caused by the bacterium 74 75 *Mycobacterium bovis* and is one of the biggest challenges facing the cattle farming industry in England and Wales. The cost of controlling bTB is the largest single 76 component of animal health related expenditure in these countries paid by the tax 77 78 payer, amounting to nearly £100 million in 2014 (Defra, 2014). The adoption of riskbased trading (RBT) has the potential to aid the management of livestock diseases by 79 80 providing those participating within schemes more accurate information when purchasing animals (Defra, 2013a). However, the performance of such schemes in 81 reducing the movement of infected cattle between farms is dependent on how well 82 83 schemes are implemented and the specific rules established to permit or prevent trade. 84 Risk scores can be implemented within assurance schemes or certification standards that are managed by industry organisations with a voluntary disclosure of the score, or 85 86 assisted by government with statutory controls whereby disclosure is compulsory in order for the legal sale of cattle. Scheme rules can dictate whether or not certain 87 batches are permitted to move between herds or zones of different risk scores, and 88 whether a herd score is affected by the purchase of animals of a lower risk status. 89 90

Discussions were facilitated with representatives from the farming community
(farmers, auctioneers, private veterinarians, government officials involved in
monitoring facilities, and farmer association representatives) at seven meetings during
2012-2013 in England and Wales to evaluate how informed cattle trading may vary
within different schemes that could be adopted. Understanding the basis of the
decisions made by farmers is crucial to the success of any functioning RBT scheme.

In order to parameterise the model, estimates on the expected level of RBT scheme 98 participation by farmers with the voluntary provision of the risk score was discussed 99 100 with stakeholders, alongside compliance levels that may be achieved within a 101 statutory scheme based on the compulsory provision of the risk score prior to purchase. From 25 interested stakeholders (farmers, valuers, and representatives from 102 103 non-government organisations) when asked whether cattle farmers would prefer a voluntary or statutory RBT scheme, 76% (19/25) expressed a preference for a 104 105 voluntary provision of the risk score, with all Welsh respondents opting for an initial voluntary scheme. However, concerns were frequently raised that without a statutory 106 scheme the system may not be effectively carried out and that there may be 107 108 differences in its application in different regions. It was felt that for farmers in clean 109 areas, or those that have not experienced a recent breakdown that a statutory system may be favoured. However, for those farms that had experienced a recent breakdown, 110 several stakeholders expressed the view that such farmers would not want to 111 participate in any scheme that reduced the price of their animals or where they had to 112 declare their bTB status. The engagement of farmers in RBT schemes by geographic 113 location, and the purchasing choices given different schemes, were explored and 114 quantitative estimates gained through a follow up questionnaire. 115

116

The aim of this research was to estimate the impact of farmers using risk scores to make more informed choices when buying cattle. The reduction in movements of infected cattle between farms over one year in England and Wales was estimated under three key scenarios: (1) cattle traded with no RBT scheme, (2) voluntary provision of the risk score, and (3) compulsory provision of the risk score in a

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122	statutory RBT scheme. Additionally, the impact of changes in calculating the risk
123	score were evaluated together with an investigating into areas of significant
124	uncertainty in input parameters.
125	
126	
127	METHODS
128	
129	A stochastic model implemented in Excel with the add on @Risk (version 6.1) was
130	used to estimate the number of infected movements under each of the three scenarios.
131	The final risk score developed using a method described in the accompanying paper,
132	that could be practically applied, is presented in Table 1.
133	
134	In this risk assessment each iteration in the model represents a random year with
135	convergence to 4% of the mean value of each output parameter achieved with 5,000
136	iterations using Latin Hypercube sampling. Each individual trading farm was included
137	in the model and separately simulated for the probability of being infected (between
138	herd infection), and if infected, the within herd prevalence was sampled for that herd
139	size. All historical trading events in England and Wales recorded on the Cattle
140	Tracing System (CTS) have been used (July 2010 to June 2011) to estimate the
141	number of total movements and infected movements in one year with no RBT scheme
142	in place. Movements to slaughter have not been included as such movements would

not spread infection to new herds. It is assumed that all remaining movements involve 143

a trade between a selling farm and a purchasing farm. The risk assessment uses 144

distributions for certain parameters to describe any known uncertainty or variability 145

associated with input parameters. Where uncertainty could not be quantified within a 146

