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Title: Microgeographic adaptation and the effect of pollen flow on the adaptive potential of a temperate tree species.

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1 Summary

In species with long-distance dispersal capacities and inhabiting a large ecological niche, local selection and gene flow are expected to be major evolutionary forces affecting the genetic
adaptation of natural populations. Yet, in species such as trees, evidence of microgeographic
adaptation and the quantitative assessment of the impact of gene flow on adaptive genetic variation are still limited.

7 • Here, we used extensive genetic and phenotypic data from European beech seedlings col8 lected along an elevation gradient, and grown in a common garden, to study the signature of
9 selection on the divergence of eleven potentially adaptive traits, and to assess the role of gene
10 flow in resupplying adaptive genetic variation.

• We found a significant signal of adaptive differentiation among plots separated by less than one kilometre, with selection acting on growth and phenological traits. Consistent with the oretical expectations, our results suggest that pollen dispersal contributes to increase genetic diversity for these locally differentiated traits.

Our results thus highlight that local selection is an important evolutionary force in natural
tree populations and suggest that management interventions to facilitate movement of gametes
along short ecological gradients would boost genetic diversity of individual tree populations,
and enhance their adaptive potential to rapidly changing environments.

19

Key words: Adaptive divergence; Climate change; Common garden ; Elevation gradient; *Fa- gus sylvatica*; Genetic diversity; Natural selection; Pollen dispersal

22 Introduction

40

23 Recent interest in microgeographic adaptation, *i.e.* adaptation at spatial scales compatible with substantial amounts of gene dispersal, suggests we must reconsider the scale at which evolution 24 occurs (Richardson et al., 2014). At this fine spatial scale, a major question that remains to be 25 answered is whether gene flow is constraining or facilitating local adaptation. Too much gene 26 flow would overwhelm the effects of natural selection that drive adaptive genetic divergence 27 among populations (Lenormand, 2002) and decrease the adaptation of populations to their lo-28 cal environment (migration load). Conversely, gene flow can also resupply the genetic diversity 29 eroded by genetic drift and selection, and thus facilitate future evolutionary responses (Pease 30 et al., 1989; Barton, 2001; Lenormand, 2002; Bridle et al., 2010; Polechova & Barton, 2015). 31 32 In many taxa, and especially long-lived organisms, gene flow is expected to be the main process maintaining the standing genetic variation within natural populations (Barton, 2001; Le Corre 33 & Kremer, 2003; Bridle et al., 2010). In the particular case of environments varying through 34 both space and time, gene flow could also spread alleles pre-adapted to future environmental 35 conditions (Davis & Shaw, 2001; Aitken et al., 2008; Kremer et al., 2012; Fitzpatrick & Reid, 36 37 2019). Although theoretical developments on this topic have been rich and provided interest-38 ing hypotheses to test in nature, empirical studies on the role and importance of gene flow in adaptation are still lacking (but see Fitzpatrick et al. 2015, 2016). 39

41 Increasing evidence of microgeographic adaptation in nature provides strong support for 42 the idea that gene flow rarely prevents the emergence of local adaptation or adaptive divergence (Anderson et al. 2015; Eckert et al. 2015; Fitzpatrick et al. 2015; Moody et al. 2015; Peterson 43 et al. 2016; for a review see Richardson et al. 2014). This evidence strengthens the conclusions 44 of studies at larger spatial scales, showing that geographic or genetic distances among popula-45 tions are not correlated with the level of local adaptation or adaptive genetic divergence (McKay 46 & Latta, 2002; Leimu & Fischer, 2008; Hereford, 2009). These results suggest that selective 47 pressures can be strong enough to counterbalance the effects of gene flow, and to shape genetic 48 variation at a fine spatial scale. However, evidence of microgeographic adaptation is mainly 49 reported for certain taxa (e.g. fish, Fitzpatrick et al. 2015; Moody et al. 2015; and annual or 50 perennial herbs Anderson et al. 2015; Peterson et al. 2016), while others, such as temperate 51

52 tree species, are underrepresented. In trees, there is a long tradition of large-scale adaptation studies, through provenance trials, following a quantitative genetic approach (Savolainen et al., 53 2007; Alberto et al., 2013). The development of genome-wide data and statistical methods has 54 considerably changed this research field, and evidence of adaptations at regional and landscape 55 56 scales are now flourishing (e.g. Csillery et al. 2014; Pluess et al. 2016). Yet, the adaptive response of phenotypic traits at the local scale is still poorly understood (but see Brousseau et al. 57 2013; Eckert et al. 2015). Methodological limitations that previously hampered the robust 58 59 testing of local phenotypic adaptation have been largely overcome (Ovaskainen et al., 2011), permitting more accurate assessments of the scale at which phenotypic adaptations occur. 60

61

62 Empirical evidence of beneficial effects of gene flow on adaptive potential in nature are still rare, and are often assessed in laboratory settings (e.g. Swindell & Bouzat 2006). Some experi-63 mental studies on inbred plant populations have demonstrated the beneficial effect of gene flow 64 on the reduction of homozygosity and reintroduction of variation for fixed deleterious alleles 65 (e.g. Costa e Silva et al. 2014; Bontrager & Angert 2019). However, for outbred species, the 66 role of gene flow on the maintenance of genetic diversity has been difficult to demonstrate, es-67 pecially when relying on natural crosses. This effect of gene flow is thought to be particularly 68 important when the populations exchanging genes are strongly genetically differentiated (Pease 69 70 et al., 1989; Barton, 2001). Long-distance dispersal events could therefore play a major role in resupplying the genetic variation of populations and favoring the spread of beneficial alle-71 les (Savolainen et al., 2007; Kremer et al., 2012). In wind-pollinated species (anemophilous), 72 these long-distance dispersal events are common, with mean pollen dispersal distances that can 73 easily reach hundreds of meters, with rare events of spread over tens of kilometres in some 74 species (Austerlitz et al., 2004; Petit & Hampe, 2006; Ashley, 2010; Kremer et al., 2012). 75 Pollen dispersal may thus be highly efficient for spreading beneficial alleles between popula-76 tions (Kremer et al., 2012). Even though a pollen grain brings half as many immigrating alleles 77 as does a seed, gene flow by pollen or seed is expected to have similar consequences on popu-78 79 lation's adaptation and migration load (Lopez et al., 2008).

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81 Wind-pollinated tree species are therefore good study systems to look for empirical ev-82 idence of the beneficial effects of long-distance gene flow on adaptive capacities in nature.

