This is the peer reviewed version of the following article:

Gutiérrez, JM, Maraun, D, Widmann, M, et al. An intercomparison of a large ensemble of statistical downscaling methods over Europe: Results from the VALUE perfect predictor cross-validation experiment. Int. J. Climatol. 2019; 39: 3750– 3785. https://doi.org/10.1002/joc.5462,

which has been published in final form at https://doi.org/10.1002/joc.5462. This article may be used for non-commercial purposes in accordance with Wiley Terms and Conditions for Use of Self-Archived Versions.

An intercomparison of a large ensemble of statistical downscaling methods over Europe: Results from the VALUE perfect predictor cross-validation experiment

J.M. Gutiérrez⁽¹⁾*, D. Maraun⁽²⁾, M. Widman⁽³⁾, R. Huth⁽⁴⁾, E. Hertig⁽⁵⁾, R. Benestad⁽⁶⁾, O. Roessler⁽⁷⁾, J. Wibig⁽⁸⁾, R. Wilcke⁽⁹⁾, S. Kotlarski⁽¹⁰⁾, D. San Martín^(1,11), S. Herrera⁽¹²⁾, J. Bedia⁽¹⁾, A. Casanueva⁽¹²⁾, R. Manzanas⁽¹⁾, M. Iturbide⁽¹⁾, M. Vrac⁽¹³⁾, M. Dubrovsky⁽¹⁴⁾, J. Ribalaygua⁽¹⁵⁾, J. Pórtoles⁽¹⁵⁾, O. Räty⁽¹⁶⁾, J. Räisänen⁽¹⁶⁾, B. Hingray⁽¹⁷⁾, D. Raynaud⁽¹⁷⁾, M. J. Casado⁽¹⁹⁾, P. Ramos⁽¹⁹⁾, T. Zerenner⁽²⁰⁾, M. Turco⁽²¹⁾, T. Bosshard⁽²²⁾, P. Štěpánek⁽²³⁾, J. Bartholy⁽²⁴⁾, R. Pongracz⁽²⁴⁾, D.E. Keller^(10,25), A.M. Fischer⁽¹⁰⁾, R.M. Cardoso⁽²⁶⁾, P.M.M. Soares⁽²⁶⁾, B. Czernecki⁽²⁷⁾, C. Pagé⁽²⁸⁾ ⁽¹⁾Meteorology Group. Instituto de Física de Cantabria, CSIC-Univ. of Cantabria, Spain. ⁽²⁾Wegener Center for Climate and Global Change, University of Graz, Austria. ⁽³⁾School of Geography, Earth and Environmental Sciences, University of Birmingham, UK. ⁽⁴⁾Dept. of Physical Geography and Geoecology, Faculty of Science, Charles University; and Institute of Atmospheric Physics, Czech Academy of Sciences, Czech Republic. ⁽⁵⁾Institute of Geography, University of Augsburg, Germany. ⁽⁶⁾The Norwegian Meteorological Institute, Norway. ⁽⁷⁾Department of Geography / Oeschger Centre for Climate Change Research, University of Bern, Switzerland. ⁽⁸⁾Department of Meteorology and Climatology, University of Lodz, Poland. ⁽⁹⁾Rossby Centre, Swedish Meteorological and Hydrological Institute, Sweden. ⁽¹⁰⁾Federal Office of Meteorology and Climatology MeteoSwiss, Switzerland. ⁽¹¹⁾Predictia Intelligent Data Solutions, SME. Spain. ⁽¹²⁾Meteorology Group. Dpto. de Matemática Aplicada y Computación. Univ. of Cantabria, Spain. ⁽¹³⁾Laboratoire des Sciences du Climat et de l'Environnement (LSCE-IPSL/CNRS), France. ⁽¹⁴⁾Institute of Atmospheric Physics, Czech Academy of Sciences, Czech Republic. ⁽¹⁵⁾Fundación para la Investigación del Clima (FIC), Spain. ⁽¹⁶⁾University of Helsinki (UHEL), Finland. ⁽¹⁷⁾Univ. Grenoble Alpes, CNRS, IRD, Grenoble INP, IGE, France ⁽¹⁹⁾Agencia Estatal de Meteorologia (AEMET), Spain. ⁽²⁰⁾Meteorological Institute. University of Bonn, Germany. ⁽²¹⁾Department of Applied Physics, University of Barcelona, Spain. ⁽²²⁾Swedish Meteorological and Hydrological Institute (SMHI), Sweden. ⁽²³⁾Global Change Research Institute CAS, Czech Republic. ⁽²⁴⁾Eötvös Loránd University (ELU), Hungary. ⁽²⁵⁾Center for Climate Systems Modeling (C2SM), ETH Zurich, Switzerland. ⁽²⁶⁾Instituto Dom Luiz, Universidade de Lisboa (IDL), Portugal. ⁽²⁷⁾Adam Mickiewicz University, Poland. ⁽²⁸⁾CECI, CERFACS - CNRS, France.

4

^{*}Corresponding author. Email: gutierjm@unican.es

Abstract

VALUE is an open European collaboration to intercompare downscaling approaches 6 for climate change research, focusing on different validation aspects (marginal, tem-7 poral, extremes, spatial, process-based, etc.). Here we describe the participating meth-8 ods and first results from the first experiment, using "perfect" reanalysis (and RCM 9 reanalysis-driven) predictors to assess the intrinsic performance of the methods for 10 downscaling precipitation and temperatures over a set of 86 stations representative of 11 the main climatic regions in Europe. This study constitutes the largest and most com-12 prehensive to date intercomparison of statistical downscaling downscaling methods, 13 covering the three common downscaling approaches (perfect prognosis, model output 14 statistics —including bias correction— and weather generators) with a total of over 15 fifty downscaling methods representative of the most common techniques. 16

Overall, most of the downscaling methods greatly improve raw model biases and 17 no approach or technique seems to be superior in general, since there is a large method-18 to-method variability. The main factors most influencing the results are the seasonal 19 calibration of the methods (e.g. using a moving window) and their stochastic nature. 20 The particular predictors used also played an important role in cases where the compar-21 ison was possible, both for the validation results and for the strength of the predictor-22 predictand link, indicating the local variability explained. However, the present study 23 cannot give a conclusive assessment of the skill of the methods to simulate regional 24 future climates, and further experiments will be soon performed in the framework of 25 the EURO-CORDEX initiative (where VALUE activities have merged and follow on). 26

Finally, research transparency and reproducibility has been a major concern and substantive steps have been taken. In particular, the necessary data to run the experiments is provided at http://www.value-cost.eu/data and data and validation results are available from the VALUE Validation Portal for further investigation: http://www.value-cost.eu/validationportal.

KEY WORDS: downscaling, bias adjustment, perfect prognosis, model output statistics, weather generators, validation, reproducibility, CORDEX.

1. Introduction

5

Global Climate Models (GCMs) are the primary tools to simulate multi-decadal climate dynamics and to 35 generate global climate change projections under different future emission scenarios (Taylor et al. 2011). 36 However, these models have a coarse resolution (typically a few hundred kilometers) and suffer from sub-37 stantial systematic biases when compared with observations (Flato et al. 2013, Sec. 9.6). Therefore, they 38 are unable to provide actionable information at the regional and local spatial scales required in impact and 39 adaptation studies. In order to bridge this gap, two main downscaling approaches have been developed 40 since the early 1990s (Leung et al. 2003; Maraun et al. 2010): Dynamical downscaling (based on Regional 41 Climate Models, RCMs) and Empirical/Statistical Downscaling (ESD, based on statistical models). The rel-42 ative merits and limitations of both dynamical and statistical downscaling —and combinations of them,— 43 have been widely discussed in the literature (see, e.g., Fowler et al. 2007; Maraun et al. 2010; Winkler et al. 44 2011; Takayabu et al. 2016) and it is nowadays recognized that they are complementary in many practical 45 applications. 46 Dynamical downscaling is carried out running one or several RCMs on a relatively fine grid (e.g. 10-47

⁴⁷ Dynamical downscaling is carried out running one of several KCMs on a relatively fine grid (e.g. 10-⁴⁸ 20 km) over a limited domain (e.g. Europe) initialized and driven at the boundaries by the coarse GCM ⁴⁹ outputs to be downscaled (Giorgi and Mearns 1991; Rummukainen 2010, for a review). These models

are able to generate regional physically-consistent predictions for a suite of climate variables (particularly 50 those less affected by model parameterizations), but still may suffer from significant biases (see, e.g. Kot-51 larski et al. 2014; Casanueva et al. 2016b) which require statistical post-processing before they can be used 52 in impact applications. Ensembles of RCMs have been extensively intercompared in the framework of a 53 series of subsequent community intercomparison initiatives considering increasing resolutions, e.g. PRU-54 DENCE (0.44°, Christensen and Christensen 2007), ENSEMBLES (0.22°, Christensen et al. 2010) and, 55 more recently, EURO-CORDEX (0.11°, Jacob et al. 2014), all focusing on Europe. These experiments 56 57 performance of the different RCMs,- and GCM driven simulations under different (historical and future) 58 scenarios. 59 ESD methods rely on statistical models linking informative GCM outputs (predictors) to the local ob-60 served predictand of interest over a particular domain (Benestad et al. 2008; Maraun and Widmann 2017). 61 These models are first trained (and tested, e.g., cross-validated) using model and observed data during a 62

representative historical period, and later applied to new (e.g. future) GCM data to obtain the downscaled local predictions. According to the nature of predictors in the training phase, three main approaches for

ESD exist (see, e.g. Maraun et al. 2010): 1) Perfect Prognosis (PP), 2) Model Output Statistics (MOS) —

⁶⁶ including the increasingly popular Bias Correction (BC) techniques,— and 3) Weather Generators (WG).

⁶⁷ On the one hand, under the PP approach, quasi-observed predictors from reanalysis are used to train the ⁶⁸ statistical models based on their temporal (e.g. daily or monthly) correspondence with observations in the

historical training period. Therefore, predictor variables well represented by both reanalyses and GCMs, and
 accounting for a major part of the variability in the predictands, are typically chosen in this approach (usually

⁷¹ large-scale variables at different vertical levels), whereas variables directly influenced by model parameteri-

zations and/or orography (e.g. precipitation) are usually discarded (Wilby et al. 2004). As a result, one of the
 most time-consuming tasks of this approach is the selection of a suitable combination of predictors, defined

⁷³ over a particular geographical domain which encompasses the main synoptic phenomena influencing the

⁷⁵ climate of the region of interest. On the other hand, under the MOS approach, model outputs from the GCM

⁷⁶ are directly used for training, thus correcting systematic biases against observations. In particular, simple

MOS alternatives based on BC techniques are becoming increasingly popular in climate change applications
 to adjust both GCM and RCM outputs (see, e.g., Themeßl et al. 2012). These techniques adjust the model

⁷⁸ to adjust both GCM and RCM outputs (see, e.g., Themeßl et al. 2012). These techniques adjust the model ⁷⁹ output distribution towards the observed one to ensure resemblance to the local climatology. The main

advantage of MOS techniques is their simplicity, since no predictor/domain screening is typically required

81 (e.g. GCM output for the target variable from the closest model gridbox is commonly considered as the

⁸² unique predictor). Finally, WG is a third approach which does not explicitly include GCM predictors in the

training phase (Wilks and Wilby 1999). The simplest form of WGs are Markov-like processess fitted to the
 local observed data, which are able to reproduce the local temporal and marginal statistical properties from

⁸⁴ local observed data, which are able to reproduce the local temporal and marginal statistical properties from ⁸⁵ a set of parameters derived from basic climatological statistics (e.g. autocorrelation, wet-day frequency,

mean and standard deviation). The global climate change signals from the GCMs are later temporally dis-

aggregated by producing daily time series from the WG with the parameters transformed according to the projected statistics.

As a result of the intensive research activity carried out in this field during the last two decades, a large 89 number of studies exist mostly describing specific ESD methods and/or applications in different regions of 90 the world, using different validation methodologies and/or experimental frameworks. There are also some 91 intercomparison studies focusing on particular approaches, either PP (Haylock et al. 2006; Frost et al. 2011; 92 Teutschbein et al. 2011; Hu et al. 2013; Gutiérrez et al. 2013; San-Martín et al. 2017), MOS (Teutschbein and 93 Seibert 2013; Gutmann et al. 2014), or WG (Semenov et al. 1998; Hartkamp et al. 2003). Moreover, a few 94 multi-approach intercomparison studies are also available, starting with the pioneering work by Wilby et al. 95 (1998) who analyzed PP and WG methods, and following with the more recent PP and MOS comparisons 96 by Bürger et al. (2012) and Vaittinada Ayar et al. (2016). However, limited comprehensive information is 97 yet available at a continental level (e.g. over Europe) for the informed application of the different ESD 98

⁹⁹ approaches for climate change impact and adaptation studies.

¹⁰⁰ The EU Cooperation in Science and Technology (COST) Action ES1102 VALUE (2012-2015, Ma-

raun et al. 2015, http://www.value-cost.eu) has been the first international initiative to create a 101 community for statistical downscaling intercomparison, providing a common experimental framework and 102 developing community validation tools for different validation aspects (marginal statistics, temporal struc-103 ture, extremes, spatial coherence, process based). Several experiments have been designed to isolate specific 104 points in the downscaling procedure where problems may occur (Maraun et al. 2015). In this paper we de-105 scribe the cross-validation experiment with "perfect" reanalysis — and RCM reanalysis driven — predictors 106 to downscale precipitation and temperatures over Europe, with over 50 different participating methods, cov-107 ering the three downscaling approaches (PP, MOS and WG) and the common techniques (quantile mapping, 108 analogs, linear and generalized linear regression, weather typing, etc.). The present paper describes in de-109 tail the contributing methods, analyzing the selection and transformation of predictors and the geographical 110 extent, and their influence on the resulting predictor-predictand relationships. This paper also focuses on 111 key method characteristics (e.g. deterministic/stochastic) and implementation details (e.g. seasonal/annual 112 train) which may be relevant for the analysis of the validation results (and is used as metadata in this work). 113 In this contribution we only focus on validation results for the marginal distributions (biases in the mean and 114 the standard deviation of the distributions), but other validation aspects are analyzed in the different papers 115 of this special issue. 116

Overall, this work constitutes the most comprehensive to date intercomparison of downscaling methods 117 on a continental scale over Europe. We want to remark that the final goal is not ranking the different 118 methods according to their performance, but providing an indication of the relative merits and limitations 119 of the different approaches and families of techniques. Thus, some clearly poor performing methods have 120 been also included to illustrate problems. We want to remark that this experiment alone is not sufficient 121 to evaluate the limitations of (MOS) bias correction techniques (see Maraun et al. 2017, for more details). 122 Moreover, it also does not fully validate PP techniques since further results using GCM predictors are needed 123 to evaluate whether well-represented predictors have been used and the PP assumption is valid. Moreover, 124 this work provides no information on the the extrapolation capabilities (to future climates) of the different 125 MOS and PP techniques (although the reproduction of reanalysis trends is analyzed in Maraun et al. 2018; 126 in this special issue). These problems will be analyzed in subsequent community-open experiments using 127 GCM predictors from historical and future scenarios, which will be open for participation in the framework 128 of the EURO-CORDEX initiative (where VALUE activities have merged and follow on). 129

Research transparency and reproducibility has been a major concern in this work and substantive steps have been taken to improve the reproducibility of the methods and results, and to promote awareness within the downscaling scientific community. In particular, the necessary data to run the experiments is provided at http://www.value-cost.eu/data, and both the downscaled data and the individual validation results are available at the VALUE validation portal http://www.value-cost.eu/ validationportal.

The paper is organized as follows. The experimental framework followed and the predictor and predic-136 tand data used are described in Sec. 2. Section 3 presents the methods contributing to this study (a brief 137 description and specific implementation details for each method are given in Annex 1). It also describes 138 the selection of the predictors and data preparation and analyzes the predictor-predictand link established 139 by the different methods. Sections 4 and 5 presents the validation results for precipitation and temperatures. 140 respectively, focusing on the biases in the mean and the standard deviation resulting from the methods. Infor-141 mation regarding transparency and reproducibility of results is given in Sec. 6. Finally, the main conclusions 142 obtained are reported in Sec. 7. 143

2. Experimental Framework and Data

In this section we briefly describe the experimental framework. In order to promote research transparency and reproducibility the data described in this section is available at http://www.value-cost.eu/ data. Further information on the VALUE experimental design is given in Maraun et al. (2015).

148 **a.** Predictands: Local observations

A subset of stations covering the different European climates and regions with a homogeneous density was 149 selected to enable a comprehensive validation revealing relative strengths and weaknesses of different meth-150 ods. To keep the exercise as open as possible, the downloadable (blended) ECA&D stations (Klein Tank 151 et al. 2002) was selected and downloaded (on September 2014). A subset of 86 stations was selected with 152 the help of local experts in the different countries, building on high-quality stations with no more than 153 5% of missing values in the analysis period (1979-2008); see http://www.value-cost.eu/WG2 154 dailystations for more details. The resulting set of stations is listed in Table 1 and graphically dis-155 played in Figure 1. The Köppen–Geiger climate type (see, e.g. Kottek et al. 2006) shown in Table 1 has been 156 calculated directly from the data using MeteoLab http://meteo.unican.es/trac/MLToolbox/ 157 wiki. Figure 1 shows the eight PRUDENCE (Christensen and Christensen 2007) sub-regions used to com-158 bine and summarize the individual validation results at a sub-regional level along the paper. These regions 159 are British Isles, Iberia, France, Central Europe, Scandinavia, the Alps, the Mediterranean, and Eastern 160 Europe. 161

The resulting dataset (including daily data for precipitation and temperatures for the analysis period) is publicly available in text (csv) format at http://www.value-cost.eu/data.

b. *Reanalysis Predictors*

ERA-Interim (Dee et al. 2011) was selected by the CORDEX initiative as the reference reanalysis for the coordinated downscaling experiments. Therefore, in order to be aligned with this initiative, VALUE also used ERA-Interim to drive the experiment with "perfect" predictors. Although reanalysis uncertainty has been recently reported as an additional source of uncertainty for statistical downscaling (Brands et al. 2012), the effect on the downscaled results is relevant only in the tropics (Manzanas et al. 2015). Therefore, this factor was not tested in VALUE.

