

RESEARCH ARTICLE

Open Access



Cost-effectiveness modelling to optimise active screening strategy for *gambiense* human African trypanosomiasis in endemic areas of the Democratic Republic of Congo

Christopher N. Davis^{1,2*} , Kat S. Rock^{1,2}, Marina Antillón^{3,4}, Erick Mwamba Miaka⁵ and Matt J. Keeling^{1,2,6}

Abstract

Background: *Gambiense* human African trypanosomiasis (gHAT) has been brought under control recently with village-based active screening playing a major role in case reduction. In the approach to elimination, we investigate how to optimise active screening in villages in the Democratic Republic of Congo, such that the expenses of screening programmes can be efficiently allocated whilst continuing to avert morbidity and mortality.

Methods: We implement a cost-effectiveness analysis using a stochastic gHAT infection model for a range of active screening strategies and, in conjunction with a cost model, we calculate the net monetary benefit (NMB) of each strategy. We focus on the high-endemicity health zone of Kwamouth in the Democratic Republic of Congo.

Results: High-coverage active screening strategies, occurring approximately annually, attain the highest NMB. For realistic screening at 55% coverage, annual screening is cost-effective at very low willingness-to-pay thresholds (\$20.4 per disability adjusted life year (DALY) averted), only marginally higher than biennial screening (\$14.6 per DALY averted). We find that, for strategies stopping after 1, 2 or 3 years of zero case reporting, the expected cost-benefits are very similar.

Conclusions: We highlight the current recommended strategy—annual screening with three years of zero case reporting before stopping active screening—is likely cost-effective, in addition to providing valuable information on whether transmission has been interrupted.

Keywords: African trypanosomiasis, African sleeping sickness, Mathematical model, Cost-effectiveness

*Correspondence: c.davis.7@warwick.ac.uk

¹Mathematics Institute, University of Warwick, CV4 7AL Coventry, UK

²Zeeman Institute (SBIDER), University of Warwick, CV4 7AL Coventry, UK

Full list of author information is available at the end of the article



© The Author(s). 2021 **Open Access** This article is licensed under a Creative Commons Attribution 4.0 International License, which permits use, sharing, adaptation, distribution and reproduction in any medium or format, as long as you give appropriate credit to the original author(s) and the source, provide a link to the Creative Commons licence, and indicate if changes were made. The images or other third party material in this article are included in the article's Creative Commons licence, unless indicated otherwise in a credit line to the material. If material is not included in the article's Creative Commons licence and your intended use is not permitted by statutory regulation or exceeds the permitted use, you will need to obtain permission directly from the copyright holder. To view a copy of this licence, visit <http://creativecommons.org/licenses/by/4.0/>. The Creative Commons Public Domain Dedication waiver (<http://creativecommons.org/publicdomain/zero/1.0/>) applies to the data made available in this article, unless otherwise stated in a credit line to the data.

Background

Despite the continued decline in the annual number of reported cases of *gambiense* human African trypanosomiasis (gHAT), accounting for fewer than 1000 new cases reported in 2019 [1, 2], the disease persists in many of the historically endemic sites in Western and Central Africa. This vector-borne disease, transmitted by a bite from a tsetse infected with the parasite *Trypanosoma brucei gambiense*, is typically—although not always—fatal when untreated [3]. Human African trypanosomiasis (HAT), which includes both *gambiense* and *rhodesiense* forms, caused an estimated 1360 deaths in 2019 and approximately 82,615 disability-adjusted life years (DALYs), of which the majority were caused by gHAT [4]. *Gambiense* HAT has been targeted for elimination by the World Health Organization (WHO); first, for elimination as a public health problem by 2020 and then for elimination of transmission (EOT) by 2030 [2, 5]. To achieve these targets, there are several recommended strategies to reduce the transmission and burden of the infection, which are constituted primarily of the medical interventions of active screening and passive surveillance.

Passive surveillance depends on the ability of fixed health centres to test for the infection and carry out treatment on self-presenting individuals, typically upon the onset of symptoms [6]. Screening and treating infected individuals both allows the infected people to be saved from a potentially fatal disease, but it also prevents further spread of infection via tsetse.

Traditionally, the most effective form of controlling gHAT infection, however, has been active screening and treatment [7–9]. Active screening is carried out by mobile teams that travel to villages in focal disease regions and target the screening of the whole population for gHAT; those determined to have the infection can then be treated at the closest health facility offering treatment. The initial screening test is typically a serological test for the presence of the antibody called the card agglutination test for trypanosomiasis (CATT) [10], although recently rapid diagnostic tests (RDTs) have also been utilised for screening [11–13]. Confirmation of the infection is then carried out via microscopic examination; traditionally, this is followed by staging of the disease, which consists of a lumbar puncture to determine whether the parasite has infected the central nervous system—considered the second stage of disease [14]. However, the recently approved drug, fexinidazole [15], can be used to treat both disease stages (except for patients with neuro-psychiatric symptoms and signs suspicious of advanced second stage disease) and so may remove the need for lumbar puncture in most cases, although retaining the requirement of parasitological confirmation [16].

Active screening has been very effective in reducing case numbers and still plays an important role in maintaining surveillance and treatments where access is problematic, yet it is an expensive intervention in terms of both time and money [11, 17]. As local elimination of gHAT occurs in focal areas, active screening will likely be scaled back and gHAT testing will become better integrated into fixed health facilities, as resources can be reallocated and it becomes unnecessary to screen entire village populations for the infection [18]. In this situation, reactive screening can be implemented, whereby after a number of successive active screenings in which no cases are detected, the screening stops unless a new case is passively reported, upon which a 'reactive' screen would occur [12]. Several active screening strategies have been proposed, including a recommendation of three repeated screening rounds with 1-year [19] or 6-month intervals [20]. WHO guidelines currently recommend annual screening for three years of zero case reporting before stopping in previously endemic villages [5].

Mathematical models of gHAT have been used for the prediction of future case numbers and evaluation of a range of plausible control strategies [21–29]. However, these have typically considered the infection dynamics and the impact of interventions without accounting for the costs of implementing such strategies. Here, we explicitly use a stochastic model of gHAT infection in a village population, developed in Davis et al. [26], to simulate different plausible active screening programmes alongside passive surveillance, allowing us to quantify the relative costs of implementation as well as the health effects compared to a baseline of passive surveillance (the comparator strategy). We use parameters matched to screening and incidence data from the health zone Kwamouth, in M'indombe province of the Democratic Republic of Congo (DRC) (formerly in Bandundu province). Kwamouth is in a historically high-endemicity gHAT area of the DRC, the country that contributes 70% of all global gHAT cases in 2019 [2]. We also present results from a moderate-endemicity health zone, Mosango (see Additional file 1: Figure S3) [5, 11, 22, 24, 26, 30–59].

The costs of gHAT interventions have been previously been evaluated [11, 46, 47, 51] and the different strategies have been considered for large populations [54, 60]. We consider the effect of active screening on individual villages in the drive for EOT, by determining how active screening can be best implemented to achieve this goal whilst providing value for money.

Methods

Mathematical modelling

To capture the effects of different active screening strategies and the underlying infection dynamics on

a village population, we use a stochastic compartmental model [26]. The stochasticity incorporates the chance events involved in infection transmission into the mechanistic model and is better suited than a deterministic model for the purpose of optimising active screening for a village population. This is because the population, and hence the number of people infected, is small and so, for the pre-elimination setting, local elimination of gHAT can be substantially affected by chance events, and the probability of those chance events. A deterministic formulation, which captures average dynamics, would be less suitable to capture the chance events associated with village level elimination of infection.

In the model, individuals in the human population are classified as either high-risk or low-risk [22], whereby the high-risk population—a small minority, which has been previously estimated to be 9.8% in the study health zone of Kwamouth [30]—have a higher exposure to tsetse and do not participate in active screening. The model structure is supported by anecdotal evidence that there is a fraction of the population, typically working males that work by the rivers, which is the habitat of tsetse [61], and who are absent for active screening in the villages [32]. Furthermore, previous modelling work indicates that humans have heterogeneous exposure to tsetse [22, 25, 62]. Rock et al. [22] fitted several risk structures for humans and, using the deviance information criterion (DIC), the model with the risk structure given here best matched to data on screening and incidence from the WHO HAT Atlas [40, 41].

The model classifies a person's infection status as susceptible S_H , exposed (or latent) E_H , stage 1 infection I_{1H} , stage 2 infection I_{2H} and hospitalised (and temporarily removed) R_H . On exposure to the parasite and upon the bite of an infected tsetse, a person will progress through these infection stages, unless detected in active screening and so treated and moved directly to the hospitalised class. Stage 2 infection is defined as the time when trypanosomes have crossed the blood–brain barrier [63]. There is an additional rate defined as the time when people change infection status from stage 1 infection to hospitalisation to simulate people being treated through passive surveillance.

In addition, tsetse are explicitly modelled as the proportion of flies that are in the states of teneral (unfed and more susceptible to infection than fed flies [64]) S_V , non-teneral yet uninfected G_V , exposed (or latent) E_V and infected I_V . The dynamics of the tsetse are modelled with proportions using ordinary differential equations, since the exact number of tsetse that interact with a given population is difficult to determine (although relative tsetse abundance is being mapped in some areas [65]).

However, the effective density ratio, the product of the number of tsetse in a population per human and the probability of human infection per single infective bite, can be inferred by model fitting to the WHO HAT Atlas data [40, 41]. A full description of the mathematical equations in the infection model and the parameters used (taken from Crump et al. [30]), as the median of the distributions inferred using Markov chain Monte Carlo (MCMC) methodology and the aggregate annual data from the WHO HAT Atlas in Kwamouth) can be found in Additional file 1: Table S1 and Table S2.

Simulating active screening

In villages targeted for active screening, we consider how the infection dynamics are affected by: the screening coverage c ; the screening interval t ; active zero-detections z_a ; reactive zero-detections z_r (see Table 1). This is modelled by taking all combinations of c, t, z_a and z_r for the values $c = 0, 1, 5, 10, 15, 20, 25, 30, 35, 40, 45, 50, 55, 60, 65, 70, 75, 80, 85, 90$, $t = 0.25, 0.33, 0.5, 0.67, 0.75, 1, 1.25, 1.5, 1.75, 2, 2.5, 3, 3.5, 4, 4.5, 5$ years, $z_a = 1, 2, 3, 4, 5$ screenings, and $z_r = 1, 2, 3$ screenings. The risk structure of the model means that the screening coverage is assumed to have a maximum of 90%, since only the low-risk human population participate (randomly) in active screening.

The specified screening procedure is then implemented in the model by using the screening coverage to randomly select a proportion c of the population from the low-risk sub-population, which move into the hospitalised class if they are correctly identified as being exposed to the infection after each screening interval. This process is stopped after number of zero-detections equal to z_a . However, if the infection has not been eliminated after the active screening has been halted, there is still a chance that a new case can be reported by the individual attending a fixed facility to be tested for the infection. The model explicitly incorporates under-reporting of passive case detections, such that only 27% of cases undetected by active screening will be detected by passive surveillance. The value of this parameter was estimated in fitting to this model structure, and is taken as the median of the distribution inferred using MCMC methodology applied to the aggregate annual data from Kwamouth from the WHO HAT Atlas [40, 41].

Reactive screening is the result of the detection of a case in passive surveillance after active screening has been stopped (see Additional file 1: Figure S5). Thus, we restart the screening procedure as reactive screening upon identification of passive cases, stopping again after the given number of consecutive reactive zero-detections, z_r .

