

# Identifiability of Finite Mixture Models

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SMTDA 2020

June 4, 2020

# Outline

## 1 Introduction

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- 2 Characterization of Identifiability

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# Identifiability

# Identifiability

## Definition:

A finite mixture model is identifiable if a given dataset leads to a uniquely determined set of model parameter estimations up to a permutation of the clusters.

Identifiability of the parameters is a necessary condition for the existence of consistent estimators for any statistical model.

Without identifiability, there might be several solutions for the parameter estimation problem.

# Literature Review

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- Yakowitz and Spragins (1968): Extension to the class of all Gaussian mixtures.
- Henning (2000): Identifiability for linear regression mixtures with Gaussian errors under certain conditions.

# Finite Mixture Models

## Data:

Variable of interest  $Y_i = y_{i_1}, y_{i_2}, \dots, y_{i_T}$

Covariants  $x_1, \dots, x_M$  and  $z_{i_1}, \dots, z_{i_T}$

$a_{i_t}$  age of subject  $i$  at time  $t$

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$a_{i_t}$  age of subject  $i$  at time  $t$

## Model:

$K$  groups of size  $\pi_k$  with trajectories

$$y_{it} = \sum_{j=0}^{s_k} \left( \beta_j^k + \sum_{m=1}^M \alpha_m^k x_m + \gamma_j^k z_{i_t} \right) a_{it}^j + \varepsilon_{it}^k, \quad (1)$$

where  $\varepsilon_{it}^k \sim \mathcal{N}(0, \sigma^k)$ .

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# Notations

Distribution  $f$  of a finite mixture model:

$$f(y_i; \Omega) = \sum_{k=1}^K \pi_k g_k(y_i; \theta^k).$$

Cumulative distribution function  $F$  of a finite mixture model:

$$F(y_i; \Omega) = \sum_{k=1}^K \pi_k G_k(y_i; \theta^k).$$

## Mixtures and mixing distributions

Let  $\mathcal{F} = \{F(y; \omega), y \in \mathbb{R}^T, \omega \in \mathbb{R}_K^{s+2}\}$  be a family of  $T$ -dimensional cdf's indexed by a parameter set  $\omega$ , such that  $F(y; \omega)$  is measurable in  $\mathbb{R}^T \times \mathbb{R}_K^{s+2}$ .

The the  $s + 2$ -dimensional cdf  $H(x) = \int_{\mathbb{R}_K^{s+2}} F(y; \omega) dG(\omega)$  is the image of the above mapping, of the  $s + 2$ -dimensional cdf  $G$ .

The distribution  $H$  is called the mixture of  $\mathcal{F}$  and  $G$  its mixing distribution.

Let  $\mathcal{G}$  denote the class of all  $s + 2$ -dimensional cdf's  $G$  and  $\mathcal{H}$  the induced class of mixtures  $H$ .

Then  $\mathcal{H}$  is identifiable if  $Q$  is a one-to-one map from  $\mathcal{G}$  onto  $\mathcal{H}$ .

# Characterization of identifiability

The set  $\mathcal{H}$  of all finite mixtures of class  $\mathcal{F}$  of distributions is the convex hull of  $\mathcal{F}$ .

$$\mathcal{H} = \left\{ H(y) : H(y) = \sum_i c_i F(y, \omega_i), c_i > 0, \sum_i c_i = 1, F(y, \omega_i) \in \mathcal{F} \right\}. \quad (2)$$

## Theorem

*A necessary and sufficient condition for the class  $\mathcal{H}$  of all finite mixtures of the family  $\mathcal{F}$  to be identifiable is that  $\mathcal{F}$  is a linearly independent family over the field of real numbers.*



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# The Model

$$Y_{it} = f(\mathbf{a}_{it}; \beta^k, \delta^k) + \varepsilon_{it}^k = \beta^k A_{it} + \delta^k W_{it} + \varepsilon_{it}^k. \quad (3)$$

We can write

$$\mathcal{L}((Y_i)_{i \in I}) = \bigotimes_{i \in I} F_{A_i, W_i, J}. \quad (4)$$

Identifiability of a model means that knowing the data distribution  $\mathcal{L}(Y_i), i \in I$ , one can uniquely identify the mixing distribution  $J$ .

That is, no two distinct sets of parameters lead to the same data distribution.

# Nagin's base model

$$\mathcal{C}_1 = \left( F_{A,J} : F_{A,J} = \bigotimes_{i \in I} F_{A_i, J} \right)_{J \in \Omega_1}$$

## Theorem

Let  $h_j = \min \{ q : \{A_{ij}, i \in I\} \subseteq \cup_{i=1}^q H_i \mid H_i \in \mathcal{H}_{n-1} \}$ .

If there exist  $j$  such that  $|S(J)| < h_j, \forall J$  then  $\mathcal{C}_1$  is identifiable.