In order to keep the experimental framework as controlled as possible and to facilitate the work of 171 the contributing groups, we generated a reference predictor dataset downloading ERA-Interim data from 172 ECMWF's MARS. The dataset includes a reduced number of commonly used predictors, degraded to a 173 common 2° grid and post-processed by computing daily means from the original 6 hourly fields when 174 required (see Table 2). This reference dataset includes most of the circulation and thermodynamic predictors 175 at different pressure levels (including some surface predictors), typically used in downscaling applications 176 in different European regions (Huth 1999; Benestad 2002; Huth 2002; Timbal et al. 2003; Huth 2005; 177 HanssenBauer et al. 2005; Gutiérrez et al. 2013; Hertig et al. 2014; San-Martín et al. 2017), excluding 178 redundancy as much as possible. For instance, vorticity and divergence have been considered as potential 179 predictors in the literature (see, e.g. Hessami et al. 2008), but they were excluded from the standard set 180 since they reported similar results to geopotential or wind directions in some studies (Gutiérrez et al. 2013). 181 However, some contributors used in-house ERA-Interim datasets instead, for convenience or because they 182 needed extra predictors (see Sec. b for more details). 183

Since MOS methods typically work with the direct model output at the nearest gridbox to the target station, we also compiled surface precipitation (PRC) and minimum (TMIN) and maximum (TMAX) temperature from the original ERA-Interim dataset at 0.75°. In order to illustrate the effect of the model resolution on the results, in the analysis we will consider raw ERA-Interim outputs at two different resolutions: 2 and 0.75° (hereafter referred to as ERAINT-200 and ERAINT-075, respectively).

c. *RCM Predictors (for MOS methods)*

Since MOS methods are typically applied to both GCM and RCM outputs, a second (optional) predictor dataset for MOS methods was produced considering daily surface precipitation (PRC) and minimum (TMIN) and maximum (TMAX) temperature from a state-of-the-art RCM (the RACMO2 model) driven in climatic mode by ERA-Interim (see Meijgaard et al. 2012, for a detailed description of the model). This simulation was produced in the framework of the EURO-CORDEX project (Jacob et al. 2014) using 40 hybrid coordinate full vertical levels on a regional 0.11° domain over Europe. RACMO2 ranked among the
 best performing RCMs over Europe in this reanalysis-driven experiment (Kotlarski et al. 2014).

The MOS techniques contributing to this additional experiment are indicated with a check mark in the second column (labelled as 'R') of Tables 3 and 4. This experiment will allow analyzing the advantages and shortcomings of these methods when downscaling to finer spatial scales. Note that a weak day-to-day correspondence with observations is expected for the RACMO2 outputs, since temporal synchrony with observations is only induced by the boundary conditions with prescribed reanalysis values. This must be taken into account when analyzing the results of non-distributional MOS methods, which are better suited for nudged climate simulations with a strong synchrony with observations (see, e.g. Eden et al. 2014).

204 **d.** Cross-validation Approach

In order to appropriately assess and compare the performance of different downscaling methods with "per-205 fect" predictor data, we applied a cross-validation approach to avoid model overfitting and artificial skill. 206 Cross-validation allows us to test whether the relationship established between predictor and predictand 207 remains valid outside the training period (e.g. in a test period). The most popular and simplest of these ap-208 proaches is data splitting, which considers independent data for training (e.g. 80% of the available data) and 209 validation (e.g. the remaining 20%). However, this can yield spurious effects due to the particular partition 210 performed. K-folding methods attempt to produce a more rigorous validation through the use of multiple 211 calibration/validation period combinations. This is done by partitioning the available data (n = 30 years 212 in our study) into k non-overlapping "folds" or subsets, each containing n/k elements. The downscaling 213 methods are then calibrated and validated k times, considering in turn each of the folds as a test set and 214 training the method with the remaining k-1 ones. The resulting k test series are typically joined and val-215 idated together in a single series spanning the whole analysis period. This approach also permits analyzing 216 the variability of the k validation results and estimating confidence intervals for model performance (see, 217 e.g. Gutiérrez et al. 2013). 218

In general, the selection of k is subject to a number of factors depending on the particular application. 219 In the present case, low k values result in longer validation periods which may be desirable to better charac-220 terize model performance, but limit at the same time the data available for training which may only capture 221 part of the climatological distribution. As a compromise, five folds (5-fold cross-validation) were consid-222 ered, each containing 6 consecutive years (1979-1984, 1985-1990, 1991-1996, 1997-2002, 2003-2008) for 223 validation. All contributed methods followed this approach and joined together the results downscaled for 224 the five test periods into a unique series —covering the whole thirty-years period— which was uploaded 225 to the VALUE validation portal and automatically validated to assess model performance. More details are 226 given in Sec. 6. 227

228 e. Validation Measures

We analyze minimum and maximum temperatures (TMIN and TMAX) and precipitation (PRC). For the 229 latter we consider separately occurrence and amount, analyzing the variables R01 (Relative wet-day fre-230 quency, $PRC \ge 1$ mm) and SDII (mean wet-day precipitation, a.k.a. Simple Day Intensity Index), although 231 we also consider the total precipitation amount (PRCTOT) in some illustrative cases. In this paper we fo-232 cus on general validation aspects involving the observed and predicted marginal distributions. In particular 233 we validate the biases in the mean and the standard deviation, although additional results on distributional 234 similarity (Kolmogorov-Smirnov and Cramer-von Mises tests) have been computed (not shown) and are 235 available through the VALUE portal http://www.value-cost.eu/validationportal for fur-236 ther research. The bias in the mean is computed as the difference (for TMIN, TMAX) or ratio (for R01, 237 SDII and PRCTOT) between the downscaled and the observed mean values, whereas the bias in the stan-238 dard devidation is always obtained as the ratio. 239

Moreover, in order to analyze the strength of the daily predictor-predictand link (informative for nondistributional MOS and PP methods), we computed the correlation of the daily downscaled and observed series (using the ranked Spearman and Pearson correlations for precipitation and temperatures, respectively).

Further validation analyses of aspects such as the representation of the temporal structure, extremes, key processes and multivariate relationships, are analyzed in detail in separate papers of this special issue.

3. Downscaling Approaches and Methods

a. Description of Contributing Methods

Tables 3 and 4 show the statistical downscaling methods contributing to this work for precipitation and (minimum and maximum) temperatures, respectively, under the same experimental framework (see Sec. 2). This constitutes the largest and most diverse ensemble of ESD methods analyzed to date, with a total of 45/49 methods for precipitation and temperatures, respectively (28 methods have been applied to both precipitation and temperatures, shaded areas). The detailed description of each of these methods and the implementation details for reproducibility (when available) are given in Annex 1.

These methods are first organized according to the three main approaches: MOS, PP and WGs — 253 conditional WGs (including some model predictor) are listed under the corresponding PP or MOS category, 254 depending on how they are calibrated.— Note that the first three rows indicate the raw model data (RAW) 255 for ERA-Interim (both at 2° and 0.75° resolution), and for the (ERA-Interim driven) RACMO2 RCM at 256 0.11° resolution. As a second categorization, the methods are organized within each approach according to 257 the families of techniques used, typically transfer functions (TF), analogs (A), and weather types (WT) for 258 both PP and MOS, and additive/multiplicative scaling (S), parametric quantile mapping (PM), and empir-259 ical quantile mapping (QM) for MOS methods. This organization groups together similar methods (same 260 approach and technique) and allows for a better intercomparison of model results (this order will be used in 261 all figures in this paper). These families are described in further detail below. 262

Tables 3 and 4 also provide some metadata information about the structural properties of the methods 263 (full metadata is available in the VALUE validation portal, http://www.value-cost.eu/validationportal). 264 In particular the column 'ST' indicates the stochastic or deterministic nature of the method ('yes' for stochas-265 tic ones, which contributed 100 realizations for the validation process). 'MS' and 'MV' indicate whether 266 the methods are suitable for multi-site and multi-variable problems, respectively; those methods using Prin-267 cipal Components (PCs) as predictors are marked as 'yes' in the 'MS' column to indicate that some spatial 268 coherence could be imprinted by the predictors. Finally, 'SE' and 'AC' indicate the explicit inclusion of 269 seasonal and autocorrelation model components, respectively. The former is typically achieved by training 270 the models separately for each of the calendar months (or with a 30-day moving window in some MOS 271 methods, see Annex 1 for details). The latter is typically achieved using first-order Markov chains (con-272 ditioning the prediction to the previous predicted value) and has only been used in the contributed WG 273 methods. As a result, the temporal structure of all PP and MOS methods in this experiment is driven by 274 275 the particular model predictors used, i.e. directly from the raw model precipitation and temperature series for MOS methods. This (metadata) information must be taken into account when comparing the evaluation 276 results of different methods, since a method can exhibit good performance for a particular aspect as a result 277 of model construction or fitting (an interesting discussion on fair comparison is given in Casanueva et al. 278 2016a). 279

The participating MOS methods (#4-25, #4-23, for precipitation and temperatures, respectively) com-280 prehensively span the range of widely used methods, from simple local scaling methods (labelled as 'S'), to 281 standard parametric ('PM') and empirical ('QM') quantile mapping techniques. More specific BC methods, 282 such as the trend preserving ISI-MIP bias correction methods, or a circulation-conditioned quantile mapping 283 method (EQM-WT) are also included in this study. These methods are usually referred to as distributional 284 MOS methods, in order to remark that they work by transforming the distribution of daily model outputs (the 285 whole distribution or some statistics) towards the observed one. Moreover, the analysis also covers some 286 more experimental recent MOS developments (#21-23, #22) such as stochastic regression (Wong et al. 2014) 287 and analog- and regression-based MOS methods (Turco et al. 2011, 2017), which exploit the (weak) tempo-288

ral correspondence existing in climatic RCM simulations to establish a link with observations. All the MOS 289 methods in this study are single-site and single-variable, with the exception of MOS-AN and DBS, respec-290 tively, the latter providing inter-variable consistency by downscaling temperature conditional on the wet/dry 291 state of the corresponding precipitation series. Finally, two particular methods for precipitation (FIC02P 292 and FIC04P, #24-25) are based on a sequential application of PP and MOS techniques —the input to these 293 methods is the output of the corresponding FIC01P and FIC03P PP results.— Note that these methods are 294 not directly comparable with the rest of MOS techniques, but provide valuable information on the potential 295 added value of mixed downscaling approaches (e.g. applying BC methods to correct systematic biases of PP 296 outputs). The same situation occurs for temperatures with FIC02T, which takes as input FIC01T PP results. 297 Empirical quantile mapping techniques ('QM') constitute the largest family in this approach, with over 298 ten contributing methods. It is important to remark that some methods are slightly different implementations 299 of the same basic technique (with different number of adjusted percentiles, or extrapolation options; see 300 Annex 1 for details). In particular the empirical methods EQM, QM-DAP, EQM-WIC658, and QMBC-301 BJ-PR are slightly different versions of the standard empirical quantile mapping approach (see, e.g. Déqué 302 2007). Parametric quantile mapping techniques ('PM') mainly differ in the distribution function(s) used 303 to calibrate the data. For instance, in the case of precipitation, a gamma distribution is used in EQM, a 304 double-gamma is used in DBS and in Ratyetal-M9 —which is a simplified version of the former,— the 305 optimum among five distributions is used in BC and, finally, gamma and generalized Pareto are used in 306 GPQM to adjust separately the extremes values. It is also important to notice that most MOS methods have 307 been trained separately for each month (or considering a moving window), with the exception of GOM, 308 GPQM, EQM, EQM-WT, and the MOS-GLM/REG/AN family. Therefore, interesting conclusions could 309 be obtained by comparing the results of QM methods taking into account the different configurations and 310 implementations. 311

The participating PP methods (#26-42,#24-46, for precipitation and temperatures, respectively) broadly 312 represent the most popular and widely used families of techniques —analogs (A), transfer function / regres-313 sion (TF) and weather-type (WT) methods— in fairly standard implementations in most of the cases. The 314 analog techniques are the only multivariate methods (multi-site and/or multi-variable). However, some of 315 the TF methods use PCs as predictors (Sec. b), which may provide some imprinted spatial inter-consistency 316 due to their spatial character. Those cases are indicated with a 'yes' multi-site code in Tables 3 and 4. 317 Nonparametric regression methods (e.g. neural networks) are among the most notorious missing families in 318 this study. In some studies these methods have shown to outperform linear models (see, e.g. Gaitan et al. 319 2014), but there are also studies showing the opposite. Therefore, the VALUE intercomparison framework 320 could provide a better understanding on the added value and limitations of these techniques. The only con-321 tributing machine learning technique is the MO-GP method, which applies genetic programming to obtain 322 general symbolic regression equations from data. Therefore, an interesting follow-on of the project would 323 be including new nonlinear machine learning methods in the intercomparison. 324

The family of Analog (A) methods includes two different variants of the standard technique, consider-325 ing raw fields with no seasonal restriction (ANALOG), and anomalies with seasonal restriction (ANALOG-326 ANOM). FIC and ANALOG-MP/SP methods are more elaborated two-step analog methods considering 327 nested global/local domains and predictors. ANALOG-MP/SP are probabilistic methods which include 328 here a stochastic component to produce the 100 realizations of the predictand from the probabilistic pre-329 diction available each day. Similarity is quantified by Euclidean distance in all cases with the exception 330 of ANALOG-MP/SP, which use the Teweless-Wobus score. WT-WG is a simple stochastic weather typ-331 ing approach simulating temperature/precipitation from gaussian/binomial-gamma distributions within each 332 weather type (obtained using only SLP in this study). 333

The contributing Transfer Function (TF) methods are different variants of Multiple Linear Regression (MLR) techniques and Generalized Linear Models (GLM) and Vectorized GLMs for precipitation. GLMs are an extension of linear regression allowing for non-normal predictand distributions (see Chandler 2005, for an introduction), which have been used for downscaling precipitation in a number of studies (see, e.g., Chandler and Wheater 2002; Abaurrea and Asín 2005). Although MLR has been applied to downscale daily precipitation in previous studies (see, e.g., Hessami et al. 2008; Chen et al. 2014; San-Martín et al. 2017), these techniques are not suitable to model daily precipitation, even after transforming precipitation —using e.g. squared or cubic root values— to make the data more normal. However, we have included them in the present work for illustrative purposes, in order to highlight the associated problems. The different MLR and GLM methods (#26-42, #24-46) have been trained on a daily basis to establish the link with local data with the exception of the ESD family (#39-42, for temperatures, in italics), trained using monthly aggregated data and providing monthly values.— ESD methods are used here for illustrative purposes and results are only shown for suitable validation scores (mean bias).

Most of the TF methods are deterministic, but there are also some stochastic implementations. A particu-347 lar method is provided in both deterministic (GLM-DET) and stochastic (GLM) variants, using the expected 348 value in the former case and simulating from the resulting binomial/gamma in the latter, including also an 349 implementation conditioned on weather types (GLM-WT). GLM-P combines logistic (binomial GLM) and 350 exponential regression to simulate occurrence and amount, but only considers a stochastic version of the for-351 mer one (i.e. only occurrence is simulated). A more sophisticated Vectorised GLM method (Vaittinada Ayar 352 et al. 2016) is used in SWG, which also simulates the predicted values from the resulting conditioned bi-353 nomial/gamma distributions. On the other hand, stochastic MLR versions (MLR-ASW/AAW) are based 354 on variance inflation using white noise. Note these methods can be compared with the simple determin-355 istic scaling variance inflation versions (MLR-ASI/AAI) in order to analyze the benefit of the stochastic 356 component. 357

The participating WG methods (#43-48, #47-52, for precipitation and temperatures, respectively) include variants of the Richardson model (Richardson 1981), simulating daily time-series of precipitation, minimum and maximum temperature using Markov chains (order one for SS-WG and one to three for MARFI) and autoregressive models. Moreover, the analysis also covers a recent non-parametric weather generator (GOMEZ) based on nearest neighbor resampling.

b. Selection of Predictors and Data Preparation

Selection of predictors and data preparation is a key task for statistical downscaling, in particular for PP 364 methods. Whereas this task is quite simple for distributional MOS methods —which operate directly with 365 model precipitation/temperature, typically on the nearest gridbox, as the single predictor,— the selection of 366 informative predictors for PP methods is a key factor both for model performance and for ability to extrapo-367 late under climate change conditions (Huth 2004; Gutiérrez et al. 2013; San-Martín et al. 2017). Therefore, 368 a region-dependent screening of suitable predictors over different (large or small) domains covering the area 369 of study is usually performed as a first step of the downscaling process. In some cases, this task is automat-370 ically performed applying some variable selection method, such as stepwise screening, which is applied in 371 most TF methods as described in Annex 1 (e.g. MLR-RSN/RAN/AAN/AAI/AAW/ASI/ASW). Therefore, 372 the final set of predictors used in these methods may change from variable to variable and from station to 373 station. 374

Several studies have shown the convenience of combining circulation and thermodynamic predictors in order to include signal-carrying predictors linked to changes in the radiation budget, avoiding to model future climate from changes in circulation alone (Wilby et al. 1998; Huth 2004). Therefore, the final decision about the predictors to be used in a particular region needs to be based on the physical understanding of the problem. The predictors must also be skillfully predicted by GCMs in terms of the statistical characteristics of the large scales (e.g., spatial and temporal structures). Ideally, they should also exhibit a strong link with the local variable in order to represent the large-scale dependency.

Table 5 shows the particular combinations of predictors used by the different PP methods in this study (Tables 3 and 4). Besides the standard variables shown in Table 2, some contributors have considered additional predictors, such as ten meter zonal and meridional wind direction (U10, V10), two meter dewpoint temperature (TD), vertical velocity (VV), relative humidity (R), or thickness between two pressure levels (TH). Only a few methods (WT-WG, FIC01P and the ESD family) use either circulation or thermodynamic predictors alone, whereas the rest of methods build on combinations of circulation predictors and middletroposphere temperature and/or humidity, which have been found among the best predictors for temperatures and precipitation, respectively. Note that the predictors are inhomogeneous (i.e. not directly related to the target variable to be downscaled) in all cases for precipitation, and in most of the cases for temperature (with the exception of those including T^2 as predictor which may be considered an homogeneous predictor for minimum and maximum temperatures).

Table 5 also shows the size of the domain used to define the predictors, ranging from continental scale, 393 to smaller national-wide domains, and to local information at the nearest gridboxes (or combinations of 394 them). Note that most of the PP methods consider national or continental-wide information since the use 395 single gridbox information for large-scale variables is not recommended due to the minimum skillful scale 396 of climate models (see, e.g. Takayabu et al. 2016, for more details). The resulting data (values of the 397 predictors for multiple gridboxes) is preprocessed in different ways before using it to train the downscaling 398 methods. Ten regression methods applied individual or combined EOF analysis to obtain PC predictors 399 in order to reduce the spatial redundancy. The rest of regression methods consider raw, standardized, or 400 anomaly point-wise values and in most of the cases they apply a step-wise procedure for predictor selection 401 (see Annex 1 for details). In this case, the models resulting for different stations may be based on different 402 (local) predictors, normally at gridboxes close to the particular station. This constitutes a key factor when 403 validating spatial aspects of the predictions (see Widmann et al. 2017, in this special issue). 404

Most analog methods consider national-wide information (with the exception of one, which is a applied at a continental level) and use raw data, anomalies, standardized values or PCs to compute the similarity of different fields. Four of the methods are two-step implementations which consider different large-scale and local predictors in nested national- and gridbox-wide domains, respectively.

Finally, we want to remark that although the ESD family of methods is based on common EOFs (both reanalysis and GCM fields are used to compute the EOFs, Benestad et al. 2008) the approach used in this paper applies standard EOFs obtained from ERA-Interim.

