

## Limit theorems for some Markov operators

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The exponential rate of convergence and the Central Limit Theorem for some Markov operators are established. The operators correspond to iterated function systems which, for example, may be used to generalize the cell cycle model given by Lasota and Mackey [12].

### I. INTRODUCTION

We are concerned with Markov operators corresponding to iterated function systems. The main goals of the paper are to prove exponential rate of convergence and establish the Central Limit Theorem (CLT). It should be indicated that the first result implies the second. The operators under consideration are more general than those used by Lasota and Mackey in [12]. The authors studied therein some cell cycle model, in which the rate of convergence is already evaluated by Wojewódka [22]. Hille et al. [9] proposed the generalization of the model and assured the existence of a unique invariant distribution in it. We have managed to evaluate the rate of convergence, which provides asymptotic stability at once, as well as allows us to show the CLT. The results bring some information important from biological point of view. To get more details on biological background of the research, see Tyson and Hannsgen [20] or Murray and Hunt [15].

In our paper we base on coupling methods introduced by Hairer in [6]. In the same spirit, exponential rate of convergence was proven in [19] for classical iterated function systems (see also [7] or [10]). However, we use coupling methods not only to evaluate the rate of convergence. It turns out that properly constructed coupling measure, if combined with the results for stationary ergodic Markov chains given by Maxwell and Woodroffe [14], is crucial in the proof of the CLT, too. If we have the coupling measure already constructed, the proof of the CLT is brief and less technical than typical proofs based on Gordin's martingale approximation. What led us to this intriguing solution was an unsuccessful attempt to follow the pattern given by Komorowski and Walczuk [11]. It is worth mentioning here that an auxiliary model, described by some non-homogenous Markov chain, is needed to take advantage of coupling methods. While reading the paper, one may see that it is a bright idea to express the Markov operator of interest by means of an auxiliary one.

Similar approach may also help to establish the Law of the Iterated Logarithm (LIL). The proof of the LIL is supposed to be provided in a future paper. Some ideas useful for proving it may be adapted from Bołt et al. [2]. However, we strongly believe that using an appropriate coupling measure will, again, make the proof much easier.

The organization of the paper goes as follows. Section 2 introduces basic notation and definitions that are needed throughout the paper. Most of them are adapted from [1], [16], [13] and [18]. Mathematical derivation of the generalized cell cycle model is provided in Section 3. The main

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theorems (Theorem 1 and Theorem 2) are also formulated there. Sections 5-7 are devoted to the construction of coupling measure for iterated function systems. Thanks to the results presented in Section 8 we are finally able to present the proofs of main theorems. Indeed, the exponential rate of convergence is established in Section 9 and the CLT in Section 10.

## II. NOTATION AND BASIC DEFINITIONS

Let  $(X, \varrho)$  be a Polish space. We denote by  $B_X$  the family of all Borel subsets of  $X$ . Let  $B(X)$  be the space of all bounded and measurable functions  $f : X \rightarrow R$  with the supremum norm and write  $C(X)$  for its subspace of all bounded and continuous functions with the supremum norm.

We denote by  $M(X)$  the family of all Borel measures on  $X$  and by  $M_{\text{fin}}(X)$  and  $M_1(X)$  its subfamilies such that  $\mu(X) < \infty$  and  $\mu(X) = 1$ , respectively. Elements of  $M_{\text{fin}}(X)$  which satisfy  $\mu(X) \leq 1$  are called sub-probability measures. To simplify notation, we write

$$\langle f, \mu \rangle = \int_X f(x) \mu(dx) \quad \text{for } f : X \rightarrow R, \mu \in M(X).$$

An operator  $P : M_{\text{fin}}(X) \rightarrow M_{\text{fin}}(X)$  is called a Markov operator if

1.  $P(\lambda_1 \mu_1 + \lambda_2 \mu_2) = \lambda_1 P\mu_1 + \lambda_2 P\mu_2$  for  $\lambda_1, \lambda_2 \geq 0$ ,  $\mu_1, \mu_2 \in M_{\text{fin}}(X)$ ;
2.  $P\mu(X) = \mu(X)$  for  $\mu \in M_{\text{fin}}(X)$ .

Markov operator  $P$  for which there exists a linear operator  $U : B(X) \rightarrow B(X)$  such that

$$\langle Uf, \mu \rangle = \langle f, P\mu \rangle \quad \text{for } f \in B(X), \mu \in M_{\text{fin}}(X)$$

is called a regular operator. We say that a regular Markov operator is Feller if  $U(C(X)) \subset C(X)$ . Every Markov operator  $P$  may be extended to the space of signed measures on  $X$  denoted by  $M_{\text{sig}}(X) = \{\mu_1 - \mu_2 : \mu_1, \mu_2 \in M_{\text{fin}}(X)\}$ . For  $\mu \in M_{\text{sig}}(X)$ , we denote by  $\|\mu\|$  the total variation norm of  $\mu$ , i.e.,

$$\|\mu\| = \mu^+(X) + \mu^-(X),$$

where  $\mu^+$  and  $\mu^-$  come from the Hahn-Jordan decomposition of  $\mu$  (see [8]). In particular, if  $\mu$  is non-negative,  $\|\mu\|$  is the total mass of  $\mu$ . For fixed  $\bar{x} \in X$  we also consider the space  $M_1^1(X)$  of all probability measures with finite first moment, i.e.,  $M_1^1(X) = \{\mu \in M_1(X) : \int_X \varrho(x, \bar{x}) \mu(dx) < \infty\}$ . The family is independent of choice of  $\bar{x} \in X$ . We call  $\mu_* \in M_{\text{fin}}(X)$  an invariant measure of  $P$  if  $P\mu_* = \mu_*$ . For  $\mu \in M_{\text{fin}}(X)$ , we define the support of  $\mu$  by

$$\mu = \{x \in X : \mu(B(x, r)) > 0 \text{ for all } r > 0\},$$

where  $B(x, r)$  is an open ball in  $X$  with center at  $x \in X$  and radius  $r > 0$ . By  $\bar{B}(x, r)$  we denote a closed ball with center at  $x \in X$  and radius  $r > 0$ .

In  $M_{\text{sig}}(X)$ , we introduce the Fortet-Mourier norm

$$\|\mu\|_{\mathcal{L}} = \sup_{f \in \mathcal{L}} |\langle f, \mu \rangle|,$$

where  $\mathcal{L} = \{f \in C(X) : |f(x) - f(y)| \leq \varrho(x, y), |f(x)| \leq 1 \text{ for } x, y \in X\}$ . The space  $M_1(X)$  with metric  $\|\mu_1 - \mu_2\|_{\mathcal{L}}$  is complete (see [5], [17] or [21]).

We say that the sequence of Borel measures  $(\mu_n)_{n \in \mathbb{N}} \subset M_{\text{fin}}(X)$  converges weakly to the measure  $\mu \in M_{\text{fin}}(X)$  if  $\lim_{n \rightarrow \infty} \langle f, \mu_n \rangle = \langle f, \mu \rangle$  for all  $f \in C(X)$ . It is known (see Theorem 11.3.3, [3]) that the following conditions are equivalent

- $(\mu_n)_{n \in \mathbb{N}}$  converges weakly to  $\mu$ ,
- $\lim_{n \rightarrow \infty} \langle f, \mu_n \rangle = \langle f, \mu \rangle$  for all  $f \in \mathcal{L}$ ,
- $\lim_{n \rightarrow \infty} \|\mu_n - \mu\|_{\mathcal{L}} = 0$ ,

where  $(\mu_n)_{n \in \mathbb{N}} \subset M_1(X)$  and  $\mu \in M_1(X)$ .

### III. MAIN IDEA AND THEOREMS

Recall that  $(X, \varrho)$  is a Polish space and let  $(\Omega, \mathcal{F}, \text{Prob})$  be a probability space. Fix  $T < \infty$ . We consider a stochastically perturbed dynamical system. The state of  $x_n$ , for every  $n \in \mathbb{N}$ , is determined by the formula

$$x_{n+1} = S(x_n, t_{n+1}).$$

We make the following assumptions.

- (I) We consider a sequence  $(t_n)_{n \in \mathbb{N}}$  of independent random variables defined on  $(\Omega, \mathcal{F}, \text{Prob})$  with values in  $[0, T]$ . Distribution of  $t_{n+1}$  conditional on  $x_n = x$  is given by

$$\text{Prob}(t_{n+1} < t | x_n = x) = \int_0^t p(x, u) du, \quad 0 \leq t \leq T, \quad (1)$$

where  $p : X \times [0, T] \rightarrow [0, \infty)$  is a measurable and non-negative function. In addition,  $p$  is normalized, i.e.,  $\int_0^T p(x, u) du = 1$  for  $x \in X$ .

- (II) Let  $S : X \times [0, T] \rightarrow X$  be a continuous function which satisfies the Lipschitz type inequality

$$\varrho(S(x, t), S(y, t)) \leq \lambda(x, t) \varrho(x, y) \quad \text{for } x, y \in X, t \in [0, T], \quad (2)$$

where  $\lambda : X \times [0, T] \rightarrow [0, \infty)$  is a Borel measurable function such that

$$a := \sup_{x \in X} \int_0^T \lambda(x, t) p(x, t) dt < 1. \quad (3)$$

- (III)  $\sup_{t \in [0, T]} \varrho(S(\bar{x}, t), \bar{x}) < \infty$  for some  $\bar{x} \in X$ .

- (IV) We assume that  $p$  satisfies the Dini condition

$$\int_0^T |p(x, t) - p(y, t)| dt \leq \omega(\varrho(x, y)) \quad \text{for } x, y \in X, \quad (4)$$

where  $\omega : \mathbb{R}_+ \rightarrow \mathbb{R}_+$ ,  $\omega(0) = 0$ , is a non-decreasing and concave function such that

$$\int_0^\sigma \frac{\omega(t)}{t} dt < +\infty \quad \text{for some } \sigma > 0.$$

We can easily check that if  $\zeta < 1$ , we have

$$\varphi(t) := \sum_{n=1}^{\infty} \omega(\zeta^n t) < +\infty \quad \text{for every } t \geq 0 \quad (5)$$

and  $\lim_{t \rightarrow 0} \varphi(t) = 0$ .

(V) Function  $p$  is bounded. We set  $\delta := \inf_{x \in X, t \in (0, T]} p(x, t)$ ,  $M := \sup_{x \in X, t \in [0, T]} p(x, t)$  and require  $\delta > 0$ .

We further assume that, for each  $A \in B_X$ ,

$$\text{Prob}(x_{n+1} \in A) := \mu_{n+1}(A) \quad \text{and} \quad P\mu_n = \mu_{n+1},$$

where

$$P\mu(A) = \int_X \left[ \int_0^T 1_A(S(x, t)) p(x, t) dt \right] \mu(dx).$$

In [12] the proof of asymptotic stability is given for the model, while the exponential rate of convergence is established thanks to some coupling methods in [22]<sup>1</sup>.

Without loss of generality, we may think of  $(X, \varrho)$  as a closed subset of some separable Banach space  $H$ . Then, trying to describe some intercellular processes more precisely, Hille et al. [9] proposed a more general dynamical system

$$x_{n+1} = S(x_n, t_{n+1}) + H_{n+1},$$

where  $(H_n)_{n \in \mathbb{N}}$ ,  $H_n \in H$ , is a family of independent random variables with the same distribution given by a measure  $\nu^\varepsilon$ , which is independent of  $S(x_n, t_{n+1})$  and its support stays in  $\bar{B}(0, \varepsilon)$ .

