

# Predicting distributions of rare species: the case of the false coral snake *Rhinobothryum bovallii* (Serpentes: Colubridae)

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## Abstract

**Predicting distributions of rare species: the case of the false coral snake *Rhinobothryum bovallii* (Serpentes: Colubridae).** Typically, the lack of enough high-quality occurrence data makes it difficult to define the geographic distribution of rare species. However, species distribution models provide a powerful tool for biodiversity management, including efforts to predict the distributions of rare species. Herein, new and historical data are used to model the distribution of the False Tree Coral snake, *Rhinobothryum bovallii*. The prediction map reveals a disjunct distribution for this species, from the Central American Isthmus to the northwestern portion of South America, with the species occupying lowlands and premontane forests below about 1500 m elevation. We identified 491,516 km<sup>2</sup> of suitable habitat for *R. bovallii* (minimum training presence threshold of 0.424) and 59,353 km<sup>2</sup> of core habitat, with concentrations in three relatively isolated core areas (10-percentile training presence threshold of 0.396), as follow: (1) a “northern core” along the Pacific and Caribbean coasts of Panama; (2) a “central core” in the Middle Magdalena Valley in Colombia; and (3) a “southern core” in the Ecuadorian Chocó. The occurrence of this species has a strong positive association with low precipitation seasonality, high precipitation in the warmest quarter, and low variability in annual temperature. Xeric and semiarid areas are unsuitable for this species and may pose environmental barriers limiting its distributional range. These results may lead to the discovery of additional populations of *R. bovallii*, identify priority survey areas, and by determining the extent of its natural habitat promote effective conservation strategies.

**Keywords:** core habitat areas, disjunct distribution, marginal habitat, species distribution models.

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## Resumo

**Prevendo distribuições de espécies raras: o caso da falsa cobra coral *Rhinobothryum bovallii* (Serpentes: Colubridae).** Normalmente, a falta de dados de ocorrência de alta qualidade suficientes torna difícil definir a distribuição geográfica de espécies raras. No entanto, os modelos de distribuição de espécies fornecem uma ferramenta poderosa para a gestão da biodiversidade, incluindo esforços para prever a distribuição de espécies raras. Aqui, dados novos e históricos são usados para modelar a distribuição da serpente *Rhinobothryum bovallii*. O mapa de previsão revela uma distribuição disjunta para essa espécie, do Istmo da América Central à porção noroeste da América do Sul, com a espécie ocupando terras baixas e florestas pré-montanas abaixo de cerca de 1.500 m de altitude. Identificamos 491.516 km<sup>2</sup> de habitat adequado para *R. bovallii* (limite de corte de presença mínima de treino de 0,424) e 59.353 km<sup>2</sup> de habitat central, com concentrações em três áreas centrais relativamente isoladas (limite de corte de 10-percentil de presença de treino de 0,396), como segue: (1) um “núcleo norte” ao longo das costas do Pacífico e do Caribe do Panamá; (2) um “núcleo central” no Vale do Médio Magdalena, na Colômbia; e (3) um “núcleo sul” no Chocó equatoriano. A ocorrência dessa espécie tem forte associação positiva com a baixa sazonalidade da precipitação, alta precipitação no trimestre mais quente e baixa variabilidade na temperatura anual. Áreas xéricas e semiáridas são inadequadas para essa espécie e podem representar barreiras ambientais que limitam sua faixa de distribuição. Esses resultados podem levar à descoberta de populações adicionais de *R. bovallii*, identificar áreas prioritárias de levantamento e, ao determinar a extensão de seu habitat natural, promover estratégias de conservação eficazes.

**Palavras-chave:** áreas centrais de habitat, distribuição disjunta, habitat marginal, modelos de distribuição de espécies.

## Introduction

A complex interaction of biotic (e.g., source-sink dynamics, habitat heterogeneity, competition, and mutualism), abiotic (e.g., climate, hydrologic and topographic changes), and evolutionary (e.g., speciation rates, extinction rates, and niche conservatism) processes have been suggested to play a key role in shaping Neotropical biodiversity (Van der Hammen and Hooghiemstra 2000, Antonelli and Sanmartín 2011, Antonelli *et al.* 2018). The Neotropical region (i.e., tropical America) is the most species-rich region on Earth and houses nearly a third of the world’s biodiversity hotspots (Mittermeier *et al.* 2004). The Chocó Biogeographic Region (located within the Tumbes-Chocó-Magdalena hotspot, formerly called the Chocó-Darién-Western Ecuador hotspot) harbors an exceptional species richness and endemism in a relatively small area (187,400 km<sup>2</sup>, Mittermeier *et al.* 2004) less than ~ 3% of

the area of the Amazon Basin. For reptiles, especially snakes, the Chocó and Amazonian ecoregions have a close biogeographic affinity at the suprageneric-level, sharing about 55% of genera and 66% of families (data analyzed from Lynch *et al.* 2016). This fact suggests that these ecoregions had an ancient common history before the Andean uplift (between 2.7–5 Ma; Gregory-Wodzicki 2000), which may have resulted in allopatric speciation when populations of lowland species became isolated on either side of the cordillera (Hernández-Camacho *et al.* 1992, Smith *et al.* 2014, Winterton *et al.* 2014).

The two species of *Rhinobothryum* Wagler, 1830 may provide the clearest example of allopatric distributions among Neotropical snakes. The distributions of *Rhinobothryum bovallii* (Andersson, 1916) and *Rhinobothryum lentiginosum* (Scopoli, 1785) suggest the separation of a once-continuous range into two geographically isolated, or allopatric, parts (Rojas-Morales 2012). *Rhinobothryum bovallii*

occurs in lowlands from southern Honduras to Costa Rica and Panama in Central America, and from western Colombia to northwestern Ecuador and northeastern Venezuela in South America (Savage 2002, Rojas-Morales 2012, McCranie 2015, Rojas-Runjaic and Infante-Rivero 2018, Turcios-Casco *et al.* 2018, Martínez-Fonseca *et al.* 2019). In contrast, *R. lentiginosum* occurs in lowlands along the Amazon Basin in South America, with records from Colombia, Ecuador, Peru, Bolivia, Paraguay, Brazil, French Guiana, Suriname, Guyana, and Venezuela (Peters and Orejas-Miranda 1970, Martins and Oliveira 1998, Miranda *et al.* 2009, Wallach *et al.* 2014, Gomes de Arruda *et al.* 2015). The geographic distributions of these species differ substantially; nevertheless, the limits of their ranges are uncertain. This is especially true of *R. bovallii*, a snake that is rarely seen and that apparently has disjunct distributions within its geographical range (Savage 2002, Wilson and McCranie 2003, Martínez-Fonseca *et al.* 2019).

Methods for estimating the distribution of a species based on fundamental ecological niche theory allow us to: (1) model the distribution of poorly known species (e.g., Raxworthy *et al.* 2003, Meza-Joya *et al.* 2018); (2) predict the potential geographic range of invasive species (e.g., Rödder and Lötters 2010, Urbina-Cardona and Castro 2010); (3) develop conservation action plans (e.g., Araújo *et al.* 2019, Botero-Delgadillo *et al.* 2012); (4) predict the effects of climate change on species distributions (e.g., Ochoa-Ochoa *et al.* 2012, Velásquez-Tibatá *et al.* 2013); (5) develop vector-borne disease spread and risk models (e.g., Peterson 2006, Fatima *et al.* 2016); and (5) identify historic refugia for biodiversity (e.g., Waltari *et al.* 2007, Vega *et al.* 2010). Species distribution models (SDMs) use predictor variables and species occurrences (either presence-only or presence-absence) to estimate objects in geographic space, referring to the actual distribution of a given species (Peterson and Soberón 2012). However, predicting the distribution of rarer species is challenging because the factors that contribute to

rarity (e.g., range size, habitat specificity, and local density) are poorly known for most taxa (Rabinowitz 1981, Karl *et al.* 2002, Franklin *et al.* 2009).

Rare species are highly sensitive to both natural and human-related disturbances, which increases their vulnerability, and therefore, their risk of extinction (Davies *et al.* 2004). Here we study the distribution of *R. bovallii*, a species that is rarely encountered and that apparently has a naturally disjunct distribution from lowest Central America to northwest South America (Savage 2002, Rojas-Morales 2012, McCranie 2015, Martínez-Fonseca *et al.* 2019). The rarity of this species has been attributed to sampling bias, low encounter rates, and the difficulty of sampling arboreal habitats (Arredondo *et al.* 2017, Rojas-Runjaic and Infante-Rivero 2018). Our primary objective was to update the distributional range of *R. bovallii* using reliable historical, as well as new records in conjunction with predictive species distribution models. We also assess the effect of predictor variables that have a major influence in shaping the current environmental niche of this species. Last, we identified areas where the predicted habitat suitability differs in order to identify core and marginal habitats for conservation purposes. Our results underscore the utility of predictive models in exploring the distribution of elusive and rare species.

