1	Quantifying similarity in animal vocal sequences: which metric performs best?
2	
3	Arik Kershenbaum ^{1*} , Ellen C. Garland ²
4	
5	¹ Department of Zoology, University of Cambridge, Cambridge, England
6	² School of Biology, University of St. Andrews, St. Andrews, Fife, KY16 9TH, Scotland
7	
8	*Author for correspondence: arik.kershenbaum@gmail.com
9	
10	Short title: Measuring sequence similarity
11	
12	Word count: 7431
13	
14	

15 SUMMARY

Many animals communicate using sequences of discrete acoustic elements which can be complex,
 vary in their degree of stereotypy, and are potentially open-ended. Variation in sequences can
 provide important ecological, behavioural, or evolutionary information about the structure and
 connectivity of populations, mechanisms for vocal cultural evolution, and the underlying drivers
 responsible for these processes. Various mathematical techniques have been used to form a
 realistic approximation of sequence similarity for such tasks.

Here, we use both simulated and empirical datasets from animal vocal sequences (rock hyrax,
 Procavia capensis; humpback whale, *Megaptera novaeangliae*; bottlenose dolphin, *Tursiops truncatus*; and Carolina chickadee, *Poecile carolinensis*) to test which of eight sequence analysis
 metrics are more likely to reconstruct the information encoded in the sequences, and to test the
 fidelity of estimation of model parameters, when the sequences are assumed to conform to
 particular statistical models.

Results from the simulated data indicated that multiple metrics were equally successful in
 reconstructing the information encoded in the sequences of simulated individuals (Markov chains,
 n-gram models, repeat distribution, and edit distance), and data generated by different stochastic
 processes (entropy rate and n-grams). However, the string edit (Levenshtein) distance performed
 consistently and significantly better than all other tested metrics (including entropy, Markov
 chains, n-grams, mutual information) for all empirical datasets, despite being less commonly used
 in the field of animal acoustic communication.

4. The Levenshtein distance metric provides a robust analytical approach that should be considered
in the comparison of animal acoustic sequences in preference to other commonly employed
techniques (such as Markov chains, hidden Markov models, or Shannon entropy). The recent
discovery that non-Markovian vocal sequences may be more common in animal communication
than previously thought, provides a rich area for future research that requires non-Markovian
based analysis techniques to investigate animal grammars and potentially the origin of human
language.

42 Keywords: Sequence, animal communication, vocal, edit distance, Markov, stochastic processes

43 INTRODUCTION

44 Many animals communicate using sequences of discrete acoustic elements, the best known example 45 being bird song, which is composed of multiple notes combined in a distinctive order. These 46 sequences are often complex, non-stereotyped, and potentially open-ended; that is, individuals may 47 use an almost unlimited repertoire of sequences by making subtle or large variations to the order of 48 notes (reviewed in Catchpole & Slater 2003). The role of such sequences varies among species. In 49 some cases, sequences appear to advertise male quality through sequence complexity, e.g., in marsh warblers, Acrocephalus palustris (Darolová et al. 2012); zebra finches, Taeniopygia guttata (Holveck 50 51 et al. 2008; Neubauer 1999; Searcy & Andersson 1986); and song sparrows, Melospiza melodia (Pfaff 52 et al. 2007). In other cases, researchers have proposed that sequences contain detailed communicative 53 information such as individual identity, e.g., bottlenose dolphins, Tursiops truncatus (Sayigh et al. 54 1999). It is also possible that in some species, acoustic sequences are essentially stochastic with little 55 significance to their precise composition.

56 Identifying the role of acoustic sequences in a particular species often involves comparing sequences 57 within and between individuals, as well as within and between populations, so that the nature of the variation can be quantified and potentially correlated to ecological or behavioural factors. The task of 58 59 comparing acoustic sequences presumes an unequivocal and globally relevant measure of sequence 60 similarity, or difference. However, in practice, no such metric exists. It could be postulated that a measure of sequence similarity should reflect the proximal processes taking place in the brains of 61 62 intended conspecific signal receivers; i.e., the best measure of sequence similarity is the one used by the animal itself (Kershenbaum et al. 2014). Given that such knowledge is essentially hidden in 63 64 practice, various mathematical techniques have been used to form a realistic approximation of signal 65 similarity (Ashby & Perrin 1988; Navarro 2001; Ranjard 2010; Young & Hamer 1994). It is possible to categorise similarity measures into two distinct approaches. Firstly, it is usually possible to 66 67 characterise a sequence by measuring a small number of metrics that are inherent to the sequence itself; examples of this include length, or entropy (Freeberg & Lucas 2012). Sequences can then be 68 69 compared by calculating the sum of square differences between each of these metrics. This is

70 equivalent to representing each sequence as a "feature vector" in some relatively compact feature 71 space, and measuring the distance between two sequences as the Euclidean distance between their two 72 feature vectors. While this method is straightforward, there is an assumption that it is possible to 73 represent every sequence in a compact way, i.e., that some sufficiently large combination of metrics 74 can "summarise" the properties of a sequence in a biologically meaningful way. However, it is far 75 from clear that there exists a compact, yet exact, mathematical representation of a sequence, short of 76 the trivial task of writing down the entire sequence of elements and attempting to measure the Euclidean distance between the full representations of two sequences, which is unlikely to produce the 77 desired results. An alternative approach is to use aggregate techniques that measure properties of a 78 79 large number of sequences, and summarise the characteristics of a corpus. For example, sequence 80 transition tables and element frequency histograms have been used in previous studies (Jin & 81 Kozhevnikov 2011). In these cases, each vector in feature space represents a collection of sequences, 82 and the Euclidean distance between vectors measures the difference between the sequences from two 83 sets of vocalisations, rather than between individual sequences. However, it is questionable whether 84 any of these techniques, individual or aggregate, can represent the nature of the sequences with 85 adequate fidelity. Since we do not know what cognitive processes an animal uses to interpret such 86 sequences, we cannot be sure that any particular summary metric accurately reflects the interpretation 87 of the sequence by the receiving individual. We refer to all of these above metrics as "unary", as they 88 are derived from measurements on each string sequence in isolation, even if distances are eventually 89 calculated on an aggregate of sequences.

Secondly, it is possible to measure the difference between a pair of sequences directly (Levenshtein 1966), thereby bypassing the construction of a feature space, and generating a series of pairwise comparisons between sequences. Analysing the sequence of elements in animal vocalisations can be considered analogous to analysing the sequence of nucleotides in DNA, and some non-aggregate techniques have been borrowed from the field of bioinformatics to capture the similarity or difference between two sequences. This approach provides a direct measure of pairwise differences, in the form of a distance matrix, but without a Euclidean feature space. We refer to these metrics as "binary", as

97 they can only be calculated as a pairwise comparison between exactly two sequences. Binary 98 difference measures are attractive, as they do not rely on the fidelity of a particular unary metric in 99 representing the properties of a sequence. Rather, binary metrics are an unequivocal measure of the 100 similarity/difference between two sequences; although it cannot be assumed that this measure of 101 similarity is the same as that used by the animal itself in distinguishing between sequences. Such 102 metrics have long been proposed for the analysis of birdsong (Bradley & Bradly 1983; Ranjard et al. 103 2010), but have not been widely adopted. One disadvantage of binary metrics is that a number of 104 common machine learning algorithms often used for clustering the results of similarity analyses (e.g., 105 k-means, neural networks) rely on data presented as a Euclidean feature space, although there are 106 exceptions, e.g. Ranjard & Ross (2008). To use such clustering techniques, it would be necessary to 107 derive a series of feature vectors from the binary metric distance matrix. This can be done using 108 techniques such as multidimensional scaling or principal component analysis to convert a distance 109 matrix to feature vectors.