412 **c.** Strength of the Predictor-Predictand Link

PP and non-distributional MOS methods build on a synchronous daily link established between predictor(s) 413 and predictand in the training phase. The strength of this link indicates the local variability explained 414 by the method as a function of the large-scale predictors. In order to provide a quick diagnostic of this 415 strength for the different methods, Figures 2 and 3 show the daily Spearman and Pearson correlations for 416 the downscaled and observed daily precipitation and maximum temperature values, respectively. The results 417 for the raw model outputs (indicated as ERAINT-200, -075 and RACMO22E) are included in the first three 418 columns of the figures and show the comparison with the local observations considering the raw model 419 values at the closest gridbox. Note that these figures are only informative for PP and non-distributional 420 MOS methods since, on the one hand, WG methods have no daily correspondence with the observed data 421 -they are purely stochastic and use no model predictors— and, on the other hand, distributional MOS 422 methods broadly preserve the temporal structure of the raw model predictor. Therefore, distributional MOS 423 and WG results are included in the figures for illustrative purposes, in order to contrast the expected results 424 and to identify potential problems. 425

As expected, distributional MOS methods closely reproduce the correlation of the corresponding model 426 predictors in most of the cases. The most notorious deviation is the EQM-WIC658 model, which in principle 427 is similar to other implementations of the empirical quantile mapping (e.g. EOM) and therefore is suspected 428 of having an error. There are also noticeable differences for the CDFt model (the case using ERA-Interim, 429 particularly for temperature), which may be due to the particular approach followed to correct the data (see 430 Annex 1 for more details) or to a problem with temporal arrangement of the downscaled data. Furthermore, 431 the ISI-MIP model exhibits smaller correlation than the raw model output for precipitation, which could be 432 explained by the two step process followed, adjusting first the monthly values and then the daily residuals. 433 Finally, on the other hand, WG methods exhibit close to zero correlations in all cases, as expected. 434

Note that the RACMO2 model (and the MOS results obtained using this predictor, with gray shading in the figures) show smaller correlations than ERA-Interim —which exhibits similar results for the two resolutions considered.— This is due to the climatic nature of the simulation, since day-to-day cor-

respondence with observations is only prescribed at the boundaries of the regional simulation domain. 438 Therefore, the non-distributional MOS analog (MOS-AN) and transfer function methods (MOS-GLM/REG. 439 VGLMGAMMA) —which exploit the (weak) temporal correspondence existing between RACMO2 outputs 440 and observations— exhibit smaller correlations than the PM and QM techniques. When applied to ERA-441 Interim, these techniques result in similar (MOS-GLM/REG), or even higher (MOS-AN), correlations when 442 compared with their PP counterparts (GLM/MLR and ANALOG, respectively). Note that in this case, the 443 methods are in fact homogeneous (using precipitation or temperature as predictor) versions of the PP meth-444 ods (considering a single gridbox instead of PCs in the case of MOS-GLM/REG). The higher correlation of 445 the MOS-AN for precipitation in this case is explained by the use of model precipitation as single predictor 446 (Widmann et al. 2003), which is superior to the predictors used by the PP ANALOG version (see Table 5). 447 However, differently to the analog MOS version, the MOS regression methods (MOS-GLM/MLR) result in 448 very small correlations when applied to the weakly synchronized RCM outputs. This may be due to the use 449 of a single gridbox, more sensitive to the weak temporal correspondence with observations. 450

The range of correlations corresponding to the PP methods are mainly due to the different predictor set-451 tings used and to the deterministic/stochastic character of the methods. For instance, the stochastic versions 452 ANALOG-MP/SP, VGLMGAMMA, GLM-P, GLM, GLM-WT, WT-WG and SWG exhibit smaller corre-453 lations due to the stochastic component. Moreover, linear regression methods using white noise variance 454 correction (MLR-ASW/AAW) exhibit smaller correlation than the standard (MLR-AAN) or the inflation 455 variance correction (MLR-ASI/AAI) implementations (see Annex 1 for details). It is noticeable that the 456 stochastic GLM method still preserves a strong correlation when compared to the deterministic implemen-457 tation (GLM-det), indicating that most of the information given by the predictors is still retained in the 458 stochastic implementation. Regarding the analog methods, the smallest correlations are obtained with the 459 method using anomalies (ANALOG-ANOM). Finally, the last two PP methods (WT-WG and SWG) exhibit 460 low correlation values, since they have been designed to have a strong stochastic component weakly forced 461 by the predictors. In particular, the correlation of the WT-WG method is similar to that of the un-conditioned 462 WGs, indicating that this method is purely stochastic (the weather types obtained solely from SLP do not 463 play a relevant conditioning role in this case). Therefore, they can be thought of as weather generators 464 weakly conditioned on circulation. 465

In the case of maximum temperature, high correlations are obtained in general in all cases. The different correlations observed in the linear regression methods are mainly explained by the different predictor settings used. In particular, those methods including the "homogeneous" predictor two-meter temperature (MO-GP, MLR-T, MLR, MLR-WT) exhibit larger correlations (particularly during winter), due to the stronger connection of this predictor with local surface temperature. Note that this is not an indication of better performance of the model for climate change applications, since upper-air predictors may be more robust.

Finally, regional and seasonal differences are observed in the link strength when looking at the results 473 aggregated over the eight Prudence regions considered (shown by the colored bars in Figs. 2 and 3). For 474 precipitation, both MOS and PP methods mostly preserve the rank of ERA-Interim (and RACMO, for MOS) 475 regional results for the different seasons. In summer the highest correlations are obtained for the Alps and the 476 weakest in Mediterranean and Iberian Peninsula regions, whereas in winter correlations are larger in Central 477 Europe and smaller in Eastern Europe and British Isles. For the case of maximum temperature, correlations 478 are higher for Eastern and Central Europe and smaller for Iberia and the British Isles (for summer) and Iberia 479 and the Alps (for winter). There exist also some cases where the PP methods show some differences with 480 respect to the ERA-Interim and MOS results. For instance, the Alps are among the regions with highest 481 winter correlations for PP methods, contrary to the ERA-Interim results. In other cases, PP methods enlarge 482 the regional variability of results. For instance, the regional correlations of summer maximum temperatures 483 have larger spread for PP methods, mainly due to the small correlation obtained for the British islands. 484 particularly for regression methods. 485

486 4. Validation Results for Precipitation

In this section we present the first validation results obtained for precipitation, focusing on general marginal
 distributional aspects. Figures 4 and 5 show the relative biases for R01 (relative wet-day frequency) and
 SDII (mean wet-day precipitation), respectively (predicted over observed mean values).

The results for the raw model outputs (indicated as ERAINT-200, -075 and RACMO22E) are included 490 in the first three columns and show the comparison with the local observations considering the raw model 491 values at the closest gridbox. Model outputs tend to overestimate wet-day frequency and underestimate 492 precipitation amount. Moreover, the resulting biases decrease when increasing the resolution, being largest 493 for the 2.0° ERA-Interim and smaller for the 0.11° RACMO2. Overall, most of the downscaling methods 494 greatly improve model biases (both for ERAINT-200 and RACMO2 predictors, the latter shaded in the 495 figures) and no downscaling approach or technique seems to be superior in general. An exception is found 496 for the four linear regression methods (from MLR-RAN to MLR-ASI), which exhibit very large biases, 497 even larger than those corresponding to the raw model outputs. A similar behavior is also found for the 498 GLM-P method, with larger biases than the rest of GLM implementations. These results clearly illustrate 499 the inadequacy of linear regression methods for downscaling precipitation values. Note that the nonlinear 500 regression method (MO-GP) presents smaller biases, though still larger than for the rest of methods. On the 501 other hand, GLMs exhibit small biases, particularly in the vectorized versions (VGLMGAMMA, SWG) and 502 the version conditioned to weather types (GLM-WT), all including some sort of seasonality, either imposed 503 by training the model separately for each month, or indirectly conditioning the model to twelve different 504 catalogues of weather types. On the other hand, the different analog implementations present similar biases, 505 with exception of FIC01P (using only geopotential fields) which exhibits larger biases for winter SDII. 506 Moreover, in this case, training the methods separately for each month/season does not clearly improve the 507 results, since the analog method trained with year around data (ANALOG) exhibits similar biases than to 508 the rest of (seasonally trained) analog implementations. This different behavior may be a consequence of 509 the fact that, as opposed to regression methods, analog methods do not explicitly calibrate the mean value 510 towards the observations. 511

Regarding the MOS methods, similar results are obtained for the different families of techniques, al-512 though there is an outstanding group of methods with very small biases, formed by the empirical quantile 513 methods (QM) including a seasonal component (see Table 3) —with the exception of CDFt, which systemat-514 ically overestimate precipitation intensity.— The key role of the seasonal component can be seen comparing 515 EOM and EOMs methods, only differing in the 31-day moving window used to train the latter. Therefore, 516 similarly to the regression techniques (TF), seasonal calibration is beneficial for QM methods. On the other 517 hand, the results are similar for the two predictor settings —ERA-Interim and RACMO2— with slightly 518 smaller (and more centered) biases for the latter, particularly for scaling and parametric quantile mapping 519 methods. Those methods with no seasonal component (e.g. GQM, GPQM, EQM) show compensating 520 DJF and JJA biases. It is also interesting to note that the particular FIC02P/04P methods (which apply the 521 parametric BC quantile method to outputs from FIC01/03, respectively) improve the performance of the cor-522 responding PP counterparts (see, e.g. wet frequency for FIC01P) and also show better results than the direct 523 application of the BC method to ERA-Interim. This indicates that the PP method (applied to ERA-Interim) 524 produces more realistic local precipitation results, well suited for a parametric correction. 525

The WG techniques exhibit small biases, with the exception of the MARFI family which systematically over- and under-estimate wet-day frequency and amount, respectively.

When looking at the regional variability of results (horizontal color bars in Figures 4 and 5), there 528 is a high method-to-method regional variability. The largest/smallest biases are found in the Mediter-529 ranean/British Isles for ERA-Interim. However, these regional differences are greatly reduced in all down-530 scaled results, although some methods still exhibits large biases in the Mediterranean region during Sum-531 mer. Figure 6 shows the individual station results for winter (DJF) and summer (JJA) for PRCTOT (total 532 precipitation) relative biases, with southernmost stations at the bottom and northernmost stations at the top. 533 This figure shows systematic bias patterns across stations for each particular methods (greeny or browny 534 vertical bars), although there are some stations where most of the downscaling methods exhibit a similar 535

systematic bias. For instance, most of the methods overestimate total precipitation for station number 12

537 (Roma-Ciampino).

Finally, Figure 7 shows the results of the relative biases for the standard deviation of daily precipitation 538 (downscaled over observed standard deviations), manifesting the deficiencies already reported for some of 539 the methods (e.g. CDFt and the linear regression family). Among the MOS methods, some techniques tend 540 to systematically underestimate (e.g. ISI-MIP) or overestimate (DBS, Ratyetal-M9) variability and, again, 541 the best performing methods are QMs including a seasonal component. As opposite to the bias in the mean, 542 larger biases are found here for RACMO2 than for ERA-Interim downscaled values for some of the MOS 543 methods. Regarding PP methods, analog techniques tend to systematically underestimate variability. It is 544 also shown that regression-based deterministic methods (including GLM-det) can only reproduce a small 545 part of the observed variance. Moreover, as opposite to the two GLM stochastic versions (GLM and GLM-546 WT), the MLR stochastic versions (GLM-P and MLR-ASW, the former simulating precipitation occurrence 547 and the latter inflating with white noise) fail to recover the observed variability. Again, FIC02P/04P methods 548 seem to correct the deficiencies of their PP counterparts. 549

550 5. Validation Results for Temperatures

Figures 8 and 9 show the results for the mean biases (downscaled minus observed mean values) of daily 551 maximum and minimum temperatures, respectively. The raw model outputs from ERA-Interim largely 552 under/over estimate maximum/minimum temperatures in almost all regions, whereas RACMO2 exhibits 553 smaller biases but still with a large regional variability (in this case, the model tend to underestimate both 554 minimum and maximum temperatures). Overall, most of the downscaling methods greatly improve model 555 biases and again no downscaling approach or technique seems to be superior in general. There are a few 556 methods exhibiting very large biases (WT-WG for both TMAX and TMIN, and CDFt, MLR-ASW for 557 TMIN) which indicate some problem with the configuration of the method or with the particular execution. 558 Moreover, the MOS-REG regression method exhibits large biases when applied to RACMO2 model (much 559 larger than when applied to ERA-Interim). Therefore, in this case the synchrony of the climatic run with 560 observations is too weak to allow for a suitable implementation of this type of MOS regression technique (at 561 least considering information only on the nearest gridbox). Moreover, the methods SB and EQM-WIC658 562 exhibit large biases (particularly during winter) which can not be explained from the definition of the method 563 (other similar techniques show small biases) and could be an indication of some problem in the application 564 of the method. 565

In general, the family of methods exhibiting larger biases are the analog techniques, but this could be 566 explained because they do not explicitly calibrate the mean during training. Among these methods, the best 567 results are obtained with ANALOG and ANALOG-SP, both using 2m temperature as predictor, which seems 568 to have an important role in this case. However, the particular choice of the predictor cannot explain the 569 differences among the regression techniques. Similarly to the case of precipitation, a key factor explaining 570 the variability of MOS results is the seasonal training of the methods (e.g. EQM vs EQMs). Note that the 571 cross-validated mean bias for simple linear scaling methods (additive and/or multiplicative; e.g. RaiRat-M6 572 and RaiRat-M7) should be zero by construction (in cases with no missing data). In this work, the seasonal 573 cross-validated results obtained for these methods are different from zero (although very small) due to the 574 two-month moving window used to compute the scaling factors (see details in Annex 1, describing the 575 methods). 576

The above figures do not show relevant regional differences for the biases of the downscaled methods with the exception of the analog methods where the regional biases observed seem to be related to the predictors used. Figure 10 gives further information showing the individual station results (sorted as in Table 1) for daily maximum temperature for winter (DJF, top) and summer (JJA, bottom). Besides of revealing the bad performing methods, these figures show the systematic biases exhibited for the methods to under/over estimate across the different stations (e.g. the analog methods).

⁵⁸³ Finally, Fig. 11 shows the results for the standard deviation (relative biases, downscaled divided by

the observed daily standard deviations) of daily minimum temperatures during winter and daily maximum 584 temperatures during summer. This figure allows to clearly differentiate those MOS scaling methods not cor-585 recting the variance of the downscaled results (RaiRat-M6/M7, SB, ISI-MIP), which show the same biases 586 as the input reanalysis or RACMO2 models. Moreover, as expected, those quantile mapping techniques 587 trained annually exhibit larger biases on the seasonal variances than those seasonally trained (see, e.g. EQM 588 and EQMs). These two factors explain most of the variability of the MOS results (together with the model 589 deficiencies already reported). Regarding the PP methods, the analog techniques tend to systematically un-590 derestimate the variance, with FIC01T being the worst of this group. Note that the results of this method are 591 later used as input for the FIC02T, which correct for this deficiency applying a quantile mapping approach. 592 In the case of the linear regression methods, all deterministic implementations underestimate the observed 593 variance, in correspondence with the daily correlation values shown in Figure 3. However, as expected, 594 those methods correcting the seasonal variance (using inflation and white noise, as in MLR-ASW and ASI, 595 respectively —note that a problem was already reported for MLR-ASW results for minimum temperature— 596) show more centered results (although there are some outlier stations where the variance is over-estimated). 597 Note that variance correction at an annual basis (e.g., MLR-AAI/AAW) yields seasonal biases comparable to 598 other deterministic linear regression implementations. Finally, the deterministic symbolic regression method 599 MO-GP is a multi-objective method optimizing several statistics, including standard deviation. As a result 600 it exhibits smaller seasonal variance biases (even though it is trained on an annual basis), and larger mean 601 biases (see Figures 9 and 8), than the regression methods. This could be beneficial for climate change appli-602 cations since it requires no postprocessing but, again, a more comprehensive assessment of other aspects is 603 needed. 604

605 6. Promoting Transparency and Reproducibility of Results

Research transparency and reproducibility is a major concern in the different experimental disciplines (see a string of freely available nature articles on reliability and reproducibility of published research at http: //go.nature.com/huhbyr). For instance, a recent survey over 1500 scientists recently reported by Baker (2016) revealed that the work published in different research fields (including Earth and Environment) were mostly not reproducible (over two-thirds). As a result, there is growing alarm about results that cannot be reproduced. In VALUE substantive steps were taken in order to improve transparency and reproducibility of results, and to promote awareness within the downscaling scientific community.

The main difficulties for research reproducibility identified include 1) access restrictions to raw input data and/or results, 2) poor experimental design information, 3) lack of code availability, and 4) incomplete documentation of the particular configuration and implementation used (data preprocessing, method configuration and specific parameter values, training options, etc.). In some cases, the steps involved in the downscaling process are very technical and they are not always appropriately documented in practical applications, thus making difficult the reproducibility of the results.

⁶¹⁹ The following actions have been undertaken in VALUE in order to avoid the above mentioned problems:

1) All the data needed for the experiments described in this paper has been collected and made available at http://www.value-cost.eu/data. Moreover, the daily downscaled data and the resulting validations for each of the methods and experiments are also publicly available under the liberal Creative Commons Attribution (CC BY) License (http://creativecommons.org/licenses/by/4.0).
 Contributors also get access to VALUE_PRIVATE published data, which is made available internally to the consortium for quality check and verification purposes before publication (more info in http://www.value-cost.eu/terms).

• 2) The experimental framework was designed and published (Maraun et al. 2015) in advance of the open call for contribution to this validation experiment.

3) The code used for the validation framework (from data loading to computing all the validation mea sures) has been coded in R and is publicly available from http://github.com/SantanderMetGroup/

R_VALUE. In addition, the packages and/or code needed to reproduce the results for some of the 631 downscaling methods are publicly available (see Annex 1). Other methods use proprietary soft-632 ware and cannot be replicated; however, we decided to also include this information to favor method 633 inter-comparability, but requiring the open publication of the results (both predictions and validation 634 results), which was mandatory for all methods contributing to this paper. 635

636

637

638

4) Furthermore, a metadata description vocabulary for statistical downscaling methods has been defined and implemented in the VALUE Validation Portal, providing information on the method characteristics and implementation details needed to properly analyze the results.