147	distribution, separate scenario simulations were carried out to investigate the impact
148	on model results of the level of participation by farmers, bTB between herd
149	prevalence and purchase behaviour by farmers as detailed in the sensitivity analysis.
150	
151	Estimating the number of infected movements per year
152	
153	The number of infected movements per year is dependent on (1) the probability each
154	farm which is selling cattle is bTB infected but the infection is undetected (farm either
155	not under restriction or with specific movement license), (2) the within herd infection
156	prevalence on that farm, (3) the proportion of animals moved from that farm in
157	batches to other farms, and (4) the sensitivity of the pre-movement test where applied.
158	The risk pathway for the movement of infected animals off farm is provided in Figure
159	1. Numerous parameter values were extracted from the National database SAM
160	RADAR bTB reception database, herein referred to as SAM.
161	
162	Probability farm infected with bTB, P_{inf}
163	
164	For each farm in the dataset the probability of the herd being bTB infected, P_{inf} was
165	estimated using a modified freedom from infection (FFI) model (AHVLA, 2011).
166	This model has been previously developed to estimate the probability that a given
167	herd was free of infection given its test and disease history, $P(free)$ (Martin et al.,
168	2007) and is described in the accompanying paper. There is considerable uncertainty
169	associated with the probability of a herd being infected with bTB which is
170	investigated in the sensitivity analysis. For each iteration, each selling farm is either

infected or not, modeled as a Bernoulli random variable, based on the probability ofinfection per year estimated for that farm.

$$P_{inf} \sim Binomial(1, 1 - P(free))$$

173

174 Number of animals infected, N_{Inf}

175 The number of infected animals in a herd is dependent on the within herd bTB

176 prevalence and the number of animals within that herd. From a review of the

177 literature, the within herd bTB prevalence applicable to undetected infected herds of

varying herd size in England and Wales was not available. To calculate, we first

179 estimated the annual number of infected animals in herds, Inf_{year} , where routine

180 whole herd testing had been carried out in 2011. Where disease is not suspected,

181 whole herd tests are conducted with the single intradermal comparative cervical

tuberculin test (SICCT) test. Given the mean sensitivity of the SICCT test, Se_{mean} ,

together with the total number of test positive reactors identified in whole herd tests S_{year} (SAM) in England and Wales, the negative binomial distribution was used to

185 describe the total annual number of infected animals in tested herds:

186
$$Inf_{year} \sim Negbin(S_{year} + 1, Se_{mean}) + S_{year},$$

187

The estimated within herd prevalence for individual herds, P_{prev} was then sampled from the surveillance dataset, representing those herds assumed to be infected, such that the cumulative estimated number of infected animals per year across herds equalled the expected number infected per year Inf_{year} . This subset included herds where no reactors had been found (*S*=0)

