A NEW LARGE DEVIATION INEQUALITY FOR U-STATISTICS OF ORDER 2

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Abstract. We prove a new large deviation inequality with applications when projecting a density on a wavelet basis.

Résumé. Nous prouvons une inégalité de grandes déviations applicable à la projection d'une densité sur une base d'ondelettes.

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1. INTRODUCTION AND MAIN RESULT

Let F be some law on \mathbb{R} . When g is a measurable function from \mathbb{R}^d to \mathbb{R} , E(g) and Var(g) denote expectation and variance with respect to the Probability $F^{\otimes d}$. Let f be a bounded and symmetric function from \mathbb{R}^2 to \mathbb{R} . Following Arcones and Giné [1], we construct its canonical projections: ξ and η being independent with law F

$$\pi_1 f(x) = Ef(x, \eta) - Ef(\xi, \eta) \cdot \pi_2 f(x, y) = f(x, y) - Ef(x, \eta) - Ef(\xi, y) + Ef(\xi, \eta) \cdot$$
(1)

Let ξ_i , $i = 1, 2, \dots, n$ be a *n*-sample of F $(n \ge 2)$. We consider the *U*-statistics (without any normalisation)

$$U_n^{(2)}(f) = \sum_{1 \le i \ne j \le n} f(\xi_i, \xi_j) \cdot U_n^{(1)}(\pi_1 f) = \sum_{1 \le i \le n} \pi_1 f(\xi_i) \cdot U_n^{(2)}(\pi_2 f) = \sum_{1 \le i \ne j \le n} \pi_2 f(\xi_i, \xi_j), \text{ thus}$$
$$U_n^{(2)}(f - Ef) = 2(n - 1)U_n^{(1)}(\pi_1 f) + U_n^{(2)}(\pi_2 f) \cdot$$
(2)

We are interested in a large deviation inequality for the latter U-statistic when f is centered and bounded.

First, if $|f| \leq c$ and $Ef^2 = \sigma^2$, the usual Bernstein type inequality is

$$P(U_n^{(2)}(f - Ef) \ge n(n-1)t) \le \exp\left(-[n/2]t^2/\{2\sigma^2 + 2ct/3\}\right).$$
(AG1)

But we can consider $U_n^{(1)}(\pi_1 f)$ (which is a sum of i.i.d. \mathbb{R} -valued random variables) as the main part and it can be interesting to bound separately the second part $U_n^{(2)}(\pi_2 f)$. Now, as $\pi_2 f$ is canonical of order 2, if $|\pi_2 f| \leq c$

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and $\sigma^2 = E(\pi_2 f)^2$, there exist two constants c_1, c_2 such that

$$P(|U_n^{(2)}(\pi_2 f)| \ge (n-1)t) \le c_1 \exp(-c_2 t / \{\sigma + c^{2/3} t^{1/3} n^{-1/3}\})$$
(AG2)

((AG1) and (AG2) can be found in Arcones and Giné [1]).

The normalized version of inequality (AG2) can be restated

$$P(|U_n^{(2)}(\pi_2 f)|) \ge a_3 \operatorname{Max}(n\sigma x, c\sqrt{n}x^{3/2})) \le a_4 \exp(-x).$$
(AG2')

The aim of the paper is to give a new large deviation inequality (Th. 1 and Cor. 1 later). To every partition D we associate two functionnals $||f||_D$ and $v_D(F)$ (see Def. 4 later) such that

$$P(|U_n^{(2)}(\pi_2 f)| \ge a_3' x ||f||_D \operatorname{Max}(n\sqrt{v_D(F)}, 1)) \le a_4' \exp(-2\sqrt{x})$$
(3)

where all a_3, a'_3 are universal constants. For a comparison between the two inequalities, see the discussion after Theorem 2 infra.

We need now some definitions: let $(I_{\lambda} \mid \lambda \in D)$ be a Borelian partition of \mathbb{R} finite or enumerable, where I_{λ} denotes the subset and its indicator. Let τ be a permutation of D. Its graph g_{τ} is $\{(\lambda, \tau(\lambda)) \mid \lambda \in D\}$, a subset of $D \times D$. A collection G is an enumerable set $(\tau_s \mid s \in G)$ of permutations such that $D \times D \subset \bigcup_{s \in G} g_{\tau_s}$. \mathbb{H} is the family of collections. For any matrix $M = (a_{\lambda,\mu} \mid \lambda \in D; \mu \in D)$ we set

$$\|M\|_G = \sum_{i \in G} \sup_{\lambda \in D} |a_{\lambda, \tau_i(\lambda)}|$$
 and $\|M\|_D = \inf_{G \in \mathbb{H}} \|M\|_G$.

Let f be a bounded real valued function defined on $\mathbb{R} \times \mathbb{R}$. We set

$$M(D, f) = (a_{\lambda,\mu}) \text{ where } a_{\lambda,\mu} = \text{ Sup } (f^2(x,y)I_\lambda(x)I_\mu(y))$$
$$\|f\|_D^2 = \|M(D,f)\|_D$$
$$v_D(F) = \Sigma_\lambda P(\xi \in I_\lambda)^2.$$

Our main result is

Theorem 1. There exists some constant C ($C = 80\pi$ holds) such that for every integer k > 0, for every partition D, for every symmetric f

$$E(U_n^{(2)}(\pi_2 f)^{2k}) \le C^{2k} \times (E\mathcal{N}^{2k})^4 \times ||f||_D^{2k} Max \ (n^{2k}(v_D(F))^k, n^2v_D(F)).$$

where \mathcal{N} denotes the standard normal distribution.

Corollary 1. There exists some constant C such that for every partition D, for every symmetric f, for every x > 0

$$P(|U_n^{(2)}(\pi_2 f)| \ge Cx ||f||_D \operatorname{Max}(n\sqrt{v_D(F)}, 1)) \le \exp(6 - 2\sqrt{x}).$$

As we will see in the following Discussion,

Remark 1: The classical inequality (AG) is strictly better whenever $x \le n\sigma^2/c^2$. **Remark 2**: Nevertheless, our inequality can work when the classical one does not. **Remark 3**: Finally, up to some logarithm, \sqrt{x} is the best possible rate.