## Materials and Methods

### Species Data

We compiled locality information for *Rhinobothryum bovalli* from museum specimens, well-supported reports from published literature, and records from our field surveys (collection permits 164-2014, 1166-2014 by Corporación Autónoma Regional de Caldas-Corpocaldas, and 0047-2015 by Autoridad Nacional de Licencias Ambientales-ANLA). We obtained additional data from the Global Biodiversity Information Facility (<http://www.gbif.org>), the VertNet data

portal (<http://www.vertnet.org>), and SiB Colombia (<http://www.sibcolombia.net>) accessed on 20 October 2019. Our analyses only include well-supported records from these electronic databases because they are prone to contain inaccuracies (Nogueira *et al.* 2011, Zaher *et al.* 2011). The accuracy of an occurrence record was considered reliable if: (1) it was published by a reliable source (e.g., herpetological collection or museum); (2) it was supported by a voucher specimen; and (3) its taxonomic diagnose was confirmed by a curator or previous studies. We also included photographic records from <http://www.inaturalist.org> because this snake is very conspicuous and easily recognized by photographs. Each record available was checked to identify uncertainty and mistakes in the geographic information. We retained localities with geographic coordinates assigned by other authors based on collection locality data (Appendix I). Although there is uncertainty associated with these coordinates, we expect them to fall near or in the correct pixel of the environmental data (pixel size is 30 arc-seconds or about 1 km<sup>2</sup>). We excluded records with incomplete, inconsistent, or imprecise geographic information.

### *Environmental Data*

To explore predictors that are likely to affect the distributions of species, we used climatic data from 19 WorldClim variables and elevation (Hijmans *et al.* 2005) at a 30 arc-seconds spatial resolution (ca. 1 km<sup>2</sup>). We restricted the selection of environmental data from “background” pixels to a region that matched areas from which of *Rhinobothryum bovalli* is known to occur and thus, might be representative of suitable environmental conditions for the species (Anderson and Raza 2010, Peterson *et al.* 2011, Galante *et al.* 2018). We used a rectangle encompassing a four-degree buffer (ca. 450 linear km) around the most extreme locality at each cardinal direction. This selection method seems appropriate because it excludes large areas that are environmentally suitable,

but where the species is likely unable to disperse—because of the presence of insurmountable physical or physiological barriers to dispersal—and/or is not known to occur (i.e., geographic areas where the species has not been reported or collected historically or inventory pseudo-absences).

### *Sampling Bias*

To reduce the likely effects of spatial autocorrelation in our initial occurrence dataset ( $N = 103$  unique localities), we used SDMtoolbox to filter localities keeping the maximum number of occurrences that were at least 20 km apart (Brown *et al.* 2017). This distance threshold was selected because *Rhinobothryum bovalli* is distributed mainly in lowlands, which are expected to have low levels of topographic heterogeneity. This procedure produced a maximum number of 65 points of occurrence. To address possible sensitivity of this procedure in resulting models (Galante *et al.* 2018), we first ran a series of starting models with the unfiltered dataset and then with the spatially filtered one. To build these models, we used the recommended feature classes to be appropriate for our sample sizes—all feature classes for the unfiltered dataset and linear plus quadratic plus hinge for the filtered one (Phillips and Dudík 2008). For each feature class combination, we used a narrow range of regularization multiplier values close to the default value—i.e., 0.5, 1 (default value), 1.5, and 2. Regularization was employed to reduce the number of variables selected for inclusion in the resulting models (Phillips and Dudík 2008, Elith *et al.* 2011). We used the occurrence dataset from the best-ranked model (i.e., unfiltered dataset) to build the suite of final candidate models.

### *Variable Reduction*

To ensure the quality of the final suite of habitat suitability models and to reduce potential over-parameterization, we followed

the process outlined by Warren *et al.* (2014). For this, we calculated the contribution scores (permutation importance and percentage contribution) for each variable in our suite of starting models (see above). To obtain alternate estimates of which variables are most important in the model, we also ran a jackknife test. Then, we calculated the spatial correlations (Pearson coefficient) between variables using the software ENMTools version 1.4.4 (Warren *et al.* 2010). We used contribution scores in conjunction with the scores from correlations to select the final set of environmental variables. First, we eliminated variables with low contribution scores (< 5%) in the starting model. Then, we deleted variables that were highly correlated ( $|r| > 0.85$ ) following Elith *et al.* (2010), keeping the variables with the highest contribution scores. We extracted from the best-ranked model the reduced set of variables to build the suite of final candidate models.

### Species Distribution Modelling

We generated SDMs using MaxEnt version 3.4.1, a presence-background algorithm that integrates environmental data with georeferenced occurrence records to model the habitat suitability of a given species (Phillips *et al.* 2006). Although many methods are available for modeling distributions, MaxEnt has high predictive accuracy for a wide range of species in diverse regions, even when available information is incomplete and the sample size is small (e.g., Pearson *et al.* 2007, Anderson and Raza 2010, Chunco *et al.* 2013, Ramos *et al.* 2018). We used recommended default values for convergence threshold ( $10^{-5}$ ), maximum number of iterations (500), maximum number of background points ( $10^4$ ), and default prevalence of the species (0.5). Last, we selected the logistic output format, which yields continuous values ranging from 0 to 1 that indicate the probability of suitable environmental conditions for the species (Phillips and Dudík 2008). We built final

candidate models randomly excluding 30% of the observation records to use as test data, with 100 bootstrapped replicates.

### Model Calibration

Final candidate models were generated using multiple combinations of distinct values of regularization multipliers and feature classes to identify the parameter settings that lead to best-fitting models. These models were built using unfiltered occurrence (Table 1) and withheld environmental data (Table 2). We created a suite of models by allowing increasing complexity of the feature classes employed, taking into account our sample size (i.e., 103 occurrence points; see Phillips and Dudík 2008, Elith *et al.* 2011): hinge (H), linear plus quadratic (LQ), linear plus quadratic plus hinge (LQH), and all features (default setting). For each feature class combination, we built models using the same range of regularization multiplier values as in starting models (i.e., 0.5, 1, 1.5, and 2).

**Table 1.** Summary of evaluation statistics of the starting models for *Rhinobothryum bovallii*. Results are provided for two model-selection techniques (AUC and OR) for two datasets (unfiltered and filtered localities) using recommended features classes (all features and linear plus quadratic plus hinge, respectively). Omission rates (OR) based on the minimum training presence threshold.

Feature	Regularization	OR	AUC
Unfiltered (103 occurrences)	0.5	0.037	0.954
	1	0.013	0.941
	1.5	0.023	0.933
	2	0.031	0.912
Filtered (65 occurrences)	0.5	0.085	0.962
	1	0.057	0.933
	1.5	0.061	0.921
	2	0.029	0.893

**Table 2.** Percentage of contribution and permutation importance of the most important variables used to build final species distribution models for *Rhinobothryum bovallii*. The four variables with the highest contributions are presented in bold.

Variable	Code	Permutation importance	Percent contribution
Precipitation seasonality	BIO15	<b>25</b>	<b>16</b>
Precipitation of warmest quarter	BIO18	<b>7.4</b>	<b>14.7</b>
Elevation	-	<b>10.6</b>	<b>14.1</b>
Precipitation of driest quarter	BIO17	<b>6.9</b>	2
Temperature annual range	BIO7	5.1	<b>13.4</b>

### Model Selection

We compared all the models generated by using the scores of two quantitative measures: (1) area under a receiver operating characteristic curve (AUC; threshold-independent), and (2) omission rate by applying the minimum training presence threshold (OR; threshold-dependent). Specifically, the AUC gives a relative measure of the overall discriminatory ability of a model (Peterson *et al.* 2011), whereas omission rates indicate whether a model is overfitted to the calibration data (Galante *et al.* 2018). By using this threshold rule, we expect 0% test omission; thus, values above zero indicate overfitting (Shcheglovitova and Anderson 2013). Hence, we used OR as the primary criterion for selecting optimal combinations of feature classes and regularization multiplier (Shcheglovitova and Anderson 2013). We first identified the models that displayed the lowest average OR, and then, from that subset of models, we chose the one with the highest average AUC score.