193
$$P_{prev} \sim \frac{Negbin(S+1,Se)+S}{h},$$

194	where S denotes the number of reactors per surveillance herd identified by the SICCT
195	test in 2011 (SAM), Se is the sensitivity of the SICCT test, and h is the total number
196	of animals tested in that surveillance herd (SAM). The negative binomial distribution
197	was truncated to ensure that the number infected in an individual herd (reactors and
198	false negatives) was not greater than the total number of animals tested in that
199	surveillance herd. The distribution of bTB within herd prevalence was generated from
200	500,000 iterations to ensure convergence to 4% of the estimated mean. Results were
201	filtered to include only those iterations where the observed 2011 England and Wales
202	reactor herds were included in the subset and are provided in Table 2.
203	
204	The distribution of the sensitivity of the SICCT test at the herd level was described
205	using the Beta distribution with values of $\alpha = 6.66$ and $\beta = 6.37$ (Downs et al., 2011).
206	At the national level, Se_{mean} , a mean sensitivity of 0.511 was used for the SICCT
207	test. The estimated prevalence of bTB on infected farms, not previously suspected of
208	disease, decreases with increasing herd size, following the same trend as the
209	prevalence of detected reactors on infected farms. Note, this is not the probability of a
210	farm being infected, but the level of infectivity on farms that are infected. Separate
211	cumulative probability distributions representing the uncertain within herd prevalence
212	by herd size were applied in the model. Given the estimated within herd prevalence, a
213	binomial distribution was used to estimate the variable number of infected animals on
214	each infected farm from the total number of animals on farm:

$N_{Inf} \sim Binomial(Herdsize, P_{prev})$

215 where *Herdsize* was the average number of animals on farm (SAM).

217 Allocation of infected animals to off movements or remaining on farm, $N_{Inf_{total}}$

218

219 Each selling farm may move animals off to a number of different locations during one year. Paired movements between all farms between July 2010 to June 2011was 220 221 extracted using the Cattle Tracing System (CTS). The estimated number of infected, N_{inf} being allocated to these different batches moved off farm, or remaining on the 222 farm, was assumed not to be dependent on animal infection status. The probability of 223 any one infected animal being allocated to a batch was therefore equal to the number 224 225 of animals sold in that batch divided by the original total number of animals in the 226 herd. For most farms there was more than one batch movement sold per year. Therefore, a multinomial distribution was implemented as a set of nested binomial 227 228 distributions to describe the between year variability for allocation of infected animals to batches or remaining on farm: 229

$$\begin{split} N_{Inf_{total}} &\sim Multinomial(N_{Inf}, \{P_{farm}, P_{batch1}, P_{batch2} \dots P_{batchn}\} \\ N_{Inf_{total}} &= N_{inf_{farm}} + N_{inf_{batch1}} + N_{inf_{batch2}} \dots + N_{inf_{batchn}} \end{split}$$

where $N_{Inf_{farm}}$ is the number of infected animals allocated to remain on farm, and 230 $N_{Inf_{batchn}}$ the number allocated to batch n. Where the selling farm is located within 231 an area subject to annual or bi-annual bTB tests (areas of high bTB incidence), all 232 cattle over 42 days of age require a pre-movement test to be taken 60 days prior to 233 movement. Within the risk assessment it is assumed that all animals originating from 234 farms located in the high risk area are tested and. This is a simplification as there are 235 236 movements which would be exempt from testing including animals under 42 days and those licensed between Approved Finishing Units (AFUs) and certain farms under 237 restriction. It was assumed that each infected animal had the same likelihood of 238

239 testing positive in the absence of any latent period included in the model. A binomial random variable with the number of infected animals in that batch and the sensitivity 240 241 of the SICCT test, Se, was sampled for the variability associated with a positive premovement test. Given any positive results it was assumed that the entire batch was not 242 sold. Detection of positive animals in the pre-movement test would result in trading 243 restrictions placed on the farm thereafter. However, given that all movements occur in 244 245 one annual time step with no chronological order, the assumption was made that batch results were independent from other batch results for that source farm. This 246 247 simplification made does not affect the comparison of RBT schemes because the entire batch is removed from all schemes for that iteration." 248

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250 Estimating the impact of a voluntary scheme

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252 This scheme was based on the risk score of the seller (S_{score}) , being made available voluntarily to auctioneers and purchasers prior to purchase by the seller. The risk 253 score of the purchaser (P_{score}) , may influence which animals they buy. The risk 254 pathway for one selling farm is shown in Figure 1 and was used to estimate the 255 256 infected and uninfected animals in each batch. This risk pathway was extended with an example batch as shown in Figure 2 to take account of whether or not the 257 purchaser participates in a scheme and, given participation, whether or not the 258 purchaser accepts the risk score of the seller. A 'failed initial movement' occurs when 259 260 the purchaser does not accept the sellers risk score.

