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Incretin-based therapies and risk of pancreatic cancer in patients with type 2 diabetes: a meta-analysis of randomised controlled trials

Short title: Incretin-based therapies and pancreatic cancer risk

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1 **ABSTRACT**

2 **Aims:** Conflicting evidence exists regarding the potential risk of pancreatic cancer with use of incretin drugs in
3 patients with type 2 diabetes (T2DM). We performed a meta-analysis of randomised controlled trials (RCTs),
4 including six recently published large-scale cardiovascular outcome trials (CVOTs), to evaluate the risk of
5 pancreatic cancer with incretin-based therapies in patients with T2DM.

6 **Materials and methods:** The PubMed, Embase, Cochrane Central Register and ClininalTrials.gov databases were
7 searched for RCTs in T2DM that compared incretin drugs with placebo or other antidiabetic drugs, with treatment
8 and follow-up durations of no less than 52 weeks, from January 1, 2007 to May 1, 2017. Two reviewers screened
9 the studies, extracted the data and assessed the risk of bias independently and in duplicate.

10 **Results:** Thirty-three studies ($n=79,971$), including the six CVOTs, with 87 pancreatic cancer events were
11 identified. Overall, the pancreatic cancer risk was not increased in patients administered with incretin drugs
12 compared to controls (Peto OR 0.67 [95%CI 0.44 to 1.02]). In the six CVOTs, 79 pancreatic cancer events were
13 identified in 55,248 subjects. Pooled estimates of the six CVOTs displayed the identical tendency (Peto OR 0.65
14 [95%CI 0.42 to 1.01]). Notably, in the subgroup of patients who received treatment and follow-up for 104 weeks
15 or more, 84 pancreatic cancer events were identified in 59,919 subjects, and a lower risk of pancreatic cancer was
16 associated with incretin-based therapies (Peto OR 0.62 [95%CI 0.41 to 0.95]).

17 **Conclusions:** Treatment with incretin drugs is not associated with an increased risk of pancreatic cancer in patients
18 with T2DM. Instead, it might protect against the pancreatic malignancy in patients treated for 104 weeks or more.
19 The major limitations of this study are that pancreatic safety was not the primary outcome of these enrolled trials,
20 and the event number and follow-up time are limited.

21 **KEYWORDS:** incretins, GLP-1 analogue, dipeptidyl peptidase-4 inhibitors, type 2 diabetes, meta-analysis

22 **1 | INTRODUCTION**

23 Incretin-based therapies include glucagon-like peptide-1 (GLP-1) receptor agonists and dipeptidyl peptidase-
24 4 (DPP-4) inhibitors. They demonstrate an anti-hyperglycaemic effect in a glucose-dependent manner and are
25 beneficial for weight control. Recently, encouraging results from two cardiovascular outcome trials (CVOTs) have
26 shown that GLP-1 receptor agonist can reduce the risk of major adverse cardiovascular events in patients with type
27 2 diabetes (T2DM) who are at high cardiovascular risk.^{1,2} Because of these favourable features, incretin drugs have
28 been recommended as important therapeutic options for patients with T2DM.³

29 However, concerns have been raised for years about the pancreatic safety of incretin drugs. In 2011, Elashoff
30 and colleagues⁴ reported that increased risks of both pancreatitis and pancreatic cancer were associated with the
31 use of incretin drugs. Thereafter, attempts have been made to investigate the safety of incretin-based therapies.
32 However, results from the preclinical reports⁵⁻⁷ and observational cohort studies⁸⁻¹¹ are conflicting. Notably, the
33 incidence of pancreatic cancer is low¹². No individual trial has enough power to assess the risk of pancreatic
34 malignancy sufficiently. Therefore, pooling data from large randomised controlled trials (RCTs) would be an
35 alternative method of investigating this safety issue.