For this reason, we need an additional assumption

(VI) Let  $\varepsilon_* < \infty$  be given. Fix  $\varepsilon \in [0, \varepsilon_*]$ . Let  $\nu^\varepsilon$  be a Borel measure on  $H$  such that its support is in  $\bar{B}(0, \varepsilon)$ . For every  $x \in X$ , we set

$$\nu_x^\varepsilon(\cdot) = \nu^\varepsilon(\cdot - x). \quad (6)$$

We assume that  $S(x, t) + h \in X$  for every  $t \in [0, T]$ ,  $x \in X$  and  $h$  from the support of  $\nu^\varepsilon$ .

The Markov chain is given by the transition function  $\Pi_\varepsilon : X \times B_X \rightarrow [0, 1]$  of the form

$$\Pi_\varepsilon(x, A) = \int_0^T p(x, t) \nu_{S(x, t)}^\varepsilon(A) dt.$$

Then, we may write the Markov operator  $P_\varepsilon : M_1(X) \rightarrow M_1(X)$  as follows

$$P_\varepsilon \mu(A) = \int_X \Pi_\varepsilon(x, A) \mu(dx).$$

The case of deterministic protein production, i.e., when  $\varepsilon = 0$ , fits to the framework presented by Lasota and Mackey [12] and the results obtained there.

Hille et al. [9] managed to show the existence of a unique invariant measure in the generalized model, described above. However, stability was not proven. We want to focus on evaluating the rate of convergence, which additionally provides asymptotic stability in the model and allows us to establish the CLT. The proof of the CLT is given in Section 10.

**Theorem 1.** *Let  $\mu \in M_1^1(X)$ . Under assumptions (I)-(VI), there exist  $C = C(\mu) > 0$  and  $q \in [0, 1)$  such that*

$$\|P_\varepsilon^n \mu - \mu_*\|_{\mathcal{L}} \leq Cq^n \quad \text{for } n \in \mathbb{N}.$$

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<sup>1</sup> In both papers the results are proven for stronger assumptions.

Now, assumption (II) is strengthened to the following condition:

(II') Let  $S : X \times [0, T] \rightarrow X$  be a continuous function which satisfies

$$\varrho(S(x, t), S(y, t)) \leq \lambda(x, t)\varrho(x, y) \quad \text{for } x, y \in X, t \in [0, T], \quad (7)$$

where  $\lambda : X \times [0, T] \rightarrow [0, \infty)$  is a Borel measurable function such that

$$\Lambda := \sup_{x \in X} \int_0^T \lambda^2(x, t)p(x, t)dt < 1. \quad (8)$$

Note that (II') implies (II), due to the Hölder inequality, and we obtain that  $a \leq \sqrt{\Lambda} < 1$ . Assuming (II') instead of (II) allows us to show that  $\mu_* \in M_1^2(X) := \{\mu \in M_1(X) : \int_X \varrho^2(x, \bar{x})\mu(dx) < \infty\}$ , which is essential to establish the CLT in the way presented in this paper. It is proven in Lemma 15 that  $\mu_*$  is indeed with finite second moment.

Now, choose an arbitrary function  $g : X \rightarrow R$  which is Lipschitz continuous, bounded and satisfies  $\langle g, \mu_* \rangle = 0$ . Let  $(x_i)_{i \in N}$  be the Markov chain with transition probability function  $\Pi_\varepsilon$  and initial distribution  $\mu \in M_1^2(X)$ . For every  $n \in N$ , put

$$\eta_n^\mu := \frac{g(x_1) + \dots + g(x_n)}{\sqrt{n}}$$

and let  $\Phi_{\eta_n^\mu}$  denote its distribution.

**Theorem 2.** *Let  $\mu \in M_1^2(X)$  be with finite second moment and let  $\Phi_{\eta_n^\mu}$  be the distribution of  $\eta_n^\mu$ , as defined above. Assuming that all conditions (I)-(VI) are fulfilled and (II) is additionally strengthened to (II'). Then  $\Phi_{\eta_n^\mu}$  converges weakly to the normal distribution, as  $n \rightarrow \infty$ .*

#### IV. AN AUXILIARY MODEL - BASIC ASSUMPTIONS

Our aim is to prove exponential rate of convergence for the model given in [9], as it is stated in Theorem 1. The idea is to use coupling methods. However, implementing these methods directly to the model given above does not give the expected results. Instead, we fix a sequence of constants  $(h_n)_{n \in N} \subset H$ , where  $h_n \in \bar{B}(0, \varepsilon)$  for all  $n \in N$ , and consider a stochastically perturbed dynamical system

$$x_{n+1} = T_{h_{n+1}}(x_n, t_{n+1}) := S(x_n, t_{n+1}) + h_{n+1}, \quad n = 0, 1, 2, \dots$$

Note that

$$x_{n+1} = T_{h_{n+1}}\left(T_{h_n}(x_{n-1}, t_n), t_{n+1}\right).$$

For abbreviation, we introduce the symbol

$$\tilde{x}_{n+1}^{x_0} := T_{h_{n+1}}\left(T_{h_n}(\dots T_{h_2}(T_{h_1}(x_0, t_1), t_2) \dots), t_{n+1}\right), \quad (9)$$

where the upper index  $x_0$  indicates the point from which the iteration begins.

Let us further assume that, for every  $A \in B_X$ ,

$$\text{Prob}(x_{n+1} \in A) := \mu_{n+1}(A) \quad \text{and} \quad P_{h_{n+1}}\mu_n = \mu_{n+1},$$

where, for arbitrary  $h \in \bar{B}(0, \varepsilon)$ ,

$$(P_h \mu)(A) := \int_X \left[ \int_0^T 1_A(T_h(x, t)) p(x, t) dt \right] \mu(dx).$$

We maintain all previous assumptions (I)-(VI). Now, for every  $h \in \bar{B}(0, \varepsilon)$ , we consider an operator

$$T_h(x, t) := S(x, t) + h.$$

Note that, as a consequence of assumption (II),  $T_h$  is continuous and satisfies the same Lipschitz type inequality as operator  $S$  satisfies.

We also set

$$c := \sup_{t \in [0, T]} \varrho(S(\bar{x}, t), \bar{x}) + \varepsilon_* > \sup_{t \in [0, T], i \in N} \varrho(T_{h_i}(\bar{x}, t), \bar{x}). \quad (10)$$

Obviously,  $c$  is finite, because of assumption (III).

## V. MEASURES ON THE PATHSPACE AND COUPLING

Set  $x \in X$  and  $(h_n)_{n \in N} \subset H$ ,  $h_n \in \bar{B}(0, \varepsilon)$  for all  $n \in N$ . One-dimensional distributions  $\Pi_{h_1, \dots, h_n}^n(x, \cdot)$ ,  $n \in N$ , are defined by induction on  $n$

$$\begin{aligned} \Pi^0(x, A) &= \delta_x(A) \\ \Pi_h^1(x, A) &= \int_0^T 1_A(T_h(x, t)) p(x, t) dt \\ &\vdots \\ \Pi_{h_1, \dots, h_n}^n(x, A) &= \int_X \Pi_{h_n}^1(y, A) \Pi_{h_1, \dots, h_{n-1}}^{n-1}(x, dy), \end{aligned} \quad (11)$$

where  $A \in B_X$ . We easily obtain two-dimensional and higher-dimensional distributions. If we assume that, for  $x \in X$ ,  $\Pi_{h_1, \dots, h_n}^{1, \dots, n}(x, \cdot)$  is a measure on  $X^n$ , generated by a sequence  $(\Pi_{h_i}^1(x, \cdot))_{i=1}^n$ , then

$$\Pi_{h_1, \dots, h_{n+1}}^{1, \dots, n+1}(x, A \times B) = \int_A \Pi_{h_{n+1}}^1(z_n, B) \Pi_{h_1, \dots, h_n}^{1, \dots, n}(x, dz), \quad (12)$$

where  $z = (z_1, \dots, z_n)$  and  $A \in B_{X^n}$ ,  $B \in B_X$ , is a measure on  $X^{n+1}$ . Note that  $\Pi_{h_1}^1(x, \cdot), \dots, \Pi_{h_1, \dots, h_n}^n(x, \cdot)$ , given by (11), are marginal distributions of  $\Pi_{h_1, \dots, h_n}^{1, \dots, n}(x, \cdot)$ , for every  $x \in X$ . Finally, we obtain a family  $\{\Pi_{h_1, h_2, \dots}^\infty(x, \cdot) : x \in X\}$  of sub-probability measures on  $X^\infty$ . This construction is motivated by [6]. The existence of measures  $\Pi_{h_1, h_2, \dots}^\infty(x, \cdot)$  is established by the Kolmogorov theorem. More precisely, for any  $x \in X$ , there exists some probability space on which we can define a stochastic process  $\xi^x$  with distribution  $\phi_{\xi^x}$  such that

$$\phi_{\xi^x}(B) = \text{Prob}(\xi^x \in B) := \Pi_{h_1, h_2, \dots}^\infty(x, B) \quad \text{for } B \in B_{X^\infty}.$$

Therefore,  $\Pi_{h_1, h_2, \dots}^\infty(x, \cdot)$  is the distribution of the non-homogeneous Markov chain  $\xi^x$  on  $X^\infty$  with sequence of transition probability functions  $(\Pi_{h_i}^1)_{i \in N}$  and  $\phi_{\xi_0^x} = \delta_x$ , for  $x \in X$ . If an initial distribution is given by any  $\mu \in M_{\text{fin}}(X)$ , not necessarily by  $\delta_x$ , we define

$$(P_{h_1, h_2, \dots}^\infty \mu)(B) = \int_X \Pi_{h_1, h_2, \dots}^\infty(x, B) \mu(dx) \quad \text{for } B \in B_{X^\infty}.$$

**Definition 3.** Let a family of probability measures  $(\{\Pi_{h_i}^1(x, \cdot) : x \in X\})_{i \in N}$  be given. For every  $i \in N$ , we can set another family of probability measures  $\{C_{h_i}^1((x, y), \cdot) : x, y \in X\}$  on  $X^2$  such that

- $C_{h_i}^1((x, y), A \times X) = \Pi_{h_i}^1(x, A)$  for  $A \in B_X$ ,
- $C_{h_i}^1((x, y), X \times B) = \Pi_{h_i}^1(y, B)$  for  $B \in B_X$ ,

where  $x, y \in X$ . For every  $i \in N$ ,  $\{C_{h_i}^1((x, y), \cdot) : x, y \in X\}$  is called coupling.

## VI. ITERATED FUNCTION SYSTEMS

We consider a continuous function  $S : X \times [0, T] \rightarrow X$  and a sequence of continuous mappings given by  $(T_{h_i})_{i \in N}$  with sequence of constants  $(h_i)_{i \in N}$  established. We assume that  $p : X \times [0, T] \rightarrow [0, \infty)$  is a non-negative and normalized function. For each  $A \in B_X$ , we build a sequence of transition operators, as we did in (11).

Let  $n \in N$ . Note that, for arbitrary  $A \in B_X$ ,  $\Pi_{h_1, \dots, h_n}^n(\cdot, A) : X \rightarrow R$  is measurable by definition. Furthermore,  $\Pi_{h_1, \dots, h_n}^n(x, \cdot) : B_X \rightarrow R$  is a probability measure, for every  $x \in X$ . Hence,  $\Pi_{h_1, \dots, h_n}^n$  is a transition probability function on the  $n$ -th marginal. Thanks to these properties (see Section 1.1, [23]), for every  $n \in N$  and a sequence of constants  $(h_i)_{i \in N}$  fixed, there exists a unique regular Markov operator  $P_{h_1, \dots, h_n}^n$ , for which  $\Pi_{h_1, \dots, h_n}^n$  is a transition probability function, and it is given by the formula

$$(P_{h_1, \dots, h_n}^n \mu)(A) = \int_X \Pi_{h_1, \dots, h_n}^n(x, A) \mu(dx),$$

where  $A \in B_X$ ,  $\mu \in M_1(X)$ . Moreover, a dual operator  $U_{h_1, \dots, h_n}^n : B(X) \rightarrow B(X)$  to  $P_{h_1, \dots, h_n}^n$  is defined as follows

$$(U_{h_1, \dots, h_n}^n f)(x) = \int_X f(y) \Pi_{h_1, \dots, h_n}^n(x, dy).$$

**Remark 4.** According to assumptions (II) and (IV), one may check, although through some tedious computations, that, for every  $n \in N$  and a sequence of constants  $(h_i)_{i \in N}$  fixed,

$$\|\Pi_{h_1, \dots, h_n}^n(x, \cdot) - \Pi_{h_1, \dots, h_n}^n(y, \cdot)\|_{\mathcal{L}} \leq a^n \varrho(x, y) + \varphi(\varrho(x, y)),$$

where  $\varphi$  is given by (5). This indicates weak continuity of the map  $X \ni x \mapsto \Pi_{h_1, \dots, h_n}^n(x, \cdot) \in M_1(X)$ . Now, this property, together with the fact that  $P_{h_1, \dots, h_n}^n$  is a regular Markov operator, implies that  $P_{h_1, \dots, h_n}^n$  is even Feller (see Chapter 6, [16]).