### Binary Model and Core Areas

Once we selected the best-ranked model, its logistic output was transformed into a binary prediction model for the suitable habitat of the species (i.e., a presence/absence map) by applying the minimum training presence threshold value for the training occurrence data

obtained by MaxEnt. This threshold rule was chosen because it has a straightforward ecological interpretation by identifying pixels predicted to be at least as suitable as those from which the species has been recorded (Pearson *et al.* 2007). Then, we evaluated the final binary model by visual examination based on our knowledge of the natural history and geographic distribution of *Rhinobothryum bovallii*. To identify core habitats, we follow the approach proposed by Molloy *et al.* (2014). We converted predictions of the best-ranked model by applying the threshold cut-off training value for the 10-percentile training presence threshold. This approach seems rational because the threshold rule excludes all regions having suitability values lower than those for the lowest 10% of occurrence records (Radosavljevic and Anderson 2014) and leads to a smaller geographical prediction of high-quality habitats in environmental terms. Following this approach, areas with environmental values containing 90% of occurrence records were designated as core habitats (i.e., pixels of the model in which test occurrence points had a high probability of falling and where the species is most likely to sustain viable populations), whereas the remaining areas correspond to marginal habitats (i.e., pixels of the model representing a marginal part of the fundamental niche of the species in which survival and reproduction are thought to be poor because of the limiting environmental conditions).

## Results

### *Models Performance*

In the starting models, the AUC values were slightly better for the models with the unfiltered dataset than those for models with filtered one (except for models calibrated to 0.5 regularization value). As expected, starting models of both datasets (spatially unfiltered and filtered) had OR values above the expected 0% for the minimum training presence threshold. However, starting models calibrated using the unfiltered dataset and all feature classes [as recommended by Phillips and Dudík (2008) for sample sizes > 80 occurrence points] produced higher average evaluation scores (mean AUC =  $0.935 \pm 0.018$ ) and lower average omission rates ( $0.026 \pm 0.010$ ) across regularization multipliers. Although the used of filtered occurrence results in the model with the highest AUC score (AUC = 0.962), its OR was six times greater (OR = 0.085) than the model with the lowest one (OR = 0.013), indicating greater overfitting in the first. The model performance measures (AUC and OR) of the suites of starting models are given in Table 1.

Observed average AUCs for our final models varied, with lower regularization multipliers leading to higher AUCs within a given feature class (Figure 1A). Default features (i.e., all features) displayed the highest average AUC values across regularization multipliers (mean AUC =  $0.930 \pm 0.019$ ), followed by LQH (mean AUC =  $0.922 \pm 0.021$ ), H (mean AUC =  $0.920 \pm 0.018$ ), and LQ features (mean AUC =  $0.863 \pm 0.009$ ; Figure 1B). The highest AUC value was observed for default features at the lowest regularization multiplier (AUC =  $0.950 \pm 0.007$ ). However, the difference between this value and that for LQH and H features at the same regularization multiplier (i.e., 0.5) was minimal (0.004 for LQH and 0.006 for H features; Figure 1A).

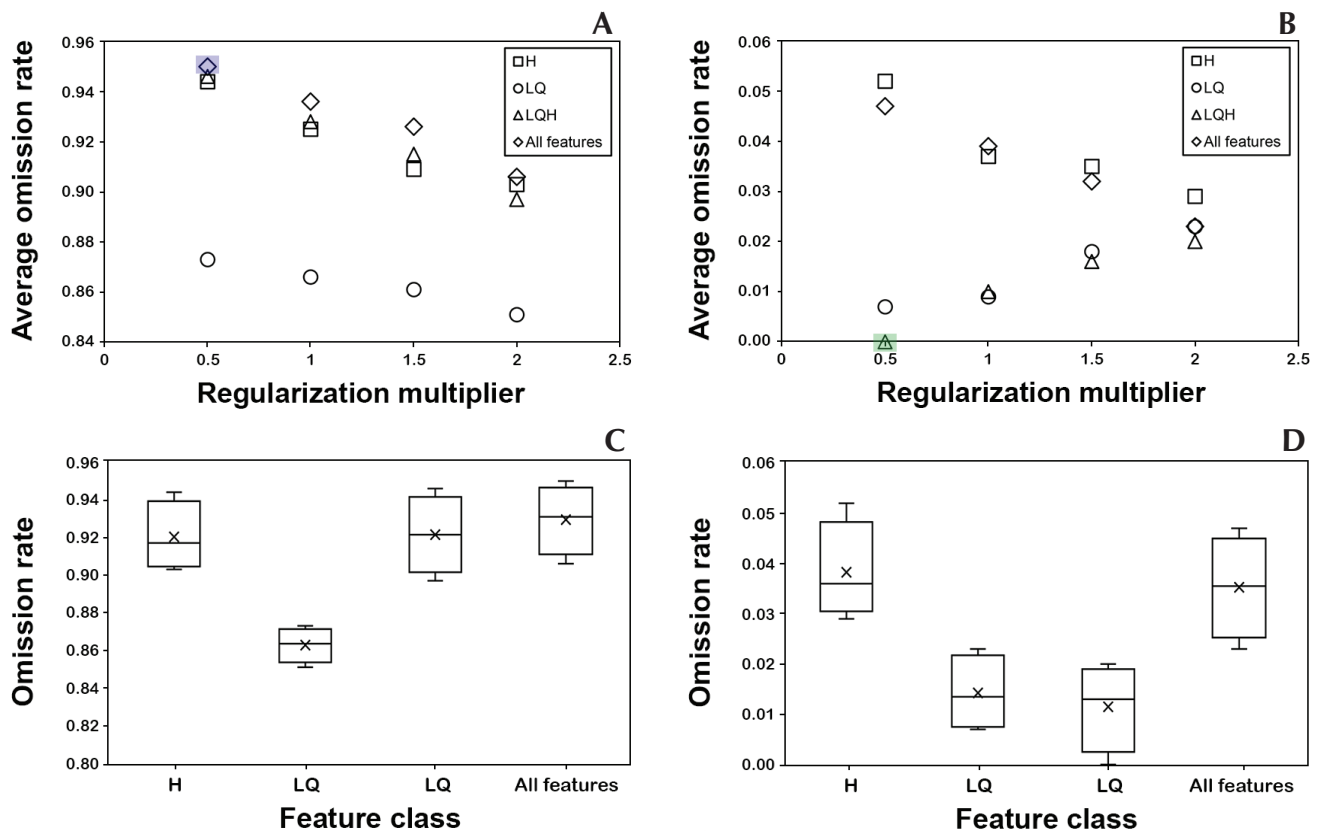
Model calibration showed great differences in average ORs among feature classes, with LQH and LQ features obtaining the lowest OR

values at low regularization multipliers and then dropping in performance at high regularization multiplier values. In contrast, default and H features showed the opposite pattern (Figure 1C). Linear plus quadratic plus hinge features displayed the lowest average OR values across regularization multipliers (mean OR =  $0.012 \pm 0.009$ ), followed by LH (mean AUC =  $0.014 \pm 0.008$ ), default (mean AUC =  $0.035 \pm 0.010$ ), and H features (mean AUC =  $0.038 \pm 0.010$ ; Figure 1D). The lowest OR value was observed for the LQH feature at the lowest regularization multiplier (OR = 0.000), which fix to the theoretical expectation of zero omission of evaluation localities using the minimum training presence threshold (Figure 1B).

Based on the criteria employed, the default settings (i.e., all feature classes and regularization multiplier of 1) were not optimal for our dataset. We identified the LQH feature class and the lowest examined regularization multiplier (i.e., 0.5) as the optimal combination for our datasets. Given the sequential nature of the selection criteria used here, we did not necessarily select the combination with the highest evaluation AUC as optimal. However, the difference between the AUC of the optimal combination (i.e., LQH features and regularization multiplier of 0.5) and that of the combination with the highest AUC (i.e., default features and regularization multiplier of 0.5) was negligible (AUC difference of 0.004). The model performance measures (AUC and OR) of the suites of final models are given in Figure 1.

### *Variable Contribution and Response Curves*

The analysis of variable contributions reveals that precipitation seasonality, precipitation of warmest quarter, elevation, precipitation of driest quarter, and temperature annual range, significantly affect the distribution of *Rhinobothryum bovalli* (Table 2). Based on both metrics (i.e., permutation importance and percent contribution), precipitation seasonality has the greatest influence on the distribution of this



**Figure 1.** Evaluation statistics resulting from calibration of Maxent models for *Rhinobothryum bovallii*. Top panels show the test Area Under the Curve (AUC) scores (A) and the omission rates (OR) at minimum training presence threshold (B). Bottom panels show the statistics (i.e., median, first and third quartiles, minimum, and maximum) for AUC (C) and ORs (D) values compared across regularization multipliers for each feature class. In each panel, the optimal model is highlighted. Model statistics are shown as feature classes (H = hinge, LQ = linear plus quadratic, LQH = linear plus quadratic plus hinge, and all features) increasing in regularization multiplier (i.e., 0.5, 1, 1.5, and 2).

