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Beneficial and detrimental effects of schema incongruence on memory for contextual events

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Abstract

Mental schemas provide a framework into which new information can easily be integrated. In a series of experiments, we examined how incongruence that stems from a prediction error (van Kesteren, Ruiters, Fernández, & Henson, 2012) modulates memory for multi-component events that instantiated pre-existing schemas. Each event consisted of four stimulus pairs with overlapping components, presented in four blocks (A-B, B-C, C-D, D-A). A-B pairs elicited contextual expectations (A: Farm, B: Tractor) that were either met by a congruent C component (C: Farmer) or violated by an incongruent one (C: Lawyer). The baseline condition included unrelated pairs, where the C component was neither congruent nor incongruent. In experiment 2, events were presented in successive trials instead of blocks, and eye movements were recorded to analyse allocation of attention. Memory was tested through old-new item recognition followed by cued recall. Across experiments, recognition and recall performance for incongruent components was reduced compared to congruent components. Incongruent items were in some cases more accurately retrieved compared to unrelated ones, depending on task demands. Additionally, better recall was observed in the incongruent D-A pairs, compared to congruent and unrelated ones, because of reduced interference from C components. Eye-tracking revealed an increased number of fixations on C components in the incongruent and unrelated conditions. These results suggest that the integration of incongruent items into an episode is impaired, compared to congruent items, despite the contextual surprise and increased attention they elicited at encoding. However, there was a beneficial effect of prediction error on memory performance, compared to a baseline, depending on the task employed.

Keywords: congruence effects, schema, prediction error, event integration

Abstract word count: 259 words

1 Introduction

2 Schemas and their role in supporting memory have been a topic of intensive research over
3 several decades (Bartlett, 1932). Schema effects can benefit memory in two ways; firstly, the
4 mere existence of a schema could enhance learning of new information. Previous research
5 suggests memory performance is improved when the items can be assimilated into an
6 existing schema (McClelland, 2013; Tse et al., 2007, 2011). Secondly, performance can be
7 assessed across congruent and incongruent information. A common observation, referred to
8 as the congruency effect, is that schema-congruent information is better recognised and
9 recalled than incongruent information (Atienza, Crespo-Garcia, & Cantero, 2010; Craik &
10 Tulving, 1975; Staresina, Gray, & Davachi, 2009; van Kesteren, Rijpkema, Ruiters, &
11 Fernandez, 2010). However, van Kesteren and colleagues (2012) proposed a model (SLIMM)
12 showing schemas can enhance memory for both congruent and incongruent information, via
13 different mechanisms. The key modulator in this model is the prediction error elicited by the
14 incongruent item (Greve, Cooper, Kaula, Anderson, & Henson, 2017). A prominent aspect
15 that remains elusive is to what extent memory is enhanced by the prediction error (e.g.
16 whether memory performance is equivalently good for congruent and incongruent items).
17 Here, we test behavioural predictions of this model by combining contextual surprise with
18 schema-incongruent items, embedded in multi-component events (Horner, Bisby, Bush, Lin,
19 & Burgess, 2015; Horner & Burgess, 2013).

20

21 SLIMM posits that incongruence could lead to superior memory when the contextual
22 schema provides a strong constraint (van Kesteren et al., 2012). In such cases, the
23 incongruent item elicits a prediction error, leading to better memory through the creation of
24 new representations. Importantly, the degree to which incongruence benefits memory,
25 compared to a schema-less, baseline level, remains unclear. The model also accounts for
26 congruency effects, as encountering congruent information results in medial prefrontal

27 cortex (mPFC) activation of the schema, which in turn facilitates encoding in the medial
28 temporal lobe (MTL; for similar ideas see Brod, Werkle-Bergner, & Shing, 2013; Preston &
29 Eichenbaum, 2013). A key aspect of the SLIMM model is that the fate of incongruent
30 information in memory is determined by the context it is embedded in. Most of the studies
31 examining schema effects build upon pre-experimental knowledge (Bayen & Kuhlmann,
32 2011; Bein, Reggev, & Maril, 2014; van Kesteren et al., 2010) and the relationship between a
33 pair of items, or their level of semantic relatedness (Bein et al., 2014; Staresina et al., 2009;
34 van Kesteren et al., 2013). For example, purple-banana would be an incongruent pair,
35 whereas yellow-banana constitutes a congruent one. However, this design does not
36 necessarily allow for predictions to develop at encoding. Subsequently, during retrieval,
37 there is no episodic contextual setting that would reinstate the schema (van Kesteren et al.,
38 2012).

39

40 Incongruence that stems from a prediction error should be accompanied by better
41 retention, supported by MTL engagement (Greve et al., 2017; van Kesteren et al., 2012).
42 Importantly, for a prediction to be wrong, it must first be elicited (Kumaran & Maguire,
43 2007). We therefore utilised interleaved learning of events comprising pairs of components,
44 to allow for predictions to be developed and violated. Previous research suggests such multi-
45 component events are well suited for this purpose, as they promote binding of components
46 into contextual events (Horner et al., 2015; Horner & Burgess, 2013). Therefore, the first pair
47 of components presented can be used to implicitly set the schema-related predictions
48 (Schlichting & Preston, 2015). Additionally, this paradigm makes it possible to test how
49 incongruent components affect the rest of the contextual event. Specifically, whether
50 incongruent components can be integrated into an otherwise congruent event (Bein et al.,
51 2015), relying on interactions between mPFC and MTL (Schlichting & Preston, 2015, 2017).
52 Thus, we can measure not only the independent recollection of incongruent components,

53 but also their indirect effect on adjacent congruent components. According to SLIMM (van
54 Kesteren et al., 2012), if incongruent representations are reactivated at retrieval (similarly to
55 congruent ones), we would expect equivalent levels of interference from incongruent and
56 congruent C components in adjacent A-B and D-A pairs.

57

58 In the studies reported here, we used events consisting of four pairs. Each pair shared a
59 common component (A-B, B-C, C-D, D-A), to promote retrieval of the previous pair during
60 study (Caplan, Rehani, & Andrews, 2014), as well as to allow for components to be
61 integrated into an event (Burton, Lek, & Caplan, 2017; Schlichting & Preston, 2017). We
62 extend previous literature on schema effects by actively eliciting contextual predictions that
63 stem from existing schemas, as opposed to relying solely on the level of relatedness of two
64 items. On this basis, we hypothesised that incongruence will modulate memory both of the
65 incongruent components and the event they are incorporated in. Whilst SLIMM predicts
66 memory enhancement of incongruent components (van Kesteren et al., 2012), it remains
67 unclear to what extent. By including an unrelated 'baseline' condition, where there is no
68 schema, we will test not only whether the presence of a prediction error supports schema-
69 related memory, but also to what extent. For example, congruent and incongruent items
70 could be equally better than unrelated ones, show a graded response (congruent >
71 incongruent > unrelated), or be equivalent to unrelated events.

72

73 To assess differences in memory performance between congruent and incongruent items,
74 compared to an unrelated baseline, we used contextual events (see Experiment 1a in the
75 Supplementary Materials for comparison between congruent-incongruent alone). The first
76 pair in the event is location-object, as it easily instantiates the schema (or lack thereof) for
77 the following items (farm-tractor immediately brings to mind other farm-related items,
78 whereas golf course-torch does not intrinsically belong to a specific context or schema, see

79 Bar & Aminoff, 2003). Both congruent and incongruent events had three components that
80 were schema-congruent (A,B,D; see Figure 1 for examples). In the schema-based conditions,
81 A-B pairs elicited contextual expectations that were either met by a congruent C component
82 or violated by an incongruent one. In unrelated events, components did not share any
83 common contextual information. Memory was tested in two steps, first a yes/no item
84 recognition task for each component, followed by a cued recall for the adjacent component
85 (e.g. B-? or ?-D, see Figure 1). Cued recall was tested only for initially recognised
86 components. Multiple retrieval trials were employed to test effects of task demand
87 (recognition vs. recall, forward vs. backward cued recall).

