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1 Title

A model relating transpiration for Japanese cedar and cypress plantations with stand
 structure

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32 Abstract

Previous studies have revealed that changes in forest structure due to management (e.g., 33 thinning, aging, and clearcutting) could affect the forest water balance. However, there 34 35 are unexplained variability in changes in the annual water balance with changing structure among different sites. This is the case even when analyzing data for specific 36 species/regions. For a more advanced and process-based understanding of changes in the 37 water balance with changing forest structure, we examined transpiration (E) observed 38 using the sap-flux method for 14 Japanese cedar and cypress plantations with various 39 40 structure (e.g., stem density and diameter) in Japan and surrounding areas and developed a model relating E with structural parameters. We expressed E using the simplified 41 Penman–Monteith equation and modeled canopy conductance  $(G_c)$  as a product of 42 reference  $G_c$  ( $G_{cref}$ ) when vapor pressure deficit is 1.0 kPa and functions expressing the 43 44 responses of  $G_c$  to meteorological factors. We determined  $G_{cref}$  and parameters of the 45 functions for the sites separately. E observed for the 14 sites was not reproduced well by the model when using mean values of  $G_{cref}$  and the parameters among the sites. However, 46 47 E observed for the sites was reproduced well when using  $G_{cref}$  determined for each site 48 and mean values of the parameters of the functions among the sites, similar to the case when using  $G_{cref}$  and the parameters of the functions determined for each site. These 49 results suggest that considering variations in  $G_{cref}$  among the sites was important to 50 reproduce variations in *E*, but considering variations in the parameters of the functions 51

52	was not. Our analysis revealed that $G_{cref}$ linearly related with the sapwood area on a stand
53	scale $(A)$ and that A linearly related with stem density $(N)$ and powers of the mean stem
54	diameter $(d_m)$ . Thus, we proposed a model relating E with A (or N and $d_m$ ), where $G_{cref}$
55	was calculated from A (or N and $d_m$ ) and the parameters of the functions were assumed to
56	be the mean values among the sites. This model estimates changes in $E$ with changing
57	structure from commonly available data ( $N$ and $d_m$ ), and therefore helps improve our
58	understanding of the underlying processes of the changes in the water balance for
59	Japanese cedar and cypress plantations.

Key words: canopy conductance; forest structure; model; sapwood area; simplified
Penman–Monteith equation; transpiration

64 1. Introduction

Changes in forest structure due to management (e.g., planting, growth, thinning, 65 aging, and clearcutting) can affect the forest water balance. Numerous studies (Scott and 66 Lesch, 1997; Cornish and Vertessy, 2001) have examined changes in the annual water 67 balance with changing forest structure on the basis of catchment water balance data. 68 Summarizing data derived from such studies and analyzing them using linear regression, 69 researchers have identified several important parameters (e.g., the ratio of the cutting area 70 71 to the total catchment area, annual rainfall, and leaf phenology) determining changes in 72 annual runoff with changing structure (Bosch and Hewlett, 1982; Brown et al., 2005; Komatsu et al., 2011). However, there remains unexplained variability in changes in the 73 annual water balance (Bosch and Hewlett, 1982; Brown et al., 2005). This is the case even 74 for data for a single species within a specific region, although the variability is less 75 76 pronounced (Adams and Fowler, 2006; Komatsu et al., 2011). This suggests the limitation of analyzing catchment water balance data simply using linear regression without 77 considering underlying processes (e.g., canopy transpiration and interception 78 evaporation). For a more advanced understanding of changes in the annual water balance 79 80 with changing structure, examining relationships of canopy transpiration and interception for given species with various structure (e.g., stem density and diameter) would be useful 81 (Komatsu et al., 2007d). Variations in stem density and diameter relate with the sapwood 82 area on a stand scale and leaf area index, which could in turn relate with canopy 83

transpiration (Granier et al., 2000a; Ewers et al., 2011). Focusing on specific species
would be useful in reducing factors to be considered, because there are factors (e.g., the
clumpling factor and stem conductivity) that could differ among different species and
affect canopy transpiration (Baldocchi and Meyers, 1998; Zwieniecki and Holbrook,
1998; Bréda, 2003).

Japanese cedar (*Cryptomeria japonica*) and cypress (*Chamaeryparis obtusa*) are 89 major plantation species in Japan and surrounding areas such as China and Taiwan (Japan 90 91 Forestry Agency, 2014). Examining the water balance of these plantations is highly 92 important from a practical viewpoint. Forest management (e.g., thinning) has been performed to increase water resources in Japan, although its effectiveness has not been 93 94 assessed sufficiently (Komatsu et al., 2010). Summarizing data for interception evaporation of Japanese cedar and cypress plantations (Hattori et al., 1982; Tanaka et al., 95 2005), Komatsu et al. (2007a) found a relationship between stem density and interception 96 97 evaporation and then developed a model relating interception evaporation with stem 98 density. However, there have been few studies examining the relationship between forest structure and canopy transpiration (E) for Japanese cedar and cypress plantations. 99

To assess changes in the water balance of Japanese cedar and cypress plantations with changing forest structure, we developed a model relating E for Japanese cedar and cypress plantations with meteorological and structural variables. The model formulates Eusing the simplified Penman–Monteith equation (McNaughton and Black, 1973; Jarvis 104 and McNaughton, 1986). Canopy conductance  $(G_c)$  in the equation was written as a series of functions expressing responses of  $G_c$  to meteorological factors (Jarvis, 1976; 105 Lohammer et al., 1980). This study comprises three steps. First, we calculated  $G_c$  using 106 107 E data derived employing the sap-flux method for 14 sites and the inverted form of the simplified Penman–Monteith equation to determine parameters of the model for the sites 108 separately. Second, we assessed the importance of each parameter in determining E on 109 the basis of sensitivity analysis. Third, we examined the relationship of the important 110 111 parameters with structural parameters. Here, we tried to relate the structural parameters 112 with commonly available data such as stem density and diameter for wide use of the 113 model.

Models estimating *E* are roughly classified into two groups. One uses many 114 empirical parameters for modeling stomatal/canopy conductance to avoid considering 115 116 internal hydraulics and keep the model structure simple (Granier et al., 2000a; Komatsu, 2004), while the other considers internal hydraulics (Domec et al., 2012; McDowell et al., 117 2013). Our model belongs to the former group. Most models of the former group focus 118 on reproducing *E* for a specific site (Cienciala et al., 1994a,b; Granier and Bréda, 1996). 119 120 Several models (Granier et al., 2000a; Komatsu, 2004) are applicable to various sites, similar to our model. Our model differs from these models in that our model focuses on 121 two species, which suggests higher predictability of the model when applied to the species. 122 Furthermore, our model differs from the models in that our model would be tested against 123

E data recorded not only during growing seasons but during winter, suggesting higher reliability of the model to predict *E* on a long time scale (e.g., one year). 125 126 2. Theory 127 The model, using the simplified Penman-Monteith equation (McNaughton and 128 Black, 1973; Jarvis and McNaughton, 1986), expresses E as 129  $E = \frac{\rho C_p G_c D}{\gamma \lambda},$ (1)130 where  $\rho$  is the air density,  $C_p$  is the specific heat of air,  $G_c$  is canopy conductance, D is 131 132 the vapor pressure deficit,  $\gamma$  is the psychrometric constant, and  $\lambda$  is the latent heat of water vaporization. This equation is derived from the Penman-Monteith equation under the 133 assumption of complete coupling between the canopy and atmosphere. 134  $G_c$  is formulated as a product of the reference value of  $G_c$  when D is 1.0 kPa 135 ( $G_{cref}$ , Oren et al., 1999) and functions expressing responses of  $G_c$  to meteorological 136 factors (Jarvis, 1976; Lohammer et al., 1980): 137  $G_c = G_{cref} \cdot f(D) \cdot g(R) \cdot h(T),$ (2)138 where f(D), g(R), and h(T) are functions expressing the responses of  $G_c$  to mean daytime 139 D, solar radiation (R), and air temperature (T), respectively. f(D), g(R), and h(T) are 140 respectively modeled as (Oren et al., 1999; Granier et al., 2000b) 141  $f(D) = 1.00 - \beta \cdot \ln(D),$ 142 (3)  $g(R) = \min\left\{ \left(\frac{R}{600}\right)^{\delta}, 1.00 \right\},\,$ 143 (4)

124

8

144 
$$h(T) = \begin{cases} 1.00 \quad (T \ge \epsilon) \\ \frac{T-\zeta}{\epsilon-\zeta} \quad (\zeta < T < \epsilon) \\ 0.00 \quad (T \le \zeta) \end{cases}$$
(5)

where  $\beta$ ,  $\delta$ ,  $\varepsilon$ , and  $\zeta$  are parameters. The model does not consider the effect of the soil water deficit on  $G_c$ . Most previous studies, making continuous measurements of E (or its substitutes) for forests in Japan including Japanese cedar and cypress plantations (Komatsu et al., 2006; Kosugi et al., 2007; Kumagai et al., 2007), did not report a clear reduction in  $G_c$  or E with a soil water deficit, although there were a few exceptions (Tanaka et al., 2002; Komatsu et al., 2007c).