However, to our knowledge, only two studies have looked for such evidence. In Pinus con-83 torta, Yeaman & Jarvis (2006) showed a positive correlation between the level of genetic vari-84 ance within populations for growth and the spatial heterogeneity of selection. This pattern is 85 consistent with the theoretical expectation that gene flow among differentiated populations is an 86 87 important source of genetic variation. Along a latitudinal gradient of genetically differentiated populations of *Pinus sylvestris*, Nilsson (1995) found that offspring harvested after one episode 88 89 of reproduction showed a shift in average genetic values as compared to their mothers. This 90 pattern is expected in the case of directional gene flow among locally differentiated populations (Figure 1). This effect of directional pollen dispersal on offspring's average genetic values for 91 92 adaptive traits could be particularly beneficial if the environmental conditions of the recipient population change toward the one of the donor population (Davis & Shaw, 2001; Aitken et al., 93 94 2008).

96 The aim of the present study is to investigate the effects of natural selection and longdistance gene flow on the adaptive potential of a major temperate tree species, the European 97 beech (Fagus sylvatica L.). This anemophilous species can disperse pollen at long-distance 98 (Piotti et al., 2011), a feature thought to explain the low level of neutral genetic differenti-99 100 ation among populations observed from the local to the European scale (Comps et al., 2001; 101 Buiteveld et al., 2007). Several phenotypic traits show significant genetic differentiation among beech provenances at the European scale (Robson et al., 2018), in particular budburst date 102 (Gomory & Paule, 2011; Kramer et al., 2017), height (Rose et al., 2009) and, to a lesser ex-103 tent, wood anatomy, hydraulic traits and physiological foliar traits (Knutzen et al., 2015; Hajek 104 et al., 2016; Stojnic et al., 2018). These patterns are usually interpreted as a long-term response 105 to selection, conferring better adaptation of beech populations to local climate and, most im-106 107 portantly, to summer drought. Our study aimed to test whether such patterns of phenotypic 108 adaptation also exist at a fine spatial scale.

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Here, we analysed the phenotypic and genotypic data of $\sim 2,300$ beech seedlings from 60 maternal families grown in a common garden. These open-pollinated families were collected from three natural beech plots, spreading along an elevation gradient at fine spatial scale (\sim 113 1.5 km in-between the most distant plots) at the southern, warm margin of this species' dis-

tribution. We focused on 11 potentially adaptive phenotypic traits, for which we previously 114 found a significant genetic component in phenotypic variation (Gauzere et al., 2016a). The 115 first question investigated was: (1) is there genetic divergence across elevations for the quan-116 titative traits of interest? We expected trees from the low-elevation, warmer and drier plot to 117 118 have evolved traits favouring adaptation to drought, and trees from the high-elevation, colder 119 plot to have evolved traits favouring adaptation to low temperatures. Our second question was 120 (2) is there a signature of local selection in the differentiation pattern of the trait studied. We 121 tested this by comparing traits differentiation to neutral differentiation shown by microsatellite markers. Our final question was: (3) does gene flow contribute to the additive variance of traits 122 123 under selection at the local scale? To that purpose, we used parentage analysis to identify the 124 offspring likely originating from local or distant pollen immigration events, and we tested how long-distance pollen dispersal affected the genetic values of offspring after one episode of re-125 126 production.

127

128 Material and methods

129 Species, sampling plots and experimental design

Fagus sylvatica L. (European beech) is a monoecious, anemophilous and predominantly out-130 crossed tree species (Merzeau et al., 1994; Wang, 2003). It is a shade-tolerant species requiring 131 132 well drained and moderately deep soils and relatively high humidity rates (Breda et al., 2006; Jump et al., 2006). Its distribution ranges from the northern Mediterranean regions to the south 133 of Scandinavia. On Mont-Ventoux, a mountain in the southeast of France, beech forests are lo-134 cated at the southern limit of their ecological range (Figure 2), in a mountainous Mediterranean 135 136 ecosystem. On the northern face of Mont-Ventoux, the beech forest ranges almost continuously from 750 to 1,700 m a.s.l.. This steep elevation gradient provides almost linear variation 137 in mean temperature and humidity with elevation (Davi et al., 2011). Three plots were defined 138 along this climatic gradient, at 1020 m (N1), 1140 m (N2) and 1340 m (N4) a.s.l (Table S1), 139 over a total distance of about 1.5 km. A previous study showed that these plots have high 140 pollen immigration rates (on average 53.3 %), moderate distances of pollen dispersal (on av-141

142 erage 45.7 m), and non-negligible long-distance pollen dispersal events (Gauzere *et al.*, 2013b).143

144 In August 2009, 20 open-pollinated families were sampled in each of the three plots (60 maternal progenies in total). Mother-trees were chosen for their high fertility and random loca-145 146 tion within the plot. In April 2010, the seeds were germinated and a subset of 46.4 seedlings on average per family were randomly planted in 25 blocks in a common garden experiment (with 147 1.8 seedlings per family in each block) at the State nursery of Aix-en-Provence (43°30'N 148 149 5°24'E). All seedlings were planted in independent pots of 1.2 L with sand substrate, fertilizer and regularly irrigated. The experiment ran for 3 years (from April 2010 to September 2013). 150 151 Details on the seed collection, germination protocol and progeny test design can be found in 152 Gauzere et al. (2016a).

All the potentially reproductive adults within the three plots, including the mother-trees, and an average of 23.9 offspring per family were genotyped at 13 microsatellites markers (690 adults and 1437 offspring in total; see Gauzere *et al.* 2013b for the genotyping details).