88

89 *Figure 1 about here*
90

91 **Experiment 1**

92 Results and Discussion

93 *Item recognition.* A three (congruence: congruent, incongruent, unrelated) by two
94 (components: A, B, C, D) repeated measures ANOVA (Figure 2c) was conducted for corrected
95 recognition responses (hits – false alarms). Greenhouse-Geisser correction for the sphericity
96 assumption of ANOVAs are reported where appropriate, and all post-hoc analyses reported
97 are Bonferroni corrected. When multiple t-tests were computed, a threshold p value is
98 reported for all of them (e.g. all p's < .05). Despite near-ceiling recognition performance
99 (overall average of 93% accuracy), a main effect of congruence was found $F(2,60) = 3.87$ $p =$
100 $.026$, $\eta_p^2 = .113$, with follow-up paired t-tests indicating congruent components were more
101 easily recognised than incongruent ones $t(30) = 2.95$, $p = .006$, Cohen's $d = .53$, and
102 unrelated cues $t(30) = .258$, $p = .015$, Cohen's $d = .464$. No differences were observed
103 between incongruent and unrelated components $t(30) = .851$, $p = .4$.

104

105 *Cued recall.* We first tested whether there were any effects of order on trials (AB - forward,
106 BA - backward) from the same pair (A-B). The three (congruence) by four (pair) by two
107 (order) repeated measures ANOVA yielded both a main effect of order $F(1,30) = 11.2, p =$
108 $.002, \eta_p^2 = .272$ and a significant three-way interaction $F(6,180) = 4.2, p = .008, \eta_p^2 = .123$. This
109 suggests there were different patterns of performance in each condition between the
110 forward and backwards trials (see Figures 2a and 2b). Therefore, we carried out two
111 separate three (congruence) by four (pair) ANOVAs, one for each order (forward X - ?, and
112 backward ? - X). The forward ANOVA (Figure 2a) revealed a significant interaction between
113 congruence and pairs $F(6,180) = 5.37, p < .001, \eta_p^2 = .152$. Similarly, the ANOVA for backward
114 pairs (Figure 4b) revealed an interaction effect $F(6,180) = 4.01, p < .001, \eta_p^2 = .118$.

115

116 Planned comparisons between the congruence conditions for each trial were carried out to
117 examine differences between conditions. Congruent BC and CD trials (forward order) were
118 better recalled than incongruent and unrelated ones (all p 's $\leq .011$). No significant
119 differences were observed between incongruent and unrelated components in these two
120 trials (all p 's $\geq .32$). In the backward order trials, a benefit of incongruent over unrelated
121 trials was found for DC $t(30) = 2.774, p = .009, \text{Cohen's } d = .5$, and a trend towards it was
122 observed in CB trials $t(30) = 1.75, p = .091, \text{Cohen's } d = .31$. These results show the order
123 effect was due to a graded pattern (congruent > incongruent > unrelated) in the backward,
124 but not forward order. Finally, in DA and AD trials, incongruent components were better
125 recalled than unrelated and congruent ones and incongruent AB were more accurately
126 retrieved compared to congruent ones (all p 's $\leq .01$).

127

128

Figure 2 about here

129

130 *Interference analysis.* To further elucidate the benefit of incongruence on D-A and A-B pairs,
131 we examined the erroneous answers for cued recall trials (Figure 2d). We inspected both
132 trials comprising the D-A pairing (AD and DA) together, to have a sufficient number of trials
133 and participants included (due to order effects in previous analysis we also carried out this
134 analysis separated by order, which showed similar results despite a lower number of trials
135 and participants included, see Supplementary Materials). We examined how many of the
136 errors were due to interference from C items, compared to a baseline (erroneous recall of
137 B/D). One out of the 31 participants had missing values for one pair, therefore data from 30
138 participants were analysed. A three (congruence: congruent, incongruent and unrelated) by
139 two (item: C and D/B) by two (pair: A-B and D-A) repeated measures ANOVA revealed
140 interaction effects of congruence by item $F(2,58) = 7.37, p = .001, \eta_p^2 = .203$, and item by pair
141 $F(1,29) = 10.36, p = .003, \eta_p^2 = .263$. The three-way interaction effect was not significant
142 $F(2,58) = 1.1, p = .34$.

143

144 Post-hoc tests revealed more C errors in congruent D-A and A-B pairs compared to
145 incongruent and unrelated ones (all p 's < .004). There were also more C errors in
146 incongruent D-A and A-B pairs compared to unrelated ones (all p 's < .038). When comparing
147 interference between C items and baseline B/D items, in D-A pairs we observed less
148 interference from incongruent C compared to incongruent B components ($t(29) = 2.14, p =$
149 $.041$, Cohen's $d = .391$). In A-B trials, on the other hand, interference to incongruent pairs is
150 equivalent between C and D items ($t(29) = .084, p = .934$), but there is more interference
151 from congruent C items compared to congruent D ones ($t(29) = 2.43, p = .021$, Cohen's $d =$
152 $.444$).

153

154 The results described above show that memory performance for congruent components is
155 superior to incongruent and unrelated ones, in all testing formats. Incongruent items are, in

156 backward cued recall, more accurately retrieved than unrelated ones. The interference
157 analysis showed better recall performance for congruence-matched pairs (A-B, D-A) of
158 incongruent events is due to reduced interference from C components, suggesting they are
159 less integrated into the event. Additionally, incongruent C items caused more interference
160 than unrelated ones, showing a similar pattern of responses as in backward cued recall.
161 Overall, these findings suggest that congruence benefits memory performance across the
162 board, but memory for incongruent components was modulated by task demands.
163 Differences in performance between incongruent and unrelated pairs, stemming from the
164 prediction error associated with incongruence, were observed in backward cued recall and
165 the interference analysis. In both cases, a graded response was observed (congruent >
166 incongruent > unrelated). In the recognition and forward recall, on the other hand, memory
167 for incongruent components was equivalent to unrelated ones.

168

169 **Experiment 2**

170 To address the integration account of reduced performance for incongruent components,
171 we devised Experiment 2 to allow for easier integration of the pairs into a cohesive event. To
172 do so, events were presented as trial-by-trial pairs, rather than across blocks. Trial-by-trial
173 presentation differs in the contextual setting of learning. In Experiment 1, each pair was
174 compared to other pairs of the same kind (i.e. B-C pairs from different events were always
175 studied together). In Experiment 2, on the other hand, the 'reference point' is the previous
176 pair from the same event (B-C will follow A-B from the same event). This change would
177 allow for a quicker build-up of predictions, as the event pairs would now be temporally
178 closer than in the previous experiments. We added an associative inference task between A
179 and C items, which were not shown together, to test differences in integration levels. To
180 examine whether unexpected pairs are processed differently at encoding, we measured eye-
181 movements during this stage. We reasoned that changes in fixation patterns observed at

182 encoding would indicate a different allocation of resources to components that are more
183 difficult to encode and integrate into the event.

