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Author(s)	Kataoka, Yuki; Yamamoto, Yosuke; Otsuki, Taiichiro; Shinomiya, Mariko; Terada, Takayuki; Fukuma, Shingo; Yamazaki, Shin; Hirabayashi, Masataka; Nakano, Takashi; Fukuhara, Shunichi
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A new prognostic index for overall survival in malignant pleural mesothelioma: the 1 rPHS (regimen, PS, Histology, or Stage) index $\mathbf{2}$ 3 Yuki Kataoka^{1,3}, Yosuke Yamamoto¹, Taiichiro Otsuki², Mariko Shinomiya³, Takayuki 4 Terada⁴, Shingo Fukuma^{1,5}, Shin Yamazaki¹, Masataka Hirabayashi³, Takashi Nakano⁴, $\mathbf{5}$ Shunichi Fukuhara^{1,5,*} 6 $\overline{7}$ 1 Department of Healthcare Epidemiology, Graduate School of Medicine and Public 8 9 Health, Kyoto University, Yoshida Konoe-cho, Sakyo-ku, Kyoto, 2 Cancer center, Hyogo College of Medicine, Mukogawa-cho, Nishinomiya, Hyogo, 3 Department of 10 Respiratory Medicine, Hyogo Prefectural Amagasaki Hospital, Higashi-Daimotsu-Cho, 11 Amagasaki, Hyogo 4 Division of Respiratory Medicine, Hyogo College of Medicine, 12Hyogo, Japan, Hyogo College of Medicine, Mukogawa-cho, Nishinomiya, Hyogo, and 135 Center for Innovative Research in Clinical Evaluative Science (CiRCLE), Fukushima 14Medical University, Hikarigaoka, Fukushima 151617*For reprints and all correspondence: Shunichi Fukuhara, MD, DMSc, Department of Healthcare Epidemiology, Graduate 18

- 19 School of Medicine and Public Health, Kyoto University, Yoshida Konoe-cho, Sakyo-ku,
- 20 Kyoto 606-8501, Japan.
- 21 Tel: +81-75-753-4646 Fax: +81-75-753-4644
- 22 Email: <u>fukuhara.shunichi.6m@kyoto-u.ac.jp</u>
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- 24 Running head:
- 25 A prognostic index in malignant mesothelioma
- 26

28 ABSTRACT

29 Background

Existing prognostic indices (PI) for malignant pleural mesothelioma (MPM) do not incorporate the recent advances in oncology care. The purpose of this study was to provide a PI for overall survival (OS) in MPM patients treated with chemotherapy with pemetrexed (PEM) or best supportive care (BSC) in the recent clinical setting.

34 Methods

A retrospective cohort study was performed in two hospitals in Japan (2007 - 2013).

36 The primary outcomes were OS. The Cox proportional hazards model was used for

37 multivariable analyses to identify prognostic factors. A final model was chosen based on

38 both clinical and statistical significance.

39 Results

A total of 283 patients (CTx: n=228, BSC: n=55) were enrolled in the study. On
multivariate analysis, regimen including platinum plus PEM, a performance status > 0,
non-epithelial histological type, and stage IV disease predicted poor OS in CTx patients.
As hazard ratios of individual risk factors were approximately similar, a prognostic
index for OS was constructed by counting the risk factors. Median OS in CTx patients
decreased by each 1-point increase in this count: 1030 days for zero; 658 days for one;

46	373 days for two; 327 days for three; 125 days for four. Internal validation using the
47	bootstrapping technique showed robustness of the model (c-index, 0.677; 95%
48	Confidence Interval [CI], 0.624-0.729). Further, the discrimination was consistent in
49	BSC patients (c-index, 0.799; 95% CI, 0.725-0.874).
50	Conclusions
51	This novel index can provide clinicians and MPM patients with a better framework for
52	discussing prognosis at the time of diagnosis.
53	
54	A mini-abstract:
54 55	A mini-abstract: We developed a new prognostic index for malignant pleural mesothelioma. The index
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63 INTRODUCTION

