



Tilburg University

Confirmatory factor analysis and differential relationships of the two subdomains of negative symptoms in chronically ill psychotic patients

Stiekema, A.P.M.; Liemburg, E.J.; van der Meer, L.; Castelein, S.; Stewart, R.; van Weeghel, J.; Aleman, A.; Bruggeman, R.

Published in: PLoS ONE

DOI: 10.1371/journal.pone.0149785

Publication date: 2016

Document Version Publisher's PDF, also known as Version of record

Link to publication in Tilburg University Research Portal

Citation for published version (APA):

Stiekema, A. P. M., Liemburg, E. J., van der Meer, L., Castelein, S., Stewart, R., van Weeghel, J., Aleman, A., & Bruggeman, R. (2016). Confirmatory factor analysis and differential relationships of the two subdomains of negative symptoms in chronically ill psychotic patients. *PLoS ONE*, *11*(2), [e0149785]. https://doi.org/10.1371/journal.pone.0149785

General rights

Copyright and moral rights for the publications made accessible in the public portal are retained by the authors and/or other copyright owners and it is a condition of accessing publications that users recognise and abide by the legal requirements associated with these rights.

- Users may download and print one copy of any publication from the public portal for the purpose of private study or research.
- You may not further distribute the material or use it for any profit-making activity or commercial gain
 You may freely distribute the URL identifying the publication in the public portal

If you believe that this document breaches copyright please contact us providing details, and we will remove access to the work immediately and investigate your claim.



Citation: Stiekema APM, Liemburg EJ, van der Meer L, Castelein S, Stewart R, van Weeghel J, et al. (2016) Confirmatory Factor Analysis and Differential Relationships of the Two Subdomains of Negative Symptoms in Chronically III Psychotic Patients. PLoS ONE 11(2): e0149785. doi:10.1371/journal. pone.0149785

Editor: Ruud van Winkel, Katholieke Universiteit Leuven, BELGIUM

Received: August 18, 2015

Accepted: February 4, 2016

Published: February 19, 2016

Copyright: © 2016 Stiekema et al. This is an open access article distributed under the terms of the <u>Creative Commons Attribution License</u>, which permits unrestricted use, distribution, and reproduction in any medium, provided the original author and source are credited.

Data Availability Statement: All relevant data are within the Supporting Information files.

Funding: AA was supported by a VICI grant from the Netherlands Organisation for Scientific Research (N. W.O. nr 453-11-004).

Competing Interests: The authors have declared that no competing interests exist.

RESEARCH ARTICLE

Confirmatory Factor Analysis and Differential Relationships of the Two Subdomains of Negative Symptoms in Chronically Ill Psychotic Patients

Annemarie P. M. Stiekema^{1,2}*, Edith J. Liemburg^{2,3,4}, Lisette van der Meer^{1,2,3}, Stynke Castelein^{2,4}, Roy Stewart⁵, Jaap van Weeghel^{6,7,8}, André Aleman³, Richard Bruggeman^{2,9}

1 Department of Rehabilitation, Lentis Center for Mental Health Care, Zuidlaren, the Netherlands, 2 Rob Giel Research Center, University of Groningen, University Medical Center Groningen, Groningen, the Netherlands, 3 Department of Neuroscience, University of Groningen, University Medical Center Groningen, Groningen, the Netherlands, 4 Research Department, Lentis Center for Mental Health Care, Groningen, the Netherlands, 5 Department of Health Sciences, Community and Occupational Medicine, University of Groningen, University Medical Center Groningen, Groningen, the Netherlands, 6 Parnassia Group, Dijk en Duin Mental Health Center, Castricum, the Netherlands, 7 Tilburg University, Tilburg School of Social and Behavioral Sciences, Tranzo Scientific center for Care and Welfare, Tilburg, the Netherlands, 8 Phrenos, Center of Expertise on severe mental illness, Utrecht, the Netherlands, 9 University Center of Psychiatry, University of Groningen, University Medical Center Groningen, Groningen, the Netherlands

* a.stiekema@lentis.nl

Abstract

Research suggests a two factor structure for negative symptoms in patients with psychotic disorders: social amotivation (SA) and expressive deficits (ED). Applying this two-factor structure in clinical settings may provide valuable information with regard to outcomes and to target treatments. We aimed to investigate 1) whether the factor structure is also supported in chronically ill patients with a psychotic disorder and 2) what the relationship is between these factors and functioning (overall functioning and living situation), depressive symptoms and quality of life. 1157 Patients with a psychotic disorder and a duration of illness of 5 years or more were included in the analysis (data selected from the Pharmacotherapy Monitoring Outcome Survey; PHAMOUS). A confirmatory factor analysis was performed using items of the Positive and Negative Syndrome Scale that were previously identified to reflect negative symptoms (N1-4, N6, G5, G7, G13, G16). Subsequently, regression analysis was performed on outcomes. The results confirmed the distinction between SA (N2, N4, G16) and ED (N1, N3, N6, G5, G7, G13) in chronically ill patients. Both factors were related to worse overall functioning as measured with the Health of the Nation Outcome Scales, ED was uniquely associated with residential living status. Higher scores for SA were associated with more depressive symptoms and worse quality of life. Thus, SA is most strongly related to level of social-emotional functioning, while ED are more related to living situation and thereby are indicative of level of everyday functioning. This subdivision may be useful for research purposes and be a valuable additional tool in clinical practice and treatment development.

Introduction

Negative symptoms, such as flattened affect, social withdrawal, apathy and avolition, are core symptoms of psychotic disorders, most notably schizophrenia. At least half of the patients with schizophrenia suffers from negative symptoms [1], which are often already present in the prodromal phase [2] and are relatively stable across the course of illness [3]. Negative symptoms have an invalidating impact on patients' functioning [4–6] and are associated with lower quality of life [7]. Despite the increased focus on negative symptoms as a subject of research, there is still a paucity of (psychosocial) interventions effective in reducing them. Many patients are left with negative symptoms after their positive symptoms have been partially or completely managed by antipsychotic medication [8]. The lack of substantial improvement in everyday functioning after antipsychotic treatment may therefore be impeded by enduring negative symptoms [9].

An accumulating body of research suggests that negative symptoms are multidimensional [10]. Factor analytic studies across different instruments consistently cause two factors to emerge, namely social amotivation (SA) and expressive deficits (ED) [11–14]. The SA subdomain encompasses social and emotional withdrawal and speaks to involvement with the environment [15]. It refects a reduction of interest in social interactions and life events, and a reduction of self-initiated or maintained behaviors with regard to social events. SA has been linked to deficits in anticipatory pleasure (i.e. failure to signal the salience of positive events), thereby losing the drive to engage in (social) situations and activities [15–17]. Thus, SA can be interpreted as a 'loss of interest' [13]. The ED subdomain involves directly observable components such as diminished facial expression, poverty of speech and blunted affect [10,15]. ED is a reduction of verbal and non-verbal emotional responsiveness, reflected by a reduction of communicative expression. ED has been associated with impaired neurocognition [11,15,18] and may reflect a 'loss of initiative' [13].

