

# Note: More Efficient Conversion of Equivalence-Query Algorithms to PAC Algorithms

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## Abstract

We present a method for transforming an Equivalence-query algorithm using  $Q$  queries into a PAC-algorithm using  $\frac{Q}{\epsilon} + O(\frac{Q^{2/3}}{\epsilon} \log \frac{Q}{\delta})$  examples in expectation. The method is a variation of that by Schuurmans and Greiner which provides, for each  $\gamma > 0$ , an algorithm using  $(1 + \gamma)\frac{Q}{\epsilon} + O(\frac{1}{\epsilon} \log \frac{Q}{\delta})$  examples in expectation. In other words, we show that the constant in front of the dominating term  $Q/\epsilon$  can be made  $1 + o(1)$ .

## 1 Introduction

In her seminal paper on learning from queries, Angluin [Ang87] showed that algorithms using Equivalence queries can be rewritten as PAC algorithms. Her simulation uses a worst-case sample  $O(\frac{Q^2}{\epsilon} \ln \frac{1}{\delta})$  to achieve  $(\epsilon, \delta)$ -confidence from an algorithm using  $Q$  Equivalence queries, but it is not difficult to show that in her same simulation, sample size  $O(\frac{Q}{\epsilon} \ln \frac{Q}{\delta})$  suffices.

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It was shown later that, with a different algorithm, that the dependence on  $n$  can be made linear. Specifically, Littlestone [Lit89] showed that there is a simulation using a worst-case sample size  $4 \frac{Q}{\epsilon} + O(\frac{1}{\epsilon} \ln \frac{Q}{\delta})$  (his simulation was phrased in terms of on-line learning rather than Equivalence queries, but the distinction is irrelevant for our purpose). Schuurmans and Greiner [SG95, Sch96] showed how to build, for every constant  $\gamma > 0$ , a simulation that uses *expected* sample size  $(1 + \gamma) \frac{Q}{\epsilon} + c(\gamma) \frac{1}{\epsilon} \ln \frac{Q}{\delta}$ . Here  $c(\gamma)$  is constant for each  $\gamma$ , but tends to infinity as  $\gamma$  tends to 0.

In this note we show that the leading constant in front of the  $Q/\epsilon$  term can be made  $1 + o(1)$ , that is, arbitrarily close to 1 as  $Q$  grows. In fact, our algorithm is essentially the same as the Schuurmans-Greiner one, except that instead of using a fixed value for  $\gamma$  a priori, we let the value of  $\gamma$  decrease at a precisely controlled rate as the algorithm progresses.

## 2 The Algorithm

We view an Equivalence query algorithm as a particular case of a strategy for generating hypothesis from sequences of labelled examples. Given such an algorithm, we build a new algorithm  $\mathbf{S}$ , given in Figure 1, which reads a sequence of example, uses the Equivalence-query strategy as a black box, and eventually outputs a hypothesis from those generated by the strategy. We will show that  $\mathbf{S}$  is a PAC-learning algorithm.

Procedure `sprt` is Wald's Sequential Probability Ratio Test, discussed below, and also used in the Schuurmans-Greiner approach. The main difference with their method is that we do not fix a constant  $\gamma$  *a priori*, but rather use a different  $\gamma_i$  that varies with  $i$ . We will fix one particular setting for the sequence of  $\gamma_i$  to obtain our bound on the sample size used by  $\mathbf{S}$ , but occasionally comment on the effect of using other values for  $\gamma_i$ .

We will argue that procedure  $\mathbf{S}$  satisfies three conditions, which we formulate as theorems: Correctness, Completeness, and Efficiency.