Finally, transparency is further promoted by the VALUE Validation Portal http://www.value-cost. 639 eu/validationportal, providing public access to the metadata and data for all contributing methods 640 and also for all validation results, as well as tools to filter and visualize (both in tabular and graphical for-641 mats) the results. 642

7. Conclusions 643

In this paper we present the ensemble statistical downscaling methods produced in the VALUE collabora-644 tion, which covers the three common downscaling approaches (perfect prognosis, model output statistics 645 ?including bias correction? and weather generators) with a total of over fifty downscaling methods. We also 646 present the first results from the inter-comparisson experiment under the same cross-validation experimental 647 framework using "perfect" predictors. Additional experiments with GCM data will follow to contribute to 648 the EURO-CORDEX initiative. Appropriate metadata on the main model characteristics (e.g. determin-649 istic or stochastic nature) and implementation details (predictors, geographical domain, monthly/seasonal 650 training, etc.) have been collected in order to properly analyze the results. 651

Overall, most of the downscaling methods greatly improve model biases and no downscaling approach 652 or technique seems to be superior in general, due to the large method-to-method variability of results. Some 653 bad performing methods have been identified as potentially failed methods due to different problems giving 654 some clues about future quality checks to be implemented in the VALUE validation portal. Our results also 655 show the inadequacy of linear regression methods for downscaling daily precipitation values, which is still 656 used in some applications (see, e.g., Jeong et al. 2012; Chen et al. 2014). Regarding the MOS methods, em-657 pirical quantile methods including a seasonal component form an outstanding group of methods with very 658 small biases. However, there are particular PP and WG methods with a similar performance. In this work 659 we found that, in agreement with previous studies (Reiter et al. 2017), introducing a seasonal component 660 (e.g. training the methods separately each calendar season, month or moving window) improves the results. 661 However, we found that all implementations (even a daily moving window) resulted in a relevant perfor-662 mance improvement, differently to Reiter et al. (2017), where seasons were recommended for calibration. 663 The deterministic or stochastic nature of the method was the most relevant factor (together with seasonal 664 training) for explaining the variability of results for biases in the standard deviation. 665

In this work we have also tested some new experimental developments, such as stochastic and analog-666 or regression-based MOS methods, applied to RCM climatic runs driven by reanalysis. The results seem 667 to be promising for precipitation but not for temperatures, apparently due to the weak synchrony between 668 RCM outputs and observations. Some promising results have been also obtained when combining PP and 669 MOS methods. In particular, parametric quantile methods are shown to produce better results when applied 670 to the outputs of an analog method (using ERA-Interim predictor data), than when applied directly to ERA-671 Interim. This indicates that the PP method produces more realistic local precipitation results than ERA-672 Interim, well suited for a parametric bias correction. A similar result is obtained with MOS methods when 673 applied to RCM climatic runs driven by reanalysis than to the reanalysis outputs directly. However, these 674 first validation results should be interpreted with caution, since a good performance in terms of bias may 675 not be an indication of a better performance of the model for climate change applications. Therefore, a 676 comprehensive validation analysis of different aspects is needed in order to properly assess the performance 677

of this technique (e.g. temporal and extreme aspects, as described in the companion papers of this special issue). Moreover, the present study cannot give a conclusive assessment of the skill of downscaling methods to simulate regional future climates, and further experiments (Maraun et al. 2015) will be soon performed in the framework of the EURO-CORDEX initiative, thus completing the analysis initiated in the present manuscript.

Finally, in order to favor research reproducibility, the experimental framework is precisely de-683 fined, all datasets needed for this experiment are publicly distributed http://www.value-cost.eu/ 684 datasets and, in some cases, the packages and/or code to reproduce the results are publicly available. 685 A metadata description vocabulary has been defined and implemented in the VALUE Validation Portal 686 http://www.value-cost.eu/validationportal, which provides metadata information for all 687 contributing methods (approach, technique, predictors, method configuration, etc.). Transparency is also 688 promoted by the VALUE Validation Portal http://www.value-cost.eu/validationportal, 689 which provides public access to data and metadata information for all contributing methods and also for all 690 validation results, as well as tools to filter and visualize (both in tabular and graphical formats) the results. 691 In particular, most of the figures of the paper can be reproduced with these tools. 692

8. Acknowledgments

This work have been performed in the framework of the VALUE is funded via the EU COST Action ES1102, 694 under FP7 programme. We thank the VALUE community for their input to this framework, helping in the 695 coding the validation routines and helping to select the representative stations at a national level. In partic-696 ular we would like to thank Andreas Gobiet, Constantin Mares, Robertas Alzbutas, Mandy Vlachogianni, 697 Adam Jaczewski, Peter Thejll, Patrick Willems, Ivan Pilaš, Meron Teferi Taye, and Fredrik Boberg. We ac-698 knowledge the data providers in the ECAD project (data and metadata are available at http://www.ecad.eu) 699 and ECMWF for allowing us to re-distribute ERA-Interim daily data (regridded at a 2° resolution) within 700 VALUE (registration is public) for the standard set of predictors. We also acknowledge KNMI for mak-701 ing publicly available the RACMO2 0.11° resolution ERA-Interim driven simulations within the EURO-702 CORDEX initiative. 703

JMG and SH acknowledge partial funding from MULTI-SDM project (MINECO/FEDER, CGL2015-66583-R). BH and DR acknowledge COMPLEX project (FP7-ENV-2012. Number: 308601). MT was supported by HOPE project (MINECO, CGL2014-52571-R). Participation of MD and RH was funded by the Ministry of Education, Youth, and Sports of the Czech Republic, contracts LD12029 and LD12059, respectively.

The authors gratefully acknowledge helpful comments by the anonymous reviewers.

710 9. ANNEX 1. Description of methods

This annex includes the detailed description of the methods used in this work (Tables 3 and 4). They are organized alphabetically within each downscaling approach (MOS, PP and WG) in the following sections.

713 **a.** MOS Methods

• **BC** (only precipitation): Parametric bias-correction method using the optimum among five theoretical distributions (Gamma, Weibull, Classical Gumbel, Reversed Gumbel and Log-logistic, all of them with four parameters) for each station on a monthly basis (Monjo et al. 2014).

717 **Implementation:** In-house R code.

• **CDFt:** The CDFt approach links the local-scale CDF of the variable of interest to the associated largescale CDF through a "quantile-quantile" approach performed between the future large- and localscale CDFs (and not between present CDFs as in the classical quantile-quantile method). To do so, the future local-scale CDF is first estimated based on the assumption of a mathematical transformation
to link the evolution of the large-scale CDF to the evolution of the local-scale one. Hence, CDFt is
a variant of quantile-quantile but CDFt accounts for the CDF changes from the calibration to the
projection (or future) time periods (Vrac et al. 2012).

- DBS: Distribution based parametric quantile mapping (Yang et al. 2010, 2015). The cumulative 725 distribution of precipitation and temperature are fitted by double-gamma and normal distributions, 726 respectively. A wet-day correction is applied to precipitation series. In case of too many wet-days in 727 the predictor data, all wet-days below a derived threshold are removed so that the wet-day frequency 728 is the same as in the predictand data. In case of too few wet-days in the predictor data, the wet-day 729 correction is done by adding wet-days to already existing wet-spell, starting with the longest ones. 730 Temperature correction is done conditional on the wet/dry state of the corresponding precipitation 731 series. The parameters were seasonally calibrated for every month in the annual cycle. 732 Implementation: In-house FORTRAN code. 733
- DBD/DBBC (only temperatures): Bias is calculated separately for all percentiles (from 1 to 99) and a polynomial function of second degree is fitted as a function of the temperature values (DBD) or the probabilities (DBBC). In the validation period the model temperature (or the corresponding percentile) is used to calculate the bias to be subtracted in the adjustment process. The difference between DBD and DBBC is that the bias is connected with temperature and percentile values, respectively. In both cases, calculations are performed for each season separately.
 Implementation: In-house Matlab code.
- EQM/EQMs/EQM-WT: Implementation of Empirical Quantile Mapping (EQM) adjusting 99 per-741 centiles and linearly interpolating inside this range every two consecutive percentiles; outside this 742 range a constant extrapolation (using the correction obtained for the 1st or 99th percentile) is applied 743 (Déqué 2007). In the case of the precipitation, when the predicted dry frequency is larger than the 744 observed one the frequency adaptation proposed by Themeßl et al. (2012) is applied. In order to 745 explicitly model the seasonal cycle, the variant EQMs considers a 31 day moving window centered 746 on every calendar day to calibrate the method. EQM-WT is a state-dependent version of EQM, con-747 ditioning the training to 12 Weather Types defined using a k-means algorithm (k=12) applied to the 748 daily SLP over Europe. For the experiment with the RACMO2 RCM predictors, SLP is taken from 749 the RCM model and smoothed to a 1° resolution. 750
- **Implementation:** EQM is implemented in the *downscaleR* (Bedia et al. 2016) R package (*bias-Correction* function) with the options *method* = "*eqm*" and *extrapolation*="*constant*", including *precipitation* = *TRUE* and *pr.threshold* = 1 for precipitation. For EQMs the extra argument *window* = c(30, 1) was included. This package is freely available without restriction.
- EQM-WIC658: Implementation of the empirical quantile mapping method (Déqué 2007) sorting the values into bins with adjustable width (e.g. 0.1°) and applying a linear interpolation between two percentiles (bins); out of range values are adjusted using constant extrapolation (using the correction obtained for the minimum or maximum). In order to cope with the seasonal cycle, a 31 day moving window centered on every calendar day is used to calibrate the method. More details in Wilcke et al. (2013).
- **FIC02P/04P** (only precipitation): Parametric bias correction technique (method BC above) applied to FIC01P/03P results. The method is applied separately for each month (Monjo et al. 2014).
- FIC02T (only temperatures): Parametric bias correction technique considering Gaussian distributions applied to FIC01T results. The method is applied separately for each station and for each month (Monjo et al. 2014).

- **GQM/GPQM:** Gamma/Gaussian parametric Quantile Mapping (GQM) to approximate the empirical distribution of precipitation intensity / temperature. In the case of the precipitation, the frequency adaptation proposed by Themeßl et al. (2012) is previously applied to calibrate precipitation occurrence (a 1mm threshold is used). The Generalized Pareto Quantile Mapping (GPQM) version considers a Generalized Pareto to adjust separately the extremes values (over the 95th percentile) as in Gutjahr and Heinemann (2013).
- 772**Implementation:** GQM/GPQM are implemented in the *downscaleR* (Bedia et al. 2016) R pack-773age (*biasCorrection* function) with the options *method* = "gqm"/"gpqm", including precipitation =774TRUE and pr.threshold = 1 for precipitation. This package is freely available without restriction.
- **ISI-MIP:** The trend preserving ISI-MIP method proposed by Hempel et al. (2013) in the framework 775 of the Inter-Sectoral Impact Model Intercomparison Project (ISI-MIP). This method works in a two-776 step approach. First, the monthly mean is adjusted with a linear/multiplicative scaling and, then, 777 the resulting daily anomalies are corrected by means of a parametric (gaussian/exponential) quantile 778 mapping, for temperatures/precipitation, respectively. In the case of precipitation also a frequency 779 adjustment is included for both the monthly and daily components. This method has been designed 780 to simultaneously adjust groups of variables (precipitation-snow, temperatures, wind speed and com-781 ponents). 782
- Implementation: ISI-MIP is implemented in the *downscaleR* (Bedia et al. 2016) R package (*isimip* function). This package is freely available without restriction.
- MOS-AN (only precipitation): MOS implementation of the analog method considering precipita-785 tion as the single predictor, and trained across different zones (similar to the Prudence regions) com-786 puting similarity using Euclidean distances of the precipitation fields. As a benchmark this method 787 has been applied directly to ERA-Interim precipitation, with "perfect" (day-to-day) synchrony with 788 observations. When applied to the ERA-Interim driven RCM simulation, this method exploits the 789 marginal temporal synchrony within the RCM domain given by the synchronous forcing at the bound-790 ary (Turco et al. 2011, 2017). Note that this method is best suited for for nudged RCM simulations. 791 Implementation: In-house Matlab code. 792
- MOS-REG/GLM (only temperatures/precipitation): MOS implementation of linear (and generalized linear) methods considering as predictor the mean of the predicted temperature (precipitation) at the four nearest gridboxes (Herrera et al. 2017).
- 796 **Implementation:** In-house Matlab code.

• QM-DAP: Implementation of the empirical quantile mapping method (Déqué 2007) smoothing the 797 final corrections (obtained for individual percentiles) with a low-pass Gaussian filter (over 20 per-798 centiles) to reduce noise in the individual percentile values. Each month was treated separately and 799 a time window including the previous and following month was applied. To preserve reasonable ex-800 trapolated values (in the tails of the distribution), changes between the last percentiles (likely to be 801 very noisy) were limited to certain values (such as a coefficient of 1.5 for maximal extrapolated value, 802 compared to the last percentile, and a ratio of 3.0 as a change between the last two percentile values). 803 More details in Štěpánek, P. et al. (2016). 804

Implementation: In-house R code, incorporated in ProClimDB software (www.climahom.eu).

• **QMm:** Equidistant empirical quantile mapping. Empirical CDFs are calculated for the observation and the calibration and validation periods. The probabilities are calculated for bins with widths set for the resolution of the observational data (e.g. 0.1°). For each day in the validation period, the probability obtained from the validation CDF is used in the observational and calibration CDFs to obtain the corresponding data values (Li et al. 2010). The difference between the observed and calibration data is used as the correction term for the validation. In the case of precipitation and in order to reduce the models drizzle effect, the percentile of the dry days of the validation period is matched to the observations, i.e. the precipitation in the validation period which corresponds
to a percentile lower than the observational percentile is set to zero. In order to account for the
seasonal cycle, the CDFs are constructed for a 31 day window centered on each day of the year.
Implementation: In-house FORTRAN code. Available upon request to R.M. Cardoso.

- QMBC-BJ-PR: Implementation of the empirical quantile mapping method (Déqué 2007) adjusting 101 percentiles (including the minimum and maximum values) and using constant interpolation (with the mean of the two correction factors) between every two consecutive percentiles. Out of range values are adjusted using constant extrapolation (using the correction obtained for the minimum or maximum). The calibration is performed separately for each month. More details in Pongrácz et al. (2014); Bartholy et al. (2015).
- 823 **Implementation:** In-house FORTRAN code.

• Ratyetal-M6-M9 (only precipitation): Monthly bias correction of daily precipitation implemented 824 as in Räty et al. (2014). Methods M6 and M7 adjust the mean and standard deviation using linear 825 and power scaling functions, respectively. M8 is a non-parametric quantile mapping with smoothing 826 tailored for precipitation. The smoothing parameter value a=0.02 in Eq. (5) of Räty et al. (2014). M9 827 is a simplified version of the DBS method monthly transfer functions are estimated by fitting separate 828 gamma distributions below and above the 95th percentile of daily precipitation (Yang et al. 2010); 829 this constitutes a simplified version of the DBS method where the wet-day correction is only applied 830 when there are too many wet days in the predictor data. No correction is done if the modeled wet 831 day frequency is smaller than the observed one. In this sense M9 is less sophisticated than the actual 832 DBS version. A 0.1 mm threshold was used to define wet-days. All methods use three-month time 833 window when deriving the monthly corrections (e.g. data from December-January-February used for 834 the correction applied in January). 835

836 **Implementation:** In-house Fortran code.

RaiRat-M6-M9 (only temperatures): Monthly varying bias correction of temperature following (Räisänen and Räty 2013). M6 adjusts only the mean value, M7 mean and standard deviation, and M8 mean, standard deviation and skewness. M9 uses a non-parametric quantile mapping approach with smoothing parameter D = 0.05 in Eq. (5) of (Räisänen and Räty 2013). A two-month data window is used in deriving the corrections (e.g. from mid-April to Mid-June for the correction applied in May) in all these methods.

- 843 **Implementation:** In-house Fortran code.
- **SB** (only temperatures) A local scaling method where mean bias calculated separately for each season is subtracted from simulations in validating period.
- 846 **Implementation:** In-house Matlab code.
- VGLMGAMMA (only precipitation): A stochastic single-site MOS approach to predict precipitation occurrence and amounts conditionally on simulated daily precipitation as predictor. Precipitation occurrence is modeled via a logistic regression; precipitation amounts on wet days based on a vector generalised linear model that expresses the rate and shape parameters of the 2-parameter gamma distribution as a function of simulated daily precipitation. Temporal dependence is not explicitly modelled but only imprinted by the predictor, i.e., individual occurrences and amounts are conditionally independent (Wong et al. 2014; Volosciuk et al. 2017).
- 854 **Implementation:** In-house R code.

b. *PP Methods*

• ANALOG: Standard analog technique using Euclidean distance considering the complete fields to compute distances (Gutiérrez et al. 2013; San-Martín et al. 2017). The method has been trained across different zones covering Europe (similar to the Prudence regions) and has no seasonal component.

- The method used raw predictor values applying a compression preprocess keeping the PCs explaining 95% of the total variance.
- Implementation: *MeteoLab* public Matlab toolbox (http://meteo.unican.es/trac/ MLToolbox/wiki) using *downTrain* function with parameters em method.type = 'analogs', *AnalogsNumber* = 1, resampling='no'.
- These results (for the case of the ERA-Interim predictors) can be also reproduced (and modified) online using the statistical downscaling portal **http://meteo.unican.es/downscaling**, which builds on MeteoLab, and includes as illustrative examples the same standard predictor data and the VALUE regions used in this study. This package is freely available without restriction.
- ANALOG-ANOM: For a given day to be downscaled, the ANALOG-ANOM (Vaittinada Ayar et al. 868 2016) determines the day in the calibration period which has the closest atmospheric situation. This 869 is determined by a similarity metric (here a Euclidean distance) between the predictor set of the day 870 to be downscaled and the predictor set of the day in the calibration period, considering the whole 871 European domain. For this method, the predictors are fields of daily anomalies with respect to the 872 annual cycle computed from cubic regression smoothing splines fitted on the empirical daily annual 873 cycle. Moreover, a seasonal restriction is applied: the selected analogs have to be in a +/-15 day-874 window around the climatological day of interest. 875
- ANALOG-MP/SP: Two versions of the analog model developed by Obled et al. (2002), optimized 876 for the multivariate prediction of weather variables over the European region (Raynaud et al. 2016). 877 For each prediction day, the probabilistic prediction is obtained from the 30 best atmospheric analogs 878 selected in the atmospheric archive (selection in a calendar window of +/- 30 days). A two-level 879 stepwise analogy is used for the analog selection: The first analogy level leads to 100 analogs from 880 which are identified the 30 best final ones thanks to the 2nd analogy level. In both MP and SP versions, 881 the first level of analogy is based on the shapes of 1000 and 500 hPa geopotential fields over a spatial 882 domain centred on the target station (or centered on the region in the case of the multisite experiment). 883 The analogy criterion is the Teweless-Wobus Score (Teweless, 1954). The 2nd level of analogy 884 relies on a thermodynamic mesoscale predictor (analogy criterion is the RMSE). In ANALOG-SP, 885 the predictor (T-Td at 2m) is the same for the three predictants (precip., Tmin, Tmax). The values of 886 local temperature obtained with each analog day are post corrected using the difference between the 887 mesoscale 2m temperature of a given target day with the one of the analog. In ANALOG-MP, the 2nd 888 analogy predictor is predictand specific (VV600 for precip., T850 for temp. variables). In the present 889 work, ANALOG-MP/SP include a stochastic process to produce the 100 required realizations of the 890 predictand from the probabilistic prediction computed for each day. 891
- ESD-EOFSLP/EOFT2/SLP/T2 (only temperatures): Multiple linear regression method using 892 monthly aggregated predictor and predictand data. It is important to remark that the ESD pack-893 age is not designed to downscale daily values, but parameters describing the seasonal distribution 894 of daily (or hourly) data, and combine this with a weather generator to produce time series. In this 895 contribution this method has been trained on a monthly basis (using monthly aggregated data) in the 896 traditional way, but this package is more flexible and it is typically calibrated differently when applied 897 to GCM data. In that case, common EOFs (representative of both reanalysis and GCMs, Benestad 898 et al. 2015b) are used as predictors and normally PCA are used as predictands for groups of stations 899 which are subject to similar weather phenomena (multi-site application), although the method can be 900 also applied to downscale more general information, such as the occurrence of intense local 24-hour 901 precipitation events over seasonal intervals (Benestad and Mezghani 2015). 902
- Implementation: ESD is implemented in the *esd* R package (Benestad et al. 2015a). This package
 is freely available without restriction.
- **FIC01P/03P** (only precipitation): FIC01P is a two-step analog methods. In a first step, the 30 closest analogs are computed for each test day based on Z1000 and Z500. Every analogue is defined

in a three-windows nested grid (for short, medium, and large scale) with different weights. For this 907 experiment we have used 42 main windows, each one with 3 nested windows, covering Europe. 908 Instead of considering the weighted (according to similarity) mean observations of the analog days 909 (p_i) , the second step performs a pooling and ranking of the analog days month by month (900 values 910 for each month) and computes the mean of consecutive blocks of 30 days q_i according to their mean 911 values. Afterwards the values p_i are substituted by the new values q_i following a rank order, i.e. 912 maximum by maximum, and so on (see, Ribalaygua et al. 2013, for more details). FIC03P is a 913 version of FIC01P using near surface Wind, Wind at 500 hPA, relative humidity at 850 hPa and 914 relative humidity at 700 hPa for computing the analogues. Moreover, we sort the n selected analogues 915 for each problem day using their relative humidities at 850 hPa values and we weight the precipitation 916 of the analogue day using the relation between the specific humidities at 700 hPa of the problem day 917 and the analogue day. Then, as in FIC01P, we reassignate the previously daily simulated precipitation 918 of a month, by using the distribution of the used analogue days for the whole month. 919