Main application: Let (Ψ_1, Ψ) be a wavelet and ε_{λ} , $\varepsilon_{\ell,\lambda}$ be the associated basis:

$$\varepsilon_{\lambda}(x) = \Psi_1(x+\lambda) / \|\Psi_1\|_2, \varepsilon_{\ell,\lambda}(x) = 2^{\ell/2} \Psi(2^\ell x + \lambda) / \|\Psi\|_2 (\lambda \in \mathbb{Z}, \ell \in \mathbb{N}).$$

Let p be some density of Probability on \mathbb{R} equipped with Lebesgue measure dx, assumed to be square integrable in applications.

Let β_{λ} be $\int p(x)\varepsilon_{\lambda}(x)dx$ and $\gamma_{\ell,\lambda}$ be $\int \varepsilon_{\ell,\lambda}(x)p(x)dx$.

We set $p_L = \Sigma_\lambda \beta_\lambda \varepsilon_\lambda + \Sigma_{\ell \leq L} \Sigma_\lambda \gamma_{\ell,\lambda} \varepsilon_{\ell,\lambda}$ (the projection of p up to level of resolution L) and want to estimate its square norm

$$\theta_L \doteq \|p_L\|_2^2 = \Sigma_\lambda \beta_\lambda^2 + \Sigma_{\ell \le L} (\Sigma_\lambda \gamma_{\ell,\lambda}^2)) \cdot$$
(4)

When $(\xi_i \mid 1 \leq i \leq n)$ is a *n*-sample with density p $(n \geq 2)$, the empirical estimators of coefficients are $\beta_{\lambda,n} = \sum_i \varepsilon_{\lambda}(\xi_i)/n$, $\gamma_{\ell,\lambda,n} = \sum_i \varepsilon_{\ell,\lambda}(\xi_i)/n$. Let $p_{L,n}$ be

$$p_{L,n} = \Sigma_{\lambda} \beta_{\lambda,n} \varepsilon_{\lambda} + \Sigma_{\ell \leq L} \Sigma_{\lambda} \gamma_{\ell,\lambda,n} \varepsilon_{\ell,\lambda}.$$

The "natural estimator" for θ_L is

$$\Sigma_{\lambda}\beta_{\lambda,n}^2 + \Sigma_{\ell \leq L}(\Sigma_{\lambda}\gamma_{\ell,\lambda,n}^2).$$

But the latter has positive bias, and the unbiased estimator is

$$\widehat{\theta}_{L,n} = (n(n-1))^{-1} \Sigma_{1 \le i \ne j \le n} \{ \Sigma_{\lambda} \varepsilon_{\lambda}(\xi_i) \varepsilon_{\lambda}(\xi_j) + \Sigma_{\ell \le L} (\Sigma_{\lambda} \varepsilon_{\ell,\lambda}(\xi_i) \varepsilon_{\ell,\lambda}(\xi_j)) \}$$
(5)

Let $\Delta_{n,L} = \hat{\theta}_{L,n} - \theta_L$. According to (5), $\Delta_{n,L}$ can be decomposed into canonical *U*-statistics in the following way: let Φ_1 be $\Psi_1/||\Psi_1||_2$ and let Φ be $\Psi/||\Psi||_2$,

$$f(x,y) = \Sigma_{\lambda} \Phi_1(x+\lambda) \Phi_1(y+\lambda), \ f_{\ell}(x,y) = \Sigma_{\lambda} \Phi(2^{\ell}x+\lambda) \Phi(2^{\ell}y+\lambda)$$
(3)

$$\delta_n = (2/n)U_n^{(1)}(\pi_1 f) + (1/n(n-1))U_n^{(2)}(\pi_2 f)$$
(4)

$$\delta_{\ell,n} = (2 \cdot 2^{\ell}/n) U_n^{(1)}(\pi_1 f_{\ell}) + (2^{\ell}/n(n-1)) U_n^{(2)}(\pi_2 f_{\ell}).$$
(6)

Then $\Delta_{n,L} = \delta_n + \Sigma_{\ell \leq L} \delta_{\ell,n}$. In the decomposition above, the sum of U-statistics of order 1 is equal to the part up to level of resolution L of $\int p(p_n - p)$. It can be bounded in Probability by classical Bernstein's inequality. The control of each U-statistic of order 2 will be performed by our inequality. This is very useful, either to estimate θ_L [3], or in model selection: in this problem, the authors consider a wide family of finite dimensional projections. Our study is quite general, but we observe that if Ψ_1 and Ψ are with compact support, for every L the family $(\beta_{\lambda}, \alpha_{\ell,\lambda} \mid \ell \leq L)$ is in fact with finite dimension, and hypothesis (H) infra holds. Thus our result can be used in adaptive estimation of quadratic functionals in a density model (see for example [4] where white noise is treated), where the theory needs good bounds up to $2^{\ell} = O(n^2)$. These bounds cannot be obtained by classical Hoeffding's bounds (see the chapter "Discussion").

We assume in the whole paper

Hypothesis (H).

$$|\Phi(x)| \le \sum_{u \in \mathbb{Z}} \omega_u \mathbf{1}_{u \le x < u+1} \text{ with } \sum_{u \in \mathbb{Z}} \omega_u < \infty$$

and we set $M^2(\Phi) = \sum_{(u,v,w) \in \mathbb{Z}^3} \omega_u \omega_v \omega_{u+w} \omega_{v+w}.$ (H)

In the particular case when the law has density p, for the normalised U-statistic, we have

Theorem 2. We assume that the function Φ with $\|\Phi\|_2 = 1$ satisfies hypothesis (H), and that the law F has a density p with $\|p\|_2 = (\int p^2(x) dx)^{1/2} < \infty$. Set

$$\|p\|_{2,\ell} = 2^{\ell/2} \left[\sum_{k \in \mathbb{Z}} \left(\int_{k2^{-\ell}}^{(k+1)2^{-\ell}} p(x) dx \right)^2 \right]^{1/2}$$

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$$\begin{split} \delta_{\ell,n}^{(2)} &= (2^{\ell}/n(n-1))U_n^{(2)}(\pi_2 f_{\ell}) \text{ where } f_{\ell}(x,y) = \Sigma_{\lambda} \Phi(2^{\ell}x+\lambda) \Phi(2^{\ell}y+\lambda) \\ &Z_{\ell,n} = \sqrt{n(n-1)} 2^{-\ell/2} \delta_{\ell,n}^{(2)}. \end{split}$$