151 We hypothesized that considering the variation in  $G_{cref}$  among sites would be important but considering variations in  $\beta$ ,  $\delta$ ,  $\varepsilon$ , and  $\zeta$  among sites would not be in 152 reproducing E for the sites. Our hypothesis was based on results of previous studies. Gcref 153 154 linearly relates with E in the simplified Penman–Monteith equation. G<sub>cref</sub> values reported previously (Granier et al., 2000a; Komatsu et al., 2012, 2013) often differ greatly among 155 156 different forest sites comprising the same species, implying that considering the variation in  $G_{cref}$  among the sites is important in estimating E. Oren et al. (1999) and succeeding 157 158 studies (Addington et al., 2004; Ewers et al., 2008) noted a fairly conservative  $\beta$  among different sites. Thus, assuming  $\beta$  as constant among sites might not introduce large errors 159 in E estimates. Komatsu et al. (2006) analyzed sap flux data on an hourly time scale for 160 a Japanese cedar plantation and pointed out that the control of  $G_c$  by R was important only 161

162 early in the morning and late afternoon, when *E* is small. This implies that assuming  $\delta$  as 163 constant among sites would not introduce large errors in estimating *E* on a daily time 164 scale.  $\varepsilon$  and  $\zeta$  affect *E* during winter, which accounts for a small portion of annual *E* for 165 Japanese cedar and cypress plantations (Suzuki, 1980; Kumagai et al., 2014). Thus, 166 assuming  $\varepsilon$  and  $\zeta$  as constant among sites might not introduce large errors in estimating 167 *E* on a longer time scale, such as one year.

We further hypothesized that  $G_{cref}$  would relate with the sapwood area on a stand 168 169 scale (A) for the following reasons. Kumagai et al. (2007) examined E for two cedar forest 170 sites located nearby but having different A. The relative difference in E between the sites was comparable to that in A. Komatsu et al. (2013) compared E values for a cedar forest 171 during two successive growing seasons just before and after thinning. The ratio of E after 172 thinning to that before thinning approximated the ratio of A after thinning to that before 173 thinning. The results of Kumagai et al. (2007) and Komatsu et al. (2013) suggest a 174 correlation between A and  $G_{cref}$ . Previous studies (Macfarlane et al., 2010; Ewers et al., 175 2011) reported relationships between A and E for forests comprising other species. These 176 results also suggest that our hypothesis is reasonable. The relationship between A and 177 178  $G_{cref}$  would allow the estimation of  $G_{cref}$  with the input of A, and then the estimation of E by assuming typical values of  $\beta$ ,  $\delta$ ,  $\varepsilon$ , and  $\zeta$  and using the simplified Penman–Monteith 179 180 equation.

181

182 3. Data used

We used E data recorded for nine Japanese cedar and five Japanese cypress 183 plantations (Table 1). Twelve of the 14 sites were located in western Japan, but the IK and 184 185 XT sites were located in eastern Japan and in Taiwan, respectively. Structural parameters differed among the sites. Stem density (N) ranged between 600 and 1575 stems ha<sup>-1</sup> for 186 cedar and between 350 and 2100 stems ha<sup>-1</sup> for cypress. The mean diameter at breast 187 height  $(d_m)$  ranged between 13.5 and 48.4 cm. The leaf area index (L), estimated using 188 optical methods (see the footnote #2 of Table 1), ranged between 0.8 and 5.9  $m^2 m^{-2}$  for 189 cedar and between 0.8 and 4.8 m<sup>2</sup> m<sup>-2</sup> for cypress. Note that L for Japanese cedar and 190 cypress forests estimated using optical methods is generally lower than that estimated 191 using destructive leaf sampling and/or allometry equations (Hasegawa et al., 2013; 192 Tsuruta et al., 2014). The sapwood area at a stand scale (A) ranged between 14.1 and 46.0 193  $m^2$  ha<sup>-1</sup> for cedar and between 8.8 and 20.4 m<sup>2</sup> m<sup>-2</sup> for cypress. A was determined by 194 measurements of sapwood thickness using core sampling and assuming the stem cross-195 section was circular. A complete description of the measurements was provided by 196 Kumagai et al. (2007) and Kume et al. (2010). 197

*E* values for all sites were measured employing the sap-flux method and Granier
sensors (Granier, 1987). A detailed description of the measurements is available in the
papers cited in Table 1. Employing the method, *E* was estimated as (Kumagai et al., 2007;
Kume et al., 2010)

202 
$$E = \frac{\sum_{i=1}^{n} F_i \cdot a_i}{S}$$
, (6)

where F is the sap flux for an individual tree averaged over its sapwood area, a is the tree 203 sapwood area, n is the number of trees in the plot, and S is the ground area. For a tree 204 205 whose sapwood thickness was much greater than the sensor length (i.e., 2.0 cm), two or three sensors were installed radially to cover the sapwood area (Kumagai et al., 2007). F 206 was calculated as the weighted average of sap flux for the sensors. Azimuthal variations 207 of sap flux were also considered for some sites (Shinohara et al., 2013a). 208 209 210 4. Methods of analysis 211 4.1. Determination of parameters 212 We calculated canopy conductance  $(G_c)$  from E estimated using the sap-flux method and meteorological factors for each site using the inverted form of the simplified 213

214 Penman–Monteith equation:

215 
$$G_c = \frac{\gamma \lambda E}{\rho C_p D}.$$
 (7)

 $G_c$  was calculated as a daily average conductance using mean daytime *T* and *D*, and *E* summed over 24 hr but divided by daylight hours (Phillips and Oren, 1998; Kumagai et al., 2008). Sap-flux sensors observe water movement through the trunk during nighttime, which may represent recharge of water into upper sections of the tree trunk and branches or transpiration during nighttime. Dividing *E* by daylight hours assumes that sap flux observed during nighttime represents recharge of water. Recent studies (Dawson et al.,

222	2007; Fisher et al., 2007; Oishi et al., 2008) have reported that transpiration during
223	nighttime accounts for a considerable portion of daily transpiration. However,
224	transpiration during nighttime would be very low in Japan possibly because of low $D$
225	during nighttime, as shown by measurements of transpiration using the leaf-chamber
226	method (Kosugi et al., 1995, 1997; Tanaka et al., 2002). The "daytime" here was assumed
227	as the period between 6 a.m. and 6 p.m. throughout the year for simplicity, i.e., the
228	daylight hours were assumed as 12. These assumptions were also made in previous
229	studies (Kumagai et al., 2008; Komatsu et al., 2012). We confirmed that these
230	assumptions did not alter our conclusions. $D$ for the period between 6 a.m. and 6 p.m.
231	differed from $D$ for the period between sunrise and sunset by no more than 8% for our
232	sites. Our preliminary analysis revealed that $G_c$ calculated using $D$ for the period between
233	6 a.m. and 6 p.m. could differ from those calculated using $D$ for the period between sunset
234	and sunrise by no more than 10%, which was not large enough to alter our conclusions.
235	We determined $G_{cref}$ , $\beta$ , and $\delta$ for the sites where E data only during the growing
236	season were available. Here, $h(T)$ was assumed to be 1.00 because of relatively high T.
237	We determined $G_{cref}$ , $\beta$ , $\delta$ , $\varepsilon$ , and $\zeta$ for the sites where year-round <i>E</i> data were available.
238	Note that analysis for the XT site did not follow this policy, as detailed later.
239	The parameters were determined in a manner similar to that employed by

Komatsu et al. (2012, 2013). We first excluded data recorded on rainy days and days with D lower than 0.2 kPa. Data for sap flux (and therefore E and  $G_c$ ) could suffer from

measurement errors on rainy days (Kumagai et al., 2008; Komatsu et al., 2012). Gc could 242 be highly affected by measurement errors in D when D is very low (Ewers and Oren, 243 2000; Komatsu et al., 2007b). We then examined the relationship between D and  $G_c$  using 244 data recorded on days with high R (>400 W m<sup>-2</sup>) and T (>15 °C). We determined  $G_{cref}$ 245 and  $\beta$  by regressing the relationship emloying the least-squares method. After 246 determining  $G_{cref}$  and  $\beta$ , we examined the relationship between R and observed  $G_c$  divided 247 by  $G_{cref} f(D)$  using data recorded on days with high T (>15 °C) to determine  $\delta$ . Here, we 248 confirmed that the relationship between soil water content and observed  $G_c$  divided by 249  $G_{cref} f(D)g(R)$  was generally unclear for sites where data for soil water content were 250 available. We finally examined the relationship between T and observed  $G_c$  divided by 251  $G_{cref} f(D)g(R)$  to determine  $\varepsilon$  and  $\zeta$ . Here, we also confirmed that the relationship between 252 T and observed  $G_c$  divided by  $G_{cref} f(D)g(R)$  was generally unclear for the sites where E 253 254 data only during the growing season were available.

We applied a different method to the determination of parameters for XT. XT was located in a mountainous region under a humid subtropical climate. If we had applied the same method as that applied to other sites, we would have only limited data with which to examine the relationship between D and  $G_c$  because of a large number of rainy days and days with low R. The method applied to XT was as follows. We first excluded data recorded on days with daily rainfall more than 5.0 mm and days with D lower than 0.2 kPa. We then examined the relationship between D and  $G_c$  using data recorded on

days with R no less than 300 W m<sup>-2</sup>. We determined  $G_{cref}$  and  $\beta$  by regressing the 262 relationship employing the least-squares method. Here, we confirmed that there was no 263 systematic difference in  $G_c$  between days without rain and days with rain no more than 264 5.0 mm. After determining  $G_{cref}$  and  $\beta$ , we examined the relationship between R and 265 observed  $G_c$  divided by  $G_{cref} f(D)$  to determine  $\delta$ . In the above analysis, we did not exclude 266 data using a criterion about T, because the range of T was narrow for XT and a large 267 portion of the data satisfied T being no less than 15 °C. We did not determine the 268 269 parameters in h(T), because we did not find a clear relationship between T and observed 270  $G_c$  divided by  $G_{cref} f(D)g(R)$  for XT.