64	Malignant pleural mesothelioma (MPM) used to be a rare malignancy of the
65	mesothelium. In recent years, the incidence of this disease has increased, and this trend
66	will likely continue worldwide over the next decade (1).
67	Despite recent advancements in treatment, surgery, radiotherapy and chemotherapy or
68	multimodality therapy has not be proven to be curative (2–4). For the majority of
69	patients, treatment options are limited to palliative chemotherapy and best supportive
70	care (BSC) (5).
71	In oncologic palliative care, early determinations of prognosis play an important role in
72	guiding end-of-life care and efforts designed to improve patients' quality of life (6, 7).
73	To determine the prognosis of patients with MPM, four prognostic indices (PI) have
74	been developed; one by the Cancer and Leukaemia Group B (CALGB) (8), and three by
75	the European Organization for Research and Treatment of Cancer (EORTC) (9–11).
76	While the first two PIs from EORTC can indicate either a favorable or an unfavorable
77	outcome, neither can predict the duration of survival, which means both are impractical
78	when discussing life expectancy with a patient. The CALGB PI is complex to use,
79	because it has various cutoffs to consider. Above all, these PIs are based on clinical trial
80	data and may not be applicable to the clinical setting. Further, they do not incorporate

81	information regarding pemetrexed, which can improve overall survival (OS), and does
82	not incorporate recent advancements in supportive care (3, 12–14). Therefore, while
83	existing PIs might be useful for researchers in deciding which patients to include in
84	clinical trials, these systems are less useful for clinicians who need to discuss prognoses
85	with their MPM patients.
86	The purpose of this study was to provide a new PI for OS in MPM patients who
87	underwent treatment with pemetrexed or best supportive care in a recent clinical setting.
88	
89	Materials and methods
90	Study design and patients
90 91	Study design and patients A retrospective cohort study was performed, covering the period between April 1 st , 2007
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91 92 93	A retrospective cohort study was performed, covering the period between April 1 st , 2007 and March 31 st , 2013. The cohort was defined as all patients with histologically proven (15) MPM at either one of two tertiary hospitals that serve the South Hanshin medical
91 92 93 94	A retrospective cohort study was performed, covering the period between April 1 st , 2007 and March 31 st , 2013. The cohort was defined as all patients with histologically proven (15) MPM at either one of two tertiary hospitals that serve the South Hanshin medical region, which is an area of high MPM incidence area in Japan (16).
 91 92 93 94 95 	A retrospective cohort study was performed, covering the period between April 1 st , 2007 and March 31 st , 2013. The cohort was defined as all patients with histologically proven (15) MPM at either one of two tertiary hospitals that serve the South Hanshin medical region, which is an area of high MPM incidence area in Japan (16). Patients who had more than one cancer, underwent autopsy, or who received palliative

99	diagnosis, trimodal therapy, or surgical therapy extra-pleural pneumonectomy or
100	pleurectomy or decortication) were excluded to avoid confounding influences (17).
101	
102	Definitions of prognostic variables
103	Potential prognostic factors that were analyzed included: histological subtypes (15),
104	International Mesothelioma Interest Group stage (18), chemotherapy regimen, age,
105	gender, Eastern Cooperative Oncology Group performance status (PS) (19), subjective
106	symptoms, smoking history, asbestos exposure history, comorbidities (Charlson score
107	(20)) and baseline blood or effusion parameters at the time of diagnosis.
108	
109	Primary outcomes measurement
110	The primary outcome endpoint was OS, as defined by the length of time from the date
111	of diagnosis to death. Patients who had not died or who were lost to follow-up were
112	censored when they were last known to be alive before September 1 st , 2013.