110 Here, we compare the performance of eight different methods for analysing animal vocal sequences, 111 using both aggregate statistical metrics and a direct pairwise distance measure. We use simulated and 112 empirical sequences to test which approach is more likely to reconstruct the information encoded in 113 the sequences, and to test the fidelity of estimation of model parameters when the sequences are 114 assumed to conform to particular statistical models. This direct comparison of a number of commonly 115 employed analytical algorithms provides a comprehensive evaluation of the utility of these 116 approaches to real-world data sets, and demonstrates the utility of comparing at least two different 117 methods when assessing novel algorithms to ensure that results are robust under a range of analytical 118 approaches.

119

120 METHODS

We performed two sets of tests (viz. artificial and empirical) to evaluate the performance of each metric. In the first tests, we generated artificial random sequences and used the different similarity

123 metrics to reconstruct the parameters used to generate these sequences, and the stochastic model 124 types. In the second set of tests, we analysed recordings of animal vocalisations and used both unary and binary difference metrics to determine contextual information known to exist in these sequences. 125 126 We used the signature whistles of the bottlenose dolphin (Kershenbaum, Sayigh & Janik 2013; Sayigh 127 et al. 2007; Sayigh et al. 2007), to reconstruct individual identity, and the songs of the rock hyrax, Procavia capensis (Kershenbaum et al. 2012), the humpback whale, Megaptera novaeangliae 128 129 (Garland et al. 2012), and the calls of the Carolina chickadee, Poecile carolinensis (Freeberg 2012), 130 to reconstruct geographical dialect. In the case of the hyrax, humpback whale, and chickadee, the calls 131 consisted of a sequence of discrete acoustic elements. In contrast, bottlenose dolphin whistles are 132 often produced in isolation (rather than as a sequence of whistles); therefore we analysed the sequence 133 of frequency modulation components (e.g., up, down, constant) within whistles, taking these 134 modulation components as the acoustic elements (for more details see Kershenbaum, Sayigh & Janik 2013). In both our analysis of artificial sequences, and empirical animal vocal sequences, we evaluate 135 136 a number of similarity metrics, both binary and unary. Humpback whale song recordings are held at 137 the University of Queensland, Australia, and by Operation Cetaces in Noumea, New Caledonia. 138 Dolphin whistle recordings are held at Woods Hole Oceanographic Institution (see Data Accessibility 139 section for contact details). Before providing details of the simulation experiments and empirical data 140 analysis, we describe each of the metrics used.

141

142 Binary metric

143 Levenshtein distance (LD)

144 The Levenshtein distance (Levenshtein 1966) is a type of string edit distance metric, as it provides a

145 quantitative measurement of the difference between two string sequences regardless of string length.

146 Specifically, the Levenshtein distance measures the minimum number of point operations (additions,

147 deletions, and substitutions) needed to convert one string into another (Levenshtein 1966). By

148 comparing the position of elements within a string and calculating the number of changes that it takes

149 to change one string into the other, this metric relies more on the sequence of elements and less on the 150 overall structural pattern. It has been used extensively in other fields, e.g., bioinformatics (Likic 2008) 151 and text search/retrieve (Reis et al. 2004), and in a small number of previous studies of animal 152 sequences (e.g., Garland et al. 2012; Garland et al. 2013; Kershenbaum et al. 2012; Krull et al. 2012), 153 and is related to the better known dynamic time warping algorithm (Buck & Tyack 1993). However, LD itself remains somewhat unknown in the field of animal acoustic communication. In practice, 154 155 string edit distances are often paired with string alignment algorithms or additional standardisations, particularly when the strings being compared are of different lengths: Figure 1; see Kershenbaum et 156 157 al. (2012) and Garland et al. (2012) for additional information on metric calculation. Importantly, the 158 Levenshtein distance forms the basis of the Needleman-Wunsch string alignment (Likic 2008; 159 Needleman & Wunsch 1970) that is used extensively in bioinformatics research to compare sections 160 of DNA. In our implementation of the LD algorithm, we assign an equal cost (of 1) to any correction 161 operation (addition, deletion, substitution), no cost (0) for a matching element, and no cost for 162 differences in sequence lengths after optimal alignment. 163 Although other binary metrics exist apart from LD, they are in general unsuitable for the task at hand.

For example, the Hamming distance requires sequences of the same length, and the most frequent kcharacters simply provides a count of the most common symbol/element. These therefore provide less information than the Levenshtein distance metric.

167

168 Unary metrics

169 **Transition table (TT)**

170 Acoustic sequences have often been modelled as a Markov chain (Berwick *et al.* 2011; Briefer *et al.*

171 2010; Briefer et al. 2010), in which the probability of a particular element occurring depends only on

the preceding element (or sometimes, more than one preceding element). These conditional

173 probabilities of each element, given the preceding element(s), can be expressed as a transition matrix

174 *T*, in which the element $T_{i,j}$ represents the probability of the element *j* occurring after the element *i*.

For a sequence consisting of *C* distinct element types, a *C* x *C* transition matrix can be estimated from empirical data. When comparing two sequences A and B, the similarity between the transition matrices T_A and T_B is an indication of the similarity between the sequences (Jin & Kozhevnikov 2011). To calculate a difference metric $D_{TT} = f(T_A, T_B)$, we can express each matrix as a C^2 dimensional feature vector *V*, where the elements of the vector are equal to the elements of the transition matrix *T*, i.e., $V = T(\cdot)$. We then calculate the Euclidean distance between the two vectors derived from sequences A and B:

$$D_{TT}(A,B) = \sqrt{\sum (V_A - V_B)^2}$$

However, such a metric would not be expected to produce a meaningful measure for sequences composed of non-overlapping element types (e.g. ABCABC, and DEFDEF). Therefore we sort vectors V_A and V_B in order of transition probability before comparison. This allows a comparison of transition probability distributions, independent of element type.

186

187 N-gram distribution (NG)

188 Researchers have previously proposed that an important property of animal sequences is the nature of
189 repeating units within the sequence (Cane 1959; Kershenbaum *et al.* 2014; Pruscha & Maurus 1979).

190 A sequence of length L consists of L-n+1 sub-sequences of length n. Thus, the five-element sequence

ABBAC consists of 5-2+1=4 two-element sub-sequences: AB, BB, BA, and AC. For a sequence

192 consisting of C distinct element types, there are a total of C^n distinct *n*-element possible sub-

193 sequences. The vector of sub-sequence frequencies, $P(i \in C^n)$ can be considered a feature vector, and

194 the distance between two strings calculated in a similar way to that shown above:

$$D_{NG}(A,B) = \sqrt{\sum (P_A - P_B)^2}$$

In the following analyses, we chose the n-gram distribution for n = 3, as this provides a good balance between coverage and diversity. For a comparison of different length n-grams in analysing birdsong, see Jin & Kozhevnikov (2011).