However, these estimates do not give us any proper result about stability of the model. That is why we still need to use some coupling methods.

Repeating the construction from the previous section, we obtain  $P_{h_1, h_2, \dots}^\infty \mu$  for  $\mu \in M_1(X)$ . Obviously, for every  $n \in N$ ,  $P_{h_1, \dots, h_n}^n \mu$  is the  $n$ -th marginal of  $P_{h_1, h_2, \dots}^\infty \mu$ .

Fix  $\bar{x} \in X$  for which assumption (III) holds. We define  $V : X \rightarrow [0, \infty)$  to be

$$V(x) = \varrho(x, \bar{x}).$$

**Lemma 5.** For every  $n \in N$  and a sequence of constants  $(h_i)_{i \in N}$  fixed, if  $\mu \in M_1^1(X)$ , then  $P_{h_1, \dots, h_n}^n \mu \in M_1^1(X)$ . Moreover,

$$\langle V, P_{h_1, \dots, h_n}^n \mu \rangle \leq a^n \langle V, \mu \rangle + \frac{1}{1-a} c,$$

where  $c$  does not depend on the sequence  $(h_i)_{i \in N}$ .



*Proof.* Recall that  $a < 1$  and  $c$  are given by (3) i (10), respectively. The state  $\tilde{x}_n^x$  is of the form (9). Following (2), we obtain

$$\begin{aligned}
& \langle V, P_{h_1, \dots, h_n}^n \mu \rangle \\
&= \int_X \left[ \int_0^T \dots \int_0^T \varrho(\tilde{x}_n^x, \bar{x}) p(\tilde{x}_{n-1}^x, t_n) p(\tilde{x}_{n-2}^x, t_{n-1}) \dots p(x, t_1) dt_n \dots dt_1 \right] \mu(dx) \\
&\leq \int_X \int_0^T \dots \int_0^T \left[ \varrho(\tilde{x}_n^x, \tilde{x}_n^{\bar{x}}) + \varrho(\tilde{x}_n^{\bar{x}}, \bar{x}) \right] p(\tilde{x}_{n-1}^x, t_n) p(\tilde{x}_{n-2}^x, t_{n-1}) \dots p(x, t_1) dt_n \dots dt_1 \mu(dx) \\
&\leq \int_X \int_0^T \dots \int_0^T \left[ \lambda(\tilde{x}_{n-1}^x, t_n) \lambda(\tilde{x}_{n-2}^x, t_{n-1}) \dots \lambda(x, t_1) \varrho(x, \bar{x}) + \varrho(\tilde{x}_n^x, \tilde{x}_{n-1}^{\bar{x}}) + \dots + \varrho(\tilde{x}_1^{\bar{x}}, \bar{x}) \right] \\
&\quad p(\tilde{x}_{n-1}^x, t_n) p(\tilde{x}_{n-2}^x, t_{n-1}) \dots p(x, t_1) dt_n \dots dt_1 \mu(dx) \\
&\leq \int_X a^n \varrho(x, \bar{x}) + c(a^n + \dots + 1) \mu(dx) \\
&\leq a^n \langle V, \mu \rangle + \frac{c}{1-a},
\end{aligned}$$

which completes the proof.  $\square$

Fix probability measures  $\mu, \nu \in M_1^1(X)$  and Borel sets  $A, B \in B_X$ . We consider  $b \in M_{\text{fin}}(X^2)$  such that

$$b(A \times X) = \mu(A), \quad b(X \times B) = \nu(B)$$

and  $b_{h_1, \dots, h_n}^n \in M_{\text{fin}}(X^2)$  such that, for every  $n \in N$ ,

$$b_{h_1, \dots, h_n}^n(A \times X) = (P_{h_1, \dots, h_n}^n \mu)(A), \quad b_{h_1, \dots, h_n}^n(X \times B) = (P_{h_1, \dots, h_n}^n \nu)(B).$$

Furthermore, we define  $\bar{V} : X^2 \rightarrow [0, \infty)$

$$\bar{V}(x, y) = V(x) + V(y) \quad \text{for } x, y \in X.$$

Note that, for every  $n \in N$ ,

$$\langle \bar{V}, b_{h_1, \dots, h_n}^n \rangle \leq a \langle \bar{V}, b_{h_1, \dots, h_{n-1}}^{n-1} \rangle + 2c \leq a^n \langle \bar{V}, b \rangle + \frac{2}{1-a} c. \quad (13)$$

For measures  $b \in M_{\text{fin}}^1(X^2)$  finite on  $X^2$  and with finite first moment, we define the linear functional

$$\phi(b) = \int_{X^2} \varrho(x, y) b(dx \times dy).$$

Following the above definitions, we easily obtain

$$\phi(b) \leq \langle \bar{V}, b \rangle. \quad (14)$$

## VII. COUPLING FOR ITERATED FUNCTION SYSTEMS

On  $X^2$  we define the transition sub-probability functions such that, for  $A, B \in B_X$ ,

$$Q_{h_i}^1((x, y), A \times B) = \int_0^T \min\{p(x, t), p(y, t)\} \delta_{(T_{h_i}(x, t), T_{h_i}(y, t))} (A \times B) dt, \quad i \in N, \quad (15)$$



and

$$Q_{h_1, \dots, h_n}^n((x, y), A \times B) = \int_{X^2} Q_{h_n}^1((u, v), A \times B) Q_{h_1, \dots, h_{n-1}}^{n-1}((x, y), du \times dv), \quad n \in N. \quad (16)$$

Measures generated by the transition functions defined above are, by convention, denoted with the same letter. Every time, the context should indicate what we mean. It is easy to check that, for every  $i \in N$ ,

$$Q_{h_i}^1((x, y), A \times X) \leq \int_0^T p(x, t) \delta_{T_{h_i}(x, t)}(A) dt = \int_0^T 1_A(T_{h_i}(x, t)) p(x, t) dt = \Pi_{h_i}^1(x, A)$$

and analogously  $Q_{h_i}^1((x, y), X \times B) \leq \Pi_{h_i}^1(y, B)$ . Similarly, for  $n \in N$ ,

$$Q_{h_1, \dots, h_n}^n((x, y), A \times X) \leq \Pi_{h_1, \dots, h_n}^n(x, A), \quad Q_{h_1, \dots, h_n}^n((x, y), X \times B) \leq \Pi_{h_1, \dots, h_n}^n(y, B).$$

For  $b \in M_{\text{fin}}(X^2)$ , let  $Q_{h_1, \dots, h_n}^n b$  denote the measure

$$(Q_{h_1, \dots, h_n}^n b)(A \times B) = \int_{X^2} Q_{h_1, \dots, h_n}^n((x, y), A \times B) b(dx \times dy) \quad \text{for } A, B \in B_X, n \in N. \quad (17)$$

Note that, for every  $A, B \in B_X$  and  $n \in N$ , we obtain

$$\begin{aligned} (Q_{h_1, \dots, h_{n+1}}^{n+1} b)(A \times B) &= \int_{X^2} Q_{h_1, \dots, h_{n+1}}^{n+1}((x, y), A \times B) b(dx \times dy) \\ &= \int_{X^2} \int_{X^2} Q_{h_{n+1}}^1((u, v), A \times B) Q_{h_1, \dots, h_n}^n((x, y), du \times dv) b(dx \times dy) \\ &= \int_{X^2} Q_{h_{n+1}}^1((u, v), A \times B) (Q_{h_1, \dots, h_n}^n b)(du \times dv) = (Q_{h_{n+1}}^1(Q_{h_1, \dots, h_n}^n b))(A \times B). \end{aligned} \quad (18)$$

Again, following (11) and (12), we are able to construct measures on products and, as a consequence, a measure  $Q_{h_1, h_2, \dots}^\infty b$  on  $X^\infty$ , for every  $b \in M_{\text{fin}}(X^2)$ . Now, we check that, for  $n \in N$  and  $b \in M_{\text{fin}}^1(X^2)$ ,

$$\phi(Q_{h_1, \dots, h_n}^n b) \leq a^n \phi(b). \quad (19)$$

Let us observe that

$$\begin{aligned} \phi(Q_{h_1, \dots, h_n}^n b) &= \int_{X^2} \int_{X^2} \varrho(u, v) Q_{h_1, \dots, h_n}^n((x, y), du \times dv) b(dx \times dy) \\ &= \int_{X^2} \int_{X^2} \int_0^T \int_{X^2} \varrho(u, v) \min\{p(\bar{u}, t), p(\bar{v}, t)\} \delta_{(T_{h_n}(\bar{u}, t), T_{h_n}(\bar{v}, t))} (du \times dv) dt \\ &\quad Q_{h_1, \dots, h_{n-1}}^{n-1}((x, y), d\bar{u} \times d\bar{v}) b(dx \times dy) \\ &\leq \int_{X^2} \int_{X^2} \int_0^T \varrho(T_{h_n}(\bar{u}, t), T_{h_n}(\bar{v}, t)) p(\bar{u}, t) dt Q_{h_1, \dots, h_{n-1}}^{n-1}((x, y), d\bar{u} \times d\bar{v}) b(dx \times dy) \\ &\leq \int_{X^2} \int_{X^2} \int_0^T \varrho(\bar{u}, \bar{v}) \lambda(\bar{u}, t) p(\bar{u}, t) dt Q_{h_1, \dots, h_{n-1}}^{n-1}((x, y), d\bar{u} \times d\bar{v}) b(dx \times dy) \\ &\leq a \int_{X^2} \int_{X^2} \varrho(\bar{u}, \bar{v}) Q_{h_1, \dots, h_{n-1}}^{n-1}((x, y), d\bar{u} \times d\bar{v}) b(dx \times dy) \leq \dots \leq a^n \phi(b). \end{aligned}$$

For every  $i \in N$ , we can find a measure  $R_{h_i}^1((x, y), \cdot)$  such that the sum of  $Q_{h_i}^1((x, y), \cdot)$  and  $R_{h_i}^1((x, y), \cdot)$  gives a new coupling measure  $C_{h_i}^1((x, y), \cdot)$ .

**Lemma 6.** Fix  $i \in N$ . There exists the family  $\{R_{h_i}^1((x, y), \cdot) : x, y \in X\}$  of measures on  $X^2$  such that we can define

$$C_{h_i}^1((x, y), \cdot) = Q_{h_i}^1((x, y), \cdot) + R_{h_i}^1((x, y), \cdot) \quad \text{for } x, y \in X$$

and, moreover,

- (i) the mapping  $(x, y) \mapsto R_{h_i}^1((x, y), A \times B)$  is measurable for every  $A, B \in B_X$ ;
- (ii) measures  $R_{h_i}^1((x, y), \cdot)$  are non-negative for  $x, y \in X$ ;
- (iii) measures  $C_{h_i}^1((x, y), \cdot)$  are probabilistic for every  $x, y \in X$  and so  $\{C_{h_i}^1((x, y), \cdot) : x, y \in X\}$  is a transition probability function on  $X^2$ ;
- (iv) for every  $A, B \in B_X$  and  $x, y \in X$ , we get  $C_{h_i}^1((x, y), A \times X) = \Pi_{h_i}^1(x, A)$  and  $C_{h_i}^1((x, y), X \times B) = \Pi_{h_i}^1(y, B)$ .