36 Recently, CVOTs of incretin drugs (e.g. EXAMINE,¹³ SAVOR,¹⁴ TECOS,¹⁵ ELIXA,¹⁶ LEADER¹ and
37 SUSTAIN-6²) have been completed or are ongoing. In these trials, a large number of patients were followed up
38 for relatively longer periods and managed with a similar glycosylated haemoglobin (HbA1c) achievement goal.
39 Accordingly, pooling data from these CVOTs might help researchers better understand the true risk of pancreatic
40 malignancy with incretin-based therapies.

41 Here, we performed a meta-analysis of large RCTs, including the six recently published CVOTs, to evaluate
42 the risk of pancreatic cancer with incretin-based therapies in patients with T2DM.

43

44 **2 | MATERIALS AND METHODS**

45 **2.1 | Data sources and searches**

46 The PubMed, Embase and Cochrane Central Register databases were searched from January 1, 2007 to May
47 1, 2017 for RCTs that involved incretin drugs and were published in English. Medical subject headings and free
48 text terms were used to identify the related articles. An endocrinologist, together with an epidemiologist, developed
49 the search strategy (S1 Text, Supporting information). The ClinicalTrials.gov was searched using the same method
50 to identify trials that were complete but unpublished. It also provided us with extended information about adverse
51 events related with the selected trials. The search was performed on May 1, 2017.

52

53 **2.2 | Study selection**

54 We selected trials that satisfied the following criteria: **1)** study type, RCTs; **2)** subjects, adult patients with
55 T2DM; **3)** intervention and comparators, trials that compared the effects of incretin drugs (GLP-1 receptor agonists
56 or DPP-4 inhibitors) with comparators (placebo or other antidiabetic drugs); **4)** duration of treatment and follow-
57 up, it is estimated that the mean time from first malignant cell to the clinical diagnosis of pancreatic cancer is 0.7
58 years in males and 0.6 years in females,¹⁷ therefore, we only included trials that had a treatment and follow-up
59 time at least 52 weeks to reduce the bias related to the undiscovered pancreatic cancer before start of intervention;
60 **5)** sample size restriction, pancreatic cancer has a low incidence (10-14 per 100,000 person years),¹² to reduce
61 sampling variation, we only included the trials with at least 500 randomised subjects; and **6)** outcome evaluation,
62 the trials were required to have clear information of pancreatic malignancy, or at least systemic reports of
63 neoplasms in the supplemental materials or in the data posted on ClinicalTrials.gov.

64

65 **2.3 | Data extraction and quality assessment**

66 Data were collected from published papers or from ClinicalTrials.gov documents (for unpublished trials).
67 Two trained reviewers screened the literature for eligible studies. A pilot format was used for the reviewers to
68 evaluate the risk of bias and to collect data independently and in duplicate for each included trial. Disputes were
69 discussed by the study group and were adjudicated by the study supervisor. For multiple reports of one trial, we
70 only documented the data from the report with the longest follow-up. For each eligible trial included in this study,
71 the characteristics of the trials, including National Clinical Trial (NCT) codes (if available), sample size, the
72 number of participants in each treatment group, duration of treatment and follow-up, percentage of male
73 participants, age and body mass index (BMI) of the participants, duration of diabetes and baseline HbA1c level
74 were recorded. As glycaemic control status may affect the risk of cancer,¹⁸ we also recorded the final HbA1c
75 difference between the groups. An HbA1c difference of more than 0.4% was regarded as clinically significant.¹⁹
76 The pancreatic cancer events in each group were recorded separately. The number of patients exposed to each
77 treatment group was recorded using intention-to-treat (ITT) data.