The WHO aims for high screening coverage in active screenings for villages in gHAT-endemic foci, but just

Table 1 Descriptions of the variables used for defining an active screening strategy

Variable	Name	Definition	Value range
c	Screening coverage	Proportion of the village population screened in a visit.	0–90%
t	Screening interval	Time between active screening visits to a village.	0.25–5 years
z_a	Active zero-detections	Number of consecutive active screenings where no cases are detected for the cessation of active screening.	1–5 screenings
z_r	Reactive zero-detections	Number of consecutive reactive screenings where no cases are detected for the cessation of reactive screening.	1–3 screenings

how high that coverage ought to be is not prescribed. Moreover, guidelines and previous modelling work suggest these screenings should continue annually until there have been three consecutive years of no new cases, followed by a further screening after three years if there are no detected cases [5, 26, 66]. Our modelling aims to provide evidence in support of this strategy or recommend how the strategy could be adapted to make more efficient use of resources.

Economic modelling

Cost-effectiveness metric

The cost of an active screening strategy is a function of several component costs: implementing the screening test, confirmation of the infection, carrying out treatments, setting up and maintaining the mobile screening teams. Moreover, active screening may impact the number of passive tests and treatments performed. In the current work, we do not consider the additional costs of passive surveillance, such as capital costs, only the costs directly affected by active screening. The costs of active screening strategies will vary depending on the type of screening test and treatment used, and also the type of mobile screening team; whilst a traditional truck team that can carry more tests and equipment, such as a generator, a motorbike team that can reach more remote villages [46, 67].

As well as considering the changes in monetary costs, we want to consider the change in the health benefit of implementing different active screening programmes; therefore, we consider the number of DALYs averted [68]. The number of DALYs are the discounted sum of the number of years of life lost (YLL) and years lived with disability (YLD), where YLL is the number of years of life lost due to premature death and YLD is the number of years of healthy years lost with a weighting for the severity of the condition [69]. Therefore, the benefit of an active screening programme is measured by averting DALYs, and we do not attempt to quantify any additional benefits of achieving elimination. We calculate the number of DALYs averted by a particular screening strategy against

a comparator consisting of passive surveillance and no active screening.

We evaluate the net monetary benefit (NMB) to assess the cost-effectiveness using a 30-year time-horizon. For each active screening strategy, the net monetary benefit (NMB) was calculated as:

$$\text{NMB} = \text{WTP} \times \text{DALYs averted} - \text{Cost of active screening strategy compared to passive surveillance only.} \quad (1)$$

The net monetary benefit is conventionally conditioned on the willingness-to-pay (WTP), or the maximum amount of money that the funder is prepared (willing) to pay to avert one additional DALY over and above the comparator, namely implementing only passive surveillance but no active screening. The change in costs is the cost of implementing the active screening strategy, including the consequential change in cost of passive surveillance, minus the cost of baseline passive surveillance. A particular strategy is more efficient than the comparator strategy at that particular WTP if the NMB is positive. In the context of uncertainty, operationalised by repeated draws of the simulation, the optimal strategy will have the highest mean NMB. Because the NMB is conditioned on the WTP, we consider a range of fixed WTP thresholds, such that decision-makers can heed recommendations according to the typical cost-effective thresholds in their programmes.

In addition to the WTP, the NMB will be affected by the population size of a village, N_H , the proportion of infections that go undetected by active screening but that are detected and treated passively p_t , and the initial level of infection in the population.

Contextualising willingness-to-pay or cost-effectiveness thresholds

The correct WTP is the cause of much debate. Typically, the WTP is taken as the product of the gross domestic product (GDP) per capita of a country and a multiplying factor. This factor, WTP_c , is traditionally set to 3 [70] for ‘cost-effective’ strategies and to 1 for ‘very cost-effective’

strategies. More recently, it has been remarked that this is too high for low-income countries [71].

The practical WTP in DRC in the past has been estimated to be between \$5–\$230 [72] per DALY, or between \$70–\$320 for Sub-Saharan African low-income countries [73]. The WHO would previously consider a strategy ‘very cost-effective’ if the WTP equals the GDP/capita (i.e. \$557 in 2018 for DRC) or as ‘cost-effective’ when a strategy was optimal at a $WTP \geq 3$ times the GDP per capita (i.e. \$1671 in 2018 in DRC).

We also note that in the context of elimination, a funder may be willing to pay more for the additional benefit of reducing the number of infections to zero, but we simply leave this choice to the funder.

In accordance with WHO recommendations, we show the optimal strategies for a range of WTP thresholds without comment, as we provide guidance, rather than prescriptions, on what the thresholds should be as that is beyond our scope as modellers [74, 75].

Economic model inputs

To evaluate the NMB, we consider all component costs and the associated health benefits of implementing an active screening strategy compared to passive surveillance alone. However, we do not consider out-of-pocket expenses, therefore framing the model from the perspective of the funder of the programme. Since we are considering the NMB of an active screening strategy, we only consider costs impacted by active screening but we do not consider costs that are fixed across all strategies, such as maintaining a fixed health centre. We note that the total treatment costs will depend on the quantity of active screening carried out and additionally on the proportion of infections that are detected passively. Fitting to the WHO data indicates that around 27% of the infections that progress to Stage 2 and are not found in active screening are detected through passive surveillance, with the remainder of infections going unreported [30]. We can make the worst-case assumption that all these reported infections are treated, but none of the unreported infections are treated and will therefore die outside the healthcare system ($p_t = 27\%$). Alternatively, we can use a higher value of the parameter p_t to assume more people are treated that are reported in passive surveillance. The parameter p_t takes values in the interval 27–100%, such that the proportion of infections being treated is between only those reported being treated and all infection being treated.

Using the formulation of the NMB given by Eq. 1, we then define the constituent parts of the equation as follows:

$$WTP = WTP_c G, \tag{2}$$

DALYs averted

$$= D_1 (1 - p_t) \sum_t (\text{Change exits from stage 2 infection}) (t) \zeta(t) \tag{3}$$

$$+ D_2 \int (\text{Change in stage 1 infections}) (t) \zeta(t) dt \tag{4}$$

$$+ D_3 \int (\text{Change in stage 2 infections}) (t) \zeta(t) dt, \tag{5}$$

Change in costs

$$= C_1 N_H (\text{Years of active screening}) \zeta(t) \tag{6}$$

$$+ C_2 N_H (\text{Active screening visits}) \zeta(t) \tag{7}$$

$$+ C_3 \sum_t (\text{Number of people in active screening}) (t) \zeta(t) \tag{8}$$

$$+ (C_4 + C_5) \sum_t (\text{Active screening true positives in Stage 1}) (t) \zeta(t) \tag{9}$$

$$+ (C_4 + C_6) \sum_t (\text{Active screening true positives in Stage 2}) (t) \zeta(t) \tag{10}$$

$$+ C_4 \sum_t (\text{Active screening false positives}) (t) \zeta(t) \tag{11}$$

$$+ C_7 \sum_t (\text{Change in stage 1 passive infections}) (t) \zeta(t) \tag{12}$$

$$+ C_8 p_t \sum_t (\text{Change in stage 2 passive infections}) (t) \zeta(t) \tag{13}$$

$$+ C_9 \sum_t (\text{Treatment of stage 1 cases}) (t) \zeta(t) \tag{14}$$

$$+ C_{10} \sum_t (\text{Treatment of stage 2 cases}) (t) \zeta(t). \tag{15}$$

The willingness-to-pay, in 2018 US dollars per DALY, is split into the GDP per capita of the DRC, G , and a multiplying factor, WTP_c as is common in the literature [70] (Eq. 2). For the health benefit of the intervention, the DALYs averted are the discounted sum of the number of years of life lost (YLL) and years lived with disability (YLD). The number of years of life lost is given by the sum

of the change in number of people that exit the Stage 2 infection class multiplied by both the proportion of these infections not treated ($1 - p_t$) and the discounted average years of life lost per death (Eq. 3). The number of years lived with disability is given by the total time spent in each infection class multiplied by the associated disability weight (Eqs. 4 and 5).

The change in costs is simply the costs incurred by implementing the active screening strategy and the effect on costs of operating passive surveillance, given as a capital cost of active screening (Eq. 6) and the recurrent cost of operating each visit (Eq. 7), the number of screening tests carried out (Eq. 8), confirmation of the infection and stage determination for the true positive (Eqs. 9 and 10) and negative confirmation for the false positives (Eq. 11), the change in testing, confirmation and stage determination for passive infections (Eqs. 12 and 13), and the treatment of detected cases (Eqs. 14 and 15).

To calculate both the health benefit costs and implementation costs we use an annual discount rate of 3%, denoted in the equation by $\zeta(t) = \exp(-\delta t)$. This is in line with standard conventions and arises through the assumption that value for society will decrease over time [42]. The time-horizon used is 30 years, which is sufficient to capture almost all the costs of the active screening programme (see Additional file 1: Figure S7). The cost parameters are fixed values composed as a sum of the cost of the test or treatment, the cost of implementation and the cost of hospitalisation [54]. These costs are also dependent on the type of treatment; we primarily consider active screening to use the CATT algorithm, whilst treatment is principally with fexinidazole with some Stage 1 and Stage 2 infections using pentamidine and NECT respectively where necessary (see Additional file 1: Table S4). We also consider the additional treatment strategies of only pentamidine and NECT, and only fexinidazole (see Additional file 1: Figure S9) and note that results for acoziborole could be computed using the accompanying app, which presents results for user inputted cost parameter values (see <https://christopherdavis.shinyapps.io/optimising-ghat-active-screening/>). We expect fexinidazole to be the standard treatment in 2020 for most gHAT patients [16], and acoziborole could replace it as a single dose cure a few years later if it passes phase III clinical trials and receives an appropriate recommendation [76] (see Additional file 1: Figure S9). Cost and health parameters are shown in Table 2 with full explanations in Additional file 1.

Results

Breakdown of costs of active screening

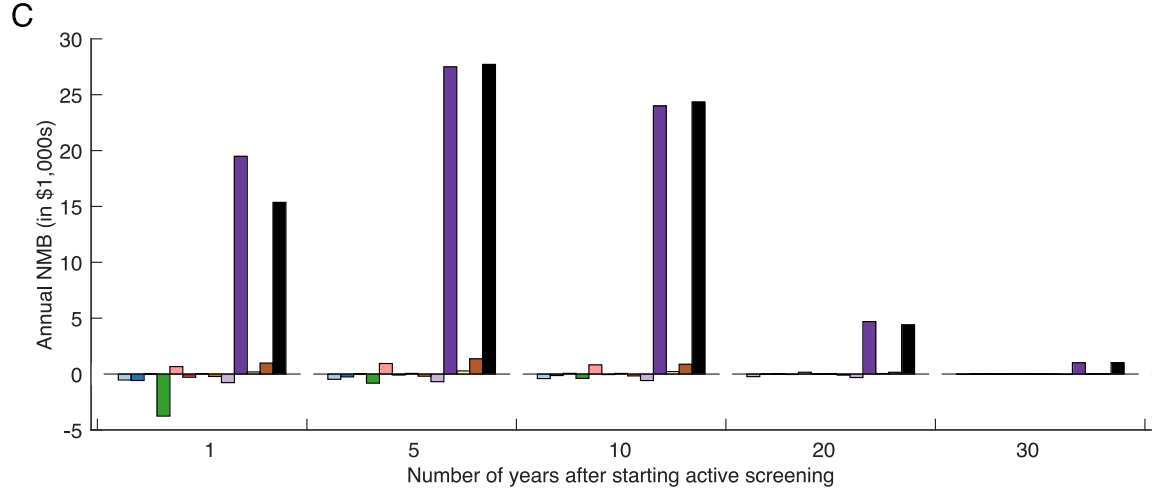
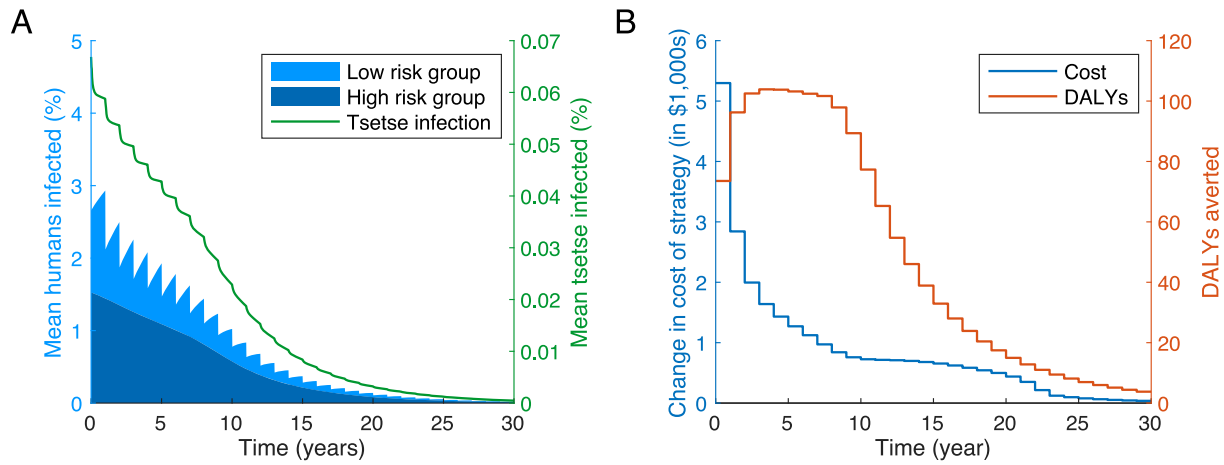
We use the stochastic compartmental model, refined from Davis et al. [26], to simulate different strategies for active screening. We first consider an active screening strategy