**Theorem 1 (Correctness)** *The probability that  $\mathbf{S}(\epsilon, \delta)$  outputs some  $h \in H$  with  $\text{error}(h) > \epsilon$  is less than  $\delta$ .*

The completeness condition can be stated in many ways, of which the following is but one example:

**Theorem 2 (Completeness)** *If for some  $i$  we have that  $\text{error}(h_i) = 0$  with probability 1, then  $\mathbf{S}(\epsilon, \delta)$  stops with probability 1.*

Algorithm  $\mathbf{S}(\epsilon, \delta)$

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1  Generate initial hypothesis  $h_1$ ;  
2   $i := 1; t := 0$ ;  
3  while TRUE  
4      do  
5           $t := t + 1$ ;  
6          get a training example  $(x_t, c(x_t))$ , labelled by the unknown target  $c$ ;  
7          if  $h_i(x_t) \neq c(x_t)$  (i.e.,  $(x_t, c(x_t))$  is a counterexample for  $h_i$ )  
8              then  
9                  use  $(x_t, y_t)$  to generate  $h_{i+1}$ ;  
10                 start testing  $error(h_i)$  on subsequent examples  
11                    using  $\mathbf{sprt}(\epsilon/(1 + \gamma_i), \epsilon, \delta/(i(i + 1)), 0)$ ;  
12                  $i := i + 1$ ;  
13          if for some  $j < i$ , the  $\mathbf{sprt}$  test for  $h_j$  rejects  
14              then  
15                  drop  $h_j$  from the list of hypothesis being tested  
16          if for some  $j < i$ , the  $\mathbf{sprt}$  test for  $h_j$  accepts  
17              then  
18                  output  $h_j$  and stop  
19  end while
```

Figure 1: Algorithm  $\mathbf{S}$

Putting both claims together, if the strategy used to generate hypothesis is an exact Equivalence-query algorithm learning with finitely many queries, with probability 1 the algorithm stops, and its output is, with probability  $1 - \delta$ , a hypothesis  $h$  having  $\text{error}(h) < \epsilon$ .

Theorem 2 in fact follows from this more general statement:

**Theorem 3 (Running time)** *Define  $\gamma_i = i^{-1/3}$ , and let the base Equivalence-query learner learn with at most  $Q$  queries. Then*

$$E[\text{running time of } \mathbf{S}(\epsilon, \delta)] \leq \frac{Q}{\epsilon} + 7 \frac{Q^{2/3}}{\epsilon} \cdot \left( \ln \frac{Q(Q+1)}{\delta} + 2 \right).$$

We do not describe here the **sprt** test. We quote, however, some relevant properties from [Sch96], appendix A:

**Theorem 4** [Sch96] *Let  $k > 1$  and suppose  $\mathbf{sprt}(\epsilon/k, \epsilon, \delta_{\text{acc}}, \delta_{\text{rej}})$  is run on a sequence  $X_1, X_2, \dots, X_i, \dots$  of i.i.d. boolean random variables. Then:*

1. *If  $E[X_i] > \epsilon$ , the probability that **sprt** accepts is at most  $\delta_{\text{acc}}$ .*
2. *If  $E[X_i] < \epsilon/k$ , the probability that **sprt** rejects is at most  $\delta_{\text{rej}}$ .*
3. *([Sch96], Lemma A.4) If  $\delta_{\text{rej}} = 0$ , the expected running time of **sprt** is*

$$\left( \frac{k}{k-1-\ln k} \right) \frac{1}{\epsilon} \left( \ln \frac{1}{\delta_{\text{acc}}} + 1 \right).$$

### 3 Proof of Theorem 1

The proof is as in [SG95, Sch96], but we reproduce it for completeness. We say that a hypothesis  $h \in H$  is  $\epsilon$ -bad iff  $\text{error}(h) \geq \epsilon$ . Observe that the **sprt** instance associated to  $h_i$  is fed boolean variables whose expected value is precisely  $\text{error}(h_i)$ . Therefore, by Theorem 4, part (1), we have the following (where probabilities are taken over infinite sequences of independently generated examples).