- **FIC01T** (only temperatures): A two-step analog method with the same first step as FIC01P with the same predictors, but considering 150 analogs for each test day. The second step consisting of a multiple linear regression using 1000-850 thickness, 1000-500 thickness and daily solar radiation (calculated as a function of the day of the year and the latitude of the station) as regressors; the regression is fitted considering the analog days. More details in Ribalaygua et al. (2013).
- GLM-DET/GLM/GLM-WT (only precipitation): Standard two-stage implementation of General-925 ized Linear Models (GLMs) for precipitation, in which a GLM with Bernoulli error distribution and 926 logit canonical link-function (also known as logistic regression) is used to downscale daily precipita-927 tion occurrence (as characterized by a threshold of 0.1mm) and a GLM with gamma error distribution 928 and log canonical link-function is applied to downscale daily precipitation amount (San-Martín et al. 929 2017). The method is trained across different zones covering Europe (similar to the PRUDENCE 930 regions) with no seasonal component. The predictors are the 20 leading PCs (15 for GLM-WT) 931 of the joined predictor fields (which account for 75-90% of the explained variance across the dif-932 ferent zones). Particular methods are provided in both deterministic (GLM-DET) and stochastic 933 (GLM) variants, using the expected value in the former case and simulating from the resulting bino-934 mial/gamma in the latter. An implementation conditioned to weather types (GLM-WT) is also used, 935 considering 12 weather types defined using a k-means algorithm (k=12) applied to the daily SLP (this 936 variable is excluded from the predictor set in this case). 937
- Implementation: MeteoLab public Matlab toolbox (http://meteo.unican.es/trac/ MLToolbox/wiki) using downTrain function with parameters type = 'glm', ThresholdPrecip = 0.1, NumberOfNearestNeighbours = 0, NumberOfPCs = 15, SimOccurrence = 'true', SimAmount = 'true', minrainydays = 5.
- These results (for the case of the ERA-Interim predictors) can be also reproduced (and modified) online using the statistical downscaling portal **http://meteo.unican.es/downscaling**, which builds on MeteoLab, and includes as illustrative examples the same standard predictor data and the VALUE regions used in this study. This package is freely available without restriction.
- MLR/MLR-WT (only temperatures): (Gutiérrez et al. 2013) Multiple linear regression trained across different zones covering Europe (similar to the Prudence regions) with no seasonal component. The predictors are the 15 leading PCs of the joined predictor fields (which account for 75-90% of the explained variance across the different zones considered). MLR-WT is a state-dependent version of MLR, conditioning the training to 12 Weather Types defined using a k-means algorithm (k=12) applied to the daily SLP over Europe (this variable is excluded from the predictor set).
- Implementation: *MeteoLab* public Matlab toolbox (http://meteo.unican.es/trac/
 MLToolbox/wiki) using *downTrain* function with parameters *type* = 'linear_regression', Num *berOfNearestNeighbours* = 0, and *NumberOfPCs* = 15.
- These results (for the case of the ERA-Interim predictors) can be also reproduced (and modified)

online using the statistical downscaling portal http://meteo.unican.es/downscaling, which builds on
 MeteoLab, and includes as illustrative examples the same standard predictor data and the VALUE
 regions used in this study. This package is freely available without restriction.

- MLR-PCA-ZRT (only temperatures): Linear regression model using s-mode PCs as predictors. The selection of predictors has been automated by iterating through all possible predictor combinations, minimizing the mean squared error and maximizing the time series correlation in calibration and validation (Hertig and Jacobeit 2008; Hertig et al. 2013; Jacobeit et al. 2014). The selection was done for each station separately and models were developed for each month separately.
- Implementation: In-house Fortran (for PC calculation) and R code (using "lm" for regression).
- MLR-RSN/RAN/AAN/AAI/AAW/ASI/ASW: Multiple linear pointwise regression (with stepwise screening) using gridpoint raw data (or anomalies), trained at an annual (or seasonal) basis and including optional variance corrections in the form of inflation or addition of white noise. The first letter of the code refers to the raw (R) or anomaly (A) data used as predictors, the second letter refers to the annual (A) or seasonal (S) training, and the third letter refers to inflation (I) or white noise (W) variance correction (N for no correction). More details in Huth (2002); Huth et al. (2015).
 Implementation: In-house Fortran code.
- MLR-T/GLM-P: These methods have been implemented following the Statistical DownScaling 972 Method SDSM, which builds on linear regression (Wilby et al. 2002). The parameters of the regres-973 sion model are obtained by the least squares method from standardized variables. The link between 974 the predictands and predictors is either an unconditional model, used for temperature (MLR-T), or a 975 conditional model, used for precipitation (GLM-P), being the conditioning variable the probability of 976 wet-day occurrence. The GLM-P method uses a logistic regression to estimate the probability of wet-977 day occurrence and an exponential regression to calculate the total daily precipitation amounts (Kilsby 978 et al. 1998). Rainfall occurs when the probability of wet-day occurrence is greater than or equal to a 979 uniform random number like in Wilby et al. (2002), thus incorporating an additional stochastic pro-980 cess. The selection of predictors changes from one site to another and from one variable to another 981 and is based on a step-wise approach building on the adjusted determination coefficient. 982 Implementation: In-house C code. 983
- MO-GP: Multi-objective Genetic Programming (MOGP) performs a symbolic regression building 984 a tree (six levels at most) with arithmetic functions and if-statements, i.e., not only the parameters 985 but also the structure of the regression models are generated by GP. The multi-objective approach 986 aims at a simultaneous optimization of RMSE, bias, standard deviation, selected quantiles and, for 987 precipitation, the number of precipitation days. MOGP is applied individually for each station and 988 variable. Except for precipitation, the predictors are interpolated from the four closest GCM grid 989 cells to the location of the respective station. Precipitation is taken at the GCM grid box closest 990 to a station. The MOGP code is based on the Strength Pareto Evolutionary Algorithm (SPEA) by 991 Zitzler and Thiele (1999) and the GPLAB by Silva and Almeida (2003). SPEA returns not one single 992 regression model for each station and variable but a set of Pareto optimal models. From each set of 993 potential downscaling models one has been selected that optimizes a trade-off between all objectives. 994 The automatic selection results in 8 predictors on average for precipitation and 6 for temperature. 995 More details can be found in Zerenner et al. (2016). 996
- ⁹⁹⁷ **Implementation:** In-house MATLAB code (based on the GPLAB).

• **SWG:** A two-step approach is implemented to model precipitation in a Vectorised Generalized Linear Models (VGLM). First, the rainfall occurrence is modeled through a logistic regression, allowing to characterize the probability of rainfall occurrence for a given day conditionally on atmospheric predictors. Then, the probability density function (pdf) of the rain intensity (given that it rains) is assumed to be a Gamma distribution whose logarithms of the shape and rate parameters are linear functions of the large-scale predictors. For temperature, a single step is used, where temperature is
 supposed to follow a Gaussian distribution with the mean and the logarithm of the standard deviation
 linearly dependent on the predictors (Vaittinada Ayar et al. 2016).

• WT-WG: Gaussian/binomial-gamma distributions are fitted to the observed temperature/precipitation values within each of the 100 weather types obtained applying k-means to the SLP fields. These distributions are obtained to simulate downscaled values. More details in Gutiérrez et al. (2013); San-Martín et al. (2017).

1010Implementation: MeteoLab public Matlab toolbox (http://meteo.unican.es/trac/1011MLToolbox/wiki) using downTrain function with parameters method.type = 'WT'. This package1012is freely available without restriction.

1013 **c.** *WG methods*

 GOMEZ-BASIC/TAD: Non-parametric weather generator based on a nearest neighbors resampling 1014 technique making no assumption on the distribution of the variables being generated. To represent the 1015 interdiurnal variability, each term (except for the first one) of the synthetic time series is derived from 1016 followers of K terms (selected from the learning observed series) closest to the previously generated 1017 term; the first term in the series is selected randomly from all available terms, which are within 10 1018 days from January 1. The distance between individual terms is based on the Mahalanobis distance, 1019 in which precipitation is a binary variable (0 stands for the dry day, 1 for the wet day). Two versions 1020 of the generator were used in the VALUE experiment. In BASIC temperature is represented by 1021 TMAX and TMIN. In temperature is represented by TAVG and DTR (defined above in description of 1022 MARFI). 1023

- MARFI-BASIC/TAD/M3: Parametric multivariate stochastic Richardson-type Richardson (1981) 1024 weather generator, which is a flexible follower of the Met&Roll generator (Dubrovský 1997; 1025 Dubrovský et al. 2004). Precipitation occurrence is modeled by Markov chain (order may vary 1026 between 1 and 3) and precipitation amount on wet day is sampled from the Gamma distribution. 1027 Standardized values of the temperature variables are modeled by the first-order bi-variate autoregres-1028 sive model, in which the means and standard deviations of the two variables are conditioned on the 1029 state (wet or dry) of the day. Three versions of the settings were used in the experiment. In BASIC the 1030 two temperature variables are TMAX and TMIN, order of the Markov chain is one. TAD is similar 1031 to BASIC, but temperature is represented by TAVG (defined as an average of TMAX and TMIN) and 1032 DTR =(TMAX-TMIN) transformed (using quantile-mapping) into normally distributed variable. M3 1033 is the same as BASIC, but a third-order Markov chain is used to model wet day occurrence. 1034
- SS-WG: Multi-variate Richardson-type (Richardson 1981) weather generator simulating daily time-1035 series of precipitation, minimum and maximum temperature (Keller et al. 2015, 2016). First, daily 1036 precipitation occurrence is modelled based on a first-order two-state Markov chain using 1mm/day 1037 as a wet threshold. Precipitation intensities are simulated from a mixture model of two exponential 1038 distributions. To ensure inter-variable consistency, the parameters of the temperature statistics are 1039 conditioned on the precipitation state. Synthetic temperature time-series are simulated using a first-1040 order autoregressive model (AR1). All WG parameters are determined for each station and each 1041 month separately. 1042

1043 References

Abaurrea, J. and J. Asín, 2005: Forecasting local daily precipitation patterns in a climate change scenario. *Climate Research*, 28 (3), 183–197, doi:10.3354/cr028183, URL http://www.int-res.
 com/abstracts/cr/v28/n3/p183-197/.

reproducibility. Baker, М., 2016: 1500 scientists lift the lid on Nature News. 1047 533 (7604). 452, doi:10.1038/533452a, URL http://www.nature.com/news/ 1048 1-500-scientists-lift-the-lid-on-reproducibility-1.19970. 1049

Pongrácz, A. Kis, 2015: Projected Bartholy, J., R. and changes of ex-1050 treme precipitation using multi-model approach. Idojaras, 119 (2),129 -1051 142, URL https://hungary.pure.elsevier.com/en/publications/ 1052 projected-changes-of-extreme-precipitation-using-multi-model-appr. 1053

Bedia, J., M. Iturbide, S. Herrera, R. Manzanas, and J. Gutiérrez, 2016: downscaleR: Climate
 data manipulation, bias correction and statistical downscaling. URL http://github.com/
 SantanderMetGroup/downscaleR/wiki, r package version 2.0-0.

Benestad, R., I. Hanssen-Bauer, and D. Chen, 2008: *Empirical-Statistical Downscaling*. World Scientific,
 URL http://www.worldscientific.com/worldscibooks/10.1142/6908.

Benestad, R., A. Mezghani, and K. Parding, 2015a: *esd: Climate analysis and empirical-statistical down-scaling (ESD) package for monthly and daily data*. URL http://rcg.gvc.gu.se/edu/esd.pdf,
 r package version 1.0.

Benestad, R. E., 2002: Empirically downscaled temperature scenarios for northern Europe based on a multi-model ensemble. *Climate Research*, **21** (2), 105–125, doi:10.3354/cr021105, URL http: //www.int-res.com/abstracts/cr/v21/n2/p105–125/.

Benestad, R. E., D. Chen, A. Mezghani, L. Fan, and K. Parding, 2015b: On using principal components to
 represent stations in empirical-statistical downscaling. *Tellus A: Dynamic Meteorology and Oceanogra- phy*, **67 (1)**, 28 326, doi:10.3402/tellusa.v67.28326, URL http://www.tandfonline.com/doi/
 abs/10.3402/tellusa.v67.28326.

Benestad, R. E. and A. Mezghani, 2015: On downscaling probabilities for heavy 24-hour precipitation events at seasonal-to-decadal scales. *Tellus A: Dynamic Meteorology and Oceanography*, 67 (1), 25 954, doi:10.3402/tellusa.v67.25954, URL http://www.tandfonline.com/doi/abs/10.
 3402/tellusa.v67.25954.

Brands, S., J. M. Gutiérrez, S. Herrera, and A. S. Cofiño, 2012: On the Use of Reanalysis Data for Downscaling. *Journal of Climate*, **25** (7), 2517–2526, doi:10.1175/JCLI-D-11-00251.1, URL http: //www.meteo.unican.es/en/node/73004.

Bürger, G., T. Q. Murdock, A. T. Werner, S. R. Sobie, and A. J. Cannon, 2012: Downscaling Extremes—An
Intercomparison of Multiple Statistical Methods for Present Climate. *Journal of Climate*, 25 (12), 4366–
4388, doi:10.1175/JCLI-D-11-00408.1, URL http://journals.ametsoc.org/doi/abs/10.
1175/JCLI-D-11-00408.1.

 Casanueva, A., S. Herrera, J. Fernández, and J. M. Gutiérrez, 2016a: Towards a fair comparison of statistical and dynamical downscaling in the framework of the EURO-CORDEX initiative. *Climatic Change*, 1– 16, doi:10.1007/s10584-016-1683-4, URL http://link.springer.com/article/10.1007/ s10584-016-1683-4. Casanueva, A., et al., 2016b: Daily precipitation statistics in a EURO-CORDEX RCM ensemble: added
 value of raw and bias-corrected high-resolution simulations. *Climate Dynamics*, 47 (3-4), 719–737,
 doi:10.1007/s00382-015-2865-x, URL https://link.springer.com/article/10.1007/
 s00382-015-2865-x.

Chandler, R. E., 2005: On the use of generalized linear models for interpreting climate variability. *Environmetrics*, **16** (7), 699–715, doi:10.1002/env.731, URL http://onlinelibrary.wiley.com/ doi/10.1002/env.731/abstract.

Chandler, R. E. and H. S. Wheater, 2002: Analysis of rainfall variability using generalized linear models: A case study from the west of Ireland. *Water Resources Research*, **38** (10), 10–1–10– 11, doi:10.1029/2001WR000906, URL http://onlinelibrary.wiley.com/doi/10.1029/ 2001WR000906/abstract.

Chen, J., F. P. Brissette, and R. Leconte, 2014: Assessing regression-based statistical approaches for down scaling precipitation over North America. *Hydrological Processes*, 28 (9), 3482–3504, doi:10.1002/hyp.
 9889, URL http://onlinelibrary.wiley.com/doi/10.1002/hyp.9889/abstract.

Christensen, J., E. Kjellstrom, F. Giorgi, G. Lenderink, and Rummukainen M, 2010: Weight assignment in
 regional climate models. *Climate Research*, 44 (2-3), 179–194, URL http://www.int-res.com/
 abstracts/cr/v44/n2-3/p179-194/.

Christensen, J. H. and O. B. Christensen, 2007: A summary of the PRUDENCE model projections of changes in European climate by the end of this century. *Climatic Change*, 81 (1), 7–30, doi:10.1007/s10584-006-9210-7, URL http://link.springer.com/article/10.1007/s10584-006-9210-7.

Dee, D. P., et al., 2011: The ERA-Interim reanalysis: configuration and performance of the data assimilation
 system. *Quarterly Journal of the Royal Meteorological Society*, 137 (656), 553–597, doi:10.1002/qj.828,
 URL http://onlinelibrary.wiley.com/doi/10.1002/qj.828/abstract.

Déqué, M., 2007: Frequency of precipitation and temperature extremes over France in an anthropogenic scenario: Model results and statistical correction according to observed values. *Global and Planetary Change*, 57 (1–2), 16–26, doi:10.1016/j.gloplacha.2006.11.030, URL http://www.sciencedirect.com/science/article/pii/S0921818106002748.

Dubrovský, M., 1997: Creating Daily Weather Series with Use of the Weather Generator. *Environmetrics*, 8 (5), 409–424, doi:10.1002/(SICI)1099-095X(199709/10)8:5(409::AID-ENV261)3.0.CO;2-0, URL http://onlinelibrary.wiley.com/doi/10.1002/(SICI)1099-095X(199709/10) 8:5<409::AID-ENV261>3.0.CO;2-0/abstract.

Dubrovský, M., J. Buchtele, and Z. Žalud, 2004: High-Frequency and Low-Frequency Variability in
 Stochastic Daily Weather Generator and Its Effect on Agricultural and Hydrologic Modelling. *Climatic Change*, 63 (1-2), 145–179, doi:10.1023/B:CLIM.0000018504.99914.60, URL http://link.
 springer.com/article/10.1023/B:CLIM.0000018504.99914.60.

Eden, J. M., M. Widmann, D. Maraun, and M. Vrac, 2014: Comparison of GCM- and RCM-simulated
 precipitation following stochastic postprocessing. *Journal of Geophysical Research: Atmospheres*,
 2014JD021732, doi:10.1002/2014JD021732, URL http://onlinelibrary.wiley.com/doi/
 10.1002/2014JD021732/abstract.

Flato, G., et al., 2013: Evaluation of Climate Models. Climate Change 2013: The Physical Science
 Basis. Contribution of Working Group I to the Fifth Assessment Report of the Intergovernmental
 Panel on Climate Change, T. Stocker, D. Qin, G.-K. Plattner, M. Tignor, S. Allen, J. Boschung,

A. Nauels, Y. Xia, V. Bex, and P. Midgley, Eds., Cambridge University Press, Cambridge, United Kingdom and New York, NY, USA, 741–866, URL www.climatechange2013.org, dOI: 10.1017/CBO9781107415324.020.