Table 2 Parameters for calculating the NMB

Parameter	Name	Value
WTP_c	Willingness-to-pay per DALY	Varies
G	GDP per capita for the DRC ¹	\$457.85
D_1	Discounted average years of life lost per death	21.03 years
D_2	Stage 1 disability weight	0.1330
D_3	Stage 2 disability weight	0.5432
C_1	Active screening capital cost per person	\$0.22
C_2	Active screening recurrent cost per person	\$0.77
C_3	Active screening test per person ²	\$1.03
C_4	Confirmation per person ³	\$10.96
C_5	Stage determination per person in stage 1 ³	\$1.59
C_6	Stage determination per person in stage 2 ³	\$17.21
C_7	Passive screening person in stage 1 ^{3,4}	\$14.17
C_8	Passive screening person in stage 2 ^{3,4}	\$29.79
C_9	Stage 1 treatment per person ⁵	\$85.23
C_{10}	Stage 2 treatment per person ⁶	\$561.78
N_H	Population size	Varies

¹2018 value used. ²We assume the use of the card agglutination test for trypanosomiasis (CATT) test. ³We assume confirmation is by microscopy using a blood sample, lymph node aspiration (LNA) and mini Anion Exchange Centrifugation Technique (mAECT). ⁴Passive screening includes the screening test (CATT), outpatient consultation, confirmation and stage determination if appropriate. ⁵We assume the use of both pentimidine and fexinidazole in proportions described in Additional file 1: Table S4. ⁶We assume the use of both nifurtimox-eflornithine combination therapy (NECT) and fexinidazole in proportions described in Additional file 1: Table S4. Values are given as the mean value. For more detailed calculations of these cost parameters see Additional file 1

with a typical screening coverage of 55% (see Additional file 1: Figure S2), carried out annually, and with three active zero-detections and one reactive zero-detection required for the cessation of screening ($c = 55\%$, $t = 1$ year, $z_a = 3$, and $z_r = 1$). We assume a village population of size $N_H = 1000$ starting from endemic equilibrium (as determined from the deterministic version of the model), and with 27% of infections undetected in active screening treated passively ($p_t = 27\%$). We calculate mean values of one million stochastic realisations of the process. For this strategy, the prevalence in both the human and tsetse populations rapidly decays towards zero (Fig. 1a). The annual cost of implementing this strategy also decreases with time (Fig. 1b); this is in part due to 3% discounting, the method of adjusting future costs to present-day values (which is applied to both costs and DALYs averted) [42], but also because decreasing the prevalence of infections in the population reduces the required number of treatments. Even with no active screening, infections may die out in the village due to 'stochastic fade out', when local disease extinction occurs purely by chance. There is a only a small annual decrease in costs around twelve years, since the difference in the number of infections treated



Contribution to the annual net monetary benefit (NMB) in year n (\$)

Bar	Benefit/cost description	$n = 1$	$n = 5$	$n = 10$	$n = 20$	$n = 30$
Light blue	Active screening tests	-530	-470	-403	-230	-27
Dark blue	Active Stage 1 treatment	-576	-251	-132	-14	-2
Light green	Passive Stage 1 treatment	+20	+31	+65	+18	+4
Dark green	Active Stage 2 treatment	-3762	-812	-391	-38	-4
Light red	Passive Stage 2 treatment	+670	+945	+825	+162	+35
Dark red	Active gHAT confirmation	-287	-84	-44	-6	-1
Light orange	Passive gHAT confirmation	+40	+57	+56	+12	+3
Dark orange	Active capital costs	-215	-191	-164	-93	-11
Light purple	Active recurrent costs	-767	-680	-582	-310	-34
Dark purple	Averted YLL	+19485	+27498	+23999	+4700	+1015
Light yellow	Averted Stage 1 YLD	+191	+280	+222	+50	+11
Dark yellow	Averted Stage 2 YLD	+983	+1370	+880	+158	+34
Black	Total	+15361	+27716	+24343	+4409	+1023

Fig. 1 The cost of an active screening strategy. The cost of active screening for a coverage of 55%, a screening interval of 1 year, stopping active screening after 3 screenings when no cases are detected and stopping reactive screening after 1 screening with no cases, under the assumption of $WTP_c = 0.5$. Mean values for all quantities are taken from one million stochastic simulations. **a** The number of infected people dramatically decreases with time for this coverage (total shaded blue area, with left axis) with the majority of these infections being in the high-risk group (darker blue fraction). The proportion of tsetse that are also infective is reduced with time (green line, with right axis). **b** The total change in costs of implementing a particular screening strategy (left axis) and the number of DALYs averted from the baseline of only passive surveillance (right axis). **c** The contribution to the cost from each component of the cost function for years 1, 5, 10, 20 and 30 after starting an active screening programme. Full costs are given in the table in the bottom row of the table. A population size of $N_H = 1,000$ is used. All costs are denominated in 2018 US dollars

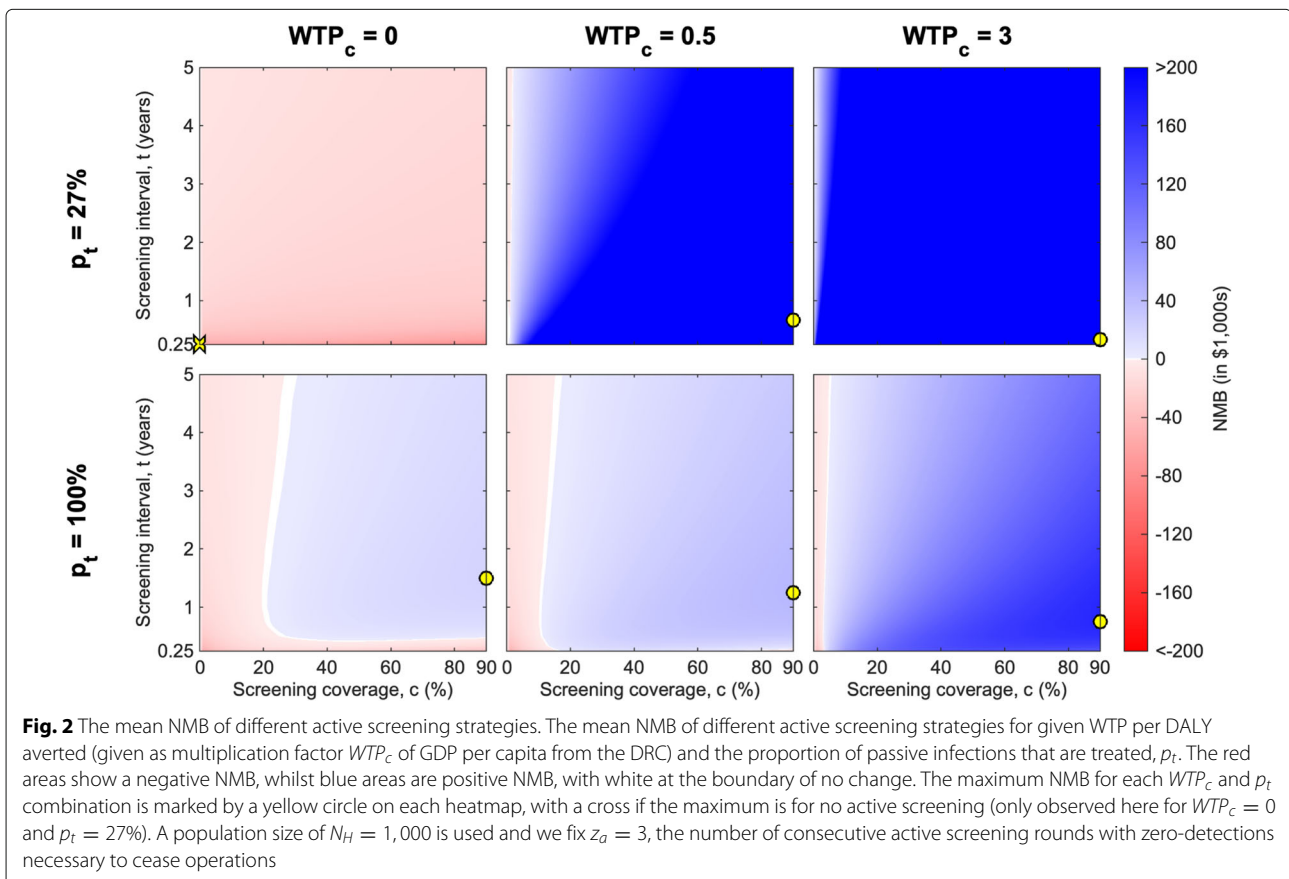
in passive surveillance is smaller in later years. However, costs decay towards zero as the probability of local gHAT elimination increases with time; whilst some recurrent costs will remain, the number of treatments will decline in time with the corresponding reduction in infections. In addition, the costs also decrease when the consecutive zero-detection threshold is reached, as the cost of the active screening is completely removed. With the assumption that $p_t = 27\%$ and $c = 55\%$, the number of DALYs averted will initially increase each year because more people are treated after detection during active screening. However, in later years, DALYs averted will decline due to fewer infections under either strategy, implying that the differential impact in the active screening strategy (versus passive surveillance alone) is most substantial in the early years of implementation because both the strategies and passive surveillance are expected to lead to elimination, albeit at a different speed (Fig. 1b).

A breakdown of the components of the NMB of implementing this active screening strategy shows that the biggest costs are the treatment from active screening, the recurrent costs of the active screening and screening populations with the CATT test (Fig. 1c). However, assuming a WTP of 50% of the GDP per capita of the DRC ($WTP_c = 0.5$ giving WTP equal to \$280.89 [77]) the monetary bene-

fit in reducing the years of life lost is dominant and the biggest factor in maximising the NMB. The total NMB (black bars) shows the full benefit of this active screening strategy is always positive with $WTP_c = 0.5$ and so, on average, this strategy is better than the comparator of no active screening. Further into the future, the NMB moves closer to zero, both because of discounting and because there is a higher probability the infection will be locally eliminated, and so active screening not required. Note that the NMB of passive surveillance is positive because the introduction of active screening and treatment means that fewer passive confirmations and treatments will need be carried out, reducing the cost. The table in Fig. 1c shows the NMB breakdown in full.

Drivers of net monetary benefit across strategies

We performed a four-way sensitivity analysis of the NMB for the WTP, the treatment coverage in passive surveillance p_t , the screening coverage c and screening interval t (Fig. 2). We considered the mean value of the NMB for one million simulations for every screening strategy considering screening coverage c and screening interval t for a large number of values (stated in Methods and Materials), but discretising WTP and p_t to three and two values respectively for the presentation here. To view how the



exact optima change with these model parameters and all the specified values for the cost and benefit parameters (including, treating WTP and p_t as continuous variables) visit the supplementary R Shiny [57] app. This app allows the sensitivity of the optimal solution to be found on any updated costings for the active screening strategies.