78 A modified Cochrane Collaboration's tool, which includes information about the randomisation process,
79 allocation concealment, blindness, adjudication of outcomes and selective reporting, was used to assess the risk of
80 bias in each trial.²⁰ We used funnel plot asymmetry to detect whether there was publication bias and Egger's
81 regression test to measure funnel plot asymmetry in Stata 11.0.²¹

82

83 **2.4 | Data synthesis and analysis**

84 We assessed heterogeneity between studies using both Chi² and I² statistics. Pancreatic cancer is rare, and the
85 Peto method is recommended and has a relatively good reputation for rare events,²⁰ therefore, pooled risk was
86 reported with the Peto odds ratio (OR) and 95% confidence interval (CI); $p < 0.05$ was considered to be significant.
87 For dichotomous outcomes, the weight for each trial was calculated based on the size of the trial and the number

88 of events.²⁰ To determine the possible factors that might affect the risk of pancreatic cancer, we performed four
89 prespecified subgroup analyses, according to the following stratifications: the duration of treatment and follow-up
90 (52 to 103 weeks or no less than 104 weeks), class of incretin drugs (GLP-1 receptor agonists or DPP-4 inhibitors),
91 type of comparators (placebo or other non-incretin antidiabetic drugs), and level of HbA1c difference between
92 treatment arms at the end of trials (more than 0.4% or not). In addition, a sensitivity analysis was performed using
93 alternative effect measures (OR vs. relative risk), pooling methods (Peto method vs. Mantel-Haenszel method),
94 and consideration of heterogeneity (fixed effect vs. random effect). We reported our results, according to the
95 PRISMA statement.²²

96

97 **3 | RESULTS**

98 Among the 5,416 potential reports from PubMed, Embase and Cochrane Central Register, and 305 reports
99 from ClinicalTrial.gov, we identified 622 reports for full-text reviews. Finally, thirty-three RCTs,^{1,2,13-16,23-49}
100 including the six recently reported CVOTs,^{1,2,13-16} fulfilled the inclusion criteria (32 from published journals and 1
101 unpublished trial from ClinicalTrials.gov). A flow diagram of the trial selection is presented in Figure 1. For all
102 included studies, the average age of the participants ranged from 51.8 to 72.6 years, and the mean BMI ranged
103 from 24.9 to 37.1 kg/m², with a mean duration of diabetes ranging from 1.0 to 13.9 years. The percentage of male
104 subjects ranged from 43% to 71%. The average baseline HbA1c level ranged from 7.2% to 9.2%. The mean or
105 median follow-up time ranged from 52 to 198 weeks (Table 1).

106 **3.1 | Quality of the included trials and publication bias**

107 Among the included trials in our analysis, randomisation was well designed in 31 studies. One trial did not
108 mention how the random sequence was generated. One trial was at high risk of bias because its randomisation was
109 stratified by different baseline treatments. For allocation concealment and blinding of the treatment, six trials

110 without treatment concealment to the investigators and participants were regarded as having a high risk of bias.

111 As for the outcome evaluation, all included studies provided safety data for the ITT population, and the six CVOTs

112 and nine non-CVOT studies had an independent adjudication committee for the cancer and pancreatitis events,

113 which were at low risk of bias (Table S1, Supporting information).

114 The funnel plot was symmetric (Egger's test $P=0.887$) (Figure S1, Supporting information). Moreover, all

115 included trials were designed to test the drug's efficacy of glucose lowering or the safety of cardiovascular

116 outcomes. Therefore, pancreatic cancer and pancreatitis events had a minor effect on the selected publications.

117

118 **3.2 | Risk of pancreatic cancer in the pooled analysis**

119 Of the 33 included RCTs, eleven studies reported pancreatic cancer events. Thirty-five events were reported

120 in 42,233 incretin group subjects, and 52 events were reported in 37,738 control group subjects. Pooled estimates

121 of the 33 trials ($n=79,971$) showed that no increased risk of pancreatic cancer was associated with the incretin

122 drugs compared to the controls (Peto OR 0.67 [95%CI 0.44 to 1.02]). In particular, in the six CVOTs, thirty-three

123 events were reported in 27,663 subjects in the incretin group, and 48 events were reported in 27,585 control group

124 subjects. The Peto OR of the pooled analysis of the six CVOTs ($n=55,248$) was 0.65 [95%CI 0.42 to 1.01] (Figure

125 2).