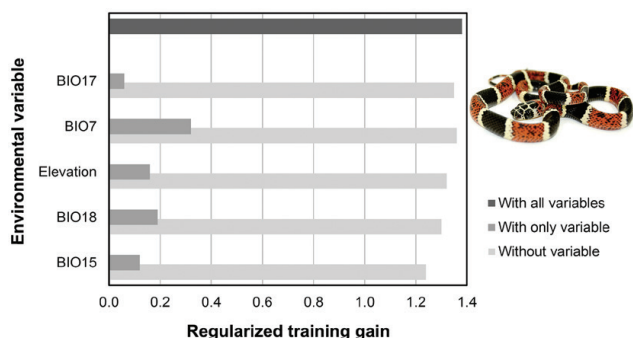
species (25% and 16%, respectively). The jackknife test of variable importance shows that the environmental variable with the highest gain when used in isolation is temperature annual range—therefore contains the most useful information by itself, whereas precipitation seasonality decreases the gain the most when it is omitted—and thus contains information not present in any other variable (Figure 2). Response curves show that suitability was positively linked with low values of precipitation seasonality (< 30%). Suitability was also maximized around 500 mm of precipitation in the warmest quarter and a temperature annual range of about 12°C.

As expected, presence probability is highest at low elevations and decreases at higher altitudes (> 1000 m a.s.l.).

### Species Distribution Range

To generate SDMs, we compiled 103 reliable locality records for *Rhinobothryum bovallii* (84 historical and 19 new). Known records of this species range from about 0° to 15° N latitude, with the archaeological site “Ciudad Blanca” in Honduras and the Caique River in Ecuador, the northernmost and southernmost known localities, respectively (Pazmiño-Otamendi 2017, Turcios-





**Figure 2.** Jackknife test of variable importance showing regularized training gain of individual predictor variables important in the development of final models for *Rhinobothryum bovallii*.

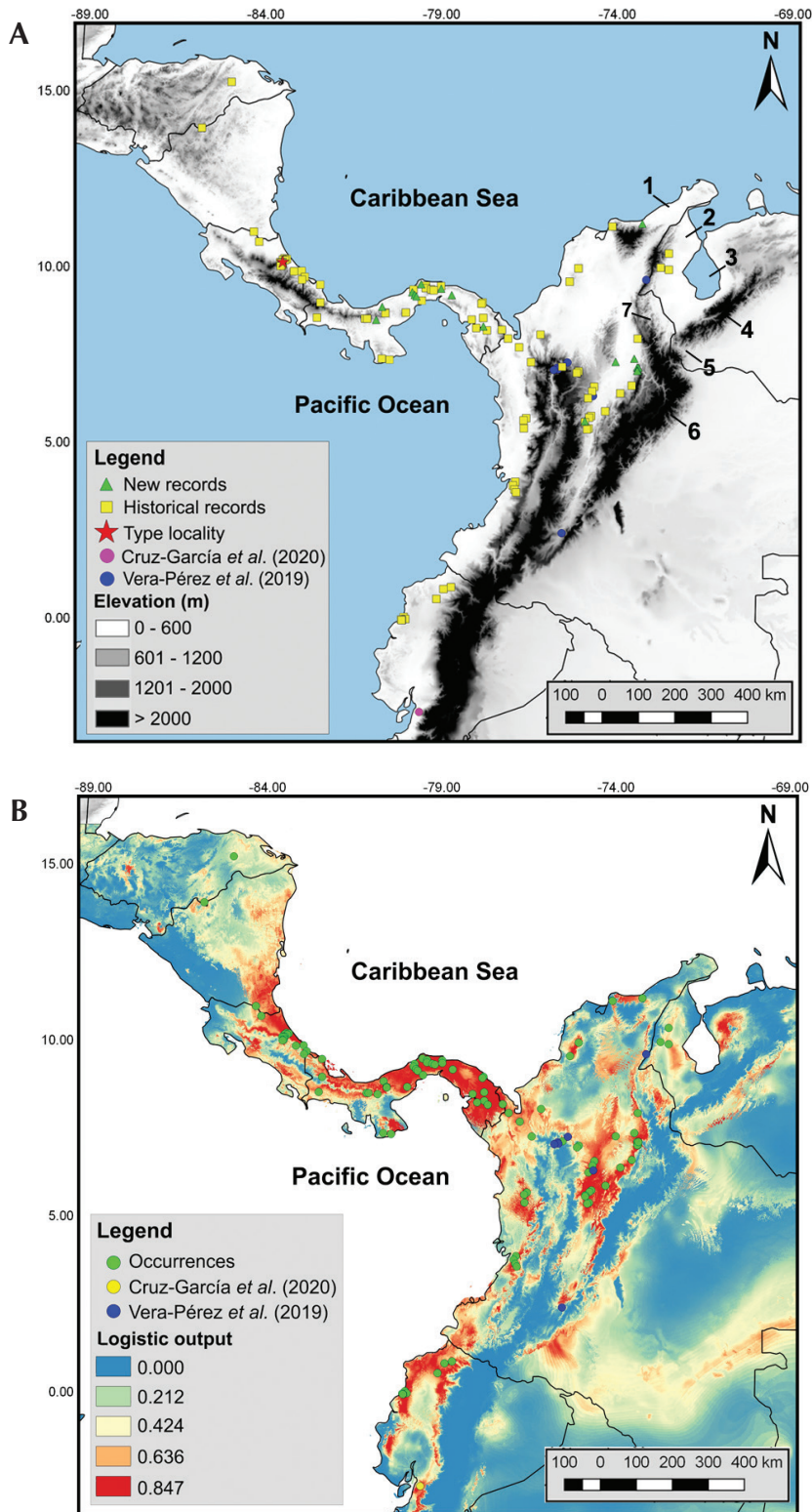
Casco *et al.* 2018). Records come from 12 of the ecoregions proposed by Dinerstein *et al.* (2017): Central American Atlantic moist forests, Talamancan montane forests, Isthmian-Atlantic moist forests, Isthmian-Pacific moist forests, Chocó-Darién moist forests, Western Ecuador moist forests, Cauca Valley montane forests, Magdalena Valley montane forests, Magdalena-Urabá moist forests, Magdalena Valley dry forests, Sinú Valley dry forests, and Cordillera Oriental montane forests. Review of digital photographs of a specimen of *R. bovallii* from Caquetá department (Solano municipality) on the Amazon versant of Colombia (IAvH-R 6453, IAvH-R 6454 in Martínez-Fonseca *et al.* 2019) confirm that it is *R. lentiginosum*. Examination of a specimen from the same locality deposited at Universidad de la Amazonia in Colombia (UAM-R uncatalogued) supports this determination. Specimens of *R. bovallii* from Suriname and Guyana (see <http://vertnet.org>) probably correspond to misidentifications of the congeneric species *R. lentiginosum* (see also Martínez-Fonseca *et al.* 2019). Based on geographic coordinates of locality records (Appendix I), the altitudinal range for *R. bovallii* ranges from sea level (QCAZ-R 4652) to 1535 m (KU 75749). The updated distribution of *R. bovallii* is summarized in Figure 3A (For additional details, see Appendix I).

The logistic output of the best-ranked model identified several regions of high prediction for *R. bovallii* in most parts of the Caribbean coast of southern Nicaragua, Costa Rica, and Panama; tenuously connected areas of the Chocó region in Panama, Colombia, and Ecuador; Middle Magdalena Valley in Colombia; and isolated portions of the Caribbean regions in Colombia and Venezuela. Furthermore, the model identified small, isolated areas of high suitability in regions from which there are no records for the species. These areas include the eastern slope of the Cordillera Oriental, the upper basin of the Magdalena River, and the Catatumbo Basin in Colombia, as well as the Machango, Misoa, and Rio Viejo basins in Venezuela, all of which drain to the east of the Maracaibo Lake. South of this area, the model predicts the potential presence of *R. bovallii* in the foothills of the Merida Cordillera (Figure 3B). According to the binary map generated using the cut-off training value for the minimum training presence threshold (i.e., 0.424), we estimated an extent of occurrence of 491,516 km<sup>2</sup> for the species (Figure 4A). After applying the cut-off value for the 10-percentile training presence threshold (i.e., 0.396) we identify three main core habitat areas, summing approximately 59,353 km<sup>2</sup> (Figure 4B): a “northern core” along the Pacific and Caribbean coasts of Panama (~ 31,529 km<sup>2</sup>); a “central core” in the Middle Magdalena Valley in Colombia (~ 20,939 km<sup>2</sup>); and a “southern core” in the Ecuadorian Chocó (~ 6,885 km<sup>2</sup>).