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158 Phenotypic measurements

159 A set of 11 different phenotypic traits were measured on the offspring grown in the common garden. These traits, already introduced in Gauzere et al. (2016a), are presented briefly be-160 161 low (see also Table 1). Height and diameter growths were measured between August 2010 and November 2011 (ΔH , ΔD). For the whole trial, phenological events were also recorded: 162 the date of budburst in 2011 and 2012 ($t_{b2\rightarrow3(2011)}$, $t_{b2\rightarrow3(2012)}$), the date of leaf senescence 163 in 2011 $(t_{s1\to 2(2011)})$ and the duration of the growing season in 2011 $(VD_{2011} = t_{s1\to 2(2011)})$ 164 165 $t_{b2\rightarrow3(2011)}$). At the end of the growing season in 2011, three light-exposed leaves were collected on the stem of each seedling to measure morphological and physiological traits. On fresh 166 leaves, we first measured the leaf area (LA), and put them to dry at 60 $^{\circ}$ C to then measure the 167 168 leaf dry mass (LM) of each seedling. The leaf mass area, related to the photosynthetic capacity and stomatal conductance of the plant (Reich *et al.*, 2003), was calculated as $LMA = \frac{LM}{LA}$. 169 Finally, for a subset of 1031 individuals (evenly sampled across families), we measured the 170 leaf carbon isotope composition ($\delta^{13}C$), as a surrogate for water use efficiency (Farquhar *et al.*, 171

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172 1989). The nitrogen content in the leaves (%*N*) was also estimated for this subset of individ-173 uals, as leaf nitrogen density is linked to the photosynthetic capacity of leaves (Kattge *et al.*, 174 2009). Note that three of these traits were transformed before analyses to limit departure from 175 linear model assumptions: ΔH with a logarithm transformation, and *LM* and *LA* with a square-176 root transformation (following Gauzere *et al.* 2016a).

178 Test for population genetic differentiation

To assess the genetic differentiation of quantitative traits among plots we ran a specific linear mixed model (called "animal model") that uses relatedness information from a pedigree to dissociate the contribution of genetic, population and environmental effects on the total phenotypic variance of traits. Thus, we analysed each trait Y measured in common garden using the following model, and focused particularly on the significance of the fixed plot effect P:

$$Y_{i,f,p,k} = \mu + P_i + B_f + M_p + a_k + \varepsilon_{i,f,p,k}$$

$$\tag{1}$$

184 with, as fixed effects, μ the intercept, P_i the effect of the plot of origin, B_f the effect of the 185 block and M_p the observer effect (only included for the analysis of the phenological traits). The 186 random terms are the additive genetic values of the individuals k, $\{a_k\} \sim N(0, AV_A)$ with A the 187 matrix of pairwise relatedness derived from a pedigree, and V_A is the additive genetic variance, 188 and the residuals, $\{\varepsilon_{i,f,p,k}\} \sim N(0, I_{de}V_R)$ with I_{de} the identity matrix and V_R the residual vari-189 ance.

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The (co)variance structure of the additive genetic effects was defined using a one-generation pedigree reconstructed in a previous study (Gauzere *et al.*, 2016a). Here, we did not model maternal effects because of the lack of power to accurately dissociate maternal and additive genetic effects (Gauzere *et al.*, 2016b). Model (1) was fitted in AsReml-R (Gilmour *et al.*, 2006). Using similar model, Gauzere *et al.* (2016a) previously showed that all traits presented moderate narrow-sense heritability, with h^2 ranging from 0.12 to 0.35.

198 Signature of natural selection accounting for genetic drift

We used the method developed by Ovaskainen *et al.* (2011) to detect signatures of natural selection in trait differentiation among populations. This method requires genotypic and the phenotypic data, that are first analysed under a neutral model that assumes that the genes coding for the trait are as divergent as the neutral markers ($Q_{st} = F_{st}$ case). This first model thus considers the population mean genetic additive values as normally distributed, with the covariance between pairs of population means being proportional to the pairwise average coancestry between these populations (*i.e.* assuming neutrality for the trait):

$$\mathbf{A}^{\mathbf{P}} \sim N(0, 2V_A \times \boldsymbol{\theta}^{\mathbf{P}}) \tag{2}$$

with $\mathbf{A}^{\mathbf{P}}$ the vector of the population-level effects, V_A the additive genetic variance of the trait, $\theta^{\mathbf{P}}$ the matrix of pairwise population-level coancestry coefficients. Note that the $\theta^{\mathbf{P}}$ matrix provides an estimation of *Fst*.

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The method then considers an animal model for each quantitative trait Y, that decomposes the additive value into a population and an individual effect. The population-level additive effect is determined by the mean genetic additive values in the parent population and the individual-level additive effects a_k are defined the same way as in model (1). The presence of selection on the trait is tested with the *S* statistics, which evaluates whether the realized pattern of population mean genetic additive values is likely under the neutral model (eq. 2).

217 To apply this method to our dataset, we first used the genotyping of the adult trees and the admixture F-model for neutral divergence among populations implemented in the R-package 218 RAFM, to estimate $\theta^{\mathbf{P}}$ and the index of neutral genetic differentiation, Fst (Karhunen & 219 220 Ovaskainen, 2012). We then analysed the traits measured in the common garden using the 221 same fixed effects as in model eq.(1) and the one-generation pedigree using the R-package DRIFTSEL, to estimate V_A , \mathbf{A}^P and S (Karhunen *et al.*, 2013). These Bayesian analyses were 222 repeated three times per trait to ensure convergence of the MCMC chains. We provide the 223 average posterior S estimates (the chains provided very similar S values; Table S2). We also 224 225 measured a low autocorrelation within the chains, indicating that the space of the parameter

226 values is efficiently explored.

227

228 Immigrant characterization based on molecular markers

229 Categorical assignments of paternity

We first tried to identify the father of the genotyped seedlings to characterize their immigration 230 231 status. To that purpose, we used the likelihood-based method of categorical paternity assign-232 ment implemented in CERVUS (Marshall et al., 1998) with allelic frequencies estimated from 233 the genotypes of all adult trees. All the reproductive adults from the three sampled plots were considered as candidate fathers, but only fathers within the mother's plot were assigned (see 234 below). We considered 0 % typing error, 100 % sampling of candidate fathers, a confidence 235 236 level of 95 %, and allowed selfing (following Gauzere et al. 2016a). These choices were made to favour assignments to genotyped fathers, despite the risk of type I errors (i.e. a wrong tree 237 is assigned while the true father is not sampled), knowing that quantitative genetic methods are 238 239 robust to pedigree errors (Charmantier & Réale, 2005; Gauzere et al., 2016b). We successfully 240 assigned the paternity for 45.8 % of the genotyped offspring. We considered that individuals 241 assigned to a father located within their mother's plot come from "local" pollen pool, while individuals with no compatible father were considered "immigrant". Non-categorised individuals 242 include individuals genotyped at less than 6 markers, offspring not assigned because of com-243 244 patibility with multiple fathers, and offspring assigned to a father outside their mother's plot. Indeed, given the large number of trees potentially contributing to median- and long-distance 245 mating events, we considered that our ability to retrieve the true father outside the mother's plot 246 was low and presented large expected type I error rate. 247