198

199 Shannon entropy (SE)

200 Information theory approaches to analysing animal vocal communication have become popular in 201 recent years. One metric that is simple to understand and easy to apply is the Shannon entropy 202 (Shannon et al. 1949), and this has been used in a number of studies to measure the complexity of 203 animal vocal sequences (Da Silva, Piqueira & Vielliard 2000; Doyle et al. 2008; McCowan, Hanser & 204 Doyle 1999; McCowan, Hanser & Doyle 1999; Suzuki, Buck & Tyack 2006). Shannon entropy 205 measures the unpredictability of a sequence, or the lack of uniformity of a sequence, so that a 206 completely predictable sequence (e.g., consisting of the same element repeated over and over) would 207 have an entropy of zero, whereas a completely unpredictable (random) sequence would have an 208 entropy of one. The equation for Shannon entropy *H* is as follows:

$$H = -\sum_{i \in 1 \dots C} P_i \log_C P_i$$

where P_i is the probability of element *i*, drawn from a set of the *C* elements occurring in the union of all sequences.

Our SE metric compares two sequences by taking the ratio of the Shannon entropies of the sequencesA and B:

$$D_{SE}(A, B) = H_A/H_B$$
 where $H_A < H_B$

Although SE is calculated as a single comparison between single measurements on two sequences (in contrast to the TT and NG metrics described above, both of which result in multiple measurements on a single sequence), SE should still be considered a unary metric, because it does not directly measure the distance between two sequences, but rather the difference in a derived metric from each. 217

218 Entropy rate (ER)

Entropy rate has been shown to be a useful metric for measuring vocal sequence complexity (Kershenbaum 2013). Entropy rate is derived from the transition table of a sequence, and can be thought of as a measure of transition table diversity, i.e., the extent to which different transitions between notes are of uniform or non-uniform probability. Given a transition table $T_{i,j}$ as described above, entropy rate ER is defined as:

$$ER = -\sum_{i \in 1...C} \pi_i \sum_{j \in 1...C} T_{i,j} \log T_{i,j}$$

where π_i is the stationary probability of element *i*, i.e., the overall probability of *i* occurring in the sequence; see Kershenbaum (2013) for additional information on metric calculation. As with Shannon entropy, we define a metric D_{ER} for the difference between sequences A and B:

$$D_{ER}(A,B) = ER_A/ER_B$$
 where $ER_A < ER_B$

227

228 Repeat distribution (RD)

229 The repeat number distribution was used in a recent study to compare the similarity between natural 230 and synthetic songs of Bengalese finches, Lonchura striata var. domestica (Jin & Kozhevnikov 231 2011). It is an aggregate measure, calculated on a corpus of sequences. For each set of sequences a histogram is generated showing the probabilities P_n that any element occurred in isolation (n = 1), was 232 repeated twice (n = 2), three times (n = 3), and so on. As with the n-gram distribution, we define a 233 234 metric that measures the difference between two such histograms, generated from sequences A and B, 235 where P_A and P_B are the feature vectors of sequences A and B, comprising the repeat distributions for 236 all the elements:

$$D_{RD}(A,B) = \sqrt{\sum (P_A - P_B)^2}$$

237

238 Mutual information (MI)

Mutual information is an information theory measure that can be applied easily to quantify the similarity of two sequences. MI combines both measures of the inherent complexity in a sequence (via Shannon entropy), and the joint entropy of the sequences, which measures the probability that a particular pair of elements will occur at the same point in two sequences; see Kershenbaum *et al.* (2012) for additional information on metric calculation. MI is defined as follows:

$$MI = H(A) + H(B) - \sum_{i} \sum_{j} p_{i,j} \log p_{i,j}$$

where H(A) is the Shannon entropy of sequence A, H(B) is the Shannon entropy of sequence B, and $p_{i,j}$ is the probability that elements *i* and *j* occur at the same point in sequences A and B. As with Shannon entropy, we define a metric D_{MI} for the difference between sequences A and B:

$$D_{MI} = MI_A / MI_B$$
 where $MI_A < MI_B$

247

248 Lempel-Ziv (LZ)

The Lempel-Ziv complexity (Lempel & Ziv 1976) is an important algorithm used for data 249 compression, as it is a measure of the number of distinct patterns in a sequence. As a metric of 250 251 sequence complexity and an approximation to Kolmogorov complexity (Evans & Barnett 2002), it is 252 potentially a useful indicator of the diversity of an animal vocal sequence. Although it has not been 253 widely used in animal studies, Suzuki, Buck & Tyack (2006) suggested the use of the LZ metric for 254 the analysis of humpback whale song, and Kershenbaum (2013) showed that the LZ metric outperformed Shannon entropy (SE) in quantifying realistic length acoustic sequences. LZ complexity 255 256 was calculated using the Applied Nonlinear Time Series Analysis library for Matlab (Small 2005).

$$LZ = \frac{c \log L}{L \log K}$$

where c is the number of distinct substrings in a sequence of length L, and K is the maximum number of possible distinct substrings.

259 Sequences for analysis

260 Artificial sequences

261 In the first test, we evaluated the utility of each of the similarity metrics by their ability to identify 262 correctly the stochastic process model from which artificial sequences were generated. We generated 263 artificial sequences using three different stochastic processes, often used to model animal vocal 264 sequences (Kershenbaum et al. 2014): the zero-order Markov process (ZOMP), the first-order 265 Markov process (FOMP), and the semi-Markov renewal process (RP). The ZOMP is an independent stochastic process, in which the probability of any particular element occurring at a particular point in 266 267 a sequence is determined solely by the prior probability of that element. In the FOMP, element 268 probabilities are determined by a transition table, where the probability of a particular element 269 depends on the immediately preceding element. The RP has been shown to be a more realistic model 270 of animal vocal sequence production (Kershenbaum et al. 2014) in which the number of repeated 271 elements is drawn from a Poisson distribution, rather than being determined by the diagonal of a 272 transition table. In each case, we examined 10 sequences of 10 elements each, drawn from five 273 possible elements (A-E). We generated 30 sequences, 10 from each of the stochastic processes, 274 ZOMP, FOMP, and RP. The ZOMP was modelled by selecting five random prior probabilities, one 275 for each element type, and renormalising to sum to unity. We then generated the sequences by 276 selecting elements according to these prior probabilities. The FOMP was modelled by generating a 277 random 5 x 5 transition table in a similar way to the ZOMP prior probabilities, so that the rows of the 278 transition matrix summed to unity. A random initial element was chosen for each 10-element 279 sequence, and the remaining nine elements in each sequence were chosen randomly according to the 280 probabilities in the transition table. The RP was modelled in a similar way to the FOMP, except that 281 for each element generated, a random number of repeats were drawn from a Poisson distribution with 282 mean five (to give 95% confidence of \leq 9 repeats). Having generated 30 sequences of 10 elements, we