## Adding covariates independent of cluster membership

$$\mathcal{C}_2 = \left( F_{A,J} : F_{A,J} = \bigotimes_{i \in I} F_{A_i, W_{i,J}} \right)_{J \in \Omega_1}, \quad (5)$$

$$\mathcal{C}_{2A} = \left( F_{A,J} : F_{A,J} = \bigotimes_{i \in I} F_{A_i, J} \right)_{J \in \Omega_1}, \quad (6)$$

$$\mathcal{C}_{2W} = \left( F_{A,J} : F_{A,J} = \bigotimes_{i \in I} F_{W_{i,J}} \right)_{J \in \Omega_1}. \quad (7)$$

### Theorem

*If  $\mathcal{C}_{2A}$  and  $\mathcal{C}_{2W}$  are identifiable and  $W_{ij}$  is not a multiple of  $A_{ij}$ , for all  $i, j$ , then  $\mathcal{C}_2$  is identifiable.*

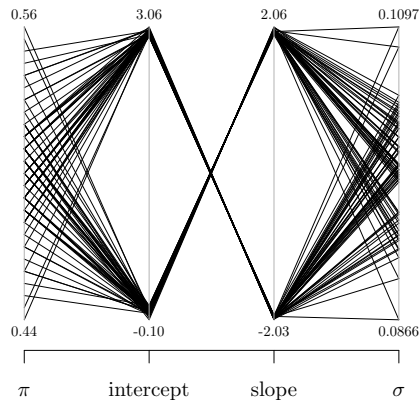
# Numerical Example

- Two clusters with sizes  $\pi_1 = \pi_2 = \frac{1}{2}$ .
- Two time-points 1 and 2.
- Same variability in both clusters  $\sigma = 0.1$

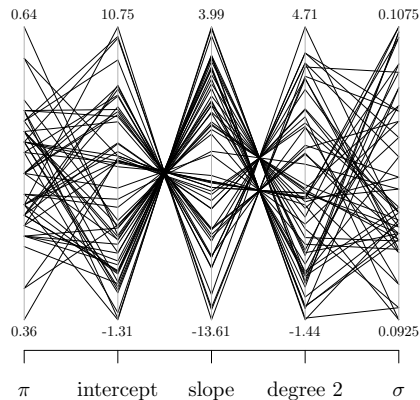
We simulate 50 samples of 100 trajectories with parameters

- $\beta^1 = (3, -2)$  and  $\beta^2 = (0, 2)$  (linear model)
- $\beta^1 = (10, -12.5, 3.5)$  and  $\beta^2 = (-2, 5, -1)$  (polynomial model).

# Parallel coordinate plots of the estimated parameter



Linear Model



Parabolic Model

# The generalized model

## Theorem

*The model is identifiable if*

- $d_k < T$  for all  $1 \leq k \leq K$  and all  $a_{it}$  are distinct, for all  $i_t$ .
- $W_k$  has full rank for all  $1 \leq k \leq K$ .
- $rk(A_k, W_k) = rk(A_k) + rk(W_k)$ , for all  $1 \leq k \leq K$ .

## Numerical Example

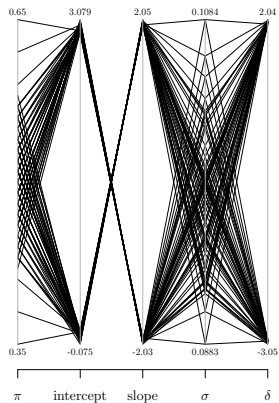
- Two clusters with sizes  $\pi_1 = \pi_2 = \frac{1}{2}$ .
- Two time-points 1 and 2.
- Same variability in both clusters  $\sigma = 0.1$
- Shape description parameters  $\beta_1 = (3, -2)$ ,  $\beta_2 = (0, 2)$ ,  $\delta_1 = 2$  and  $\delta_2 = -3$ .

We simulate 50 samples of 100 trajectories for 3 types of models:

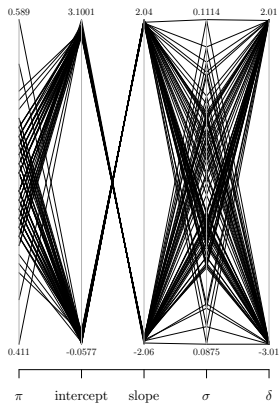
- The covariate is independent of time and only takes values 0 or 1
- The covariate is time dependent but in a nonlinear way
- The covariate is time dependent in a linear way



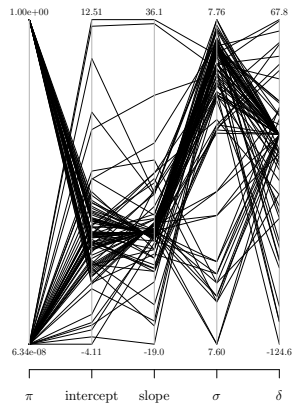
# Parallel coordinate plots of the estimated parameter



Model 1



Model 2



Model 3

# Bibliography

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