The interaction between WTP and p_t .

The NMB for some strategies is highly dependent on both the WTP and the proportion of passive infections treated (p_t), with $(1 - p_t)$ suffering disease-induced mortality, but the impacts of these factors on NMB are neither linear nor consistent. For instance, at low values of p_t ($p_t = 27\%$), the adoption of active screening strategies (of any frequency or coverage) is substantially different in the WTP value range of 0–0.5, but increasingly high WTP values would not yield different strategies. However, at high values of p_t ($p_t = 100\%$), the adoption of active screening strategies at any frequency or coverage level are incumbent on very high WTP values.

The impact of WTP and p_t on optimal values of screening coverage and interval.

When assumptions about WTP and p_t are fixed, the screening coverage has a greater impact on whether the NMB is positive than the screening interval. Low screening coverage levels (< 20%) can be insufficient to obtain a positive NMB, whilst the screening interval does not change the sign of the NMB for most coverage levels, unless the interval is very small (< 0.5 years). The question of what value to fix for WTP and p_t has important implications on the optimal choice of screening coverage and optimal screening interval.

First, the optimal screening interval will be approximately one year under all assumptions of passive surveillance treatment coverage (p_t) and WTP (yellow dots in all panels of Fig. 2). For $WTP_c = 0.5$ and $p_t = 27\%$ (our standard assumption), the maximum mean NMB is found when the screening coverage is 90% and the screening interval is 0.67 years (Fig. 2; the yellow dot in top centre panel). For $WTP_c = 0.5$ and $p_t = 100\%$, the maximum mean NMB is also at the maximum screening coverage, but the higher treatment coverage (p_t) indicates that the optimal screening interval is of 1.25 years. The optimal screening interval is the same for all values of the number of zero-detections, but the minimum NMB was found at $z_a = 1$ and $z_r = 1$ (see accompanying app). It is notable that a very high WTP (3 times the GDP per capita) lends strong support for shorter screening intervals, favouring screenings in a village multiple times a year.

Second, and turning the attention from the active screening interval to the coverage, we found that in terms of NMB, the screening coverage has an inverse relationship with the treatment coverage under passive surveil-

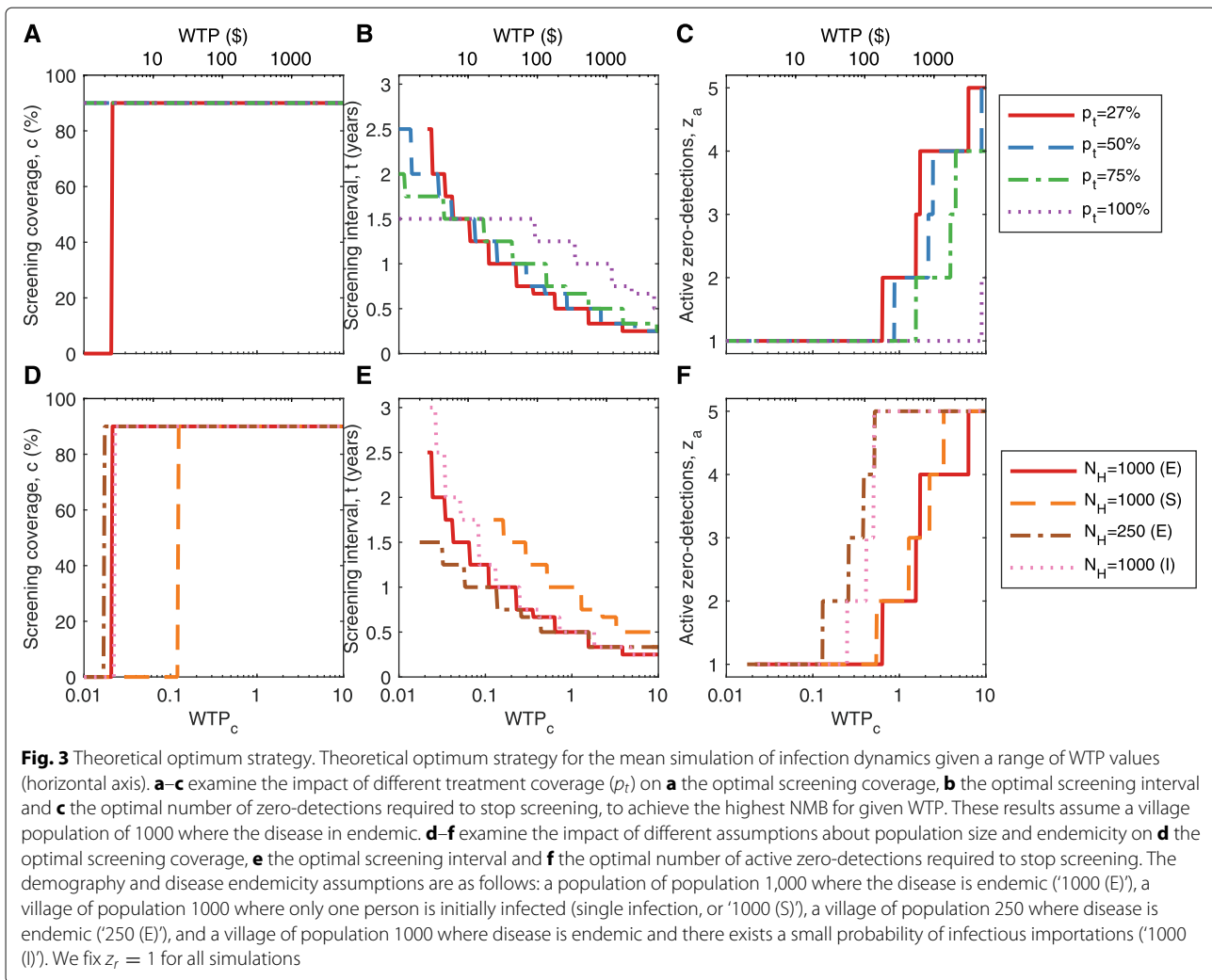
lance (p_t). When $p_t = 100\%$, the assumption is that all infections are eventually treated, implying no loss of life, which would otherwise be a large component of the change in NMB. Thus, under the assumption of a high p_t , active screening strategies with low coverage will be preferable, indicating that high costs of screening implementation will hardly be justified by DALYs averted, as passive surveillance is already very effective (Fig. 2). In contrast, under an assumption of a lower p_t , high active screening coverage is needed to compensate for lower treatment coverage in passive surveillance.

Sensitivity analysis of village characteristics and maximum net monetary benefits

We present the active screening strategy that gives the maximum mean NMB for a range of WTP values, examining the role of the treatment coverage under passive surveillance, the village population size, and the endemicity status of the village. To do this, we apply costs (see Additional file 1: Table S3) to the mean simulation outputs to see which strategy provides the largest NMB. Each line in Fig. 3a–c shows the value of c , t and z_a that together give the optimal strategy for given values of p_t and WTP. For instance, assuming that $WTP_c = 0.5$ and $p_t = 27\%$, the maximum NMB is obtained for $c = 90\%$, $t = 0.67$ years, and $z_a = 1$ (Fig. 3a–c). Figure 3d–f considers the same results for villages of different sizes (N_H) and with different initial conditions of the simulation (endemic, disease-free, and endemic with importations), all with $p_t = 27\%$. We fix $z_r = 1$ for all simulations.

Screening coverage, c , is optimal at 0% (not doing any screening) or at a very high coverage (the maximum of 90%) (Fig. 3a). For $p_t = 27\%$, active screening at the maximum coverage is optimal for $WTP_c > 0.02$ (Fig. 3a). This means if there is no WTP to avert DALYs (the decision-maker wants to remain cost-neutral), it is best not to incur any screening costs, since the NMB will be negative, however, if the WTP is above threshold 0.02 of GDP per capita, it is optimal to screen entire village populations to reduce the prevalence and prevent further transmission. The threshold WTP where maximum screening coverage is optimal is also influenced by p_t ; the WTP threshold decreases for larger values of p_t (above 27%) and active screening campaigns are always optimal for higher p_t values. Therefore, at high p_t values (treatment coverage), the additional costs of active screening are justified in order to shorten disease duration and expedite elimination (recovering screening costs via averted treatment costs).

The optimal screening interval, t , and the number of active zero-detection before screening cessation are more sensitive to the WTP, but have similar patterns across values of treatment coverage (p_t). The optimal screening interval, t , decreases with increasing WTP, since DALYs



are valued more highly and more frequent screening averts more DALYs (Fig. 3b). We also show that for any value of p_t , the screening interval must be two years or shorter, typically approximately annually. For most WTP values, a single active zero-detection is enough to justify cessation of active screening, but we note that if the funder is willing to pay more, there is a benefit in repeated active zero-detection campaigns before cessation in order to ensure no resurgence of transmission (Fig. 3c). To consider the full range of p_t values a heatmap of the change in cost with respect to both p_t and WTP is given in Additional file 1: Figure S8.

Additionally, we have considered three other scenarios: a smaller village population $N_H = 250$ at endemic equilibrium; a disease-free population starting with a single infection rather than endemic equilibrium, therefore mimicking a local post-elimination reintroduction of the infection; and a population with a small chance of an

imported infection (Fig. 3d–f). For scenarios with the population size $N_H = 1000$, the results are qualitatively similar: lower active screening coverage is optimal when there are fewer infections (due to the single reintroduction). For a smaller population size $N_H = 250$, it is more effective to have a shorter screening interval and more campaigns that yield active zero-detections to ensure local elimination, since the cost of active screening is smaller when there are fewer people to screen. When there is a small rate of importation of infection, the higher probability of sustaining local infection means that a higher number of active zero-detections z_a are optimal for any given WTP, and therefore active screening must continue for a longer period of time.

Limiting analysis to practical strategies

According to other literature, a high coverage in active screening with a minimum number of visits is desirable

[78], which agrees with our results (Fig. 3). However, unlike the screening interval and the number of active zero-detections before cessation, which can be designated by district managers, it is not always possible to achieve a desired screening coverage by either decree or investment because screening coverage depends on the availability and consent of the population [32]. In fact, attendance at active screening is often low [79]; for instance, in Kwamouth during 2000–2016, a median screening coverage of 55% was achieved for all village-level active screenings taken from the WHO HAT Atlas [40, 41] (see Additional file 1: Figure S2).

Therefore, since a high screening coverage cannot be guaranteed, we optimise active screening when we have imposed a maximum on the screening coverage consistent with historic (obtainable) coverage in DRC. A higher maximum level of screening coverage allows for a large screening interval and a small number of zero-detections before cessation of active screening. For the very low minimum screening coverage of 5%, the optimum is screening four times a year $t = 0.25$ and five zero-detections to stop ($z_a = 5$), whilst for a high coverage, we see the expected result of $t \approx 1$ year and $z_a = 1$ (Fig. 4). For the median screening coverage in Kwamouth of 55%, the optimal strategy is an active screening every four months, with two active zero-detection required for cessation. A lower screening coverage can be compensated for by an increase in screening frequency thereby reducing the screening interval t .

Cost-effectiveness of realistic strategies

Whilst we have determined which strategy, on average, maximises NMB for achievable levels of screening coverage, we now consider the cost-effectiveness of select strategies, restricting the number of strategies under consideration to a smaller number of options. For this process, we have selected seven options: doing no active screening and six realistic proposal schemes for active screening including biennial and annual screening with different cessation criteria. These active screening strategies are shown in Table 3. We assume $z_r = 1$, $p_t = 27\%$ and $N_H = 1,000$, and we initialise simulations with conditions consistent with endemic equilibrium.

For our comparator strategy (passive surveillance only) the total cost of implementation is the cost of testing and treating self-presenting individuals infected with gHAT. As previously stated, we do not include the fixed costs of continually operating a passive surveillance network, as we assume that implementing a strategy does not change this cost. Thus, Table 3 shows the average cost of only treating self-presenting patients is \$25,754 with 2488.8 DALYs. Since we are considering this our baseline strategy, zero DALYs are averted from this process.