## Discussion

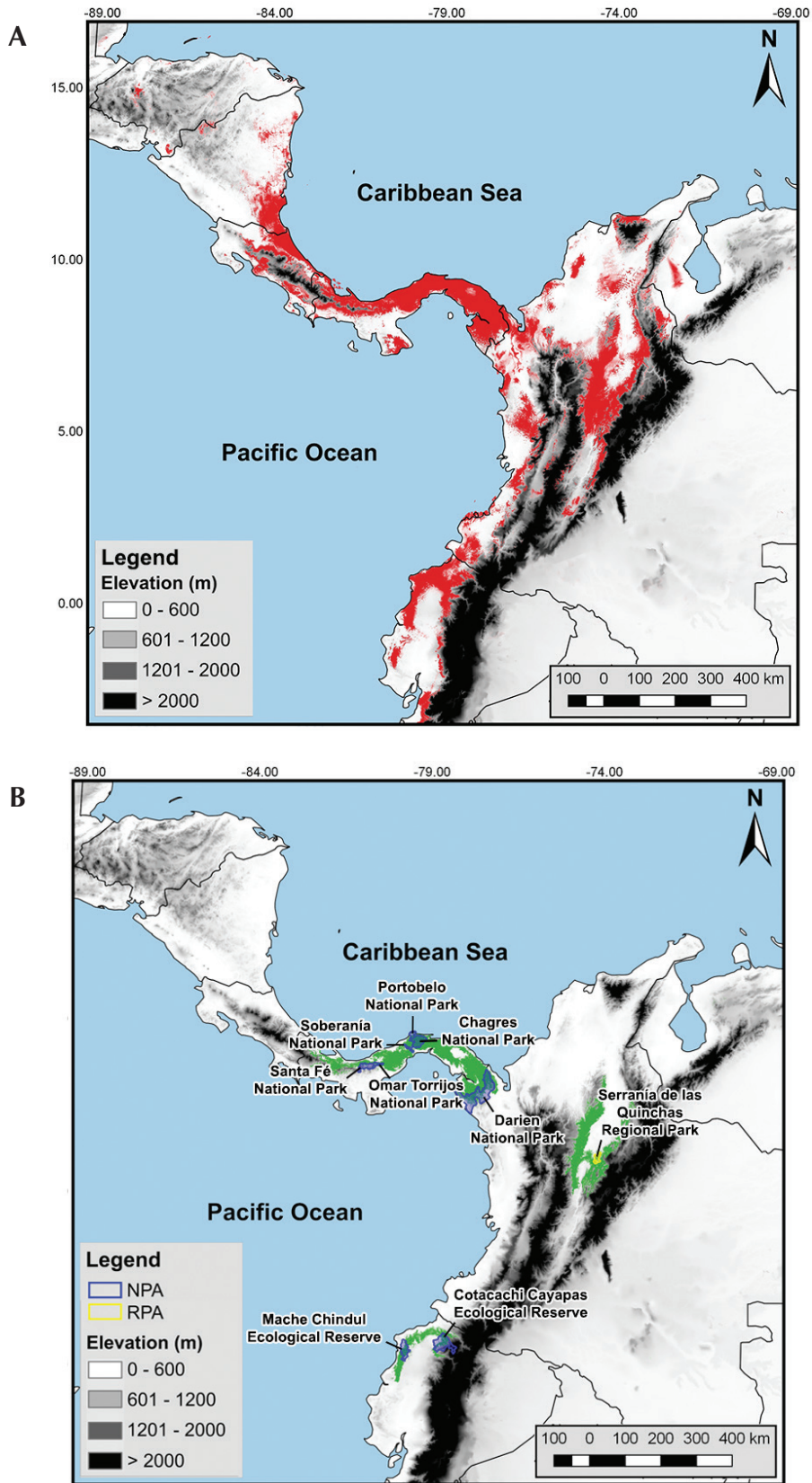
### Models Performance

The starting models using unfiltered versus filtered datasets yielded substantially higher OR values when the minimum training presence threshold was applied to both datasets. However, the unfiltered dataset leads to models with higher AUC scores and lower ORs across regularization multipliers. Use of the unfiltered dataset led to models with areas of higher prediction



**Figure 3.** Updated distributional map (A) and logistic output for the best-ranked model (B) for *Rhinobothryum bovallii*. Map numbers indicate: 1 = Guajira-Barranquilla xeric scrub ecoregion (Colombia), 2 = Paraganá xeric scrubs ecoregion (Venezuela), 3 = Maracaibo Lake (Venezuela), 4 = Cordillera de Mérida (Venezuela), 5 = Táchira depression (Venezuela), 6 = Cordillera Oriental (Colombia), 7 = Catatumbo basin (Colombia). Geodetic datum = WGS-84.

Predicting the distribution of a rare snake



**Figure 4.** Binary distribution map (A) and main core habitat areas (B) for *Rhinobothryum bovallii*. NPA = National protected areas. RPA = Regional protected areas. Geodetic datum = WGS-84.

concentrated around known records, suggesting low sampling bias, as well as more realistic geographic predictions, in contrast to the filtered dataset. Although spatially filtered data is most likely to match the assumption of unbiased sampling (Galante *et al.* 2018), it is probable that the rarity of *Rhinobothryum bovalli* throughout its range may contribute to reduced sampling bias in our unfiltered dataset, and as a consequence, to best geographic predictions. Although the unfiltered dataset used here seems to be minimally affected by bias problems, it is critical to conduct filtering analysis to assess whether the resulting models are affected by spatial sampling bias (Radosavljevic and Anderson 2014, Boria *et al.* 2014, Galante *et al.* 2018).

In the model calibration, lower regularization multipliers lead to higher AUC values within a given feature class. Although OR scores were variable across the feature classes, low regularization multipliers (0.5) were necessary to reduce overfitting to the lowest levels (Figure 1). Although the highest point in AUC occurs when default features (i.e., all features) are used, ORs indicate much better performance with LQH features. The fact that optimal performance corresponds to more complex sets of feature classes supports the findings of previous studies (e.g., Shcheglovitova and Anderson 2013, Galante *et al.* 2018, Ramos *et al.* 2018). Notably, these studies examined species with few occurrence records (< 23 localities), whereas *R. bovalli* has 103 occurrence points. Thus, our finding that overfitting decreases at regularization multipliers lower than default when used with more complex sets of feature classes contrasts with the results from other studies using both small (< 23 localities) and large (> 208 localities) datasets (e.g., Shcheglovitova and Anderson 2013, Radosavljevic and Anderson 2014, Ramos *et al.* 2018). While low regularization values tend to result in overly complex models (Radosavljevic and Anderson 2014), it is probable that the removal of highly correlated variables in our study provides additional protection against overfitting that leads to low

regularization scores and an increase in performance.

#### *Variable Contribution and Response Curves*

The predicted habitat suitability for *Rhinobothryum bovalli* in environmental space is represented by low variability of monthly precipitation values during the year (i.e., precipitation seasonality), high values of precipitation during the warmest three months of the year (i.e., precipitation in the warmest quarter), and low variability in annual temperature (i.e., temperature annual range). In general, these results show that climatic habitat suitability for this species is highest in most equatorial regions and decreases toward more northerly and southern latitudes. Areas with high habitat suitability coincide largely with tropical moist forests, which are characterized by high levels of rainfall and low variability in annual temperature (Rangel-Ch and Arellano-P 2004). Chocó-Darién moist forests, Northern Andean montane forest, and Isthmian-Atlantic moist forests are the best-represented ecoregions in the distribution of this species. In addition, our model produced low habitat suitability scores for higher elevations (i.e., above ~ 1000 m a.s.l.) in the Andes of northern South America and the mountain ranges of Central America (Figure 3B). The effect of elevation on species distributions is realized by changes in ecological variables, such as temperature, rainfall, and solar radiation (Elith and Leathwick 2009) that determine vegetation structure and habitat quality (Pokharel *et al.* 2016). Investigation of the point at which environmental variables diverge to distinguish core versus marginal habitats may result in a better understanding of conditions that favor adaptation to lower-quality habitats.