$$\begin{aligned} & \Pr[\mathbf{S}(\epsilon, \delta) \text{ outputs an } \epsilon\text{-bad hypothesis}] \\ & \leq \sum_{i=1}^{\infty} \Pr[h_i \text{ is } \epsilon\text{-bad yet } \mathbf{S}(\epsilon, \delta) \text{ outputs } h_i] \end{aligned}$$

$$\begin{aligned}
&\leq \sum_{i=1}^{\infty} \Pr[\text{sprt}(\epsilon/(1 + \gamma_i), \epsilon, \delta/(i(i + 1)), 0) \text{ accepts } h_i \mid h_i \text{ is } \epsilon\text{-bad}] \\
&\leq \sum_{i=1}^{\infty} \frac{\delta}{i(i + 1)} = \delta.
\end{aligned}$$

## 4 Proof of Theorem 3

For every  $i$ , we define the following random variables and expected values:

- $h_i$  is the random variable representing the  $i$ th generated hypothesis,
- $\epsilon_i$  is such that  $1/\epsilon_i = E[1/\text{error}(h_i)]$ ,
- $T_i$  is the number of examples read from the moment in which  $h_i$  is generated until either  $h_{i+1}$  is generated (if  $h_{i+1}$  is ever generated; otherwise,  $T_i = \infty$ )
- $R_i$  is the running time of the `sprt` test run on  $h_i$ , and
- $T$  is the running time of the algorithm.

Proving Theorem 3 is thus bounding  $E[T]$ . Let  $i$  be the first index such that  $\epsilon_i(1 + \gamma_i) < \epsilon$ . Note that if the base Equivalence learner uses at most  $Q$  queries, we have  $i \leq Q$ . Observe also that

$$T \leq \sum_{j < i} T_j + R_i \tag{1}$$

because, by definition of  $T_j$  and  $R_i$ , by this time  $h_i$  has been generated and the `sprt` test for  $h_i$  has stopped. Since the test is run with parameter  $\delta_{rej}$ , it rejects  $h_i$  with probability 0, i.e., it accepts  $h_i$ . Therefore, by this time either  $\mathbb{S}$  stops outputting  $h_i$ , unless it has stopped before due to another  $h_j$ .

Taking expectations in Equation (1), we have

$$E[T] \leq \sum_{j < i} E[T_j] + E[R_i]. \tag{2}$$

We first bound  $E[T_j]$ ; the proof of the lemma is given later.

**Lemma 1**  $E[T_j] = 1/\epsilon_j$ .

Taking  $k = (1 + \gamma_i)$  in Theorem 4, part (3), provides the following bound on  $E[R_i]$ :

$$E[R_i] \leq \frac{1 + \gamma_i}{\gamma_i - \ln(1 + \gamma_i)} \frac{1}{\epsilon} \left( \ln \frac{i(i+1)}{\delta} + 1 \right). \quad (3)$$

As a detour, let us note how to get the result in [SG95, Sch96]. Since  $i$  is the first index such that  $\epsilon_i(1 + \gamma_i) < \epsilon$ , for  $j < i$  we have  $\epsilon_j \geq \epsilon/(1 + \gamma_j)$ , that is,  $E[T_j] = 1/\epsilon_j \leq (1 + \gamma_j)/\epsilon$ . Fix  $\gamma_i = \gamma$  for every  $i$ . Then from Equation (2) we get

$$\begin{aligned} E[T] &\leq \sum_{j < i} \frac{1 + \gamma}{\epsilon} + \frac{1 + \gamma}{\gamma - \ln(1 + \gamma)} \frac{1}{\epsilon} \left( \ln \frac{i(i+1)}{\delta} + 1 \right) \\ &= (1 + \gamma) \frac{i}{\epsilon} + c(\gamma) \frac{1}{\epsilon} \left( \ln \frac{i(i+1)}{\delta} + 1 \right). \end{aligned}$$

Now, take instead  $\gamma_i = i^{-1/3}$ . We have the following two lemmas, whose proofs are given later:

**Lemma 2** For  $\gamma_j = j^{-1/3}$ ,

$$\sum_{j < i} (1 + \gamma_j) \leq i + \frac{3}{2} i^{2/3}.$$

**Lemma 3** Define  $c(\gamma) = (1 + \gamma)/(\gamma - \ln(1 + \gamma))$ . Then  $c(\gamma) \leq 7/\gamma^2$  for every  $\gamma \in (0, 1]$ , and  $c(\gamma)$  tends to  $2/\gamma^2$  as  $\gamma$  tends to 0.

From Equations (2) and (3) and Lemmas 2 and 3, and using again that for all  $j < i$  we have  $E[T_j] = 1/\epsilon_j \leq (1 + \gamma_j)/\epsilon$ , we obtain