114 Statistical analyses

115	We developed the PI in those who were treated with chemotherapy to minimize the bias
116	due to confounding by indication (21). We also evaluated the applicability of the PI in
117	those that received BSC.
118	In derivation, step continuous and nominal prognostic variables were dichotomized
119	according to previous studies (8, 9, 11, 21–27). OS was estimated using the
120	Kaplan-Meier method. The log-rank tests for each prognostic factor were used for
121	univariate analyses. The Cox proportional hazards model was used for multivariate
122	analyses. The Akaike's information criterion (AIC), Schwartz's Bayesian information
123	criterion (BIC), and Harrell's c index (c-index) were used for the discrimination of the
124	model. A final model was chosen based on both clinical and statistical significance. We
125	compared the discrimination of our index with the EORTC prognostic index (9) and the
126	progression-free index of EORTC (11).
127	Calibration curves showing agreement between observed and predicted outcomes over a
128	range of predicted probabilities were drawn. We also drew Cox-Snell residuals and
129	measured Moreau, O'Quigley, and Lellouch statistics (28). We drew log-log hazards
130	curves and tested the proportional hazard assumption. The bootstrapping technique was
131	used for the internal validation (for 500 replications (29)).

132	We carried out sensitivity analysis using multiple imputation for variants with clinically
133	significance. Two-sided p values < 0.05 were considered to indicate statistical
134	significance. We used Stata® ver. 13.0 (Stata Corp., College Station, TX).
135	
136	Ethical considerations
137	This study was performed according to the Declaration of Helsinki and the Ethical
138	Guidelines for Epidemiological Research by the Japanese Ministry of Health, Labour
139	and Welfare. The protocol for the study was approved by the Ethics Committee of
140	Kyoto University Graduate School and Faculty of Medicine (E1883). The protocol was
141	registered in the University Hospital Medical Information Network Clinical Trials
142	Registry with the number: UMIN000011733.
143	
144	Results
145	This study included 228 patients who were treated with chemotherapy with pemetrexed
146	and 55 patients who received BSC (Figure 1). Patient characteristics are shown in Table

147 1. Survival curves for each group are shown in the Figure 2.

148	The median lengths of follow-up were 345.5 days for the chemotherapy group and 250
149	days for the BSC group. During the follow-up period, 161 patients (70.6%) died in
150	chemotherapy group, and 40 patients (72.7%) died in the BSC group, respectively.
151	Univariate survival analyses are also shown in Table 1. Fifteen parameters were
152	significantly correlated with OS according to univariate analyses: asbestos exposure, PS,
153	dyspnea, anorexia, chest pain, body weight (BW) loss, fever, histological type, Stage,
154	Regimen, white blood cell (WBC), platelet (Plt) count, C-reactive protein (CRP),
155	Lactate dehydrogenase (LDH), and cytokeratin-19 fragment (CYFRA).
156	Because of the theoretical collinearity of symptom variables, we chose only PS with
157	respect to clinical relevance. We repeated the multivariate analysis while analyzing
158	WBC, Plt, and CRP, separately, because of the collinearity of inflammatory variables.
159	The discrimination for PS, Asbestos Exposure, Histology, Stage, Regimen, LDH, and
160	CYFRA were 823 (AIC), 844 (BIC), and 0.714 (c-index). The discrimination for seven
161	variables with WBC were 821 (AIC), 845 (BIC), and 0.726 (c-index). The
162	discrimination for six variables with CRP were 825 (AIC), 849 (BIC), and 0.715
163	(c-index). The discrimination for six variables with Plt were 824 (AIC), 848 (BIC), and
164	0.711 (c-index). We entered WBC into a stepwise backward Cox proportional hazards
165	model (Table 2). PS, histology, stage, and regimen remained significant after the

166	multivariate analysis. Hazard ratios of individual risk factors were 1.82-2.25. Therefore,
167	a PI for the OS was constructed using a simple count of the number of risk factors
168	(Table 3). The median OS of each category is shown in Table 4.
169	We calculated the discrimination of the rPHS (regimen, PS, Histology, or Stage) index.
170	The c-index was 0.677. After 500 bootstrap replications from the original patients, the
171	95% confidence interval (CI) of the c-index of the PHS score was 0.624-0.729.
172	We calculated the c-index for the EORTC prognostic index (9), which was 0.569. The
173	difference between the two indices persisted after bootstrap replications (0.108; 95% CI,
174	0.053-0.163). We also calculated the c-index for the progression-free index of the
175	EORTC (11), which was 0.552. The difference between the two indices persisted after
176	bootstrap replications (0.125, 95%CI, 0.082-0.166).
177	There was good calibration of the model, with close agreement between observed and
178	predicted OS (Figure S1), and also with close agreement between Cox-Snell residuals
179	and the 45-degree slope (Figure S2). The Moreau, O'Quigley, and Lellouch test showed
180	that the model fit of the Cox regression model was adequate $(p = 0.38)$.
181	We drew log-log hazards curves for the CTx group which were parallel (Figure S3). The
182	p value of the test for the proportional hazard assumption was 0.07.