261

262 Probability of participating in a trading scheme, $P_{scheme(P_{score}, P_{region})}$

The percentage of farmers that would be likely to purchase through a voluntary RBT 264 scheme was discussed at seven meetings with stakeholders during 2012-2013 in 265 266 England and Wales, with a follow up questionnaire (available from corresponding author). There were 17 quantitative estimates received. Stakeholders felt that there are 267 many dependencies to be factored into estimates generated including the individual 268 bTB status and circumstances of the purchaser and how successfully the scheme was 269 270 rolled out. For Wales, it was deemed that the level of uptake of an RBT scheme would differ by region. Therefore, different estimates for uptake were calculated for regions 271 272 defined as Low risk and High risk. Estimates were also stratified by purchasers risk score as it was thought that incurring a breakdown in recent years would influence the 273 purchasing farmers' behaviour. The effect of differences in the purchasing relating to 274 farm herd type was also raised. For example finishing farms (animals fattened for 275 slaughter) were considered less likely to be concerned about the bTB risk of animal 276 than breeding farms, however, insufficient data were available to include stratification 277 by farm type in the model. The opinion elicited is provided in Table 3. The probability 278 of farmers purchasing through a voluntary scheme was associated with significant 279 unquantified uncertainty which was further investigated in the sensitivity analysis. 280 Over one year it was assumed that each batch purchaser elected either to participate in 281 the scheme or not for all batches destined for that farm represented by a Bernoulli 282 283 random variable.

284

Probability of purchase given risk score, $P_{buy(P_{score}, S_{score})}$

286

For those farmers participating in the scheme, the probability that farmers will buy certain animals will depend on their own farm status, their risk appetite, and also on

289 the information provided by the score regarding the animals for sale. As with the percentage of farmers using the scheme, there will be considerable variability between 290 291 farmer needs (breeding farmer purchasing versus farmer restocking large numbers), other factors, such as the price of the animal, and on the overall 'trust' a farmer places 292 in the risk scores and on the local implementation of the RBT scheme including the 293 amount of educational activities rolled out with schemes. Stakeholders were asked to 294 295 consider a hypothetical farmer that was interested in using risk scores. For each risk score pairing (seller score - purchaser score), respondents were asked to select a 296 297 probability ranging from "Will" to "Will not" divided into six increments. Each of the boxes was associated with a probability, with a maximum of 100% representing 298 "Will" and minimum of 0% for "Will not" with 1%-25%, 26%-50%, 51%-75%, and 299 300 76%-99% for the middle four boxes. There were 12 quantitative responses provided 301 with 5 unknowns (5 stakeholders did not answer this question in the questionnaire). A discrete distribution was then simulated until convergence for each pairing to estimate 302 303 the combined expert opinion mean, maximum and standard deviation of the associated uncertainty. The uncertain probability of purchase for each pairing of risk 304 score between purchaser and seller was applied in the risk assessment using a fitted 305 lognormal distribution using the key statistics of the distribution shown in Table 4. A 306 Bernoulli random variable with the given probability was sampled for the variability 307 308 associated with the decision to purchase given the risk score.

309

Estimating the number of infected movements within a statutory scheme

310

The statutory scheme was based on the compulsory provision of the risk score to auctioneers and purchasers prior to purchase. In a perfect system this would imply that all purchasers would be involved in the scheme with $P_{scheme(P_{score}, P_{region})}=1$.

However, the potential for purchase of animals from farmers not using the system was discussed with stakeholders and a minimum of 5% and maximum 15% thought to be plausible bounds for the uncertain probability of not complying with the scheme, with a mean value of 10%. For those farmers participating in the statutory scheme, the probability of the purchase being made ($P_{buy(P_{score}, S_{score})}$), given that the risk score was made available was assumed to be the same as that estimated within the voluntary scheme.