Both factors seem to affect functional and psychosocial outcomes differently [10,19-21]. This has important implications, because many (treatment) studies use total negative symptoms scores, which could average out relationships that are mainly driven by one of the subdomains. That is, when subjects demonstrate different scores on each subdomain (high on SA and low on ED or vice versa), their total of negative symptoms may be similar, while their relationship with outcomes could be different as this may be driven by one factor. Therefore, a distinction in subdomains, and more importantly the understanding of possible differential correlates of these factors, could be of importance for clinical diagnosis, therapeutic decisionmaking and research on treatment development [10,11]. Literature suggests that SA is most strongly associated with functional outcomes such as employment, number of hospitalizations, instrumental role performance and family functioning [10,14,22,23] and that males may score higher on SA than ED [14]. However, the role of ED is less clear. Components of ED such as blunted or flat affect have been associated with poor social functioning and quality of life as well [24,25], but ED shows weaker associations with outcomes than SA [14] or has no additional predictive value after controlling for SA [16]. Therefore, SA is often seen as the key contributor to the relationship between negative symptoms and functional outcomes [20]. The majority of studies investigating the correlates of both domains have focused on functional outcomes and less on other aspects, such as depressive symptoms, which are common in psychotic disorder [26], and quality life. Investigating whether the subdomains differentially relate to quality of life, can guide treatment strategies to more specific targets. For depressive symptoms, a differential relationship of the subdomains could clarify the inconsistencies with regard to the association between global negative symptoms and depressive symptoms in the literature.

However, correlational analysis in one study showed no relationship of either factor with depressive symptoms, but did show an association between SA and quality of life [13].

The two factors have been mostly established in samples of patients with recent onset psychosis [10,11,13], but one study established the factor structure in patients with chronic psychosis and a longer duration of illness [22]. As a consequence, little is known about this factor structure in patients with a psychotic disorder with a longer duration of illness. Considering the paucity of studies investigating the subdomains in chronic populations, replication of the factor structure in this population is needed. And, if the factor structure is replicated, the relationship between these factors and functional outcomes and psychosocial well-being should be examined. Therefore, we aimed to investigate 1) whether the factor structure of negative symptoms can be replicated in chronically ill patients with a psychotic disorder and 2) the relationship between these factors and functioning (overall functioning and living situation), depressive symptoms, and quality of life.

Methods

Participants

Data were selected from the Pharmacotherapy Monitoring Outcome Survey (PHAMOUS). PHAMOUS is an annual screening of mental and physical health of patients using antipsychotics and receiving mental health care in the North of the Netherlands. We included all patients between 2011 and 2013, diagnosed with a psychotic disorder, with a duration of illness of more than 5 years and of whom the Positive and Negative Syndrome Scale (PANSS) items N1-N4, N6, G7, G13 and G16 were available (items previously identified [13]). When multiple screenings were available of the same patient, the most recent record was selected unless an older record was more complete. Data were collected in accordance with the latest version of the Declaration of Helsinki. Data were collected for diagnostic purposes, no interventions outside standard care were performed. The procedures were in accordance with local and international rules, as confirmed by the local ethical committee of the University Medical Center of Groningen, who stated that use of anonymized data from the PHAMOUS protocol for research purposes does not fall under the scope of the Medical Research Involving Human Subjects Act and therefore does not need to undergo a prior review by the medical ethical committee.

Assessment measures

The interviews and clinician-rated scales used in this study were assessed and rated by a trained research nurse, each patient was rated by one nurse.

Functional outcome. Functional outcome was measured with the Health of the Nation Outcome Scales (HoNOS) [27]. The items of this clinician-rated instrument were scored on a five point scale ranging from 'no problem' to 'severe to very severe problem'. The HoNOS consist of 4 subscales: behavioral problems, impairment, symptomatic and social problems. The HoNOS has shown moderately high internal consistency and moderate interrater reliability [27].

Furthermore, living situation (living in the community versus residential living) was used as a second measure of functional outcome. Patients who were living on their own, with family, friends or other housemates were characterized as 'living in the community', whereas patients who were living in sheltered or clinical care facilities fell into the 'residential living' category.

Symptom assessment. Symptomatology was measured with the Positive And Negative Syndrome Scale (PANSS), a commonly used semi-structured interview including three subscales, namely positive symptoms, negative symptoms and general pathology [28], on a seven

point scale ranging from ranging from 'absent' to 'severe'. The PANSS has shown high internal consistency and good construct validity [28].

Depressive symptoms were measured with the Calgary Depression Scale for Schizophrenia (CDSS) [29], a structured interview with nine items on a four point scale, ranging from 'absent' to 'severe'. Depression as measured with the CDSS can predict outcomes differentially from negative symptoms.

Quality of Life. Quality of life was measured using the Manchester Short Assessment of Quality of Life (MANSA). The MANSA is a self-report questionnaire and addresses patients' satisfaction within several psychosocial domains, including satisfaction with life as a whole, job (or sheltered employment), training/education, or unemployment/retirement), financial situation, number and quality of friendships, leisure activities, accommodation, personal safety, people that the patient lives with (or living alone), sex life, relationship with family, physical health, and mental health [30]. The twelve items that are rated on a seven point scale ('could not be worse' to 'could not be better') were used for analysis (the other four items are dichotomous (yes/no) and were excluded for methodological reasons). The MANSA has good construct validity and internal consistency [30].

Statistical analysis

Confirmatory factor analysis. Based on previous work [13], the presupposed two factor structure of SA and ED was evaluated through confirmatory factor analysis (CFA) with the computer program Mplus version 7 [31]. SA (factor 1) and ED (factor 2) were entered as latent variables of the nine PANSS items. Because of violation of the multivariate normality assumption, the items were entered according to an ordinal scale using a polychoric correlation matrix. Furthermore, a robust weighted least squares estimator (WLSMV) was used, as recommended by the literature [31-34]. To measure the goodness-of-fit (GOF) of the factor structure, the following indices and cut-off criteria were used: the Comparative Fit index (CFI > .95), the Goodness-of-Fit index (GFI > .95), the Tucker-Lewis index (TLI > .95), the Root Mean Square Error of Approximation (RMSEA < 0.06), and the Weighted Root Mean Square of Residuals (WRMR < 0.90) [31]. Significantly correlated residuals were introduced into the model.