By employing active screening the total costs increase to the benefit of health outcomes; annual screening costs more than biennial screening, but a correspondingly larger number of DALYs are averted under strategies that use annual screening. On the other hand, increasing the number of active zero-detections increases

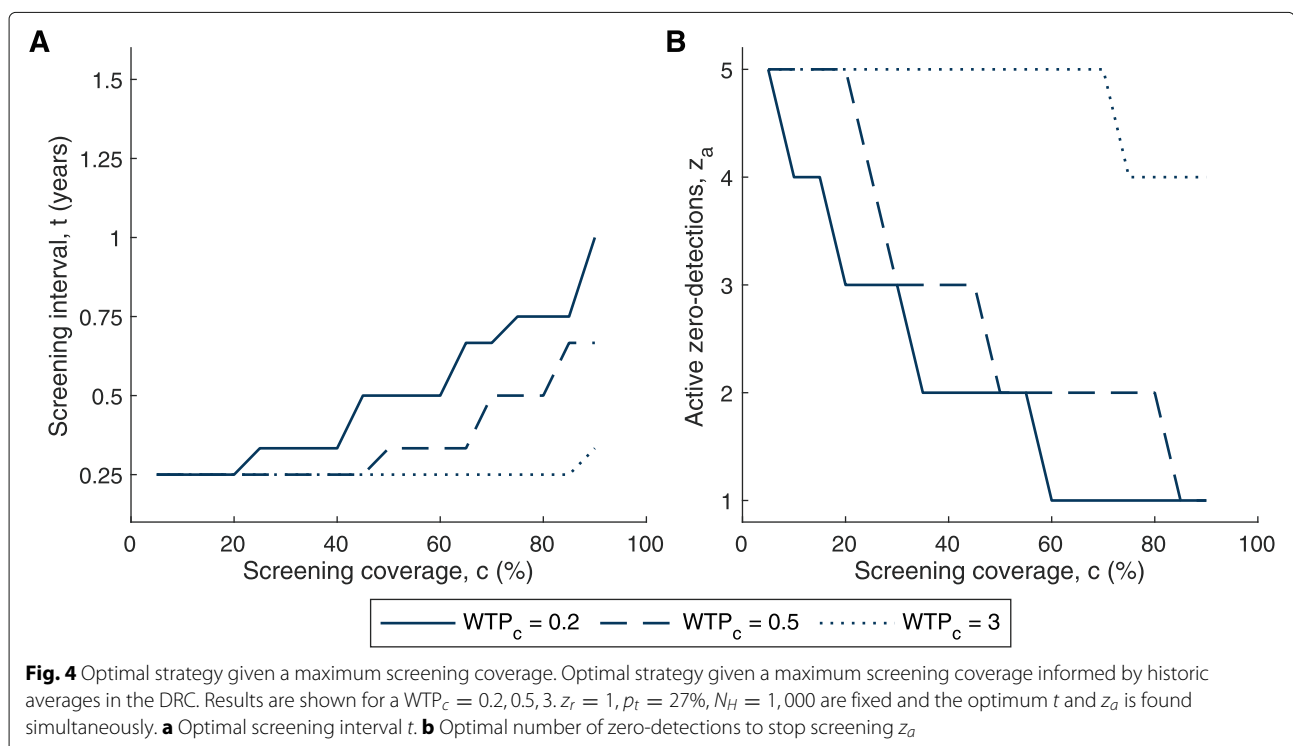


Table 3 Active screening strategies considered in the probability of cost-effectiveness calculations

No.	Strategy	Screening coverage, c (%)	Screening interval, t (years)	Active zero-detections, z_a	Total cost (\$)	Total DALYs	ACER (\$/DALY)	ICER (\$/DALY)
1	Passive surveillance only ¹	0	N/A	N/A	25754 [12267, 40807]	2488.8 [1207.8, 3912.6]	Minimum cost	Minimum cost
2	Biennial screening with one zero for cessation	55	2	1	42517 [23546, 62968]	1338.9 [437.8, 1755.7]	14.6	14.6 ²
3	Biennial screening with two zeros for cessation	55	2	2	43066 [24432, 63133]	1339.4 [436.3, 1755.5]	15.1	Dominated
4	Biennial screening with three zeros for cessation	55	2	3	43501 [25092, 63266]	1340.1 [434.8, 1756.5]	15.4	Dominated
5	Annual screening with one zero for cessation	55	1	1	48885 [26628, 72263]	1027.9 [570.3, 2262.7]	15.8	20.4 ³
6	Annual screening with two zeros for cessation	55	1	2	49571 [27896, 72428]	1027.7 [570.6, 2264.3]	16.3	2501.4 ⁴
7	Annual screening with three zeros for cessation	55	1	3	50163 [28670, 72806]	1027.4 [572.9, 2261.9]	16.7	2772.3 ⁵

¹The comparator strategy. ²Relative to Strategy 1. ³Relative to Strategy 2. ⁴Relative to Strategy 5. ⁵Relative to Strategy 6.

We show the mean total cost (to nearest dollar) and the total number of DALYs (to the nearest 0.1 DALYs) for each strategy with the 95% prediction intervals across all stochastic realisations. The ACER is the change in cost over the change in DALYs averted as compared to the baseline strategy, whilst the ICER is compared to the next best strategy (given in the table footnotes). Costs are denominated in 2018 US dollars

costs but yields few additional DALYs under a regimen of annual screening and even fewer DALYs under biennial screening. Given the 95% prediction intervals for the total DALYs averted for varying just the active zero-detections greatly overlap, there is little basis on which to choose between these strategies other than lowering costs, but the screening interval is much more significant in terms of health benefits conferred (Table 3).

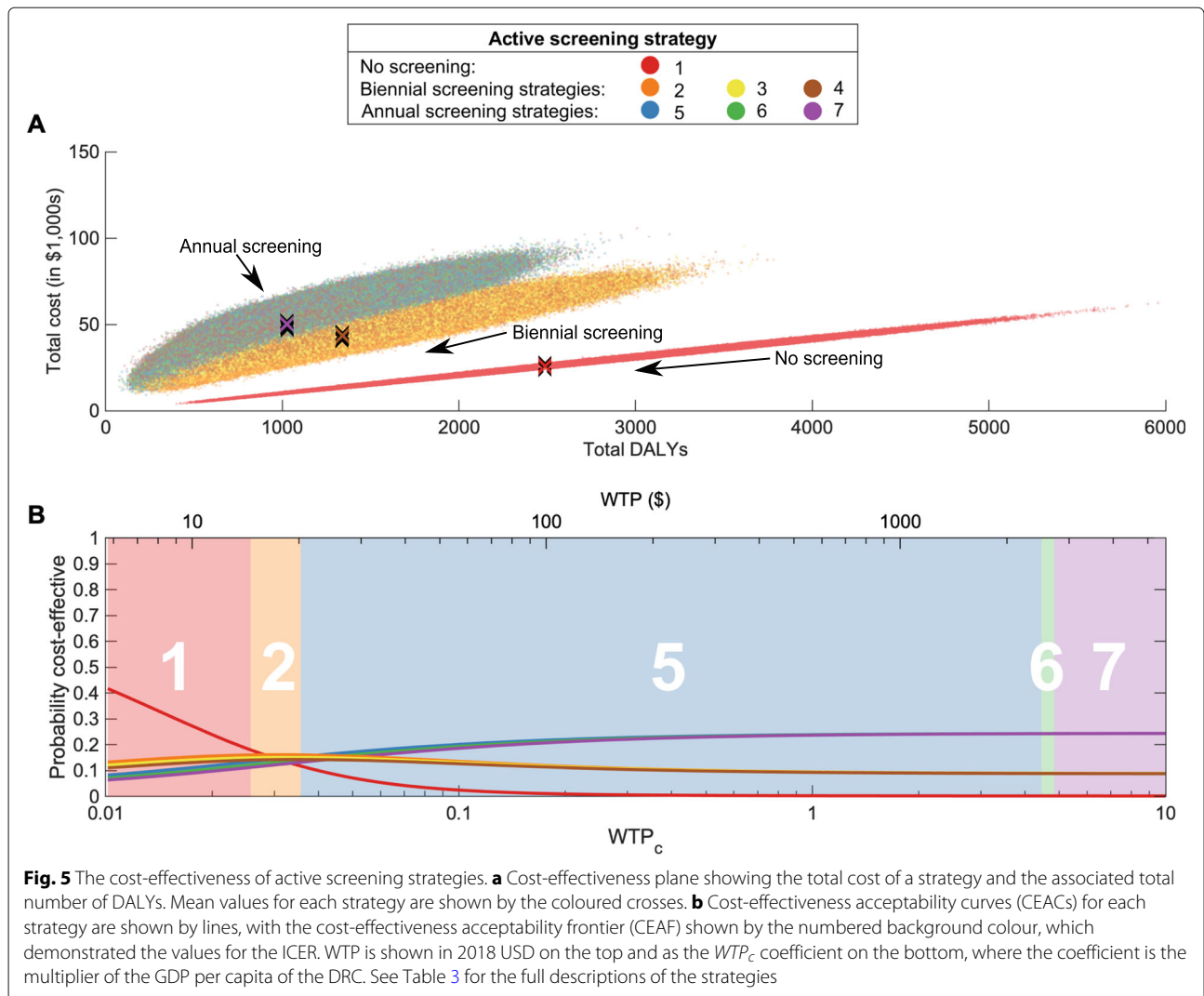
We calculate the average cost-effectiveness ratio (ACER) as the ratio of the change in cost to change in DALYs averted relative to the comparator strategy, whilst the incremental average cost-effectiveness ratio (ICER) is the ratio of the change in cost to change in DALYs averted relative to the next best option (see Table 3). The ICER is the conventional metric for cost-effectiveness: a strategy is cost-effective compared to the next best strategy if $ICER < WTP$. Active screening at 55% coverage is cost-effective at extremely low WTP values (\$14–17 per DALY averted), and the costs-per-DALY of biennial and annual screening strategies are so similar as to suggest that there is little loss in efficiency (few diminishing returns) in undertaking more frequent (yearly) campaigns. Notably, however, a higher WTP (\$2501 per DALY) would be needed to support the choice of strategies with higher values of z_a (multiple active zero-detections), but such strategies should not be implemented at the expense of longer screening intervals. Strategies with higher values of z_a are primarily payment for certainty that transmission chains have been broken, rather purchasing any substantial health burden averted (in terms of DALYs).

Cost-effectiveness analysis in the presence of parameter uncertainty

Not every village will experience gHAT infection as depicted by the mean infection profile. Hence, we aim to account for uncertainty and present the probability that a strategy is cost-effective. Using the full range of possibilities for the infection dynamics is particularly important for gHAT infection in a village as we know there is potential for large differences between seemingly identical villages, due to the focal nature of the infection [8]. Therefore, we have simulated the infection dynamics of each strategy one million times to compare how the costs and number of DALYs averted can vary.

When there is no active screening, the variations in the cost and DALYs incurred arises from uncertainty in the transmission model (since cost parameters are held constant). By the same token, costs and DALYs are positively correlated because both arise from transmission, passive detection and treatment. The more infections there are in the village, the more treatments will be performed, increasing costs and incurring DALYs. Inversely, when there are few infections, DALYs and costs are lower. Moreover, the correlation between DALYs and costs is not perfect, there is large variation in both measures (Fig. 5a).

There is a similar pattern when active screening exists, but the range of introduced costs and the number of occurring treatments increases, expanding the variance in outcomes and weakening the correlation between costs and DALYs. However, reducing the screening interval both increases the costs and reduces the number of DALYs incurred, all other things held equal. Results for



alternative numbers of zero-detections remain marginal over all parameter values (Fig. 5a). There are significant differences in the outcomes of strategies with different screening intervals, but we reiterate that the differences are robust to different numbers of active zero-detections (Table 3 and Additional file 1: Figure S3).