184

185 Eye-tracking results

186 We first subjected the number of fixations to a three (congruence) by four (pair) by 2 (area
187 of interest, AOI) ANOVA (Figure 3a), which showed a significant main effect of pair
188 $F(2,1,56.73) = 73.6, p < .001, \eta_p^2 = .732$, a congruence by pair interaction $F(6,162) = 5.63, p$
189 $< .001, \eta_p^2 = .173$ and a significant AOI by pair interaction $F(2.23,60.32) = 25.7, p < .001, \eta_p^2 =$
190 $.49$. Post-hoc comparisons showed the least amount of fixations on the B item of unrelated
191 A-B pairs, when compared to the congruent and incongruent conditions (all p 's $\leq .008$).
192 Additionally, unrelated B-C pairs were associated with fewer fixations on B, when compared
193 to congruent and incongruent pairs (all p 's $< .001$). For the C components in the B-C pairs,
194 congruent components were associated with the lowest number of fixations compared to
195 incongruent and unrelated ones (all p 's $\leq .008$). A similar analysis for fixation durations
196 (Figure 3b) revealed significant main effects of pair $F(1.53, 42.91) = 76.33, p < .001, \eta_p^2 = .732$
197 and AOI $F(1,28) = 13.21, p = .001, \eta_p^2 = .321$, as well as a pair by congruence interaction
198 $F(6,168) = 5.9, p < .001, \eta_p^2 = .174$, and a pair by AOI interaction $F(3.84) = 19.98, p < .001,$
199 $\eta_p^2 = .416$. Planned comparisons indicated that in unrelated B-C pairs, fixations on the B
200 components were shortest and those on C components were longest, compared to
201 congruent and incongruent pairs (all p 's $\leq .003$).

202

203 *Figure 3 about here*

204

205 Behavioural Results

206 *Item recognition.* A three (congruence: congruent, incongruent, unrelated) by four
207 (component: A, B, C, D) repeated measures ANOVA was conducted for corrected recognition

208 responses. Again, a main effect of congruence was found $F(2,54) = 14.83$ $p < .001$, $\eta_p^2 = .355$,
209 with follow-up paired t-tests indicating congruent components were more easily recognised
210 than incongruent and unrelated ones (all p 's $\leq .002$). No differences were observed between
211 incongruent and unrelated components ($p > .2$).

212

213 *Cued recall.* We first tested whether there were any effects of order. Similar to Experiment
214 1, the three (congruence) by four (pair) by two (order) repeated measures ANOVA yielded a
215 significant three-way interaction $F(6,162) = 4.39$, $p < .001$, $\eta_p^2 = .14$. This suggests there were
216 different patterns of performance in each condition between the forward and backwards
217 orders. Therefore, we performed two separate three (congruence) by four (pair) ANOVAs,
218 one for each presentation order (forward $X - ?$, and backward $? - X$). The forward ANOVA
219 (Figure 4a) revealed a significant congruence by pair interaction $F(4.14, 112) = 5.67$, $p < .001$,
220 $\eta_p^2 = .174$. The ANOVA for backward trials (Figure 4b) was akin to that for the forward trials,
221 with an interaction between congruence and pair $F(6,162) = 9.37$, $p < .001$, $\eta_p^2 = .258$.

222

223 Planned comparisons between the congruence conditions for each trial were then carried
224 out. For AB and DA trials a similar pattern emerged, whereby there were no significant
225 differences between congruent trials and any of the other conditions (all p 's $\geq .02$), but a
226 significant benefit of incongruent trials over unrelated ones was observed (all p 's $\leq .002$).
227 Unrelated BA trials were associated with reduced recall performance compared to
228 congruent and incongruent ones (all p 's $\leq .006$). For BC, CB, CD and DC trials there were
229 again similar findings, with congruent trials associated with better recall compared to
230 incongruent and unrelated trials (all p 's $\leq .013$), but no difference was observed between
231 incongruent and unrelated trials (all p 's $\geq .668$). These results suggest that performance for
232 trials that are part of incongruent events tracks their level of relatedness.

233

234 *Figure 4 about here*

235

236 *Associative inference retrieval task (A-C)*. A three (congruence) by two (component: A vs. C)

237 ANOVA yielded a significant main effect of congruence $F(1.57, 40.97) = 105.46, p < .001, \eta_p^2$

238 = .802 (Figure 4c). The main effect of component was not significant $F(1,26) = .016, p = .901,$

239 as was the interaction effect $F(1.43, 37.23) = .58, p = .508$. Post-hoc tests revealed that

240 congruent components were associated with better performance compared to incongruent

241 and unrelated ones, and a benefit of incongruent components over unrelated ones (all p 's <

242 .001).

243

244 *Interference analysis*. . A three (congruence: congruent, incongruent and unrelated) by two

245 (component: C and D/B) by two (pair: A-B and D-A) repeated measures ANOVA (Figure 4d)

246 revealed a significant interaction between congruence and item $F(2,54) = 34.8, p < .001, \eta_p^2$

247 = .563). The other two-way interactions, as well the three-way interaction were not

248 significant (all p 's > .248). Planned comparisons showed increased interference from

249 congruent C components in A-B and D-A pairs, compared to incongruent and unrelated

250 components (all p 's < .001). There were no significant differences between incongruent and

251 unrelated C components (all p 's > .528). Additionally, there was more interference from

252 incongruent B/D (baseline) components, compared to incongruent C components, in D-A

253 pairs ($t(27) = 4.45, p < .001, \text{Cohen's } d = .841$) and A-B pairs, respectively ($t(27) = 3.44, p =$

254 .002, Cohen's $d = .778$).

255

256 General Discussion

257 In a series of experiments, we found that unexpected incongruent components were less

258 likely to be recognised and recalled, compared to congruent ones. This pattern was observed

259 both when events were presented across blocks, to allow for predictions to develop
260 gradually, and when events were constructed trial-by-trial, to promote easier integration.
261 Examination of fixation patterns during encoding of such events revealed increased fixations
262 on the first unexpected incongruent component, compared to their congruent counterparts,
263 suggesting they were more difficult to encode. Our results suggest the presence of a
264 prediction error in incongruent pairs did not enhance memory to the level observed in
265 congruent events. However, depending on task demands, it did benefit memory compared
266 to unrelated components (showing a graded pattern of responses). Finally, a beneficial
267 effect of incongruence was observed in A-B and D-A pairs, which were only indirectly related
268 to the incongruent C component.

269

270 The advantageous role of schema congruence on memory performance, compared to both
271 incongruent and unrelated items, is in line with the congruency effect. Previous findings, and
272 predictions from SLIMM, suggest congruent items benefit from the existing strong
273 representation of the schema (Craik & Tulving, 1975; Staresina et al., 2009; van Kesteren et
274 al., 2012). Previous experiences of similar associations (e.g. farm-tractor) are reactivated by
275 the mPFC (Brod et al., 2013; Preston & Eichenbaum, 2013; Schlichting & Preston, 2015), and
276 proposed to be more readily available during retrieval (Hemmer & Steyvers, 2009;
277 Moscovitch & Craik, 1976; Steyvers & Hemmer, 2012). This notion is in line with our finding
278 that congruent A-B and D-A pairs are more prone to interference from C components,
279 showing reduced recall accuracy. Incongruent components, and unrelated ones even more
280 so, are more difficult to integrate into the event (Bein et al., 2015; Craik & Tulving, 1975),
281 resulting in less associative competition (Caplan et al., 2014), or a more constrained search
282 space (Anderson, 1981) for retrieving A-B and D-A.