*Proof.* Fix  $A, B \in B_X$ . Let

$$\begin{aligned} & R_{h_i}^1((x, y), A \times B) \\ &= (1 - Q_{h_i}^1((x, y), X^2))^{-1} (\Pi_{h_i}^1(x, A) - Q_{h_i}^1((x, y), A \times X)) (\Pi_{h_i}^1(y, B) - Q_{h_i}^1((x, y), X \times B)) \end{aligned}$$

if  $Q_{h_i}^1((x, y), X^2) < 1$  and  $R_{h_i}^1((x, y), A \times B) = 0$  if  $Q_{h_i}^1((x, y), X^2) = 1$ . Obviously, the formula may be extended to the measure. The mapping has all desirable properties (i) – (iv).  $\square$

Lemma 6 shows that, for every  $i \in N$ , we may construct the coupling  $\{C_{h_i}^1((x, y), \cdot) : x, y \in X\}$  for  $\{\Pi_{h_i}^1(x, \cdot) : x \in X\}$  such that  $Q_{h_i}^1((x, y), \cdot) \leq C_{h_i}^1((x, y), \cdot)$ , whereas measures  $R_{h_i}^1((x, y), \cdot)$  are non-negative. Following the rules given in (11), (12), as well as the whole construction from Section V, we easily obtain the family of probability measures  $\{C_{h_1, h_2, \dots}^\infty((x, y), \cdot) : x, y \in X\}$  on  $(X^2)^\infty$  with marginals  $\Pi_{h_1, h_2, \dots}^\infty(x, \cdot)$  and  $\Pi_{h_1, h_2, \dots}^\infty(y, \cdot)$ . This construction appears in [6]. Note that, for every  $n \in N$  and  $x, y \in X$ ,  $C_{h_1, \dots, h_n}^n((x, y), \cdot)$ , constructed as in (11), is the  $n$ -th marginal of  $C_{h_1, h_2, \dots}^\infty((x, y), \cdot)$ . Additionally,  $\{C_{h_1, \dots, h_n}^n((x, y), \cdot) : x, y \in X\}$  fulfills the role of coupling for  $\{\Pi_{h_1, \dots, h_n}^n(x, \cdot) : x \in X\}$ . Indeed, for  $A \in B_X$ ,

$$\begin{aligned} C_{h_1, \dots, h_n}^n((x, y), A \times X) &= \int_{X^2} C_{h_n}^1((u, v), A \times X) C_{h_1, \dots, h_{n-1}}^{n-1}((x, y), du \times dv) \\ &= \int_{X^2} \Pi_{h_n}^1(u, A) C_{h_1, \dots, h_{n-1}}^{n-1}((x, y), du \times dv) = \dots = \Pi_{h_1, \dots, h_n}^n(x, A) \end{aligned}$$

and, similarly,  $C_{h_1, \dots, h_n}^n((x, y), X \times B) = \Pi_{h_1, \dots, h_n}^n(y, B)$ .

Fix  $(x_0, y_0) \in X^2$  and  $(h_n)_{n \in N} \subset [0, \varepsilon]$ . The sequence of transition probability functions  $\left( \{C_{h_1, \dots, h_n}^n((x, y), \cdot) : x, y \in X\} \right)_{n \in N}$  defines the non-homogenous Markov chain  $\Psi$  on  $X^2$  with starting point  $(x_0, y_0)$ , while the sequence of transition probability functions  $\left( \{\hat{C}_{h_1, \dots, h_n}^n((x, y, \theta), \cdot) : x, y \in X, \theta \in \{0, 1\}\} \right)_{n \in N}$  defines the Markov chain  $\hat{\Psi}$  on the augmented space  $X^2 \times \{0, 1\}$  with initial distribution  $\hat{C}^0((x_0, y_0), \cdot) = \delta_{(x_0, y_0, 1)}(\cdot)$ . If  $\hat{\Psi}_n = (x, y, i)$ , where  $x, y \in X, i \in \{0, 1\}$ , then

$$\text{Prob}(\hat{\Psi}_{n+1} \in A \times B \times \{1\} \mid \hat{\Psi}_n = (x, y, i), i \in \{0, 1\}) = Q_{h_1, \dots, h_n}^n((x, y), A \times B),$$

$$\text{Prob}(\hat{\Psi}_{n+1} \in A \times B \times \{0\} \mid \hat{\Psi}_n = (x, y, i), i \in \{0, 1\}) = R_{h_1, \dots, h_n}^n((x, y), A \times B),$$

where  $A, B \in B_X$ . Once again, we refer to (11), (12) and the Kolmogorov theorem to obtain the measure  $\hat{C}_{h_1, h_2, \dots}^\infty((x_0, y_0), \cdot)$  on  $(X^2 \times \{0, 1\})^\infty$  which is associated with the Markov chain  $\hat{\Psi}$ .

From now on, we assume that processes  $\Psi$  and  $\hat{\Psi}$  taking values in  $X^2$  and  $X^2 \times \{0, 1\}$ , respectively, are defined on  $(\Omega, F, \mathbf{P})$ . The expected value of measures  $C_{h_1, h_2, \dots}^\infty((x_0, y_0), \cdot)$ ,  $\hat{C}_{h_1, h_2, \dots}^\infty((x_0, y_0), \cdot)$  is denoted by  $E_{x_0, y_0}$ .

### VIII. AUXILIARY THEOREMS

Recall that  $a$  is given by (3). Fix  $\varkappa \in (0, 1 - a)$ . Set

$$K_\varkappa = \{(x, y) \in X^2 : \bar{V}(x, y) < \varkappa^{-1}2c\},$$

where  $c$  is given by (10). Let  $d : (X^2)^\infty \rightarrow N$  denote the time of the first visit in  $K_\varkappa$ , i.e.

$$d((x_n, y_n)_{n \in N}) = \inf\{n \in N : (x_n, y_n) \in K_\varkappa\}.$$

As a convention, we put  $d((x_n, y_n)_{n \in N}) = \infty$ , if there is no  $n \in N$  such that  $(x_n, y_n) \in K_\varkappa$ .

**Theorem 7.** *For every  $\zeta \in (0, 1)$  there exist positive constants  $C_1, C_2$  such that*

$$E_{x_0, y_0}[(a + \varkappa)^{-\zeta d}] \leq C_1 \bar{V}(x_0, y_0) + C_2.$$

*Proof.* Fix  $(x_0, y_0) \in X^2$ . Let  $\Psi = (x_n, y_n)_{n \in N}$  be the Markov chain with starting point  $(x_0, y_0)$  and sequence of transition probability functions  $(\{C_{h_i}^1((x, y), \cdot) : x, y \in X\})_{i \in N}$ . Let  $F_n \subset F$ ,  $n \in N$ , be the natural filtration in  $\Omega$  associated with  $\Psi$ . We define

$$A_n = \{\omega \in \Omega : \Psi_i = (x_i(\omega), y_i(\omega)) \notin K_\varkappa \text{ for } i = 1, \dots, n\}, \quad n \in N.$$

Obviously,  $A_{n+1} \subset A_n$  and  $A_n \in F_n$ , for  $n \in N$ . In consequence of (13), as well as the definitions of  $A_n$  and  $K_\varkappa$ , the following inequalities are  $\mathbf{P}$ -a.s. satisfied in  $\Omega$ :

$$1_{A_n} E_{x_0, y_0}[\bar{V}(x_{n+1}, y_{n+1}) | F_n] \leq 1_{A_n} (a \bar{V}(x_n, y_n) + 2c) \leq 1_{A_n} (a + \varkappa) \bar{V}(x_n, y_n).$$

Accordingly, we obtain

$$\begin{aligned} \int_{A_n} \bar{V}(x_n, y_n) d\mathbf{P} &\leq \int_{A_{n-1}} \bar{V}(x_n, y_n) d\mathbf{P} = \int_{A_{n-1}} E[\bar{V}(x_n, y_n) | F_{n-1}] d\mathbf{P} \\ &\leq \int_{A_{n-1}} [a \bar{V}(x_{n-1}, y_{n-1}) + 2c] d\mathbf{P} \leq (a + \varkappa) \int_{A_{n-1}} \bar{V}(x_{n-1}, y_{n-1}) d\mathbf{P}. \end{aligned}$$

On applying these estimates finitely many times, we obtain

$$\int_{A_n} \bar{V}(x_n, y_n) d\mathbf{P} \leq (a + \varkappa)^{n-1} \int_{A_1} \bar{V}(x_1, y_1) d\mathbf{P} \leq (a + \varkappa)^{n-1} [a \bar{V}(x_0, y_0) + 2c].$$

Note that

$$\mathbf{P}(A_n) \leq \int_{A_n} \varkappa(2c)^{-1} \bar{V}(x_n, y_n) d\mathbf{P} \leq \varkappa [2c(a + \varkappa)]^{-1} (a + \varkappa)^n [a \bar{V}(x_0, y_0) + 2c].$$

Set  $\hat{c} := \varkappa [2c(a + \varkappa)]^{-1} [a \bar{V}(x_0, y_0) + 2c]$ . Then,  $\mathbf{P}(A_n) \leq (a + \varkappa)^n \hat{c}$ . Fix  $\zeta \in (0, 1)$ . Since  $d$  takes natural values  $n \in N$ , we obtain

$$\sum_{n=1}^{\infty} (a + \varkappa)^{-\zeta n} \mathbf{P}(A_n) \leq \sum_{n=1}^{\infty} (a + \varkappa)^{-\zeta n} (a + \varkappa)^n \hat{c} = \sum_{n=1}^{\infty} (a + \varkappa)^{(1-\zeta)n} \hat{c},$$

which implies convergence of the series. The proof is complete by the definition of  $\hat{c}$  and with properly chosen  $C_1, C_2$ .  $\square$

For every positive  $r > 0$ , we define the set

$$C_r = \{(x, y) \in X^2 : \varrho(x, y) < r\}.$$

**Lemma 8.** Fix  $\tilde{a} \in (a, 1)$ . Let  $C_r$  be the set defined above and suppose that  $b \in M_{fn}(X^2)$  is such that  $\text{supp } b \subset C_r$ . There exists  $\bar{\gamma} > 0$  such that

$$(Q_{h_1, \dots, h_n}^n b)(C_{\tilde{a}^n r}) \geq \bar{\gamma}^n \|b\|$$

for  $\delta$  and  $M$  defined in assumption (V) (see Section III).