Regression analysis. Hierarchical multiple regression models were used to investigate the associations between SA and ED scores on the one hand (independent variables) and HoNOS, CDSS and MANSA total scores and HoNOS subscale scores on the other hand (dependent variables), while controlling for positive symptoms (total score of PANSS positive symptoms subscale), age, gender and antipsychotic medication (expressed in chlorpromazine equivalents [35]). SA (total score of PANSS items N2, N4, G16) and ED (total score of PANSS items N1, N3, N6, G5, G7 and G13) were entered in the first block, positive symptoms, age, gender, and antipsychotic medication were entered in the second block. A logistic regression model was used to examine the relationship between the negative symptom factors (independent variables) and living situation (dependent variable; 0 = non-residential, 1 = residential) controlling for the same confounders. All statistical analyses were performed with IBM SPSS Statistics version 22 (IBM Corp, Armonk, NY).

Results

Patient characteristics

In total, 1157 patients fulfilled the inclusion criteria. Baseline demographic and clinical characteristics are presented in <u>Table 1</u>. Patients were mostly male with a mean age of 45 years and they had been ill for 18,5 years on average (time since first psychotic episode). The majority of the patients were diagnosed with schizophrenia.

Table 1. Baseline clinical and demographic characteristics (N = 1157).

| | Total sample | | Non-residentia | I | Residential | | |
|---|--|----------|---|--------|---|----------|---------------------|
| | Mean ± SD, N (%) or median [25 th ; 75 th percentile] (n = 1157) | Range | Mean ± SD, N (%) or median [25 th ; 75 th percentile] (n = 693 ^a) | Range | Mean ± SD, N (%) or median [25 th ; 75 th percentile] (n = 390 ^a) | Range | <i>p</i> - value |
| Demographics | | | | | | | |
| Age | 44.3 ± 10.8 | 19–72 | 43.6 ± 10.5 | 19–71 | 46.0 ± 11.2 | 21–72 | < .001 |
| Duration of illness | 18.5 ± 9.7 | 5–55 | 17.21 ± 8.9 | 5–55 | 21.1 ± 10.7 | 5–54 | < .001 |
| Gender, % Male | 775 (67.0) | | 435 (62.8) | | 295 (75.6) | | < .001 |
| Living situation | | | | | | | |
| Independent without partner | 474 (41.0) | | 474 (68.4) | | - | | |
| Independent with partner | 122 (10.5) | | 122 (17.6) | | - | | |
| With family/others | 97 (8.4) | | 97 (14.0) | | - | | |
| Sheltered living/ social pension | 247 (21.3) | | - | | 247 (63.3) | | |
| Long-stay clinical facilities | 143 (12.4) | | - | | 143 (36.7) | | |
| Other/unknown | 74 (6.4) | | - | | - | | |
| Diagnosis | | | | | | | |
| Schizophrenia | 848 (73.3) | | 466 (67.2) | | 322 (82.6) | | < .001 |
| Schizoaffective disorder | 180 (15.6) | | 127 (18.3) | | 44 (11.3) | | .002 |
| Psychotic disorder NOS | 98 (8.5) | | 78 (11.3) | | 16 (4.1) | | < .001 |
| Schizophreniform disorder | 18 (1.6) | | 15 (2.2) | | 3 (0.8) | | .135 |
| Delusional disorder | 13 (1.1) | | 7 (1.0) | | 5 (1.3) | | .765 |
| Psychiatric comorbidity ^b | 340 (29.4) | | 170 (24.5) | | 145 (62.8) | | < .001 |
| Substance abuse | 137 (11.8) | | 71 (10.2) | | 57 (14.6) | | .081 |
| Developmental disorder | 24 (2.1) | | 13 (1.9) | | 10 (2.6) | | .537 |
| Anxiety disorder | 25 (2.2) | | 17 (2.5) | | 6 (1.5) | | .411 |
| Somatoform disorder | 1 (0.1) | | - | | 1 (0.3) | | .360 |
| Personality disorder | 155 (13.4) | | 78 (11.3) | | 64 (16.4) | | .025 |
| Intellectual disability | 37 (3.2) | | 15 (2.2) | | 19 (4.9) | | .011 |
| Medication | | | | | | | |
| Antipsychotic medication | | | | | | | |
| None | 76 (6.6) | | 55 (8.0) | | 20 (5.1) | | .082 |
| Clozapine | 335 (29.0) | | 140 (20.2) | | 177 (45.4) | | < .001 |
| Risperidone | 194 (16.8) | | 76 (11.0) | | 37 (9.5) | | .542 |
| Olanzapine | 207 (17.9) | | 120 (17.3) | | 68 (17.4) | | .934 |
| Aripiprazol | 157 (13.6) | | 105 (15.2) | | 41 (10.5) | | .036 |
| Quetiapine | 107 (9.2) | | 73 (10.5) | | 27 (6.9) | | .039 |
| Haloperidol | 75 (6.5) | | 32 (4.6) | | 17 (4.4) | | .880 |
| Other ^c | 308 (26,6) | | 195 (28.1) | | 166 (42.6) | | < .001 |
| Nr of antipsychotics | 1.2 ± 0.6 | 0–4 | 1,1 ± 0.5 | 0–4 | 1.4 ± 0.7 | 0–4 | < .001 |
| CPZ equivalent (mg/d) ^d | 350 [150; 600] | 0–3037.5 | 300 [115; 525] | 0–1800 | 480 [225; 750] | 0–3037.5 | < .001 |

(Continued)

Table 1. (Continued)