184	multiply imputed datasets. We imputed only PS with regards to clinical significance.
185	These estimates and their standard errors were combined using Rubin's rules (30). The
186	results showed consistency (Table 4). The discrimination was also consistent in the BSC
187	group (c-index, 0.799; 95%CI, 0.725-0.874).
188	
189	Discussion
190	We developed a new PI for patients with MPM that predicts median OS, incorporates
191	pemetrexed information, and incorporates recent advancements in supportive care in the
192	normal clinical setting. The rPHS index is obtained by a simple count of the risk factors
193	(regimen including platinum plus PEM, PS>0, non-epithelial histology, and stage>3).
194	The index can stratify patients into four different prognostic groups with different
195	median survivals. The index has good discrimination for those treated with pemetrexed
196	group as well as those treated with BSC.
197	Patients with advanced cancer often want to know their prognosis (31). One study (32)
198	reported that patients with advanced cancer have an overwhelming preference for an
199	opportunity to prepare for the end of life. They want to know that their families are
200	prepared for their death, which often includes having finances in order, and for patients,

We carried out sensitivity analysis using multiple imputation to create and analyze 10

183

201	having funeral arrangements planned. They want to have the opportunity to resolve
202	unfinished business, remember personal accomplishments, and to say goodbye to
203	important people. In order to allow these patients to direct their energies to these matters,
204	it is important to provide them with accurate information regarding their prognosis. In
205	fact, early palliative care, including early accurate perceptions of prognosis, has
206	improved the quality of life and possibly the OS of patients with advanced cancer (6).
207	We believe that the present findings will influence the usual care of MPM patients for
208	several reasons. When one patient diagnosed with MPM and decided to treat with
209	pemetrexed-regimen, the patient and their physician can discuss based on the
210	median OS of the rPHS index. Without the index we discussed the prognosis based
211	on the median survival time from the trial or the cohort study. Our PI consists of
212	variables frequently used in usual care of MPM patients. Indeed, PS, histology, and
213	stage are well-known prognostic factors in previous studies (8-11, 33) and are
214	components of the evaluation at the time of initial diagnosis (34). Further, our PI can be
215	calculated easily by simple counting; calculators are not necessary, and our PI has more
216	discriminatory power than the EORTC PI (9), which is one of the best-known clinical
217	PIs. We note that the distribution of median age and OS were different when comparing
218	previous reports (8-11) and our CTx cohort; our study included more elderly patients