248

249 Probability of long-distance pollination

To refine the "immigrant" status, we also used the whole genetic dataset and a non-categorical paternity analysis to assess the likelihood for each genotyped seedlings to belong to longdistance immigrant pollen pool. To that purpose, we estimated the Mendelian transition probabilities (*T*) that each offspring has the observed genotype g_0 while originating from cross-

fertilization of female j with an ungenotyped male from a population with the allelic frequen-254 cies BAF: $T(g_0|g_i, BAF)$. We estimated this transition probability using (1) the global allelic 255 frequencies estimated from a larger-scale study on the whole north-face of the Mont-Ventoux 256 $(AF_{global}; Lander et al. 2011)$, and (2) the local allelic frequencies estimated from the genotyp-257 258 ing of all the reproductive trees in the three study plots (AF_{local}) . This allowed us to compute 259 the probabilities that each offspring has the observed genotype g_0 given that it originates from long-distance and local pollination event, $T(g_0|g_j, AF_{global})$ and $T(g_0|g_j, AF_{local})$ respectively. 260 261 From these two probabilities, we derived for each offspring the LOD-score of long-distance immigration vs. local immigration: 262

$$\mathbf{v}_{g_0} = \log\left(\frac{T(g_0|g_j, AF_{global})}{T(g_0|g_j, AF_{local})}\right) \tag{3}$$

263 $v_{g_0} > 0$ indicating that offspring g_0 is more likely originating from a long-distance pollina-264 tion event and $v_{g_0} < 0$ that it is more likely originating from a short-distance pollination event. 265 These transitions probabilities were estimated using the MEMMi model (Gauzere *et al.*, 2013b). 266

Impact of pollen immigration on quantitative traits

We explicitly tested whether (i) offspring originating from distant fathers presented more ge-268 269 netic variance for quantitative traits than offspring originating from local fathers, which is expected in the case of migration from genetically differentiated populations in an island model, 270 and whether (ii) offspring originating from distant fathers had shifted average genetic values 271 for quantitative traits as compared to the ones originating from local fathers, which is expected 272 in the case of directional pollen immigration from genetically differentiated populations along 273 the gradient (Figure 1). In both cases (i) and (ii), pollen dispersal would result in increasing 274 275 genetic variance and adaptive capacities. We ran two different models depending on how the immigration status was characterized (categorical or quantitative variable). 276

277

First, for each phenotypic trait *Y*, we looked for differences in the mean and variance of additive genetic values (a_k) between "local" and "immigrant" groups, using the following animal model:

$$Y_{i,f,p,k} = \mu + P_i + B_f + M_p + a_k + \varepsilon_{i,f,p,k}$$

$$\tag{4}$$

with the same fixed effects than in model (1), but here the (co)variance structure of the additive genetic effects depend on the immigration classes, with l = "local" and m = "immigrant": $\begin{bmatrix} V_{Al} & 0 \\ 0 & V_{Am} \end{bmatrix}$. Note that the effect of "local" and "immigrant" categories on the average trait values was tested and found to be never significant. This fixed effect is thus absent in the final model. From model (4), we thus estimated two variance components: V_{Al} and V_{Am} .

286

287 We then performed a random regression model to investigate the effect of long-distance pollen dispersal, v, on the variation in additive genetic variance. Random regression models 288 are commonly used in evolutionary biology and breeding to allow additive genetic effects (or 289 any random effect) to vary with a covariate (e.g. environment, age; Nussey et al. 2008). This 290 291 model defines a random intercept and slope at the additive genetic level, as deviation from the mean linear effect of v on Y, to model the variation of V_A with v as: $V_A(v) = V_A$ intercept + 292 v^2 . $V_{A \ slope} + 2.v.cov_A(intercept, slope)$. The effects of v on V_A was tested using the following 293 models: 294

$$Y_{i,f,p,k} = \mu + P_i + B_f + M_p + \alpha v_k + a_k + \varepsilon_{i,f,p,k}$$
(5)

$$Y_{i,f,p,k} = \mu + P_i + B_f + M_p + \alpha \nu_k + \tilde{a_k} + \varepsilon_{i,f,p,k}$$
(6)

where $\tilde{a}_k = a_k + \beta_k v_k$ describes the random intercept and slope of v on Y at the additive genetic level in model (6). In both models (5) and (6) the term αv_k is the fixed effect of the LOD-score of long-distance vs. local migration on the trait.

From the model (6), we estimated the variances of the random intercept and slope, and the covariance between them. Note that for each individual k, we only have one value of v and Y, but the model estimates these (co)variance components by comparing phenotypic similarities and the effects of v on Y between relatives.

302

The significance of the qualitative and quantitative effects of pollen immigration on V_A were tested by comparing the two hierarchical models (1) and (4), and the two models (5) and (6), with a log-likelihood ratio test. We performed these analyses with AsReml-R, considering only 306 the information about maternal relatednesses, as we (often) ignore the father's identity of "im-307 migrant" individuals.

308

309 Relatedness and Nep estimates

310 We used the genetic dataset to estimate the average paternal relatedness $(\bar{\rho_p})$ and effective 311 number of pollen donors ($Nep = 1/\bar{\rho_p}$) in the "local" and "immigrant" pollen clouds. We first extracted the haplotype corresponding to the paternal contribution to each diploid genotype 312 following Gauzere et al. (2016a). These paternal genotypes were then used to estimate the 313 314 pairwise paternal relatednesses among each pair of offspring k and k' ($\rho_{k,k'}$) using the kinship coefficient of Loiselle et al. (1995) implemented in SPAGEDI (Hardy & Vekemans, 2002), and 315 316 the local allelic frequencies (AF_{local}). These coefficients were estimated removing all the selfed individuals to only compare the outcrossed pollen pools. 317

318

319 **Results**

320 Neutral genetic differentiation among plots and pollen immigration

The admixture F-model of Karhunen & Ovaskainen (2012) estimated a posterior Fst value 321 of 0.026 [0.024; 0.029], consistent with previous estimates of neutral genetic differentiation 322 among populations measured with a comparable method but from 51 populations over a much 323 wider area of Mont-Ventoux (Fst = 0.026; Lander *et al.* 2011). The coancestry estimates within 324 plots (diagonal elements of the θ^P matrix) indicated that plot N1 tended to have lower effec-325 326 tive population size than plot N2 and N4 (Table 2). Plots N1 and N2 exhibited slightly higher 327 between-plots coancestry coefficient, suggesting that the two closest plots exchanged more gene flow ($\theta_{N1N2} > \theta_{N1N4} \sim \theta_{N2N4}$). Overall, the coancestry estimates were low (maximum 328 $\theta_{N1} = 0.057$; Table 2) indicating that effective population sizes were globally large. This result 329 330 is consistent with the large effective number of pollen donors Nep estimated for these sites using the same dataset (Table S1; Gauzere et al. 2013b). 331