Fowler, H. J., S. Blenkinsop, and C. Tebaldi, 2007: Linking climate change modelling to impacts studies:
 recent advances in downscaling techniques for hydrological modelling. *International Journal of Clima- tology*, 27 (12), 1547–1578, doi:10.1002/joc.1556, URL http://onlinelibrary.wiley.com/
 doi/10.1002/joc.1556/abstract.

Frost, A. J., et al., 2011: A comparison of multi-site daily rainfall downscaling techniques under Australian conditions. *Journal of Hydrology*, **408** (1–2), 1–18, doi:10.1016/j.jhydrol.2011.06.021, URL http:// www.sciencedirect.com/science/article/pii/S0022169411004525.

Gaitan, C. F., W. W. Hsieh, A. J. Cannon, and P. Gachon, 2014: Evaluation of Linear and Non-Linear Downscaling Methods in Terms of Daily Variability and Climate Indices: Surface Temperature in Southern
Ontario and Quebec, Canada. *Atmosphere-Ocean*, 52 (3), 211–221, doi:10.1080/07055900.2013.857639,
URL http://dx.doi.org/10.1080/07055900.2013.857639.

Giorgi, F. and L. O. Mearns, 1991: Approaches to the simulation of regional climate change: A review.
 Reviews of Geophysics, 29 (2), 191–216, doi:10.1029/90RG02636, URL http://onlinelibrary.
 wiley.com/doi/10.1029/90RG02636/abstract.

Gutiérrez, J. M., D. San-Martín, S. Brands, R. Manzanas, and S. Herrera, 2013: Reassessing Statistical
 Downscaling Techniques for Their Robust Application under Climate Change Conditions. *Journal of Climate*, 26 (1), 171–188, doi:10.1175/JCLI-D-11-00687.1, URL http://journals.ametsoc.
 org/doi/abs/10.1175/JCLI-D-11-00687.1.

Gutjahr, O. and G. Heinemann, 2013: Comparing precipitation bias correction methods for high-resolution
 regional climate simulations using COSMO-CLM. *Theoretical and Applied Climatology*, **114** (3-4),
 511–529, doi:10.1007/s00704-013-0834-z, URL http://link.springer.com/article/10.
 1007/s00704-013-0834-z.

Gutmann, E., T. Pruitt, M. P. Clark, L. Brekke, J. R. Arnold, D. A. Raff, and R. M. Rasmussen, 2014: An
intercomparison of statistical downscaling methods used for water resource assessments in the United
States. *Water Resources Research*, 50 (9), 7167–7186, doi:10.1002/2014WR015559, URL http://
onlinelibrary.wiley.com/doi/10.1002/2014WR015559/abstract.

HanssenBauer, I., C. Achberger, R. E. Benestad, D. Chen, and E. J. Frland, 2005: Statistical downscaling
 of climate scenarios over Scandinavia. *Climate Research*, 29 (3), 255–268, doi:10.3354/cr029255, URL
 http://www.int-res.com/abstracts/cr/v29/n3/p255-268/.

Hartkamp, A. D., J. W. White, and G. Hoogenboom, 2003: Comparison of three weather generators for crop modeling: a case study for subtropical environments. *Agricultural Systems*, 76 (2), 539–560, doi:10.1016/S0308-521X(01)00108-1, URL http://www.sciencedirect. com/science/article/pii/S0308521X01001081.

Haylock, M. R., G. C. Cawley, C. Harpham, R. L. Wilby, and C. M. Goodess, 2006: Downscaling heavy precipitation over the United Kingdom: a comparison of dynamical and statistical methods and their future scenarios. *International Journal of Climatology*, 26 (10), 1397–1415, doi:10.1002/joc.1318, URL http://onlinelibrary.wiley.com/doi/10.1002/joc.1318/abstract.

Hempel, S., K. Frieler, L. Warszawski, J. Schewe, and F. Piontek, 2013: A trend-preserving bias correction
- the ISI-MIP approach. *Earth Syst. Dynam.*, 4 (2), 219–236, doi:10.5194/esd-4-219-2013, URL http:

1169 //www.earth-syst-dynam.net/4/219/2013/.

Herrera, S., M. Turco, and J. M. Gutiérrez, 2017: A MOS-Regression Technique for Temporally-Coherent
 Bias Correction of Regional Climate Model Simulations. *Climate Dynamics*, submitted.

Hertig, E. and J. Jacobeit, 2008: Downscaling future climate change: Temperature scenarios for the Mediterranean area. *Global and Planetary Change*, 63 (2–3), 127–131, doi:10.1016/j.
gloplacha.2007.09.003, URL http://www.sciencedirect.com/science/article/pii/
S0921818107001749.

Hertig, E., S. Seubert, A. Paxian, G. Vogt, H. Paeth, and J. Jacobeit, 2013: Changes of total versus extreme precipitation and dry periods until the end of the twenty-first century: statistical assessments for the Mediterranean area. *Theoretical and Applied Climatology*, 111 (1-2), 1–20, doi:10.1007/s00704-012-0639-5, URL http://link.springer.com/article/10.1007/s00704-012-0639-5.

 Hertig, E., S. Seubert, A. Paxian, G. Vogt, H. Paeth, and J. Jacobeit, 2014: Statistical modelling of extreme
 precipitation indices for the Mediterranean area under future climate change. *International Journal of Climatology*, 34 (4), 1132–1156, doi:10.1002/joc.3751, URL http://onlinelibrary.wiley.
 com/doi/10.1002/joc.3751/abstract.

Hessami, M., P. Gachon, T. B. M. J. Ouarda, and A. St-Hilaire, 2008: Automated regression-based
statistical downscaling tool. *Environmental Modelling & Software*, 23 (6), 813–834, doi:10.1016/
j.envsoft.2007.10.004, URL http://www.sciencedirect.com/science/article/pii/
S1364815207001867.

Hu, Y., S. Maskey, and S. Uhlenbrook, 2013: Downscaling daily precipitation over the Yellow River source
 region in China: a comparison of three statistical downscaling methods. *Theoretical and Applied Clima- tology*, **112** (3-4), 447–460, doi:10.1007/s00704-012-0745-4, URL http://link.springer.com/
 article/10.1007/s00704-012-0745-4.

Huth, R., 1999: Statistical downscaling in central Europe: evaluation of methods and potential predictors. *Climate Research*, **13** (2), 91–101, doi:10.3354/cr013091, URL http://www.int-res.com/ abstracts/cr/v13/n2/p91-101/.

Huth, R., 2002: Statistical Downscaling of Daily Temperature in Central Europe. Journal of Climate, 15 (13), 1731–1742, doi:10.1175/1520-0442(2002)015(1731:SDODTI)2.0.CO;
 URL http://journals.ametsoc.org/doi/abs/10.1175/1520-0442(2002)015%
 3C1731%3ASDODTI%3E2.0.CO%3B2.

Huth, R., 2004: Sensitivity of Local Daily Temperature Change Estimates to the Selection of Down scaling Models and Predictors. *Journal of Climate*, 17 (3), 640–652, doi:10.1175/1520-0442(2004)
 017(0640:SOLDTC)2.0.CO;2, URL http://journals.ametsoc.org/doi/abs/10.1175/
 1520-0442%282004%29017%3C0640%3ASOLDTC%3E2.0.CO%3B2.

Huth, R., 2005: Downscaling of humidity variables: a search for suitable predictors and predictands. *International Journal of Climatology*, **25** (2), 243–250, doi:10.1002/joc.1122, URL http: //onlinelibrary.wiley.com/doi/10.1002/joc.1122/abstract.

Huth, R., J. Mikšovský, P. Štěpánek, M. Belda, A. Farda, Z. Chládová, and P. Pišoft, 2015: Comparative validation of statistical and dynamical downscaling models on a dense grid in central Europe: temperature. *Theoretical and Applied Climatology*, **120** (3-4), 533–553, doi:10.1007/s00704-014-1190-3, URL http://link.springer.com/article/10.1007/s00704-014-1190-3.

Jacob, D., et al., 2014: EURO-CORDEX: new high-resolution climate change projections for European impact research. *Regional Environmental Change*, **14** (**2**), 563–578, doi:10.1007/s10113-013-0499-2. Jacobeit, J., E. Hertig, S. Seubert, and K. Lutz, 2014: Statistical downscaling for climate change projections in the Mediterranean region: methods and results. *Regional Environmental Change*, **14** (5), 1891–1906, doi:10.1007/s10113-014-0605-0, URL http://link.springer.com/article/ 10.1007/s10113-014-0605-0.

Jeong, D. I., A. St-Hilaire, T. B. M. J. Ouarda, and P. Gachon, 2012: Comparison of transfer functions in statistical downscaling models for daily temperature and precipitation over Canada. *Stochastic Environmental Research and Risk Assessment*, **26** (5), 633–653, doi:10.1007/s00477-011-0523-3, URL https://link.springer.com/article/10.1007/s00477-011-0523-3.

Keller, D. E., A. M. Fischer, C. Frei, M. A. Liniger, C. Appenzeller, and R. Knutti, 2015: Implementation and validation of a Wilks-type multi-site daily precipitation generator over a typical Alpine river catchment. *Hydrol. Earth Syst. Sci.*, **19** (5), 2163–2177, doi:10.5194/hess-19-2163-2015, URL http://www.hydrol-earth-syst-sci.net/19/2163/2015/.

Keller, D. E., A. M. Fischer, M. A. Liniger, C. Appenzeller, and R. Knutti, 2016: Testing a weather generator
 for downscaling climate change projections over Switzerland. *International Journal of Climatology*, n/a–
 n/a, doi:10.1002/joc.4750, URL http://dx.doi.org/10.1002/joc.4750.

 Kilsby, C. G., P. S. P. Cowpertwait, P. E. O'Connell, and P. D. Jones, 1998: Predicting rainfall statistics in England and Wales using atmospheric circulation variables. *International Journal of Climatology*, 18 (5), 523–539, doi:10.1002/(SICI)1097-0088(199804)
 18:5(523::AID-JOC268)3.0.CO;2-X, URL http://onlinelibrary.wiley.com/doi/10.
 1002/(SICI)1097-0088(199804)18:5<523::AID-JOC268>3.0.CO;2-X/abstract.

Klein Tank, A. M. G., et al., 2002: Daily dataset of 20th-century surface air temperature and precipita tion series for the European Climate Assessment. *International Journal of Climatology*, 22 (12), 1441–
 1453, doi:10.1002/joc.773, URL http://onlinelibrary.wiley.com/doi/10.1002/joc.
 773/abstract.

Kotlarski, S., et al., 2014: Regional climate modeling on European scales: a joint standard evaluation of the EURO-CORDEX RCM ensemble. *Geosci. Model Dev.*, 7 (4), 1297–1333, doi:10.5194/
 gmd-7-1297-2014, URL http://www.geosci-model-dev.net/7/1297/2014/.

Kottek, M., J. Grieser, C. Beck, B. Rudolf, and F. Rubel, 2006: World Map of the Köppen-Geiger climate classification updated. *Meteorologische Zeitschrift*, **15** (**3**), 259–263, doi:10.1127/0941-2948/2006/0130.

Leung, L. R., L. O. Mearns, F. Giorgi, and R. L. Wilby, 2003: Regional Climate Research. *Bulletin* of the American Meteorological Society, **84** (1), 89–95, doi:10.1175/BAMS-84-1-89, URL http: //journals.ametsoc.org/doi/abs/10.1175/bams-84-1-89.

Li, H., J. Sheffield, and E. F. Wood, 2010: Bias correction of monthly precipitation and temperature fields from Intergovernmental Panel on Climate Change AR4 models using equidistant quantile matching. *Journal of Geophysical Research: Atmospheres (1984–2012)*, **115 (D10)**, doi:10.1029/2009JD012882, URL http://onlinelibrary.wiley.com/doi/10.1029/2009JD012882/abstract.

Manzanas, R., S. Brands, D. San-Martín, A. Lucero, C. Limbo, and J. M. Gutiérrez, 2015: Statistical
 Downscaling in the Tropics Can Be Sensitive to Reanalysis Choice: A Case Study for Precipitation in
 the Philippines. *Journal of Climate*, 28 (10), 4171–4184, doi:10.1175/JCLI-D-14-00331.1, URL http:
 //journals.ametsoc.org/doi/abs/10.1175/JCLI-D-14-00331.1.

Maraun, D. and M. Widmann, 2017: *Statistical Downscaling and Bias Correction for Climate Research by Douglas Maraun.* Cambridge University Press, URL /core/books/ statistical-downscaling-and-bias-correction-for-climate-research/ 4ED479BAA8309C7ECBE6136236E3960F. Maraun, D., et al., 2010: Precipitation downscaling under climate change: Recent developments to
 bridge the gap between dynamical models and the end user. *Reviews of Geophysics*, 48 (3), n/a n/a, doi:10.1029/2009RG000314, URL http://onlinelibrary.wiley.com/doi/10.1029/
 2009RG000314/abstract.

Maraun, D., et al., 2015: VALUE: A framework to validate downscaling approaches for climate change studies. *Earth's Future*, **3** (1), 2014EF000259, doi:10.1002/2014EF000259, URL http:// onlinelibrary.wiley.com/doi/10.1002/2014EF000259/abstract.

Maraun, D., et al., 2017: Towards process-informed bias correction of climate change simulations. *Nature Climate Change*, in press.

Meijgaard, E. v., L. H. v. Ulft, G. Lenderink, S. R. d. Roode, E. L. Wipfler, R. Boers, and
 R. M. A. Timmermans, 2012: *Refinement and Application of a Regional Atmospheric Model for Climate Scenario Calculations of Western Europe*. Programme Office Climate changes Spatial
 Planning, URL http://www.wur.nl/de/Publicatie-details.htm?publicationId=
 publication-way-343237303937.

Monjo, R., G. Chust, and V. Caselles, 2014: Probabilistic correction of RCM precipitation
 in the Basque Country (Northern Spain). *Theoretical and Applied Climatology*, 117 (1-2),
 317–329, doi:10.1007/s00704-013-1008-8, URL http://link.springer.com/article/10.
 1007/s00704-013-1008-8.

Obled, C., G. Bontron, and R. Garçon, 2002: Quantitative precipitation forecasts: a statistical adaptation of model outputs through an analogues sorting approach. *Atmospheric Research*, **63** (3-4), 303-324, doi:10.1016/S0169-8095(02)00038-8, URL http://www.sciencedirect.
com/science/article/pii/S0169809502000388.

Estimation of future precipitation con-Pongrácz. R., J. Bartholy, and A. Kis, 2014: 1279 ditions for Hungary with special focus on dry periods. Idojaras, 118 (4), 305 -1280 321. URL https://hungary.pure.elsevier.com/hu/publications/ 1281 estimation-of-future-precipitation-conditions-for-hungary-with-sp. 1282

Räisänen, J. and O. Räty, 2013: Projections of daily mean temperature variability in the future: cross-validation tests with ENSEMBLES regional climate simulations. *Climate Dynamics*, 41 (5-6), 1553–1568, doi:10.1007/s00382-012-1515-9, URL http://link.springer.com/article/ 10.1007/s00382-012-1515-9.

Räty, O., J. Räisänen, and J. S. Ylhäisi, 2014: Evaluation of delta change and bias correction methods for
 future daily precipitation: intermodel cross-validation using ENSEMBLES simulations. *Climate Dynamics*, 42 (9), 2287–2303, doi:10.1007/s00382-014-2130-8, URL http://dx.doi.org/10.1007/
 s00382-014-2130-8.

Raynaud, D., B. Hingray, I. Zin, S. Anquetin, S. Debionne, and R. Vautard, 2016: Atmospheric analogues for
 physically consistent scenarios of surface weather in Europe and Maghreb. *International Journal of Cli- matology*, doi:10.1002/joc.4844, URL http://onlinelibrary.wiley.com/doi/10.1002/
 joc.4844/abstract.

Reiter, P., O. Gutjahr, L. Schefczyk, G. Heinemann, and M. Casper, 2017: Does applying quantile mapping to subsamples improve the bias correction of daily precipitation? *International Journal of Climatology*, n/a-n/a, doi:10.1002/joc.5283, URL http://onlinelibrary.wiley.com/doi/10.
 1002/joc.5283/abstract.

- Ribalaygua, J., L. Torres, J. Pórtoles, R. Monjo, E. Gaitán, and M. R. Pino, 2013: Description and validation
- of a two-step analogue/regression downscaling method. *Theoretical and Applied Climatology*, **114** (1-2),
- 1301 253-269, doi:10.1007/s00704-013-0836-x, URL http://link.springer.com/article/10. 1007/s00704-013-0836-x.

Richardson, C. W., 1981: Stochastic simulation of daily precipitation, temperature, and solar radiation. Water Resources Research, 17 (1), 182–190, doi:10.1029/WR017i001p00182, URL http:// onlinelibrary.wiley.com/doi/10.1029/WR017i001p00182/abstract.

Rummukainen, M., 2010: State-of-the-art with regional climate models. Wiley Interdisciplinary Reviews: *Climate Change*, 1 (1), 82–96, doi:10.1002/wcc.8, URL http://onlinelibrary.wiley.com/ doi/10.1002/wcc.8/abstract.

San-Martín, D., R. Manzanas, S. Brands, S. Herrera, and J. M. Gutiérrez, 2017: Reassessing Model Uncertainty for Regional Projections of Precipitation with an Ensemble of Statistical Downscaling Methods. *Journal of Climate*, **30** (1), 203–223, doi:10.1175/JCLI-D-16-0366.1, URL http://journals.
 ametsoc.org/doi/10.1175/JCLI-D-16-0366.1.

Semenov, M. A., R. J. Brooks, E. M. Barrow, and C. W. Richardson, 1998: Comparison of the WGEN and
 LARS-WG stochastic weather generators for diverse climates. *Climate Research*, 10 (2), 95–107, doi:
 10.3354/cr010095, URL http://www.int-res.com/abstracts/cr/v10/n2/p95-107/.

Silva, S. and J. Almeida, 2003: GPLAB - A Genetic Programming Toolbox for MATLAB. Nor3dic MAT LAB Conference (NMC-2003), URL https://www.cisuc.uc.pt/publication/show/1290.

Takayabu, I., H. Kanamaru, K. Dairaku, R. Benestad, H. v. Storch, and J. H. Christensen, 2016: Reconsider ing the Quality and Utility of Downscaling. *Journal of the Meteorological Society of Japan. Ser. II*, 94A, 31–45, doi:10.2151/jmsj.2015-042.

Taylor, K. E., R. J. Stouffer, and G. A. Meehl, 2011: An Overview of CMIP5 and the Experiment Design.
 Bulletin of the American Meteorological Society, 93 (4), 485–498, doi:10.1175/BAMS-D-11-00094.1,
 URL http://journals.ametsoc.org/doi/abs/10.1175/BAMS-D-11-00094.1.