Then, if $n\|p\|_{2,\ell} \ge 2^{\ell/2}$, we have, for C as in Theorem 1

$$P(|Z_{\ell,n}| \ge 2C \times M(\Phi) \times x) \le \exp(6 - 2\sqrt{x}).$$

In this formula, $Var(Z_{\ell,n})$ does not depend on n. Moreover, there exists positive and finite constants $a(\Phi)$, $b(\Phi)$ depending only on Φ such that

$$\lim_{\ell \to \infty} \|p\|_{2,\ell} = \|p\|_{2} \cdot$$
$$\lim_{\ell \to \infty} \|\pi_2 f_\ell\|_{\infty} = b(\Phi) \text{ and } a(\Phi) \|p\|_2^2 \leq \lim_{\ell \to \infty} \operatorname{Var}(Z_{\ell,n}) \cdot$$
$$\|\pi_2 f_\ell\|_{\infty} \leq 4M(\Phi) \text{ and } \operatorname{Var}(Z_{\ell,n}) \leq \|p\|_2^2 M^2(\Phi) \cdot$$

Remark 4: Such a result is interesting only if it works, for given n, ℓ , uniformly for large classes of densities. Obviously we need some uniform control of $||p||_2$, but this is not sufficient in view of condition $n||p||_{2,\ell} \ge 2^{\ell/2}$.

Thus we need some extra condition. If for example we assume that the support of p is contained in some interval [x, x + M] we get $||p||_{2,\ell} \ge 1/(M+1)$ and the bound works if $n^2 \ge (M+1)2^{\ell}$.

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2. Discussion

a) About Corollary 1.

We consider inequality stated in Corollary 1 and assertion (AG2"). We assume $||f||_{\infty} = 1$. Without further knowledge about the law F, we can only bound $\sigma^2 = Var(\pi_2(f))$ by Ef^2 and $c = ||\pi_2 f||_{\infty}$ by 4. Up to some change of a_3 , (AG2') is restated

$$P\left(|U_n^{(2)}(\pi_2 f)| \ge a_3 \operatorname{Max}\left(nx\sqrt{Ef^2}, \sqrt{nx^{3/2}}\right)\right) \le a_4 \exp(-x).$$

On the other hand, we have the obvious inequality

$$Ef^2 \le ||f||_D^2 v_D(F).$$

Proof of Remark 1: Let us assume $x \leq n\sigma^2$. Thus $\operatorname{Max}(n\sigma x, c\sqrt{n}x^{3/2}) = n\sigma x \leq n ||f||_2 x \leq x ||f||_D$ Max $(n\sqrt{v_D(F)}, 1)$ and, up to constants (AG2') is always better than the bound of Corollary 1 whenever $x \leq n\sigma^2$.

Nevertheless, our inequality provides a possible alternative if there exists some partition D such that $n^2 v_D \ge 1$ and $||f||_D^2 v_D(F) \approx Ef^2$ (up to constant). We exhibit two extremal cases in the case when the law is the uniform one:

1) $f(x,y) = \mathbb{1}_{0 \le x < p; 0 \le y < p}$, where 0 . We set <math>q = 1 - p. We have $||f||_{\infty} = 1$, $Ef^2 = p^2$. Choosing $I_1 = (0,p) \ I_{\lambda} = p + (\lambda - 2)q/K$, $p + (\lambda - 1)q/K$) for $1 < \lambda \le K + 1$, we get $v_D = p^2 + q^2/K$ and, for a convenient choice of G, $||f||_D^2 = 1$. Moreover $Var(\pi_2 f) = p^2 q^2$ and $||\pi_2 f||_{\infty} = max(p^2, q^2)$.

When p is small, up to constants, in the classical inequality we can use either precise true parameters or rough estimates $(||f||_{\infty} \text{ and } Ef^2 \text{ for } ||\pi_2 f||_{\infty} \text{ and } Var(\pi_2 f))$ and we get

$$P(|U_n^{(2)}(\pi_2 f)| \ge a \operatorname{Max}(nxp, \sqrt{nx^{3/2}})) \le a_4 \exp(-x)$$
(AG2')

Our result is

$$P(|U_n^{(2)}(\pi_2 f)| \ge Cx \text{ Max } (np, 1)) \le \exp(6 - 2\sqrt{x}).$$
(Cor 1)

In this setting, we can assume x > 1, n large.

If $1 < x < np^2$, the classical inequality is better, but (up to constants) only with respect to the exponent of x.

A contrario the classical result does not work in the case when $np^2 = o(1)$ but not our one provided that np is large, and this justify the Remark 2.

We will see that it is a quite general result in the main application.

Remark 3: Assuming np = 1 and denoting f_n the corresponding function, when $n \to \infty$, it is easy to prove that $U_n^{(2)}(\pi_2 f_n)$ converges in law to $Y^2 - 3Y + 1$ where the law of Y is the Poisson law with parameter 1. Thus

$$\liminf_{x \to \infty} \lim_{n \to \infty} \frac{1}{1} \exp\left(\frac{U_n^{(2)}(\pi_2 f_n) \ge x}{\sqrt{x} \log\left(\sqrt{x}\right)}\right) \ge 1$$

proving that the power 1/2 is the best possible.

2) Let be $g(x,y) = \mathbb{1}_{0 \le x \le 1; \ 0 \le y \le 1} - \mathbb{1}_{p \le x \le 1-p; \ p \le y \le 1-p}$ (p small) and $\varepsilon(x) = \mathbb{1}_{x \le 1/2} - \mathbb{1}_{1/2 < x}$ and finally $f(x,y) = \varepsilon(x)\varepsilon(y)g(x,y)$. Then, for the uniform law, $f = \pi_2 f$ and $Var(\pi_2 f) = Var(f) = 4p(1-p)$. Obviously, for every D, $||f||_D^2 = |D|$ and $v_D(F) \ge 1/D$, thus $||f||_D \max(n\sqrt{v_D(F)}, 1) \ge n$ (and n is obtained by the partition with one element).