126

127 **3.3 | Risk of pancreatic cancer in the subgroup analysis**

128 When evaluating the effect of incretin drugs on the risk of pancreatic cancer, the exposure time is an important

129 factor. Among the trials that followed subjects for 52-103 weeks, three pancreatic cancer events in 11,765 incretin

130 users and no events in 8,287 non-incretin users were observed (Peto OR 5.63 [95%CI 0.52 to 60.4]) (Figure S2,

131 Supporting information). For the subjects who received incretin treatment and were followed for 104 weeks or

132 more, the risk of pancreatic cancer decreased significantly, compared to subjects who received the control
133 treatment (Peto OR 0.62 [95%CI 0.41 to 0.95]) (Figure 3).

134 No significant difference in the risk of pancreatic cancer was observed in the subgroup analysis of the class
135 of incretin drugs (Peto OR 0.77 [95%CI 0.42 to 1.42] for GLP-1 receptor agonists and 0.59 [95%CI 0.33 to 1.05]
136 for DPP4 inhibitors). In the subgroup analysis of the type of comparators, we did not find any significant difference
137 in the risk of pancreatic cancer in the subgroups of incretins vs. active controls (Peto OR 1.12 [95%CI 0.25 to
138 5.06]), such as metformin, thiazolidinediones, sulfonylureas, insulin and sodium-dependent glucose transporters 2
139 inhibitors. A decreased risk of pancreatic cancer was observed in the incretin group compared to the placebo group
140 (Peto OR 0.63 [95%CI 0.40 to 0.97]). In addition, there was no significant difference in the risk of pancreatic
141 cancer between the incretin-based and control therapies in the subgroup stratified by the level of the final HbA1c
142 difference (Peto OR 0.70 [95%CI 0.45 to 1.09] and 0.45 [95%CI 0.12 to 1.70] for the subgroups with HbA1c
143 differences $\leq 0.4\%$ and $>0.4\%$, respectively) (FigureS3-5, Supporting information).

144

145 **3.4 | Risk of pancreatitis in the pooled analysis**

146 The overall pancreatitis risk was not increased in the incretin group compared with the control group (Peto
147 OR 1.12 [95%CI 0.85 to 1.47]). Pooled analysis of the six CVOTs did not show an increased risk of pancreatitis
148 associated with incretin-based therapies (Peto OR 1.06 [95%CI 0.80 to 1.42]) (Figure S6, Supporting information).

149

150 **3.5 | Sensitivity analysis**

151 The sensitivity analysis of pancreatic cancer risk using an alternative pooling method (Mantel-Haenszel OR
152 0.67 [95%CI 0.44 to 1.02]), effect measure (relative risk 0.67 [95%CI 0.44 to 1.02]), and consideration of

153 heterogeneity (random effects OR 0.68 [95%CI 0.44 to 1.06]) did not show any important change in the pooled
154 effects.

155

156 **4 | DISCUSSION**

157 Overall, we screened 5,721 studies and included 33 eligible RCTs reporting 87 pancreatic cancer events
158 among 79,971 patients. We found that compared with the controls, treatment with incretin drugs was not associated
159 with an increased risk of pancreatic cancer in patients with T2DM. Instead, use of incretin drugs for 104 weeks or
160 more might even decrease the risk of pancreatic malignancy by 38% compared with controls.