283

284 Further evidence for the reduced integration of incongruent C components can be found in
285 forward-order trials in Experiment 1, as well as in all retrieval trials in Experiment 2.
286 Incongruent components show comparable results, in memory performance and eye-
287 fixation patterns, to those of their unrelated counterparts. The eye-tracking results indicate
288 increased effort invested in encoding incongruent items, but this effort does not come into
289 fruition later in retrieval performance. The encoding part of this effect can be attributed to
290 MTL-driven function, showing the prediction error associated with an incongruent
291 component promotes more elaborate encoding. During retrieval, congruent information
292 seems to be dominating, especially in trials where a congruent item (B or D) cues retrieval of
293 an incongruent item. In this case, mPFC involvement would potentially direct retrieval
294 towards congruent representations (Preston & Eichenbaum, 2013; van Kesteren et al.,
295 2012). Although previous studies have shown similar encoding and retrieval effects
296 independently (van Kestern et al., 2010, 2013), future neuroimaging studies could examine
297 the co-occurrence of such effects. Additionally, due to the use of bidirectional cued-recall
298 tests, we could not directly correlate fixations at encoding with later retrieval performance,
299 but this would be an interesting effect to examine.

300

301 Incongruent components also served as worse cues for their associates (Schulman, 1974),
302 showing impaired recall performance (equivalent to unrelated pairs). This finding is more
303 difficult to interpret in light of SLIMM's predictions (van Kesteren et al., 2012), as successful
304 retrieval in this case requires reactivation of a congruent item. We argue this finding points
305 to difficulty in binding the incongruent component into the event (Craik & Tulving, 1975;
306 Bein et al., 2015; Packard et al., 2017), as also indicated by increased fixations. Given that
307 schemas facilitate gist extraction and abstraction of commonalities (Gilboa & Marlatte,
308 2017), the presence of an incongruent component in the event could interrupt this process.
309 This would therefore result in reduced schema instantiation to support binding of the

310 incongruent C component to its congruent pairwise associates, explaining impaired
311 performance on C-B and C-D pairs.

312

313 Interestingly, in the backward order retrieval trials and interference analysis (Experiment 1),
314 as well as in the A-C inference task (Experiment 2), a graded pattern of responses was
315 observed (congruent > incongruent > unrelated). To our knowledge, this is the first study to
316 show such pattern. Although van Kesteren and colleagues (2012) postulated memory for
317 both congruent and incongruent items would be enhanced, the extent to which this effect
318 varies between conditions was unclear. Previous findings pertaining to congruency effects
319 have not used an unrelated baseline condition, making it difficult to account for such
320 differences. Here, we find that whilst components from congruent events were
321 unequivocally better recalled, incongruent components showed better performance
322 compared to unrelated ones. This pattern of results suggests a prediction error can enhance
323 memory performance (compared to unrelated items), though not to the same extent as
324 congruence does. Interestingly, this result was observed only under specific circumstances,
325 suggesting this effect could be susceptible to task demands (Ghosh & Gilboa, 2014).

326

327 In Experiment 1, events were created across blocks, thus online comparisons were between
328 B-C pairs from different events. Conversely, in Experiment 2 comparisons were made with A-
329 B pairs from the same event. This difference in temporal context during encoding, could
330 have biased processing of incongruent pairs in Experiment 1 compared to Experiment 2
331 (Howard & Kahana, 2002). Although overall memory performance in Experiment 1 was
332 symmetric between forward and backward trials, graded responses were observed only in
333 backward trials. Asymmetry in memory recall has been suggested to depend on the
334 relationship between paired associates (Greene & Tussing, 2001; Li & Lewandowsky, 1995;
335 Yang et al., 2013) and to engage the anterior hippocampus (Giovanello, Schnyer, &

336 Verfaellie, 2009). Forward recall is believed to be schema-driven (Geiselman & Callot, 1990)
337 and more susceptible to disruptions during formation of associations at encoding (Li &
338 Lewandowsky, 1995). In line with this, we find superior memory for congruent components,
339 with no differences between incongruent and unrelated pairs in the forward order.
340 Backward recall, on the other hand, is suggested to be more data-driven (Geiselman &
341 Callot, 1990) and thus more susceptible to contextual details at encoding. We therefore
342 suggest backward retrieval in our task promoted the beneficial effect of prediction error,
343 mediated by hippocampal engagement (van Kesteren et al., 2012).

344

345 Graded responses were also obtained in the associative inference task in Experiment 2.
346 Successful performance on such tasks is often used as a marker for schematic organisation
347 of representations in memory, as it supports novel integration of indirectly related items
348 (Kumaran, Summerfield, Hassabis, & Maguire, 2009; Preston & Eichenbaum, 2013; Tse et al.,
349 2007; Zeithamova, Dominick, & Preston, 2012). In this task, incongruent and unrelated pairs
350 were matched on relatedness and differed only on the build-up of expectations from the
351 event's schema. Thus, the presence of a prediction error here could have mediated
352 enhanced inference in this task. Alternatively, the mere existence of a schema in the
353 incongruent condition, as opposed to the unrelated one, could have supported this
354 inference (Kumaran, 2013; McClelland, 2013; Preston & Eichenbaum, 2013; Tse et al., 2011;
355 Zeithamova et al., 2012). Future research on schema effects would benefit from further
356 exploring these effects, specifically in relation to how task demands can modulate memory
357 for incongruent information.

358

359 Our main aim was to test behavioural predictions outlined by SLIMM (van Kesteren et al.,
360 2012). We thus utilised a paradigm that allows expectations to gradually develop by using
361 interleaved learning of paired associates. The findings reported above provide some support

362 to the notion prediction errors can enhance memory for incongruent items. An alternative
363 interpretation is that the amount of prediction error associated with incongruent
364 components in our studies was not large enough to result in conclusively improved
365 performance. Critically, our task was designed to implicitly set participants' predictions. This
366 was done in order to capture the inherent aspect of predictions as they arise in daily life, and
367 to avoid any artificial allocation of attention towards this manipulation. As a result of this
368 manipulation, we could not quantify the amount of prediction error elicited by incongruent
369 components, but only indirectly assert contextual predictions were elicited by the stimuli
370 used (Bar & Aminoff, 2003).

371

372 In conclusion, our results provide further evidence for the notion that schemas aid memory
373 by providing a structured representation into which congruent information can easily fit. The
374 findings reported here also shed light on the extent to which prediction errors in
375 incongruent items support its presence in memory. Although it requires more effortful
376 encoding, retrieval success of incongruent items is always reduced compared to congruent
377 components. The extent to which incongruent items are better remembered compared to
378 unrelated components, on the other hand, is modulated by task demands. Future research
379 looking into schema-mediated memory may build on the approach and findings highlighted
380 above to better understand factors contributing to these effects.

381

382 Methods

383 **Experiment 1**

384 *Participants.* 35 participants (12 males) gave informed consent to take part in the
385 experiment. Four participants whose recognition performance was either above or below
386 three times the IQR were excluded from any further analysis. Thus, data from 31 participants
387 between the ages 18-27 ($M = 19.8$, $SD = 2.91$) were analysed.

388

389 *Materials.* The experiment was controlled using E-Prime 1 (Psychology Software Tools).
390 Stimuli were 30 four-components events (10 congruent, 10 incongruent 10 unrelated
391 events). Each event contained a location (component A, e.g. farm), two objects (components
392 B and D, haystack and a tractor) and a person's profession (component C, farmer). Items
393 were presented as images with labels above them (see Figure 1 for examples). Congruent
394 and incongruent events were constructed to elicit strong contextual predictions, meaning
395 that their components are most likely to appear in the given context, as established by
396 previous work (Bar & Aminoff, 2003). Ten of the events were assigned to be in the
397 incongruent condition, such that the person (C component) was unexpected in the context
398 (a lawyer in the context of a farm with a haystack and a tractor). Another 10 were
399 congruent, meaning the person was expected given the context (a farmer in a farm). The
400 final 10 events included objects with low contextual value (Bar & Aminoff, 2003), meaning
401 they can be found in in a variety of contexts. The allocation of events to conditions was
402 counterbalanced across participants. Images were obtained from freely available online
403 resources labelled with a Creative Commons License.