#### *Species Distribution Range*

The geographic information presented here indicates that the distribution of *Rhinobothryum bovalli* is a disjunct within its range between the

Central American Isthmus and the northwestern portion of South America, where the snake occupies lowlands and premontane forests below ~ 1500 m elevation (Figure 3A; Appendix I). The most eastern reported localities for *R. bovalli* lie on the eastern flank of the Serranía del Perijá (or Sierra de Perijá) in Venezuela (Figure 3A; Rojas-Runjaic and Infante-Rivero 2018). Although the species has been reported to occur in moist forest on the western slope of the Cordillera Oriental, inter-Andean Valleys, and Sierra Nevada de Santa Marta in Colombia (Figure 3A), its presence in the Venezuelan Sierra de Perijá is unanticipated because there is no known basimontane moist corridor along which the species could disperse. However, *R. bovalli* might have colonized the area during the Pliocene (~ 3.30 Mya) prior to an intense global glaciation that occurred after the major known Andean uplift (2000–3000 m a.s.l.) at the end of this period (Lazala 2007). Validation of this hypothesis would require a dated phylogeny for snakes that includes several populations of both species of *Rhinobothryum*.

Our distribution model includes small areas of high probability of occurrence east of the Andes. Included are the eastern slope of the Cordillera Oriental in Colombia (with the Catatumbo Basin), as well as areas around the Maracaibo Lake and the foothills of the Cordillera de Mérida in Venezuela (Figure 4A). Given that the potential habitat for *R. bovalli* is characterized by low seasonal precipitation, high precipitation in the warmest quarter, and low variability in annual temperature (Table 2, Figure 2), it is logical to suppose that arid and semiarid conditions would restrict the distribution of this snake. Therefore, xeric areas in the Guajira-Barranquilla and Paraguaná xeric scrubs ecoregions (Dinerstein *et al.* 2017) and semiarid areas of the Táchira Depression (La Marca 1997) may represent environmental barriers limiting the distribution of this species south Cordillera de Mérida. Unexplored or under-sampled areas must be surveyed to generate more robust SDMs for *R. bovalli*.

Two studies reporting new locality records for *R. bovalli* were published recently—including the first record from the Colombian slope of the Serranía del Perijá (Vera-Pérez *et al.* 2019) and the second from Ecuador (Cruz-García *et al.* 2020) (Figure 3A, Appendix I). These appeared coincident with the submission of this manuscript for publication; thus, neither was included in our species distribution modeling. However, these records coincide with areas of high habitat suitability predicted by our models, especially the noteworthy records from the upper Magdalena Valley (Huila Department) and Serranía del Perijá (Cesar Department) in Colombia, and the southernmost record from Guayas province in Ecuador (Figure 3B). This suggests that our models have a high predictive performance and underscores the importance of this methodological approach for making predictions about the distribution of elusive and rare species (Raxworthy *et al.* 2003, Pearson *et al.* 2007, Anderson and Raza 2010, Chunco *et al.* 2013). Last, it is important to note that other factors not incorporated in our models (e.g., land-cover change and interspecies interactions) also may limit the presence of this species in otherwise suitable areas.

### Historical Biogeography

The phylogenetic relationships of Neotropical colubrids are unresolved. However, recent studies suggest that *Rhinobothryum* may be closely related to Nearctic colubrids such as *Drymarchon*, *Drymoluber*, *Mastigodryas*, and *Palusophis* (Montingelli *et al.* 2019, Zaher *et al.* 2019), with the divergence between *Rhinobothryum* and the clade formed by *Drymoluber*, *Mastigodryas*, and *Palusophis* occurring in the middle Miocene, at ~ 16 Mya (Zaher *et al.* 2019). These results support Duellman's proposal (1990) that the *Rhinobothryum* clade is old and originated in the north long before the lineage arrived in South America shortly after the final closure of the Isthmus of Panama during the Pliocene, at ~ 3.5–4.5 Mya. The allopatric ranges of current species

of *Rhinobothryum* are separated by major Andean mountain ranges, thereby suggesting an ancestral vicariance event (Rojas-Morales 2012). Our models indicate small areas of high habitat suitability for *R. bovalli* in the southeastern foothills of the Cordillera Oriental in Colombia (Figure 4A). This region represents an Andes-Amazon-Orinoquia transition area, where biotic elements of these three large units converge (Hernández-Camacho and Sánchez 1992). Although this area contains pixels of high environmental suitability, it is unlikely that *R. bovalli* occurs there because environmental barriers would limit the distribution of the species south Cordillera de Mérida (discussed above).

It is improbable that an area of sympatry exists between the two species of *Rhinobothryum*, owing to the presence of environmental and topographical barriers limiting their distributions. Pliocene uplift events (i.e., final uplift of the Northern Andes and Huancabamba Depression) may have promoted the divergence between *R. bovallii* and *R. lentiginosum* by isolating eastern from western populations of the genus (Rojas-Morales 2012). The most western known locality for *R. lentiginosum* in Colombia is Araracuara (Solano municipality) at the Caquetá River (IAvH-R 6453–54)—this site is about 400 km (airline) from the Cordillera Oriental in Colombia. This distance suggests that there may be additional dispersal barriers, such as Amazonian rivers that are thought to have played a key role as barriers to dispersal and prompting diversification and speciation events (Pomara *et al.* 2014). Moreover, expansion and retraction of dry tropical forest and grassland habitats in the Amazon and Orinoco basins since the Last Glacial Maximum (Anhuf *et al.* 2006) also might promote diversification.

#### *Implications for Conservation*


Species conservation often must be prioritized to achieve the best outcomes from the resources invested (Hughey *et al.* 2003). Core areas are a priority for the conservation of *Rhinobothryum*

*bovalli*, because these areas possess a unique combination of environmental features that is likely to ensure the survival and maintenance of viable populations through time. Relative to the predicted range of *R. bovalli* (i.e., 491,516 km<sup>2</sup>), the sum of its core areas is minimal (59,353 km<sup>2</sup>), and the cores are fragmented and relatively isolated from one another (Figure 4B). These areas largely coincide with lowland (below ~ 500 m a.s.l.) tropical moist forests, which represent primary habitat for this species. Core areas in Panama and Ecuador coincide with more extensive and least-fragmented forest, and they also possess an important number of national protected areas with strict categories of management (Category I and II *sensu* Dudley 2008)—viz., six National Parks in Panama and two Ecological Reserves in Ecuador. In contrast, the core area in Colombia coincides with areas suffering extensive deforestation primarily resulting from commercial agriculture, cattle ranching, hydroelectric generation, and extractive industry (Garzón and Gutiérrez 2013, Fagua *et al.* 2019); only a regional protected area for strict conservation exists in this region—Serranía de Las Quinchas Regional Natural Park (Figure 4B).

As for many other species, the range of *R. bovalli* comprises a few, large core habitats surrounded by numerous, smaller marginal ones. However, these peripheral populations differ from those inhabiting core areas because often they are patchily distributed, less dense, more isolated, and subject to limited resources and greater environmental variability, all of which contributes to their risk of extinction (Eckert *et al.* 2008, Hardie and Hutchings 2010, Peterman *et al.* 2013). We do not favor the conservation of marginal habitats at the expense of core ones, but we do think it is crucial to determine which marginal habitats are likely to be most important for conservation purposes, and protect them before they are lost (e.g., those inhabited by populations with high evolutionary potential or/and acting as links between existing metapopulations). From a conservation perspective, key areas that can ensure the long-term viability of

populations (e.g., priority habitats within core areas) and connectivity (e.g., marginal “stepping stone” areas connecting metapopulations) should be protected.

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**Appendix I.** Confirmed locality records for *Rhinobothryum bovallii*. Records are ordered in a latitudinal gradient. Institutional codes for museum collections follow those of Sabaj-Perez (2016), except for Circulo Herpetológico de Panamá, Panama (CH), Colección Herpetológica of the Museo de Historia Natural at Universidad de Caldas, Colombia (MHN-UCa), Colección Zoológica Dr. Eustorgio Méndez del Instituto Conmemorativo Gorgas de Estudios de la Salud, Panama (COZEM), Museo de Historia Natural de la Universidad del Cauca, Colombia (MHNUC), University of Kansas Biodiversity Institute Herpetological Collection, USA (KU), University of Texas at Arlington Digital Collection, USA (UTADC). Elevations (m a.s.l.) in bold were estimated from geographic coordinates. Geographic coordinates are provided in decimal degrees (as used to build the models) in the WGS-84 datum. \*Type locality, coordinates from Martínez-Fonseca et al. (2019). <sup>1</sup>Coordinates by M. A. Turcios-Casco. <sup>2</sup>Coordinates from Martínez-Fonseca et al. (2019). <sup>3</sup>Coordinates from Helgen et al. (2013). <sup>4</sup>Coordinates from Cadle (2012). <sup>5</sup>Coordinates estimated or corrected based on the locality provided by the source. <sup>6</sup>Locality corrected based on the coordinates provided by the source. <sup>7</sup>Coordinates from the online catalog of the herpetological collection where the specimen voucher is deposited.