$$\begin{aligned} E[T] &\leq \sum_{j < i} \frac{1 + \gamma_j}{\epsilon} + \frac{7}{\gamma_i^2} \frac{1}{\epsilon} \left( \ln \frac{i(i+1)}{\delta} + 1 \right) \\ &\leq \frac{1}{\epsilon} \left( i + \frac{3}{2} i^{2/3} \right) + 7 \frac{i^{2/3}}{\epsilon} \left( \ln \frac{i(i+1)}{\delta} + 1 \right) \\ &\leq \frac{i}{\epsilon} + 7 \frac{i^{2/3}}{\epsilon} \left( \ln \frac{i(i+1)}{\delta} + 2 \right) \end{aligned}$$

i.e., the statement of Theorem 3.

**Proof of Lemma 1.** Suppose that in a particular run of the algorithm the random variable  $h_j$  takes a particular value  $h \in H$ . Conditioned to  $h_j = h$ , the expected number of examples that have to be read to produce a counterexample for  $h_j$  is an exponential distribution with base  $error(h)$ , and therefore,

$$E[T_j | h_j = h] = \sum_{\ell=1}^{\infty} (1 - error(h))^{\ell-1} \cdot error(h) \cdot \ell = 1/error(h).$$

So  $E[T_j] = E[1/error(h_j)]$  (where the expectation is taken over  $h$  on the right-hand side), which is  $1/\epsilon_j$  by definition of  $\epsilon_j$ . ■ (Lemma 1)

**Proof of Lemma 2.** We show by induction on  $i$  the following inequality, which implies the lemma:

$$\sum_{j \leq i} (1 + j^{-1/3}) \leq \frac{i}{\epsilon} + \frac{3}{2} \frac{i^{2/3}}{\epsilon}.$$

For  $i = 1$  it is obvious. Assume true for  $i$ , then

$$\sum_{j=1}^{i+1} j^{-1/3} \leq \frac{3}{2} i^{2/3} + (i+1)^{-1/3}$$

and observe that

$$\frac{3}{2} i^{2/3} + (i+1)^{-1/3} \leq \frac{3}{2} (i+1)^{2/3}$$

iff (multiplying on both sides by  $(i+1)^{1/3}$ )

$$\frac{3}{2} (i^2(i+1))^{1/3} + 1 \leq \frac{3}{2} (i+1)$$

iff (taking cubes on both sides)

$$\left(\frac{3}{2}\right)^3 (i^2(i+1)) \leq \left(\frac{3}{2}(i+1) - 1\right)^3$$

which is verified to be true by simple algebra. ■ (Lemma 2)

**Proof of Lemma 3.** We have  $c(1)1^2 = 2/(1 - \ln(2)) < 7$ , and studying the Taylor expansion of  $c(\gamma)\gamma^2$  shows that it is strictly increasing with  $\gamma$ , so  $c(\gamma)\gamma^2 < 7$  for all  $\gamma < 1$ . Also, for small enough  $\gamma$  we have  $\ln(1+\gamma) \cong \gamma - \gamma^2/2$ , from which  $c(\gamma) \cong 2/\gamma^2$  follows. ■ (Lemma 3)

## 5 Final Remarks

Observe that Theorem 3 does not strictly require that the algorithm produces an hypothesis with 0 error within the first  $Q$  queries. It is enough to assume that within the first  $Q$  queries it generates a hypothesis  $h_i$  with  $\epsilon_i(1+\gamma_i) < \epsilon$ .

Note also that a variety of bounds on the sample size are possible by taking other definitions for  $\gamma_i$ . In particular, with essentially the same proof, if we take  $\gamma_i = 1/i^\beta$  for  $\beta < 1$ , we obtain (approximately)

$$E[T] \leq \frac{Q}{\epsilon} + \frac{1}{1-\beta} \frac{Q^{1-\beta}}{\epsilon} + 7 \frac{Q^{2\beta}}{\epsilon} \ln \frac{Q(Q+1)}{\delta}.$$

We just chose  $\beta = 1/3$  to make  $1 - \beta = 2\beta$ , but if the values of  $Q$  and  $\delta$  are known in advance, other values of  $\beta$  may give better bounds.

Finally, as indicated by Lemma 3, the factor 7 in front of the second term is actually a decreasing function of  $Q$  that tends to 2 as  $Q$  grows.

## References

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