The classical inequality gives for some universal a

$$P(|U_n^{(2)}(\pi_2 f)| \ge a \operatorname{Max}(n\sqrt{px}, \sqrt{nx^{3/2}})) \le a_4 \exp(-x)$$

and, whatever be p, our inequality provides only

$$P(|U_n^{(2)}(\pi_2 f)| \ge Cxn) \le \exp(6 - 2\sqrt{x})$$

a very poor result!

b) About the main application:

b1) We consider firstly the case of the Haar basis $(\Phi(x) = \mathbf{1}_{0 \le x < 1} \text{ or } \Phi(x) = \mathbf{1}_{0 \le x < 1/2} - \mathbf{1}_{1/2 \le x < 1})$, with uniform law on the interval]0,1]. In the first case, at the level ℓ , setting $D = ([\lambda 2^{-\ell}, (\lambda + 1)2^{-\ell}] | \lambda \in \mathbb{Z})$, $f_{\ell} = \sum_{\lambda \in \mathbb{Z}} \Phi(2^{\ell}x + \lambda) \Phi(2^{\ell}y + \lambda)$, we have $M(\Phi) = ||f_{\ell}||_D = 1$, $||\pi_2(f_{\ell})||_{\infty} = (1 - 2^{-\ell})$, $v_D(F) = 2^{-\ell}$ and $Var(\pi_2(f_{\ell})) = (1 - 2^{-\ell})^2 2^{-\ell}$.

Thus for every $\ell \geq 1$ we have

$$2^{\ell} v_D(F) = \|f\|_D = 1.$$

$$1/2 \le Var(Z_{\ell,n}) = 2^{\ell} Var(\pi_2 f_{\ell}) \le 1$$

$$1/2 \le \|\pi_2 f_{\ell}\|_{\infty} \le 1$$

and, whenever $2^{\ell} \leq n^2$, Theorem 2 provides

$$P\left(|Z_{\ell}|/\sqrt{Var(Z_{\ell,n})} \ge 2Cx\right) \le \exp(6 - 2\sqrt{x}).$$

The classical one provides

$$P\left(|Z_{\ell}|/\sqrt{Var(Z_{\ell,n})} \ge a_3(x + (2^{\ell}/n)^{1/2}x^{3/2})\right) \le a_4 \exp(-x)$$

and does not work if $n = o(2^{\ell})$.

Remark 5: Massart (private communication) thinks that using Talagrand's inequality the best possible bandwith is $2^{\ell} = O(n^{3/2})$ and this is the principal motivation of this work. b2) Finally, let $\kappa_{n,\ell}^2$ the chi-square (with $2^{\ell} - 1$ degrees of freedom) associated to the partition: with $N_{\lambda} = \sum_{1 \le i \le n} \mathbb{1}_{\lambda 2^{-\ell} < \xi_i \le (\lambda+1)2^{-\ell}}$

$$\kappa_{n,\ell}^2 = \Sigma_\lambda (N_\lambda - EN_\lambda)^2 / EN_\lambda.$$

The centered and normalised $\kappa_{n,\ell}^2$ is equal to $Z_{\ell}/\sqrt{Var(Z_{\ell,n})}$. Thus our result provides a large deviation inequality for $\kappa_{n,\ell}^2$ even in the case when $n \approx 2^{\ell/2}$. Remark that the mean number of visits EN_{λ} can be O(1/n)!

The second case is the best possible: for every ℓ we have $||f_{\ell}||_{\infty} = 1$, f_{ℓ} is canonical and $||f||_D^2 v_D(F) = Var(Z_{\ell})$.

b3) We consider now the general case in the main application:

Using the final assertions of Theorem 2, we see that we have asymptotically the same conclusion as in the case of b1: whenever ℓ is large, our Z, up to constants depending only on Φ and the law p, is the normalised U-statistic of order 2 corresponding to some canonical function the Sup norm of which is equivalent to 1.

Thus we get a large deviation inequality which cannot be obtained using the classical result for $2^{\ell/2} \ll n \ll 2^{\ell}$.

3. Proofs

The proof is based on *De la Peña's* inequalities [2]. As all bounds are continuous with respect to $||f||_D$, $v_D(F)$, it suffices to prove that, if G is a collection such that $\sum_{i \in G} \sup_{\lambda \in D} \sup (f^2(x,y)I_\lambda(x)I_{\tau_i(\lambda)}(y)) = 1$, then

$$E(V_n^{2k}) \le C^{2k} \times (\mathbb{E}\mathcal{N}^{2k})^4 \times \operatorname{Max}\left(n^2 v, (n^2 v)^k\right)$$
(8)

where, to simplify notations, we set

$$V_n = \sum_{1 \le i \ne j \le n} \pi_2 f(\xi_i, \xi_j) \quad \text{and} \quad v = v_D(F) \cdot$$
(9)

1: Symmetrization

Let $\varepsilon_i, \varepsilon'_i, \mathcal{N}_i, \mathcal{N}'_i, \xi_i, \eta_i$ be six independent n-samples: the common law of ε 's is the law of the centered sign, the common law of the \mathcal{N} 's is the normal $\mathcal{N}(0, 1)$, the common law of ξ 's and η 's is the law F.

Using the first Theorem of $De \ la \ Pe\tilde{n}a$, as $\pi_2 f$ is canonical, we get:

For every Γ even, increasing on \mathbb{R}^+ and convex

$$E\Gamma(V_n) \le E\Gamma(4\Sigma_{1\le i,j\le n;i\ne j}\pi_2 f(\xi_i,\eta_j))$$

Using the classical symmetrization inequalities (see [2] again), we have

$$E\Gamma(V_n) \le E\Gamma(16\Sigma_{1 \le i,j \le n; i \ne j}\varepsilon_i\varepsilon'_j\pi_2 f(\xi_i,\eta_j)).$$

As the ε_i , ε'_i can be viewed as conditional expectations of $\sqrt{\pi/2}\mathcal{N}_i$, $\sqrt{\pi/2}\mathcal{N}'_i$, using convexity again we get

$$E\Gamma(V_n) \le E\Gamma(8\pi\Sigma_{1\le i,j\le n; i\ne j}\mathcal{N}_i\mathcal{N}'_j\pi_2f(\xi_i,\eta_j)).$$

We set now

$$W_n = \sum_{1 \le i, j \le n; i \ne j} \pi_2 f(\xi_i, \eta_j)^2.$$
(10)

In law, $(\Sigma_{1 \leq i,j \leq n; i \neq j} \mathcal{N}_i \mathcal{N}'_j \pi_2 f(\xi_i, \eta_j))^2 = \mathcal{N}^2 \Sigma_i (\Sigma_j \pi_2 f(\xi_i, \eta_j) \mathcal{N}_j)^2)$ or $\mathcal{N}^2 \Sigma_k \lambda_k \mathcal{N}_k^2$, where $\Sigma_k \lambda_k = W_n$ with $\lambda_k \geq 0$, thus, by convexity:

Lemma 1. For every $k \in \mathbb{N}$, we have

$$EV_n^{2k} \le (8\pi)^{2k} (E\mathcal{N}^{2k})^2 EW_n^k.$$

2: Bounds for functions

k is a natural integer. The current indexes i,j of the sample belongs to [1,n]. The current s belongs to G, other current indexes as λ, μ, \dots belong to D. We have $|f| \leq \sqrt{h}$ where

$$h(x,y) = \sum_{\lambda,\mu} a_{\lambda,\mu} I_{\lambda}(x) I_{\mu}(y).$$
(11)

Thus $|\pi_2 f(x,y)| \leq \int \int (\sqrt{h}(x,y) + \sqrt{h}(x,t) + \sqrt{h}(z,y) + \sqrt{h}(z,t))F(dz)F(dt)$ and finally

Thus by convexity

Lemma 2. For h defined in (11) and natural integer k we have

$$EW_n^k \le (16)^k E((\sum_{1 \le i \ne j \le n} h(\xi_i, \eta_j))^k).$$

$$\tag{12}$$

3: Bounds for moments

We define the numbers of visits of I_{λ} by each of the two samples as

$$X_{\lambda} = \sum_{1 \le i \le n} I_{\lambda}(\xi_i) \quad \text{and} \quad Y_{\lambda} = \sum_{1 \le i \le n} I_{\lambda}(\eta_i) \cdot$$
(13)

Let τ be the current permutation of G and π_{τ} be Sup $\lambda a_{\lambda,\tau(\lambda)}$. We have obviously

$$\sum_{1 \le i \ne j \le n} h(\xi_i, \eta_j) \le \sum_{\tau} \pi_{\tau}(\sum_{\lambda} X_{\lambda} Y_{\tau(\lambda)})$$

As $\Sigma_{\tau} \pi_{\tau} = 1$, by convexity again and (13) we obtain

$$EW_n^k \leq (16)^k \operatorname{Sup}_{\tau} E((\Sigma_{\lambda} X_{\lambda} Y_{\tau(\lambda)})^k)$$

Appendix 1 contains the proof of the main technical result, namely:

Lemma 3. With previous notations, for every τ , for every integer $k \geq 1$, we have

$$E(\Sigma_{\lambda} X_{\lambda} Y_{\tau(\lambda)})^k \le 6^k \operatorname{Max} (n^2 v, (n^2 v)^k) (E\mathcal{N}^{2k})^2.$$

Collecting the previous bounds, proof of Theorem 1 is achieved.

4: Proof of Corollary 1

A) We assume that $n^2 v \ge 1$. Let X be $U_n^{(2)}(f)/C ||f||_D n \sqrt{v}$. Appendix 2 contains the proof of the quite obvious

Lemma 4. If for every natural integer k we have $\mathbb{E}X^{2k} \leq (\mathbb{E}\mathcal{N}^{2k})^4$, then

$$\mathbb{P}(|X| \ge x) \le \exp\left(6 - 2\sqrt{x}\right).$$

B) Now, if $n^2 v \leq 1$, let Y be $U_n^{(2)}(f)/C||f||_D$. For the same reason we have

$$\mathbb{P}(|Y| \ge x) \le \exp \left(6 - 2\sqrt{x}\right).$$

This achieves the proof.

5: Proof of Theorem 2

Let Δ be some positive integer and D be the partition $(I_{\lambda,D} = (\lambda/\Delta, (\lambda+1)/\Delta(|\lambda \in \mathbb{Z})))$. In what follows indexes λ, u, v, s belong to \mathbb{Z} .

Let $p_{\lambda,D}$ be $P(\xi \in I_{\lambda,D})$ and p_D be the density $p_D(x) = \sum \Delta p_{\lambda,D} I_{\lambda,D}$. Expectation with respect to p_D is denoted E_D . We have

 $\lim_{\Delta \to \infty} \|p_D - p\|_2 = 0, \ \Delta v_D = \|p_D\|_2^2 \text{ and thus } \lim_{\Delta \to \infty} \Delta v_D = \|p\|_2^2.$

We set $\Phi(x) = \sum_{u \in \mathbb{Z}} \gamma_u(x-u) \mathbb{1}_{u \le x < u+1}$, where the support of γ_u is included in [0,1[(thus defining the γ_u 's); we have $\|\gamma_u\|_{\infty} \le \omega_u$. As $\|\Phi\|_2 = 1$, there exists some u_o with $\|\gamma_{u_o}\|^2 > 0$. We set $f_D(x, y) = \sum_{\lambda} \Phi(\Delta x + \lambda) \Phi(\Delta y + \lambda)$.

a) Bounds for $Ef_D^2(\xi, \eta)$ and $||f_D||_D$:

We begin by bounding from below the quantity $Ef_D^2(\xi, \eta)$.

We have $Ef_D^2(\xi,\eta) = \sum_{\lambda,\mu} (E\Phi(\Delta\xi + \lambda)\Phi(\Delta\eta + \mu))^2 \geq \sum_{\lambda} (E\Phi^2(\Delta x + \lambda))^2$, thus $Ef_D^2(\xi,\eta) \geq \sum_{\lambda} (E\gamma_{u_o}^2(\Delta\xi + \lambda - u_o)I_{u_o-\lambda,D})^2$. A classical computation gives $(EgI_{\mu,D})^2 - (E_DgI_{\mu,D})^2 \geq -2 ||gI_{\mu,D}||_{\infty} ||gI_{\mu,D}||_2 p_{\mu,D} (\int_{I_{\mu,D}} (p - p_D)^2 dx)^{1/2}$ then $Ef_D^2(\xi,\eta) \geq \sum_{\lambda} (E_D\gamma_{u_o}^2(\Delta\xi + \lambda - u_o)I_{u_o-\lambda,D})^2 - 2\omega_{u_o} (||\gamma_{u_o}^2||_2/\sqrt{\Delta})\sqrt{v_D} ||p - p_D||_2$. As $(E_D\gamma_{u_o}^2(\Delta\xi + \lambda - u_o)I_{u_o-\lambda,D})^2 = p_{u_o-\lambda,D}^2 ||\gamma_{u_o}||_2^4$, we get

$$\Delta E f_D^2(\xi, \eta) \ge \Delta v_D \|\gamma_{u_o}\|_2^4 - 2\omega_{u_o}^3 \|p - p_D\|_2 \sqrt{\Delta v_D}$$
 and

$$\lim \inf_{\Delta \to \infty} \Delta E f_D^2(\xi, \eta) \ge \|p_2\|^2 a(\Phi) := \|p\|_2^2 \|\gamma_{u_o}\|_2^4 > 0.$$
(15)

On the other hand, $|f_D(x,y)| \leq g(x,y) := \Sigma_{\lambda} |\Phi(\Delta x + \lambda) \Phi(\Delta y + \lambda)|$. Using (H), we have $g(x,y) \leq \Sigma_{\lambda,u,v} \omega_{u+\lambda} \omega_{v+\lambda} I_u(x) I_v(y)$.