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## 4.2. Assessing the importance of parameters

We hypothesized that considering the variation in  $G_{cref}$  among the sites would be 273 274 important but considering variations in  $\beta$ ,  $\delta$ ,  $\varepsilon$ , and  $\zeta$  among the sites would not be in reproducing E for the 14 sites (detailed in the theory). To verify this hypothesis, we 275 calculated E for days with no rain and T no less than 15  $^{\circ}$ C (corresponding to a growing 276 season) for the sites using three different parameterizations: (P1G) the mean values of 277  $G_{cref}$ ,  $\beta$ , and  $\delta$  among the sites, (P2G)  $G_{cref}$  determined for each site and the mean values 278 of  $\beta$  and  $\delta$  among the sites, and (P3G)  $G_{cref}$ ,  $\beta$ , and  $\delta$  determined for each site. To evaluate 279 the reproducibility of E by P1G, P2G, and P3G, we examined the relationships of the 280 mean E for the days calculated using P1G, P2G, and P3G respectively with the mean 281

observed *E*. If the difference in reproducibility between P1G and P2G was found to be important, considering the variation in  $G_{cref}$  among the sites would be important in calculating the mean *E* for a growing season. If the difference in reproducibility between P2G and P3G was found to be unimportant, considering variations in  $\beta$  and  $\delta$  would not be important.

Similarly, we calculated E for days with no rain for sites having data recorded in 287 winter as well as those recorded in a growing season using three different 288 parameterizations: (P1W) the mean values of  $G_{cref}$ ,  $\beta$ ,  $\delta$ ,  $\varepsilon$ , and  $\zeta$  among the sites, (P2W) 289 290  $G_{cref}$  determined for each site and the mean values of  $\beta$ ,  $\delta$ ,  $\varepsilon$ , and  $\zeta$  among the sites, and (P3W)  $G_{cref}$ ,  $\beta$ ,  $\delta$ ,  $\varepsilon$ , and  $\zeta$  determined for each site. To evaluate the reproducibility of E 291 by P1W, P2W, and P3W, we examined the relationships of the mean E for the days 292 calculated using P1W, P2W, and P3W respectively with the mean observed E. If the 293 difference in reproducibility between P1W and P2W was found to be important, 294 considering the variation in  $G_{cref}$  would be important in calculating the mean E for a period 295 including a growing season and winter. If the difference in reproducibility between P2W 296 and P3W was found to be unimportant, considering variations in  $\beta$ ,  $\delta$ ,  $\varepsilon$ , and  $\zeta$  among the 297 298 sites would not be important.

When evaluating the model reproducibility, we primarily focused on the mean Eon a longer time scale, because our model was intended to be used to improve our understanding of the forest water balance on a long time scale (see Section 1). For

assessing the importance of the difference in reproducibility between different 302 parameterizations, we did not examine whether there was a statistically significant 303 difference between errors in *E* estimates using different parameterizations. What we need 304 305 to know is not whether there is a statistical difference, but whether the magnitude of the difference is important in a practical context (Bakan, 1966; Thompson, 1996, 1998; 306 Nuzzo, 2014). We thus examined whether errors in estimates of the mean E made using 307 the  $G_c$  model were less than potential errors in observed E. Kumagai et al. (2007) and 308 309 Kume et al. (2010) examined uncertainty in *E* estimates made using the sap-flux method 310 for Japanese cedar and cypress plantations. Sampling data for sap flux on a sensor scale repeatedly using the Monte-Carlo technique, Kumagai et al. (2007) and Kume et al. 311 (2010) revealed that the standard variation of E estimates was generally 10% or more of 312 the mean E. When considering that 95% of data theoretically fall in the range of two 313 314 standard deviations above and below the mean, potential errors in observed E would be less than 20% of the value in most cases. If errors in E estimates made using different 315 parameterizations were more than 20% of observed *E*, we would regard the difference in 316 reproducibility as important. 317

In addition, we examined whether errors in estimates of the mean E made using the  $G_c$  model were less than those in interception evaporation estimates made using the model developed by Komatsu et al. (2007a) or its revised form developed by Komatsu et al. (submitted), because our  $G_c$  model was intended to be used with the interception evaporation model to assess the water balance for Japanese cedar and cypress plantations (Komatsu et al., submitted). The typical errors in interception evaporation estimates for the period with *T* no less than 15 °C and for a year would be 0.22 and 0.20 mm day<sup>-1</sup>, respectively (Appendix A).

326

4.3. Relating important parameters with species and structures

We examined the relationship between A and  $G_{cref}$  to model  $G_{cref}$  using A. After 328 329 confirming correlation between the two variables, we regressed the relationship to obtain 330 a linear equation relating  $G_{cref}$  with A. The intercept of the equation was assumed to be zero. The slope of the equation was determined employing the least-squares method. To 331 examine the stability of the correlation and the slope, we calculated 95% confidence 332 intervals of these variables employing the bootstrapping method (Efron, 1979; Diaconis 333 and Efron, 1983). These intervals were "bias-corrected, accelerated" percentile intervals 334 calculated in the manner described by Efron (1987) and Fox (2008). 335

As data for *A* are not usually available for most Japanese cedar and cypress plantations, we tried to relate *A* with the mean diameter at breast height  $(d_m)$  and stem density (*N*) data in the following way. Tsuruta et al. (2011) summarized data for the sapwood area at a tree scale (*a*) for 81 cedar trees from six sites and 109 cypress trees from nine sites and examined the relationships between diameter at breast height (*d*) and *a* for these species. They did not find clear differences in the relationship among sites. They regressed the data to develop general equations for the relationships:  $a = a_{ref} \cdot d^k$ , where *a* and *d* were respectively in units of cm<sup>2</sup> and cm and  $a_{ref}$  and *k* were parameters. *a<sub>ref</sub>* was 1.40 and 1.96 cm<sup>2</sup> for cedar and cypress, respectively. *k* was 1.55 and 1.42 for cedar and cypress, respectively. Approximating this equation by the tangential line at *d* being equal to  $d_m$ , *A* was expressed as

 $347 A = N \cdot a_{\text{ref}} \cdot d_m^{\ k}. (8)$ 

To examine the validity of this method of estimating *A* from *N* and *d<sub>m</sub>*, we investigated the relationship between *A* estimated employing this method and observed *A*. For this validation, we used *A* data for YB, YA, IS, KL, KU, KS, XT, and IH, and another cypress site of Sun et al. (2014). Data for the other six sites (i.e., IK, IR, OL, OU, SK, and HW) were not used, because *a* data for the sites were used by Tsuruta et al. (2011) to develop the relationship between *d* and *a*.

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355 5. Results

356 5.1. Determination of parameters

Figure 1 shows the relationships between *D* and *G*<sub>c</sub>, between *R* and *G*<sub>c</sub> divided by  $G_{cref} f(D)$ , and between *T* and *G*<sub>c</sub> divided by  $G_{cref} f(D) g(R)$ . Here, we show data only for OL as an example. Regressing these relationships and relationships for the other sites, parameters for the sites were determined as listed in Table 2. The maximum  $G_{cref}$  (9.77 mm s<sup>-1</sup> for IK) among the sites was more than five times the minimum  $G_{cref}$  (1.76 mm s<sup>-1</sup> 362

<sup>1</sup> for IH). We observed variations in f(D), g(R), and h(T) among sites (Figure 2).

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365 Figure 3 shows the relationships of the mean E for days with no rain and T no less than 15 °C calculated using P1G, P2G, and P3G with observed E. E calculated using 366 the model was not strongly correlated (r = 0.454) with observed E for P1G, but was 367 strongly correlated (r = 0.993 and 0.997, respectively) and fell along the 1:1 line for P2G 368 369 and P3G. Errors for P1G were often greater than potential errors in observed E and interception evaporation estimates, but errors for P2G and P3G were not. Figure 4 shows 370 the relationships of mean E for days with no rain calculated using P1W, P2W, and P3W 371 with observed E for sites having data recorded in winter as well as those recorded in a 372 growing season. E calculated using the model was not strongly correlated (r = 0.597) with 373 observed E for P1W, but was strongly correlated (r = 0.999 and >0.999, respectively) and 374 fell along the 1:1 line for P2W and P3W. Errors for P1W were often greater than potential 375 376 errors in observed E and interception evaporation estimates, but errors for P2W and P3W were not. These results suggest that considering the variation in  $G_{cref}$  among the sites was 377 378 important for reproducing the mean E on a long time scale, but considering variations in  $\beta$ ,  $\delta$ ,  $\varepsilon$ , and  $\zeta$  was not. 379

380 Besides the mean *E* on a long time scale, we also examined model reproducibility 381 of seasonal and day-to-day variations in *E* using data on a daily time scale for sites where

382	year-round data were available. Figure 5 shows time series of E calculated using PIW,
383	P2W, and P3W and observed E. Here, we show data only for OU as an example. The
384	slope of the regression line for the relationship between calculated and observed $E$ often
385	fell outside the range between 0.8 and 1.2 for P1W, where the range corresponded to the
386	uncertainty in observed $E$ . However, the slope always fell in the range for P2W and P3W
387	(Table 3). The determination coefficients for P1W and P2W were often lower than that
388	for P3W. However, the difference was relatively small. These results suggest that
389	considering the variation in $G_{cref}$ among the sites was important for reproducing seasonal
390	and day-to-day variations in <i>E</i> , but considering variations in $\beta$ , $\delta$ , $\varepsilon$ , and $\zeta$ was not.