283 then calculated a 30 x 30 distance matrix for each of the similarity metrics. We then used an Adaptive 284 Resonance Theory (ART) artificial neural network to cluster these 30 points into natural groupings, 285 setting a maximum of 100 possible clusters. ART networks have been used in a number of previous 286 studies to cluster data derived from animal vocalisations (Deecke & Janik 2006; Janik 1999; Quick & 287 Janik 2012). We then calculated the normalised mutual information (NMI) as a metric of goodness of 288 clustering (Zhong & Ghosh 2005), by comparing the composition of the generated clusters H(Y) with the true generation process of each $H(\hat{Y})$. Thus, NMI indicates the proportion of uncertainty predicted 289 290 by the metric. We then repeated this process 100 times using new random transition matrices, generating 3000 sequences in total. 291

In the second test using artificial sequences, we simulated "individuals" by generating 100 random RP 292 293 transition matrices, and from each of them producing a set of 10 sequences of 10 elements each. We 294 used the RP generation process, rather than a Markovian ZOMP or FOMP, as the RP more reliably 295 describes many types of animal vocal sequences (Kershenbaum et al. 2014). Each sequence generated 296 from a single transition matrix would be expected to be more similar to other sequences from the 297 same transition matrix, than sequences generated by a different random transition matrix, therefore we 298 used a similar clustering approach as in the stochastic process analysis above. We calculated the 100 x 299 100 distance matrix for each similarity metric, obtained by comparing the sequences from each of the 300 100 transition matrices, and clustered the results as before, measuring the NMI as an indication of 301 clustering success.

For a final test using artificial sequences, we examined the effect of typical sample sizes (number of sequences) on each of the similarity metrics. Using the sequences generated in the individual simulation above, we varied the number of sequences analysed from one to ten, recalculated the distance matrices and clustering, and measured the NMI.

306

307 Animal sequences

308 We tested the performance of the above metrics using empirical sequences of animal vocalisations, 309 where those sequences are thought to contain information that is known a priori. Very few examples 310 exist where contextual information is objectively known to exist in animal vocal sequences. However, 311 the signature whistles of bottlenose dolphins have been shown to encode individual identity in the 312 sequence of up-down frequency shifts, known as a Parsons code (Kershenbaum, Sayigh & Janik 313 2013). We used a data set consisting of 400 signature whistles, 20 from each of 20 individual 314 dolphins, recorded during capture-release events; see Sayigh et al. (2007) and Kershenbaum, Sayigh 315 & Janik (2013) for additional details. We converted each whistle into a 9-element Parsons code, with seven possible element values ("large drop", "medium drop", "small drop", "no change", "small 316 rise", "medium rise", and "large rise"). We then calculated distance matrices using each of the 317 318 similarity metrics described above, and clustered using an ART network. For the calculation of NMI, 319 we compared the generated clusters to the known clusters of individual identity. As empirical data do 320 not allow the generation of unlimited data sets as with artificial sequences, we estimated confidence 321 intervals for each of the empirical data sets by randomly selecting 80% of the calls for clustering and 322 calculation of NMI, and repeated this process 100 times.

We analysed three further empirical data sets for which contextual information in vocal sequences has been proposed. The first data set used recordings of humpback whales (for details see Garland *et al.* 2012), the second data set used recordings of rock hyraxes (see Kershenbaum *et al.* 2012), and the third set Carolina chickadees (see Freeberg 2012). Previous studies have shown that in the humpback whale, rock hyrax, and Carolina chickadee, song syntax varies according to the geographical origin of the population. For example, not only does chickadee song syntax vary between locations, but there appear to be different functional use of certain sequences in the different populations (Freeberg 2012).

330 The humpback whale data set consisted of 202 songs composed of 20 different element types

331 (themes), recorded from 42 individuals. Humpback whale song is a complex, stereotyped, repetitive,

long, male display that has multiple levels of hierarchy in its organisation (Herman & Tavolga 1980;

Payne & Payne 1985; Payne & McVay 1971). A few sounds (units) are arranged in a stereotyped

334 phrase which is repeated multiple times to make a theme (Payne & McVay 1971). A number of

335 themes, sung in a particular order, are combined to form a song. The order and content of the themes 336 are highly stereotyped, and all males within a population adhere to the same arrangement and content 337 of the song at any given time as the display is constantly changing (Frumhoff 1983; Payne, Tyack & 338 Payne 1983; Payne & Payne 1985). This analysis focused on the theme level in the hierarchical 339 arrangement of humpback whale song. Each string therefore represented the sequence of themes 340 (elements) that comprised a song; e.g., theme 1, theme 2, theme 3, theme 4, theme 5; see Garland et 341 al. (2012) for further information and example sequences. This level within the hierarchy takes into account information on the sequence of units and the repetition of phrases at a higher level, but does 342 not examine these lower levels explicitly. Strings were classified according to their geographical 343 344 location: New Caledonia, Vanuatu, or eastern Australia, and this geographical origin was compared to the clusters generated by the ART network. Humpback whale song is constantly changing, and has 345 346 been shown to undergo complete song revolutions in this region (Garland et al. 2011; Garland et al. 347 2011; Noad et al. 2000). The current analysis incorporates two different song types (lineages) that 348 contain different themes (vocabulary), and are present in these populations at various points over the 349 four years of recording. Therefore, each metric must be robust to the underlying transmission 350 dynamics of this display.

The hyrax data consisted of 1130 song sequences composed of five different element types, recorded from a single individual at each of 18 different locations in Israel. The Carolina chickadee data consisted of 1184 sequences of calls, recorded from 60 sites in the states of Tennessee and Indiana, USA. Links to these data sets are available in the supplemental information.

355

356 RESULTS

357 Artificial sequences

For sequences generated by different stochastic processes, the entropy rate (ER) metric provided the best clustering, with a NMI value of 0.518 ± 0.005 (standard error) (Figure 2a), while the binary

360 Levenshtein distance (LD) metric gave a NMI of 0.476 ± 0.006 . A post-hoc Tukey test following

ANOVA showed significant differences between the NMI scores of these two metrics. All other
 metrics produced significantly lower NMI values.

363 Results from clustering sequences of simulated "individuals" (sequences generated by stochastic

364 processes with similar parameters), indicated that NG produced the highest NMI score 0.751 ± 0.001 ,

while the LD, RD, and TT metrics all produced high but slightly lower NMI scores (greater than 0.7;

366 Figure 2b), with no significant differences among the NMI values of these three metrics.

367 Both the LD and NG metrics that performed well on the above clustering tasks were also robust to

368 sample size (Figure 3). Most other metrics were also relatively unaffected by sample size. However,

the RD performed poorly at smaller sample sizes (≤ 4), and the MI declined with increasing corpus

371

370

372 Animal sequences

size (> 2).

373 When clustering to reconstruct the individual identity from bottlenose dolphin signature whistles, the 374 Levenshtein distance (LD) performed significantly better than all other tested metrics, with an NMI of 375 0.661 ± 0.001 (Figure 4a). The n-gram distribution (NG) also performed well, with an NMI of $0.63p \pm$ 0.001. Clustering of the humpback whale song data to indicate population (geographic) origin, 376 377 showed the LD again performed significantly better than all other tested metrics (NMI of $0.491 \pm$ 378 0.005; Figure 4b). The NG provided the second best, although significantly poorer, metric (NMI of 379 0.367 ± 0.005). All metrics performed poorly in clustering the geographical origin of hyrax songs; 380 however, the LD metric was again significantly better than all others tested (NMI 0.1684 \pm 0.001, 381 compared to the next best NMI of 0.130 ± 0.001 for TT; Figure 4c). Clustering of the chickadee data 382 to distinguish between birds recorded in Tennessee and those recorded in Indiana, showed the LD 383 performed significantly better than all other metrics (NMI of 0.450 ± 0.001 ; Figure 4d), followed by 384 NG (NMI 0.369 ± 0.001).