404

405 *Procedure.* The experiment consisted of three phases: encoding, distraction and retrieval
406 (see Figure 1). The encoding phase was interleaved and took place over 4 blocks, one block
407 for each pairwise association (A-B, B-C, C-D, D-A). Critically, A-B and D-A pairs are not directly
408 associated with the C component (which defines whether an event is assigned to a
409 congruent or incongruent condition). These pairs are congruent in both congruent and
410 incongruent conditions, thus providing an opportunity for testing the effect of incongruence
411 on the rest of the event. Each block consisted of 30 randomised-order trials, resulting in a
412 total of 120 trials in the encoding phase. Each trial began with a one-second fixation cross,
413 followed by presentation of one pair of labelled images for three seconds. Participants were

414 instructed to imagine the components interacting together, as vividly as they could, while
415 being aware of their respective location on the screen (right and left).

416

417 To prevent participants from actively rehearsing the encoded information, a distractor task
418 involving solving arithmetic problems was used for five minutes. Participants were instructed
419 to be as accurate as possible and were informed that if they failed to reach a certain
420 performance threshold their data would be excluded from further analysis. Following this
421 task, the retrieval phase began, where items were presented in a pseudo-randomised order,
422 based on 10 pre-made lists. Two retrieval tasks were employed, a recognition task for each
423 component, followed by a cued-recall task only for the recognised components.

424

425 Participants were first presented with the yes/no recognition task. They had a maximum of
426 10 seconds to complete this task. If they responded 'yes', a second cued-recall task took
427 place immediately. For this task, the recognised component was coded as 'cue' and the
428 recalled component as 'target'. Participants were asked to recall which item appeared with
429 the previously recognised item (a source recall task, retrieve the item in the location
430 indicated on the screen). Each pairwise association from each event was tested in both
431 directions in a randomised order (for example, forward A-? and backward ?-B). A cued-recall
432 answer was scored as correct if it was identical to the item presented at encoding, if it was a
433 specific case of the same category (for example, 'car mechanic' instead of 'mechanic') or
434 semantically similar ('gymnast' instead of 'acrobat'). If the participant failed to give an
435 answer within 20 seconds of the cue being presented, the trial was scored as incorrect. In
436 addition to the 120 items that were encoded, 30 additional items, making up six events,
437 were used as foils.

438

439 **Experiment 2**

440 Experiment 2 was identical to Experiment 1, with the following modifications:

441 *Participants.* 35 participants (9 males) gave informed consent to take part in the experiment.

442 Data from seven participants were removed from any further analysis due to technical

443 failure of the eye-tracker (one participant), poor eye-tracking data (three participants;

444 inclusion of these participants in the behavioural analysis did not change the results) and

445 poor memory performance, above or below three times the IQR (three participants). Thus,

446 data from 28 participants between the ages 18-28 ($M = 20.8$, $SD = 2.78$) were analysed.

447

448 *Materials and apparatus.* The materials were identical to those used in Experiment 1. To

449 examine fixation patterns during encoding, eye movements were recorded during encoding,

450 using an ASL infrared eye tracker (Eye-Trac 6000, Applied Science Laboratories) at a sampling

451 rate of 60 Hz. The desktop-mounted camera was placed under the presentation screen,

452 70cm away from the participant. A chin-rest was used to minimise participants' movement.

453

454 *Procedure.* Before the experiment started, eye calibration was performed using a 9-point

455 matrix. During the encoding phase in Experiment 2, events were presented sequentially,

456 trial-by-trial, rather than across blocks as was done in Experiment 1. Each of the four pairs

457 comprising an event was presented for three seconds, with a 1s fixation cross between

458 them. Following the last pair of the event, a fixation cross was displayed for 2s, until a new

459 event had started. To examine whether reduced performance for the incongruent C

460 components was due to a difficulty in integrating them as part of the events, an additional

461 retrieval inference task was employed. Following the cued-recall task, participants were

462 presented with a location or a person (A or C components) for a maximum of 10 seconds

463 and were asked to recall its counterpart person or location from the same event,

464 respectively.

465

466 *Eye tracking analysis.* Eyeneal software (Applied Science Laboratories) was used to convert
467 the raw gaze coordinates to fixation points. The start of a fixation point was defined as six
468 sequential gaze points with a standard deviation smaller than 0.5 visual degrees. The end of
469 a fixation was marked when three consecutive gaze points were at least one visual degree
470 away from the initial fixation location. The fixation points reported below are the average
471 point of the start and end fixation locations. Two areas of interest (AOI) were defined in the
472 Fixplot software (Applied Science Laboratory), one for each of the images displayed on the
473 screen. Statistical analyses were conducted on the number of fixations and fixation duration
474 for each AOI.

475

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478

479 *Author Contributions Statement:*

480 D.F., D.M., B.W., and D.T. designed the experiment, D.F collected and analysed data, D.F.,
481 D.M., B.W. and D.T. wrote the manuscript.

482

483 *Competing financial interest statement:*

484 The author(s) declare no competing financial interests.

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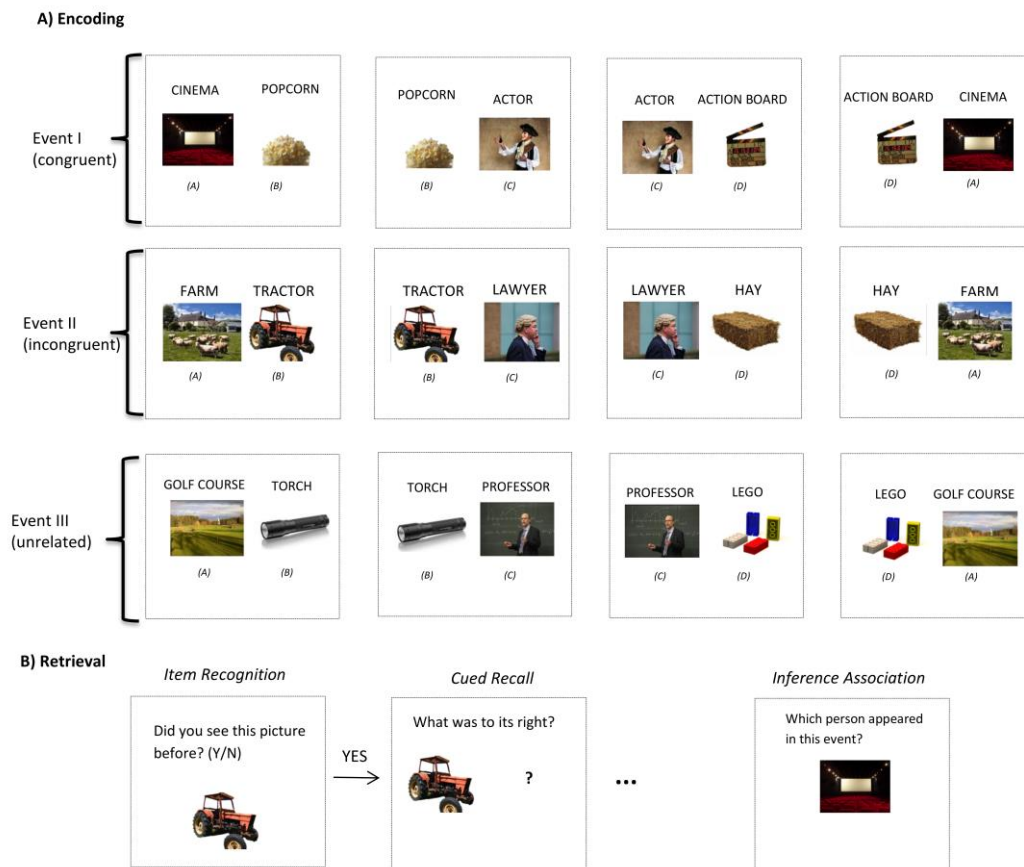
486 References