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## 392 5.3. Relating important parameters with species and structures

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 $G_{cref}$  tended to increase with A (Figure 6). The correlation coefficient (r) was 393 0.698 and the 95% confidence interval was (0.446, 0.951), when pooling data for cedar 394 and cypress. The correlation between A and  $G_{cref}$  was stronger than that between L and 395  $G_{cref}$  (r = 0.479). The regression line was determined as  $G_{cref}$  = 0.157 A, where A and  $G_{cref}$ 396 were in units of  $m^2 ha^{-1}$  and  $m s^{-1}$ , respectively. The 95% confidence interval of the slope 397 was (0.127, 0.195). Data for both cedar and cypress were located along the regression 398 line, suggesting no clear difference in G<sub>cref</sub> for a given A between cedar and cypress. Thus, 399 the regression equation could be used to predict  $G_{cref}$  from A for both cedar and cypress 400 plantations and then to predict E. Figure 7 shows the normal probability plot (Fujii, 2005; 401

Peck and Devore, 2005) for differences between  $G_{cref}$  estimated using the regression 402 equation and observed  $G_{cref}$ , where the normal score in the y-axis indicates the difference 403 divided by the standard deviation of the difference for the 14 sites. Data for all sites except 404 405 SK and IK were approximated by a line, indicating that most of these data followed a normal distribution. However, data for SK and IK were located far from the line, 406 suggesting that observed  $G_{cref}$  for these sites was higher than expected from the regression 407 equation and a normal distribution of the differences between estimated and observed 408 Gcref. 409

Figure 8 compares *A* estimated from *N* and  $d_m$  and observed *A*. Data were generally located around the 1:1 line. The mean relative error (i.e., the ratio of the difference between estimated and observed *A* to observed *A*) was 26%. The magnitude of this error is discussed in Section 6.3.

414

415 6. Discussion

416 6.1. Roles of f(D), g(R), and h(T)

Our model succeeded in reproducing *E* observed for the 14 sites, suggesting validity of our hypothesis. Our model includes three functions, i.e., f(D), g(R), and h(T). These functions influence on calculated *E* differently. Omitting f(D) from Eq. (2) (i.e., assuming f(D) being 1.0) causes overestimation of *E* during May–August (Figure 9a), resulting in a relatively low determination coefficient ( $R^2 = 0.740$ ) for the relationship 422

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between calculated and observed E. This suggests that considering reduction in  $G_c$  with increasing D is important in reproducing E when D is high.

- Omitting g(R) causes overestimation of E during October–January when R is relatively low. Omitting h(T) causes overestimation of E during January–February when T is relatively low. However, the determination coefficients for these cases ( $R^2 = 0.792$ and 0.798, respectively) are higher than that for the case of omitting f(D). This suggests that effects of omitting g(R) and h(T) on reproducing E on a daily time scale are less important than that of omitting f(D).
- 430

431 6.2. Variability in 
$$G_{cref}$$

 $G_{cref}$  values for SK and IK were higher than expected from the regression 432 equation (Figure 7). There are technical factors that potentially affect variations in  $G_{cref}$ 433 among sites. An insufficient number of sensors for sap flux measurements could result in 434 large errors in E estimates. Shinohara et al. (2013a) compared errors in E estimates 435 introduced by ignoring tree-to-tree variations, radial variations, and circumference 436 variations in sap flux for a Japanese cedar plantation. They concluded that tree-to-tree 437 438 variations in sap flux are the primary factor of errors in *E* estimates. Previous studies (Kumagai et al., 2005, 2007; Kume et al., 2010; Shinohara et al., 2013b) reported that 439 errors in E estimates for Japanese cedar and cypress were serious when the number of 440 trees in which sensors were installed was low, especially when there were less than 441

approximately 10. Note that this threshold number would be species specific, because 442 different threshold numbers were reported for other species (Oren et al., 1998; Mackay et 443 al., 2010). There were approximately 10 in which sensors were installed at each site (10 444 445 for SK and 9 for IK). Furthermore, E for IK estimated from sap flux data derived for nine trees differed less than 5% from that estimated from sap flux data derived for all 18 trees 446 according to intensive measurements performed during May 16–18 and June 1–4, 2010 447 for the site. Thus, the numbers of sensors at these sites are not likely a main factor 448 449 explaining the higher  $G_{cref}$ .

The use of the simplified Penman-Monteith equation could also be a factor 450 causing higher G<sub>cref</sub>. Aerodynamic conductance, which is assumed as infinite in the 451 simplified Penman–Monteith equation, generally ranges between 70 and 400 mm s<sup>-1</sup> for 452 coniferous forests including Japanese cedar and cypress plantations (Stewart and Thom, 453 1973; Yamanoi and Ohtani, 1992; Loustau et al., 1996; Tanaka et al., 1996). When 454 assuming aerodynamic conductance as 70 mm s<sup>-1</sup> and calculating  $G_{cref}$  using the original 455 Penman–Monteith equation (see the methods used by Komatsu et al., 2012), G<sub>cref</sub> values 456 for SK and IK are determined as 3.53 and 6.22 mm s<sup>-1</sup>, respectively. These  $G_{cref}$  data are 457 458 located close to the regression line in Figure 6, implying that the use of the simplified Penman–Monteith equation might be a potential factor explaining the higher  $G_{cref}$  for the 459 sites. Unfortunately, there have been no data for aerodynamic conductance observed at 460 these sites. However, data for wind speed for IK are available. The relationship between 461

462 D and  $G_c$  differed only slightly according to wind speed, suggesting that low aerodynamic 463 conductance would not be the primary factor causing low  $G_{cref}$  for IK. Data for wind speed 464 for SK were unavailable.

Besides the factors discussed above, age might be another possible factor explaining high  $G_{cref}$  for SK, which is younger than the other sites (Table 1). There have been several studies reporting or suggesting changes in sap flux on a stand scale with age for mono-specific forests (Tsuruta et al., 2008; Forrester et al. 2010). We do not have data for other sites of similar age to SK. We thus recommend examining  $G_{cref}$  for young Japanese cedar and cypress plantations.

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## 472 6.3. Errors in *A* estimates

We observed a clear correlation between estimated and observed *A* (Figure 8). However, the error in *A* estimates was relatively large. The mean relative error in *A* estimates (26%) was comparable to that in  $G_{cref}$  estimates obtained using the regression equation in Figure 5 (24%). Thus, it would be better to use observed *A* to calculate  $G_{cref}$ , if observed *A* is available.

The error in *A* estimates would be primarily caused by errors in *a* estimates using Tsuruta et al.'s (2011) equation, and only secondarily caused by errors due to tangential approximation of the equation (Eq. (8)). The tangential approximation could cause underestimation of *A*, because *a* values calculated using the tangential line are no more than those calculated using the original equation. In fact, *A* estimated employing the above
method did not always underestimate observed *A* (Figure 8).

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# 485 6.4. Possible applications and implications

Our model would be useful as a research tool for hydrologists. There have been 486 many studies (Dung et al., 2012; Kubota et al., 2013) examining changes in the annual 487 water balance with changing structure of Japanese cedar and cypress plantations due to 488 489 forest management. There are large variations in the change in the annual water balance among studies. Our model, accompanied with the interception evaporation model 490 developed by Komatsu et al. (2007a), could be used by hydrologists to calculate changes 491 in E and interception evaporation for catchments and improve our understanding of 492 underlying processes of the variations. Our model and the interception model are also 493 useful in estimating spatial variations in E and interception evaporation on a landscape 494 scale, when these models are used with forest inventory data. 495

As described above, our model has great potential for application. This is because our model can be used only with the inputs of commonly available data for forest structure (i.e., N and  $d_m$ ) and meteorology (see Appendix B). Our model is specific to Japanese cedar and cypress plantations. However, our model is important in that it demonstrates how to relate E with commonly available data. The concept of our model would be useful in developing similar models to estimate E for mono-specific forests 502 comprising other species.

On the other hand, the model needs to be tested further. The sites for cedar and 503 cypress plantations (Table 1) were located mainly in western Japan, where temperature is 504 505 intermediate or higher in the distribution areas of cedar and cypress (Japan Forestry Agency, 2013). We do not have enough E data for cedar and cypress plantations recorded 506 in regions where temperature is lower, except data for IK. The response of  $G_c$  to 507 temperature for stands in this region might be different from that observed in this study. 508 509 We thus recommend testing the applicability of the model using data derived for sites 510 located in this region. Furthermore, it is preferable to test the model applicability using data for sites with various age classes. Ages of the sites (Table 1) ranged between 19 and 511 512 99 years, which almost covers the age range for most Japanese cedar and cypress plantations in Japan (Japan Forestry Agency, 2013). However, many of the sites fall in 513 the range of 40–60 years. Thus, applicability of our model has been tested sufficiently for 514 stands within this age class, but has not been fully tested outside the age class. A major 515 516 portion of Japanese cedar and cypress plantations in Japan falls in this age class, suggesting the practical usefulness of the model. On the other hand, the portion of older 517 518 stands is expected to increase, because forestry in Japan has stagnated and cedar and cypress plantations have not been actively harvested (Komatsu et al., 2010). Therefore, 519 we recommend testing the applicability of the model to stands outside the age class. 520