We have calculated the probability a strategy is cost-effective by taking the proportion of our one million simulations for each strategy that has the largest NMB at each specific value of WTP (Fig. 5b), thereby producing cost-effectiveness acceptability curves (CEACs) [80]. The strategy with no active screening has the highest probability of being cost-effective for low WTP; however, the probability is approximately 50% for the lowest WTP (at cost-neutral values), indicating that performing any active screening might still be cost-effective. The strategy with no active screening has negligible probability of being the cost-effective (< 1%) for approximately $WTP_c > 0.23$. For

the highest WTP values, strategy 7 (annual screening with three zero-detections), has the highest probability, with similar probabilities for the other strategies with annual screening. Given the probability of being cost-effective for high WTP is so similar for strategies 5, 6 and 7, it is difficult to make a recommendation about which is better, and so the most cautious strategy s should likely be favoured, given the minimal change in costs.

It is notable that for certain values of WTP, the strategy with the highest probability of cost-effectiveness does not always correspond with the preferred strategy according to the ICER. These calculations are demonstrated by the cost-effectiveness acceptability frontier (CEAF), which is shown as the shaded and numbered background (Fig. 5b). The CEAF is the optimal strategy according to the maximum expected NMB (rather than the highest probability of the NMB). Discrepancies between the highest mean NMB and the probability of the highest

NMB occur because of asymmetric distributions in our parameters, but it is the CEAF which is used to provide strategy recommendations for a given WTP threshold for risk-neutral decision-makers [81].

In the lower-prevalence health zone of Mosango (Additional file 1: Figure S3) we find similar results, with a recommendation for annual screening over biennial screening at moderate WTP values or no active screening at lower WTP values.

Discussion

To achieve the goal of eliminating gHAT it is useful to have robust models that can inform policy makers about the potential of different intervention strategies [82]. As such, we have followed the five principles of the Neglected Tropical Diseases Modelling Consortium (see Additional file 1: Table S7), which were proposed to improve the quality of communication between modellers and stakeholders [58]. Furthermore, the addition of economic analysis will further develop the use of this work, as to not only evaluate which strategies are able to decrease in infection in the population, but which are cost-effective.

We have presented a stochastic model for individual villages that demonstrates how active screening should be considered, by determining costs of implementing the screening for different screening coverage levels c , screening intervals t , and the number of zero-detections observed to stop screening z_a and z_r . Using $WTP_c = 0.5$ we find on average that ideally screening would be done approximately yearly with maximal screening coverage and ceased when no infection is found in a single screening ($c = 90\%$, $t = 0.67$ years, $z_a = 1$) (Fig. 3). Whilst the optimum for the screening interval is found to be 0.67 years, if there is a higher proportion of infection eventually treated in the population than the assumed $p_t = 27\%$, the optimal interval is larger (1.25 years for $p_t = 100\%$). Practically these intervals might be difficult to implement, so we believe that the current work supports the implementation of yearly screening. These results have been specifically calibrated to the high-endemicity health zone of Kwamouth, DRC, however similar results for the lower-prevalence health zone of Mosango, DRC (Additional file 1: Figure S3) support that our recommendations are generalisable to other regions with low to high-endemicity (further analysis would be required to apply these results to very low-endemicity regions or regions with historically very low screening coverage). It is noted that such high screening coverage will rarely be able to be achieved, and so multiple visits where no infection is observed may be necessary to optimise control, although the model shows no significant differences in cost-effectiveness. This is in line with WHO guidelines of annual active screenings until there have been three

consecutive years of no new cases, followed by a further screening with no cases three years after cessation of activities [5].

In particular, we note that whilst we assumed that reactive screening should immediately resume upon identification of an infection through passive surveillance, the time interval for reactive screening to begin has little effect on the results (see Additional file 1: Figure S6). Therefore, we conclude that practical concerns about the feasibility of reactive screening do not impact our conclusions, as long as reactive surveillance is deployed within two years of finding a new case through passive surveillance.

In fact, the choice of a low z_a has a high probability of triggering reactive screening ($>70\%$), and therefore, we recommend that logistics for reactive screening are put in place (see Additional file 1: Figure S5). The time-horizon of 30 years is sufficient in the village context to capture the dynamics; when we expanded the horizon to 100 years, we found that roughly 99.1% of costs and 99.8% of DALYs are attributable to the first 30 years (see Additional file 1: Figure S7).

As new treatments and active screening modes are introduced, the costs of the model will change, however, the biggest effect is that of the number of DALYs averted, assuming the WTP threshold is set to a reasonable level. Details of the variation in NMB for different screening diagnostics and medical treatments can be found in Additional file 1: Figure S9.

We also note that from the perspective of a single village (and from the perspective of a risk-neutral payer), we do not put any weight on local elimination beyond that captured by expected DALYs averted, favouring an optimal screening strategy that terminates the programme after a single active zero-detection ($z_a = 1$), rather than repeated zero-detections to ensure elimination. On the other hand, when it is assumed that a village is susceptible to importations of infection, we find that more active zero-detections are required to maximise the NMB (Fig. 3f). Other work has shown that at least three zero-detections for villages of this size ($N_H = 1000$) [26] to ensure elimination, but it is unclear how much monetary value we should attribute to meeting EOT targets.

Our finding that a single zero-detection is optimal in Fig. 3 is particularly notable when the screening coverage c is at the maximum (90%). In this case, there is higher confidence that local gHAT elimination has been achieved, as almost all the population is screened and there are no cases left to be detected, and even if infection temporarily persists after this, there is a large probability it will die out due to stochastic fade out [26]. However, a regular 90% coverage is probably unfeasible, and more realistic screening coverage will require more zero-detections to terminate active screening (Fig. 4).

We note that there may be also be additional costs in restarting active screening as reactive screening, in particular if regional cessation results in disbanded trained mobile teams. We have not accounted for this in our model, but it may lend support to a higher number of active zero-detections, which lower the probability of reactive screening once routine active surveillance has ceased (see Additional file 1: Figure S5). We cannot make a recommendation for the number of reactive zero-detections as the impact is negligible on the cost-effectiveness (even less so than the active zero-detections) and the effect is completely outweighed by the stochasticity of the infection dynamics (see Additional file 1: Figure S6).

Whilst many other diseases have established procedures for active case finding and evaluated cost-effectiveness (i.e. tuberculosis) few have done it for elimination and no studies have evaluated active screening properties with this level of detail. Bessell et al. [11] found that RDTs can be cost-effective; Sutherland et al. [54] took active screening programme properties for granted and instead focused on the combination of active screening, passive surveillance, and vector control. Therefore, this is the first cost-effectiveness paper that examines in detail the relative efficiency of active screening strategies. Furthermore, we have provided the tools for the reader to adapt the analysis to their specific chosen costs (see accompanying app).

Unsurprisingly, we find that a big factor in choosing a strategy to implement is how much the programme funder, ministry of health or external donor is willing to pay to avert an additional DALY beyond other acceptable configurations of the me. WTP is not a metric about the total cost of the me, it is a metric of comparative efficiency, considering incremental costs and incremental effects between two or more strategies. To calculate what WTP gHAT programmes were acting on in the past, one would have to know what alternative strategies they were considering, which is beyond the scope of our current work. In addition, contributions to gHAT programmes are complicated, since much of the activity has historically been funded by donors; indeed, without external funding there would be substantial harm to the programme [83]. Therefore, we present our results across a range of WTP values. We have used a WTP value of 0.5 of the GDP per capita of the DRC ($WTP_c = 0.5$), which is commonly used in the literature [72, 84], but note there may be a higher WTP to achieve the additional aim of gHAT elimination.

Future research is warranted to evaluate the specific characteristics of each village, how villages and health zones (or districts) share costs and the impact it makes on relative efficiency. Moreover, the risk of importation, the impact of potential sero-negative skin-infected cases, and the risk of animal reservoirs would

have to be further explored in in-depth epidemiological modelling.

Conclusions

Using a dynamic cost-effectiveness framework we have performed analysis to examine the optimal use of medical resources for strategies against gHAT as it approaches the end-game. With a limited number of active screening teams and resources for them to carry out their duties, it is important to have guidelines that optimise their activities with the aim of driving towards elimination. We considered a range of different active screening interventions, including changing the coverage, frequency and cessation criterion of screening, on top of a baseline passive surveillance system. The results indicate that active screening is effective in reducing case numbers and hence the infection in the population, with approximately annual screening representing good use of resources and being cost-effective at very low WTP ($> \$20.4/\text{DALY}$ averted) when realistic coverage of 55% was assumed. Biennial screening at the same coverage each visit would avert fewer DALYs at a slightly lower cost and was only cost-effective for a narrow range of WTP values ($\$14.6\text{--}20.4$). The calculated costs and DALYs averted were very similar for all the cessation criteria considered here (one, two or three years of zero cases before stopping active screening), however the additional information gained about progress towards elimination of transmission, leads us to suggest cessation after three years of zeros is appropriate. We recommend that active screening should be carried out annually with cessation after three years of detecting no cases for a moderate-to high-endemicity health zone in the DRC—in line with current WHO guidance.

Abbreviations

gHAT: *Gambian* human African trypanosomiasis; NMB: Net monetary benefit; HAT: Human African trypanosomiasis; DALY: Disability-adjusted life year; WHO: World Health Organization; EOT: Elimination of transmission; CATT: Card agglutination test for trypanosomiasis; RDT: Rapid diagnostic test; DRC: Democratic Republic of Congo; DIC: Deviance information criterion; YLL: Years of life lost; YLD: Years lived with disability; WTP: Willingness-to-pay; GDP: Gross domestic product; ACER: Average cost-effectiveness ratio; ICER: Incremental cost-effectiveness ratio; CEAC: Cost-effectiveness acceptability curve; CEAF: Cost-effectiveness acceptability frontier

Supplementary Information

The online version contains supplementary material available at <https://doi.org/10.1186/s12916-021-01943-4>.

Additional file 1: Figures S1–S10 and Tables S1–S7. Supporting information for the manuscript including further details of the gHAT modelling, explanations and derivations of all net monetary benefit parameters, a short analysis of a different health zone, details on reactive screening, some additional results, and a table of how we satisfied the five principles of the Neglected Tropical Diseases Modelling Consortium. Figure S1 – Schematic of the compartmental model for gHAT infection dynamics in humans (low- and high-risk) and tsetse. Figure S2 – Histograms of screening coverage for (A) Kwamouth and (B) Mosango. Figure S3 – The cost-effectiveness of active screening strategies for the health zone of Mosango. Figure S4 – The effect of the number of zero-detections in

reactive screening on the cost effective acceptability curves (CEACs). Figure S5 – Active and reactive screenings information. Figure S6 – The time interval between observing a passive detection and resuming reactive screening for different values of active zero-detections. Figure S7 – Distributions of (A) the change in cost of an active screening strategy and (B) the number of DALYs averted for 30- and 100-year time-horizons. Figure S8 – Theoretical optimum strategy for mean simulation of infection dynamics given WTP and p_r . Figure S9 – Mean change in costs relative to the assumed interventions for a village with 55% screening coverage annually with 3 zero-detections. Figure S10 – Linear fits through the results of different active screening strategies from Fig. 5. Table S1 – The transition rates of the Markov-chain gHAT-infection model and additional ordinary differential equation component of the model. Table S2 – Parameter notation and values for the gHAT infection model. Table S3 – Net monetary benefit parameters that determine our manuscript parameters. Table S4 – Proportion of people receiving each type of treatment. Table S5 – Formulation of manuscript parameters from cost data. Table S6 – The values that determine the linear fits for each strategy in Figure S10. Table S7 – PRIME-NTD criteria fulfillment.

Acknowledgements

Not applicable.

Authors' contributions

CND, KSR and MJK developed the mathematical model, performed simulations, analysed results; CND, KSR, MA and MJK performed the economic analysis; EMM provided knowledge of gHAT in DRC and help with the original data curation; CND and KSR wrote the first draft of the manuscript, and all co-authors contributed to the final draft. All authors read and approved the final manuscript.