Locality	Record type	Latitude	Longitude	Elevation (m a.s.l.)	Source
<b>HONDURAS</b>					
Gracias a Dios, The Lost City of the Monkey God	UVS-V-01080	15.245244	-84.965346	204	Turcios-Casco <i>et al.</i> 2018
El Paraíso, Trojes, near Arenales <sup>1</sup>	LACM 20488	13.934110	-85.812119	<b>467</b>	Campbell and Howell 1965
<b>NICARAGUA</b>					
Río San Juan, Reserva de la Biósfera Río San Juan, Refugio Bartola	UMMZ 247120	10.970686	-84.333780	65	Martínez-Fonseca <i>et al.</i> 2019
<b>COSTA RICA</b>					
Alajuela, Pital, Laguna Lagarto	Photograph	10.687000	-84.180900	<b>38</b>	iNaturalist 2648512
Limón, Siquirres, near El Carmen <sup>2</sup>	Distribution map	10.210000	-83.460000	<b>20</b>	Savage 2002
Limón, Siquirres, near Encanto <sup>2</sup>	Distribution map	10.190000	-83.390000	<b>10</b>	Savage 2002
Limón, Siquirres, near El Peje de Cairo <sup>2</sup>	Distribution map	10.170000	-83.510000	<b>23</b>	Savage 2002
Limón, Siquirres, near La Lucha <sup>2</sup>	Distribution map	10.160000	-83.440000	<b>19</b>	Savage 2002
Limón, Siquirres, near La Florida <sup>2</sup>	Distribution map	10.100000	-83.580000	<b>266</b>	Savage 2002
Limón, Siquirres*	GNM 1221	10.100000	-83.510000	<b>81</b>	Andersson 1916
Limón, Siquirres, near El Coco	Photograph	10.090437	-83.541017	<b>228</b>	iNaturalist 1562800
Limón, Siquirres, near Guayacán <sup>2</sup>	Photograph	10.040000	-83.530000	<b>565</b>	Savage 2002
Limón, Siquirres, Pacuare River	NCSM 84544	9.998700	-83.542400	<b>243</b>	Norton 2018
Limón, Siquirres, 3 km NE Tres Equis	LACM 154272	9.990613	-83.570140	<b>500</b>	Feeney 2018
Limón, Valle de La Estrella, near Bananito Sur <sup>2</sup>	Distribution map	9.850000	-82.970000	<b>20</b>	Savage 2002
Limón, Banano River Basin <sup>2</sup>	Distribution map	9.840000	-83.190000	<b>1025</b>	Savage 2002
Limón, Talamanca, near Bribri <sup>2</sup>	Distribution map	9.680000	-82.880000	<b>162</b>	Savage 2002
Limón, Talamanca, near Shiroles <sup>2</sup>	Distribution map	9.600000	-82.960000	<b>188</b>	Savage 2002
<b>PANAMA</b>					
Colón, Portobelo, Cerro Bruja	LACM-R 2380-84	9.471734	-79.567519	946	This study
Bocas del Toro, Changuinola, La Gloria <sup>2</sup>	Reported	9.460000	-82.440000	<b>9</b>	Lotzkat 2014
Panamá, Parque Nacional Chagres <sup>2</sup>	Distribution map	9.430000	-79.430000	<b>350</b>	Ray and Ruback 2015

Appendix I. Continued.

Locality	Record type	Latitude	Longitude	Elevation (m a.s.l.)	Source
Guna Yala, Llano Carti, Burbayar <sup>2</sup>	Photograph	9.420000	-79.000000	54	McConnell 2014
Panamá, Parque Nacional Chagres <sup>2</sup>	MCZ R-50220	9.360000	-79.440000	496	Ray and Ruback 2015
Guna Yala, Narganá, footpath Nusagandi	CH 08759	9.344417	-78.996390	335	This study
Guna Yala, Cerro Brewster, Parque Nacional Chagres	CH 06095	9.321585	-79.282010	872	EMBL-EBI 2019
Colón, Boyd-Roosevelt Highway, near Rita River	UMMZ 147770	9.320000	-79.790000	151	Martínez-Fonseca <i>et al.</i> 2019
Panamá, Chepo, Zahinas, Cocobolo Nature Reserve <sup>2</sup>	Photograph	9.300000	-79.210000	444	Leenders 2017
Colón, 19 Km NW Gamboa, Canal Zone (navy Pipeline Road)	COZEM-REP 0372	9.242696	-79.822077	61	This study
Colón, Pipeline Road	CH 07631	9.170808	-79.753484	136	This study
Panamá, Road to Tortí	CH 10437	9.160108	-78.698850	90	This study
Colón, Gamboa, Canal Zone	COZEM-REP 0371	9.125603	-79.704555	156	This study
Panamá, Panamá city, Canal Zone <sup>2</sup>	MCZ R-45404	9.000000	-79.550000	63	Morris 2018
Bocas del Toro, Changuinola River basin <sup>2</sup>	Reported	8.950000	-82.439000	950	Lotzkat 2014
Guna Yala, San Blás, Sasardi camp <sup>3</sup>	KU 112463–64	8.950000	-77.817000	41	Brown 2018
Guna Yala, San Blás, Summit camp <sup>4</sup>	KU 112462	8.916000	-77.850000	361	Brown 2018
Colón, Donoso, Petaquiya stream	Photograph	8.826253	-80.678855	180	This study
Coclé, La Pintada, El Copé, near Omar Torrijos National Park	CH 05938	8.662691	-80.590435	712	EMBL-EBI 2019
Panamá Oeste, near Altos del María <sup>5</sup>	Observation	8.662151	-80.003867	276	Ray and Ruback 2015
Coclé, La Pintada, El Copé, near Omar Torrijos National Park	Photograph	8.644611	-80.589944	926	iNaturalist 1309822
Chiriquí, Boquerón, Bágala, Cerro Colorado, Escopeta camp	USNM 297732	8.518700	-82.537100	900	USNM
Darién, Ucurganti River, ca 7 km above mouth <sup>2</sup>	KU 112461	8.510000	-77.800000	46	Brown 2018
Veraguas, Rasca River, above El Paredón <sup>6</sup>	SMF 91577	8.491000	-81.169400	775	Lotzkat 2014
Veraguas, Santa Fé, Cerro Tute	SMF 90022	8.488140	-81.109840	1116	Carrizo 2010
Darién, Río Tuirá at Río Mono <sup>2</sup>	KU 112460	8.460000	-78.130000	96	Brown 2018
Veraguas, Calobre, La Laguna	COZEM-REP 0370	8.455487	-80.851749	651	This study
Darién, Pinogana, Reserva Forestal Canglón	CH 00535	8.266388	-77.796801	17	This study
Darién, Serranía del Bagre, Reserva Chepigana de Darién <sup>2</sup>	Reported	8.220000	-78.000000	20	Martínez-Fonseca <i>et al.</i> 2019
Darién, Pinogana, Yaviza	MCZ R-38236	8.152900	-77.697300	10	Morris 2018