161 The association between incretin-based therapies and pancreatic cancer has drawn a great concern recently.
162 Unfortunately, neither preclinical studies^{5-7,50} nor the following cohort studies have answered this question
163 consistently.^{8-11,51,52} In a recent large multinational cohort study, the risk of pancreatic cancer even seemed to be
164 lower with longer incretin-based therapy durations (HR 1.53 [95%CI 0.93 to 2.51], 1.07 [95%CI 0.82 to 1.39] and
165 0.62 [95%CI 0.36 to 1.07] for duration of use <1 year, 1-1.9 years and ≥ 2 years, respectively), although the
166 difference was not statistically significant.⁸ Additionally, the United Kingdom clinical practice research datalink
167 (UK-CPRD) cohort study has reported that the minor increase of pancreatic cancer risk in new incretin users
168 [adjusted HR 1.67 (1.01–2.77)] was likely caused by protopathic bias because of the lack of a duration of use and
169 dosage effect for incretin agents on pancreatic cancer risk.¹¹ Inconsistency and methodological limitations
170 undermined the strength of those results. In the cohort studies, baseline characteristics and metabolic control levels
171 could not be well matched between groups. Even in the nested case-control study, the incretin group still differed
172 from the control group in the parameters that could affect the incidence of malignancy, including age, duration of
173 diabetes, BMI and HbA1c levels.⁸ Furthermore, in most cohort studies, the report of pancreatic cancer events was
174 based on medical or insurance records, which may have led to the inaccurate definition of the events.

175 Nevertheless, cohort studies are from the real world. It is rational to take the results from cohort studies and
176 meta-analyses of RCTs together into consideration when evaluating the risk of pancreatic malignancy. Recently,
177 two meta-analyses by Monami and colleagues have suggested that there is no increased risk of pancreatic cancer
178 associated with the use of incretins.^{53,54} In the 2014 report, the primary outcome was pancreatitis, and the data
179 collection was not based on pancreatic cancer.⁵³ In addition, the sample size of each trial varied greatly (from 24
180 to 9340). The follow-up durations of the enrolled trials were not long enough (more than 70% of the trials had
181 follow-up durations of 12-51 weeks),^{53,54} which did not take the latent period of cancer into consideration.

182 In our study, we enrolled qualified RCTs with baseline characteristics that were balanced between the groups.
183 The drug exposure and follow-up were clear and well managed. All patients treated with incretin drugs were new
184 users, thereby avoiding the bias in cohort studies caused by combining new users with prevalent users and the
185 possible protopathic bias¹¹. Note that pancreatic cancer is insidious and rare. The estimated time from first
186 malignant cell to the clinical diagnosis of pancreatic cancer is 0.6-0.7 year.¹⁷ Therefore, we only included trials
187 with treatment and follow-up durations of at least 52 weeks to reduce the possibility of occult pancreatic
188 malignancy at the start of trials, and we excluded studies with fewer than 500 subjects in case that cancer events
189 reported in small trials by chance could dramatically affect the incidence. Recently, several large-scale CVOTs of
190 incretin drugs have been completed or are ongoing. A large number of patients (from 3,927 to 16,492) were
191 enrolled, and the duration of follow-up was much longer (median duration ranged from 1.5 to 3 years) in these
192 trials. Moreover, the primary endpoint of these CVOTs was drug safety rather than efficacy, and patients in
193 different intervention groups were managed under a similar glycaemic goal. Therefore, the differences in the
194 HbA1c level achieved between the incretin and control groups were relatively small, providing a more parallel
195 metabolic status. Accordingly, pooling data from these CVOTs might help us to better understand the pancreatic
196 safety issue of long-term incretin-based therapies.⁵⁵⁻⁵⁷ Nauck and colleagues have remarked that CVOT studies

197 could provide us with valuable information about pancreatic safeness, and they provided good evidence against
198 previous estimates of the increased pancreatic cancer risk.⁵⁷ We collected all available RCTs with follow-up times
199 of at least 52 weeks, thus, our results could further support their conclusion.