387 We analysed the performance of eight different techniques from two broad approaches, to investigate 388 the utility of each approach in the comparison of animal sequences. The unary and binary metrics 389 performed similarly well in the artificial sequence tests, with the entropy rate (ER) metric slightly 390 better than the Levenshtein distance binary metric (LD), in distinguishing between data generated by 391 different stochastic processes, and n-gram (NG) slightly better in distinguishing simulated individuals. 392 However, the LD metric performed significantly better than all other tested metrics when presented 393 with empirical animal sequences. This result emphasises that caution should be used when using 394 artificially generated sequences based on simple stochastic models to simulate animal vocal 395 sequences. Recent work has shown that assumptions of simple models for animal vocal production are 396 likely to be inaccurate (Kershenbaum et al. 2014), and similar conclusions have been indicated for 397 cetacean song (Miksis-Olds et al. 2008). The difference between metric performance on artificial and 398 on empirical data is striking. Little is known of the cognitive mechanisms by which animals encode 399 and decode information in vocalisations (Thornton, Clayton & Grodzinski 2012); researchers must 400 rely on isolated examples where information content is known a priori to draw conclusions about 401 which analytical techniques are best suited for vocal sequence data. Our results clearly show that the 402 LD metric outperforms other metrics on empirical data, despite performing less effectively on 403 simulated data. This indicates that the sequential order of the sequences varied across 404 location/individual while the level of complexity is similar. The Levenshtein distance was the metric 405 of choice for clustering dolphin signature whistles into individuals, humpback whale song into populations, hyrax songs into geographical region, and chickadee calls into state of origin. Analysis of 406 407 the sensitivity of the different metrics to sample size showed that most of the metrics that performed 408 well across the data sets (LD, NG, LZ), were also robust to sample size.

409 Results from the current paper in combination with previous work (Eriksen *et al.* 2005; Garland *et al.*

410 2012; Garland et al. 2013; Helweg et al. 1998; Tougaard & Eriksen 2006), highlight the success of

411 the Levenshtein distance (LD) metric in the analysis of sequence content and comparison of

412 humpback whale song. A large body of work has previously shown that song differences among

413 humpback whale populations can indicate geographic origin of a singer (e.g., Garland et al. 2015; 414 Helweg et al. 1998; Payne & Guinee 1983). Despite dynamic song transmission in the South Pacific 415 region, fine-scale song differences allow the identification of population origin (Garland et al. 2011; 416 Garland et al. 2012; Garland et al. 2013; Garland et al. 2015). The current paper examined the theme 417 sequences (i.e., a set of phrases under a single label) as part of the largest analysis to date of sequence 418 comparison algorithms for humpback whale song (Garland et al. 2013), which indicated the LD out 419 performed all other tested metrics. We suggest when *comparing* song sequences, the LD metric 420 should be employed preferentially, while if the complexity or information content of each song is the 421 focus of study, the researcher should employ other techniques such as entropy.

422 Previous studies of sequence comparison in hyrax song (Kershenbaum et al. 2012) have shown 423 geographical variation in sequence structure using the LD metric, as these findings were supported by 424 application of an unrelated (unary) metric, mutual information (MI). In the current study, MI 425 performed very poorly on both simulated and empirical data, although MI performance was somewhat 426 better on the hyrax data than on the other data sets. This implies that the aspect of the sequences that 427 is measured by MI does not vary in correlation with geographic location or individual. While not all 428 studies can compare large numbers of analytical algorithms, this emphasises the utility of comparing 429 at least two different techniques when assessing novel algorithms, to ensure that results are robust 430 under a range of analytical approaches.

431 Despite all tested metrics performing poorly in the assessment of geographic origin in hyrax song, the 432 LD metric was significantly better than all others. In previous work, (Kershenbaum et al. 2012) 433 measured the correlation between sequence similarity and the distance between populations, rather 434 than classification success, and the latter suggests that distinct dialects are not present in the hyrax. 435 Rather, small but significant differences are present between all pairs of populations, depending on 436 geographic isolation. In contrast, humpback whales, chickadees, and bottlenose dolphins show strong 437 discrimination between in-group and out-group sequences, indicating that the differences between the 438 vocal sequences of different individuals or populations are much more marked. This may indicate an 439 adaptive role to distinctive vocalisations in dolphins and whales, such as individual identification

440 (Janik & Slater 1998; Janik, Sayigh & Wells 2006; Quick & Janik 2012), while in chickadees 441 adaptive, developmental, and phylogenetic explanations for regional dialects have been suggested 442 (Freeberg 2012). Humpback whale song is hypothesised to contain information about the reproductive 443 fitness and population origin of the signaller (Helweg et al. 1992; Helweg et al. 1992; Payne & 444 Guinee 1983). Hyrax song complexity is not thought to contain contextual information beyond male fitness (Demartsev et al. 2014; Koren & Geffen 2009), although this assumption is currently untested. 445 446 In contrast, dolphin signature whistles are known to be individually distinctive whistles that can be 447 identified by the unique pattern of frequency modulations (Janik, Sayigh & Wells 2006). The 448 characterisation of signature whistles based on a 7-element Parsons code in a previous study 449 (Kershenbaum, Sayigh & Janik 2013) allows individual identification of the whistler. The LD 450 significantly outperformed all other models in clustering to reconstruct not only the individual identity 451 from signature whistles, but the geographic origin for humpback whale song, chickadee calls, and 452 hyrax song, highlighting the importance of evaluating different metrics with *a priori* information. 453 One likely explanation for the higher performance of the LD metric is that it alone among the metrics 454 analysed uses a direct comparison of the vocal sequences between samples, thereby using more 455 information about the sequences than the other metrics. The LD metric by design can solely be 456 employed to *compare* two strings and it excels at this task; it does not provide an understanding of the 457 information content within each string, or the sequence structure. By necessity this means that LD 458 also compares the vocabularies of a pair of sequences, and therefore two sequences that are based on 459 the same set of sequence elements are likely to have a lower LD value than two sequences that are 460 composed of different elements, but have similar sequence structure. Regional differences in the 461 vocabulary (e.g., humpback song themes) provide important information on the connectivity of 462 populations at a broad-scale despite an overall similarity in song structure (hierarchical arrangement). 463 To establish the influence of overlapping vocabulary is beyond the scope of this paper (although two 464 of the three humpback populations switched between two vocabularies – song types – over the course of this study), but we present as supplemental information (Figure S1) the element distributions of the 465 466 different data sets, which in most cases were quite consistent.

467 Sample sizes can be constrained in the study of wild animals and particularly in marine mammal 468 studies. Samples may be collected infrequently and with a patchy distribution due to the challenging 469 conditions presented in collecting such data. Understanding how a metric reacts to a small sample size 470 is invaluable in metric choice. The robust nature of the LD and NG to smaller sample sizes and their 471 high performance in the comparison task makes them appealing for analysis. The data presented here 472 indicated that LD and NG performed well with a sample size of three or less, while TT and RD should 473 not be considered as a metric for analysis until a sample size of four or more is available.