321

322 Estimating the impact of changes to calculating the risk score

323 The baseline risk score for each farm, as described in the accompanying paper, was based on selected risk factors from a full model identified by a logistic regression. The 324 impact of including some of the removed risk factors (region risk West England and 325 326 Wales, and breakdowns > 10 years previously) on the performance of the score was investigated together with a more simplified scheme (only 0-2 years since breakdown 327 and breakdown information without high risk movements), and finally the impact of 328 implementing a rule whereby farmers are not permitted to purchase animals of higher 329 risk status than their own herd. 330

331

332 Parameter uncertainty and sensitivity analysis

333 During development of the risk assessment several key parameters were identified as

being uncertain with little available information to describe that uncertainty.

335 Therefore, upper and/or lower limits of parameters were identified and implemented

in separate simulations of the risk assessment:

338	(1) The between herd prevalence of bTB $p(inf)$, calculation uses a value from the
339	literature that herds cannot achieve a probability of freedom greater than 62%
340	for 24 months post breakdown (detailed in the accompanying paper). The
341	uncertainty associated with this value is not known. To estimate the impact of
342	this uncertainty, the probability of infection for each farm was increased and
343	decreased by 5% and separately simulated.
344	(2) The level of participation of farmers in a voluntary RBT scheme,
345	$P_{scheme(P_{score}, P_{region})}$ was acknowledged in discussions as being highly
346	uncertain - relating to farmer trust in that RBT scheme and ease of use and
347	accessibility. Model scenarios were run at levels of 20%, 40%, 60%, 80% and
348	100% farmer participation to evaluate the relationship between participation
349	and performance of the scheme.
350	(3) The probability farmers would still purchase high risk animals once bTB
351	information was provided, $P_{buy(P_{score}, S_{score})}$, was associated with the
352	purchasers status and the amount of education and explanation that
353	accompanied the roll out of any scheme, which at present is uncertain. To
354	investigate the impact of RBT schemes that change the baseline probability of
355	buying higher risk animals, a scenario was simulated where all purchasing
356	farmers were 10% more likely and 10% less likely to purchase higher risk
357	animals than the values elicited for the baseline model.
358	
359	A sensitivity analysis based on Analysis of Variance (ANOVA) was undertaken. An
360	ANOVA was selected as it has previously provided robust insights regarding
361	identification of key inputs in probabilistic risk assessments, for example, Mokhtari
362	and Frey, 2005. The reduction in infected movements comparing no RBT scheme and

a statutory RBT scheme at 90% compliance per farm was used as the response
variable. Predictor variables were values of each input parameter for that farm
represented by a range. The ANOVA was populated with half a million randomly
selected farms.

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368 RESULTS

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The trade in cattle between farms without a RBT scheme, with a voluntary RBT 370 371 scheme, and with a statutory RBT scheme, were simulated over one year for each farm. The number of infected movements that, if pre-movement tested, batch tested 372 clear was summed and stratified by country and area. It was assumed in the baseline 373 and each scenario that all movements from herds in the high risk area were pre-374 movement tested. Uncertainty and variability considered in the model was represented 375 by 5th and 95th percentiles (within parentheses), which indicate the range within which 376 90% of the results lie. Uncertainty was also considered in separate scenario runs of the 377 risk assessment. It should be emphasised that not all variability and uncertainty has 378 been estimated in the calculations and scenarios, as not all can be quantified. 379 Therefore results describe the amount of quantified variability and uncertainty 380 included in the assessment. Results stratified by region and by farm risk score, are 381 presented in the supplementary materials. 382

383

Results with no RBT scheme

For trade in cattle with no RBT scheme there were 379,951 batches of animals moved

off farm in England and Wales to another farm in England and Wales where the risk

387 score and region of the seller and purchaser was determined. As shown in Table 5,