PLOS ONE

| | Total sample | | Non-residential | | Residential | | |
|-------------------------------|--|--------|---|--------|---|--------|---------------------|
| | Mean ± SD, N (%) or median [25 th ; 75 th percentile] (n = 1157) | Range | Mean ± SD, N (%) or median [25 th ; 75 th percentile] (n = 693 ^a) | Range | Mean ± SD, N (%) or median [25 th ; 75 th percentile] (n = 390 ^a) | Range | <i>p</i> - value |
| Nr of concomitant medications | 2.4 ± 2.7 | 0–16 | 1.8 ± 2.1 | 0–15 | 3.7 ± 3.1 | 0–16 | < .001 |
| Outcomes | | | | | | | |
| PANSS total | 51.9 ± 15.7 | 30–132 | 48.2 ± 13.7 | 30–100 | 57.3 ± 16.7 | 30–132 | < .001 |
| PANSS positive | 12.1 ± 4.8 | 7–38 | 11.4 ± 4.4 | 7–31 | 13.0 ± 5.3 | 7–38 | < .001 |
| PANSS negative | 13.8 ± 6.0 | 7–42 | 12.4 ± 5.1 | 7–32 | 16.1 ± 6.5 | 7–42 | < .001 |
| PANSS general | 25.9 ± 7.9 | 16–69 | 24.4 ± 6.8 | 16–50 | 28.1 ± 8.6 | 16–69 | < .001 |
| PANSS social amotivation | 5 [3;8] | 3–20 | 5 [3; 7] | 3–17 | 6 [4; 9] | 3–16 | < .001 |
| PANSS expressive deficits | 9 [7;13] | 6–34 | 8 [6;12] | 6–27 | 11 [8; 15] | 6–34 | < .001 |
| HoNOS total | 9.5 ± 5.7 | 0–37 | 8.1 ± 5.1 | 0–26 | 11.7 ± 5.1 | 0–37 | < .001 |
| CDSS total | 2.5 ± 3.1 | 0–17 | 2.5 ± 3.2 | 0–17 | 2.3 ± 2.7 | 0–12 | .384 |
| MANSA total | 59.2 ± 12.1 | 14–84 | 59.6 ± 11.4 | 26–84 | 59.0 ± 13.2 | 14–84 | .457 |

Abbreviations: CPZ: chlorpromazine; PANSS: Positive and Negative Syndrome Scale; HoNOS: Health of the Nation Outcome Scales (subtotal of items 4, 7, 8, 9 and 10); CDSS: Calgary Depression Scale for Schizophrenia; MANSA: Manchester Short Assessment of Quality of Life.

a Of 74 patients the living situation was unknown

b Nr of patients with one or more comorbid psychiatric disorder, most comorbid disorders were personality disorders (19.0%) and substance abuse disorders (16.8%).

c Other medication included: zuclopentixol (22.7%), paliperidon (13.7%), flupentixol (9.3), pimozide (7.6%), miscellaneous (46,7%).

d Chlorpromazine equivalents of antipsychotic dosage were calculated based on Gardner and colleagues [35].

doi:10.1371/journal.pone.0149785.t001

Factor analysis

In <u>Table 2</u> we present the results of standardized factor loadings with significant correlated residuals. The goodness of fit indices for the CFA are good according the criteria given in the literature [<u>36</u>]. The RMSEA is 0.06 (CI 90%: 0.05–0.07), the WRMR is 0.86 [<u>31</u>], CFI is 0.99 and TLI is 0.98. All factor loadings are above 0.5.

Hierarchical regression

Data distributions were examined for linearity and normality. CDSS scores and HoNOS behavioral problems scores were positively skewed. The distribution was improved after applying square root transformations. Furthermore, there was no evidence for multicollinearity in the regression models. Hierarchical regression analyses were performed to investigate the relationship between both SA and ED and the outcome measures.

The analyses revealed that higher SA was significantly related to worse overall functioning (HoNOS total score), more depressive symptoms (CDSS) and worse quality of life (MANSA) (results of the final models are shown in <u>Table 3</u>). For the HoNOS subscales, higher SA was associated with symptomatic problems and social problems (see <u>S1 Table</u>). The observed associations remained significant after controlling for positive symptoms, age, gender and antipsychotic medication.

With regard to ED, higher scores were related to significantly worse overall functioning (HoNOS total score) and depressive symptoms. A positive relationship between ED and quality



Table 2. Results of confirmatory factor analysis: univariate proportions of the items and factor loadings of items N1-N4, N6, G5, G7, G13 and G16 of the PANSS (N = 1157).

| | Univariate proportions of the items | | | | | | | |
|--------------------------------|-------------------------------------|------------|------------|------------|------------|------------|------------|----------------|
| PANSS item* | Category 1 | Category 2 | Category 3 | Category 4 | Category 5 | Category 6 | Category 7 | Factor loading |
| Factor 1 (social amotivation) | | | | | | | | |
| N2 Emotional withdrawal | 0.408 | 0.274 | 0.167 | 0.114 | 0.027 | 0.01 | 0.001 | 0.938 |
| N4 Passive/apathetic | 0.367 | 0.243 | 0.208 | 0.088 | 0.072 | 0.02 | 0.001 | 0.872 |
| G16 Active social avoidance | 0.593 | 0.207 | 0.135 | 0.035 | 0.024 | 0.004 | 0.002 | 0.674 |
| Factor 2 (expressive deficits) | | | | | | | | |
| N1 Flat affect | 0.361 | 0.22 | 0.199 | 0.118 | 0.095 | 0.003 | 0.005 | 0.821 |
| N3 Poor rapport | 0.58 | 0.171 | 0.183 | 0.041 | 0.016 | 0.006 | 0.003 | 0.847 |
| N6 Lack of spontaneity | 0.596 | 0.145 | 0.161 | 0.064 | 0.022 | 0.01 | 0.002 | 0.793 |
| G5 Mannerisms and posturing | 0.707 | 0.144 | 0.124 | 0.014 | 0.004 | 0.004 | 0.003 | 0.504 |
| G7 Motor retardation | 0.649 | 0.152 | 0.147 | 0.046 | 0.004 | 0.002 | 0 | 0.651 |
| G13 Avolition | 0.709 | 0.128 | 0.114 | 0.04 | 0.009 | 0.001 | 0 | 0.585 |

Abbreviations: PANSS: Positive and Negative Syndrome Scale

* significant correlated residuals are (N1 with N2,N6,G7,G13,G16); (N2 with N4); (N6 with N3,G7); (G5 with G7,G16).

doi:10.1371/journal.pone.0149785.t002

of life was found, indicating that higher ED was associated with higher quality of life. Higher ED scores were associated with higher scores on the HoNOS impairment subscale (cognitive and psychical or disability problems), the behavioral problems subscale and the social problems subscale. Logistic regression analyses revealed that ED was associated with residential living status (living in sheltered or clinical care facilities), which remained significant after controlling for confounders (Table 4).