521

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531

## 532 Appendix. A. Errors in interception evaporation estimates

The interception evaporation model developed by Komatsu et al. (2007a) and its 533 534 revised form developed by Komatsu et al. (submitted) typically have an error of 4% of incident rainfall. The period with T no less than 15  $^{\circ}$ C is typically April–October in 535 regions where Japanese cedar and cypress are distributed and incident rainfall during the 536 period is typically 1200 mm (National Astronomical Observatory, 2013). Annual rainfall 537 538 in the regions is typically 1800 mm. Thus, the typical error in interception evaporation estimates for the period with T no less than 15 °C would be 48 mm, which was equivalent 539 to 0.22 mm day<sup>-1</sup>. The typical error for a year would be 72 mm, which was equivalent to 540  $0.20 \text{ mm day}^{-1}$ . 541

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#### Appendix. B. Methods of preparing meteorological inputs 543

544	The model developed in this study requires daytime (6 a.m6 p.m.) mean solar
545	radiation ( $R$ ), temperature ( $T$ ), and vapor pressure deficit ( $D$ ) as meteorological inputs. $R$
546	data are not always available throughout Japan (Shinohara et al., 2007). It is often the
547	case that only daily maximum $(T_x)$ and minimum temperatures $(T_n)$ are available for
548	historical data (The University of Tokyo Forests, 2014).
549	The former problem can be solved by substituting $R$ for the target area by data
550	for solar radiation recorded at a meteorological observatory in surrounding areas. Data

for daily solar radiation are recorded at main meteorological observatories in Japan (Japan Meteorological Agency, 2014). R in units of W  $m^{-2}$  can be approximated by daily solar 552 radiation in units of MJ  $m^{-2} day^{-1}$  multiplied by an index of 23.1, which converts the units. 553 E during January 1–December 31, 2008 calculated using the model with the input of R, 554 T, and D observed at OL is 234 mm. Here,  $G_{cref}$  is assumed to be 0.00322 m s<sup>-1</sup> on the 555 basis of observed A and the relationship between A and  $G_{cref}$  (Figure 6). E values during 556 the same period calculated with the input of solar radiation observed at meteorological 557

558 observatories at Fukuoka, Saga, Oita, and Hiroshima are 238, 234, 237, and 240 mm, respectively. These meteorological observatories are located 15, 50, 110, 200 km from 559 OL, respectively. All these values approximate the value calculated using R for OL. 560

Qualitatively, the same results are available for other sites. Thus, accurate estimates of R561

562	are not important for estimating $E$ using the model. This is attributed to $E$ being
563	insensitive to $R$ when $R$ is high (Figure 2b) and $E$ on days with high $R$ accounting for a
564	relatively large portion of annual <i>E</i> .
565	The latter problem (i.e., only $T_x$ and $T_n$ are available) can be solved making the
566	following assumptions. First, the diurnal trend in temperature is approximated using a
567	sine function minimized at 6 a.m. and maximized at 2 p.m. A similar approximation has
568	been commonly used to produce hourly temperature data from $T_x$ and $T_n$ (Campbell and
569	Norman, 1998). Under this assumption, $T$ (i.e., temperature during 6 a.m6 p.m.) is
570	analytically written as $T = (T_x + T_n)/2 + (T_x - T_n)/(3\pi)$ . Second, vapor pressure deficit at
571	6 a.m. is zero and daytime vapor pressure deficit is caused by a temperature rise during
572	the day. Note that the assumption of vapor pressure deficit being zero in the morning is
573	generally valid except in arid and semi-arid regions (Running et al., 1987; Kimball et al.,
574	1994). $D$ (i.e., the mean vapor pressure deficit during 6 a.m.–6 p.m.) is thus approximated
575	by $D = e_s(T) - e_s(T_n)$ , where $e_s$ is the saturation vapor pressure. E during January 1–
576	December 31, 2008 calculated using the model with the input of $R$ , $T$ , and $D$ observed at
577	OL is 234 mm. <i>E</i> during the same period calculated with the input of <i>R</i> , $T_x$ and $T_n$ observed
578	at OL is 225 mm, which approximates $E$ calculated with the input of $T$ and $D$ .
579	Qualitatively, the same results are available for other sites.

- 580 References
- 581 Adams, K.N., Fowler, A.M., 2006. Improving empirical relationships for predicting the
- effect of vegetation change on annual water yield. J. Hydrol. 321, 90–115.
- 583 Addington, R.N., Mitchell, R.J., Oren, R., Donovan, L.A., 2004. Stomatal sensitivity to
- vapor pressure deficit and its relationship to hydraulic conductance in *Pinus palustris*.
- 585 Tree Physiol. 24, 561–569.
- Bakan, D., 1966. The test of significance in psychological research. Psychol. Bull. 66,
  423–437.
- 588 Baldocchi, D., Meyers, T.P., 1998. On using eco-physiological, micrometeorological and
- biogeochemical theory to evaluate carbon dioxide, water vapor and trace gas fluxes
  over vegetation: a perspective. Agric. For. Meteorol. 90, 1–25.
- 591 Bosch, J.M., Hewlett, J.D., 1982. A review of catchment experiments to determine the
- effect of vegetation changes on water yield and evapotranspiration. J. Hydrol. 55, 3–
  23.
- 594 Bréda, N.J.J., 2003. Ground-based measurements of leaf area index: a review of methods,
- instruments and current controversies. J. Exp. Bot. 54, 2403–2417.
- Brown, A., Zhang, L., McMahon, T.A., Western, A.W., Vertessy, R.A., 2005. A review of
- 597 paired catchment studies for determining changes in water yield resulting from
- alterations in vegetation. J. Hydrol. 310, 28–61.
- 599 Campbell, G.S., Norman, J.M., 1998. An Introduction to Environmental Biophysics.

- 600 Springer-Verlag, New York.
- 601 Cienciala, E., Eckersten, H., Lindroth, A., Hallgren, J., 1994a. Simulated and measured
- water uptake by *Picea abies* under non-limiting soil water conditions. Agric. For.
- 603 Meteorol. 71, 147–164.
- 604 Cienciala, E., Lindroth, A., Cermak, J., Hallgren, J., Kucera, J., 1994b. The effects of
- water availability on transpiration, water potential and growth of *Picea abies* during a
- 606 growing season. J. Hydrol. 155, 57–71.
- 607 Cornish, P.M., Vertessy, R.A., 2001. Forest age-induced changes in evapotranspiration
- and water yield in a eucalypt forest. J. Hydrol. 242, 43–63.
- Dawson, T.E., Burgess, S.S.O., Tu, K.P., Oliveira, R.S., Santiago, L.S., Fisher, J.B.,
- 610 Simonin, K.A., Ambrose, A.R., 2007. Nighttime transpiration in woody plants from
- contrasting ecosystems Tree Physiol. 27, 561–575.
- Diadonis, P., Efron, B., 1983. Computer-intensive methods in statistics. Sci. Am. 248,
  116–131.
- Domec, J.C., Ogée, J., Noormets, A., Jouangy, J., Gavazzi, M., Treasure, E., Sun, G.,
- 615 McNulty, S.G., King, J.S., 2012. Interactive effects of nocturnal transpiration and
- climate change on the root hydraulic redistribution and carbon and water budgets of
- southern United States pine plantations. Tree Physiol. 32, 707–723.
- Dung, B.X., Gomi, T., Miyata, S., Sidle, R.C., Kosugi, K., Onda, Y., 2012. Runoff
- 619 responses to forest thinning at plot and catchment scales in a headwater catchment

draining Japanese cypress forest. J. Hydrol. 444–445, 51–62.

- Efron, B., 1979. Bootstrap methods: another look at the Jacknife. Ann. Stat. 7, 1–26.
- Efron, B., 1987. Better bootstrap confidence intervals. J. Amer. Statis. Assoc. 82, 171–
- 623 185.
- Ewers, B.E., Oren, R., 2000. Analyses of assumptions and errors in the calculation of
  stomatal conductance from sap flux measurements. Tree Physiol. 20, 579–589.
- Ewers, B.E., Mackay, D.S., Tang, J., Bolstad, P.V., Samanta, S., 2008. Intercomparison of
- 627 sugar maple (*Acer saccharum* Marsh.) stand transpiration responses to environmental
- 628 conditions from the Western Great Lakes Region of the United States. Agric. For.
- 629 Meteorol. 148, 231–246.
- 630 Ewers, B.E., Bond-Lamberty, B., Machay, D.S., 2011. Consequences of stand age and
- 631 species' functional trait changes on ecosystem water use of forests, in Meinzer, F.C.,
- Lachenbruch, B., Dawson, T.E. (Eds.), Size- and Age-Related Changes in Tree
- 633 Structure and Function. Springer, Dordrecht, pp. 481–505.
- Fisher, J.B., Baldocchi, D.D., Misson, L., Dawson, T.E., Goldstein, A.H., 2007. What the
- towers don't see at night: nocturnal sap flow in trees and shrubs at two AmeriFlux sites
- 636 in California. Tree Physiol. 27, 597–610.
- 637 Forrester, D.I., Collopy, J.J., Morris, J.M., 2010. Transpiration along an age series of
- 638 *Eucalyptus globulus* plantations in southeastern Australia. Forest Ecol. Manage. 259,

639 1754–1760.