Setting $\sqrt{a_s} = \sum_u \omega_u \omega_{u+s}$, $g(x, y) \leq \sum_{\lambda, s} \sqrt{a_s} I_\lambda(x) I_{\lambda+s}(y)$, then

$$g^{2}(x,y) \leq \Sigma_{s} a_{s}(\Sigma_{\lambda} I_{\lambda}(x) I_{\lambda+s}(y)) \cdot$$
(16)

But $\Sigma_s a_s = \Sigma_{u,v,s} \omega_u \omega_v \omega_{u+s} \omega_{v+s} = M(\Phi)^2 \ge \Sigma_{u,v} \omega_u \omega_v \omega_u \omega_v = (\Sigma_u \omega_u^2)^2 \ge 1$ because $\|\Phi\|_2 = 1$. Taking for G the collection of $\lambda :\to \lambda + s$, we get

$$||f_D||_D^2 \le M^2(\Phi) \text{ and } M^2(\Phi) \ge 1.$$
(17)

We recall the obvious upper bound

$$Ef_D^2 \le v_D M^2(\Phi) \le \|p\|_2^2 M^2(\Phi) / \Delta \cdot$$
 (18)

b) Bounds for $||f_D||_{\infty}$:

Obviously, $||f_D||_{\infty}$ does not depend on D, and is less than $M(\Phi)$:

There exists some
$$b(\Phi)$$
 with $0 < b(\Phi) = ||f_D||_{\infty} \le M(\Phi)$. (19)

d) Bounds for $\|\pi_2 f_D\|_2$ and $\|\pi_2 f_D\|_{\infty}$:

As for Δ large, $\sup_{\lambda} p_{\lambda,D} = o(1/\sqrt{\Delta})$ we have $|E\Phi(\Delta\eta + \lambda)| = o(1/\sqrt{\Delta})$. Thus $||Ef(x,\eta)||_{\infty} := \sup_{x} |Ef(x,\eta)| \to 0$. We have $\pi_2 f_D(x,y) = f_D(x,y) - Ef(x,\eta) - Ef(\xi,y) + Ef(\xi,\eta)$, and asymptotically we have

$$\lim_{\Delta \to \infty} \|\pi_2 f_D\|_{\infty} = b(\Phi) \text{ (and obviously by (19) } \|\pi_2 f_D\|_{\infty} \le 4M(\Phi).$$
(20)

We have $E(f_D)^2 \ge Var(\pi_2 f_D) \ge \Sigma_{\lambda} (Var\Phi^2(\Delta\xi + \lambda))^2 = \Sigma_{\lambda} (E(\Phi^2(\Delta\xi + \lambda) - E(\Phi(\Delta\xi + \lambda)^2)^+)^2)$. Using the fact that $\Delta E(\Phi(\Delta\xi + \lambda)^2)$ goes uniformly to 0 and (15), we get

$$a(\Phi) \|p\|_2^2 \le \lim \inf_{\Delta \to \infty} \Delta Var(\pi_2 f_D)$$
(21)

$$\Delta Var(\pi_2 f_D) \le \|p\|_2^2 M^2(\Phi).$$

e) Proof of Theorem 2

We set now $\Delta = 2^{\ell}$. Using $nC/\sqrt{n/(n-1)} \leq 2C$, $n^2v_D \geq 1$, $||f_D||_D \leq M(\Phi)$ and $\Delta v_D \leq ||p||_2^2$, Corollary 1 gives the exponential upper bound.

Appendix 1: Proof of Lemma 3

In this appendix where partition D is fixed, we use notation

$$v_D = v, \ p_{\lambda} = P(\xi \in I_{\lambda}), \ \eta_{\lambda} = \mathbb{E}(X_{\lambda}) = np_{\lambda} \text{ thus } \Sigma \eta_{\lambda}^2 = n^2 v \cdot$$

We recall that τ is some permutation of D, and that λ is the current point of D.

We consider two laws on $\mathbb{N}^D \times \mathbb{N}^D$, the current point of which is (\mathbf{X}, \mathbf{Y}) . In all cases, \mathbf{X} and \mathbf{Y} are independent with the same law.

In the first case, the law of **X** is $\mathcal{M}(n, \mathbf{p})$ $(n \geq 2)$, the Multinomial where $\mathbf{p} = (p_{\lambda})$, with associated expectation E.

In the second one, the X_{λ} are independent, with Poisson law, and mean value $E_{\eta}(\mathbf{X}) = \eta$, where $\eta = (\eta_{\lambda}) = n\mathbf{p}$. The associated expectation is \mathbb{E}_{η} .

We consider the mapping U from $\mathbb{N}^D \times \mathbb{N}^D$ to \mathbb{R} defined by $U(\mathbf{X}, \mathbf{Y}) = \Sigma_\lambda X_\lambda Y_{\sigma(\lambda)}$. We will first prove that for every positive integer $k \ E(U^k)$ is less than $E_\eta(U^k)$ and then furnish an upper bound for this moment.

1: Reduction to the Poisson case

In what follows, E_{μ} denotes the expectation associated to the Poisson law with parameter μ . $X^{[k]}$ is the Polynomial $X(X-1) \cdot (X-k+1)$, for which $E_{\mu}X^{[k]} = \mu^k$.