Discussion

In this study we established that the negative symptoms factor structure consisting of social amotivation (SA) and expressive deficits (ED) also holds in a chronic population with psychotic disorders. It thereby extends previous reports demonstrating two separate factors of negative symptoms in patients in the early phase of their psychotic illness [13] and factor analytic

Table 3. Hierarchical multiple regression models for overall functioning, quality of life and depressive symptoms.

| | | HoNOS (N = 715) ^a | | | CDSS (N = 588) ^b | | | MANSA (N = 777) ^c | | | | | |
|------|---------------------|------------------------------|-------|--------|-----------------------------|------|--------|------------------------------|---------------------|------|--------|--------|---------------------|
| Step | Variable added | β | t | р | Adj. R ² | β | t | Р | Adj. R ² | β | t | Р | Adj. R ² |
| 1 | Social amotivation | .173 | 3.925 | < .001 | .158 | .227 | 4.419 | < .001 | .066 | 184 | -3.908 | < .001 | .032 |
| | Expressive deficits | .127 | 2.914 | .004 | | .037 | .733 | .464 | | .096 | 2.047 | .041 | |
| 2 | PANSS positive | .341 | 9.911 | < .001 | .272 | .150 | 3.624 | < .001 | .102 | 200 | -5.416 | < .001 | .086 |
| | CPZ eq | .065 | 1.936 | .053 | | 048 | -1.177 | .240 | | .031 | .866 | .387 | |
| | Age | .029 | .906 | .365 | | 079 | -1.964 | .050 | | .137 | 3.886 | < .001 | |
| | Gender | .023 | .699 | .485 | | 130 | -3.236 | .001 | | 066 | -1.855 | .064 | |

Abbreviations: PANSS: Positive and Negative Syndrome Scale; CPZ: chlorpromazine; HoNOS: Health of the Nation Outcome Scales (subtotal of items 4, 7, 8, 9 and 10); CDSS: Calgary Depression Scale for Schizophrenia; MANSA: Manchester Short Assessment of Quality of Life.

^a Overall adjusted model R^2 = .278, F(6.708) = 45.357, p < .001

^b Overall adjusted model R^2 = .111, F(6.576) = 12.035, p < .001

^c Overall adjusted model R^2 = .093, F(6.751) = 12.858, p < .001.

doi:10.1371/journal.pone.0149785.t003



| Step ^a | Variables | В | OR | 95% C.I. for OR |
|-------------------|---------------------|------|-------|-----------------|
| 1 | Social amotivation | 053 | .948 | .893–1.007 |
| | Expressive deficits | .118 | 1.126 | 1.080–1.174* |
| | PANSS positive | .049 | 1.050 | 1.018–1.083* |
| | CPZ equivalent | .001 | 1.001 | 1.001-1.002* |
| | Age | .019 | 1.019 | 1.006–1.033* |
| | Gender | 531 | .588 | .428–.808* |

Table 4. Logistic regression model for living situation: admission to sheltered or clinical facility (N = 1018).

Abbreviations: PANSS: Positive and Negative Syndrome Scale; CPZ: chlorpromazine; OR = Odds ratio; C.I. = confidence interval

^a reference category: non-residential living

* p < .001.

doi:10.1371/journal.pone.0149785.t004

studies using the Scale for Assessment of Negative Symptoms (SANS) [37] or the Schedule for Deficit Syndrome (SDS) [38] (see for an overview [39]). Furthermore, the SA factor was associated with more depressive symptoms and worse quality of life, while the ED factor was most importantly related to residential living status. These relationships were not affected by positive symptoms, age, gender or antipsychotic dosage.

The replication of the dimensional structure of negative symptoms provides good support for the subdomains across the course of illness, which was not yet firmly established in chronic samples. Furthermore, the dimensional structure of the PANSS is an important addition to the factor analytic studies using the SANS and SDS, because the PANSS is widely used in clinical trials as well as in clinical practice and recognized as an appropriate tool for assessing negative symptoms [40]. As such, subdomains SA and ED can be used to assess differences in treatment response and eventually guide clinical practice in choosing a treatment strategy. There are a few notable differences in the PANSS factor analytic results compared to other instruments that should not go without mention. The main difference between the SANS studies and our results, is that PANSS avolition item (G13) loads on ED, while the SANS avolition items load on SA. PANSS ratings for avolition are merely based on observed behavior and could therefore be rated as a disturbance in willful initiation of behavior or facial expression, whereas the SANS avolition items may be rated more of a social motivational deficit [13]. Furthermore, the avolition item of the PANSS (G13) and the mannerisms and posing item (G5), which also loads on the ED factor in our study, have previously been reported as part of the disorganized factor of the PANSS. These items were nevertheless included in our analysis, because previous work showed that the factor loadings warranted inclusion in ED and that removal of these items did not improve the model fit [13]. The current factor loadings of G5 and G13 were comparable to this previous study.

The value of the distinction in subdomains is its relationship with functional and clinical outcomes [41]. Most importantly we found that higher ED was related to residential living (i.e. living in a sheltered or clinical care facility), while SA was related to more depressive symptoms and lower quality of life. Residential patients generally have a more severe course of illness and the poorest outcomes. This suggests that ED is more strongly associated with a more severe course of illness and poorer functional outcomes, contradicting evidence for SA as the key predictor of functioning [14,22]. A possible explanation for the relationship with residential living is that patients with ED seem more 'ill'. That is, family, friends or health care workers may more often interpret SA as for example demoralization, indifference or laziness; extremes of 'normal' behavior. ED on the other hand, is more difficult to place within the frames of normal behavior and can seem more deviant and therefore lead to seeking help, for example in the

form of admission to a residential care facility. Indeed, ED has been linked to neurocognitive deficits before [11,13,15,18] and was related to the impairment subscale of the HoNOS in this study (measuring cognition and disability) confirming higher disability and higher need for intensive (residential) care. Interestingly, these patients do not report lower quality of life than less disabled patients (given that higher ED was related to better quality of life; see also De Heer-Wunderink and colleagues [42]). The experience of a good 'person-environment fit' by the residential group may in part explain these findings.

Previous findings on the relationship between negative symptoms and depression have been inconsistent (i.e. some studies have reported an association [43–45], while others have not [46–48]). Our findings indeed suggest a relationship between depressive symptoms and negative symptoms. However, this relationship seems to be limited to SA. This suggests that the inconsistency in the relationship between depression and negative symptoms may (in part) be explained by the subdomain structure of negative symptoms. That is, when subjects demonstrate different scores on each subdomain (high on SA and low on ED or vice versa), their total of negative symptoms may be similar, while their relationship with depression is different as this is driven by SA.

Quality of life was significantly associated with SA. This is in line with previous work [13]. In addition, we found a relationship between SA and the subscale social problems of the HoNOS (S1 Table). Since quality of life has also been associated with social functioning [49,50], this leads us to suggest that the relationship between SA and quality of life is of an indirect nature. That is, SA causes problems with social functioning, which in turn has an effect upon the subjective quality of life. A mediation analysis demonstrated that the relationship between SA and quality of life was indeed influenced by social problems (partial mediation) (S1 Fig). However, since we did not explicitly state any hypothesis with regard to this relationship, this interpretation should be treated with caution. Higher ED was associated with better quality of life. The direction of this association is surprising and the strength of the association increased upon including positive symptoms in the regression model. This suggests that other factors influence this relationship, which makes this result difficult to interpret with the current data and deserves further investigation.