*Proof.* Recall that  $\tilde{x}_n^x$  is given by (9). Directly from (17), (16) and (15) we obtain

$$\begin{aligned} & (Q_{h_1, \dots, h_n}^n b)(C_{\tilde{a}^n r}) \\ &= \int_{X^2} \int_{X^2} \int_0^T \min\{p(u, t_n), p(v, t_n)\} \delta_{(T_{h_n}(u, t_n), T_{h_n}(v, t_n))} (C_{\tilde{a}^n r}) dt_n Q_{h_1, \dots, h_{n-1}}^{n-1}((x, y), du \times dv) b(dx \times dy) \\ &= \int_{X^2} \left[ \int_{(0, T)^n} 1_{C_{\tilde{a}^n r}}(\tilde{x}_n^x, \tilde{x}_n^y) \min\{p(\tilde{x}_{n-1}^x, t_n), p(\tilde{x}_{n-1}^y, t_n)\} \dots \min\{p(x, t_1), p(y, t_1)\} dt_n \dots dt_1 \right] b(dx \times dy). \end{aligned}$$

Note that  $1_{C_{\tilde{a}^n r}}(\tilde{x}_n^x, \tilde{x}_n^y) = 1$  if and only if  $(t_1, \dots, t_n) \in \mathcal{T}_n$ , where

$$\mathcal{T}_n := \{(t_1, \dots, t_n) \in (0, T)^n : \varrho(\tilde{x}_n^x, \tilde{x}_n^y) < \tilde{a}^n r\}.$$

Set  $\mathcal{T}'_n := (0, T)^n \setminus \mathcal{T}_n$ . Note that, according to assumption (II), we have

$$\int_{\mathcal{T}'_n} \varrho(\tilde{x}_n^x, \tilde{x}_n^y) p(\tilde{x}_{n-1}^x, t_n) \dots p(x, t_1) dt_n \dots dt_1 \leq a^n \varrho(x, y) < a^n r$$

for  $(x, y) \in C_r$ . Comparing this with the definition of  $\mathcal{T}'_n$ , we obtain

$$\tilde{a}^n r \int_{\mathcal{T}'_n} p(\tilde{x}_{n-1}^x, t_n) \dots p(x, t_1) dt_n \dots dt_1 < a^n r,$$

which implies

$$\int_{\mathcal{T}'_n} p(\tilde{x}_{n-1}^x, t_n) \dots p(x, t_1) dt_n \dots dt_1 < \frac{a^n}{\tilde{a}^n} < 1.$$

We then obtain that the integral over  $\mathcal{T}_n$  is not less than  $1 - \left(\frac{a}{\tilde{a}}\right)^n \geq (1 - \frac{a}{\tilde{a}})^n =: \gamma^n$ , for sufficiently big  $n \in N$ , which provides, using assumption (V), that  $|\mathcal{T}_n| \geq \left(\frac{\gamma}{M}\right)^n$ . Finally,

$$(Q_{h_1, \dots, h_n}^n b)(C_{\tilde{a}^n r}) \geq \int_{X^2} \delta^n |\mathcal{T}_n| b(dx, dy) \geq \delta^n \left(\frac{\gamma}{M}\right)^n \|b\|.$$

If we set  $\bar{\gamma} := \delta M^{-1} \gamma$ , the proof is complete.  $\square$

**Theorem 9.** For every  $\varkappa \in (0, 1 - a)$ , there exists  $n_0 \in N$  such that

$$\|Q_{h_1, h_2, \dots}^\infty((x, y), \cdot)\| \geq \frac{1}{2} \bar{\gamma}^{n_0} \quad \text{for } (x, y) \in K_\varkappa,$$

where  $\bar{\gamma} > 0$  is given in Lemma 8.

*Proof.* Note that, for every real numbers  $u, v \in R$ , there is a general rule:  $\min\{u, v\} + |u - v| - u \geq 0$ . Hence, for every  $(x, y) \in X^2$  and  $i \in N$ , we obtain

$$\int_0^T \left[ \min\{p(x, t), p(y, t)\} + |p(x, t) - p(y, t)| - p(x, t) \right] dt \geq 0$$

and therefore, due to (15),

$$\|Q_{h_i}^1((x, y), \cdot)\| + \int_0^T |p(x, t) - p(y, t)| dt \geq 1.$$

For every  $b \in M_{\text{fin}}(X^2)$ , due to the Dini condition (see assumption (IV)) and the Jensen inequality, we get

$$\begin{aligned} \|Q_{h_i}^1 b\| &= \int_{X^2} Q_{h_i}^1((x, y), X^2) b(dx \times dy) = \int_{X^2} \|Q_{h_i}^1((x, y), \cdot)\| b(dx \times dy) \\ &\geq \|b\| - \int_{X^2} \omega(\varrho(x, y)) b(dx \times dy) \geq \|b\| - \omega(\phi(b)). \end{aligned}$$

Then, by (18),

$$\begin{aligned} \|Q_{h_1, \dots, h_n}^n b\| &= \int_{X^2} Q_{h_n}^1((x, y), \cdot) (Q_{h_1, \dots, h_{n-1}}^{n-1} b)(dx \times dy) \geq \|Q_{h_1, \dots, h_{n-1}}^{n-1} b\| - \omega(\phi(Q_{h_1, \dots, h_{n-1}}^{n-1} b)) \\ &\geq \|b\| - \omega(\phi(Q_{h_1}^1 b)) - \dots - \omega(\phi(Q_{h_1, \dots, h_n}^n b)). \end{aligned}$$

Following (19) and recalling that  $\omega$  is non-decreasing, we obtain

$$\|Q_{h_1, \dots, h_n}^n b\| \geq \|b\| - \sum_{i=1}^n \omega(a^{i-1} \phi(b)).$$

See (5) to recall the definition of  $\varphi$ . Thanks to assumption (IV), we know that  $\lim_{t \rightarrow 0} \varphi(t) = 0$ . Hence, we may choose  $r > 0$  such that if  $\varrho(x, y) < r$  and therefore  $a^{-1} \phi(b) \leq r a^{-1} \|b\|$ , then  $\sum_{i=1}^n \omega(a^{i-1} \phi(b)) \leq \varphi(a^{-1} \phi(b)) < \frac{1}{2} \|b\|$ .

If  $\text{supp } b \subset C_r$ , then we obtain

$$\|Q_{h_1, h_2, \dots}^\infty b\| \geq \frac{\|b\|}{2}. \quad (20)$$

Fix  $\varkappa \in (0, 1 - a)$ . It is clear that  $K_\varkappa \subset C_{\varkappa^{-1}2c}$ . If we define  $n_0 := \min\{n \in N : a^n(\varkappa)^{-1}2c < r\}$ , then  $C_{a^{n_0}\varkappa^{-1}2c} \subset C_r$ . Remembering that  $Q_{h_1, \dots, h_n, h_{n+1}, \dots, h_{n+m}}^{n+m}((x, y), \cdot) = (Q_{h_{n+1}, \dots, h_{n+m}}^m Q_{h_1, \dots, h_n}^n)((x, y), \cdot)$  and using the Markov property, we obtain

$$Q_{h_1, h_2, \dots}^\infty((x, y), X^2) = (Q_{h_{n_0+1}, \dots}^\infty Q_{h_1, \dots, h_{n_0}}^{n_0})((x, y), X^2).$$

Then, according to (20) and Lemma 8, we obtain

$$\begin{aligned} \|Q_{h_1, h_2, \dots}^\infty((x, y), \cdot)\| &= \|(Q_{h_{n_0+1}, \dots}^\infty Q_{h_1, \dots, h_{n_0}}^{n_0})((x, y), \cdot)\| \geq \frac{\|Q_{h_1, \dots, h_{n_0}}^{n_0}((x, y), \cdot)|_{C_r}\|}{2} \\ &= \frac{Q_{h_1, \dots, h_{n_0}}^{n_0}((x, y), C_r)}{2} \geq \frac{Q_{h_1, \dots, h_{n_0}}^{n_0}((x, y), C_{a^{n_0}\varkappa^{-1}2c})}{2} \geq \frac{\bar{\gamma}^{n_0}}{2} \end{aligned}$$

for  $(x, y) \in K_\varkappa$ . This finishes the proof.  $\square$

**Definition 10.** Coupling time  $\tau : (X^2 \times \{0, 1\})^\infty \rightarrow N$  is defined as follows

$$\tau((x_n, y_n, \theta_n)_{n \in N}) = \inf\{n \in N : \theta_k = 1 \text{ for } k \geq n\}.$$

As a convention, we put  $\tau((x_n, y_n, \theta_n)_{n \in N}) = \infty$ , if there is no  $n \in N$  such that  $\theta_k = 1$  for every  $k \geq n$ .

**Theorem 11.** There exist  $\tilde{q} \in (0, 1)$  and  $C_3 > 0$  such that

$$E_{x,y}[\tilde{q}^{-\tau}] \leq C_3(1 + \bar{V}(x, y)) \quad \text{for } (x, y) \in X^2.$$

*Proof.* Fix  $\varkappa \in (0, 1 - a)$  and  $(x, y) \in X^2$ . To simplify notation, we write  $\alpha = (a + \varkappa)^{-\frac{1}{2}}$ . Let  $d$  be the random moment of the first visit in  $K_\varkappa$ . Suppose that

$$d_1 = d, \quad d_{n+1} = d_n + d \circ \Gamma_{d_n},$$

where  $n \in N$  and  $\Gamma_n$  are shift operators on  $(X^2 \times \{0, 1\})^\infty$ , i.e.  $\Gamma_n((x_k, y_k, \theta_k)_{k \in N}) = (x_{k+n}, y_{k+n}, \theta_{k+n})_{k \in N}$ . Theorem 7 implies that every  $d_n$  is  $C_{h_1, h_2, \dots}^\infty((x, y), \cdot)$ -a.s. finished. The strong Markov property shows that

$$E_{x,y}[\alpha^d \circ \Gamma_{d_n} | F_{d_n}] = E_{(x_{d_n}, y_{d_n})}[\alpha^d] \quad \text{for } n \in N,$$

where  $F_{d_n}$  denotes the  $\sigma$ -algebra on  $(X^2 \times \{0, 1\})$  generated by  $d_n$  and  $\Psi = (x_n, y_n)_{n \in N}$  is the non-homogenous Markov chain with sequence of transition probability functions  $(\{C_{h_i}^1((x, y), \cdot) : x, y \in X\})_{i \in N}$ . By Theorem 7 and the definition of  $K_\varkappa$ , we obtain

$$E_{x,y}[\alpha^{d_{n+1}}] = E_{x,y}[\alpha^{d_n} E_{(x_{d_n}, y_{d_n})}[\alpha^d]] \leq E_{x,y}[\alpha^{d_n}](C_1 \varkappa^{-1} 2c + C_2).$$

Fix  $\eta = C_1 \varkappa^{-1} 2c + C_2$ . Consequently,

$$E_{x,y}[\alpha^{d_{n+1}}] \leq \eta^n E_{x,y}[\alpha^d] \leq \eta^n [C_1 \bar{V}(x, y) + C_2]. \quad (21)$$

We define  $\hat{\tau}((x_n, y_n, \theta_n)_{n \in N}) = \inf\{n \in N : (x_n, y_n) \in K_\varkappa, \theta_k = 1 \text{ for } k \geq n\}$  and  $\sigma = \inf\{n \in N : \hat{\tau} = d_n\}$ . By Theorem 9, there is  $n_0 \in N$  such that

$$\hat{C}_{h_1, h_2, \dots}^\infty((x, y), \{\sigma > n\}) \leq (1 - \frac{\bar{\gamma}^{n_0}}{2})^n \quad \text{for } n \in N. \quad (22)$$

Let  $p > 1$ . By the Hölder inequality, (21) and (22), we obtain

$$\begin{aligned} E_{x,y}[\alpha^{\frac{\hat{\tau}}{p}}] &\leq \sum_{k=1}^{\infty} E_{x,y}[\alpha^{\frac{d_k}{p}} 1_{\sigma=k}] \leq \sum_{k=1}^{\infty} \left( E_{x,y}[\alpha^{d_k}] \right)^{\frac{1}{p}} \left( \hat{C}_{h_1, h_2, \dots}^\infty((x, y), \{\sigma = k\}) \right)^{(1-\frac{1}{p})} \\ &\leq [C_1 \bar{V}(x, y) + C_2]^{\frac{1}{p}} \eta^{-\frac{1}{p}} \sum_{k=1}^{\infty} \eta^{\frac{k}{p}} (1 - \frac{1}{2} \bar{\gamma}^{n_0})^{(k-1)(1-\frac{1}{p})} \\ &= [C_1 \bar{V}(x, y) + C_2]^{\frac{1}{p}} \eta^{-\frac{1}{p}} (1 - \frac{1}{2} \bar{\gamma}^{n_0})^{-(1-\frac{1}{p})} \sum_{k=1}^{\infty} \left[ \left( \frac{\eta}{1 - \frac{1}{2} \bar{\gamma}^{n_0}} \right)^{\frac{1}{p}} (1 - \frac{1}{2} \bar{\gamma}^{n_0}) \right]^k. \end{aligned}$$

For  $p$  sufficiently large and  $\tilde{q} = \alpha^{-\frac{1}{p}}$ , we get

$$E_{x,y}[\tilde{q}^{-\hat{\tau}}] = E_{x,y}[\alpha^{\frac{\hat{\tau}}{p}}] \leq (1 + \bar{V}(x, y)) C_3$$

for some  $C_3$ . Since  $\tau \leq \hat{\tau}$ , we finish the proof.  $\square$

**Lemma 12.** *Let  $f \in \mathcal{L}$ . Then, there exist  $q \in (0, 1)$  and  $C_5 > 0$  such that*

$$\int_{X^2} |f(u) - f(v)| (\Pi_{X^2}^* \Pi_n^* \hat{C}_{h_1, h_2, \dots}^\infty((x, y), \cdot)) (du \times dv) \leq q^n C_5 (1 + \bar{V}(x, y)) \text{ for every } x, y \in X, n \in N,$$

where  $\Pi_n^* : (X^2 \times \{0, 1\})^\infty \rightarrow X^2 \times \{0, 1\}$  are the projections on the  $n$ -th component and  $\Pi_{X^2}^* : X^2 \times \{0, 1\} \rightarrow X^2$  is the projection on  $X^2$ .