- 487 Anderson, J. R. (1981). Effects of prior knowledge on memory for new information. *Memory*
488 & *Cognition*, 9(3), 237–246. Retrieved from
489 [http://www.scopus.com/inward/record.url?eid=2-s2.0-](http://www.scopus.com/inward/record.url?eid=2-s2.0-0002794789&partnerID=40&md5=7587d2d916c8d019e0564dfc47322d62%5Cn/Users/Fernando/Documents/Project Based learning/PBL Articles/Anderson1981.pdf)
490 [0002794789&partnerID=40&md5=7587d2d916c8d019e0564dfc47322d62%5Cn/Users/](http://www.scopus.com/inward/record.url?eid=2-s2.0-0002794789&partnerID=40&md5=7587d2d916c8d019e0564dfc47322d62%5Cn/Users/Fernando/Documents/Project Based learning/PBL Articles/Anderson1981.pdf)
491 [Fernando/Documents/Project Based learning/PBL Articles/Anderson1981.pdf](http://www.scopus.com/inward/record.url?eid=2-s2.0-0002794789&partnerID=40&md5=7587d2d916c8d019e0564dfc47322d62%5Cn/Users/Fernando/Documents/Project Based learning/PBL Articles/Anderson1981.pdf)
492 Atienza, M., Crespo-Garcia, M., & Cantero, J. L. (2010). Semantic congruence enhances
493 memory of episodic associations: role of theta oscillations. *Journal of Cognitive*
494 *Neuroscience*, 23(1), 75–90. <http://doi.org/10.1162/jocn.2009.21358>
495 Bar, M., & Aminoff, E. (2003). Cortical analysis of visual context. *Neuron*, 38(2), 347–358.
496 [http://doi.org/10.1016/S0896-6273\(03\)00167-3](http://doi.org/10.1016/S0896-6273(03)00167-3)
497 Bartlett, F. C. (1932). *Remembering: an experimental and social study*. Cambridge:
498 Cambridge University.
499 Bayen, U. J., & Kuhlmann, B. G. (2011). Influences of source-item contingency and schematic
500 knowledge on source monitoring: Tests of the probability-matching account. *Journal of*
501 *Memory and Language*, 64(1), 1–17. <http://doi.org/10.1016/j.jml.2010.09.001>
502 Bein, O., Livneh, N., Reggev, N., Gilead, M., Goshen-Gottstein, Y., & Maril, A. (2015).
503 Delineating the Effect of Semantic Congruency on Episodic Memory: The Role of
504 Integration and Relatedness. *Plos One*, 10(2).
505 <http://doi.org/10.1371/journal.pone.0115624>
506 Bein, O., Reggev, N., & Maril, A. (2014). Prior knowledge influences on hippocampus and
507 medial prefrontal cortex interactions in subsequent memory. *Neuropsychologia*,
508 64(October 2015), 320–330. <http://doi.org/10.1016/j.neuropsychologia.2014.09.046>
509 Brod, G., Werkle-Bergner, M., & Shing, Y. L. (2013). The Influence of Prior Knowledge on
510 Memory: A Developmental Cognitive Neuroscience Perspective. *Frontiers in Behavioral*
511 *Neuroscience*, 7(October), 1–13. <http://doi.org/10.3389/fnbeh.2013.00139>
512 Burton, R. L., Lek, I., & Caplan, J. B. (2017). Associative independence revisited: competition
513 between conflicting associations can be resolved or even reversed in one trial. *The*
514 *Quarterly Journal of Experimental Psychology*, 70(4), 832–857.
515 <http://doi.org/10.1080/17470218.2016.1171886>
516 Caplan, J. B., Rehani, M., & Andrews, J. C. (2014). Associations compete directly in memory.
517 *Quarterly Journal of Experimental Psychology (2006)*, 67(5), 955–78.
518 <http://doi.org/10.1080/17470218.2013.838591>
519 Craik, F. I. M., & Tulving, E. (1975). Depth of processing and the retention of words in
520 episodic memory. *Journal of Experimental Psychology: General*, 104(3), 268–294.
521 <http://doi.org/10.1037/0096-3445.104.3.268>
522 Geiselman, R. E., & Callot, R. (1990). Reverse versus forward recall of script-based texts.
523 *Applied Cognitive Psychology*, 4(2), 141–144. <http://doi.org/10.1002/acp.2350040206>
524 Ghosh, V. E., & Gilboa, A. (2014). What is a memory schema? A historical perspective on
525 current neuroscience literature. *Neuropsychologia*, 53, 104–14.
526 <http://doi.org/10.1016/j.neuropsychologia.2013.11.010>
527 Gilboa, A., & Marlatte, H. (2017). Neurobiology of Schemas and Schema-Mediated Memory.
528 *Trends in Cognitive Sciences*, 1–14. <http://doi.org/10.1016/j.tics.2017.04.013>
529 Giovanello, K. S., Schnyer, D., & Verfaellie, M. (2009). Distinct hippocampal regions make
530 unique contributions to relational memory. *Hippocampus*, 19(2), 111–117.
531 <http://doi.org/10.1002/hipo.20491>
532 Greene, R. L., & Tussing, A. A. (2001). Similarity and associative recognition. *Journal of*
533 *Memory and Language*, 45(4), 573–584. <http://doi.org/10.1006/jmla.2001.2795>
534 Greve, A., Cooper, E., Kaula, A., Anderson, M. C., & Henson, R. N. (2017). Does prediction
535 error drive one-shot declarative learning? *Journal of Memory and Language*, 94, 149–
536 165. <http://doi.org/10.1016/j.jml.2016.11.001>