Teutschbein, C. and J. Seibert, 2013: Is bias correction of regional climate model (RCM) simulations possible for non-stationary conditions? *Hydrol. Earth Syst. Sci.*, **17** (**12**), 5061–5077, doi:10.5194/ hess-17-5061-2013, URL http://www.hydrol-earth-syst-sci.net/17/5061/2013/.

 Teutschbein, C., F. Wetterhall, and J. Seibert, 2011: Evaluation of different downscaling techniques for hydrological climate-change impact studies at the catchment scale. *Climate Dynamics*, 37 (9-10), 2087–2105, doi:10.1007/s00382-010-0979-8, URL http://link.springer.com/article/ 10.1007/s00382-010-0979-8.

Themeßl, M. J., A. Gobiet, and G. Heinrich, 2012: Empirical-statistical downscaling and error correction of regional climate models and its impact on the climate change signal. *Climatic Change*, 112 (2), 449–468, doi:10.1007/s10584-011-0224-4, URL http://link.springer.com/article/10.
 1007/s10584-011-0224-4.

Timbal, B., A. Dufour, and B. McAvaney, 2003: An estimate of future climate change for west ern France using a statistical downscaling technique. *Climate Dynamics*, 20 (7-8), 807–823,
 doi:10.1007/s00382-002-0298-9, URL https://link.springer.com/article/10.1007/
 s00382-002-0298-9.

1339 Turco, M., M. C. Llasat, S. Herrera, and J. M. Gutiérrez, 2017: Bias correction and downscaling of future

RCM precipitation projections using a MOS-Analog technique. *Journal of Geophysical Research: Atmo-*

spheres, 122 (5), 2631-2648, doi:10.1002/2016JD025724, URL http://onlinelibrary.wiley.

1342 com/doi/10.1002/2016JD025724/abstract.

Turco, M., P. Quintana-Seguí, M. C. Llasat, S. Herrera, and J. M. Gutiérrez, 2011: Testing MOS pre cipitation downscaling for ENSEMBLES regional climate models over Spain. *Journal of Geophysical Research*, 116 (D18), doi:10.1029/2011JD016166, URL http://www.meteo.unican.es/en/
 node/73023.

 Štěpánek, P., Zahradníček, P., Farda, A., Skalák, P., Trnka, M., Meitner, J., and Rajdl, K., 2016: Projection of drought-inducing climate conditions in the Czech Republic according to Euro-CORDEX models. *Climate Research*, **70** (2-3), 179–193, URL http://www.int-res.com/abstracts/cr/v70/n2-3/ p179–193/.

 Vaittinada Ayar, P., M. Vrac, S. Bastin, J. Carreau, M. Déqué, and C. Gallardo, 2016: Intercomparison of statistical and dynamical downscaling models under the EURO- and MED-CORDEX initiative framework:
 present climate evaluations. *Climate Dynamics*, 46 (3-4), 1301–1329, doi:10.1007/s00382-015-2647-5, URL http://link.springer.com/article/10.1007/s00382-015-2647-5.

 Volosciuk, C., D. Maraun, M. Vrac, and M. Widmann, 2017: A combined statistical bias correction and stochastic downscaling method for precipitation. *Hydrol. Earth Syst. Sci.*, 21 (3), 1693–1719, doi: 10.5194/hess-21-1693-2017, URL http://www.hydrol-earth-syst-sci.net/21/1693/ 2017/.

Vrac, M., P. Drobinski, A. Merlo, M. Herrmann, C. Lavaysse, L. Li, and S. Somot, 2012: Dynam ical and statistical downscaling of the French Mediterranean climate: uncertainty assessment. *Nat. Hazards Earth Syst. Sci.*, 12 (9), 2769–2784, doi:10.5194/nhess-12-2769-2012, URL http://www.
 nat-hazards-earth-syst-sci.net/12/2769/2012/.

Widmann, M., C. S. Bretherton, and E. P. Salathé, 2003: Statistical Precipitation Downscaling over the Northwestern United States Using Numerically Simulated Precipitation as a Predictor. Journal of Climate, 16 (5), 799–816, doi:10.1175/1520-0442(2003)016(0799:SPDOTN)2.0.CO;
2, URL http://journals.ametsoc.org/doi/abs/10.1175/1520-0442(2003)016%
3C0799%3ASPDOTN%3E2.0.CO%3B2.

Wilby, R. L., S. P. Charles, E. Zorita, B. Timbal, P. Whetton, and L. O. Mearns, 2004: Guidelines for Use
 of Climate Scenarios Developed from Statistical Downscaling Methods. Supporting Material, Intergov ernmental Panel on Climate Change. URL http://www.ipcc-data.org/guidelines/dgm_
 no2_v1_09_2004.pdf, accessed on 28 Aug 2013.

Wilby, R. L., C. W. Dawson, and E. M. Barrow, 2002: sdsm — a decision support tool for the assessment of regional climate change impacts. *Environmental Modelling & Software*, 17 (2), 145–157, doi:10.1016/S1364-8152(01)00060-3, URL http://www.sciencedirect. com/science/article/pii/S1364815201000603.

Wilby, R. L., T. M. L. Wigley, D. Conway, P. D. Jones, B. C. Hewitson, J. Main, and D. S. Wilks, 1998:
 Statistical downscaling of general circulation model output: A comparison of methods. *Water Resources Research*, 34 (11), 2995–3008, doi:10.1029/98WR02577, URL http://onlinelibrary.wiley.
 com/doi/10.1029/98WR02577/abstract.

Wilcke, R. A. I., T. Mendlik, and A. Gobiet, 2013: Multi-variable error correction of regional climate
 models. *Climatic Change*, 120 (4), 871–887, doi:10.1007/s10584-013-0845-x, URL http://link.
 springer.com/article/10.1007/s10584-013-0845-x.

Wilks, D. S. and R. L. Wilby, 1999: The weather generation game: a review of stochastic weather models.
 Progress in Physical Geography, 23 (3), 329–357, doi:10.1177/030913339902300302, URL http://
 ppg.sagepub.com/content/23/3/329.

Winkler, J. A., et al., 2011: Climate Scenario Development and Applications for Local/Regional Climate Change Impact Assessments: An Overview for the Non-Climate Scientist. *Geography Compass*, 5 (6), 275–300, doi:10.1111/j.1749-8198.2011.00425.x, URL http://onlinelibrary.wiley. com/doi/10.1111/j.1749-8198.2011.00425.x/abstract.

Wong, G., D. Maraun, M. Vrac, M. Widmann, J. M. Eden, and T. Kent, 2014: Stochastic Model Output
 Statistics for Bias Correcting and Downscaling Precipitation Including Extremes. *Journal of Climate*,
 27 (18), 6940–6959, doi:10.1175/JCLI-D-13-00604.1, URL http://journals.ametsoc.org/
 doi/abs/10.1175/JCLI-D-13-00604.1.

Yang, W., J. Andréasson, L. P. Graham, J. Olsson, J. Rosberg, and F. Wetterhall, 2010: Distribution-based
 scaling to improve usability of regional climate model projections for hydrological climate change impacts studies. *Hydrology Research*, 41 (3-4), 211–229, doi:10.2166/nh.2010.004, URL http://hr.
 iwaponline.com/content/41/3-4/211.

 Yang, W., M. Gardelin, J. Olsson, and T. Bosshard, 2015: Multi-variable bias correction: application of forest fire risk in present and future climate in Sweden. *Nat. Hazards Earth Syst. Sci.*, 15 (9), 2037–2057, doi:10.5194/nhess-15-2037-2015, URL http://www.nat-hazards-earth-syst-sci.net/
 15/2037/2015/.

Zerenner, T., V. Venema, P. Friederichs, and C. Simmer, 2016: Downscaling near-surface atmospheric fields with multi-objective Genetic Programming. *Environmental Modelling & Software*, 84, 85–98, doi:10.1016/j.envsoft.2016.06.009, URL http://www.sciencedirect.com/science/
article/pii/S1364815216302122.

Zitzler, E. and L. Thiele, 1999: Multiobjective evolutionary algorithms: a comparative case study and
the strength Pareto approach. *IEEE Transactions on Evolutionary Computation*, 3 (4), 257–271, doi:
10.1109/4235.797969.

| Table 1: | List | of statio | ons indicati | ng the or | der (sorted | l by latitu | ide), |
|----------|-------|-----------|--------------|-----------|-------------|-------------|-------|
| ECA&D | IDs, | name, | longitude, | latitude, | elevation, | country, | and |
| Köppen- | -Geig | er clima | ate type. | | | • | |

| # | ID | Name | Lon. | Lat. | Elev. | Country | Köppen |
|----|------|----------------------------|-------|-------|-----------|------------------------|--------|
| 1 | 231 | Malaga | -4.49 | 36.67 | 7 | Spain | Csa |
| 2 | 63 | Methoni | 21.70 | 36.83 | 51 | Greece | Csa |
| 3 | 214 | Lisboa-Geofisica | -9.15 | 38.72 | 77 | Portugal | Csa |
| 4 | 229 | Badajoz/Talavera-La-Real | -6.83 | 38.88 | 185 | Spain | Csa |
| 5 | 175 | Cagliari | 9.05 | 39.23 | 21 | Italy | Csa |
| 6 | 3919 | Palma-De-Mallorca | 2.74 | 39.56 | 8 | Spain | BSk |
| 7 | 59 | Corfu | 19.92 | 39.62 | 11 | Greece | Csa |
| 8 | 62 | Larissa | 22.45 | 39.65 | 72 | Greece | BSk |
| 9 | 3946 | Madrid-Barajas | -3.56 | 40.47 | 609 | Spain | BSk |
| 10 | 232 | Navacerrada | -4.01 | 40.78 | 1894 | Spain | Csb |
| 11 | 236 | Tortosa-Observatorio-Ebro | 0.49 | 40.82 | 44 | Spain | Csa |
| 12 | 176 | Roma-Ciampino | 12.58 | 41.78 | 105 | Italy | Csa |
| 13 | 212 | Braganca | -6.73 | 41.80 | 690 | Portugal | Csb |
| 14 | 1394 | Santiago-De-Compostela | -8.41 | 42.89 | 370 | Spain | Cfb |
| 15 | 1686 | Hvar | 16.45 | 43.17 | 20 | Croatia | Csa |
| 16 | 234 | San-Sebastian-Igueldo | -2.04 | 43.31 | 251 | Spain | Cfb |
| 17 | 39 | Marseille-Marignane | 5.23 | 43.44 | 5 | France | Csa |
| 18 | 800 | Toulouse-Blagnac | 1.38 | 43.62 | 151 | France | Cfa |
| 19 | 355 | Mont-Aigoual | 3.58 | 44.12 | 1567 | France | Cfb |
| 20 | 2062 | Constanta | 28.63 | 44.22 | 13 | Romania | Cfa |
| 21 | 219 | Bucuresti-Baneasa | 26.08 | 44.52 | 90 | Romania | Cfa |
| 22 | 1684 | Gospic | 15.37 | 44.55 | 564 | Croatia | Cib |
| 23 | 168/ | Zavizan | 14.98 | 44.82 | 1594 | Croatia | Dic |
| 24 | 1// | Verona-Villafranca | 10.8/ | 45.38 | 68 150 | Italy | Cia |
| 25 | 1/3 | Nilan Sibir | 9.19 | 45.47 | 150 | Italy | Cia |
| 20 | 450 | Sibiu Zaarah Cria | 24.15 | 45.80 | 444 | Romania | CID |
| 21 | 242 | Zagred-Gric | 13.98 | 45.82 | 100 | Croatia | Cla |
| 20 | 242 | Arad | 0.97 | 40.00 | 300 | Switzeriand Domenie | Cla |
| 29 | 1662 | Arau Sion 2 | 21.33 | 40.15 | 110 | Komania Switzerland | CID |
| 31 | 1002 | Sonnblick | 12.05 | 40.22 | 3106 | Austria | ET |
| 32 | 32 | Bourges | 2 37 | 47.03 | 161 | France | Cfb |
| 33 | 12 | Graz | 15 45 | 47.07 | 366 | Austria | Cfb |
| 34 | 951 | Iasi | 27.63 | 47.00 | 102 | Romania | Cfa |
| 35 | 243 | Saentis | 9 35 | 47 25 | 2502 | Switzerland | ET |
| 36 | 13 | Innsbruck | 11 40 | 47 27 | 577 | Austria | Cfb |
| 37 | 244 | Zueriswitzerland | 8 57 | 47 38 | 556 | Switzerland | Cfb |
| 38 | 4002 | Oberstdorf | 10.28 | 47 40 | 806 | Germany | Cfb |
| 39 | 58 | Zugspitze | 10.99 | 47.42 | 2964 | Germany | ET |
| 40 | 239 | Basel-Binningen | 7.58 | 47.55 | 316 | Switzerland | Cfb |
| 41 | 14 | Salzburg | 13.00 | 47.80 | 437 | Austria | Cfb |
| 42 | 48 | Hohenpeissenberg | 11.01 | 47.80 | 977 | Germany | Cfb |
| 43 | 322 | Rennes | -1.73 | 48.07 | 36 | France | Cfb |
| 44 | 16 | Wien | 16.35 | 48.23 | 198 | Austria | Cfb |
| 45 | 38 | Paris-14e | 2.34 | 48.82 | 75 | France | Cfb |
| 46 | 2762 | Rheinstetten | 8.33 | 48.97 | 116 | Germany | Cfb |
| 47 | 4004 | Regensburg | 12.10 | 49.04 | 365 | Germany | Cfb |
| 48 | 3991 | Giessen-Wettenberg | 8.65 | 50.60 | 203 | Germany | Cfb |
| 49 | 17 | Uccle | 4.37 | 50.80 | 100 | Belgium | Cfb |
| 50 | 483 | Dresden-Klotzsswitzerlande | 13.76 | 51.13 | 227 | Germany | Cfb |
| 51 | 274 | Oxford | -1.27 | 51.77 | 63 | UK | Cfb |
| 52 | 2006 | Brocken | 10.62 | 51.80 | 1142 | Germany | Dfc |
| 53 | 333 | Siedlce | 22.25 | 52.25 | 152 | Poland | Dfb |
| 54 | 54 | Potsdam | 13.06 | 52.38 | 81 | Germany | Cfb |
| 55 | 42 | Bremen | 8.80 | 53.05 | 4 | Germany | Cfb |
| 56 | 351 | Waddington | 0.52 | 53.17 | 68 | UK | Cfb |
| 57 | 350 | Valley | -4.53 | 53.25 | 11 | UK | Cfb |
| 58 | 468 | Helgoland | 7.89 | 54.18 | 4 | Germany | Cfb |
| 59 | 1020 | Lazdijai | 23.52 | 54.23 | 133 | Lithuania | Dfb |

| Table 1: | List | of static | ons indicati | ng the or | der (sorted | l by latitu | ıde), |
|----------|-------|-----------|--------------|-----------|-------------|-------------|-------|
| ECA&D | IDs, | name, | longitude, | latitude, | elevation, | country, | and |
| Köppen- | -Geig | er clima | ate type. | | | 5, | |

| # | m | Name | Lon | Lat | Elev | Country | Können |
|----|------|--------------------------|-------|-------|--------|-----------|--------|
| π | m | 1 and | LUII. | Lat. | Eac v. | Country | robhen |
| 60 | 3994 | Arkona | 13.44 | 54.68 | 42 | Germany | Cfb |
| 61 | 332 | Leba | 17.53 | 54.75 | 2 | Poland | Dfb |
| 62 | 200 | Kaunas | 23.83 | 54.88 | 77 | Lithuania | Dfb |
| 63 | 272 | Eskdalemuir | -3.20 | 55.32 | 242 | UK | Cfb |
| 64 | 201 | Klaipeda | 21.07 | 55.73 | 6 | Lithuania | Cfb |
| 65 | 113 | Tranebjerg | 10.60 | 55.85 | 11 | Denmark | Cfb |
| 66 | 1009 | Birzai | 24.77 | 56.20 | 60 | Lithuania | Dfb |
| 67 | 107 | Vestervig | 8.32 | 56.77 | 18 | Denmark | Cfb |
| 68 | 465 | Visby | 18.33 | 57.67 | 42 | Sweden | Cfb |
| 69 | 462 | Goteborg | 11.99 | 57.72 | 5 | Sweden | Cfb |
| 70 | 349 | Stornoway | -6.32 | 58.33 | 9 | UK | Cfb |
| 71 | 275 | Wick | -3.08 | 58.45 | 36 | UK | Cfb |
| 72 | 192 | Faerder | 10.53 | 59.03 | 6 | Norway | Cfb |
| 73 | 194 | Utsira-Fyr | 4.88 | 59.31 | 55 | Norway | Cfb |
| 74 | 28 | Helsinki-Kaisaniemi | 24.95 | 60.18 | 4 | Finland | Dfb |
| 75 | 708 | Jokioinen-Jokioisten | 23.50 | 60.81 | 104 | Finland | Dfb |
| 76 | 5585 | Salen | 13.26 | 61.17 | 360 | Sweden | Dfc |
| 77 | 191 | Kjoeremsgrende | 9.05 | 62.10 | 626 | Norway | Dfc |
| 78 | 330 | Fokstua | 9.28 | 62.12 | 952 | Norway | Dfc |
| 79 | 1051 | Tafjord | 7.42 | 62.23 | 15 | Norway | Csb |
| 80 | 29 | Jyvaskyla-Lentoasema | 25.68 | 62.40 | 139 | Finland | Dfc |
| 81 | 7682 | Siikajoki-Revonlahti | 25.09 | 64.68 | 48 | Finland | Dfc |
| 82 | 339 | Haparanda | 24.14 | 65.83 | 5 | Sweden | Dfc |
| 83 | 1427 | Jackvik | 17.00 | 66.38 | 430 | Sweden | Dfc |
| 84 | 30 | Sodankyla-Lapin-Ilmatiet | 26.63 | 67.37 | 179 | Finland | Dfc |
| 85 | 190 | Karasjok | 25.50 | 69.47 | 129 | Norway | Dfc |
| 86 | 195 | Vardoe | 31.08 | 70.37 | 14 | Norway | ET |

| Variable | Code | Levels | Units | Temporal Aggregation |
|--------------------------|------|-------------------------|---------|----------------------|
| Minimum Temperature | TMIN | - | Κ | Daily minimum |
| Maximum Temperature | TMAX | - | Κ | Daily maximum |
| Total Precipitation | PRC | - | m | Daily accumulated |
| Mean Sea Level Pressure | MSL | - | Pa | Daily Mean |
| 2m Temperature | 2T | 2m | Κ | Daily mean |
| Geopotential | Ζ | 250 500 700 850 1000 mb | m2 s-2 | Daily Mean |
| Temperature | Т | 250 500 700 850 1000 mb | Κ | Daily Mean |
| westerly wind component | U | 250 500 700 850 1000 mb | m s-1 | Daily Mean |
| southerly wind component | V | 250 500 700 850 1000 mb | m s-1 | Daily Mean |
| Specific humidity | Q | 250 500 700 850 1000 mb | kg kg-1 | Daily Mean |

Table 2: Description of the variables, pressure levels, units and temporal aggregation of the common set of predictors used in the reference VALUE dataset.