219	(67.7 versus 58-62 years), and our study included patients with relatively better
220	prognoses (11.5 versus 5-12.6 months). The cohort of our study ensures the
221	generalizability of the findings, because the two hospitals cover the South Hanshin
222	medical region and any patients with MPM in this region will visit one of these two
223	hospitals. So, the participants in the present study are a good representation of patients
224	with MPM. We included only patients with histologically proven MPM and not those
225	with only cytologically proven MPM. Because there is morphologic overlap between
226	benign reactive mesothelial cells and malignant cells of mesothelioma (15), it is not
227	recommended to make a diagnosis of mesothelioma based on cytology alone (34). We
228	think that this restriction ensured our study result.
229	Our cohort consisted of patients treated with BSC. For the small number of BSC
229 230	Our cohort consisted of patients treated with BSC. For the small number of BSC patients we didn't develop another index for BSC patients, but validated PHS index.
229 230 231	Our cohort consisted of patients treated with BSC. For the small number of BSC participants we didn't develop another index for BSC patients, but validated PHS index. The discrimination was good (c-index, 0.799; 95% CI, 0.725-0.874). No previous study
229 230 231 232	Our cohort consisted of patients treated with BSC. For the small number of BSC participants we didn't develop another index for BSC patients, but validated PHS index. The discrimination was good (c-index, 0.799; 95% CI, 0.725-0.874). No previous study has validated a PI in patients treated with BSC. This information will be useful for
 229 230 231 232 233 	Our cohort consisted of patients treated with BSC. For the small number of BSC participants we didn't develop another index for BSC patients, but validated PHS index. The discrimination was good (c-index, 0.799; 95% CI, 0.725-0.874). No previous study has validated a PI in patients treated with BSC. This information will be useful for discussions regarding prognosis between clinicians and their patients.
 229 230 231 232 233 234 	Our cohort consisted of patients treated with BSC. For the small number of BSC participants we didn't develop another index for BSC patients, but validated PHS index. The discrimination was good (c-index, 0.799; 95% CI, 0.725-0.874). No previous study has validated a PI in patients treated with BSC. This information will be useful for discussions regarding prognosis between clinicians and their patients. Since 1998, several PIs have been described. In contrast to our PHS index, other PIs
 229 230 231 232 233 	Our cohort consisted of patients treated with BSC. For the small number of BSC participants we didn't develop another index for BSC patients, but validated PHS index. The discrimination was good (c-index, 0.799; 95% CI, 0.725-0.874). No previous study has validated a PI in patients treated with BSC. This information will be useful for discussions regarding prognosis between clinicians and their patients.

237	of Bottomley (23) because we did not evaluate patients with the EORTC LC13 or
238	QLQ-C30. Their PI's c-index was 0.66. The point estimation was similar to that of our
239	PI. Pass (27) reported stage, histology, sex, age, treatment, adjuvant treatment, platelets
240	and WBC are clinical prognostic factor except for PS. We think this discrepancy may
241	reflect the difference of target population. We excluded those received surgery, but
242	Pass's target population is those received either palliative or potentially curative surgery.
243	There are several limitations in the study. First, this was a retrospective study with a
244	substantial number of missing PS data, so we performed sensitivity analysis using
245	multiple imputation. The result confirms the robustness of our model. Second, we were
246	not able to know the reason why each patient treated with the modality because this is a
247	retrospective study and treatment allocations were not protocol based. To clarify the
248	preferences for treatment in MPM patients prospective qualitative and quantitative
249	studies will be needed (36). But this limitation reflects the normal clinical setting. Third,
250	we assessed internal validation with the bootstrap method, but the sample size of this
251	study did not allow for external validation, so validation studies are needed.
252	We developed a new PI using PS, histology, and stage for MPM patients treated with
253	chemotherapy or BSC. This PI will allow better discussion between clinicians and

254	patients	with re	gards to	prognosis.	Further	prospective	e studies	using	this I	PI	are

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261

262 **Conflict of interest statement**

None declared.

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Clinical factors	Chemotherapy	Median OS		BSC		
	(n = 228), N (%)	(days)	95%CI	p value	(n=55), n (%)	
Age (years)	67.7±8.2	-			74.5±9.6	
mean±SD	07.7±0.2				74.3±9.0	
Age (years)						
75>	181 (79.4)	512	375-562	0.2000	24 (43.6)	
75≤	47 (20.6)	366	190-441		31 (56.4)	
Gender						
Female	39 (17.1)	514	314-699	0.4700	13 (23.6)	
Male	189 (82.9)	432	359-524		42 (76.4)	
Smoke						
Never	65 (30.0)	524	366-624	0.4100	19 (38.0)	
Current / Ever	152 (70.0)	425	327-524		31 (62.0)	
Missing	11				5	
Charlson						
comorbidity index						
<2	205 (89.9)	461	372-533	0.5100	42 (76.4)	
2≤	23 (10.1)	366	224-1213		13 (23.6)	
Asbestos exposure						
No	28 (12.4)	710	327-1213	0.0480		
Yes	197 (87.6)	397	353-511		14 (26.4)	
Missing	3				39 (73.6)	
PS					2	
0	37 (22.4)	926	524-1372	0.0014		
1≤	128 (77.6)	434	362-562		19 (41.3)	
Missing	63				37 (58.7)	
Dyspnea					9	
No	52 (33.8)	658	524-1030	0.0003	13 (31.7)	
Yes	102 (66.2)	425	319-512		28 (68.3)	
Missing	74				14	
Anorexia						
No	145 (82.4)	524	432-654	0.0001	29 (55.8)	