388	this represented a total of 1.2 million animals with 18.4%, 43.3%, 5.6%, 10.8%, and
389	22.0% of animals sold by farms scoring 1, 2, 3, 4, and 5 respectively, where a score of
390	1 is the lowest risk and a score of 5 is the highest risk score. An estimated mean of
391	35,588 infected animals were on farms from which off movements occurred (farms
392	not under restriction or those restricted but with a specific licence to move to another
393	restricted facility) with 5 th and 95 th percentiles that this varied between 32,881 and
394	38,369. Of these infected animals, approximately 11% or 3,981 (5 th 2,775, 95 th 5,288)
395	were sold to purchasers in England and Wales with the majority remaining on farm.
396	Of those 3,981 infected animals per year, an estimated mean of 41.8% infected
397	animals were sold by farms scoring 5, rising to an average 60.2% for farms scoring 4
398	or 5, whilst 6.1% were estimated to be sold from the lowest risk farms scoring 1.

400 Voluntary RBT scheme

401 Uptake by farmers for a voluntary RBT scheme was estimated to vary between 40% 402 to 81%, as shown in Table 3, dependent on location and purchaser bTB status. Table 5 presents the estimated results from implementation of a voluntary RBT scheme with 403 approximately 17% of animals that were traded with no RBT scheme being rejected 404 by the initial buyer. It can be seen that the estimated trade from lower risk sellers was 405 found to be less affected, with trade from higher risk sellers being most affected to 406 low risk purchasers. The estimated trade was most affected in the high risk areas in 407 England and Wales (regional differences shown in supplementary materials). There 408 was an estimated mean rejection of 23% (5th 22%, 95th 25%) of infected animals by 409 purchasers based on sellers providing the risk score voluntarily. 410

412 Statutory RBT scheme

Under a statutory RBT scheme with an estimated mean compliance of 90% of 413 414 purchasers having access to the risk score of the seller an estimated mean of 26% of animals were rejected once the risk scores were made available. The majority of 415 estimated trade to low risk purchasers (score 1) from high risk sellers (score 4 or 5) 416 417 was affected by the implementation of a statutory scheme. Of the estimated number of infected animals on farm a mean of 37% (5th 35%, 95th 39%) of infected animals were 418 rejected by purchasers. Of those infected animals rejected from sellers, the majority 419 420 are estimated to be those sold by high risk farms (score 4 or 5), with on average a 53% reduction in infected movements from those farms. 421

422

423 Alternative schemes

Figure 3 displays the boxplot of different RBT schemes according to the estimated 424 425 mean percentage reduction of infected movements. Results using the baseline risk score are presented in dark green and highlights the linear relationship between the 426 percentage uptake by farmers and the percentage reduction achieved by that scheme. 427 428 The dark green dashed line through the simulation results represents the uncertainty regarding the level of uptake for each scheme. The dashed black vertical lines through 429 each box plot represent the between year variability and uncertainty about the mean 430 431 simulation result and terminate at the estimated minimum and maximum value. Variations on the baseline risk score used in an RBT scheme, adding or subtracting 432 certain risk factors from the scoring system (as described in the accompanying paper) 433 at 90% compliance has been provided together with an extrapolation of how those 434

435 schemes would perform. From the results it can be seen that there are only marginal

increases in the performance of the scheme given the addition of risk factors selected from the logistic regression (region risk West England and Wales, and breakdowns > 10 years previously). The impact of a ban on farmers purchasing below their farm risk score, assumed to be implemented with 100% compliance yields a 49% reduction the initial purchase of infected animals (5th 47%, 95th 51%).