Table 3: Table of ESD methods contributing to VALUE Experiment 1a for precipitation using ERA-Interim predictors (and RACMO2 RCM predictors additionally, for those methods with a cross in the second column). CODE is the public code of the method as shown in (http://www.value-cost.eu/ validationportal). APPRO. and TECH. indicate the approach and techniques used, respectively. The codes used for the approaches are: RAW (raw data), MOS (Model Output Statistics), PP (Perfect Prognosis), WG (Weather Generators), and the families of techniques: S (additive/multiplicative scaling), PM (parametric quantile mapping), QM (empirical quantile mapping), WT (weather types), A (analogs), TF (transfer function), WG (Markov-type WGs). ST indicates the stochastic nature of the method (yes for stochastic ones, providing 100 realizations); MS and MV indicate whether the methods are multi-site and multi-variable, respectively (methods using PCs as predictors are indicated with a *yes* in the MS column). Finally, SE and AC indicate the explicit inclusion of seasonal and autocorrelation components, respectively. All methods provide daily data for the 86 stations. The shading indicates the subset of methods applied also for temperatures. (*) Only occurrence is randomized, amounts are based on inflated regression (in this case, a single realization was provided and used for validation).

| # | R | INSTITUTION | CODE | APPRO. | TECH. | ST | MS | MV | SE | AC |
|-----------------|--------|------------------|--------------------|--------|---------|--------|-----|-----|-----------|-----|
| 1 | - | ECMWF | ERAINT-200 | RAW | - | - | - | - | - | - |
| 2 | - | ECMWF | ERAINT-075 | RAW | - | - | - | - | - | - |
| 3 | X | KNMI | RACMO22E | RAW | - | - | - | - | - | - |
| 4 | X | UHEL | Ratyetal-M6 | MOS | S | no | no | no | yes | no |
| 5 | X | UHEL | Ratyetal-M7 | MOS | S | no | no | no | yes | no |
| 6 | Х | UCAN/CSIC | ISI-MIP | MOS | SPM | no | no | no | yes | no |
| 7 | X | SMHI | DBS | MOS | PM | no | no | ves | ves | no |
| 8 | Х | UHEL | Ratvetal-M9 | MOS | PM | no | no | no | ves | no |
| 9 | X | FIC | BC | MOS | PM | no | no | no | ves | no |
| 10 | X | UCAN/CSIC | ĞŌM | MOS | PM | no | no | no | no | no |
| ĨĬ | X | UCAN/CSIC | GPOM | MOS | PM | no | no | no | no | no |
| 12 | X | UCAN/CSIC | EOM | MÕŠ | OM | no | no | no | no | no |
| 13 | X | UCAN/CSIC | EÒMs | MOS | ÒМ | no | no | no | ves | no |
| 14 | X | UCAN/CSIC | EOM-WT | MOS | Òм∣wт | no | no | no | no | no |
| 15 | X | IDL | 0Mm | MOS | δM | no | no | no | ves | no |
| 16 | X | FLU | OMBC-BI-PR | MOS | δM | no | no | no | ves | no |
| 17 | X | LEC LSCE/IPSL | CDFt | MOS | ÔM | no | no | no | ves | no |
| 18 | X | GCRLCAS | OM-DAP | MOS | ÔM | no | no | no | ves | no |
| 10 | X | SMHI | FOM-WIC658 | MOS | ÔM | no | no | no | ves | no |
| $\frac{1}{20}$ | X | UHEI | Ratvetal_M8 | MOS | ÔM | no | no | no | Ves | no |
| $\frac{20}{21}$ | X X | UB | MOS AN | MOS | | no | | no | yes no | no |
| $\frac{21}{22}$ | X X | UCAN/CSIC | MOS-AN MOS GI M | MOS | | | yes | no | no | no |
| $\frac{22}{23}$ | Λ | UNICEAT | VGLMGAMMA | MOS | | yes | no | no | | no |
| 23 | - | EIC | FICO2D | | | yes | 110 | no | yes | no |
| 24 | - | FIC | FIC02P | MOS PP | PMAIF | по | no | по | yes | no |
| 25 | - | FIC | FIC04P | MOS | PM A TF | no | no | no | yes | no |
| 26 | - | FIC | FIC01P | PP | A TF | no | yes | no | yes | no |
| 27 | - | FIC | FIC03P | PP | ATF | no | yes | no | yes | no |
| 28 | - | LSCE/IPSL | ANALOG-ANOM | PP | A | no | ves | yes | ves | no |
| 29 | - | UCAN/CSIC | ANALOG | PP | А | no | ves | ves | no | no |
| 30 | - | CNRS/IGE | ANALOG-MP | PP | А | ves | ves | ves | ves | no |
| 31 | - | CNRS/IGE | ANALOG-SP | PP | А | ves | ves | ves | ves | no |
| 32 | _ | MIUB | MO-GP | PP | TF | no | no | no | no | no |
| 33 | - | AEMET | GLM-P | PP | ŤĒ | ves(*) | no | no | no | no |
| 34 | - | CUNI | MLR-RAN | PP | TF | no | no | no | no | no |
| 35 | - | CUNI | MLR-RSN | PP | TF | no | no | no | ves | no |
| 36 | - | CUNI | MLR-ASW | PP | TF | ves | no | no | ves | no |
| 37 | - | CUNI | MLR-ASI | PP | ŤĒ | no | no | no | ves | no |
| - 38 | - | UCAN/CSIC | GLM-DET | PP | TF | no | ves | no | no | no |
| 39 | _ | UCAN/CSIC | GLM | PP | TF | ves | ves | no | no | no |
| 40 | _ | UCAN/CSIC | GLM-WT | PP | TFWT | ves | ves | no | no | no |
| 41 | | UCAN/CSIC | WT-WG | PP | WT | ves | no | no | no | no |
| $\frac{1}{42}$ | | I SCE/IPSI | SWG | PP | TE | ves | ves | no | Ves | no |
| 12 | - | METEOSWICS | STU | WG | WG | yes | yes | NOC | yes | NOC |
| 43 | - | MELEO2W122 | MADEL DAGIC | WG | WG | yes | no | yes | yes | yes |
| 44 | - | IAF-CAS | MADEL TAD | WG | WG | yes | no | yes | yes | yes |
| 43 | - | IAP-CAS | MADEL M2 | WG | WC | yes | 110 | yes | yes | yes |
| 40 | - | IAP-CAS | MAKFI-MIS | WG | WG | yes | no | yes | yes | yes |
| 4/ | - | IAP-CAS | GOMEZ-BASIC | WG | WG | yes | no | yes | yes | yes |
| 48 | - | IAP-CAS | GOMEZ-TAD | WG | WG | yes | no | yes | yes | yes |

Table 4: As Table 3 but for minimum and maximum temperatures. All methods provide daily data for the 86 stations, except the ESD family (#39-42, in *italics*) which provide monthly data. The shading indicates the subset of methods applied also for precipitation.

| # | R | INSTITUTION | CODE | APPRO. | TECH. | ST | MS | MV | SE | AC |
|------------------|---|-------------|--------------|--------|---------------|------------|------------|------------|-----|-----|
| 1 | - | ECMWF | ERAINT-200 | RAW | - | - | - | - | - | - |
| 2 | - | ECMWF | ERAINT-075 | RAW | - | - | - | - | - | - |
| 3 | X | KNMI | RACMO22E | RAW | - | - | - | - | - | - |
| 4 | Х | UHEL | RaiRat-M6 | MOS | S | no | no | no | yes | no |
| 5 | X | UHEL | RaiRat-M7 | MOS | S | no | no | no | yes | no |
| 6 | X | UHEL | RaiRat-M8 | MOS | S | no | no | no | yes | no |
| 7 | Х | UL | SB | MOS | S | no | no | no | yes | no |
| 8 | Х | UCAN/CSIC | ISI-MIP | MOS | SPM | no | no | no | yes | no |
| 9 | Х | SMHI | DBS | MOS | PM | no | no | yes | ves | no |
| 10 | Х | UCAN/CSIC | GPOM | MOS | PM | no | no | no | no | no |
| 11 | Х | UCAN/CSIC | EQŇ | MOS | QM | no | no | no | no | no |
| 12 | Х | UCAN/CSIC | EQMs | MOS | QМ | no | no | no | yes | no |
| 13 | Х | UCAN/CSIC | EQM-WT | MOS | Q M WT | no | no | no | no | no |
| 14 | Х | IDL | OMm | MOS | ÔΜ΄ | no | no | no | ves | no |
| 15 | Х | ELU | ÒMBC-BJ-PR | MOS | ÒМ | no | no | no | ves | no |
| 16 | Х | LSCE/IPSL | ČDFt | MOS | ÒМ | no | no | no | ves | no |
| 17 | X | GCRI-CAS | OM-DAP | MOS | ÒМ | no | no | no | ves | no |
| 18 | X | SMHI | EOM-WIC658 | MOS | ÔΜ | no | no | no | ves | no |
| 19 | X | UHEL | RaiRat-M9 | MOS | δM | no | no | no | ves | no |
| 20 | X | ŬL. | DBBC | MOS | δM | no | no | no | ves | no |
| $\overline{21}$ | X | ŬĹ | DBD | MOS | δM | no | no | no | ves | no |
| $\bar{2}2$ | X | ŬĈAN/CSIC | MOS-REG | MOS | ŤF | no | no | no | no | no |
| $\bar{2}\bar{3}$ | - | FIC | FIC02T | MOS | PMATE | no | no | no | ves | no |
| $\frac{23}{24}$ | _ | FIC | FIC01T | PP | | no | Ves | no | ves | no |
| 25 | _ | I SCE/IPSI | ANALOG-ANOM | DD | Δ | no | Ves | Vec | Ves | no |
| $\frac{23}{26}$ | | LICAN/CSIC | ANALOG | DD | | no | Ves | ves | no | no |
| $\frac{20}{27}$ | | CNRS/IGE | ANALOG-MP | DD DD | | | yes ves | yes ves | | no |
| $\frac{27}{28}$ | | CNRS/IGE | ANALOG-SP | DD DD | | yes ves | yes ves | yes | yes | no |
| $\frac{20}{20}$ | | MILIR | MO-GP | DD DD | TE | no | no | no | no | no |
| $\frac{29}{30}$ | _ | AFMET | MIR-T | PP | TF | no | n0 | no | no | no |
| 31 | _ | CUNI | MIR-RAN | PP | TF | no | no | no | no | no |
| 32 | _ | CUNI | MLR-RSN | PP | ŤĒ | no | no | no | ves | no |
| 33 | _ | CUNI | MLR-ASW | PP | TF | ves | no | no | ves | no |
| 34 | _ | CUNI | MIR-ASI | PP | TF | no | no | no | ves | no |
| 35 | _ | CUNI | MLR-AAN | PP | TF | no | no | no | | no |
| 36 | - | CUNI | MLR-AAI | PP | TF | no | no | no | no | no |
| 37 | - | CUNI | MLR-AAW | PP | TF | ves | no | no | no | no |
| 38 | - | IGUA | MLR-PCA-ZTR | PP | TF | no | ves | no | ves | no |
| 39 | _ | AMU | ESD-EOFSLP | PP | TEWT | no | ves | no | ves | no |
| 10 | _ | AMU | ESD = EOF SD | DD | TEWT | no | vas | no | Ves | no |
| 11 | _ | | | | | no | yes | no | yes | no |
| 41 | - | | ESD-SLF | | | по | 110 | 110 | yes | 110 |
| 42 | - | AMU | ESD-12 | PP | | no | no | no | yes | no |
| 45 | - | UCAN/CSIC | | | | no | yes | no | no | no |
| 44 | - | UCAN/CSIC | MLK-WT | PP | IF WT | no | yes | no | no | no |
| 45 | - | UCAN/CSIC | WI-WG | PP | WI | yes | no | no | no | no |
| 40 | - | LSCE/IPSL | 200 | PP | | yes | yes | no | yes | no |
| 4/ | - | METEOSWISS | SS-WG | WG | WG | yes | no | yes | yes | yes |
| 48 | - | IAP-CAS | MADEL TAD | WG | WG | yes | no | yes | yes | yes |
| 49 | - | IAP-CAS | MARFI-IAD | WG | WG | yes | no | yes | yes | yes |
| 50 | - | IAP-CAS | MAKFI-M3 | WG | WG | yes | no | yes | yes | yes |
| 51 | - | IAP-CAS | GOMEZ-BASIC | WG | WG | yes | no | yes | yes | yes |
| 52 | - | IAP-CAS | GOMEZ-IAD | WG | WG | yes | no | yes | yes | yes |

Table 5: Details about the predictors, geographical domains and preprocessing transformations used in the different MOS and PP statistical downscaling methods (note that distributional MOS and WG methods are not included since they use precipitation/temperatures at the closest gridbox, or use no predictor, respectively). The first column refers to the codes given in Tables 3 and 4. The last two columns indicate the transformations applied to the predictors (standardization, anomalies over the annual cycle, EOF/PC computation) and the size of the domain used: 'cont' for a single continental domain, 'nat' for multiple nation-wide domains, and 'gb'for information from the closest gridbox (or four gridboxes, '4 gb'). *Two-step* methods are indicated by including a '>' symbol between the two predictor/domain configurations used.

| CODE | PREDICTORS | TRANSFORM | DOMAIN |
|-------------|---|--------------------|----------|
| MOS-GLM | Precip./Temp. | standardized | 4 gb |
| MOS-REG | Precip./Temp. | standardized | 4 gb |
| MOS-AN | Precip. | raw data | nat |
| VGLMGAMMA | Precip. | standardized | gb |
| FIC01P | Z1000+500 | standardized | nat |
| FIC03P | U+V10, U+V500, R850+700 > R850, Q700 | standardized | nat > gb |
| FIC01T | Z1000-500 > TH1000-850 + 1000-500 | standardized | nat > gb |
| ANALOG-ANOM | SLP, TD, T2, U850, V850, Z850 | anomalies | cont |
| ANALOG | SLP, T2, T500+700+850, Q500+850, Z500 | PCs (95% variance) | nat |
| ANALOG-MP | Z1000+500 > VV600, T850 | raw data | nat > gb |
| ANALOG-SP | Z1000+500 > T2-TD, T2 | raw data | nat > gb |
| MO-GP | Standard set | raw data | gb |
| GLM-P | SLP, U+V10, T+Q+U+V850+700+500 | standardized | gb |
| GLM-DET | SLP, T2, T500+700+850, Q500+850, Z500 | 20 joined PCs | nat |
| GLM | SLP, T2, T500+700+850, Q500+850, Z500 | 20 joined PCs | nat |
| GLM-WT | T2, T500+700+850, Q500+850, Z500 (SLP for WT) | 15 joined PCs | nat |
| MLR-RAN | Z500, T850 | raw data | cont |
| MLR-RSN | Z500, T850 | raw data | cont |
| MLR-ASW | Z500, T850 | anomalies | cont |
| MLR- ASI | Z500, T850 | anomalies | cont |
| MLR-AAN | Z500, T850 | anomalies | cont |
| MLR-AAI | Z500, <u>T850</u> | anomalies | cont |
| MLR-AAW | Z500, T850 | anomalies | cont |
| MLR-PCA-ZTR | Z850, T850, R850 | s-mod PCs | cont |
| MLR-T | T2, SLP, U+V10, T+Q+U+V850-700-500 | standardized | gb |
| MLR | SLP, T2, T500+700+850, Q500+850, Z500 | 15 joined PCs | nat |
| MLR-WT | SLP, T2, T500+700+850, Q500+850, Z500 | 15 joined PCs | nat |
| ESD-EOFSLP | SLP | 20 PCs | cont |
| ESD-EOFT2 | T2 | 20 PCs | cont |
| ESD-SLP | SLP | raw data | cont |
| ESD-T2 | <u>T2</u> | raw data | cont |
| WT-WG | SLP | 15 PCs | nat |
| SWG | SLP, TD, T2, U850+V850+Z850 | 2 PCs each | cont |



Figure 1: Location of the 86 stations used in the paper, sorted according to latitude (see Table 1). Colors represent the orography (for the EURO-CORDEX 0.11° resolution grid, in meters). The colored boxes (and circles) show the eight PRUDENCE sub-regions (and the corresponding stations); the legend at the bottom of the figure indicates the names of the different regions.



Figure 2: Spearman correlation of downscaled and observed daily precipitation for winter (DJF, top) and summer (JJA, bottom). For each method, the box-whisker-plot summarizes the results of the 86 stations. Boxes span the 25-75% range and the whiskers the minimum/maximum value (within 1.5 times the interquartile range); outliers are plotted individually. Average results over the different Prudence regions are indicated by a colored horizontal bar for each method (see the colors in the bottom legend). Shading indicates the MOS results using RACMO2 predictors (all others use ERA-Interim). The methods are sorted as in Table 3.



Figure 3: As Figure 2 but for Pearson correlation of downscaled and observed daily maximum temperature.



Figure 4: Observed climatological R01 (relative wet-day frequency) values for winter (DJF, top left) and summer (JJA, top right). The biases of the downscaling methods (Table 3) are shown in the middle and bottom panels, for winter and summer, respectively. For each method, the box-whisker-plot summarizes the results of the 86 stations. Boxes span the 25-75% range and the whiskers the maximum value (within 1.5 times the interquartile range); outliers are plotted individually. A red asterisk indicates that values lie outside the plotted range. Average results over the different Prudence regions are indicated for each method (see the colors in the bottom legend). Shades indicate the MOS results using RACMO2 predictors (all others use ERA-Interim). The methods are sorted as in Table 3 (first the raw model outputs, followed by MOS, PP and WG methods).



Figure 5: As figure 4, but for SDII (mean wet-day precipitation).



Figure 6: Individual station results (sorted as in Table 1) for total precipitation (PRCTOT) biases for winter (DJF, top) and summer (JJA, bottom). Vertical dashed lines separate the different approaches and techniques (see Table 3).



Figure 7: As figure 4, but for the standard deviation of daily precipitation for winter (DJF, top left) and summer (JJA, top right). Relative standard deviation biases (predicted over observed deviations) are shown in the middle and bottom panels, for winter and summer, respectively.



Figure 8: Observed mean climatologies (deg C) of daily maximum temperature for winter (DJF, top left) and summer (JJA, top right). The biases of the downscaling methods (Table 4) are shown in the middle and bottom panels, for winter and summer, respectively. For each method, the box-whisker-plot summarizes the results of the 86 stations. Boxes span the 25-75% range and the whiskers the maximum value (within 1.5 times the interquartile range); outliers are plotted individually. A red asterisk indicates that values lie outside the plotted range. Average results over the different Prudence regions are indicated for each method (see the colors in the bottom legend). Shades indicate the MOS results using RACMO2 predictors (all others use ERA-Interim).



Figure 9: As Figure 8, but for daily minimum temperature.



Figure 10: Individual station results (sorted as in Table 1) for daily maximum temperature biases for winter (DJF, top) and summer (JJA, bottom). Vertical dashed lines separate the different approaches and techniques (see Table 4).



Figure 11: As Figure 8, but for the standard deviation of daily minimum temperature for winter (DJF, top left) and maximum temperature for summer (JJA, top right). Relative standard deviation biases (predicted over observed values) are shown in the middle and bottom panels, for winter and summer, respectively.