*Proof.* For  $n \in N$  we define sets

$$A_{\frac{n}{2}} = \{t \in (X^2 \times \{0, 1\})^\infty : \tau(t) \leq \frac{n}{2}\},$$

$$B_{\frac{n}{2}} = \{t \in (X^2 \times \{0, 1\})^\infty : \tau(t) > \frac{n}{2}\}.$$

Note that  $A_{\frac{n}{2}} \cap B_{\frac{n}{2}} = \emptyset$  and  $A_{\frac{n}{2}} \cup B_{\frac{n}{2}} = (X^2 \times \{0, 1\})^\infty$ , so, for  $n \in N$ , we have

$$\hat{C}_{h_1, h_2, \dots}^\infty((x, y), \cdot) = \hat{C}_{h_1, h_2, \dots}^\infty((x, y), \cdot)|_{A_{\frac{n}{2}}} + \hat{C}_{h_1, h_2, \dots}^\infty((x, y), \cdot)|_{B_{\frac{n}{2}}}.$$

Hence,

$$\begin{aligned} & \int_{X^2} |f(u) - f(v)| (\Pi_{X^2}^* \Pi_n^* \hat{C}_{h_1, h_2, \dots}^\infty((x, y), \cdot)|_{A_{\frac{n}{2}}}) (du \times dv) \\ & \quad + \int_{X^2} |f(u) - f(v)| (\Pi_{X^2}^* \Pi_n^* \hat{C}_{h_1, h_2, \dots}^\infty((x, y), \cdot)|_{B_{\frac{n}{2}}}) (du \times dv) \\ & \leq \int_{X^2} \varrho(u, v) (\Pi_{X^2}^* \Pi_n^* \hat{C}_{h_1, h_2, \dots}^\infty((x, y), \cdot)|_{A_{\frac{n}{2}}}) (du \times dv) + 2\hat{C}_{h_1, h_2, \dots}^\infty((x, y), B_{\frac{n}{2}}). \end{aligned}$$

Note that, by iterative application of (19), we obtain

$$\begin{aligned} \int_{X^2} \varrho(u, v) (\Pi_{X^2}^* \Pi_n^* \hat{C}_{h_1, h_2, \dots}^\infty((x, y), \cdot)|_{A_{\frac{n}{2}}}) (du, dv) &= \phi(\Pi_{X^2}^* \Pi_n^* (\hat{C}_{h_1, h_2, \dots}^\infty((x, y), \cdot)|_{A_{\frac{n}{2}}})) \\ &\leq a^{\lfloor \frac{n}{2} \rfloor} \phi(\Pi_{X^2}^* \Pi_{\lfloor \frac{n+1}{2} \rfloor}^* (\hat{C}_{h_1, h_2, \dots}^\infty((x, y), \cdot)|_{A_{\frac{n}{2}}})) \end{aligned}$$

Then, it follows from (13) and (14) that

$$\phi(\Pi_{X^2}^* \Pi_{\lfloor \frac{n+1}{2} \rfloor}^* (\hat{C}_{h_1, h_2, \dots}^\infty((x, y), \cdot)|_{A_{\frac{n}{2}}})) \leq a^{\lfloor \frac{n+1}{2} \rfloor} \bar{V}(x, y) + \frac{2c}{1-a}.$$

We obtain

$$\begin{aligned} & \int_{X^2} |f(u) - f(v)| (\Pi_{X^2}^* \Pi_n^* \hat{C}_{h_1, h_2, \dots}^\infty((x, y), \cdot)) (du \times dv) \\ & \leq a^{\lfloor \frac{n}{2} \rfloor} \left[ a^{\lfloor \frac{n+1}{2} \rfloor} \bar{V}(x, y) + \frac{2c}{1-a} \right] + 2\hat{C}_{h_1, h_2, \dots}^\infty((x, y), B_{\frac{n}{2}}). \end{aligned}$$

It follows from Theorem 11 and the Chebyshev inequality that

$$\begin{aligned} \hat{C}_{h_1, h_2, \dots}^\infty((x, y), B_{\frac{n}{2}}) &= \hat{C}_{h_1, h_2, \dots}^\infty((x, y), \{\tau > \frac{n}{2}\}) = \hat{C}_{h_1, h_2, \dots}^\infty((x, y), \{\tilde{q}^{-\tau} \geq \tilde{q}^{-\frac{n}{2}}\}) \\ &\leq \frac{E_{x,y}[\tilde{q}^{-\tau}]}{\tilde{q}^{-\frac{n}{2}}} \leq \tilde{q}^{\frac{n}{2}} C_3 (1 + \bar{V}(x, y)), \end{aligned}$$

for some  $\tilde{q} \in (0, 1)$  and  $C_3 > 0$ . Finally,

$$\int_{X^2} |f(u) - f(v)| (\Pi_{X^2}^* \Pi_n^* \hat{C}_{h_1, h_2, \dots}^\infty((x, y), \cdot)) (du \times dv) \leq a^{\lfloor \frac{n}{2} \rfloor} C_4 (1 + \bar{V}(x, y)) + 2\tilde{q}^{\frac{n}{2}} C_3 (1 + \bar{V}(x, y)),$$

where  $C_4 = \max\{a^{\frac{1}{2}}, (1-a)^{-1}2c\}$ . Setting  $q := \max\{a^{\frac{1}{2}}, \tilde{q}^{\frac{1}{2}}\}$  and  $C_5 := C_4 + 2C_3$ , gives our claim.  $\square$



**Remark 13.** If  $g : X \rightarrow R$  is an arbitrary bounded and Lipschitz function with constant  $L_g$ , then, there are  $q \in (0, 1)$  and  $C_5 > 0$ , exactly the same as in Lemma 12, for which we obtain

$$\int_{X^2} |g(u) - g(v)| (\Pi_{X^2}^* \Pi_n^* \hat{C}_{h_1, h_2, \dots}^\infty((x, y), \cdot)) (du \times dv) \leq Gq^n C_5 (1 + \bar{V}(x, y)) \text{ for every } x, y \in X, n \in N,$$

where  $G := \max\{L_g, \sup_{x \in X} |g(x)|\}$ .

**Theorem 14.** There exist  $q \in (0, 1)$  and  $C_5 > 0$  such that

$$\|\Pi_{h_1, \dots, h_n}^n(x, \cdot) - \Pi_{h_1, \dots, h_n}^n(y, \cdot)\|_{\mathcal{L}} \leq q^n C_5 (1 + \bar{V}(x, y)) \text{ for } x, y \in X \text{ and } n \in N.$$

*Proof.* The theorem is a consequence of Lemma 12. It is enough to observe that

$$\begin{aligned} \|\Pi_{h_1, \dots, h_n}^n(x, \cdot) - \Pi_{h_1, \dots, h_n}^n(y, \cdot)\|_{\mathcal{L}} &= \sup_{f \in \mathcal{L}} \left| \int_X f(z) (\Pi_{h_1, \dots, h_n}^n(x, \cdot) - \Pi_{h_1, \dots, h_n}^n(y, \cdot))(dz) \right| \\ &= \sup_{f \in \mathcal{L}} \left| \int_{X^2} (f(z_1) - f(z_2)) (\Pi_{X^2}^* \Pi_n^* \hat{C}_{h_1, h_2, \dots}^\infty((x, y), \cdot))(dz_1 \times dz_2) \right|. \end{aligned}$$

Hence, using the argument of Lemma 12, we obtain

$$\|\Pi_{h_1, \dots, h_n}^n(x, \cdot) - \Pi_{h_1, \dots, h_n}^n(y, \cdot)\|_{\mathcal{L}} \leq q^n C_5 (1 + \bar{V}(x, y)),$$

which finishes the proof.  $\square$

## IX. EXPONENTIAL RATE OF CONVERGENCE - PROOF OF THEOREM 1

Note that we may write

$$P_\varepsilon^n \mu(\cdot) = \int_X \int_{B(0, \varepsilon)} \dots \int_{B(0, \varepsilon)} \Pi_{h_1, \dots, h_n}^n(x, \cdot) \nu^\varepsilon(dh_1) \dots \nu^\varepsilon(dh_n) \mu(dx).$$

Comparing this approach with Remark 1 and Lemma 1, we see that  $P_\varepsilon$  is Feller and, for every  $n \in N$ , it satisfies the following property

$$\langle V, P_\varepsilon^n \mu \rangle \leq a^n \langle V, \mu \rangle + \frac{c}{1 - a}. \quad (23)$$

We present the proof of Theorem 1 below.