Taken together, our results seem to indicate that both factors differentially relate to distinct aspects of functioning. SA seems to be most strongly related to social-emotional aspects of functioning, reflected in associations with depressive and psychological symptoms (HoNOS subscale) and quality of life. ED on the other hand, seems to be more strongly related to aspects of everyday functioning and behavioral problems, as reflected by its associations with living situation, cognitive and disability problems (HoNOS impairment subscale) and the behavioral problems subscale of the HoNOS.

Keeping in mind that replication is needed in both chronic and other samples, some clinical implications of these findings could be cautiously suggested. Considering that SA has been linked to deficits in anticipatory pleasure [15–17], the individually oriented Cognitive Behavioral Therapy (CBT) model constructed by Staring and colleagues [51] that specifically aims at reducing negative symptoms by targeting dysfunctional beliefs including experiencing pleasure, could be particularly suited to target SA. For ED, the loss of initiative factor, personalized rehabilitation approaches aimed at examining each patient's wishes and strengths, and accepting and working around the impairments, may be most suitable. These could include the rehabilitation approach by Anthony and colleagues [52], or compensatory strategies such as Cognitive Adaptation Training [53] or Cognitive Compensatory Training [54]. Some pharmacological treatments have shown to selectively impact SA and ED. For example, add-on mirtazapine or selegiline showed a selective effect on SA, while add-on galantamine showed specific effects on ED, and amisulpride affected both subdomains [21]. However, further research into the effects of drugs on the specific subdomains is needed.

Future research should focus on the distinction between social-emotional functioning and everyday activities to further disentangle the differential clinical correlates of both factors and to elucidate the inconsistencies in the literature. Longitudinal studies should investigate whether early interventions are useful in preventing the development of subdomain related functional problems. Intervention studies that take the subdivision of negative symptoms into account are still rare. Further, it would be useful to investigate whether the subdomains retrieved from the PANSS and SANS are interchangeable (which one would expect based on the high correlation between the PANSS negative subscale and the SANS [55]), in order to examine whether inconsistencies with regard to the functional correlates can be explained by the scale that is used. Efforts have been made in developing scales which reliably measure both subdomains of negative symptoms [56, 57].

Strengths of this study are its large sample size and the fact that the data were derived from a Routine Outcome Monitoring database for which patients were not selected for research purposes and therefore are representative of the real-world population. Another strength of our study is that we did not only focus on functional outcomes but on depression and subjective quality of life as well. A limitation is the relatively low negative symptom scores (on average a rating of 'minimal' on each item). The lack of inclusion criteria with regard to negative symptom severity may have biased our results. Future research with patients with more profound negative symptoms is necessary to further investigate whether the relationships that we found are also applicable to those with severe negative symptoms. Furthermore, we were not able to explore proposed underlying mechanisms of SA and ED because cognitive measures and measures of anticipatory pleasure were not part of the standard PHAMOUS screening. Different neurobiological correlates have been proposed for lack of interest versus lack of initiative [58], concepts related to the present two negative factors, which deserve further investigation.

In conclusion, this study replicates the multidimensionality of negative symptoms and showed unique correlates of these two factors. Our results suggest that SA is predominantly related to social-emotional aspects of functioning, and that ED is particularly related to aspects of everyday functioning. Better understanding of the negative symptom subdomains is of value in developing treatments targeting negative symptoms in schizophrenia, which still represent an unmet need in this patient population.

Supporting Information

S1 Data. Selected variables phamous study.

(SAV)

S1 Fig. Standardized regression coefficients, standard errors and p-values for the relationship between social amotivation and quality of life (Fig a) as mediated by social problems (Fig b). Analysis were conducted in Mplus and corrected for age, gender and chlorpromazine equivalents.

(DOCX)

S1 Table. Hierarchical multiple regression models for HoNOS subscales (results of final models are shown). (DOCX)

Acknowledgments

We would like to thank dr. Ellen Visser for assistance with use of the database. Data collection was made possible by the Rob Giel Research Center. The authors thank patients and staff of the Mental Health Organizations: GGZ Friesland, Lentis, GGZ Drenthe and the University Center of Psychiatry of the University Medical Center Groningen.

Author Contributions

Conceived and designed the experiments: APMS EJL LM RB. Analyzed the data: RS APMS. Wrote the paper: APMS EJL LM SC RS JW AA RB.