474 Here, we have presented a robust understanding of which metric should be preferentially employed in 475 studies involving the comparison of individual- or group-specific vocalisations, such as signature 476 whistles. The success in identifying individual/geographic variations in vocal sequences has 477 implications for assessing population structure, song transmission, and dialect similarity, particularly 478 for populations where rapid song changes occur. For example, the analysis of humpback whale song 479 presented here was able to identify population origin despite rapid song dynamics (Garland et al. 480 2011; Garland et al. 2012; Garland et al. 2013). We suggest that the LD can be applied to any level 481 within a complex display, but suggest future studies strive for the lowest level sequence within the 482 hierarchy (i.e., sequence of units or phrases), to increase the amount of information directly compared 483 and thus encapsulated by the sequence.

484 The LD method provides a metric to compare sequence content and organisation (and thus songs) 485 within and among multiple individuals, populations, years, and locations. In particular, transmission 486 of humpback whale song is largely cultural, and the level and rate of change remains unparalleled in 487 any other non-human animal as complete population-wide changes are replicated in multiple 488 populations at a vast geographic scale (Garland et al. 2011). Thus, fundamental questions in animal 489 culture, vocal learning, and cultural evolution can be explored using humpback whale song as a 490 model, and with the help of the LD metric. Further, the evolution of complex vocal labels (i.e., 491 signature whistles) and the underlying cognitive abilities required for such evolution, are extremely important in understanding the evolution of vocal complexity (Janik 2014). Robust metrics that 492 493 capture the information encoded in the sequences with the highest fidelity are thus required to address

these far-reaching evolutionary questions. We suggest the LD should be utilised in such comparisonstudies in preference to Markov and information theory based models.

496

497 Conclusions

498 The Levenshtein distance (LD; binary metric) significantly outperformed all other tested metrics in 499 our comparative analysis of animal acoustic sequences. It provides a direct measure of pairwise 500 differences among sequences, instead of a comparison of aggregate similarity. N-grams (Markov 501 chains) were the second most successful metric; the underlying issue that the tested species' 502 vocalisations may be governed by non-Markovian dynamics and the consistent success of the LD 503 metric, suggests n-grams should always be a second choice. Given the inherent interest in the origins 504 of human language and the evolution of signalling complexity, robust and reliable metrics that can 505 capture the content and arrangement of the signal are essential to address these fundamental questions 506 in animal communication and cultural evolution.

507

508 Acknowledgements

509 We thank Melinda Rekdahl, Todd Freeberg and his graduate students, Amiyaal Ilany, Elizabeth 510 Hobson, and Jessica Crance for providing comments of on a previous version of this manuscript. We 511 thank Mike Noad, Melinda Rekdahl, and Claire Garrigue for assistance with humpback whale song 512 collection and initial categorisation of the song, Vincent Janik and Laela Sayigh for assistance with 513 signature whistle collection, Todd Freeberg with chickadee recordings, and Eli Geffen and Amiyaal 514 Ilany for assistance with hyrax song collection and analysis. E.C.G is supported by a Newton 515 International Fellowship. Part of this work was conducted while E.C.G. was supported by a National 516 Research Council (National Academy of Sciences) Postdoctoral Fellowship at the National Marine 517 Mammal Laboratory, AFSC, NMFS, NOAA. The findings and conclusions in this paper are those of 518 the authors and do not necessarily represent the views of the National Marine Fisheries Service. We

519 would also like to thank Randall Wells and the Sarasota Dolphin Research Program for the

- 520 opportunity to record the Sarasota dolphins, where data were collected under a series of National
- 521 Marine Fisheries Service Scientific Research Permits issued to Randall Wells. A.K. is supported by

522 the Herchel Smith Postdoctoral Fellowship Fund. Part of this work was conducted while A.K. was a

523 Postdoctoral Fellow at the National Institute for Mathematical and Biological Synthesis, an Institute

524 sponsored by the National Science Foundation through NSF Award #DBI-1300426, with additional

- 525 support from The University of Tennessee, Knoxville.
- 526

527 DATA ACCESSIBILITY

528 Hyrax and chickadee data:

http://rspb.royalsocietypublishing.org/highwire/filestream/47311/field_highwire_adjunct_files/1/rspb
20141370supp2.xls

Authors do not own all of the data and have not been granted permission to archive it. Access isavailable as follows:

533 Dolphin data: Dolphin whistle data are held by Dr Sayigh (Woods Hole Oceanographic Institution;

534 lsayigh@whoi.edu) and Dr Janik (University of St Andrews; vj@st-andrews.ac.uk). Please contact the

535 PI's directly for access to their recordings.

536 Humpback whale data: Humpback whale song data are held by Dr. Noad (University of Queensland,

537 Australia; mnoad@uq.edu.au) and Dr. Garrigue (Operation Cetaces, New Caledonia;

538 op.cetaces@lagoon.nc). Please contact the PI's directly for access to their recordings.

540 REFERENCES

```
541
```

- 543 Ashby, F.G. & Perrin, N.A. (1988). Toward a unified theory of similarity and recognition.
- 544 *Psychological review*, **95**, 124.
- Berwick, R.C., Okanoya, K., Beckers, G.J.L. & Bolhuis, J.J. (2011). Songs to syntax: the linguistics
 of birdsong. *Trends in Cognitive Sciences*, 15, 113-121.
- 547 Bradley, D.W. & Bradly, R. (1983). Application of sequence comparison to the study of bird songs.
- 548 *Time warps, string edits, and macromolecules: the theory and practice of sequence comparison/edited*
- 549 by David Sankoff and Joseph B.Krustal,.
- 550 Briefer, E., Osiejuk, T.S., Rybak, F. & Aubin, T. (2010). Are bird song complexity and song sharing
- shaped by habitat structure? An information theory and statistical approach. *Journal of Theoretical Biology*, 262, 151-164.
- 553 Buck, J.R. & Tyack, P.L. (1993). A quantitative measure of similarity for *Tursiops truncatus*
- signature whistles. *The Journal of the Acoustical Society of America*, **94**, 2497-2506.
- Cane, V.R. (1959). Behaviour sequences as semi-Markov chains. *Journal of the Royal Statistical Society.Series B (Methodological)*, 21, 36-58.
- 557 Catchpole, C.K. & Slater, P.J.B. (2003). *Bird song: biological themes and variations*. Cambridge
 558 Univ Press, Cambridge.
- Da Silva, M.L., Piqueira, J.R.C. & Vielliard, J.M.E. (2000). Using Shannon entropy on measuring the
 individual variability in the rufous-bellied thrush *Turdus rufiventris* vocal communication. *Journal of Theoretical Biology*, 207, 57-64.