Table 1 Patient characteristics and results of univariate analyses of OS

Vac	21(176)	206	166 272		22(44.2)
Yes	31 (17.6)	296	166-373		23 (44.2)
Missing	52				3
Chest pain					
No	58 (39.2)	648	511-926	0.0007	16 (43.2)
Yes	90 (60.8)	353	263-432		21 (56.8)
Missing	80				18
BW loss					
No	96 (70.7)	566	512-804	0.0001	20 (60.6)
Yes	41 (30.0)	299	177-425		13 (39.4)
Missing	91				22
Fever					
No	92 (76.7)	524	397-648	0.0280	38 (92.7)
Yes	28 (23.3)	353	223-518		3 (7.3)
Missing	108				14
Histological type					
Epithelial	149 (65.4)	545	493-640	0.0000	17 (30.9)
Non-epithelial	79 (34.7)	277	221-330		38 (69.1)
Stage					
I-III	133 (58.3)	549	461-658	0.0000	30 (54.5)
IV	95 (41.7)	327	242-375		25 (45.5)
Regimen					
Platinum plus PEM	205 (89.9)	221	373-547	0.0007	
PEM monotherapy	23 (10.1)	499	86-425		
WBC (/µl)					
8300>	160 (70.5)	512	391-598	0.0400	36 (65.5)
8300≤	67 (29.5)	359	238-501		19 (34.5)
Missing	1				
Neutro/lymph					
5>	182 (82.4)	445	368-549	0.0600	35 (64.8)
5≤	39 (17.7)	362	188-514		19 (35.2)
Missing	7	002	100 011		1
Hb (g/dL)	,				Ĩ
10 (g/uL)	218 (96)	445	372-544	0.0700	45 (81.8)
10 - 10 - 10 -	9 (4.0)	224	66-526	0.0700	43 (81.8) 10 (18.2)
		<i>22</i> 4	00-320		10 (10.2)
Missing	1				

Plt (10^5/	/μl)					
	40>	188 (82.8)	461	373-549	0.0100	42 (76.4)
	40≤	39 (17.2)	327	176-526		13 (23.6)
	Missing	1				
ALP (IU/	I)					
	Abnormal	32 (14.9)	397	228-562	0.93	0 (0.0)
	Normal	183 (85.1)	441	362-544		53 (100.0)
	Missing	13				2
LDH (IU/	'L)					
	Abnormal	26 (11.6)	493	375-544	0.011	11 (20.4)
	Normal	198 (88.4)	242	87-603		43 (79.6)
	Missing	4				1
CRP (mg/	/dl)					
	5>	189 (83.3)	461	373-549	0.0076	40 (72.7)
	5≤	38 (16.7)	359	167-518		15 (27.3)
	Missing	1				
CEA (ng/	ml)					
5>		200 (94.3)	338	156-NE	0.7800	45 (95.7)
	5≤	12 (5.7)	441	366-526		2 (4.3)
	Missing	16				8
CYFRA (ng/ml)					
	3.5>	162 (75)	512	375-598	0.0090	22 (48.9)
	3.5≤	54 (25)	368	242-445		23 (51.1)
	Missing	12				10
Pleural	glucose					
(mg/dl)						
	40>	21(22.3)	511	156-710	0.2200	10 (30.3)
	40≤	73(77.7)	373	319-547		23 (69.7)
	Missing	134				22

Abbreviations: N, number; OS, overall survival; SD, standard deviation; CI, confidence interval; BSC, best supprotive care; PS, Eastern Cooperative Oncology Group performance status; BW, body weight; PEM, pemetrexed; WBC, white blood cell; Neutro, neutrocyte; Lymph, lymphocyte; Hb, hemoglobin; Plt, platelet; ALP, alkaly phosphatase; LDH, lactate dehydrogenase; CRP, C-reactive protein; CEA, carcinoembryonic antigen; CYFRA, cytokeratin-19 fragment.