## Appendix I. Continued.

Locality	Record type	Latitude	Longitude	Elevation (m a.s.l.)	Source
Darién, Tacarcuna <sup>2</sup>	KU 75749	8.170000	-77.270000	1535	Brown 2018
Veraguas, Mariato, Cerro Hoya National Park	UTADC 8654	7.345690	-80.691330	533	Flores <i>et al.</i> 2016
Los Santos, Tonosí, Ave María, Guanico River	KU 107839	7.316700	-80.466700	104	Brown 2018
<b>COLOMBIA</b>					
La Guajira, Dibulla, Pozo Azul, Santa Rita de la Sierra	Specimen	11.193333	-73.267000	124	This study
Magdalena, Santa Marta, Minca, Arimaca	Photograph	11.115820	-74.124465	843	iNaturalist 16216599
Bolivar, San Juan Nepomuceno, Miraflores	Photograph	9.925884	-75.093211	167	Barreto-Martínez and Rojas-Morales 2020
Cesar, La Jagua de Ibirico, La Victoria de San Isidro	Observation	9.595222	-73.158861	623	Vera-Pérez <i>et al.</i> 2019
Sucre, Colosó, El Sereno Alto, El Sereno stream	IAvH-R-5486	9.542939	-75.340000	300	Borja-Acosta 2017
Córdoba, Tierralta, Palmira, La Plumilla	ICN 053929	8.033056	-76.170556	162	Rojas-Morales 2012
Chocó, Unguía, Peye	UTCH 1172	7.920430	-77.096770	40	Rengifo and Pino 2018
Cesar, San Alberto, Miramar <sup>5</sup>	ICN 054087a	7.917869	-73.401721	715	Rojas-Morales 2012
Antioquia, Chigorodó, near Turbo	USNM 154026	7.670000	-76.780000	10	Martínez-Fonseca <i>et al.</i> 2019
Santander, Sabana de Torres, Reserva Cabildo Verde	Photograph	7.344955	-73.500397	168	This study
Antioquia, Mutatá	MHUA-R 14276	7.247222	-76.439167	140	Rojas-Morales 2012
Bolívar, Cantagallo, Caguí	Photograph	7.252528	-74.033500	98	This study
Antioquia, Valdivia, La Habana	Observation	7.238667	-75.403556	642	Vera-Pérez <i>et al.</i> 2019
Antioquia, Briceño, between Ticuita and Capitan stream <sup>7</sup>	MHUA-R 14703	7.112200	-75.554200	1350	Rojas-Morales 2012
Santander, Girón, Sogamoso	Specimen	7.095583	-73.382694	563	This study
Santander, Betulia, La Putana	Photograph	7.094102	-73.416165	398	This study
Santander, Girón, Sogamoso	Photograph	7.090167	-73.384389	457	This study
Antioquia, Toledo, La Cascarela	Observation	7.085917	-75.694333	381	Vera-Pérez <i>et al.</i> 2019
Santander, Girón, Sogamoso, Las Cruces	UIS-R 2776	7.083218	-73.383432	323	Ramírez-Pinilla and Meneses-Pelayo 2020
Antioquia, Toledo, El Valle	MHNUC-He-0713	7.041222	-75.670694	523	Vera-Pérez <i>et al.</i> 2019

Appendix I. Continued.

Locality	Record type	Latitude	Longitude	Elevation (m a.s.l.)	Source
Antioquia, Sabanalarga, Orobajo	Observation	7.025000	-75.793194	394	Vera-Pérez <i>et al.</i> 2019
Santander, Betulia, Sogamoso, Corintios	Photograph	7.000917	-73.393676	349	This study
Santander, Betulia, Sogamoso, Corintios	Photograph	6.996707	-73.416801	761	This study
Antioquia, Anorí, El Retiro <sup>7</sup>	MHUA-R 14559	6.985000	-75.089722	<b>871</b>	Rojas-Morales 2012
Antioquia, Anorí, Dam wall at Porce III dam	Photograph	6.938611	-75.140219	<b>633</b>	iNaturalist 41131
Antioquia, Amalfí, María Teresa	Photograph	6.924417	-75.142528	657	Vera-Pérez <i>et al.</i> 2019
Santander, El Carmen de Chucurí, Riosucio	IAvH-ACX566	6.577139	-73.573111	760	Acosta-Galvis and Borja-Acosta 2018
Antioquia, Maceo, Las Brisas, Santa Barbara farm <sup>7</sup>	MHUA-R 14583	6.546900	-74.643600	<b>577</b>	Rojas-Morales 2012
Antioquia, Puerto Berrío, Hermilda	Photograph	6.411665	-74.701710	<b>687</b>	iNaturalist 8360221
Santander, Landazurí <sup>7</sup>	ICN 054086	6.358611	-73.899444	<b>202</b>	Rojas-Morales 2012
Antioquia, Puerto Nare, Canteras	Observation	6.278500	-74.675908	214	Vera-Pérez <i>et al.</i> 2019
Antioquia, San Carlos, El Jordán, Juanes, Casino Viejo	ISAGEN 46-2978	6.219139	-74.817139	<b>717</b>	Martínez-Fonseca <i>et al.</i> 2019
Santander, Sucre, La Tipa	UIS-R 4278	5.956132	-74.065343	<b>564</b>	Ramírez-Pinilla and Meneses-Pelayo 2020
Boyacá, Puerto Boyacá, Las Pavas, Puerto Romero <sup>7</sup>	ICN 054085	5.843056	-74.326667	<b>359</b>	Rojas-Morales 2012
Antioquia, La Miel River <sup>2</sup>	IAvH-R-5880–81	5.710000	-74.730000	<b>169</b>	Borja-Acosta 2017
Caldas, Norcasia, Quebra Roque, Reserva Río Manso	Observation	5.664167	-74.785833	280	Rojas-Morales 2012
Chocó, Atrato, San Martín de Purré	UTCH 0299	5.649520	-76.579380	80	Rengifo and Pino 2018
Chocó, Atrato, Samurindó	UTCH 0339	5.587500	-76.654167	36	Rengifo and Pino 2018
Caldas, Norcasia, road to Berlín (Samaná)	MHN-UC 0306	5.572690	-74.904020	648	This study
Caldas, Norcasia, Puerto Norcasia, Amani dam	MHN-UC 0238	5.564766	-74.907133	550	This study
Caldas, Samaná, Cañaverál <sup>7</sup>	MHUA-R 14142	5.537028	-74.908694	<b>623</b>	Rojas-Morales 2012
Chocó, Unión Panamericana, Salero	UTCH 1240, 1244	5.360278	-76.645833	115	Rengifo and Pino 2018

Appendix I. *Continued.*

Locality	Record type	Latitude	Longitude	Elevation (m a.s.l.)	Source
Caldas, La Dorada, Guarinocito, Purnio	Photograph	5.355227	-74.798606	305	Flórez-Jaramillo and Barona-Cortés 2016
Caldas, Victoria, El Llano <sup>7</sup>	MHUA-R 14785	5.329388	-74.849305	289	Rojas-Morales 2012
Valle del Cauca, Buenaventura, San Cipriano	Observation	3.834167	-76.890556	115	Rojas-Morales 2012
Valle del Cauca, Buenaventura, Guamía	UVC-13662, 13685	3.731667	-76.958333	41	Rojas-Morales 2012
Valle del Cauca, Buenaventura, Bajo Anchicayá	UVC-9591	3.613741	-76.910950	536	Salinas and Valencia-Aguilar 2018
Valle del Cauca, Buenaventura, Bajo Anchicayá dam	UVC-5266	3.533263	-76.867508	702	Salinas and Valencia-Aguilar 2018
Huila, Gigante, La Honda	MHNUC-He-0693	2.369611	-75.571722	725	Vera-Pérez <i>et al.</i> 2019
<b>VENEZUELA</b>					
Zulia, Rosario de Perijá, Puerto Nuevo, Cogollo River	MBLUZ R-486	10.346944	-72.516111	350	Rojas-Runjaic and Infante-Rivero 2018
Zulia, Kasmera, Yasa River basin, Sierra de Perijá	MBLUZ R-182	9.943333	-72.749167	270	Pons 1965
Zulia, Ipika, Tokuko River, Sierra de Perijá	MHNLS 20940	9.881667	-72.516111	560	Rojas-Runjaic and Infante-Rivero 2018
<b>ECUADOR</b>					
Esmeraldas, Eloy Alfaro, Playa de Oro, Reserva Cotacachi Cayapas	QCAZ-R 10703	0.828464	-78.722010	567	Pazmiño-Otamendi 2019
Esmeraldas, Zapallo Grande	QCAZ-R 1595	0.773994	-78.936000	487	Pazmiño-Otamendi 2019
Esmeraldas, Eloy Alfaro, Telembí, Reserva Tesoro Escondido	QCAZ-R 15012	0.493760	-79.136080	675	Pazmiño-Otamendi 2019
Manabí, Cuaque River, 10 min from Pedernales, Canoa way	QCAZ-R 8962	-0.021260	-80.068930	25	Pazmiño-Otamendi 2019
Manabí, 15 km Pedernales, Jama way	QCAZ-R 5757	-0.078821	-80.021630	36	Pazmiño-Otamendi 2019
Manabí, Jama, Reserva Ecológica Jama-Coaque	QCAZ-R 4652	-0.090978	-80.147190	3	Pazmiño-Otamendi 2019
Manabí, Jama, Reserva Ecológica Jama-Coaque	QCAZ-R 5889	-0.109405	-80.117529	392	Pazmiño-Otamendi 2019
Manabí, Jama, Reserva Ecológica Jama-Coaque	No voucher	-0.110156	-80.128183	278	Steinke 2016
Guayas, Naranjal, Molleturo-Mollepungo, Cerro de Hayas	ZSFQ 4081	-2.730800	-79.628600	110	Cruz-García <i>et al.</i> 2020