332

Using paternity assignment, we categorized n = 658 offspring as "local" (45.5 %) and n334 = 492 as "immigrant" (34 %). A total of 295 (20 %) offspring were outside of these two 335 categories (*i.e.* ambiguous origin). Note that this "immigrant" status includes offspring with 336 non-sampled fathers which could be located either at a short- or long-distance to the mother-337 tree's plot. We checked that the LOD-score of long-distance vs. local immigration was indeed 338 higher on average for the "immigrant" than "local" individuals, with $\overline{v} = -0.33$ and 0.18 for 339 local and immigrant offspring respectively (*p-value* < 0.001 ANOVA test; Figure S1). 340

The analysis of the pairwise coefficient of paternal relatedness showed that within the immigrant pollen pool fathers were significantly less related than within the local pollen pool, with $\overline{\rho}_p = 0.011$ and $\overline{\rho}_p = 0.007$ within the local and immigrant pollen pools respectively (*pvalue* < 0.001 ANOVA test). Consequently, we estimated a higher effective number of pollen donors (*Nep*) in the immigrant than in the local pollen pool (*Nep* = 26.4 and 31.3 for local and immigrant pollen pools respectively).

347 Patterns of genetic trait differentiation and test for selection

348 Significant genetic differentiation across elevations was found for ΔH_{log} , ΔD , $t_{b2 \rightarrow 3(2011)}$ and $t_{b2\to3(2012)}$, while marginal differentiation was found for $t_{s1\to2(2011)}$ (as revealed by P effect 349 in model (1); Figure 3). Using the method of Karhunen & Ovaskainen (2012), we detected 350 a signal of natural selection for three of these four traits: ΔH_{log} , ΔD and $t_{b2 \rightarrow 3(2011)}$ (*i.e.* $S > t_{b2 \rightarrow 3(2011)}$) 351 352 0.90; Figure 3). For growth traits, plots at intermediate and high elevations presented the highest height and diameter increments ($\Delta H_{logN2} = 0.88$, $\Delta H_{logN4} = 1.03$ and $\Delta D_{N2} = 0.82$, ΔD_{N4} 353 = 0.72), while the site at the lowest elevation showed the lowest growth ($\Delta H_{logN1} = 0.56$ and 354 $\Delta D_{N1} = 0.50$; Figure 3). For the phenological traits, buds flushed earlier for plots from high 355 356 elevations as compared to low elevation plots ($t_{b2\to3(2011)N1} = 101.9$; $t_{b2\to3(2011)N2} = 100.3$; 357 $t_{b2\rightarrow3(2011)N4}$ = 99.8; Figure 3). Interestingly, contrary to growth and phenological traits, morphological and physiological traits (except LMA) tended to be slightly less differentiated then 358 expected under neutral processes (with $S \in [0.24; 0.39]$), with a strong conservation of the ge-359 netic values between plots (e.g. for the leaf mass $LM_{1/N1} = 0.44$, $LM_{1/N2} = 0.45$, $LM_{1/N4} = 0.44$; 360 361 Figure 3).

362

363 Effect of pollen immigration on within-plot genetic variation

364 Using the categorical assignment and the model (4), we only estimated marginally significant 365 difference in V_A between "local" and "immigrant" individuals for $t_{b2\rightarrow3(2012)}$ ($p_{value} = 0.088$; for all the other traits $p_{value} > 0.1$). In that case, we predicted higher V_A within the "immigrant" 366 than the "local" group ($V_{A_{immig}}$ = 16.57 and $V_{A_{local}}$ = 3.16; Table 3). Using the random regression 367 model (6), we found significant variation in additive genetic variance with the likelihood of 368 provenance from long-distance pollen pool for ΔH_{log} ($p_{value} = 0.005$), and a marginally signifi-369 cant effect for $t_{b2\rightarrow3(2012)}$ ($p_{value} = 0.078$; for all the other traits $p_{value} > 0.1$; Table 3). The esti-370 mated variance-covariance terms showed increasing V_A with higher likelihood of long-distance 371 pollination event for both ΔH_{log} and $t_{b2 \rightarrow 3(2012)}$ (Figure 4). Nevertheless, for these traits, the 372 fixed effect of v was not significant, which indicated no global trend in the variation of breed-373 ing values with the likelihood that an offspring originates from long-distance pollen dispersal. 374 375 Interestingly, these two different analyses highlighted the same trend towards a higher genetic variance for offspring originating from long-distance dispersal than from local dispersal. 376 377

378 **Discussion**

379 The scale of phenotypic adaptation

This study provides evidence that divergent selection can act on phenotypic traits at fine spatial 380 381 scale in temperate tree species. The study beech plots showed important dispersal capacities, with pollen dispersal that can easily reach 100 m, a fat-tailed pollen dispersal kernel and high 382 383 rates of immigration (Gauzere et al., 2013b; Oddou-Muratorio et al., 2018), similarly to other 384 anemophilous tree species (see Ashley 2010; Kremer et al. 2012 for reviews). Since dispersal 385 distances overlap the spatial scale of the selective environment, our results provide strong evidence that adaptive genetic differentiation can occur despite high gene flow (see Fitzpatrick 386 387 et al. 2015; Peterson et al. 2016 for recent evidence in other taxa). This also suggests that 388 selective pressures on seedling growth (in diameter and height) and on budburst timing were strong enough to counteract the homogenizing effect of gene flow on structuring the genetic 389 390 variation at the scale of a few hundred meters up to one kilometre. We found that the seedlings 391 derived from the plots at high elevation, inhabiting the lowest temperature conditions, flushed earlier and had a higher height and diameter growth in the common garden than the plot at 392 393 low elevation. For the budburst date, this pattern of genetic differentiation is opposed to the 394 pattern of phenotypic variation measured in situ, where seedlings from plots at low elevations 395 flushed earlier than plots at high elevations (Davi et al. 2011). Beech populations originating 396 from higher longitude or elevation have also been shown to be genetically earlier in provenance 397 tests (Teissier du Cros et al., 1988; Gomory & Paule, 2011; Kramer et al., 2017), suggesting 398 that these populations evolved phenological traits promoting a longer growing season length. In forestry, growth is considered a good approximation of the performance of trees in a given 399 400 condition (e.g. Rehfeldt et al. 1999, 2002). In situ and ex situ, the highest diameter growth values were measured at intermediate to high elevations (Cailleret & Davi 2011 and present 401 402 study), suggesting that the genetic and environmental influences on growth traits are in the same direction. At this southern edge of the ecological conditions tolerated by beech (Figure 403 2), the reduced allocation to stem growth at the low elevation plot is likely an adaptive response 404 to drought, which has previously been described by comparing marginal vs central beech pop-405 ulations (Rose et al., 2009). These studies, including ours, shed light on the phenotypic traits 406 under selection and underlying adaptation at different spatial scales. However, they do not pro-407 vide a direct assessment of the strength of natural selection or the ecological bases of selection, 408 409 nor do they quantify phenotypic mismatch and maladaptation, which would require selection analyses (MacColl, 2011). 410