References

- Bobes J, Arango C, Garcia-Garcia M, Rejas J. Prevalence of negative symptoms in outpatients with schizophrenia spectrum disorders treated with antipsychotics in routine clinical practice: findings from the CLAMORS study. J Clin Psychiatry. 2010; 71: 280–6. doi: <u>10.4088/JCP.08m04250yel</u> PMID: <u>19895779</u>
- 2. Harvey PD, Koren D, Reichenberg A, Bowie C. Negative symptoms and cognitive deficits: What is the nature of their relationship? Schizophr Bull. 2006; 32: 250–258. PMID: <u>16221995</u>
- Ventura J, Subotnik KL, Gitlin MJ, Gretchen-Doorly D, Ered A, Villa KF, et al. Negative symptoms and functioning during the first year after a recent onset of schizophrenia and 8 years later. Schizophr Res. Elsevier B.V.; 2015; 161: 407–13. doi: 10.1016/j.schres.2014.10.043 PMID: 25499044
- Hunter R, Barry S. Negative symptoms and psychosocial functioning in schizophrenia: neglected but important targets for treatment. Eur Psychiatry. 2012; 27: 432–6. doi: <u>10.1016/j.eurpsy.2011.02.015</u> PMID: <u>21602034</u>
- Rabinowitz J, Levine SZ, Garibaldi G, Bugarski-Kirola D, Berardo CG, Kapur S. Negative symptoms have greater impact on functioning than positive symptoms in schizophrenia: analysis of CATIE data. Schizophr Res. 2012; 137: 147–50. doi: 10.1016/j.schres.2012.01.015 PMID: 22316568
- Ventura J, Helleman G, Thames AD, Koellner V, Nuechterlein KH. Symptoms as mediators of the relationship between neurocognition and functional outcome in schizophrenia: a meta-analysis. Schizophr Res. 2009; 113: 189–199. doi: 10.1016/j.schres.2009.03.035 PMID: 19628375
- Fitzgerald PB, De Castella RA, Filia K, Collins J, Brewer K, Williams CL, et al. A longitudinal study of patient- and observer-rated quality of life in schizophrenia. Psychiatry Res. 2003; 119: 55–62. PMID: 12860360
- Millan MJ, Fone K, Steckler T, Horan W. Negative symptoms of schizophrenia: Clinical characteristics, pathophysiological substrates, experimental models and prospects for improved treatment. Eur Neuropsychopharmacol. Elsevier; 2014; 24: 645–692. doi: <u>10.1016/j.euroneuro.2014.03.008</u> PMID: <u>24820238</u>
- Robinson DG, Woerner MG, McMeniman M, Mendelowitz A, Bilder RM. Symptomatic and functional recovery from a first episode of schizophrenia or schizoaffective disorder. Am J Psychiatry. 2004; 161: 473–479. PMID: <u>14992973</u>
- Messinger JW, Trémeau F, Antonius D, Mendelsohn E, Prudent V, Stanford AD, et al. Avolition and expressive deficits capture negative symptom phenomenology: implications for DSM-5 and schizophrenia research. Clin Psychol Rev. 2011; 31: 161–8. doi: 10.1016/j.cpr.2010.09.002 PMID: 20889248
- 11. Blanchard JJ, Cohen AS. The structure of negative symptoms within schizophrenia: implications for assessment. Schizophr Bull. 2006; 32: 238–45. PMID: <u>16254064</u>
- Kirkpatrick B, Fischer B a. Subdomains within the negative symptoms of schizophrenia: commentary. Schizophr Bull. 2006; 32: 246–9. PMID: <u>16492798</u>
- Liemburg E, Castelein S, Stewart R, van der Gaag M, Aleman A, Knegtering H. Two subdomains of negative symptoms in psychotic disorders: established and confirmed in two large cohorts. J Psychiatr Res. 2013; 47: 718–25. doi: <u>10.1016/j.jpsychires.2013.01.024</u> PMID: <u>23472837</u>
- Strauss GP, Horan W, Kirkpatrick B, Fischer B a, Keller WR, Miski P, et al. Deconstructing negative symptoms of schizophrenia: avolition-apathy and diminished expression clusters predict clinical presentation and functional outcome. J Psychiatr Res. 2013; 47: 783–90. doi: <u>10.1016/j.jpsychires.2013</u>. <u>01.015</u> PMID: <u>23453820</u>
- Foussias G, Remington G. Negative symptoms in schizophrenia: avolition and Occam's razor. Schizophr Bull. 2010; 36: 359–69. doi: <u>10.1093/schbul/sbn094</u> PMID: <u>18644851</u>
- Foussias G, Mann S, Zakzanis KK, van Reekum R, Agid O, Remington G. Prediction of longitudinal functional outcomes in schizophrenia: the impact of baseline motivational deficits. Schizophr Res. Elsevier B.V.; 2011; 132: 24–7. doi: 10.1016/j.schres.2011.06.026 PMID: 21771567
- Gard DE, Kring AM, Gard GM, Horan W, Green MF. Anhedonia in schizophrenia: distinctions between anticipatory and consummatory pleasure. Schizophr Res. 2007; 93: 253–260. PMID: <u>17490858</u>

- Keefe RS, Harvey PD, Lenzenweger MF, Davidson M, Apter SH, Schmeidler J, et al. Empirical assessment of the factorial structure of clinical symptoms in schizophrenia: Negative symptoms. Psychiatry Res. 1992; 44: 153–165. PMID: <u>1480680</u>
- Malaspina D, Walsh-messinger J, Gaebel W, Morris L, Gorun A, Prudent V, et al. Negative symptoms, past and present: a historical perspective and moving to DSM-5. Eur Neuropsychopharmacol. Elsevier; 2014; 24: 710–24. doi: <u>10.1016/j.euroneuro.2013.10.018</u> PMID: <u>24314851</u>
- Foussias G, Siddiqui I, Fervaha G, Agid O, Remington G. Dissecting negative symptoms in schizophrenia: Opportunities for translation into new treatments. J Psychopharmacol. 2014; 1–11.
- Azorin J-M, Belzeaux R, Adida M, Hospital SM, Marguerite B Sainte. Negative Symptoms in Schizophrenia: Where We have been and Where We are Heading. CNS Neurosci Ther. 2014; 20: 801–8.
- Fervaha G, Foussias G, Agid O, Remington G. Motivational and neurocognitive deficits are central to the prediction of longitudinal functional outcome in schizophrenia. Acta Psychiatr Scand. 2014; 1–10.
- Fervaha G, Agid O, Takeuchi H, Foussias G, Remington G. Effect of antipsychotic medication on overall life satisfaction among individuals with chronic schizophrenia: Findings from the NIMH CATIE study. Eur Neuropsychopharmacol. Elsevier; 2014; 24: 1078–1085. doi: <u>10.1016/j.euroneuro.2014.03.001</u> PMID: 24726579
- Gur R, Kohler CG, Ragland JD, Siegel SJ, Lesko K, Bilker WB, et al. Flat affect in schizophrenia: Relation to emotion processing and neurocognitive measures. Schizophr Bull. 2006; 32: 279–287. PMID: 16452608
- Faerden A, Friis S, Agartz I, Barrett EA, Nesvåg R, Finset A, et al. Apathy and functioning in first-episode psychosis. Psychiatr Serv. 2009; 60: 1495–1503. doi: <u>10.1176/appi.ps.60.11.1495</u> PMID: <u>19880468</u>
- Siris SG. Depression in schizophrenia: perspective in the era of "Atypical" antipsychotic agents. Am J Psychiatry. 2000; 157: 1379–1389. PMID: <u>10964850</u>
- 27. Wing JK, Beevor AS, Curtis RH, Park SBG, Hadden S, Burns A. Health of the nation outcome scales (HoNOS): Research and development. British Journal of Psychiatry. 1998. pp. 11–18.
- Kay SR, Fiszbein A, Opler L a. The positive and negative syndrome scale (PANSS) for schizophrenia. Schizophr Bull. 1987; 13: 261–76. PMID: <u>3616518</u>
- Addington D, Addington J, Schissel B. A depression rating scale for schizophrenics. Schizophr Res. 1990; 3: 247–51. PMID: <u>2278986</u>
- Priebe S, Huxley P, Knight S, Evans S. Application and results of the Manchester Short Assessment of Quality of Life (MANSA). Int J Soc Psychiatry. 1999; 45: 7–12. PMID: <u>10443245</u>
- Muthén LK, Muthén BO. Mplus User's Guide. Seventh Edition. (1998–2012). 7th ed. Los Angeles, CA: Muthén & Muthén; 2012.
- Jackson DL, Gillaspy JA, Purc-Stephenson R. Reporting practices in confirmatory factor analysis: an overview and some recommendations. Psychol Methods. 2009; 14: 6–23. doi: <u>10.1037/a0014694</u> PMID: <u>19271845</u>
- Schreiber JB, Stage FK, King J, Nora A, Barlow EA. Reporting structural equation modeling and confirmatory factor analysis results: A review. J Educ Res. 2006; 99: 323–337.
- Morata-Ramírez M, Holgado-Tello FP. Construct Validity of Likert Scales through Confirmatory Factor Analysis: A Simulation Study Comparing Different Methods of Estimation Based on Pearson and Polychoric Correlations. Int J Soc Sci Stud. 2013; 1: 54–61.
- Gardner DM, Murphy AL, O'Donnell H, Centorrino F, Baldessarini RJ. International consensus study of antipsychotic dosing. Am J Psychiatry. 2010; 167: 686–93. doi: <u>10.1176/appi.ajp.2009.09060802</u> PMID: 20360319
- Hu L, Bentler P. Cutoff criteria fo fit indexes in covariance structure analysis: conventional criteria versus new alternatives. Struct Equ Model. 1999; 6: 1–55.
- Andreasen NC. Scale for the assessment of negative symptoms (SANS). University of Iowa, Iowa City; 1983.
- Kirkpatrick B, Buchanan RW, McKenney PD, Alphs L, Carpenter WT. The Schedule for the Deficit syndrome: an instrument for research in schizophrenia. Psychiatry Res. 1989; 30: 119–123. PMID: 2616682
- Foussias G, Agid O, Fervaha G, Remington G. Negative symptoms of schizophrenia: Clinical features, relevance to real world functioning and specificity versus other CNS disorders. Eur Neuropsychopharmacol. Elsevier; 2014; 24: 693–709.
- Alphs L. An industry perspective on the NIMH consensus statement on negative symptoms. Schizophr Bull. 2006; 32: 225–230. PMID: <u>16469940</u>