- Fox, J., 2008. Applied Regression Analysis and Generalized Linear Models. Sage
  Publications, Thousand Oaks, Carfornia.
- 642 Franzer, G.W., Ganha, C.D., Lerzman, K.P., 1999. Gap Light Analyzer (GLA) Version
- 643 2.0: Image software to Extract Canopy Structure and Gap Light Transmission Indices
- 644 from True-colour Fisheye Photographs. Simon Franser University, British Columbia,
- and the Institute of Ecosystem Studies, Millbrook, NewYork, pp. 1–36.
- 646 Fujii, H., 2005. Practical Methods of Data Analysis for Engineers. Tokyo-Kagaku-Dojin,
- 647 Tokyo.
- Granier, A., 1987. Evaluation of transpiration in a Douglas-fir stand by means of sap flow
- measurements. Tree Physiol. 3, 309–320.
- 650 Granier, A., Bréda, N., 1996. Modelling canopy conductance and stand transpiration of
- an oak forest from sap flow measurements. Ann. Sci. For. 53, 537–546.
- 652 Granier, A., Loustau, D., Bréda, N., 2000a. A generic model of forest canopy conductance
- dependent on climate, soil water availability and leaf area index. Ann. For. Sci. 57,
  755–765.
- Granier, A., Biron, P., Lemoine, D., 2000b. Water balance, transpiration and canopy
- conductance in two beech stands. Agric. For. Meteorol. 100, 291–308.
- Hasegawa, K., Omi, H., Hiruma, Y., Kumagai, S., Yamamoto, R., Izumi, T., Matsuyama,
- 658 H., 2013. Estimation of leaf area index of Cryptomeria japonica using various
- 659 methods : A case study of Aso District, Kumamoto Prefecture. J. Geograph. 122, 875-

660 891.

661	Hattori, S.,	Chikaarashi, H.	, Takeuchi, N.	, 1982. Measurement	of the rainfall	l interception
	,,		,	, _, _,, _,, _		

- and its micrometeorological analysis in a Hinoki stand. Bull. FFPRI 318, 79–102.
- 663 Ichihashi, R., Komatsu, H., Kume, T., Onozawa, Y., Shinohara, Y., Tsuruta, K., Otsuki,
- 664 K., Stand-scale transpiration of two Moso bamboo stands with different culm densities.
- 665 Ecohydrol. In press.
- Japan Forestry Agency, 2013. White Paper for Forest and Forestry. Japan Forestry Agency,
- 667 Tokyo. (Available at: http://www.rinya.maff.go.jp/j/kikaku/hakusyo/index.html).
- Japan Forestry Agency, 2014. Data for Japanese cedar and cypress plantations.
   http://www.rinya.maff.go.jp/j/sin\_riyou/kafun/data.html.
- 670 Japan Meteorological Agency, 2014. Meteorological statistics.
  671 http://www.jma.go.jp/jma/menu/report.html.
- Jarvis, P.G., 1976. The interpretation of the variations in leaf water potential and stomatal
- conductance found in canopies in the field. Phil. Trans. R. Soc. Lond. B 273, 593–610.
- Jarvis, P.G., McNaughton, K.G., 1986. Stomatal control of transpiration: scaling up from
- leaf to region. Adv. Ecol. Res. 15, 1–49.
- 676 Kimball, J.S., Running, S.W., Nemani, R. 1994. An improved method for estimating
- surface humidity from daily minimum temperature. Agric. For. Meteorol. 85, 87–98.
- 678 Komatsu, H., 2004. A general method of parameterizing the big-leaf model to predict the
- dry-canopy evaporation rate of individual coniferous forest stands. Hydrol. Process.

680 18, 3019–3036.

- Komatsu, H., Kang, Y., Kume, T., Yoshifuji, N., Hotta, N., 2006. Transpiration from a
- 682 Cryptomeria japonica plantation, part 2: responses of canopy conductance to
- meteorological factors. Hydrol. Process. 20, 1321–1334.
- Komatsu, H., Tanaka, N., Kume, T., 2007a. Do coniferous forests evaporate more water
- than broad-leaved forests in Japan? J. Hydrol. 336, 361–375.
- 686 Komatsu, H., Hotta, N., Kume, T., 2007b. What is the best way to represent surface
- conductance for a range of vegetated sites? Hydrol. Process. 21, 1142–1147.
- 688 Komatsu, H., Katayama, A., Hirose, S., Kume, A., Higashi, N., Ogawa, S., Otsuki, K.,
- 689 2007c. Reduction in soil water availability and tree transpiration in a forest with
- 690 pedestrian trampling. Agric. Forest Meteorol. 146, 107–114.
- 691 Komatsu, H., Kume, T., Otsuki, K., 2007d. Contemporary role of catchment water
- balance data for forest evapotranspiration research. J. Jpn. For. Soc. 89, 346–359.
- 693 Komatsu, H., Kume, T., Otsuki, K., 2010. Water resource management in Japan-forest
- management or dam reservoirs? J. Environ. Manage. 91, 814–823.
- 695 Komatsu, H., Kume, T., Otsuki, K., 2011. Increasing annual runoff-broadleaf or
- 696 coniferous forests? Hydrol. Process. 25, 302–318.
- 697 Komatsu, H., Onozawa, Y., Kume, T., Tsuruta, K., Shinohara, Y., Otsuki, K., 2012.
- 698 Canopy conductance for a Moso bamboo (*Phyllostachys pubescens*) forest in western
- 699 Japan. Agric. For. Meteorol. 156, 111–120.

700	Komatsu, H., Shinohara, Y., Nogata, M., Tsuruta, K., Otsuki, K., 2013. Changes in canopy
701	transpiration due to thinning of a Cryptomeria japonica plantation. Hydrol. Res. Lett.
702	7, 60–65.
703	Komatsu, H., Shinohara, Y., Otsuki, K., Models to predict changes in annual runoff with
704	thinning and clearcutting of Japanese cedar and cypress plantations in Japan.
705	Submitted to Hydrol. Process.
706	Kosugi, Y., Kobashi, S., Shibata, S., 1995. Modeling stomatal conductance on leaves of
707	several temperate evergreen broad-leaved trees. J. Jap. Reveget. Tech. 20, 158–167.
708	Kosugi, Y., Shibata, S., Matsui, K., Kobashi, S., 1997. Differences between deciduous
709	and evergreen broad-leaved trees in the pattern of seasonal change of leaf-scale
710	photosynthetic net assimilation rate and transpiration rate. J. Jap. Reveget. Tech. 22,
711	205–215.
712	Kosugi, Y., Takanashi, S., Tanaka, H., Ohkubo, S., Tani, M., Yano, M., Katayama, T.,
713	2007. Evapotranspiration over a Japanese cypress forest. I. Eddy covariance fluxes
714	and surface conductance characteristics for 3 years. J. Hydrol. 337, 269-283.
715	Kubota, T., Tsuboyama, Y., Nobuhiro, T., Sawano, S., 2013. Change of evapotranspiration
716	due to stand thinning in the Hitachi Ohta Experimental Watershed. J. Jpn. For. Soc. 95,
717	37–41.
718	Kumagai, T., Aoki, S., Nagasawa, H., Mabuchi, T., Kubota, K., Inoue, S., Utsumi, Y. and

719 Otsuki, K., 2005. Effects of tree-to-tree and radial variations on sap flow estimates of

- transpiration in Japanese cedar. Agric. For. Meteorol. 135, 110–116.
- 721 Kumagai, T., Aoki, S., Shimizu, T., Otsuki, K., 2007. Sap flow estimates of stand
- transpiration at two slope positions in a Japanese cedar forest watershed. Tree Physiol.
- 723 27, 161–168.
- 724 Kumagai, T., Tateishi, M., Shimizu, T., Otsuki, K., 2008. Transpiration and canopy
- conductance at two slope positions in a Japanese cedar forest watershed. Agric. For.
- 726 Meteorol. 148, 1444–1455.
- 727 Kumagai, T., Tateishi, M., Miyazawa, Y., Kobayashi, M., Yoshifuji, N., Komatsu, H.,
- Shimizu, T., 2014. Estimation of annual forest evapotranspiration from a coniferous
- plantation watershed in Japan (1): Water use components in Japanese cedar stands. J.
- 730 Hydrol. 508, 66–76.
- 731 Kume, T., Tsuruta, K., Komatsu, H., Kumagai, T., Higashi, N., Shinohara, Y., Otsuki, K.,
- 2010. Effects of sample size on sap flux-based stand-scale transpiration estimates Tree
  Physiol. 30, 129–138.
- Kume, T., Tsuruta, K., Komatsu, H., Shinohara, Y., Katayama, A., Ide, J., Otsuki, K.,
- 735 Differences in sap flux based stand transpiration between upper and lower slope
- positions in a Japanese cypress plantation watershed. Submitted to J. Hydrol.
- 737 Laplace 2013. Study on Transpiration in a Taiwanese Moso Bamboo Forest using Sap
- Flow Measurement. Master thesis, National Taiwan University.
- Comparison Comparison

- in Scots. Pine. Ecol. Bull. (Stockholm) 32, 505–523.
- 741 Loustau, D., Berbigier, P., Roumagnac, P., Arruda-Pacheco, C., David, J.S., Ferreira, M.I.,
- 742 Pereira, J.S., Tavares, R., 1996. Transpiration of a 64-year-old maritime pine stand in
- Portugal. 1. Seasonal course of water flux through maritime pine. Oecologia 107, 33–
- 744 42.
- Macfarlane, C., Bond, C., White, D.A., Grigg, A.H., Ogdena, G.N., Silberstein, R., 2010.
- Transpiration and hydraulic traits of old and regrowth eucalypt forest in southwestern
- 747 Australia. Forest Ecol. Manage. 260, 96–105.
- 748 Mackay, D.S., Ewers, B.E., Loranty, M.M., Kruger, E.L., 2010. On the representativeness
- of plot size and location for scaling transpiration from trees to a stand. J. Geophys.
- 750 Res. 115, G02016.
- 751 McDowell, N.G., Fisher, R.A., Xu, C., Domec, J.C., Hölttä, T., Mackay, D.S., Sperry, J.S.,
- 752 Boutz, A., Dickman, L., Gehres, N., Limousin, J.M., Macalady, A., Martínez-Vilata,
- J., Mencuccini, M., Plaut, J.A., Ogée, J., Pangle, R.E., Rasse, D.P., Ryan, M.G.,
- Sevanto, S., Waring, R.H., Williams, A.P., Yepez, E.A., Pockman, W.T., 2013.
- 755 Evaluating theories of drought-induced vegetation mortality using a multimodel-
- experiment framework. New Phytol. 200, 304–321.
- 757 McNaughton, K.G., Black, T.A., 1973. A study of evapotranspiration from a Douglas fir
- forest using the energy balance approach. Water Resour. Res. 9, 1579–1590.
- 759 National Astronomical Observatory, 2013. Chronological Scientific Tables 2014.