411

The weak adaptive divergence of morphological and physiological traits as compared to 412 growth and phenological traits along ecological gradients remains an open question. Meta-413 analyses highlighted that growth and phenological traits are the most genetically differentiated 414 415 phenotypic traits at large spatial scale in temperate and boreal tree species (Savolainen et al., 416 2007; Alberto et al., 2013). Here, our study suggests that this conclusion also holds at a narrower spatial scale. The existing literature also tend to indicate that morphological and physi-417 418 ological traits generally exhibit a lower fraction of genetically controlled phenotypic differentiation, and a higher within-population genetic variation than other traits (Hajek et al., 2016). 419 420 A putative explanation is that latitudinal or elevation gradients do not fully capture the spatial scale of variation for drought stress, even if they capture a strong variation in temperature. 421

422 In particular in our study, the important micro-heterogeneity in soil water capacity on Mont-Ventoux (Nourtier et al., 2013) may generate micro-local differentiation of morphological and 423 424 physiological traits, blurring the signal of divergent selection for morphological and physio-425 logical traits at a larger spatial scale (Cubry et al. unpublished). Another possible explanation 426 is that morphological and physiological traits could be under stabilizing selection (e.g. Lamy 427 et al. 2011). Unfortunately, the power to detect the signal of stabilizing selection is low for 428 species with low neutral genetic differentiation, such as trees (Savolainen et al., 2007), and pat-429 terns of weak quantitative trait differentiation are difficult to interpret (Whitlock, 2008). With the development of genomic resources, this question of microgeographic adaptation is now 430 431 mainly addressed using genome-wide association studies (e.g. Eckert et al. 2015). However, 432 associations between genotype, phenotype and environment often provide weaker evidence of 433 adaptive differentiation at the phenotypic levels (Le Corre & Kremer, 2012), likely because of 434 the complex multi-locus determinism of the traits underlying phenotypic adaptation. Using a more precise characterisation of ecological gradients, e.g. by considering the spatial-scale of 435 water availability, and new methods to compare the similarity of habitats and phenotypes of-436 fer promising prospects to better understand the spatial scale of phenotypic adaptation and the 437 ecological drivers of adaptive divergence (Karhunen et al., 2014; Csillery et al., 2020). 438 439

440 Effects of pollen dispersal on quantitative trait variation at a local scale

For populations locally adapted and experiencing strong selective pressures that vary through 441 space, theory suggests that moderate gene flow can maintain quantitative genetic variation 442 within these populations (Barton, 2001; Le Corre & Kremer, 2003; Bridle et al., 2010). Al-443 though adaptive genetic differentiation is widespread, only a few studies have provided evi-444 445 dence that gene flow is an important process to maintain the high levels of genetic variation usually detected in nature (e.g. Yeaman & Jarvis 2006). Here, consistent with this theoretical 446 expectation, we detected an effect of pollen dispersal on increased genetic variance for two 447 locally differentiated traits, *i.e.* height growth and budburst date. However, we likely lacked 448 power to detect this effect for the third heritable trait under selection, diameter growth. Be-449 450 cause female flowers are receptive before the emission of local pollen (protogyny) and lower populations flower earlier than higher populations, we hypothesized a directional pollen flow 451

across elevations in Gauzere *et al.* (2013a), which would have resulted in shifted genetic values
between immigrant and local individuals (Figure 1; Nilsson 1995). Our results did not support
this expectation.

455 456 We used two different approaches to characterize the pollen immigration status. Both analyses suggested that immigrant pollen brings more genetic variance for locally differentiated 457 458 traits than local pollen. These effects were quantitatively high, with genetic variance for immi-459 grants more than twice as high as for local offspring. However, this result had large standard errors around the estimates, and the significance changed depending on how the immigration 460 status was characterized. We checked that "immigrants" individuals originated from more di-461 verse fathers (higher Nep), and had paternal genotypes more similar to global than local allelic 462 463 frequencies (higher v). Nevertheless, our immigrant status still likely suffers some inaccuracy. In particular, the low neutral genetic differentiation among the study beech plots (Fst = 0.026) 464 465 and, overall, on the northern face of the Mont-Ventoux (Lander et al., 2011), may reduce our ability to retrieve the origin of a pollen grain using molecular markers. Moreover, the low levels 466 of quantitative genetic differentiation measured along the elevation gradient suggest that the ex-467 pected effects of pollen flow on genetic values after one episode of reproduction is necessarily 468 469 moderate. Here, this effect may thus be difficult to detect, even knowing the true immigration 470 status of all offspring. 471

The approach used here provides a direct quantification of the impact of contemporary 472 pollen immigration on the genetic variance of quantitative traits. Although using molecular 473 markers to distinguish migrants and local offspring could have been done on numerous ex-474 isting datasets, to our knowledge no study before ours had reiterated the approach first used 475 476 by Nilsson (1995). However, we cannot estimate the immigration load associated with this 477 increase in genetic variance for the locally differentiated traits. Although this genetic load is expected to be beneficial under rapidly changing environmental conditions (Yeaman & Jarvis, 478 479 2006; Savolainen et al., 2007; Kremer et al., 2012), this remains to be tested. Applying this 480 approach to other species and populations may improve our ability to both understand and 481 quantify the impact of gene flow on the adaptive potential of quantitative traits under natural conditions. This effect should be more detectable for populations departing from migration-482

selection equilibrium, such as isolated (or marginal) populations, or transplanted plant populations (see Fitzpatrick *et al.* 2015, 2016 for examples in guppies).