200 It is known that patients with T2DM are at high risk of developing pancreatic cancer, with high mortality
201 rates.⁵⁸ Unfortunately, pancreatic cancer still lacks an effective management strategy and even presents with
202 increasing incidence and mortality.¹² Here, our subgroup analysis showed that treatment with incretin drugs for 2
203 years or more significantly reduced the risk of pancreatic cancer, compared to the controls, by 38%. Notably, a
204 similar trend was also observed in the large nested case-control cohort study mentioned above (HR 0.62 [95%CI
205 0.36 to 1.07]).⁸ Signals from laboratory studies have also suggested an anti-tumour effect of incretin drugs. In our
206 previous studies, we found that GLP-1 receptor levels were lower in cancer tissues than in tumour adjacent
207 pancreatic tissues, and a lower GLP-1 receptor level was associated with poorer prognoses in patients with
208 pancreatic cancer. Moreover, GLP-1 receptor activation with liraglutide inhibited growth and promoted apoptosis
209 of human pancreatic cancer cells in a GLP-1 receptor-dependent manner *in vitro*, and attenuated pancreatic tumour
210 growth in a mouse xenograft model *in vivo*.^{59,60} In agreement with our previous findings, it has been shown that
211 GLP-1 receptor agonist exendin-4 can inhibit cell growth in colon cancer cells⁶¹ and breast cancer cells^{62,63} *in vitro*
212 and *in vivo*. Furthermore, exendin-4 also counteracts the invasive potential of human neuroblastoma cells.⁶⁴ These
213 consistent signals from clinical and laboratory studies suggest that long-term incretin-based therapies might shed
214 some light on how to prevent the development and progression of pancreatic malignancy, which usually has a poor
215 outcome.

216 GLP-1 receptor agonists and DPP-4 inhibitors are both incretin drugs, however, they act differently. GLP-1
217 receptor agonists directly and intensively stimulate GLP-1 receptor and its downstream signalling pathways, while
218 DPP-4 inhibitors can increase the levels of endogenous incretin hormones by inhibiting DPP-4-mediated incretin

219 degradation. Hence, it has been suggested to analyse them separately.⁵⁶ Here, we found no difference between
220 GLP-1 receptor agonists and DPP-4 inhibitors, in terms of their association with pancreatic cancer risk. It is well
221 known that diabetes itself is an independent risk factor for pancreatic cancer and its high mortality.^{12,58} Although
222 patients had parallel baseline HbA1c levels when they entered a trial, they received different antidiabetic therapies
223 and might significantly vary in their glycaemic control. Therefore, we conducted a prespecified subgroup analysis
224 based on a 0.4% difference in the final HbA1c level, which is usually considered to be the non-inferiority margin,¹⁹
225 to clarify whether our result would be altered by glycaemic control variations. Again, we could not find any
226 increased risk of pancreatic cancer associated with incretin drugs in the subgroups stratified by level of HbA1c
227 difference, suggesting the consistency of the observations that incretin drugs were not the promoter of pancreatic
228 cancer.

229 It has been argued that incretin drugs were a potential inducer of acute and chronic pancreatitis,⁶⁵⁻⁶⁷ thus in
230 the long run promoting the development of pre-neoplastic lesions and increasing the risk of pancreatic cancer.
231 However, increasing reports from clinical trials,⁶⁸ cohort studies^{69,70} and systemic reviews^{53,71} have shown no
232 increased risk of pancreatitis associated with incretin-based therapies. Here, we also found results similar to those
233 reports, providing additional evidence for the pancreatic safety of incretin-based therapies.

234 Several limitations should be considered in our study. First, pancreatic safety was not the primary outcome
235 of the included trials, and the number of pancreatic cancer events in our study was relatively smaller than that in
236 some observational cohort studies.^{8, 9,11} This limitation is primarily attributed to the nature of RCTs because it is
237 not practical for a RCT to enrol such a large population as the cohort studies performed using databases.
238 Nevertheless, signals from our study and most cohort studies consistently suggest that incretin drugs were not the
239 carcinogen of pancreatic malignancy. Second, we noticed that 91% of pancreatic cancer events were reported in
240 the six CVOTs, whereas the number of the events was small (8 cases in 24,723 subjects) which led to a wide range