Clinical factors	HR	95%CI
PS		
0	1	
1≤	2.40	1.36-4.23
Asbestos exposure		
no	1	
yes	1.64	0.75-3.58
Histological type		
Epithelial	1	
Non-epithelial	2.16	1.40-3.32
Regimen		
Platinum doublet	1	
Pemetrexed only	3.18	1.59-6.39
Stage		
I-III	1	
IV	1.57	1.03-2.39
LDH		
Normal	1	
Abnormal	1.46	0.71-2.99
CYFRA (ng/ml)		
3.5>	1	
3.5≤	1.10	0.69-1.76
WBC (/µl)		
8300>	1	
8300≤	1.56	0.99-2.45

Table 2 Backward Cox proportional hazards model

Abbreviations: HR, hazard ratio; CI, confidence interval; PS, Eastern Cooperative Oncology Group performance status; LDH, lactate dehydrogenase; CYFRA, cytokeratin-19 fragment; WBC, white blood cell. Table 3 Final model

Clinical factors	HR	95% CI	Score
PS			
0	1		1
1≤	2.06	1.22-3.44	
Histological type			
Epithelial	1		1
Non-epithelial	2.15	1.41-3.26	
Stage			
I-III	1		1
IV	1.82	1.23-2.69	
Regimen			
Platinum plus PEM	1		1
PEM monotherapy	2.25	1.16-4.36	

Abbreviations: HR, hazard ratio; CI, confidence interval; PS, Eastern

Cooperative Oncology Group performance status; PEM, pemetrexed.

Chemotherapy			Best supportive care			
Score	Ν	Median OS	95% CI	Ν	Median OS	95%CI
		(days)			(days)	
0	24 (28)	1030 (926)	661-1399(598-1253)			
1	57 (76)	658 (603)	444-872 (458-678)	6 (7)	573 (573)	530-616 (477-669)
2	56 (79)	373 (367)	223-522 (305-429)	15 (20)	408 (402)	178-638 (221-583)
3	22 (39)	327 (240)	189-465 (133-347)	11 (20)	250 (94)	11-489 (0-228)
4	5(6)	125(48)	16-234(0-184)	6 (8)	26 (34)	0-103 (0-126)

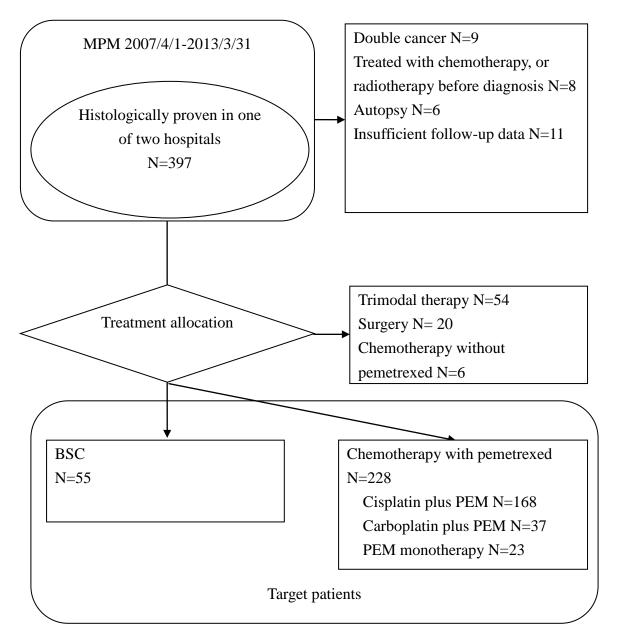
 Table 4 The rPHS index for overall survival (sensitivity analysis)

rPHS index = (if platinum + PEM 0, otherwise 1) + (if PS 0<, otherwise 0) + (if Histology non-epithelial,

otherwise 0) + (if Stage=4, otherwise 0)

Abbreviations: N, number; OS, overall survival; CI, confidence interval; NE, not estimable; PS, Eastern Cooperative Oncology Group performance status; PEM, pemetrexed.