485

486 The present results about trait differentiation and genetic variance of migrants and non-487 migrants extend previous studies about pollen dispersal kernels and migration rates (Gauzere 488 et al., 2013a), trait heritabilities (Gauzere et al., 2016a) and variance of fecundities (Oddou-489 Muratorio et al., 2018). Pulling all effects together suggests building a predictive mechanistic 490 model to both (i) check whether the observed levels of genetic variance in migrant pollen are consistent with the amounts of migrant genetic variance that would be predicted mechanisti-491 492 cally and (ii) further investigate the adaptive dynamics at the scale of few generations (following e.g. Kuparinen et al. 2010). This work is beyond the scope of the present study because 493 494 it still requires a map of beech density over the whole northern face of Mont-Ventoux and an 495 extrapolation model to propose a spatially explicit distribution of adaptive diversity. Point (ii) 496 would require information about the variation in selection gradients or adaptive landscapes with elevation. Unfortunately, empirical estimates of selection need specific approaches in natural 497 498 populations of long-lived species and, hence, remain scarce in forest trees (but see Bontemps et al. 2017). 499

500

501 Gene flow in the context of climate change

502 One of the most documented consequences of climate change on tree ecosystem functioning is probably the impact of climate warming on the advancement of spring phenology (Menzel 503 et al., 2006; Fu et al., 2015). Understanding the adaptive or maladaptive value of the response 504 505 of phenology to climate has become a key focus of evolutionary studies (Tansey et al. 2017; 506 Gauzere et al. in press). An underrated issue is that these phenological changes are also likely to affect gene flow among populations, and efficiency of long-distance pollen dispersal. For many 507 508 temperate tree species, we now have strong evidence that climate warming reduces spring phenology synchrony among individuals experiencing the same environmental conditions (Zohner 509 510 et al., 2018), but also leads to more uniform spring phenology across climatic gradients (Vitasse et al., 2018). The evolutionary consequences of pollen dispersal may therefore change with in-511 creasing climate warming. In this study, we detected an effect of local selection and pollen 512

dispersal on the adaptive potential of traits that are known to affect dispersal or reproduction in 513 anemophilous tree species (tree height and phenology; Klinkhamer et al. 1997; Gauzere et al. 514 2013a; Oddou-Muratorio et al. 2018). In particular, long-distance pollen dispersal events are 515 only effective if the donor and recipient trees have synchronous phenology (Aitken et al., 2008). 516 517 Our results thus suggest that refining the models of phenotypic adaptation by incorporating the 518 feedbacks between traits under divergent selection and dispersal-reproductive capacities may 519 be important to thoroughly understand the evolutionary consequences of gene flow in plant 520 species (e.g. Soularue & Kremer 2012).

521

522 With the concern that populations could be maladaptated and lag behind with increasing 523 climate change, new conservation and management strategies are needed to mitigate these effects. Assisted gene flow, *i.e.* the managed movement of individuals or gametes between pop-524 525 ulations within species ranges, is one of the tools proposed to help populations adapt to new 526 climatic conditions (Aitken & Whitlock, 2013; Aitken & Bemmels, 2016). However, assisted gene flow among distant populations, at different latitudes, may have unexpected consequences 527 for the response of individuals or new hybrids, because of outbreeding depression (Aitken & 528 Whitlock, 2013), but also because other environmental cues than temperature can govern physi-529 530 ological responses in trees (see for instance the role of photoperiod in bud development; Way & 531 Montgomery 2015). Here, our results suggest some effect of local pollen dispersal on increasing genetic variation for locally differentiated traits after solely one episode of reproduction. 532 Therefore, programs considering the movements of gametes along short ecological gradients 533 could be efficient to boost genetic diversity and introduce genetic innovation within natural or 534 535 managed tree populations. Our results also encourage the development of evolution-oriented forest management and silviculture, whereby thinning could be done to favour the spread of 536 537 long-distance gene flow (e.g. by reducing local densities; Lefèvre et al. 2014). Nonetheless, 538 contrarily to assisted gene flow, natural gene flow can also lead to increased maladaptation 539 through migration load (Aitken & Whitlock, 2013).

540

541 Conclusion

By combining population genetics and quantitative genetics, we provided a first comprehen-542 sive analysis of the effects of gene flow on the adaptive potential of a tree species in nature. We 543 found that gene flow among trees at different elevations did not prevent adaptive differentiation 544 among plots separated by less than one kilometre, and that divergent selection at this fine spa-545 546 tial scale acted mainly on growth and phenological traits, likely as an adaptation to drought and temperature variation. For two genetically differentiated traits, our analyses suggested an effect 547 of pollen dispersal on increasing genetic diversity after one episode of reproduction. However, 548 our power to detect and estimate these effects was probably quite low. Our results thus suggest 549 550 that local selection is an important evolutionary force in natural tree populations and that, under 551 rapidly changing environmental conditions, gene flow can increase the adaptive capacities of 552 natural populations.

553

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572 Author contributions

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578

J.G., E.K.K. and S.O.M. conceived and designed the study. J.G collected the genetic and phenotypic data, performed the analyses, wrote the first draft of the text and contributed the revisions.
S.O.M. conceived and coordinated the beech experiment. J.G., H.D., O.B., E.K.K. and S.O.M.
discussed and interpreted the findings. E.K.K., H.D., O.B. and S.O.M. contributed to the analyses and the revisions of preliminary drafts.

579 Data archiving statement

580 Data available upon request from the authors

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821 Supporting Information

- 822 Fig. S1 Distribution of the LOD-score of long-distance vs local immigration between 'immi-
- 823 grant' and 'local' groups.
- 824 **Table S1** Location, climate and ecological characteristics of the study plots.
- 825 **Table S2** Test for plot differentiation and signal of selection in quantitative traits.
- 826 Table S3 Differences in additive genetic variance between 'immigrant' and 'local' groups esti-
- 827 mated from model eq. (4).
- Table S4 Effect of pollen immigration on the variation in additive genetic variance estimatedfrom model eq. (6).