- 562 Darolová, A., Krištofik, J., Hoi, H. & Wink, M. (2012). Song complexity in male marsh warblers:
 563 does it reflect male quality? *Journal of Ornithology*, **153**, 431-439.
- Deecke, V.B. & Janik, V.M. (2006). Automated categorization of bioacoustic signals: avoiding
 perceptual pitfalls. *The Journal of the Acoustical Society of America*, **119**, 645-653.
- 566 Demartsev, V., Kershenbaum, A., Ilany, A., Barocas, A., Ziv, E.B., Koren, L. et al. (2014). Male
- 567 hyraxes increase song complexity and duration in the presence of alert individuals. *Behavioral*568 *Ecology*.
- 569 Doyle, L.R., McCowan, B., Hanser, S.F., Chyba, C., Bucci, T. & Blue, J.E. (2008). Applicability of
- 570 information theory to the quantification of responses to anthropogenic noise by southeast Alaskan
- 571 humpback whales. *Entropy*, **10**, 33-46.
- 572 Eriksen, N., Miller, L.A., Tougaard, J. & Helweg, D.A. (2005). Cultural change in the songs of
- 573 humpback whales (*Megaptera novaeangliae*) from Tonga. *Behaviour*, **142**, 305-328.
- 574 Evans, S.C. & Barnett, B. (2002). *Network security through conservation of complexity*. Proceedings
 575 of IEEE MILCOM 2002.
- 576 Freeberg, T.M. (2012). Geographic variation in note composition and use of chick-a-dee calls of
 577 Carolina chickadees (*Poecile carolinensis*). *Ethology*, **118**, 555-565.
- 578 Freeberg, T.M. & Lucas, J.R. (2012). Information theoretical approaches to chick-a-dee calls of
- 579 Carolina chickadees (*Poecile carolinensis*). Journal of Comparative Psychology, **126**, 68-81.
- 580 Frumhoff, P. (1983). Aberrant songs of humpback whales (Megaptera novaeangliae): Clues to the
- 581 structure of humpback songs. *Communication and Behavior of Whales* (ed R. Payne), pp. 81-127.
- 582 Westview Press, Boulder, Colorado.

- 583 Garland, E.C., Goldizen, A.W., Lilley, M.S., Rekdahl, M.L., Garrigue, C., Constantine, R. et al.
- 584 (2015). Population structure of humpback whales in the western and central South Pacific Ocean as
- 585 determined by vocal exchange among populations. *Conservation Biology*,.
- 586 Garland, E.C., Goldizen, A.W., Rekdahl, M.L., Constantine, R., Garrigue, C., Hauser, N.D. et al.
- 587 (2011). Dynamic horizontal cultural transmission of humpback whale song at the ocean basin scale.
- 588 *Current Biology*, **21**, 687-691.
- 589 Garland, E.C., Lilley, M.S., Goldizen, A.W., Rekdahl, M.L., Garrigue, C. & Noad, M.J. (2012).
- 590 Improved versions of the Levenshtein distance method for comparing sequence information in
- animals' vocalisations: tests using humpback whale song. *Behaviour*, **149**, 1413-1441.
- 592 Garland, E.C., Noad, M.J., Goldizen, A.W., Lilley, M.S., Rekdahl, M.L., Garrigue, C. et al. (2013).
- 593 Quantifying humpback whale song sequences to understand the dynamics of song exchange at the
- 594 ocean basin scale. *The Journal of the Acoustical Society of America*, **133**, 560-569.
- 595 Helweg, D.A., Cato, D.H., Jenkins, P.F., Garrigue, C. & McCauley, R.D. (1998). Geographic
- variation in South Pacific humpback whale songs. *Behaviour*, **135**, 1-27.
- 597 Helweg, D.A., Frankel, A.S., Mobley Jr, J.R. & Herman, L.M. (1992). Humpback whale song: Our
- 598 current understanding. *Marine Mammal Sensory Systems*, pp. 459-483. Springer.
- 599 Herman, L.M. & Tavolga, W.N. (1980). The communication systems of cetaceans. *Cetacean*
- 600 *behavior: Mechanisms and functions,*, 149-209.
- Holveck, M., de Castro, Ana Catarina Vieira, Lachlan, R.F., ten Cate, C. & Riebel, K. (2008).
- 602 Accuracy of song syntax learning and singing consistency signal early condition in zebra finches.
- 603 Behavioral Ecology, **19**, 1267-1281.
- Janik, V.M. (1999). Pitfalls in the categorization of behaviour: a comparison of dolphin whistle
- 605 classification methods. *Animal Behaviour*, **57**, 133-143.

- Janik, V.M., Sayigh, L. & Wells, R. (2006). Signature whistle shape conveys identity information to
- 607 bottlenose dolphins. *Proceedings of the National Academy of Sciences*, **103**, 8293-8297.
- 508 Janik, V.M. & Slater, P.J.B. (1998). Context-specific use suggests that bottlenose dolphin signature

609 whistles are cohesion calls. *Animal Behaviour*, **56**, 829-838.

- Janik, V.M. (2014). Cetacean vocal learning and communication. *Current Opinion in Neurobiology*, **28**, 60-65.
- Jin, D.Z. & Kozhevnikov, A.A. (2011). A compact statistical model of the song syntax in Bengalese
 finch. *PLoS Computational Biology*, 7, e1001108.
- Kershenbaum, A. (2013). Entropy rate as a measure of animal vocal complexity. *Bioacoustics*, 23,
 195-208.
- Kershenbaum, A., Ilany, A., Blaustein, L. & Geffen, E. (2012). Syntactic structure and geographical
 dialects in the songs of male rock hyraxes. *Proceedings of the Royal Society B: Biological Sciences*,
 279, 2974-2981.
- Kershenbaum, A., Sayigh, L.S. & Janik, V.M. (2013). The encoding of individual identity in dolphin
 signature whistles: how much information is needed? *PLoS One*, **8**, e77671.
- 621 Kershenbaum, A., Blumstein, D.T., Roch, M.A., Akçay, Ç, Backus, G., Bee, M.A. et al. (2014).
- 622 Acoustic sequences in non-human animals: a tutorial review and prospectus. *Biological Reviews*,.
- 623 Kershenbaum, A., Bowles, A.E., Freeberg, T.M., Jin, D.Z., Lameira, A.R. & Bohn, K. (2014). Animal
- 624 vocal sequences: not the Markov chains we thought they were. *Proceedings of the Royal Society B:*
- 625 *Biological Sciences*, **281**, 20141370.
- 626 Koren, L. & Geffen, E. (2009). Complex call in male rock hyrax (Procavia capensis): a multi-
- 627 information distributing channel. *Behavioral Ecology and Sociobiology*, **63**, 581-590.