241 of 95%CI in the non-CVOT studies. There might be underreporting in some of the non-CVOTs with no pancreatic
242 malignancies reported. However, the nature of RCTs could partially balance the possibility of underreporting.
243 Moreover, the total number of pancreatic cancer events in our pooling data was 87 in 79,971 subjects
244 (approximately 108 per 100,000 persons). The estimated overall incidence was approximately 47 per 100,000
245 person years in all trials and 55 per 100,000 person years in trials with follow-ups of 104 weeks or more, which
246 was much higher than that in the general population (10-14 per 100,000 person years)¹² and similar to that reported
247 in a previous large cohort study in patients with T2DM (60 per 100,000 person years).⁸ Therefore, there might be
248 no obvious underreporting in the trials included in our study, and the limited number of the cancer events may not
249 substantially undermine our results. Third, we did not have the primary time-to-event data for all included trials,
250 and it is possible that a risk of immortal time bias might exist in our subgroup analysis. However, in the nested
251 case-control cohort study, the declined tendency of pancreatic cancer risk was also found among the subjects
252 treated with incretin drugs for 2 years or more.⁸ Furthermore, incretins are not carcinogens, and they may influence
253 the rate of neoplasm progression and affect the time period from the first malignant cell to the clinical diagnosis
254 of pancreatic cancer. Theoretically, the longer the exposure, the more significant effect could be found. Therefore,
255 pooling data from more large-scale trials with long-term incretin-based therapies, particularly the CVOTs with the
256 primary time-to-event data, might provide us with a clearer picture on this topic. Fourth, we did not identify
257 pancreatitis as acute or chronic, partly because some trials did not define or report pancreatitis in detail. After all,
258 it was not the primary outcome in this analysis.

259

260 In conclusion, our meta-analysis of 33 RCTs involving 79,971 subjects suggests that treatment with incretin
261 drugs for no less than 52 weeks is not associated with an increased risk of pancreatic cancer in patients with T2DM.
262 Instead, treatment with incretin drugs might protect against the risk of pancreatic malignancy, particularly in

263 patients with T2DM who received the treatment for 104 weeks or more. Even so, it is difficult to verify this issue
264 in a single study because pancreatic cancer is rare and occult. Accordingly, pooling more data from large-scale
265 RCTs, particularly long-term CVOTs, may help us to find the true answer to the question of whether incretin-
266 based therapies are safe and might even protect patients with T2DM against pancreatic malignancy.

267

268 **Figure and table legends.**

269 **Figure 1. Flow diagram of trial selection**

270

271 **Table 1. Characteristics of randomised controlled trials of incretin-based therapies and pancreatic cancer**
272 **events in patients with type 2 diabetes**

273 NCT, National Clinical Trial. NR, not reported. * Final HbA1c (%) difference: incretin group vs. control group at
274 the end of the trial. † There was no publication of the studies, and presented here is the time of the last data update
275 on ClinicalTrial.gov website. § Sitagliptin/metformin fixed-dose combination. ‡ No report of pancreatic cancer in
276 the article, but there were systemic reports of neoplasms in the supplemental materials or in the data posted on the
277 ClinicalTrial.gov website.

278

279

280 **Figure 2. Risk of pancreatic cancer in patients with type 2 diabetes who were treated with incretin drugs or**
281 **controls.**

282 CVOT, cardiovascular outcome trial.

283

284 **Figure 3. Risk of pancreatic cancer in patients with type 2 diabetes who were treated with incretin drugs or**
285 **controls and followed up for 104 weeks or more.**

286

287

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290

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293 any other potential conflicts of interest relevant to this article.

294

295 **Author Contributions:**

296 T.H. was involved in designing the meta-analysis and data collection and analysis, as well as the writing, editing
297 and revising of the manuscript. H.W and Y.L were involved in the data collection and analysis, as well as writing
298 the manuscript. Q.T, J.Y, R.L were involved in the data collection, quality evaluation of the trials and manuscript
299 editing. Z.S and J.H were involved in the study methodology and data analysis, as well as reviewing and editing
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306

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