830 Figure legends

831 Figure 1. Expected effects of random and directional gene flow between locally differentiated populations on individuals genetic values. (a) In the case of random gene flow, migrant gene 832 flow (orange arrows) tends to homogenise the frequencies of the red, grey and blue genotypes 833 in each environment. Hence, in the core population (e.g. grey habitat), individuals and gametes 834 835 from migrant gene flow should exhibit more genetic variance for the traits under local selection $(v_1 > v_2)$ than individuals and gametes from local gene flow (*i.e.* gene flow within population, 836 represented by the green arrows), while the mean genetic value (dotted line) should be identical 837 (assuming a gradual shift of the optimum). (b) In the case of directional gene flow (e.g. only the 838 red genotypes migrate to the grey habitat), migrant gene flow should introduce individuals and 839 gametes with shifted mean genetic values (dotted lines; $m_1 \neq m_2$), while the variance in genetic 840 values between migrant and local gene flow should be similar (assuming stabilising selection 841 is the same in each habitat). 842

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Figure 2. Climatic space explored by the three study plots in comparison to the whole species range of *Fagus sylvatica*. This climatic space is represented as the variation of temperature and summer precipitation over beech distribution area, extracted for the SAFRAN database for the period 1958-2015 (collected on a 8 km² grid represented by grey dots). Each black triangle represents the average climate recorded since 2007 with HOBO weather stations for the plot N1 (1020 m), N2 (1140 m) and N4 (1340 m a.s.l.).

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Figure 3. Level of trait adaptive divergence between the three plots located at 1020 m (N1), 851 852 1140 m (N2) and 1340 m (N4) a.s.l. for the growth, phenological, physiological and morpho-853 logical traits measured on seedlings of Fagus sylvatica. Histograms represent the a posteriori 854 distributions of additive genetic values estimated for each plot. The dotted lines represent the 855 median of these distributions. We also provide the test for population differentiation (P) and signal of selection (S) for these quantitative traits. S coefficient close to 1 and 0 indicates signal 856 of divergent and stabilizing selection respectively. With "***" p < 0.001, "**" 0.001 ,857 "*" 0.01 0.09. 858

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Figure 4. Change in additive genetic variance (V_A) with the LOD-score of long-distance im-

861 migration vs. local immigration (v), estimated using model (6) in Fagus sylvatica. This relationship is represented only for the two traits, out of 11 tested, for which we found that 862 863 this effect was significant or marginally significant: (left) the log-transformed height growth (ΔH_{log}) and (right) the budburst date in 2012 $(t_{b2\rightarrow3(2012)})$. The solid line represents the es-864 865 timate, and the dashed lines the 95 % confidence intervals around this estimate. From the random regression model (6), we can estimate the change in V_A with a covariate x as: $V_A(x) =$ 866 $var(trait_{intercept} + trait_{slope}.x) = V_{A intercept} + x^2 V_{A slope} + 2.x.cov_A(intercept, slope)$, using the 867 868 estimates provided in Table 3.

Abbreviation	Trait	Category	N_w	Unit	Transformation
ΔH_{log}	Total height increase	Growth	2380	rate	log
ΔD	Total diameter increase	Growth	2380	rate	-
$t_{b2\to 3(2011)}$	Budburst date in 2011	Phenology	2334	DOY	-
$t_{b2 \rightarrow 3(2012)}$	Budburst date in 2012	Phenology	1835	DOY	-
$t_{s1\rightarrow 2}$	Senescence date in 2011	Phenology	2386	DOY	-
<i>VD</i> ₂₀₁₁	Vegetation season duration in 2011	Phenology	2193	DOY	-
LM_{\checkmark}	Dry leaf mass	Morphology	2491	g	square-root
LA_{\checkmark}	Leaf Area	Morphology	2491	cm ²	square-root
LMA	Leaf Mass Area	Morphology	2491	$g.m^{-2}$	-
%N	Nitrogen content in leaves	Physiology	1031	%	-
$\delta^{13}C$	Water use efficiency	Physiology	1031	%0	-

Table 1: Description and descriptive statistics for the traits measured on seedlings of *Fagus* sylvatica. With N_w the number of phenotyped individuals. DOY: day of the year.

Table 2: Matrix of the coancestry coefficients estimated by the admixture F-model of Karhunen & Ovaskainen (2012) for the three plots of *Fagus sylvatica*. The diagonal correspond to the within-plot coancestry coefficient, with its median value and confidence intervals (*i.e.* values of the estimates that bound 95 % of the posterior distribution). Lower diagonal elements are the median coancestry coefficients among-plots, with confidence intervals in brackets.

	N1	N2	N4
N1	0.057 [0.033; 0.085]		
N2	0.018 [0.005; 0.035]	0.029 [0.017; 0.045]	
N4	0.012 [0.00; 0.024]	0.013 [0.002; 0.023]	0.035 [0.02; 0.052]

Table 3: Effect of pollen immigration on the variation in additive genetic variance for traits measured in seedlings of Fagus sylvatica. This effect was found significant or marginally significant for two genetically differentiated traits: using model (4) for $t_{b2\rightarrow3(2012)}$ with *pvalue* = 0.088; using model (6) for ΔH_{log} with *pvalue* = 0.005 and for $t_{b2\rightarrow3(2012)}$ with *pvalue* = 0.078. Model (4) estimates an additive genetic variance (V_A) for each "immigrant" and "local" group (with standard error in brackets). Model (6) estimates the additive genetic variance-covariance terms for the intercept and slope of the trait with the LOD-score of long-distance immigration vs. local immigration v (with standard error in brackets). The diagonal elements thus provide $V_{A intercept}$ and $V_{A slope}$, and the off-diagonal element $cov_A(intercept, slope)$. See Supporting Information Tables S3 and S4 for the model outputs for all phenotypic traits.

Trait	V _{Aimmig}	V _{Alocal}
$t_{b2\to 3(2012)}$	16.57 (8.2)	3.16 (5.6)

Model 6

additive genetic effects for ΔH_{log}						
	$\Delta H_{log_{intercept}}$	$\Delta H_{log_{slope}}$				
$\Delta H_{log_{intercept}}$	0.16 (0.041)					
$\Delta H_{log_{slope}}$	0.029 (0.010)	0.0092 (0.008)				
additive genetic effects for $t_{b2\rightarrow3(2012)}$						
	$t_{b2 \rightarrow 3(2012)_{intercept}}$	$t_{b2 \rightarrow 3(2012)_{slope}}$				
$t_{b2 \rightarrow 3(2012)_{intercept}}$	7.36 (4.8)					
$t_{b2 \rightarrow 3(2012)_{slope}}$	3.44 (2.12)	2.47 (2.18)				

34

Summer precipitation (mm) 200 300 400 N2 **N1** 100 2 12 8 10 6 Temperature (°C)