- 628 Krull, C., Ranjard, L., Landers, T., Ismar, S., Matthews, J. & Hauber, M. (2012). Analyses of sex and
- 629 individual differences in vocalizations of Australasian gannets using a dynamic time warping
- algorithm. The Journal of the Acoustical Society of America, 132, 1189.
- 631 Lempel, A. & Ziv, J. (1976). On the complexity of finite sequences. *IEEE Transactions on*
- 632 *Information Theory*, **22**, 75-81.
- Levenshtein, V.I. (1966). *Binary codes capable of correcting deletions, insertions and reversals. Soviet Physics Doklady*, **10**, 707-710
- 635 Likic, V. (2008). The Needleman-Wunsch algorithm for sequence alignment,
- 636 http://www.ludwig.edu.au/course/lectures2005/likic.pdf edn. Lecture given at the 7th Melbourne
- 637 Bioinformatics Course, University of Melbourne.
- 638 McCowan, B., Hanser, S.F. & Doyle, L.R. (1999). Quantitative tools for comparing animal
- 639 communication systems: information theory applied to bottlenose dolphin whistle repertoires. *Animal*640 *Behaviour*, **57**, 409-419.
- 641 Miksis-Olds, J.L., Buck, J.R., Noad, M.J., Cato, D.H. & Stokes, M.D. (2008). Information theory
- analysis of Australian humpback whale song. *The Journal of the Acoustical Society of America*, **124**,
 2385-2393.
- Navarro, G. (2001). A guided tour to approximate string matching. *ACM computing surveys (CSUR)*,
 33, 31-88.
- Needleman, S.B. & Wunsch, C.D. (1970). A general method applicable to the search for similarities
 in the amino acid sequence of two proteins. *Journal of Molecular Biology*, 48, 443-453.
- 648 Neubauer, R.L. (1999). Super-normal length song preferences of female zebra finches (*Taeniopygia*
- 649 *guttata*) and a theory of the evolution of bird song. *Evolutionary Ecology*, **13**, 365-380.

- Noad, M.J., Cato, D.H., Bryden, M., Jenner, M. & Jenner, K.C.S. (2000). Cultural revolution in whale
 songs. *Nature*, 408, 537-537.
- Payne, K. & Payne, R. (1985). Large scale changes over 19 years in songs of humpback whales in
 Bermuda. *Zeitschrift für Tierpsychologie*, 68, 89-114.
- 654 Payne, K., Tyack, P. & Payne, R. (1983). Progressive changes in the songs of humpback whales
- 655 (Megaptera novaeangliae): a detailed analysis of two seasons in Hawaii. Communication and
- 656 Behavior of Whales, (ed. R. Payne), Westview Press, Boulder, pp 9-57.
- 657 Payne, R.S. & McVay, S. (1971). Songs of humpback whales. Science, 173, 585-597.
- 658 Payne, R. & Guinee, L.N. (1983). Humpback whale (Megaptera novaeangliae) songs as an indicator
- of "stocks". *Communication and Behavior of Whales*, (ed. R. Payne), Westview Press, Boulder, pp,
 333-358.
- 661 Pfaff, J.A., Zanette, L., MacDougall-Shackleton, S.A. & MacDougall-Shackleton, E.A. (2007). Song
- repertoire size varies with HVC volume and is indicative of male quality in song sparrows (Melospiza
- 663 melodia). Proceedings of the Royal Society B: Biological Sciences, 274, 2035-2040.
- Pruscha, H. & Maurus, M. (1979). Analysis of the temporal structure of primate communication. *Behaviour*, **69**, 118-134.
- Quick, N.J. & Janik, V.M. (2012). Bottlenose dolphins exchange signature whistles when meeting at
 sea. *Proceedings of the Royal Society B: Biological Sciences*, 279, 2539-2545.
- 668 Ranjard, L. (2010). Computational biology of bird song evolution. PhD, University of Auckland.
- 669 Ranjard, L., Anderson, M.G., Rayner, M.J., Payne, R.B., McLean, I., Briskie, J.V. et al. (2010).
- 670 Bioacoustic distances between the begging calls of brood parasites and their host species: a
- 671 comparison of metrics and techniques. *Behavioral Ecology and Sociobiology*, **64**, 1915-1926.

- 672 Ranjard, L. & Ross, H.A. (2008). Unsupervised bird song syllable classification using evolving neural
- networks. *The Journal of the Acoustical Society of America*, **123**, 4358-4368.
- Reis, D.d.C., Golgher, P.B., Silva, A. & Laender, A. (2004). *Automatic web news extraction using tree edit distance*. ACM.
- 676 Sayigh, L.S., Esch, H.C., Wells, R.S. & Janik, V.M. (2007). Facts about signature whistles of
- 677 bottlenose dolphins, *Tursiops truncatus*. *Animal Behaviour*, **74**, 1631-1642.
- 678 Sayigh, L.S., Tyack, P.L., Wells, R.S., Solow, A.R., Scott, M.D. & Irvine, A.B. (1999). Individual
- recognition in wild bottlenose dolphins: a field test using playback experiments. *Animal Behaviour*,
 57, 41-50.
- Searcy, W.A. & Andersson, M. (1986). Sexual selection and the evolution of song. *Annual Review of Ecology and Systematics*, 17, 507-533.
- 683 Shannon, C.E., Weaver, W., Blahut, R.E. & Hajek, B. (1949). The Mathematical Theory of
- 684 Communication. University of Illinois Press, Urbana.
- 685 Small, M. (2005). Applied nonlinear time series analysis: applications in physics, physiology and
- 686 *finance*. World Scientific Publishing Company Incorporated, Singapore.
- 687 Suzuki, R., Buck, J.R. & Tyack, P.L. (2006). Information entropy of humpback whale songs. *The*
- *Journal of the Acoustical Society of America*, **119**, 1849-1866.
- 689 Thornton, A., Clayton, N.S. & Grodzinski, U. (2012). Animal minds: from computation to evolution.
- 690 Philosophical Transactions of the Royal Society B: Biological Sciences, 367, 2670-2676.
- 691 Tougaard, J. & Eriksen, N. (2006). Analysing differences among animal songs quantitatively by
- means of the Levenshtein distance measure. *Behaviour*, **143**, 239-252.

- 693 Young, F. & Hamer, R. (1994). Theory and applications of multidimensional scaling. *Hillsdale, NJ:*
- 694 Eribaum Associates,.
- 695 Zhong, S. & Ghosh, J. (2005). Generative model-based document clustering: a comparative study.
- 696 *Knowledge and Information Systems*, **8**, 374-384.

698 FIGURES

699

WQSQSQS xxxxxx QSQSQS (a)	TCQQQQSCQCSCSC xxxx xx xxxxx TTTTCQQQQWWWQQ (c)
WQSQSQS x QSQSQS (b)	WQQQQQQQQQQQQQQQ x WSQQQQQQQQQQQQQ (d)

Figure 1. Examples of string alignment and edit distance. (a) Two unaligned strings with a LD of 7.

(b) After aligning the strings to minimise the difference, LD = 1. (c) Two hyrax bouts which are

highly different, LD = 11. (d) Two bouts which are very similar, LD = 1. Reproduced from

704 (Kershenbaum *et al.* 2012).

705



Figure 2. Results of the normalised mutual information (NMI) scores for each metric using a)

709 synthetic processes, and b) synthetic individuals. Metric labels: Levenshtein distance (LD), Repeat

710 distribution (RD), Transition table (TT), Shannon entropy (SE), Lempel-Ziv (LZ), N-gram (NG),

711 Mutual information (MI), and entropy rate (ER). A-F indicate post-hoc Tukey groupings.

712



Figure 3. Results of the effect of sample (corpus) size on the NMI scores (± standard error) for each
similarity metric. Metric labels are the same as Figure 2.



Figure 4. Results of the NMI (normalised mutual information) scores for each metric using a)
bottlenose dolphin signature whistles, b) humpback whale songs, c) rock hyrax songs, and d) Carolina
chickadee calls. Metric labels are the same as Figure 2. A-F indicate post-hoc Tukey groupings.