Funding

This work was supported by the Bill and Melinda Gates Foundation (www.gatesfoundation.org) in partnership with the Task Force for Global Health through the NTD Modelling Consortium [OPP1184344] (CND, KSR and MJK), the Bill and Melinda Gates Foundation through the Human African Trypanosomiasis Modelling and Economic Predictions for Policy (HAT MEPP) project [OPP1177824] (KSR, MA, and MJK), and EPSRC/MRC via the MathSys Centre for Doctoral Training (CND and MJK). The funders had no role in study design, data collection and analysis, decision to publish, or preparation of the manuscript.

Availability of data and materials

The authors thank PNLTHA for original data collection and WHO for data access (in the framework of the WHO HAT Atlas [2]). Results from the manuscript using user-chosen cost parameter values can be viewed in the accompanying app at <https://christopherdavis.shinyapps.io/optimising-ghat-active-screening/>.

Declarations

Ethics approval and consent to participate

Not applicable.

Consent for publication

Not applicable.

Competing interests

The authors declare that they have no competing interests.

Author details

¹Mathematics Institute, University of Warwick, CV4 7AL Coventry, UK. ²Zeeman Institute (SBIDER), University of Warwick, CV4 7AL Coventry, UK. ³Swiss Tropical and Public Health Institute, Socinstrasse 57, 4051 Basel, Switzerland.

⁴University of Basel, Petersplatz 1, 4051 Basel, Switzerland. ⁵Programme National de Lutte contre la Trypanosomiase Humaine Africaine (PNLTHA), Ave Coisement Liberation et Bd Triomphal No 1, Commune de Kasavubu, Kinshasa, Democratic Republic of Congo. ⁶School of Life Sciences, University of Warwick, CV4 7AL Coventry, UK.

Received: 18 August 2020 Accepted: 16 February 2021

Published online: 01 April 2021

References

1. World Health Organization. WHO Global Health Observatory. <https://www.who.int/data/gho>. Accessed 17 Nov 2020.
2. Franco JR, Cecchi G, Priotto G, Paone M, Diarra A, Grout L, Simarro PP, Zhao W, Argaw D. Monitoring the elimination of human african trypanosomiasis at continental and country level: Update to 2018. *PLoS Negl Trop Dis*. 2020;14(5):0008261.
3. Jamonneau V, Ilboudo H, Kaboré J, Kaba D, Koffi M, Solano P, Garcia A, Courtin D, Laveissière C, Lingue K, et al. Untreated human infections by trypanosoma brucei gambiense are not 100% fatal. *PLoS Negl Trop Dis*. 2012;6(6):1691.
4. Institute for Health Metrics and Evaluation (IHME). GBD Compare Data Visualization. Seattle, WA: IHME, University of Washington; 2020. Available from <http://vizhub.healthdata.org/gbd-compare>. (Accessed 2020).
5. World Health Organization. Control and surveillance of human African trypanosomiasis. *World Health Organ Tech Rep Ser*. 2013;(984):1–237. PMID: 24552089.
6. Franco J, Simarro P, Diarra A, Ruiz-Postigo J, Jannin J. The journey towards elimination of gambiense human african trypanosomiasis: not far, nor easy. *Parasitology*. 2014;141(6):748–60.
7. Robays J, Bilengue MMC, Stuyft P. V. d., Boelaert M. The effectiveness of active population screening and treatment for sleeping sickness control in the democratic republic of congo. *Trop Med Int Health*. 2004;9(5):542–50.
8. Büscher P, Cecchi G, Jamonneau V, Priotto G. Human African trypanosomiasis. *Lancet*. 2017;390(10110):2397–409.
9. Lumbala C, Simarro PP, Cecchi G, Paone M, Franco JR, Mesu VKBK, Makabuza J, Diarra A, Chansy S, Priotto G, et al. Human african trypanosomiasis in the democratic republic of the congo: disease distribution and risk. *Int J Health Geogr*. 2015;14(1):20.
10. Magnus E, Van Meirvenne N, Vervoort T, Le Ray D, Wéry M, Others. Use of freeze-dried trypanosomes in the indirect fluorescent antibody test for the serodiagnosis of sleeping sickness. *Ann Soc Belg Méd Trop*. 1978;58: 103–9.
11. Bessell PR, Lumbala C, Lutumba P, Baloji S, Bieler S, Ndung'u JM. Cost-effectiveness of using a rapid diagnostic test to screen for human african trypanosomiasis in the democratic republic of the congo. *PLoS ONE*. 2018;13(9):e0204335.
12. Wamboga C, Matovu E, Bessell PR, Picado A, Biéler S, Ndung'u JM. Enhanced passive screening and diagnosis for gambiense human african trypanosomiasis in north-western uganda—moving towards elimination. *PLoS ONE*. 2017;12(10):e0186429.
13. Lee SJ, Palmer JJ. Integrating innovations: a qualitative analysis of referral non-completion among rapid diagnostic test-positive patients in uganda's human african trypanosomiasis elimination programme. *Infect Dis Poverty*. 2018;7(1):1–16.
14. Mumba Ngoyi D, Menten J, Pyana PP, Büscher P, Lejon V. Stage determination in sleeping sickness: Comparison of two cell counting and two parasite detection techniques. *Trop Med Int Health*. 2013;18(6): 778–82.
15. Chappuis F. Oral fexinidazole for human african trypanosomiasis. *The Lancet*. 2018;391(10116):100–2.
16. WHO interim guidelines for the treatment of gambiense human African trypanosomiasis. Geneva: World Health Organization; 2019. PMID: 31449367.
17. Koffi M, N'Djetchi M, Ilboudo H, Kaba D, Coulibaly B, N'Gouan E, Kouakou L, Bucheton B, Solano P, Courtin F, et al. A targeted door-to-door strategy for sleeping sickness detection in low-prevalence settings in Côte d'Ivoire. *Parasite*. 2016;23:51.
18. Mitashi P, Hasker E, Mbo F, Van Geertruyden J, Kaswa M, Lumbala C, Boelaert M, Lutumb a. P. Integration of diagnosis and treatment of sleeping sickness in primary healthcare facilities in the democratic republic of the congo. *Trop Med Int Health*. 2015;20(1):98–105.
19. Van Nieuwenhove S. Present strategies in the treatment of human African trypanosomiasis. In: *Progress in Human African Trypanosomiasis, Sleeping Sickness*. Paris: Springer; 1999. p. 253–80.
20. Simarro P, Franco J, Asumu PN. Has the focus of human African trypanosomiasis in Luba, Equatorial Guinea been eradicated?. *Med Trop revue du Corps de sante colonial*. 2001;61(4-5):441–4.
21. Pandey A, Atkins KE, Bucheton B, Camara M, Aksoy S, Galvani AP, Ndeffo-Mbah ML. Evaluating long-term effectiveness of sleeping sickness control measures in Guinea. *Parasites Vectors*. 2015;8(1):550.