Definitions: A mapping ψ from N to N is strongly positive if

$$\psi(X) = \sum_{k} a_k X^{[k]}, \text{ with } a_k \ge 0 \text{ for every } k.$$
(d1)

A mapping Ψ from \mathbb{N}^D to \mathbb{N} is strongly positive if there exist some enumerable I, a family $(\psi_{\lambda,i} \mid \lambda; i \in I)$, a family $(a_i \mid i \in I)$, where each $\psi_{\lambda,i}$ is strongly positive and each a_i is positive, such that

$$\Psi(\mathbf{X}) = \sum_{i} a_{i} \Pi_{\lambda} \psi_{\lambda,i}(X_{\lambda}) \cdot \tag{d2}$$

Lemma 5. If Ψ is strongly positive, then

$$E(\Psi(\mathbf{X})) \le E_{\eta}(\Psi(\mathbf{X})). \tag{a1}$$

For every k, X^k and $\{X(X-1)\}^k$ are strongly positive. Moreover for k > 0, we have

$$E_{\mu} \{ X(X-1) \}^k \le \operatorname{Max} (\mu^{2k}, \mu^2) \times (E\mathcal{N}^{2k})^2.$$
 (a2)

Remark: The upper bound in (a2) is increasing of k and μ ($k \ge 1$ and $\mu > 0$).

Proof of (a1): By d2, it suffices to prove the formula when $\Psi(\mathbf{X}) = \Pi X_{\lambda}^{[k_{\lambda}]}$. For such a Ψ , $E\Psi(\mathbf{X}) = 0$ if $\Sigma_{\lambda}k_{\lambda} > n$, and $(n!/(n - \Sigma_{\lambda}k_{\lambda})!)\Pi_{\lambda}p_{\lambda}^{k_{\lambda}}$ else, obviously less than $\Pi_{\lambda}(np_{\lambda})^{k_{\lambda}} = E_{\eta}\psi(\mathbf{X})$.

Proof of (a2): The fact that X^k is strongly positive (in our sense) is well-known.

Let T_k be $\{X(X-1)\}^k$. T_1 is $X^{[2]}$ and $E_{\mu}T_1 = \mu^2$. Assume k > 1; with x = X - 2, $T_k = X^{[2]}\{(x+2)(x+1)\}^{k-1}$. But $\{(x+2)(x+1)\}^{k-1}$ is polynomial with respect to x, with positive coefficients, thus strongly positive

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with respect to x: $\{(x+2)(x+1)\}^{k-1} = \sum_{0 \le j \le 2k-2} \gamma_{j,k} x^{[j]}$; finally $X^{[2]} x^{[j]} = X^{[j+2]}$ and T^k is strongly positive. Moreover, $E_{\mu}T_k = E_{\mu}\sum_{0 \le j \le 2k-2} \gamma_{j,k} X^{[2+j]} = \mu^2 \sum_{0 \le j \le 2k-2} \gamma_{j,k} EX^{[j]} = \mu^2 E_{\mu} \{(X+2)(X+1)\}^{k-1}$. Let g_k be $g_k(X) = T_k(X) \mathbb{1}_{X>1}$. g_k is convex, and if X is a Poisson r.v, almost surely $g_k(X) = T_k$. For k = 1, $E_{\mu}g_1(X) = \mu^2$; let k be > 1. We have obtained $E_{\mu} \{X(X-1)\}^k = \mu^2 E_{\mu}g_{k-1}(X+2)$. Let Y be independent of X, Poisson with parameter 2; by Jensen, conditionally on X = x, $g_{k-1}(x+2) \le E_2 g_{k-1}(x+Y)$, thus, as the law of X + Y is Poisson with parameter $2 + \mu$, we get $E_{\mu} \{X(X-1)\}^k \le \mu^2 E_{\mu+2} g_{k-1}(X)$, thus, recursively

if
$$k > 0$$
, then $E_{\mu} \{ X(X-1) \}^k \le [\mu(\mu+2)\cdots(\mu+2k-2)]^2$

The product $\mu(\mu+2)\cdots(\mu+2k-2)$ is bounded by Max $(\mu^k,\mu)\times(1\cdot3\cdot5\cdots(2k-1)) = Max (\mu^k,\mu)\times E\mathcal{N}^{2k}$ and the proof is achieved for a2.

Now we return to the proof. U^k being a sum with positive coefficients of products of powers of the almost surely positive X's and the Y's is obviously strongly positive with respect to the X's and Y's; by independence and (a1), we obtain for every $k \ge 0$:

$$EU^k \le E_\eta U^k \,. \tag{a3}$$

2: The Poisson case

For every pair x, y of natural integers, we have easily

$$xy \le x(x-1) + y(y-1) + \mathbf{1}_{x=1} \times \mathbf{1}_{y=1}$$
 (a4)

Let us define now

$$Z = \Sigma_{\lambda} X_{\lambda} (X_{\lambda} - 1), \ Z' = \Sigma_{\lambda} Y_{\lambda} (Y_{\lambda} - 1)$$

$$T = \Sigma_{\lambda} \mathbf{1}_{X_{\lambda} = 1} \times \mathbf{1}_{Y_{\tau}(\lambda)} = 1$$
(d3)

Using the fact that τ is a permutation, by a4 we have $U \leq Z + Z' + T$, then, as the laws of Z and Z' are the same

$$E_{\eta}U^{k} \leq 3^{k} \operatorname{Max}\left(E_{\eta}Z^{k}, E_{\eta}T^{k}\right) \cdot \tag{a5}$$

3: Bound for the first term

As $E_n Z = n^2 v$, we assume that k > 1. We set

$$E(k,\mu) = Max \ (\mu^{2k},\mu^2) (E\mathcal{N}^{2k})^2 \ if \ k > 0 \ and \ 1 \ else$$

By a2, we have

$$E_{\eta}Z^{k} \leq \Sigma_{k_{\lambda} \geq 0; \Sigma_{\lambda}k_{\lambda} = k} \{k! / \Pi_{\lambda}k_{\lambda}!\} \Pi_{\lambda}E(k_{\lambda}, \eta_{\lambda}) \cdot$$
(a6)

First case: If for each λ , $\eta_{\lambda} \geq 1$, then, as $\Pi_{\lambda} E \mathcal{N}^{2k_{\lambda}} \leq E \mathcal{N}^{2\Sigma_{\lambda}k_{\lambda}}$, we have

$$E_{\eta}Z^{k} \leq (E\mathcal{N}^{2k})^{2} \Sigma_{k_{\lambda} \geq 0; \Sigma_{\lambda}k_{\lambda} = k} \{k! / \Pi_{\lambda}k_{\lambda}!\} \eta_{\lambda}^{2k_{\lambda}}$$

$$\tag{5}$$

thus, in the first case, for every $k \ge 0$

$$E_{\eta}Z^k \le (n^2 v)^k (E\mathcal{N}^{2k})^2 \cdot \tag{A1}$$

Second case: For each λ , $\eta_{\lambda} \leq 1$:

Let A be a non-void subset of $[0, D[\cap \mathbb{Z}, \text{ and } M(A, k)]$ be the subset of \mathbb{N}^A given by $(k_i \mid i \in A; k_i > 0$ for each $i \in A$; $\Sigma_{i \in A} k_i = k$). We set

$$S(A,k) = \sum_{M(A,k)} \{k! / \prod_i k_i! \} \prod_i E(k_i, \eta_i) \cdot$$

The general term of S(A, k) is $\{k!/\Pi_i k_i!\}\Pi_i \eta_i^2 \Pi_i E(\mathcal{N}^{2k_i})^2$.