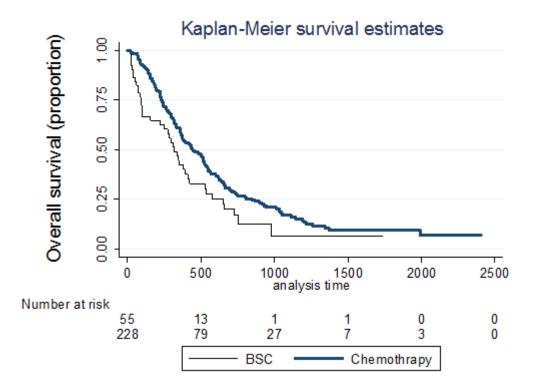
Figure 1 Patient flow diagram



Abbreviations: N, number; MPM, malignant pleural mesothelioma; BSC, best

supportive care.

Figure 2 Survival curve (days)



Abbreviations: BSC, best supportive care;

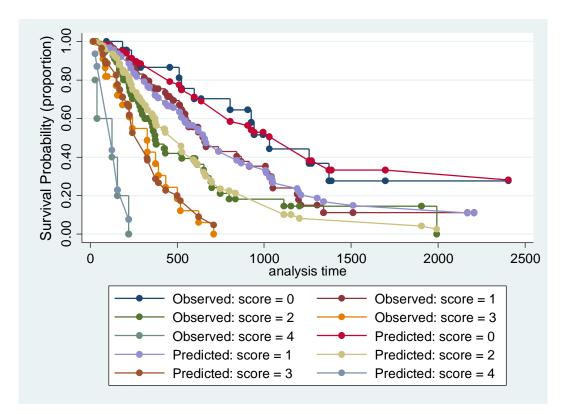
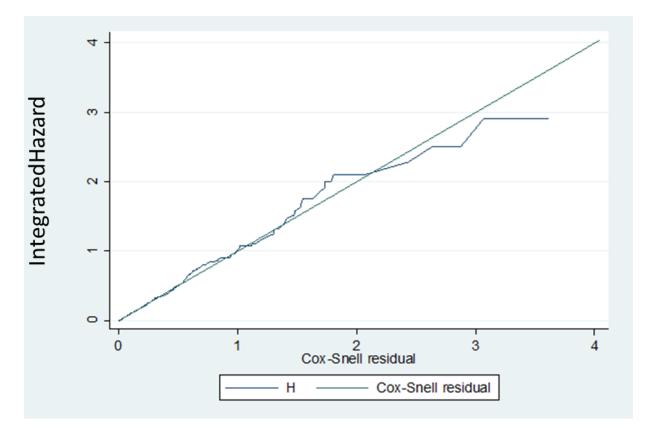


Figure S1 Calibration Kaplan-Meire curve of the rPHS index for chemotherapy group

Abbreviations: BSC, best supportive care;

Figure S2 Cox-Snell Residuals Graph



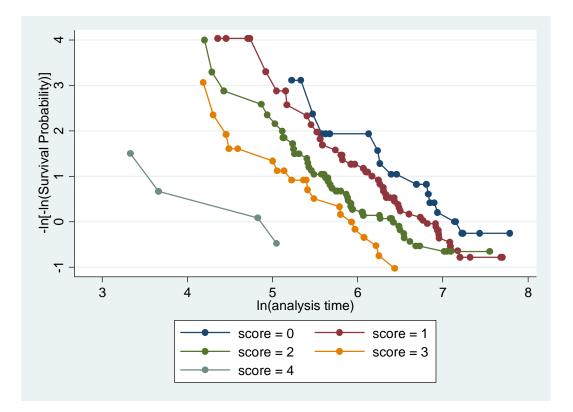


Figure S3 Cumulative hazards curves for the pemetrexed (CTx) group