22. Rock KS, Torr SJ, Lumbala C, Keeling MJ. Quantitative evaluation of the strategy to eliminate human African trypanosomiasis in the DRC. *Parasites Vectors*. 2015;8(1):532.
23. De Vries H, Wagelmans AP, Hasker E, Lumbala C, Lutumba P, De Vlas SJ, van de Klundert J. Forecasting human african trypanosomiasis prevalences from population screening data using continuous time models. *PLoS Comput Biol*. 2016;12(9):e1005103.
24. Rock KS, Torr SJ, Lumbala C, Keeling MJ. Predicting the impact of intervention strategies for sleeping sickness in two high-endemicity health zones of the Democratic Republic of Congo. *PLoS Negl Trop Dis*. 2017;11:0005162.
25. Mahamat MH, Peka M, Rayaïsse J-b., Rock KS, Toko MA, Darnas J, Brahim GM, Alkatib AB, Yoni W, Tirados I, Courtin F, Brand SPC, Nersy C, Alfaroukh O, Torr SJ, Lehane MJ, Solano P. Adding tsetse control to medical activities contributes to decreasing transmission of sleeping sickness in the Mandoul focus (Chad). *PLoS Negl Trop Dis*. 2017;11(7):0005792.
26. Davis CN, Rock KS, Miaka EM, Keeling MJ. Village-scale persistence and elimination of gambiense human african trypanosomiasis. *PLoS Negl Trop Dis*. 2019;13(10):e0007838.
27. Ndeffo-Mbah ML, Pandey A, Atkins KE, Aksoy S, Galvani AP. The impact of vector migration on the effectiveness of strategies to control gambiense human african trypanosomiasis. *PLoS Negl Trop Dis*. 2019;13(12):e0007903.
28. Castaño MS, Ndeffo-Mbah ML, Rock KS, Palmer C, Knock E, Miaka EM, Ndung'u JM, Torr S, Verlé P, Spencer SE, et al. Assessing the impact of aggregating disease stage data in model predictions of human african trypanosomiasis transmission and control activities in bandundu province (drc). *PLoS Negl Trop Dis*. 2020;14(1):0007976.
29. Huang C-I, Crump RE, Brown P, Spencer SE, Miaka EM, Shampa C, Keeling MJ, Rock KS. Shrinking the ghat map: identifying target regions for enhanced control of gambiense human african trypanosomiasis in the democratic republic of congo. medRxiv. 2020;07(03):20145847. <https://doi.org/10.1101/2020.07.03.20145847>.
30. Crump RE, Huang C-I, Knock ES, Spencer SE, Brown PE, Mwamba Miaka E, et al. Quantifying epidemiological drivers of gambiense human African Trypanosomiasis across the Democratic Republic of Congo. *PLoS Comput Biol*. 2021;17(1):e1008532. <https://doi.org/10.1371/journal.pcbi.1008532>.
31. The World Bank. Data: Democratic Republic of Congo. 2015. <http://data.worldbank.org/country/congo-democratic-republic>. Accessed 2015.
32. Mpanya A, Hendrickx D, Vuna M, Kanyinda A, Lumbala C, Tshilombo V, Mitashi P, Luboya O, Kande V, Boelaert M, Lefèvre P, Lutumba P. Should I get screened for sleeping sickness? A qualitative study in Kasai province. *PLoS Negl Trop Dis*. 2012;6(1):1467.
33. Rogers D. A general model for the African trypanosomiasis. *Parasitology*. 1988;97(1):193–212.
34. Checchi F, Filipe JAN, Haydon DT, Chandramohan D, Chappuis F. Estimates of the duration of the early and late stage of gambiense sleeping sickness. *BMC Infect Dis*. 2008;8(1):16.
35. Checchi F, Funk S, Chandramohan D, Haydon DT, Chappuis F. Updated estimate of the duration of the meningo-encephalitic stage in gambiense human African trypanosomiasis. *BMC Res Notes*. 2015;8(1):292.
36. Davis S, Aksoy S, Galvani A. A global sensitivity analysis for African sleeping sickness. *Parasitology*. 2011;138(4):516–26.
37. Ravel S, Grébaud P, Cuisance D, Cuny G. Monitoring the developmental status of *Trypanosoma brucei* gambiense in the tsetse fly by means of PCR analysis of anal and saliva drops. *Acta Trop*. 2003;88(2):161–5.
38. Clausen P, Adeyemi I, Bauer B, Breloer M, Salchow F, Staak C. Host preferences of tsetse (Diptera: Glossinidae) based on bloodmeal identifications. *Med Vet Entomol*. 1998;12(2):168–80.
39. Checchi F, Chappuis F, Karunakara U, Priotto G, Chandramohan D. Accuracy of five algorithms to diagnose gambiense human african trypanosomiasis. *PLoS Negl Trop Dis*. 2011;5(7):e1233.
40. Simarro PP, Cecchi G, Franco JR, Paone M, A D, Priotto G, R M, J J. Monitoring the progress towards the elimination of gambiense human African trypanosomiasis. *PLoS Negl Trop Dis*. 2015;9(6):0003785.
41. Simarro PP, Cecchi G, Jannin JG. The Atlas of human African trypanosomiasis: a contribution to global mapping of neglected tropical diseases. *Int J Health Geogr*. 2010;9:57.
42. World Health Organization, Baltussen RMPM, Adam T, Tan-Torres Edejer T, Hutubessy RCW, et al. Making choices in health : WHO guide to costeffectiveness analysis / edited by T. Tan-Torres Edejer ... [et al]. World Health Organization. 2003. <https://apps.who.int/iris/handle/10665/42699>.
43. Lutumba P, Makieya E, Shaw A, Meheus F, Boelaert M. Human african trypanosomiasis in a rural community, democratic republic of congo. *Emerg Infect Dis*. 2007;13(2):248.
44. Indicators WD. World bank. Technical report. 2015. <https://datacatalog.worldbank.org/dataset/world-development-indicators>. Accessed 2020.
45. Kyu HH, Abate D, Abate KH, Abay SM, Abbafati C, Abbasi N, Abbastabar H, Abd-Allah F, Abdela J, Abdelalim A, et al. Global, regional, and national disability-adjusted life-years (daly) for 359 diseases and injuries and healthy life expectancy (hale) for 195 countries and territories, 1990–2017: a systematic analysis for the global burden of disease study 2017. *Lancet*. 2018;392(10159):1859–922.
46. Snijders R, Fukinsia A, Claeys Y, Mpanya A, Hasker E, Meheus F, Miaka EM, Boelaert M. Cost of a new method of active screening for human African trypanosomiasis in the Democratic Republic of the Congo. *PLoS Negl Trop Dis*. 2020;14(12):0008832.
47. Lutumba P, Meheus F, Robays J, Miaka C, Kande V, Büscher P, Dujardin B, Boelaert M. Cost-effectiveness of algorithms for confirmation test of human african trypanosomiasis. *Emerg Infect Dis*. 2007;13(10):1484.
48. Iruzun-Lopez M, Erondou NA, Djibo A, Griffiths U, Stuart JM, Fernandez K, Ronveaux O, Le Gargasson J-B, Gessner BD, Colombini A. The actual and potential costs of meningitis surveillance in the african meningitis belt: Results from chad and niger. *Vaccine*. 2016;34(8):1133–8.
49. Bertram MY, Stenberg K, Brindley C, Li J, Serje J, Watts R, Edejer TT-T. Disease control programme support costs: an update of who-choice methodology, price databases and quantity assumptions. *Cost Eff Resour Allocation*. 2017;15(1):21.
50. Stenberg K, Lauer JA, Gkountouras G, Fitzpatrick C, Stanciole A. Econometric estimation of who-choice country-specific costs for inpatient and outpatient health service delivery. *Cost Eff Resour Allocation*. 2018;16(1):11.
51. Keating J, Yukich JO, Sutherland CS, Woods G, Tediosi F. Human African trypanosomiasis prevention, treatment and control costs: A systematic review. *Acta Trop*. 2015;150:4–13.
52. Shaw A, Cattand P. Analytical tools for planning cost-effective surveillance in gambiense sleeping sickness. *Med Trop revue du Corps de sante colonial*. 2001;61(4-5):412–21.
53. Mesu VKBK, Kalonji WM, Bardonneau C, Mordt OV, Blesson S, Simon F, Delhomme S, Bernhard S, Kuziena W, Lubaki J-PF, et al. Oral fexinidazole for late-stage african trypanosoma brucei gambiense trypanosomiasis: a pivotal multicentre, randomised, non-inferiority trial. *Lancet*. 2018;391(10116):144–54.
54. Sutherland CS, Stone CM, Steinmann P, Tanner M, Tediosi F. Seeing beyond 2020: an economic evaluation of contemporary and emerging strategies for elimination of *Trypanosoma brucei* gambiense. *Lancet Glob Health*. 2017;5(1):69–79.
55. Antillon M, Huang C-I, Crump RE, Brown PE, Snijders R, Miaka EM, Keeling MJ, Rock KS, Tediosi F. Economic evaluation of gambiense human African trypanosomiasis elimination campaigns in five distinct transmission settings in the Democratic Republic of Congo. medRxiv. 2020. 2020.08.25.20181982; <https://doi.org/10.1101/2020.08.25.20181982>.
56. Hackett F, Ford LB, Fevre E, Simarro P. Incorporating scale dependence in disease burden estimates: the case of human african trypanosomiasis in uganda. *PLoS Negl Trop Dis*. 2014;8(2):e2704.
57. Chang W, Cheng J, Allaire J, Xie Y, McPherson J. shiny: Web Application Framework for R. R package version 1.5.0. 2020. <https://CRAN.R-project.org/package=shiny>.
58. Behrend MR, Basáñez M-G, Hamley JI, Porco TC, Stolk WA, Walker M, de Vlas SJ, NTD Modelling Consortium. Modelling for policy: The five principles of the neglected tropical diseases modelling consortium. *PLoS Negl Trop Dis*. 2020;14(4):0008033.
59. Franco JR, Cecchi G, Priotto G, Paone M, Diarra A, Grout L, Simarro PP, Zhao W, Argaw D. Monitoring the elimination of human African trypanosomiasis: Update to 2016. *PLoS Negl Trop Dis*. 2018;12(12):0006890.
60. Sutherland CS, Tediosi F. Is the elimination of sleeping sickness' affordable? Who will pay the price? Assessing the financial burden for the elimination of human African trypanosomiasis *Trypanosoma brucei* gambiense in sub-Saharan Africa. *BMJ Glob Health*. 2019;4(2):1–11.

61. Franco JR, Simarro PP, Diarra A, Jannin JG. Epidemiology of human african trypanosomiasis. *Clin Epidemiol*. 2014;6:257.
62. Stone CM, Chitnis N. Implications of heterogeneous biting exposure and animal hosts on trypanosomiasis brucei gambiense transmission and control. *PLoS Comput Biol*. 2015;11(10):1004514.
63. Kennedy PG. Clinical features, diagnosis, and treatment of human african trypanosomiasis (sleeping sickness). *Lancet Neurol*. 2013;12(2):186–94.
64. Haines LR. Examining the tsetse teneral phenomenon and permissiveness to trypanosome infection. *Front Cell Infect Microbiol*. 2013;3:84.
65. Stanton MC, Esterhuizen J, Tirados I, Betts H, Torr SJ. The development of high resolution maps of tsetse abundance to guide interventions against human African trypanosomiasis in northern Uganda. *Parasites Vectors*. 2018;11(1):340.
66. Castaño MS, Aliee M, Mwamba Miaka E, Keeling MJ, Chitnis N, Rock KS. Screening strategies for a sustainable endpoint for gambiense sleeping sickness. *J Infect Dis*. 2020;221(Supplement_5):539–45.
67. Steinmann P, Stone CM, Sutherland CS, Tanner M, Tediosi F. Contemporary and emerging strategies for eliminating human African trypanosomiasis due to *Trypanosoma brucei gambiense*: Review. *Trop Med Int Health*. 2015;20(6):707–18.
68. Lea RA. World Development Report 1993: 'Investing in Health'. *Forum Dev Stud*. 1993;20(1):114–7.
69. Murray C. Quantifying the burden of disease : the technical basis for disability-adjusted life years. *Bull World Health Organ*. 1994;72(3):429–45.
70. Hutubessy R, Chisholm D, Edejer TT-T. Generalized cost-effectiveness analysis for national-level priority-setting in the health sector. *Cost Effect Resour Allocation*. 2003;1(1):8.
71. Leech AA, Kim DD, Cohen JT, Neumann PJ. Use and misuse of cost-effectiveness analysis thresholds in low-and middle-income countries: trends in cost-per-daly studies. *Value Health*. 2018;21(7):759–61.
72. Woods B, Revill P, Sculpher M, Claxton K. Country-level cost-effectiveness thresholds: Initial estimates and the need for further research. *Value Health*. 2016;19(8):929–35.
73. Ochalek J, Lomas J, Claxton K. Estimating health opportunity costs in low-income and middle-income countries: a novel approach and evidence from cross-country data. *BMJ Glob health*. 2018;3(6):e000964.
74. Bertram MY, Lauer JA, Joncheere KD, Edejer T, Hutubessy R, Kieny P, Hill SR, Bertram MY. Cost – effectiveness thresholds: pros and cons Use and misuse of thresholds. *Bull World Health Organ*. 2016;94(July):925–30.
75. Marseille E, Larson B, Kazi DS, Kahn JG, Rosen S. Thresholds for the cost–effectiveness of interventions: Alternative approaches. *Bull World Health Organ*. 2015;93(2):118–24.
76. Wall RJ, Rico E, Lukac I, Zuccotto F, Elg S, Gilbert IH, Freund Y, Alley MRK, Field MC, Wyllie S, Horn D. Clinical and veterinary trypanocidal benzoxaboroles target CPSF3. *Proc Natl Acad Sci USA*. 2018;115(38):9616–21.
77. The World Bank. GDP per capita (current US\$) - Congo, Dem. Rep. World Bank national accounts data, and OECD National Accounts data files. 2018. <https://data.worldbank.org/indicator/NY.GDP.PCAP.CD?locations=CD> (Accessed 2020).
78. Louis F, Kohagne LT, Ebo'OVE, Simarro P. Organizing an active screening campaign for human African trypanosomiasis due to *Trypanosoma brucei gambiense*. *Méd Trop revue du Corps de santé colonial*. 2008;68(1):11–16.
79. Welburn S, Maudlin I, Simarro P. Controlling sleeping sickness—a review. *Parasitology*. 2009;136(14):1943–9.
80. Fenwick E, Claxton K, Sculpher M. Representing uncertainty: The role of cost-effectiveness acceptability curves. *Health Econ*. 2001;10(8):779–87.
81. Barton GR, Briggs AH, Fenwick EAL. Optimal cost-effectiveness decisions: The role of the cost-effectiveness acceptability curve (CEAC), the cost-effectiveness acceptability frontier (CEAF), and the expected value of perfection information (EVPI). *Value Health*. 2008;11(5):886–97.
82. Hollingsworth TD, Adams ER, Anderson RM, Atkins K, Bartsch S, Basáñez M-g, Behrend M, Blok DJ, Chapman LAC, Coffeng L, Courtenay O, Crump RE, Vlas SJD, Dobson A, Dyson L, Farkas H, Galvani AP, Gambhir M, Gurarie D, Irvine MA, Jervis S, Keeling MJ, Kelly-hope L, King C, Lee BY, Rutte EAL, Lietman TM, Ndeffo-mbah M, Medley GF, Michael E, Pandey A, Peterson JK, Pinsent A, Porco TC, Richardus JH, Reimer L, Rock KS, Singh BK, Stolk W, Swaminathan S, Torr SJ, Townsend J, Truscott J, Walker M, Zoueva A, Consortium NTDM. Quantitative analyses and modelling to support achievement of the 2020 goals for nine neglected tropical diseases. *Parasites Vectors*. 2015;8(1):630.
83. Hasker E, Lutumba P, Chappuis F, Kande V, Potet J, De Weggheleire A, Kambo C, Depoortere E, Pecoul B, Boelaert M. Human african trypanosomiasis in the democratic republic of the congo: a looming emergency?. *PLoS Negl Trop Dis*. 2012;6(12):1950.
84. Ochalek JM, Lomas J, Claxton KP. Cost per DALY averted thresholds for low- and middle-income countries: evidence from cross country data. York: Centre for Health Economics, University of York; 2015, pp. 1–50. (CHE Research Paper; 122).

Publisher's Note

Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.

Ready to submit your research? Choose BMC and benefit from:

- fast, convenient online submission
- thorough peer review by experienced researchers in your field
- rapid publication on acceptance
- support for research data, including large and complex data types
- gold Open Access which fosters wider collaboration and increased citations
- maximum visibility for your research: over 100M website views per year

At BMC, research is always in progress.

Learn more biomedcentral.com/submissions