760 Maruzen, Tokyo.

- 761 Nuzzo, R., 2014. Statistical errors. Nature 506, 150–152.
- 762 Oishi, A.C., Oren, R., Stoy, P.C., 2008. Estimating components of forest
- revapotranspiration: A footprint approach for scaling sap flux measurements. Agric.
- For. Meteorol. 248, 1719–1732.
- Oren, R., Phillips, N., Katul, G., Ewers, B.E., Pataki, D.E., 1998. Scaling xylem sap flux
- and soil water balance and calculating variance: a method for partitioning water flux
- 767 in forests. Ann. Sci. For. 55, 191–216.
- 768 Oren, R., Sperry, J.S., Katul, G.G., Pataki, D.E., Ewers, B.E., Phillips, N., Schäfer, K.V.R.,
- <sup>769</sup> 1999. Survey and synthesis of intra- and interspecific variation in stomatal sensitivity

to vapour pressure deficit. Plant Cell Environ. 22, 1515–1526.

- Peck, R., Devore, J.L., 2005. Statistics: the Exploration and Analysis of Data, Second
- 772Edition. Brooks/Cole, Boston, Massachusetts.
- Phillips, N., Oren, R., 1998. A comparison of daily representations of canopy conductance
- based on two conditional time averaging methods and the dependence of daily
- conductance on environmental factors. Ann. Sci. For. 55, 217–235.
- Running, S.W., Nemani, R.R. and Hungerford, R.D., 1987. Extrapolation of synoptic
- meteorological data in mountainous terrain, and its use for simulating forest
  evapotranspiration. Can. J. For. Res. 17, 472–483.
- 779 Scott, D.F., Lesch, W., 1997. Streamflow responses to afforestation with Eucalyptus

- *grandis* and *Pinus petula* and to felling in the Mokobulaan experimental catchments,
- 781 South Africa. J. Hydrol. 199, 360–377.
- 782 Shinohara, Y., Komatsu, H., Otsuki, K., 2007. A method for estimating global solar
- radiation from daily maximum and minimum temperatures: Its applicability to Japan.
- 784 J. Jpn. Soc. Hydrol. Water Resour. 20, 462–469.
- 785 Shinohara, Y., Tsuruta, K., Ogura, A., Noto, F., Komatsu, H., Otsuki, K., Maruyama, T.,
- 786 2013a. Azimuthal and radial variations in sap flux density and effects on stand-scale
- transpiration estimates in a Japanese cedar forest. Tree Physiol. 33, 550–558.
- 788 Shinohara, Y., Tsuruta, K., Kume, T., Otsuki, K., 2013b. An overview of stand-scale
- transpiration measurements using the sap flow technique for evaluating the effects of

forest management practices. J. Jpn. For. Soc. 95, 321–331.

- Stewart, J.B., Thom, A.S., 1973. Energy budgets in pine forest. Quart. J. Roy. Met. Soc.
  99, 154–170.
- Sun, X., Onda, Y., Kato, H., Otsuki, K., Gomi, T., 2014. Partitioning of the total
  evapotranspiration in a Japanese cypress plantation during the growing season.
  Ecohydrol. 7, 1042–1053.
- 796 Suzuki, M., 1980. Evapotranspiratoin from a small catchment in hilly mountains (I)
- 797 Seasonal variations in evapotranspiration, rainfall interception and transpiration. J. Jpn.
- 798 For. Soc. 62, 46–53.
- 799 Tanaka, K., Tanaka, H., Nakamura, A., Ohte, N., Kobashi, S., 1996. Conductance at a

- community level and characteristics of CO<sub>2</sub> exchange in hinoki (*Chamaecyparis obtusa*) stand. J. Jpn. For. Soc. 78, 266–272.
- 802 Tanaka, K., Kosugi, Y., Nakamura, A., 2002. Impact of leaf physiological characteristics
- 803 on seasonal variation in CO<sub>2</sub>, latent and sensible heat exchanges over a tree plantation.
- 804 Agric. For. Meteorol. 114, 103–122.
- Tanaka, N., Kuraji, K., Shiraki, K., Suzuki, Y., Suzuki, M., Ohta, T., Suzuki, M., 2005.
- 806 Throughfall, stemflow and rainfall interception at mature *Cryptomeria japonica* and
- 807 Chamaecyparis obtusa stands in Fukuroyamasawa watershed. Bull. Tokyo Univ.
- 808 Forest 113, 197–240.
- 809 The University of Tokyo Forests, 2014. Meteorological data. http://www.uf.a.u-810 tokyo.ac.jp/eri/public.html.
- 811 Thompson, B., 1996. AERA editorial policies regarding statistical significance testing:
- three suggested reforms. Educational Res. 25, 26–30.
- Thompson, B., 1998. Statistical significance and effect size reporting: Portrait of a
  possible future. Res. Sch. 5, 33–38.
- Tsuruta, K., Kume, T., Komatsu, H., Higashi, N., Kumagai, T., Otsuki, K., 2008.
- 816 Relationship between tree height and transpiration for individual Japanese Cypress
- 817 (*Chamaecyparis obtusa*). J. Jpn. Soc. Hydrol. Water Resour. 21, 414–422.
- Tsuruta, K., Komatsu, H., Shinohara, Y., Kume, T., Ichihashi, R., Otsuki, K., 2011.
- Allometric equations between stem diameter and sapwood area of Japanese cedar and

820

Japanese cypress for stand transpiration estimates using sap flow measurement. J. Jpn.

- 821 Soc. Hydrol. Water Resour. 24, 261–270.
- Tsuruta, K., Nogata, M., Shinohara, Y., Komatsu, H., Otsuki, K., 2014. The correction
- coefficient for leaf area index measurement based on the optical method in a Japanese

824 cedar (*Cryptomeria japonica*) forest. Bull. Kyushu Univ. For. Accepted.

- 825 Tsuruta, K., Komatsu, H., Kume, T., Shinohara, Y., Otsuki, K., Canopy transpiration in
- two Japanese cypress forests with contrasting structures. Submitted to J. For. Res.
- Xiang Y, Tateishi M, Saito T, Otsuki K, Kasahara T., Changes in canopy transpiration of
- Japanese cypress and Japanese cedar plantations due to selective thinning. Submittedto Hydrol. Process.
- 830 Yamanoi, K., Ohtani, Y., 1992. Eddy correlation measurements of energy budget and
- characteristics of evapotranspiration above a hinoki stand. J. Jpn. For. Soc. 74, 221–
  228.
- Zwieniecki, M.A., Holbrook, N.M., 1998. Diurnal variation in xylem hydraulic
  conductivity in white ash (*Fraxinus americana* L.), red maple (*Acer rubrum* L.) and
- red spruce (*Picea rubens* Sarg.). Plant, Cell Environ. 21, 1173–1180.

836 Figure captions

- Figure 1. Relationships (a) between vapor pressure deficit (*D*) and canopy conductance
- 838 ( $G_c$ ), (b) between solar radiation (R) and  $G_c$  divided by  $G_{cref} f(D)$ , and (c) between
- temperature (*T*) and  $G_c$  divided by  $G_{cref}f(D)g(R)$  for the OL site. Solid lines in Figures 1a,
- 1b, and 1c indicate f(D), g(R), and h(T) for OL, respectively.  $G_{cref}$  is the reference value
- of canopy conductance. f(D), g(R), and h(T) are functions expressing the responses of  $G_c$
- 842 to *D*, *R*, and *T*.

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Figure 2. Functions expressing the responses of canopy conductance to (a) vapor pressure deficit (f(D)), (b) solar radiation (g(R)), and (c) temperature (h(T)) for each site.

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Figure 3. Relationships of the mean transpiration (E) for days with no rain and 847 temperature being no less than 15 °C (i.e., corresponding to a growing season) calculated 848 using (a) P1G, (b) P2G, and (c) P3G with the mean observed E. Here, P1G, P2G, and 849 P3G respectively use the mean values of  $G_{cref}$ ,  $\beta$ , and  $\delta$  among the sites,  $G_{cref}$  determined 850 for each site and the mean values of  $\beta$  and  $\delta$  among the sites, and  $G_{cref}$ ,  $\beta$ , and  $\delta$  determined 851 852 for each site. The solid line indicates the 1:1 relationship. Dotted lines indicate conditions that errors in *E* estimates equal potential errors in observed *E*. Error bars indicate standard 853 854 deviations.