*Proof of Theorem 1.* Let  $\mu_1, \mu_2 \in M_1^1(X)$ . We first want to evaluate  $\|P_\varepsilon^n \mu_1 - P_\varepsilon^n \mu_2\|_{\mathcal{L}}$ . Let  $f \in \mathcal{L}$ . We obtain

$$\begin{aligned} &|\langle f, P_\varepsilon^n \mu_1 - P_\varepsilon^n \mu_2 \rangle| \\ &= \left| \int_X \int_{B(0, \varepsilon)} \dots \int_{B(0, \varepsilon)} \int_X f(z) \Pi_{h_1, \dots, h_n}^n(x, dz) \nu^\varepsilon(dh_1) \dots \nu^\varepsilon(dh_n) \mu_1(dx) \right. \\ &\quad \left. - \int_X \int_{B(0, \varepsilon)} \dots \int_{B(0, \varepsilon)} \int_X f(z) \Pi_{h_1, \dots, h_n}^n(y, dz) \nu^\varepsilon(dh_1) \dots \nu^\varepsilon(dh_n) \mu_2(dy) \right| \\ &= \left| \int_X \left[ \int_X \int_{B(0, \varepsilon)} \dots \int_{B(0, \varepsilon)} \int_X f(z) \Pi_{h_1, \dots, h_n}^n(x, dz) \nu^\varepsilon(dh_1) \dots \nu^\varepsilon(dh_n) \mu_1(dx) \right] \mu_2(dy) \right. \\ &\quad \left. - \int_X \left[ \int_X \int_{B(0, \varepsilon)} \dots \int_{B(0, \varepsilon)} \int_X f(z) \Pi_{h_1, \dots, h_n}^n(y, dz) \nu^\varepsilon(dh_1) \dots \nu^\varepsilon(dh_n) \mu_2(dy) \right] \mu_1(dx) \right| \end{aligned}$$

Now, following the result from Theorem 14,

$$\|\Pi_{h_1, \dots, h_n}^n(x, \cdot) - \Pi_{h_1, \dots, h_n}^n(y, \cdot)\|_{\mathcal{L}} \leq q^n C_5 (1 + \bar{V}(x, y)),$$

where  $q$  and  $C_5$  are independent of choice of  $h_1, h_2, \dots$ , we obtain the inequality

$$\begin{aligned} & |\langle f, P_\varepsilon^n \mu_1 - P_\varepsilon^n \mu_2 \rangle| \\ &= \int_X \int_X \left[ \int_{B(0, \varepsilon)} \cdots \int_{B(0, \varepsilon)} \left| \int_X f(z) \Pi_{h_1, \dots, h_n}^n(x, dz) - \int_X f(z) \Pi_{h_1, \dots, h_n}^n(y, dz) \right| \right. \\ & \quad \left. \nu^\varepsilon(dh_1) \cdots \nu^\varepsilon(dh_n) \right] \mu_1(dx) \mu_2(dy) \\ &\leq \int_X \int_X \left[ \int_{B(0, \varepsilon)} \cdots \int_{B(0, \varepsilon)} \|\Pi_{h_1, \dots, h_n}^n(x, \cdot) - \Pi_{h_1, \dots, h_n}^n(y, \cdot)\|_{\mathcal{L}} \nu^\varepsilon(dh_1) \cdots \nu^\varepsilon(dh_n) \right] \mu_1(dx) \mu_2(dy) \\ &\leq q^n C_5 \int_X \int_X \left[ \int_{B(0, \varepsilon)} \cdots \int_{B(0, \varepsilon)} (1 + \bar{V}(x, y)) \nu^\varepsilon(dh_1) \cdots \nu^\varepsilon(dh_n) \right] \mu_1(dx) \mu_2(dy) \\ &= q^n C_5 \int_X \int_X (1 + \bar{V}(x, y)) \mu_1(dx) \mu_2(dy), \end{aligned}$$

where measures  $\mu_1, \mu_2 \in M_1^1(X)$ .

Now, set  $\mu_1 := \mu \in M_1^1(X)$  and  $\mu_2 := P_\varepsilon^m \mu \in M_1^1(X)$ , for arbitrary  $m \in N$ . Note that it follows from Lemma 1 that  $P_\varepsilon^m \mu$  is with finite first moment if  $\mu \in M_1^1$ . We obtain

$$\begin{aligned} \|P_\varepsilon^n \mu - P_\varepsilon^{n+m} \mu\|_{\mathcal{L}} &\leq q^n C_5 \int_X \int_X (1 + \bar{V}(x, y)) \mu(dx) P_\varepsilon^m \mu(dy) \\ &= q^n C_5 \left[ 1 + \int_X V(x) \mu(dx) + \int_X V(y) P_\varepsilon^m \mu(dy) \right] \leq q^n C_6 \end{aligned}$$

for some constant  $C_6$ . Hence,  $(P_\varepsilon^n \mu)_{n \in N}$  is a Cauchy sequence in  $(M_1(X), \|\cdot\|_{\mathcal{L}})$ . It is then proven, because of completeness of the space, that  $(P_\varepsilon^n \mu)_{n \in N}$  converges in  $(M_1(X), \|\cdot\|_{\mathcal{L}})$ . Put  $\mu_*(\mu) = \lim_{n \rightarrow \infty} P_\varepsilon^n \mu$ . As mentioned before, we know that  $P_\varepsilon$  is a Feller operator and this implies that the measure  $\mu_*(\mu)$  is invariant. Then, for  $\mu_1, \mu_2 \in M_1^1(X)$  and every  $\varepsilon > 0$ , we have

$$\|\mu_*(\mu_1) - \mu_*(\mu_2)\|_{\mathcal{L}} \leq \|\mu_*(\mu_1) - P_\varepsilon^n \mu_1\|_{\mathcal{L}} + \|P_\varepsilon^n \mu_1 - P_\varepsilon^n \mu_2\|_{\mathcal{L}} + \|\mu_*(\mu_2) - P_\varepsilon^n \mu_2\|_{\mathcal{L}} < \varepsilon$$

for sufficiently large  $n \in N$ . Hence, we have the invariant measure  $\mu_* := \mu_*(\mu)$  which is unique in  $M_1^1(X)$ . We should make it clear that  $\mu_* \in M_1^1(X)$ . Note that we can take a non-decreasing sequence  $(V_k)_{k \in N}$  such that  $V_k(y) = \min\{k, V(y)\}$ , for every  $k \in N$  and  $y \in X$ . Fix  $x \in X$ . From the first part of the proof, we know that  $\langle f, P_\varepsilon^n \delta_x \rangle$  converges to  $\langle f, \mu_* \rangle$  for every  $f \in \mathcal{L}$ , which means, by the Aleksandrov theorem (see Theorem 11.3.3 in [3]), that  $P_\varepsilon^n \delta_x$  converges weakly to  $\mu_*$ , since both measures are probabilistic. Hence, for all  $k \in N$ ,  $V_k \in C(X)$  and we obtain

$$\lim_{n \rightarrow \infty} \int_X V_k(y) P_\varepsilon^n \delta_x(dy) = \int_X V_k(y) \mu_*(dy).$$

Note that, according to (23), for every  $n \in N$ ,

$$\langle V_k, P_\varepsilon^n \delta_x \rangle = a^n \langle V_k, \delta_x \rangle + (1 - a)^{-1} c \leq a^n V_k(x) + (1 - a)^{-1} c$$

and, additionally,

$$\langle V_k, \mu_* \rangle = \lim_{n \rightarrow \infty} \langle V_k, P_\varepsilon^n \delta_x \rangle \leq (1 - a)^{-1} c,$$

so the sequence  $(\langle V_k, \mu_* \rangle)_{k \in N}$  is bounded. Because  $(V_k)_{k \in N}$  is non-negative and non-decreasing, we may use the Monotone Convergence Theorem to obtain

$$\int_X V(y) \mu_*(dy) = \lim_{k \rightarrow \infty} \int_X V_k(y) \mu_*(dy).$$

Then,  $V$  is integrable with respect to  $\mu_*$ , so  $\mu_*$  is with finite first moment.

Keeping in mind that  $\bar{V}(x, y) = V(x) + V(y)$ , the exponential rate of convergence to the unique invariant measure  $\mu_* \in M_1^1(X)$  derives from the following estimates

$$\|P_\varepsilon^n \mu - \mu_*\|_{\mathcal{L}} \leq \int_X \int_X q^n C_5 (1 + \bar{V}(x, y)) \mu_*(dy) \mu(dx) \leq q^n C,$$

where  $C := \int_X \int_X C_5 (1 + \bar{V}(x, y)) \mu_*(dy) \mu(dx) < \infty$  for  $\mu \in M_1^1(X)$ . Finally, since  $C$  is dependant only on  $\mu$ , the proof is complete.  $\square$

## X. CENTRAL LIMIT THEOREM - PROOF OF THEOREM 2

Let us first make the follownig observation.

**Lemma 15.** *If  $\mu \in M_1^2(X)$  is with finite second moment, then  $\langle V^2, P_\varepsilon^n \mu \rangle < \infty$  and therefore  $\mu_* := \mu_*(\mu) = \lim_{n \rightarrow \infty} P_\varepsilon^n \mu$  has finite second moment.*

*Proof.* Let  $\mu \in M_1^2(X)$ . Fix  $x \in X$ ,  $n \geq 1$ . Recall that  $\Lambda < 1/2$  and  $c$  are given by (8) and (10), respectively. Moreover,  $\tilde{x}_n^x$  is of the form (9). Reasoning as in Lemma 1, we obtain

$$\begin{aligned} & \langle V^2, P_{h_1, \dots, h_n}^n \mu \rangle \\ &= \int_X \left[ \int_0^T \dots \int_0^T \varrho(\tilde{x}_n^x, \bar{x}) p(\tilde{x}_{n-1}^x, t_n) \dots p(x, t_1) dt_n \dots dt_1 \right] \mu(dx) \\ &\leq \int_X \int_0^T \dots \int_0^T \left[ 2\varrho^2(\tilde{x}_n^x, \tilde{x}_n^{\bar{x}}) + 2\varrho^2(\tilde{x}_n^{\bar{x}}, \bar{x}) \right] p(\tilde{x}_{n-1}^x, t_n) \dots p(x, t_1) dt_n \dots dt_1 \mu(dx) \\ &\leq 2 \int_X \int_0^T \dots \int_0^T \left[ \lambda^2(\tilde{x}_n^x, t_n) \dots \lambda^2(x, t_1) \varrho^2(x, \bar{x}) + 2\varrho^2(\bar{x}, \tilde{x}_1^{\bar{x}}) + 4\varrho^2(\tilde{x}_1^{\bar{x}}, \tilde{x}_2^{\bar{x}}) + 4\varrho^2(\tilde{x}_2^{\bar{x}}, \tilde{x}_n^{\bar{x}}) \right] \\ &\quad p(\tilde{x}_{n-1}^x, t_n) \dots p(x, t_1) dt_n \dots dt_1 \mu(dx) \\ &\leq 2\Lambda^n \langle V^2, \mu \rangle + 2^2 c^2 (1 + 2\Lambda + \dots + 2^n \Lambda^n) \\ &\leq 2\Lambda^n \langle V^2, \mu \rangle + 4 \frac{c^2}{1 - 2\Lambda}, \end{aligned}$$

Estimates are independent of choice of sequence  $(h_n)_{n \in N}$  and therefore  $\langle V^2, P_\varepsilon^n \mu \rangle < 2\Lambda^n \langle V^2, \mu \rangle + 4c^2(1 - 2\Lambda)^{-1}$ . We take a non-decreasing sequence  $(V_k^2)_{k \in N}$  such that  $V_k^2(y) = \min\{k, V^2(y)\}$ , for every  $k \in N$  and  $y \in X$ . We know that  $P_\varepsilon^n \mu$  converges weakly to  $\mu_*$ . Hence, for all  $k \in N$ ,  $V_k^2 \in C(X)$  and

$$\lim_{n \rightarrow \infty} \langle V_k^2, P_\varepsilon^n \mu \rangle = \langle V_k^2, \mu_* \rangle < 4c^2(1 - 2\Lambda)^{-1},$$

so the sequence  $(\langle V_k^2, \mu_* \rangle)_{k \in N}$  is bounded. Because  $(V_k^2)_{k \in N}$  is non-negative and non-decreasing, we may use the Monotone Convergence Theorem to obtain

$$\langle V^2, \mu_* \rangle = \lim_{k \rightarrow \infty} \langle V_k^2, \mu_* \rangle$$

so, indeed,  $\mu_*$  is with finite second moment.  $\square$

Let  $\eta_n^\mu$  and  $\Phi\eta_n^\mu$  be as in Section III. In particular,  $\eta_n^*$  and  $\eta_n^x$  are defined for the Markov chains with the same transition probability function  $\Pi_\varepsilon$  and initial distributions  $\mu_*$  and  $\delta_x$ , respectively. Further, let  $g : X \rightarrow R$  be a bounded and Lipschitz continuous function, with constant  $L_g$ , which satisfies  $\langle g, \mu_* \rangle = 0$ .

Central Limit Theorems for ergodic stationary Markov chains have already been proven in many papers. See, for example, Theorem 1 and the subsequent Corollary 1 in [14] by Maxwell and Woodroffe. The following lemma implies that assumptions of Theorem 1 and Corollary 1 ([14]) are satisfied.