441

442 Parameter uncertainty and sensitivity analysis

443

There were three important parameters identified by the ANOVA: (1) the uncertain probability of the purchaser buying the animal once the sellers score was shown (derived from expert opinion) $P_{buy(P_{score}, S_{score})}$, (2) the variable risk score of the seller, S_{score} , and (3) the variable risk score of the purchaser, P_{score} . It should be noted that the uncertain level of compliance for the statutory scheme, $P_{scheme(P_{score}, P_{region})}$ was significant but less significant than the top three. For the voluntary scheme, the uncertainty associated with the probability of participating in

451 the scheme was also highly important.

452

453 In addition to the sensitivity analysis, scenarios were identified during model

454 development and parameterisation where there was limited information on parameter

uncertainty with results shown in Table 6 and displayed in the boxplot in Figure 3.

456 The true between herd prevalence of bTB infection, P_{inf} , the proportion of herds that

457 have at least one infected animal, is associated with considerable uncertainty from the

458 freedom from infection model (AHVLA, 2011) which is heavily reliant on input

459 assumptions. Using alternative parameterisations, the performance of RBT schemes

460 was within the convergence values for the original parameterised simulations. This is

due to the fact that the percentage change in infected movements is not dependent on the scale of the true prevalence, only the pattern of the true prevalence across English and Welsh farms. However, the absolute number of infected movements per year was significantly affected. Decreasing the between herd prevalence by 5% decreased the number of infected movements by a mean of 22%, whilst increasing by 5% increased the average number of infected movements by 21%.

467

Simulations were carried out varying the percentage uptake by farmers and the percentage reduction achieved. For every 10% of farmers that participated in the baseline scheme there was an additional 3.8% reduction in the initial purchase of infected animals until the mean estimated maximum of 38% was reached at the maximum of 100% participation.

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The greatest increase in performance of the score arose from a 10% decrease in the
baseline estimates for risky farmer behaviour (purchasing cattle at higher risk than
their own farms) with a 45% mean reduction in the initial purchase of infected
animals (5th 43%, 95th 47%). This result concurs with the identification in the
ANOVA of this parameter as having the highest impact on the RBT performance
output considering the associated quantified uncertainty and variability.

481 DISCUSSION

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Cattle trading patterns are complex and dynamic due to seasonal factors, economic
factors and changes in Government controls. Nevertheless a quantitative approach to
estimating the impact of a RBT scheme was possible for England and Wales. It was

486 possible to estimate with a reasonable amount of confidence the impact of a specific 487 scheme over one year and show that a significant impact could be achieved with the 488 reduction of movements from high risk areas or high risk farms.

489

One of the major reasons for adopting a quantitative approach was the need to account 490 for the dynamic movement patterns between farm types and farm areas and regional 491 492 differences in the application of control measures. Historic paired movements were used which linked direct farm to farm animal movements and those via markets to 493 494 farms. This allowed a comparison between high and low risk areas and different trading schemes. The absolute results for the number of animals infected and traded 495 was dependent on the scale of the between herd and within herd prevalence. The 496 497 between herd prevalence was associated with uncertainty not quantified in the model. 498 However, the comparison between cattle traded with no RBT and the different RBT schemes was not dependent on the magnitude of prevalence – only the regional or 499 farm characteristic pattern. It was apparent that changes in the calculation of the 500 between herd prevalence could have a significant effect on the absolute number of 501 infected movements predicted. The provision of values for the number of infected 502 animals with associated uncertainty is, however, provided as such values are 503 important for economic analyses when considering the cost benefits of establishing 504 and maintaining a RBT scheme. Before consideration could be made of a statutory 505 scheme, a cost-benefit analysis would be required estimating the full costs of 506 implementing a scheme, such as impacts on trade and adjustments of the market, 507 together with the benefits of reduced disease spread. 508