855

Figure 4. Relationships of the mean transpiration (E) for days with no rain calculated 856 using (a) P1W, (b) P2W, and (c) P3W with the mean observed E for sites having data 857 recorded in winter as well as those recorded in a growing season. Here, P1W, P2W, and 858 P3W respectively use the mean values of  $G_{cref}$ ,  $\beta$ ,  $\delta$ ,  $\varepsilon$ , and  $\zeta$  among the sites,  $G_{cref}$ 859 determined for each site and the mean values of  $\beta$ ,  $\delta$ ,  $\varepsilon$ , and  $\zeta$  among the sites, and  $G_{cref}$ , 860  $\beta$ ,  $\delta$ ,  $\varepsilon$ , and  $\zeta$  determined for each site. The solid line indicates the 1:1 relationship. Dotted 861 lines indicate conditions that errors in *E* estimates equal potential errors in observed *E*. 862 863 Error bars indicate standard deviations. 864

Figure 5. Time series of daily transpiration (*E*) calculated using P1W, P2W, and P3W and
observed *E* for OU. Lines for P2W and P3W overlap each other. *E* observed on rainy days
is not plotted in this figure.

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Figure 6. Relationship between the sapwood area at a stand scale (*A*) and the reference value of canopy conductance ( $G_{cref}$ ). The regression line for the relationship between *A* and  $G_{cref}$ , determined using the least-squares method, is written as y = 0.157 x.

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Figure 7. Normal probability plot for the difference between estimated and observed reference values of canopy conductance ( $G_{cref}$ ). The solid line is the regression line, determined using the least-squares method, for all data except those for SK and IK. 876

Figure 8. Comparison between estimated and observed sapwood areas on a stand scale

- 878 (*A*). The solid line indicates the 1:1 relationship.
- 879
- Figure 9. Time series of daily transpiration (E) calculated using P2W by omitting
- functions expressing the responses of  $G_c$  to (a) vapor pressure deficit (f(D)), (b) solar
- radiation (g(R)), and (c) temperature (h(T)).







Figure 1



Figure 2





Figure 4



 $\bigcirc$ 

Figure 5





Estimated minus observed  $G_{cref}$  (mm s<sup>-1</sup>)



 $\bigcirc$ 



Site	Location	Р	T	Age	N	d <sub>m</sub>	$L (m^2)$	A	Plot	Number	Period	Reference
*1		(mm)	(°C)	(yr)	(stems	(cm)	$m^{-2})^{*2}$	(m <sup>2</sup>	area	of the		
		( )		0 /	ha <sup>-1</sup> )	· /	,	ha <sup>-1</sup> )	(m <sup>2</sup> )	trees in		
										the plot		
Cedar												
IK	36°N, 137°E	2814	12.7	55	600	48.4	3.2	36.6	300	18 (9) <sup>*3</sup>	May 15, 2010–May 24, 2011	Shinohara et al. (2013a)
YB	34°N, 131°E	1790	17.2	39	1100	28.9	4.7	22.4	100	11 (11)	Aug 1-Sept 31, 2010	Komatsu et al. (2013)
YA	34°N, 131°E	1790	17.2	40	600	31.6	2.3	14.1	100	6 (6)	Aug 1-Sept 31, 2011	Komatsu et al. (2013)
IS	34°N, 131°E	1807	15.9	43	658	28.2	0.8	21.5	213	14 (11)	May 22, 2012–Feb 28, 2013	Xiang et al. (submitted)
IR	34°N, 131°E	1790	17.2	60	1400	30.3	3.3	28.4	200	28 (14)	July 1-Aug 31, 2010	Ichihashi et al. (in press)
KL	33°N, 131°E	2150	15.0	50	904	40.3	5.7	46.0	321	29 (15)	Mar 3, 2007–July 3, 2008	Kumagai et al. (2007)
KU	33°N, 131°E	2150	15.0	50	1575	23.8	5.4	36.3	318	50 (23)	Mar 3, 2007–July 3, 2008	Kumagai et al. (2007)
KS	33°N, 131°E	2150	15.0	50	1330	26.6	5.9	37.3	203	27 (19)	Mar 3, 2007–July 3, 2008	Kumagai et al. (2014)
XT	24°N, 121°E	2635	17.0	60	625	39.0	2.6	18.4	400	25 (15)	Apr 14, 2012–Mar 19, 2013	Laplace (2013)
Cypres	55											
IH	34°N, 131°E	1807	15.9	43	863	20.2	0.8	8.8	197	17 (17)	May 22, 2012–Feb 28, 2013	Xiang et al. (submitted)
OL	34°N, 131°E	1790	17.2	49	1450	21.0	3.2	20.4	400	58 (14)	Jan 1-Dec 31, 2008	Kume et al. (2010)
OU	34°N, 131°E	1790	17.2	49	1700	14.9		16.8	100	58 (17)	Jan 1-Dec 31, 2008	Kume et al. (submitted)
SK	34°N, 131°E	1790	17.2	19	2100	13.5	4.8	20.4	100	21 (10)	Apr 1-Aug 15, 2009	Tsuruta et al. (submitted)
HW	34°N, 131°E	1790	17.2	99	350	44.6	3.1	15.2	200	7 (7)	Apr 1–Aug 31, 2009	Tsuruta et al. (submitted)

Table 1. Location, meteorology (the mean annual rainfall P and temperature T), structure (stem density N, the mean diameter at breast height  $d_m$ , leaf area index L, and sapwood area A), and description of sap-flux measurements for the plantation sites

<sup>\*1</sup> The original names of the sites are given as follows. IK: Ishikawa-ken Forest Experimental Station, YB: Yamanokami site (before thinning), YA: Yamanokami site (after thinning), IS: Yayama Experimental Catchment (cedar) in Iizuka, IR: Ichirinpan Plot in the Kasuya Research Forest, KU: UP of Kahoku Experimental Watershed, KL: LP of Kahoku Experimental Watershed, KS: SP of Kahoku Experimental Watershed, XT: Xitou Experimental Forest, IH: Yayama Experimental Catchment (cypress) in Iizuka, OL: Riparian Plot of Ochozu Experimental Watershed, OU: Ridge Plot of Ochozu Experimental Watershed, SK: Sakuta Plot in Kasuya Research Forest, HW: Hiawada Plot in Kasuya Research Forest <sup>\*2</sup> *L* for all sites except OL and IK was measured using a plant-canopy analyzer (LAI-2000, Li-Cor Inc., Lincoln, Nebraska, USA). *L* for OL and IK was measured using a digital non-

spherical color photograph and Gap Light Analyzer software (Franzer et al., 1999).

\*3 Numerals in the parentheses indicate the number of trees in which sensors were installed.

Site	$G_{cref} (\mathrm{mm} \mathrm{s}^{-1})$	$\beta \inf f(D)$	$R^2$	$\delta \inf g(R)$	$R^2$	$\varepsilon$ (°C) in $h(T)$	$\zeta$ (°C) in <i>h</i> ( <i>T</i> )	$R^2$
Cedar								
IK	9.77	0.750	0.782	0.287	0.389	7.75	0.313	0.469
YB	3.46	0.441	0.493	$0.0893^{*1}$	0.00793			
YA	1.89	0.427	0.365	$0.0625^{*1}$	0.00247			
IS	2.84	0.569	0.580	0.238	0.339	20.7	-48.8	0.164
IR	2.81	0.351	0.214	0.288	0.219			
KL	6.07	0.783	0.605	0.347	0.316	12.3	0.065	0.587
KU	4.22	0.749	0.438	0.410	0.240	15.1	-5.22	0.343
KS	5.19	0.667	0.628	0.378	0.392	14.8	-2.49	0.544
XT	2.84	0.412	0.213	0.308	0.310			
Cypress								
IH	1.76	0.472	0.622	$0.000^{*1}$	-0.0329	21.0	-2.84	0.740
OL	3.22	0.453	0.628	0.285	0.330	18.2	-5.19	0.636
OU	1.86	0.591	0.738	0.417	0.561	17.2	-5.78	0.584
SK	5.65	0.468	0.412	0.432	0.252			
HW	2.81	0.688	0.436	0.441	0.270			
Mean	3.89	0.556		0.285		15.9	-8.74	
Standard deviation	2.19	0.146		0.142		4.43	16.4	

Table 2. Parameter values optimized for each site .

<sup>\*2</sup> A small  $\delta$  value does not suggest that there was no effect of solar radiation on transpiration, but suggests that the effect was not detectable by an analysis on a daily time scale.

Site	P1W		P2W		P3W	
	Slope <sup>*1</sup>	$R^2$	Slope	$R^{2 * 2}$	Slope	$R^2$
Cedar						
IK	2.43	0.329 *3	0.946	0.329	1.02	0.576
IS	0.776	0.867	1.06	0.867	1.03	0.874
KL	1.50	0.845	0.964	0.845	1.04	0.818
KU	1.05	0.757	0.970	0.757	1.06	0.726
KS	1.32	0.848	0.988	0.848	1.05	0.845
Cypress						
IH	0.510	0.927	1.13	0.927	1.02	0.937
OL	0.879	0.843	1.06	0.843	1.04	0.883
OU	0.466	0.891	0.974	0.891	1.05	0.911

Table 3. Slope and coefficient of determination  $(R^2)$  for the relationship between calculated and observed transpiration.

 \*<sup>1</sup> The intercept of the regression equation was assumed to be zero.
 \*<sup>2</sup> Determination coefficients for P2W are identical to those for P1W, because the responses of canopy conductance to meteorological factors for P2W are same as those for P1W.

\*3 Low determination coefficients for IK were primarily caused by the small number of data with low *E* due to rejection of data recorded on rainy days. When including data recorded on rainy days in the analysis, the determination coefficients were 0.829, 0.829, and 0.858 for P1W, P2W, and P3W, respectively.