**Lemma 16.** *Let  $g : X \rightarrow R$  be a bounded and Lipschitz continuous function with constant  $L_g$ . Additionally,  $\langle g, \mu_* \rangle = 0$ . Then,*

$$\sum_{n=1}^{\infty} n^{-3/2} \left[ \int_X \left( \sum_{k=0}^{n-1} \langle g, P_\varepsilon^k \delta_x \rangle \right)^2 \mu_*(dx) \right]^{1/2} < \infty \quad (24)$$

*Proof.* Note that, by Lemma 12 and Remark 13,

$$\begin{aligned} \sum_{k=0}^{n-1} \langle g, P_\varepsilon^k \delta_x \rangle &= \sum_{k=0}^{n-1} \left( \langle g, P_\varepsilon^k \delta_x \rangle - \langle g, \mu_* \rangle \right) \\ &= \sum_{k=0}^{n-1} \int_X \left[ \int_X g(z) (\Pi_\varepsilon^k(x, \cdot) - \Pi_\varepsilon^k(y, \cdot))(dz) \right] \mu_*(dy) \\ &= \sum_{k=0}^{n-1} \int_X \left[ \int_{X^2} (g(z_1) - g(z_2)) (\Pi_{X^2}^* \Pi_k^* \hat{C}_{h_1, h_2, \dots}^\infty((x, y), \cdot))(dz_1 \times dz_2) \right] \mu_*(dy) \\ &\leq \sum_{k=0}^{n-1} Gq^n C_5 \int_{X^2} (1 + \bar{V}(x, y)) \mu_*(dy). \end{aligned}$$

Then, for every  $x \in X, n \in N$ ,

$$\sum_{k=0}^{n-1} \langle g, P_\varepsilon^k \delta_x \rangle \leq GC_5 \frac{1 - q^n}{1 - q} \int_{X^2} (1 + \bar{V}(x, y)) \mu_*(dy) \leq C_9(1 + V(x)),$$

where  $C_9 := GC_5(1 - q)^{-1}(1 + \int_X V(y) \mu_*(dy))$ . Keeping in mind that  $\mu_*$  is with finite second moment, we obtain that (24) is not bigger than

$$\sum_{n=1}^{\infty} n^{-3/2} [C_9^2(1 + 2V + V^2, \mu_*)]^{1/2} < \infty$$

and the proof is complete.  $\square$

Hence, by applying Corollary 1, we obtain that  $\Phi\eta_n^*$  converges to the normal distribution in Levy metric, as  $n \rightarrow \infty$ , which equivalently means that the distributions converge weakly to each other (see [4] for proofs).

Now, the idea of the proof is based on the following remark.

**Remark 17.** *Note that, to complete the proof of Theorem 2, it is enough to establish that  $\Phi\eta_n^\mu$  converges weakly to  $\Phi\eta_n^*$ , as  $n \rightarrow \infty$ . Equivalently, it is enough to show that  $\lim_{n \rightarrow \infty} \|\Phi\eta_n^\mu - \Phi\eta_n^*\|_{\mathcal{L}} = 0$ , since weak convergence is metrised by the Fourtet-Mourier norm.*

*Proof of Theorem 2.* Set  $x, y \in X$  and choose arbitrary  $f \in \mathcal{L}$ . Suppose that we know that the following convergence is satisfied, as  $n \rightarrow \infty$ ,

$$\left| \int_R f(u) \Phi \eta_n^x(du) - \int_R f(v) \Phi \eta_n^y(dv) \right| \rightarrow 0. \quad (25)$$

Then, by the Dominated Convergence Theorem, we obtain

$$\begin{aligned} & \left| \int_R f(u) \Phi \eta_n^\mu(du) - \int_R f(v) \Phi \eta_n^*(dv) \right| \\ & \leq \int_X \int_X \left| \int_R f(u) \Phi \eta_n^x(du) - \int_R f(v) \Phi \eta_n^y(dv) \right| \mu(dx) \mu_*(dy) \rightarrow 0, \end{aligned} \quad (26)$$

as  $n \rightarrow \infty$ . Note that, by Theorem 11.3.3 in [3], (26) implies that  $\Phi \eta_n^\mu$  converges weakly to  $\Phi \eta_n^*$ , as  $n \rightarrow \infty$ , which, according to Remark 17, completes the proof of the CLT in the model. Now, it remains to show (25). Note that

$$\begin{aligned} \left| \int_R f(u) \Phi \eta_n^x(du) - \int_R f(v) \Phi \eta_n^y(dv) \right| &= \left| \int_{X^n} f\left(\frac{g(u_1) + \dots + g(u_n)}{\sqrt{n}}\right) \Pi_\varepsilon^{1, \dots, n}(x, du_1 \times \dots \times du_n) \right. \\ & \quad \left. - \int_{X^n} f\left(\frac{g(v_1) + \dots + g(v_n)}{\sqrt{n}}\right) \Pi_\varepsilon^{1, \dots, n}(y, dv_1 \times \dots \times dv_n) \right|, \end{aligned} \quad (27)$$

where  $\Pi_\varepsilon^{1, \dots, n}(x, \cdot) = \int_{B(0, \varepsilon)} \dots \int_{B(0, \varepsilon)} \Pi_{h_1, \dots, h_n}^{1, \dots, n}(x, \cdot) \nu^\varepsilon(dh_1) \dots \nu^\varepsilon(dh_n)$  is a measure on  $X^n$ . We may write

$$\begin{aligned} & \left| \int_{X^n} \int_{X^n} \left[ f\left(\frac{g(u_1) + \dots + g(u_n)}{\sqrt{n}}\right) - f\left(\frac{g(v_1) + \dots + g(v_n)}{\sqrt{n}}\right) \right] \right. \\ & \quad \left. \Pi_{h_1, \dots, h_n}^{1, \dots, n}(x, du_1 \times \dots \times du_n) \Pi_{h_1, \dots, h_n}^{1, \dots, n}(y, dv_1 \times \dots \times dv_n) \right| \\ & \leq \int_{(X^2)^n} \left| f\left(\frac{g(u_1) + \dots + g(u_n)}{\sqrt{n}}\right) - f\left(\frac{g(v_1) + \dots + g(v_n)}{\sqrt{n}}\right) \right| \\ & \quad \left( \Pi_{X^{2n}}^* \Pi_{1, \dots, n}^* \hat{C}_{h_1, h_2, \dots}^\infty((x, y), \cdot) \right) (du_1 \times \dots \times du_n \times dv_1 \times \dots \times dv_n), \end{aligned} \quad (28)$$

where  $\Pi_{1, \dots, n}^* : (X^2 \times \{0, 1\})^\infty \rightarrow (X^2 \times \{0, 1\})^n$  are the projections on the first  $n$  components and  $\Pi_{X^{2n}}^* : (X^2 \times \{0, 1\})^n \rightarrow X^{2n}$  is the projection on  $X^{2n}$ . Since  $f$  is Lipschitz with constant  $L_f$ , we may further estimate (28) from above

$$\begin{aligned} & \frac{L_f}{\sqrt{n}} \int_{X^{2n}} \left[ |g(u_1) - g(v_1)| + \dots + |g(u_n) - g(v_n)| \right] \left( \Pi_{X^{2n}}^* \Pi_{1, \dots, n}^* \hat{C}_{h_1, h_2, \dots}^\infty((x, y), \cdot) \right) ((du_i \times dv_i)_{i=1}^n) \\ & = \frac{L_f}{\sqrt{n}} \sum_{i=1}^n \int_{X^2} |g(u_i) - g(v_i)| \left( \Pi_{X^2}^* \Pi_i^* \hat{C}_{h_1, h_2, \dots}^\infty((x, y), \cdot) \right) (du_i \times dv_i). \end{aligned}$$

Now, for every  $1 \leq i \leq n$ , we refer to Lemma 12 and Remark 13 to observe that (28) is not bigger than

$$\frac{L_f G}{\sqrt{n}} \sum_{i=1}^n q^i C_5 (1 + \bar{V}(x, y)) = n^{-\frac{1}{2}} L_f G C_5 q \frac{1 - q^n}{1 - q} (1 + \bar{V}(x, y)).$$

Note that the expression above is independent of choice of sequence  $(h_n)_{n \in \mathbb{N}}$  and, thanks to this, is also the upper bound of (27). We go with  $n$  to infinity and obtain (25). The proof is complete.  $\square$

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- [1] P. Billingsley, *Convergence of probability measures*, Wiley, New York (1968).
  - [2] W. Bołt, A.A. Majewski & T. Szarek, *An invariance principle for the law of the iterated logarithm for some Markov chains*, *Studia Math.* 212 (2012), 41-53.
  - [3] R.M. Dudley, *Real analysis and probability*, Cambridge University Press (2004)
  - [4] S.N. Ethier & T.G. Kurtz, *Markov processes. Characterization and convergence*, John Wiley & Sons, New York (1986).
  - [5] R. Fortet & B. Mourier, *Convergence de la répartition empirique vers la répartition théorique*, *Ann. Sci. École Norm. Sup.* 70 (1953), 267-285.
  - [6] M. Hairer, *Exponential mixing properties of stochastic PDEs through asymptotic coupling*, *Probab. Theory Relat. Fields* 124 (2002), 345-380.
  - [7] M. Hairer, & J.C. Mattingly, *Spectral gaps in Wasserstein distances and the 2D stochastic Navier-Stokes equations*, *The Annals of Probability*, Vol. 36, No. 6 (2008), 2050-2091.
  - [8] P.R. Halmos, *Measure Theory*, Graduate Texts in Mathematics 18, Springer-Verlag, 117-136 (1974).
  - [9] S.C. Hille, K. Horbacz, & T. Szarek, *Unique steady-state molecular distribution for a regulatory network with random bursting*, submitted to *Ann. Math. Blaise Pascal*.
  - [10] R. Kapica & M. Ślęczka, *Random iteration with place dependent probabilities*, arXiv:1107.0707 [math.PR] (2012).
  - [11] T. Komorowski & A. Walczuk, *Central limit theorem for Markov processes with spectral gap in the Wasserstein metric*, *Stoch. Proc. Appl.* 122 (2012), 2155-2184.
  - [12] A. Lasota & M.C. Mackey, *Cell division and the stability of cellular populations*, *J. Math. Biology* 38 (1999), 241-261.
  - [13] A. Lasota & J.A. Yorke, *Lower bound technique for Markov operators and iterated function systems*, *Random Comput. Dynam.* 2(1) (1994), 41-77.
  - [14] M. Maxwell & M. Woodroffe, *Central limit theorems for additive functionals of Markov chains*, *The Annals of Probability*, Vol. 28, No. 2 (2000), 713-724.
  - [15] A. Murray & T. Hunt, *The Cell Cycle*, Oxford University Press (1993).
  - [16] S.P. Meyn, & R.L. Tweedie, *Markov Chains and Stochastic Stability*, Springer, London (1993).
  - [17] S.T. Rachev, *Probability Metrics and the Stability of Stochastic Models*, John Wiley, New York (1991).
  - [18] T. Szarek, *Invariant measures for nonexpansive Markov operators on Polish spaces*, *Dissertationes Math.* 415 (2003), 1-62.
  - [19] M. Ślęczka, *The rate of convergence for iterated function systems*, *Studia Math* 205, no. 3 (2011), 201-214.
  - [20] J.J. Tyson, & K.B. Hannsgen, *Cell growth and division: a deterministic/probabilistic model of the cell cycle*, *J. Math. Biology* 26 (1988), 465-475.
  - [21] C. Villani, *Optimal transport: old and new*, Grundlehren der mathematischen Wissenschaften, Vol. 338, Springer (2009).
  - [22] H. Wojewódka, *Exponential rate of convergence for some Markov operators*, *Statistics and Probability Letters* 83 (2013), 2337-2347.
  - [23] R. Zaharopol, *Invariant probabilities of Markov-Feller operators and their supports*, Boston, MA, Birkhauser Verlag (2005).