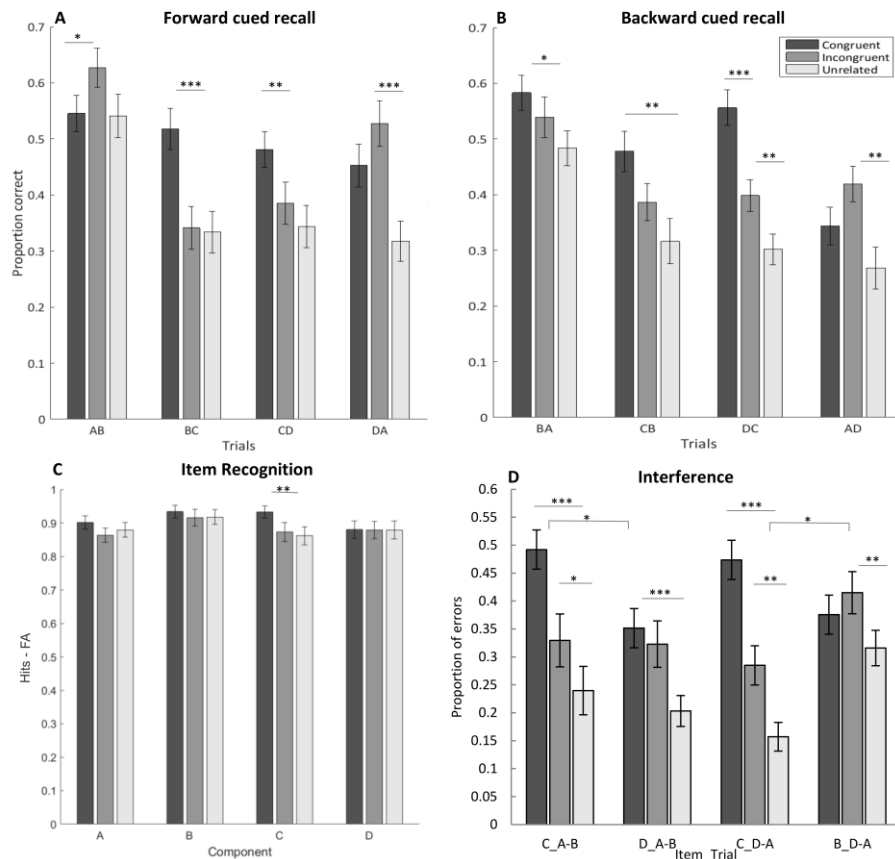
- 537 Hemmer, P., & Steyvers, M. (2009). Integrating Episodic and Semantic Information in
538 Memory for Natural Scenes. *Proceedings of the 31st Annual Conference of the*
539 *Cognitive Science Society*, 1557–1562. Retrieved from
540 <http://csjarchive.cogsci.rpi.edu/proceedings/2009/papers/337/paper337.pdf>
- 541 Horner, A. J., Bisby, J. A., Bush, D., Lin, W.-J., & Burgess, N. (2015). Evidence for holistic
542 episodic recollection via hippocampal pattern completion. *Nature Communications*,
543 *6*(7462), 7462. <http://doi.org/10.1038/ncomms8462>
- 544 Horner, A. J., & Burgess, N. (2013). The associative structure of memory for multi-element
545 events. *Journal of Experimental Psychology. General*, *142*(4), 1370–83.
546 <http://doi.org/10.1037/a0033626>
- 547 Howard, M. W., & Kahana, M. J. (2002). When does semantic similarity help episodic
548 retrieval? *Journal of Memory and Language*, *46*(1), 85–98.
549 <http://doi.org/10.1006/jmla.2001.2798>
- 550 Kahana, M. J., & Caplan, J. B. (2002). Associative asymmetry in probed recall of serial lists.
551 *Memory and Cognition*, *30*(6), 841–849. <http://doi.org/10.3758/BF03195770>
- 552 Kumaran, D. (2013). Schema-driven facilitation of new hierarchy learning in the transitive
553 inference paradigm. *Learning & Memory*, 388–394.
- 554 Kumaran, D., & Maguire, E. A. (2007). Which Computational Mechanisms Operate in the
555 Hippocampus During Novelty Detection ?, *748*, 735–748. <http://doi.org/10.1002/hipo>
- 556 Kumaran, D., Summerfield, J. J., Hassabis, D., & Maguire, E. a. (2009). Tracking the
557 Emergence of Conceptual Knowledge during Human Decision Making. *Neuron*, *63*(6),
558 889–901. <http://doi.org/10.1016/j.neuron.2009.07.030>
- 559 Li, S.-C., & Lewandowsky, S. (1995). Forward and Backward Recall: Different Retrieval
560 Processes. *Journal of Experimental Psychology: Learning, Memory and Cognition*, *21*(4),
561 837–847.
- 562 McClelland, J. L. (2013). Incorporating rapid neocortical learning of new schema-consistent
563 information into complementary learning systems theory. *Journal of Experimental*
564 *Psychology: General*, *142*(4), 1190–1210. <http://doi.org/10.1037/a0033812>
- 565 Moscovitch, M., & Craik, F. I. M. (1976). Depth of processing, retrieval cues and
566 uniqueness of encoding as factors in recall. *Journal of Verbal Learning and Verbal*
567 *Behavior*, *15*, 447–458.
- 568 Packard, P. A., Rodríguez-Fornells, A., Bunzeck, N., Nicolás, B., de Diego-Balaguer, R., &
569 Fuentemilla, L. (2017). Semantic Congruence Accelerates the Onset of the Neural
570 Signals of Successful Memory Encoding. *The Journal of Neuroscience*, *37*(2), 291–301.
571 <http://doi.org/10.1523/JNEUROSCI.1622-16.2017>
- 572 Preston, A. R., & Eichenbaum, H. (2013). Interplay of Hippocampus and Prefrontal Cortex in
573 Memory. *Current Biology*, *23*(17), R764–R773.
574 <http://doi.org/10.1016/j.cub.2013.05.041>
- 575 Schlichting, M. L., & Preston, A. R. (2015). Memory integration: neural mechanisms and
576 implications for behavior. *Current Opinion in Behavioral Sciences*, *1*, 1–8.
577 <http://doi.org/10.1016/j.cobeha.2014.07.005>
- 578 Schlichting, M. L., & Preston, A. R. (2017). The Hippocampus and Memory Integration:
579 Building Knowledge to Navigate Future Decisions. In D. E. Hannula & M. C. Duff (Eds.),
580 *The Hippocampus from Cells to Systems: Structure, Connectivity, and Functional*
581 *Contributions to Memory and Flexible Cognition* (pp. 405–437). Cham: Springer
582 International Publishing. http://doi.org/10.1007/978-3-319-50406-3_13
- 583 Schulman, A. I. (1974). Memory for words recently classified. *Memory & Cognition*, *2*(1), 47–
584 52. <http://doi.org/10.3758/BF03197491>
- 585 Staresina, B. P., Gray, J. C., & Davachi, L. (2009). Event Congruency Enhances Episodic
586 Memory Encoding through Semantic Elaboration and Relational Binding. *Cerebral*
587 *Cortex*, *19*(5), 1198–1207. <http://doi.org/10.1093/cercor/bhn165>

- 588 Steyvers, M., & Hemmer, P. (2012). Reconstruction from Memory in Naturalistic
589 Environments. In B. H. Ross (Ed.), *The Psychology of Learning and Motivation* (Vol. 56,
590 pp. 125–144). <http://doi.org/10.1016/B978-0-12-394393-4.00004-2>
- 591 Tse, D., Langston, R. F., Kakeyama, M., Bethus, I., Spooner, P. A., Wood, E. R., ... Morris, R. G.
592 M. (2007). Schemas and Memory Consolidation. *Science*, *316*(5821), 76–82.
593 <http://doi.org/10.1126/science.1135935>
- 594 Tse, D., Takeuchi, T., Kakeyama, M., Kajii, Y., Okuno, H., Tohyama, C., ... Morris, R. G. M. M.
595 (2011). Schema-Dependent Gene Activation. *Science*, *333*(August), 891–895.
596 <http://doi.org/10.1126/science.1205274>
- 597 van Kesteren, M. T. R., Beul, S. F., Takashima, A., Henson, R. N., Ruiters, D. J., & Fernández, G.
598 (2013). Differential roles for medial prefrontal and medial temporal cortices in schema-
599 dependent encoding: From congruent to incongruent. *Neuropsychologia*, *51*(12), 2352–
600 2359. <http://doi.org/10.1016/j.neuropsychologia.2013.05.027>
- 601 van Kesteren, M. T. R., Rijpkema, M., Ruiters, D. J., & Fernandez, G. (2010). Retrieval of
602 Associative Information Congruent with Prior Knowledge Is Related to Increased
603 Medial Prefrontal Activity and Connectivity. *Journal of Neuroscience*, *30*(47), 15888–
604 15894. <http://doi.org/10.1523/JNEUROSCI.2674-10.2010>
- 605 van Kesteren, M. T. R., Ruiters, D. J., Fernández, G., & Henson, R. N. (2012). How schema and
606 novelty augment memory formation. *Trends in Neurosciences*, *35*(4), 211–219.
607 <http://doi.org/10.1016/j.tins.2012.02.001>
- 608 Yang, J., Zhao, P., Zhu, Z., Mecklinger, A., Fang, Z., & Li, H. (2013). Memory asymmetry of
609 forward and backward associations in recognition tasks. *Journal of Experimental*
610 *Psychology: Learning Memory and Cognition*, *39*(1), 253–269.
611 <http://doi.org/10.1037/a0028875>
- 612 Zeithamova, D., Dominick, A. L., & Preston, A. R. (2012). Hippocampal and Ventral Medial
613 Prefrontal Activation during Retrieval-Mediated Learning Supports Novel Inference.
614 *Neuron*, *75*(1), 168–179. <http://doi.org/10.1016/j.neuron.2012.05.010>
- 615
616