- Kirkpatrick B, Fenton WS, Carpenter W, Marder SR. The NIMH-MATRICS consensus statement on negative symptoms. Schizophr Bull. 2006; 32: 214–9. PMID: <u>16481659</u>
- 42. De Heer-Wunderink C, Visser E, Caro-Nienhuis A, Sytema S, Wiersma D. Supported housing and supported independent living in the Netherlands, with a comparison with England. Community Ment Health J. 2012; 48: 321–7. doi: 10.1007/s10597-011-9381-1 PMID: 21246274
- Markou P. Depression in schizophrenia: A descriptive study. Aust N Z J Psychiatry. 1996; 30: 354–357. PMID: <u>8839947</u>
- Kitamura T, Suga R. Depressive and negative symptoms in major psychiatric disorders. Compr Psychiatry. 1991; 32: 88–94. PMID: <u>2001626</u>
- Sax KW, Strakowski SM, Keck PE, Upadhyaya VH, West S a, McElroy SL. Relationships among negative, positive, and depressive symptoms in schizophrenia and psychotic depression. Br J Psychiatry. 1996; 168: 68–71. PMID: 8770431
- Norman R, Malla A. Correlations Over Time Between Dysphoric Mood and Symptomatology in Schizophrenia. Compr Psychiatry. 1994; 37: 34–38.
- Newcomer JW, Faustman WO, Yeh W, Csernansky JG. Distinguishing depression and negative symptoms in unmedicated patients with schizophrenia. Psychiatry Res. 1990; 31: 243–50. PMID: 2333356
- Lindenmayer J, Grochowski S, Kay SR. Schizophrenic patients with depresion: psychopathological profiles and relationship with negative symptoms. Compr Psychiatry. 1991; 32: 528–533. PMID: <u>1778080</u>
- Bell MD, Corbera S, Johannesen JK, Fiszdon JM, Wexler BE. Social cognitive impairments and negative symptoms in schizophrenia: are there subtypes with distinct functional correlates? Schizophr Bull. 2013; 39: 186–96. doi: <u>10.1093/schbul/sbr125</u> PMID: <u>21976710</u>
- Rocca P, Montemagni C, Zappia S, Piterà R, Sigaudo M, Bogetto F. Negative symptoms and everyday functioning in schizophrenia: a cross-sectional study in a real world-setting. Psychiatry Res. 2014; 218: 284–9. doi: 10.1016/j.psychres.2014.04.018 PMID: 24814140
- Staring ABP, Ter Huurne M-AB, van der Gaag M. Cognitive Behavioral Therapy for negative symptoms (CBT-n) in psychotic disorders: a pilot study. J Behav Ther Exp Psychiatry. 2013; 44: 300–6. doi: <u>10.</u> <u>1016/j.jbtep.2013.01.004</u> PMID: <u>23454550</u>
- 52. Anthony WA, Cohen M, Farkas M, Gagne C. Psychiatric Rehabilitation. Boston: Center for Psychiatric Rehabilitation; 2002.
- Velligan D, Bow-thomas CC, Huntzinger C, Ritch J, Ledbetter N, Prihoda TJ, et al. Randomized Controlled Trial of the Use of Compensatory Strategies to Enhance Adaptive Functioning in Outpatients With Schizophrenia. Am J Psychiatry. 2000; 157: 1317–1323. PMID: <u>10910797</u>
- Mendella PD, Burton CZ, Tasca G a, Roy P, St Louis L, Twamley EW. Compensatory cognitive training for people with first-episode schizophrenia: results from a pilot randomized controlled trial. Schizophr Res. 2015; 162: 108–11. doi: 10.1016/j.schres.2015.01.016 PMID: 25631454
- Lyne JP, Kinsella A, O'Donoghue B. Can we combine symptom scales for collaborative research projects? J Psychiatr Res. 2012; 46: 233–238. doi: <u>10.1016/j.jpsychires.2011.10.002</u> PMID: <u>22056401</u>
- 56. Kirkpatrick B, Strauss GP, Nguyen L, Fischer B a, Daniel DG, Cienfuegos A, et al. The brief negative symptom scale: Psychometric properties. Schizophr Bull. 2011; 37: 300–305. doi: <u>10.1093/schbul/sbq059</u> PMID: <u>20558531</u>
- 57. Kring AM, Gur R, Blanchard JJ, Horan W, Reise SP. The Clinical Assessment Interview for Negative Symptoms (CAINS): Final development and validation. Am J Psychiatry. 2013; 170: 165–172. doi: <u>10.</u> <u>1176/appi.ajp.2012.12010109</u> PMID: <u>23377637</u>
- Aleman A. Neurocognitive basis of schizophrenia: information processing abnormalities and clues for treatment. Adv Neurosci. 2014;