Let ν_j be $(E\mathcal{N}^{2j})^2/j!$ $(j \in \mathbb{N})$. Elementary computation gives

If
$$1 \le j \le k$$
, then $\nu_j \nu_k \le \nu_{j-1} \nu_{k+1}$. (a7)

Thus the general term of S(A, k) is bounded by $\prod_i \eta_i^2 E(\mathcal{N}^{2k})^2$ (obtained outside of M(A, k), when all k_i are 0 except one).

On the other hand, it is well-known that $|M(A,k)| = \binom{k-1}{a-1}$, where a = |A|. Finally

$$E_{\eta}(Z^k) \leq \Sigma_{\text{non void}} S(A,k) \leq (E(\mathcal{N}^{2k})^2 \times \{\Sigma_{a>0} \binom{k-1}{a-1} (\Sigma_{|A|=a} \prod_{i \in A} \eta_i^2)\}$$

As $\Sigma_{|A|=a} \prod_{i \in A} \eta_i^2 \leq (n^2 v)^a$ and $\Sigma_{a>0} \binom{k-1}{a-1} = 2^{k-1}$, we get in the second case, for any $k \geq 1$

$$E_{\eta}Z^k \le 2^{k-1} \operatorname{Max} (n^2 v, (n^2 v)^k) (E\mathcal{N}^{2k})^2$$
 (A2)

General case: We divide [0, D[into two (non void) subsets: $\Delta_1 = (\lambda \mid \eta_{\lambda} < 1)$ and $\Delta_2 = (\lambda \mid \eta_{\lambda} \ge 1)$, and set $v_i = \sum_{\lambda \in \Delta_i} \eta_{\lambda}^2$. Using (A1, A2), we obtain

$$E(Z^k) \le (E\mathcal{N}^{2k})^2 (n^2 v_2)^k + \sum_{j>0} {k \choose j} 2^{j-1} (E\mathcal{N}^{2j})^2 (E\mathcal{N}^{2k-2j})^2 \operatorname{Max} (n^2 v_1, (n^2 v_1)^j) (n^2 v_2)^{k-j}.$$

The latter bound is increasing of v_i , each bounded by v. Thus finally, in any case, for $k \ge 1$

$$E_{\eta}Z^k \le 2^k \operatorname{Max} (n^2 v, (n^2 v)^k) (E\mathcal{N}^{2k})^2$$
 (A3)

4: Bound for the second term

We can bound $E_{\eta} \mathbb{1}_{X_{\lambda}=1} \times \mathbb{1}_{Y_{\tau(\lambda)}=1}$ by $b_{\lambda}^2 := \eta_{\lambda} \eta_{\tau(\lambda)}$. We take notations of Second case of previous paragraph. Setting now

$$S'(A,k) = \sum_{M(A,k)} \{k! / \prod_i k_i! \} \prod_i b_i^2, \text{ we have } E_\eta T^k = \sum_{A \text{ non void}} S'(A,k).$$

The current term of S'(A, k) is bounded by $k! \prod_i b_i^2$. Thus we obtain here

 $E_{\eta}T^k \leq k! 2^{k-1} \operatorname{Max}(w, w^k)$ where $w = \Sigma_{\lambda} b_{\lambda}^2$.

As by Cauchy-Schwartz $w \leq n^2 v$ and $k! \leq (E\mathcal{N}^{2k})^2$, we get again for $k \geq 1$

$$E_{\eta}T^{k} \le 2^{k} \operatorname{Max} (n^{2}v, (n^{2}v)^{k})(E\mathcal{N}^{2k})^{2} \cdot$$
 (A4)

Using (a3, a5, A3) and (A4), the proof is finished.

Appendix2: Proof of Lemma 4

For $k \in \mathbb{N}$, we set $u_k = e^k E \mathcal{N}^{2k} (2k+1)^{-k}$, $r_k = u_{k+1}/u_k = e\{(2k+1)/(2k+3)\}^{k+1}$ and finally $\varphi(x) = 1 + (x+1) \log((2x+1)/(2x+3))$ for $x \ge 0$, then $\varphi(k) = \log(r_k)$. We have

$$\varphi'(x) = \operatorname{Log}\left((2x+1)/(2x+3)\right) + 1/(2x+1) + 1/(2x+3).$$

$$\varphi''(x)/2 = 1/(2x+1) - 1/(2x+3) - 1/(2x+1)^2 - 1/(2x+3)^2 \le 0.$$

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As φ' goes to 0 when x goes to ∞ , we have $\varphi' \ge 0$. As φ goes to 0 when x goes to ∞ , $\varphi \le 0$. Thus, for k > 0, $u_{k+1} \le u_k \le u_0 = 1$: we have proved that

for every
$$k \in \mathbb{N}$$
, $E\mathcal{N}^{2k} \le e^{-k}(2k+1)^k$.
every $k \in \mathbb{N}$, $E\mathcal{N}^{2k} \le (E\mathcal{N}^{2k})^4$. If $2k+3 \ge \sqrt{r} \ge 2k+1$, via Markov's inequality and

Let us assume that for every $k \in \mathbb{N}$, $EX^{2k} \leq (E\mathcal{N}^{2k})^4$. If $2k+3 \geq \sqrt{x} \geq 2k+1$, via Markov's inequality and assertion *, $P(|X| \geq x) \leq e^{-4k} \leq e^{-2\sqrt{x}+6}$. Then the result is proved for $x \geq 1$ and obvious for $0 \leq x \leq 1$.

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