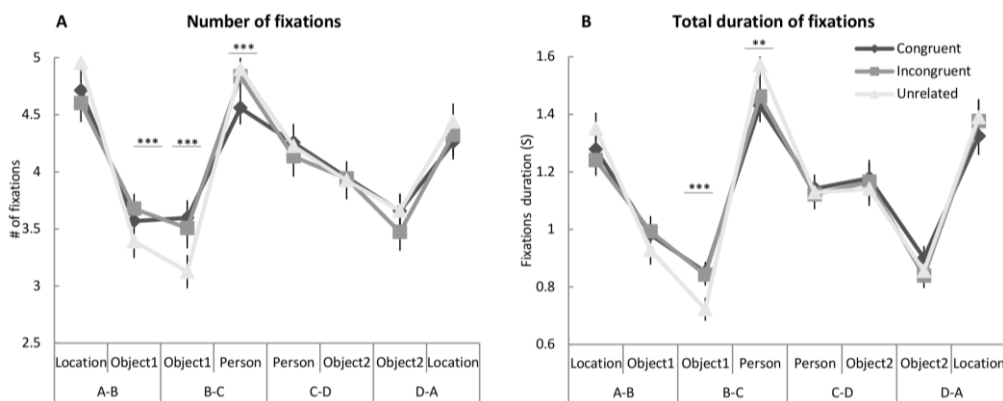
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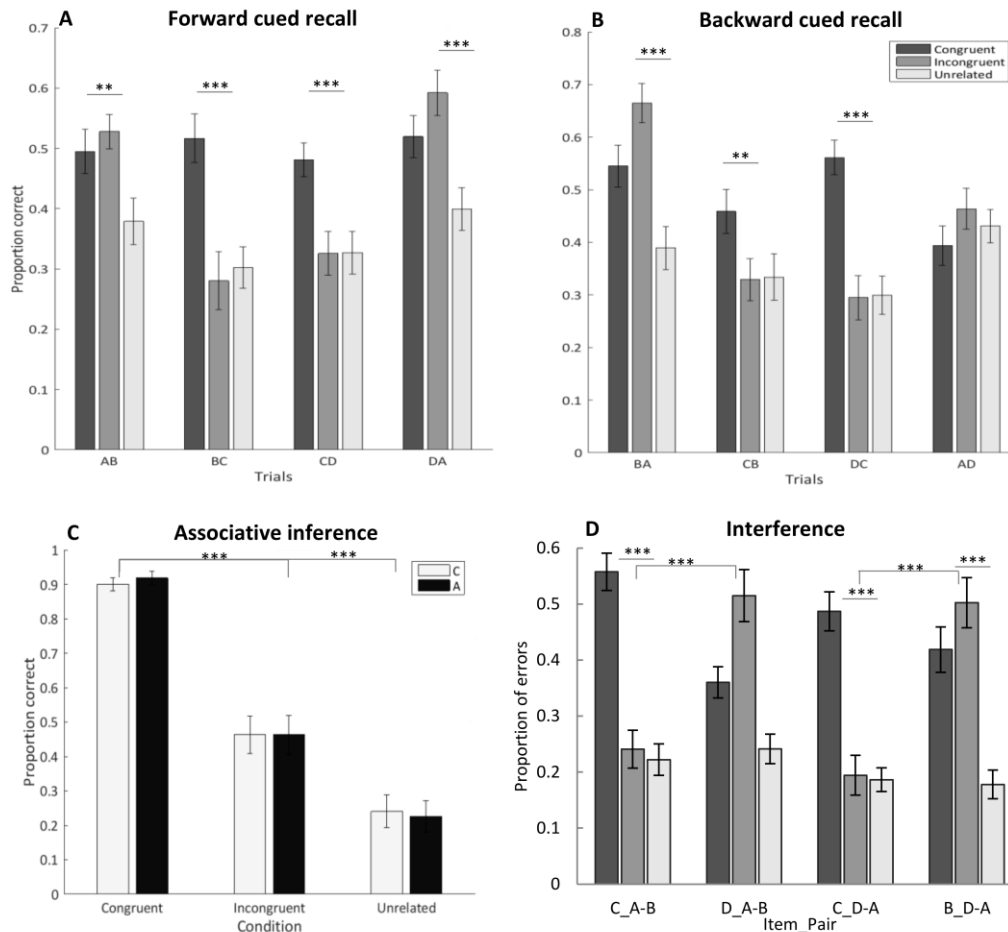
618
 619 **Figure 1 | Experimental Design.** A) Encoding phase. Participants 120 paired
 620 associates, over four blocks, one for each pairwise association from every event.
 621 They imagined each pair interacting in a meaningful way for 3s. Each pair was
 622 preceded by a 1s fixation cross. B) Retrieval phase. Participants were presented with
 623 a cue and asked to indicate whether they remember seeing it at encoding. If they
 624 responded ‘yes’ they were asked to recall one of the other components from the
 625 same event, based on the spatial location of the cue. Inference association task was
 626 used in Experiment 2. Labels in parentheses are for illustration and were not
 627 presented during the experiment.
 628



629
 630 **Figure 2 | Results Experiment 1.** A) Cued recall performance forward pairs (e.g. A-?).
 631 B) Cued recall performance backward pairs (e.g. ?-B). In both orders performance
 632 tracks levels of relatedness of pairs, such that most incongruent and unrelated pairs
 633 are equivocal. C) Item recognition, Congruent C components show better accuracy
 634 compared to incongruent ones. D) Interference analysis, percentage of erroneously
 635 recalled C items in the cued recall task. Most interference from congruent items,
 636 followed by incongruent and then unrelated. Unless otherwise states, error bars
 637 represent standard error of mean. * $p \leq .05$, ** $p \leq .01$, *** $p \leq .001$



638 **Figure 3 | Eye tracking results Experiment 2.** A) Number of fixations on each item,
 639 per pair. Increased fixations on first presentation of incongruent and unrelated
 640 items. B) Total time spent fixating on each item of the pair during the 3s encoding
 641 trial. More time spent fixating on unrelated C items. * $p \leq .05$, ** $p \leq .01$, *** $p \leq$
 642 $.001$
 643



644 **Figure 4 | Behavioural results Experiment 2.** A) Cued recall performance forward
 645 pairs (e.g. A-?), incongruent items on par with unrelated ones. B) Cued recall
 646 performance backward pairs (e.g. ?-B), similarly to forward order, incongruent and
 647 unrelated items associated with reduced performance. C) Associative inference,
 648 recall of A cued by C and vice versa. Near-ceiling performance for congruent items,
 649 followed by incongruent items associated with better performance than unrelated
 650 ones D) Interference analysis, higher percentage of erroneously recalled congruent C
 651 items in the cued recall task. * $p \leq .05$, ** $p \leq .01$, *** $p \leq .001$
 652
 653