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510 Analysis of the results from the risk assessment demonstrated the importance of encouraging maximal uptake of schemes. The sensitivity analysis and parameter 511 512 uncertainty scenarios demonstrated the importance of farmer purchase behaviour, $P_{buy(P_{score},S_{score})}$ on the performance of any RBT scheme. The quantified 513 514 uncertainty associated with this parameter could be reduced from gathering 515 appropriate data from any pilots conducted. In addition, careful consideration should be given to any programme of education of farmers which could result in reducing 516 risky purchase behaviour, thereby considerably improving the performance of RBT 517 schemes. Importantly, we repeatedly heard at stakeholder meetings that many farmers 518 519 believed that if an animal had been tested for bTB, then that animal was not infected, i.e. they considered that the bTB test applied was 100% sensitive. This may lead to 520 the conclusion that further effective education of farmers may be warranted. The 521 522 England TB RBT group also identified that a voluntary scheme will only succeed if a critical mass of farmers participate (Defra, 2013a). This will depend on how well any 523 scheme is rolled out, ease of use, trust, the level of understanding achieved of the risk 524 posed by purchasing cattle to herds and sufficient information being made available to 525 farmers to make an informed choice. 526

527

In the absence of any RBT scheme being piloted in England and Wales during the lifetime of this research project, the values elicited by expert opinion represented a 'best guess', however, it is the only data currently available. Should any schemes be piloted, it would be advisable to monitor initial uptake and changes in farmer behaviour to update the risk assessment. For example, Gates and colleagues monitored the change brought about by cattle movement restrictions on Scottish farms (Gates et al., 2013). Such data would be invaluable to reduce the uncertainty

associated with model estimates. Given sufficient data, further work could investigate 535 the most likely fate of those infected movements that initially fail from high scoring 536 537 sellers. The England RBT group commented that a short research project be conducted after an introductory period to investigate engagement and behavioural 538 change. This may include a survey of auctioneers as to whether any risk-based trading 539 data has been included in catalogues or on screen/boards at point of sale and how 540 541 many buyers are asking for the risk score prior to purchase. Statutory databases could also be queried as to whether any significant changes had occurred to paired 542 543 movements (particularly those deemed the most risky) between/into/out of selected geographical/incidence based/score based categories. An alternative would be a check 544 on the average distance travelled for movements from holdings of certain categories. 545

546

A RBT scheme would reduce infection transmission attributable to cattle movements 547 which is one transmission pathway contributing to the bTB epidemic (Gopal et al., 548 2006). This would reduce the between herd prevalence (the proportion of farms with 549 at least one infected animal). In the risk assessment, historical movements are either 550 accepted or rejected; the model makes no attempt to reallocate the movement to 551 another farm or area once the original trade is declined. However, at the market or 552 sale, another farmer may purchase the rejected batch at a lower price. Alternatively 553 554 farmers with high scores may seek out other purchasing farmers with the same risk status for trade, for example, with the development of 'orange' markets. The model 555 indicates that, given the introduction of a RBT scheme, there would be significantly 556 less infected animals purchased by low scoring farms, particularly for those low risk 557 farm that are located in the high risk area (HRA). If those rejected movements were 558 sold to high risk farms, which may already be harbouring undetected infection, this 559

560 may, in the long term, increase the bTB within herd prevalence of those herds engaging in this risky behaviour. Unfortunately, the risk assessment is simulated only 561 over one year and therefore cannot quantify the long-term changes that may eventuate 562 from implementation of risk-based schemes, however, if such farms resided in an area 563 of higher testing frequency, such as the HRAs in England and Wales, detection of 564 those infected animals may occur earlier due to a higher prevalence of infection on the 565 566 test farm, and increased frequency of testing in the form of pre-movement tests and annual whole herd tests thus complementing and potentially improving the sensitivity 567 568 of the current regional controls in place.

569

570 CONCLUSIONS

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In conclusion, this paper details the design of the first risk assessment to measure the impact of theoretical risk-based animal trading schemes based on a given farm risk score for bTB. If a voluntary or statutory RBT scheme was in place, a significant impact could be achieved with the reduction of infected movements from high risk areas or high risk farms. Key to reducing infected movements through a risk-based trading scheme is promoting maximal uptake in schemes and on reducing risky farmer purchase behaviour.

579

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